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RESEARCH ARTICLE

Impact of Adjuvant Chemotherapy on Breast Cancer Survival: A Real-World Population

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Abstract

Background

The impact of adjuvant chemotherapy on breast cancer prognosis has been demonstrated in randomized trials, but its impact is unknown in real-world populations. The aim of this study was to evaluate the effect of adjuvant chemotherapy on the survival of breast cancer patients in an unselected population.

Methods

This prospective cohort study included 32,502 women treated at the Institut Curie between 1981 and 2008 for a first invasive breast cancer without metastasis. The patients were matched based on their propensity score to receive adjuvant chemotherapy.

Results

The matching generated a subsample of 9,180 patients with an overlapping propensity score. In the group without chemotherapy, the overall survival (OS) rates at 5 and 10 years of follow-up were 87.6% (95% CI [86.7–88.6]) and 75.0% (95% CI [73.6–76.5]), respectively, versus 92.1% (95% CI [91.3–92.9]) and 81.9% (95% CI [80.6–83.2]), respectively, in the chemotherapy group. Distant disease-free survival (DDFS) was significantly improved in the five first years (absolute benefit of 3.5%). In a multivariate analysis, adjuvant chemotherapy was associated with better OS (HR = 0.75, 95% CI [0.69–0.83], p<0.0001) and DDFS (HR = 0.82, 95% CI [0.75–0.90], p<0.0001).

Conclusion

Adjuvant chemotherapy significantly improves OS and DDFS rates in an unselected population, in accordance with previous results reported by randomized trials.

Introduction

Over the past thirty years, the incidence of breast cancer has greatly increased by 2-3% each year in numerous countries until the mid-2000's, including several in Europe and North America [1-3]. However, breast cancer mortality has not similarly increased; in fact, it has slightly decreased by approximately 2% each year in most of these countries. This stability and the more recent decline in mortality (despite the increasing incidence rate) reflect an improvement in breast cancer patient survival. This improvement could be explained by treatment advancements in the late 1990s and the increasing proportion of cancers detected at an early stage due to the development of mammographic screening practices.

The advent of chemotherapy most likely contributed significantly to the therapeutic progress observed in the treatment of breast cancer. Indications and prescriptions increased during the 1980s concurrent with the use of cyclophosphamide, methotrexate, and 5-fluorouracil (CMF)-like and anthracycline regimens for patients with metastases. Anthracyclines were later employed in the initial treatment of high-risk breast cancers in the mid-1990s, as were taxane regimens for metastatic patients. Taxanes were used in adjuvant primary treatment in the mid-2000s [4]. Chemotherapy reduced the risk of death due to invasive breast cancers by between 7% and 33% in randomized trials and large meta-analyses; this varied according to tumor characteristics, patient age, and the type and duration of treatment[5]. However, there are very few data on the effects of chemotherapy in real-world populations; additionally, the sparse data for these patients were acquired using limited sub-populations, such as those composed of metastatic patients.

Importantly, an apparent overall effect should not be confused with the actual treatment effect. In experimental conditions, investigators typically achieve this result through the randomization of groups that differ only in treatment allocation [6]. Several methods are feasible for determining the therapeutic effect of chemotherapy, particularly multivariate models and propensity scores [7-8]. These methods are intended to reduce confounding bias and thus clarify the influence attributed to treatment. The propensity score is the probability of being exposed to treatment. This probability is calculated by introducing the main factors that are significantly related to the allocation of treatment into a multivariate regression model. The coefficients for each factor in the regression equation enable score calculation. A propensity score has never been used to determine the impact of breast chemotherapy in real-world populations, with the exception of very specific patient subgroups [9-10].

The aim of this study was to assess the specific impact of chemotherapy on the survival of breast cancer patients in an unselected population.

Materials and Methods

Population and data source

Eligible patients included women treated for a primary invasive breast cancer without distant metastasis in the Institut Curie between 1981 and 2008. Patients with a history of previous cancer, with primary chemotherapy, or whose follow-up was less than 3 months were not included.

Data were collected prospectively in this cohort study and were analyzed anonymously. The hospital database was created by knowledgeable workers and validated by a physician. This database was also regularly updated to collect the health status or recent vital status of the patients via phone calls to their doctors and via contacting the governmental records. Data including clinical patient data (age, height, weight, medical history, menopausal status, date of diagnosis, etc.) and tumor characteristics (size, nodal status, Scarff Boom and Richardson

(SBR) grade, hormone receptor overexpression (the hormone receptor status was considered positive if the proportion of stained tumor nuclei was \geq 10%), human epidermal growth factor receptor-2 (HER2) status, etc.) were gathered from medical records and pathology reports. The treatment sequence was also recorded: dates of surgery, chemotherapy, radiotherapy, hormone therapy or the administration of targeted therapies such as trastuzumab, as well as adjuvant or neoadjuvant administration. Information regarding the recurrence and vital status of the patients was collected and updated every 12 months.

Study endpoints

Overall survival (OS) was the primary endpoint. We also evaluated the distant disease-free survival (DDFS) and the interval without distant metastasis.

Statistical analysis

Evolution of patients' characteristics and treatments. To analyze the data by time period, we grouped the dates of diagnosis into 6 periods: a period of 3 years from 1981 to 1983 and 5 periods of 5 years from 1984 to 2008. The analyses of changes in clinical and histological data were performed using the Chi-square test or Fisher's exact test for qualitative variables and Student's t-test or analysis of variance (ANOVA) for the comparison of means.

Survival analysis. The date of diagnosis of breast cancer was considered the start of follow-up. The study cut-off point was November 13th 2013. The OS was measured from the date of diagnosis to the date of death (any cause) or to the date of last contact. DDFS was measured from the date of diagnosis until the date of the distant metastasis of breast cancer or last patient contact. In a univariate analysis, OS and DDFS were assessed using the Kaplan-Meier method, and log-rank tests were used to compare the differences between the resulting curves. Hazard ratios, 95% confidence intervals (CI) and multivariate survival analyses were performed using the Cox proportional-hazards model (adjusted for time period, age, menopausal status, body mass index (BMI), histological type, histological tumor size, histological nodal status, SBR grade and hormone receptor status). A stepwise backward approach was used to select the covariates that were significantly associated with survival.

Primary survival analyses were performed on the population with untreated invasive breast cancer. However, a treatment selection bias was evident in this cohort; this bias was related to multiple factors influencing the decision to administer adjuvant chemotherapy. As a second step, we did a case-control analysis, nested in the cohort. We used a propensity score analysis to minimize the effect of this confounding factor [11]. We derived the propensity score from a logistic regression model using variables associated with the indication of chemotherapy (age, histological type, histological tumor size, histological nodal status, SBR grade, hormone receptor status and time period of treatment) to achieve maximal group similarity for these parameters rather than on the basis of statistical significance. Thus, each patient was assigned a propensity score corresponding to the likelihood of receiving chemotherapy. We matched patients (1 to 1) on their propensity score +/- 0.05 using the "nearest-neighbor" matching method. We compared the survival of the group treated with chemotherapy with that of the non-treated group (matched using propensity scores).

All statistical tests are reported with two-tailed p-values and 95% CIs at an alpha level of 0.05 or lower. All statistical analyses were performed using R software (<u>http://cran.r-project.</u> org, The R Foundation for Statistical Computing, Vienna, Austria, version 0.98.978, 2009–2013 RStudio, Inc.) with the *rms*, *survival*, and *MatchIt* packages.

Ethics Statement. The registration of patients of the Institut Curie in this cohort received a favorable agreement of the french National Committee on Computers and Liberties (CNIL,

Commission nationale de l'informatique et des libertes). Patients gave informed written consent prior to be registered in the cohort. The study was approved by the breast cancer study group and the comity of clinical research study of the Institut Curie.

Results

Patients characteristics over time

Between January 1981 and December 2008, 48,469 patients were identified in the Institut Curie database. Among these patients, 32,502 women had a previously untreated invasive cancer and met the inclusion criteria. The number of patients increased with time, and their characteristics are detailed in <u>Table 1</u>. The proportion of patients between 50 and 65 years of age increased over time. Grade 1 lesions, hormone receptor positive cancers, and non-palpable (T0) tumors or tumors less than 2 cm (T1) without lymph-node involvement were more prevalent in the most recent periods. The rates of adjuvant chemotherapy, radiotherapy and hormonal therapy increased over time and were applied to a mean of 28.2%, 78.5% and 67.2% of patients, respectively.

The median follow-up was 100 months. Eighty-five per cent of women were followed for at least 5 years or until death. Over the 28 years of the study, 8,119 (25%) deaths and 5,946 distant metastasis (18%) were observed among the 32,502 patients with invasive breast cancer.

Matching on propensity scores

From the full sample of 32,502 women with an invasive untreated breast cancer, 29,382 patients without missing data could be included in the model, and a subsample was created composed of 9,180 women with overlapping propensity scores to receive chemotherapy. Propensity score matching improved the similarities in each factor distribution and resulted in overall propensity scores that were not significantly different after matching (Table 2, Fig 1A–1D). This approach resulted in the selection of all patients with low propensity scores who received chemotherapy and all patients with high propensity scores who did not receive chemotherapy (Fig 1E and 1F).

OS

In a univariate analysis, the OS of the 9,165 patients who received adjuvant chemotherapy was significantly better than that of the 20,354 patients who did not receive chemotherapy (raw HR = 0.89, 95% CI [0.84–0.96], p<0.0001) (Fig 2, Table 3). However, this significant difference did not reflect the clinical impact of chemotherapy. The patients who received chemotherapy were younger than the patients treated without chemotherapy (Table 2) with better late OS (selection bias). After matching via propensity scores, a univariate analysis showed that the patients who received adjuvant chemotherapy exhibited a better OS than those who did not receive chemotherapy (HR = 0.69, 95% CI [0.63-0.75], p< 0.0001). The 5-year survival rate in the patients treated with chemotherapy was 92.1% with a 95% CI [91.3-92.9] versus 87.6% with a 95% CI [86.7-88.6] in the untreated group (absolute benefit: 4.5%). At 10 years, the survival rate in the group treated with chemotherapy was 81.9% with a 95% CI [80.6-83.2] versus 75.0% with a 95% CI [73.6-76.5] in the untreated group (absolute benefit: 6.9%). In multivariate analysis of the matched population, the difference of OS attributable to the adjuvant chemotherapy remained significant (HR = 0.75, 95% CI [0.69-0.83], p< 0.0001) (Table 4). The hormone receptor overexpression status, the pathological tumor size, the number of invaded nodes, and the SBR grade were also associated with prognosis.

Table 1. Patient characteristics according to diagnosis periods.

	Years of diagnosis			1984	1984–1988		1989–1993		1994–1998		1999–2003		2004–2008		Total	
		N	(%)	N	(%)	N	(%)	N	(%)	N	(%)	N	(%)	N	(%)	p-value
	All	2331	(7.2)	4401	(13.6)	5886	(18.1)	6407	(19.7)	4912	(15.1)	8565	(26.3)	32502	(100)	
	< 35	64	(2.7)	102	(2.3)	134	(2.3)	110	(1.7)	69	(1.4)	136	(1.6)	615	(1.9)	<0.0001
	35 to 49	685	(29.4)	1229	(27.9)	1604	(27.2)	1666	(26.0)	1058	(21.5)	1919	(22.4)	8161	(25.1)	
Age (years)	50 to 65	883	(37.9)	1896	(43.1)	2590	(44.0)	2930	(45.7)	2342	(47.7)	3901	(45.5)	14542	(44.7)	
	> 65	647	(27.8)	1149	(26.1)	1558	(26.5)	1701	(26.5)	1443	(29.4)	2609	(30.5)	9107	(28.0)	
	NA	52	(2.2)	25	(0.6)	0	(0)	0	(0)	0	(0)	0	(0)	77	(0.2)	
	post- menopausal	1287	(55.2)	2523	(57.3)	3323	(56.5)	3331	(52.0)	3198	(65.1)	5577	(65.1)	19239	(59.2)	<0.0001
Menopausal status	pre- menopausal	799	(34.3)	1443	(32.8)	1683	(28.6)	1699	(26.5)	1163	(23.7)	1946	(22.7)	8733	(26.9)	
	NA	245	(10.5)	435	(9.9)	880	(14.9)	1377	(21.5)	551	(11.2)	1042	(12.2)	4530	(13.9)	
Clinical tumor size	Т0	62	(2.7)	175	(4.0)	564	(9.6)	875	(13.7)	770	(15.7)	1222	(14.3)	3668	(11.3)	< 0.0001
	T1	636	(27.3)	1505	(34.2)	2097	(35.6)	2590	(40.4)	2446	(49.8)	4980	(58.1)	14254	(43.9)	
	T2	1287	(55.2)	2272	(51.6)	2548	(43.3)	2357	(36.8)	1377	(28.0)	1808	(21.1)	11649	(35.8)	
	Т3	253	(10.8)	269	(6.1)	408	(6.9)	352	(5.5)	159	(3.2)	235	(2.7)	1676	(5.2)	
	Т4	86	(3.7)	147	(3.3)	257	(4.4)	216	(3.4)	108	(2.2)	111	(1.3)	925	(2.8)	
	NA	7	(0.3)	33	(0.7)	12	(0.2)	17	(0.3)	52	(1.1)	209	(2.4)	330	(1.0)	
	0	1100	(47.2)	2339	(53.1)	3309	(56.2)	3600	(56.2)	2981	(60.7)	5493	(64.1)	18822	(57.9)	<0.0001
	1 to 3	535	(22.9)	1028	(23.4)	1169	(19.9)	1420	(22.2)	1186	24.1)	2060	(24.0)	7398	(22.8)	
Involved axillary nodes	4 to 9	214	(9.2)	335	(7.6)	429	(7.3)	494	(7.7)	323	(6.6)	550	(6.4)	2345	(7.2)	
axinary nodes	≥ 10	91	(3.9)	157	(3.6)	180	(3.1)	191	(3.0)	94	(1.9)	184	(2.1)	897	(2.8)	
	NA	391	(16.8)	542	(12.3)	799	(13.6)	702	(11.0)	328	(6.7)	278	(3.2)	3040	(9.3)	
	1	352	(15.1)	1048	(23.8)	1450	(24.6)	1775	(27.7)	1664	(33.9)	2543	(29.7)	8832	(27.2)	<0.0001
CDD arredo	2	1027	(44.1)	1828	(41.5)	2136	(36.3)	2126	(33.2)	1904	(38.8)	3682	(43.0)	12703	(39.1)	
SBR grade	3	853	(36.6)	1226	(27.9)	1784	(30.3)	2050	(32.0)	1051	(21.4)	1953	(22.8)	8917	(27.4)	
	NA	99	(4.2)	299	(6.8)	516	(8.8)	456	(7.1)	293	(6.0)	387	(4.5)	2050	(6.3)	
Hormone	Hormone receptor +	1497	(64.2)	3029	(68.8)	4204	(71.4)	4282	(66.8)	3723	(75.8)	7181	(83.8)	23916	(73.6)	<0.0001
receptor status	Hormone receptor -	394	(16.9)	689	(15.6)	917	(15.6)	834	(13.0)	562	(11.4)	1108	(12.9)	4504	(13.9)	
	NA	440	(18.9)	683	(15.5)	765	(13.0)	1291	(20.1)	627	(12.8)	276	(3.2)	4082	(12.6)	
Adjuvant	no	2080	(89.2)	3630	(82.5)	4452	(75.6)	4616	(72.0)	3164	(64.4)	5395	(63.0)	23337	(71.8)	<0.0001
chemotherapy	yes	251	(10.8)	771	(17.5)	1434	(24.4)	1791	(28.0)	1748	(35.6)	3170	(37.0)	9165	(28.2)	
Hormone	no	1575	(67.6)	2814	(63.9)	3929	(66.7)	3982	(62.1)	1720	(35.0)	2812	(32.8)	16832	(51.8)	<0.0001
therapy	yes	756	(32.4)	1587	(36.1)	1957	(33.3)	2425	(37.9)	3192	(65.0)	5753	(67.2)	15670	(48.2)	
Dedietherrow	no	1156	(49.6)	1533	(34.8)	1495	(25.4)	1060	(16.5)	741	(15.1)	991	(11.6)	6976	(21.5)	<0.0001
Radiotherapy	yes	1175	(50.4)	2868	(65.2)	4391	(74.6)	5347	(83.5)	4171	(84.9)	7574	(88.4)	25526	(78.5)	

NA: not assessable. T0: no palpable tumor. T1: clinical tumor size \leq 2 cm. T2: tumor > 2 cm and \leq 5 cm. T3: >5 cm. T4: extension to skin and/or chest wall. SBR: Scarff Bloom and Richardson.

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DDFS

Prior to matching, the patients who received adjuvant chemotherapy had a worse DDFS than those who did not receive chemotherapy, as found in a univariate analysis (raw HR = 1.41, 95% CI [1.34–1.49], p<0.0001) (Fig 2, Table 3); this reflected an indication bias (patients with worse prognosis received chemotherapy more often). After matching for propensity scores, the

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		PRIOR TO MATCHING					AFTER MATCHING					
		no chemotherapy		chemo	otherapy		no chemotherapy		chemotherapy			
		N	(%)	N	(%)	p-value	N	(%)	N	(%)	p-value	
	All	23337	(71.8)	9165	(28.2)	•	4590	(50)	4590	(50)		
	< 35	257	(1.1)	358	(3.9)	<0.0001	102	(2.2)	186	(4.0)	<0.0001	
	35 to 49	4630	(19.8)	3531	(38.5)		1349	(29.4)	1346	(29.3)		
Age (years)	50 to 65	10172	(43.6)	4370	(47.7)		2274	(49.5)	2281	(49.7)		
	> 65	8201	(35.1)	906	(9.9)		865	(18.8)	777	(16.9)		
	NA	77	(0.3)	0	(0)		0	(0)	0	(0)		
	Ductal	15704	(67.3)	7594	(82.9)	< 0.0001	3593	(78.3)	3558	(77.5)	0.68	
Histological type	Lobular	2925	(12.5)	926	(10.1)		549	(12.0)	569	(12.4)		
	Other	4708	(20.2)	645	(7.0)		448	(9.8)	463	(10.1)		
	pT1	13989	(59.9)	4578	(49.9)	< 0.0001	2526	(55.0)	2593	(56.5)	0.01	
	pT2	4940	(21.2)	3761	(41.0)		1629	(35.5)	1517	(33.0)		
Histological tumor size	рТ3	363	(1.5)	351	(3.8)		133	(2.9)	167	(3.6)		
	pT4	45	(0.2)	42	(0.4)		10	(0.2)	22	(0.5)		
	NA	4000	(17.1)	433	(4.7)		292	(6.4)	291	(6.4)		
	0	15885	(68.1)	2937	(32.0)	<0.0001	2218	(48.3)	2207	(48.1)	<0.0001	
	1 to 3	3420	(14.7)	3978	(43.4)		1706	(37.2)	1526	(33.2)		
Involved axillary nodes	4 to 9	826	(3.5)	1519	(16.6)		473	(10.3)	611	(13.3)		
	\geq 10	297	(1.3)	600	(6.6)		193	(4.2)	246	(5.3)		
	NA	2909	(12.5)	131	(1.4)		0	(0)	0	(0)		
	1	7553	(32.4)	1279	(14.0)	<0.0001	929	(20.2)	1001	(21.8)	0.003	
	2	9000	(38.6)	3703	(40.4)		1934	(42.1)	2031	(44.2)		
SBR grade	3	4903	(21.0)	4014	(43.8)		1727	(37.7)	1558	(34.0)		
	NA	1881	(8.0)	169	(1.8)		0	(0)	0	(0)		
	Hormone receptor +	17424	(74.7)	6492	(70.8)	< 0.0001	3349	(75.1)	3446	(75.1)	0.82	
Hormone receptor status	Hormone receptor -	2349	(10.0)	2155	(23.5)		763	(16.7)	779	(17.0)		
	NA	3564	(15.3)	518	(5.7)		378	(8.2)	365	(7.9)		
	1981–1983	2080	(89.1)	251	(2.7)	<0.0001	160	(3.5)	171	(3.7)	<0.0001	
	1984–1988	3630	(15.6)	771	(8.4)		516	(11.2)	493	(10.7)		
	1989–1993	4452	(19.1)	1434	(15.6)		978	(21.3)	950	(20.7)		
Years of diagnosis	1994–1998	4616	(19.8)	1791	(19.5)		956	(20.8)	1060	(23.1)		
	1999–2003	3164	(13.6)	1748	(19.1)		705	(15.3)	818	(17.8)		
	2003-2008	5395	(23.1)	3170	(34.6)		1275	(27.8)	1098	(23.9)		
Propensity score	Mean (sd)	0.28			(0.20)	<0.0001		(0.25)		(0.25)	0.22	

Table 2. Characteristics of the chemotherapy and no chemotherapy groups prior to and after matching via propensity scores.

NA: not assessable. pT1: pathological tumor size \leq 2 cm. pT2: tumor > 2 cm and \leq 5 cm. pT3: > 5 cm. pT4 = extension to skin and/or chest wall. SBR: Scarff Bloom and Richardson. sd: standard deviation.

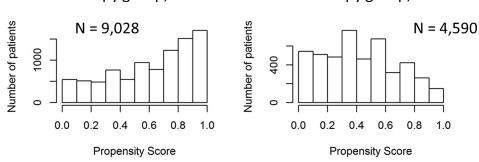
doi:10.1371/journal.pone.0132853.t002

risk of distant metastasis was lower in the patients who were treated with chemotherapy (HR = 0.82, 95% CI [0.75–0.90], p<0.0001), as found in a univariate analysis. At 5 and 10 years of follow-up, the women who were treated with chemotherapy had an absolute benefit of 3.5% with respect to DDFS. In multivariate analysis of the matched population, the HR was 0.82, 95% CI [0.75–0.90] (p<0.0001) (Table 4).

Discussion

An improvement in patient prognosis was observed as a result of the administration of adjuvant chemotherapy in a real-world setting (32,502 women). After patients were matched based



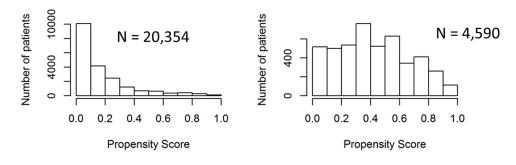


a. Chemotherapy group, raw

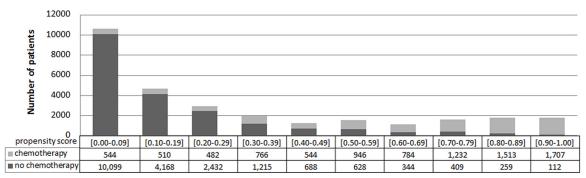
b. Chemotherapy group, matched

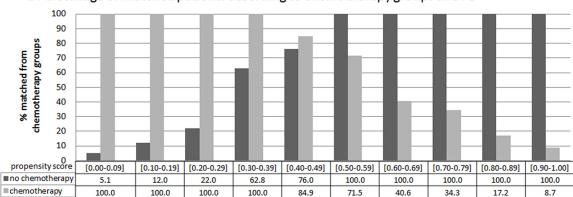


d. No chemotherapy group, matched

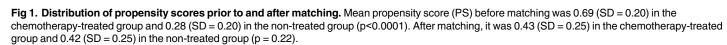


e. Distribution of PS in the whole population





f. Percentage of matched patients according to chemotherapy groups and PS

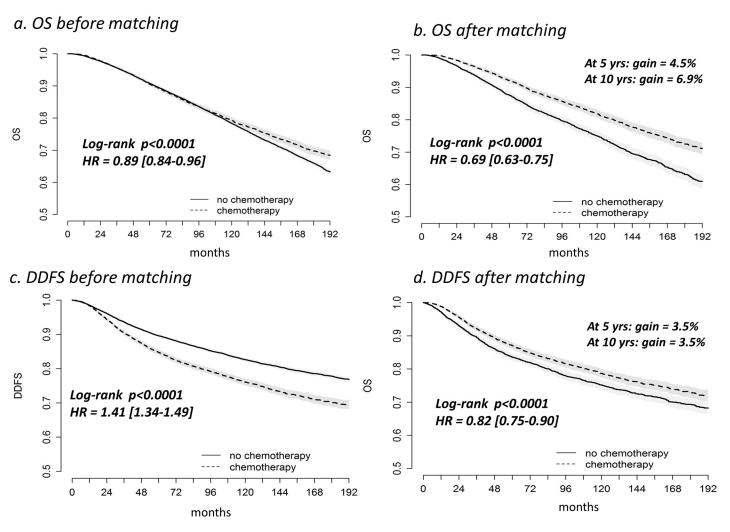


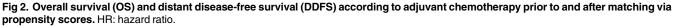
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on their propensity score to receive chemotherapy, the treated group in the period between 1981 and 2008 exhibited a 25% reduction in the relative risk of death compared with the untreated group (HR = 0.75 [0.69–0.83], p<0.0001). This reduction resulted in an absolute survival benefit of 4.5% at 5 years and 6.9% at 10 years. Chemotherapy was also associated with an 18% relative reduction in the risk of distant metastasis (HR = 0.82 [0.75–0.90], p<0.0001). The absolute benefit in terms of DDFS was 3.5% at both 5 and 10 years.

Prior to matching, OS was significantly better in the treatment group; this difference appeared on the survival curve at 10 years of follow-up (Fig 2). This result can be explained by the fact that the group that did not receive chemotherapy included many more patients over 65 years of age as well as more patients from earlier periods. Using death from any cause results in the accumulation of events in older patients. This factor explains the better absolute benefit in





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Table 3. Survival prior to matching: prognosis factors in univariate and multivariate analyses.

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			C	S		DDFS					
		Univariat	e	Multivaria	ate	Univariate		Multivariate			
		HR [CI 95%]	p-value	HR [CI 95%]	p-value	HR [CI 95%]	p-value	HR [CI 95%]	p-value		
	1981–1983	1		1		1		1			
	1984–1988	0.85 [0.80–0.91]	<0.0001	0.89 [0.70–1.13]	0.33	0.89 [0.82–0.97]	0.009	1.01 [0.76–1.34]	0.95		
Veere of diamonia	1989–1993	0.75 [0.71–0.81]	< 0.0001	0.87 [0.69–1.10]	0.25	0.86 [0.79–0.93]	0.0002	0.97 [0.74–1.28]	0.86		
Years of diagnosis	1994–1998	0.67 [0.62–0.72]	<0.0001	0.77 [0.61–0.98]	0.03	0.75 [0.69–0.82]	<0.0001	0.82 [0.62–1.08]	0.16		
	1999–2003	0.51 [0.47–0.56]	<0.0001	0.68 [0.51092]	0.01	0.56 [0.51–0.62]	<0.0001	0.61 [0.44–0.84]	0.0002		
	2004–2008	0.38 [0.35–0.42]	<0.0001	0.42 [0.32–0.55]	<0.0001	0.41 [0.37–0.46]	<0.0001	0.44 [0.32–0.59]	<0.0001		
	< 35	1.47 [1.31–1.65]	<0.0001	1.42 [1.00–2.02]	0.05	2.01 [1.88–2.35]	<0.0001	1.64 [1.24–2.16]	0.0005		
• ()	35 to 49	0.86 [0.82-0.91]	<0.0001	1.15 [0.96–1.37]	0.13	1.21 [1.14–1.27]	<0.0001	1.14 [1.01–1.29]	0.03		
Age (years)	50 to 65	1		1		1		1			
	> 65	2.25 [2.15–2.37]	<0.0001	2.06 [1.85–2.30]	< 0.0001	1.00 [0.94–1.06]	0.99	1.09 [0.9–1.25]	0.17		
Manager and adaptive	post-menopausal	1		1	p-value n 1 0.33 0 13 0.33 0 10 0.25 0 20 0.03 0 21 0.01 0 22 0.001 0 22 0.05 2 377 0.13 1 301 <0.0001	1		1			
Menopausal status	pre-menopausal	0.62 [0.59–0.65]	< 0.0001	0.68 [0.57–0.80]	< 0.0001	1.23 [1.17–1.29]	<0.0001		NS		
	< 18.5	0.96 [0.83–1.11]	0.57	1.25 [1.02–1.52]	0.03	0.92 [0.79–1.09]	0.38	1.16 [0.93–1.45]	0.20		
	[18.5–25[1		1		1		1			
BMI (kg/m²)	[25–30[1.29 [1.20–1.39]	<0.0001	1.13 [1.01–1.25]	0.03	1.18 [1.09–1.27]	<0.0001	1.09 [0.97–1.22]	0.16		
	\geq 30	1.49 [1.35–1.65]	<0.0001	1.22 [1.05–1.41]	0.008	1.31 [1.18–1.46]	<0.0001	1.22 [1.04–1.43]	0.01		
	Ductal	1		1		1		1			
Histological type	Lobular	1.07 [1.00–1.14]	0.04		NS	1.08 [1.00–1.16]	0.04		NS		
	Other	1.05 [0.99–1.10]	0.1		NS	0.95 [0.88–1.01]	0.11		NS		
	pT1	1		1		1		1			
	pT2	2.17[2.06–2.28]	<0.0001	1.47 [1.34–1.63]	<0.0001	2.54 [2.39–2.70]	<0.0001	1.66 [1.49–1.84]	<0.0001		
Histological tumor size	рТ3	3.18 [2.81–3.59]	<0.0001	2.26 [1.83–2.80]	<0.0001	3.94 [3.44–4.50]	<0.0001	2.20 [1.76–2.75]	<0.0001		
	pT4	2.53 [1.61–3.97]	<0.0001	1.65 [0.73–3.73]	0.23	2.93 [1.89–4.56]	<0.0001	Multivaria HR [Cl 95%] 1 1.01 [0.76–1.34] 0.97 [0.74–1.28] 0.82 [0.62–1.08] 0.61 [0.44–0.84] 0.44 [0.32–0.59] 1.64 [1.24–2.16] 1.14 [1.01–1.29] 1 1.09 [0.9–1.25] 1 1.09 [0.9–1.25] 1 1.16 [0.93–1.45] 1 1.09 [0.97–1.22] 1.22 [1.04–1.43] 1 2.20 [1.76–2.75] 2.34 [1.27–4.29] 1 1.47 [1.31–1.67] 2.93 [2.52–3.42] 5.52 [4.54–6.72] 1 1.61 [1.37–1.89] 2.68 [2.31–3.12] 1 1.26 [1.10–1.44] 1	0.006		
	0	1		1	-	1	•	1			
	1 to 3	1.54 [1.47–1.63]	<0.0001	1.22 [1.09–1.36]	0.0006	1.96 [1.85–2.08]	<0.0001	1.47 [1.31–1.67]	<0.0001		
Involved axillary nodes	4 to 9	2.94[2.76-3.13]	<0.0001	2.14 [1.85–2.48]	<0.0001	4.47 [4.18–4.78]	<0.0001	2.93 [2.52–3.42]	<0.0001		
	≥ 10	4.81 [4.43–5.22]	<0.0001	3.38 [2.75–4.15]	<0.0001	7.75 [7.12–8.44]	<0.0001	5.52 [4.54–6.72]	<0.0001		
	1	1		1	-	1	•				
SBR grade	2	1.56 [1.47–1.66]	<0.0001	1.09 [0.96–1.24]	0.20	2.16 [2.00–2.33]	<0.0001	1.61 [1.37–1.89]	<0.0001		
	3	2.41 [2.27–2.56]	<0.0001	1.84 [1.63–2.08]	< 0.0001	3.69 [3.43–3.98]	<0.0001	2.68 [2.31–3.12]	<0.0001		
	Hormone receptor +	1		1		1		1			
Hormone receptor status	Hormone receptor -	1.55 [1.47–1.64]	<0.0001	1.34 [1.18–1.52]	< 0.0001	1.65 [1.56–1.74]	<0.0001	1.26 [1.10–1.44]	0.0007		
• P	No	1		1		1		1.64 [1.24–2.16] 1.14 [1.01–1.29] 1 1.09 [0.9–1.25] 1 1.16 [0.93–1.45] 1 1.09 [0.97–1.22] 1.22 [1.04–1.43] 1 1 1.66 [1.49–1.84] 2.20 [1.76–2.75] 2.34 [1.27–4.29] 1 1.47 [1.31–1.67] 2.93 [2.52–3.42] 5.52 [4.54–6.72] 1 1.61 [1.37–1.89] 2.68 [2.31–3.12] 1 1.26 [1.10–1.44] 1			
Adjuvant chemotherapy	Yes	0.89 [0.84-0.96]	< 0.0001	0.88 [0.78–1.00]	0.05	1.41 [1.34–1.49]	< 0.0001	0.85 [0.74–0.97]	0.01		

OS: overall survival. DDFS: Distant disease-free survival. BMI: body mass index. pT1: pathological tumor size \leq 2 cm. pT2: tumor > 2 cm and \leq 5 cm. pT3: > 5 cm. pT4: extension to skin and/or chest wall. SBR: Scarff Bloom and Richardson. HR: hazard ratio. NS: not significant in multivariate analysis and therefore not included in the final multivariate model.

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terms of OS compared with DDFS in favor of chemotherapy after matching. This absolute benefit could also reflect more co-morbidities in patients who did not receive chemotherapy. Matching via propensity score created more comparable groups; however, co-morbidities were not included in the model and were therefore not taken into account when matching. This propensity score approach reduced indication bias $[\underline{7}-\underline{8}]$, enabling a more accurate estimate of the effect of chemotherapy. We could not take into account the effects of HER2 status, the type of chemotherapy regimen, the type of hormonal therapy, comorbidities or screening programs



Table 4. OS and DDFS after matching: prognosis factors in multivariate analysis.

		OS		DDFS		
		HR [CI 95%]	p-value	HR [CI 95%]	p-value	
	1981–1983	1		1		
	1984–1988	0.92 [0.77–1.09]	0.33	0.98 [0.82–1.18]	0.84	
Verse of diamonais	1989–1993	0.81 [0.68–0.97]	0.020	0.93 [0.78–1.11]	0.43	
Years of diagnosis	1994–1998	0.75 [0.63–0.90]	0.002	0.79 [0.66–0.95]	0.01	
	1999–2003	0.56 [0.46–0.72]	<0.0001	0.64 [0.52–0.80]	<0.0001	
	2004–2008	0.37 [0.29–0.48]	<0.0001	0.46 [0.36–0.58]	<0.0001	
	< 35	1.32 [0.99–1.76]	0.05	1.50 [1.20–1.88]	0.0004	
	35 to 49	0.97 [0.82–1.15]	0.73	1.09 [0.98–1.21]	0.12	
Age (years)	50 to 65	1		1		
	> 65	1.64 [1.45–1.85]	<0.0001	1.04 [0.92–1.18]	0.52	
	post-menopausal	1		1		
Menopausal status	pre-menopausal	0.79 [0.67–0.93]	0.004		NS	
	< 18.5	1		1	NS	
	[18.5–25[NS			
BMI (kg/m²)	[25–30[NS		NS	
	\ge 30		NS		NS	
	Ductal	1		1		
Histological type	Lobular		NS		NS	
	Other		NS		NS	
	pT1	1		1		
	pT2	1.43 [1.29–1.59]	<0.0001	1.59 [1.44–1.77]	<0.0001	
Histological tumor size	рТЗ	1.54 [1.23–1.92]	<0.0001	1.79 [1.46–2.21]	<0.0001	
	рТ4	1.18 [0.44–3.19]	0.74	1.54 [0.69–3.47]	0.29	
	0	1		1		
	1 to 3	1.40 [1.24–1.58]	<0.0001	1.51 [1,34–1.69]	<0.0001	
involved axillary nodes	4 to 9	2.56 [2.22-2.94]	<0.0001	3.02 [2.64-3.46]	<0.0001	
	≥ 10	3.95 [3.34–4.67]	<0.0001	5.41 [4.61–6.35]	<0.0001	
	1	1		1		
SBR grade	2	1.48 [1.27–1.74]	<0.0001	1.73 [1.48–2.03]	<0.0001	
•	3	2.00 [1.71–2.35]	<0.0001	2.34 [1.99–2.74]	<0.0001	
	Hormone receptor +	1		1		
Hormone receptor status	Hormone receptor -	1.29 [1.15–1.45]	<0.0001	1.24 [1.10–1.39]	0.0003	
	No	1		1		
Adjuvant chemotherapy	Yes	0.75 [0.69–0.83]	<0.0001	0.82 [0.75–0.90]	<0.0001	

OS: overall survival. DDFS: Distant disease-free survival. BMI: body mass index. pT1: pathological tumor size \leq 2 cm. pT2: tumor > 2 cm and \leq 5 cm. pT3: > 5 cm. pT4 = extension to skin and/or chest wall. SBR: Scarff Bloom and Richardson. HR: hazard ratio. NS: not significant in multivariate analysis and therefore not included in the final multivariate model.

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because we lacked prospectively collected data for these parameters. Only 1.6% of our entire population received trastuzumab (data not shown).

In 1998, a meta-analysis by the Early Breast Cancer Trialists' Collaborative Group [12] showed that chemotherapy was associated with a relative risk reduction of death that was estimated to be between 8% and 27% depending on the age of the patients and the type of breast cancer (RR = 0.85 in a population of 17,723 patients). The benefit of chemotherapy was independent from menopausal status, lymph node involvement and the use of hormone therapy.

This study included trials primarily testing the CMF (cyclophosphamide, methotrexate, 5-fluorouracil) combination. Thus, adjuvant chemotherapy was more widely administered. Later, anthracyclines [13] and taxanes [14] demonstrated superior efficacy [5] and were integrated into chemotherapy regimens. Substantial differences in age and comorbidities have been reported between patients in randomized clinical trials and real-world registries because of explicit exclusion criteria and subtle recruitment biases [15]. However, there are no data on the impact of chemotherapy on breast cancer survival in the real world, with the exception of a few population-based studies reviewing specific sub-groups such as metastatic cancer, HER2 + tumors, triple negative breast cancer or elderly patients [9] [10] [16]. Our results, which were obtained from an unselected population, are consistent with those from randomized controlled trials. In a real clinical setting, older patients are less likely to receive chemotherapy, similar to randomized trials. Therefore, in a real-world population, these patients decrease the overall prognosis of individuals who do not receive chemotherapy, which may have inflated the difference in overall survival between the chemotherapy-treated and non-treated groups in this study. We consider our population as unselected while our patients were treated in a referral center. The mean age of patients treated the Institut Curie was 57.9 years (S.D. = 12.2), which is younger that mean age of breast cancer in USA or France (61 years) [17] [18]. However, the distribution of tumors' characteristics in those patients treated in a referral center was representative of that observed in the breast cancer population as a whole [19] [20].

OS and DDFS improved over time (Table 3). It is difficult to determine the role of advances in therapeutics and screening (increasing the incidence of tumors discovered at an early stage) in a real-world population. Some studies attempted to estimate the impact of treatments and screening on breast cancer survival. Using 7 different modeling approaches, Berry *et al.* found that screening could explain 28% to 65% of the improvement in survival; 35% to 72% of the improvements were attributable to adjuvant treatments [21]. Another before/after study (1981–1984/1990–1994) estimated that one-third of the improvement in survival may be due to earlier tumors diagnosis, and the remaining two-thirds may be due to the adjuvant treatments [22]. However, the former study was based on models; the second was based on a small sample of patients. Additionally, the authors did not specify the role of chemotherapy alone. Our population-based study assessed the benefit of chemotherapy and confirmed that it provides a DDFS advantage primarily in the first years of follow-up (no additional benefit between 5 and 10 years, Fig.2) [23].

There is no consensus on the optimal method to assess the impact of treatment in a clinical setting. We chose propensity scores to minimize confounding and indication bias between the treated and non-treated groups. Other methods are employed to avoid bias (such as multivariable regression restricted to a high-risk population or the instrumental variable method [24]); however, none of these fully resolve the problem [25]. Our results obtained with the Cox multivariate model and with the propensity score matching method were similar, but the latter method could be used to assess whether the benefit occurs in the first years of treatment or thereafter. However, with the propensity score matching method, most of the patients who are excluded from the analysis exhibit low or high treatment probability. Our results reflect the average benefit of adjuvant chemotherapy in a population; however, this population likely excluded the patients who were unlikely to benefit or would certainly benefit from this treatment. We could not clarify the effect of chemotherapy in breast cancer subgroups because subgroups were not defined over the whole period at the time when patients were treated.

Conclusions

In invasive breast cancer, adjuvant chemotherapy reduced the relative risk of death by 25% and the relative risk of distant metastasis by 18% in this study. After randomized trials, this study

confirms the magnitude of the impact of chemotherapy in a real-world population and demonstrates the level of overtreatment. It advocates for better tumor evaluation and personalized treatment. The effect of chemotherapy in subgroups related to patient characteristics, tumor burden or tumor biological features remains to be clarified.

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Author Contributions

Conceived and designed the experiments: LR DS FR MR BA RR. Performed the experiments: LR DS MR BA RR. Analyzed the data: LR FR MR BA RR. Contributed reagents/materials/analysis tools: LR DS JYP FL FR MR BA RR. Wrote the paper: LR FR MR BA RR.

References

- 1. Héry C, Ferlay J, Boniol M, Autier P. Changes in breast cancer incidence and mortality in middle-aged and elderly women in 28 countries with Caucasian majority populations. Ann Oncol Off J Eur Soc Med Oncol ESMO 2008. 19(5):1009–18.
- Youlden DR, Cramb SM, Dunn NAM, Muller JM, Pyke CM, Baade PD. The descriptive epidemiology of female breast cancer: an international comparison of screening, incidence, survival and mortality. Cancer Epidemiol 2012. 36(3):237–48. doi: 10.1016/j.canep.2012.02.007 PMID: 22459198
- Ban KA, Godellas CV. Epidemiology of breast cancer. Surg Oncol Clin N Am 2014. 23(3):409–22. doi: 10.1016/j.soc.2014.03.011 PMID: 24882341
- Hassan MSU, Ansari J, Spooner D, Hussain SA. Chemotherapy for breast cancer (Review). Oncol Rep 2010. 24(5):1121–31. PMID: <u>20878101</u>
- Early Breast Cancer Trialists' Collaborative Group (EBCTCG), Peto R, Davies C, Godwin J, Gray R, Pan HC, et al. Comparisons between different polychemotherapy regimens for early breast cancer: meta-analyses of long-term outcome among 100,000 women in 123 randomised trials. Lancet 2012. 379(9814):432–44. doi: 10.1016/S0140-6736(11)61625-5 PMID: 22152853
- Sørensen HT, Lash TL, Rothman KJ. Beyond randomized controlled trials: a critical comparison of trials with nonrandomized studies. Hepatol Baltim Md 2006. 44(5):1075–82.
- Drake C, Fisher L. Prognostic models and the propensity score. Int J Epidemiol 1995. 24(1):183–7. PMID: <u>7797341</u>
- Glynn RJ, Schneeweiss S, Stürmer T. Indications for propensity scores and review of their use in pharmacoepidemiology. Basic Clin Pharmacol Toxicol 2006. 98(3):253–9. PMID: <u>16611199</u>
- Olszewski AJ, Migdady Y, Boolbol SK, Klein P, Boachie-Adjei K, Sakr BJ, et al. (2013) Effects of adjuvant chemotherapy in HER2-positive or triple-negative pT1ab breast cancers: a multi-institutional retrospective study. Breast Cancer Res Treat 138(1):215–23. doi: <u>10.1007/s10549-013-2423-3</u> PMID: <u>23354365</u>
- Schneider M, Zuckerman IH, Onukwugha E, Pandya N, Seal B, Gardner J, et al. (2011) Chemotherapy treatment and survival in older women with estrogen receptor-negative metastatic breast cancer: a population-based analysis. J Am Geriatr Soc 59(4):637–46. doi: <u>10.1111/j.1532-5415.2011.03351.x</u> PMID: <u>21453377</u>
- Austin PC (2011) An Introduction to Propensity Score Methods for Reducing the Effects of Confounding in Observational Studies. Multivar Behav Res 46(3):399–424.
- Early Breast Cancer Trialists' Collaborative Group (1998) Polychemotherapy for early breast cancer: an overview of the randomised trials. Lancet 352(9132):930–42. PMID: <u>9752815</u>
- Poole CJ, Earl HM, Hiller L, Dunn JA, Bathers S, Grieve RJ, et al. (2006) Epirubicin and cyclophosphamide, methotrexate, and fluorouracil as adjuvant therapy for early breast cancer. N Engl J Med 355 (18):1851–62. PMID: <u>17079759</u>
- Henderson IC, Berry DA, Demetri GD, Cirrincione CT, Goldstein LJ, Martino S, et al. (2003) Improved outcomes from adding sequential Paclitaxel but not from escalating Doxorubicin dose in an adjuvant chemotherapy regimen for patients with node-positive primary breast cancer. J Clin Oncol Off J Am Soc Clin Oncol 21(6):976–83.

- Kemeny MM, Peterson BL, Kornblith AB, Muss HB, Wheeler J, Levine E, et al. (2003) Barriers to clinical trial participation by older women with breast cancer. J Clin Oncol Off J Am Soc Clin Oncol 21 (12):2268–75.
- 16. Elkin EB, Hurria A, Mitra N, Schrag D, Panageas KS (2006) Adjuvant chemotherapy and survival in older women with hormone receptor-negative breast cancer: assessing outcome in a population-based, observational cohort. J Clin Oncol Off J Am Soc Clin Oncol 24(18):2757–64.
- 17. Cancer of the Breast—SEER Stat Fact Sheets. Available: <u>http://seer.cancer.gov/statfacts/html/breast.</u> <u>html</u>. Accessed 2014 Sep 16.
- 18. Les données—Institut National Du cancer. Available: <u>https://lesdonnees.e-cancer.fr/</u>. Accessed 2014 Apr 9.
- Kwan ML, Kushi LH, Weltzien E, Maring B, Kutner SE, Fulton RS, et al. (2009) Epidemiology of breast cancer subtypes in two prospective cohort studies of breast cancer survivors. Breast Cancer Res BCR 11(3):R31. doi: <u>10.1186/bcr2261</u> PMID: <u>19463150</u>
- 20. SEER Cancer Statistics Review 1975–2001—Previous Version—SEER Cancer Statistics. Available: http://seer.cancer.gov/archive/csr/1975_2001/. Accessed 2014 Sep 16.
- Berry DA, Cronin KA, Plevritis SK, Fryback DG, Clarke L, Zelen M, et al. (2005) Effect of screening and adjuvant therapy on mortality from breast cancer. N Engl J Med 353(17):1784–92. PMID: 16251534
- 22. Webb PM, Cummings MC, Bain CJ, Furnival CM (2004) Changes in survival after breast cancer: improvements in diagnosis or treatment? Breast Edinb Scotl 13(1):7–14.
- Early Breast Cancer Trialists' Collaborative Group (EBCTCG) (2005) Effects of chemotherapy and hormonal therapy for early breast cancer on recurrence and 15-year survival: an overview of the randomised trials. Lancet 365(9472):1687–717. PMID: <u>15894097</u>
- 24. Brookhart MA, Schneeweiss S (2007) Preference-based instrumental variable methods for the estimation of treatment effects: assessing validity and interpreting results. Int J Biostat 3(1):Article 14.
- Bosco JLF, Silliman RA, Thwin SS, Geiger AM, Buist DSM, Prout MN, et al (2010) A most stubborn bias: no adjustment method fully resolves confounding by indication in observational studies. J Clin Epidemiol 63(1):64–74. doi: <u>10.1016/j.jclinepi.2009.03.001</u> PMID: <u>19457638</u>