

# Causal relationship between type I diabetes mellitus and atrial fibrillation: A Mendelian randomization study

Yongkai Li <sup>a,1</sup>, Shasha Liu <sup>b,1</sup>, Yiming Dong <sup>b,1</sup>, Jianzhong Yang <sup>c,\*</sup>, Yingping Tian <sup>a,\*\*</sup>

<sup>a</sup> Emergency Department, The Second Hospital of Hebei Medical University, Shijiazhuang, Hebei, China

<sup>b</sup> Emergency Department, The Fourth Hospital of Shijiazhuang, Shijiazhuang, China

<sup>c</sup> Emergency Trauma Center, The First Affiliated Hospital of Xinjiang Medical University, Urumqi, China

## ARTICLE INFO

### Keywords:

Type 1 diabetes mellitus  
Atrial fibrillation  
Genome-wide association  
Mendelian randomization

## ABSTRACT

**Background:** Patients with type 1 diabetes mellitus have been at heightened risk for developing atrial fibrillation. We aimed to investigate whether this association is causal using Mendelian randomization.

**Methods:** Using publicly available genome-wide association studies data, we selected single nucleotide polymorphisms significantly associated with type 1 diabetes mellitus as instrumental variables. We employed inverse variance-weighted, weighted median, MR-Egger regression, simple mode, and weighted mode methods within a two-sample Mendelian randomization framework to assess the causal relationship between type 1 diabetes mellitus and atrial fibrillation. We evaluated the pleiotropy and heterogeneity levels of the included genetic instruments using MR-PRESSO, MR-Egger intercept test, Cochran's Q test, funnel plots, and leave-one-out plots.

**Results:** Causal impact of type 1 diabetes mellitus on atrial fibrillation: Inverse variance weighted (odds ratio [OR] = 0.996, 95 % confidence interval [CI]: 0.985–1.007,  $P = 0.498$ ). MR-Egger (OR = 1.000, 95 % CI: 0.985–1.016,  $P = 0.963$ ). Weighted median (OR = 0.985, 95 % CI: 0.973–0.998,  $P = 0.022$ ). Simple mode (OR = 1.007, 95 % CI: 0.974–1.040,  $P = 0.698$ ). Weighted mode (OR = 0.995, 95 % CI: 0.984–1.005,  $P = 0.298$ ). MR-Egger intercept test ( $P = 0.437$ ). There was no evidence of pleiotropy among the genetic instrumental variables included in the analysis.

**Conclusions:** In Mendelian randomization analysis, we did not find evidence of a causal relationship between genetically determined type 1 diabetes mellitus in European ancestry populations and atrial fibrillation.

## 1. Introduction

Type 1 diabetes mellitus (T1DM) is a common chronic autoimmune disease characterized by beta cells destruction leading to insufficient insulin secretion. However, its pathogenesis remains incompletely understood, likely involving interactions among susceptibility genes, autoantigens, and environmental factors [1]. Atrial fibrillation (AF) is a rapid and irregular atrial rhythm that is believed to result from chaotic re-entry involving multiple wavelets within the atria. Its incidence and prevalence are increasing both in the United States and globally [2]. In 2020, atrial fibrillation was estimated globally to affect 50 million individuals and is a leading cause of stroke-related disability and mortality [3]. In recent years, the incidence of T1DM has been steadily rising,

imposing significant burdens on families and economies. Most existing data on the vicious cycle between AF and diabetes pertain to type 2 diabetes mellitus (T2DM), highlighting it as a key contributing factor to the occurrence of atrial fibrillation [2,4]. However, controversy remains regarding the relationship between T1DM and atrial fibrillation. Only a few clinical studies have demonstrated the independent role of T1DM as a risk factor for the development of atrial fibrillation, and this remains a topic of discussion. These findings are inconclusive, possibly confounded by other abnormalities associated with T1DM, including obesity and hypertension [5–7]. While the long-term macrovascular and microvascular complications of diabetes have been well-established, the relationship between T1DM and AF remains unclear, and causality has not been established.

\* Corresponding author at: Emergency Trauma Center, The First Affiliated Hospital of Xinjiang Medical University, Urumqi, China.

\*\* Corresponding author at: Emergency Department, The Second Hospital of Hebei Medical University, Shijiazhuang, Hebei, China.

E-mail addresses: [yongkaiyaya@sina.com](mailto:yongkaiyaya@sina.com), [29405814@hebmu.edu.cn](mailto:29405814@hebmu.edu.cn) (Y. Li), [sjhlssdyx@sina.com](mailto:sjhlssdyx@sina.com) (S. Liu), [958111368@qq.com](mailto:958111368@qq.com) (Y. Dong), [yjz6542@126.com](mailto:yjz6542@126.com) (J. Yang), [tianyingping2021@163.com](mailto:tianyingping2021@163.com) (Y. Tian).

<sup>1</sup> These authors have contributed equally to this work.

Mendelian randomization (MR) is a study design where random allocation of genetic variants at conception is used as instrumental variables to estimate the unconfounded causal effects of risk factors on disease[8]. Therefore, this study aims to investigate the causal association between T1DM and AF using Mendelian randomization, aiming to provide insights for clinical prevention and treatment strategies regarding T1DM and AF.

## 2. Methods

### 2.1. Study design

This study considers T1DM as the exposure factor and atrial fibrillation as the outcome variable, utilizing publicly available genome-wide association studies (GWAS) as the data source. We select single nucleotide polymorphisms (SNPs) significantly associated with T1DM as instrumental variables (IVs) for a two-sample Mendelian randomization analysis of the causal relationship between T1DM and AF. We employ Cochran's Q test to assess heterogeneity in the results and conduct sensitivity analyses to validate their reliability. To ensure effective IVs for causal association analysis in two-sample Mendelian randomization studies, this study establishes three core assumptions: 1) IVs are significantly associated with the exposure factor T1DM; 2) IVs are unrelated to all confounding factors of T1DM and the outcome variable AF; 3) IVs influence the outcome variable AF solely through the exposure factor T1DM and not through other pathways (Fig. 1), (Additional file 1: Fig. S1).

Fig. 1 Study overview. GWAS, genome-wide association studies; T1DM, type 1 diabetes mellitus; AF; atrial fibrillation; SNPs, single nucleotide polymorphisms; MR, mendelian randomization; IVW, Inverse variance weighted; MR-PRESSO, Mendelian randomization pleiotropy residual sum and outlier; MR-Egger, Mendelian randomization-Egger.

### 2.2. Data sources

We obtained summary statistical data for T1DM and AF from the IEU Open GWAS project (<https://gwas.mrcieu.ac.uk/>). The GWAS ID for T1DM is ebi-a-GCST90014023, comprising 59,999,551 SNPs, with

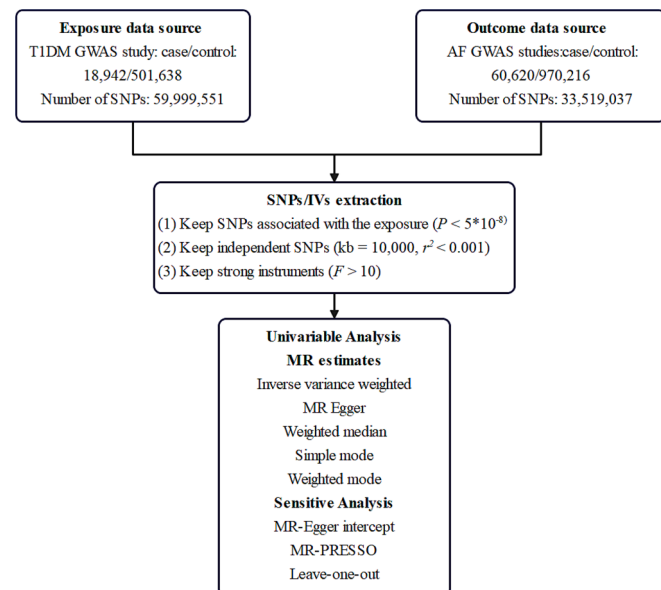


Fig. 1. Study overview. GWAS, genome-wide association studies; T1DM, type 1 diabetes mellitus; AF; atrial fibrillation; SNPs, single nucleotide polymorphisms; MR, mendelian randomization; IVW, Inverse variance weighted; MR-PRESSO, Mendelian randomization pleiotropy residual sum and outlier; MR-Egger, Mendelian randomization-Egger.

summary statistics for European ancestry including 18,942 cases and 520,580 controls[9]. This represents a comprehensive, interdisciplinary data resource spanning continents, still publicly accessible in the life sciences field. It is the largest and most recent GWAS study on T1DM. For atrial fibrillation, the GWAS ID is ebi-a-GCST006414, comprising 33,519,037 SNPs, with summary statistics for European ancestry including 60,620 cases and 970,216 controls[10]. All data were sourced from public databases and do not require additional ethical review.

The selection of instrumental variables proceeded as follows: Firstly, SNPs closely associated with T1DM were screened with a threshold of  $P < 5e-08$ . Secondly, linkage disequilibrium (LD) was adjusted for; LD refers to non-random associations between alleles at different loci, indicating some degree of genetic linkage. Parameters were set at  $r^2 = 0.001$  and  $kb = 10000$  to ensure independence among genetic instruments. To assess the potential impact of weak instrumental variables, F-statistics were computed using the formula  $F = R^2(N-2)/(1-R^2)$ [11]. SNPs with an F-statistic greater than 10 were deemed strongly associated with atrial fibrillation, while those with an F-statistic  $< 10$  indicated possible weak instrumental variable bias and were therefore excluded to mitigate potential effects on results[12]. Finally, the PhenoScanner database was utilized to investigate secondary phenotypes associated with each SNP, removing any confounding factors related to the exposure and outcome variables.

### 2.3. Statistical analysis

We conducted a two-sample Mendelian randomization analysis using the TwoSampleMR package in R (v0.6.1; <https://mrcieu.github.io/TwoSampleMR/>) to assess the relationship between T1DM and AF. This study employed five Mendelian randomization methods to determine the causal relationship between T1DM and AF, including inverse variance weighted (IVW), weighted median (WM), Mendelian randomization-Egger (MR-Egger), simple mode, and weighted mode. The Mendelian randomization-Egger method assesses pleiotropy and adjusts for intercept presence in regression; convergence with IVW indicates minimal horizontal pleiotropy among instruments[13]. WM complements MR-Egger regression by providing consistent estimates if at least 50 % of weights originate from valid IVs[14]. Heterogeneity was assessed using Cochran's Q test, with a significance threshold of  $P < 0.05$  indicating presence treated via random effects models; fixed effects models were used otherwise. Leave-one-out analysis evaluated individual SNP impacts on overall results, while scatter plots visualized SNP effects on exposure and outcomes. Forest plots depicted causal associations between exposure and outcomes, and funnel plots assessed selection bias per SNP. For sensitivity analysis, Mendelian randomization pleiotropy residual sum and outlier (MR-PRESSO) test identified horizontal pleiotropic outliers in multi-instrument summary-level Mendelian randomization[15]. All analyses were conducted using R 4.4.0.

Power calculations for Mendelian randomization were performed according to the formula proposed by Brion[16]. We estimated the minimum effect size to be able to observe the measured outcome (atrial fibrillation) using a total sample size of 1,030,836 with 60,620 cases of atrial fibrillation, an  $\alpha$  threshold of 0.05, variance explained for each specific instrumental variable, and power  $> 80\%$ .

## 3. Results

### 3.1. Selection of SNPs

We ultimately selected 89 SNPs that met the criteria, all with F-values  $> 10$ , strongly associated with atrial fibrillation, and these instrumental variables only influence the outcome event of AF through the exposure factor of T1DM. Following harmonization with AF, the final Mendelian randomization analysis was conducted using 86 SNPs (Table 1).

**Table 1**

Instrumental variables associated with type I diabetes mellitus and atrial fibrillation.

| SNPs       | CHr | EA | NEA | $\beta$ value | S. E. | F value  |
|------------|-----|----|-----|---------------|-------|----------|
| rs9517712  | 13  | C  | T   | -0.102        | 0.741 | 41.701   |
| rs9468618  | 6   | T  | C   | -0.272        | 0.082 | 106.730  |
| rs9385401  | 6   | T  | C   | 0.120         | 0.454 | 65.961   |
| rs9260802  | 6   | G  | A   | -0.356        | 0.035 | 72.925   |
| rs855330   | 1   | C  | T   | 0.111         | 0.259 | 43.219   |
| rs8046043  | 16  | C  | G   | -0.085        | 0.392 | 31.070   |
| rs78325861 | 10  | G  | C   | -0.282        | 0.039 | 44.692   |
| rs7795896  | 7   | T  | C   | -0.135        | 0.692 | 68.065   |
| rs7776597  | 7   | G  | A   | 0.244         | 0.959 | 45.156   |
| rs7752257  | 6   | G  | T   | -0.680        | 0.740 | 1677.550 |
| rs7668577  | 4   | C  | A   | 0.094         | 0.312 | 37.947   |
| rs7511678  | 1   | A  | G   | 0.094         | 0.229 | 30.650   |
| rs73432769 | 6   | T  | C   | 0.335         | 0.020 | 53.693   |
| rs7301381  | 12  | C  | T   | -0.094        | 0.464 | 43.085   |
| rs7237497  | 18  | C  | T   | -0.220        | 0.839 | 139.966  |
| rs722988   | 10  | C  | T   | 0.083         | 0.353 | 32.887   |
| rs7130222  | 11  | G  | T   | -0.092        | 0.308 | 32.324   |
| rs7068821  | 10  | T  | G   | -0.165        | 0.251 | 102.182  |
| rs6908626  | 6   | T  | G   | 0.203         | 0.166 | 120.055  |
| rs6908236  | 6   | C  | A   | 0.166         | 0.514 | 138.290  |
| rs689      | 11  | T  | A   | 0.712         | 0.731 | 1476.890 |
| rs6679677  | 1   | A  | C   | 0.642         | 0.114 | 939.332  |
| rs663743   | 11  | A  | G   | -0.100        | 0.349 | 43.872   |
| rs6434435  | 2   | A  | G   | -0.123        | 0.160 | 41.421   |
| rs61839660 | 10  | T  | C   | -0.357        | 0.085 | 188.999  |
| rs61759532 | 17  | T  | C   | 0.118         | 0.235 | 40.563   |
| rs607703   | 11  | T  | C   | 0.092         | 0.484 | 41.520   |
| rs601338   | 19  | A  | G   | 0.127         | 0.479 | 77.695   |
| rs574384   | 1   | A  | C   | -0.134        | 0.895 | 31.311   |
| rs57209021 | 17  | T  | C   | 0.101         | 0.226 | 30.286   |
| rs56994090 | 14  | C  | T   | -0.134        | 0.431 | 84.627   |
| rs55993634 | 16  | G  | C   | 0.219         | 0.085 | 80.970   |
| rs55893453 | 2   | G  | A   | 0.095         | 0.202 | 29.866   |
| rs4820827  | 22  | C  | T   | -0.130        | 0.621 | 82.590   |
| rs4548024  | 6   | C  | T   | -0.096        | 0.234 | 32.853   |
| rs4490209  | 2   | G  | C   | -0.085        | 0.360 | 29.901   |
| rs41295159 | 10  | G  | C   | -0.700        | 0.009 | 60.079   |
| rs3802214  | 8   | C  | T   | -0.107        | 0.799 | 30.735   |
| rs35327136 | 17  | A  | C   | -0.119        | 0.156 | 39.448   |
| rs34593439 | 15  | A  | G   | -0.218        | 0.108 | 81.748   |
| rs34536443 | 19  | C  | G   | -0.385        | 0.043 | 100.073  |
| rs3184504  | 12  | C  | T   | -0.231        | 0.533 | 270.096  |
| rs3087243  | 2   | A  | G   | -0.199        | 0.422 | 196.595  |
| rs3024493  | 1   | A  | C   | -0.164        | 0.154 | 69.597   |
| rs28752526 | 6   | G  | A   | 0.630         | 0.357 | 1569.786 |
| rs28648882 | 2   | A  | G   | 0.099         | 0.224 | 30.686   |
| rs2611211  | 4   | T  | C   | -0.144        | 0.824 | 59.248   |
| rs2543537  | 22  | T  | C   | -0.083        | 0.460 | 33.971   |
| rs2493411  | 1   | C  | T   | 0.127         | 0.132 | 32.364   |
| rs2429557  | 6   | A  | T   | 1.021         | 0.011 | 308.356  |
| rs238873   | 6   | G  | A   | 0.793         | 0.015 | 221.481  |
| rs238265   | 13  | G  | T   | -0.091        | 0.695 | 35.903   |
| rs231972   | 16  | C  | A   | 0.171         | 0.118 | 65.750   |
| rs229527   | 22  | A  | C   | 0.104         | 0.416 | 54.192   |
| rs2188962  | 5   | T  | C   | 0.079         | 0.404 | 31.776   |
| rs2111485  | 2   | G  | A   | 0.128         | 0.604 | 77.960   |
| rs202535   | 20  | A  | C   | -0.141        | 0.828 | 58.748   |
| rs1947178  | 8   | G  | A   | -0.103        | 0.792 | 36.325   |
| rs1881146  | 2   | T  | A   | -0.095        | 0.311 | 29.892   |
| rs1808094  | 18  | C  | T   | -0.114        | 0.524 | 58.180   |
| rs1794269  | 6   | T  | C   | 1.594         | 0.432 | 8073.918 |
| rs17623914 | 1   | C  | T   | -0.135        | 0.100 | 33.282   |
| rs17323934 | 7   | G  | C   | -0.130        | 0.223 | 59.440   |
| rs17106304 | 14  | G  | C   | 0.115         | 0.656 | 60.648   |
| rs1701704  | 12  | G  | T   | 0.244         | 0.339 | 281.025  |
| rs1611236  | 6   | A  | G   | -0.260        | 0.299 | 223.172  |
| rs1574285  | 9   | T  | G   | -0.127        | 0.591 | 79.746   |
| rs1350275  | 14  | G  | T   | -0.094        | 0.698 | 37.561   |
| rs13259300 | 8   | C  | A   | -0.092        | 0.598 | 39.498   |
| rs13147049 | 4   | G  | A   | -0.110        | 0.641 | 55.593   |
| rs12927355 | 16  | T  | C   | -0.204        | 0.316 | 180.198  |
| rs12742756 | 1   | G  | A   | -0.083        | 0.428 | 30.384   |
| rs12644686 | 4   | G  | C   | -0.108        | 0.194 | 31.104   |

**Table 1 (continued)**

| SNPs        | CHr | EA | NEA | $\beta$ value | S. E. | F value |
|-------------|-----|----|-----|---------------|-------|---------|
| rs12464462  | 2   | G  | A   | -0.088        | 0.410 | 37.615  |
| rs12257077  | 10  | T  | C   | 0.231         | 0.032 | 39.156  |
| rs12128789  | 1   | C  | T   | 0.127         | 0.132 | 34.762  |
| rs114378220 | 5   | T  | C   | 0.178         | 0.070 | 34.147  |
| rs113374757 | 19  | T  | C   | -0.171        | 0.163 | 68.009  |
| rs112733823 | 6   | T  | C   | 0.383         | 0.142 | 362.648 |
| rs11203203  | 21  | A  | G   | 0.144         | 0.349 | 99.657  |
| rs10844597  | 12  | A  | G   | -0.090        | 0.506 | 41.461  |
| rs10801128  | 1   | G  | A   | 0.096         | 0.717 | 37.536  |
| rs10751776  | 1   | C  | A   | 0.078         | 0.510 | 30.935  |
| rs1050979   | 6   | G  | A   | 0.106         | 0.515 | 56.493  |
| rs10275896  | 7   | C  | T   | -0.121        | 0.232 | 53.488  |
| rs10224046  | 7   | G  | T   | 0.086         | 0.324 | 30.904  |

### 3.2. Analysis of the Mendelian randomization method

Our analysis revealed the following results: The Inverse variance weighted (IVW) method showed no statistically significant difference (OR = 0.996, 95 % CI: 0.985–1.007,  $P = 0.498$ ). Similarly, the MR Egger method also showed no statistically significant difference (OR = 1.000, 95 % CI: 0.985–1.016,  $P = 0.963$ ). In contrast, the Weighted median method demonstrated a statistically significant difference (OR = 0.985, 95 % CI: 0.973–0.998,  $P = 0.022$ ). Both the Simple mode (OR = 1.007, 95 % CI: 0.974–1.040,  $P = 0.698$ ) and Weighted mode methods (OR = 0.995, 95 % CI: 0.984–1.005,  $P = 0.298$ ) showed no statistically significant differences. Among these methods, the IVW method, which assumes no pleiotropy and does not account for intercept presence in regression, is considered the primary analysis method[17]. Therefore, it can still be concluded that there is no causal relationship between the exposure factor and the outcome, indicating that T1DM does not increase the risk of AF occurrence (Table 2, Figs. 2, 3).

### 3.3. Heterogeneity and pleiotropy

Heterogeneity assessed by IVW method analysis indicated Cochran's  $Q = 175.747$ ,  $P < 0.05$ . Similarly, MR-Egger method analysis showed Cochran's  $Q = 174.482$ ,  $P < 0.05$ . Both methods revealed heterogeneity among individual SNPs. MR-PRESSO results in Global Test  $P < 0.001$ ; however, MR-PRESSO results in the Distortion test,  $P = 0.922$ , indicated no difference in outlier values affecting the outcomes, thus a random effects model was employed. Pleiotropy analysis showed MR-Egger intercept = -0.001,  $P = 0.437$ , indicating no horizontal pleiotropy and robust results. Funnel plot results indicated a symmetric distribution of causal effect estimates of included SNPs, suggesting no bias and stable outcomes (Fig. 4).

### 3.4. Sensitivity analysis

Using the “Leave-one-out plot” to assess the influence of each SNP on the Mendelian randomization analysis results, no sensitive SNPs were identified (Fig. 5).

**Table 2**

Causal relationship between T1DM and AF.

| Method                          | Odds Ratio (OR) | 95 % Confidence Interval (CI) | P value |
|---------------------------------|-----------------|-------------------------------|---------|
| Inverse Variance Weighted (IVW) | 0.996           | 0.985–1.007                   | 0.498   |
| MR Egger                        | 1.000           | 0.985–1.016                   | 0.963   |
| Weighted Median                 | 0.985           | 0.973–0.998                   | 0.022   |
| Simple Mode                     | 1.007           | 0.974–1.040                   | 0.698   |
| Weighted Mode                   | 0.995           | 0.984–1.005                   | 0.298   |

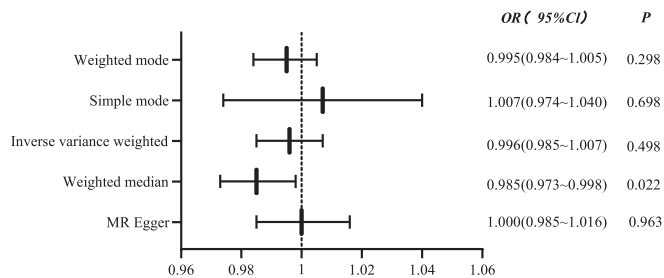


Fig. 2. MR results between T1DM and AF.

4. Discussion

In current studies, T1DM has been identified as a potential independent risk factor for AF. However, in the current Mendelian randomization study involving over one million individuals of European ancestry, there is no convincing evidence to suggest a causal relationship between T1DM and AF. Our study findings indicate no evidence of a causal relationship between T1DM and AF, underscoring the need for further clinical research to consolidate existing conclusions.

AF and diabetes mellitus (DM) are two common comorbidities in Western countries. The prevalence of AF among adults in Europe, the United States, and Australia is reported to range from 1 % to 4 % [18]. The association between diabetes and AF occurrence is primarily observed in populations diagnosed with diabetes, suggesting that prolonged exposure to high blood glucose may increase the risk of AF. While the mechanisms underlying this association and other potential

pathways remain hypothetical, there is scarce supportive functional data. However, our study findings indicate that genetic T1DM is not associated with AF, suggesting that these disruptions in glucose metabolism are not causative factors for AF. A Mendelian randomization study has indicated no causal relationship between T2DM or glucose abnormalities and the risk of AF [19]. Some studies have shown an association between the two, with one potential explanation being that diabetes-related comorbidities such as hypertension and obesity contribute to the causal link between diabetes and AF. Indeed, recent Mendelian randomization analyses have demonstrated a causal relationship between BMI and the incidence of AF [20]. A plausible mechanism may involve the comorbid conditions associated with Type 1 diabetes (T1D), specifically hypertension and obesity, which could mediate the pathophysiological link between T1D and atrial fibrillation (AF).

A Swedish population-based cohort study employed Cox proportional hazards models to assess the relationship between blood glucose status and atrial fibrillation. In models adjusted for age and sex, significant correlations were observed between blood glucose status and AF incidence across all groups except for impaired glucose tolerance, with the diabetes group showing the strongest correlation ( $P < 0.001$ ). However, in models further adjusted for factors including gender, age, systolic blood pressure, BMI, antihypertensive medications, cholesterol, alcohol consumption, smoking, education level, marital status, and physical activity, no significant association was found between blood glucose status and AF. These findings suggest that after adjusting for potential confounders, the association between blood glucose status and AF diminishes. Diabetes and prediabetes do not appear to be

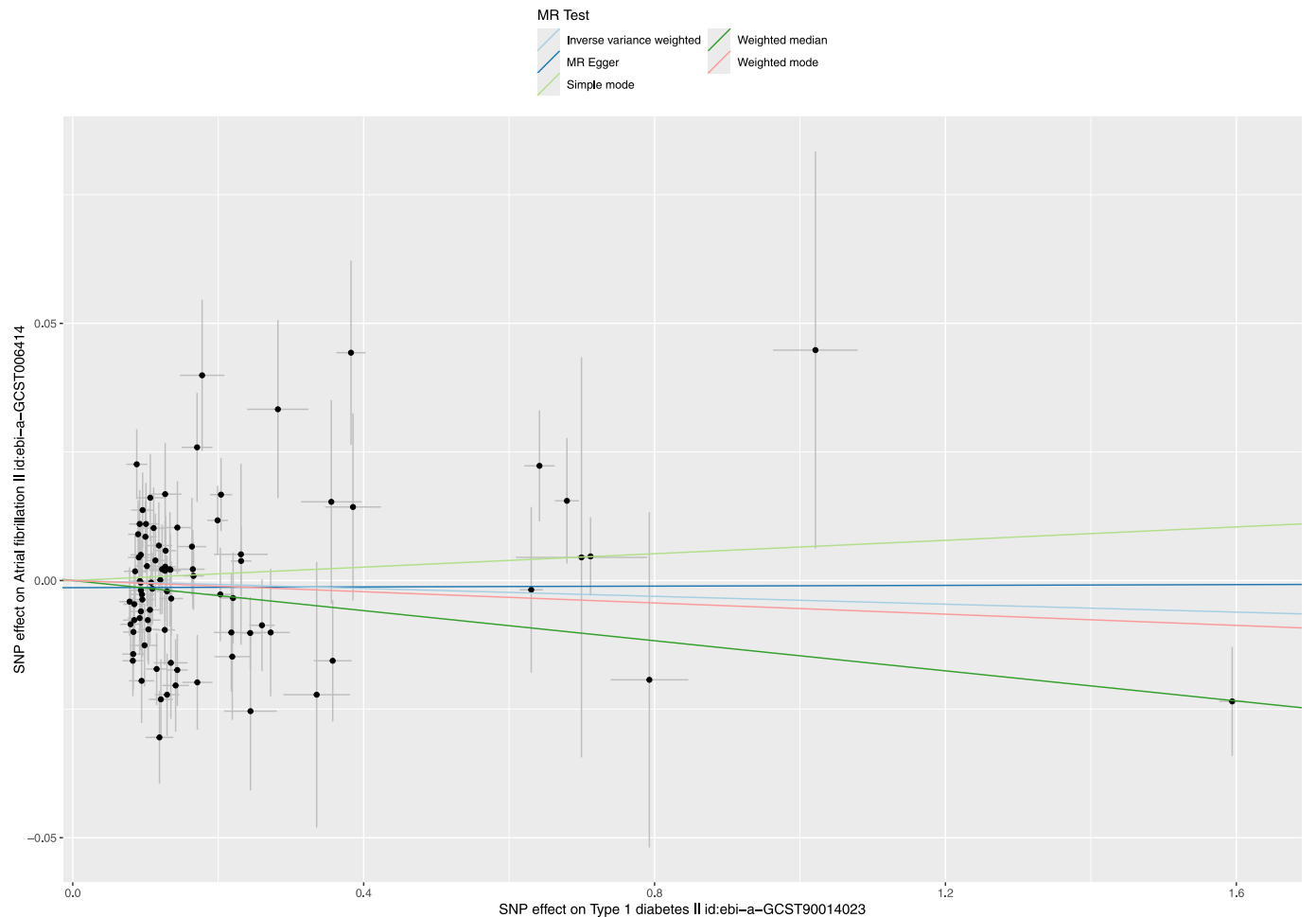


Fig. 3. Causal relationship between T1DM and AF.

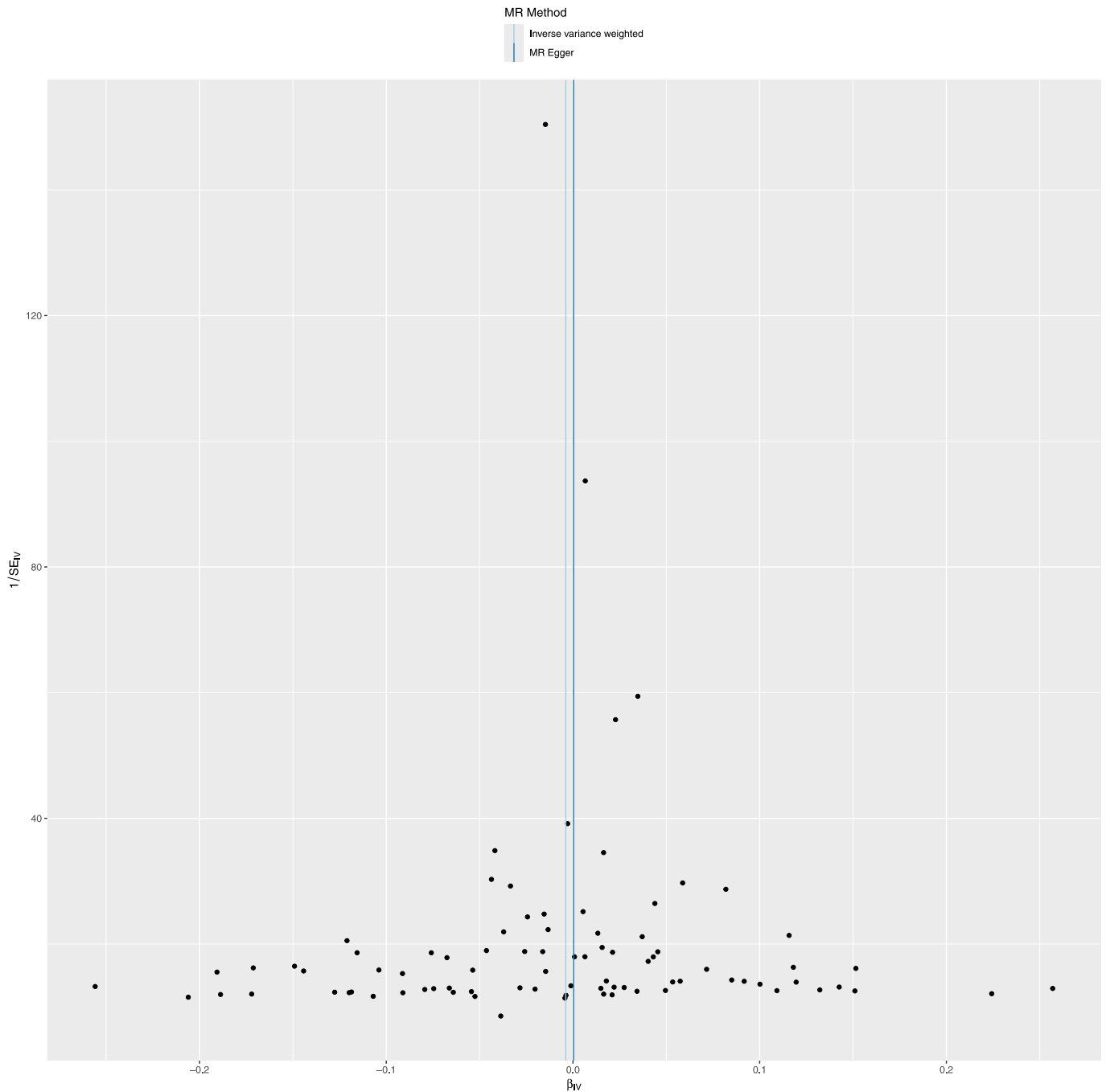


Fig. 4. Funnel plot of the two-sample mendelian randomization.

independent risk factors for AF. The increased risk of AF observed in patients with diabetes and prediabetes may be related to the high prevalence of other risk factors[21]. Conversely, a study in Korea reported results contradicting those of the Swedish study. They found that the hazard ratio (HR, 95 % CI) for AF occurrence in the T1DM group compared to the non-diabetes group was 1.748 (1.534–1.993). However, the article also acknowledged the limitations of the study, noting that the observed HRs ranging from 1.608 (for AF) to 2.105 (for heart failure hospitalization) in the T1DM group compared to the T2DM group could only be explained by an unmeasured confounding factor related to diabetes type, which poses a higher risk than or equal to the confounding factors measured in their study. This confounding factor may influence the final results, highlighting the study's limitations. Finally, the article cautioned about extrapolating results to different racial groups due to all

subjects being Korean, underscoring the study's inherent limitations in generalizability[7].

A meta-analysis has indicated that diabetes is associated with an increased likelihood of non-paroxysmal atrial fibrillation rather than paroxysmal atrial fibrillation[22]. This systematic review provides a comprehensive summary of the evidence linking diabetes with different types of atrial fibrillation. However, it does not establish causality between diabetes and atrial fibrillation types. It underscores the need for future studies to appropriately adjust for confounding factors, use Cox models with time-varying covariates to consider the onset of diabetes, and conduct high-quality research to examine causality, elucidate the exact mechanisms linking diabetes with non-paroxysmal atrial fibrillation, and assess the potential predictive value of diabetes in predicting non-paroxysmal atrial fibrillation. In a Mendelian randomization study

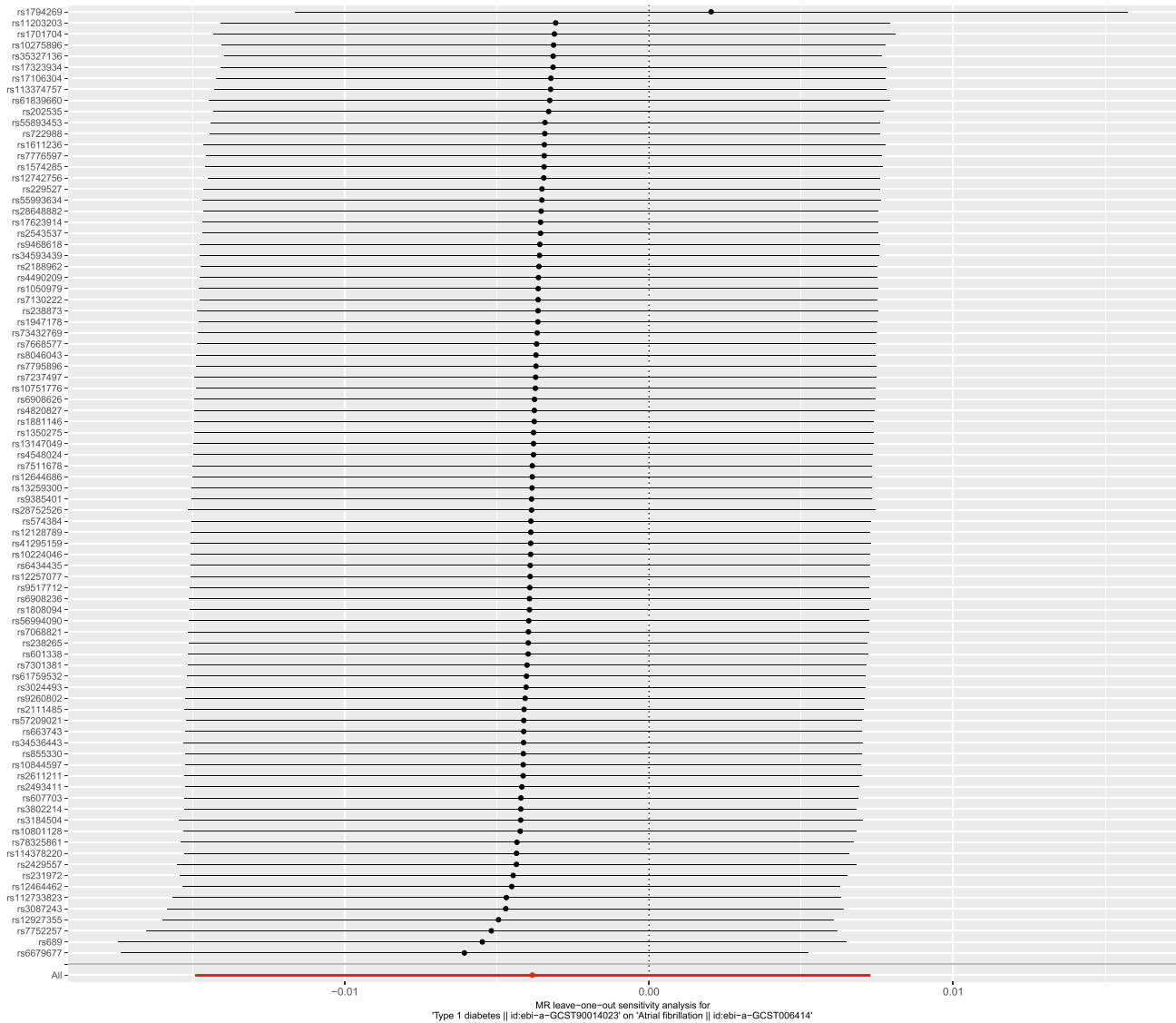


Fig. 5. The leave-one-out plots of the sensitivity analysis.

investigating the causal relationship between T1DM and cardiovascular diseases, it was noted that no significant causal relationship was found between T1DM and AF. The study only provided evidence supporting a causal relationship between T1DM and peripheral arterial disease and coronary artery disease[23].

Compared to traditional epidemiological studies, this study has several strengths: it utilizes two-sample Mendelian randomization to explore the causal relationship between T1DM and AF. This method leverages genotypes randomly allocated in nature as instrumental variables to infer the impact of exposures on outcomes, thereby largely mitigating issues related to reverse causation and confounding factors. We validated our hypothesis using the latest AF genome-wide association study data. Moreover, the study employed multiple Mendelian randomization analysis methods to ensure the reliability and stability of the results. However, this study also has its limitations: In the present dataset, the diagnostic information for atrial fibrillation (AF) was acquired from registry databases, the accuracy and reliability of which may be subject to multiple confounding factors. It relies on summary data from European populations, thus the conclusions may lack generalizability and should be complemented by data from diverse ethnic groups. Finally, the study solely explores the relationship between T1DM and AF from a genetic perspective, leaving the specific

mechanisms requiring further validation through additional research. Future studies should incorporate longitudinal data from prospective cohort studies to elucidate the mechanistic pathways through which disease progression trajectories influence the clinical course of atrial fibrillation. Although MR analysis revealed no direct causal association between T1DM and AF, preclinical evidence suggests that chronic hyperglycemia and oxidative stress—hallmarks of T1DM—may promote progressive atrial structural remodeling. Such remodeling likely initiates decades prior to AF symptom onset, a dynamic pathophysiological process inherently undetectable by MR methodologies, which rely on genetic variants as static proxies. Future investigations should integrate advanced imaging biomarkers (e.g., left atrial strain quantification) and mediator analyses (e.g., inflammatory cytokine trajectories) to delineate the temporal cascade linking T1DM-driven subclinical changes to AF manifestation.

## 5. Conclusion

In the current Mendelian randomization study, we did not find evidence of a causal relationship between genetically determined T1DM in European ancestry populations and atrial fibrillation. The question of whether T1DM leads to atrial fibrillation remains difficult to answer

definitively to date, given the challenges posed by cardiovascular complications, active management of risk factors, and the relatively lower incidence rate of T1DM. In clinical research, emphasis should be placed on understanding the impact of confounding factors on study outcomes, enhancing study rigor, and improving accuracy.

### CRediT authorship contribution statement

**Yongkai Li:** Writing – original draft, Software, Resources, Project administration, Methodology, Formal analysis, Data curation, Conceptualization. **Shasha Liu:** Data curation. **Yiming Dong:** Data curation. **Jianzhong Yang:** Writing – review & editing, Writing – original draft. **Yingping Tian:** Writing – review & editing, Writing – original draft.

### Funding

None.

### Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

### Acknowledgments

We would like to thank the researchers and study participants for their contributions.

### Ethics approval and consent to participate

All data were sourced from public databases and do not require additional ethical review.

### Appendix A. Supplementary material

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.ijcha.2025.101643>.

### Data availability

Data will be made available on request.

All data generated or analyzed during this study are included in this article and its [supplementary information](#) files.

### References

- [1] F.Z. Syed, Type 1 Diabetes Mellitus, *Ann. Intern. Med.* 175 (3) (2022) ITC33-ITC48, <https://doi.org/10.7326/aitc202203150>.
- [2] J.A. Joglar, M.K. Chung, A.L. Armbruster, et al., 2023 ACC/AHA/ACCP/HRS Guideline for the Diagnosis and Management of Atrial Fibrillation: A Report of the American College of Cardiology/American Heart Association Joint Committee on Clinical Practice Guidelines, *Circulation* 149 (1) (2024) e1–e156, <https://doi.org/10.1161/cir.0000000000001193>.
- [3] C.W. Tsao, A.W. Aday, Z.I. Almarazoo, et al., Heart Disease and Stroke Statistics—2023 Update: A Report From the American Heart Association, *Circulation* 147 (8) (2023) e93–e621, <https://doi.org/10.1161/cir.0000000000001123>.
- [4] A.M. Gillis, A Sober Reality? Alcohol, Abstinence, and Atrial Fibrillation, *N. Engl. J. Med.* 382 (1) (2020) 83–84, <https://doi.org/10.1056/NEJMe1914981>.
- [5] S. Dahlqvist, A. Rosengren, S. Gudbjörnsdóttir, et al., Risk of atrial fibrillation in people with type 1 diabetes compared with matched controls from the general population: a prospective case-control study, *Lancet Diabetes Endocrinol.* 5 (10) (2017) 799–807, [https://doi.org/10.1016/s2213-8587\(17\)30262-0](https://doi.org/10.1016/s2213-8587(17)30262-0).
- [6] S.C. Larsson, A. Wallin, N. Håkansson, et al., Type 1 and type 2 diabetes mellitus and incidence of seven cardiovascular diseases, *Int. J. Cardiol.* 262 (2018) 66–70, <https://doi.org/10.1016/j.ijcard.2018.03.099>.
- [7] Y.B. Lee, K. Han, B. Kim, et al., Risk of early mortality and cardiovascular disease in type 1 diabetes: a comparison with type 2 diabetes, a nationwide study, *Cardiovasc. Diabetol.* 18 (1) (2019) 157, <https://doi.org/10.1186/s12933-019-0953-7>.
- [8] C.J. O'donnell, M.S. Sabatine, Opportunities and Challenges in Mendelian Randomization Studies to Guide Trial Design, *JAMA Cardiol.* 3 (10) (2018) 967, <https://doi.org/10.1001/jamacardio.2018.2863>.
- [9] J. Chiou, R.J. Geusz, M.L. Okino, et al., Interpreting type 1 diabetes risk with genetics and single-cell epigenomics, *Nature* 594 (7863) (2021) 398–402, <https://doi.org/10.1038/s41586-021-03552-w>.
- [10] J.B. Nielsen, R.B. Thorolfsdóttir, L.G. Fritsche, et al., Biobank-driven genomic discovery yields new insight into atrial fibrillation biology, *Nat. Genet.* 50 (9) (2018) 1234–1239, <https://doi.org/10.1038/s41588-018-0171-3>.
- [11] N. Papadimitriou, N. Dimou, K.K. Tsilidis, et al., Physical activity and risks of breast and colorectal cancer: a Mendelian randomisation analysis, *Nat. Commun.* 11 (1) (2020) 597, <https://doi.org/10.1038/s41467-020-14389-8>.
- [12] E. Sanderson, W. Spiller, J. Bowden, Testing and correcting for weak and pleiotropic instruments in two-sample multivariable Mendelian randomization, *Stat. Med.* 40 (25) (2021) 5434–5452, <https://doi.org/10.1002/sim.9133>.
- [13] J. Bowden, G. Davey Smith, S. Burgess, Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression, *Int. J. Epidemiol.* 44 (2) (2015) 512–525, <https://doi.org/10.1093/ije/dyv080>.
- [14] S. Burgess, J. Bowden, T. Fall, et al., Sensitivity Analyses for Robust Causal Inference from Mendelian Randomization Analyses with Multiple Genetic Variants, *Epidemiology* 28 (1) (2017) 30–42, <https://doi.org/10.1097/ede.0000000000000559>.
- [15] M. Verbanck, C.Y. Chen, B. Neale, et al., Detection of widespread horizontal pleiotropy in causal relationships inferred from Mendelian randomization between complex traits and diseases, *Nat. Genet.* 50 (5) (2018) 693–698, <https://doi.org/10.1038/s41588-018-0099-7>.
- [16] M.J. Brion, K. Shakhbuzov, P.M. Visscher, Calculating statistical power in Mendelian randomization studies, *Int. J. Epidemiol.* 42 (5) (2013) 1497–1501, <https://doi.org/10.1093/ije/dyt179>.
- [17] S. Burgess, F. Dudbridge, S.G. Thompson, Combining information on multiple instrumental variables in Mendelian randomization: comparison of allele score and summarized data methods, *Stat. Med.* 35 (11) (2016) 1880–1906, <https://doi.org/10.1002/sim.6835>.
- [18] F. Rahman, G.F. Kwan, E.J. Benjamin, Global epidemiology of atrial fibrillation, *Nat. Rev. Cardiol.* 11 (11) (2014) 639–654, <https://doi.org/10.1038/nrcardio.2014.118>.
- [19] H. Harati, D. Zanetti, A. Rao, et al., No evidence of a causal association of type 2 diabetes and glucose metabolism with atrial fibrillation, *Diabetologia* 62 (5) (2019) 800–804, <https://doi.org/10.1007/s00125-019-4836-y>.
- [20] N.A. Chatterjee, F. Giulianini, B. Geelhoed, et al., Genetic Obesity and the Risk of Atrial Fibrillation: Causal Estimates from Mendelian Randomization, *Circulation* 135 (8) (2017) 741–754, <https://doi.org/10.1161/circulationaha.116.024921>.
- [21] C. Johansson, L. Örtengren, M.M. Lind, et al., Diabetes, prediabetes, and atrial fibrillation—A population-based cohort study based on national and regional registers, *J. Intern. Med.* 294 (5) (2023) 605–615, <https://doi.org/10.1111/joim.13688>.
- [22] F. Alijla, C. Buttia, T. Reichlin, et al., Association of diabetes with atrial fibrillation types: a systematic review and meta-analysis, *Cardiovasc. Diabetol.* 20 (1) (2021) 230, <https://doi.org/10.1186/s12933-021-01423-2>.
- [23] Z. Liu, H. Wang, Z. Yang, et al., Causal associations between type 1 diabetes mellitus and cardiovascular diseases: a Mendelian randomization study, *Cardiovasc. Diabetol.* 22 (1) (2023) 236, <https://doi.org/10.1186/s12933-023-01974-6>.