

Citation: Angulo JC, Schönburg S, Giammò A, Abellán FJ, Arance I, Lora D (2019) Systematic review and meta-analysis comparing Adjustable Transobturator Male System (ATOMS) and Adjustable Continence Therapy (ProACT) for male stress incontinence. PLoS ONE 14(12): e0225762. https://doi.org/10.1371/journal.pone.0225762

Editor: Peter F.W.M. Rosier, University Medical Center Utrecht, NETHERLANDS

Received: September 17, 2019

Accepted: November 12, 2019

Published: December 2, 2019

Peer Review History: PLOS recognizes the benefits of transparency in the peer review process; therefore, we enable the publication of all of the content of peer review and author responses alongside final, published articles. The editorial history of this article is available here: https://doi.org/10.1371/journal.pone.0225762

Copyright: © 2019 Angulo et al. This is an open access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Data Availability Statement: All relevant data are within the paper and its Supporting Information files.

RESEARCH ARTICLE

Systematic review and meta-analysis comparing Adjustable Transobturator Male System (ATOMS) and Adjustable Continence Therapy (ProACT) for male stress incontinence

Javier C. Angulo^{1,2®}*, Sandra Schönburg³, Alessandro Giammò⁴, Francisco J. Abellán¹, Ignacio Arance^{1,2}, David Lora^{5,6,7®}

1 Departamento Clínico, Universidad Europea de Madrid, Madrid, 2 Servicio de Urología, Hospital Universitario de Getafe, Madrid, Spain, 3 Department of Urology and Kidney Transplantation, Martin Luther University, Halle (Saale), Germany, 4 Department of Neuro-Urology, CTO-Spinal Unit, Città della Salute e della Scienza di Torino, Turin, Italy, 5 Instituto de Investigación Sanitaria Hospital "12 de Octubre" (i+12), Madrid, Spain, 6 CIBER de Epidemiología y Salud Pública (CIBERESP), Madrid, Spain, 7 Universidad Complutense de Madrid, Madrid, Spain

So These authors contributed equally to this work.

* javier.angulo@salud.madrid.org

Abstract

Background and purpose

Urinary incontinence is one of the most serious complications of prostate cancer treatment. The objective of this study was to assess efficacy and safety of Adjustable Transobturator Male System (ATOMS) compared to Adjustable Continence Therapy (proACT) for male stress urinary incotinence according to literature findings.

Material and methods

A systematic review and meta-analysis on adjustable devices ATOMS and ProACT is presented. Studies on female or neurogenic incontinence were excluded. Differences between ATOMS and proACT in primary objective: dryness status (no-pad or one safety pad/day) after initial device adjustment, and in secondary objectives: improvement, satisfaction, complications and device durability, were estimated using random-effect model. Statistical heterogeneity among studies included in the meta-analysis was assessed using tau2, Higgins's I2 statistics and Cochran's Q test.

Results

Combined data of 41 observational studies with 3059 patients showed higher dryness (68 vs. 55%; p = .01) and improvement (91 vs. 80%; p = .007) rate for ATOMS than ProACT. Mean pad-count (-4 vs. -2.5 pads/day; p = .005) and pad-test decrease (-425.7 vs. -211.4 cc; p < .0001) were also significantly lower. Satisfaction was higher for ATOMS (87 vs. 56%; p = .002) and explant rate was higher for proACT (5 vs. 24%; p < .0001). Complication rate for ProACT was also higher, but not statistically significant (17 vs. 26%; p = .07). Mean **Funding:** The authors received no specific funding for this work.

Competing interests: The authors have declared that no competing interests exist.

follow-up was 25.7 months, lower for ATOMS than ProACT (20.8 vs. 30.6 months; p = .02). The rate of working devices favoured ATOMS at 1-year (92 vs. 76; p < .0001), 2-years (85 vs. 61%; p = .0008) and 3-years (81 vs. 58%; p = .0001). Significant heterogeneity was evidenced, due to variable incontinence severity baseline, difficulties for a common reporting of complications, different number of adjustments and time of follow-up and absence of randomized studies.

Conclusions

Despite the limitations that studies available are exclusively descriptive and the follow-up is limited, literature findings confirm ATOMS is more efficacious, with higher patient satisfaction and better durability than ProACT to treat male stress incontinence.

Introduction

Urinary incontinence can severely impact the quality of life of prostate cancer survivors after radical prostatectomy and/or radiotherapy [1,2]. The artificial urinary sphincter (AUS) has been the main therapy after failed conservative treatment but reposition of the posterior ure-thra using retrobulbar slings has been a great revolution [3]. Also development of new devices usig postoperative adjustment has allowed a more personalized approach. Systematic reviews on the topic tend to annalyze a variety of devices with different modes of action [4–7].

Adjustable continence therapy (ProACT[®], Uromedica, Minneapolis, MN) using balloons placed periurethraly near the bladder neck was the first adjustable implant to treat stress incontinence. It consists of two volume-adjustable balloons, that increase urethral resistance during voiding. Adjustable transobturator male system (ATOMS[®], A.M.I., Feldkirch, Austria) is another adjustable device that compresses the bulbar urethra ventrally but, unlike ProACT, its two-arm mesh allows firm transobturator fixation on both sides of the ischiopubic bone. Both devices can be adjusted postoperatively without need for aneasthesia. Besides, they open a new perspective of treatment for patients with cognitive impairment and/ or limited dexterity, as no manipulation is needed for voiding [8]. Unfortunately, prospective controlled studies comparing anti-incontinence devices are lacking [9].

ATOMS and ProACT share two important peculiarities. Their mode of action is through urethral compression, and both are easily adjusted postoperatively by percutaneous injection of sterile saline solution until continence or a maximun filling volume is accomplished. Their main differences stand in that compression is ventral in ATOMS and lateral (both sides) in ProACT, and also that ATOMS is fixed to the pelvis while ProACT is a non-fixed device. ProACT has longer been used as its first description for male stress incontinence was in 2005 [10] and ATOMS started in 2011 [8]. ProACT remains unchanged but several modifications have been undertaken in ATOMS port location and design, evolving from inguinal to scrotal and silicone-covered.

Randomized comparative studies between ProACT or ATOMS have not been performed to date but a number of individual or multi-center studies with same or very similar endpoints covered are available for both devices. With this study we aim to evaluate the current evidence on the efficacy and safety ATOMS compared to ProACT using a systematic review approach and meta-analysis.

Materials and methods

A systematic review of the scientific literature was carried out in September 2019. Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) checklist is included (S1 Table). The search included studies published between January 2005 and August 2019. The search was undertaken in PubMed, Embase, Web of Science and Scopus. Search strategy was designed according to PICOS criteria (Population, Intervention, Comparator, Outcomes and Study Design) (Table 1) for the identification of studies using free and controlled terminology. Search strategy included the terms: "Urinary incontinence" AND "male" AND "ATOMS" OR "ProACT". A manual revision of the bibliographic references was also carried out. We included prospective and retrospective, multi-centre or single-centre case series, published in English, Spanish, German or Italian on patients treated with ATOMS or ProACT devices for SUI after prostate surgery. No clinical trial on the topic could be found. Duplicate studies, editorial comments, letters to the editor or expert opinions in non-systematic reviews, a report after cystoprostatectomy and a meta-analysis centered only on ATOMS were excluded.

The identification, selection and extraction of data from the studies were carried out exhaustively and independently by two reviewers. A first selection of the studies was performed by reading the title and the abstract, and those that met the inclusion criteria were reviewed by reading the full text. Disagreements were resolved by consensus or in collaboration with another member of the research team. Consecutive publications from the same groups were evaluated to avoid overlapping of patients in different studies. Whenever repetition was detected we only used the last publication with larger number of patients. The evidence was summarized with the data extracted from the studies that included bibliographic information, characteristics of the study and patients, characteristics of the intervention, and outcome measures in relation to the efficacy and safety of the procedure. The definitions of postoperative dryness and improvement were comparable among studies. Differential pad-count or differential 24h pad-test was calculated as after adjustment minus baseline when available, assuming both values are non-dependent. Safety was defined according to rate of explantation and complications.

Primary efficacy indicator was percentage of dry patients after adjustment, defined as patients using no pad or only one safety pad/day (PPD). When dry rate change evolution in time was described in a study only the initial report was used. Main secondary efficacy outcome was percentage of overall improvement (defined as \geq 50% decrease in pad-count and/or in 24h pad-test). Differential pad-count and differential 24h pad-test between baseline and after adjustment were investigated to evaluate the magnitude effect. The proportion of patients satisfied with the intervention, complication rate, explantation rate, number of adjustments, follow-up after implantation and durability of the devices were also evaluated as secondary

Population	Males with mild, moderate or severe stress urinary incontinence after prostate surgery, eithr previously radiated or not and treated primarily or after failure of other surgical devices
Intervention	Placement of ATOMS ® device
Comparison	Placement of ProACT (R) device
Outcomes	Primary: Overall dryness rate (no pad or one security pad per day) Secondary: Overall improvement rate, differential pad-count and/or pad-test (after adjustment with respects to baseline), complication rate and durability
Study Design	Retrospective and prospective case series

Table 1. PICOS criteria to guide the systematic review.

outcomes. We also included an analysis of manuscript quality assessment using the Newcastle-Ottawa scale.

The heterogeneity assumption was evaluated by the chi-square-based Cochran's Q test (which was considered significant at p-value < .05), and quantified with the I2 statistic (with values <25%, 25% to 75%, and >75% interpreted as representing low, moderate and high levels of heterogeneity, respectively), and with tau2 (the between-study variance) that was obtained with the DerSimonian and Laird method. Due to the presence of heterogeneity, random-effects model with the inverse variance method was used for pooling of single proportions, single means or mean differences from primary studies. The pooled effect was described with 95% confidence interval (CI). Publication bias was assessed by visual inspection of funnel plots and quantified by the Egger's linear regression test. Statistical analysis and figures were performed with the meta package of R software version 3.4.1 (R Foundation for Statistical Computing, Vienna, Austria) [11].

Results

Literature search

The search initially provided 174 references, 61 of which dealt with the topic evaluated. Additional records identified 6 further references through other sources. The reason for exclusion was multiple. Three studies were duplicated and 13 were not considered eligible (11 did not meet inclusion criteria and 2 were previous non-updated meta-analysis). Besides, 10 full-text articles were excluded for several reasons: 2 described the operative technique and urodynamic data but did not provide information to be included, 1 used different devices simultaneously, 1 was a case report with an unusual complication, 3 were pilot studies published before the definite article included and 3 were post-hoc studies of primary reports already included that focused on specific results (patient reported outcomes, stratification of results according to different generation of devices and a subgroup of patients with longer follow-up) (Fig 1).



Fig 1. Flow chart with studies included in the systematic review.

Characteristics of the studies

Tables 2 and **3** show the characteristics of the 41 selected studies, published between 2005 and 2019, that include 3059 patients in total. Twenty-one articles were published using ProACT [12–32] and 20 using ATOMS [33–52]. No study was randomized controlled, and there was no comparison reported among these or other devices. The main etiology of incontinence was radical prostatectomy but other causes were included. Also the percentage of radiated patients and the degree of baseline incontinence was variable. Dryness rate and improvement rate were very homogeneously reported among the studies. The proportion of patients self-considering satisfied with the device was also often specified; however, 24-h pad-count and, more interestingly, pad-test (baseline, after adjustment and differential) was absent in some studies. The Freeman-Tukey arcsine transform was developed (**S1 Fig**) and pooled proportion (dryness, improvement, satisfaction, explant and complication rates) and mean (number of fillings, differential pad-count and differential pad-test) with 95% confidence intervals were used as summary effect measure.

Tables 4 and 5 summarize the different domains to evaluate study quality. Observational cohorts available have been analyzed according to the Newcastle-Ottawa Scale for quality assessment. None of the studies included in this meta-analysis was a randomized clinical trial and studies generally lack a nonexposed cohort or control. However, the comparision of the intervention is possible with respects to baseline situation before use of device. Attrition bias seems low and the cohorts appear comparable for outcome assessment.

Baseline characteristics

Similarities between the populations evaluated for each device were investigated to confirm the groups are comparable. Firstly, regarding baseline incontinence severity baseline pad-count for ATOMS and ProACT were equivalent (5.05 (95% CI 4.8;5.3) PPD compared to 4.6 (95% CI 3.9;5.3) PPD; p = .27). Secondly, the proportion of patients with previous surgery for incontinence was similar between devices (16.1% (95% CI 9.7;23.5) for ATOMS vs. 11% (95% CI 5.05;18.6) for ProACT; p = .29). Finally, patient age at the time of surgery was higher for ATOMS (70.8 (95% CI 69.9;71.6) years vs. 68.9 (95% CI 68.1;69.6); p = .0009), but this difference is not clinically significant.

Efficacy

Regarding the primary outcome, mean dryness rate was higher with ATOMS than ProACT (68% (95% CI 62;73) compared to 55% (95% CI 47;63); p = .01) (Fig 2). Also regarding the main secondary objective of the study, percentage of overall improvement was also higher for ATOMS than proACT (91% (95% CI 87;94) vs. 80% (95% CI 72;87; p = .007) (Fig 3). A higher proportion of patients self-declared satisfied with ATOMS as well (87% (95% CI 81;92) vs 56% (95% CI 36;76); p = .002) (Fig 4). Conversely, mean number of fillings for adjustment was lower for ATOMS than proACT (2.4 (95% CI 1.9;2.9) vs. 3.5 (95% CI 3;3.9); p = .001) (Fig 5).

Results revealing the magnitude of effect for improving urine loss are quantitatively notorious. Mean baseline pad-count was equivalent in the studies evaluated (5 PPD (95% CI 4.8;5.3) for ATOMS and 4.6 PPD (95% CI 3.9;5.3) for proACT; p = .27) but postoperative pad-count was lower for ATOMS (1.1 PPD (95% CI .9;1.3) vs. 2.1 PPD (95% CI 1.7;2.5); p < .0001). Therefore, mean differential pad-count was also lower for ATOMS (-4 PPD (95% CI -4.4;-3.7) vs. -2.5 PPD (95% CI (3.5;-1.5); p = .005) (**Fig 6**). Studies revealing mean pad-test were less numerous, especially for proACT; however, interesting data were also evidenced. Mean baseline pad-test reported was significantly higher for ATOMS (505 mL (95% CI 461;549) vs. 312 mL (95% CI 231;394); p < .0001) but postoperative pad-test became equivalent (22 mL (95%

Table 2. Stu	idies v	with P	roACT inclu	ided in the	meta-analysis	and variab	les evalua	ted.								
Author Year (Reference)	z	Dry rate (%)	Definition of dryness	Improved rate (%)	Number of adjustments	Satisfied rate (%)	Explant rate (%)	Complication rate (%)	Major complication rate (%)	Baseline pad count (PPD)	Postoperative pad count (PPD)	Baseline pad test (mL/day)	Postoperative pad test (mL/ day)	Previous surgery for incontinence (%)	Mean age (yrs)	Mean follow- up (mo)
Hübner & Schlarp 2005 (12)	117	67	0-1 PPD	88	3(1-15)	NR	27.4	46.2	NR	5.6±3.8	2.5±2.5	NR	NR	NR	70 (50- 89)	13(3- 54)
Trigo- Rocha et al 2006 (13)	25	65.2	0-1 PPD	78.2	4.6(1-7)	NR	17.3	NR	NR	4.7±1.7	1.8±1.6	NR	NR	NR	68.6 (61- 72)	22.4(6- 48)
Kocjancic et al 2007 (14)	64	67	0-1 PPD	82	3(0-9)	NR	14	17.2	NR	5.2	3.6±3.3	NR	NR	NR	65.4 (25- 79)	19.5 (12- 62)
Lebret et al 2008 (15)	62	30	0 PPD	89	NR	NR	30.6	NR	NR	4.6(1-10)	1.1(0-6)	NR	NR	NR	71.1 (52- 87)	9
Gilling et al 2008 (16)	37	62	0 PPD	82	3.3(0-7)	NR	13	48.6	16.2	2.8±2	1.6±1.5	NR	NR	NR	69.9 (59- 79)	51.5 (24- 60)
Crivellaro et al 2008 (17)	46	68	0-1 PPD	85	3(1-7)	NR	14	NR	NR	5.1	2.5	NR	NR	NR	67 (45- 82)	19
Martens et al 2009 (18)	29	31%	0-1 PPD	NR	NR	56	44.8	69	27.6	4.8	3.1	NR	NR	NR	65	41
Gregori et al 2010 (19)	79	66.1	0-1 PPD	91.9	3.6(0-14)	NR	NR	10.1	NR	3.5(1-10)	NR	407.5 (40– 1300)	NR	NR	68 (51- 82)	25
García- Matres et al 2009 (20)	69	69.8	0-1 PPD	84	3	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	22(3- 48)
Giammò et al 2010 (21)	18	NR	NR	61	3(1-6)	NR	NR	NR	NR	NR	NR	NR	NR	NR	NR	24(12- 38)
Rouprêt et al 2011 (22)	128	66.4	0-1 PPD	75	2.3(0-5)	NR	18	25	NR	4.2(1- 20)	1.5	NR	NR	10	71 (52- 87)	56.3 (24- 95)
Kjaer et al 2012 (23)	114	50	0-1 PPD	80	4(0-14)	53	20.2	20.2	NR	4.75(1- 26)	2.25(0–26)	352.5 (16– 2800)	11(0-3000)	NR	NR	58(1- 80)
Crivellaro et al 2012 (24)	42	71	0-1 PPD	92	NR	7	9.5	NR	NR	NR	NR	NR	NR	23.8	65.2 (21- 80)	12(3- 19)
Gatti et al 2012 (25)	28	60.7	0-1 PPD	85.7	NR	NR	28.6	21	NR	NR	NR	NR	NR	NR	NR	NR
Utomo et al 2013 (26)	49	75.5	0-1 PPD	83.7	4	NR	16.3	NR	NR	NR	NR	NR	NR	22.5	NR	NR
Venturino et al 2015 (27)	22	18	0 PPD	82	4.5(0-15)	45	55	NR	NR	5.9(3- 12)	3.9(0-12)	242.3 (12- 1200)	NR	NR	70.2 (53- 80)	57
Baron et al 2017 (28)	14	57	0-1 PPD	88	NR	77	28	NR	NR	NR	NR	95±130	34±83	100	69 (70- 79)	34(4- 89)
															(Cor	tinued)

PLOS ONE

Table 2. (C	Contin	(pənu														
Author Year (Reference)	z	Dry rate (%)	Definition of dryness	Improved rate (%)	Number of adjustments	Satisfied rate (%)	Explant rate (%)	Complication rate (%)	Major complication rate (%)	Baseline pad count (PPD)	Postoperative pad count (PPD)	Baseline pad test (mL/day)	Postoperative pad test (mL/ day)	Previous surgery for incontinence (%)	Mean age (yrs)	Mean follow- up (mo)
Nash et al 2018 (29)	123	41	90–100% pad-weight reduction	61	NR	NR	24.2	25.2	NR	4.1±2.3	2.8±1.8	399±437	216±322	22.5	69.7 ±7.9	18
Nestler et al 2019 (30)	134	NR	NR	82.6	NR	NR	52.7	8.2	2.2	6(4-7)	1(1–2)	NR	NR	NR	71 (67– 75)	26(9- 59)
Noordhoff et al 2018 (31)	143	47.4	0-1 PPD	72.9	4(2-6)	88.3	30.1	21.7	NR	3.5±3	1±1	NR	NR	14.7	69 (66- 73)	46(21- 76)
Finazzi Agrò et al 2019 (32)	240	29.6	<8gr pad weight	37.5	3.78(1-10)	66.3	12.5	22.5	7.9	NR	NR	367±145	113±145	7.5	68.3 ±7.5	NR
PPD: pads I	er da	iy; mL:	: mililiter; yrs:	years; mo: 1	nonths; NR: 1	ot reported	-									

PLOSONE

Table 3. Sti	udies v	with A	TOMS inclu	ıded in the	meta-analysis	and varia	bles evalu:	ated.								
Author Year (Reference)	z	Dry rate (%)	Definition of dryness	Improved rate (%)	Number of adjustments	Satisfied rate (%)	Explant rate (%)	Complication rate (%)	Major complication rate (%)	Baseline pad count (PPD)	Postopertive pad count (PPD)	Baseline pad test (mL/ day)	Postoperative pad test (mL/ day)	Previous surgery for incontinence (%)	Mean age (yrs)	Mean follow- up (mo)
Hoda et al 2012 (33)	124	61.6	0-1 PPD	93.8	4.3±1.8	NR	4	8.9	NR	8.8±3.8	1.8±1.2	725±372	NR	NR	71.2±5.5	19.2 ±2.2
Seweryn et al 2012 (34)	38	60.5	0-1 PPD <15mL	84.2	3.97(0–9)	NR	10.5	NR	15.8	6.78(2- 10)	1.36(0-10)	747 (230- 1600)	115(0–1500)	28.9	70(60- 83)	16.9
Hoda et al 2013 (35)	66	63	0-1 PPD <10mL	92	3.8±1.3	NR	4	NR	NR	7.1±2.3	1.3±1.1	681±466	79.7±210	34.3	70.4±6.2	17.8 ±-1.6
Krause et al 2014 (36)	36	39	0-1 PPD	50	NR	61.8	30.5	44.4	NR	8.33	4.4	NR	NR	30.5	70.4(50- 79)	NR
González- Pérez et al 2014 (37)	13	92.3	0 PPD	87.1	NR	100	0	NR	NR	NR	NR	NR	NR	15.4	63(59- 87)	16(4- 32)
Friedl et al 2016 (38)	34	56	0-1 PPD	88	NR	NR	11.8	17.6	NR	3.5±0.2	1.5 ± 0.3	NR	NR	29.4	70.7(55- 83)	5.7±0.5
Mühlstädt et al 2016 (39)	54	48.1	0 PPD	77.7	4.5±-2.3	NR	7.4	25.9	NR	7.7±4.8	1.6±1.7	NR	NR	20.4	67.5 ± 7.3	27.5 ±18.4
Friedl et al 2016 (40)	62	61.3	0-1 PPD	06	1.5±1.2	NR	14.5	6.45	NR	4(3-5)	1(0-2)	350 (300– 542)	5(0-135)	27.4	71.3(69- 75)	17.7 (1.7- 55.5)
Hüsch et al 2016 (41)	49	NR	NR	NR	NR	NR	2.0	14.2	2.0	NR	NR	NR	NR	NR	NR	NR
Buresova et al 2017 (42)	35	62.9	0-1 PPD	100	4.3(1–15)	NR	2.9	20	NR	5	1	NR	NR	NR	66.7(51- 3	21.2(3- 63)
Friedl et al 2017 (43)	287	64	0-1 PPD <10mL	NR	3(2-4)	NR	19.5	7	2	4(3-5)	1(0-2)	400 (300– 700)	18(0-105)	NR	74) 74)	31(10- 54)
Friedl et al 2017 (44)	49	57.1	0-1 PPD <10mL	89.7	2±1	NR	19.5	34.7	16.3	4(3-6)	1(0-2)	458 (310- 630)	10(0-90)	16.3	73 (68– 76)	32±8.5
Angulo et al 2017 (45)	34	85.3	0-1 PPD	95	1±3	97	0	14.7	NR	5±3	0∓0	510±500	0±15	11.8	70.5 (48-79)	18.5 ±10
Manso et al 2018 (46)	25	64	0-1 PPD	100	1.54±1.3	84	0	NR	4	4.84 ±2.95	1.6 ± 2.02	NR	NR	12	71.4±6.6	21.56 ±8.89
Angulo et al 2018 (47)	215	80.5	0-1 PPD <10mL	85,1	1.4±1.9	85.1	3.25	15.3	3.7	3.9±2	0.9±1.5	484±372	63.5±201	5.6	69.7±6.8	24.3 ±15
Esquinas et al 2018 (48)	60	81.7	0-1 PPD	93.3	1±2	93.2	1.7	18.6	NR	5±3	0±1	465±450	0±20	6.7	72±7	21±22
Angulo et al 2018 (49)	20	75	0-1 PPD	91.5	1±3	80	0	15	NR	4±3	0±1.5	375 ±-855	10±32.25	10	76.5±9	38.5 ±19.5
Angulo et al 2018 (50)	30	76.7	0-1 PPD	100	1±1	83.3	3.3	13.3	NR	4±3	0±1	435±393	10±30	100	73±10	21±22
Giammò et al 2019 (51)	52	73.1	0-1 PPD	98.1	1.55±1	NR	0	19	0	4.23(2- 8)	NR	412 (180– 1100)	100(0-440)	57.7	77.7(58– 84)	20.1 ±20.7
															(Con	tinued)

Ø PLOS ONE

(Continued
<i>.</i> е
uble
E

Author Year (Reference)	z	Dry rate (%)	Definition of dryness	Improved rate (%)	Number of adjustments	Satisfied rate (%)	Explant rate (%)	Complication rate (%)	Major complication rate (%)	Baseline pad count (PPD)	Postopertive pad count (PPD)	Baseline pad test (mL/ day)	Postoperative pad test (mL/ day)	Previous surgery for incontinence (%)	Mean age (yrs)	Mean follow- up (mo)
Doiron et al 2019 (52)	160	80	0-1 PPD	87.8	2.4	86.3	NR	22.3	4.4	4(3-5)	0.5(0-1-9	NR	NR	16.3	70.5±6.6	9(4.5- 13.5)

PPD: pads per day; mL: mililiter; yrs: years; mo: months; NR: not reported

		Selection					Outcome	
Author Year (Reference)	Representativeness of the exposed (interventional cohort) (a)	Selection of the nonexposed cohort (b)	Ascertainment of exposure (intervention) (c)	Incident disease (d)	Comparability (e)	Assessment of outcome (f)	Length of follow-up (g)	Adequacy of follow-up (h)
Hübner & Schlarp 2005 (12)	А	С	А	A	A	В	А	A
Trigo-Rocha et al 2006 (13)	А	С	А	A	A	В	А	В
Kocjancic et al 2007 (14)	А	С	А	A	А	В	A	А
Lebret et al 2008 (15)	А	С	А	A	А	В	В	А
Gilling et al 2008 (16)	А	С	А	A	А	В	А	А
Crivellaro et al 2008 (17)	А	С	А	A	А	В	A	А
Martens et al 2009 (18)	А	С	А	A	А	В	А	А
Gregori et al 2010 (19)	А	С	А	A	В	D	А	А
García-Matres et al 2009 (20)	А	С	С	A	В	D	A	D
Giammò et al 2010 (21)	А	С	D	A	В	D	А	D
Rouprêt et al 2011 (22)	А	С	А	A	A	В	А	А
Kjaer et al 2012 (23)	А	С	А	A	А	В	A	А
Crivellaro et al 2012 (24)	A	С	А	A	В	С	A	A
Gatti et al 2012 (25)	А	С	D	A	В	D	В	D
Utomo et al 2013 (26)	А	С	D	A	В	D	В	D
Venturino et al 2015 (27)	А	С	А	A	А	В	A	А
Baron et al 2017 (28)	С	С	А	A	A	В	А	А
Nash et al 2018 (29)	А	С	А	A	A	В	А	А
Nestler et al 2018 (30)	А	С	А	A	A	В	А	А
Noordhoff et al 2018 (31)	А	С	А	A	A	В	A	A
Finazzi Agrò et al 2019 (32)	А	С	A	A	A	В	A	A

Table 4. Newcastle-Ottawa scale for assessing the quality of cohort studies with ProACT included in the meta-analysis.

(a) A, truly representative of the average patient at risk for male stress incontinence; B, somewhat representative of the average patient at risk; C, selected group; D, no description.

(b) A, drawn from the same source as the intervention cohort (concurrent controls); B, drawn from a different source (historical controls); C, no description of the derivation of the nonexposed control

(c) A, secure record; B, structures review; C, written self-report; D, no description.

(d) Demonstration that outcome of interest was not present at the start of the study: A, yes; B, no.

(e) Comparability of cohorts on the basis of the design or analysis: A, study controls for the most important factor (conditioning regimen), B, study controls for any additional factor; C, not carried out or not reported.

(f) A, independent blind assessment; B, record linkage; C, self-report; D, no description.

(g) Was follow-up long enough for outcomes to occur? A, yes; B, no.

(h) A, complete follow-up (all subjects were accounted for); B,Subject lost to follow-up were unlikely to introduce bias because small numbers were lost (>90% had follow-up, or description was provided of those lost); C, follow-up rate <90%, and there was no description of those lost; D, no statement.

		Selection			Comparability		Outcome	
Author Year (Reference)	Representativeness of the exposed (interventional cohort) (a)	Selection of the nonexposed cohort (b)	Ascertainment of exposure (intervention) (c)	Incident disease (d)	(e)	Assessment of outcome (f)	Length of follow-up (g)	Adequacy of follow-up (h)
Hoda et al 2012 (33)	А	С	А	A	A	В	A	А
Seweryn et al 2012 (34)	А	С	А	A	A	В	A	А
Hoda et al 2013 (35)	А	С	А	A	A	В	A	A
Krause et al 2014 (36)	В	С	А	A	A	С	A	D
González-Pérez et al 2014 (37)	В	С	А	A	В	С	A	A
Friedl et al 2016 (38)	A	С	A	A	A	В	A	A
Mühlstädt et al 2016 (39)	А	С	А	A	A	В	A	А
Friedl et al 2016 (40)	А	С	А	A	A	В	A	А
Hüsch et al 2016 (41)	А	А	А	A	С	В	В	D
Buresova et al 2017 (42)	А	С	А	A	A	В	A	А
Friedl et al 2017 (43)	А	С	А	A	A	В	A	A
Friedl et al 2017 (44)	С	С	В	A	В	В	A	A
Angulo et al 2017 (45)	А	С	А	A	A	В	A	A
Manso et al 2018 (46)	А	С	А	A	A	В	A	A
Angulo et al 2018 (47)	А	С	А	A	A	В	A	A
Esquinas et al 2018 (48)	А	С	А	A	A	В	A	А
Angulo et al 2018 (49)	С	С	А	A	A	В	A	А
Angulo et al 2018 (50)	С	С	А	A	A	В	A	A
Giammò et al 2019 (51)	В	С	A	A	A	В	A	A
Doiron et al 2019 (52)	A	С	A	A	A	В	A	A

Table 5. Newcastle-Ottawa scale for assessing the quality of cohort studies with ATOMS included in the meta-analysis.

(a) A, truly representative of the average patient at risk for male stress incontinence; B, somewhat representative of the average patient at risk; C, selected group; D, no description.

(b) A, drawn from the same source as the intervention cohort (concurrent controls); B, drawn from a different source (historical controls); C, no description of the derivation of the nonexposed control

(c) A, secure record; B, structures review; C, written self-report; D, no description.

(d) Demonstration that outcome of interest was not present at the start of the study: A, yes; B, no.

(e) Comparability of cohorts on the basis of the design or analysis: A, study controls for the most important factor (conditioning regimen), B, study controls for any additional factor; C, not carried out or not reported.

(f) A, independent blind assessment; B, record linkage; C, self-report; D, no description.

(g) Was follow-up long enough for outcomes to occur? A, yes; B, no.

(h) A, complete follow-up (all subjects were accounted for); B,Subject lost to follow-up were unlikely to introduce bias because small numbers were lost (>90% had follow-up, or description was provided of those lost); C, follow-up rate <90%, and there was no description of those lost; D, no statement.

Study	Events Total		Proportion	95%-CI	Weight
rupo = ATOMS Hoda et al 2012 Seweryn et al 2012 Hoda et al 2013 (rause et al 2014 Pérez et al 2014 Friedl et al 2016 Wibhistädt et al 2016 Buresova et al 2017 Friedl et al 2017 Friedl et al 2017 Angulo et al 2017 Angulo et al 2018 Angulo et al 2018 Cangulo et al 2018 Angulo et al 2018 Cangulo et al 2018 Cangulo et al 2018 Cangulo et al 2018 Angulo et al 2018 Cangulo et al 2019 Coiron et al 2019 Canguno et al 2019	76 124 23 38 62 99 14 36 12 13 19 34 26 54 38 62 22 35 184 287 28 49 29 34 16 25 173 215 49 60 15 20 23 30 38 52 128 160 1427 40.0121, p<0.01		0.61 0.63 0.39 0.92 0.56 0.48 0.61 0.63 0.64 0.57 0.85 0.64 0.80 0.82 0.77 0.73 0.80 0.68	$\begin{matrix} [0.52; \ 0.70] \\ [0.43; \ 0.76] \\ [0.52; \ 0.72] \\ [0.23; \ 0.57] \\ [0.44; \ 1.00] \\ [0.34; \ 0.73] \\ [0.44; \ 0.73] \\ [0.45; \ 0.70] \\ [0.58; \ 0.70] \\ [0.42; \ 0.71] \\ [0.68; \ 0.95] \\ [0.75; \ 0.86] \\ [0.70; \ 0.90] \\ [0.55; \ 0.90] \\ [0.55; \ 0.90] \\ [0.55; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.53; \ 0.90] \\ [0.54; \ 0.73] \\ \end{matrix}$	2.9% 2.5% 2.5% 1.8% 2.7% 2.7% 2.7% 2.7% 2.7% 2.3% 3.0% 2.4% 2.4% 2.4% 2.4% 49.6%
Trigo- PROACT Hübner & Schlarp 2005 Trigo-Rocha et al 2006 (ocjancic et al 2007 Lebret et al 2008 Gilling et al 2008 Martens et al 2009 Grupori et al 2010 García-Martres et al 2009 Rouprêt et al 2011 Giarcia-Martres et al 2009 Rouprêt et al 2011 Giarcia et al 2012 Crivellaro et al 2012 Johno et al 2012 Johno et al 2013 Venturino et al 2015 Saron et al 2017 Nash et al 2018 Noordhoff et al 2018 Agrò et al 2018 Random effects model Heterogeneity: / ² = 88%, r ² =	78 117 16 25 43 64 19 62 23 37 31 46 9 29 52 79 48 69 85 128 57 114 30 42 17 28 37 49 4 22 - 8 143 50 123 68 143 71 240 1431 \$0.0257, p<0.01		0.67 0.64 0.67 0.31 0.62 0.67 0.31 0.66 0.50 0.70 0.70 0.66 0.50 0.71 0.61 0.77 0.41 0.41 0.41 0.41 0.41	$\begin{matrix} [0.57; \ 0.75] \\ [0.43; \ 0.82] \\ [0.54; \ 0.78] \\ [0.20; \ 0.44] \\ [0.52; \ 0.80] \\ [0.52; \ 0.80] \\ [0.52; \ 0.80] \\ [0.53; \ 0.75] \\ [0.44; \ 0.76] \\ [0.55; \ 0.84] \\ [0.41; \ 0.78] \\ [0.61; \ 0.87] \\ [0.61; \ 0.87] \\ [0.29; \ 0.82] \\ [0.32; \ 0.50] \\ [0.32; \ 0.50] \\ [0.34; \ 0.36] \\ [0.24; \ 0.36] \\ [0.24; \ 0.36] \\ \hline \end{matrix}$	2.9% 2.3% 2.7% 2.5% 2.6% 2.4% 2.8% 2.8% 2.9% 2.9% 2.4% 2.4% 2.4% 2.7% 2.9% 3.0% 3.0% 50.4%
Random effects model Heterogeneity: / ² = 88%, T ² =	2858 = 0.0255, p< 0.01	0.2 0.4 0.6 0.8	0.61	[0.56; 0.67]	100.0%



CI 13;31) for ATOMS and 91 mL (95% CI 15;186) for ProACT; p < .08). Not strikingly, mean differential pad-test was lower for ATOMS (-426 mL (95% CI -462;-389) vs. -211 mL (95% CI -312;-111); p < .0001) (Fig 7).

Risk of publication bias was not identified for primary efficacy outcome, and variables evaluating the effect magnitude (differential pad-test and differential pad-count) (S2 Table). However, significant heterogeneity was detected in all variables evaluated (S3 Table), what possibly depends on different baseline severity of incontinence and patient profile, use of differrent generation devices for ATOMS and also different technique to appropriately place ProACT as described in the literature. Only randomization could reduce these biases.

Safety and device durability

Mean follow-up was 25.7 months (95% CI 22.7;28.7), lower for ATOMS than ProACT (20.8 months (95% CI 17.1;24.5) vs. 30.6 months (95% CI 23;38.1); p = .02). The studies presented interesting safety results, both on explatation and complications. Mean explant rate was significantly lower for ATOMS (5% (95% CI 2;9) vs. 25% (95% CI 19;31); p < .0001) (Fig 8). Complication rate was also lower for ATOMS, but without reaching statistical significance (17% (95% CI 13;22) vs. 26% (95% CI 18;34); p = 0.067) (Fig 9). Similarly, major complication rate

Study	Events	Total		Proportion	95%-CI	Weight
grupo = ATOMS Hoda et al 2012 Seweryn et al 2012 Hoda et al 2013 Krause et al 2014 Pérez et al 2014 Friedl et al 2016 Friedl et al 2016 Buresova et al 2017 Friedl et al 2017 Angulo et al 2017 Manso et al 2017 Angulo et al 2018 Angulo et al 2018 Angulo et al 2018 Angulo et al 2018 Giammò et al 2018 Giammò et al 2019 Doiron et al 2019 Random effects model Heterogenetly: / ² = 77%, r ² =	116 32 91 18 11 30 42 56 35 44 32 25 183 56 18 30 51 140 0.0136,	124 38 99 36 13 34 54 62 35 49 34 20 20 20 30 52 160 1140 p<0.01		0.94 0.84 0.92 0.50 0.85 0.88 0.78 0.90 1.00 0.90 1.00 0.94 1.00 0.85 0.93 0.90 1.00 0.93 0.90 1.00 0.98 0.93 0.90	$\begin{matrix} [0.88; 0.97] \\ [0.69; 0.94] \\ [0.85; 0.96] \\ [0.33; 0.67] \\ [0.55; 0.98] \\ [0.73; 0.97] \\ [0.64; 0.86] \\ [0.90; 1.00] \\ [0.78; 0.97] \\ [0.80; 0.99] \\ [0.86; 1.00] \\ [0.86; 0.99] \\ [0.84; 1.00] \\ [0.84; 0.98] \\ [0.68; 0.99] \\ [0.84; 1.00] \\ [0.90; 1.00] \\ [0.91; 0.94] \\ \end{matrix}$	2.9% 2.6% 2.5% 2.0% 2.7% 2.7% 2.5% 2.7% 2.5% 2.4% 3.0% 2.5% 2.3% 2.5% 2.3% 2.5% 2.9% 46.9%
grupo = PROACT Hübner & Schlarp 2005 Trigo-Rocha et al 2006 Kocjancic et al 2007 Lebret et al 2008 Gilling et al 2008 Gregori et al 2010 García-Martres et al 2010 García-Martres et al 2010 García-Martres et al 2010 García-Martres et al 2010 García et al 2011 Kjaer et al 2012 Crivellaro et al 2012 Gatti et al 2012 Utomo et al 2013 Venturino et al 2015 Baron et al 2017 Nash et al 2018 Noordhoff et al 2018 Agrò et al 2019 Random effects model Heterogeneity: I ² = 91%, T ² =	103 20 52 55 30 39 73 58 11 96 91 39 24 41 18 12 75 111 104 90 0.0353,	117 254 62 37 46 79 69 128 114 42 28 49 22 14 123 134 123 134 123 134 123 134 123 134 1254 p<0.01		0.88 0.80 0.81 0.89 0.81 0.85 0.92 0.84 0.61 0.75 0.80 0.93 0.86 0.84 0.84 0.82 0.86 0.61 0.83 0.73 0.38 0.88	$\begin{matrix} [0.81; 0.93]\\ [0.59; 0.93]\\ [0.70; 0.90]\\ [0.78; 0.95]\\ [0.65; 0.92]\\ [0.71; 0.94]\\ [0.84; 0.97]\\ [0.73; 0.92]\\ [0.36; 0.83]\\ [0.67; 0.82]\\ [0.71; 0.87]\\ [0.71; 0.86]\\ [0.70; 0.93]\\ [0.67; 0.96]\\ [0.57; 0.98]\\ [0.52; 0.70]\\ [0.55; 0.80]\\ [0.55; 0.80]\\ [0.65; 0.80]\\ [0.31; 0.44]\\ [0.72; 0.87] \end{matrix}$	2.9% 2.7% 2.7% 2.6% 2.8% 2.8% 2.9% 2.9% 2.9% 2.9% 2.4% 2.7% 2.9% 2.9% 2.9% 3.0% 53.1%
Random effects model Heterogeneity: I^2 = 91%, τ^2 =	0.0347,	2694 p< 0.01	0.4 0.5 0.6 0.7 0.8 0.9	0.85	[0.80; 0.90]	100.0%

Fig 3. Forest plot of studies analyzed for improve rate.

https://doi.org/10.1371/journal.pone.0225762.g003

Study	Events Total		Proportion	95%-CI Weight
grupo = ATOMS Krause et al 2014 Pérez et al 2014 Angulo et al 2017 Manso et al 2018 Angulo et al 2018 Angulo et al 2018 Angulo et al 2018 Doiron et al 2019 Random effects mode Heterogeneity: $l^2 = 67\%$, *	22 36 13 13 33 34 21 25 183 215 56 60 16 20 25 30 138 160 1593 2= 0.0089, p < 0.0		0.61 1.00 0.97 0.84 0.85 0.93 0.80 0.83 0.80 0.83	
grupo = PROACT Martens et al 2009 Kjaer et al 2012 Crivellaro et al 2012 Venturino et al 2015 Baron et al 2017 Noordhoff et al 2018 Agrò et al 2019 Random effects mode Heterogeneity: / ² = 95%,*	$\begin{array}{cccccccccccccccccccccccccccccccccccc$	*	0.55 0.53 0.07 0.45 0.79 0.88 0.66 0.56	
Random effects mode Heterogeneity: <i>I</i> ² = 93%,	1197 ² = 0.0509, p< 0.0		0.75	[0.64; 0.85] 100.0%

Fig 4. Forest plot of studies analyzed for satisfaction rate.

Study	Mean	MRAW	95%-CI	Weight
Study grupo = ATOMS Hoda et al 2012 Seweryn et al 2012 Hoda et al 2013 Mühlstädt et al 2016 Friedl et al 2016 Buresova et al 2017 Friedl et al 2017 Angulo et al 2017 Manso et al 2017 Manso et al 2018 Angulo et al 2018 Angulo et al 2018 Angulo et al 2018 Angulo et al 2018	Mean	MRAW 4.30 3.97 3.80 - 4.50 1.50 - 4.30 3.00 2.00 1.00 1.54 1.40 1.00 1.00 1.00	95%-CI [3.98; 4.62] [3.49; 4.45] [3.54; 4.06] [3.89; 5.11] [1.20; 1.80] [3.54; 5.06] [2.97; 3.03] [1.72; 2.28] [-0.01; 2.01] [1.03; 2.05] [1.15; 1.65] [0.49; 1.51] [-0.31; 2.31] [-0.31; 2.31]	Weight 3.5% 3.3% 3.6% 3.5% 2.9% 3.6% 3.6% 3.3% 2.5% 3.3% 3.6% 3.3% 2.1%
Giammò et al 2019		1.55	[0.64; 1.36] [1.28; 1.82]	3.5% 3.5%
Doiron et al 2019 Random effects model Heterogeneity: l^2 = 98%, τ^2 = 0.8576,	p< 0.01 ↔	2.40 2.42	[2.20; 2.60] [1.95; 2.89]	3.6% 52.5%
grupo = PROACT Hübner & Schlarp 2005 Trigo-Rocha et al 2006 Kocjancic et al 2007 Gilling et al 2008 Crivellaro et al 2008 Gregori et al 2010 García-Martres et al 2009 Giammò et al 2010 Rouprêt et al 2011 Kjaer et al 2012 Utomo et al 2013 Venturino et al 2015 Noordhoff et al 2018 Agrò et al 2019 Random effects model Heterogeneity: I^2 = 97%, τ^2 = 0.6199,	p< 0.01	3.00 4.60 3.00 3.30 3.00 3.00 2.30 4.00 4.00 4.50 4.00 3.78 3.48	[2.58; 3.42] [4.21; 4.99] [2.63; 3.37] [2.91; 3.69] [2.71; 3.29] [3.09; 4.11] [2.62; 3.38] [2.63; 3.37] [2.16; 2.44] [3.58; 4.42] [3.55; 4.45] [3.46; 5.54] [3.89; 4.11] [3.59; 3.97] [3.05; 3.91]	3.4% 3.5% 3.4% 3.5% 3.3% 3.4% 3.6% 3.4% 3.6% 3.6% 3.6% 47.5%
Random effects model Heterogeneity: $I^2 = 98\%$, $T^2 = 0.60\overline{65}$,	p ₹ 0.01 →	2.93	[2.63; 3.22]	100.0%
-4	-2 0 2 4			



was similar between devices 4.2% (95% CI 1.7;7.7) for ATOMS vs. 10.4% (95% CI 3.15;20.7) for ProACT; p = 0.15). Risk of publication bias was not identified for these safety outcomes, but significant heterogeneity was probably related to the variability in reporting complications among studies (S2 Table, S3 Table),

Only 4 studies reveal strong data on durability (2 for each device) and specific analysis of Kaplan-Meier curves given and number of patients at risk allows to compare the proportio of devices that continue working in place, was higher for ATOMS on 1-year (92% (95% CI 87;96) vs. 76% (95% CI 69;83); p < .0001), 2-years (85% (95% CI 74;93) vs. 61% (95% CI 52;70); p = .0008) and 3-years (81% (95% CI 73;87) vs. 58% (95% CI 48;67); p < .0001) (Fig 10).

Discussion

Despite male SUI therapy is an area of increasing interest the overall risk of bias must be reported as the majority of studies available about all male incontinence devices including AUS, slings and adjustable devices are retrospective or quasi prospective and mostly single center uncontrolled cohorts. In the absence of randomized controlled trials systematic reviews and meta-analysis can provide better evidence than simple prospective cohorts. Even though,

Study	Ex Total	perimental Mean SD	Total	Control Mean SD	Mean Difference	MD	95%-CI	Weight
,								
grupo = ATOMS								
Hoda et al 2012	124	1.80 1.20	124	8.80 3.80	*	-7.00 [-7.70; -6.30]	3.4%
Seweryn et al 2012	38	1.36 1.70	38	6.78 1.30	-	-5.42 [-6.10; -4.74]	3.4%
Hoda et al 2013	99	1.30 1.10	99	7.10 2.30		-5.80 [-6.30; -5.30]	3.6%
Krause et al 2014	36	4.40 1.40	36	8.33 2.10	<u> </u>	-3.93 [-4.75; -3.11]	3.3%
Friedl et al 2016	34	1.50 0.30	34	3.50 0.20		-2.00 [-2.12; -1.88]	3.8%
Mühlstädt et al 2016	54	1.60 1.70	54	7.70 4.80		-6.10 [-7.46; -4.74]	2.7%
Friedl et al 2016	62	1.00 0.30	62	4.00 0.30		-3.00 [-3.11; -2.89]	3.8%
Buresova et al 2017	35	1.00 1.40	35	5.00 2.10	<u>+</u>	-4.00 [-4.84; -3.16]	3.3%
Friedl et al 2017	287	1.00 0.30	287	4.00 0.30		-3.00 [-3.05; -2.95]	3.8%
Friedl et al 2017	49	1.00 0.30	49	4.00 0.50		-3.00 [-3.16; -2.84]	3.8%
Angulo et al 2017	34	0.00 1.00	34	5.00 3.00		-5.00 [-6.06; -3.94]	3.0%
Manso et al 2018	25	1.60 2.02	25	4.84 2.95		-3.24 [-4.64; -1.84]	2.6%
Angulo et al 2018	215	0.90 1.50	215	3.90 2.00		-3.00 [-3.33; -2.67]	3.7%
Esquinas et al 2018	60	0.00 1.00	60	5.00 3.00		-5.00 [-5.80; -4.20]	3.3%
Angulo et al 2018	20	0.00 1.50	20	4.00 3.00		-4.00 [-5.47; -2.53]	2.5%
Angulo et al 2018	30	0.00 1.00	30	4.00 3.00		-4.00 [-5.13; -2.87]	2.9%
Doiron et al 2019	160	0.50 1.48	160	4.00 0.33	-	-3.50 [-3.73; -3.27]	3.7%
Random effects model	1362		1362		♦	-4.04 [-4.40; -3.68]	56.5%
Heterogeneity: 12= 98%, T	= 0.44	34, p< 0.01						
- PROACT								
grupo = PROACT	447	0.50.0.50	447	5 00 0 00	<u> </u>	0.40	0.00 0.001	0.00/
Hubner & Schlarp 2005	117	2.50 2.50	117	5.60 3.80	吉	-3.10 [-3.92; -2.28]	3.3%
Trigo-Rocha et al 2006	25	1.80 1.60	25	4.70 1.70	<u> </u>	-2.90 [-3.82; -1.98]	3.2%
Kocjancic et al 2007	64	3.60 3.30	64	5.20 2.10	<u> </u>	-1.60 [-2.56; -0.64]	3.1%
Lebret et al 2008	62	1.10 1.00	62	4.60 1.50		-3.50	-3.95; -3.05]	3.6%
Gilling et al 2008	37	1.60 1.50	37	2.80 2.00	_=	-1.20 [-2.01; -0.39]	3.3%
Crivellaro et al 2008	46	2.50 1.40	46	5.10 2.10	<u> </u>	-2.60	-3.33; -1.87]	3.4%
Martens et al 2009	29	3.10 1.40	29	4.80 2.10		-1.70	-2.62; -0.78]	3.2%
Roupret et al 2011	128	1.50 1.40	128	4.20 3.20	11 III III III III III III III III III	-2.70	-3.31; -2.09]	3.5%
Kjaer et al 2012	114	2.25 4.30	114	4.75 4.20		-2.50 [-3.60; -1.40]	3.0%
Venturino et al 2015	22	3.90 2.00	22	5.90 1.50		-2.00 [-3.04; -0.96]	3.0%
Nash et al 2018	123	2.80 1.80	123	4.10 2.30	_ =	-1.30 [-1.82; -0.78]	3.6%
Nestler et al 2018	134	1.00 0.15	134	6.00 0.50	· · · · · · · · · · · · · · · · · · ·	-5.00	-5.09; -4.91]	3.8%
Noordhoff et al 2018	143	1.00 1.00	143	3.50 1.50		-2.50 [-2.80; -2.20]	3.7%
Random effects model	1044		1044		~	-2.52 [-3.51; -1.53]	43.5%
Heterogeneity: I'= 98%, T	= 3.16	64, p< 0.01						
Random effects model	2406		2406		4	-3.43 [-3.84; -3.031	100.0%
Heterogeneity: /2= 99%, T	= 1.13	73, p = 0						
					-6 -4 -2 0 2 4 6	6		



systematic reviews on the topic are scarce and confusing as tend to compile multiple devices with very different mode of action [4]. Also reviews are often biased to compare AUS and other devices, sometimes all named under the misleading term "slings". Adjustable slings developed to achieve urethral tensioning and results can be satisfactory even in severe incontinence and/or previous irradiation. However, current evidence is not firm to consider they are more efficacious than fixed male slings [6,7].

Both ProACT and ATOMS are really adjustable devices and unlike the AUS exert a noncircumferential periurethral compression that can be increased postoperatively in an office-

	Total	Experi Mean	mental SD	Total	C Mean	ontrol SD	Mean Di	fference	MD	95%-CI
grupo = ATOMS Seweryn et al 2012 Hoda et al 2013 Friedl et al 2013 Friedl et al 2017 Friedl et al 2017 Angulo et al 2017 Angulo et al 2018 Angulo et al 2018 Giarmò et al 2019 Random effects model Heterogeneity: /² = 95%, r	38 99 62 287 49 34 215 60 20 30 52 946 22= 216	115.0 79.7 5.0 18.0 10.0 63.5 0.0 10.0 10.0 10.0 100.0	250.0 210.0 22.3 17.5 15.0 201.0 20.0 32.2 30.0 74.0	38 99 62 287 49 34 215 60 20 30 52 946	747.0 681.0 350.0 400.0 510.0 484.0 465.0 375.0 435.0 412.0	228.0 466.0 23.5 53.0 500.0 372.0 450.0 855.0 393.0 153.0	*****		-632.00 -601.30 -345.00 -382.00 -448.00 -510.00 -420.50 -465.00 -365.00 -425.00 -312.00 -425.72	[-739.58; -524.42] [-701.98; -500.62] [-353.06; -336.94] [-389.96; -374.04] [-463.42; -432.58] [-477.02; -363.98] [-778.96; -351.02] [-778.96; -351.02] [-779.93, 93; 9.98] [-566.04; -283.96] [-358.19; -265.81] [-461.99; -389.46]
grupo = PROACT Kjær et al 2012 Baron et al 2017 Nash et al 2018 Agrò et al 2019 Random effects model Heterogeneity: I ² = 89%, T Random effects model Heterogeneity: I ² = 96%, T	114 14 123 240 491 ² = 903 1437 ² = 346	11.0 34.0 216.0 113.0 4.2500, 7	50.0 83.0 322.0 145.0 0< 0.01	114 14 123 240 491 1437	352.5 95.0 399.0 367.0	464.0 130.0 437.0 145.0	***		-341.50 -61.00 -183.00 -254.00 -211.44 -369.37	[-427.17; -255.83] [-141.79; 19.79] [-278.93; -87.07] [-279.94; -228.06] [-311.90; -110.98] [-406.00; -332.73]



Study	Events	Total	Proportion	95%-CI	Weight
arupo = ATOMS					
Hoda et al 2012	5	124 +	0.04	[0.01: 0.09]	3.0%
Seweryn et al 2012	4	38	0.11	[0.03; 0.25]	2.6%
Hoda et al 2013	4	99 +	0.04	[0.01; 0.10]	2.9%
Krause et al 2014	11	36	0.31	[0.16; 0.48]	2.6%
Pérez et al 2014	0	13	0.00	[0.00; 0.25]	2.0%
Friedl et al 2016	4	34	0.12	[0.03; 0.27]	2.6%
Mühlstädt et al 2016	4	54	0.07	[0.02; 0.18]	2.8%
Friedl et al 2016	9	62	0.15	[0.07; 0.26]	2.8%
Husch et al 2016	1	49	0.02	[0.00; 0.11]	2.7%
Eriod et al 2017	56	297	0.03	[0.00; 0.15]	2.0%
Friedl et al 2017	10	19	0.20	[0.15, 0.25]	2 7%
Angulo et al 2017	õ	34	0.00	[0.10, 0.04]	2.6%
Manso et al 2018	ŏ	25	0.00	[0.00: 0.14]	2.4%
Angulo et al 2018	7	215 +	0.03	[0.01; 0.07]	3.1%
Esquinas et al 2018	1	60	0.02	[0.00; 0.09]	2.8%
Angulo et al 2018	0	20	0.00	[0.00; 0.17]	2.3%
Angulo et al 2018	1	30	0.03	[0.00; 0.17]	2.5%
Giammò et al 2019	0	52 —	0.00	[0.00; 0.07]	2.8%
Random effects model		1316 🗢	0.05	[0.02; 0.09]	50.9%
Heterogeneity: I ² = 83%, T	= 0.0181	1, p< 0.01			
grupo = PROACT					
Hübner & Schlarp 2005	32	117	0.27	[0.20; 0.36]	3.0%
Trigo-Rocha et al 2006	4	25	0.16	[0.05; 0.36]	2.4%
Kocjancic et al 2007	10	64	0.14	[0.07; 0.25]	2.8%
Cilling of al 2008	19	37	0.31	[0.20; 0.44]	2.0%
Crivellaro et al 2008	6		0.14	[0.05; 0.29]	2.0%
Martens et al 2000	13	29	0.15	[0.00, 0.20]	2.7%
Rouprêt et al 2011	23	128	0.18	[0.12: 0.26]	3.0%
Kiaer et al 2012	23	114	0.20	[0.13: 0.29]	3.0%
Crivellaro et al 2012	4	42	0.10	[0.03; 0.23]	2.7%
Gatti et al 2012	8	28 +	0.29	[0.13; 0.49]	2.5%
Utomo et al 2013	8	49 — • — –	0.16	[0.07; 0.30]	2.7%
Venturino et al 2015	12	22		[0.32; 0.76]	2.3%
Baron et al 2017	4	14	0.29	[0.08; 0.58]	2.0%
Nash et al 2018	30	123	0.24	[0.17; 0.33]	3.0%
Nestler et al 2018	/1	134	0.53	[0.44; 0.62]	3.0%
Agrò et al 2010	43	240	0.30	[0.23; 0.38]	3.0%
Random effects model	30	1417	0.12	[0.09, 0.17]	49 1%
Heterogeneity: $I^2 = 85\%$, τ^2	= 0.0182	2, p< 0.01	J.Z4	[0.10, 0.50]	-10.170
Pandom offecto model		2722	0.43	10 00. 0 401	100.0%
Heterogeneity: $l^2 = 0.0\%$	= 0.030		7 0.13	[0.09; 0.18]	100.0%
10.010geneity. 7 = 30 %, 1	- 0.000	0 0.1 0.2 0.3 0.4 0.5 0.6 0).7		



based setting without re-intervention. The patient carries out urination without manipulation and the risk of mechanical failure is minimal, due to simplicity. Recent systematic reviews and meta-analysis confirm both ATOMS and ProACT are efficacious and safe alternatives to treat male stress incontinence of different degree but provide no comparison between them [53,54]. Interestingly, individual data reported for each device in those meta-analysis is very consistent with data here presented.

ATOMS has been implanted in Europe for a decade. It is used in mild and moderate-tosevere SUI, also in patients with previous radiotherapy. However, results are better in nonradiated patients and non-severe SUI [43,47,53]. Factors determinant of best patient perception with ATOMS are postoperative dryness, severe baseline incontinence severity, non-radiation, less pain at discharge and absence of postoperative complications [55]. ProACT has been implanted for two decades in Europe and more recently in the United States, often in outpatient basis under X-ray or transrectal ultrasound guidance. Based on this minimally invasiveness it has been recommended as first-line treatment in non-irradiated patients with mild-tomoderate incontinence [6,54,56], although revision and explantation rates are high [18,27,29,30] because a correct positioning of the balloons is mandatory to achieve good results [21]. Based on our findings, ProACT does not appear to be an ideal device for durable

Study	Events Total		Proportion	95%-CI	Weight
grupo = ATOMS Hoda et al 2012 Krause et al 2014 Friedl et al 2016 Mühlstädt et al 2016 Hüsch et al 2016 Buresova et al 2017 Friedl et al 2017 Friedl et al 2017 Angulo et al 2017 Angulo et al 2018 Angulo et al 2018 Angulo et al 2018 Giammò et al 2019 Cammò et al 2019 Random effects model Heterogeneity: /²= 78%, T	11 124 16 36 6 34 14 54 4 62 7 49 7 35 20 287 17 49 5 34 33 215 11 60 3 20 4 30 10 52 36 160 1301 \sim \sim \sim		0.09 0.44 0.18 0.26 0.06 0.14 0.20 0.07 0.35 0.15 0.15 0.15 0.15 0.15 0.15 0.15 0.1	$ \begin{bmatrix} 0.05; \ 0.15 \\ 0.28; \ 0.62 \\ 0.07; \ 0.35 \\ 0.15; \ 0.40 \\ 0.02; \ 0.16 \\ 0.06; \ 0.27 \\ 0.08; \ 0.37 \\ 0.04; \ 0.11 \\ 0.22; \ 0.50 \\ 0.05; \ 0.31 \\ 0.05; \ 0.31 \\ 0.10; \ 0.33 \\ 0.03; \ 0.38 \\ 0.04; \ 0.31 \\ 0.10; \ 0.33 \\ 0.16; \ 0.30 \\ 0.13; \ 0.22 \\ \end{bmatrix} $	$\begin{array}{c} 4.0\%\\ 3.2\%\\ 3.1\%\\ 3.5\%\\ 3.6\%\\ 3.4\%\\ 4.2\%\\ 3.1\%\\ 4.2\%\\ 3.1\%\\ 4.1\%\\ 3.6\%\\ 2.6\%\\ 3.0\%\\ 3.0\%\\ 3.5\%\\ 4.1\%\\ 55.5\%\\ \end{array}$
grupo = PROACT Hübner & Schlarp 2005 Kocjancic et al 2007 Gilling et al 2008 Martens et al 2010 Rouprêt et al 2010 Rouprêt et al 2011 Kjare et al 2012 Gatti et al 2012 Nash et al 2018 Nestler et al 2018 Noordhoff et al 2018 Agrò et al 2019 Random effects model Heterogeneity: / ² = 89%, T	54 117 11 64 18 37 20 29 8 79 32 128 23 114 6 28 31 123 11 134 54 240 1236 2= 0.0201, p< 0.01		0.46 0.17 0.69 0.10 0.25 0.20 0.21 0.25 0.08 0.22 0.22 0.22	[0.37; 0.56] [0.09; 0.29] [0.32; 0.66] [0.49; 0.85] [0.04; 0.19] [0.18; 0.33] [0.18; 0.33] [0.18; 0.34] [0.18; 0.34] [0.14; 0.28] [0.17; 0.28] [0.17; 0.26]	3.9% 3.6% 3.2% 3.7% 4.0% 3.9% 2.9% 4.0% 4.0% 4.0% 4.0% 4.2%
Random effects model Heterogeneity: $I^2 = 86\%$, T	2537 ² = 0.0175, p< 0.01 0.2	0.4 0.6 0.8	0.21	[0.17; 0.26]	100.0%

Fig 9. Forest plot of studies analyzed for complication rate.

continence or patient satisfaction. Also it is not so minimally invasive procedure as the complication rate, both total and major complications, surmounts that of ATOMS implant. A very recent experience uses ATOMS after retrieval of ProACT and indirectly supports ATOMS is superior to ProACT in the short-term [51]. What is more ATOMS can be used after radiation while ProACT is not recommended in this situation [6,31].

Pooled data in this meta-analysis reveals ATOMS is superior to ProACT in all the items evaluated for efficacy: dryness (68% vs 55%), improvement (91% vs 80%) and patient satisfaction (87% vs 56%). Definitions used as outcome measures are equivalent in most of the series analyzed (**Tables 2** and **3**). The magnitude of effect is also higher for ATOMS than ProACT, regarding both pad-count change (-4 vs -2.5 PPD) and pad-test change (-426 vs -211 mL). Regarding safety, ATOMS is again superior to ProACT with less explant rate (5% vs 24%) and a higher proportion of devices working in place during the first three years of follow-up.

The main limitations of this meta-analysis stand in the short follow-up, especially in the ATOMS arm, and in the very high heterogeneity observed between studies; probably reflecting a variable severity of sphincteric damage included and the absence of controlled randomized studies. Also the critria to report complications appear variable between the studies analyzed. The limitations highlighted are in consonance with the publication bias identified according to Egger's linear regression. Despite the existing limitations, we consider this meta-analysis can be of great help both for physicians and health care providers.

	Study	Events To	otal		Proportion	95%-CI	Weight
	grupo = ATOMS						
	Friedl et al 2017	258 2	287		0.90	[0.86; 0.93]	28.3%
	Angulo et al 2018	164	173	-	- 0.95	[0.90; 0.98]	27.3%
	Random effects model Heterogeneity: I ² = 71%, т	² = 0.0029, p	160 = 0.06		• 0.92	[0.87; 0.96]	55.6%
	grupo = PROACT						
	Venturino et al 2015	15	22 ——		0.68	[0.45; 0.86]	17.5%
	Noordhoff et al 2018	110	143		0.77	[0.69; 0.84]	26.9%
	Random effects model Heterogeneity: $I^2 = 0\%$, τ^2	= 0, <i>p</i> = 0.36	165		0.76	[0.69; 0.83]	44.4%
	Random effects model	2 0.0150	3 2 5		0.85	[0.75; 0.94]	100.0%
•	Heterogeneity: $T = 90\%$, T	= 0.0158, p	0.5 0.6	0.7 0.8 0.9			
	Study	Events To	otal		Proportion	95%-CI	Weight
	grupo = ATOMS			1			
	Friedl et al 2017	166 2	208	· · ·	0.80	[0.74; 0.85]	28.1%
	Angulo et al 2018	115	128		- 0.90	[0.83; 0.94]	27.3%
	Heterogeneity: I^2 = 84%, 1	r ² = 0.0080, p	= 0.01		0.85	[0.74; 0.93]	55.5%
	grupo = PROACT						
	Venturino et al 2015	8	15 ———		0.53	[0.27; 0.79]	17.5%
	Noordhoff et al 2018	68	10		0.62	[0.52; 0.71]	27.0%
	Heterogeneity: $I^2 = 0\%$, τ^2	= 0, p = 0.52	20	~>	0.61	[0.52; 0.70]	44.5%
3	Random effects model Heterogeneity: $l^2 = 91\%$, 1	1 ² = 0.0235, p	161 o< 0.01		0.75	[0.59; 0.88]	100.0%
			0.3 0.4 0	.5 0.6 0.7 0.8 0.9)		
	Study	Events T	otal		Proportion	95%-CI	Weight
	grupo = ATOMS			:			
	Friedl et al 2017	130	167	÷	0.78	[0.71; 0.84]	29.8%
	Angulo et al 2018	53	62		- 0.85	[0.74; 0.93]	26.5%
	Random effects mode	I I	229	\sim	0.81	[0.73; 0.87]	56.4%
	Heterogeneity: $I^2 = 37\%$,	r ² = 0.0016, _j)= 0.21				
	grupo = PROACT	6	12		0.50	[0 21: 0 70]	15.6%
	Noordboff et al 2018	52	89	i	0.50	[0.47: 0.69]	28.0%
	Random effects mode Heterogeneity: $l^2 = 0\%$, τ^2	p = 0, p = 0.58	101		0.58	[0.48; 0.67]	43.6%
	Random effects mode	L	330		0.71	[0.56; 0.84]	100.0%
ر	Heterogeneity: $I^2 = 84\%$,	τ ² = 0.0188, μ	0.01 0.3 0.4	0.5 0.6 0.7 0.8 0	.9		

Fig 10. Forest plot of studies analyzed with durability of the devices at 12-months (A), 24-months (B) and 36-months (C).

Conclusion

1

The decision upon which is the best device to treat male stress incontinence in a particular patient is influenced by patient status, manual dexterity, incontinence severity and previous radiotherapy. According to this systematic review and meta-analysis both the ProACT and the ATOMS system appear efficacious and safe procedures to treat male stress incontinence. However, taking into account the statistical summary of effect size ATOMS is a more efficacious alternative compared to ProACT with higher dryness, improvement and patient satisfaction rates, lower explant rate and higher durability. It must however be noted that the ATOMS studies have shorter follow-up than the ProACT studies.

Supporting information

S1 Table. Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) 2009 Checklist.

(DOC)

S2 Table. Assessment of the risk of publication bias by means of Egger's linear regression test for the study outcomes evaluated. (DOCX) S3 Table. Random effect model (proportion estimate, 95% Confidence Interval) and quantifying heterogeneity (I2, p-value) for the study outcomes evaluated. (DOCX)

S1 Fig. Freeman-Tukey double arcsine transform proportion for dryness (A), improvement (B), satisfaction (C), differential pad-count (D), differential pad-test (E), number of fillings (F), explant (G) and complication (H). (TIF)

Author Contributions

Conceptualization: Javier C. Angulo.

Data curation: Javier C. Angulo, Sandra Schönburg, Alessandro Giammò, Francisco J. Abellán, David Lora.

Formal analysis: Javier C. Angulo, Sandra Schönburg, Francisco J. Abellán, David Lora.

Investigation: Javier C. Angulo, David Lora.

Methodology: Javier C. Angulo, David Lora.

Project administration: Javier C. Angulo.

Software: Javier C. Angulo, David Lora.

Supervision: Javier C. Angulo, Sandra Schönburg.

Validation: David Lora.

Visualization: Javier C. Angulo, Alessandro Giammò, Francisco J. Abellán, Ignacio Arance, David Lora.

Writing - original draft: Javier C. Angulo, Sandra Schönburg, David Lora.

Writing – review & editing: Javier C. Angulo, Sandra Schönburg, Alessandro Giammò, Francisco J. Abellán, Ignacio Arance, David Lora.

References

- Marzorati C, Monzani D, Mazzocco K, Pavan F, Cozzi G, De Cobelli O, et al. Predicting trajectories of recovery in prostate cancer patients undergone Robot-Assisted Radical Prostatectomy (RARP). PLoS One. 2019 Apr 4; 14(4):e0214682. https://doi.org/10.1371/journal.pone.0214682 PMID: 30946773
- Schaake W, van der Schaaf A, van Dijk LV, van den Bergh ACM, Langendijk JA. Development of a prediction model for late urinary incontinence, hematuria, pain and voiding frequency among irradiated prostate cancer patients. PLoS One. 2018 Jul 17; 13(7):e0197757. <u>https://doi.org/10.1371/journal.</u> pone.0197757 PMID: 30016325
- Van der Aa F, Drake MJ, Kasyan GR, Petrolekas A, Cornu JN. Young Academic Urologists Functional Urology Group. The artificial urinary sphincter after a quarter of a century: a critical systematic review of its use in male non-neurogenic incontinence. Eur Urol. 2013; 63:681–9. https://doi.org/10.1016/j. eururo.2012.11.034 PMID: 23219375
- Crivellaro S, Morlacco A, Bodo G, Agrò EF, Gozzi C, Pistolesi D, et al. Systematic review of surgical treatment of post radical prostatectomy stress urinary incontinence. Neurourol Urodyn. 2016; 35:875– 881. https://doi.org/10.1002/nau.22873 PMID: 26397171
- Chen YC, Lin PH, Jou YY, Lin VC. Surgical treatment for urinary incontinence after prostatectomy: A meta-analysis and systematic review. PLoS One. 2017 May 3; 12(5):e0130867. https://doi.org/10. 1371/journal.pone.0130867 PMID: 28467435

- Kretschmer A, bner W, Sandhu JS, Bauer RM. Evaluation and management of postprostatectomy incontinence: A systematic review of current literature. Eur Urol Focus. 2016; 2: 245–59. <u>https://doi.org/ 10.1016/j.euf.2016.01.002</u> PMID: 28723370
- Meisterhofer K, Herzog S, Strini KA, Sebastianelli L, Bauer R, Dalpiaz O. Male slings for postprostatectomy incontinence: A systematic review and meta-analysis. Eur Urol Focus. 2019 Feb 1. pii: S2405-4569(19)30008-2. [Epub ahead of print]
- Bauer W, Brössner C. Adjustable transobturator male system (ATOMS) for the treatment of post-prostatectomy urinary incontinence: the surgical technique. Pelviperineology 2011; 30: 10–8.
- Comiter CV, Dobberfuhl AD. The artificial urinary sphincter and male sling for postprostatectomy incontinence: Which patient should get which procedure? Investig Clin Urol. 2016; 57:3–13 https://doi.org/10. 4111/icu.2016.57.1.3 PMID: 26966721
- Hübner WA. ProACT for the treatment of male SUI. In Post-prostatectomy incontinence. Evaluation and management. Singla Ajay & Comiter Craig, Editors. pp 117–123. Springer International Publishing AG 2017
- 11. R Core Team. R:A language and envionment for statistical computing. R Foundation for Statistical Computing. Vienna, Austria; 2016. Available from https://www.R-project.org/.
- Hübner WA, Schlarp OM. Treatment of incontinence after prostatectomy using a new minimally invasive device: adjustable continence therapy. BJU Int. 2005; 96:587–94. <u>https://doi.org/10.1111/j.1464-410X.2005.05689.x PMID: 16104915</u>
- Trigo-Rocha F, Gomes CM, Pompeo AC, Lucon AM, Arap S. Prospective study evaluating efficacy and safety of Adjustable Continence Therapy (ProACT) for post radical prostatectomy urinary incontinence. Urology. 2006; 67(5):965–9. https://doi.org/10.1016/j.urology.2005.11.011 PMID: 16698356
- Kocjancic E, Crivellaro S, Ranzoni S, Bonvini D, Gontero P, Frea B. Adjustable Continence Therapy for the treatment of male stress urinary incontinence: a single-centre study. Scand J Urol Nephrol. 2007; 41:324–8. https://doi.org/10.1080/00365590601153488 PMID: 17763225
- Lebret T, Cour F, Benchetrit J, Grise P, Bernstein J, Delaporte V, et al. Treatment of postprostatectomy stress urinary incontinence using a minimally invasive adjustable continence balloon device, ProACT: results of a preliminary, multicenter, pilot study. Urology. 2008; 71:256–60. https://doi.org/10.1016/j. urology.2007.08.062 PMID: 18308096
- Gilling PJ, Bell DF, Wilson LC, Westenberg AM, Reuther R, Fraundorfer MR. An adjustable continence therapy device for treating incontinence after prostatectomy: a minimum 2-year follow-up. BJU Int. 2008; 102:1426–30. https://doi.org/10.1111/j.1464-410X.2008.07816.x PMID: 18564132
- Crivellaro S, Singla A, Aggarwal N, Frea B, Kocjancic E. Adjustable continence therapy (ProACT) and bone anchored male sling: Comparison of two new treatments of post prostatectomy incontinence. Int J Urol. 2008; 15:910–4. https://doi.org/10.1111/j.1442-2042.2008.02161.x PMID: 18761534
- Martens FM, Lampe MI, Heesakkers JP. ProACT for stress urinary incontinence after radical prostatectomy. Urol Int. 2009; 82:394–8. https://doi.org/10.1159/000218526 PMID: 19506404
- Gregori A, Romanò AL, Scieri F, Pietrantuono F, Incarbone GP, Salvaggio A, et al. Transrectal ultrasound-guided implantation of Adjustable Continence Therapy (ProACT): surgical technique and clinical results after a mean follow-up of 2 years. Eur Urol. 2010; 57:430–6. https://doi.org/10.1016/j.eururo. 2009.11.031 PMID: 19942340
- 20. García Matres MJ, Cansino Alcaide JR, Monasterio S, Rodríguez de Bethencourt F, Hidalgo Togores L, de la Peña Barthel J. Paraurethral adjustable continence therapy device (Pro-ACT) in the treatment of urinary incontinence after prostatic surgery. Arch Esp Urol. 2009; 62:871–81. https://doi.org/10.4321/s0004-06142009001000012 PMID: 20068264
- Giammò A, Bodo G, Castellano S, Borrè A, Carone R. Spiral multidetector computerized tomography evaluation of adjustable continence therapy implants. J Urol. 2010; 183:1921–6. https://doi.org/10. 1016/j.juro.2010.01.019 PMID: 20303116
- 22. Rouprêt M, Misraï V, Gosseine PN, Bart S, Cour F, Chartier-Kastler E. Management of stress urinary incontinence following prostate surgery with minimally invasive adjustable continence balloon implants: functional results from a single center prospective study. J Urol. 2011; 186:198–203. https://doi.org/10. 1016/j.juro.2011.03.016 PMID: 21575974
- Kjær L, Fode M, Nørgaard N, Sønksen J, Nordling J. Adjustable continence balloons: clinical results of a new minimally invasive treatment for male urinary incontinence. Scand J Urol Nephrol. 2012; 46:196– 200. https://doi.org/10.3109/00365599.2012.660986 PMID: 22364390
- Crivellaro S, Tosco L, Palazzetti A, Abbinante M, Martinez G, Kocjancic E, et al. Geometrical stepperguided navigation system for ProACT implant under transrectal ultrasound control: preliminary data. Urol Int. 2012; 89:473–9. https://doi.org/10.1159/000341991 PMID: 22965196

- Gatti L, Moroni A, Cristinelli L, Benedusi F, Antonelli A, Zani D, et al. ProACT (Adjustable Continence Therapy) implants in case of failure of other techniques for urinary incontinence. Urologia. 2012; 79 Suppl 19:46–9.
- Utomo E, Groen J, Vroom IH, van Mastrigt R, Blok BF. Urodynamic effects of volume-adjustable balloons for treatment of postprostatectomy urinary incontinence. Urology. 2013; 81:1308–14. <u>https://doi.org/10.1016/j.urology.2013.01.020 PMID: 23465144</u>
- Venturino L, Dalpiaz O, Pummer K, Primus G. Adjustable continence balloons in men: Adjustments do not translate into long-term continence. Urology. 2015; 85:1448–52. https://doi.org/10.1016/j.urology. 2015.01.045 PMID: 26099887
- Baron MG, Delcourt C, Nouhaud FX, Gillibert A, Pfister C, Grise P, et al. Sequential treatment with ProACTTM device implantation after male sling failure for male urinary incontinence. Progrès en urologie 2017; 27:1098–1103. https://doi.org/10.1016/j.purol.2017.08.001 PMID: 28847446
- Nash S, Aboseif S, Gilling P, Gretzer M, Samowitz H, Rose M, et al. Treatment with an adjustable longterm implant for post-prostatectomy stress incontinence: The ProACT[™] pivotal trial. Neurourol Urodyn. 2018; 37:2854–2859. https://doi.org/10.1002/nau.23802 PMID: 30178536
- Nestler S, Thomas C, Neisius A, Rubenwolf P, Roos F, Hampel C, et al. Long-term results of ProACT primary and repeat implantation for treatment of stress urinary incontinence in men. World J Urol. 2019; 37:1173–1179. https://doi.org/10.1007/s00345-018-2483-4 PMID: 30225800
- Noordhoff TC, Scheepe JR, Blok BFM. Outcome and complications of adjustable continence therapy (proACTTM) after radical prostatectomy: 10 years' experience in 143 patients. Neurourol Urodyn. 2018; 37:1419–1425. https://doi.org/10.1002/nau.23463 PMID: 29266406
- 32. Finazzi Agrò E, Gregori A, Bianchi D, Carone R, Giammò A, Ammirati E, et al. Efficacy and safety of adjustable balloons (Proact[™]) to treat male stress urinary incontinence after prostate surgery: Medium and long-term follow-up data of a national multicentric retrospective study. Neurourol Urodyn. 2019; 38:1979–1984. https://doi.org/10.1002/nau.24103 PMID: 31302928
- Hoda MR, Primus G, Schumann A, Fischereder K, von Heyden B, Schmid N, et al. Treatment of stress urinary incontinence after radical prostatectomy: adjustable transobturator male system—results of a multicenter prospective observational study. Urologe A. 2012; 51:1576–83. <u>https://doi.org/10.1007/</u> s00120-012-2950-2 PMID: 22836942
- Seweryn J, Bauer W, Ponholzer A, Schramek P. Initial experience and results with a new adjustable transobturator male system for the treatment of stress urinary incontinence. J Urol. 2012; 187:956–61. https://doi.org/10.1016/j.juro.2011.10.138 PMID: 22264469
- 35. Hoda MR, Primus G, Fischereder K, Von Heyden B, Mohammed N, Schmid N, et al. Early results of a European multicentre experience with a new self-anchoring adjustable transobturator system for treatment of stress urinary incontinence in men. BJU Int. 2013; 111:296–303. https://doi.org/10.1111/j.1464-410X.2012.11482.x PMID: 23186285
- Krause J, Tietze S, Behrendt W, Nast J, Hamza A. Reconstructive surgery for male stress urinary incontinence: Experiences using the ATOMS([®]) system at a single center. GMS Interdiscip Plast Reconstr Surg DGPW. 2014 Dec 17; 3:Doc15. <u>https://doi.org/10.3205/iprs000056</u> eCollection 2014. PMID: 26504726
- González SP, Cansino JR, Portilla MA, Rodriguez SC, Hidalgo L, De la Peña J. First experience with the ATOMS(®) implant, a new treatment option for male urinary incontinence. Cent European J Urol. 2014; 67:387–91. https://doi.org/10.5173/ceju.2014.04.art14 PMID: 25667760
- Friedl A, Bauer W, Rom M, Kivaranovic D, Lüftenegger W, Brössner C. Sexuality and erectile function after implantation of an Adjustable Transobturator Male System (ATOMS) for urinary stress incontinence. A multi-institutional prospective study. Arch Ital Urol Androl. 2016; 87:306–11. https://doi.org/10. 4081/aiua.2015.4.306 PMID: 26766803
- 39. Mühlstädt S, Friedl A, Mohammed N, Schumann A, Weigand K, Kawan F, et al. Five-year experience with the adjustable transobturator male system for the treatment of male stress urinary incontinence: a single-center evaluation. World J Urol. 2017; 35:145–151. https://doi.org/10.1007/s00345-016-1839-x PMID: 27156092
- 40. Friedl A, Mühlstädt S, Rom M, Kivaranovic D, Mohammed N, Fornara P, et al. Risk factors for treatment failure with the adjustable transobturator male system incontinence device: Who will succeed, Who will fail? Results of a multicenter study. Urology. 2016; 90:189–94. https://doi.org/10.1016/j.urology.2015. 12.044 PMID: 26773347
- Hüsch T, Kretschmer A, Thomsen F, Kronlachner D, Kurosch M, Obaje A, et al. Debates on Male Incontinence (DOMINO)-Project. Risk factors for failure of male slings and artificial urinary sphincters: Results from a large middle European cohort study. Urol Int. 2017; 99:14–21. https://doi.org/10.1159/ 000449232 PMID: 27598774

- **42.** Buresova E, Vidlar A, Grepl M, Student V Jr, Student V. Single-centre experience in using the adjustable transobturator male system in treatment of stress urinary incontinence in patients after radical prostatectomy. Journal of Clinical Urology 1–7. https://doi.org/10.1177/2051415817701054
- 43. Friedl A, Mühlstädt S, Zachoval R, Giammò A, Kivaranovic D, Rom M, et al. Long-term outcome of the adjustable transobturator male system (ATOMS): results of a European multicentre study. BJU Int. 2017; 119:785–792. https://doi.org/10.1111/bju.13684 PMID: 27868328
- Friedl A, Schneeweiss J, Stangl K, Mühlstädt S, Zachoval R, Hruby S, et al. The adjustable transobturator male system in stress urinary incontinence after transurethral resection of the prostate. Urology. 2017; 109:184–189. https://doi.org/10.1016/j.urology.2017.07.004 PMID: 28712889
- 45. Angulo J, Arance I, Esquinas C, Dorado J, Marcelino J, Martins F. Outcome measures of adjustable transobturator male system with pre-attached scrotal port for male stress urinary incontinence after radical prostatectomy. Adv Ther. 2017; 34: 1173–83. https://doi.org/10.1007/s12325-017-0528-5 PMID: 28405960
- 46. Manso M, Alexandre B, Antunes-Lopes T, Martins-da-Silva C, Cruz F. Is the adjustable transobturator system ATOMS® useful for the treatment of male urinary incontinence in low to medium volume urological centers? Actas Urol Esp. 2018; 42:267–272 <u>https://doi.org/10.1016/j.acuro.2017.07.011</u> PMID: 29174630
- Angulo JC, Cruz F, Esquinas C, Arance I, Manso M, Rodríguez A, et al. Treatment of male stress urinary incontinence with the adjustable transobturator male system: Outcomes of a multi-center Iberian study. Neurourol Urodyn. 2018; 37:1458–1466. https://doi.org/10.1002/nau.23474 PMID: 29315765
- Esquinas C, Arance I, Pamplona J, Moraga A, Dorado JF, Angulo JC. Treatment of stress urinary incontinence after prostatectomy with the adjustable transobturator male system (ATOMS®) with preattached scrotal port. Actas Urol Esp 2018; 42:473–82. https://doi.org/10.1016/j.acuro.2018.02.005 PMID: 29642999
- Angulo JC, Fonseca J, Esquinas C, Ojea A, Rodríguez A, Rabassa M, et al. Adjustable transobturator male system (ATOMS®) as treatment of stress urinary incontinence secondary to transurethral resection of the prostate. Actas Urol Esp. 2018; 42:567–573. https://doi.org/10.1016/j.acuro.2018.05.002 PMID: 29929737
- Angulo JC, Esquinas C, Arance I, Rodríguez A, Pereira J, Rabassa M, et al. Adjustable transobturator male system after failed surgical devices for male stress urinary incontinence: A feasibility study. Urol Int. 2018; 101:106–113 https://doi.org/10.1159/000489316 PMID: 29953998
- Giammò A, Ammirati E, Tullio A, Bodo G, Manassero A, Gontero P, et al. Implant of ATOMS® system for the treatment of postoperative male stress urinary incontinence: results of a single centre. Int Braz J Urol.2019; 45:127–136. https://doi.org/10.1590/S1677-5538.IBJU.2018.0171 PMID: 30521175
- 52. Doiron RC, Saavedra A, Haines T, Nadeau G, Tu LM, Morisset J, et al. Canadian Experience with the Adjustable Transobturator Male System for Post-Prostatectomy Incontinence: A Multicenter Study. J Urol. 2019 Jun 28:101097JU0000000000420.
- Esquinas C, Angulo JC. Effectiveness of adjustable transobturator male system (ATOMS) to treat male stress incontinence: A systematic review and meta-analysis. Adv Ther. 2019; 36:426–441. <u>https://doi.org/10.1007/s12325-018-0852-4</u> PMID: 30560525
- Larson T, Jhaveri H, Yeung LL. Adjustable continence therapy (ProACT) for the treatment of male stress urinary incontinence: A systematic review and meta-analysis. Neurourol Urodyn. 2019 Aug 20. https://doi.org/10.1002/nau.24135 [Epub ahead of print] PMID: 31429982
- 55. Angulo JC, Arance I, Ojea A, Carballo M, Rodríguez A, Pereira J, et al. Patient satisfaction with adjustable transobturator male system in the Iberian multicenter study. World J Urol. 2019 Jan 16. <u>https://doi.org/10.1007/s00345-019-02639-4</u> [Epub ahead of print] PMID: 30649591
- 56. Nash S, Aboseif S, Gilling P, Gretzer M, Samowitz H, Rose M, et al. Four-year follow-up on 68 patients with a new post-operatively adjustable long-term implant for post-prostatectomy stress incontinence: ProACT[™]. Neurourol Urodyn. 2019; 38:248–253. https://doi.org/10.1002/nau.23838 PMID: 30311667