

BMJ Open Electroacupuncture as an adjunctive therapy for motor dysfunction in acute stroke survivors: a systematic review and meta-analyses

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ABSTRACT

Objectives To assess the effectiveness and safety of electroacupuncture (EA) combined with rehabilitation therapy (RT) and/or conventional drugs (CD) for improving poststroke motor dysfunction (PSMD).

Design Systematic review and meta-analysis.

Methods The China National Knowledge Infrastructure, Chinese Biological Medicine Database, Chinese Scientific Journal Database, Cochrane Library, Medline and Embase were electronically searched from inception to December 2016. The methodological quality of the included trials was assessed using the Cochrane risk of bias assessment tool. Statistical analyses were performed by RevMan V.5.3 and Stata SE V.11.0.

Results Nineteen trials with 1434 participants were included for qualitative synthesis and meta-analysis. The methodological quality of the included trials was generally poor. The meta-analysis indicated that the EA group might be benefiting more than the non-EA group in terms of the changes in the Fugl-Meyer Assessment Scale (FMA) (weighted mean difference (WMD): 10.79, 95% CI 6.39 to 15.20, $P < 0.001$), FMA for lower extremity (WMD: 5.16, 95% CI 3.78 to 6.54, $P < 0.001$) and activities of daily living (standardised mean difference: 1.37, 95% CI 0.79 to 1.96, $P < 0.001$). However, there was no difference between EA and non-EA groups in terms of the effective rate (relative risk: 1.13, 95% CI 1.00 to 1.27, $P = 0.050$). Moreover, there were not any reports of side effects due to EA combined with RT and/or CD in the included trials.

Conclusions This review provides new evidence for the effectiveness and safety of EA combined with RT and/or CD for PSMD. However, the results should be interpreted cautiously because of methodological weakness and publication bias. Further clinical trials with a rigorous design and large sample sizes are warranted.

PROSPERO registration number CRD42016037597.

INTRODUCTION

Stroke is part of the world's leading causes of death and disability,^{1,2} causing heavy burdens to patients' families, communities and health-care systems.³ Motor dysfunction is a frequent and widely recognised complication that often follows stroke. Approximately 85% of patients

Strengths and limitations of this study

- This systematic review with a comprehensive search of three English and three Chinese databases up to December 2016 was focused especially on assessing the adjunctive effects of electroacupuncture (EA) for motor dysfunction in acute stroke survivors within 14 days.
- Although the included trials in this review have methodological weakness, meta-regression analyses to explain the potential influence of the duration of treatment were performed, and sensitivity analyses with different risk of bias showed that the results were robust.
- Built on the low quality of included trials, we anticipate considerable difficulties in identifying the effectiveness of EA for motor dysfunction in acute stroke.

with stroke suffer from haemiparesis immediately after their stroke, and between 55% and 75% of stroke survivors may experience incomplete recovery with lingering motor dysfunction.⁴ Poststroke motor dysfunction (PSMD), which has a negative impact on the independence in functional activities, can reduce quality of life and limit activities of daily living (ADL). Therefore, effective treatment of PSMD is necessary to promote neurological function recovery and to alleviate the social and familial burdens of stroke.

Motor function recovery after stroke requires multidisciplinary treatment team and involves various approaches, such as conventional drugs (CD), rehabilitation therapy (RT) and nursing care. RT plays an important role in comprehensive stroke rehabilitation programme aimed at recovering the function so as to reduce disabilities. Previous studies have demonstrated that neurological deficits due to stroke can benefit from RT.⁵ However, the effects of current RT for motor dysfunction caused by stroke are not optimal.⁶ Over

the last decade, an increasing number of researchers have focused on alternative therapies for stroke rehabilitation, such as acupuncture.

In recent years, acupuncture, as one of the best known complementary medicines, has also been increasingly applied in China and other regions of the world.⁷ Electroacupuncture (EA), derived from the integration of traditional acupuncture and modern electrical stimulation, is another kind of acupuncture. EA is generally accepted because it is a relatively straightforward, safe and cheap therapy, compared with other conventional therapies.⁸ Additionally, EA has become more and more widely used in clinical practice because of its repeatability and standardisation of frequency, intensity and duration.^{9,10} After the needles are inserted into the acupuncture points, the electrodes are attached to the pairs of needles, and then a small electric current, usually with a pulse frequency of 1–100 Hz and pulse amplitude of 2–3 mA, is passed through the needles into the subject for 15–60 min.^{11,12} The efficacy of EA with strongly continued stimulation to treat peripheral facial paralysis is preferable to manual acupuncture.¹² Furthermore, EA has been commonly recommended for use in clinic and research on acupuncture in China and in other countries.^{10,13} EA may improve functional recovery after stroke by inhibiting cell apoptosis, regulating the miR-9-mediated nuclear factor kappa-light-chain-enhancer of activated B cells downstream pathway and the miR-181b/paired immunoglobulin-like receptor B/ras homolog family member A/growth-associated protein 43 axis, and dynamically maintaining the balance of matrix metalloproteinase-9 and tissue inhibitor of metalloproteinase-1.^{14–16}

A systematic review had suggested that EA was helpful in the treatment of acute ischaemic stroke, but its search time was up to June 2013 and was not focused on PSMD.¹⁷ Another newly published systematic review had revealed that EA had the potential to reduce spasticity within 180 days post stroke.¹⁸ However, these systematic reviews did not focus specifically on the effects of EA as an adjunctive therapy for motor dysfunction of acute stroke (within 14 days of onset¹⁹). Up to now, no clear evidence has been found that EA is more effective than non-EA in improving motor dysfunction within 14 days after stroke. Therefore, this study was conducted to assess the effectiveness and safety of EA combined with RT for PSMD in acute period, and to provide the best available evidence of clinical practice. The systematic review was registered on PROSPERO (<https://www.crd.york.ac.uk/PROSPERO/>) (trial registration number: CRD42016037597).

METHODS

Types of studies

We included all randomised controlled trials (RCTs) evaluating the effectiveness and safety of EA combined with RT for PSMD. The comparators or controls in the trials were any other therapy modalities. Also, we only included trials with outcomes measuring changes in

motor function. All eligible trials published in Chinese or English were included, regardless of publication status.

Types of participants

We considered trials that included patients in the acute stage being the onset of their first stroke, with motor dysfunction measured by validated instruments or by a decrease in the level of movement activity. Patients were to be more than 18 years old and from any ethnicity. Stroke diagnosis had to meet the WHO criteria or the corresponding diagnostic criteria adopted in China,^{20–23} and had to be confirmed by CT or MRI. Trials involving participants with subarachnoid haemorrhages or cerebrovascular tumours, as well as those in which patients did not have acute stroke (within 14 days of onset¹⁹), were excluded.

Types of interventions

Patients in the experimental groups of the included trials had been treated with EA combined with RT and/or CD, at any frequency, intensity or duration. Patients in the control group of the trials had been treated by other therapies such as CD, RT and sham acupuncture, or had no treatments. However, trials that did not provide a detailed description or explanation of intervention, or those that compared different acupoint prescriptions or acupuncture types, were excluded.

Primary outcome assessments

The primary outcome for the systematic review was motor function. There are many types of motor function scales, including but not limited to the Fugl-Meyer Assessment Scale (FMA),²⁴ the modified Rankin Scale,²⁵ the Motor Assessment Scale²⁶ and Brunnstrom Stages.²⁷

Secondary outcome assessments

The secondary outcomes included measures of ADL, such as the Barthel Index (BI),²⁸ the Functional Independence Measure (FIM),²⁹ and the response or effective rate (ER). The ER was a standard of therapeutic effect recommended by the Fourth National Cerebrovascular Diseases Conference in China. The ER classified disability of stroke into five grades: cure (the scores of functional deficit were decreased up to 91%–100%), significant improvement (the scores of functional deficit were decreased at 46%–90%), improvement (the scores of functional deficit were decreased at 18%–45%), no improvement (the scores of functional deficit were decreased less than 18%) and deterioration (the scores of functional deficit were increased over 18%).³⁰ Safety assessments involved adverse events reporting and evaluation of causality. Adverse events as reported in the included trials were also recorded. The time of the outcome assessments was at the end of the intervention phase.

Electronic searches

We electronically searched databases from their respective inception up to December 2016. Databases included the China National Knowledge Infrastructure,

Chinese Biological Medicine Database, Chinese Scientific Journal Database (VIP), Cochrane Library, Medline and Embase. We combined the PubMed search with the Cochrane highly sensitive search strategy for identifying randomised trials, and adapted this strategy to search the other databases.

We also searched other resources to identify potentially relevant trials. For example, we screened the reference lists of included trials and contacted the trial authors. The detailed search steps are illustrated in online supplementary appendix 1.

Data extraction and quality assessment

Two review authors (JZ and MZ) independently scanned the titles and abstracts of the articles derived from the search and kept all potentially relevant articles. Then, they retrieved the full texts of these articles, and another two authors (ZH and RP) independently examined them to confirm that the trials met the inclusion criteria. We also recorded the reasons for exclusion of trials. If the same trial had more than one report, we only kept one originally published version. If necessary, we acquired additional information from trial authors by email or

telephone. Moreover, we discussed any disagreements to decide whether a trial should be included or excluded, and if necessary we consulted with another author (ZW).

Two authors (JZ and MZ) independently extracted information from the included trials. The information was entered into an Excel-formatted table (JZ) and the accuracy of the information entered was also checked (RP). The information extracted was as follows: trial design (eg, sample size, randomisation method, blinding method), participants (eg, gender, age), the details of intervention, outcomes (primary and secondary outcomes), adverse events, the name of the author, publication year and so on. Trial selection details are shown in a flow chart that complied with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (figure 1).

Risk of bias assessment in the included studies

Two review authors (JZ and RP) independently used the risk of bias assessment tool in the *Cochrane Handbook for Systematic Reviews of Interventions*³¹ to assess the methodological quality of each included trial. The specific domains were assessed as follows: random sequence generation, allocation sequence concealment, blinding,

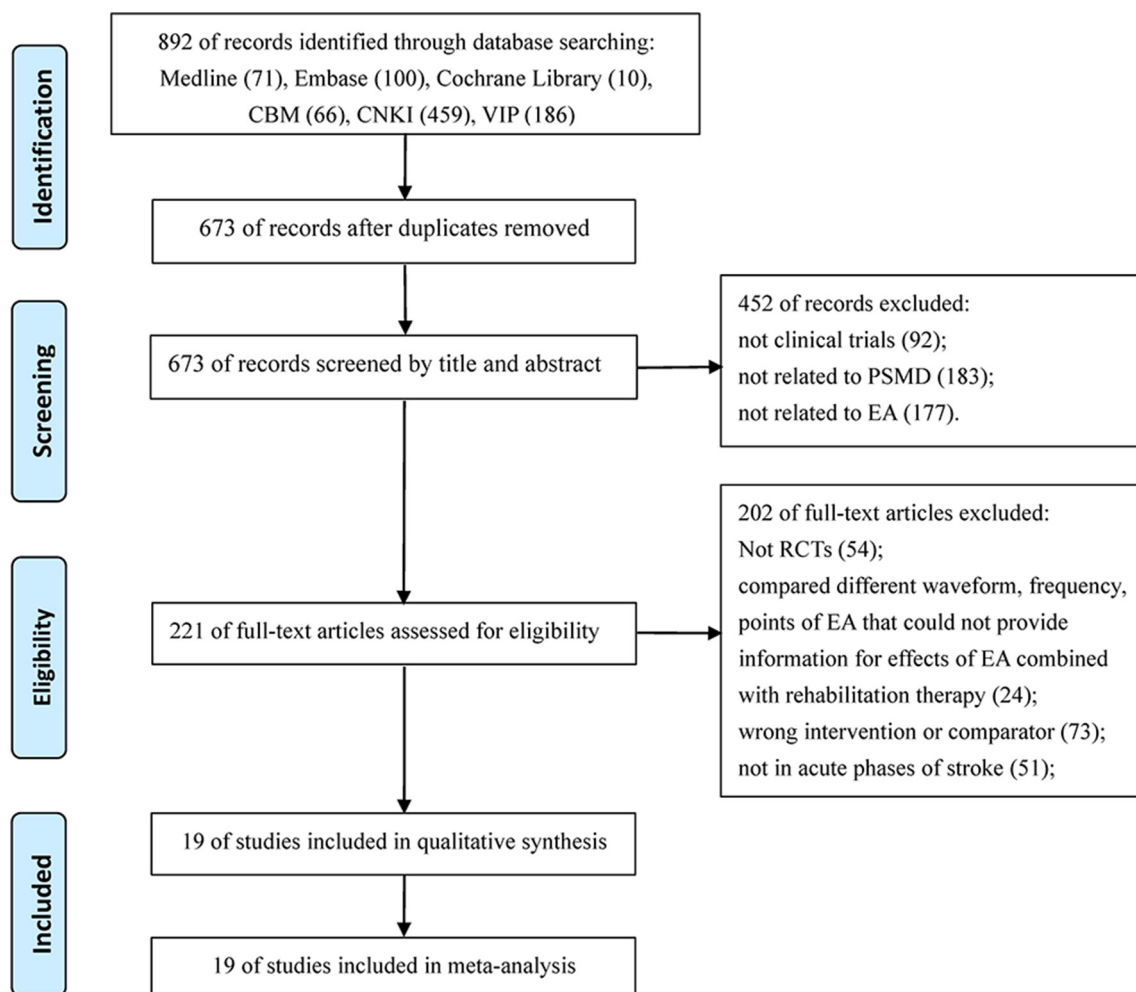


Figure 1 Study flow diagram. CBM, Chinese Biological Medicine Database; CNKI, China National Knowledge Infrastructure; EA, electroacupuncture; PSMD, poststroke motor dysfunction; RCT, randomised controlled trials; VIP, Chinese Scientific Journal Database.

incomplete outcome data, selective outcome reporting and other sources of bias. We graded the risk of bias for each domain as low risk of bias, high risk of bias or unclear. We settled quality assessment disagreements by discussion with a third reviewer (ZW).

Data analysis

We performed all statistical analyses using RevMan V.5.3 (The Nordic Cochrane Centre, The Cochrane Collaboration 2014) and Stata SE V.11.0. For continuous data, we calculated the weighted mean difference (WMD) with corresponding 95% CI. If the outcomes were measured by different scales in the included trials, we calculated the standardised mean difference (SMD) with 95% CI instead of WMD. For dichotomous data, we calculated the relative risk (RR) with 95% CI. We tested clinical and statistical heterogeneity among trials by comparing the characteristics of the trials, and used the I^2 statistics to test heterogeneity. If heterogeneity was not significant, we chose a fixed-effects model to pool the data; otherwise, we used a random-effects model after considering clinical homogeneity. When heterogeneity was substantial, we examined trials for potential explanations, or else conducted a qualitative summary rather than a meta-analysis.

A meta-regression analysis was used to explain the potential trial-level covariates such as the duration of treatment. Subgroup analyses were carried out as follows: EA and RT plus CD compared with RT plus CD, and EA plus RT compared with RT alone. Sensitivity analysis was performed to explore the robustness of our analysis, excluding studies from the overall analysis of high risk of bias due to lack of allocation concealment and blinding of assessors for the primary outcome. If the number of included trials was greater than 10, funnel plots and Egger's test were employed for publication bias analysis.

RESULTS

Trial description

We initially identified 892 relevant articles according to the search strategy, while 219 duplicates from different databases were excluded. In total, 19 trials met the eligibility criteria after being screened by title, abstract or full texts, and were included in the meta-analysis (figure 1). We did not find additional trials for this review by examining the reference lists of included trials. All trials were published between 2004 and 2016. One trial³² was published in English, while the others were all published in Chinese. A total of 1434 participants were enrolled in these trials. All trials were conducted in China. Sixteen trials³²⁻⁴⁷ compared EA plus CD and RT with CD plus RT. Three trials⁴⁸⁻⁵⁰ gave EA and RT to the experimental groups, while the control groups only received RT. CD was used according to the Chinese national guidelines, including general supportive care and specialised care such as antiplatelet agents, anticoagulants, fibrinogen-depleting agents, volume expansion and vasodilators. Because the

patients enrolled in the trials were all within 14 days after stroke, CD was used similarly in each included trial. The characteristics of the trials included in this review are shown in table 1.

Assessing risk of bias in the included trials

In general, the methodological quality of the included trials was poor. In random sequence generation, seven trials^{32 34 35 38 44 48 50} used proper generation methods with a low risk of bias, and the random number sequences were produced by either a random number table, computer software or drawing lots. One trial used an incorrect sequence generation method.⁴⁹ Eleven trials^{33 36 37 39-43 45-47} did not describe the randomisation procedure clearly. Two trials^{32 34} used concealed envelopes, and the other trials did not report allocation concealment. Two trials^{32 34} reported that outcome assessors were blind to group allocation. Two trials^{32 34} mentioned that investigators were unknown for allocation. One trial³² reported dropouts and conducted intention-to-treat analyses. In other sources of bias, 11 trials^{36 37 39-41 43 45-48 50} had a high risk because of inadequate statistical methods. These trials had not described the specific steps and methods of statistical analysis. The results of the assessments are presented in figure 2.

Primary outcomes

Fugl-Meyer Assessment Scale

The primary outcome, FMA score, was mentioned in 13 trials with 1010 patients.^{32-41 48-50} The effect of EA on FMA between the EA and non-EA groups was evaluated by a random-effects model, owing to significant heterogeneity. A meta-regression analysis was used to explain the potential covariates. The treatment duration was included as a potential covariable in the meta-regression model because the duration was from 2 to 12 weeks. However, there were no significant differences in treatment duration in the meta-regression model (adjusted R^2 : 0.124, $t=-1.57$, $P=0.144$). The FMA score in the EA group increased more than those in the non-EA group, and there was a significant difference (WMD 10.79, 95% CI 6.39 to 15.20, $P<0.001$) (figure 3).

FMA for upper extremity

One trial⁴⁷ with 98 participants used the FMA for upper extremity to evaluate the function of the upper extremity, and the difference between the EA group and the non-EA group was obvious ($P<0.050$).

FMA for lower extremity

The function of the lower extremity was assessed by FMA for lower extremity (FMA-L) in four trials⁴²⁻⁴⁵ with 234 participants. The effect on FMA-L was analysed using a fixed-effects model, and there was a significant difference between the EA group and the non-EA group in the FMA-L (WMD 5.16, 95% CI 3.78 to 6.54, $P<0.001$) (table 2). A meta-regression analysis was also carried out to explain the potential impact of treatment duration, and there was no significance in treatment duration in

Table 1 Characteristics of included trials

| Study | Sample size | Age (year) | Sex (male/female) | T regimen vs C regimen | Electrical stimulation frequency | Power | Needle retention duration | Treatment duration | Main outcomes |
|----------------------------------|----------------|--------------------------------------|----------------------|------------------------|----------------------------------|---------|---------------------------|--------------------|---------------|
| Dai <i>et al</i> ⁴⁸ | T: 32 C: 31 | T: 66.4±3.2 C: 65.7±4.1 | T: 18/14 C: 17/14 | EA+RT vs RT | NR | 2 mA | 30 min | 4 weeks | FMA, BI |
| Zhang ³³ | T: 40 C: 40 | T: 60.50 (51–69) C: 58.75 (47–72) | T: 22/18 C: 24/16 | EA+RT+CD vs RT+CD | NR | PCT | 30 min | 4–12 weeks | FMA, BI |
| Zhang <i>et al</i> ⁴⁹ | T: 41 C: 30 | T: 58.5±12.27 C: 52.43±12.24 | T: 24/17 C: 17/13 | EA+RT vs RT | NR | PCT | 20 min | 4 weeks | FMA, BI, ER |
| Li <i>et al</i> ⁵⁴ | T: 50 C: 50 | T: 64±5 C: 65±4 | T: 27/23 C: 29/21 | EA+RT+CD vs RT+CD | 15–30 Hz | 1–2 mA | 30 min | 4 weeks | FMA, BI, ER |
| Luo ⁴² | T: 30 C: 26 | T: 60.83±9.58 C: 62.47±8.72 | T: 21/9 C: 19/7 | EA+RT+CD vs RT+CD | NR | PCT | 30 min | 4 weeks | FMA-L |
| Fu <i>et al</i> ³⁵ | T: 60 C: 60 | T: 62.05±6.25 C: 63.07±5.10 | T: 37/23 C: 36/24 | EA+RT+CD vs RT+CD | NR | PCT | 30 min | 4 weeks | FMA, BI, BBS |
| Zhang <i>et al</i> ³⁶ | T: 30 C: 30 | 60.7±8.6 | 36/24 | EA+RT+CD vs RT+CD | NR | MCS | 20 min | 20 days | FMA |
| Peng <i>et al</i> ³⁷ | T: 34 C: 34 | 55 (37–77) | 41/27 | EA+RT+CD vs RT+CD | NR | PCT | 30 min | 8 weeks | FMA, BI |
| Luo ⁴³ | T: 30 C: 30 | 55 (37–78) | 39/21 | EA+RT+CD vs RT+CD | NR | PCT | 25 min | 8 weeks | FMA-L |
| Zhang ⁴⁴ | T: 29 C: 29 | T: 67.14±10.17 C: 68.14±10.93 | T: 16/13 C: 14/15 | EA+RT+CD vs RT+CD | 20 Hz | PCT | 30 min | 2 weeks | FMA-L |
| Liu ⁴⁶ | T: 46 C: 46 | T: 65 C: 62 | T: 29/17 C: 30/16 | EA+RT+CD vs RT+CD | 100 Hz | PCT | 30 min | 4 weeks | BI |
| Zhou <i>et al</i> ³⁸ | T: 35 C: 37 | T: 63.7±10.5 C: 64.5±10.6 | T: 17/18 C: 20/17 | EA+RT+CD vs RT+CD | 2.7 Hz | PCT | 30 min | 4 weeks | FMA |
| Liu <i>et al</i> ⁴⁷ | T: 49 C: 49 | 55±0.24 | T: 26/23 C: 24/25 | EA+RT+CD vs RT+CD | 10–15 Hz | PCT | 20 min | 8 weeks | FMA-U |
| Peng <i>et al</i> ⁶⁰ | T: 40 C: 40 | 58.5 (35–78) | 48/32 | EA+RT vs RT | 1.0–2.0 Hz | PCT | 30 min | 6 weeks | FMA, BI |
| Zhang and Li ³⁹ | T: 49 C: 49 | T: 51.49±8.35 C: 54.73±6.75 | T: 26/23 C: 24/25 | EA+RT+CD vs RT+CD | 10–15 Hz | PCT | 20 min | 8 weeks | FMA, BI |
| Hsieh <i>et al</i> ³² | T: 30 C: 33 | T: 68.8 C: 70.7 | T: 12/18 C: 23/10 | EA+RT+CD vs RT+CD | 3–15 Hz | 1–10 mA | 20 min | 4 weeks | FMA, FIM |
| Liu and Zou ⁴⁰ | T: 38 C: 37 | T: 59.4±10.2 C: 56.4±10.5 | T: 25/13 C: 14/23 | EA+RT+CD vs RT+CD | 3.3 Hz | PCT | 30 min | 20 days | FMA, BI |
| Peng <i>et al</i> ⁴¹ | T: 30 C: 30 | 55 (37–78) | 39/21 | EA+RT+CD vs RT+CD | NR | PCT | 25 min | 8 weeks | FMA |
| Xie <i>et al</i> ⁴⁵ | T: 34 C: 26 | T: 53.0±9.3 C: 56.5±6.4 | T: 22/12 C: 17/9 | EA+RT+CD vs RT+CD | 2.7 Hz | MCS | 20 min | 30–40 days | FMA-L, BI |

BBS, Berg Balance Scale; BI, Barthel Index; C, control group; CD, conventional drugs; EA, electroacupuncture; ER, effective rate; FIM, Functional Independence Measure; FMA, Fugl-Meyer Assessment Scale; FMA-L, Fugl-Meyer Assessment Scale for lower extremity; FMA-U, Fugl-Meyer Assessment Scale for upper extremity; MCS, muscles contract slightly; NR, not referred; PCT, patient can tolerate; RT, rehabilitation therapy; T, treatment group.

| | Random sequence generation (selection bias) | Allocation concealment (selection bias) | Blinding of participants and personnel (performance bias) | Blinding of outcome assessment (detection bias) | Incomplete outcome data (attrition bias) | Selective reporting (reporting bias) | Other bias |
|--------------------|---|---|---|---|--|--------------------------------------|------------|
| [32] Hsieh 2007 | + | + | + | + | + | + | + |
| [33] Zhang AH 2015 | ? | - | - | - | + | + | ? |
| [34] Li XJ 2014 | + | + | + | + | + | + | + |
| [35] Fu Y 2013 | + | - | - | - | + | + | ? |
| [36] Zhang C 2013 | ? | - | - | - | + | + | - |
| [37] Peng L 2011 | ? | - | - | - | + | + | - |
| [38] Zhou HY 2009 | + | - | - | - | + | + | ? |
| [39] Zhang H 2008 | ? | - | - | - | + | + | - |
| [40] Liu Y 2007 | ? | - | - | - | + | + | - |
| [41] Peng L 2004 | ? | - | - | - | + | + | - |
| [42] Luo XP 2014 | ? | - | - | - | + | + | ? |
| [43] Luo Y 2011 | ? | - | - | - | + | + | - |
| [44] Zhang X 2011 | + | - | - | - | + | + | ? |
| [45] Xie DL 2004 | ? | - | - | - | + | + | - |
| [46] Liu H 2010 | ? | - | - | - | + | + | - |
| [47] Liu HH 2009 | ? | - | - | - | + | + | - |
| [48] Dai R 2016 | + | - | - | - | + | + | - |
| [49] Zhang SY 2015 | - | - | - | - | + | + | ? |
| [50] Peng LH 2008 | + | - | - | - | + | + | - |

Figure 2 Risk of bias assessments of included studies.

the meta-regression model (adjusted R^2 : -0.198, t =-0.86, P =0.482).

Secondary outcomes

Activities of daily living

The effect of EA on ADL was analysed using a random-effects model, due to significant heterogeneity in 12

trials^{32-35 37 39 40 45 46 48-50} with 970 participants. We calculated the SMD with 95% CI as the outcomes were measured by different scales (FIM and BI) in the included trials. The improvement of ADL in the EA group was better than that in the non-EA group (SMD 1.37, 95% CI 0.79 to 1.96, P <0.001) (figure 4).

Response or ER

Two trials^{34 49} with a total of 171 participants showed that there was no significant difference in ER between EA and non-EA groups (RR 1.13, 95% CI 1.00 to 1.27, P =0.050; fixed-effects model) (table 2).

Berg Balance Scale

Berg Balance Scale (BBS) was assessed in one trial³⁵ with 120 participants. The improvement of BBS in the EA group was preferable to that in the non-EA group (P <0.050).

Safety assessment

None of the included trials mentioned adverse events due to both EA combined with RT and/or CD and RT and/or CD alone.

Subgroup analysis

EA plus RT and CD versus RT plus CD

Ten trials³²⁻⁴¹ used FMA to measure the motor function of 796 participants with PSMD. A random-effects model was used to analyse the effect on FMA and ADL in this subgroup analysis due to significant heterogeneity. There was a significant difference between EA combined with RT and CD versus RT plus CD (WMD 8.03, 95% CI 5.17 to 10.90, P <0.001) (figure 3). Nine trials^{32-35 37 39 40 45 46} used BI or FIM to measure the ADL of 756 patients following PSMD. EA plus RT and CD for the improvement of ADL was better than that of RT plus CD (SMD 1.29, 95% CI 0.55 to 2.02, P <0.001) (figure 4).

EA plus RT versus RT alone

Three trials⁴⁸⁻⁵⁰ with 214 participants applied FMA to compare the effectiveness of EA plus RT against RT alone. Meta-analyses with a random-effects model were performed to evaluate the effect on FMA and ADL in this subgroup analysis owing to statistical heterogeneity. There was a significant difference in these three trials (WMD 20.90, 95% CI 18.61 to 23.19, P <0.001) (figure 3). In the comparison of EA plus RT versus RT alone in the three trials, the difference in ADL was obvious (SMD 1.63, 95% CI 1.01 to 2.25, P <0.001) (figure 4).

Sensitivity analysis

We used the method of removing item by item to test the stability of meta-analysis, and the results showed that there had been no noticeable change on any of the outcomes. The difference between the random-effects and fixed-effects models may have influenced the outcomes. Therefore, we used different statistical models to pool the data for the FMA, FMA-L, ADL and ER. No observable change in any of the outcomes was found (table 2).

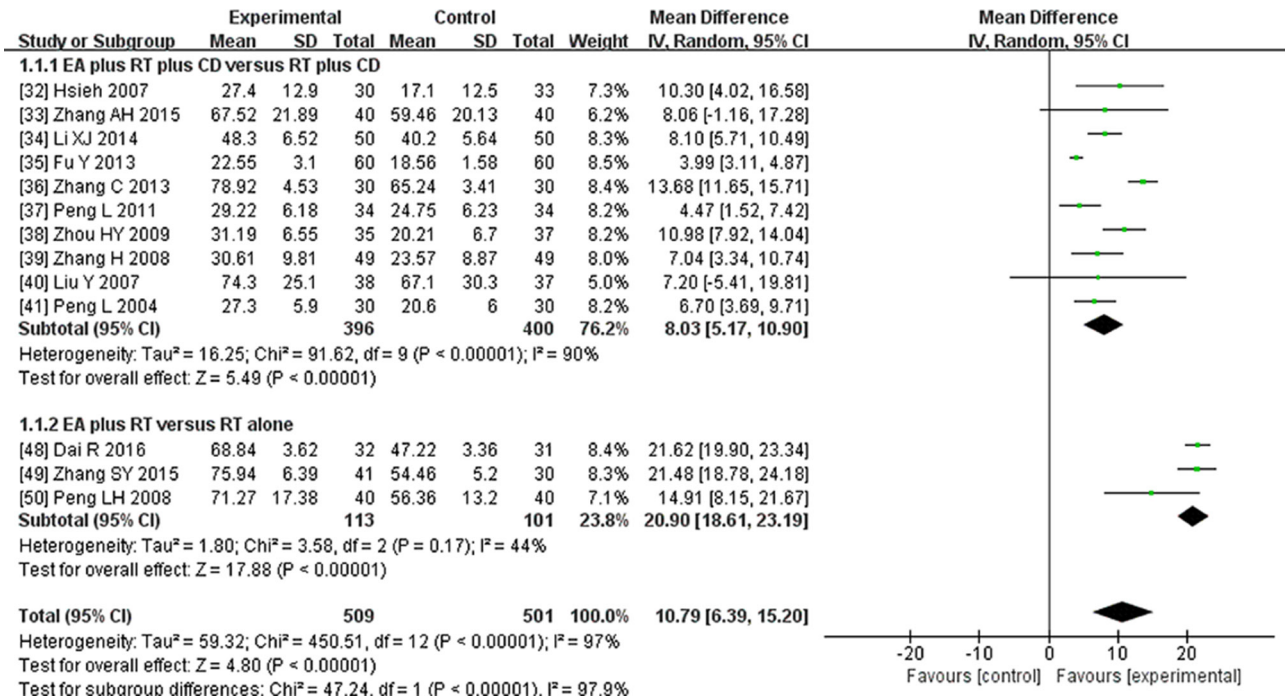


Figure 3 Forest plot and meta-analysis of Fugl-Meyer Assessment Scale. CD, conventional drugs; EA, electroacupuncture; RT, rehabilitation therapy.

Furthermore, sensitivity analysis was carried out to explore the robustness of our analysis, excluding trials from the overall analysis of high risk of bias due to lack of adequate sequence generation, allocation concealment and blinding of assessors for primary outcomes (table 2). The effects on FMA, ADL and ER were robust with random-effects and fixed-effects models with adequate sequence generation, with the exception of the comparison of EA plus RT versus RT alone and the trial subgroups (table 2).

Publication bias

Thirteen trials^{32-41 48-50} and 12 trials^{32-35 37 39 40 45 46 48-50} respectively showed a difference in FMA and ADL between the EA and the non-EA groups. Egger’s tests showed that there were publication biases for the included trials of FMA (t=5.21, P<0.001) or ADL (t=3.61, P=0.005). The funnel plots showed that some trials did not lie inside the 95% CI and these were asymmetric. This may indicate potential publication bias (figure 5 and figure 6).

DISCUSSION

This systematic review included 19 RCTs with 1434 participants comparing the effectiveness and safety of EA therapy and non-EA therapy. The meta-analysis of four RCTs with 234 patients showed that adjunctive EA was better in improving the motor function of the lower extremity. One RCT with 98 patients demonstrated that added EA was beneficial for upper extremity motor control. The meta-analysis of 13 RCTs with 1010 patients indicated that adjunctive EA had greater advantage in the recovery of overall motor function, and the

pooled results of 12 RCTs with 970 patients revealed that adjunctive EA was beneficial in the improvement of ADL. There was no any difference in the ER between EA and non-EA. However, it should be noted that the review included add-on designed trials of EA plus RT and/or CD, which suggests that EA is a complementary therapy for PSMD. Meanwhile, there were insufficient data to assess the safety of EA plus RT, EA plus CD, and EA plus both CD and RT because none of the included trials reported adverse events. Considering the pooled effects on FMA and FMA-L with significant heterogeneity, a meta-regression analysis was conducted to explain the impact of treatment duration as a covariable. The result showed that the treatment duration was not significant in the meta-regression model. This means the heterogeneity could not be explained by the trial level’s treatment duration.

One should be cautious when interpreting the results of the review due to some limitations. In this review, most of the included trials had small sample sizes. Sixty-three per cent of the included trials used a high risk of bias method or did not describe the generation of a random sequence. Eighty-nine per cent of the trials did not report allocation concealment and had inadequate blinding of outcome assessments. Fifty-eight per cent of the trials used inadequately designed statistical methods or did not fully describe the statistical methods. Only two trials^{32 34} were well-designed to assess the effect of EA combined with RT for PSMD. Additionally, all trials included in this review were conducted in China and most were published in Chinese, which likely lead to a potential bias and therefore limit their representativeness.

Table 2 Results of sensitivity analysis

| Study type | Participants (n) | | Study heterogeneity | | | | Analysis model | WMD (95% CI) | P value |
|---|------------------|------------------|---------------------|----------|----|--------------------|----------------|------------------------|---------|
| | Studies (n) | Experiment group | Control group | χ^2 | df | I ² , % | | | |
| FMA | | | | | | | | | |
| EA vs non-EA | 13 | 509 | 501 | 450.51 | 12 | 97 | Random | 10.79 (6.39 to 15.20) | <0.001 |
| | | | | | | | Fixed‡ | 8.97 (8.35 to 9.58) | <0.001 |
| EA plus RT plus CD vs RT plus CD | 10 | 396 | 400 | 91.62 | 9 | 90 | Random | 8.03 (5.17 to 10.90) | <0.001 |
| | | | | | | | Fixed‡ | 6.13 (5.45 to 6.82) | <0.001 |
| EA plus RT vs RT alone | 3 | 113 | 101 | 3.58 | 2 | 44 | Random | 20.90 (18.61 to 23.19) | <0.001 |
| | | | | | | | Fixed‡ | 21.29 (19.86 to 22.71) | <0.001 |
| FMA for lower extremity | | | | | | | | | |
| EA vs non-EA | 4 | 123 | 111 | 1.76 | 3 | 0 | Random‡ | 5.16 (3.78 to 6.54) | <0.001 |
| | | | | | | | Fixed | 5.16 (3.78 to 6.54) | <0.001 |
| Activities of daily living | | | | | | | | | |
| EA vs non-EA | 12 | 494 | 476 | 176.60 | 11 | 94 | Random | 1.37 (0.79 to 1.96)† | <0.001 |
| | | | | | | | Fixed‡ | 1.20 (1.06 to 1.35)† | <0.001 |
| EA plus RT plus CD vs RT plus CD | 9 | 381 | 375 | 162.20 | 8 | 95 | Random | 1.29 (0.55 to 2.02)† | <0.001 |
| | | | | | | | Fixed‡ | 1.10 (0.94 to 1.27)† | <0.001 |
| EA plus RT vs RT alone | 3 | 113 | 101 | 7.78 | 2 | 74 | Random | 1.63 (1.01 to 2.25)† | <0.001 |
| | | | | | | | Fixed‡ | 1.57 (1.25 to 1.88)† | <0.001 |
| Response or effective rate | | | | | | | | | |
| EA vs non-EA | 2 | 91 | 80 | 2.80 | 1 | 64 | Random‡ | 1.13 (0.90 to 1.42)* | 0.290 |
| | | | | | | | Fixed‡ | 1.13 (1.00 to 1.27)* | 0.050 |
| Trials with adequate sequence generation: FMA | | | | | | | | | |
| EA vs non-EA | 6 | 247 | 251 | 327.63 | 5 | 98 | Random‡ | 11.61 (4.17 to 19.04) | 0.002 |
| | | | | | | | Fixed‡ | 7.98 (7.26 to 8.69) | <0.001 |
| Trials with adequate concealed allocation and blinding of assessors: FMA | | | | | | | | | |
| EA vs non-EA | 2 | 80 | 83 | 0.41 | 1 | 0 | Random‡ | 8.38 (6.14 to 10.61) | <0.001 |
| | | | | | | | Fixed‡ | 8.38 (6.14 to 10.61) | <0.001 |

*Presented as relative risk.

†Presented as standardised mean difference.

‡Represents the meta-analysis results not shown in the figures.

CD, conventional drugs; EA, electroacupuncture; FMA, Fugl-Meyer Assessment Scale; RT, rehabilitation therapy; WMD, weighted mean difference.

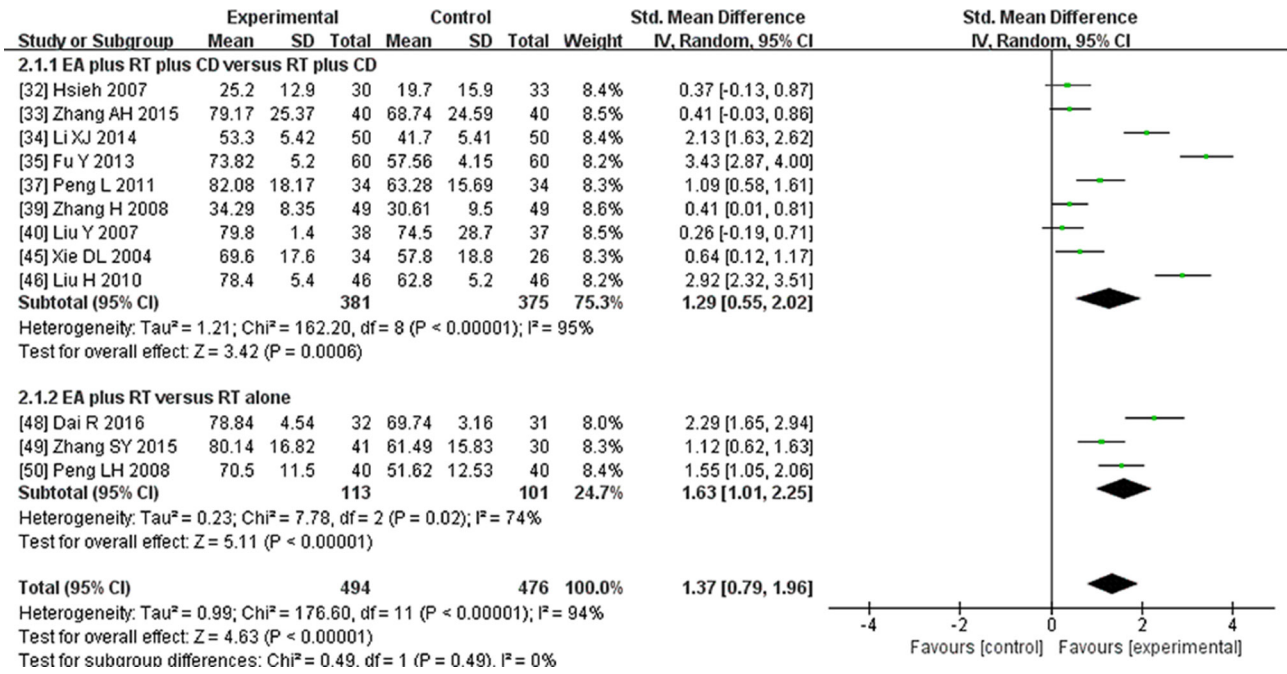


Figure 4 Forest plot and meta-analysis of activities of daily living. CD, conventional drugs; EA, electroacupuncture; RT, rehabilitation therapy.

Two previous systematic reviews were reported to assess the effects of EA on the treatment of ischaemic stroke.^{17,18} However, there were noticeable differences in the study characteristics of participants, comparison and outcomes, and both reviews did not focus on evaluating EA as an adjunctive therapy for PSMD of acute stage within 14 days of onset. We still need to gain a clear evidence for PSMD in this critical stage, which will influence the prognosis of stroke.

Most of the included trials had methodological defects, and the funnel plots of FMA and ADL suggested a potential publication bias. These issues potentially lead to low quality of evidence, over-reporting of positive results and under-reporting of adverse events. Also, the use of diverse RT as an add-on basis in the included trials makes it difficult to pool data using a fixed-effect model to interpret the clinical significance of EA. Therefore, the potential benefits of EA as an adjunctive therapy for PSMD evident

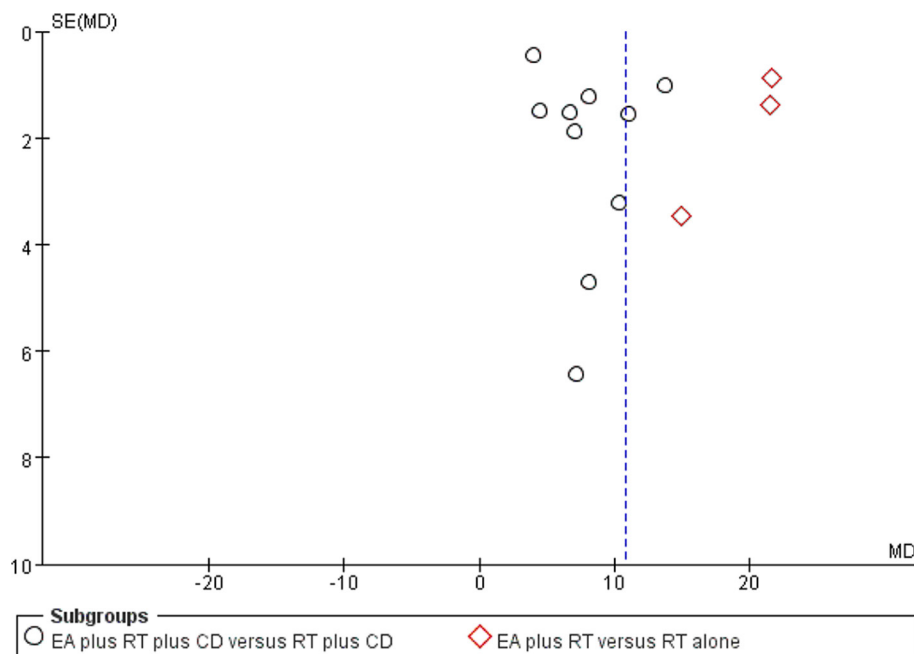


Figure 5 Funnel plots illustrating meta-analysis of Fugl-Meyer Assessment Scale. CD, conventional drugs; EA, electroacupuncture; MD, mean difference; RT, rehabilitation therapy.

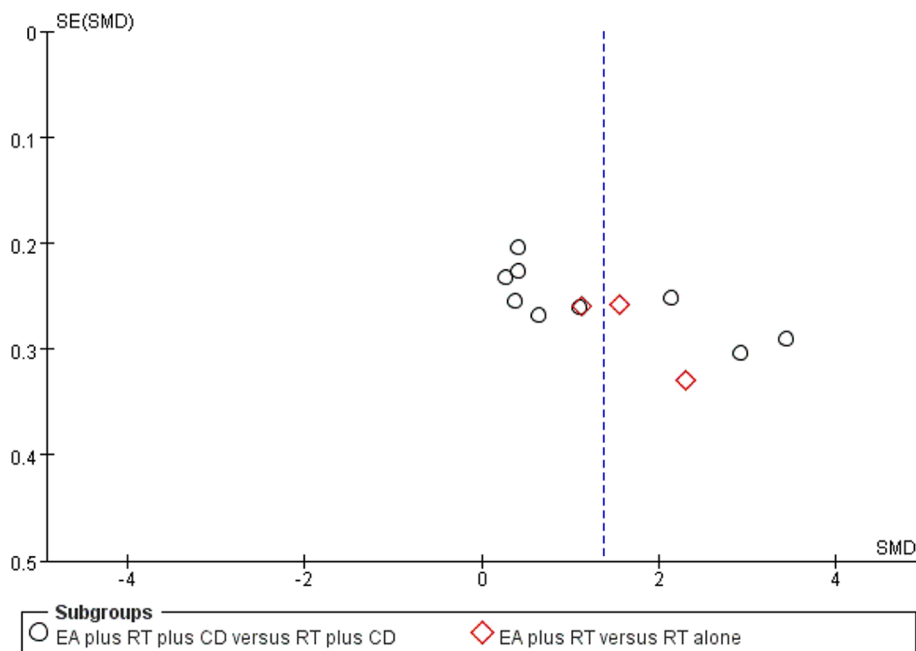


Figure 6 Funnel plots illustrating meta-analysis of activities of daily living. CD, conventional drugs; EA, electroacupuncture; RT, rehabilitation therapy; SMD, standardised mean difference.

in this review need to be further appraised through well-designed, large-scale, multicentre RCTs.

CONCLUSIONS

EA as a complementary therapy seems to have clinical benefits in terms of improving the function of extremities, ADL and balance function. However, these apparent benefits require further evaluation through well-designed multicentre trials with large sample sizes. The safety of EA combined with RT and/or CD is still uncertain.

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