Meta-Analysis

Homeobox protein CDX2 as a prognostic biomarker in solid malignancies: a meta-analysis

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ABSTRACT

Background: CDX2 is a caudal-homeobox gene and its expression is abnormal in numerous tumour cell types. Nevertheless, its prognostic value for solid tumours requires further investigation. Hence, we conducted a meta-analysis to determine the significance of CDX2 as a prognostic biomarker in solid malignancies systematically.

Materials and Methods: We performed a systematic literature search in PUBMED and EMBASE up to May 2017. Retrospective studies comparing the prognostic value of different CDX2 levels in human malignancies were included. Data extractions and methodological assessments were performed separately by two investigators using a standard procedure. The statistical procedures were performed using Review Manager 5.3 and STATA/MP 14.0.

Results: A total of 26 retrospective studies met the inclusion criteria and comprised 5008 participants. Patients with CDX2 overexpression had significantly better 3-year, 5-year, 10-year and disease-free survival outcomes in solid malignancies, regardless of the cancer type, mean age, and source region. Nevertheless, there was no significant difference in the patients from Europe. The expression level of CDX2 was not statistically associated with cancer relapse. Moreover, our analysis showed that CDX2 overexpression is correlated to better responses to chemotherapy in patients with TNM IV stage cancers. The stability of the pooled outcomes was verified by sensitivity analysis. The funnel plots, Egger's test and Begg's test jointly confirmed that there was no publication bias.

Conclusions: Overexpression of CDX2 is a reliable biomarker of a better prognosis in solid malignancies.

INTRODUCTION

Over the past decade, steady improvements in the overall survival (OS) and disease-free survival (DFS) outcomes have been observed in cancer patients, depending on the availability of new adjuvant treatment regimens. However, no reliable biomarker exists that enables stratification of patients with high-risk diseases for appropriate therapy. Patients with early stage tumours are often cured with surgery alone, but many patients will relapse and eventually succumb to their diseases [1]. Identifying which molecules can predict the cancer prognosis and adjuvant therapeutic benefit remains a research priority. *CDX2* is a caudal-homeobox gene which is mostly expressed in intestinal epithelial cells. CDX2 expression is essential to the intestinal epithelial proliferation and the differentiation and maintenance of the intestinal phenotypes [2, 3]. Considerable studies have confirmed that CDX2 is expressed not only in normal intestinal epithelial cells but also in different cancer cell types and functions as a tumour suppressor [4, 5]. Moreover, various clinical studies suggest that CDX2 expression is often lost in cancers with high tumour grade and advanced tumour stage [6, 7].

A previous study evaluated the expression levels of CDX2 and its association with the clinicopathological characteristics of gastric cancers [6], but the independent prognostic value of CDX2 remains controversial. Braak et al reported that there is no association between CDX2 expression and the prognosis of colon cancer patients [8], whereas Li et al suggested that CDX2 expression would be useful for predicting the prognosis of gallbladder carcinomas [9]. In our study, we aimed to perform a more comprehensive quantitative assessment with a meta-analysis and to determine the value of CDX2 as a prognostic and predictive tumour biomarker.

RESULTS

Study and patient characteristics

Among 1313 studies that were identified in our literature search, 26 retrospective studies (consisting of 28 cohorts) met the inclusion criteria (a flow diagram is provided in Figure 1). The total sample size comprised 5008 participants, individually ranging from 40 to 713 per study with a median value of 108 patients. Most of the 26 studies were carried out in Asia (n = 14) and Europe (n = 7), and the remaining five were performed in the USA; these studies mainly focused on colorectal cancers (n = 9) and gastric cancers (n = 8). Other detailed features are included in Table 1 [10–35].

Methodological assessment

The Newcastle-Ottawa Quality Assessment Scale (NOS) quality score ranged from 5 to 9, and the assessment included two 5-score studies, eight 6-score studies, ten 7-score studies, five 8-score studies and one 9-score study. The 5-score studies included Jamieson 2013 and Wong 2011, the low score was mainly due to their unrepresentative study cohorts (Supplementary Table 1).

Correlations between CDX2 levels and 3-year overall survival

23 retrospective studies offered original data on 3-year OS in terms of different CDX2 expressions. Compared with lower CDX2 expression, patients with CDX2 positivity had significantly better 3-year OS outcomes, along with a moderate heterogeneity of undefined source (OR: 0.29, 95%CI: [0.22,0.37], P < 0.00001, $l^2 = 49\%$) (Figure 2).

Subgroup analysis by cancer types showed that CDX2 overexpression was a prognostic biomarker of a favourable 3-year OS for participants with gastric cancer (OR: 0.34, 95%CI: [0.20,0.58], P < 0.0001, $I^2 = 67\%$), colorectal cancer (OR: 0.32, 95%CI: [0.20,0.49], P < 0.00001, $I^2 = 55\%$) and other cancer types (OR: 0.22, 95%CI: [0.16,0.31], P < 0.00001, $I^2 = 0\%$) (Figure 3).

According to the subgroup analysis of different mean age ranges, regardless of whether the mean age was < 60 (OR: 0.26, 95%CI: [0.16,0.41], P < 0.00001,

 $I^2 = 18\%$), between 60 and < 65 (OR: 0.26, 95%CI: [0.20,0.34], P < 0.00001, $I^2 = 8\%$), > 65 (OR: 0.18, 95%CI: [0.08,0.39], P < 0.0001, $I^2 = 0\%$) or unknown (OR: 0.41, 95%CI: [0.23,0.73], P = 0.002, $I^2 = 73\%$), high CDX2 expression was invariably associated with a better 3-year OS for patients with solid malignancies (Supplementary Figure 1).

All included studies were then divided into three subgroups according to the source regions of the cancer patients. For patients from Asia (OR: 0.26, 95%CI: [0.21,0.33], P < 0.00001, $I^2 = 0\%$) and other regions (OR: 0.22, 95%CI: [0.12,0.41], P < 0.00001, $I^2 = 52\%$), excessive CDX2 expression was linked to increased 3-year OS. Nevertheless, there was no significant difference in patients from Europe (OR: 0.50, 95%CI: [0.22,1.13], P = 0.10, $I^2 = 78\%$).

In the subgroup analysis by sample sizes, elevated CDX2 expression was confirmed to play a favourable prognostic role in terms of 3-year OS in solid malignancies, no matter whether the number of participants was smaller (< 100) (OR: 0.38, 95%CI: [0.22,0.67], P = 0.0008, $I^2 = 60\%$) or larger (> 100) (OR: 0.26, 95%CI: [0.21,0.32], P < 0.00001, $I^2 = 23\%$).

Correlations between CDX2 levels and 5-year overall survival

The original data of 5-year OS in terms of different CDX2 expressions was extracted from 24 retrospective studies. Our pooled results indicated that CDX2 overexpression played a favourable role on the 5-year OS rate in solid malignancies (OR: 0.31, 95%CI: [0.23,0.41], P < 0.00001) (Figure 4). A high heterogeneity was observed ($I^2 = 62\%$). We stratified the original data for subgroup analysis to further explore the potential sources of heterogeneity across studies.

With respect to subgroups by different cancer types, a high CDX2 level reflected a favourable 5-year OS for patients with gastric cancer (OR: 0.32, 95%CI: [0.18,0.55], P < 0.0001, $I^2 = 71\%$), colorectal cancer (OR: 0.37, 95%CI: [0.25,0.56], P < 0.0001, $I^2 = 52\%$) and other cancer types (OR: 0.20, 95%CI: [0.11,0.36], P < 0.00001, $I^2 = 53\%$) (Figure 5).

High CDX2 level in different mean-age subgroups also suggested a better 5-year OS for cancer patients with a mean age that either was < 60 (OR: 0.32, 95%CI: [0.18,0.56], P < 0.0001, $I^2 = 41\%$), between 60 and < 65 (OR: 0.26, 95%CI: [0.18,0.37], P < 0.00001, $I^2 = 42\%$), > 65 (OR: 0.09, 95%CI: [0.03,0.26], P < 0.0001, $I^2 = 10\%$) or unknown (OR: 0.47, 95%CI: [0.28,0.78], P = 0.003, $I^2 = 69\%$) (Supplementary Figure 1).

As for the source regions of solid malignancies, excessive CDX2 positivity was significantly correlated with a better 5-year OS in patients from Asia (OR: 0.27, 95%CI: [0.21, 0.34], P < 0.00001, $I^2 = 12\%$) and other regions (OR: 0.21, 95%CI: [0.08, 0.53], P = 0.001, $I^2 = 73\%$). However,

Reference	Country	Cancer type	No.	Mean age(y)	Male/Female (The percentage of females)	TNM stage	Median follow- up(m)	CDX2 (-/+)	NOS score
Baba et al. 2009	USA	Colorectal cancer	598	NA	NA	I-IV	NA	172/426	9
Bae et al. 2015	Korea	Colorectal cancer	713	62	434/279 (39.13)	I-IV	56.5	42/671	8
Bai et al. 2013	China	Gastric cancer	228	60.72 ± 12.95	170/58 (25.44)	I-IV	NA	129/99	7
Camilo et al. 2014	Portugal	Gastric cancer	94	64.6	124/77 (38.31)	I-IV	NA	55/39	6
Dalerba et al. 2016	USA	Colon cancer-C1	314	NA	NA	II/III	NA	38/276	8
Dalerba et al. 2016	USA	Colon cancer-C2	466	NA	NA	II/III	NA	32/434	8
Fan et al. 2005	China	Gastric cancer	109	59	75/34 (31.19)	I-IV	NA	69/40	7
Hansel et al. 2005	USA	Periampullary cancer	53	68	30/23 (44.40)	I-IV	NA	39/14	6
Hong et al. 2013	Korea	Colorectal cancer	207	62.3±11.4	119/88 (42.51)	I-IV	65.8	11/196	7
Huang et al. 2012	China	Ovarian cancer	182	NA	NA	I-IV	36.2	129/53	7
Jamieson et al. 2013	USA	Pancreatic ductal cancer	60	NA	NA	NA	NA	47/13	5
Jun et al. 2014	Korea	Small intestinal cancer	189	59±14	118/71 (37.57)	I-IV	28.1	107/82	7
Kim et al. 2013	Korea	Colorectal cancer	109	NA	66/43 (39.45)	I-IV	NA	15/94	6
Kumari et al. 2013	India	Periampullary cancer	108	57.2	74/34 (31.48)	I-III	NA	65/43	6
Lundberg et al. 2016	Sweden	Colorectal cancer	431	NA	231/200 (46.40)	I-IV	NA	62/369	8
Matsuda et al. 2010	Japan	Colorectal cancer	97	68	NA	I-IV	NA	9/88	6
Mizoshita et al. 2003	Japan	Gastric cancer	177	63.2±11.1	105/72 (40.68)	I-IV	NA	88/89	7
Perysinakis et al. 2016	Greece	Ampullary cancer	47	66.3±12.5	27/20 (42.55)	I-IV	52	19/28	6
Pilati et al. 2017	France	Colon cancer-C1	79	NA	NA	II/III	NA	39/40	8
Pilati et al. 2017	France	Colon cancer-C2	99	NA	NA	II/III	NA	23/76	8
Qin et al. 2012	China	Gastric cancer	85	61.75	60/25 (29.41)	I-IV	NA	44/41	7
Schildberg et al. 2014	Germany	Gastric cancer	79	NA	NA	NA	88	52/27	7
Seno et al. 2002	Japan	Gastric cancer	40	61	26/14 (35.00)	I-IV	NA	22/18	6
Treese et al. 2016	Germany	Gastro- Oesophageal cancer	65	62.9	81/48 (37.21)	I-IV	NA	28/37	6

Table 1: Demographic information for included studies with CDX2 expression

Wong et al. 2011	China	Colorectal cancer	64	NA	NA	I-IV	NA	32/32	5
Xiao et al. 2014	USA	Pancreatic ductal cancer	61	64	32/29 (47.54)	NA	NA	39/22	7
Zhang et al. 2009	Japan	Gastric cancer	109	62.43 ± 10.12	63/46 (42.20)	I-IV	NA	52/57	8
Zhang et al. 2016	USA	Metastatic colorectal cancer	145	63.6	69/76 (52.41)	IV	NA	66/79	7

No.: number; y: year; m: month; NOS: Newcastle-Ottawa

Scale; NA: not available; C1: cohort 1; C2: cohort 2.

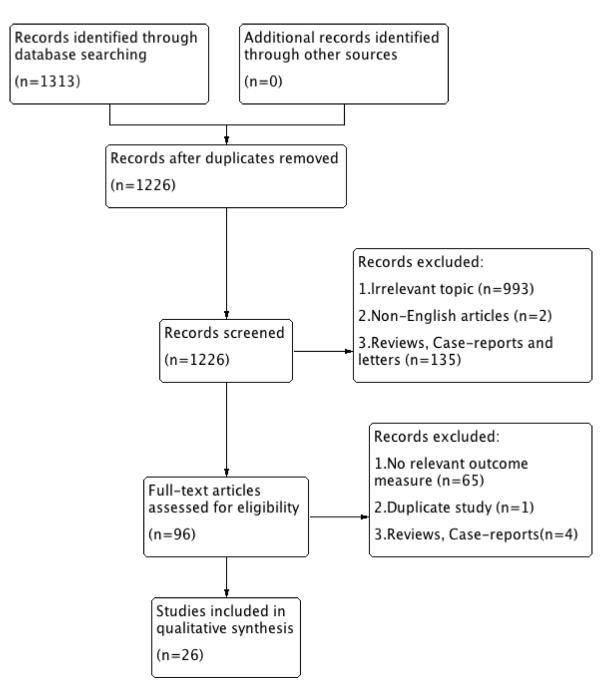


Figure 1: Selection flow diagram of this meta-analysis.

no correlation between CDX2 expression and 5-year OS was evident for European patients (OR: 0.54, 95%CI: [0.26,1.12], P = 0.1, P = 74%) (Supplementary Figure 4).

For the subgroup determined by sample size of the studies, the pooled analysis showed that CDX2 positivity had a strong connection with a better 5-year OS in solid malignancies, regardless of whether the sample size was smaller (< 100) (OR: 0.30, 95%CI: [0.15,0.60], P = 0.0007, $I^2 = 73\%$) or larger (> 100) (OR: 0.30, 95%CI: [0.23,0.38], P < 0.00001, $I^2 = 40\%$) (Supplementary Figure 2).

Correlations between CDX2 levels and 10-year overall survival

Concerning the 10-year OS in solid malignancies, only 7 studies offered original data in terms of different CDX2 expressions. In our pooling analysis, a high CDX2 positivity in solid malignancies corresponded to a significantly better 10-year OS rate (OR: 0.36, 95%CI: [0.20,0.65], P = 0.0007) (Figure 6). A high heterogeneity was observed across studies ($I^2 = 77\%$).

Correlations between CDX2 levels and Diseasefree survival

The merged outcome showed that a better prognosis was observed concerning 3-year (OR: 0.50, 95%CI: [0.25,0.98], P = 0.04, $I^2 = 65\%$) and 5-year DFS (OR:

0.45, 95%CI: [0.25,0.81], P = 0.008, $I^2 = 56\%$) among the malignancies with CDX2 overexpression, along with a high heterogeneity (Supplementary Figure 5).

Correlations between CDX2 levels and cancer relapse

A comparable number of patients with relapsed disease was observed regardless of whether the CDX2 levels (OR: 0.93, 95%CI: [0.39,2.25], P = 0.87, $I^2 = 70\%$) (Figure 7).

Stratified by cancer types of the included studies, the CDX2 expression levels in colorectal cancer (OR: 0.80, 95%CI: [0.20,3.25], P = 0.75, $I^2 = 79\%$) and ovarian cancer (OR: 1.23, 95%CI: [0.65,2.34], P = 0.52) were irrelevant to disease relapse (Figure 8).

According to the subgroup analysis of different TNM stages, the numbers of relapsed patients with TNM II/III (OR: 1.36, 95%CI: [0.28,6.68], P = 0.70, $I^2 = 80\%$) and TNM I-IV (OR: 0.60, 95%CI: [0.12,3.08], P = 0.54, $I^2 = 78\%$) stage diseases were statistically equivalent regardless of the CDX2 expression levels (Figure 8).

Correlations between CDX2 levels and chemotherapeutic effects

The original data on the correlations between CDX2 levels and the effects of adjuvant chemotherapy

	CDX2	2-	CDX2	2+		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% Cl
Baba 2010	119	172	359	426	7.5%	0.42 [0.28, 0.64]	
Bae 2015	22	42	537	671	5.9%	0.27 [0.15, 0.52]	
Bai 2013	46	129	71	99	6.4%	0.22 [0.12, 0.39]	
Camilo 2014	18	55	23	39	4.5%	0.34 [0.14, 0.79]	
Fan 2017	21	69	30	40	4.4%	0.15 [0.06, 0.35]	
Hansel 2005	14	39	11	14	2.3%	0.15 [0.04, 0.64]	
Hong 2013	5	11	175	196	2.8%	0.10 [0.03, 0.36]	
Huang 2012	82	129	47	53	4.2%	0.22 [0.09, 0.56]	
Jamieson 2013	8	47	8	13	2.5%	0.13 [0.03, 0.50]	
Jun 2014	40	107	56	82	6.1%	0.28 [0.15, 0.51]	
Kim 2013	8	15	80	94	3.1%	0.20 [0.06, 0.64]	
Kumari 2013	22	65	25	43	4.9%	0.37 [0.17, 0.82]	
Lundberg 2016	26	62	278	369	6.5%	0.24 [0.14, 0.41]	
Matsuda 2010	5	9	74	88	2.3%	0.24 [0.06, 0.99]	
Mizoshita 2003	46	88	64	89	6.0%	0.43 [0.23, 0.80]	
Perysinakis 2016	6	19	21	28	2.7%	0.15 [0.04, 0.56]	
Pilati 2017–C1	35	39	32	40	2.7%	2.19 [0.60, 7.97]	
Pilati 2017–C2	12	23	56	76	4.0%	0.39 [0.15, 1.02]	
Qin 2012	26	44	31	41	4.1%	0.47 [0.18, 1.18]	
Schildberg 2014	35	52	14	27	4.0%	1.91 [0.74, 4.95]	
Seno 2002	15	22	16	18	1.7%	0.27 [0.05, 1.50]	
Xiao 2014	4	39	7	22	2.5%	0.24 [0.06, 0.96]	
Zhang 2009	27	52	49	57	4.1%	0.18 [0.07, 0.44]	
Zhang 2016	10	66	44	79	4.8%	0.14 [0.06, 0.32]	
Total (95% CI)		1395		2704	100.0%	0.29 [0.22, 0.37]	◆
Total events	652		2108				
Heterogeneity: Tau ² =	= 0.17; Cł	$ni^2 = 4!$	5.38, df =	= 23 (P	= 0.004)	; $I^2 = 49\%$	0.02 0.1 1 10 50
Test for overall effect	:: Z = 9.72	2 (P < 0).00001)				Favours [CDX+] Favours [CDX-]

Figure 2: Forest plot of the association between CDX2 expression and 3-year overall survival in solid malignancies.

Study or Subarous	CDX2		CDX2		Woight	Odds Ratio M-H, Random, 95% CI	Odds Ratio M–H, Random, 95% Cl
Study or Subgroup 6.3.1 Gastric cancer			Events	TOLAI	weight	M-H, Kanuolii, 95% Cl	M-H, Kandolli, 95% Cl
	-		71	00	C 40/	0 22 [0 12 0 20]	
Bai 2013	46	129 55	71 23	99	6.4%	0.22 [0.12, 0.39]	
Camilo 2014	18 21		23	39	4.5%	0.34 [0.14, 0.79]	
Fan 2017		69		40	4.4%	0.15 [0.06, 0.35]	
Mizoshita 2003	46	88	64	89	6.0%	0.43 [0.23, 0.80]	
Qin 2012	26	44	31	41	4.1%	0.47 [0.18, 1.18]	-
Schildberg 2014	35	52	14	27	4.0%	1.91 [0.74, 4.95]	
Seno 2002	15	22	16	18	1.7%	0.27 [0.05, 1.50]	
Zhang 2009 Subtotal (95% CI)	27	52 511	49	57 410	4.1% 35.3%	0.18 [0.07, 0.44] 0.34 [0.20, 0.58]	
	234	511	298	410	33.3/0	0.34 [0.20, 0.38]	
Total events Heterogeneity: Tau ²		2 - 2 ·		7 (D	0.003)	2 670/	
Test for overall effect				= 7 (P =	0.005), 1	= 07%	
rest for overall effec	L = 5.96	5 (F < C	.0001)				
6.3.2 Colorectal can	cer of 3-y	/ear O	5				
Baba 2010	119	172	359	426	7.5%	0.42 [0.28, 0.64]	
Bae 2015	22	42	537	671	5.9%	0.27 [0.15, 0.52]	
Hong 2013	5	11	175	196	2.8%	0.10 [0.03, 0.36]	
Kim 2013	8	15	80	94	3.1%	0.20 [0.06, 0.64]	
Lundberg 2016	26	62	278	369	6.5%	0.24 [0.14, 0.41]	
Matsuda 2010	5	9	74	88	2.3%	0.24 [0.06, 0.99]	
Pilati 2017-C1	35	39	32	40	2.7%	2.19 [0.60, 7.97]	
Pilati 2017-C2	12	23	56	76	4.0%	0.39 [0.15, 1.02]	
Subtotal (95% CI)		373		1960	34.8%	0.32 [0.20, 0.49]	◆
Total events	232		1591				-
Heterogeneity: Tau ²	= 0.20; Ch	$ni^2 = 15$	5.73, df =	= 7 (P =	0.03); I ²	= 55%	
Test for overall effect	t: Z = 5.09) (P < C	.00001)				
6.3.3 Other types of							
Hansel 2005		39	11	14	2.3%	0.15 [0.04, 0.64]	
	14					• • •	
Huang 2012	82	129	47	53	4.2%	0.22 [0.09, 0.56]	
Jamieson 2013	82 8	129 47	47 8	13	2.5%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50]	
Jamieson 2013 Jun 2014	82 8 40	129 47 107	47 8 56	13 82	2.5% 6.1%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51]	
Jamieson 2013 Jun 2014 Kumari 2013	82 8 40 22	129 47 107 65	47 8 56 25	13 82 43	2.5% 6.1% 4.9%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016	82 8 40 22 6	129 47 107 65 19	47 8 56 25 21	13 82 43 28	2.5% 6.1% 4.9% 2.7%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014	82 8 40 22 6 4	129 47 107 65 19 39	47 8 25 21 7	13 82 43 28 22	2.5% 6.1% 4.9% 2.7% 2.5%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016	82 8 40 22 6	129 47 107 65 19 39 66	47 8 56 25 21	13 82 43 28 22 79	2.5% 6.1% 4.9% 2.7% 2.5% 4.8%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI)	82 8 40 22 6 4 10	129 47 107 65 19 39	47 8 56 25 21 7 44	13 82 43 28 22	2.5% 6.1% 4.9% 2.7% 2.5%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI) Total events	82 8 40 22 6 4 10 186	129 47 107 65 19 39 66 511	47 8 56 25 21 7 44 219	13 82 43 28 22 79 334	2.5% 6.1% 4.9% 2.7% 2.5% 4.8% 29.9%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32] 0.22 [0.16, 0.31]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI) Total events Heterogeneity: Tau ²	82 8 40 22 6 4 10 186 = 0.00; Ch	129 47 107 65 19 39 66 511 hi2 = 4.	47 8 56 25 21 7 44 219 48, df =	13 82 43 28 22 79 334	2.5% 6.1% 4.9% 2.7% 2.5% 4.8% 29.9%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32] 0.22 [0.16, 0.31]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI) Total events Heterogeneity: Tau ² Test for overall effect	82 8 40 22 6 4 10 186 = 0.00; Ch	$129 47 107 65 19 39 66 511 m2 = 4. \Theta (P < C$	47 8 56 25 21 7 44 219 48, df =	13 82 43 28 22 79 334 7 (P =	2.5% 6.1% 4.9% 2.7% 2.5% 4.8% 29.9% 0.72); I ² =	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32] 0.22 [0.16, 0.31]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI) Total events Heterogeneity: Tau ² Test for overall effect Total (95% CI)	82 8 40 22 6 4 10 186 = 0.00; Ch t: Z = 8.89	129 47 107 65 19 39 66 511 hi2 = 4.	47 8 56 25 21 7 44 219 48, df =	13 82 43 28 22 79 334 7 (P =	2.5% 6.1% 4.9% 2.7% 2.5% 4.8% 29.9%	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32] 0.22 [0.16, 0.31]	
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI) Total events Heterogeneity: Tau ² Test for overall effect Total (95% CI) Total events	82 8 40 22 6 4 10 186 = 0.00; Ch t: Z = 8.89 652	129 47 107 65 19 39 66 511 m2 = 4. 9 (P < C 1395	47 8 56 25 21 7 44 219 48, df = 0.00001) 2108	13 82 43 28 22 79 334 7 (P = 2704	2.5% 6.1% 4.9% 2.7% 2.5% 4.8% 29.9% 0.72); I ² =	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32] 0.22 [0.16, 0.31] 0% 0.29 [0.22, 0.37]	• •
Jamieson 2013 Jun 2014 Kumari 2013 Perysinakis 2016 Xiao 2014 Zhang 2016 Subtotal (95% CI) Total events Heterogeneity: Tau ² Test for overall effect Total (95% CI)	82840226410186= 0.00; Cht: Z = 8.89652= 0.17; Ch	129 47 107 65 19 39 66 511 ni2 = 4. 0 (P < C 1395 ni2 = 4. 9 (P < C 1395 ni2 = 4. 9 (P < C) 19 19 19 19 19 19 19 19	47 8 56 25 21 7 44 219 48, df = 000001) 2108 5.38, df =	13 82 43 28 22 79 334 7 (P = 2704	2.5% 6.1% 4.9% 2.7% 2.5% 4.8% 29.9% 0.72); I ² =	0.22 [0.09, 0.56] 0.13 [0.03, 0.50] 0.28 [0.15, 0.51] 0.37 [0.17, 0.82] 0.15 [0.04, 0.56] 0.24 [0.06, 0.96] 0.14 [0.06, 0.32] 0.22 [0.16, 0.31] 0% 0.29 [0.22, 0.37]	

Figure 3: The correlation between CDX2 expression and 3-year overall survival based on different cancer types.

	CDX2	2-	CDX2	2+		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% Cl
Baba 2010	108	172	324	426	6.6%	0.53 [0.36, 0.78]	-
Bae 2015	19	42	497	671	5.5%	0.29 [0.15, 0.54]	
Bai 2013	42	129	64	99	5.9%	0.26 [0.15, 0.46]	- - -
Camilo 2014	14	55	19	39	4.4%	0.36 [0.15, 0.86]	
Fan 2017	16	69	23	40	4.6%	0.22 [0.10, 0.52]	
Hansel 2005	11	39	11	14	2.6%	0.11 [0.03, 0.46]	
Hong 2013	5	11	164	196	3.1%	0.16 [0.05, 0.57]	
Huang 2012	65	129	44	53	4.8%	0.21 [0.09, 0.46]	_ —
lamieson 2013	4	47	4	13	2.3%	0.21 [0.04, 1.00]	
lun 2014	28	107	47	82	5.6%	0.26 [0.14, 0.49]	
Kim 2013	8	15	77	94	3.4%	0.25 [0.08, 0.79]	
Kumari 2013	22	65	20	43	4.8%	0.59 [0.27, 1.30]	
Lundberg 2016	25	62	244	369	5.9%	0.35 [0.20, 0.60]	
Matsuda 2010	2	9	60	88	2.2%	0.13 [0.03, 0.68]	
Mizoshita 2003	31	88	55	89	5.6%	0.34 [0.18, 0.62]	
Perysinakis 2016	0	19	19	28	0.9%	0.01 [0.00, 0.23]	·
Pilati 2017–C1	18	39	16	40	4.3%	1.29 [0.53, 3.14]	_
Pilati 2017–C2	6	23	40	76	3.8%	0.32 [0.11, 0.89]	
Qin 2012	21	44	30	41	4.3%	0.33 [0.13, 0.83]	
Schildberg 2014	34	52	13	27	4.1%	2.03 [0.79, 5.24]	—
Seno 2002	3	22	16	18	1.7%	0.02 [0.00, 0.13]	
Treese 2016	23	28	31	37	2.9%	0.89 [0.24, 3.28]	
Xiao 2014	3	39	5	22	2.4%	0.28 [0.06, 1.33]	
Zhang 2009	19	52	39	57	4.8%	0.27 [0.12, 0.59]	
Zhang 2016	4	66	36	79	3.6%	0.08 [0.03, 0.23]	
Total (95% CI)		1423		2741	100.0%	0.31 [0.23, 0.41]	•
Total events	531		1898				
Heterogeneity: Tau ² =	= 0.28; Cl	ni² = 62	2.35, df =	= 24 (P	< 0.0001	1); $I^2 = 62\%$	0.001 0.1 1 10 100
Test for overall effect							0.001 0.1 1 10 100 Favours [CDX2+] Favours [CDX2-]

Figure 4: Forest plot of the association between CDX2 expression and 5-year overall survival in solid malignancies.

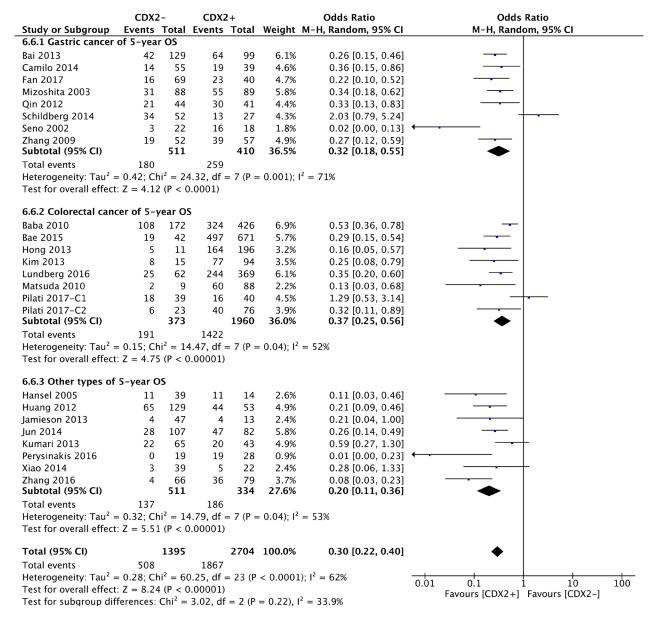


Figure 5: The correlation between CDX2 expression and 5-year overall survival based on different cancer types.

	CDX2-	-	CDX2	!+		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Baba 2010	90	172	268	426	19.9%	0.65 [0.45, 0.93]	-
Camilo 2014	8	55	16	39	13.4%	0.24 [0.09, 0.65]	_ - -
Huang 2012	35	129	40	53	16.0%	0.12 [0.06, 0.25]	
Jun 2014	22	107	32	82	17.0%	0.40 [0.21, 0.77]	
Lundberg 2016	25	62	220	369	18.1%	0.46 [0.26, 0.79]	
Perysinakis 2016	0	19	17	28	3.4%	0.02 [0.00, 0.31]	·
Treese 2016	21	28	26	37	12.2%	1.27 [0.42, 3.85]	
Total (95% CI)		572		1034	100.0%	0.36 [0.20, 0.65]	◆
Total events	201		619				
Heterogeneity: Tau ² =	= 0.42; Chi	i ² = 26	5.22, df =	= 6 (P =	= 0.0002)	; $I^2 = 77\%$	0.001 0.1 1 10 1000
Test for overall effect	: Z = 3.40	(P = 0	0.0007)				Favours [CDX2+] Favours [CDX2-]

Figure 6: Forest plot of the association between CDX2 expression and 10-year overall survival in solid malignancies.

with first-line drugs were extracted. The pooled analysis of retrospective studies suggested that there was no significant difference in the chemotherapeutic effects between patients with low CDX2 expression and higher CDX2 expression in solid malignancies (OR: 0.92, 95%CI: [0.45, 1.90], P = 0.83, $I^2 = 82\%$) (Figure 7).

Included studies were divided into two subgroups according to the cancer types, CDX2 expression was not associated with the chemotherapeutic effects in solid malignancies, no matter whether the cancer was colorectal (OR: 0.78, 95%CI: [0.33,1.82], P < 0.56, $I^2 = 83\%$) or ovarian (OR: 1.76, 95%CI: [0.91, 3.44], P < 0.09) (Figure 9).

As for different TNM stages of solid malignancies, in participants with TNM IV (OR: 0.20, 95%CI: [0.09,0.43], P < 0.0001) stage diseases, higher CDX2 levels were linked to a better chemotherapeutic effects. Nevertheless, regarding patients with I-IV (OR: 1.76, 95%CI: [0.91,3.44], P = 0.09) and II/III (OR: 1.10, 95%CI: [0.71,1.69], P = 0.68, $I^2 = 25\%$) stage diseases, an equivalent chemotherapeutic effect was obtained between lower CDX2 expression and higher CDX2 expression in solid malignancies (Figure 9).

Sensitivity analysis

First, we switched the statistical model from the random-effects to the fixed-effects model, and the OS and DFS remained unchanged. A similar result was also observed for studied concerned with cancer relapses and chemotherapeutic effects (Supplementary Table 2).

Second, when included trials were randomly removed, the outcome stability for OS, DFS, cancer relapse and chemotherapeutic effect was graphically confirmed (Supplementary Figures 3, 6–9).

Third, we excluded the low-quality trials conducted by Jamieson 2013 and Wong 2011, and the results of CDX2 expressions that were associated with OS and cancer relapse remained stable (Supplementary Figures 10-12).

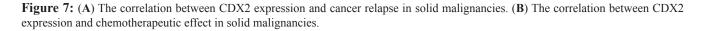
Publication bias

The funnel plots, Egger's test and Begg's test jointly confirmed that there was no publication bias concerning the pooled results except for the results for 5-year OS (*P* values were 0.042 and 0.038). Thus, we performed a sensitivity analysis using the trim-and-fill method, which suggested that there was a low risk of publication bias (*P* = 0.000) (Supplementary Figure 13–23).

DISCUSSION

Individual studies have partially revealed the favourable prognostic roles of CDX2 expression in gastric cancers [6]. Nevertheless, from the clinical perspective, the prognostic significance of CDX2 remains unconvincing because the experimental cohorts and participants were too scarce. Whether CDX2 expression is associated with tumour prognosis is worthy of further pursuit. To our knowledge, this meta-analysis is the first comprehensive exploration of the possible prognostic roles of CDX2 in solid malignancies.

Α	CDX2	!-	CDX2	!+		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% CI
Huang 2012	70	129	26	53	30.9%	1.23 [0.65, 2.34]	
Pilati 2017–C1	22	39	27	40	26.4%	0.62 [0.25, 1.56]	
Pilati 2017–C2	18	23	40	75	23.6%	3.15 [1.06, 9.37]	
Wong 2011	3	32	10	32	19.1%	0.23 [0.06, 0.93]	
Total (95% CI)		223		200	100.0%	0.93 [0.39, 2.25]	-
Total events	113		103				
Heterogeneity: Tau ² =	• 0.55; Cł	$11^2 = 9.$	94, df =	3 (P =	0.02); I ² :	= 70%	0.01 0.1 1 10 100
Test for overall effect:	Z = 0.16	5 (P = 0)).87)				Favours [CDX2+] Favours [CDX2-]
B			65 1/2				
	CDX2		CDX2			Odds Ratio	Odds Ratio
Study or Subgroup			Events		Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
					Weight 14.6%		
Study or Subgroup	Events	Total	Events	Total	-	M-H, Random, 95% Cl	
Study or Subgroup Dalerba 2016-C1	Events 20	Total 23	Events 290	Total 389	14.6%	M-H, Random, 95% Cl 2.28 [0.66, 7.82]	
Study or Subgroup Dalerba 2016-C1 Dalerba 2016-C2	Events 20 37	Total 23 60	Events 290 570	Total 389 1003	14.6% 22.3%	M-H, Random, 95% Cl 2.28 [0.66, 7.82] 1.22 [0.72, 2.09]	
Study or Subgroup Dalerba 2016-C1 Dalerba 2016-C2 Huang 2012	Events 20 37 92	Total 23 60 129	Events 290 570 31	Total 389 1003 53	14.6% 22.3% 20.9%	M-H, Random, 95% Cl 2.28 [0.66, 7.82] 1.22 [0.72, 2.09] 1.76 [0.91, 3.44]	
Study or Subgroup Dalerba 2016-C1 Dalerba 2016-C2 Huang 2012 Pilati 2017-C1	Events 20 37 92 28	Total 23 60 129 71	Events 290 570 31 175	Total 389 1003 53 396 72	14.6% 22.3% 20.9% 22.5%	M-H, Random, 95% Cl 2.28 [0.66, 7.82] 1.22 [0.72, 2.09] 1.76 [0.91, 3.44] 0.82 [0.49, 1.38]	
Study or Subgroup Dalerba 2016-C1 Dalerba 2016-C2 Huang 2012 Pilati 2017-C1 Zhang 2016	Events 20 37 92 28	Total 23 60 129 71 51	Events 290 570 31 175	Total 389 1003 53 396 72	14.6% 22.3% 20.9% 22.5% 19.6%	M-H, Random, 95% Cl 2.28 [0.66, 7.82] 1.22 [0.72, 2.09] 1.76 [0.91, 3.44] 0.82 [0.49, 1.38] 0.20 [0.09, 0.43]	
Study or Subgroup Dalerba 2016-C1 Dalerba 2016-C2 Huang 2012 Pilati 2017-C1 Zhang 2016 Total (95% Cl)	Events 20 37 92 28 15 192	Total 23 60 129 71 51 334	Events 290 570 31 175 49 1115	Total 389 1003 53 396 72 1913	14.6% 22.3% 20.9% 22.5% 19.6% 100.0%	M-H, Random, 95% CI 2.28 [0.66, 7.82] 1.22 [0.72, 2.09] 1.76 [0.91, 3.44] 0.82 [0.49, 1.38] 0.20 [0.09, 0.43] 0.92 [0.45, 1.90]	



Overall, our quantitative analysis indicates that a beneficial impact of CDX2 redundancy is correlated with better 3-year, 5-year, and 10-year OS and DFS, which does not account for subgroup confounding factors. Additionally, this positive prognostic role was confirmed to be independent of cancer types, mean ages, and sample sizes. Unlike other nuclear transcription factors, CDX2 has been experimentally confirmed to have carcinostatic and carcinogenic roles in various solid malignancies. Multiple researchers have reported that CDX2 could inhibit cell growth which is associated with significant cell cycle arrest at the G0/G1 phase [36]. Several studies have shown that excessive expression of CDX2 could prompt cell apoptosis in malignancies [37, 38]. However, other studies have suggested that CDX2 can function as an oncogene as well, promoting the proliferation of cells beyond their normal constraints [39, 40]. Our pooled results suggest that despite the presence of its carcinogenic role, the anticancer activity of CDX2 has stronger effects on the clinical prognoses of cancer patients. Additionally, in-depth molecular-level investigations have demonstrated that CDX2 can reduce the expression of downstream target genes, including axis inhibition protein 2 (*AXIN2*) and lung lineage transcription factor *Nkx2-1*, which subsequently inhibit tumour invasion and metastasis [41, 42]. This is a potential explanation for the better prognosis that is associated with CDX2 overexpression. Regarding the source regions, European patients displayed

Α	CDX2-	. (DX2+		Odds Ratio	Odds Ratio
Study or Subgroup	Events 1	Fotal Eve	ents To	al Weight	M-H, Random, 95% Cl	M–H, Random, 95% Cl
8.3.1 Colorectal canc						
Pilati 2017–C1	22	39	27	40 26.4%	0.62 [0.25, 1.56]	
Pilati 2017–C2	18	23	40	75 23.6%	3.15 [1.06, 9.37]	
Wong 2011	3	32	10	32 19.1%	0.23 [0.06, 0.93]	_
Subtotal (95% CI)		94	1	47 69.1%	0.80 [0.20, 3.25]	
Total events	43		77			
Heterogeneity: Tau ² =	,	,		= 0.009);	$r^{2} = 79\%$	
Test for overall effect:	Z = 0.31	(P = 0.75))			
8.3.2 Ovarian Cancer						
Huang 2012	70	129 129		53 30.9% 5 3 30.9 %		
Subtotal (95% CI)	70	129		55 50.9%	1.23 [0.65, 2.34]	
Total events	70		26			
Heterogeneity: Not ap Test for overall effect:	•	(D _ O E 2)				
Test for overall effect:	Z = 0.64	(P = 0.52))			
Total (95% CI)		223	2	00 100.0%	0.93 [0.39, 2.25]	-
Total events	113		103			
Heterogeneity: Tau ² =	0.55; Chi ²	$^{2} = 9.94,$	df = 3 (P	= 0.02); I ²	= 70%	0.01 0.1 1 10 100
Test for overall effect:	Z = 0.16	(P = 0.87))			Favours [CDX2+] Favours [CDX2-]
Test for subgroup diff	erences: C	$hi^2 = 0.3$	0, df = 1	(P = 0.58),	$I^2 = 0\%$	
Test for subgroup diff ${f B}$	CDX2-	. (DX2+		Odds Ratio	Odds Ratio
B Study or Subgroup	CDX2-	. (DX2+			
B Study or Subgroup 8.2.1 I-IV	CDX2- Events T	. (Fotal Eve	CDX2+ ents To	al Weight	Odds Ratio M-H, Random, 95% CI	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012	CDX2- Events 1 70	. (Fotal Eve 129	26	al Weight	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011	CDX2- Events T	. (<u> Total Eve</u> 129 32	26 10	al Weight 53 30.9% 32 19.1%	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI)	CDX2- Events 1 70 3	. (Fotal Eve 129	26 10	al Weight	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events	CDX2- Events 1 70 3 73	. (<u>Fotal Eve</u> 129 32 161	26 10 36	al Weigh 53 30.9% 32 19.1% 35 50.0 %	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² =	CDX2- Events 7 70 3 73 1.12; Chi ²	$ \begin{array}{c} \hline $	26 10 36 df = 1 (F	al Weigh 53 30.9% 32 19.1% 35 50.0 %	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events	CDX2- Events 7 70 3 73 1.12; Chi ²	$ \begin{array}{c} \hline $	26 10 36 df = 1 (F	al Weigh 53 30.9% 32 19.1% 35 50.0 %	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² =	CDX2- Events 7 70 3 73 1.12; Chi ²	$ \begin{array}{c} \hline $	26 10 36 df = 1 (F	al Weigh 53 30.9% 32 19.1% 35 50.0 %	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: 8.2.2 II/III	CDX2- Events 1 70 3 1.12; Chi ² Z = 0.62 ($\begin{array}{c} & () \\ \hline \textbf{Total} & \textbf{Eve} \\ 129 \\ 32 \\ \textbf{161} \\ \end{array}$ $\begin{array}{c} 2 \\ 2 \\ 2 \\ 2 \\ 4 \\ (P = 0.54) \end{array}$	$\frac{26}{10}$ $\frac{36}{df} = 1 (F)$	al Weight 3 30.9% 32 19.1% 35 50.0% = 0.03); 1 ²	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78%	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect:	CDX2- Events 7 70 3 73 1.12; Chi ²	$ \begin{array}{c} \hline $	26 10 36 df = 1 (F	al Weigh 53 30.9% 32 19.1% 35 50.0 %	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78% 0.62 [0.25, 1.56]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: 8.2.2 II/III Pilati 2017-C1	CDX2- Events 1 70 3 1.12; Chi ² Z = 0.62 (22	$\begin{array}{c} & & & & \\ \hline \textbf{Total} & \textbf{Eve} \\ 129 \\ 32 \\ \textbf{161} \\ \\ 161 \\ \\ 161 \\ \\ 162 \\ 162 \\ 161 \\ \\ 161 \\ 161 \\ \\ 161$	26 10 36 df = 1 (F 27 40	al Weight 53 30.9% 32 19.1% 35 50.0% = 0.03); 1² 40 26.4%	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78% 0.62 [0.25, 1.56] 3.15 [1.06, 9.37]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: 8.2.2 II/III Pilati 2017-C1 Pilati 2017-C2	CDX2- Events 1 70 3 1.12; Chi ² Z = 0.62 (22	$\begin{array}{c} & & & & \\ \hline \textbf{Total} & \textbf{Eve} \\ 129 \\ 32 \\ \textbf{161} \\ \\ 161 \\ \\ 2^2 = 4.62, \\ (P = 0.54) \\ \\ 39 \\ 23 \end{array}$	26 10 36 df = 1 (F 27 40	al Weight 53 30.9% 32 19.1% 35 50.0% = 0.03); 1 ² 40 26.4% 75 23.6%	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78% 0.62 [0.25, 1.56] 3.15 [1.06, 9.37]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: 8.2.2 II/III Pilati 2017-C1 Pilati 2017-C2 Subtotal (95% CI) Total events Heterogeneity: Tau ² =	CDX2- Events 1 70 3 1.12; Chi ² Z = 0.62 18 40 1.05; Chi ²	$\begin{array}{c} \text{Total Evolution}\\ 129\\ 32\\ 161\\ ^2 = 4.62,\\ (P = 0.54)\\ 39\\ 23\\ 62\\ ^2 = 5.00,\\ \end{array}$	$\begin{array}{c} 2DX2 + \\ ents & To \\ \hline 26 \\ 10 \\ \hline 36 \\ df = 1 (F) \\ 27 \\ 40 \\ \hline 1 \\ 67 \\ df = 1 (F) \end{array}$	al Weight 53 30.9% 32 19.1% 35 50.0% = 0.03); 1 ² 40 26.4% 75 23.6% 15 50.0%	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78% 0.62 [0.25, 1.56] 3.15 [1.06, 9.37] 1.36 [0.28, 6.68]	Odds Ratio
B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: 8.2.2 II/III Pilati 2017-C1 Pilati 2017-C2 Subtotal (95% CI) Total events	CDX2- Events 1 70 3 1.12; Chi ² Z = 0.62 18 40 1.05; Chi ²	$\begin{array}{c} \text{Total Evolution}\\ 129\\ 32\\ 161\\ ^2 = 4.62,\\ (P = 0.54)\\ 39\\ 23\\ 62\\ ^2 = 5.00,\\ \end{array}$	$\begin{array}{c} 2DX2 + \\ ents & To \\ \hline 26 \\ 10 \\ \hline 36 \\ df = 1 (F) \\ 27 \\ 40 \\ \hline 1 \\ 67 \\ df = 1 (F) \end{array}$	al Weight 53 30.9% 32 19.1% 35 50.0% = 0.03); 1 ² 40 26.4% 75 23.6% 15 50.0%	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78% 0.62 [0.25, 1.56] 3.15 [1.06, 9.37] 1.36 [0.28, 6.68]	Odds Ratio
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B Study or Subgroup 8.2.1 I-IV Huang 2012 Wong 2011 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: 8.2.2 II/III Pilati 2017-C1 Pilati 2017-C2 Subtotal (95% CI) Total events Heterogeneity: Tau ² = Test for overall effect: Total (95% CI) Total events	$\begin{array}{c} \text{CDX2-} \\ \hline \text{Events} & 1 \\ \hline 70 \\ 3 \\ \hline 73 \\ 1.12; \text{Chi}^2 \\ Z = 0.62 \\ \hline 22 \\ 18 \\ 40 \\ 1.05; \text{Chi}^2 \\ Z = 0.38 \\ \hline 113 \end{array}$	$\begin{array}{c} & & & & \\ \hline \textbf{Total} & \textbf{Eve} \\ 129 \\ 32 \\ \textbf{161} \\ 2^2 = 4.62, \\ (P = 0.54) \\ 23 \\ \textbf{62} \\ 2^2 = 5.00, \\ (P = 0.70) \\ \textbf{223} \end{array}$	$\begin{array}{c} \text{2DX2} + \\ \text{ents} & \text{To} \\ 26 \\ 10 \\ 36 \\ \text{df} = 1 (P \\ 40 \\ 1 \\ 67 \\ \text{df} = 1 (P \\ 1 \\ 0 \\ 2^{1} \\ 1 \\ 1 \\ 1 \\ 1 \\ 3 \end{array}$	al Weight 53 30.9% 32 19.1% 35 50.0% = 0.03); 1 ² 40 26.4% 75 23.6% 15 50.0% = 0.03); 1 ² 00 100.0%	Odds Ratio M-H, Random, 95% Cl 1.23 [0.65, 2.34] 0.23 [0.06, 0.93] 0.60 [0.12, 3.08] = 78% 0.62 [0.25, 1.56] 3.15 [1.06, 9.37] 1.36 [0.28, 6.68] = 80% 0.93 [0.39, 2.25]	Odds Ratio M-H, Random, 95% CI
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Figure 8: The correlation between CDX2 expression and cancer relapse. (A) based on different cancer types. (B) based on different TNM stages.

no obvious connection between CDX2 overexpression and a better prognosis in solid malignancies. The included studies of CDX2 expression and prognosis from the European population are insufficient. Thus, increasing the number of studies may result in consistency with other regions.

Based on our pooled evidence, the CDX2 expression level may not be associated with cancer

relapse, and the subgroup analysis for TNM stages and cancer types is consistent with this finding. Additionally, chemotherapeutic resistance is another challenge for therapeutic efficacy and patient prognosis, and our analysis showed that CDX2 overexpression in cancer patients with TNM IV stage diseases had a better response to adjuvant chemotherapy with first-line drugs. This conclusion is reasonable based on the current studies,

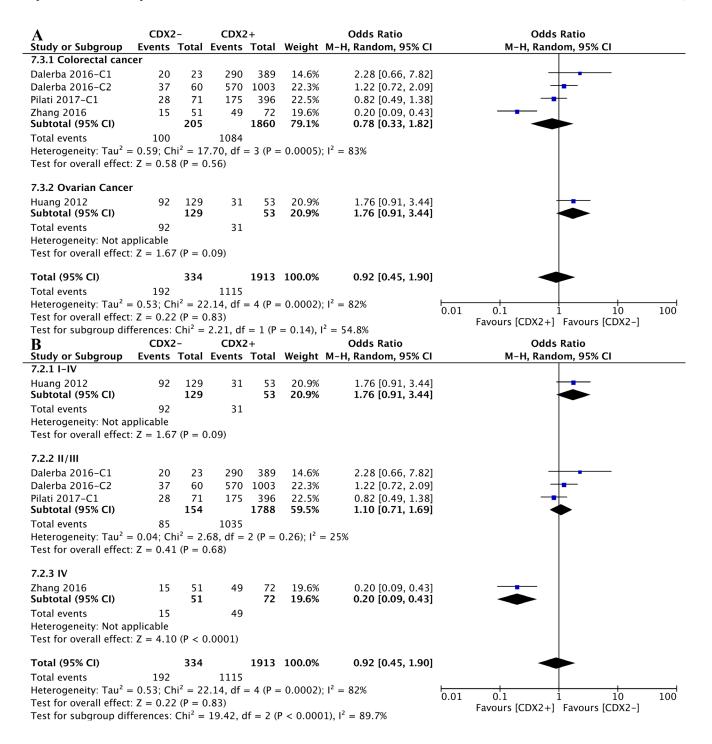


Figure 9: The correlation between CDX2 expression and chemotherapeutic effects. (A) based on different cancer types. (B) based on different TNM stages.

and CDX2 positivity is associated with microsatellite instability (MSI)-low malignancies, which respond better to chemotherapies than MSI-high malignancies [21, 27]. However, paradoxically, previous studies have confirmed that CDX2 induces the expression of the multidrug resistance protein 1 (MDR1) gene by binding to MDR1 promoter elements [39]. This finding illustrates that CDX2 may affect the response to cancer chemotherapy through multiple mechanisms. Meanwhile, similar results have not been obtained for CDX2 overexpression in cancer patients with TNM I-III stage diseases. In Stage I-II cancers, the adjuvant chemotherapy provides an unsatisfied benefit, which is usually not worth the toxic effects of the drugs. Stage III cancers have regional lymph node metastases, and multiple clinical studies have demonstrated significant increases in survival with the administration of adjuvant chemotherapy [43]. Our analysis lacks individual data for CDX2 expression that is associated with stage III malignancies. An elucidation of how CDX2 expression is associated with MSI status and TNM stage requires further study. Additionally, this result also suggests that chemotherapy may prolong the OS of high-CDX2-level cancer patients with TNM IV stage diseases.

There are several limitations in our meta-analysis. First, although random-effects models and subgroup analyses were used, heterogeneity could not be completely eliminated. This may have led to a specific bias in the results and may suggest that certain clinical elements were not fully analysed, such as tissue types and pathological grades. Moreover, Pilati et al-C1 and Schildberg et al were responsible for most of the internal heterogeneity across the studies because removal of both trials cleared the I2 value completely to 0% [27]. Schildberg et al attributed the opposite outcomes in the CDX2-positive groups to the unbalanced patient age distribution (i.e., no significant impact on survival with downregulated CDX2 concerning all patients less than 50 years of age) [29]. Nevertheless, our subgroup analysis did not show consistent results regarding the mean ages. Therefore, more pertinent studies are still necessary for further analysis. Second, although the total participant sample size exceeded 5000, the number of included studies was insufficient, especially for the analysis of CDX2 expression with the cancer relapse and the chemotherapeutic effects. Sex and ethnicity differences in biomarkers and response are currently the hot topics, our subgroup analysis failed to obtain these results due to the lack of data in the original study. Therefore, further studies are necessary to address these shortcomings to obtain more convincing conclusions.

Taken together, this meta-analysis demonstrates the beneficial effects of CDX2 overexpression on the life expectancy and prognosis of cancer patients, except for the European populations. In addition, CDX2 overexpression in cancer patients with TNM IV stage diseases performed a better response to adjuvant chemotherapy with first-line drugs. Therefore, we conclude that CDX2 is likely to be an important biomarker for guiding evaluations of clinical practice and prognosis.

MATERIALS AND METHODS

Literature search

We performed a systematic literature search using PUBMED and EMBASE up to May 2017 with the following terms as the search strategy: "CDX2 AND (tumor OR cancer OR malignancy OR neoplasm OR adenoma)". Furthermore, we manually screened the references in relevant articles to identify additional articles.

Study selection

Studies in accordance with the following criteria were included: 1. articles were written in English; 2.the expression of CDX2 was detected by immunohistochemistry; and 3. human studies addressing the correlation between CDX2 expression and clinical prognosis in solid malignancies.

Studies were excluded due to the following reasons: 1. inadequate survival data for further statistical analysis; 2. a follow-up duration that was shorter than 3 years; 3. duplicated or overlapping studies; 4. inappropriate article types such as reviews or case-reports; and 5. studies based on a sample size that comprised less than 10 participants.

Data extraction

Using predefined standardized extraction forms, two investigators independently extracted data from each qualified study. The general information, namely, 3-year OS, 5-year OS, 10-year OS, DFS, cancer relapse and chemotherapeutic effect was documented. The original survival data were obtained from the text, tables or Kaplan-Meier curves for both comparative groups.

Methodological assessment

The NOS was used to quantitatively evaluate the included study quality. Studies graded with more than six scores using the methodology with a maximum score of nine were identified as high-quality trials.

Statistical analysis

We summarized the pooled outcomes using forest plots (Review Manager version 5.3) and calculated the I^2 as a measure of heterogeneity. I^2 values were prespecified to indicate low (< 25%), moderate (25%–50%), and high (> 50%) heterogeneity. P values less than 0.05 were considered significant. Moreover, a sensitivity analysis was applied to examine the stability of the results. We

visually assessed the funnel plots and used the Egger's test and Begg's test to assess the publication bias (STATA/MP version 14.0). We performed the trim-and-fill procedure to further evaluate the possible publication bias on our meta-analysis. Two investigators separately reviewed the eligible studies and extracted the data, and a joint decision was made in cases of disagreement.

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CONFLICTS OF INTEREST

The authors declare no conflicts of interest.

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