#### RESEARCH



# Myosteatosis and the survival of patients with hepatocellular carcinoma: a meta-analysis

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#### Abstract

Myosteatosis, characterized by fat infiltration into skeletal muscle, is increasingly recognized as a prognostic factor in hepatocellular carcinoma (HCC), although the results were not consistent. This meta-analysis aimed to summarize impact on overall survival (OS) and progression-free survival (PFS) in patients with HCC. A systematic search of PubMed, Embase, and Web of Science was conducted to identify observational studies reporting survival outcomes in HCC patients with and without myosteatosis. Pooled hazard ratios (HRs) with 95% confidence intervals (CIs) were calculated using a random-effects model. A total of 24 retrospective cohort studies involving 7436 HCC patients were included. Myosteatosis was significantly associated with poorer OS (HR: 1.60, 95% CI: 1.40–1.83, p < 0.001,  $l^2 = 65\%$ ) and PFS (HR: 1.53, 95% CI: 1.33–1.76, p < 0.001,  $I^2 = 36\%$ ). Subgroup analysis revealed a stronger association in Asian studies (HR: 1.74 for OS; 1.57 for PFS) compared to European studies (HR: 1.08 for OS; 1.05 for PFS). The prognostic impact remained significant regardless of anticancer treatment type, myosteatosis assessment method, sex-specific or universal cutoff values, and follow-up duration (p for subgroup differences all > 0.05). The results remained significant in studies adjusting for sarcopenia (HR: 1.89 for OS; 1.50 for PFS). Meta-regression analyses did not suggest any of the following variables may affect the results, including sample size, mean ages of the patients, proportions of men, follow-up durations, and study quality scores (p all > 0.05). Myosteatosis is independently associated with worse survival in HCC patients, particularly in Asian populations. These findings highlight the significance of assessing muscle quality as a prognostic factor in HCC.

**Keywords** Hepatocellular carcinoma · Myosteatosis · Survival · Progression · Meta-analysis

Abbreviation	ons	TAE	Transarterial embolization
HCC	Hepatocellular carcinoma	TKI	Tyrosine kinase inhibitor
OS	Overall survival	PRISMA	Preferred Reporting Items for Systematic
PFS	Progression-free survival		Reviews and Meta-Analyses
HR	Hazard ratio	NOS	Newcastle-Ottawa Scale
CI	Confidence interval	PROSPERO	International Prospective Register of Sys-
CT	Computed tomography		tematic Reviews
MRI	Magnetic resonance imaging	BMI	Body mass index
IMAC	Intramuscular adipose content	MASLD	Metabolic dysfunction-associated steatotic
SMD	Skeletal muscle density		liver disease
TACE	Transarterial chemoembolization	RFA	Radiofrequency ablation
		CAR-T	Chimeric antigen receptor T-cell therapy

Yongjuan Wu and Guangyuan Cheng have contributed equally to this work.

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#### Introduction

Hepatocellular carcinoma (HCC) is the most common primary liver malignancy and a leading cause of cancerrelated mortality worldwide [1, 2]. The global incidence of HCC continues to rise, particularly in regions with high



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hepatitis B and C virus prevalence, alcohol-related liver disease, and metabolic dysfunction-associated steatotic liver disease (MASLD) [3]. Despite advances in surveillance and therapeutic strategies, the prognosis of HCC remains poor, with a five-year survival rate below 20% in most populations [4]. Treatment options for HCC are diverse and depend on tumor stage, liver function, and patient performance status [5]. Curative approaches, such as surgical resection and liver transplantation, are feasible in early-stage disease [6], while locoregional therapies, including transarterial chemoembolization (TACE) and radiofrequency ablation (RFA), are commonly used for intermediate-stage tumors [7]. Systemic treatments, such as tyrosine kinase inhibitors (TKIs), immune checkpoint inhibitors, and combination regimens, have expanded options for patients with advanced disease [8]. However, treatment response varies, underscoring the need for reliable prognostic markers to refine risk stratification and optimize patient management [9].

In this context, the identification of novel predictors of survival is crucial for improving clinical outcomes in HCC. Myosteatosis, defined as pathological fat infiltration within skeletal muscle, has emerged as a potential prognostic factor in cancer patients [10, 11]. Unlike sarcopenia, which refers to a loss of muscle mass and strength, myosteatosis reflects alterations in muscle composition and quality [12]. It is commonly assessed using computed tomography (CT) or magnetic resonance imaging (MRI) at the level of the third lumbar vertebra, with specific parameters such as skeletal muscle density (SMD) and intramuscular adipose content (IMAC) used to quantify fat infiltration [13, 14]. The presence of myosteatosis is often associated with systemic inflammation, insulin resistance, and metabolic dysfunction, which may contribute to cancer progression and poor survival outcomes [15–17].

In patients with HCC, myosteatosis may promote an unfavorable prognosis through several mechanisms, such as chronic inflammation, hyperinsulinemia, increased circulating free fatty acids, and oxidative stress, which can create a tumor-promoting microenvironment [18]. Despite emerging evidence linking myosteatosis to poor outcomes in HCC, findings across studies remain inconsistent. Some reports suggest a strong association between myosteatosis and reduced overall survival (OS) and progression-free survival (PFS) in patients with HCC [19-36], while others have found no significant impact [37–42]. Differences in study populations, imaging modalities, cutoff values, and adjustment for confounders such as sarcopenia may account for these discrepancies. Given these uncertainties, in this study, we performed a meta-analysis aiming to systematically evaluate the association between myosteatosis and survival outcomes in patients with HCC.



#### Methods

The study adhered to PRISMA 2020 [43, 44] and the Cochrane Handbook for Systematic Reviews and Meta-analyses [45] guidelines for conducting this meta-analysis, including for the study protocol design, data extraction, statistical analysis, and results presentation. The protocol of the meta-analysis has been registered at PROSPERO with the identifier CRD42025637129.

#### Literature search

To identify studies pertinent to this meta-analysis, we searched PubMed, Embase, and Web of Science databases using an extensive array of search terms, which included: (1)"myosteatosis"OR"muscle density"OR"muscle attenuation"OR"intramuscular adipose tissue content"OR"intramuscular adipose tissue infiltration"OR"intramuscular aditissue deposition"OR"intramuscular pose content" OR "intramuscular fat infiltration"OR"intramuscular fat deposition"; and (2)"hepatocellular cancer"OR"hepatocellular tumor"OR"hepatocellular carcinoma"OR"hepatocellular neoplasm"OR"liver cancer"OR"liver tumor"OR"liver carcinoma"OR"liver neoplasm"OR"HCC"OR"hepatic cancer"OR"hepatic tumor"OR"hepatic carcinoma."The search was restricted to studies conducted on human subjects and included only full-length articles published in English in peer-reviewed journals. Additionally, the references of relevant original and review articles were manually screened to identify any additional eligible studies. The literature search covered the period from the inception of the databases to January 18, 2025. The detailed search strategy for each database is shown in Supplemental File 1.

# Inclusion and exclusion criteria

The inclusion criteria for potential studies were defined according to the PICOS framework:

P (patients): Adult patients (aged 18 years or older) with confirmed diagnosis of HCC, regardless of the cancer etiology, stage, or main anticancer treatments.

I (exposure): Patients with myosteatosis. The methods, parameters, and cutoffs for the diagnosis of myosteatosis were consistent with those used in the original studies.

C (comparison): Patients without myosteatosis.

O (outcome): Survival outcomes, including OS and PFS, compared between patients with and without myosteatosis. In general, OS is defined as the time from treatment

initiation to death from any cause, while PFS is defined as the time from treatment initiation to disease progression or death, whichever occurs first.

S (study design): Observational studies with longitudinal follow-up, such as cohort studies, nested case—control studies, or post hoc analyses of clinical trials; no minimum follow-up duration or specific covariate adjustment in multivariate analyses was required for inclusion, in order to comprehensively capture all relevant observational evidence.

Studies were excluded if they were reviews, editorials, meta-analyses, preclinical research, or involved patients without HCC, lacked myosteatosis as exposure, or did not report the survival outcomes of interest. In cases where multiple publications appeared to involve overlapping cohorts (e.g., from the same institution with similar enrollment periods), we included only the study with the largest sample size or most comprehensive data to avoid double counting of patients.

# Study quality assessment and data extraction

The literature search, study selection, quality assessment, and data extraction were independently performed by two authors. Key data—including hazard ratios (HRs), confidence intervals (CIs), and study characteristics were cross-checked for consistency. Discrepancies were resolved through discussion and, when needed, consultation with the corresponding author. Study quality was assessed using the Newcastle-Ottawa Scale (NOS) [46], which evaluates selection, control of confounding factors, and outcome measurement and analysis, with scores ranging from 1 to 9, where a score of 9 indicates the highest quality. Studies with the NOS scores of 7 or above were generally considered as high-quality studies [46]. For this meta-analysis, we considered the exposed cohort representative if patients were consecutively or randomly selected. Adjustment for other confounding factors was credited if variables beyond age and sex were included in the multivariate models. A follow-up duration of at least 36 months was considered sufficient for the outcome to occur, and adequacy of follow-up was defined as having less than 10% loss to follow-up. Data extracted for analysis included study characteristics (author, year, country, and design), participant details (number of patients, mean age, sex, and main anticancer treatments), images (CT or MRI), parameters (IMAC or SMD), or cutoffs (sex-specified or not) for evaluating myosteatosis, number of patients with myosteatosis, mean follow-up durations, and variables adjusted when the association between myosteatosis and the survival outcomes of patients with HCC was analyzed.

#### Statistical analyses

The associations between myosteatosis and OS/PFS of patients with HCC were summarized as HRs and corresponding 95% CIs, compared between patients with and without myosteatosis. The HRs and their standard errors were derived from 95% CIs or p values and subsequently log-transformed to stabilize variance and achieve a normalized distribution [45]. To assess heterogeneity, we used the Cochrane Q test and  $I^2$  statistics [47], with  $I^2 < 25\%$ ,  $25 \sim 75\%$ , and > 75% indicating low, moderate, and high heterogeneity. A random-effects model was applied to integrate the results, accounting for study variability [45]. Via excluding individual studies sequentially, a sensitivity analysis was performed to evaluate the robustness of the findings. In addition, subgroup analyses were performed to evaluate study characteristics on the outcomes, such as study country (Asian versus non-Asian), main anticancer treatments, parameters (IMAC or SMD) or cutoffs (sexspecified or not) for evaluating myosteatosis, mean followup durations, or whether sarcopenia was adjusted when the association between myosteatosis and the survival outcomes of patients with HCC was analyzed. For continuous variables such as follow-up duration, the median value across included studies was used to define subgroups (e.g., short vs. long follow-up), a common practice in meta-analyses when no standardized thresholds exist. Regarding diagnostic criteria for myosteatosis, studies were classified as using'sexspecific'cutoffs if different thresholds were applied for men and women, and universal'if a single cutoff was applied to all participants, in accordance with the methods described in the original studies. In addition, univariate meta-regression analyses were performed to evaluate the influence of study characteristics in continuous variables on the outcome, such as sample sizes of the studies, mean ages of the patients, proportions of men, follow-up durations, and study quality scores in NOS [45]. Publication bias was assessed by visually inspecting funnel plot symmetry, where substantial asymmetry may suggest small-study effects or selective publication. Egger's regression test was also performed to statistically detect funnel plot asymmetry, with p values < 0.05 indicating potential bias [48]. Analyses were performed using RevMan (version 5.1; Cochrane Collaboration, Oxford, UK) and Stata software (version 12.0; Stata Corporation, College Station, TX, USA).

#### Results

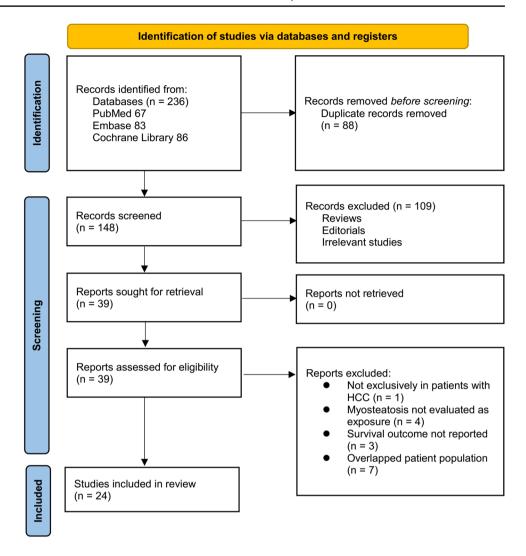
#### Study identification

The study selection process is summarized in Fig. 1. A total of 236 potentially relevant records were initially identified



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Fig. 1 Flowchart of database search and study inclusion



from the three databases searched and screening of citations of related articles, with 88 duplicates removed. Screening of titles and abstracts resulted in the exclusion of 109 articles that did not meet the objectives of the meta-analysis. The full texts of the remaining 39 articles were independently reviewed by two authors, leading to the exclusion of 15 studies for various reasons detailed in Fig. 1. Ultimately, 24 studies were included in the quantitative analysis [19–42].

### **Overview of the study characteristics**

Table 1 shows the summarized characteristics of the available studies included in the meta-analysis. Overall, 24 retrospective cohorts, published between 2015 and 2025, and performed in Japan, Indonesia, Korea, German, China, Italy, and Thailand, were involved in the meta-analysis [19–42]. A total of 7436 patients with HCC were included in these studies. The mean ages of the patients varied from 50.0 to 74.0 years, and the proportion of men varied from 65.9 to 100.0%. The main anticancer treatments were surgical

resection in seven studies [20, 21, 23, 27, 28, 35, 39], liver transplant in two studies [30, 32], TACE or transarterial embolization (TAE) in four studies [22, 25, 38, 42], immunotherapy in four studies [24, 26, 31, 33], and TKIs in five studies [29, 34, 36, 40, 41]. The main anticancer treatments were not specified in one study [19], and a comprehensive treatment involving TACE, RFA, surgical resection, chemoradiotherapy, and supportive treatment was used in another study [37]. All the included studies evaluated myosteatosis using CT imaging at the level of the third lumbar vertebra except one study [41], in which the MRI imaging at the same level was used. Myosteatosis was evaluated using IMAC in six studies [20, 21, 27, 28, 35, 38] and using SMD in the other 18 studies [19, 22-26, 29-34, 36, 37, 39-42]. A sexspecified cutoff for the diagnosis of myosteatosis was used in 11 studies [19–22, 25, 27, 32, 33, 35, 38, 42], and a universal cutoff was used in the other 13 studies [23, 24, 26, 28–31, 34, 36, 37, 39–41]. Accordingly, 3260 patients (43.8%) had myosteatosis. The median follow-up durations were 10 to 71 months. The outcome of OS was reported in 21 studies



 Table 1
 Characteristics of the included studies

Outcomes Variables  up reported adjusted	OS and PFS Age, sex, BMI, sarcopenia, ICGR15, albumin, total bilirubin, prothrombin time, AST, platelet count, AFP, PIVKA-II, operative blood loss, tumor stage, and tumor size	OS Age, sex, viral status, BCLC stage, treatment methods for BCLC stage 0/A, prothrom- bin time, presence of previous treatment, and presence of CKD	OS Age, sex, BMI, sarcopenia, visceral adiposity, viral satus, Child-Phoh
No. of Median patients with follow-up myosteatosis duration (months)	71 48	30	65 13
Cutoff values of myosteatosis parameters	Previously defined cutoffs: -0.44 for men and -0.31 for women	Maximally selected rank statistics determined cutoffs: men: \$\leq 44.4 \text{ HU; women: }\leq 39.3 \text{ HU}	Previously defined: men: ≤44.4 HU; women: ≤39.3 HU
r Param- s eter for myosteatosis evaluation	IMAC	SMD	SMD
Methods for myosteatosis evaluation	CT, L3	d CT, L3	CT, L3
Main anticancer treatment	Surgical resection	Not specified CT, L3	TACE or supportive treatment
Men (%)	75.9	65.9	47
Mean age (years)	NR R	8.8	55
Design No. of patients included	141	1257	100
Desig	RC	RC	ı RC
Country	Japan	Japan	] Indonesia
Study	Kaibori [20] Japan	Fujiwara [19]	Mardian [22] Indonesia



Table 1 (continued)	tinued)												
Study	Country	Design	Design No. of patients included	Mean age (years)	Men (%) Main antica treatm	Main anticancer treatment	Methods for myosteatosis evaluation	Param- eter for myosteatosis evaluation	Cutoff values of myosteatosis parameters	No. of patients with myosteatosis	Median follow-up duration (months)	Outcomes	Variables adjusted
Hamaguchi [21]	Japan	RC	909	89	79.9	Surgical	CT, L3	IMAC	Previously defined cutoffs: -0.36 for men and -0.23 for women	258	43	OS and PFS	Age, sex, BMI, etiology, previous treatment, comorbidi- ties, Child- Pugh Class, tumor size, stage, num- ber, surgical character- istics, and sarcopenia
Sano [37]	Japan	RC	187	6:69	r.27	TACE, RFA, surgical resection, chemoradiotherapy, and supportive treatment	CT, L3	SMD	Previously defined cutoffs: <41 HU in patients with a BMI <25 kg/m² and <33 HU in those with BMI ≥25 kg/m²	110	90	so	Age, sex, Child-Pugh Class, etiology, and cancer stage
Jang [23]	Когеа	RC	091	55.2	75	Surgical resection	CT, L3	SMD	ROC curve analysis determined cutoffs: men: ≤46.1 HU; women: ≤48.4 HU	XX	71	OS and PFS	Age, sex, obesity, eti- ology, PLT, albumin, MELD-Na score, AFP, cancer stage, and sarco- penia



Table 1 (continued)

Age, sex, and hemoglobin rhosis, tumor size, surgery BČLC stage, BMI, Child-Pugh Class, surgery characteris-Age, sex, BMI, ASA Age, sex, BMI, AFP, sarcopenia sarcopenia tics, tumor sarcopenia Class, ciristics, and metastatic status, and albumin, cirrhosis, character-Age, sex, etiology, size, and Variables adjusted OS and PFS OS and PFS OS and PFS Outcomes reported PFS SO Median follow-up duration (months) 30 10 56 13 28 patients with myosteatosis No. of 115 170 36 32 9 Median: 32.2 HU <33 HU in myosteatosis with a BMI determined  $BMI \ge 25$  $kg/m^2$ in patients those with ROC curve specified -0.44 for -0.31 for <25 kg/ The lowest parameters <41 HU Previously men and Previously cutoffs: defined analysis  $m^2$  and values of defined cutoffs: women cutoffs by sex tertile Cutoff myosteatosis evaluation eter for IMAC IMAC SMD SMD SMD myosteatosis Methods for evaluation CT, L3 CT, L3 CT, L3 CT, L3 CT, L3 Lenvatinib resection resection anticancer Immunotherapy treatment Surgical Surgical Men (%) Main TAE 84.6 8.9/ 6.97 80.8 72 age (years) 73.2 20 74 99 67 patients included No. of 100 245 151 52 65 Design RCRCRCRCRCCountry Germany China China Japan Italy Meister [39] Masetti [38] Yamamoto Shi [27] Yi [24] Study [40]



Table 1 (continued)	tinued)												
Study	Country	Design No. of patient include	No. of patients included	Mean age (years)	Men (%) Main antica treatm	Main anticancer treatment	Methods for myosteatosis evaluation	Parameter for myosteatosis evaluation	Cutoff values of myosteatosis parameters	No. of patients with myosteatosis	Median follow-up duration (months)	Outcomes	Variables adjusted
Yoshikawa [41]	Japan	RC	65	47	76.9	Lenvatinib	MRI, L3	SMD	Median: 32.2 HU	32	30	PFS	Age, sex, tumor num- ber, size, ICGR15, tumor stage, differen- tiation, AFP, PIVKA-II, and sarco- penia
Chen [26]	Taiwan (China)	RC	Ξ	59	4.78	therapy	CT, L3	SMD	Previously defined cutoffs: <41 HU in patients with a BMI <25 kg/ m² and <33 HU in those with BMI >25 kg/kg/m² kg/m²	91	31	OS and PFS	Age, sex, HBsAg, anti-HCV, AFP, Child- Pugh Class, BCLC stage, lines of therapy, and sarcopenia
Bannang- koon [25]	Thailand	RC	611	61.4	72.8	TACE	CT, L3	SMD	Previously defined: men: ≤44.4 HU; women: ≤39.3 HU	237	42	SO	Age, sex, chronic lung disease, and CKD



Table 1 (continued)

prothrombin time, AST, Age, sex, obesity, DM, Age, sex, ECOG PS, Child–Pugh Class, BCLC size, meta-static status, cancer stage, and sarcohyperten-sion, tumor Child-Pugh static status, operative blood loss, Class, AFP, stage, meta-PLT, albu-min, total and surgery sarcopenia, ALT, AFP, sarcopenia PIVKA-II, characterbilirubin, number, Variables Age, sex, adjusted penia istics OS and PFS OS and PFS Outcomes reported PFS dn-wolloj duration (months) Median 9 15 20 patients with myosteatosis No. of Ř 59 22 in patients with a BMI myosteatosis <33 HU in <33 HU in with a BMI  $BMI \ge 25$   $kg/m^2$ determined  $BMI \ge 25$   $kg/m^2$ in patients those with those with cutoffs: <41 HU ROC curve Previously < 25 kg/ m<sup>2</sup> and parameters <41 HU analysis Previously <25 kg/ defined defined values of cutoffs:  $m^2$  and cutoff: -0.46Cutoff myosteatosis evaluation eter for IMAC SMD SMD Methods for myosteatosis evaluation CT, L3 CT, L3 CT, L3 resection anticancer Sorafenib therapy treatment Immuno-Surgical Men (%) Main 84.6 86.1 9.77 age (years) 68.9 54 67 patients included Design No. of 116 188 245 RCRCRCCountry China Korea Japan [Shida [28] Kang [29] Liu [31] Study



Age, sex, BMI, MELD Age, sex, total bilirubin, BCLC stage, number, dif-ferentiation, AFP, and score, diabetes status, hepatitis B, Child-Pugh BCLC stage. Age, sex, etiology, cirrhosis, ECOG PS, ALT, AST, AFP, and tumor numtumor size, sarcopenia rubin level, and tumor Age, BMI, cirrhosis, sarcopenia ber, size, albumin, etiology total biliburden Variables Class, scores adjusted OS and PFS Outcomes reported OS OS OS follow-up duration (months) Median 18 9 17 31 patients with myosteatosis No. of 160 187 82 37 myosteatosis <33 HU in with a BMI  $BMI \ge 25$   $kg/m^2$ P Previously determined determined ≤31.2 HU; in patients those with ≤23.8 HU ≤37.5 HU ≤28 HU; ROC curve <41 HU <25 kg/ Previously Maximally parameters ≤27 HU statistics analysis defined: selected cutoffs: women: women: values of defined cutoffs:  $m^2$  and cutoffs: men: men: Cutoff myosteatosis evaluation eter for Param-SMD SMD SMD SMD Methods for myosteatosis evaluation CT, L3 CT, L3 CT, L3 CT, L3 anticancer therapy treatment Immuno-TACE Men (%) Main  $\Gamma$ LT 85.3 82.3 87.9 100 (years) 55.6 66.4 53.4 age 55 patients included Design No. of 784 673 305 224 RCRCRCRCGermany Country China China Ouyang [33] China Table 1 (continued) Surov [34] Li [31] Lu [32] Study



Table 1 (continued)

Study	Country	Design	Design No. of patients included	Mean age (years)	Men (%) Main antica treatn	Main anticancer treatment	Methods for myosteatosis evaluation	Parameter for myosteatosis evaluation	Cutoff values of myosteatosis parameters	No. of patients with myosteatosis	Median follow-up duration (months)	Outcomes reported	Variables adjusted
Surov [42]	10 European countries	RC	363	66.1	87.1	sorafenib	CT, L3	SMD	Previously defined cutoffs: <41 HU in patients with a BMI <25 kg/m² and <33 HU in those with BMI ≥25 kg/m²	148	30	so	Age, sex, ECOG PS, albumin, PLT, BCLC stage, and extrahepatic metastases
Luo [36]	Taiwan (China)	RC .	<del></del>	68.4	82.7	Lenvatinib	CT, L3	SMD	Previously defined cutoffs: <41 HU in patients with a BMI <25 kg/ m² and <33 HU in those with BMI ≥25 kg/m²	36	20	OS and PFS	Age, sex, underweight, HbsAg, HCV status, AFP, other therapy, tumor size, BCLC stage, ALBI grade, and sarco- penia
Lee [35]	Japan	RC	909	89	79.9	Surgical resection	CT, L3	IMAC	Previously defined cutoffs: -0.36 for men and -0.23 for women	258	36	OS and PFS	Age, sex, underweight, HbsAg, HCV status, AFP, SCr, ALT, AST, PT, PLT, ALBI grade, tumor features, and sarcopenia

K absence-II; PLT, platelet count; PT, prothrombin time; RC, retrospective cohort; RFA, radiofrequency ablation; ROC, receiver operating characteristic; SCr, serum creatinine; SIRT, selective internal radiation therapy; SMD, skeletal muscle density; TACE, transarterial chemoembolization; TAE, transarterial embolization MELD, Model for End-Stage Liver Disease; NR, not reported; MRI, magnetic resonance imaging; OS, overall survival; PFS, progression-free survival; PIVKA-II, protein induced by vitamin AFP, alpha-fetoprotein; ALBI, albumin-bilirubin; ALT, alanine aminotransferase; ASA, American Society of Anesthesiologists; AST, aspartate aminotransferase; BCLC, Barcelona Clinic Liver Cancer; BMI, body mass index; CKD, chronic kidney disease; CT, computed tomography; DM, diabetes mellitus; ECOG PS, Eastern Cooperative Oncology Group performance status; HBsAg, hepatitis B surface antigen; HCV, hepatitis C virus; HU, Hounsfield unit; ICGR15, indocyanine green retention rate at 15 min; IMAC, intramuscular adipose content; LT, liver transplantation;



[19-26, 28-30, 32-40, 42], and the outcome of PFS was reported in 15 studies [20, 21, 23, 24, 26–29, 31, 32, 35, 36, 39–41]. Since one of the included studies reported the outcome in men and women separately [35], these datasets were independently included in the meta-analysis, making 22 and 16 datasets available for the outcome of OS and PFS. The multivariate analyses were performed in all studies except one study, in which the univariate analysis was performed [38]. In total, 16 of the 24 included studies [20–23, 26–29, 31–33, 35, 36, 39–41] adjusted for sarcopenia in their multivariate analyses. In addition to sarcopenia, commonly adjusted variables included age, sex, BMI, liver function markers (e.g., Child-Pugh Class, albumin, bilirubin), tumorrelated characteristics (e.g., size, stage, AFP), and treatment modality. The specific covariates adjusted for in each study are detailed in Table 1. The included studies achieved NOS scores ranging from six to nine, reflecting a generally moderate to high quality of methodology and reporting (Table 2).

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# Association between myosteatosis and OS in patients with HCC

Overall, 21 studies [19–26, 28–30, 32–40, 42] with 22 datasets evaluated the association between myosteatosis and OS in patients with HCC. A moderate heterogeneity was observed among these studies (p for Cochrane Q test < 0.001;  $I^2 = 65\%$ ). The pooled results showed that overall, myosteatosis was associated with poorer OS in patients with HCC (HR: 1.60, 95% CI: 1.40 to 1.83, p < 0.001; Fig. 2A). Sensitivity analyses by excluding one study at a time showed similar results (HR: 1.55 to 1.65, p all < 0.05). Specifically, excluding the only study with univariate meta-analysis showed consistent results (HR: 1.63, 95% CI: 1.42 to 1.87, p < 0.001;  $I^2 = 65\%$ ). In addition, the sensitivity analysis limited to studies with NOS score of 7 or higher also showed consistent results (HR: 1.63, 95% CI: 1.43 to 1.85, p <0.001;  $I^2 = 55\%$ ). Moreover, the subgroup analysis showed that the association between myosteatosis and poor OS of HCC was significant in Asian studies but not in European studies (HR: 1.74 versus 1.08; p for subgroup difference < 0.001; Table 3). Further subgroup analysis suggested that the association between myosteatosis and poor OS of HCC may not be significantly affected by main anticancer treatments, parameters for evaluating myosteatosis, whether the cutoffs for the diagnosis of myosteatosis were sex-specific or universal, or the median follow-up durations of the studies (p for subgroup difference all > 0.05; Table 3). In addition, a stronger association between myosteatosis and poor OS was observed in studies with adjustment of myosteatosis (HR: 1.89, 95% CI: 1.53 to 2.33, p < 0.001; Table 3). Finally, meta-regression analyses did not suggest any of the following variables may significantly affect the results, including sample size, mean ages of the patients, proportions of men,

follow-up durations, and study quality scores (p all >0.05; Table 4).

# Association between myosteatosis and PFS in patients with HCC

The pooled results of 16 datasets from 15 studies [20, 21, 23, 24, 26–29, 31, 32, 35, 36, 39–41], all with multivariate analyses, showed that myosteatosis also was associated with poorer PFS in patients with HCC (HR: 1.53, 95% CI: 1.33 to 1.76, p < 0.001; Fig. 2B) with moderate heterogeneity (p for Cochrane Q test = 0.07;  $I^2$  = 36%). Further sensitivity analysis by excluding one study at a time did not significantly change the results (HR: 1.45 to 1.58, p all < 0.05). Specifically, the sensitivity analysis including only studies with NOS score of 7 or higher also showed consistent results (HR: 1.54, 95% CI: 1.34 to 1.76, p < 0.001;  $I^2 = 36\%$ ). Subsequent subgroup analyses indicated that the association between myosteatosis and poor PFS in HCC patients was significant in Asian studies but not in European studies (HR: 1.57 vs. 1.05). However, the difference between subgroups was not statistically significant (p = 0.08; Table 3). Similar to the findings for OS, further subgroup analyses indicated that the association between myosteatosis and poor PFS was not significantly affected by the primary anticancer treatments, the parameters used to evaluate myosteatosis, whether the diagnostic cutoffs were sex-specific or universal, or the median follow-up durations of the studies (p for subgroup difference all > 0.05; Table 3). Moreover, the association between myosteatosis and poor PFS remains significant in subgroup of studies with adjustment of myosteatosis (HR: 1.50, 95% CI: 1.31 to 1.72, p < 0.001; Table 3). Lastly, metaregression analyses did not suggest any of the following variables could significantly modify the results, such as sample size, mean ages of the patients, proportions of men, followup durations, and study quality scores (p all > 0.05; Table 4).

#### **Publication bias**

The funnel plots for the meta-analyses assessing the association between myosteatosis and OS/PFS of patients with HCC are shown in Fig. 3A and B. Visual inspection of the plots reveals symmetry, indicating a low risk of publication bias. These findings are further supported by Egger's regression analyses (for OS: p = 0.42; for PFS: p = 0.35).

# **Discussion**

This meta-analysis provides comprehensive evidence that myosteatosis is significantly associated with worse survival outcomes in patients with HCC. By pooling data from 24 retrospective cohort studies with 7436 patients, we found



Table 2 Study quality evaluation via the Newcastle-Ottawa Scale

Study	Representative- ness of the exposed cohort	Selection of the non-exposed cohort	Ascertain- ment of exposure	Outcome not present at baseline	Control for age and sex	Control for other confounding factors	Assessment of outcome	Enough long follow-up dura- tion	Adequacy of follow-up of cohorts	Total
Kaibori [20]	0	1	1	1	1	1	1	1	1	∞
Fujiwara [19]	1	1	1	1	1	1	1	0	1	8
Mardian [22]	0	1	1	1	1	1		0	1	7
Hamaguchi [21]	1	1	1	1	1	1		1	1	6
Sano [37]	0	1	1	1	1	1		0	1	7
Jang [23]	0	1	1	1		1		1	1	8
Meister [39]	0	1	1	1	1	1		0	1	7
Yi [24]	0	1	1	1	1	0		0	1	9
Masetti [38]	1	1	1	0	0	1	1	0	1	9
Yamamoto [40]	0	1	1	1	1	1		0	1	7
Shi [27]	1	1	1	1	1	1	1	0	1	8
Yoshikawa [41]	0	1	1	1	1	1	1	0	1	7
Chen [26]	0	1	1	1	1	1	1	0	1	7
Bannangkoon [25]	0	1	1	1	1	1	1	1	1	8
Ishida [28]	1	1	1	1	1	1	1	1	1	6
Kang [29]	0	1	1	1	1	1	1	0	1	7
Liu [31]	0	1	1	1	1	1	1	0	1	7
Li [31]	0	1	1	1	1	1	1	0	1	7
Surov [34]	0	1	1	1	1	1	1	1	1	8
Lu [32]	0	1	1	1	1	1	1	1	1	8
Ouyang [33]	0	1	1	1	1	1	1	0	1	7
Surov [42]	1	1	1	1	1	1	1	1	1	∞
Luo [36]	0	1	1	1	1	1	1	0	1	7
Lee [35]	1	1	1	1	1	1	1	1	1	6



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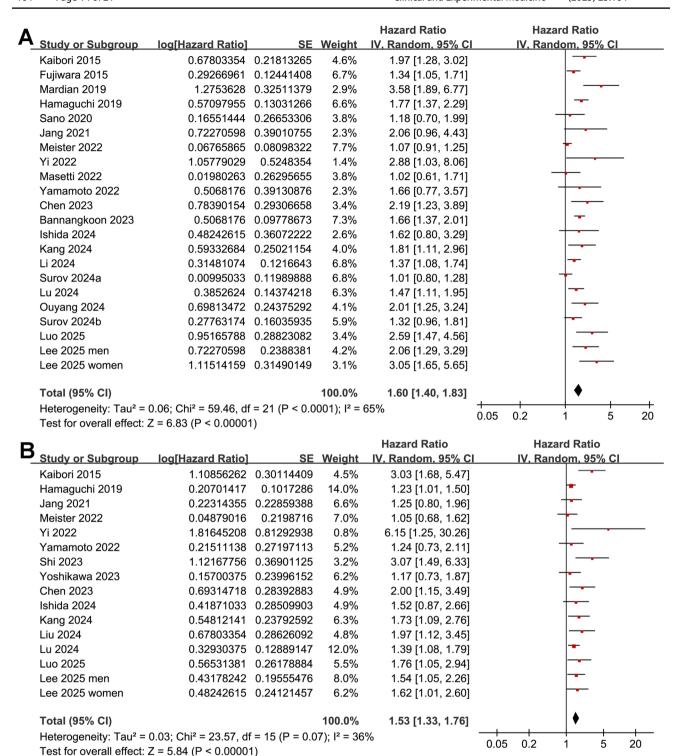


Fig. 2 Forest plots for the meta-analyses of the association between myosteatosis and survival outcomes of patients with HCC; A, forest plots for the meta-analysis of OS; and B, forest plots for the meta-analysis of PFS

that myosteatosis was independently linked to poorer OS and PFS. Subgroup analyses demonstrated that the association remained robust across different study populations and methodologies, with a stronger effect observed in Asian

studies compared to European studies. The results also remained significant regardless of anticancer treatment type, myosteatosis assessment methods, sex-specific or universal cutoffs, and follow-up duration. Notably, studies adjusting



Table 3 Results of subgroup analyses

	•									
	SO					PFS				
Subgroup	No. of studies	HR (95% CI)	$I^2$	p for subgroup effects	p for subgroup difference	No. of studies	HR (95% CI)	$I^2$	p for subgroup effects	p for subgroup difference
Location of study										
Asia	18	1.74 [1.55, 1.95]	30%	< 0.001		15	1.57 [1.36, 1.81]	34%	< 0.001	
Europe	4	1.08 [0.96, 1.22]	%0	0.18	< 0.001	1	1.05[0.68, 1.62]	ı	0.82	80.08
Anticancer treatments										
Surgical resection	7	1.78 [1.30, 2.43]	<i>%LL</i>	< 0.001		8	1.53 [1.22, 1.92]	54%	< 0.001	
Liver transplant	2	1.41 [1.18, 1.69]	%0	< 0.001		1	1.39[1.08, 1.79]	12%	0.01	
TACE or TAE	4	1.49 [0.97, 2.28]	85%	0.07		0	ı	ı	ı	
Immunotherapy	3	2.16 [1.53, 3.05]	%0	< 0.001		3	2.12 [1.44, 3.11]	%0	< 0.001	
TKIs	4	1.69 [1.25, 2.29]	33%	< 0.001	0.25	4	1.45 [1.13, 1.86]	%0	0.003	0.32
Parameter for myosteatosis										
IMAC	9	1.80 [1.42, 2.29]	38%	< 0.001		9	1.72 [1.29, 2.30]	62%	< 0.001	
SMD	16	1.53 [1.32, 1.78]	%99	< 0.001	0.26	10	1.45 [1.23, 1.70]	14%	< 0.001	0.30
Cutoffs for myosteatosis										
Sex-specific	111	1.66 [1.37, 2.00]	%02	< 0.001		9	1.64 [1.28, 2.10]	62%	< 0.001	
Not sex-specific	111	1.52 [1.25, 1.85]	53%	< 0.001	0.54	10	1.47 [1.23, 1.76]	14%	< 0.001	0.50
Mean follow-up duration										
< 30 months	~	1.84 [1.39, 2.43]	25%	< 0.001		9	1.85 [1.40, 2.45]	22%	< 0.001	
$\geq 30 \text{ months}$	14	1.52 [1.30, 1.77]	%89	< 0.001	0.23	10	1.41 [1.22, 1.64]	30%	< 0.001	0.09
Adjusted for sarcopenia										
Yes	14	1.89 [1.53, 2.33]	%69	< 0.001		15	1.50[1.31, 1.72]	31%	< 0.001	
No	∞	1.32 [1.13, 1.54]	49%	< 0.001	0.007	1	6.15 [1.25, 30.26]	ı	0.03	80.0

CI, confidence interval; HR, hazard ratio; I², heterogeneity index; IMAC, intramuscular adipose content; OS, overall survival; PFS, progression-free survival; SMD, skeletal muscle density; TACE, transarterial chemoembolization; TAE, transarterial embolization; TKIs, tyrosine kinase inhibitors



Table 4 Results of univariate meta-regression analysis

Variables	HR for OS			HR for PFS		
	Coefficient	95% CI	P values	Coefficient	95% CI	P values
Sample size	-0.00031	- 0.00076 to 0.00014	0.17	-0.00031	- 0.00096 to 0.00035	0.33
Mean age (years)	- 0.017	- 0.039 to 0.006	0.14	-0.018	- 0.039 to 0.004	0.10
Men (%)	-0.0040	- 0.0128 to 0.048	0.35	- 0.00016	- 0.00777 to 0.00745	0.97
Follow-up duration (months)	-0.0048	- 0.0150 to 0.0054	0.34	-0.0036	- 0.0147 to 0.0076	0.51
NOS	0.063	- 0.119 to 0.246	0.47	-0.048	- 0.237 to 0.140	0.59

HR, hazard ratio; OS, overall survival; PFS, progression-free survival; CI, confidence interval; NOS, Newcastle-Ottawa Scale;

for sarcopenia demonstrated an even stronger association between myosteatosis and poor survival, suggesting that muscle quality deterioration may play an important role in the prognosis of HCC beyond muscle mass loss alone. Further meta-regression analyses did not suggest that the results were significantly modified by sample sizes of the included studies, mean ages of the patients, proportions of men, follow-up durations, and study quality scores. These findings support the emerging recognition of myosteatosis as an adverse prognostic factor in HCC and highlight the need for further research to refine its clinical relevance.

To the best of our knowledge, two previous meta-analyses have evaluated the association between myosteatosis and prognosis of patients with HCC. An early meta-analysis included two studies published before 2024 showed that myosteatosis was not associated poor OS in patients with HCC who were treated with TACE [49]. A subsequent metaanalysis included six studies published before 2023 suggested that myosteatosis may be associated with an increased risk of mortality in patients with HCC [50]. Compared to these meta-analyses, our study has several methodological strengths. First, our study is the most updated, incorporating multiple newly published studies that were not included in these previous reviews, thereby increasing statistical power and generalizability. With a larger sample size and broader geographic representation, our meta-analysis provides a more comprehensive assessment of the prognostic role of myosteatosis in HCC. Moreover, our study included a wider range of treatment modalities, reflecting real-world clinical diversity, and performed multiple subgroup analyses to confirm the robustness of the findings. These methodological strengths enhance the reliability of our conclusions and contribute to a deeper understanding of the prognostic significance of myosteatosis in HCC.

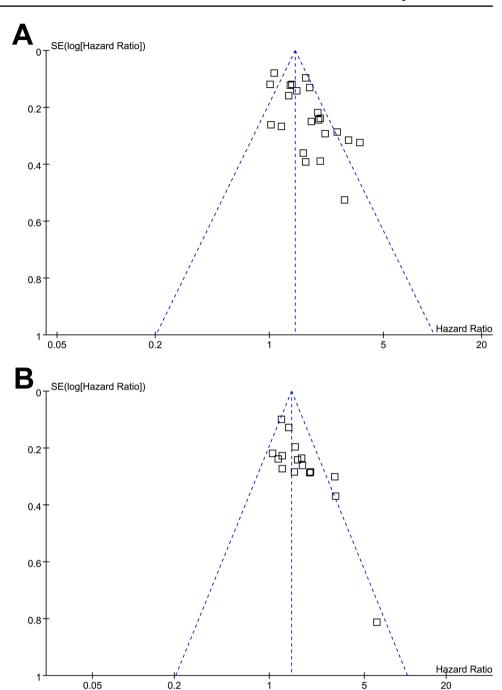
The underlying mechanisms linking myosteatosis to poor OS in HCC are multifactorial and involve both pathophysiological and clinical pathways. From a biological perspective, myosteatosis is associated with chronic systemic inflammation, oxidative stress, and mitochondrial dysfunction, all of which contribute to muscle deterioration and impaired metabolic homeostasis [51, 52]. Fat infiltration within skeletal

muscle has been shown to disrupt muscle fiber integrity, impair glucose utilization, and promote insulin resistance, creating a tumor-promoting microenvironment [11, 53]. Additionally, myosteatosis may lead to immune system dysregulation, reducing the host's ability to mount an effective antitumor response [54]. Clinically, myosteatosis is often linked to poor nutritional status, sarcopenia, and reduced functional reserve, which may compromise a patient's ability to tolerate aggressive cancer therapies [55]. Furthermore, myosteatosis may influence the efficacy [56] and toxicity of anticancer treatments [25]. Patients with myosteatosis may experience altered drug metabolism and increased toxicity due to impaired liver and muscle function [51], leading to suboptimal treatment responses and higher rates of treatment-related complications. These factors collectively contribute to worse survival outcomes in HCC patients with myosteatosis [18]. In addition, metabolic syndrome and MASLD—conditions increasingly recognized as key drivers of hepatocarcinogenesis—are frequently associated with chronic inflammation, hyperinsulinemia, and ectopic fat deposition, including within skeletal muscle [57]. This fat infiltration contributes to insulin resistance, mitochondrial dysfunction, and immune dysregulation, all of which can impair the host's antitumor response and promote cancer progression [10]. These mechanisms may be more pronounced in certain populations, particularly in regions with a higher prevalence of MASLD or obesity-related HCC, thereby contributing to geographic variation in the prognostic impact of myosteatosis [58].

Results of the subgroup analysis showed that the association between myosteatosis and poor survival outcomes of patients with HCC was significant in studies from Asian countries (n = 18), but not in those from European countries (n = 4). This geographic variation may be attributed to differences in HCC etiology, such as a higher prevalence of hepatitis B virus infection in Asia compared to hepatitis C virus or metabolic dysfunction-associated steatotic liver disease (MASLD) in Europe [59, 60]. Additionally, genetic background, body composition profiles, lifestyle factors, and access to or choice of anticancer treatments (e.g., use of curative vs. palliative therapies) may contribute to



Fig. 3 Funnel plots for evaluating the publication bias underlying the meta-analyses; A, funnel plots for the meta-analysis of the association between myosteatosis and OS of patients with HCC; and B, funnel plots for the meta-analysis of the association between myosteatosis and PFS of patients with HCC



differential prognostic implications of myosteatosis [61]. Nonetheless, the limited number of European studies may have reduced statistical power and contributed to the observed lack of significance. Further multicenter studies with balanced geographic representation are needed to validate these findings.

Despite its strengths, this study has several limitations. First, all included studies were retrospective, introducing the potential for selection bias and recall bias [62]. Second, the retrospective design of all included studies may introduce inherent biases, including selection bias and residual

confounding. Important variables such as detailed HCC etiology (e.g., HBV vs. HCV vs. MASLD), tumor stage at diagnosis, and comorbidity profiles were not uniformly reported or adjusted for across studies, which may have influenced the pooled estimates. Specifically, while tumor stage at diagnosis is a key prognostic factor and influences treatment decisions in HCC, stratified data by tumor stage were rarely reported across the included studies. As a result, we were unable to perform subgroup analyses based on this variable, and the lack of stage-specific adjustment may introduce residual confounding. Moreover, the majority



of included studies were conducted in Asian populations, potentially limiting the generalizability of our findings to other regions with different patient demographics and clinical practices. Furthermore, optimal methods and cutoff values for defining myosteatosis remain uncertain. While most studies used CT-based SMD or IMAC at the third lumbar vertebra, heterogeneity in measurement techniques and thresholds may affect comparability across studies. Third, given the study-level nature of this meta-analysis, we were unable to conduct an individual patient data analysis, which could provide more precise risk stratification and adjustment for confounding variables. Additionally, while HCC has diverse etiologies, including hepatitis B, hepatitis C, alcohol-related liver disease, and MASLD, we were unable to determine whether the impact of myosteatosis on survival varies by HCC etiology due to limited data. Furthermore, the presence of moderate heterogeneity suggests underlying differences in study populations, assessment methods, and treatment strategies that could not be fully accounted for in this analysis. Moreover, as with all observational studies, causality cannot be established, and prospective studies are needed to validate our findings. Finally, while no significant publication bias was detected, the possibility of small-study effects or unpublished negative studies cannot be fully excluded. Therefore, the pooled estimates should still be interpreted with some caution.

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The clinical implications of this study are important but should be interpreted with caution. While myosteatosis is a promising prognostic marker in HCC, it is premature to consider it a direct therapeutic target. Instead, these findings underscore the importance of assessing muscle quality in routine clinical practice and incorporating myosteatosis evaluation into risk stratification models for HCC. Future research should focus on refining the criteria for diagnosing myosteatosis, exploring its predictive value for treatment response, and investigating potential interventions to mitigate its impact on survival. Large-scale prospective studies with standardized imaging protocols and predefined cutoff values are needed to further elucidate the clinical utility of myosteatosis in HCC management. Additionally, integrating myosteatosis assessment with other prognostic factors, such as liver function parameters, immune status, and genetic biomarkers, may improve patient stratification and guide personalized treatment approaches. Emerging evidence also suggests that myosteatosis may reflect disease severity in MASLD and may be influenced by BMI and body fat percentage [63]. Given the growing burden of MASLD-associated HCC, future studies should explore the prognostic relevance of myosteatosis in this subgroup, while controlling for metabolic parameters, adiposity, and hepatic fibrosis severity. Such investigations may help clarify the interplay between metabolic dysregulation, muscle quality, and cancer outcomes. On the other hand, to enhance clinical

applicability, future efforts should focus on standardizing the assessment of myosteatosis in routine oncology and hepatology practice. This includes developing consensus guidelines on imaging protocols (e.g., CT acquisition parameters at the L3 vertebral level), establishing universally accepted cutoff values (potentially stratified by sex, BMI, or ethnicity), and validating reproducibility across software platforms and institutions. Integrating myosteatosis measurement into radiology reports or risk stratification tools could help guide treatment planning and follow-up strategies for patients with HCC.

### **Conclusions**

In conclusion, this meta-analysis provides up-to-date evidence that myosteatosis may be an independent predictor of poor survival in patients with HCC. The findings remain robust across different treatment modalities and assessment methods, and after adjustment of sarcopenia, with a particularly pronounced effect observed in Asian populations. While the mechanisms underlying this association are complex and multifactorial, myosteatosis likely contributes to poor outcomes through both metabolic and clinical pathways. Despite some limitations, this study represents the most comprehensive analysis to date, reinforcing the need for further research into the prognostic and clinical significance of myosteatosis in HCC. Future studies should focus on optimizing diagnostic criteria, validating findings in prospective cohorts, and investigating strategies to mitigate the adverse effects of myosteatosis on cancer outcomes.

**Supplementary Information** The online version contains supplementary material available at https://doi.org/10.1007/s10238-025-01671-6.

Author contributions Yongjuan Wu, Guangyuan Cheng, and Qingsong Yang designed the study. Yongjuan Wu and Guangyuan Cheng performed database search, study identification, study quality evaluation, and data extraction. Yongjuan Wu, Guangyuan Cheng, and Jun Han performed statistical analyses. Yongjuan Wu and Qingsong Yang interpreted the results. Yongjuan Wu and Guangyuan Cheng wrote the manuscript. All authors revised the manuscript and approved the submission.

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**Data availability** The authors confirm that the data supporting the findings of this study are available within the article. Further inquiries can be directed to the corresponding author.

#### **Declarations**

**Conflict of interest** The authors declare no competing interests.

**Ethics approval** Institutional Review Board approval was not required because this is a meta-analysis.



Consent to participate Not applicable.

Consent to publish Not applicable.

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