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**Citation:** Li R, Tang Z, Liu F, Yang M (2021) Efficacy and safety of trimethoprimsulfamethoxazole for the prevention of pneumocystis pneumonia in human immunodeficiency virus-negative immunodeficient patients: A systematic review and meta-analysis. PLoS ONE 16(3): e0248524. https://doi.org/ 10.1371/journal.pone.0248524

Editor: Carlos M. Isales, Augusta University, UNITED STATES

Received: September 28, 2020

Accepted: February 27, 2021

Published: March 25, 2021

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Data Availability Statement: All relevant data are within the paper and its <u>Supporting Information</u> files.

**Funding:** Prof. Ming Yang is funded by Sichuan science and technology project (No.2018JY0415; http://kjt.sc.gov.cn/) and Mrs. Rui Li is funded by affiliated hospital of north sichuan medical college project (No.2019JC038; https://www.nsmc.edu.cn/). The funders had no role in study design,

RESEARCH ARTICLE

Efficacy and safety of trimethoprimsulfamethoxazole for the prevention of pneumocystis pneumonia in human immunodeficiency virus-negative immunodeficient patients: A systematic review and meta-analysis

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# Abstract

# Background

Pneumocystis pneumonia (PCP) has a significant impact on the mortality of immunocompromised patients. It is not known whether the prophylactic application of trimethoprim-sulfamethoxazole (TMP-SMZ) can reduce the incidence of PCP and mortality in the human immunodeficiency virus (HIV)-negative immunodeficient population. The safety profile is also unknown. There have been few reports on this topic. The aim of this study was to systematically evaluate the efficacy and safety of the use of TMP-SMZ for the prevention of PCP in this population of patients from the perspective of evidence-based medicine.

# Methods

A comprehensive search without restrictions on publication status or other parameters was conducted. Clinical randomized controlled trials (RCTs) or case-control trials (CCSs) of TMP-SMZ used for the prevention of PCP in HIV-negative immunocompromised populations were considered eligible. A meta-analysis was performed using the Mantel-Haenszel fixed-effects model or Mantel-Haenszel random-effects model, and odds ratios (ORs) with 95% confidence intervals (CIs) were calculated and reported.

# Results

Of the 2392 records identified, 19 studies (n = 4135 patients) were included. The efficacy analysis results indicated that the PCP incidence was lower in the TMP-SMZ group than in the control group (OR = 0.27, 95% CI (0.10, 0.77), p = 0.01); however, the rate of drug discontinuation was higher in the TMP-SMZ group than in the control group (OR = 14.31, 95%)

data collection and analysis, decision to publish, or preparation of the manuscript.

**Competing interests:** The authors have declared that no competing interests exist.

CI (4.78, 42.91), p<0.00001). In addition, there was no statistically significant difference in the rate of mortality between the two groups (OR = 0.54, 95% CI (0.21, 1.37), p = 0.19). The safety analysis results showed that the rate of adverse events (AEs) was higher in the TMP-SMZ group than in the control group (OR = 1.92, 95% CI (1.06, 3.47), p = 0.03).

## Conclusions

TMP-SMZ has a better effect than other drugs or the placebo with regard to preventing PCP in HIV-negative immunocompromised individuals, but it may not necessarily reduce the rate of mortality, the rate of drug discontinuation or AEs. Due to the limitations of the research methodologies used, additional large-scale clinical trials and well-designed research studies are needed to identify more effective therapies for the prevention of PCP.

# Introduction

Pneumocystis jirovecii was originally believed to be a protozoan. However, Stringer later discovered that it is an atypical fungus taxonomically located between Ascomycota and Basidiomycota that is resistant to most antifungal drugs [1]. P. jirovecii in the lungs can cause pneumocystis pneumonia (PCP), which is a severe and potentially fatal disease [2]. PCP has a significant impact on the mortality of immunocompromised patients, especially those with acquired immune deficiency syndrome (AIDS). PCP is the most common opportunistic infection and is the primary complication and cause of death in AIDS patients [3]. Because the clinical manifestations of PCP, such as shortness of breath, hypoxia, tachycardia, etc., are often the first symptoms in AIDS patients, PCP has been regarded as a hallmark disease signaling human immunodeficiency virus (HIV) infection. Consequently, PCP has received a great deal of attention. In recent years, the incidence of PCP and PCP-related mortality in HIV-positive patients have gradually decreased due to increasingly mature diagnostic technology, enabling an early diagnosis; refined intensive care management; and active prevention and treatment measures [3]. However, compared to those in HIV-positive patients, the early clinical symptoms of PCP in HIV-negative immunocompromised patients are not obvious, and delayed diagnosis may eventually lead to death due to sudden respiratory failure. Therefore, the mortality rate due to PCP in HIV-negative immunocompromised patients is higher than that in HIV-positive immunocompromised patients [4, 5]. Consequently, it is particularly important for HIV-negative immunocompromised patients to use antibacterial drugs to prevent the occurrence of PCP.

Trimethoprim-sulfamethoxazole (TMP-SMZ), the preferred preventive application, has been shown to significantly reduce the incidence of PCP in AIDS patients, and its clinical application at home and abroad is relatively mature. However, it is not known whether the prophylactic application of TMP-SMZ can reduce the incidence of PCP and PCP-related mortality in an HIV-negative immunodeficient population or what the safety profile is in that population. There have been few reports on this subject. A retrospective analysis showed that patients with rheumatic diseases who used high-dose glucocorticoids for a long time should also use sulfonamide drugs to prevent the occurrence of PCP and that the PCP incidence and mortality rates were lower in the TMP-SMZ group than in the control group [6]. In addition, some other studies indicated that up to 40% of patients with acute lymphoblastic leukemia or lymphoproliferative disease can develop PCP if they do not take prophylactic medications, and approximately 50% of PCP patients experience acute lung injury. The preventive use of TMP-SMZ in HIV-negative immunodeficient patients was found to significantly reduce the incidence of adverse events (AEs) in these patients compared with the control group (33:1) [7, 8]. Furthermore, patients with immunodeficiency due to causes other than infection with HIV, such as those undergoing allogeneic stem cell transplantation or solid organ transplantation and those with multiple myeloma, are all recommended to take TMP-SMZ to prevent PCP [9–11]. However, there is still a lack of clinical evidence regarding the effectiveness and safety of the prophylactic use of TMP-SMZ in this population.

To address this knowledge gap, we performed a systematic analysis of the efficacy and safety of the prophylactic use of TMP-SMZ to prevent PCP in HIV-negative immunocompromised patients from the perspective of evidence-based medicine to provide a reference for clinical decision-making and promote the rational use of drugs in clinical practice.

## Methods

Our study protocol and analysis were planned in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines, Parts of the methodology can be found in the article by Rui Li et al. [12].

#### Search strategy

A systematic search of the following electronic databases was performed to identify relevant literature published in English before December 24, 2020: PubMed, EMBASE, Web of Science, and the Cochrane Library. The search strategy included the following medical index terms: "pneumocystis pneumonia," "pneumocystis infections," "pneumocystis jirovecii," "prophylactic," "prophylaxis," "prevention," "trimethoprim sulfamethoxazole," "HIV," "AIDS," "acquired immunodeficiency syndrome," and "human immunodeficiency virus". Furthermore, the references in the initially identified articles, including relevant reviews, were manually searched and reviewed to ensure that no relevant study was missed (up to December 24, 2020).

### Study selection

Studies were included if they met the following inclusion criteria: (1) they were randomized controlled trials (RCTs) or case-control studies (CCSs), (2) they contained data regarding the preventative use of TMP-SMZ by HIV-negative immunodeficient patients, and (3) they contained data on the incidence of PCP, the rate of drug discontinuation (discontinuation due to AEs or patient intolerance), the rate of mortality or the rate of AEs. Studies were excluded if they (1) were unrelated to PCP, (2) were duplicate reports, (3) had irrelevant data, or (4) were not RCTs or CCSs.

All retrieved studies were scanned by two reviewers (RL and ZYT), who independently assessed all potentially relevant studies and then reached a consensus. Then, relevant studies were examined to obtain data on the incidence of PCP and the rates of drug discontinuation, mortality and AEs. In the case of disagreement between the two reviewers, the senior coauthor (MY) was consulted, and the disagreement was resolved by consensus.

#### Data extraction

Two reviewers (RL and ZYT) extracted the relevant data from each eligible study independently, and discrepancies were resolved through discussion with the senior coauthor (MY). A predefined form was used to record the following information: (1) first author; (2) year of publication; (3) research type; (4) mean age; (5) proportion of males; (6) follow-up period; (7) numbers of patients in the treatment group and the control group; (8) therapeutic regimen; (9) incidence rate of PCP; (10) rate of drug discontinuation; (11) mortality rate; (12) AEs related to the study medications; (13) overall risk of bias; and (14) quality of the evidence.

### Quality appraisal and assessment of the risk of bias

The quality of each included study was evaluated according to the modified Jadad score [13]. Two reviewers (RL and ZYT) independently assessed the methodological quality of all included studies without blinding regarding the source journal or authorship. Disagreements were resolved by discussion or consultation with the third reviewer (MY) if required. The risk of bias in each included study was also evaluated [14]. Potential publication bias was assessed by the visual inspection of asymmetry in Begg's funnel plots, and Egger's test was then used to provide statistical evidence of funnel plot symmetry (p<0.05 indicating bias and p>0.05 indicating no bias) [15, 16].

### Data analysis and statistical methods

All statistical analyses, except for the publication bias analysis (which was performed using STATA software: version 12.0, StataCorp, College Station, TX, USA), were performed using Review Manager software (RevMan, version 5.1, Oxford, UK; The Cochrane Collaboration, 2008). The heterogeneity among studies was initially assessed graphically by examining forest plots and was subsequently assessed statistically with the chi-square test for homogeneity, and both I<sup>2</sup> statistics and p-values were considered [17]. A p-value <0.1 or I<sup>2</sup>>50% indicated high heterogeneity among studies. An I<sup>2</sup> value between 25% and 50% indicated moderate heterogeneity, and a p-value >0.1 or I<sup>2</sup><25% signified low heterogeneity [14]. Pooled odds ratios (ORs; calculated by adding 0.5 to each cell of the 2×2 table for the trial when one arm of the study contained no events [18]) and 95% confidence intervals (CIs) were also used in the meta-analysis. The meta-analysis was performed using the Mantel-Haenszel fixed-effects model (FEM) or Mantel-Haenszel random-effects model (REM).

## Outcomes analyzed

In this meta-analysis, regarding the outcome measures used to assess efficacy, the incidence of PCP was used as the primary outcome measure, and the rate of drug discontinuation and the rate of mortality were used as the secondary outcome measures. We performed subgroup analyses according to the therapeutic intervention, e.g., TMP-SMZ prevention group vs nonprevention group and TMP-SMZ vs other drugs. The rate of AEs was used as the primary safety outcome measure in this meta-analysis. We compared the TMP-SMZ group with the control group. According to the clinical manifestations of AEs (i.e., rash, hematologic system effects, infection, abnormal liver and renal function and other AEs), we divided the data into five subgroups for analysis and comparison.

# Results

### Search results

The complete search strategy used for each database is described in S1 Table. The detailed process of the literature search and article screening process is described in Fig 1. A total of 2392 records were identified from the English databases, and 0 records were identified through other searches. After excluding duplicates and screening the titles of the studies, 950 articles remained for further review. An additional 918 articles were excluded after reading the abstracts of the potentially relevant articles, and 32 were subjected to a full-text review based on their relevance to the study topic. Finally, a total of 19 articles with 4135 patients were

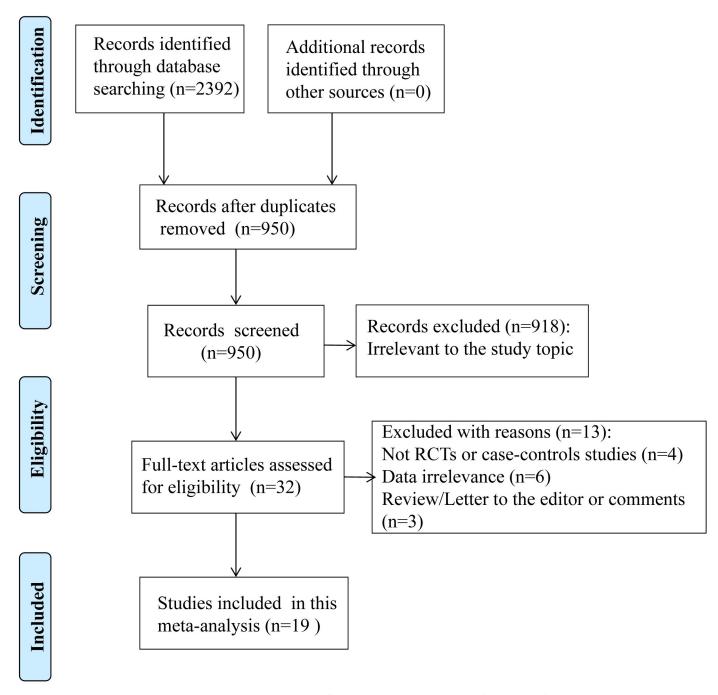


Fig 1. Flow diagram of studies included in the meta-analysis. PRISMA flow diagram showing the number of articles identified and evaluated during the review. https://doi.org/10.1371/journal.pone.0248524.g001

included after the exclusion of 13 articles that were not RCTs or CCSs, did not report the outcomes of interest or were reviews, letters to the editor or comments [6, 10, 19-32].

## Study quality assessment and risk of bias assessment

The quality of each included study was assessed with the modified Jadad score (S2 Table). The results indicated that five studies were of high quality [23–25, 34, 35]. The majority of the

Study	Research Type	Mean Age	Male (%)	Follow-up (months)	Enrolled Patients		Drug Regimen		Assessment Index	Evidence Quality	Risk of Bias
					Т	С	Т	С	-		
TMP-SMZ vs. Nor	prevention		-							,	
Park, 2017 [ <u>6</u> ]	CCS	43.7	22	12	262	1260	TMP-SMZ	Nonprevention	(1)(3)	Moderate	Unclear
Neofytos, 2018 [10]	CCS	65	36	12	41	2801	TMP-SMZ	Nonprevention	-1	Moderate	Unclear
Katsuyama, 2014 [19]	CCS	58.5	19	16	141	561	TMP-SMZ	Nonprevention	(1)(4)	Moderate	Unclear
Ogawa, 2005 [20]	CCS	56.2	52	6	49	75	TMP-SMZ	Nonprevention	(1)(4)	Moderate	Unclear
Okada, 1999 [ <u>21</u> ]	CCS	38	21	2.5	37	47	TMP-SMZ	Nonprevention	(1)(4)	Moderate	Unclear
Colby,1999 [22]	RCT	45.5	28	3	18	16	TMP-SMZ	Nonprevention	(2)(4)	Moderate	Unclear
Vananuvat, 2011 [23]	RCT	34.5	91	3	59	79	TMP-SMZ	Nonprevention	(1)(4)	High	Low
Levinsen, 2011 [24]	RCT	4	43	12	112	250	TMP-SMZ	Nonprevention	(1)(3)	High	Low
Ward, 1993 [25]	RCT	55.4	92.9	1	22	20	TMP-SMZ	Nonprevention	(1)(3)	High	Low
TMP-SMZ vs. Oth	er drugs										
Evans, 2015 [26]	CCS	56.2	58	12	79	79	TMP-SMZ	Dapsone	-4	Moderate	Unclear
Nazir, 2017 [ <u>27</u> ]	CCS	7	55	12	24	34	TMP-SMZ	Dapsone	(1)(2)(4)	Moderate	Unclear
Schmajuk, 2018 [ <u>28]</u>	CCS	43	20	6	129	28	TMP-SMZ	Dapsone	-4	Moderate	Unclear
Redjoul, 2018 [29]	CCS	56	60	12	113	13	TMP-SMZ	Atovaquone	(1)(2)(3)(4)	Moderate	Unclear
Gabardi, 2012 [ <u>30]</u>	CCS	52	64	12	160	25	TMP-SMZ	Atovaquone	(2)(4)	Moderate	Unclear
Zmarlicka, 2015 [ <u>31</u> ]	CCS	50.7	63	12	67	11	TMP-SMZ	Atovaquone	(3)(4)	Moderate	Unclear
Kimura, 2008 [ <u>32</u> ]	CCS	49.4	33	12	27	19	TMP-SMZ	Pentamidine	(1)(4)	Moderate	Unclear
Sangiolo, 2005 [ <u>33]</u>	CCS	44	42	6	310	155	TMP-SMZ	Dapsone	-1	Moderate	Unclear
Kitazawa, 2019 [ <u>34]</u>	CCS	66.5	22	12	55	28	TMP-SMZ	Pentamidine	(2)(4)	High	Low
Hughes, 1977 [35]	RCT	<18	Unclear	Unclear	80	80	TMP-SMZ	Atovaquone	-1	High	Low

#### Table 1. Main characteristics of the studies included in the meta-analysis.

Characteristics of the included RCTs or CCSs comparing TMP-SMZ with a placebo or other drugs for the prevention of PCP in HIV-negative immunocompromised patients.

Abbreviations: CCS: case-control study; RCT: randomized controlled trial; T: treatment (TMP-SMZ); C: control (nonprevention or other drugs); (1): PCP incidence; (2): the rate of drug discontinuation; (3): mortality; (4): AEs.

https://doi.org/10.1371/journal.pone.0248524.t001

studies (fourteen studies) were of moderate quality (Table 1, S2 Table) [6, 10, 19–22, 26–33]. The results of the assessment of the overall risk of bias for each included study indicated that five reports exhibited a low risk of bias [23–25, 34, 35] and that the remaining fourteen reports exhibited an unclear risk of bias (Table 1, S1 Fig) [6, 10, 19–22, 26–33]. The summaries and characteristics of the included articles are presented in Table 1.

### **Efficacy outcomes**

**PCP incidence.** There were thirteen studies concerning the PCP incidence [6, 10, 19–21, 23–25, 27, 29, 32, 33, 35]. We categorized them into two subgroups: TMP-SMZ prevention vs. nonprevention and TMP-SMZ vs. other drugs (that prevent PCP infection), for statistical analysis (Fig 2). A total of 1185 patients were included in the TMP-SMZ group, and 2255 patients

	Experimental		Control		Odds Ratio		Odds Ratio		
Study or Subgroup	Events	Total	<b>Events</b>	Total	Weight	M-H, Random, 95% C	M-H, Random, 95% Cl		
1.1.1 TMP-SMZ vs Non-prevention									
Hughes 1977	0	80	17	80	7.2%	0.02 [0.00, 0.38]	← ■		
Katsuyama 2014	0	141	9	561	7.1%	0.21 [0.01, 3.55]			
Levinsen 2012	0	112	10	250	7.1%	0.10 [0.01, 1.75]	• • • • • • • • • • • • • • • • • • •		
Neofytos 2018	10	12	11	29	10.7%	8.18 [1.50, 44.49]			
Ogawa 2005	2	46	7	78	11.0%	0.46 [0.09, 2.32]			
Okada 1999	0	37	2	47	6.6%	0.24 [0.01, 5.21]			
Park 2017	1	262	29	830	9.7%	0.11 [0.01, 0.78]	· · · · · · · · · · · · · · · · · · ·		
Vananuvat 2011	0	59	6	79	7.0%	0.10 [0.01, 1.72]	• • • • • • • • • • • • • • • • • • •		
Ward 1993	1	22	0	20	6.1%	2.86 [0.11, 74.31]			
Subtotal (95% CI)		771		1974	72.5%	0.32 [0.08, 1.22]			
Total events	14		91						
Heterogeneity: Tau <sup>2</sup> = 2	2.60; Chi² =	= 22.23,	df = 8 (P	= 0.00	5); l <sup>2</sup> = 649	%			
Test for overall effect: 2	Z = 1.67 (P	= 0.09)							
4 4 9 THE 947 94									
1.1.2 TMP-SMZ vs Oth				10	0.00/				
Kimura 2008	0	27	4	19	6.8%	0.06 [0.00, 1.24]			
Nazir 2017	1	24	1	34	7.2%	1.43 [0.09, 24.13]			
Redjoul 2018	0	53	7	73	7.0%	0.08 [0.00, 1.48]			
Sangiolo 2005	0	310	2	155	6.6%	0.10 [0.00, 2.07]			
Subtotal (95% CI)		414		281	27.5%	0.17 [0.04, 0.76]			
Total events	1		14						
Heterogeneity: Tau <sup>2</sup> = (				= 0.39);	l <sup>2</sup> = 1%				
Test for overall effect: 2	Z = 2.33 (P	= 0.02)							
Total (95% CI)		1185		2255	100.0%	0.27 [0.10, 0.77]			
Total events	15		105						
Heterogeneity: $Tau^2 = T$		= 26.18.		P = 0.0	1): l <sup>2</sup> = 54 <sup>o</sup>	%			
Test for overall effect: 2			•	210	.,		0.01 0.1 1 10 100		
Test for subgroup differ	•			P = 0.5	6), l² = 0%	Fa	avours experimental Favours control		
i i i i i i i i i i i i i i i i i i i	Chees. On	0.0		0.0	0 070				

Fig 2. Forest plot of the incidence rate of PCP incidence. The vertical line indicates no difference between the groups. ORs are represented by diamonds, and 95% CIs are depicted by horizontal lines. Squares indicate point estimates, and the size of each square indicates the weight of the given study in the meta-analysis. M-H, Mantel-Haenszel random-effects model.

https://doi.org/10.1371/journal.pone.0248524.g002

were included in the control group (nonprevention or other drugs). The meta-analysis results suggested that there was significant heterogeneity among these studies (p = 0.01,  $I^2 = 54\%$ ). With regard to comparison between the TMP-SMZ group and the control group, the results of the meta-analysis showed a significant difference in the incidence of PCP (OR = 0.27, 95% CI (0.10, 0.77), p = 0.01), and the incidence of PCP in the TMP-SMZ group was significantly lower than that in the control group (Fig 2).

**Rate of drug discontinuation.** Five studies were included in the rate of drug discontinuation analysis [22, 27, 29, 30, 34]. A total of 403 patients were included in the TMP-SMZ group, and 164 patients were included in the control group (other drugs to prevent PCP). The meta-analysis results indicated that there was significant heterogeneity among these studies (p = 0.09,  $I^2 = 50\%$ ). With regard to the comparison between the TMP-SMZ group and the control group, the results of the meta-analysis showed a significant difference in the drug discontinuation rate (OR = 14.31, 95% CI (4.78, 42.91, p<0.00001); the rate of drug discontinuation in the TMP-SMZ group was significantly higher than that in the control group (Fig 3).

	Experimental		Control		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95%	CI M-H, Random, 95% CI
Colby 1999	8	18	0	16	10.5%	26.71 [1.39, 513.09	
Gabardi 2012	41	160	0	25	11.3%	17.71 [1.05, 297.45	j] <b>→</b>
Kitazawa 2019	34	55	8	28	32.5%	4.05 [1.51, 10.83	s]
Nazir 2017	33	57	1	35	17.3%	46.75 [5.98, 365.67	·]
Redjoul 2018	60	113	3	60	28.5%	21.51 [6.36, 72.74	-] — <b>—</b>
Total (95% Cl)		403		164	100.0%	14.31 [4.78, 42.91]	
Total events	176		12				
Heterogeneity: Tau <sup>2</sup> =	0.71; Chi <sup>2</sup> :	= 7.97, c	lf = 4 (P =	= 0.09);	l² = 50%		
Test for overall effect:	Z = 4.75 (P	< 0.000	001)			I	0.01 0.1 1 10 100 Favours experimental Favours control

Fig 3. Forest plot of the rate of drug discontinuation. The vertical line indicates no difference between the groups. ORs are represented by diamonds, and 95% CIs are depicted by horizontal lines. Squares indicate point estimates, and the size of each square indicates the weight of the given study in the meta-analysis. M-H, Mantel-Haenszel random-effects model.

https://doi.org/10.1371/journal.pone.0248524.g003

**Rate of mortality.** In the analysis of the rate of mortality, five studies were included [6, 24, 25, 29, 31], which were categorized into two subgroups (TMP-SMZ vs. nonprevention and TMP-SMZ vs. other drugs to prevent PCP) (Fig 4). A total of 524 patients were included in the TMP-SMZ group, and 1221 patients were included in the control group (nonprevention or other drugs). The meta-analysis results did not show significant heterogeneity among these studies ( $I^2 = 5\%$ , p = 0.38). With regard to the comparison of the TMP-SMZ group and the

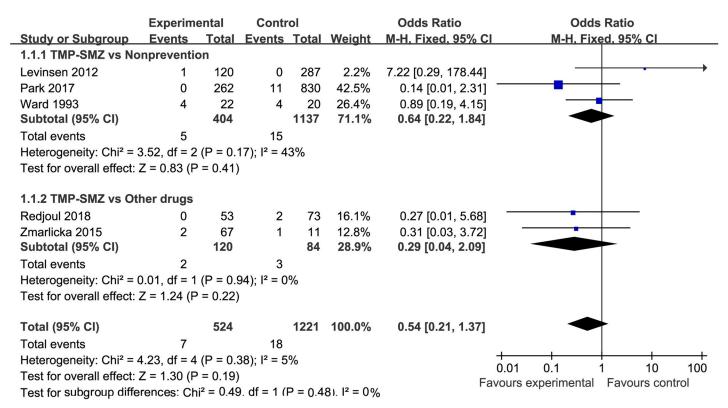


Fig 4. Forest plot of the rate of mortality. The vertical line indicates no difference between the groups. ORs are represented by diamonds, and 95% CIs are depicted by horizontal lines. Squares indicate point estimates, and the size of each square indicates the weight of the given study in the meta-analysis. M-H, Mantel-Haenszel fixed-effects model.

https://doi.org/10.1371/journal.pone.0248524.g004

control group, there was no statistically significant difference in the rate of mortality between the two groups (OR = 0.54, 95% CI (0.21, 1.37), p = 0.19) (Fig 4).

#### Safety outcomes

**Rate of AEs.** The reported AEs in the included studies were rash, hematologic system effects, infections, liver and kidney dysfunction, and other AEs [19–23, 26–32, 34]. We performed subgroup analyses for each category of AEs. The meta-analysis results indicated that there was significant heterogeneity among these studies ( $I^2 = 60\%$ , p<0.00001) and that the rate of AEs was higher in the TMP-SMZ group than in the control group. The results from thirteen studies were separated into five subgroups (OR = 1.92, 95% CI (1.06, 3.47), p = 0.03) (Fig 5).

### Meta-analysis stratified by study design

Most of the included articles were CCSs. Therefore, to verify whether the study design had a significant impact on the outcome measures, we also conducted meta-analyses of the RCTs and CCSs separately for each outcome measure. Because only one RCT reported the rate of drug discontinuation, no comparative analysis was performed. The results of the meta-analysis showed that the incidence of PCP (p = 0.05 vs. p = 0.01) and the incidence of AEs (p = 0.0006 vs. p = 0.03) were statistically significant, while the mortality rate was not statistically significant (p = 0.63 vs. p = 0.19) when comparing RCTs with RCTs and CCSs (the efficacy and safety outcomes above). The summary data of the meta-analysis for each outcome measure stratified by study design are presented in Table 2.

#### Publication bias and sensitivity analyses

In this article, we performed publication bias analyses for the incidence of PCP, rate of drug discontinuation and rate of mortality, and Begg's funnel plots were drawn using STATA 12.0 software. In S2 Fig, each small circle represents a study. Visual inspection of the three funnel plots showed that the majority of small circles were roughly symmetrically distributed above, below and to the left and right on the Begg's funnel diagram, except for the individual small circles distributed outside the funnel diagram, which may be due to the significant heterogeneity among these studies. Egger's test was performed for each outcome, and the results are shown in S3 Table: all p-values were greater than 0.05, which further proved that the above three funnel plots were symmetrical, and the individual small outliers may be due to heterogeneity among these studies rather than to publication bias. Therefore, both the shape of Begg's funnel plot and Egger's test results (all p>0.05) demonstrate that there was no publication bias in the included studies.

The meta-analysis showed significant heterogeneity among the studies for the incidence of PCP ( $I^2 = 54\%$ ) and the rate of drug discontinuation ( $I^2 = 50\%$ ). On the one hand, age, follow-up duration and research type may have had substantial impacts on these two indicators. Thus, we conducted sensitivity analyses with stratification by these two indicators to explore the underlying sources of heterogeneity. The final analysis showed that the primary results were not influenced by age, follow-up duration or research type. The detailed data are shown in <u>S4 Table</u>. On the other hand, the included studies were excluded one by one for each outcome indicator (the incidence of PCP and the rate of drug discontinuation) to analyze the sources of heterogeneity. The results are shown in <u>S5 Table</u>. The heterogeneity among the studies reporting the incidence of PCP was mainly due to the study by Neofytos, and the heterogeneity among the studies reporting the drug discontinuation rate was mainly due to the study by Kitazawa [10, 34].

0	Experim		Contro			Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.1.1 Rash							
Katsuyama 2014	1	82	0	12	2.1%	0.46 [0.02, 11.93]	
Kimura 2008	2	27	0	19	2.3%	3.82 [0.17, 84.29]	
Kitazawa 2019	6	96	0	33	2.5%	4.81 [0.26, 87.77]	
Ogawa 2005	1	49	0	75	2.2%	4.67 [0.19, 116.99]	
Okada 1999	4	26	0	47	2.4%	19.00 [0.98, 368.31]	
Redjoul 2018	4	113	0	73	2.4%	6.04 [0.32, 113.89]	
Schmajuk 2018	4	129	1	28	3.3%	0.86 [0.09, 8.04]	
Vananuvat 2011	4	59	0	79	2.4%	12.89 [0.68, 244.30]	
Subtotal (95% CI)	7	581	U	366	19.6%	3.53 [1.27, 9.80]	
Total events	26	001	1	000	10.070	0.00 [1.21, 0.00]	
				0.04)	12 - 00/		
Heterogeneity: Tau <sup>2</sup> = Test for overall effect:			п – 7 (Р –	0.61),	, I <sup>-</sup> – 0%		
1.1.2 Hematological	system cor	relation					
Colby 1999	4	18	0	16	2.4%	10.24 [0.51, 206.88]	
Gabardi 2012	56	160	8	25	5.4%	1.14 [0.46, 2.82]	_ <del></del>
Katsuyama 2014	2	82	0	12	2.3%	0.78 [0.04, 17.14]	
Kimura 2008	4	27	0	19	2.3%	7.47 [0.38, 147.44]	
Kitazawa 2019	15	96	0	33	2.4%	12.74 [0.74, 219.13]	
				33 29			
Nazir 2017	2	19	11		4.1%	0.19 [0.04, 1.00]	
Ogawa 2005	1	49	0	75	2.2%	4.67 [0.19, 116.99]	
Redjoul 2018	17	113	2	73	4.4%	6.29 [1.41, 28.09]	I
Schmajuk 2018	0	129	4	28	2.4%	0.02 [0.00, 0.40]	
Subtotal (95% CI)		693		310	28.1%	1.56 [0.46, 5.29]	
Total events	101		25				
Heterogeneity: Tau <sup>2</sup> = Test for overall effect:			df = 8 (P	= 0.00	3); l² = 66	%	
		,					
1.1.3 Infection	~	70	05	70	E 70/	0 55 10 00 4 053	
Evans 2015	24	79	35	79	5.7%	0.55 [0.29, 1.05]	
Gabardi 2012	36	160	1	25	3.5%	6.97 [0.91, 53.29]	
Zmarlicka 2015	13	67	4	11	4.6%	0.42 [0.11, 1.66]	
Subtotal (95% CI)		306		115	13.8%	0.90 [0.24, 3.44]	
Total events	73		40				
Heterogeneity: Tau <sup>2</sup> = Test for overall effect:		and there is	f = 2 (P =	0.04);	; I² = 69%		
1.1.4 Abnormal liver	and kidney	/ functio	'n				
Colby 1999	and kidney 1	18	0	16	2.1%	2.83 [0.11, 74.46]	
Katsuyama 2014	5	82	0	12	2.4%	1.77 [0.09, 34.10]	· · · · · · · · · · · · · · · · · · ·
Kimura 2008	2	27	0	19	2.3%	3.82 [0.17, 84.29]	
Kitazawa 2019	6	96	0	33	2.5%	4.81 [0.26, 87.77]	
Ogawa 2005	1	49	0	75	2.2%	4.67 [0.19, 116.99]	· ·
Redjoul 2018	40	113	0	73	2.6%	81.00 [4.89, 1342.10]	
Subtotal (95% CI)		385		228	14.0%	6.16 [1.67, 22.68]	
Total events	55		0				
Heterogeneity: Tau <sup>2</sup> = Test for overall effect:				0.35);	l² = 10%		
	- 2.73 (F	0.000	/				
1.1.5 Other	-		_				
Colby 1999	3	18	0	16	2.3%	7.45 [0.36, 156.28]	
Katsuyama 2014	1	82	0	12	2.1%	0.46 [0.02, 11.93]	
Kitazawa 2019	15	96	5	33	5.0%	1.04 [0.35, 3.11]	
Ogawa 2005	1	49	0	75	2.2%	4.67 [0.19, 116.99]	
Okada 1999	2	26	0	47	2.3%	9.69 [0.45, 209.92]	
Redjoul 2018	1	113	6	73	3.4%	0.10 [0.01, 0.85]	<b>←−−−−</b>
Schmajuk 2018	4	129	5	28	4.6%	0.15 [0.04, 0.59]	———
Vananuvat 2011	6	59	0	79	2.5%	19.32 [1.07, 350.12]	
Subtotal (95% CI)		572	•	363	24.4%	1.20 [0.31, 4.67]	
	33		16		,0		Г
		- 20 77		- 0.00	4), 12 - 00	0/	
Total events	2.26; Chi2 :		at = 7 (P	= 0.00	4); I <sup>2</sup> = 66	70	
Heterogeneity: Tau <sup>2</sup> =	Z = 0.27 (P	r = 0.79					
Heterogeneity: Tau² = Test for overall effect:	Z = 0.27 (P	,		1382	100 በ%	1.92 [1.06.3.47]	
Heterogeneity: Tau² = Test for overall effect: Total (95% CI)	,	2537 2537		1382	100.0%	1.92 [1.06, 3.47]	<b>•</b>
Heterogeneity: Tau <sup>2</sup> = Test for overall effect: Total (95% CI) Total events	288	2537	82				▲
Heterogeneity: Tau² = Test for overall effect: Total (95% CI)	288 1.49; Chi² :	<b>2537</b> = 82.09,					0.02 0.1 1 10

**Fig 5. Forest plot of the rate of AEs.** The vertical line indicates no difference between the groups. ORs are represented by diamonds, and 95% CIs are depicted by horizontal lines. Squares indicate point estimates, and the size of each square indicates the weight of the given study in the meta-analysis. M-H, Mantel-Haenszel random- effects model.

https://doi.org/10.1371/journal.pone.0248524.g005

Outcome Measure	<b>Research</b> Type	<b>Included Studies</b>	OR	95% CI	р
PCP incidence	RCTs	4	0.14	[0.02, 0.97]	0.05*
	CCSs	9	0.36	[0.10, 1.20]	0.10
	RCTs+CCSs	13	0.27	[0.10, 0.77]	0.01
Rate of mortality	RCTs	2	1.39	[0.37, 5.22]	0.63
	CCSs	3	0.20	[0.04, 1.03]	0.05*
	RCTs+CCSs	5	0.54	[0.21, 1.37]	0.19
Rate of AEs	RCTs	2	10.11	[2.69, 38.02]	0.0006
	CCSs	11	1.55	[0.83, 2.88]	0.17
	RCTs+CCSs	13	1.92	[1.06, 3.47]	0.03

Table 2. Summary of the meta-analysis of each outcome measure stratified by study design.

Abbreviations: PCP: pneumocystis pneumonia; AEs: adverse events; RCTs: randomised controlled trials; CCSs: case-control studies; OR: odds ratio; CI: confidence interval.

\*: marginal p-value, there was a significant difference, but further analysis is needed with a larger sample size.

https://doi.org/10.1371/journal.pone.0248524.t002

#### Discussion

The aim of this study was to systematically evaluate the efficacy and safety of the prophylactic application of TMP-SMZ to prevent PCP in HIV-negative immunocompromised patients from the perspective of evidence-based medicine. This study has revealed that TMP-SMZ has a better effect than other drugs or placebo on preventing PCP in HIV-negative immunocompromised people. The prophylactic use of TMP-SMZ compared to a placebo or other drugs can significantly reduce the incidence of PCP in such patients, resulting in a lower rate of mortality. However, the rate of drug discontinuation for TMP-SMZ was significantly higher than that for other drugs. Moreover, the meta-analysis showed that the rate of AEs was also higher in the TMP-SMZ group than in the control group.

Although RCTs generally produce higher-quality evidence than CCSs because they eliminate selection bias, the Cochrane handbook also states that consolidation is acceptable when observational studies have large sample sizes or are of high quality or when RCTs have a small sample size or mediocre quality [14]. Considering the small total number of RCTs and the meta-analysis indicated that there was no significant impact of the study design on the outcome measures. Therefore, in the results section, we present the meta-analysis of the RCTs and CCSs combined.

Regarding efficacy, the incidence of PCP in the TMP-SMZ group was significantly lower than that in the control group, but the rate of discontinuation of TMP-SMZ was significantly higher than that in the control group. While there was no statistically significant difference in the rate of mortality between the two groups, the mortality in the TMP-SMZ group was lower than that in the control group. The Cochrane meta-analysis showed that TMP-SMZ prophylaxis resulted in a 91% reduction in the incidence of PCP and an 83% reduction in the rate of mortality compared to the control group [7], which was similar to the results reported in this study. Although TMP-SMZ is the first choice for the prevention of PCP, TMP-SMZ often has to be replaced with second-line drugs such as dapsone, atomized pentamidine, and atovaquone due to AEs associated with TMP-SMZ [7, 27, 30], and drug intolerance caused by G6PD deficiency and neutropenia. Therefore, the rate of discontinuation of TMP-SMZ was significantly higher than that in the control group. Regarding safety, the meta-analysis showed that the rate of AEs was higher in the TMP-SMZ group than in the control group (23.86% vs. 34.78%; Fig 5), the number of AEs was higher in the TMP-SMZ group than in the control

group, including rash, hematologic system effects (including methemoglobinemia, thrombocytopenia and agranulocytosis, etc.), liver and kidney dysfunction and other AEs (including asthma, nausea, vomiting, hyponatremia, etc.). Most of these AEs can be eliminated or improved after drug withdrawal, but the AEs associated with the blood system (e.g., agranulocytosis) can lead to serious opportunistic infections accompanied by fever in these immunocompromised patients. Hence, appropriate drugs should be selected for inclusion in the regimen according to the type and severity of the patient's disease, and the occurrence of AEs should be closely monitored.

The sensitivity analysis suggested that age, follow-up duration and research type may not be sources of heterogeneity. We excluded the included studies one by one from the main outcome analysis and found that the heterogeneity was primarily due to the clinical trial methodology in the included literature, such as in the studies by Neofytos and Kitazawa [10, 34]. The diagnosis of PCP in the study by Neofytos was entirely based on polymerase chain reaction (PCR), and one of the main limitations of that study was that the PCR cycle threshold could not be used to diagnose PCR-positive cases. Although only 4 cases were diagnosed based on a positive PCR for Pneumocystis in that study, it ultimately affected the diagnosis and prevention of PCP. In the study by Kitazawa, there were significant differences in the baseline characteristics of the population using the prophylaxis. Although the same drugs were used, the heterogeneity may have been caused by the differences in baseline conditions and the characteristics of the patients, and false negative or false positive results might have occurred.

Our study has several limitations. First, the nineteen studies included populations from different countries, so the results could have been affected by region, ethnicity, or language, leading to bias. Second, although the quality of the majority of the studies was moderate, the small overall sample size may have led to false positive or false negative results. Third, most of the studies were CCSs, and the method of randomization was unclear or not mentioned in most studies, indicating the possibility of selection bias. Fourth, the heterogeneity among some of the included studies could have resulted in bias. Fifth, although the results of the sensitivity analyses stratified by age, follow-up duration and research type did not show any significant effects, the large range in follow-up durations (1–12 months) may have resulted in bias.

In conclusion, these findings have important clinical implications for the prevention of PCP in HIV-negative immunocompromised people. However, due to the limitations of the research methodology, these conclusions need to be further verified in large-scale, prospective RCTs to provide reasonable guidance with regard to the prevention of PCP.

## Supporting information

**S1 Checklist. PRISMA checklist for the meta-analysis.** (DOC)

**S1 Table. Search strategy.** (DOC)

S2 Table. Assessment of the quality of evidence in each included study with the modified JADAD score.

(DOC)

**S3 Table.** Egger's tests for the three Begg's funnel plots. (DOC)

**S4 Table.** Sensitivity analyses of the incidence of PCP and rate of drug discontinuation. (DOC)

S5 Table. Heterogeneity analysis for the incidence of PCP and rate of drug discontinuation.

(DOC)

**S1 Fig. Risk of bias summary.** Each risk of bias item for each included study was reviewed. (TIF)

**S2 Fig.** Begg's funnel plots for the meta-analysis of the incidence of PCP (A), rate of drug discontinuation (B) and rate of mortality (C). (TIF)

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