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# Cancer Epidemiology

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# Mortality in patients with cancer and SARS-CoV-2 infection: Results from the Argentinean Network of Hospital-Based Cancer Registries

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ARTICLE INFO	A B S T R A C T				
Keywords: COVID-19 Cancer Mortality RITA Argentina	<i>Background:</i> Cancer is an important risk factor in patients with COVID-19. We aimed to describe the clinical and demographic characteristics associated with mortality in patients with cancer who were infected with SARS-CoV-2. <i>Methods:</i> We conducted a retrospective longitudinal study of 1206 patients with confirmed SARS-CoV-2 infection and cancer, registered in the Argentinean Network of Hospital-Based Cancer Registries (RITA) from March 31, 2020 to January 31, 2021. Demographic and clinical differences between survivors and non-survivors were summarized using descriptive statistics. The primary endpoint was all-cause mortality within 30 days of COVID-19 diagnosis. Risk factors for mortality were identified using logistic regression models. <i>Results:</i> 1206 patients with cancer and confirmed SARS-CoV-2 infection were included, median age was 54 years (interquartile range: 42–65); 793 (65.8%) were female. 1101 (91.3%) had solid tumors and 105(8.7%) had hematological malignancies. The most frequent solid tumor was breast (278, 23.1%), while lymphoma was the main hematological one (59, 4.9%). Cervical cancer was more frequent in survivors, while lung cancer predominated in non-survivors. 275 (22.8%) patients were diagnosis, with a case fatality rate of 15.2% (16/105) for hematologic malignancies and 10.3% (113/1101) for solid tumors. Multivariable regression analysis showed that age 60–79 (odds ratio [OR]: 4.69, 95% confidence interval [CI]: 2.72–9.70), age ≥ 80 (OR: 12.86, 95%CI: 5.08–32.54), time since cancer diagnosis < 1 year (OR: 2.49, 95%CI: 1.57–3.93) and 1–2 years (OR: 2.20, 95%CI: 1.36–3.57), and lung cancer (OR: 4.35, 95%CI: 2.02–9.36) were risk factors for death. <i>Conclusion:</i> Patients with cancer and SARS-CoV-2 infection had a high case-fatality rate. Identified risk factors (older age, recent diagnosis and lung type) could guide prevention strategies aimed at reducing the risk of dying from COVID-19 in cancer ratients				

## 1. Introduction

Since the beginning of the SARS-CoV-2 pandemic in Wuhan in December 2019, many studies have shown that the risk of dying from COVID-19 is not uniform among infected patients, but depends on individual characteristics such as sex, age and presence and type of comorbidities. Cancer, a major cause of mortality worldwide, also increases the risk of COVID-19 [1]. There is evidence showing that individuals living with cancer have an increased risk of contracting SARS-CoV-2 infection and developing the disease [1–3]. This can be

explained because of their systemic immunosuppressive status, caused by many factors such as the disease itself, the anticancer treatments, the increased immune response to infection secondary to immunomodulatory drugs, [4,5] and the frequent visits to hospitals [3]. In addition, cancer patients are often older and have one or more comorbidities [1, 6], which put them at risk of worse serious outcomes, additional intensive care hospitalization [7] and, eventually, an increased risk of death.

Although there is evidence suggesting that all-cause mortality is higher in COVID-19 patients with cancer than in those without cancer

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[2,8], mortality risk differs among cancer patients and results are sometimes contradictory. On the one hand, because of the physiological aging process and the greater prevalence of comorbidities in older patients, cancer patients aged 60 and above may have an increased risk of severe outcomes [5,7,9,10–12]. Nevertheless, results from a retrospective cohort [13] and a large meta-analysis [14] showed no increased risk of death in elderly people. On the other hand, while several studies have shown higher mortality from COVID-19 in patients with lung and hematological cancers [2,3,10,11,13,15–17], data from the Cancer Consortium (CCC19) registry [5] and from a cohort of cancer inpatients of the Brazilian National Cancer Institute (INCA) [18], reported no

increased in mortality risk among leukemia patients. Moreover, an increased risk was seen in cancer patients diagnosed < 1 year before COVID-19 diagnosis [12,19]; for patients under cancer treatment in the past 3–12 months [18], and with recent chemotherapy [5,20]; contradicting Robilotti and colleagues [7] results, which showed no increased mortality risk among cancer patients treated with chemotherapy or surgery within 30 days before COVID-19 diagnosis.

Because of these heterogeneous results coming from studies with several limitations (small sample sizes, short follow-up period, only hospitalized patients), it is difficult to extract any solid conclusions. In Argentina, studies evaluating COVID-19 mortality in cancer patients are



Fig. 1. Study profile.

scarce and involve a small number of patients (75) hospitalized in an oncology institution [21]. Better evidence about COVID-19 infection in cancer patients using local data is a priority for the management of these patients. Therefore, this study aims to describe the clinical and demographic characteristics associated with mortality in a large group of patients with cancer who was infected with SARS-CoV-2.

#### 2. Methods

This retrospective longitudinal study analyzed data from cancer patients with confirmed SARS-CoV-2 infection between March 31, 2020, and January 31, 2021. Data on cancer were obtained from the Argentinean Network of Hospital-Based Cancer Registries (Registro Institutional de Tumores de Argentina - RITA), dependent of the National Institute of Cancer (INC). RITA is a national hospital-based cancer registry system created in 2011 and implemented in several public hospitals from 19 (out of 24) provinces. The registry contains administrative and clinical data from all cancer types in a standardized way, used for administration and improvement of quality of care purposes. Data provided by RITA come from hospitals adhered to the registry, which mainly belong to the public health subsystem. Mostly, these hospitals are reference centers within their jurisdiction. As other institutional cancer registries, RITA collects information of all new patients with cancer who are assisted in selected hospitals, which are not representative of the cases that occur in the entire population. Therefore, it is not possible to estimate population incidence. Furthermore, the frequency of registered tumors may reflect the epidemiological profile of the population assisted at a given institution.

Data on all confirmed COVID-19 positive patients were extracted from the National Health Surveillance System (SNVS 2.0), a registry based on mandatory electronic reporting of notifiable events. Linkages were performed using the National Identity Document (DNI) and surname. Recruitment procedure is shown in Fig. 1. Patients came from 34 hospitals distributed in 19 Argentinian provinces in which RITA was implemented. Contributing institutions independently report data from cancer patients through the online electronic database housed at the INC according to the procedures defined in RITA manuals. The cases had follow-up information, and we were able to obtain information on vital status.

Patients eligible for inclusion were people aged sixteen years or older, diagnosed with cancer between January 1, 2015, and January 31, 2021, and with laboratory-confirmed SARS-CoV-2 infection (n = 1880). Exclusion criteria were that SARS-CoV-2 diagnosis preceded the cancer diagnosis by at least 10 days and that tumor site was poorly defined or unknown. In cases of multiple primary tumors and, considering that our unit of analysis was the person (and not the tumor), we decided to exclude the first tumor diagnosed and keep the most recent tumor information. The number of people with more than one diagnosed tumor in the sample was too small (n = 9) and the study did not have enough power to identify differences between groups with and without a history of multiple primary tumors. The vital status of each patient was reviewed 30 days after the diagnosis of SARS-CoV-2 infection. This information was provided by SNVS 2.0 and was confirmed or completed with RITA records.

All procedures were conducted in accordance with the Helsinki Declaration, the International Ethical Guidelines for Health-related Research Involving Humans (2016), and the Personal Data Protection Law No. 25,326, ant the resolution of the Ministry of Health of the Nation No. 1480/11. The study was approved by the Research Ethics Committee of the National Hospital Prof. Alejandro Posadas, under registration number 600 EUPeS0/22/CEIHP.

# 3. Data collection

We obtained information about demographics, clinical and tumor characteristics, treatments, and vital status from electronic records. We

found four cases of inconsistencies between sex and tumor site that were corrected in the system by contacting the registry. Duplicates tumor or COVID-19 event records were eliminated. The variables analyzed were sex, age (categorized as <40, 40–59, 60–79, and  $\geq$ 80 years old), time since cancer diagnosis (categorized as <1, 1–2, and >2 years), Eastern Cooperative Oncology Group (ECOG) performance status score at cancer diagnosis, cancer stage, cancer type (hematological and solid), and cancer histology. Cancer stage was defined as the early (I-II) or late (III-IV) for solid tumors at cancer diagnosis according to the TNM staging system. The staging for hematological malignancies was not considered because the recommended staging system is different. Histology was classified using the International Classification of Diseases for Oncology, Third Edition (ICD-O3). Tumor sites in the sample included breast (C50), cervix uteri (C53), corpus uteri (C54), colorectal (C18-C20), lung (C34), skin (C44), prostate (C61), urinary tract (C64-C68), testis (C62). Hematological cancer was classified using the morphological codes as follows: lymphoma (959-972), leukemia (980-994), and other hematological (973-976, 998). History of treatment was limited to the first treatment at the institution registering the case.

Variables related to COVID-19 symptoms and clinical manifestations were obtained from the SNVS 2.0 mandatory notification form and included fever, dysgeusia, anosmia, cough, diarrhea, vomiting, pneumonia, respiratory insufficiency, dyspnea, headache, and myalgia.

Data on comorbidities were obtained from RITA records and SNVS 2.0 notification forms. They were grouped as cardiovascular, metabolic (which included diabetes and obesity), respiratory, neurological, chronic kidney, and chronic liver disease, immunosuppression, smoking and pregnancy. The number of comorbidities was grouped as  $(0, 1-2, \geq 3)$ .

The primary endpoint was all-cause mortality, defined as death among patients with cancer and COVID-19 within 30 days of diagnosis of COVID-19.

# 3.1. Statistical analysis

The hypothesis was that there were differences in demographic and clinical characteristics in patients with cancer history and SARS-CoV-2 infection between survivors and non-survivors. Quantitative variables were presented as medians (IQR), and qualitative variables were presented as frequencies and percentages (only available data were calculated). The median test, Fisher's exact test,  $\chi^2$  test, were applied to analyze the differences between groups according to the type of data.

To explore potential risk factors associated with death due to COVID-19 infection in cancer patients, odds ratio (OR) and 95% confidence interval (95%CI) were analyzed using bivariate logistic regression. For the multivariable logistic regression analysis, only those variables with complete data and statistical significance from the bivariate logistic regression analysis (p < 0.05) were included. Interaction terms to evaluate the interaction between covariates on mortality risk were also included. The tests used were all two-sided with less than 5% type I error. The differences between groups were considered to be significant when the p-value was less than 0.05. We used IBM SPSS Statistics 21.0 software for statistical analysis.

# 4. Results

## 4.1. Description of the population

From March 31 of 2020 to January 31 of 2021, 1880 patients with cancer registered in RITA were notified to the SNVS 2.0 with confirmed diagnoses of SARS-CoV-2 infection. Among these patients, 1206 met the criteria for inclusion and were enrolled (Fig. 1).

4.2. Demographic, clinical, and tumor characteristics for the enrolled population are described in Table 1

Of the 1206 patients with cancer included, 129 (10.7%) had died as of March 03, 2021. 793 (65.8%) patients were female. Median age was 54 years (IQR: 42–65) and 470 (38.9%) patients were aged 60 or older, whereas in the non-survivors group that age range represented 64.4%. The case-fatality rate increased with increasing age, reaching 23.5% in men and 30.3% in women aged 80 or older (Fig. 2).

The median time from cancer diagnosis to the onset of COVID-19 symptoms was 873 days (IQR: 1463–395). 275 (22.8%) patients had been diagnosed with cancer within the past year. Solid tumors were more frequent than hematological ones. The most prevalent solid tumor was breast (278, 23.1%) and cervical (228, 18.9%). Among hematological malignancies, lymphoma (59, 4.9%) and leukemia (27, 2.2%) predominated. Cervical cancer was more frequent in survivors, while lung cancer was more common in non-survivors.

Data of COVID-19 symptoms and clinical manifestation at SARS-CoV-2 diagnosis were available for 664 (55.0%) patients, the most common were cough and fever, followed by headache, myalgia, and fatigue. Compared with survivors, non-survivors were more likely to have dyspnea, chest pain, and pneumonia. Key COVID-19 symptoms like anosmia and dysgeusia were more common in survivors.

Comorbidities data were available in 748 (62%) patients. 477 (63.8%) of them had 1 or 2 comorbidities besides cancer, and 94 (12.6%) had three or more. Cardiovascular disease was the most frequent, followed by metabolic and smoking history (Table 1).

Clinical stage data were available for 453 (41.1%) of the 1101 solid tumors reported. Cancer stage I-II was the most frequent in the cohort (51.0%), while stage III-IV predominated in non-survivors (57.5%). ECOG performance at cancer diagnoses was missing for 462 (38.3%) patients. There were no differences in ECOG scores between survivors and non-survivors. Regarding the history of treatment, data were available for 711 cases (58.9%), with surgery (44.0%) and chemotherapy (32.8%) being the more frequent ones. History of chemotherapy seemed to predominate in non-survivors (47.1%).

We also compared the demographic, clinical and tumor characteristics between patients with solid and hematological cancer (Supplementary material). Case fatality rate was higher in patients with hematological malignancies (15.2%) than in those with solid tumors



Fig. 2. Case fatality rate in patients with cancer and COVID-19 by sex and age group.

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Table 1

Demographic	and	clinical	characteristics	in	patients	with	cancer	and	COVID	
infection.										

Variables	All patients	Non-	Survivor	p-value
	(n = 1206)	survivor	(n = 1077)	
		(n = 129)		
Sor				0.004
Mala	41.0		254	0.004
Male	413	59 (45.7%)	354	
De se al s	(34.2%)	70 (54 00/)	(32.9%)	
Female	793	70 (54.3%)	723	
	(65.8%)		(67.1%)	
Age (median, IQR)	54 (42–65)	64 (56–71)	53 (41–64)	< 0.001
Age group, years				< 0.001
< 40	249	9 (7.0%)	240	
	(20.6%)		(22.3%)	
40–59	487	37 (28.7%)	450	
	(40.4%)		(41.8%)	
60–79	420	69 (53.5%)	351	
	(34.8%)		(32.6%)	
$\geq$ 80	50 (4.1%)	14 (10.9%)	36 (3.3%)	
Cancer type				0.115
Solid tumor	1101	113	988	
	(91.3%)	(87.6%)	(91.7%)	
Hematological	105 (8 7%)	16(12.4%)	89 (8 3%)	
Tumor localization	100 (0.770)	10 (12.170)	09 (0.070)	
Solid tumor				
Breest	070	25 (10 40/)	252	0.005
DIEGSL	2/0 (00.10/)	23 (19.4%J	200	0.295
O minut	(23.1%)	0 (7 00/)	(23.5%)	
Cervical	228	9 (7.0%)	219	< 0.001
	(18.9%)		(20.3%)	
Colorectal	97 (8.0%)	11 (8.5%)	86 (8.0%)	0.831
Skin	90 (7.5%)	6 (4.7%)	84 (7.8%)	0.198
Prostate	44 (3.6%)	2 (1.6%)	42 (3.9%)	0.221
Kidney and urinary	37 (3.1%)	7 (5.4%)	30 (2.8%)	0.105
tractor				
Testicle	30 (2.5%)	1 (0.8%)	29 (2.7%)	0.361
Lung	34 (2.8%)	13 (10.1%)	21 (1.9%)	< 0.001
Uterus body	33 (2.7%)	6 (4.7%)	27 (2.5%)	0.155
Other solid†	230	33 (25.6%)	197	0.046
other oone	(19.1%)	00 (201070)	(18.3%)	01010
Hematological	(1).170)		(10.070)	
Lymphoma	50 (4 0%)	8 (6 2%)	51 (6 2%)	0.466
Lymphoma	37(9.9%)	5 (0.2%)	31(0.270)	0.400
Others of DEC	27 (2.2%)	5 (3.9%)	22 (2.0%)	0.199
Others of RES	19 (1.6%)	3 (2.3%)	16 (1.5%)	0.447
Time since cancer				< 0.001
diagnosis				
$\leq 1$ year	275	50 (38.8%)	225	
	(22.8%)		(20.9%)	
1–2 year	248	31 (24.0%)	217	
	(20.6%)		(20.1%)	
> 2 year	683	48 (37.2%)	635	
	(56.6%)		(59.0%)	
Cancer Stage*				0.025
In situ	53/453	2/40	51/413	
	(11.7%)	(5.0%)	(12.3%)	
I-II	231/453	15/40	21/413	
	(51.0%)	(37.5%)	(52.3%)	
III-IV	169/453	23/40	146/413	0 100
	(37 30/)	(57 5%)	(35.4%)	0.190
No. information	(37.3%)	(37.3%)	(33.4%)	
Non information	648	/3(64.6%)	5/5	
	(58.9%)		(58.2%)	
ECOG performance status				0.011
score*				
0	382/744	30/73	352/671	
	(51.3%)	(41.1%)	(52.5%)	
1	307/744	32/73	275/671	
	(41.3%)	(43.8%)	(41.0%)	
2	37/744	5/73	32/671	
	(5.0%)	(6.8%)	(4.8%)	
3	16/744	5/73	11/671	
	(2.2%)	(6.8%)	(1.6%)	
4	2/744	1/73	1/671	
	(0.3%)	(1.4%)	(0.1%)	
Non information	462	56 (49 404)	406	0.207
	(20.20/)	JU (43.4%)	(27 70/)	0.207
ECOC monforms	(38.3%)		(3/./%)	0.071
ECOG performance status				0.0/1
score in last year*				

(continued on next page)

#### Table 1 (continued)

able I (continued)				
Variables	All patients	Non-	Survivor	p-value
	(n = 1206)	survivor	(n = 1077)	-
		(n = 129)		
0	87/162	11/26	76/136	
0	(53.7%)	(42.3%)	(55.9%)	
1	61/162	10/26	51/136	
-	(37.7%)	(38.5%)	(37.5%)	
2	5/162	3/26	2/136	
	(3.1%)	(11.5%)	(3.1%)	
3	8/162	2/26	6/136	
	(4.9%)	(7.7%)	(4.9%)	
4	1/162	0/26 (0%)	1/136	
	(0.6%)		(0.6%)	
Non information	113	24 (48.0%)	89 (39.6%)	0.272
	(41.1%)			
Comorbidities*				
Cardiovascular disease	308/748	46/100	262/648	0.292
	(41.2%)	(46.0%)	(40.4%)	
Metabolic disease	226/748	36/100	190/648	0.176
	(30.1%)	(36.0%)	(29.3%)	
Smoking	216/748	32/100	188/648	0.439
	(28.8%)	(32.0%)	(28.2%)	
Respiratory disease	73/748	15/100	58/648	0.058
	(9.8%)	(15.0%)	(9.0%)	0.000
Immunosuppression	50/748	7/100	43/648	0.892
Donal diasas	(6.7%)	(7.0%)	(6.6%)	0.005
Reliai disease	40/748	(10,00)	30/048	0.085
Nourological disease	(0.1%)	(10.0%)	(5.0%)	0.160
Neurological disease	(2 504)	5/100 (E 004)	(2, 204)	0.100
Liver disease	(2.3%)	(3.0%)	9/648	0.083
LIVEI UISEASC	(1.7%)	4/100	(1.4%)	0.085
Pregnancy	5/748	1/100	4/648	0.513
regulatey	(0.7%)	(1.0%)	(0.6%)	0.010
Non information	458	29 (22.5%)	429	< 0.001
	(38.0%)		(39.8%)	
Number of	<b>( ·</b> · · · <b>· )</b>		<b>(</b> ,	0.003
comorbidities*				
0	177/748	20/100	157/648	
	(23.7%)	(20.0%)	(24.2%)	
1–2	477/7484	57/100	420/648	
	(63.8%)	(57.0%)	(64.8%)	
> 3	94/748	23/100	71/648	
	(12.6%)	(23.0%)	(11.0%)	
COVID-19				
manifestations*				
Cough	376/664	48/78	328/586	0.351
_	(56.6%)	(61.5%)	(56.0%)	
Fever	365/664	46/78	319/586	0.449
	(55.0%)	(59.0%)	(54.4%)	
Headache	276/664	13/78	263/586	< 0.001
Marala in	(41.6%)	(16.7%)	(44.9%)	0.104
Myaigia	231/664	22/78	209/586	0.194
General discomfort	(34.0%) 213/664	(20.2%) 20/78	(33.7%) 193/586	0 1 9 5
General disconnort	(32.1%)	(25.6%)	(32.9%)	0.195
Odvnonhagia	210/664	10/78	200/586	0.001
ouynophugiu	(31.6%)	(12.8%)	(34.1%)	0.001
Anosmia	133/664	2/78	131/586	< 0.001
	(20.0%)	(2.6%)	(22,4%)	
Dyspnea	111/664	33/78	78/586	< 0.001
· ·	(16.7%)	(42.3%)	(13.3%)	
Diarrhea	93/664	10/78	83/586	0.748
	(14.0%)	(12.8%)	(14.2%)	
Dysgeusia	84/664	3/78	81/586	0.013
	(12.7%)	(3.8%)	(13.8%)	
Arthralgia	62/664	8/78	54/586	0.766
	(9.3%)	(10.3%)	(9.2%)	
Chest pain	50/664	11/78	39/586	0.019
	(7.5%)	(14.1%)	(6.7%)	
Respiratory	48/664	16/78	32/586	< 0.001
insufficiency	(7.2%)	(20.5%)	(5.5%)	
Vomiting	44/664	7/78	37/586	0.375
	(6.6%)	(9.0%)	(6.2%)	
Abdominal pain	31/664	5/78	26/586	0.395
	(4.7%)	(6.4%)	(4.4%)	0.007
Pneumonia				0.021

Variables	All patients (n = 1206)	Non- survivor (n = 129)	Survivor (n = 1077)	p-value
Non information	31/664 (4.7%) 543	8/78 (10.3%) 47 (38.2%)	23/586 (3.9%) 496	0.115
	(44.9%)		(45.7%)	
History of treatments*				
Surgery	313/711	28/70	285/641	0.475
	(44.0%)	(40.0%)	(44.5%)	
Chemotherapy	233/711	33/70	200/641	0.007
	(32.8%)	(47.1%)	(31.2%)	
Radiotherapy	91/711	5/70	86/641	0.136
	(12.8%)	(7.1%)	(13.4%)	
Hormonotherapy	52/711	3/70	49/641	0.305
	(7.3%)	(4.3%)	(7.6%)	
Immunotherapy	4 /711	0/70 (0%)	4/641	1.000
	(0.6%)		(0.6%)	
Other treatments	18/711	1/70	17/641	1.000
	(2.5%)	(1.4%)	(2.7%)	
Non information	495	59 (45.7%)	436	0.252
	(41.0%)		(40.5%)	
Treatments within 4				
weeks before				
symptoms onset*				
All	23/700	7/69	16/631	0.005
	(3.3%)	(10.1%)	(2.5%)	
Surgery	14/23	6/7	8/16	0.116
	(60.9%)	(85.7%)	(50.0%)	
Chemotherapy	6/23	1/7	5/16	0.612
	(26.1%)	(14.3%)	(31.3%)	
Radiotherapy	3/23	0/7 (0%)	3/16	0.526
	(13.0%)		(18.8%)	
Treatments within last				
year before symptoms				
onset*				
All	159/700	30/69	129/631	< 0.001
	(22.7%)	(43.5%)	(20.4%)	
Surgery	77/159	13/30	64/129	0.535
0 7	(48.4%)	(43.3%)	(49.6%)	
Chemotherapy	57/159	14/30	43/129	0.170
1.7	(35.8%)	(46.7%)	(33.3%)	
Radiotherapy	13/159	1/30	12/129	0.465
	(8.2%)	(3.3%)	(9.3%)	
Hormonotherapy	8/159	1/30	7/129	1.000
· · · · · · · · · · · · · · · · · · ·	(5.0%)	(3.3%)	(5.4%)	
Other treatments	4/159	1/30	3/129	0.571
	(2.5%)	(3.3%)	(2.3%)	

Data are n (%) or IQR = interquartile rate. RES= reticular endothelial system. † Other solids: thyroid, brain, stomach, ovary, melanoma, bladder, pancreas, unspecified uterus, soft tissue malignant tumor liver and intrahepatic bile ducts, larynx, gallbladder and intrahepatic bile ducts, bone, esophagus, other thoracic organs, anus, nostril, sinuses, middle ear, mouth, parotid, salivary glands, tongue, amygdala, other malignant tumors of the female genitalia, nasopharynx, oropharynx, pharynx and poorly defined lip and mouth, small intestine, other digestives, eye, penis. \*From available data.

(10.3%), although the difference was not statistically significant. Patients with solid tumor were more frequently females (751/1101, 68.2%), while hematological malignancies predominated in males (63/105, 60.0%). Cardiovascular disease history was more frequent in patients with solid tumors than in those with hematological malignancies (294/680, 43.2% vs 14/68, 20.6%), as well as metabolic disease history (214/680, 31.5% vs 12/68, 17.6%).

Conversely, immunosuppression history predominated in hematological malignancies (13/68, 19.1% vs 37/680, 5.4%). From 666 patients with solid tumors and information about treatment, 313 (44.0%) had history of surgery, while 42 (93.3%) of 45 with hematological malignancies had history of chemotherapy. Regarding history of treatment within the last year before the onset of COVID-19 symptoms, 77 (52.0%) of 148 patients with solid tumor had surgery compared with none of the patients with hematological malignancies, and 11 (100%) of patients with hematological malignancies received chemotherapy compared with 46 (31.1%) of 148 with solid tumors.

The association between demographic and clinical factors and death among patients with cancer and COVID-19 is summarized in Table 2. Results show that male sex, age older than 40, type of cancer (lung), and time since cancer diagnosis (less than two years) increased the risk of death; while cancer stage, number of comorbidities, respiratory symptoms and ECOG performance status did not show a statistically significant difference. Statistically significant variables with completed data were included in the multivariable regression analysis (Table 2).

Results showed that age 60–79 (OR: 4.69 [95%CI: 2.72–9.70]) and  $\geq$  80 (OR: 12.86 [95%CI: 5.08–32.54]), time since cancer diagnosis less than 1 year (OR: 2.49 [95%CI: 1.57–3.93]) and 1–2 years (OR: 2.20 [95%CI: 1.36–3.57]), and lung cancer (OR: 4.35 [95%CI: 2.02–9.36]) were associated with increased odds of death. There was no interaction effect between the selected variables (Supplementary material).

# 5. Discussion

To our knowledge, this is the first large report describing the clinical features and risk factors for mortality among patients with cancer diagnosed with SARS-Cov-2 in Argentina during the first year of the pandemic. By integrating data from cancer registries (RITA, INC) with data on COVID-19 from the national surveillance system (SNVS 2.0), we found that patients with cancer are at higher risk of mortality once diagnosed with COVID-19. Furthermore, the risk of dying increased with

age and was greater in patients older than 60 years, when time since cancer diagnosis was more recent (less than two years), and in those with lung cancer, but not with hematological malignancies.

Our findings show an overall case-fatality rate of 10.2%, nearly fourfold higher than the CFR observed in the Argentine general population (2.7%) [22], and similar to those reported by a Brazilian cancer center (12.4%) [10], by the Cancer Consortium database (CCC19) (13%) [5], and by the PRE-COVID-19 study (10%) [23]. In contrast, it is lower than the observed in Hubei (20%) [13], and by the UK Coronavirus Cancer Monitoring Project (UKCCMP) (28%) [24]. Reasons for these findings are unclear, but these two latter studies are based on hospitalized patients, enrolling more severe cases. Patients from the UKCCMP tend to be older, with a greater probability to require hospitalization and having adverse events. Conversely, our case-fatality is considerably higher than that found by Williams [25], who used the infection fatality rate in order to correct for the ascertainment bias (when only more ill patients are tested for the disease). Our data demonstrate that the risk of death in patients with cancer and COVID-19 increases substantially with age. This result is consistent with previous studies showing that age is a key determinant of the prognosis among patients with COVID-19 and cancer [5,7,10–12,24]. An early report from China did not find significant differences in age between survivors and non- survivors, probably because the study already included an elderly population [13]. Interestingly, in comparison with the general population, elderly patients with cancer may not be at increased risk of death when infected with COVID-19 [14], which implies that the presence of cancer may not

## Table 2

Logistic regression models. Risk factors associated with death in cancer patients with SARS-CoV-2 infection (n = 1206).

Variables	Univariable OR (IC95%)	p value	Multivariable OR (IC95%)	p value
Sex				
Female	1 (Ref)			
Male	1.72 (1.19–2.49)	0.004	1.34 (0.89–1.98)	0.152
Age range, years				
$\leq 40$	1 (Ref)			
40–59	2.19 (1.04-4.62)	0.049	2.13 (1.01-4.53)	0.054
60–79	5.24 (2.57-10.70)	< 0.001	4.69 (2.72–9.70)	< 0.001
> 80	10.37 (4.18–25.70)	< 0.001	12.86 (5.08-32.54)	< 0.001
Time since cancer diagnosis				
> 2 years	1 (Ref)			
1–2 years	1.89 (1.17-3.04)	0.009	2.20 (1.36-3.57)	0.001
$\leq 1$ year	2.94 (1.92-4.49)	< 0.001	2.49 (1.57-3.93)	< 0.001
Tumor localization				
Other cancer	1 (Ref)			
Lung	5.63 (2.75–11.55)	< 0.001	4.35 (2.02–9.36)	< 0.001
Cancer stage				
In situ	1 (Ref)			
I-II	1.77 (0.39–7.99)	0.457		
III-IV	4.02 (0.91-17.64)	0.065		
Number of comorbidities				
0	1 (Ref)			
1–2	1.07 (0.62–1.83)	0.819		
> 3	2.54 (1.31-4.93)	0.006		
Symptoms				
Dyspnea	4.78 (2.87–7.94)	< 0.001		
Respiratory insufficiency	4.47 (2.32-8.60)	< 0.001		
Chest pain	2.30 (1.13-4.71)	0.022		
Pneumonia	2.80 (1.21-6.49)	0.017		
Anosmia	0.09 (0.02–0.38)	0.001		
Dysgeusia	0.25 (0.08-0.81)	0.021		
Odynophagia	0.28 (0.14-0.56)	< 0.001		
Headache	0.25 (0.13-0.45)	< 0.001		
ECOG performance status score				
0	1 (Ref)			
1	1.37 (0.81–2.30)	0.243		
2	1.83 (0.66–5.05)	0.241		
3	5.33 (1.73–16.36)	0.003		
4	11.73 (0.72–192.33)	0.084		
Chemotherapy				
No	1(Ref)			
Yes	1.97 (1.19–3.24)	0.008		

further increase the already poor prognosis among elderly people.

As other studies indicate, mortality is significantly affected by the type of tumor [2,3,10,11,13,15–17]. From our analysis, patients with lung cancer have the highest death rates among all patients. Decreased lung function and severe infection in patients may contribute to the worse outcome in this subgroup [2]. As other studies have shown [5,18, 21], we did not find increased risk of death in patients with leukemia or other hematological malignancies. Patients with hematologic cancer, especially leukemia and myeloma, are more often treated with more myelosuppressive therapy and are severely immunocompromised because of underlying disease, so they may potentially be more susceptible to cytokine-mediated inflammation [11]. Despite the fact that the case-fatality rate was higher in patients with hematological malignancies, our study did not have enough power to show statistical significance. As Mehta et al. [26] have shown, results may reflect the prognosis of main cancer types at the country. In Argentina, the annual mortality (ratio of annual deaths/new diagnosis) is the highest for lung cancer (88.6%) and is also high for hematological malignancies (55.8%) [27], suggesting that COVID-19 infection may increase the mortality risk associated with the type of cancer itself. Consistent with other studies [12,18,25], we observed an increased risk of death for patients recently diagnosed with cancer, which supports the fact that patients with active or progressive disease, as they have increased levels of immunosuppression (intrinsic or by cancer treatment), have a worse outcome [5].

The strengths of this study include sample size, study length and database linkage. The use of COVID-19 data from the National Health Surveillance System allowed us to match laboratory-confirmed SARS-CoV-2 cases from our cancer registry over a sustained period of time and obtain a considerable number of cases for analysis. During this period, several waves occurred, which determined the availability of hospital technical and human resources, as well as the ICU ventilation capacity that could have affected the risk of mortality [28]. We used laboratory-confirmed cases to avoid possible confounding by other infections. However, it depends on the sensitivity of the PCR diagnostic test, and it is likely that the number of cancer patients affected with COVID has been underestimated [23]. Also, reported cases varied depending on the availability of testing, and it is expected that patients with mild or asymptomatic disease had not been tested and were, therefore, not included in this study. Other limitations include the incomplete documentation for many variables of interest, which does not allow us to evaluate the effect of treatment, comorbidities, and other measures of severity of disease such as hospitalization or need for intensive care.

Our results have to be interpreted in the health care context of Argentina. Despite Argentina has universal access to health, the health system is fragmented in terms of financing and service delivery into three subsectors: the public subsector (national, provincial, and municipal), the social security subsector, and the private health subsector. Such fragmentation has been proposed as one of the causes of inequalities in health care [29]. In this line, studies have shown that the lack of health insurance is one of the factors explaining inequalities in cancer mortality [30] and cancer screening underuse in Argentina [31, 32]. As RITA is mainly implemented in public hospitals, which assist mostly low-income people with (only) the public coverage, it may result in a selection bias and cases registered could experience worse health outcomes than insured cases. However, in our sample, the proportion of people without formal health coverage was 39%, and with social security 44%, similar to the general population (according to the 2010 Population Census).

During the first year of the pandemic, isolation measures in Argentina were strict, and professional societies discouraged unnecessary visits to hospitals to prevent cancer patients from developing COVID. In turn, changes in cancer care could have caused delays in treatment, and a reduction in early cancer diagnosis, which might result in excess deaths from cancer in the future [33].

## 6. Conclusions

In conclusion, patients with cancer who develop COVID-19 have a high case-fatality rate. Clinicians should pay special attention to those subgroups of cancer patients that showed a major vulnerability, such as those over 60 years of age, with lung cancer, and recently diagnosed. The risk factors identified emphasize the need to develop specific strategies aimed at reducing the risk of dying from COVID-19. Also, imperative efforts should be made to improve the quality of cancer registries.

# CRediT authorship contribution statement

Gisel L. Fattore: Project administration, Conception and study design, Methodology, Interpretation of data, Writing – original draft. Natalia S. Aráoz Olivos: Methodology, Formal analysis and interpretation of data, Writing – review & editing. Jose E. Carrizo Olalla: Interpretation of data, Writing – review & editing, Visualization. Lara Gomez: Conception and study design, Data collection, Writing – review & editing. Agustina Flamenco Marucco: Data collection, Formal analysis and interpretation of data, Writing – review & editing. María Paz Rojas Mena: Data collection, Formal analysis and interpretation of data, Writing – review & editing.

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#### Appendix A. Supplementary material

Supplementary data associated with this article can be found in the online version at doi:10.1016/j.canep.2022.102200.

#### References

- [1] K.A. Lee, W. Ma, D.R. Sikavi, D.A. Drew, L.H. Nguyen, R. Bowyer, et al., COPE consortium. Cancer and risk of COVID-19 through a general community survey, Oncologist 26 (1) (2021) e182–e185, https://doi.org/10.1634/ theoncologist.2020-0572.
- [2] M. Dai, D. Liu, M. Liu, F. Zhou, G. Li, Z. Chen, et al., Patients with cancer appear more vulnerable to SARS-CoV-2: a multicenter study during the COVID-19 outbreak, Cancer Discov. 10 (6) (2020) 783–791, https://doi.org/10.1158/2159-8290.CD-20-0422.
- [3] J. Yu, W. Ouyang, M. Chua, C. Xie, SARS-CoV-2 transmission in patients with cancer at a tertiary care hospital in Wuhan, China, JAMA Oncol. 6 (7) (2020) 1108–1110, https://doi.org/10.1001/jamaoncol.2020.0980.
- [4] C. Blimark, E. Holmberg, U.H. Mellqvist, O. Landgren, M. Björkholm, M. Hultcrantz, et al., Multiple myeloma and infections: a population-based study on 9253 multiple myeloma patients, Haematologica 100 (1) (2015) 107–113, https://doi.org/10.3324/haematol.2014.107714.
- [5] N.M. Kuderer, T.K. Choueiri, D.P. Shah, Y. Shyr, S.M. Rubinstein, D.R. Rivera, et al., Clinical impact of COVID-19 on patients with cancer (CCC19): a cohort study, Lancet 395 (10241) (2020) 1907–1918, https://doi.org/10.1016/S0140-6736(20) 31187-9.
- [6] I.D.L. Cavalcanti, J.C.S. Soares, Impact of COVID-19 on cancer patients: a review, Asia Pac. J. Clin. Oncol. 17 (3) (2021) 186–192, https://doi.org/10.1111/ ajco.13445.
- [7] E.V. Robilotti, N.E. Babady, P.A. Mead, T. Rolling, R. Perez-Johnston, M. Bernardes, et al., Determinants of COVID-19 disease severity in patients with cancer, Nat. Med. 26 (2020) 1218–1223, https://doi.org/10.1038/s41591-020-0979-0.
- [8] K.S. Saini, M. Tagliamento, M. Lambertini, R. McNally, M. Romano, M. Leone, et al., Mortality in patients with cancer and coronavirus disease 2019: a systematic review and pooled analysis of 52 studies, Eur. J. Cancer 139 (2020) 43–50, https:// doi.org/10.1016/j.ejca.2020.08.011.
- [9] C. Bonanad, S. García-Blas, F. Tarazona-Santabalbina, J. Sanchis, V. Bertomeu-González, L. Fácila, et al., The effect of age on mortality in patients with COVID-19: a meta-analysis with 611,583 subjects, J. Am. Med. Dir. Assoc. 21 (7) (2020) 915–918, https://doi.org/10.1016/j.jamda.2020.05.045.

- [10] G.A. Fernandes, D. Feriani, I. França, E. Silva, D.R. Mendonça, E. Silva, P. E. Arantes, J. Canteras, et al., Differences in mortality of cancer patients with COVID-19 in a Brazilian cancer center, Semin. Oncol. 48 (2) (2021) 171–180, https://doi.org/10.1053/j.seminoncol.2021.01.003.
- [11] V. Mehta, S. Goel, R. Kabarriti, D. Cole, M. Goldfinger, A. Acuna-Villaorduna, et al., Case fatality rate of cancer patients with COVID-19 in a New York hospital system, Cancer Discov. 10 (7) (2020) 935–941, https://doi.org/10.1158/2159-8290.CD-20-0516.
- [12] J. Tian, X. Yuan, J. Xiao, Q. Zhong, C. Yang, B. Liu, et al., Clinical characteristics and risk factors associated with COVID-19 disease severity in patients with cancer in Wuhan, China: a multicentre, retrospective, cohort study, Lancet Oncol. 21 (7) (2020) 893–903, https://doi.org/10.1016/S1470-2045(20)30309-0.
- [13] K. Yang, Y. Sheng, C. Huang, Y. Jin, N. Xiong, K. Jiang, et al., Clinical characteristics, outcomes, and risk factors for mortality in patients with cancer and COVID-19 in Hubei, China: a multicentre, retrospective, cohort study, Lancet Oncol. 21 (7) (2020) 904–913, https://doi.org/10.1016/S1470-2045(20)30310-7.
- [14] V.G. Giannakoulis, E. Papoutsi, I.I. Siempos, Effect of cancer on clinical outcomes of patients with COVID-19: a meta-analysis of patient data, JCO Glob. Oncol. 6 (2020) 799–808, https://doi.org/10.1200/GO.20.00225.
- [15] A. Desai, R. Gupta, S. Advani, L. Ouellette, N.M. Kuderer, G.H.A. Li, Mortality in hospitalized patients with cancer and coronavirus disease 2019: a systematic review and meta-analysis of cohort studies, Cancer 127 (9) (2021) 1459–1468, https://doi.org/10.1002/cncr.33386.
- [16] H. Zhang, H. Han, T. He, K.E. Labbe, A.V. Hernandez, H. Chen, et al., Clinical characteristics and outcomes of COVID-19-infected cancer patients: a systematic review and meta-analysis, J. Natl. Cancer Inst. 113 (4) (2021) 371–380, https:// doi.org/10.1093/jnci/djaa168.
- [17] M. Zorzi, S. Guzzinati, F. Avossa, U. Fedeli, A. Calcinotto, M. Rugge, SARS-CoV-2 infection in cancer patients: a population-based study, Front. Oncol. 11 (2021), 730131, https://doi.org/10.3389/fonc.2021.730131.
- [18] A.C. de Melo, L.C.S. Thuler, J.L. da Silva, L.Z. de Albuquerque, A.C. Pecego, L. Rodrigues, et al., Cancer inpatients with COVID-19: a report from the Brazilian National Cancer Institute, PLoS One 15 (10) (2020), e0241261, https://doi.org/ 10.1371/journal.pone.0241261.
- [19] T.B. Johannesen, S. Smeland, S. Aaserud, E.A. Buanes, A. Skog, G. Ursin, A. Helland, COVID-19 in cancer patients, risk factors for disease and adverse outcome, a population-based study from Norway, Front. Oncol., 11, 652535. (https://doi. org/10.3389/fonc.2021.65253).
- [20] L.Y. Lee, J.B. Cazier, V. Angelis, R. Arnold, V. Bisht, N.A. Campton, et al., COVID-19 mortality in patients with cancer on chemotherapy or other anticancer treatments: a prospective cohort study, Lancet 395 (10241) (2020) 1919–1926, https://doi.org/10.1016/S0140-6736(20)31173-9.
- [21] M. Zylberman, F.A. Díaz-Couselo, C. Irrazabal, S. Flagel, R. Custidiano, A. Racciopi, et al., Evolución clínica de pacientes internados con cáncer e infección por COVID-19, Medicina 81 (5) (2021) 695–702.

- [22] Ministerio de Salud Argentina, Boletín Integrado de Vigilancia, No. 528, SE 52, 2020. (https://bancos.salud.gob.ar/recurso/boletin-integrado-de-vigilancia-n528se52-2020), (Accessed 18 February 2022).
- [23] S. Assaad, V. Avrillon, M.L. Fournier, B. Mastroianni, B. Russias, A. Swalduz, et al., High mortality rate in cancer patients with symptoms of COVID-19 with or without detectable SARS- COV-2 on RT-PCR, Eur. J. Cancer 135 (2020) 251–259, https:// doi.org/10.1016/j.ejca.2020.05.028.
- [24] L. Lee, J.B. Cazier, T. Starkey, S. Briggs, R. Arnold, V. Bisht, et al., COVID-19 prevalence and mortality in patients with cancer and the effect of primary tumour subtype and patient demographics: a prospective cohort study, Lancet Oncol. 21 (10) (2020) 1309–1316, https://doi.org/10.1016/S1470-2045(20)30442-3.
- [25] M. Williams, E. Mi, K. Le Calvez, J. Chen, L. Pakzad-Shahabi, S. Dadhania, et al., Estimating the risk of death from COVID-19 in adult cancer patients, Clin. Oncol. 33 (3) (2021) e172–e179, https://doi.org/10.1016/j.clon.2020.10.021.
- [26] V. Mehta, S. Goel, R. Kabarriti, D. Cole, M. Goldfinger, A. Acuna-Villaorduna, et al., Case fatality rate of cancer patients with COVID-19 in a New York hospital system, Cancer Discov. 10 (7) (2020) 935–941, https://doi.org/10.1158/2159-8290.cd-20-0516.
- [27] International Agency for Research of Cancer, The Global Cancer Observatory, Argentina, 2021. (https://gco.iarc.fr/today/data/factsheets/populations/32-a rgentina-fact-sheets.pdf), (Accessed 18 February 2022).
- [28] G. Curigliano, Cancer patients and risk of mortality for COVID-19, Cancer Cell 38 (2) (2020) 161–163, https://doi.org/10.1016/j.ccell.2020.07.006.
- [29] A. Palacios, N. Espinola, C. Rojas-Roque, Need and inequality in the use of health care services in a fragmented and decentralized health system: evidence for Argentina, Int. J. Equity Health 19 (2020) 67, https://doi.org/10.1186/s12939-020-01168-6.
- [30] M.L. Martinez, C.G. Guevel, Social inequalities in cervical cancer mortality in the Autonomous City of Buenos Aires, 1999-2003 and 2004-2006, Salud Colect. 9 (2) (2013) 169–182, https://doi.org/10.1590/s1851-82652013000200004.
- [31] B. Nuche-Berenguer, D. Sakellariou, Socioeconomic determinants of participation in cancer screening in argentina: a cross-sectional study, Front. Public Health 9 (2021), 699108, https://doi.org/10.3389/fpubh.2021.699108.
- [32] S. Arrossi, S. Ramos, M. Paolino, R. Sankaranarayanan, Social inequality in Pap smear coverage: identifying under-users of cervical cancer screening in Argentina, Reprod. Health Matters 16 (32) (2008) 50–58, https://doi.org/10.1016/S0968-8080(08)32410-0.
- [33] A.G. Lai, L. Pasea, A. Banerjee, G. Hall, S. Denaxas, W.H. Chang, et al., Estimated impact of the COVID-19 pandemic on cancer services and excess 1-year mortality in people with cancer and multimorbidity: near real-time data on cancer care, cancer deaths and a population-based cohort study, BMJ Open 10 (11) (2020), e043828, https://doi.org/10.1136/bmjopen-2020-043828.