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Received 17 Jul 2015 | Accepted 29 Nov 2015 | Published 20 Jan 2016

DOI: 10.1038/ncomms10313

OPEN

Quantifying the origin of metallic glass formation

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The waiting time to form a crystal in a unit volume of homogeneous undercooled liquid exhibits a pronounced minimum τ_X^* at a 'nose temperature' T^* located between the glass transition temperature T_g and the crystal melting temperature, T_L . Turnbull argued that τ_X^* should increase rapidly with the dimensionless ratio $t_{rg} = T_g/T_L$. Angell introduced a dimensionless 'fragility parameter', m , to characterize the fall of atomic mobility with temperature above T_g . Both t_{rg} and m are widely thought to play a significant role in determining τ_X^* . Here we survey and assess reported data for T_L , T_g , t_{rg} , m and τ_X^* for a broad range of metallic glasses with widely varying τ_X^* . By analysing this database, we derive a simple empirical expression for $\tau_X^*(t_{rg}, m)$ that depends exponentially on t_{rg} and m , and two fitting parameters. A statistical analysis shows that knowledge of t_{rg} and m alone is therefore sufficient to predict τ_X^* within estimated experimental errors. Surprisingly, the liquid/crystal interfacial free energy does not appear in this expression for τ_X^* .

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It is widely believed that any liquid will form a glass if cooled sufficiently rapidly. Even elemental liquid metals can be vitrified^{1,2} if quenched to low temperature at ultrahigh cooling rates of 10^{12} – 10^{14} K s⁻¹. At the opposite extreme are certain molten eutectic alloys that form bulk metallic glass at cooling rates of ~ 1 K s⁻¹ or less^{3,4}. The glass forming ability (GFA) of a liquid is defined by the temperature-dependent waiting time, $\tau_X(T)$, for a detectable fraction of crystal(s) to nucleate and grow in a unit volume of liquid undercooled to a temperature $T < T_L$, where T_L is the melting temperature or more specifically the liquidus temperature of an alloy^{5–8}. Far below T_L , liquids undergo configurational freezing at the glass transition temperature T_g and crystal nucleation is kinetically arrested. Between T_g and T_L , the $\tau_X(T)$ -curve, or so-called time–temperature–transformation diagram (TTT diagram), exhibits a sharp minimum τ_X^* at an experimentally measurable nose temperature T^* (refs 5–7). To form glass, the liquid must roughly be cooled from T_L to below T^* in a time less than τ_X^* . This defines a critical cooling rate $R_C \sim [T_L - T^*]/\tau_X^*$ that must be exceeded to avoid crystallization. Transient cooling, or quenching, is often governed by time-dependent heat conduction. Heat is extracted from the external sample boundary; and the cooling history is a function of location within the sample, being highest near the surface and lowest near the sample center of symmetry (for example, along the centerline of a rod or mid-plane of a plate). For a uniform shape with characteristic sample dimension d (for instance, a rod diameter or plate thickness), a Fourier time scale or thermal relaxation time can be defined as $\tau_Q \sim d^2/D_t$, where D_t is the liquid thermal diffusivity and is roughly constant among the various metallic glass alloy compositions (typically 2 – 4 mm² s⁻¹). This time scale characterizes cooling at the center of symmetry. The proportionality constant depends on the geometry. Requiring $\tau_Q < \tau_X^*$ implies a maximum sample dimension d_{\max} for forming a glass referred to as the critical casting thickness of the alloy. The parameters τ_X^* , R_C and d_{\max} are interchangeably used in the literature as alternative measures of GFA.

Turnbull and others^{5–8} showed generally that the nucleation rate can be expressed as the product of a pre-factor and two thermally activated rate factors:

$$\tau_X^{-1} = \nu \exp\left[\frac{-W(T)}{kT}\right] \exp\left[\frac{-\Delta G(T)}{kT}\right], \quad (1)$$

where the pre-factor ν is taken to be a typical atomic vibration frequency. The first exponential factor is the thermally activated atomic rearrangement rate in the liquid. It describes atomic mobility, fluidity or inverse viscosity $\eta^{-1} \sim \exp[-W(T)/kT]$. The atomic rearrangement barrier $W(T)$ falls with increasing T above T_g . To characterize the rate of fall of $W(T)$, Angell introduced the concept of liquid fragility that he quantified using a fragility parameter, m (refs 9,10). Viscosity is modelled by various empirical laws^{11–15}. The most common is the Vogel-Fulcher-Tamann law, though recently a more reliable law based on the cooperative shear flow model has been introduced that is directly relatable to the definition of m (refs 13,14) (Supplementary Information).

The second factor $\exp[-\Delta G/kT]$ in equation (1) is the probability of a fluctuation leading to the formation of a critical crystalline nucleus. In Classical Nucleation Theory (CNT), $\Delta G(T)$, arises from the excess free energy per unit area, γ_{XL} , required to form the liquid/crystal interface^{5–8}. As the temperature T of an undercooled liquid increases approaching T_L , $\Delta G(T)$ increases rapidly with T and diverges to infinity as $(T_L - T)^{-2}$ so that the nucleation rate becomes immeasurably small^{5,6} near T_L . The temperature dependence of the sum $W(T) + \Delta G(T)$ is dominated by the rapid drop of $W(T)$ for

$T > T_g$, and by the divergence of ΔG as $T \rightarrow T_L$. This sum therefore exhibits a sharp minimum at some intermediate temperature T^* , where $\tau_X(T^*)$ exhibits a pronounced minimum. CNT, therefore, correctly predicts the expected ‘C’-shaped TTT diagram.

Turnbull⁵ suggested that if $t_{rg} > 2/3$, then τ_X^* should exceed typical experimental time scales (for example, 1 – 10^3 s) thereby resulting in easy glass formation. Metallic glasses were first synthesized by rapid quenching (cooling rates $\sim 10^6$ K s⁻¹) of low-melting eutectic Au–Si and Pd–Si alloys by Duwez and colleagues^{16,17}. For a eutectic alloy^{5,6,18}, the composition-dependent T_L curve forms a cusp-like minimum at the eutectic composition and t_{rg} exhibits a corresponding cusp-like maximum. For a particularly low melting ternary near-eutectic alloy, Pd₄₀Ni₄₀P₂₀, with $t_{rg} \approx 0.6$, Turnbull’s group used fluxing methods to purify the liquid and cast bulk glassy ingots with $d_{\max} \approx 1$ cm (refs 19,20).

It has been argued that strong liquids (with low m) should exhibit greater GFA than fragile liquids (with high m values)^{21–24}. For example, Mukherjee *et al.* demonstrated that within a limited group of Zr-based alloys, the variation in viscosity at the measured T^* of the TTT diagram scaled with the measured variation in τ_X^* (ref. 22). Senkov attempted to quantify the role of m (ref. 23) in GFA of metallic glasses by assuming τ_X^* to be proportional to liquid viscosity at T^* . Estimating T^* to be roughly the average of T_g and T_L , he expressed $\eta(T^*)$, in terms of t_{rg} and m and further assumed that crystal nucleation times for metallic glasses are simply proportional to $\eta(T^*)$ or equivalently to a parameter $F_l = 2[(m/16)(t_{rg}^{-1} - 1) + 2]^{-1}$, introduced by Senkov to estimate $\eta(T^*)$. Na *et al.*²⁴ recently reported detailed maps of d_{\max} (or equivalently τ_X^*) as a function of composition c of a five-component Ni–Cr–Nb–P–B alloy where c represented a vector composition variation near a eutectic composition. They demonstrated that these maps could be quantitatively understood in terms of measured composition variations $t_{rg}(c)$ and $m(c)$. They suggested that their analysis might be generalized to other alloys²⁴.

In the present work, we conduct a broad survey and critical assessment of the published literature on metallic glasses. We have compiled a database that includes values of d_{\max} , τ_X^* , T_g , T_L , t_{rg} , and fragility parameter m for broad range of metallic glasses with widely varying GFA. Based on an analysis of this database, we introduce a simple empirical expression that assumes $\log(\tau_X^*)$ to be a bilinear function of t_{rg} and m . Fitting this expression to our database captures the systematics of τ_X^* in real metallic glasses within estimated experimental uncertainties in the relevant parameters alone. Our result shows that the maximum crystallization rate in undercooled metallic liquids can be rationalized using a remarkably simple picture. We discuss the implications of this finding in the context of traditional nucleation theory.

Results

Compiled database for metallic glasses. Supplementary Table 1 contains a summary of experimental data for the critical casting diameter, d_{\max} , along with data for τ_X^* , T_g , T_L , t_{rg} and m , for ~ 40 diverse but well-characterized metallic glass-forming alloys that exhibit widely varying GFA. The tabulated data are obtained from a comprehensive and critical assessment of published literature on metallic glasses. References and details of methods used to evaluate published data are provided in the Supplementary Information together with a discussion of experimental errors in the data. The alloys included in this database were selected based on the availability of reliable, reproducible and consistently determined values for GFA and the other relevant parameters. The criteria used to assess the reported experimental data are described. The discussion includes a comparison of alternate

experimental measures of GFA, for example, τ_X^* versus d_{\max}^2 . A standardized approach was used to determine consistent values for t_{rg} and m from calorimetric and rheological data. For instance, the tabulated values of T_g are based on the rheological definition of the glass transition, $\eta(T_g) = 10^{12}$ Pa-s. Determining reliable m values requires equilibrium liquid viscosity data in the vicinity of the rheological T_g (typically between $\eta \sim 10^7$ and 10^{13} Pa s⁻¹). Data at high temperatures (near or above T_L where $\eta \sim 0.001$ – 1 Pa s⁻¹) were used where available and combined with the low temperature data to obtain a best value of m for each alloy. Viscosity analysis based solely on high-temperature data is found to yield consistently large errors in and overestimates of m (compared with low temperature data) and is, therefore, not used for quantitative analysis (as discussed in the Supplementary Information).

Correlation of GFA with Turnbull's t_{rg} or Angell's m alone.

The correlation between Turnbull's t_{rg} and GFA (defined by either τ_X^* or d_{\max}^2), is illustrated by plotting either $\log(\tau_X^*)$ or $\log(d_{\max}^2)$ versus t_{rg} for the database (Supplementary Table I in Supplementary Information). This is shown in Fig. 1. The plot shows an expected trend that can be described by a linear regression with an average slope of $A_t = d[\log(d_{\max}^2)]/dt_{rg} = 28.5$, as illustrated. The linear form implies an exponential dependence of τ_X^* or equivalently d_{\max}^2 on t_{rg} of the form $\tau_X^* \sim \exp(A_t t_{rg})$, as might be anticipated from equation (1). While a trend is clear, the scatter in the plot is large. The relative scatter is quantified by the coefficient of determination $R^2 = 0.598$ for the linear fit. The assumed linear correlation accounts for $\sim 60\%$ of the total variance in the $\log(d_{\max}^2)$ values, or equivalently the $\log(\tau_X^*)$ values. The mean square misfit (per data point) is ~ 0.96 yielding a standard error or s.e. of ± 0.98 in $\log(d_{\max}^2)$. Apparently, t_{rg} alone predicts τ_X^* within roughly plus/minus one order of magnitude. This uncertainty should be contrasted with the overall variation of τ_X^* over ~ 7 – 8 orders of magnitude over the database. One concludes that t_{rg} , while useful, is an inadequate quantitative predictor of GFA.

Figure 2 plots the variation of $\log(d_{\max}^2)$, or equivalently $\log(\tau_X^*)$, versus Angell's parameter m for the systems in Supplementary Table I. As was the case for Fig. 1, a trend is clearly visible. Linear regression yields an average slope $A_m = -0.0572$ with a coefficient of determination of $R^2 = 0.458$. The correlation of GFA with m alone accounts for $\sim 46\%$ of the variance in the experimental values of $\log(d_{\max}^2)$. The standard

error of the misfit in $\log(d_{\max}^2)$ is ~ 1.26 , similar to but somewhat greater than that obtained from the correlation with t_{rg} in Fig. 1. Angell's parameter alone actually correlates with observed GFA nearly as well as t_{rg} alone. Summarizing, either m or t_{rg} alone can be used to predict τ_X^* within roughly plus/minus one order of magnitude. The two parameters, t_{rg} and m , are apparently of roughly equal utility in predicting GFA. For comparison with Figs 1 and 2, we have included a plot of Senkov's F_I versus $\log(d_{\max}^2)$ for our database. This is shown in the Supplementary Fig. 2 for reference. Senkov's parameter depends on both t_{rg} and m and provides an improved correlation with experimental GFA compared with either Figs 1 or 2. A linear regression gives $R^2 = 0.879$, a significantly improved description of GFA than provided by either t_{rg} or m alone. As will shortly be seen, one can do much better without making arbitrary assumptions regarding the location of T^* with respect to T_g and T_L .

A bilinear expression for $\log(\tau_X^*)$ in terms of both t_{rg} and m . In the work of Na *et al.*²⁴, it was demonstrated that the GFA-composition variation around a deep eutectic composition takes the form of exponential hyper-cusps in the four-dimensional composition space of the five-component Ni-Cr-Nb-P-B alloy. Using the measured composition dependences of t_{rg} and m , they introduced a bilinear expression for $\ln[\tau_X^*(t_{rg}, m)]$ to interpret the experimental GFA-composition variation. The fitting parameters $A_t = d[\ln(d_{\max}^2)]/dt_{rg}$ and $A_m = d[\ln(d_{\max}^2)]/dm$ were taken as characteristic of the Ni-Cr-Nb-P-B alloy system. Their analysis yielded values $A_t = 89 \pm 20\%$ and $A_m \approx -0.2 \pm 40\%$. (errors are estimated). The authors noted that their expression for $\tau_X^*(t_{rg}, m)$ might be more generally applicable. The similarity of their A_t with that obtained from the linear fit in Fig. 1, $A_t = 28.5 \times \ln(10) \sim 66$ of is suggestive (Note, the factor of $\ln(10)$ arises from the present use of the $\log(d_{\max}^2)$ versus $\ln(d_{\max}^2)$ in ref. 24).

Following ref. 24, we shall assume that $\log(\tau_X^*)$, or equivalently $\log(d_{\max}^2)$, is a bilinear function of the two independent variables t_{rg} and m . Whereas Na *et al.* applied their equation to near-eutectic alloys of a single system (Ni-Cr-Nb-P-B), we apply it to all glass forming alloys in our database. We assume, quite generally, that $\log(d_{\max}^2)$ is some continuous and differentiable function of the independent variables t_{rg} and m . In principle, this function might also depend on other independent material parameters besides t_{rg} and m . For instance, the interfacial free energy of a liquid/crystal interface, γ_{XL} , is a natural third

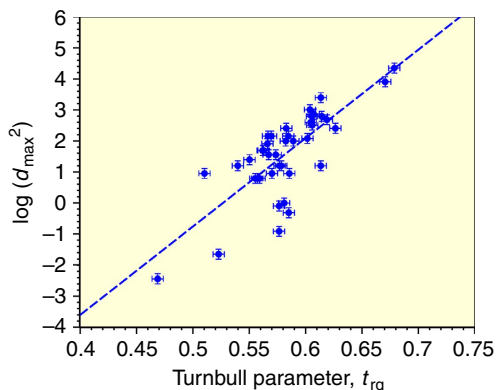


Figure 1 | Plot of $\log(d_{\max}^2)$ versus the dimensionless parameter t_{rg} .

Dashed line is the result of a linear regression least squares fit with an average slope of $A_t = 28.5$. The coefficient of determination for the linear fit is R^2 is 0.598. Error bars are estimated experimental errors as described in the text.

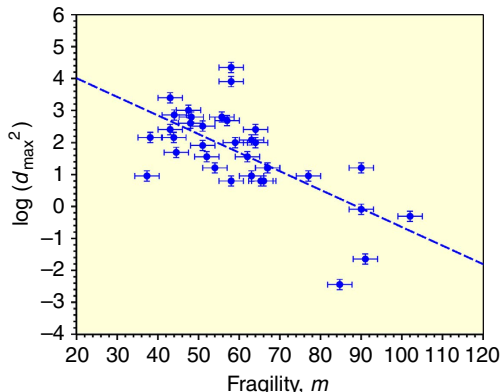


Figure 2 | Plot of $\log(d_{\max}^2)$ as a function of Angell's fragility parameter m .

Dashed line is the result of a linear regression least squares fit with an average slope of $A_m = -0.0572$. The coefficient of determination R^2 is 0.458. Error bars are estimated experimental errors in fragility and casting diameter as described in the text.

parameter arising in CNT. Here we test the assumption that $\log(d_{\max}^2)$ is a universal function of t_{rg} and m alone. Consider the expansion of $\log(d_{\max}^2)$ in a Taylor series around some reference values ($t_{\text{rg},0}$, m_0). To lowest order, one has:

$$\log(d_{\max}^2) = \log(d_{\max,0}^2) + \Lambda_t t_{\text{rg}} + \Lambda_m m + [\text{higher order terms}], \quad (2)$$

where the reference values have, without loss of generality, been set equal to zero, and the higher order terms are assumed to be small relative to the leading linear terms.

Fitting the data in Supplementary Table I using equation (2) and retaining only the linear terms in the Taylor series, we obtain best values for the fitting parameters $\log(d_{\max,0}^2) = -10.36$, $\Lambda_t = 25.6$ and $\Lambda_m = -0.0481$. The quality of the fit is displayed by plotting the experimental data for $\log(d_{\max}^2)$ versus the optimized prediction, $\log(d_{\text{calc}}^2)$, of equation (2) as shown in Fig. 3. The coefficient of determination for this fit is $R^2 = 0.980$. Equation (2) accounts for a remarkable $\sim 98\%$ of the variance in $\log(d_{\max}^2)$. Using both t_{rg} and m versus either alone increases R^2 from $\sim 0.46/0.60$ to 0.980 . This statistically compelling result establishes the relevance of both parameters and yields a useful quantitative prediction of GFA. Experimental uncertainties in d_{\max} , t_{rg} and m can be estimated. The errors in these values are taken to be random and will contribute to the observed misfit between the model prediction of equation (2) and the experimental GFA data. Uncertainties of $\sigma_{d_{\max}}/d_{\max} \sim 0.15$, $\sigma_t \sim 0.006$ and $\sigma_m \sim 3$ are estimated to be representative errors in the experimental determination of d_{\max} , t_{rg} and m , respectively, for the present database. The basis for these error estimates is discussed in the Supplementary Information. An analysis of the variance σ^2 for the misfit between $\log(d_{\text{calc}}^2)$ and $\log(d_{\max}^2)$ contributed by these estimated experimental errors yields:

$$\begin{aligned} \sigma^2 &= \sigma_{d_{\max}}^2 \left[\frac{\partial(\log(d_{\max}^2))}{\partial(d_{\max})} \right]^2 + \sigma_m^2 \left[\frac{\partial \log(d_{\text{calc}}^2)}{\partial m} \right]^2 + \sigma_t^2 \left[\frac{\partial \log(d_{\text{calc}}^2)}{\partial t_{\text{rg}}} \right]^2 \\ &= \left[\frac{\sigma_{d_{\max}}^2}{d_{\max}^2} + \Lambda_m^2 \sigma_m^2 + \Lambda_t^2 \sigma_t^2 \right] \approx 0.0225 + 0.0236 + 0.0208 \\ &= 0.0669. \end{aligned} \quad (3)$$

The estimated experimental errors yield a s.e. of $\sigma \approx 0.26$ for the misfit between the experimental $\log(d_{\max}^2)$ and the predicted

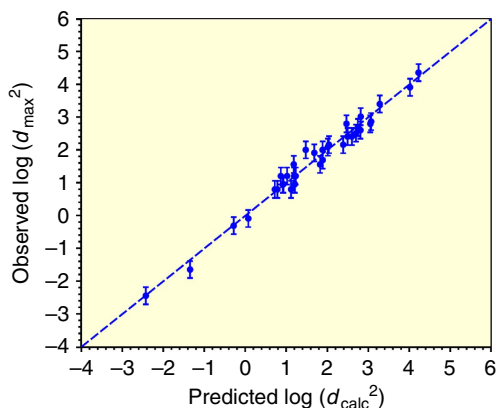


Figure 3 | Correlation between observed $\log(d_{\max}^2)$ and predicted $\log(d_{\text{calc}}^2)$. Values of $\log(d_{\text{calc}}^2)$ are obtained from t_{rg} and m using equation (2). Dashed line is the result of a linear regression least squares fit with fitting parameters $\log(d_{\max,0}^2) = -10.36$, $\Lambda_t = 25.6$ and $\Lambda_m = 0.0481$. The coefficient of determination for the linear fit is R^2 is 0.980 . Error bars given by equation (3) in text.

$\log(d_{\text{calc}}^2)$. The error bars displayed in Fig. 3 were chosen to have this value.

Discussion

The actual variance of the misfit in Fig. 3 gives a s.d. of ~ 0.23 , very close to that expected from experimental error alone (0.26). Adding a third input parameter such as γ_{XL} to the expression for GFA will not be useful since experimental errors will mask any improvement in the prediction of GFA versus the simple two-parameter model of equation (2). In other words, the prediction of a better model could not be empirically distinguished from that of equation (2).

From a modelling perspective, we have assumed t_{rg} and m to be uncorrelated independent variables. Analysing an m versus t_{rg} plot for our database can test this assumption. A weak anti-correlation (alloys with higher t_{rg} have statistically slightly lower m) with corresponding coefficient of determination of $R^2 \sim 0.08$ is obtained from such a plot. For the present database, t_{rg} and m are apparently very weakly correlated, if at all. One consequence of a weak anti-correlation will be a small systematic over-estimate of the magnitude of Λ -parameters obtained in Figs 1 and 2 versus those obtained from the bilinear fit of Fig. 3. For the bilinear fit, we obtained $\Lambda_t \approx 25.6$ and $\Lambda_m \approx -0.0481$. For the single parameter analysis of Figs 1 and 2, we obtained $\Lambda_t \approx 28.5$ and $\Lambda_m \approx -0.057$.

In summary, we have demonstrated that knowledge of t_{rg} and m is sufficient to predict GFA within experimental uncertainties in the relevant variables. This result is unexpected and must be rationalized in the context of traditional nucleation theory. We proceed to discuss the implications of this result and to examine the practical utility of equation (2).

Perhaps the most practical result of the present work lies in clarifying the relative roles played by Turnbull's parameter and Angel's fragility concept in achieving very high ($d_{\max} > 1$ cm) GFA. While bulk glass formation has commonly been associated with low melting eutectic alloys, it is the combination of a deep eutectic with strong-liquid rheology that underlies superior glass-forming ability. This is best illustrated by quantifying the relative contributions of t_{rg} and m to the GFA of superior eutectic glass formers. As an example, consider the Pd-Ni(Cu)-P alloy system. This system includes a ternary Pd-Ni-P bulk glass forming alloy with $d_{\max} \sim 3$ cm (see Supplementary Table I) and a related quaternary alloy of Pd-Ni-Cu-P with observed $d_{\max} > 8$ cm (refs 25–30). The simple binary eutectic alloys $\text{Ni}_{81}\text{P}_{19}$ and $\text{Pd}_{81}\text{P}_{19}$ ($T_L = 1,143$ and $1,044$ K, respectively) form rapidly quenched glassy ribbons of thickness ~ 50 μm . The equivalent rod diameter is $d_{\max} \sim 140$ μm (Supplementary Information). Ternary alloys along the $(\text{Pd}_{1-x}\text{Ni}_x)_{80}\text{P}_{20}$ composition line comprise a pseudobinary eutectic system forming bulk glass rods at $x = 0.5$ having an observed $d_{\max} > 25$ mm (refs 25,26). As a function of composition x , d_{\max}^2 exhibits a broad maximum that varies by 4–5 orders of magnitude along the pseudobinary line^{25,26,28,30}. Chen^{31,32} carried out thermal characterization and systematic creep studies for glasses along this line and reported highly reliable viscosity data well above T_g . Accurate m values can be obtained using his Vogel-Fulcher-Tamann viscosity fits (see Supplementary Table I and discussion in Supplementary Methods). Using equation (2), one may separate the enhancement of GFA from increasing t_{rg} from that attributable to decreasing m . The logarithmic form of equation (2) means that the variation of d_{\max}^2 with x is a product of two factors, attributable respectively to the variation in t_{rg} and m (see Supplementary Methods and Supplementary Figs 3–5 for details of the t_{rg} and m variation with x). Figure 4 illustrates this separation of factors and demonstrates that t_{rg} and m play

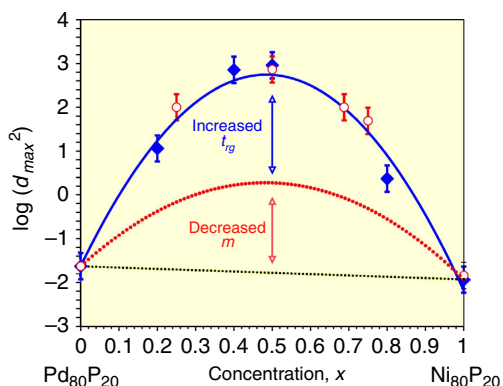


Figure 4 | GFA versus composition diagram for the pseudobinary

(Pd_{1-x}Ni_x)₈₀P₂₀ system. Input data are taken from refs 25–32. Solid blue square symbols are calculated GFA from experimental t_{rg} and m values^{31,32} using equation (2). The solid blue line is a fit to the calculated GFA values. Dotted red line is a parabolic fit to the calculated contribution from fragility. Red open circles are experimental GFA data from refs 25–30. The reader is referred to the Supplementary Information for details. Error bars represent experimental uncertainty of 15% in the maximum casting diameter as discussed in text.

roughly comparable roles in maximizing $\log(d_{max}^2)$ at $x = 0.5$ as one traverses the pseudobinary series. The nucleation nose time τ_x^* of the ternary (Pd_{0.5}Ni_{0.5})₈₀P₂₀ exceeds that of either binary alloy by over four orders of magnitude! Two orders of magnitude are attributable to an increase in t_{rg} , while another two orders of magnitude arise from a decrease in m . Interestingly, the addition of Cu to the ternary alloy (see entry No. 26 in Supplementary Table I) to obtain the precise quaternary eutectic alloy Pd_{42.5}Cu₃₀Ni_{7.5}P₂₀ increases d_{max}^2 by an additional factor of ~ 15 yielding the best glass forming alloy known. The quaternary eutectic alloy actually has a larger m (58 versus 48) than the ternary, but a substantially higher t_{rg} (0.679 versus 0.589) that more than offsets the higher fragility and is responsible for the elevated GFA of the quaternary eutectic versus the ternary alloy. By applying such analysis, it can be generally shown that superior glass formers (for example, Vitreloy 1 with $d_{max} \sim 4\text{--}5$ cm, or Mg–Gd–Ag–Cu–Al with $d_{max} \sim 3$ cm), owe their elevated GFA (compared with the simpler binary or ternary basis alloys) to a confluence of two roughly comparable multiplicative factors arising from low-lying eutectic melting and strong-liquid rheology, respectively. This type of analysis provides a quantitative means to rationalize the underlying factors responsible for the variation of GFA in complex multi-component systems.

The success of equation (2) in predicting GFA over a broad range of metallic alloys would not be possible if heterogeneous nucleation effects played a substantial role in limiting experimentally measured GFA. Heterogeneous nucleants such as oxide inclusions ought to exhibit widely varying catalytic potency that depends on their size, crystal structure and effectiveness as a template for nucleation of each competing crystalline phase. Large variations in catalytic activity would mask any systematic dependence of intrinsic GFA on t_{rg} and m . Apparently, common experimental measures used to suppress heterogeneous nucleation are relatively effective. These measures include: (1) using high purity starting materials^{11,25}, (2) melt overheating above the oxide phase liquidus temperature to dissolve oxide inclusions in the liquid^{33–40}, (3) use of non-crystalline containers¹¹, (4) container-less processing combined with overheating^{33–39} and (5) fluxing methods to remove oxide inclusions^{11,20,21,25,26,40,41}. Apparently, these

measures are effective in achieving near homogeneous nucleation conditions during melt undercooling and glass formation.

The absence of the liquid/crystal interfacial energy in equation (2) is unexpected and raises fundamental questions. Assuming that the crystal nucleation rate controls GFA, then what determines the nucleation barrier $\Delta G(T^*)$ in equation (1) at temperature T^* ? Our result might be understood if $\Delta G(T^*)$ were uniquely determined by t_{rg} and m . In the context of CNT, this might imply that γ_{XL} is some unique function of t_{rg} and m and thus already implicitly included in equation (2). This seems somewhat implausible since γ_{XL} should depend on the crystal structure of the nucleating phase, its composition (and that of the parent liquid), temperature and so on. Alternatively, if transient nucleation is important, then the incubation time to establish a steady state population of crystalline embryos may control GFA. This time might be determined by t_{rg} and m . The incubation time should be related to the time required for development of local chemical fluctuations on a spatial scale comparable to that of the critical nucleus. For instance, the early work of Borelius *et al.*⁴² suggests that the free energy cost of chemical fluctuations in multicomponent alloys might dominate the topological contribution to the nucleation barrier. Other recent work by Desre *et al.*^{43,44} and the work of Wu (refs 45,46) emphasize the role of composition fluctuations in crystal nucleation for multicomponent liquids. A nucleation rate limited by such chemical fluctuations might explain the success of equation (2). However, there remains unanswered the question of what determines the relevant spatial scale of the critical fluctuation, that is, the critical nucleus size?

The transformation from a liquid to a fully crystallized solid generally involves both the crystal nucleation rate and crystal growth velocity^{47,48}. A sluggish nucleation process is a sufficient, but not necessary condition for glass formation. Orava and Greer⁴⁹ have argued that sluggish crystal growth is the likely rate-limiting factor that controls the apparent GFA of many silicates and other molecular glasses. As noted by Greer⁵⁰, growth velocities are independent of γ_{XL} . Thus, growth controlled crystallization would explain the absence of a role played by interfacial energy in determining GFA. However, the analysis of ref. 49 included only one metallic system, a low-melting Cu₅₀Zr₅₀ alloy with $t_{rg} = 0.55$, $m = 58$, $d_{max} = 2$ mm and corresponding $\tau_x^* \sim 0.04$ s (Supplementary Table I). This alloy crystallizes to a single crystalline phase having the same composition as the parent liquid, that is, by polymorphic crystallization, not typical of most metallic glass formers. Recent measurements of the multiphase eutectic growth velocity near T^* in several easy glass formers by time-resolved infrared imaging⁵¹ give relatively high velocities ranging from several cm s^{-1} for Zr-based bulk glasses up to several m s^{-1} for high-GFA eutectic Ni-alloys (see ref. 51 and also J.P. Schramm, G. Kaltenboeck, M.D. Demetriou, and W.L. Johnson, manuscript in preparation). Such high-growth velocities indicate that the observed GFA of these eutectic glass formers must be determined by nucleation rates. The case of elemental metals alluded to in the introduction^{1,2} is of interest in this context. Zhong *et al.*¹ experimentally demonstrated that cooling rates of 10^{14}K s^{-1} are sufficient to suppress crystal growth and produce glass during rapid transient melting of crystalline Ta and V nano-bridges. In this case, the crystalline phase is present, but cannot regrow due to ultra-rapid transient cooling. By contrast, An *et al.*² showed that lower cooling rates of $\sim 10^{12} \text{K s}^{-1}$ are sufficient to form glass in simulations where nano-droplets of liquid Cu are quenched onto an amorphous substrate (with no crystalline nucleus present). Further investigation is clearly required to clarify the relative roles of nucleation versus growth in specific cases.

The glass-forming ability of metallic alloys is alternatively defined by either the crystallization nose time of the TTT diagram or by a critical casting thickness for avoiding detectable crystallinity. We have compiled and critically assessed available data for a diverse set of ~ 40 of metallic glass-forming alloy systems where reliable experimental GFA, liquid rheology and thermodynamic data are available. This database is used to develop a universal expression that quantitatively predicts GFA based only on Turnbull's parameter t_{rg} and Angell's liquid fragility parameter m as independent variables. The analysis yields two fitting parameters Λ_t and Λ_m characterizing the intrinsic variation of GFA with t_{rg} and m , respectively. The expression quantitatively predicts the critical casting thickness or nucleation nose time for forming glassy alloys with the uncertainty expected from experimental errors in the relevant parameters alone. We interpret this universal GFA expression as a description of the crystal nucleation rate in the undercooled liquid state. Evidence suggests that for typical multicomponent eutectic glass formers, crystal growth velocities (specifically eutectic growth velocities) at temperatures near T^* are sufficiently high to justify neglecting the influence of sluggish growth in limiting observed GFA. From the present results, one may conclude that:

The conditions for homogeneous nucleation must be generally approached in experimental studies where common practices are employed to suppress heterogeneous nucleation. The crystal-liquid interfacial energy does not explicitly appear to play a significant role in determining GFA. A relatively simple theory of glass formation and crystallization ought to be possible for metallic systems. Traditional nucleation theory offers no obvious explanation for why this turns out to be the case.

Methods

Metallic glass database construction. The present article is based on a broad survey and critical assessment of published experimental data on metallic glasses. This assessed data have been compiled into a metallic glass database that is presented as Supplementary Table I. The database includes selected data for 42 separate alloys. The compiled data includes values of the rheological glass transition temperature T_g , alloy liquidus temperature T_L , reduced glass transition temperature $t_{rg} = T_g/T_L$, Angell fragility parameter m , critical casting diameter d_{max} of a metallic glass rod, the calculated critical casting thickness d_{calc} based on equation (2) in the text, the estimated nucleation nose time τ_X^* based on a scaling relation with d_{max} , and an experimentally measured nucleation nose time τ_{X^*TTT} obtained directly from a measured TTT diagram. The detailed assessment of all parameters is described in detail in the Supplementary Information.

The selection of the alloy entries in the metallic glass database was determined by the availability of consistent and reliable data for the various parameters. For instance, values of T_g in the database are based on the rheological definition of the glass transition, $\eta(T_g) = 10^{12}$ Pa-s. This requires the availability of equilibrium liquid viscosity data in the vicinity of the glass transition. Determination of Angell's m parameter requires accurate equilibrium viscosity data over a typical range from 10^8 to 10^{13} Pa s $^{-1}$ surrounding the glass transition. Determination of d_{max} for a given alloy requires systematic and controlled quenching experiments. Direct determination of τ_X^* requires a measured experimental TTT diagram. Such diagrams are available only for a limited number of alloy systems that exhibit sufficiently large d_{max} . Based on this limited number of systems, an empirical relationship is established between τ_X^* and d_{max} .

The Supplementary Information includes discusses the methods and criteria used to assess each parameter in the database. In addition, uncertainty in each parameter arising from experimental error and the propagation of experimental errors in the determination of the parameter are discussed. The estimated errors in the parameters have been used in the main article to assess the statistical significance of the correlations developed in the paper. For example, the reader is referred to the Supplementary Information for a discussion of the analysis of variance presented in equation (3) of the main article.

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Acknowledgements

W.L.J. and M.D.D. acknowledge Glassmetal Technology, Inc., for financial support of portions of this research activity carried out at Caltech.

Author contributions

W.L.J. participated in compiling the database and in carrying out analysis of the data. W.L.J. wrote the manuscript. J.H.N. and M.D.D. both participated in compiling the database, analysing the data and editing of the manuscript.

Additional information

Supplementary Information accompanies this paper at <http://www.nature.com/naturecommunications>

Competing financial interests: All the authors have an ownership interest in Glassmetal Technology, Inc.

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How to cite this article: Johnson, W. L. *et al.* Quantifying the origin of metallic glass formation. *Nat. Commun.* **7**:10313 doi: 10.1038/ncomms10313 (2016).



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