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Cardiovascular disease risk and androgen deprivation therapy in patients with localised prostate cancer: a prospective cohort study

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Background: As androgen deprivation therapy (ADT) is increasingly being used in men with localised prostate cancer, our goal was to examine the association between ADT and the risk of cardiovascular disease (CVD).

Methods: We conducted a prospective cohort study using records of a large health-care system in California. The study included men with newly diagnosed localised prostate cancer (1998–2008) who initially underwent active surveillance ($N=7637$) and were followed through 2010. We examined 10 individual CVD outcomes. Cox proportional hazard models incorporated time-varying treatment variables and controlled for race/ethnicity, age, and tumour characteristics, recurrence risk, CVD medication use, and CVD risk factors.

Results: Of the 7637 subjects, nearly 30% were exposed to ADT. In the multivariable analyses, ADT was associated with an increased risk of heart failure (adjusted HR = 1.81, 95% CI 1.40–2.32) in men without preexisting CVD. Elevated risks of arrhythmia (adjusted HR = 1.44, 95% CI 1.02–2.01), and conduction disorder (adjusted HR = 3.11, 95% CI 1.22, 7.91) were only observed among patients with preexisting CVD.

Conclusions: In men with clinically localised prostate cancer who were initially under active surveillance, ADT was associated with a greater risk of heart failure in men without any preexisting CVD. We also found an increased risk of arrhythmia and conduction disorder in men with preexisting CVD. This study provides the basis for identifying high-risk men treated with ADT who might benefit from regular cardiac monitoring and lifestyle modification recommendations.

Androgen deprivation therapy (ADT), a hormone treatment, is recommended for advanced prostate cancer but is increasingly being used for localised prostate cancer treatment (Sharifi *et al*, 2005). In parallel, localised prostate cancer diagnosis has tremendously increased, reaching 90% in the United States, given the widespread prostate screening antigen

(PSA) testing (Thompson *et al*, 2007). ADT lowers male hormone to castration testosterone levels within 3 weeks after treatment (Burton *et al*, 2013). Several randomised clinical trials suggest that radiation therapy plus ADT improves overall survival among patients with locally advanced disease (i.e., extracapsular or node positive) *vs* radiation therapy alone, but

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findings remain controversial (Bolla *et al*, 1997, 2002, 2009; Horwitz *et al*, 2008).

Despite certain benefits of ADT, it is associated with problematic side effects similar to those experiencing testosterone deficiency due to other causes, such as gynaecomastia and erectile dysfunction (Sharifi *et al*, 2005; Cattabiani *et al*, 2012; Nguyen *et al*, 2015). Additionally, androgen deficiency has been linked to cardiovascular disease (CVD) risk factors. For example, low bioavailable testosterone may increase atherosclerosis risk (Hak *et al*, 2002; Jones *et al*, 2005; Khaw *et al*, 2007). However, these studies mainly included much older men.

Conflicting evidence exists regarding the association between ADT and CVD, and prior studies have not been able to sufficiently adjust for a comprehensive set of confounders (Bolla *et al*, 2002; Keating *et al*, 2006; D'Amico *et al*, 2007; Van Hemelrijck *et al*, 2010; Nguyen *et al*, 2011; O'Farrell *et al*, 2015). A meta-analysis of observational studies found a 40% increased risk of non-fatal CVD in men with prostate cancer who used ADT, specifically gonadotropin-releasing hormone (GnRH) agonists (Bosco *et al*, 2015). By contrast, ADT was not associated with CVD mortality in an earlier report (Punnen *et al*, 2011). The cardiotoxicity of ADT was confirmed in a population-based study suggesting a slightly elevated myocardial infarction risk regardless of existing CVD history (Keating *et al*, 2013), though other studies suggested that CVD-specific and all-cause mortality only occurred in patients with preexisting CVD (Nanda *et al*, 2014; Ziehr *et al*, 2015). Thus, given the mixed cohorts of both patients with and without existing CVD, distinguishing whether ADT is associated with the development of CVD or accelerating the progression of the existing disease remains unclear. Although studies based on both SEER-Medicare and VA databases found moderate associations between ADT and non-fatal CVD events, they did not have access to a broad set of covariates, such as CVD medications, which are commonly used by men with localised prostate cancer (Keating *et al*, 2006, 2010).

Given this background, our goal was to assess the association of ADT and important incident CVD outcomes in a cohort that also included younger men with localised prostate cancer initially conservatively managed. We accounted for important confounders, including prior CVD history, PSA levels, CVD medications, and CVD risk factors. Additionally, we assessed whether ADT only has effect on new-onset CVD or also on the progression of preexisting cardiovascular conditions. To our knowledge, this is the first study to comprehensively account for CVD medications, including antiarrhythmic medications, anticoagulants, antihypertensive medications, calcium channel blockers, digoxin, nitrates, and antidiabetics.

MATERIALS AND METHODS

Study design and setting. We conducted a prospective cohort study of men with newly diagnosed localised prostate cancer using electronic health records from Kaiser Permanente Southern California (KPSC) health-care delivery system. KPSC includes a network of 6600 physicians who serve 14 hospitals, over 220 outpatient clinics, and nearly 4.3 million members. The health plan's database includes comprehensive information from prospectively recorded inpatient and outpatient diagnoses, clinical encounters, laboratory test values, pharmacy data and cancer registry data.

Participants. We identified 30 092 men diagnosed with prostate cancer between 1998 and 2008 from the health plan's NCI SEER-affiliated tumour registry (Potosky *et al*, 2014). Men were excluded if they had advanced stage prostate cancer at diagnosis (T4/nodal involvement/distant metastasis, $N=3352$); received other

therapies, including radiation, radical prostatectomy, or chemotherapy, within 1 year after prostate cancer diagnosis ($N=18\,904$); underwent orchiectomy within 1 year after diagnosis ($N=59$); received neoadjuvant ADT ($N=120$); or were missing date of death and other data errors ($N=16$). Four men who died on the cohort entry date were further excluded from the study, resulting in a final cohort of 7637 men with clinically localised prostate cancer who initially were under active surveillance.

Androgen deprivation therapy. The main independent variable was exposure to any ADT after prostate cancer diagnosis. ADT was ascertained from electronic pharmacy dispensing records and outpatient procedures. This included men who might have received ADT as their sole primary therapy for clinically localised disease or as salvage therapy anytime thereafter. ADT was defined as a GnRH analogist (leuprolide, goserelin, or triptorelin) with or without an oral antiandrogen (flutamide, bicalutamide, or nilutamide for combined androgen blockade). ADT was treated as dichotomous (ever/never) time-dependent variable. During the study period, ADT antagonists (abarelix and degarelix) were not available in the KPSC formulary. Men contributed information to the exposed group only after the drug initiation and were categorised as exposed to ADT once initiated until reaching one of the study's end points.

Study outcomes. The cohort was followed through death or 31 December 2010 for the study endpoints. We identified first occurrence of CVD outcomes based on hospitalisation with a primary diagnoses of a CVD event (using ICD9 and ICD10 discharge diagnosis codes; see Supplementary Appendix 1). We focussed on CVD outcomes that were serious enough to require hospitalisation using a previously published algorithm with demonstrated validity of diagnosis codes in the electronic health records (Haque *et al*, 2011). Men with multiple primary CVD diagnoses on the same date were assigned to one outcome group according to this algorithm: (1) acute myocardial infarction, (2) cardiac arrest, (3) stroke, (4) heart failure (HF), (5) hypertensive heart disease with HF, (6) cardiomyopathy, (7) arrhythmias; (8) valvulopathy, (9) angina pectoris; and (10) conduction disorder. Furthermore, we combined the individual events and grouped them into three major categories based on the mechanism of disease: cardiac ischaemia, stroke, and other heart disease (Haque *et al*, 2011). The CVD outcomes were treated as a binary variable. Information on the date of death for health plan members were derived from a combination of clinical databases, linkages with California death certificate records, and linkage with the national Social Security Administration data. There were 178 patients who died owing to CVD but with no specified CVD diagnoses during their health plan membership. We handled such patients in two ways. First, patients were combined with those who died owing to any other causes and were censored at the death date. Second, we excluded the 178 men in sensitivity analysis and compared the HR results with those of the full cohort.

Covariates. Because ADT users and non-users could have been systematically different in prognostic factors of prostate cancer that may also affect the risk of CVD, we compared tumour characteristics (year of diagnosis, Gleason score, prostate cancer risk group, baseline PSA level) between the two groups and adjusted those factors in the multivariable analyses. In addition, we extracted other covariates, including patients' demographics (age, race/ethnicity), prior history of CVD, risk factors for CVD (hypertension, diabetes), Charlson comorbidity score, and CVD medications. The CVD medications included antiarrhythmics, anticoagulants, antilipemics, calcium channel blockers, digoxin/lanoxin, nitrates, and antidiabetic drugs. As the antihypertensive medication use did not differ by exposure status, we excluded it

from the covariate set. The demographic information and tumour characteristics were gathered from electronic health records around the prostate cancer diagnosis date while CVD risk factors and comorbidity scores were captured 1 year before diagnosis. CVD medications were extracted from pharmacy dispensing and coded as time-dependent variables. History of preexisting CVD conditions (at least two records) were ascertained from inpatient and outpatient electronic health records up to 2 years before prostate cancer diagnosis. Men must have had at least two consecutive prescriptions within 90 days during the follow-up period to be counted as exposed to a particular CVD medication.

Statistical analyses. Differences in demographic characteristics, tumour status, CVD risk factors, and medications were first examined by comparing frequency distributions by ADT use status. Because subjects had varying lengths of follow-up, we examined the person-year rates of each of the CVD outcomes. To estimate the associations between ADT and each of the 10 cardiovascular outcomes, we fit 10 separate time-dependent Cox-proportional hazards models. For each model, the follow-up years began on prostate cancer diagnosis date and ended on the date of the incident CVD as the outcome event or censoring event corresponding to termination of health plan membership, death, or end of the study period (31 December 2010), whichever occurred first. In the Cox models, ADT and the prescription medications were treated as time-dependent variables (i.e., 0 up to start date; 1 after drug initiation). The multivariable Cox proportional hazards models were adjusted for all potential confounders deemed to be clinically or statistically significant ($P < 0.05$). These confounders included demographics, tumour characteristics (stage at prostate cancer diagnosis, Gleason score, risk group), and CVD medications. In sensitivity analyses, we further adjusted for BMI and smoking. The proportional hazards assumption was tested via graphic plots and Schoenfeld residuals. Because a proportion of cases were missing at least one or more of the key clinical prognostic variables (clinical stage, Gleason score, or baseline PSA), we performed multiple imputations using all other covariates. We constructed five imputed data sets, each having estimates for the missing values for PSA, Gleason score, and T-stage. We then pooled the estimates and corresponding s.e. across the five imputations using Rubin's method (Rubin, 2008). All model results used these imputed data sets; multivariable models using only the complete cases did not show any significant deviations from the results presented here. To distinguish the effect on newly diagnosed CVD from the effect on the progression on preexisting CVDs, we also conducted a sensitivity analysis to assess whether preexisting CVD was an important confounder by performing analyses stratified by the presence or absence of preexisting CVD. In another sensitivity analysis, we repeated the Cox model excluding the men who died of CVD without having a prior CVD diagnosis ($N = 178$). Because these outcomes were studied in prior reports, we decided not to make adjustments in the P -values for multiple comparisons per Rothman (1990). All statistical analyses were conducted using SAS (Version 9.4, SAS institute Inc., Cary, NC, USA).

RESULTS

Of the 7637 men with localised prostate cancer who did not undergo curative intent therapy (initially conservatively managed), nearly 30% ($n = 2170$) were exposed to ADT during a median of 3.4 years of follow-up (interquartile range: 1.6–6.0 years). The cohort includes both men with ($N = 1665$, 22%) and without ($N = 5972$, 78%) preexisting CVD. Of the 2170 exposed to ADT, 36% began treatment within 6 months after

prostate cancer diagnosis, 32% between 6–24 months, and 32% after 2 years. Thus ADT might have been prescribed as primary or salvage therapy. There were 2061 men who experienced CVD events over 31255 person-years, which translated to 65.9 event per 1000 PY. The average time to a CVD event was 3.2 years (s.d. 2.6 years). The ADT group had significantly longer follow-up compared with the non-user group (average of 3.6 vs 5.3 years).

Comparison of patients by ADT status. Demographics, comorbidities, CVD risk factors, and tumour characteristics are displayed in Table 1 by ADT status. The ADT group was more likely to be older, have hypertension, diabetes, and more comorbid conditions 1 year before prostate cancer diagnoses. Prevalence of preexisting CVD was 22.5% among men who used ADT and 21.5% among non-users. ADT use was strongly associated with patients' tumour characteristics. Compared with the ADT group, a larger proportion of the non-user group was diagnosed after 2006 (31.9 vs 20.5%), had lower Gleason scores (69.0 vs 42.3% for Gleason score < 7), at less advanced tumour stage (69.1 vs 55.8% for T1), and were in the lower-risk group (47.7 vs 19.2% for low risk). In the ADT group, a larger proportion of patients had higher baseline PSA level (17.3 vs 36.1% for PSA > 10 units).

Use of CVD and antidiabetic medications by ADT status. Table 2 shows patients' CVD medication use by ADT use. Significantly more patients in the ADT group used anticoagulants, antilipemics, calcium channel blockers, digoxin/lanoxin, nitrates, and antidiabetic drugs than patients in the non-user group. Thus ADT users were more likely to be treated for CVD.

CVD occurrence by ADT status. The person-year incidence rates, adjusted HRs and 95% CIs of each CVD event, and HRs stratified by history of preexisting CVD are shown in Table 3 and Figure 1. Generally, the unadjusted person-year absolute CVD rates were higher for men treated with ADT than for non-users. The crude CVD rates were higher among men with more advanced age at cancer diagnosis and with worse tumour characteristics, that is, higher T-stage, Gleason score, and recurrence risk group (Supplementary Appendix 2). After adjusting for all other covariates including CVD medications in the multivariate models, men exposed to ADT were 27% more likely to develop HF than non-exposed men (adjusted HR = 1.27, 95% CI: 1.06–1.51); this association was seen in men without preexisting CVD (adjusted HR = 1.81, 95% CI: 1.40–2.32) but not in men with preexisting CVD (adjusted HR = 1.00, 95% CI = 0.78–1.29). ADT users with preexisting CVD were at increased risk of developing arrhythmia (adjusted HR = 1.44, 95% CI: 1.02–2.01) and conduction disorders (adjusted HR = 3.11, 95% CI: 1.22–7.91; Figure 1). We did not find increased risks for ischaemic heart disease, stroke, or the other conditions. To assess the potential confounding effect of BMI and smoking, we performed a sensitivity analysis comparing models with and without additional adjustment for these factors in a subset of the cohort ($N = 1983$) with complete information of the two lifestyle factors; there was little change in the hazard ratios for the three major CVD groups (Table 4). After multiple imputations for missing data (PSA level: ($N = 1270$, 15.8%), Gleason score: ($N = 709$, 9.3%), and T-stage ($N = 96$, 1.3%), the HR results did not substantially change. In the sensitivity analysis, excluding those 178 patients who died of CVD, the HRs were approximately the same as in the full cohort (data not shown). In stratified analysis among African-American patients (data not shown), the magnitude of association was greater in arrhythmia (HR = 1.46, 95% CI: 0.87–2.45) and HF (HR = 1.65, 95% CI: 1.09–2.50). We also found increased risks in angina (HR = 4.37, 95% CI: 0.99–13.36) and cardiomyopathy (HR = 7.86, 95% CI: 1.22–50.62) in African-American men, although confidence intervals were wide.

Table 1. Demographic and tumour characteristics of prostate cancer patients at baseline by ADT exposure

	ADT not used (N = 5467)	ADT used (N = 2170)	Total (N = 7637)	P-value
Age, years				
≤65	2072 (37.9)	430 (19.8)	2502 (32.8)	<0.0001
>65	3395 (62.1)	1740 (80.2)	5135 (67.2)	
Charlson score 1 year before PC diagnoses				
0	2716 (60.3)	826 (52.7)	3542 (58.4)	<0.0001
1–2	1426 (31.7)	584 (37.3)	2010 (33.1)	
3+	361 (8.0)	157 (10.0)	518 (8.5)	
Missing	964	603	1567	
Race/ethnicity				
Hispanic	793 (14.5)	312 (14.4)	1105 (14.5)	0.0097
White	3356 (61.4)	1292 (59.5)	4648 (60.9)	
African-American	856 (15.7)	405 (18.7)	1261 (16.5)	
Asian	462 (8.5)	161 (7.4)	623 (8.2)	
Hypertension				
No	2237 (40.9)	789 (36.4)	3026 (39.6)	0.0002
Yes	3230 (59.1)	1381 (63.6)	4611 (60.4)	
Diabetes				
No	4572 (83.6)	1729 (79.7)	6301 (82.5)	<0.0001
Yes	895 (16.4)	441 (20.3)	1336 (17.5)	
Year of PC diagnosis				
1998–2000	1256 (23.0)	826 (38.1)	2082 (27.3)	<0.0001
2001–2005	2465 (45.1)	899 (41.4)	3364 (44.0)	
2006+	1746 (31.9)	445 (20.5)	2191 (28.7)	
Stage at diagnosis				
T1	3726 (69.1)	1196 (55.8)	4922 (65.3)	<0.0001
T2	1600 (29.7)	837 (39.0)	2437 (32.3)	
T3	69 (1.3)	112 (5.2)	181 (2.4)	
Missing	72	25	97	
Gleason score				
<7	3388 (69.0)	851 (42.3)	4239 (61.2)	<0.0001
7	1088 (22.1)	631 (31.4)	1719 (24.8)	
>7	437 (8.9)	528 (26.3)	965 (13.9)	
Missing	554	160	714	
Risk groups				
Low	2218 (47.7)	389 (19.2)	2607 (39.1)	<0.0001
Intermediate	1398 (30.1)	637 (31.5)	2035 (30.5)	
High	1030 (22.2)	997 (49.3)	2027 (30.4)	
Missing	821	147	968	
Preexisting CVD				
No	4290 (78.5)	1682 (77.5)	5972 (78.2)	0.3598
Yes	1177 (21.5)	488 (22.5)	1665 (21.8)	
Smoking status^a				
Current	451 (12.8)	190 (12.9)	641 (12.8)	0.0078
Former	1472 (41.7)	680 (46.1)	2152 (43.0)	
Never	1611 (45.6)	605 (41.0)	2216 (44.2)	
Missing	1933	695	2628	
BMI^b				
Underweight or normal	498 (30.0)	172 (38.0)	670 (31.7)	0.0098
Overweight	729 (43.9)	181 (40.0)	910 (43.0)	
Obese	435 (26.2)	100 (22.1)	535 (25.3)	

Abbreviations: ADT = androgen deprivation therapy; BMI = body mass index; CVD = cardiovascular disease; PC = prostate cancer. P-value based on Pearson's chi-square test, two-sided.
^a35% was missing, smoking history was not available in EMR until 2004.
^b72% missing, BMI was not available in EMR until 2006.

DISCUSSION

In this prospective population-based study of insured men with localised prostate cancer who were initially conservatively managed

Table 2. Simultaneous medications related with CVD by ADT exposure

	ADT not used (N = 5467)	ADT used (N = 2170)	Total (N = 7637)	P-value
Antiarrhythmics				
No	5457 (99.8)	2162 (99.6)	7619 (99.8)	0.1311
Yes ^a	10 (0.2)	8 (0.4)	18 (0.2)	
Anticoagulants				
No	5182 (94.8)	1978 (91.2)	7160 (93.8)	<0.0001
Yes ^a	285 (5.2)	192 (8.8)	477 (6.2)	
Antihypertensives				
No	5201 (95.1)	2064 (95.1)	7265 (95.1)	0.9719
Yes ^a	266 (4.9)	106 (4.9)	372 (4.9)	
Antilipemics				
No	3653 (66.8)	1187 (54.7)	4840 (63.4)	<0.0001
Yes ^a	1814 (33.2)	983 (45.3)	2797 (36.6)	
Calcium channel blockers				
No	5004 (91.5)	1902 (87.6)	6906 (90.4)	<0.0001
Yes ^a	463 (8.5)	268 (12.4)	731 (9.6)	
Antidiabetic drugs				
No	5090 (93.1)	1928 (88.8)	7018 (91.9)	<0.0001
Yes ^a	377 (6.9)	242 (11.2)	619 (8.1)	
Digoxin/lanoxin				
No	5364 (98.1)	2092 (96.4)	7456 (97.6)	<0.0001
Yes ^a	103 (1.9)	78 (3.6)	181 (2.4)	
Nitrates				
No	5232 (95.7)	1982 (91.3)	7214 (94.5)	<0.0001
Yes ^a	235 (4.3)	188 (8.7)	423 (5.5)	

Abbreviations: ADT = androgen deprivation therapy; CVD = cardiovascular disease. P-value based on Pearson's chi-square test, two-sided.
^aYes: there are at least two consecutive prescriptions within 90 days.

and followed a maximum of 13 years, we found associations between ADT and certain CVD outcomes after adjustment for multiple confounder variables, including hypertension, diabetes, age and stage of prostate cancer diagnosis, Gleason score, prostate cancer risk score, race/ethnicity, and covariate CVD medications. The increased risk of other heart diseases (e.g., arrhythmia, HF, and conduction disorder) among ADT users persisted even in the sensitivity analysis that controlled for BMI and smoking. ADT use was associated with an increased risk of HF; this finding was similar to a study in UK General Practice Research Database but they observed a slightly smaller effect size (Martin-Merino *et al*, 2011). ADT was also associated with arrhythmia and conduction disorders but only among men with preexisting CVD. A major advantage of our study was that we were able to consider the influence of cardiovascular medications and examine a comprehensive set of outcomes.

Given the pharmacological functions of ADT, our findings may be explained by possible mechanisms pertaining to testosterone deficiency and CVD risk factors. For example, testosterone deficiency increases fat mass, a risk factor of CVDs, such as HF and arrhythmia. Men with androgen deficiency also have abnormal lipid profiles, elevated pro-inflammatory factors, endothelial function, and hypertension (Traish *et al*, 2009). In the Caperhilly Study, Smith *et al* (2005) suggested that the cortisol/testosterone ratio was related with ischaemic heart disease, mediated through insulin resistance, which may exacerbate conduction disorders (Traish *et al*, 2009; Keating *et al*, 2010). Therefore, the negative impact of testosterone deficiency may derive from its influencing CVD risk factors. The observed increased risk of arrhythmia among ADT users in our study is consistent with the confirmed inverse association between testosterone level and prolonged QT interval, a strong risk factor of fatal arrhythmias (van Noord *et al*,

Table 3. Crude disease rates and hazard ratios of incident CVD outcomes

CVD	Total, PY: 31 255		ADT used, PY: 7527		ADT not used, PY: 23 728		All patients (n = 7637)		Patients without preexisting CVD (n = 5972)		Patients with preexisting CVD (n = 1665)	
	N	Crude rate ^a	N	Crude rate	N	Crude rate	Adjusted HR	95% CI	Adjusted HR	95% CI	Adjusted HR	95% CI
Cardiac ischaemia												
AMI	431	13.79	139	18.47	292	12.31	1.19	0.95, 1.49	1.24	0.93, 1.64	1.14	0.78, 1.65
Angina	61	1.95	18	2.39	43	1.81	0.95	0.52, 1.75	0.64	0.24, 1.67	1.31	0.58, 2.93
Cardiac arrest	58	1.86	19	2.52	39	1.64	1.15	0.62, 2.12	1.29	0.61, 2.76	0.85	0.25, 2.83
Total	550	17.68	176	23.38	374	15.76	1.15	0.95, 1.41	1.18	0.91, 1.52	1.13	0.82, 1.57
Stroke												
Total	71	2.27	23	3.06	48	2.02	1.09	0.63, 1.88	0.93	0.46, 1.88	1.62	0.64, 4.12
Other heart diseases												
Arrhythmia	564	18.05	176	23.38	388	16.35	1.17	0.96, 1.43	1.07	0.84, 1.37	1.44	1.02, 2.01
Heart failure	689	22.04	244	32.42	445	18.75	1.27	1.06, 1.51	1.81	1.40, 2.32	1.00	0.78, 1.29
Cardiomyopathy	36	1.15	16	2.13	20	0.84	1.84	0.87, 3.88	2.25	0.74, 6.78	1.51	0.51, 4.46
Conduction disorder	81	2.59	27	3.59	54	2.28	1.40	0.84, 2.35	1.08	0.57, 2.03	3.11	1.22, 7.91
Hypertensive heart disease	3	0.1	1	0.13	2	0.08	NA	NA	NA	NA	NA	NA
Valvulopathy	67	2.14	22	2.92	45	1.90	1.22	0.69, 2.17	1.08	0.53, 2.21	1.59	0.59, 4.26
Total	1440	45.64	486	64.56	954	40.21	1.25	1.11, 1.42	1.34	1.14, 1.58	1.22	1.01, 1.47

Abbreviations: ADT = androgen deprivation therapy; AMI = acute myocardial infarction; CI = confidence interval; CVD = cardiovascular disease; HR = hazard ratio; NA = not applicable; PY = person-years. HRs adjusted for: ADT (time dependent), hypertension (Y/N), diabetes (Y/N), Gleason scores (<7/7/>7), stage at diagnosis (T1/T2/T3), grading for PC (1/2/3/4), year of PC diagnosis (<2000/2001–2005/≥2006), race (White/African/Asian/Hispanic), age (>65/≤65 years), calcium channel blockers (time dependent), anticoagulants (time dependent), antilipemics (time dependent), digoxin (time dependent), nitrates (time dependent), antidiabetic drugs (time dependent), baseline PSA level (≤4/4–10/10–20/>20), risk group (low/intermediate/high), and Charlson score 1 year before diagnose (0/1–2/≥3). Bold values: 95% confidence intervals do not include 1.00.

^aCrude rate: rate per 1000 person-years.

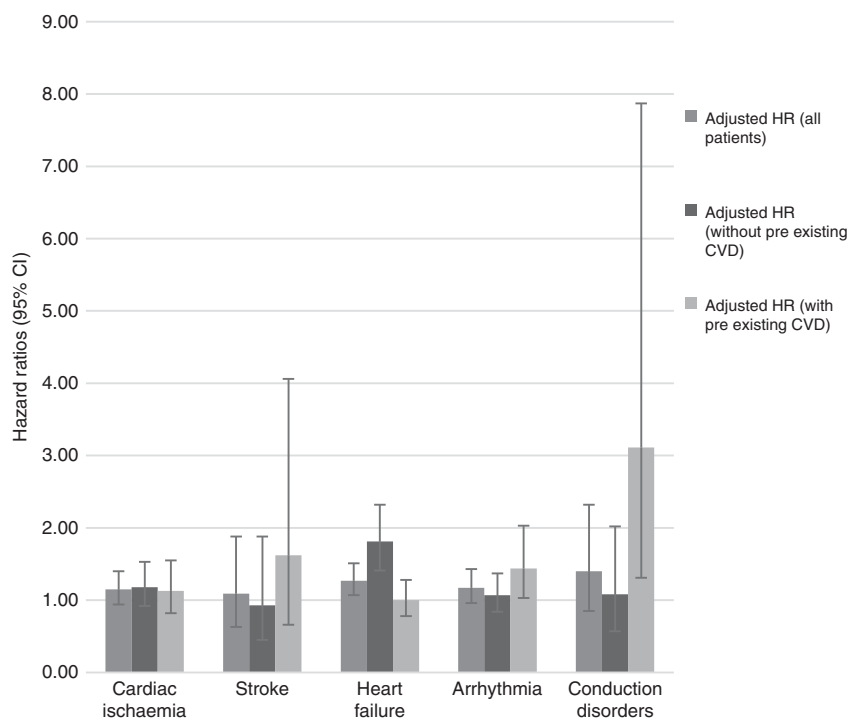


Figure 1. Adjusted HRs of CVD events following ADT exposures.

2011; Lubart *et al*, 2015). We also found an increased risk of conduction disorder among patients with preexisting CVD (HR = 3.11, 95% CI: 1.22–7.91) but not in those without CVD history (HR = 1.08, 95% CI: 0.57, 2.03). Reasons for this are not clear, but it is possible that those with preexisting CVD may be more susceptible to developing conduction disorders or that individuals with certain preexisting conditions are more likely to have greater health-care utilisation, be monitored more frequently,

and hence, more likely to be diagnosed with other CVD conditions. This observation requires additional confirmation.

The adjusted hazard ratios of ischaemic heart disease and stroke were >1.00 (1.15 and 1.09) but not significantly different between the ADT use and non-use. The lack of elevated risks in our study may be attributed to higher proportion of younger men in our study compared with the rates in a SEER-Medicare study (33% of men in our cohort were aged ≤65 years; Keating *et al*, 2006), the

Table 4. Hazard ratio estimates of incident CVD outcomes from different models (N = 1983)

CVD	Multivariable models ^a		Additional adjustment for smoking		Additional adjustment for BMI + smoking	
	Adjusted HR	95% CI	Adjusted HR	95% CI	Adjusted HR	95% CI
Cardiac ischaemia ^b	1.15	0.94, 1.40	1.14	0.93, 1.39	1.12	0.92, 1.37
Stroke	1.04	0.60, 1.80	1.01	0.58, 1.76	1.01	0.58, 1.76
Other heart diseases ^c	1.24	1.10, 1.40	1.24	1.10, 1.40	1.23	1.09, 1.39

Abbreviations: ADT = androgen deprivation therapy; CI = confidence interval; CVD = cardiovascular disease; HR = hazard ratio; PY = person-years.

^aHRs adjusted for: ADT (time dependent), hypertension (Y/N), diabetes (Y/N), Gleason scores (<7/7/>7), stage at diagnosis (T1/T2/T3), grading for PC (1/2/3/4), year of PC diagnosis (<2000/2001–2005/≥2006), race (White/African/Asian/Hispanic), age (>65/≤65 years), calcium channel blockers (time dependent), anticoagulants (time dependent), antilipemics (time dependent), digoxin (time dependent), nitrates (time dependent), antidiabetic drugs (time dependent), baseline PSA level (≤4/4–10/10–20/>20), risk group (low/intermediate/high), and Charlson score 1 year before diagnose (0/1–2/≥3).

^bIncludes acute myocardial infarction, angina, and cardiac arrest.

^cIncludes arrhythmia, heart failure, cardiomyopathy, conduction disorder, hypertensive heart disease, and valvulopathy.

health plan's management of CVD risk factors, and our ascertainment of outcomes from inpatient databases only (i.e., those conditions that were serious enough to require hospitalisation). Importantly, we were able to ascertain oral antiandrogen use, thereby minimising exposure misclassification. If we had also included patients who died of CVD as part of the outcome (as opposed to excluding them in the sensitivity analysis), the magnitude of our observed would have been even greater as those patients were more likely to be exposed to ADT.

Our study has several strengths. The KPSC health plan data enabled us to adjust for a comprehensive set of covariates, enabled long-term follow-up, and provided a large population. Moreover, the oral medications were extracted from dispensed prescriptions, ensuring that patients were indeed exposed to such medications. Unlike other studies, we were able to account for the confounding effects of antidiabetic medication use as well as other cardiovascular medications. For example, as nearly 40% of the men used antilipemics, we were able to account for this drug and other medications that were strongly related with CVD outcomes. In the SEER-Medicare studies, the higher risk of CVD might have been attributed to older study populations and their inability to account for CVD medication use, although other studies that included younger populations have also demonstrated associations of ADT with CVD (Keating *et al*, 2010). With the prospectively collected utilisation data, we were able to assess if history of preexisting CVD and known CVD risk factors (hypertension, diabetes, race/ethnicity, and other comorbidities) accounted for all the elevated risk of ADT on CVD. Furthermore, the time-dependent Cox proportional models allowed us to address immortal time bias. Additionally, the health plan's members' socioeconomic status, cancer incidence, and race/ethnicity reflect the general southern California population. Thus our results are generalisable to the region the health plan serves. However, the study is based in a managed care setting where patients with risk factors for CVD are closely monitored. Finally, we examined 10 individual CVD events rather than a single composite outcome by doing 10 statistical tests, which allows us to distinguish potential differential effect on specific CVD subtypes. We do not believe that the observed increased risks were simply due to chance, considering prior knowledge of the associations between lower hormone level and CVD incidences. Thus no adjustment for multiple comparison was conducted (Rothman, 1990).

Certain limitations of our study should also be considered. First, there may be residual confounding by indication. Although we adjusted for numerous potential confounders in the model and addressed missing data, we cannot rule out the possibility that some factors that may influence ADT use and the outcomes were not captured. Well-conducted randomised controlled trials are needed to verify the study results. Second, owing to data limitations, we could not examine the duration of ADT use and treated the exposure as permanent use once prescription was initiated even if they

discontinued afterwards; the rationale of handling ADT exposure in this way is based on the prior study by Keating *et al* (2006). In that study, a sensitivity analysis was conducted comparing two definitions of ADT exposure: (1) as continuously treated for 6 months after each dose of GnRH agonist, and (2) as permanently used on GnRH agonist therapy. Similar results were obtained with the two definitions, because in some men, hypogonadism persisted for >6 months after discontinuing GnRH agonist therapy (Keating *et al*, 2006). Another limitation is that we did not distinguish intermittent and continuous ADT administration. However, a recent clinical trial comparing intermittent and continuous ADT treatment indicated an increased cumulative incidence of ischaemic events following intermittent ADT (Hershman *et al*, 2016). Third, disenrollment from the health plan among ADT group *vs* the non-user group might contribute to selection bias. Fourth, we did not have BMI status for 72% of the subjects, nor smoking status for 35% of the subjects, important predictors of CVDs as these factors were not electronically available until 2006 in our electronic health records. However, in sensitivity analyses restricted to 1983 subjects with complete BMI and smoking status, the additional adjustment for these two covariate minimally changed the hazard ratio estimates (Table 4), probably because our adjustment of CVD drugs, diabetes, hypertension, statins, and antidiabetic medications have indirectly controlled for smoking and potentially other metabolomic disorders. Fifth, our study only examined one subtype of ADT, GnRH agonists; however, prior studies show that most men are treated with such agonists while GnRH antagonists (which were not available in the KPSC formulary) were less commonly used. Separate studies that focus on cardiovascular events following GnRH antagonist treatment are needed (Albertsen *et al*, 2014). Finally, we may have been underpowered to find differences for some of the rarer outcomes.

Future observational studies need to examine CVD outcomes in prostate cancer survivors who have been diagnosed in more recent years when the newer GnRH antagonist drugs were launched into the market and follow them for sufficient time to compare the long-term safety of different subtypes of ADT. Additionally, intervention studies addressing primary or secondary prevention of CVD in men with localised prostate cancer are needed.

CONCLUSIONS

In this population-based study, ADT was associated with greater risk of HF and arrhythmia and conduction disorder in men with preexisting CVD. These conditions affect patients' quality of life and morbidity and can possibly compromise survival. This study provides the basis for identifying susceptible individuals for regular cardiac check-up, including identifying subclinical signs, monitoring hypertension, diabetes, and encouraging physical activity

and healthy dietary habits. Indeed, emerging evidence exists that recommending diet and exercise to patients receiving ADT for prostate cancer treatment benefits the cardiometabolic profile (Galvao *et al*, 2009); such efforts may indirectly help lower the risk of non-fatal CVD in men with localised prostate cancer treated with ADT.

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CONFLICT OF INTEREST

The authors declare no conflict of interest.

ETHICS APPROVAL AND CONSENT TO PARTICIPATE

The Kaiser IRB approved this study and written informed consent was waived as we were using secondary data.

AUTHOR CONTRIBUTIONS

Study concept and design, acquisition of data and drafting of the manuscript: RH, XX. Analysis and interpretation of data and critical revision of the manuscript for important intellectual content: all authors. All authors read and approved the final manuscript.

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