# **Research Article**

# Ten-Year Survival after Liver Resection for Colorectal Metastases: Systematic Review and Meta-Analysis

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Received 23 February 2011; Accepted 30 March 2011

Academic Editors: M. Mandala, G. Pecher, and K. Sonoda

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*Background*. Liver resection in metastatic colorectal cancer is proved to result in five-year survival of 25–40%. Several factors have been investigated to look for prognostic factors stratifications such as resection margins, node involvement in the primary disease, and interval between the primary disease and liver metastases. *Methods*. We searched MEDLINE and EMBASE for studies that reported ten-year survival. Metaanalysis was performed to analyse the effect of recognised prognostic factors on cure rate for colorectal metastases. The meta-analysis was performed according to Ottawa-Newcastle method of analysis for nonrandomised trials and according to the guidelines of the PRISMA. *Results*. Eleven studies were included in the analysis, which showed a ten-year survival rate of 12–36%. Factors that have favourable impact are clear resection margin, low level of CEA, single metastatic deposit, and node negative disease. The only factor that excluded patients from cure is the positive status of the resection margin. *Conclusion*. Predicted ten-year survival after liver resection for colorectal metastases varies from 12 to 36%. Only positive resection margins resulted in no 10-year survivors. No patient can be excluded from consideration for liver resection so long the result is negative margins.

# 1. Introduction

It is known that 60–70% of recurrent colorectal cancer involves the liver and that the liver is the only involved organ in 35% of cases. Colorectal liver metastases used to be thought of as systemic disease that involves many organs and systems until recently when local liver therapy in the form of liver resection has been reported to results in 5-year survival of 27–39% and even longer survival and cure [1, 2]. Patients with liver metastases that are resectable but left untreated have average survival of 6–12 months and rarely longer than 2 years [3].

Over the last 20 years liver resection for colorectal metastases has seen many refinements; the improvement in anaesthesia and postoperative care have reduced the morbidity and mortality with subsequent more aggressive surgical approach. Strategies that have widened the indications for liver resection includes portal vein embolisation, staged liver resection, neoadjuvant chemotherapy, ablative procedures, and locoregional chemotherapy [4–7]. Many studies have developed and validated scoring systems to predict prognosis and recurrence of colorectal metastases based on clinical and pathological data of large number of hepatic resections for colorectal metastases. These scores are based on variety of factors that include stage of the primary disease, time interval between diagnosis of the primary lesion and occurrence of liver metastases, level of CEA immediately prior to liver resection, size and number of resected liver lesions, surgical resection margin, blood transfusion, and bilateral distribution of liver disease [1, 8–10].

Various factors were found to have influence on disease recurrence and overall survival; resection margins and lymph nodes involvement are common predictors of recurrence. Other controversial factors are number and size of lesions, blood transfusion, and disease-free interval. Other factors are extrahepatic disease and portal nodes metastases [11–14].

This systematic review is conducted to evaluate the risk factors influence of overall long-term survival following hepatic resection for colorectal metastases.

# 2. Methods

2.1. Search and Study Identification. Electronic search was performed and relevant reports were identified using electronic databases (MEDLINE 1950–2010 and EMBASE 1980–2010), the search was restricted to human adults and English language literature. The search terms used were colorectal neoplasm, overall survival, and disease-free survival. All the references used in the published original and review studies were searched to identify more studies.

*2.2. Criteria for Study Selection.* To be included, studies had to meet the following criteria.

- (1) Design: prospective or retrospective cohort studies.
- (2) Population: patients with liver metastases from colorectal cancer who had liver resection as a curative treatment.
- (3) Exposure: surgical liver resection for metastases whether anatomic resection or segmental nonanatomic resection regardless of whether they had or did not have adjuvant chemotherapy. That included studies that evaluated patients' survival following reresection following primary liver resection.
- (4) Outcome: overall ten-year survival following liver resection for colorectal liver metastases.

Duplicate publications were excluded and wherever publications that evaluated the same population group were encountered, the report with the most relevant and comprehensive data was selected.

2.3. Data Extraction. Articles that met all the inclusion criteria were retrieved as full text articles. Two independent reviewers using standard data collection form extracted all relevant data from the full text articles. Inconsistencies were resolved by discussion to reach a reasonable consensus. Whenever missing data were encountered, the authors were contacted to request the data required to be included in the meta-analysis. One study in non-English language was encountered and was excluded.

2.4. Quality Assessment. Methodological quality of the studies was evaluated independently by two reviewers using the Newcastle-Ottawa Scale [15]. A quality score was calculated on the basis of the following components: selection of the study groups (0–4 points), quality of the adjustment for confounding variables (0–2), and outcome of interest in the study population (0–3 points). A higher score represents better methodological quality.

2.5. Statistical Analysis. Odds ratio (OR) of overall survival was used as the primary effect estimate in this meta-analysis. From the eligible studies that met the inclusion criteria, estimates of the OR and its associated 95% confidence interval (CI) were calculated using the Review Manager software (Version 5 for Windows, Copenhagen, Denmark; The Nordic Cochrane Centre, The Cochrane Collaboration,

2008). Data that could not be extracted directly were reconstructed indirectly by two reviewers when required.

Prespecified factors that was thought to affect the overall survival after liver resection for colorectal metastases were analysed. Those included resection margin, tumour size, number of metastases, bilateral versus unilateral disease, T stage of the primary, lymph nodes positive primary versus lymph nodes negative, disease-free interval, CEA level, and blood transfusion. Sensitivity analysis on the included studies was conducted on the Review Manager.

Heterogeneity between the included studies was appraised using the *Q* measure for statistical significance and the  $I^2$  measure for the amount of heterogeneity, with P < .1 being statistically significant and  $I^2 > 25\%$  showing important heterogeneity. A random effect model based on Der-Simonian-Laird estimator was used wherever there was significant heterogeneity, and fixed effect model based on Mantel-Haenszel estimator was used when there was no significant heterogeneity [16]. We conducted Begg's test and the Harbord modified test to identify publication bias for small study effect, with P > .5 being statistically significant [17]. The results of this systematic review were reported using the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines [18].

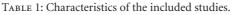
### 3. Results

3.1. Study Characteristics and Methodological Quality. The initial search revealed 164 titles, abstracts for those articles (Figure 1) were reviewed, 28 articles were considered to be potentially useful for inclusion, and their full text was retrieved and reviewed. Seventeen of these 28 articles were subsequently excluded from the meta-analysis as they did not meet the inclusion criteria. Eleven original reports (all retrospective cohort studies) had enough data to investigate the role of different variables on overall long-term survival after liver resection for colorectal metastases. Two studies were excluded due to the fact that they were a duplication of the same study population [19, 20]. One further study was excluded because it had no data that can be used in the metaanalysis [21]. The included studies were published between 1995 and 2009 (Table 1). The methodological quality of the included studies is shown in (Table 2). There was no statistical evidence of publication bias between the included studies.

There was no statistical evidence of publication bias among the included studies based on the funnel plot used in Review Manager.

3.2. Patients Characteristics. The eleven studies reviewed 3442 patients who had liver resection for colorectal metastases, the eight included studies had a total number of 2387 patients, all studies reported five- and ten-year survival [22–29]. Overall 5-year survival was 21–51% and overall ten-year survival was 12–36%. These studies had variably reported the impact of different factors on overall survival; these factors included resection margins, size of the largest liver lesion, number of liver lesions, distribution of lesions, CEA levels

Study Study design Total number of patients 5-year survival 10-year survival Giuliante et al. 2009 [22] Retrospective 251 38.9% 24.2% Hamady et al. 2006 [23] Retrospective 293 44% 36% Jamison et al. 1997 [24] Retrospective 280 27% 20% Minagawa et al. 2000 [25] Retrospective 235 38% 26% Scheele et al. 1995 [26] 469 39.3% Retrospective 23.6% Shimizu et al. 2007 [27] Retrospective 164 51.8% 36.6% Tomlinson et al. 2007 [28] 612 17% Retrospective 21% Wanebo et al. 1996 [29] Retrospective 74 24% 12%



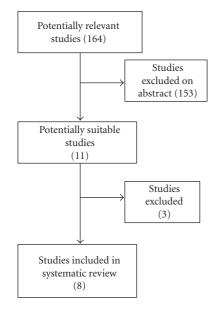


FIGURE 1: Flow diagram of study selection process.

 TABLE 2: Quality assessment of included studies (Newcastle-Ottawa Scale).

Study	Selection	Comparability	Outcome
Giuliante et al. 2009 [22]	4	1	3
Hamady et al. 2006 [23]	4	2	3
Jamison et al. 1997 [24]	4	2	3
Minagawa et al. 2000 [25]	4	1	3
Scheele et al. 1995 [26]	4	2	3
Shimizu et al. 2007 [27]	4	1	3
Tomlinson et al. 2007 [28]	4	2	3
Wanebo et al. 1996 [29]	4	2	3

prior to liver resection, lymph node status of the primary, satellite configuration of liver metastases, type of resection, extrahepatic disease, and whether the liver metastases were presented in a synchronous or metachronous to the diagnosis of the primary colorectal cancer.

*3.3. Disease-Free Survival.* In the study of Tomlinson et al. [28] they reported that 34% of patients who are disease-free at 5 years after hepatic resection did experience recurrence.

This figure is high compared with other reports by Minagawa et al., Giuliante et al., and Scheele et al. who reported 50and 10-year disease-free survival of 26 and 23%, 28.2 and 25.4%, and 33.6 and 25.4%, respectively [22, 25, 26]. From these data we conclude that 5-year survival does not equate cure. Since recurrence after 10-year of disease-free survival is rare; ten-year survival could be considered as definitive cure.

Factors that influence disease-free survival and may predict that has been reported only in the study by Minagawa et al. [25], hence it was not suitable to conduct meta-analysis on this category of outcome.

3.4. Influence of Resection Margins. Four studies reported adequate data to determine the relationship between positive and negative resection margin and long-term survival. The average overall survival for positive margin was 29%, which is significantly better compared with 20% of negative resection margin (P = .03) with odds ratio of 0.41 (95% CI 0.18–0.9). There was moderate degree of heterogeneity  $I^2$ 52%; however, this heterogeneity was eliminated when prespecified sensitivity analysis was performed by elimination of the study by Tomlinson et al. [28] that had a wide 95% CI and that has little overlap with other studies. The result of the meta-analysis after this elimination showed OR of 0.55 (0.95% CI 0.36–0.84) with P = .005, which ensures that the results of the meta-analysis are robust (Figure 2).

Four studies reported the influence of wider negative margin of more than 1cm compared with negative margin of 0–10 mm [23–26]. Wider resection margin had no beneficial effect on overall survival. Pooled analysis for the likelihood of survival is shown in Figure 3 (OR = 1.11; 95% CI: 0.59–2.08; P = .75). Moderate heterogeneity was seen. The pooled estimate was robust: omission of individual study at a time did not change the statistical results (data not shown).

Six studies reported the effect of tumour size on overall survival [22–26, 28]; meta-analysis of the studies that reported analysis data on patients liver metastases more than 5 cm compared with patients who had liver lesions of 5 cm or less, showed no prognostic relationship between size of the resected tumour and overall patients' survival. Figure 4 shows the results of pooled estimate for survival in relation to tumour size (OR = 0.73, 95% CI: 0.46–1.16; P = .18). There was moderate heterogeneity among the included studies  $I^2 = 75\%$ , the results were robust, omission of the study by

	Positive resect	on margin	Negative resection	on margin		Odds ratio	Odds	ratio	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% CI	M-H, rand	dom, 95% CI	
Giuliante 2009	2	16	95	235	18.1%	0.21 [0.05, 0.95]		-	
Hamady 2006	34	96	94	197	43.4%	0.6 [0.36, 0.99]		-	
Jamison 1997	7	41	40	154	31.6%	0.59 [0.24, 1.43]		-	
Tomlinson 2007	0	59	102	553	7%	0.04 [0, 0.6]	•		
Total (95% CI)		212		1139	100%	0.41 [0.18, 0.9]	•		
Total events	43		331						
Heterogeneity: $\tau^2$ =	$= 0.31; \chi^2 = 6.2$	9, df = 3 (P)	$= .10); I^2 = 52\%$			<u>├</u>			
Test for overall effe	ct: $Z = 2.23$ (P	= .03)				0.01	0.1	1 10	100
						Favours ex	perimental	Favours con	trol

	Less that	n 1 cm	More tha	ın 1 cm		Odds ratio		(	Odds rati	0	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% (	CI	M-H	random	n, 95% CI	
Hamady 2006	72	138	23	59	29.1%	1.71 [0.92, 3.18]			₽		
Jamison 1997	30	123	10	31	23.1%	0.68 [0.29, 1.6]		-			
Minagawa 2000	31	118	3	27	15.3%	2.85 [0.8, 10.13]			+	-	
Scheele 1995	43	204	41	146	32.5%	0.68 [0.42, 1.12]					
Total (95% CI)		583		263	100%	1.11 [0.59, 2.08]			•		
Total events	176		77								
Heterogeneity: $\tau^2 = 0.2$ Test for overall effect: Z			(P = .04); I	$^{2} = 65\%$	Ď	(	0.01	0.1	1	10	100
							Favou	's experin	nental	Favours con	ntrol

FIGURE 2: Positive and negative margins.

FIGURE 3: Resection margin less than 1 cm or more than 1 cm.

Giuliante et al. [22] resulted in very close odds ratio and removed the heterogeneity.

Four studies reported data on the effect of time interval between the diagnoses of the primary colorectal cancer and the occurrence of liver metastases on the overall patients' survival. Pooled estimate of the survival time after liver resection for colorectal metastases showed no significant prognostic relationship (OR = 1.22; 95% CI: 0.75–1.99; P = .42) Figure 5. There was moderate degree of heterogeneity ( $I^2 = 68\%$ ). The analysis results were robust, exclusion of the study by Scheele et al. [26] removed the heterogeneity and the results and resulted in OR = 1.03; 95% CI: 0.64–1.65, data not shown.

Six studies reported data to determine the relationship between lymph nodes metastasis status of the primary colorectal cancer and survival after liver resection for colorectal metastases [22, 23, 25, 26, 28, 29]. Figure 6 shows the forest plot with pooled estimates of the odds ratio of survival in patients who had node positive primary disease compared with those who had node negative disease. The overall tenyear survival for node negative disease was 32% compared with 22% for nodes positive primary disease (OR = 0.46; 95% CI: 0.26–0.79; P = .006), the pooled analysis showed significant heterogeneity  $I^2 = 81\%$ ; omission of the study by Tomlinson et al. [28] reduced the heterogeneity to insignificant level with results of odds ratio of 0.38; 95% CI: 0.23–6; P < .0001, which ensures the robustness of the pooled estimates of the effect.

Three studies reported data for determination of the relationship between types of liver resection (segmental or anatomic) [22, 25, 29]. Figure 7 shows the Forest plot of the pooled estimates of long-term survival for patients who had segmental resection compared with patients who had anatomic resection (right hepatectomy, left hepatectomy, or trisectionectomy). The type of resection did not have any significant impact on overall survival (OR = 2.60; 95% CI: 0.88–7.63; P = .08), there was significant heterogeneity among the included studies. Omission of individual studies did not change the results of the analysis.

3.5. CEA Levels. Four studies provided data for determining the relationship of CEA levels, prior to resection of colorectal liver metastases, and the overall patients' survival [22, 23, 25, 26]. Figure 8 shows the Forest plot with the pooled estimate for likelihood of ten-year survival for patients with CEA levels more than 50 ng/mL and those with CEA levels less than 50 ng/mL. Overall pooled survival for those with CEA level less than 50 ng/mL is 37% compared with 19% for those with CEA greater than 50 ng/mL. The results show statistically significant better survival in those with low CEA levels (OR = 2.27; 95% CI: 1.03–5.02; P = .04). There is

	More tha	n 5 cm	Five centime	eter or less		Odds ratio	Odds	ratio	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% CI	M-H, rand	dom, 95% CI	
Giuliante 2009	10	60	88	191	14.5%	0.23 [0.11, 0.49]			
Hamady 2006	39	94	88	189	18%	0.81 [0.49, 1.34]			
Jamison 1997	28	132	26	140	16.5%	1.18 [0.65, 2.14]		_	
Minagawa 2000	48	167	14	68	15.3%	1.56 [0.79, 3.06]	+	-	
Scheele 1995	18	117	67	233	16.8%	0.45 [0.25, 0.8]			
Tomlinson 2007	36	236	66	376	18.8%	0.85 [0.54, 1.32]			
Total (95% CI)		806		1197	100%	0.73 [0.46, 1.16]	•		
Total events	179		349						
Heterogeneity: $\tau^2 = 0$	$0.25; \chi^2 = 19$	9.75, df =	5 (P = .001);	$I^2 = 75\%$		<u> </u>		+	
Test for overall effect:	$Z = 1.34 \ (H$	P = .18)				0.01	0.1 1	10	100
	L = 1.54 (110)				Favor	urs experimental	Favours co	ntrol	

#### FIGURE 4: Tumour size.

	12 months or les		More than 12 months		s or less More than 12 months			Odds ratio	Odds ratio
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% CI	M-H, random, 95% CI		
Giuliante 2009	28	52	36	95	21.5%	1.91 [0.96, 3.79]	-8		
Hamady 2006	85	197	44	89	26.8%	0.78 [0.47, 1.28]	-8-		
Scheele 1995	27	77	27	131	23%	2.08 [1.11, 3.91]			
Tomlinson 2007	37	238	65	374	28.7%	0.88 [0.56, 1.36]	+		
Total (95% CI)		564		689	100%	1.22 [0.75, 1.99]			
Total events	177		172				•		

Heterogeneity:  $\tau^2 = 0.17$ ;  $\chi^2 = 9.29$ , df = 3 (P = .03);  $I^2 = 68\%$ Test for overall effect: Z = 0.81 (P = .42)

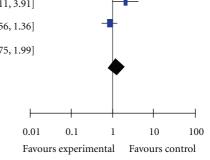


FIGURE 5: Disease-free interv	al.
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	Node po	ositive	Node r	egative		Odds ratio		Odds ration	0	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% CI	M-	H, random	ı, 95% C	CI
Giuliante 2009	55	154	34	74	17.6%	0.65 [0.37, 1.15]	-	┡┿		
Hamady 2006	63	165	60	103	18.3%	0.44 [0.27, 0.73]		-		
Minagawa 2000	34	163	22	55	16.5%	0.4 [0.2, 0.76]	-	-		
Scheele 1995	21	216	45	116	17.4%	0.17 [0.09, 0.3]				
Tomlinson 2007	51	301	51	311	19.1%	1.04 [0.68, 1.59]		+		
Wanebo 1996	12	57	7	17	11.1%	0.38 [0.12, 1.21]		+		
Total (95% CI)		1056		676	100%	0.46 [0.26, 0.79]	•	•		
Total events	236		219							
Heterogeneity: $\tau^2 = 0.3$	37; $\chi^2 = 26$ .	22, df = 5	( <i>P</i> < .0001)	); $I^2 = 81$	%	H				
Test for overall effect: Z	Z = 2.77 (P)	= .006)				0.01	0.1	1	10	100
						Fav	tal Favor	irs cont	rol	

FIGURE 6: Nodal disease of the primary cancer.

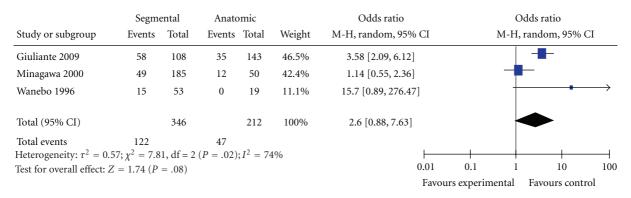


FIGURE 7: Type of resection.

	Less than 5	50 ng/mL	More than	50 ng/mL		Odds ratio	Odds ratio	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% CI	M-H, random, 95% CI	
Giuliante 2009	74	170	11	43	29.2%	2.24 [1.06, 4.74]		
Hamady 2006	81	172	33	79	33.4%	1.24 [0.72, 2.13]		
Minagawa 2000	42	130	15	83	30.8%	2.16 [1.11, 4.22]		
Scheele 1995	60	205	0	74	6.7%	61.96 [3.78, 1015.99]		
Total (95% CI)		677		279	100%	2.27 [1.03, 5.02]	•	
Total events	257		59					
Heterogeneity: $\tau^2 = 0$	$0.41; \chi^2 = 10.3$	88, df = 3 (.	$P = .01); I^2 =$	72%			+ + +	
Test for overall effect:						0.01	0.1 1 10	100
	,	<i>,</i>				Favours e	experimental Favours o	ontrol

FIGURE 8: CEA level.

significant heterogeneity ( $I^2 = 72\%$ ); omission of the study by Scheele et al. [26], which looks like an outlier has removed the heterogeneity, and still the results were significant (OR = 1.71; 95% CI: 1.15–2.55; P = .009; data not shown), which confirms the robust results of the analysis.

3.6. Distribution of Liver Lesions. Five studies provided data to determine the effect of bilateral distribution of resected colorectal metastases on the ten-year survival [22, 25, 26, 28, 29]. Ten-year survival was 36% compared with 18% in patients who had bilateral disease resected. Figure 9 shows the Forest plot of the pooled estimate for ten-year survival, the presence of bilateral disease in the liver leads to significantly reduced long-term survival (OR = 1.64; 95% CI: 1.19–2.27; P = .003. There was no significant heterogeneity among the included studies  $I^2 = 0\%$ .

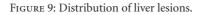
3.7. Number of Liver Lesions. Seven studies reported data to determine the effect of the number of the resected liver lesions on long-term survival [22–26, 28, 29]; these studies compared survival in patients with four or less lesions versus patients who had more than four lesions and found that ten-year survival for patients with four or less lesion was 38% compared with 20% of those who had more than four lesions (OR = 1.75; 95% CI: 0.87–3.51; P = .11); there was a significant heterogeneity  $I^2 = 73\%$ ; omission of individual

studies made no statistical difference; however, omission of three studies, Scheele et al., Minagawa et al., and Tomlinson et al. [25, 26, 28] removes the heterogeneity and results in significantly better survival in patients who had four or less lesions (OR = 2.26; 95% CI: 1.36–3.75; P = .002), data not shown. This made it not possible to make any valid conclusion about the effect of lesion's number on overall survival (Figure 10).

3.8. Synchronous versus Metachronous Metastases. Six studies provided data to determine the relationship of timing of liver metastases; whether it was found at the time of diagnosis of the primary colorectal cancer or afterward [22, 23, 25–28]. Figure 11 shows the Forest plot with pooled estimate of likelihood of survival for patients who had liver metastases diagnosed synchronous with the primary colorectal cancer compared with those who developed liver metastases afterward (metachronous; OR = 0.77; 95% CI: 0.59–1.01; P = .06). The timing of liver metastases has no significant effect on long-term survival.

3.9. Effect of Blood Transfusion. Four studies provided data for determining the relationship of blood transfusion, after liver resection, and long-term survival [22–24, 29]. Figure 12 shows the Forest plot with the pooled estimate for long-term survival. Patients who received two or less units had

Study or subgroup	Unila Events	teral Total	Bila Events	teral Total	Weight	Odds ratio M-H, random, 95% (	CI		s ratio om, 95% CI	
Giuliante 2009	84	192	15	59	24.7%	2.28 [1.19, 4.38]				
Minagawa 2000	8	40	18	85	12.1%	0.93 [0.37, 2.37]				
Scheele 1995	32	90	14	57	19.1%	1.69 [0.81, 3.56]		-		
Tomlinson 2007	66	333	25	181	41.9%	1.54 [0.93, 2.55]				
Wanebo 1996	17	64	1	7	2.2%	2.17 [0.24, 19.36]				-
Total (95% CI)		719		389	100%	1.64 [1.19, 2.27]			•	
Total events	207		73							
Heterogeneity: $\tau^2 = 0$ ;	$\chi^2 = 2.53, c$	df = 4 (P)	$= .64); I^2$	= 0%			<b> </b>			
Test for overall effect: 2							0.01	0.1	1 10	100
							Favou	rs experimental	Favours co	ntrol



	Four or le	ess More t	han four		Odds ratio	Odds ratio
Study or subgroup	Events Te	otal Events	Total	Weight	M-H, random, 95% CI	M-H, random, 95% CI
Giuliante 2009	92 2	218 5	33	15.1%	4.09 [1.52, 10.99]	
Hamady 2006	110 2	234 19	53	18.6%	1.59 [0.86, 2.94]	+
Jamison 1997	62 2	257 2	23	11%	3.34 [0.76, 14.64]	+
Minagawa 2000	46 1	182 15	53	18%	0.86 [0.43, 1.7]	
Scheele 1995	70 3	318 12	32	17.3%	0.47 [0.22, 1.01]	
Tomlinson 2007	95 4	450 5	84	15.7%	4.23 [1.67, 10.74]	
Wanebo 1996	18	68 0	5	4.4%	4.03 [0.21, 76.5]	
Total (95% CI)	1	727	283	100%	1.75 [0.87, 3.51]	•
Total events	493	58				
Heterogeneity: $\tau^2 = 0$	.58; $\chi^2 = 22.3$	37, df = 6 (P = .00)	$(1); I^2 =$	73%	H	
Test for overall effect:	Z = 1.58 (P =	= .11)			0.01	0.1 1 10 100
					Fa	vours experimental Favours control



significantly better survival than patients who had more than two units (OR = 3.69; CI: 1.79–7.60; P = .0004). There was significant heterogeneity among the included studies  $I^2 =$ 63%. Omission of the study of Giuliante et al. [22] removed the heterogeneity, and the pooled estimate remained valid (OR = 2.51; CI: 1.63–3.85; P < .0001), data not shown.

Three studies provided data that compared the difference in survival between patients who had a single lesion resected and those who had more than a single lesion [22, 25, 29]. Figure 13 shows the Forest plot of the pooled estimates of survival for patients with single lesion compared to patients who had multiple lesions. Patients with a single lesions had significantly better prognosis than those with multiple lesions (OR = 1.56; 95% CI: 1.08–2.25; P = .02).

Three studies provided data that compared survival of patients who had satellite lesions along with larger lesion or lesions compared with those who had no satellite lesions [23, 24, 26]. Figure 14 shows the Forest plot of the pooled estimate of survival, the estimate shows significantly better survival for patients who had no satellite lesions compared with those who had satellite lesions (OR = 0.37; CI: 0.18–0.77; P = .008), there was significant heterogeneity among the included studies.

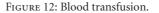
#### 4. Discussion

This systematic review and meta-analysis showed that factors that affect long-term survival following hepatic resection, for colorectal cancer metastases, include clear resection margins, advanced primary colorectal cancer with nodal metastases, CEA levels, distribution of liver lesions, timing of diagnosis of liver metastases (synchronous or metachronous), quantity of blood transfusion, single lesion compared with multiple lesions, and presence or absence of satellite nodules close to the main lesion. Patients who had clear resection margins had significantly better long-term survival than those with positive resection margins. Patients who had early stage colorectal cancer with no lymph nodes metastases had better survival than those with lymph nodes metastases. It also showed that metachronous presentation of liver metastases is a good prognostic factor compared with synchronous presentation. A single liver lesion particularly in the absence of satellite nodules has better outcome than multiple lesions with or without satellite nodules; however, when comparison was made between patients with more than four nodules and those with four or less lesions, there was no difference in overall long-term survival. CEA levels less than 50 ng/mL,

Study or subgroup	Synchr Events	onous Total	Metach: Events		Weight	Odds ratio M-H, random, 95%	CI	C M-H, ra	dds rati andom,		
Giuliante 2009	29	104	65	147	18.6%	0.49 [0.28, 0.84]		-			
Hamady 2006	67	146	61	147	23%	1.2 [0.75, 1.9]				-	
Minagawa 2000	23	106	36	129	15.7%	0.72 [0.39, 1.31]					
Scheele 1995	27	142	52	208	19.3%	0.7 [0.42, 1.19]					
Shimizu 2007	24	70	35	94	14%	0.88 [0.46, 1.68]					
Tomlinson 2007	7	54	95	558	9.4%	0.73 [0.32, 1.65]					
Total (95% CI)		622		1283	100%	0.77 [0.59, 1.01]					
Total events	177		344								
Heterogeneity: $\tau^2 = 0$	$0.03; \chi^2 = 0$	5.59, df =	= 5 (P = .2)	$5); I^2 = 2$	4%						
Test for overall effect:							0.01	0.1	1	10	100
······································					Favor	urs experin	nental	Favours con	ntrol		

FIGURE 11: Synchronous versus metachronous metastases.

	0–2 u	inits	More that	n two units		Odds ratio	Odds ra	ntio	
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% CI	M-H, random	n, 95% CI	
Wanebo 1996	5	16	2	15	11.5%	2.95 [0.48, 18.34]			
Jamison 1997	32	84	41	188	34%	2.21 [1.26, 3.86]			
Hamady 2006	118	246	11	47	30%	3.02 [1.47, 6.2]			
Giuliante 2009	97	193	5	58	24.4%	10.71 [4.1, 27.95]			
Total (95% CI)		539		308	100%	3.69 [1.79, 7.6]		•	
Total events	252		59						
Heterogeneity: $\tau^2 = 0$	.32; $\chi^2 = 8$ .	.11, df =	= 3 (P = .04)	$;I^2 = 63\%$					
Test for overall effect:	Z = 3.53 (1	P = .000	04)			0.01	0.1 1	10	100
						Favour	s experimental	Favours con	trol



unilateral liver disease, and two units or less of perioperative blood transfusion were found to be favourable prognostic factors. There was no single prognostic factor of sufficient power to predict long-term survival and cure.

Other factors that were analysed in this meta-analysis and found not to have significant influence on long-term survival included a the width of resection margin, it was found that if the resection margin is clear there is no survival benefit from a wider resection margin more than 1 cm. Whether the resected largest lesion was more or less than 5 cm did not have effect on survival. The interval between the diagnosis of the primary and liver recurrence less or more than 12 months did not seem to affect long term survival. Whether the lesion removed in anatomic resection technique or segmental resection, number of resected lesions, and synchronous versus metachronous metastases, all had no effect on longterm survival.

During the past two decades, liver resection for colorectal liver metastases has been increasing the standard of care whenever the disease is limited to the liver and is technically possible by leaving adequate liver remnants. There is overwhelming evidence to support the survival benefit with reports of actuarial 5-year survival of 25–40% compared with patients who are treated only with chemotherapy who rarely survive up to five years [30, 31]. With technical advances and improved perioperative the mortality of liver resection is less than 5%, Wei et al. reported 1.7% mortality in a large series with morbidity of 19% [21]. This improvement makes liver resection a standard treatment for colorectal metastases [12, 20, 26, 28].

Tomlinson et al. had investigated risk factors for 10-year survival and redesigned the original score devised by the same investigators. The original score had five components with one point for each component that includes tumour number more than four, size more than 5 cm, CEA level more than 200 ng/mL, and disease-free interval less than 12 months and positive resection margin with one point for each. After analysing 10-year survival, patients fell in two groups; the low risk group (0-2 points) who had 10-year survival of 21% and high-risk group (3-5 points) with 10year survival of 10% (P < .0001) [1, 28]. No patients with positive resection margin had survived for 10 years, which seems to be the only factor that ruled out any possibility of cure [28]. This raises the question of the benefit of neoadjuvant chemotherapy particularly now there are more effective agents that may increase the rate of complete resection, the effect of that approach on survival is not known.

	Single lesions		Multiple lesions		Odds ratio	Odds ratio
Study or subgroup	Events Total		Events Total	Weight	M-H, random, 95% CI	M-H, random, 95% CI
Giuliante 2009	57	137	40 114	51.3%	1.32 [0.79, 2.2]	-
Minagawa 2000	35	110	26 125	38.8%	1.78 [0.99, 3.2]	
Wanebo 1996	7	23	8 49	9.9%	2.24 [0.7, 7.21]	
Total (95% CI)		270	288	100%	1.56 [1.08, 2.25]	•
Total events	99		74			
Heterogeneity: $\tau^2 =$	$0; \chi^2 = 0.9$	7, df = 2	$(P = .61); I^2 = 0\%$		<b>⊢</b>	
Test for overall effec	t: $Z = 2.37$	(P = .02)			0.01	0.1 1 10 100
					Favo	ours experimental Favours control

FIGURE 13: Single versus multiple lesions.

	Present		Abscent			Odds ratio		Odds ratio			
Study or subgroup	Events	Total	Events	Total	Weight	M-H, random, 95% (	CI M-H, rand			lom, 95% CI	
Hamady 2006	25	72	102	213	49%	0.58 [0.33, 1.01]		-			
Jamison 1997	4	32	54	182	26.9%	0.34 [0.11, 1.01]					
Scheele 1995	3	49	84	301	24.1%	0.17 [0.05, 0.56]		-	-		
Total (95% CI)		153		696	100%	0.37 [0.18, 0.77]					
Total events	32		240								
Heterogeneity: $\tau^2 = 0$	$\lambda_{2}; \chi^{2} = 3.8$	, df = 2 (	P = .15); I	$1^2 = 47\%$	6						
Test for overall effect: $Z = 2.66 (P = .008)$					0.	.01	0.1	1	10	100	
							Favours experimental			Favours control	

FIGURE 14: Satellite lesions.

It was not possible in this review to analyse the effect of adjuvant chemotherapy on survival after hepatic resection for colorectal metastases. However, most of the study patients had been treated in the era when the standard chemotherapy was 5-flourouracil, which has limited effect, compared with the modern chemotherapeutic agents like irinotecan, oxaliplatin, and bevacizumab [28]. This makes the evaluation of the pure curative benefit from surgical resection relatively difficult to assess.

The improvement of outcome in liver resection has been attributed to a combination of factors such as aggressive surgical treatment, improved chemotherapy, and improvement of preoperative imaging and patient selection [32–34].

Repeat hepatectomy for recurrent hepatic disease following initial liver resection is being increasingly used, also staged liver resection and resection of isolated extrahepatic disease is being more and more utilised with encouraging results that lead to improvement in survival [4, 34–37]. Despite all the advances and improvements, it remains not possible to discriminate with reasonable certainty which subset of patients is likely to be cured or live more than 5 years. Repeat liver resection gives overall survival results of up to 52% in one series and is comparable to the results of initial liver resection [37].

Improvement of preoperative staging by liberal use of helical CT scans and MRI and the introduction of PET scan and PET CT have allowed for better patients selection by early detection of extrahepatic and bilobar disease [33, 38–40].

Many authors have reported using radiofrequency ablation in conjunction with surgery for nonresectable liver metastases either intraoperative or postoperative with variable results [21, 32], the impact of this technique on overall longterm survival remains questionable.

The emergence of new chemotherapeutic agents such as oxaliplatin, irinotecan, and bevacizumab has increased the treatment option available for clinicians to deal with metastatic colorectal disease. These new agents have been used in conjunction with liver resection either as neoadjuvant or adjuvant manner in many studies but the effect of this approach on survival has not been test in a randomised controlled trial [5, 41, 42]. Chemotherapy is being commonly used as an adjuvant agents following liver resection particularly in patients who had no previous chemotherapy, the survival benefit of this approach remains to be proved [21]. Another potential use is to downstage potentially resectable liver metastases, in this situation patients with advanced disease are offered resection if they show good response to chemotherapy, this strategy has been used with limited success [5, 34].

In conclusion this review defines 10-year survival and cure to be between 12% and 28%; we described the factors that affect survival in this meta-analysis. There is no single factor that was of sufficient power to rule out cure with the possible exception of positive resection margin. This leads to the fact that patients' selection for resection with disease limited to the liver or liver disease with resectable extrahepatic metastases remains a matter of trial and error particularly for patients who have marginal suitability for resection. This review also indicates that we need newer prognostic factors perhaps based on tumour biology that may discriminate between curable and noncurable metastatic colorectal cancer. An important limitation of this study that reflects the quality of the available data is the fact that raw data was not available to all the studies' patients and the presence of heterogeneity among the included studies and the fact that those studies are generally retrospective reviews. Also patients who had 10-year survival are likely to have been treated prior to the era of PET scan routine use and likely have been treated with old and less effective chemotherapeutic agents.

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