



The different colorectal tumor risk related to GLP-1 receptor agonists and SGLT2 inhibitors use: a network meta-analysis of 68 randomized controlled trials

Chao-Ming Hung, MD, PhD^{a,b}, Bing-Yan Zeng, MD^{c,d}, Chih-Wei Hsu, MD, PhD^e, Po-Huang Chen, MD, PhD^{f,g}, Cheuk-Kwan Sun, MD, PhD^{h,i}, Andre F. Carvalho, MD, PhD^j, Brendon Stubbs, PhD^{k,l}, Yen-Wen Chen, MD^m, Tien-Yu Chen, MD, PhD^{g,n,o}, Wei-Te Lei, MD, PhD^{p,q}, Jiann-Jy Chen, MD^{m,r}, Yow-Ling Shiue, PhD^{c,s,*}, Kuan-Pin Su, MD, PhD^{t,u,v,*}, Chih-Sung Liang, MD^{w,x,*}, Ping-Tao Tseng, MD, PhD^{c,m,s,y,*}

Background: Recent evidence has raised concerns about potential pro-oncogenic effects associated with glucagon-like peptide-1 (GLP-1) receptor agonists and sodium-glucose cotransporter 2 (SGLT2) inhibitors, particularly regarding gastrointestinal malignancies. Colorectal tumors, the third most diagnosed cancer globally, already show increased incidence in patients with metabolic disorders who typically require these medications. For example, obese subjects had a significantly higher risk of colorectal tumors than healthy subjects. However, existing evidence on this association remains inconsistent. This network meta-analysis (NMA) evaluated the comparative incidence of colorectal tumors associated with specific GLP-1 receptor agonists and SGLT2 inhibitors.

Materials and Methods: We conducted a confirmatory NMA focused specifically on colorectal tumor incidence as an adverse effect, following Cochrane methodological recommendations. We performed a frequentist-based NMA of randomized controlled trials (RCTs) evaluating GLP-1 receptor agonists or SGLT2 inhibitors. Our primary outcome was colorectal tumor incidence, with safety profiles assessed by dropout rates as a secondary outcome.

Results: This NMA encompassing 68 RCTs with 207 200 participants found that only semaglutide was associated with increased incidence of colorectal tumors compared to controls. A dose-stratified analysis revealed that high-dose injectable semaglutide (2.4 mg/week) was the only regimen associated with increased incidence. Furthermore, when focusing on RCTs of obese subjects, the increased colorectal tumor rate related to semaglutide still existed. Neither other GLP-1 receptor agonists nor any SGLT2 inhibitors demonstrated significant associations with colorectal tumor development.

Conclusion: Our study provides the first comprehensive NMA addressing the incidence of colorectal tumors related to individual GLP-1 receptor agonists and SGLT2 inhibitors, suggesting a dose-dependent relationship to semaglutide, particularly in its high-dose injectable form (2.4 mg/week). These findings represent a potential risk signal that requires further validation, given the already elevated baseline colorectal tumor risk in the target population, especially in subjects with obesity. Future research should focus on long-term follow-up studies to better characterize the mechanisms and clinical implications of this semaglutide-specific risk signal.

What this adds to the existing literature: This network meta-analysis, based on 68 randomized controlled trials, suggested that only semaglutide, especially in high-dose injectable form (2.4 mg/week), was associated with increased incidence of colorectal tumors compared to controls, especially in obese patients. Neither other GLP-1 receptor agonists nor any SGLT2 inhibitors demonstrated significant associations with colorectal tumor development.

Learning points: Our study provides evidence regarding the increased incidence of colorectal tumors related to semaglutide in a dose dependent way (2.4 mg/week).

Keywords: colorectal malignancy, colorectal tumor, GLP-1 receptor agonist, network meta-analysis, semaglutide, SGLT2 inhibitor

^aDivision of General Surgery, Department of Surgery, E-Da Cancer Hospital, I-Shou University, Kaohsiung, Taiwan, ^bSchool of Medicine, College of Medicine, I-Shou University, Kaohsiung, Taiwan, ^cInstitute of Biomedical Sciences, National Sun Yat-sen University, Kaohsiung, Taiwan, ^dDepartment of Internal Medicine, E-Da Dachang Hospital, I-Shou University, Kaohsiung, Taiwan, ^eDepartment of Psychiatry, Kaohsiung Chang Gung Memorial Hospital and Chang Gung University College of Medicine, Kaohsiung, Taiwan, ^fDivision of Hematology and Oncology, Department of Internal Medicine, Tri-Service General Hospital; School of Medicine, National Defense Medical University, Taipei, Taiwan, ^gSchool of Medicine, National Defense Medical Center, Taipei, Taiwan, ^hDepartment of Emergency Medicine, E-Da Dachang Hospital, I-Shou University, Kaohsiung, Taiwan, ⁱSchool of Medicine for International Students, I-Shou University, Kaohsiung, Taiwan, ^jInnovation in Mental and Physical Health and Clinical Treatment (IMPACT) Strategic Research Centre, School of Medicine, Barwon Health, Deakin University, Geelong, VIC, Australia, ^kDepartment of Psychological Medicine, Institute of Psychiatry, Psychology and Neuroscience, King's College London, London, UK, ^lDepartment

of Sport, University of Vienna, Vienna, Austria, ^mProspect Clinic for Otorhinolaryngology & Neurology, Kaohsiung, Taiwan, ⁿDepartment of Psychiatry, Tri-Service General Hospital, Taipei, Taiwan, ^oDepartment of Psychiatry, College of Medicine, National Defense Medical University, Taipei, Taiwan, ^pSection of Immunology, Rheumatology, and Allergy Department of Pediatrics, Hsinchu Municipal MacKay Children's Hospital, Hsinchu, Taiwan, ^qCenter for Molecular and Clinical Immunology, Chang Gung University, Taoyuan, Taiwan, ^rDepartment of Otorhinolaryngology, E-Da Cancer Hospital, I-Shou University, Kaohsiung, Taiwan, ^sInstitute of Precision Medicine, National Sun Yat-sen University, Kaohsiung, Taiwan, ^tMind-Body Interface Research Center (MBI-Lab), China Medical University Hospital, Taichung, Taiwan, ^uCollege of Medicine, China Medical University, Taichung, Taiwan, ^vAn-Nan Hospital, China Medical University, Tainan, Taiwan, ^wDepartment of Psychiatry, Beitou Branch, Tri-Service General Hospital; School of Medicine, National Defense Medical Center, Taipei, Taiwan, ^xDepartment of Psychiatry, National Defense Medical Center, Taipei, Taiwan and ^yDepartment of Psychology, College of Medical and Health Science, Asia University, Taichung, Taiwan

Introduction

Glucagon-like peptide-1 (GLP-1) receptor agonists and sodium-glucose cotransporter 2 (SGLT2) inhibitors have emerged as novel glucose-lowering agents, featuring mechanisms of action distinct from those of conventional treatments^[1]. However, recently accumulating evidence raised concerns about their potential association with various tumor risks^[2,3]. This pro-carcinogenic concerns resulted from their anti-apoptotic effects via activation of protein kinase C-epsilon (PKCε) and extracellular signal-regulated kinase 1/2 (ERK1/2) pathways^[4], which played important roles in cell proliferation and anti-apoptosis^[5,6]. GLP-1 receptor signaling has been implicated in regulating mucosal growth in the small intestine and colon, indicating its potential involvement in intestinal tumor development^[7]. Echoing these concerns, several traditional pairwise meta-analyses were conducted to address the association between various tumors and prescription of such medications, such as breast tumor^[8], thyroid tumor^[9], and prostate tumor^[10].

On the other hand, seldom reports directly addressed the relationship between gastrointestinal tract tumor risks and GLP-1 receptor agonists and SGLT2 inhibitors use, which tumors accounted for 26% of the global cancer incidence and 35% of all cancer-related deaths^[11]. Wang *et al*, by retrospective database search without randomized procedure to minimize the potential difference across experimental and control groups, noticed that the overall GLP-1 receptor agonists were associated with a lower risk of colorectal tumors, especially in obese participants^[12,13]. Zhu *et al*, through immunological analyses of GLP-1 signaling-related genes in colorectal cancer, suggested that semaglutide might have protective roles in colorectal tumors^[14]. Lin *et al* had addressed a phenomenon regarding the different cancer risk profiles related to individual GLP-1 receptor agonists in their review article^[15], which might reflect the hypothesis of “biased agonism and polymorphic variation” regarding the different activation of downstream signal transduction pathways in response to different extrinsic GLP-1 receptor agonists^[16]. Furthermore, Flausino *et al* demonstrated a beneficial effect of SGLT2 inhibitors on the survival outcomes in gastrointestinal cancer patients undergoing chemotherapy/radiotherapy in their retrospective cohort study^[17]. When a clinical trial is unable to reach a conclusion due to limitations in experimental design, meta-analyses can provide an alternative and important insight. By deriving data from numerous large-scale clinical studies or databases, they would be more reflective of clinical experience. Many scholars believe that meta-analyses based on large databases or numerous clinical trials can serve as a benchmark for guiding future research directions^[18,19]. In 2024,

HIGHLIGHTS

- Recent evidence has raised concerns about potential pro-oncogenic effects associated with glucagon-like peptide-1 (GLP-1) receptor agonists and sodium-glucose cotransporter 2 (SGLT2) inhibitors, particularly regarding colorectal tumors, which were the third most commonly diagnosed cancer worldwide.
- This network meta-analysis of 68 randomized controlled trials (RCTs), including 207 200 participants, found that only semaglutide was associated with a significantly increased risk of colorectal tumors compared to controls.
- Dose-stratified analysis revealed that high-dose injectable semaglutide (2.4 mg/week) carried the highest risk.
- Furthermore, when focusing on RCTs of obese subjects, this increased colorectal tumor risk related to semaglutide still existed.
- Importantly, no other GLP-1 receptor agonists or SGLT2 inhibitors demonstrated significant associations with colorectal tumor development, suggesting this effect may be specific to semaglutide rather than a class effect.

Figlioli *et al* published a traditional pairwise meta-analysis, which suggested that GLP-1 receptor agonists did not alter overall gastrointestinal tumor risk but did not provide detailed information regarding site-specific tumors^[20]. Similarly, Xu *et al* conducted another traditional pairwise meta-analysis addressing the association between SGLT2 inhibitors and several types of cancers and found that SGLT2 inhibitors did not alter gastrointestinal cancer risk^[21]. However, as we know, the gastrointestinal tumor includes various kinds of tumors with different risk factors^[22] and epidemiologic data^[23]. Among those gastrointestinal tumors, colorectal tumors were the third most diagnosed cancer in the world, with increasing rates in emerging economies^[24]. The rationale of importance of exploring the association between colorectal tumors and GLP-1 receptor agonists and SGLT2 inhibitors prescription relied on two reasons. First, given the fact that subjects who needed such medications already had increased risks of colorectal tumors at baseline^[25], any medications with potential for pro-oncogenesis should be used with caution. Finally, the anti-apoptotic effects related to such medications were specifically noticed in colorectal tumors^[26,27]. Therefore, the exploration of association between different risks of colorectal tumors and such medications was clinically relevant and important. However, to date, there have been few meta-analyses that addressed this issue. Specifically, Bushi and colleagues, by pooling different medications into one whole group, suggested that GLP-1 receptor

Chao-Ming Hung, Bing-Yan Zeng, and Chih-Wei Hsu contributed equally as first authors.

Sponsorships or competing interests that may be relevant to content are disclosed at the end of this article.

*Corresponding author. Address: Institute of Precision Medicine, National Sun Yat-sen University, 70 Lienhai Rd., Kaohsiung 80424, Taiwan. Tel.: +886-7-525-2000 ext. 5818; +886-915-515-971. E-mail: shirley@imst.nsysu.edu.tw (Y.-L. Shiu); Department of Psychiatry, China Medical University Hospital, No. 2 Yuh-Der Road, North District, Taichung 404332, Taiwan. Tel.: +886-422062121 ext. 4126. E-mail: cobolsu@gmail.com (K.-P. Su); Department of Psychiatry, Beitou Branch, Tri-Service General Hospital; School of Medicine, National Defense Medical Center, No. 60, Xinmin Road, Beitou District, Taipei 112003, Taiwan. Tel.: +886-2-2895-9808. E-mail: lcsyfw@gmail.com (C.-S. Liang); Prospect Clinic for Otorhinolaryngology & Neurology, No. 252, Nanxin Road, Nanzi District, Kaohsiung City 81166, Taiwan. Tel.: +886-7-3524100. E-mail: ducktseng@gmail.com (P.-T. Tseng).

Copyright © 2025 The Author(s). Published by Wolters Kluwer Health, Inc. This is an open access article distributed under the Creative Commons Attribution License 4.0 (CCBY), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

International Journal of Surgery (2026) 112:443–459

Received 14 April 2025; Accepted 27 August 2025

Supplemental Digital Content is available for this article. Direct URL citations are provided in the HTML and PDF versions of this article on the journal's website, www.ijournal.com/international-journal-of-surgery.

Published online 11 September 2025

<http://dx.doi.org/10.1097/JS9.0000000000003450>

agonists overall did not increase the risk of colorectal tumors^[28]. However, given the fact that different medications would have different pharmacokinetic properties, the procedure of pooling different medications into one group would ignore the potential different tumor risk across individual medications.

To our knowledge, there was no network meta-analysis (NMA) in this field. Actually, a well-designed NMA, which allows for direct comparisons among different medications, could enhance the ability to make multiple treatment efficacy comparisons and assess the potential superiority of specific interventions at various dosages^[29]. This approach offers a more detailed and evidence-based framework for guiding future clinical practices. Our current study follows the rationale of previously published studies on GLP-1 receptor agonists and SGLT2 inhibitors in the incidence of neurodegenerative disease^[30] and hearing loss^[31]. Specifically, we conducted the current NMA to evaluate the potentially different risk of colorectal tumors related to GLP-1 receptor agonists and SGLT2 inhibitors prescription in subjects without baseline colorectal tumors.

Methods

In this NMA, we used a confirmatory approach to specifically focus on particular adverse effects of interest (i.e., incidence of colorectal tumor here) based on the Cochrane recommendation^[32]. This NMA adhered to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines^[33] (Supplementary Digital Content eTable 1, available at: <http://links.lww.com/JS9/F134>) and A Measurement Tool to Assess Systematic Reviews 2 guidelines^[34] (AMSTAR-2 checklist in Supplementary Digital Content Materials, available at: <http://links.lww.com/JS9/F74>). The current manuscript followed the TITAN Guidelines 2025 and the authors declared that there was no AI applied in the current work^[35] (Supplementary Digital Content eTable 9, available at: <http://links.lww.com/JS9/F134>).

Database searches and study identification

We conducted comprehensive searches across multiple databases, including PubMed, Embase, ClinicalKey, Cochrane CENTRAL, ProQuest, ScienceDirect, Web of Science, and ClinicalTrials.gov (Supplementary Digital Content eTable 2, available at: <http://links.lww.com/JS9/F73>), covering publications up to 10 March 2025. Two independent researchers carried out these searches, screened titles and abstracts, and resolved disagreements through discussion. We also manually reviewed reference lists from relevant review articles and meta-analyses for additional studies^[20,28,36,37]. No language restrictions were applied.

Inclusion and exclusion criteria

This NMA was guided by the PICOS framework (Population, Intervention, Comparison, Outcome, Study Design): **Population:** Individuals without documented colorectal tumor at baseline; **Intervention:** Use of GLP-1 receptor agonists or SGLT2 inhibitors at any dose; **Comparison:** Placebo, standard care, or active comparator; **Outcome:** Incidence of colorectal tumor; **Study design:** Randomized controlled trials (RCTs).

RCTs exclusively enrolling participants with documented baseline colorectal tumors were excluded to maintain focus on

causative or prophylactic evaluation. Furthermore, to minimize selective reporting bias, only RCTs with systematic adverse event assessments or targeted outcome evaluations were included^[38]. Inclusion criteria were: (1) RCTs with participants free from documented colorectal tumor at baseline; (2) RCTs evaluating GLP-1 receptor agonists or SGLT2 inhibitors; (3) studies involving human subjects; and (4) RCTs with systematic adverse event monitoring or direct evaluation of target outcomes.

Exclusion criteria included: (1) non-RCT studies; (2) RCTs specifically enrolling participants with documented baseline colorectal tumors; (3) RCTs lacking direct comparisons of GLP-1 receptor agonists or SGLT2 inhibitors; (4) RCTs lacking information on target outcomes; and (5) animal studies.

Methodological quality appraisal

Two independent reviewers assessed the risk of bias using the Cochrane Risk of Bias Tool 1.0^[39], achieving an inter-rater reliability score of 0.86. Discrepancies were resolved by consulting a third reviewer.

Outcome definition

We recognized the fact that most RCTs were not specifically designed for detecting the incidence of colorectal tumors, so that they might not necessarily differentiate benign and malignant tumors clearly. Furthermore, the diagnostic criteria and cut-off points of benign tumors versus malignant cancers varied across countries and years. Finally, the colorectal tumor had a high percentage of transforming from a benign one into a malignant one depending on various clinical factors^[40,41]. Therefore, we did not separate benign tumors and malignant cancers because of the limitation. Rather, we counted them together. Safety was evaluated through dropout rates, indicating study withdrawals for any reason.

Data extraction, management, and conversion

Two independent authors extracted relevant data, including demographics, study design, treatment details, primary outcomes, and safety data. The corresponding authors were contacted for missing information. Data extraction followed protocols from the Cochrane Handbook for Systematic Reviews of Interventions and other medical literature standards^[42].

Statistical analyses

A random-effects model was applied for analyses involving multiple treatment arms^[43], using MetaInsight (version 4.0.2, Complex Reviews Support Unit, National Institute for Health Research, London, UK) in a frequentist framework. MetaInsight uses the *netmeta* package in R for NMAs^[44].

For categorical data, single-zero-event studies received continuity correction, while studies with zero events in both treatment and control arms were excluded to avoid bias^[45,46]. Forest plots were created to present odds ratios (ORs) with 95% confidence intervals (95% CIs) for effect size calculation^[47]. We then generated treatment rankings and effect sizes for direct and indirect comparisons, tabulated accordingly. The “node-splitting” method was applied to test consistency between direct and indirect evidence, a method particularly beneficial in NMA

when trial-level data are available^[44,48]. A two-tailed *P*-value less than 0.05 indicated statistical significance.

Heterogeneity, inconsistency, and sensitivity analyses

We evaluated heterogeneity among the included studies with the tau value, I^2 , and corresponding *P* value. The inconsistencies were explored by the loop-specific approach, node-splitting, and design-by-treatment model^[49]. Furthermore, we used funnel plots and Egger's regression to evaluate potentially publication bias. Because MetaInsight did not provide heterogeneity test, inconsistency test of loop-specific approach and design-by-treatment model, funnel plots, and Egger's regression, we did these analyses with the *mvmeta* command in STATA (version 18.0)^[50]. To assess the robustness of our findings, we conducted a sensitivity analysis of subgroup analysis by subgrouping treatments according to different dosages of GLP-1 receptor agonists or SGLT2 inhibitors. The dose definitions followed original RCT classifications^[51–116]:

- Canagliflozin: Low: 100 mg; High: 300 mg.
- Dapagliflozin: Low: 2.5 mg; Medium: 5 mg; High: 10 mg.
- Dulaglutide: Low: <1.5 mg; Medium: 1.5 mg; High: >1.5 mg.
- Efpeglenatide: Low: 2 mg; Medium: 4 mg; High: 6 mg.
- Empagliflozin: Low: 1–10 mg; High: 25–50 mg.
- Ertugliflozin: Low: 5 mg; High: 15 mg.
- Injectable semaglutide: Low: 0.05–0.5 mg; Medium: 1.0 mg; High: 2.4 mg.
- Tirzepatide: Low 5 mg; Medium: 10 mg; High: 15 mg.

Finally, regarding the aforementioned different effects of GLP-1 receptor agonists on colorectal tumor risk in subjects with obesity^[12,13], we arranged a sensitivity analysis of subgroup analysis focusing on RCTs with obese subjects. Additionally, in consideration of the potential effect of medication exposure time on the incidence of colorectal tumor risks, we arranged a sensitivity analysis of subgroup analysis of RCTs with (1) less than 2-year treatment and (2) at least 2-year treatment. Finally, we evaluated the overall quality of evidence in our NMA with the GRADE^[117].

General declaration

This study complies with the principles outlined in the Declaration of Helsinki.

Results

Eligibility of the studies

Figure 1 illustrates the flowchart summarizing the literature search and screening process for this NMA. After excluding 82 articles for various reasons (Supplementary Digital Content eTable 3, available at: <http://links.lww.com/JS9/F134>), a total of 64 articles encompassing 68 RCTs were included in the analysis (Supplementary Digital Content eTable 4, available at: <http://links.lww.com/JS9/F134>)^[51–116]. The selected studies involved 207200 participants (mean age = 63.0 years, range:

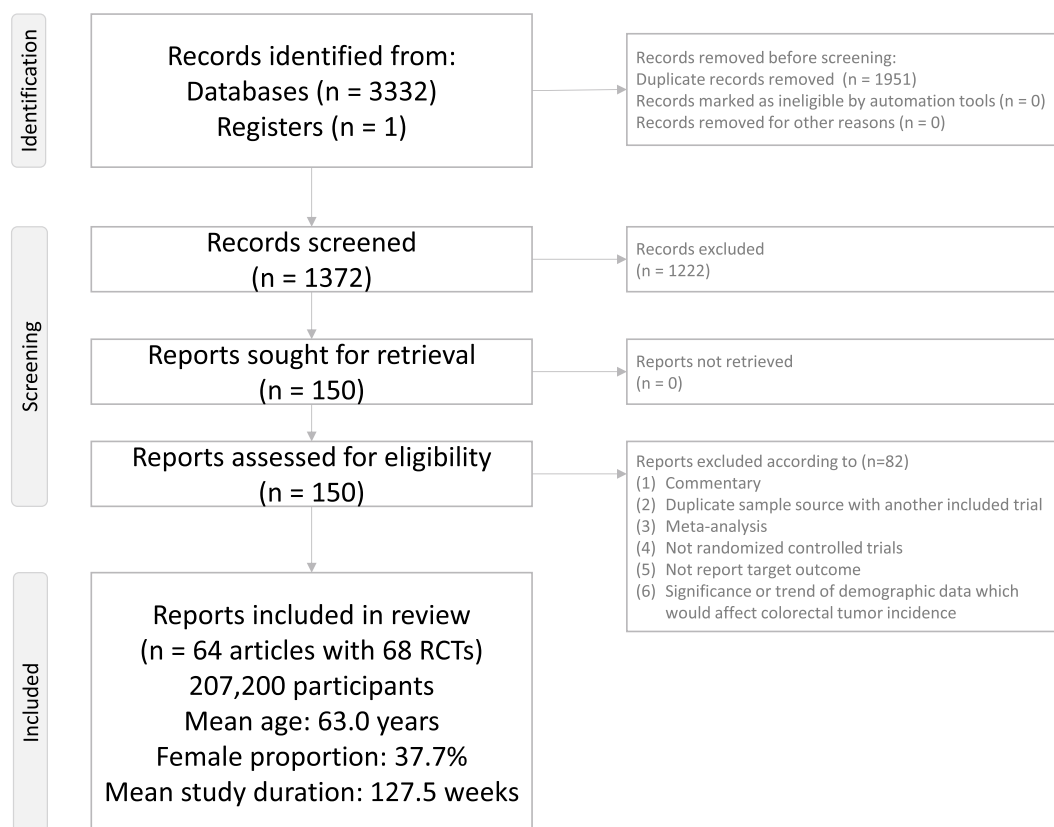


Figure 1. PRISMA2020 flowchart of current network meta-analysis.

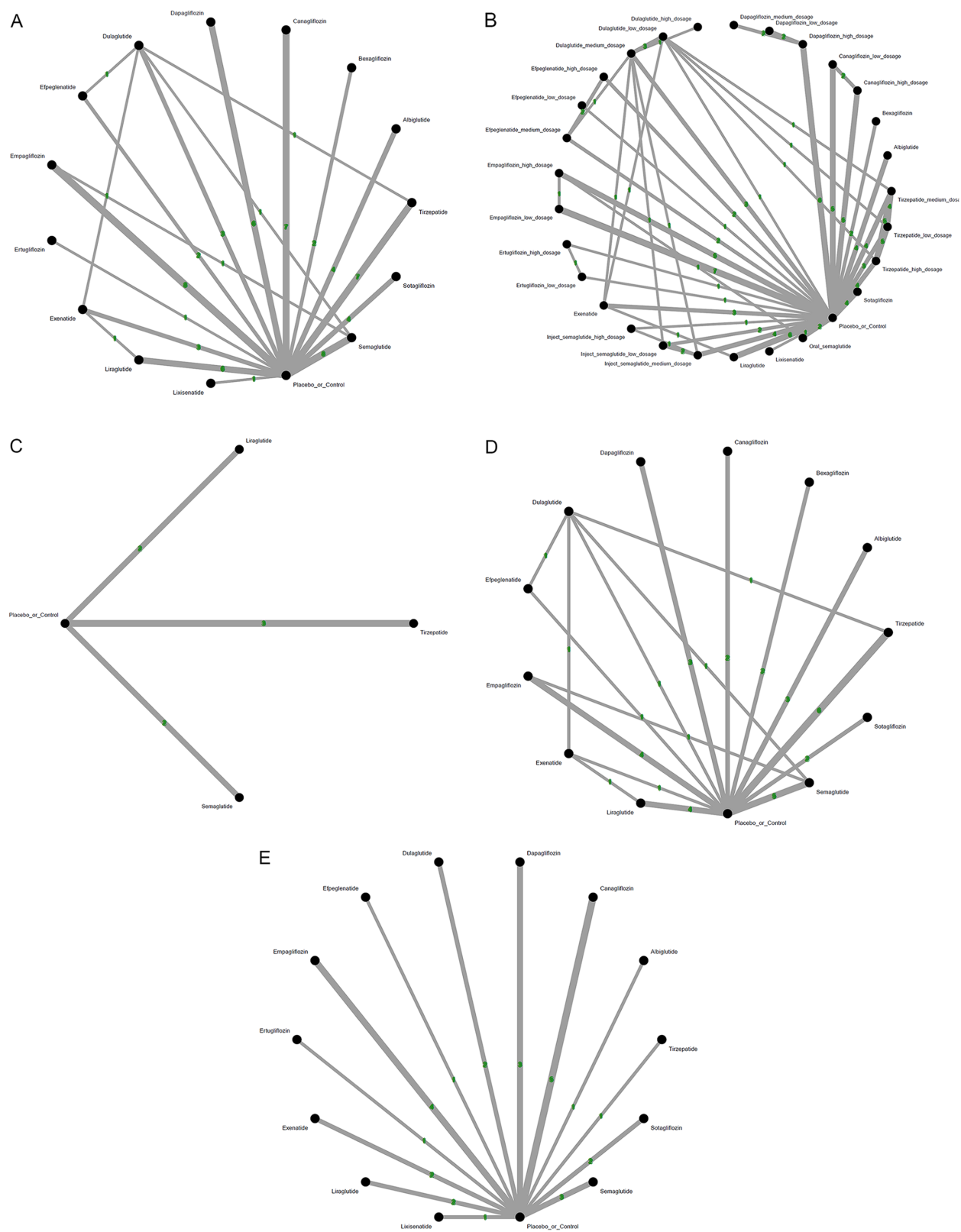


Figure 2. Network structure of the primary outcome: (A) overall colorectal tumor events, (B) colorectal tumor events in aspect of various dosage subgroups, (C) colorectal tumor events in RCTs of subjects with obesity, (D) colorectal tumor events in RCTs of treatment duration less than 2 years, and (E) colorectal tumor events in RCTs of treatment duration at least 2 years. The overall structure of the network meta-analysis. The lines between nodes represent direct comparisons from various trials, with the numbers over the lines indicating the number of trials providing these comparisons for each specific treatment. The thickness of the lines corresponds to the number of trials linked to the network.

45.1–71.9 years; mean female proportion = 37.7%, range: 20.8–78.5%). The average study duration was 127.5 weeks (range: 24–281 weeks). The GLP-1 receptor agonists investigated included tirzepatide, efpeglenatide, liraglutide, albiglutide, dulaglutide, exenatide, semaglutide, and lixisenatide. The SGLT2 inhibitors investigated included bexagliflozin, canagliflozin, empagliflozin, ertugliflozin, dapagliflozin, and sotagliflozin.

Primary outcome: overall colorectal tumor events

The main results of the primary outcome revealed that semaglutide was the only treatment associated with a statistically significant increased incidence of colorectal tumor [OR = 1.49, 95% CIs = 1.01–2.19, absolute risk difference (ARD) = 0.13%, number needed to harm (NNH) = 772] in comparison with the controls. On the other hand, bexagliflozin was associated with significantly less incidence of colorectal tumor than semaglutide (OR = 0.14, 95% CIs = 0.02–0.91) and sotagliflozin (OR = 0.09, 95% CIs = 0.01–0.73), respectively (Fig. 2A, Fig. 3A, Supplementary Digital Content eFigure 3A, available at: <http://links.lww.com/JS9/F73>, and Table 1).

Sensitivity analysis of subgroup analysis of colorectal tumor events in aspect of various dosage subgroups

In our analysis of subgroups according to various dosages, high-dose injectable semaglutide (2.4 mg/week) was the only regimen associated with a significantly increased incidence of colorectal tumors (OR = 1.61, 95% CIs = 1.02–2.54, ARD = 0.23%, NNH = 426) compared to control. On the other hand, bexagliflozin was associated with significantly less incidence of colorectal tumor than high-dose injectable semaglutide (2.4 mg/week) (OR = 0.13, 95% CIs = 0.02–0.85), sotagliflozin (OR = 0.09, 95% CIs = 0.01–0.73), and oral semaglutide (OR = 0.07, 95% CIs = 0.01–0.69), respectively (Fig. 2B, Fig. 3B, Supplementary Digital Content eFigure 3B, available at: <http://links.lww.com/JS9/F134>, and Table 2).

Sensitivity analysis of subgroup analysis of colorectal tumor events in RCTs of obese subjects

In our analysis of subgroups focusing on RCTs on obese subjects, only semaglutide was associated with a significantly increased incidence of colorectal tumors (OR = 1.66, 95% CIs = 1.05–2.61, ARD = 0.19%, NNH = 513) compared to control (Fig. 2C, Fig. 3C, and Table 3).

Sensitivity analysis of subgroup analysis of colorectal tumor events in RCTs of “less than 2-year treatment” and “at least 2-year treatment”

In the subgroup analysis of “less than 2-year treatment,” none of the investigated medications was associated with a significantly different risk of colorectal tumor in comparison with controls (Fig. 2D, Fig. 3D, and Table 4). Rather, in subgroup analysis of “at least 2-year treatment,” only semaglutide achieved a trend of increased risk of colorectal tumor (OR = 1.50, 95% CIs = 0.99–2.28) compared to control (Fig. 2E, Fig. 3E, and Table 5).

Safety profile: dropout rate

The tirzepatide (OR = 0.61, 95% CIs = 0.50–0.74), canagliflozin (OR = 0.73, 95% CIs = 0.60–0.89), dulaglutide (OR = 0.77, 95% CIs = 0.63–0.95), liraglutide (OR = 0.80, 95% CIs = 0.68–0.95), semaglutide (OR = 0.81, 95% CIs = 0.68–0.97), and empagliflozin (OR = 0.85, 95% CIs = 0.74–0.97) were associated with significantly less

dropout rates than the control group did. Among these interventions, tirzepatide ranked the best (Supplementary Digital Content eFigure 1, available at: <http://links.lww.com/JS9/F134>, Supplementary Digital Content eFigure 2, available at: <http://links.lww.com/JS9/F134>, Supplementary Digital Content eFigure 3C, available at: <http://links.lww.com/JS9/F134>, and Supplementary Digital Content eTable 5, available at: <http://links.lww.com/JS9/F134>).

Publication bias, risk of bias, inconsistency, heterogeneity, and quality of evidence

Funnel plots of publication bias across the included studies revealed a general symmetry, and no significant publication bias was detected among the articles included in NMA by using the Egger test (Supplementary Digital Content eFigures 4A–4L, available at: <http://links.lww.com/JS9/F134>). We identified that 77.1% (367/476 items), 15.5% (74/476 items), and 7.4% (35/476 items) of the included studies had low, unclear, and high risks of bias, respectively (Supplementary Digital Content eFigures 5A–5B, available at: <http://links.lww.com/JS9/F134>). The inconsistency test, measured by loop-specific approach, node-splitting, and design-by-treatment model, showed no significant inconsistencies in the present NMA (Supplementary Digital Content eTable 6A–6G, available at: <http://links.lww.com/JS9/F134>). There was no significant heterogeneity detected within the NMA (Supplementary Digital Content eTable 7A–7F, available at: <http://links.lww.com/JS9/F134>). The overall quality of evidence of this NMA falls within moderate-high (Supplementary Digital Content eTable 8A–8C, available at: <http://links.lww.com/JS9/F134>).

Discussion

To our knowledge, this is the first NMA to directly address the potential impact on colorectal tumor risk related to GLP-1 receptor agonists and SGLT2 inhibitors use in a dose-dependent way. In our NMA, we noticed that overall semaglutide within the recommended dosage was associated with a significantly increased incidence of colorectal tumors in a dose-dependent way, especially high-dose injectable semaglutide (2.4 mg/week) in subjects with obesity. Furthermore, even after subgrouping the included studies based on a treatment duration of at least 2 years, a trend toward an increased risk of colorectal tumors associated with semaglutide use remained, despite the reduced number of studies. Conversely, other investigated regimens were not associated with significantly different risks of colorectal tumors compared with controls.

The most important finding was that semaglutide was associated with a significantly increased risk of colorectal tumors in a dose-dependent way, which was different from the findings of previous traditional pairwise meta-analyses. In an earlier meta-analysis by Figlioli and colleagues, the authors pooled different kinds of gastrointestinal tract tumors into a whole group, ignoring the different site-specific tumor risk factors, and achieved results that overall GLP-1 receptor agonists would not alter the risk of overall gastrointestinal tract tumors^[20]. Following that, although another traditional pairwise meta-analysis specifically focused on colorectal tumors, the authors were still limited to a methodology restriction, which was to pool all medications with different mechanisms into one group, and achieved results that overall GLP-1 receptor agonists did not significantly

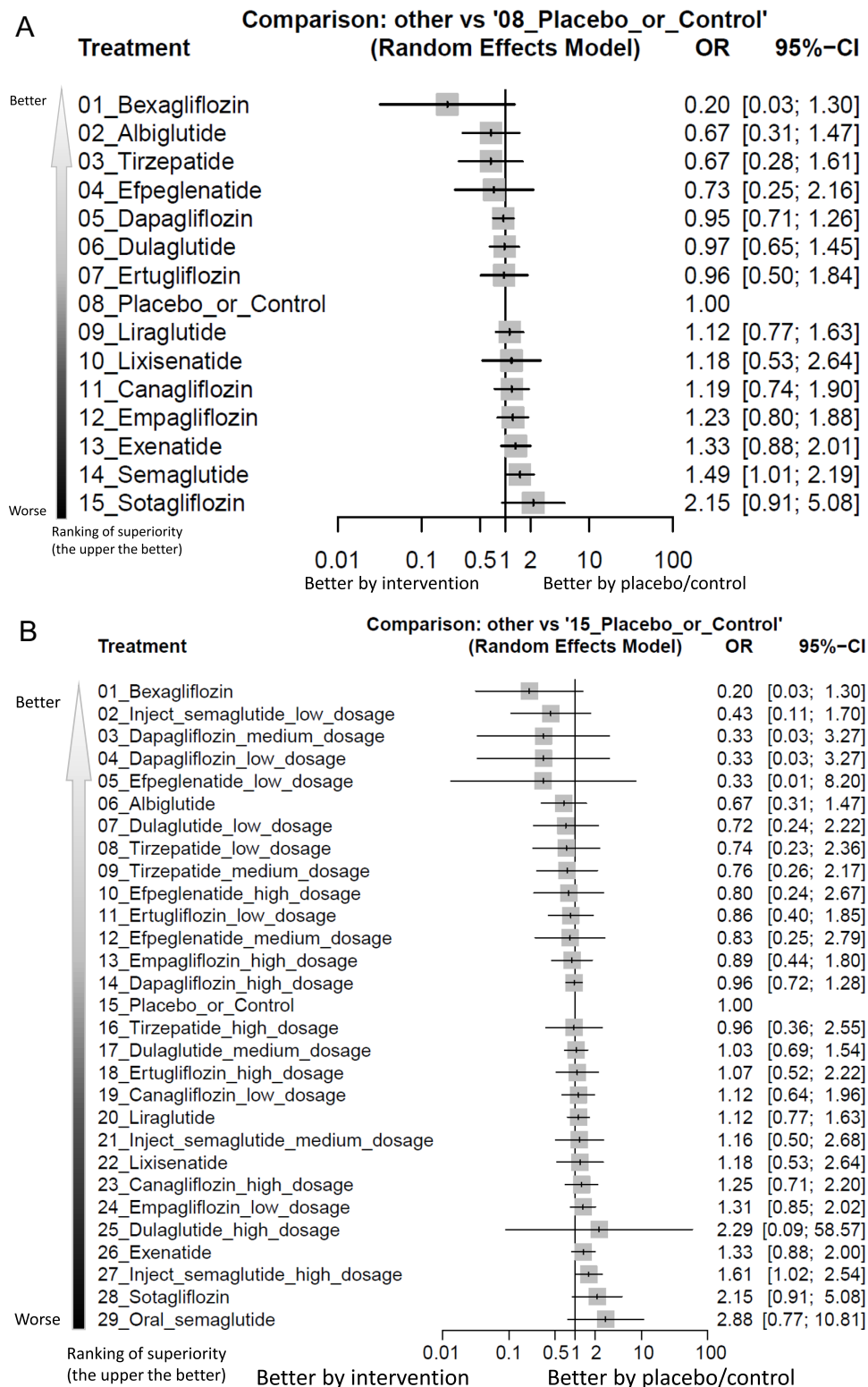


Figure 3. Forest plot of primary outcome: (A) overall colorectal tumor events, (B) colorectal tumor events in aspect of various dosage subgroups, (C) colorectal tumor events in RCTs of subjects with obesity, (D) colorectal tumor events in RCTs of treatment duration less than 2 years, and (E) colorectal tumor events in RCTs of treatment duration at least 2 years. When the effect size (expressed as odds ratio) is less than 1, the specified treatment is associated with fewer colorectal tumor events compared to placebo/controls. 95% CIs, 95% confidence intervals; GLP-1 agonist, glucagon-like peptide-1 agonist; NMA, network meta-analysis; OR, odds ratio; RCT, randomized controlled trial; SGLT2 inhibitor, sodium-glucose cotransporter 2 inhibitor.

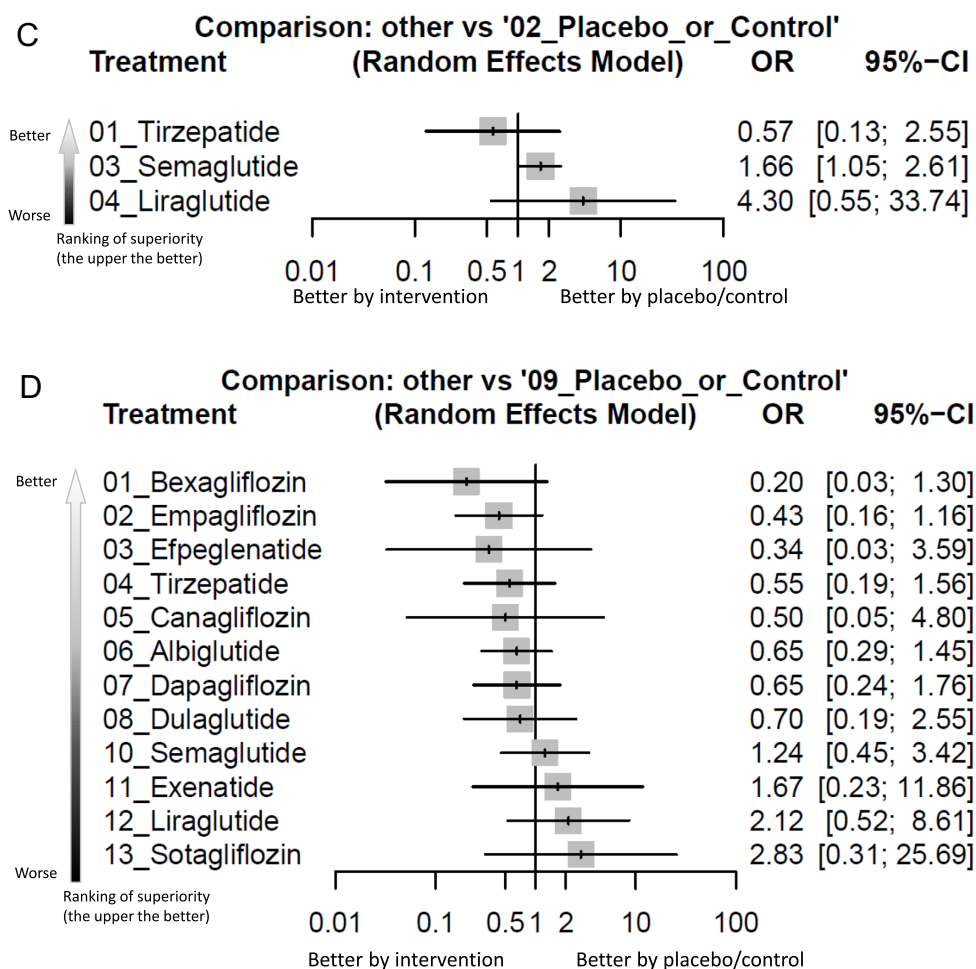


Figure 3.

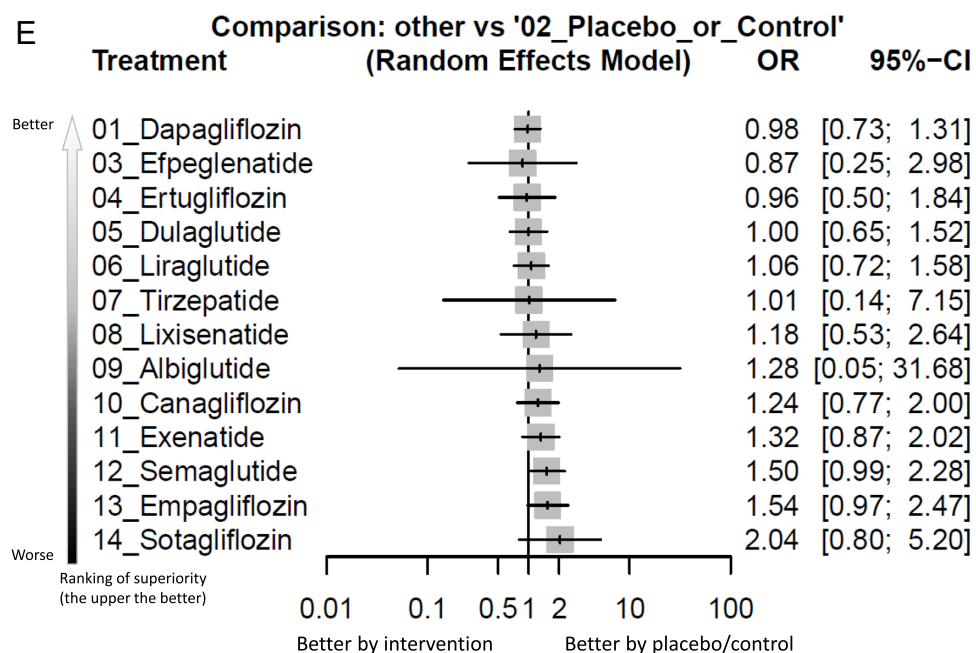


Figure 3.

Table 1
League table of the primary outcome: overall colorectal tumor events

[illegible]

95% CIs, 95% confidence intervals; GLP-1 agonist, glucagon-like peptide-1 agonist; NMA, network meta-analysis; OR, odds ratio; RCT, randomized controlled trial; SGLT2 inhibitor, sodium-glucose cotransporter 2 inhibitor.

League table of the primary outcome: colorectal tumor events in aspect of various dosage subgroups

Longitude	Latitude	Altitude	Temperature	Humidity	Wind Speed	Wind Direction	Cloud Cover	Visibility	Pressure	Soil Moisture	Soil Temperature	Plant Growth	Animal Activity	Human Presence	Time of Day	Date	Location
6.6100, 4.410	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-	-
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300
6.6100, 4.410	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0100, 1.300	0.0

95% CIs, 95% confidence intervals; GLP-1 agonist, glucagon-like peptide-1 agonist; NMA, network meta-analysis; OR, odds ratio; RCT, randomized controlled trial; SGLT2 inhibitor, sodium-glucose cotransporter 2 inhibitor.

Table 3
League table of the primary outcome: colorectal tumor events in RCTs of subjects with obesity

Tirzepatide	0.57 [0.13; 2.55]		
0.57 [0.13; 2.55]	Placebo_or_Control	*0.60 [0.38; 0.95]	0.23 [0.03; 1.83]
0.34 [0.07; 1.65]	*0.60 [0.38; 0.95]	Semaglutide	
0.13 [0.01; 1.69]	0.23 [0.03; 1.83]	0.39 [0.05; 3.18]	Liraglutide

95% CIs, 95% confidence intervals; GLP-1 agonist, glucagon-like peptide-1 agonist; NMA, network meta-analysis; OR, odds ratio; RCT, randomized controlled trial; SGLT2 inhibitor, sodium-glucose cotransporter 2 inhibitor.
Data present as OR [95% CIs]. Pairwise (upper-right portion) and network (lower-left portion) meta-analysis results are presented as estimated effect sizes for the outcome of events of colorectal tumors in RCTs of subjects with obesity. Interventions are reported in order of mean ranking of beneficially prophylactic effect on events of colorectal tumors in RCTs of subjects with obesity, and outcomes are expressed as OR (95% CIs). For the pairwise meta-analyses, OR of less than 1 indicates that the treatment specified in the row obtained more beneficial effect than that specified in the column. For the NMA, OR of less than 1 indicates that the treatment specified in the column obtained more beneficial effect than that specified in the row. Bold results marked with * indicate statistical significance.

increase the colorectal tumor risk without specific information regarding individual medications^[28]. Finally, Nagendra *et al* specifically focused on one medication (i.e., semaglutide) and achieved the result that semaglutide did not increase the risk of pancreatic tumor or thyroid tumor. However, they did not evaluate the changes in colorectal tumor risk^[37]. Different from those previous traditional pairwise meta-analyses, the current NMA, by exhaustively dividing different medications with various dosages into different groups, found the significantly increased risk of colorectal tumor by a specific regimen under a specific dosage. The NMA achieved its results based on direct evidence (i.e., semaglutide-control direct comparison) and indirect evidence [i.e., semaglutide–dulaglutide (another GLP-1 receptor agonist)–control and semaglutide–empagliflozin (one SGLT2 inhibitor)–control indirect comparison] in Figure 2A. Therefore, the significant difference between semaglutide and controls existed not only in direct pairwise comparison (right-upper part in Table 1) but also in overall NMA comparison (left-lower part in Table 1). This finding would provide insight into the clinical practice that (1) not every GLP-1 receptor agonist would be safe regarding the risk of colorectal tumor and (2) semaglutide should be used with caution about colorectal tumor risk when the dosage reaches an upper limit (i.e., 2.4 mg/week) in inject form. In addition to the main finding, the NMA had extra merits providing multiple comparisons between various GLP-1 receptor agonists and SGLT2 inhibitors in the left-lower part of Table 1.

The potential pro-tumorigenic effect of semaglutide might be supported by some *in vivo* and *in vitro* studies. As addressed before, semaglutide exerted its anti-apoptotic effect via activation of PKC ϵ and ERK1/2-MAPK pathway^[4]. Consequently, activated PKC ϵ could promote the survival of lung cancer cells by suppressing apoptosis^[5]. The overexpressed PKC ϵ has been demonstrated in numerous tumor-derived cell lines and histopathology tumor specimens^[118]. The overexpression of PKC ϵ was also one of the key features of metastatic colonic tumor cells and may be linked to ras-modulated signal transduction leading to neoplastic transformation in colonic epithelium^[26]. On the other hand, the activation of ERK1/2 would generally promote cell proliferation, in which dysregulation was a hallmark of many cancers^[6]. Specifically, the dysregulated

ERK1/2 and downstream mitogen-activated protein kinase (MAPK) activity was frequently seen in colorectal tumors^[27]. The molecular structural modification on amino acid chain positions 2 and 28, where alanine and lysine are replaced by 2-aminoisobutyric acid and arginine, improved semaglutide’s better combination with albumin, enabled longer presence in the blood circulation, and prevented chemical breakdown by dipeptidyl peptidase-4^[119]. This molecular-level difference could extend its half-life and allow longer stimulation on the MAPK/ERK pathway in colorectal cells. On the other hand, the increased dosage of injection-form semaglutide would increase the potential pro-tumorigenic signals in the targeted cells through its constant high concentration related to the injection-form designation^[120]. Finally, different from the uniform reaction to intrinsic GLP-1, the GLP-1 receptor would, under different extrinsic GLP-1 receptor agonist stimulation, activate different downstream signal transduction pathways (the so-called biased agonism and polymorphic variation)^[16]. That is one of the main reasons why various GLP-1 receptor agonists would have different adverse effect profiles. However, as addressed before, the hypothesis remained theoretical and lacked direct evidence to support. Future direct human trials should be warranted to provide more information regarding the risk of colorectal tumors related to semaglutide prescription, especially in high-dose inject form.

This NMA offers several methodological strengths that enhance the reliability and clinical utility of our findings. First, the NMA designation could enable statistical comparisons between different GLP-1 receptor agonists and SGLT2 inhibitors, providing more comprehensive evidence than traditional pairwise meta-analyses. Our rigorous methodology included exclusively focusing on RCTs, ensuring high-quality evidence while minimizing potential bias. By specifically excluding participants with pre-existing colorectal tumors, we were able to isolate true causative or prophylactic effects. Finally, our detailed subgroup analysis across various dosages and focusing on RCTs of subjects with obesity offer clinicians granular evidence to inform a dose-dependent risk related to a specific regimen.

Some limitations warrant consideration. First, the primary limitation relates to study duration; although the included trials averaged 127.5 weeks (range: 24–281 weeks) for study duration, this timeframe may be insufficient to fully capture the development of colorectal tumors. Although we arranged subgroup analysis based on study duration and noticed a trend of increased colorectal tumor risk in the subgroup of “at least 2-year treatment,” the concerns of insufficient medication exposure still existed in a relatively short follow-up duration. Second, our stringent focus on RCTs, while ensuring methodological rigor, potentially excluded valuable observational data from long-term cohort studies. Third, the most included RCTs were not specifically designed for detecting the incidence of colorectal tumors, so that the variation in diagnostic approaches across multi-country trials would present a notable limitation. The designation of original RCTs may have introduced heterogeneity in case identification of colorectal tumors, so that it might lead to challenges in establishing causality from secondary safety endpoints. Although meta-analyses may raise concerns about “inconsistencies in diagnostic tools,” the nature of these studies is rooted in large databases or numerous clinical studies, making them more reflective of clinical experience. Therefore,

Table 4													
League table of the primary outcome: colorectal tumor events in RCTs of treatment duration less than 2 years													
Bexagliflozin													
0.47 [0.06; 3.90]	Empagliflozin												
0.60 [0.03; 12.02]	1.26 [0.10; 16.28]	Efpeglenatide											
0.37 [0.04; 3.12]	0.78 [0.18; 3.30]	0.62 [0.05; 7.41]	Tirzepatide										
0.41 [0.02; 7.63]	0.86 [0.07; 10.25]	0.68 [0.03; 17.96]	1.10 [0.09; 13.36]	Canagliflozin									
0.31 [0.04; 2.38]	0.66 [0.18; 2.40]	0.53 [0.04; 6.36]	0.85 [0.23; 3.18]	0.77 [0.07; 8.55]	Albiglutide								
0.31 [0.04; 2.59]	0.66 [0.16; 2.73]	0.53 [0.04; 6.81]	0.85 [0.20; 3.61]	0.77 [0.06; 9.18]	1.00 [0.28; 3.63]	Dapagliflozin							
0.29 [0.03; 2.79]	0.61 [0.12; 3.10]	0.49 [0.05; 5.13]	0.78 [0.21; 2.98]	0.71 [0.05; 9.67]	0.92 [0.20; 4.24]	0.92 [0.18; 4.74]	Dulaglutide						
0.20 [0.03; 1.30]	0.43 [0.16; 1.16]	0.34 [0.03; 3.59]	0.55 [0.19; 1.56]	0.50 [0.05; 4.80]	0.65 [0.29; 1.45]	0.65 [0.24; 1.76]	0.70 [0.19; 2.55]	Placebo_or_Control					
0.16 [0.02; 1.36]	0.35 [0.09; 1.34]	0.27 [0.02; 3.45]	0.44 [0.11; 1.83]	0.40 [0.03; 4.82]	0.52 [0.14; 1.91]	0.52 [0.13; 2.18]	0.56 [0.12; 2.60]	0.81 [0.29; 2.23]	Semaglutide				
0.12 [0.01; 1.82]	0.26 [0.03; 2.32]	0.20 [0.01; 3.98]	0.33 [0.04; 2.81]	0.30 [0.01; 5.99]	0.39 [0.05; 3.24]	0.39 [0.04; 3.52]	0.42 [0.05; 3.37]	0.60 [0.08; 4.27]	0.74 [0.08; 6.58]	Exenatide			
*0.10 [0.01; 0.98]	0.20 [0.04; 1.13]	0.16 [0.01; 2.44]	0.26 [0.05; 1.46]	0.23 [0.02; 3.37]	0.30 [0.06; 1.54]	0.30 [0.05; 1.71]	0.33 [0.05; 2.10]	0.47 [0.12; 1.92]	0.58 [0.10; 3.27]	0.79 [0.10; 6.46]	Liraglutide		
0.07 [0.00; 1.28]	0.15 [0.01; 1.70]	0.12 [0.00; 3.03]	0.19 [0.02; 2.22]	0.18 [0.01; 4.15]	0.23 [0.02; 2.39]	0.23 [0.02; 2.57]	0.25 [0.02; 3.18]	0.35 [0.04; 3.20]	0.44 [0.04; 4.95]	0.59 [0.03; 11.25]	0.75 [0.05; 10.20]	Sotagliflozin	

95% CIs, 95% confidence intervals; GLP-1 agonist, glucagon-like peptide-1 agonist; NMA, network meta-analysis; OR, odds ratio; RCT, randomized controlled trial; SGLT2 inhibitor, sodium-glucose cotransporter 2 inhibitor. Data present as OR [95% CIs]. Pairwise (upper-right portion) and network (lower-left portion) meta-analysis results are presented as estimated effect sizes for the outcome of events of colorectal tumors in RCTs of subjects with obesity. Interventions are reported in order of mean ranking of beneficially prophylactic effect on events of colorectal tumors in RCTs of subjects with obesity, and outcomes are expressed as OR (95% CIs). For the pairwise meta-analyses, OR of less than 1 indicates that the treatment specified in the row obtained more beneficial effect than that specified in the column. For the NMA, OR of less than 1 indicates that the treatment specified in the column obtained more beneficial effect than that specified in the row. Bold results marked with * indicate statistical significance.

Table 5

[illegible]

95% CIs, 95% confidence intervals; GLP-1 agonist, glucagon-like peptide-1 agonist; NMA, network meta-analysis; OR, odds ratio; RCT, randomized controlled trial; SGLT2 inhibitor, sodium-glucose cotransporter 2 inhibitor.

many meta-analyses on drug side effects^[121,122] have become a benchmark for guiding future research^[18,19]. Finally, all the recruited RCTs were designed based on various underlying diseases (i.e., diabetes or obesity) but not specifically to examine colorectal tumor occurrence, so that they might not necessarily distinguish between benign and malignant tumors clearly. Rather, most of them provided data on “colorectal tumors” but did not specifically address “benign or malignant tumors.” Therefore, although we recognized the importance of distinguishing benign and malignant tumors, we could not arrange a specific analysis focusing on malignant colorectal tumors. These limitations suggest the need for longer-term studies specifically designed to assess the risk of colorectal tumors related to semaglutide use.

Conclusion

This NMA, the first to assess colorectal tumor risk across GLP-1 agonists and SGLT2 inhibitors, suggested that semaglutide, particularly the high-dose injectable formulation (2.4 mg/week) in individuals with obesity – may be associated with a dose-dependent increase in colorectal tumor incidence. This is notable given patients’ elevated baseline risk. Accordingly, the signal should be interpreted cautiously and confirmed in further studies.

Ethical approval

The study protocol was approved by the Institutional Review Board of the Tri-Service General Hospital, National Defense Medical Center (TSGHIRB E202516007).

Consent

The current study did not directly involve individual participants, so that we did not have the chance to approach individual participants or explore individual participants’ information. Therefore, it is impossible to obtain consent to participate in the current study.

Sources of funding

There is no funding for this study. This paper presents independent research. The views expressed in this publication are those of the authors and not necessarily those of the acknowledged institutions.

Author contributions

Chao-Ming Hung, Bing-Yan Zeng, and Chih-Wei Hsu, who contributed equally as first authors, took the whole responsibility of literature search, data extraction, data analysis, and manuscript drafting.

Po-Huang Chen, Cheuk-Kwan Sun, Andre F. Carvalho, Brendon Stubbs, Yen-Wen Chen, Tien-Yu Chen, Wei-Te Lei, and Jiann-Jy Chen contributed to study design, concept formation, and manuscript revision.

Yow-Ling Shiue, Kuan-Pin Su, Chih-Sung Liang, and Ping-Tao Tseng, who contributed equally as corresponding authors, took the whole responsibility of collecting information from the

other authors, manuscript’s major revisions, and manuscript submission.

Conflicts of interest disclosure

The authors report no financial interests or potential conflicts of interest.

Guarantor

Yow-Ling Shiue, Kuan-Pin Su, Chih-Sung Liang, and Ping-Tao Tseng.

Research registration unique identifying number (UIN)

PROSPERO CRD420251007952.

Provenance and peer review:

Not commissioned, externally peer-reviewed.

Data availability statement

All the data of the current study were available upon reasonable request to the corresponding authors.

Declaration of the usage of Artificial Intelligence (AI)

The authors declared that there was no AI applied in the current work.

References

- [1] Avogaro A, de Kreutzenberg SV, Morieri ML, Fadini GP, Del Prato S. Glucose-lowering drugs with cardiovascular benefits as modifiers of critical elements of the human life history. *Lancet Diabetes Endocrinol* 2022;10:882–89.
- [2] Tang H, Zhang B, Lu Y, *et al.* Assessing the benefit-risk profile of newer glucose-lowering drugs: a systematic review and network meta-analysis of randomized outcome trials. *Diabetes Obes Metab* 2024;27:1444–55.
- [3] Lin C, Cai X, Yang W, Lv F, Nie L, Ji L. Glycemic control and the incidence of neoplasm in patients with type 2 diabetes: a meta-analysis of randomized controlled trials. *Endocrine* 2020;70:232–42.
- [4] Zhu Q, Luo Y, Wen Y, Wang D, Li J, Fan Z. Semaglutide inhibits ischemia/reperfusion-induced cardiomyocyte apoptosis through activating PKG/PKCepsilon/ERK1/2 pathway. *Biochem. Biophys. Res. Commun.* 2023;647:1–8.
- [5] Ding L, Wang H, Lang W, Xiao L. Protein kinase C-epsilon promotes survival of lung cancer cells by suppressing apoptosis through dysregulation of the mitochondrial caspase pathway. *J Biol Chem* 2002;277:35305–13.
- [6] Sugiura R, Satoh R, Takasaki T. ERK: a double-edged sword in cancer. ERK-dependent apoptosis as a potential therapeutic strategy for cancer. *Cells* 2021;10:2509.
- [7] Koehler JA, Baggio LL, Yusta B, *et al.* GLP-1R agonists promote normal and neoplastic intestinal growth through mechanisms requiring Fgf7. *Cell Metab* 2015;21:379–91.
- [8] Piccoli GF, Mesquita LA, Stein C, *et al.* Do GLP-1 receptor agonists increase the risk of breast cancer? A systematic review and meta-analysis. *J Clin Endocrinol Metab* 2021;106:912–21.
- [9] Silverii GA, Monami M, Gallo M, *et al.* Glucagon-like peptide-1 receptor agonists and risk of thyroid cancer: a systematic review and

- meta-analysis of randomized controlled trials. *Diabetes Obes Metab* 2024;26:891–900.
- [10] Sharma N, Khatib MN, Balaraman AK, *et al.* Effect of GLP-1 receptor agonists on prostate cancer risk reduction: a systematic review and meta-analysis. *Int Urol Nephrol* 2024;57:1039–49.
 - [11] Arnold M, Abnet CC, Neale RE, *et al.* Global burden of 5 major types of gastrointestinal cancer. *Gastroenterology* 2020;159:335–349e315.
 - [12] Wang L, Wang W, Kaelber DC, Xu R, Berger NA. GLP-1 receptor agonists and colorectal cancer risk in drug-naïve patients with Type 2 diabetes, with and without overweight/obesity. *JAMA Oncol* 2024;10:256–58.
 - [13] Wang L, Xu R, Kaelber DC, Berger NA. Glucagon-like peptide 1 receptor agonists and 13 obesity-associated cancers in patients with Type 2 diabetes. *JAMA Network Open* 2024;7:e2421305.
 - [14] Zhu C, Lai Y, Liu C, *et al.* Comprehensively prognostic and immunological analyses of GLP-1 signaling-related genes in pan-cancer and validation in colorectal cancer. *Front Pharmacol* 2024;15:1387243.
 - [15] Lin A, Ding Y, Li Z, *et al.* Glucagon-like peptide 1 receptor agonists and cancer risk: advancing precision medicine through mechanistic understanding and clinical evidence. *Biomark Res* 2025;13:50.
 - [16] El Eid L, Reynolds CA, Tomas A, Ben J. Biased agonism and polymorphic variation at the GLP-1 receptor: implications for the development of personalised therapeutics. *Pharmacol Res* 2022;184:106411.
 - [17] Flausino LE, Carrasco AGM, Furuya TK, Tuan WJ, Chammas R. Impact of SGLT2 inhibitors on survival in gastrointestinal cancer patients undergoing chemotherapy and/or radiotherapy: a real-world data retrospective cohort study. *BMC Cancer* 2025;25:542.
 - [18] Wang SD. Opportunities and challenges of clinical research in the big-data era: from RCT to BCT. *J Thorac Dis* 2013;5:721–23.
 - [19] Zhang Z. Big data and clinical research: perspective from a clinician. *J Thorac Dis* 2014;6:1659–64.
 - [20] Figlioli G, Piovani D, Peppas S, *et al.* Glucagon-like peptide-1 receptor agonists and risk of gastrointestinal cancers: a systematic review and meta-analysis of randomized controlled trials. *Pharmacol Res* 2024;208:107401.
 - [21] Xu B, Kang B, Li S, Chen J, Zhou J. Association between sodium-glucose cotransporter 2 inhibitors and cancer: a systematic review and meta-analysis of randomized active-controlled trials. *Int J Clin Pharm* 2025 Apr 28. doi:10.1007/s11096-025-01924-0
 - [22] Huang YX, Wu JH, Zhao YQ, *et al.* An atlas on risk factors for gastrointestinal cancers: a systematic review of Mendelian randomization studies. *Prev Med* 2024;189:108147.
 - [23] Kayali S, Marabotto E, Giannini E. Gastrointestinal tract cancers, an increasing burden of the modern era: epidemiology and prevention. *Cancers (Basel)* 2023;15:15.
 - [24] Eng C, Yoshino T, Ruiz-Garcia E, *et al.* Colorectal cancer. *Lancet* 2024;404:294–310.
 - [25] Yu GH, Li SF, Wei R, Jiang Z. Diabetes and colorectal cancer risk: clinical and therapeutic implications. *J Diabetes Res* 2022;2022:1747326.
 - [26] Perletti GP, Folini M, Lin HC, Mischak H, Piccinini F, Tashjian AH Jr. Overexpression of protein kinase C epsilon is oncogenic in rat colonic epithelial cells. *Oncogene* 1996;12:847–54.
 - [27] Fang JY, Richardson BC. The MAPK signalling pathways and colorectal cancer. *Lancet Oncol* 2005;6:322–27.
 - [28] Bushi G, Gaidhane S, Ballal S, *et al.* Association between GLP1 RAs use and risk of colorectal cancer: a systematic review and meta-analysis. *Health Sci Rep* 2025;8:e70490.
 - [29] Higgins JP, Welton NJ. Network meta-analysis: a norm for comparative effectiveness? *Lancet* 2015;386:628–30.
 - [30] Tseng PT, Zeng BY, Hsu CW, *et al.* The pharmacodynamics-based prophylactic benefits of GLP-1 receptor agonists and SGLT2 inhibitors on neurodegenerative diseases: evidence from a network meta-analysis. *BMC Med* 2025;23:197.
 - [31] Chen JJ, Hsu CW, Hung CM, *et al.* Risk of hearing loss in patients treated with exendin-4 derivatives: a network meta-analysis of glucagon-like peptide-1 receptor agonists and sodium-glucose cotransporter 2 inhibitors. *Pharmaceuticals (Basel)* 2025;18:18.
 - [32] Peryer G, Golder S, Junqueira D, Vohra S, Loke YK, Group CAEM. Chapter 19: adverse effects. In: Higgins JPT, Thomas J, Chandler J, *et al.* eds. *Cochrane Handbook for Systematic Reviews of Interventions*. Cochrane, Cochrane; 2023. <https://www.cochrane.org/authors/handbooks-and-manuals/handbook/current/chapter-19>
 - [33] Page MJ, McKenzie JE, Bossuyt PM, *et al.* The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. *Int J Surg* 2021;88:105906.
 - [34] Shea BJ, Reeves BC, Wells G, *et al.* AMSTAR 2: a critical appraisal tool for systematic reviews that include randomised or non-randomised studies of healthcare interventions, or both. *Bmj* 2017;358:j4008.
 - [35] Agha RA, Mathew G, Rashid R, *et al.* Transparency in the reporting of Artificial Intelligence – the TITAN guideline. *Prem J Sci* 2025;10:100082.
 - [36] Wilson RB, Lathigara D, Kaushal D. Systematic review and meta-analysis of the impact of bariatric surgery on future cancer risk. *Int J Mol Sci* 2023;24:24.
 - [37] Nagendra L, Bg H, Sharma M, Dutta D. Semaglutide and cancer: a systematic review and meta-analysis. *Diabetes Metab Syndr* 2023;17:102834.
 - [38] Phillips R, Hazell L, Sauzet O, Cornelius V. Analysis and reporting of adverse events in randomised controlled trials: a review. *BMJ Open* 2019;9:e024537.
 - [39] Higgins J, Green S. *Cochrane Handbook for Systematic Reviews of Interventions* Version 5.0.2. The Cochrane Collaboration; 2009. <https://www.cochrane.org/authors/handbooks-and-manuals/handbook>
 - [40] Contedduca V, Sansonno D, Russi S, Dammacco F. Precancerous colorectal lesions (Review). *Int J Oncol* 2013;43:973–84.
 - [41] Gandhi SK, Reynolds MW, Boyer JG, Goldstein JL. Recurrence and malignancy rates in a benign colorectal neoplasm patient cohort: results of a 5-year analysis in a managed care environment. *Am J Gastroenterol* 2001;96:2761–67.
 - [42] Chaimani A, Caldwell DM, Li T, Higgins JPT, Salanti G. Chapter 11: undertaking network meta-analyses. In: Higgins JPT, Thomas J, Chandler J, *et al.* eds. *Cochrane Handbook for Systematic Reviews of Interventions*. Cochrane, London; 2018. 7.
 - [43] Borenstein M, Hedges LV, Higgins JP, Rothstein HR. A basic introduction to fixed-effect and random-effects models for meta-analysis. *Res Synth Methods* 2010;1:97–111.
 - [44] Owen RK, Bradbury N, Xin Y, Cooper N, Sutton A. MetaInsight: an interactive web-based tool for analyzing, interrogating, and visualizing network meta-analyses using R-shiny and netmeta. *Res Synth Methods* 2019;10:569–81.
 - [45] Cheng J, Pullenayegum E, Marshall JK, Iorio A, Thabane L. Impact of including or excluding both-armed zero-event studies on using standard meta-analysis methods for rare event outcome: a simulation study. *BMJ Open* 2016;6:e010983.
 - [46] Brockhaus AC, Bender R, Skipka G. The Peto odds ratio viewed as a new effect measure. *Stat Med* 2014;33:4861–74.
 - [47] Borenstein M. Converting among effect sizes. [Cited 2009] <https://www.meta-analysis.com/downloads/Meta-analysis%20Converting%20among%20effect%20sizes.pdf>.
 - [48] Dias S, Welton NJ, Caldwell DM, Ades AE. Checking consistency in mixed treatment comparison meta-analysis. *Stat Med* 2010;29:932–44.
 - [49] Higgins JP, Del Giovane C, Chaimani A, Caldwell DM, Salanti G. Evaluating the quality of evidence from a network meta-analysis. *Value Health* 2014;17:A324.
 - [50] White IR. Network meta-analysis. *Stata J* 2015;15:951–85.
 - [51] Anker SD, Butler J, Filippatos G, *et al.* Empagliflozin in heart failure with a preserved ejection fraction. *N Engl J Med* 2021;385:1451–61.
 - [52] Aroda VR, Frias JP, Ji L, *et al.* Efficacy and safety of once-weekly efpeglenatide in people with suboptimally controlled type 2 diabetes: the AMPLITUDE-D, AMPLITUDE-L and AMPLITUDE-S randomized controlled trials. *Diabetes Obes Metab* 2023;25:2084–95.
 - [53] Aronne LJ, Sattar N, Horn DB, *et al.* Continued treatment with tirzepatide for maintenance of weight reduction in adults with obesity: the SURMOUNT-4 randomized clinical trial. *Jama* 2024;331:38–48.
 - [54] Bailey CJ, Gross JL, Pieters A, Bastien A, List JF. Effect of dapagliflozin in patients with type 2 diabetes who have inadequate glycaemic control with metformin: a randomised, double-blind, placebo-controlled trial. *Lancet* 2010;375:2223–33.
 - [55] Bhatt DL, Szarek M, Pitt B, *et al.* Sotagliflozin in patients with diabetes and chronic kidney disease. *N Engl J Med* 2021;384:129–39.
 - [56] Bhatt DL, Szarek M, Steg PG, *et al.* Sotagliflozin in patients with diabetes and recent worsening heart failure. *N Engl J Med* 2021;384:117–28.
 - [57] Blonde L, Jendle J, Gross J, *et al.* Once-weekly dulaglutide versus bedtime insulin glargine, both in combination with prandial insulin

- lispro, in patients with type 2 diabetes (AWARD-4): a randomised, open-label, phase 3, non-inferiority study. *Lancet* 2015;385:2057–66.
- [58] Buse JB, Nordahl Christensen H, Harty BJ, *et al.* Study design and baseline profile for adults with type 2 diabetes in the once-weekly subcutaneous SEMaglutide randomized PRagmatic (SEPRa) trial. *BMJ Open Diabetes Res Care* 2023;11:11.
- [59] Buse JB, Rosenstock J, Sesti G, *et al.* Liraglutide once a day versus exenatide twice a day for type 2 diabetes: a 26-week randomised, parallel-group, multinational, open-label trial (LEAD-6). *Lancet* 2009;374:39–47.
- [60] Cannon CP, Pratley R, Dagogo-Jack S, *et al.* Cardiovascular outcomes with Ertugliflozin in Type 2 Diabetes. *N Engl J Med* 2020;383:1425–35.
- [61] Cherney DZI, Ferrannini E, Umpierrez GE, *et al.* Efficacy and safety of sotagliflozin in patients with type 2 diabetes and stage 3 chronic kidney disease. *Diabetes Obes Metab* 2023;25:1646–57.
- [62] Davies M, Faerch L, Jeppesen OK, *et al.* Semaglutide 2.4 mg once a week in adults with overweight or obesity, and type 2 diabetes (STEP 2): a randomised, double-blind, double-dummy, placebo-controlled, phase 3 trial. *Lancet* 2021;397:971–84.
- [63] Davies MJ, Bain SC, Atkin SL, *et al.* Efficacy and safety of liraglutide versus placebo as add-on to glucose-lowering therapy in patients with Type 2 diabetes and moderate renal impairment (LIRA-RENAL): a randomized clinical trial. *Diabetes Care* 2016;39:222–30.
- [64] Davies MJ, Bergenstal R, Bode B, *et al.* Efficacy of liraglutide for weight loss among patients with Type 2 diabetes: the SCALE diabetes randomized clinical trial. *Jama* 2015;314:687–99.
- [65] Del Prato S, Kahn SE, Pavo I, *et al.* Tirzepatide versus insulin glargine in type 2 diabetes and increased cardiovascular risk (SURPASS-4): a randomised, open-label, parallel-group, multicentre, phase 3 trial. *Lancet* 2021;398:1811–24.
- [66] Gallwitz B, Guzman J, Dotta F, *et al.* Exenatide twice daily versus glimepiride for prevention of glycaemic deterioration in patients with type 2 diabetes with metformin failure (EUREXA): an open-label, randomised controlled trial. *Lancet* 2012;379:2270–78.
- [67] Gao L, Lee BW, Chawla M, *et al.* Tirzepatide versus insulin glargine as second-line or third-line therapy in type 2 diabetes in the Asia-Pacific region: the SURPASS-AP-Combo trial. *Nat Med* 2023;29:1500–10.
- [68] Garber A, Henry R, Ratner R, *et al.* Liraglutide versus glimepiride monotherapy for type 2 diabetes (LEAD-3 Mono): a randomised, 52-week, phase III, double-blind, parallel-treatment trial. *Lancet* 2009;373:473–81.
- [69] Garvey WT, Frias JP, Jastreboff AM, *et al.* Tirzepatide once weekly for the treatment of obesity in people with type 2 diabetes (SURMOUNT-2): a double-blind, randomised, multicentre, placebo-controlled, phase 3 trial. *Lancet* 2023;402:613–26.
- [70] Gerstein HC, Colhoun HM, Dagenais GR, *et al.* Dulaglutide and cardiovascular outcomes in type 2 diabetes (REWIND): a double-blind, randomised placebo-controlled trial. *Lancet* 2019;394:121–30.
- [71] Gerstein HC, Sattar N, Rosenstock J, *et al.* Cardiovascular and renal outcomes with Efglenatide in Type 2 diabetes. *N Engl J Med* 2021;385:896–907.
- [72] Heerspink HJL, Stefansson BV, Correa-Rotter R, *et al.* Dapagliflozin in patients with chronic kidney disease. *N Engl J Med* 2020;383:1436–46.
- [73] Hernandez AF, Green JB, Janmohamed S, *et al.* Albiglutide and cardiovascular outcomes in patients with type 2 diabetes and cardiovascular disease (Harmony Outcomes): a double-blind, randomised placebo-controlled trial. *Lancet* 2018;392:1519–29.
- [74] Holman RR, Bethel MA, Mentz RJ, *et al.* Effects of once-weekly exenatide on cardiovascular outcomes in Type 2 diabetes. *N Engl J Med* 2017;377:1228–39.
- [75] Home PD, Ahren B, Reusch JEB, *et al.* Three-year data from 5 HARMONY phase 3 clinical trials of albiglutide in type 2 diabetes mellitus: long-term efficacy with or without rescue therapy. *Diabet Res Clin Pract* 2017;131:49–60.
- [76] Husain M, Birkenfeld AL, Donsmark M, *et al.* Oral semaglutide and cardiovascular outcomes in patients with Type 2 diabetes. *N Engl J Med* 2019;381:841–51.
- [77] Inagaki N, Takeuchi M, Oura T, Imaoka T, Seino Y. Efficacy and safety of tirzepatide monotherapy compared with dulaglutide in Japanese patients with type 2 diabetes (SURPASS J-mono): a double-blind, multicentre, randomised, phase 3 trial. *Lancet Diabetes Endocrinol* 2022;10:623–33.
- [78] Januzzi JL Jr, Butler J, Jarolim P, *et al.* Effects of canagliflozin on cardiovascular biomarkers in older adults with Type 2 diabetes. *J Am Coll Cardiol* 2017;70:704–12.
- [79] Ji L, Lu Y, Li Q, *et al.* Efficacy and safety of empagliflozin in combination with insulin in Chinese patients with type 2 diabetes and insufficient glycaemic control: a phase III, randomized, double-blind, placebo-controlled, parallel study. *Diabetes Obes Metab* 2023;25:1839–48.
- [80] Kadowaki T, Chin R, Ozeki A, Imaoka T, Ogawa Y. Safety and efficacy of tirzepatide as an add-on to single oral antihyperglycaemic medication in patients with type 2 diabetes in Japan (SURPASS J-combo): a multicentre, randomised, open-label, parallel-group, phase 3 trial. *Lancet Diabetes Endocrinol* 2022;10:634–44.
- [81] Kaku K, Yamada Y, Watada H, *et al.* Safety and efficacy of once-weekly semaglutide vs additional oral antidiabetic drugs in Japanese people with inadequately controlled type 2 diabetes: a randomized trial. *Diabetes Obes Metab* 2018;20:1202–12.
- [82] Kellner M, Kaltoft MS, Lawson J, *et al.* Effect of once-weekly semaglutide versus thrice-daily insulin aspart, both as add-on to metformin and optimized insulin glargine treatment in participants with type 2 diabetes (SUSTAIN 11): a randomized, open-label, multinational, phase 3b trial. *Diabetes Obes Metab* 2022;24:1788–99.
- [83] Laval-Gonzalez FJ, Januszewicz A, Davidson J, *et al.* Efficacy and safety of canagliflozin compared with placebo and sitagliptin in patients with type 2 diabetes on background metformin monotherapy: a randomised trial. *Diabetologia* 2013;56:2582–92.
- [84] Lincoff AM, Brown-Frandsen K, Colhoun HM, *et al.* Semaglutide and cardiovascular outcomes in obesity without diabetes. *N Engl J Med* 2023;389:2221–32.
- [85] Lock JP. Bexagliflozin Efficacy and Safety Trial (BEST). [Cited 28 October 2024] <https://clinicaltrials.gov/study/NCT02558296?cond=NCT02558296&rank=1>.
- [86] Ludvik B, Giorgino F, Jodar E, *et al.* Once-weekly tirzepatide versus once-daily insulin degludec as add-on to metformin with or without SGLT2 inhibitors in patients with type 2 diabetes (SURPASS-3): a randomised, open-label, parallel-group, phase 3 trial. *Lancet* 2021;398:583–98.
- [87] Marso SP, Bain SC, Consoli A, *et al.* Semaglutide and cardiovascular outcomes in patients with Type 2 diabetes. *N Engl J Med* 2016;375:1834–44.
- [88] Marso SP, Daniels GH, Brown-Frandsen K, *et al.* Liraglutide and cardiovascular outcomes in Type 2 diabetes. *N Engl J Med* 2016;375:311–22.
- [89] McMurray JJV, Solomon SD, Inzucchi SE, *et al.* Dapagliflozin in patients with heart failure and reduced ejection fraction. *N Engl J Med* 2019;381:1995–2008.
- [90] Mellander A, Billger M, Johnsson E, Traff AK, Yoshida S, Johnsson K. Hypersensitivity events, including potentially hypersensitivity-related skin events, with dapagliflozin in patients with Type 2 diabetes mellitus: a pooled analysis. *Clin. Drug Invest.* 2016;36:925–33.
- [91] Natale P, Tunnicliffe DJ, Toyama T, *et al.* Sodium-glucose co-transporter protein 2 (SGLT2) inhibitors for people with chronic kidney disease and diabetes. *Cochrane Database Syst Rev* 2024;5:5.
- [92] Nauck M, Frid A, Hermansen K, *et al.* Efficacy and safety comparison of liraglutide, glimepiride, and placebo, all in combination with metformin, in type 2 diabetes: the LEAD (liraglutide effect and action in diabetes)-2 study. *Diabetes Care* 2009;32:84–90.
- [93] Nauck MA, Stewart MW, Perkins C, *et al.* Efficacy and safety of once-weekly GLP-1 receptor agonist albiglutide (HARMONY 2): 52 week primary endpoint results from a randomised, placebo-controlled trial in patients with type 2 diabetes mellitus inadequately controlled with diet and exercise. *Diabetologia* 2016;59:266–74.
- [94] Neal B, Perkovic V, Mahaffey KW, *et al.* Canagliflozin and cardiovascular and renal events in Type 2 diabetes. *N Engl J Med* 2017;377:644–57.
- [95] Packer M, Anker SD, Butler J, *et al.* Cardiovascular and renal outcomes with empagliflozin in heart failure. *N Engl J Med* 2020;383:1413–24.
- [96] Perkovic V, Jardine MJ, Neal B, *et al.* Canagliflozin and renal outcomes in Type 2 diabetes and nephropathy. *N Engl J Med* 2019;380:2295–306.
- [97] Pfeffer MA, Claggett B, Diaz R, *et al.* Lixisenatide in patients with Type 2 diabetes and acute coronary syndrome. *N Engl J Med* 2015;373:2247–57.

- [98] Pieber TR, Bode B, Mertens A, *et al.* Efficacy and safety of oral semaglutide with flexible dose adjustment versus sitagliptin in type 2 diabetes (PIONEER 7): a multicentre, open-label, randomised, phase 3a trial. *Lancet Diabetes Endocrinol* 2019;7:528–39.
- [99] Pi-Sunyer X, Astrup A, Fujioka K, *et al.* A randomized, controlled trial of 3.0 mg of liraglutide in weight management. *N Engl J Med* 2015;373:11–22.
- [100] Polidori D, Mari A, Ferrannini E. Canagliflozin, a sodium glucose co-transporter 2 inhibitor, improves model-based indices of beta cell function in patients with type 2 diabetes. *Diabetologia* 2014;57:891–901.
- [101] Pratley RE, Aroda VR, Lingvay I, *et al.* Semaglutide versus dulaglutide once weekly in patients with type 2 diabetes (SUSTAIN 7): a randomised, open-label, phase 3b trial. *Lancet Diabetes Endocrinol* 2018;6:275–86.
- [102] Ridderstrale M, Andersen KR, Zeller C, *et al.* Comparison of empagliflozin and glimepiride as add-on to metformin in patients with type 2 diabetes: a 104-week randomised, active-controlled, double-blind, phase 3 trial. *Lancet Diabetes Endocrinol* 2014;2:691–700.
- [103] Rodbard HW, Rosenstock J, Canani LH, *et al.* Oral semaglutide versus empagliflozin in patients with Type 2 diabetes uncontrolled on metformin: the PIONEER 2 trial. *Diabetes Care* 2019;42:2272–81.
- [104] Roden M, Weng J, Eilbracht J, *et al.* Empagliflozin monotherapy with sitagliptin as an active comparator in patients with type 2 diabetes: a randomised, double-blind, placebo-controlled, phase 3 trial. *Lancet Diabetes Endocrinol* 2013;1:208–19.
- [105] Rosenstock J, Fonseca VA, Gross JL, *et al.* Advancing basal insulin replacement in type 2 diabetes inadequately controlled with insulin glargine plus oral agents: a comparison of adding albiglutide, a weekly GLP-1 receptor agonist, versus thrice-daily prandial insulin lispro. *Diabetes Care* 2014;37:2317–25.
- [106] Rosenstock J, Frias JP, Rodbard HW, *et al.* Tirzepatide vs insulin lispro added to basal insulin in Type 2 diabetes: the SURPASS-6 randomized clinical trial. *Jama* 2023;330:1631–40.
- [107] Solomon SD, McMurray JJV, Claggett B, *et al.* Dapagliflozin in heart failure with mildly reduced or preserved ejection fraction. *N Engl J Med* 2022;387:1089–98.
- [108] The E-KCG. Herrington WG, Staplin N, Wanner C, *et al.* Empagliflozin in patients with chronic kidney disease. *N Engl J Med* 2023;388:117–27.
- [109] Tuttle KR, Levin A, Nangaku M, *et al.* Safety of empagliflozin in patients with Type 2 diabetes and chronic kidney disease: pooled analysis of placebo-controlled clinical trials. *Diabetes Care* 2022;45:1445–52.
- [110] Wada T, Mori-Anai K, Takahashi A, *et al.* Effect of canagliflozin on the decline of estimated glomerular filtration rate in chronic kidney disease patients with type 2 diabetes mellitus: a multicenter, randomized, double-blind, placebo-controlled, parallel-group, phase III study in Japan. *J Diabetes Investig* 2022;13:1981–89.
- [111] Wadden TA, Chao AM, Machineni S, *et al.* Tirzepatide after intensive lifestyle intervention in adults with overweight or obesity: the SURMOUNT-3 phase 3 trial. *Nat Med* 2023;29:2909–18.
- [112] Wason S Efficacy and bone safety of sotagliflozin 400 and 200 mg versus placebo in participants with type 2 diabetes mellitus who have inadequate glycemic control (SOTA-BONE). [Cited 28 October 2024] <https://clinicaltrials.gov/study/NCT03386344?cond=NCT03386344&rank=1>.
- [113] Weinstock RS, Guerci B, Umpierrez G, Nauck MA, Skrivaneck Z, Milicevic Z. Safety and efficacy of once-weekly dulaglutide versus sitagliptin after 2 years in metformin-treated patients with type 2 diabetes (AWARD-5): a randomized, phase III study. *Diabetes Obes Metab* 2015;17:849–58.
- [114] Wiviott SD, Raz I, Bonaca MP, *et al.* Dapagliflozin and cardiovascular outcomes in Type 2 diabetes. *N Engl J Med* 2019;380:347–57.
- [115] Wysham C, Blevins T, Arakaki R, *et al.* Efficacy and safety of dulaglutide added onto pioglitazone and metformin versus exenatide in type 2 diabetes in a randomized controlled trial (AWARD-1). *Diabetes Care* 2014;37:2159–67.
- [116] Zinman B, Wanner C, Lachin JM, *et al.* Empagliflozin, cardiovascular outcomes, and mortality in Type 2 diabetes. *N Engl J Med* 2015;373:2117–28.
- [117] Brignardello-Petersen R, Izcovich A, Roehweg B, *et al.* GRADE approach to drawing conclusions from a network meta-analysis using a partially contextualised framework. *Bmj* 2020;371:m3907.
- [118] Gorin MA, Pan Q. Protein kinase C ϵ : an oncogene and emerging tumor biomarker. *Mol Cancer* 2009;8:1.
- [119] Lau J, Bloch P, Schaffer L, *et al.* Discovery of the once-weekly Glucagon-Like Peptide-1 (GLP-1) analogue semaglutide. *J Med Chem* 2015;58:7370–80.
- [120] Ibrahim SS, Ibrahim RS, Arabi B, Brockmueller A, Shakibaei M, Busselberg D. The effect of GLP-1R agonists on the medical triad of obesity, diabetes, and cancer. *Cancer Metastasis Rev* 2024;43:1297–314.
- [121] Hinkle JT, Graziosi M, Nayak SM, Yaden DB. Adverse events in studies of classic psychedelics: a systematic review and meta-analysis. *JAMA Psychiatry* 2024;81:1225–35.
- [122] Pillinger T, McCutcheon RA, Vano L, *et al.* Comparative effects of 18 antipsychotics on metabolic function in patients with schizophrenia, predictors of metabolic dysregulation, and association with psychopathology: a systematic review and network meta-analysis. *Lancet Psychiatry* 2020;7:64–77.