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# Survival and neurodevelopmental impairment of outborn preterm infants at 5.5 years of age: an EPIPAGE-2 prospective, matched study using multiple imputation

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### **ABSTRACT**

**Objective** To determine whether birth outside a level-3 centre (outborn) is associated with a difference in the combined outcome of mortality or moderate-to-severe neurological impairment at 5.5 years of age compared with birth in a level-3 centre (inborn) when antenatal steroids and gestational age (GA) are accounted for.

**Design** Individual matched study nested within a prospective cohort. Each outborn infant was matched using GA and antenatal steroids with a maximum of four inborns. Conditional logistic regression was used to calculate ORs before being adjusted using maternal and birth characteristics. Analyses were carried out after multiple imputation for missing data.

**Setting** EPIPAGE-2 French national prospective cohort including births up to 34 weeks GA inclusive.

Patients Outborn and inborn control infants selected between 24 and 31 weeks GA were followed in the neonatal period and to 2 and 5.5 years. 3335 infants were eligible of whom all 498 outborns and 1235 inborn infants were included—equivalent to 2.5 inborns for each outborn. Main outcome measure Survival without moderate-to-severe neurodevelopmental impairment at 5.5 years. Results Chorioamnionitis, pre-eclampsia, caesarian birth and small-for-dates were more frequent among inborns, and spontaneous labour and antepartum haemorrhage among outborns. There was no difference in the main outcome measure at 5.5 years of age (adjusted OR 1.09, 95% CI 0.82 to 1.44); sensitivity analyses suggested improved outcomes at lower GAs for inborns.

**Conclusion** In this GA and steroid matched cohort, there was no difference in survival without moderate-to-severe neurodevelopmental impairment to 5.5 years of age between inborn and outborn very preterm children. This suggests steroids might be important in determining outcomes.

### INTRODUCTION

Excess mortality and increased neurological complications have been demonstrated in preterm infants born in non-tertiary (level 1 and level 2) hospitals without neonatal

### WHAT IS ALREADY KNOWN ON THIS TOPIC

- ⇒ Excess mortality and/or neurological complications have been demonstrated in preterm infants born outside neonatal intensive care units compared with inborns in many countries.
- ⇒ Pre-eclampsia, premature rupture of membranes and small for gestational age are frequent in inborns whereas spontaneous birth and placental abruption are frequent in outborns.
- ⇒ Administration of antenatal steroids is lower in outborn than inborn pregnancies, and is independently associated with worse outcomes.

### WHAT THIS STUDY ADDS

- ⇒ There were no differences in survival to 5.5 years without moderate/severe neurodevelopmental impairment between inborns and outborns when gestational age and antenatal steroids were accounted for.
- ⇒ There were improved outcomes at 5.5 years with inborn delivery for births at 24–27 weeks of gestation, but not below 24 weeks.
- ⇒ There were no differences in outcomes at 2 years of age or at hospital discharge—other than for necrotising enterocolitis, which was more frequent among outborns.

# HOW THIS STUDY MIGHT AFFECT RESEARCH, PRACTICE OR POLICY

- Improved protocols should be implemented to enhance antenatal transfer of women delivering at extremely preterm gestations.
- Administration of antenatal steroids to women presenting to non-level 3 units who risk premature delivery should be enhanced.
- ⇒ Further research about the impact of place of birth on outcomes to school age for children born very preterm (below 32 weeks of gestation) is required.

intensive care facilities in the USA, <sup>1-3</sup> Finland, <sup>4</sup> Spain, <sup>5</sup> Australia, <sup>6 7</sup> Canada <sup>8</sup> and England. <sup>9 10</sup> Possible reasons that are not mutually exclusive are that outborn infants were more unwell





prior to delivery, that perinatal care was suboptimal for those outborn or that postnatal transfer is deleterious to the infants' well-being. Previous studies have investigated the longer-term neurological consequences of outborn very preterm infants up to 2–3 years of age. <sup>7 11–13</sup>

Non-level-3 hospitals are often able to transfer mothers with 'stable' pathologies (where birth is not imminent) such as pre-eclampsia, spontaneous rupture of membranes without labour and small for gestation age. Hospitals 'pathologies such as spontaneous labour and antepartum haemorrhage are more frequent in outborn births. This creates distortion in caseload between transferring and receiving units which introduces selection bias when trying to understand differences in outcomes according to place of birth. In addition, the use of antenatal steroids is more prevalent in inborn births 7 9 10 14-17 and represents an important source of confounding between the outcomes of outborn and inborn infants; steroids being independently associated with reduced mortality and neuromorbidity. We used the EPIPAGE-2<sup>19-21</sup> national, population-based

We used the EPIPAGE-2<sup>19-21</sup> national, population-based prospective study to examine the hypothesis that, when the influences of gestational age (GA) and antenatal steroids are controlled for, the combined outcome of survival without moderate or severe neurodevelopmental impairment at 5.5 years is not different between outborn and inborn infants.

### **METHODS**

### Study population

EPIPAGE-2 prospectively collected data on all births occurring in France in 2011 under 35 weeks GA. <sup>19–21</sup> Infants born in mainland France between 24 and 31 weeks of GA were eligible for inclusion in this study. Infants born at home, who had lethal congenital abnormalities or for whom antenatal steroids data were missing were excluded.

### **Exposure**

In France, perinatal centres are categorised into level-3 units, which provide neonatal intensive care for infants requiring ongoing invasive ventilation,and non-level-3 units, which may provide some short-term invasive ventilation. National guidelines in 2011 recommended that births at less than 32 weeks of gestation or less than 1500 g occur in level-3 units, <sup>22</sup> building on the regionalised system of care first mandated in 1998. <sup>23</sup>

### **Outcomes**

The primary outcome was survival without moderate or severe neurodevelopmental impairment at 5.5 years of age. Evaluations were performed by trained paediatricians and developmental child psychologists, and included cerebral palsy (CP), sensory (vision and hearing) and cognitive abilities. CP was classified using the Global Motor Function Classification System (GMFCS), with level 1 being mild, levels 2–3 moderate and levels 4–5

severe impairment.<sup>24</sup> Moderate or severe visual impairment was defined as bilateral binocular visual acuity 3.2/10, and moderate or severe hearing impairment as unilateral or bi-lateral hearing loss >40 dB not corrected or only partially corrected by hearing aids. Cognitive deficiency was measured using the full scale intelligence quotient (FSIQ) from the French version of the Wechsler Preschool and Primary Scale of Intelligence fourth Edition,<sup>25</sup> with moderate-to-severe impairment defined as an FSIQ 2 SD below the mean of a term-born reference group.<sup>21</sup> We present results for the primary outcome and survival, and for moderate-to-severe neurodevelopmental impairment among survivors.

We also investigated earlier outcomes. At discharge, we examined survival without severe neonatal morbidities (intraventricular haemorrhage grades III-IV, or cystic periventricular leukomalacia; necrotising enterocolitis, Bell stages 2–3; severe bronchopulmonary dysplasia, requiring  $\geq 30\%$  oxygen and/or ventilatory support at 36 weeks' GA; retinopathy of prematurity stage 3 or higher) as well as individual morbidities.<sup>20</sup> At age 2, we examined survival without moderate or severe sensorimotor impairment defined as CP (GMFCS levels 2-5) and/or unilateral or bilateral deafness and/or blindness, as reported by attending paediatricians. We additionally report the number of children with a parentally reported Ages and Stages Questionnaire (ASQ) score below the reference threshold in at least one of the five domains (communication abilities, gross motor skills, fine motor skills, problem solving abilities and personal-social skills).<sup>26</sup>

### Statistical analysis

We compared potential confounding factors between outborn and inborn groups of infants, then performed conditional logistic regression to explore outcome differences using a matched cohort design to eliminate differences in GA and antenatal steroid administration between the two groups. A greedy matching algorithm without replacement was used, with exact matching on week of GA and antenatal steroid administration in the 2 weeks prior to birth (three categories: full course, defined as two doses more than 24 hours before birth, partial course, or no steroids); neither variable had missing data. As there were missing data among outcomes and covariates, all analyses were performed after multiple imputation; hence unique identifiers of matched infants were noted for subsequent use with the imputed data. The imputation model included the exposure, outcomes and matching variables, as well as variables potentially predicting non-response or the outcome. Categorical variables were imputed using logistic or multinomial regression and continuous variables using predictive mean matching. Full details are shown in online supplemental table 1. Using the entire population of infants eligible for study inclusion; we generated 50 independent imputed datasets with 30 iterations each. Data for infants previously selected by the matching process were then retained from each imputed data set for use in analyses.



Data were also weighted to account for imbalances due to varying numbers of matched inborn children; outborns were assigned a weight of one and each control subject had a weight proportional to the number of cases in its matched set divided by the number of controls in the set. Conditional logistic regression was performed with adjustment for maternal age (years), parents' socioeconomic status (using the highest level of either parent, or the mother only if she was single), nulliparous, use of infertility treatment, multiple pregnancy, tocolysis, antenatal administration of magnesium sulfate, premature rupture of membranes (PROM), spontaneous onset of labour, chorioamnionitis, pre-eclampsia, antepartum haemorrhage, caesarean section, cephalic presentation, small-for-GA and sex (all variables binary unless stated).

Analyses were performed in SAS using 'proc psmatch' for matching and 'proc MI' for imputation. Results were pooled according to Rubin's rule.<sup>27</sup> All statistical tests

were two sided and the probability of a type 1 error ( $\alpha$ ) was set at <0.05.

### **Sensitivity analyses**

We assessed the primary outcome in restricted populations of 24–27 weeks and 28–31 weeks of gestation to counter for potential effects from infants who were not transferred and from differences in attitude to extreme preterm birth; we also provide estimates by week of GA. Because reasons for delivery may vary between fetuses in multiple pregnancies, we repeated these analyses using the complete GA range (24–31 weeks) but restricted to singleton births only. Finally, to account for the influence of obstetric decision-making, we examined the population of 24–31 week GA fetuses alive at maternal admission to hospital. We also repeated the principal analyses using only subjects with complete data.

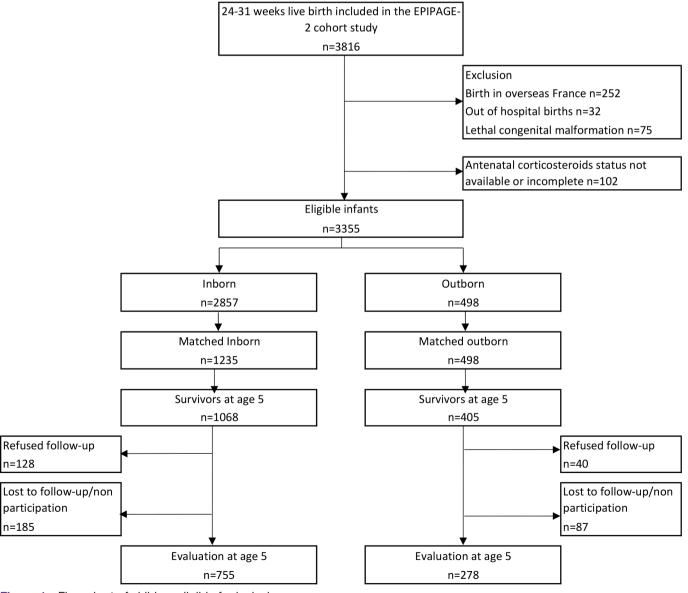


Figure 1 Flow chart of children eligible for inclusion.



Table 1 Perinatal characteristics of the initial cohort and matched outborn infants with inborn controls

	Initial cohort* Matched cohort†								
	Inborn		Outborn		Inborn 1:n (up to 4)		Outborn		
	n=2857	%	n=498	%	n=1235	%	n=498		P value:
Gestational age, week									
24	130	4.6	31	6.2	66	6.2	31	6.2	>0.999
25	233	8.2	40	8.0	91	8.0	40	8.0	
26	308	10.8	49	9.8	105	9.8	49	9.8	
27	297	10.4	41	8.2	99	8.2	41	8.2	
28	361	12.6	47	9.4	121	9.4	47	9.4	
29	401	14.0	51	10.2	108	10.2	51	10.2	
30	528	18.5	81	16.3	235	16.3	81	16.3	
31	599	21.0	158	31.7	410	31.7	158	31.7	
Maternal characteristics at birth									
Maternal age									
<25 years	507	17.8	120	24.1	224	17.9	120	24.1	0.002
25–34 years	1693	59.3	299	60.0	743	59.8	299	60.0	
>35 years	655	22.9	79	15.9	267	22.2	79	15.9	
Missing	2		0		_		-		
Mother born in France	2092	74.9	393	82.4	905	72.5	407	81.7	<0.001
Missing	63		21		_		_		
Parents' socioeconomic status§									
Professional	611	22.8	77	17.3	253	19.2	84	17.0	0.42
Intermediate	548	20.4	91	20.4	256	20.5	100	20.1	
Administrative, public service, self-employed, students	733	27.3	111	24.9	338	27.2	125	25.2	
Shop assistants, service workers	374	14.0	88	19.8	176	15.5	97	19.5	
Manual workers	338	12.6	59	13.3	159	12.8	68	13.8	
Unemployed	77	2.9	19	4.3	51	4.9	22	4.5	
Missing	176		53		_		_		
Obstetric factors									
Primiparous	1555	54.9	255	52.1	662	51.8	259	52.1	0.88
Missing	22		9		_		_		
Multiple pregnancy	973	34.1	129	25.9	399	29.4	129	25.9	0.16
Infertility treatment	524	18.5	39	8.0	199	14.5	39	8.0	<0.001
Missing	31		12		_		_		
Antenatal steroids									
No	337	11.8	244	49.0	337	49.0	244	49.0	1
Incomplete course	453	15.9	131	26.3	406	26.3	131	26.3	
Complete course	2067	72.3	123	24.7	492	24.7	123	24.7	
Tocolysis	1657	58.2	224	45.3	652	48.1	225	45.3	0.32
Missing	1037	30.2	3	.0.0	-	.0.1	_	.0.0	J.UL
Antenatal administration of magnesium sulfate	266	9.4	6	1.2	105	7.0	6	1.3	<0.001
Missing	32	0.7	9	1.4	_	7.0	_	1.0	\0.001
Premature rupture of membranes	952	33.6	88	18.3	362	27.8	93	18.7	<0.001
Missing	20	55.0	18	10.0	-	27.0	-	10.1	\J.001
Spontanous labour	1369	49.2	314	65.6	625	52.3	320	64.4	<0.001
Missing	75	43.2	19	03.0	-	52.5	- -	04.4	₹0.001
Chorioamnionitis	103	3.7	9	1.9	47	4.3	11	2.2	0.062
Missing	51	0.1	15	1.0	-	7.0	-	۷.۷	0.002

Continued



Table 1 Continued

	Initial co	hort*			Matched	• .			
	Inborn		Outborn		Inborn 1:n (up to 4)		Outborn		
	n=2857	%	n=498	%	n=1235	%	n=498	%	P value‡
Antepartum haemorrhage	150	5.3	53	10.8	81	6.9	54	10.9	0.011
Missing	28		8		_		_		
Pre-eclampsia	513	18.4	51	10.6	234	18.2	54	10.9	<0.001
Missing	65		15		_		_		
Caesarean	1891	66.4	260	53.7	782	61.5	264	53.1	0.002
Missing	9		14		_		_		
Neonatal factors									
Cephalic presentation	1750	62.9	313	67.7	799	66.7	334	67.3	0.81
Missing	76		36		_		_		
Male sex	1503	52.6	265	53.2	651	51.4	265	53.2	0.51
Missing					_		_		
Small-for-gestational age¶	1032	36.1	124	25.0	418	32.7	124	25.0	0.002
Missing	1		2		_		_		

<sup>\*</sup>Denominators vary according to the number of missing data for each variable.

### Patient and public involvement

There was no patient involvement in this study. However, EPIPAGE-2 maintains contact with the cohort through letters, newsletters and its website, and follow-up is ongoing. National parents' associations assist with the dissemination of results.

### **RESULTS**

Of 3816 live born infants enrolled between 24 and 31 weeks of gestation in EPIPAGE-2, 3355 were eligible for inclusion in the study (figure 1); no important differences were seen with the 102 children missing information about antenatal steroid administration (online supplemental table 2). A total of 498 outborn infants were matched with 1235 of the 2857 inborn infants (2.48 inborns per outborn). A comparison of the characteristics of the matched and unmatched inborn infants is contained in online supplemental table 3.

Baseline maternal and newborn characteristics are presented in table 1. Administration of a complete course of steroids was lower for outborns (24.7%) than inborns (72.3%) in the whole population. After matching, important differences remained for maternal age, mother's country of birth, use of infertility treatment, maternal pregnancy complications and obstetric treatments, and in fetal size at birth.

### **Primary outcome**

There was no difference in the primary outcome of survival without moderate or severe neurodevelopmental

impairment at 5.5 years of age using imputed data (adjusted OR 1.09, 95% CI 0.82 to 1.44, p=0.56), as shown in table 2.

### **Secondary outcomes**

Table 2 also shows results for the secondary outcomes. Prior to adjustment, there was a suggestion of improved survival at 5.5 years for inborns, with an unadjusted OR 1.36 (95% CI 0.99 to 1.88). This association was mitigated after accounting for potential confounders. At 2 years corrected age, survival without moderate or severe sensorimotor impairment was higher in the inborns in unadjusted analysis (OR 1.45, 95% CI 1.09 to 1.91), but this was no longer apparent after adjustment (adjusted OR 1.27, 95% CI 0.93 to 1.72). Among survivors at 2 years, there were no differences in CP (adjusted OR 0.66, 95% CI 0.21 to 1.34) or for the ASQ (adjusted OR 1.05,  $95\%\,\mathrm{CI}$  0.78 to 1.42). Prior to hospital discharge, there were no statistically significant differences in survival without severe neonatal morbidity or for individual pathologies except necrotising enterocolitis which was less frequent among inborns (3.6% compared with 5.8%, adjusted OR 0.56, 95% CI 0.32 to 0.98).

### **Sensitivity analyses**

No statistically significant effects were seen in sensitivity analyses among those alive at the beginning of labour or when the population was restricted to singletons only. Among those born at 24–27 weeks' GA, there was evidence of higher survival without moderate-to-severe impairment at 5.5 years of age for inborns compared with

<sup>†</sup>Result after multiple imputation for missing data and weighted for the variable matching ratio.

<sup>‡</sup>P values calculated using conditional logistic regression.

<sup>§</sup>Defined as the highest occupational status of the mother and father, or mother only if living alone.

<sup>¶</sup>Small-for-gestational age was defined as birth weight less than the 10th percentile for gestational age and sex based on French intrauterine 'EPOPé' growth curves (Ego 2016).



**Table 2** Outcomes at age 5.5 years in a matched study of inborn and outborn children from the EPIPAGE-2 cohort according to gestational age at birth (analysis using data after multiple imputation)

	Matched cohort 1:n (up to 4)*								
	Inborn		Outbor	n	Inborn versus outborn				
	n=1235	%	n=498	%	Unadjusted OR (95% CI)†	P value	Adjusted OR (95% CI)†‡	P value	
Primary outcomes									
Survival without severe or moderate disabilities at 5.5 years§	876	69.1	322	64.8	1.21 (0.94 to 1.57)	0.140	1.09 (0.82 to 1.44)	0.560	
Secondary outcomes									
Survival at 5.5 years	1068	85.6	405	81.3	1.36 (0.99 to 1.88)	0.056	1.15 (0.83 to 1.58)	0.400	
Among survivors	n=1068		n=405						
Severe or moderate disabilities at 5.5 years§	191	19.3	82	20.3	0.94 (0.66 to 1.35)	0.730	0.97 (0.66 to 1.42)	0.870	
Cerebral palsy GMFCS-2/5	46	4.8	25	6.4	0.74 (0.41 to 1.36)	0.340	0.78 (0.42 to 1.46)	0.440	
FSIQ <79 (<-2 SD¶)	157	16.0	66	16.5	0.97 (0.65 to 1.45)	0.870	0.98 (0.63 to 1.53)	0.940	
Moderate-to-severe visual impairment	24	2.7	10	2.5	1.09 (0.40 to 2.96)	0.870	1.11 (0.41 to 3.02)	0.840	
Moderate-to-severe hearing impairment	9	1.0	4	1.1	0.86 (0.24 to 3.13)	0.820	1.11 (0.29 to 4.28)	0.880	
Survival at discharge without severe neonatal morbidity**	910	71.5	333	67.0	1.24 (0.97 to 1.58)	0.080	1.08 (0.83 to 1.40)	0.570	
Among survivors	n=1072		n=407						
Severe cerebral lesion	65	7.0	26	6.6	1.07 (0.65 to 1.75)	0.800	1.22 (0.72 to 2.07)	0.470	
Necrotising enterocolitis	41	3.6	23	5.8	0.60 (0.35 to 1.04)	0.069	0.56 (0.32 to 0.98)	0.043	
Severe bronchopulmonary dysplasia	61	6.7	27	6.7	1.02 (0.61 to 1.69)	0.950	0.99 (0.57 to 1.70)	0.960	
Severe retinopathy of prematurity	10	1.4	4	1.0	1.35 (0.40 to 4.60)	0.630	1.34 (0.30 to 5.90)	0.700	
Survival without moderate or severe neuromotor or sensory disabilities at 2 years CA††	1026	81.9	377	75.8	1.45 (1.09 to 1.91)	0.010	1.27 (0.93 to 1.72)	0.130	
Among survivors	n=1068		n=405						
Cerebral palsy GMFCS-2/5	30	3.2	19	4.9	0.64 (0.32 to 1.30)	0.220	0.66 (0.21 to 1.34)	0.250	
Moderate-to-severe visual impairment	10	1.0	3	0.9	1.11 (0.13 to 9.14)	0.920	1.28 (0.15 to 10.85)	0.820	
Moderate-to-severe hearing impairment	4	0.5	7	1.8	0.27 (0.06 to 1.22)	0.089	0.28 (0.04 to 2.05)	0.210	
ASQ below threshold‡‡	409	41.0	149	40.7	1.01 (0.76 to 1.35)	0.930	1.05 (0.78 to 1.42)	0.750	

<sup>\*</sup>Result after multiple imputation for missing data and weighted for the variable matching ratio.

outborns (adjusted OR 1.71, 95% CI 1.05 to 2.81) but not for those born at 28–31 weeks' gestation (table 3). Between 26 and 31 weeks gestation, there was a trend towards increasing ORs with decreasing GA although all of the 95% CI included 1.0 except for at 26 weeks where the 95% CI in adjusted analysis was 1.01 to 8.24 (p=0.048). Below 26 weeks, there was no evidence of a difference

between inborns and outborns (online supplemental table 4).

Results among cases with complete data (table 4) were broadly the same as the main analysis other than for CP, which showed a markedly lower odds at ages 2 (adjusted OR 0.22, 95% CI 0.10 to 0.47) and at age 5.5 (adjusted OR 0.41, 95% CI 0.19 to 0.87) years and for hearing

<sup>†</sup>ORs were calculated using conditional logistic regression.

<sup>‡</sup>OR adjusted for maternal age, mother born in France, parents' socioeconomic status, primiparous, infertility treatment, multiple pregnancy, tocolysis, antenatal administration of magnesium sulfate, context of preterm birth, caesarean section, cephalic presentation, small-for-gestational age and sex.

<sup>§</sup>Defined as at least one of: severe or moderate cerebral palsy (GMFCS level 2-5), visual (bilateral binocular visual acuity <3.2/10) or hearing (unilateral-bilateral hearing loss > 40dB not corrected or partially corrected with hearing aid) impairment, or FSIQ <2 SDs below the mean of the reference group born at term (Pierrat et al<sup>21</sup> 2021).

<sup>¶</sup>Cut-off of the distribution related to the reference group born at term (Pierrat et al<sup>21</sup> 2021).

<sup>\*\*</sup>Severe bronchopulmonary dysplasia or necrotising enterocolitis stage 2–3 or severe retinopathy of prematurity stage >3 or any of the following severe cerebral abnormalities on cranial ultrasonography: intraventricular haemorrhage grade III or IV or cystic periventricular leukomalacia (Ancel et al<sup>20</sup> 2015).

<sup>††</sup>Cerebral palsy GMFCS levels 2–5 and/or unilateral or bilateral deafness and/or unilateral or bilateral blindness.

<sup>‡‡</sup>Below threshold of the United States ASQ-3 reference. Among children without cerebral palsy, deafness or blindness.

ASQ, Ages and Stages Questionnaire; CA, corrected age; FSIQ, Full Scale Intelligence Quotient; GMFCS, Gross Motor Function Classification System.



**Table 3** Sensitivity analyses for survival without moderate-to-severe impairment at age 5.5 years in a matched study of inborn and outborn children from the EPIPAGE-2 cohort according to gestational age at birth (analysis using data after multiple imputation)

	Matched	cohort	1:n (up to	4)*						
	Inborn		Outborn		Inborn versus outborn					
	n/N	%	n/N	%	Unadjusted OR (95% CI)†	P value	Adjusted OR (95% CI)†‡	P value		
Survival without severe or m	noderate dis	abilities	at 5.5 year	's§						
24-27 weeks	162/361	46.7	56/161	35.0	1.63 (1.06 to 2.51)	0.026	1.71 (1.05 to 2.81)	0.032		
28-31 weeks	711/874	79.8	266/337	79.1	1.04 (0.70 to 1.54)	0.840	0.95 (0.62 to 1.45)	0.810		
24–31 SA singleton	614/868	68.9	240/369	65.1	1.19 (0.88 to 1.60)	0.260	1.05 (0.75 to 1.46)	0.780		
24–31 SA including alive at beginning of labour	877/1282	64.9	322/536	60.2	1.22 (0.96 to 1.56)	0.110	1.11 (0.85 to 1.45)	0.460		

<sup>\*</sup>Result after multiple imputation for missing data and weighted for the variable matching ratio. Matching were performed in each subgroup. †ORs were calculated using conditional logistic regression.

impairment at 2 years of age (adjusted OR 0.17, 95% CI 0.07 to 0.44).

### **DISCUSSION**

### **Principal results**

When GA and antenatal steroids are accounted for, in a large national cohort with neurological follow-up to 5.5 years of age, no differences were identified in survival without moderate or severe neurological impairment for very preterm outborn infants compared with inborns, nor were any differences seen in individual components. This was also the case at 2 years of age. With the exception of necrotising enterocolitis, no differences were seen in neonatal morbidities.

### **Strengths and limitations**

The major strength of the EPIPAGE 2 study is that it is a large national cohort with detailed, prospective, longterm follow-up to preschool age carried out by trained investigators. Other recent publications<sup>2 6 9 10 28</sup> have looked exclusively at mortality and neonatal neurological complications, which do not necessarily extrapolate to later functional ability. The follow-up rate—around 70% at 5.5 years—means our study compares favourably with other studies of outborns, for example, 62% of outborns and 69% of inborns at 1 year of age in Western Australia<sup>7</sup>; 57% of outborns and 77% of inborns were seen at 2–3 years in a cohort from New South Wales<sup>11</sup>; and under 50% were seen at 3 years of age in a Japanese study.<sup>29</sup> However, the loss to follow-up is potentially a limitation: not all children completed all of the tests, meaning that there were up to 40% missing data for some components of the outcome; this was addressed

using multiple imputation, which has been shown to be robust even when the missing-at-random assumption is violated.<sup>30</sup>

Another strength is the inclusion of all live births, plus the sensitivity analysis accounting for the entire population of at risk fetuses alive at maternal admission to hospital. Results from studies not including these data are difficult to interpret. <sup>2</sup> 10 12 13 For instance, one recent study which did not include pretransfer deaths used proportional hazards regression after propensity score matching and reported increased in-hospital mortality for inborns.<sup>2</sup> Our analyses showed no differences between the inborn and outborn groups regardless of which baseline population was used. We were, however, limited by an inability to consider factors like the Apgar score or cord gas results that reflect status at birth and which may be less than ideal in non-level 3 units due to suboptimal obstetric care. We also could not evaluate the potential impact of the presence of paediatric transport teams in non-level 3 units prior to delivery as these data were not collected. Furthermore, despite the importance of studying long-term outcomes, management of infants in the EPIPAGE-2 cohort may no longer be representative of current care standards. For instance, there is an increasing emphasis on the neonatal 'Golden Hour', 31 which may not have been uniformly implemented at the time in non-level-3 units. We unfortunately did not have the data to examine this, nor do we have more recent data with which to evaluate evolution of practices over time. However, long-term follow-up can only be based on historic practices; hence, the importance of trying to understand which factors might affect later outcomes.

<sup>‡</sup>Adjusted OR on maternal age, mother born in France, parents' socioeconomic status, primiparous, infertility treatment, multiple pregnancy, tocolysis, antenatal administration of magnesium sulfate, context of preterm birth, caesarean section, cephalic presentation, small-forgestational age and sex.

<sup>§</sup>Severe or moderate cerebral palsy (Gross Motor Function Classification System level-2/5), visions (bilateral binocular visual acuity<3.2/10), hearing (unilateral-bilateral hearing loss >40 dB not corrected or partially corrected with hearing aid) and full-scale IQ<2 SD below the mean of the reference sample born at term (See Pierrat et al 2021).



**Table 4** Outcomes at age 5.5 years in a matched study of inborn and outborn children from the EPIPAGE-2 cohort according to gestational age at birth (analysis using only subjects with complete data)

	Matched cohort 1:n (up to 4)*								
	Inborn		Outborn		Inborn vs outborn				
	n=1235	%	n=498	%	Unadjusted OR (95% CI)†	P value	Adjusted OR (95% CI)†‡	P value	
Primary outcomes									
Survival without severe or moderate disabilities at 5.5 years§	551/820	64.5	188/327	57.5	1.34 (1.03 to 1.76)	0.030	1.16 (0.84 to 1.61)	0.380	
Secondary outcomes									
Survival at 5.5 years	1068/1235	85.6	405/498	81.3	1.36 (0.99 to 1.88)	0.056	1.18 (0.82 to 1.71)	0.380	
Among survivors	n=1068		n=405						
Severe or moderate disabilities at 5.5 years§	102/653	17.7	46/234	19.7	0.86 (0.62 to 1.21)	0.390	0.90 (0.59 to 1.37)	0.610	
Cerebral palsy GMFCS-2/5	23/748	3.2	15/274	5.5	0.58 (0.33 to 1.00)	0.052	0.41 (0.19 to 0.87)	0.021	
FSIQ<79 (<-2 SD¶)	82/650	14.5	33/230	14.3	1.01 (0.68 to 1.50)	0.960	1.00 (0.60 to 1.67)	0.990	
Moderate-to-severe visual impairment	6/655	0.8	2/230	0.9	0.98 (0.20 to 4.35)	0.980	2.32 (0.20 to 26.4)	0.500	
Moderate-to-severe hearing impairment	6/731	1.0	4/265	1.5	0.64 (0.20 to 2.03)	0.450	1.29 (0.30 to 5.57)	0.730	
Secondary exploratory outcomes									
Survival at discharge without severe neonatal morbidity**	866/1235	68.1	314/498	63.1	1.25 (0.97 to 1.61)	0.086	1.17 (0.87 to 1.58)	0.300	
Among survivors	n=1072		n=407						
Severe cerebral lesion	64/1057	7.0	26/396	6.6	1.06 (0.70 to 1.61)	0.770	1.34 (0.75 to 2.41)	0.320	
Necrotising enterocolitis	40/1059	3.5	23/403	5.7	0.60 (0.35 to 1.03)	0.066	0.45 (0.24 to 0.84)	0.013	
Severe bronchopulmonary dysplasia	57/1042	6.6	24/395	6.1	1.08 (0.67 to 1.76)	0.740	0.85 (0.48 to 1.50)	0.570	
Severe retinopathy of prematurity	10/1064	1.4	4/404	1.0	1.37 (0.61 to 3.05)	0.440	2.02 (0.52 to 7.75)	0.310	
Survival without severe or moderate neuromotor or sensory disabilities at 2 years CA††	1047/1235	83.8	385/498	77.3	1.51 (1.12 to 2.05)	0.008	1.55 (1.10 to 2.20)	0.014	
Among survivors	n=1068		n=405						
Cerebral palsy GMFCS-2/5	15/890	2.0	15/320	4.7	0.41 (0.22 to 0.74)	0.003	0.22 (0.10 to 0.47)	<0.001	
Moderate-to-severe visual impairment	3/838	0.3	1/292	0.3	0.74 (0.23 to 2.33)	0.600	1.43 (0.58 to 3.48)	0.440	
Moderate-to-severe hearing impairment	3/874	0.4	6/309	1.9	0.22 (0.10 to 0.50)	<0.001	0.17 (0.07 to 0.44)	<0.001	
ASQ below threshold‡‡	321/801	40.1	105/269	39.0	1.04 (0.80 to 1.37)	0.760	1.02 (0.74 to 1.42)	0.880	

<sup>\*</sup>Result after multiple imputation for missing data and weighted for the variable matching ratio.

### Interpretation

The decision to match for antenatal steroids was based on their association with reductions in neonatal death and intraventricular haemorrhage. <sup>18</sup> Previous studies of outborn preterm infants have noted similar patterns of

increased neonatal deaths and neurological complications and the disproportional use of steroids in inborn versus outborn births.<sup>3 5–10 15 17</sup> Our two-step methodology, matching then adjustment, has the advantage of initially neutralising confounding by matching steroids

<sup>†</sup>ORs were calculated using conditional logistic regression.

<sup>‡</sup>OR adjusted for maternal age, mother born in France, parents' socioeconomic status, primiparous, infertility treatment, multiple pregnancy, tocolysis, antenatal administration of magnesium sulfate, context of preterm birth, caesarean section, cephalic presentation, small-for-gestational age and sex.

<sup>§</sup>Defined as at least one of: severe or moderate cerebral palsy (GMFCS level 2-5), visual (bilateral binocular visual acuity <3.2/10) or hearing (unilateral-bilateral hearing loss > 40 dB not corrected or partially corrected with hearing aid) impairment, or FSIQ <2 SDs below the mean of the reference group born at term (Pierrat et al<sup>21</sup> 2021).

<sup>¶</sup>Cut-off of the distribution related to the reference group born at term (Pierrat et al<sup>21</sup> 2021).

<sup>\*\*</sup>Severe bronchopulmonary dysplasia or necrotising enterocolitis stage 2–3 or severe retinopathy of prematurity stage >3 or any of the following severe cerebral abnormalities on cranial ultrasonography: intraventricular haemorrhage grade III or IV or cystic periventricular leukomalacia (Ancel et al<sup>20</sup> 2015).

<sup>††</sup>Cerebral palsy GMFCS levels 2–5 and/or unilateral or bilateral deafness and/or unilateral or bilateral blindness.

<sup>‡‡</sup>Below threshold of the United States ASQ-3 reference. Among children without cerebral palsy, deafness or blindness.

ASQ, Ages and Stages Questionnaire; CA, corrected age; FSIQ, Full Scale Intelligence Quotient; GMFCS, Gross Motor Function Classification System.



at the expense of tending to harmonise the characteristics of inborns and outborns. However, despite the matching process, the inborn group retained a similar predominance of 'stable' pathologies such as pre-eclampsia, <sup>8</sup> intrauterine growth retardation6, <sup>8</sup> <sup>10</sup> preterm PROM<sup>6</sup> <sup>10</sup> and chorioamnionitis <sup>10</sup> as seen in other studies. Similarly, the outborn deliveries featured more 'unstable' pathologies such as placental abruption6 and delivery was more frequent following spontaneous labour. <sup>6</sup>

The disadvantage of matching is the undesirable, and foreseeable, effect of potentially introducing a selection bias because of the assumption that there is no random error attributable to sampling variability, which is unlikely to be true. Differences exist between the matched and unmatched inborns for several antenatal characteristics—particularly, the use of antenatal steroids which was higher for inborns who were not matched. There were also higher levels of tocolysis, small for GA and PROM in unmatched inborns, but more pre-eclampsia among those who were matched.

The second step of our method therefore was to adjust the point estimate using maternal and pregnancy characteristics to reduce the potential selection bias induced by matching. Both before and after adjustment, we found no difference in outcome between the groups. This suggests that, after accounting for other factors, differences in antenatal steroid administration between inborn and outborn births play an important role in influencing long-term outcomes, and that public health policies should promote greater administration of antenatal steroids for women presenting with threatened preterm delivery.

Our sensitivity analyses looked at excluding both the lowest and highest GA infants. The high level of mortality observed among outborn infants of the lowest GAs has previously been reported. The 24-31 weeks' GA population had an estimated OR for the principal outcome not far from one. When we looked at only the 24-27 week infants, the adjusted estimate showed improved outcomes for inborns. Analyses performed by week of GA suggested a trend favouring inborn infants at lower GAs, although not at the GAs of 24 and 25 weeks for which attitudes in France in 2011 were known to be more mixed.<sup>33 34</sup> These results are consistent with other studies that suggest benefits of antenatal transfer are greater with decreasing GA.<sup>1 7 35</sup> Moreover, while over 80% of children were inborn and a regionalised system of care has existed in France since the 1990s,<sup>23</sup> there remain a subset of outborns that could potentially be avoided.<sup>34</sup>

At 2 years of age, there were no differences in outcomes between our inborn and outborn groups, similar to two smaller Australian studies that looked at infants of 23–28 weeks' GA.<sup>11</sup> and 23–25 weeks' GA.<sup>7</sup> Similarly, there were no differences in neonatal outcomes other than for necrotising enterocolitis which was of borderline statistical significance (p=0.043) and may be a chance finding due to multiple testing; this warrants further investigation.

### Generalisability

There is no consensus regarding which GAs to include in studies of outborns, with other studies using different GA ranges, <sup>2 6 8-11 14 16 17 28</sup> and some studies instead using a birth weight cut-off. <sup>3-5 36</sup> We included infants of 24–31 weeks due to the recommendation that births of less than 32 weeks should occur in a Level-3 unit<sup>22</sup> and because of the known poor outcomes of births at <24 weeks in France. <sup>19</sup> Many European and other high-income countries have similar guidelines and approaches to those of France in 2011. <sup>37</sup> Our results should, therefore, be relevant for such countries too.

### **Conclusions**

Overall, there was no difference in longer-term survival to 5.5 years of age without moderate or severe neurological impairment between inborn and outborn very preterm children when GA and maternal antenatal steroids were accounted for. This suggests that antenatal steroids may be an important determinant of differences in outcome found between inborns and outborns in other studies. For infants born at the lowest GAs, however, an improved outcome was associated with birth in level 3 hospitals; this should motivate the implementation of improved protocols to enhance antenatal transfer of women delivering at these gestations.

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### **REFERENCES**

- 1 Lasswell SM, Barfield WD, Rochat RW, et al. Perinatal regionalization for very low-birth-weight and very preterm infants: a meta-analysis. JAMA 2010;304:992–1000.
- 2 Fang JL, Mara KC, Weaver AL, et al. Outcomes of outborn extremely preterm neonates admitted to a NICU with respiratory distress. Arch Dis Child Fetal Neonatal Ed 2020;105:33–40.
- 3 Mohamed MA, Aly H. Transport of premature infants is associated with increased risk for intraventricular haemorrhage. Arch Dis Child Fetal Neonatal Ed 2010;95:F403-7.
- 4 Rautava L, Lehtonen L, Peltola M, et al. The effect of birth in secondary- or tertiary-level hospitals in Finland on mortality in very preterm infants: a birth-register study. Pediatrics 2007;119:e257–63.
- 5 Moro M, Figueras-Aloy J, Fernández C, et al. Mortality for newborns of birthweight less than 1500 G in Spanish neonatal units (2002-2005). Am J Perinatol 2007;24:593–601.
- 6 Boland RA, Dawson JA, Davis PG, et al. Why birthplace still matters for infants born before 32 weeks: Infant mortality associated with



- birth at 22-31 weeks' gestation in non-tertiary hospitals in Victoria over two decades. *Aust N Z J Obstet Gynaecol* 2015;55:163–9.
- 7 Thompson K, Gardiner J, Resnick S. Outcome of outborn infants at the borderline of viability in Western Australia: a retrospective cohort study. J Paediatr Child Health 2016;52:728–33.
- 8 Chien LY, Whyte R, Aziz K, et al. Improved outcome of preterm infants when delivered in tertiary care centers. Obstet Gynecol 2001;98:247–52.
- 9 Marlow N, Bennett C, Draper ES, et al. Perinatal outcomes for extremely preterm babies in relation to place of birth in England: the EPICure 2 study. Arch Dis Child Fetal Neonatal Ed 2014;99:F181–8.
- Helenius K, Longford N, Lehtonen L, et al. Association of early postnatal transfer and birth outside a tertiary hospital with mortality and severe brain injury in extremely preterm infants: observational cohort study with propensity score matching. BMJ 2019;367:I5678.
- 11 Mahoney K, Bajuk B, Oei J, et al. Risk of neurodevelopmental impairment for outborn extremely preterm infants in an Australian regional network. J Matern Fetal Neonatal Med 2017;30:96–102.
- Hirata K, Kimura T, Hirano S, et al. Outcomes of outborn very-low-birth-weight infants in Japan. Arch Dis Child Fetal Neonatal Ed 2021:106:131–6.
- 13 Amer R, Moddemann D, Seshia M, et al. Neurodevelopmental outcomes of infants born at <29 weeks of Gestation admitted to Canadian neonatal Intensive care units based on location of birth. J Pediatr 2018;196:31–7.
- 14 Boland RA, Davis PG, Dawson JA, et al. Outcomes of infants born at 22-27 weeks' gestation in Victoria according to outborn/inborn birth status. Arch Dis Child Fetal Neonatal Ed 2017;102:F153–61.
- 15 Lee SK, McMillan DD, Ohlsson A, et al. The benefit of preterm birth at tertiary care centers is related to gestational age. Am J Obstet Gynecol 2003;188:617–22.
- 16 Lui K, Abdel-Latif ME, Allgood CL, et al. Improved outcomes of extremely premature outborn infants: effects of strategic changes in perinatal and retrieval services. *Pediatrics* 2006;118:2076–83.
- 17 Palmer KG, Kronsberg SS, Barton BA, et al. Effect of inborn versus outborn delivery on clinical outcomes in ventilated preterm neonates: secondary results from the NEOPAIN trial. J Perinatol 2005;25:270–5.
- 18 Roberts D, Brown J, Medley N, et al. Antenatal corticosteroids for accelerating fetal lung maturation for women at risk of preterm birth. Cochrane Database Syst Rev 2017.
- 19 Ancel P-Y, Goffinet F, EPIPAGE 2 Writing Group. EPIPAGE 2: a preterm birth cohort in France in 2011. BMC Pediatr 2014;14:97.
- 20 Ancel P-Y, Goffinet F, et al, EPIPAGE-2 Writing Group. Survival and morbidity of preterm children born at 22 through 34 weeks' gestation in France in 2011: results of the EPIPAGE-2 cohort study. JAMA Pediatr 2015;169:230–8.
- 21 Pierrat V, Marchand-Martin L, Marret S, et al. Neurodevelopmental outcomes at age 5 among children born preterm: EPIPAGE-2 cohort study. BMJ 2021;373:n741.
- 22 Hébart T, Serfaty A. Haute Autorité de Santé / Service des bonnes pratiques professionnelles. Grossesses risque: Orientation des femmes enceintes entre les maternités en vue de l'accouchement. In: Haute Autorité de Santé; Haute Autorité de Santé, 2009.

- 23 Ministère de l'emploi et de la solidarité. Décret no 98-899 et no 98-900 du 9 octobre 1998. Journal Officiel de la République Française 1998;235:15344 https://www.legifrance.gouv.fr/jorf/id/ JORFTEXT000000756322/
- 24 Ghassabian A, Sundaram R, Bell E, et al. Gross motor milestones and subsequent development. *Pediatrics* 2016;138.
- 25 Wechsler D. WPPSI-IV Échelle d'intelligence de Wechsler pour enfants, 4th ed, 2014. Pearson Clinical & Talent Assessment. Available: https://www.pearsonclinical.fr/wppsi-iv-echelledintelligence-de-wechsler-pour-la-periode-pre-scolaire-et-primairequatrieme-edition [Accessed 12 December 2022].
- 26 Pierrat V, Marchand-Martin L, Arnaud C, et al. Neurodevelopmental outcome at 2 years for preterm children born at 22 to 34 weeks' gestation in France in 2011: EPIPAGE-2 cohort study. BMJ 2017;358:j3448.
- 27 Rubin DB. *Multiple imputation for nonresponse in surveys*. USA: John Wiley & Sons, Inc, 1987. ISBN: 978-0-471-65574-9.
- 28 Hossain S, Shah PS, Ye XY, et al. Outborns or inborns: where are the differences? A comparison study of very preterm neonatal intensive care unit infants cared for in Australia and New Zealand and in Canada. Neonatology 2016;109:76–84.
- 29 Sasaki Y, Ishikawa K, Yokoi A, et al. Short- and long-term outcomes of extremely preterm infants in Japan according to outborn/inborn birth status. Pediatr Crit Care Med 2019;20:963–9.
- 30 Piedvache A, van Buuren S, Barros H, et al. Strategies for assessing the impact of loss to follow-up on estimates of neurodevelopmental impairment in a very preterm cohort at 2 years of age. BMC Med Res Methodol 2021;21:118.
- 31 Hodgson KA, Owen LS, Lui K, et al. Neonatal golden hour: a survey of Australian and New Zealand neonatal network units' early stabilisation practices for very preterm infants. J Paediatr Child Health 2021;57:990–7.
- 32 Rose S, Laan MJvander, van der Laan MJ. Why match? Investigating matched case-control study designs with causal effect estimation. Int J Biostat 2009;5:Article 1.
- 33 Morgan AS, Foix L'Helias L, Diguisto C, et al. Intensity of perinatal care, extreme prematurity and sensorimotor outcome at 2 years corrected age: evidence from the EPIPAGE-2 cohort study. BMC Med 2018;16:227.
- 34 Desplanches T, Morgan AS, Jones P, et al. Risk factors for very preterm delivery out of a level III maternity unit: the EPIPAGE-2 cohort study. Paediatr Perinat Epidemiol 2021;35:694–705.
- 35 Watson SI, Arulampalam W, Petrou S, et al. The effects of designation and volume of neonatal care on mortality and morbidity outcomes of very preterm infants in England: retrospective population-based cohort study. BMJ Open 2014;4:e004856.
- 36 Chung M-Y, Fang P-C, Chung C-H, et al. Comparison of neonatal outcome for inborn and outborn very low-birthweight preterm infants. *Pediatr Int* 2009;51:233–6.
- 37 Guillén Úrsula, Weiss EM, Munson D, et al. Guidelines for the management of extremely premature deliveries: a systematic review. Pediatrics 2015;136:343–50.