#### ORIGINAL RESEARCH

# Risks and Clinical Predictors of Hepatocellular Carcinoma in Chinese Populations: A Real-World Study of 10,359 Patients in Six Medical Centers

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**Purpose:** Early detection of hepatocellular carcinoma (HCC) through surveillance could reduce this cancer-associated mortality. We aimed to develop and validate algorithms using panel serum biomarkers to identify HCC in a real-world multi-center study in China. **Patients and Methods:** A total of 10,359 eligible subjects, including HCCs and benign liver diseases (BLDs), were recruited from six Chinese medical centers. The three nomograms were built using logistic regression and their sensitivities and specificities were carefully assessed in training and validation cohorts. HCC patients after surgical resection were followed to determine the prognostic values of these algorithms. Prospective surveillance performance was assessed in a cohort of chronic hepatitis B patients during 144 weeks follow-up.

**Results:** Independent risk factors such as alpha-fetoprotein (AFP), lens cuinaris agglutinin-reactive fraction of AFP (AFP-L3), desgamma-carboxy prothrombin (DCP), albumin (ALB), and total bilirubin (TBIL) obtained from train cohort were used to construct three nomograms (LAD, C-GALAD, and TAGALAD) using logistic regression. In the training and two validation cohorts, their AUCs were all over 0.900, and the higher AUCs appeared in TAGALAD and C-GALAD. Furthermore, the three nomograms could effectively stratify HCC into two groups with different survival and recurrence outcomes in follow-up validation. Notably, TAGALAD could predict HCC up to 48 weeks (AUC: 0.984) and 24 weeks (AUC: 0.900) before clinical diagnosis.

**Conclusion:** The proposed nomograms generated from real-world Chinese populations are effective and easy-to use for HCC surveillance, diagnosis, as well as prognostic evaluation in various clinical scenarios based on data feasibility.

Keywords: hepatocellular carcinoma, real-world data, diagnostic, prognosis, algorithm

#### Introduction

Hepatocellular carcinoma (HCC) is a growing public health problem and the third leading cause of cancer mortality worldwide,<sup>1</sup> about half of which occur in China.<sup>2</sup> Chronic hepatitis B virus (HBV) infection is closely associated with the development of HCC in China.<sup>3</sup> The lack of surveillance and early diagnosis tools is blamed for the poor prognosis and high mortality of HCC.<sup>4,5</sup> The imaging technology [such as computed tomography (CT) and ultrasonography (US)], together with serum  $\alpha$ -fetoprotein (AFP), has been recommended by clinical guidelines for the screening of HCC.<sup>6</sup> However, ultrasonography screening has been challenged when distinguishing small HCC from cirrhosis with an accuracy ranging from 47% to 63%.<sup>7.8</sup> Similarly, the serum

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AFP as a vital tumor marker for HCC is elevated to varying degrees compared to healthy negative controls, and sensitivity is limited to 65%.<sup>7</sup> Therefore, there is an unmet need for a single serum marker in detecting early-stage HCC.

As the amount of available high-throughput clinical data in medical centers continues to grow, an open question has arisen, ie how to make appropriate use of these large-scale data to achieve a comprehensive identification of cancer. Machine learning algorithms are powerful scientific mathematical tools and have provided valuable early diagnosis and prognosis evaluation methods for several human diseases.<sup>9</sup> Recently, several algorithms, such as GALAD [derived from Gender, Age, AFP-L3, AFP, and Des-carboxy-prothrombin (DCP)] and HCC early detection screening (HES), have been introduced as important alternatives for HCC diagnosis and prognosis prediction.<sup>10–13</sup> The diagnostic efficacy of the GALAD algorithm had been validated in international cohorts including the UK, Germany, and Japan.<sup>11–13</sup> Previous studies indicated that the performances of biomarkers varied with geographical region and were affected by etiology in different areas.<sup>14</sup> For example, HBV viral load could significantly influence the predictive value of AFP applied solely.<sup>15</sup> Our previous study confirmed that the GALAD efficacy was diluted due to the low diagnostic efficacy of AFP in HBV DNA-positive populations.<sup>14</sup> Thus, GALAD might require further optimization if applied to the Chinese population.<sup>14</sup>

Additionally, most of these designed studies were selective with inevitable sample bias and thus might not represent the "real-world" situation in most Chinese populations.<sup>16</sup> In this context, we conducted a comprehensive multi-center study using the real-world setting, aiming to both train and validate the  $\beta$  coefficients for each variable of GALAD in the Chinese population. In addition, while accounting for differences in the accessibility of clinical laboratory indicators, we identified factors associated with the presence of HCC and proceeded to develop as well as validate the algorithms. The diagnostic, prognostic, and screening values of the models were also validated in a retrospective multicenter cohort, a longitudinal follow-up cohort of HCC, and a prospective follow-up cohort of chronic hepatitis B (CHB) patients with liver fibrosis/cirrhosis, respectively. Both the results from the cross-sectional and longitudinal real-world data indicate that our newly developed nomograms are feasible and effective in HCC predicting and monitoring. Clinical physicians can utilize digital outputs from model scenarios to assess patient conditions, thereby facilitating informed clinical decision-making.

#### **Materials and Methods**

#### **Study Populations**

The study cohort comprised a total of 10,359 eligible participants, with 8,222 individuals diagnosed with HCC and 2,137 patients with benign liver diseases (BLDs). These participants were recruited between June 2007 and December 2019 at various medical institutions, including Shanghai Eastern Hepatobiliary Surgery Hospital (EHBH), Yueyang Hospital of Integrated Traditional Chinese and Western Medicine of Shanghai University of Traditional Chinese Medicine (YY), Shanghai Changzheng Hospital (CZ), Mengchao Hepatobiliary Hospital at Fujian Medical University (FJ), First Affiliated Hospital of Kunming Medical University (KM), and Jing'an District Central Hospital of Shanghai (JA), within the framework of the Multi-center Cooperation Platform of Molecular Diagnostics in China (http://www.multico.com.cn/).

The enrollment of HCC and BLDs was based on a comprehensive range of etiologies, including chronic viral liver disease, chronic non-alcoholic fatty liver disease, and autoimmune liver disease. The inclusion criteria of HCC were as follows: (1) histological confirmation of HCC diagnosis, (2) absence of macroscopic hepatic vein tumor thrombus, extrahepatic spread, or distant metastases, (3) no concurrent malignancies, and (4) no prior anticancer treatment.

BLDs consisted of individuals previously diagnosed with chronic liver disease (CLD), liver fibrosis (LF), and cirrhosis (LC) with different etiology. Since 68.98% (1474/2137) BLD subjects were attributed to HBV infection, the CHB group was set up for comparison in addition to BLDs with broad etiology.

LC was diagnosed when one or more of following criteria<sup>17</sup> were met: (1) histopathological evidence of cirrhosis; (2) platelet counts  $< 100 \times 10^{9}$ /L; (3) evidence of portal hypertension that included liver cirrhosis upon US, CT, or MRI examinations and associated splenomegaly but no vascular thrombosis, thrombocytopenia with a platelet counts  $< 120 \times 10^{9}$ /L, and/or presence of esophagogastric varices on endoscopic examination.

CHB was defined as having positive serum hepatitis B surface antigen (HBsAg) for at least 6 months, or HBV DNA titer exceeding 20,000 IU/mL in hepatitis B e antigen (HBeAg) positive patients, or HBV DNA titer surpassing 2000 IU/mL in HBeAg-negative patients.<sup>17</sup>

The study was approved by the institutional ethics committee at the leading study center (EHBHKY2018-1-010).

#### Study Design

We conducted a three-phase, real-world multi-center study to discover and validate the panel of serum biomarkers for HCC. Figure S1 showed the pipeline of the study.

#### Phase I (Training and Cross-Sectional Validation Phase)

During the training phase, we retrospectively gathered serum clinical laboratory test results from a case–control cohort of 7,729 Chinese subjects, comprising 5,998 HCC patients and 1,731 matched control subjects in EHBH, between January 2015 and December 2017 to identify candidate serum biomarkers. Subsequently, using the rms software package (<u>http://www.r-project.org/</u>), version 4.0 of R, we developed nomograms based on candidate serum biomarkers. The diagnostic performance among different algorithms was compared using the DeLong test.

The serum biomarkers AFP, AFP-L3, DCP, albumin (ALB), and total bilirubin (TBIL) had previously been proposed as biomarkers for HCC screening.<sup>12,14,18</sup> Here, the AFP level was determined using the electro-chemiluminescence immunoassay (ECLIA) method on Cobas e601 (Roche Diagnostics, Alameda, CA, USA). Lens culinaris agglutinin (LCA) affinity-binding fraction solutions were obtained from all serum samples using the Hotgen Biotech glycosyl capture spin column (Hotgen Biotech Co., Ltd, Beijing, China). Both the serum samples and corresponding LCA-reactive fraction affinity-binding solutions were measured using the ECLIA method on Cobas e601 (Roche Diagnostics, Alameda, CA, USA) to determine the AFP-L3/AFP index (the ratio of AFP-L3 to total AFP). The DCP quantitation was performed using chemiluminescence enzyme immunoassay on the LUMIPULSE G1200 (FUJIREBIO, Tokyo, Japan). The concentrations of ALB and TBIL were detected using the bromocresol green colorimetry method and the vanadate microplate method on the Cobas c701 (Roche Diagnostics, Alameda, CA, USA), respectively.

We assessed the diagnostic performances of the nomogram in independent internal (HCC: n=2064; BLDs: n=253) and external (HCC: n=148; BLDs: n=129) validation cohorts. The data for internal validation were collected between January 2018 and December 2018 in EHBH, and this dataset was referred to as independent internal validation. The external validation dataset primarily comprised patients from four medical centers (CZ, KM, FJ, and YY) as described above collected between January 2016 and December 2019, defining it as a multi-center cohort or external validation cohort. To evaluate the diagnostic performance, we employed the receiver operating characteristic (ROC) curve and determined the cutoff value using the maximum Youden index (sensitivity + specificity – 1). Sensitivity, specificity, positive predictive value (PPV), and negative predictive value (NPV) were used to assess the diagnostic accuracies at the cutoff values.

## Phase II (Follow-Up Validation Phase)

The prognostic value of the nomograms was verified in a follow-up study cohort comprising 204 HCC patients between January 2017 and December 2018 in EHBH. Each enrolled patient underwent follow-up evaluations every 3 months for the first 2 years after hepatectomy, and subsequently every 3–6 months until either death or withdrawal from the program. Overall survival (OS) was defined as the duration between hepatectomy and either death or the final follow-up program date, while recurrence-free survival (RFS) was defined as the period between hepatectomy and either recurrence, death, or the final follow-up program date. During this phase, the median postoperative follow-up time (interquartile range) was 33.50 (25.73–42.45) months. Kaplan–Meier curves were utilized to compare OS and RFS among different HCC subgroups as defined by the nomograms.

## Phase III (Prospective Validation Phase)

Upon completing Phases I and II, nomograms based on clinical serum biomarkers were developed and validated. Subsequently, the efficacy of these nomograms was assessed in a prospective validation cohort. This cohort originated from JA, a center with which we have previously collaborated.<sup>14</sup>

This study followed specific inclusion and exclusion criteria: (1) inclusion of individuals aged 30–60 years, both men and women; (2) capacity to comprehend the study, provide informed consent, and a willingness to adhere to research requirements; (3) positivity for HBsAg for a minimum duration of 1 year; (4) at the time of enrollment, the participants

could be either HBeAg positive or HBeAg negative, with HBV-DNA levels of at least  $\geq 10^4$  copies/mL; (5) clinical indications of cirrhosis had to be evident at the time of enrollment (detailed criteria provided), and the Child-Pugh grade had to be A. Alternatively, if cirrhosis was not apparent during enrollment, a liver biopsy conducted within the last 2 years should have revealed an Ishak fibrosis grade of at least 4 (on a scale of 0–6). Moreover, (6) exclusion criteria encompassed patients with HCC, decompensated cirrhosis, autoimmune hepatitis, severe physical and mental disorders (eg, alcoholism, uncontrolled diabetes, or tumors), and those with concurrent HCV, HDV, and HIV infections.

The study involved a structured follow-up schedule, which included the following: (1) clinical assessments every 12 weeks to monitor the patient's condition; (2) comprehensive evaluations such as blood routine, urine routine, prothrombin time, and biochemical tests conducted every 12 weeks; (3) semi-annual assessments including serum AFP levels and B-ultrasound liver examinations; (4) monitoring of liver fibrosis indicators at 6-month intervals; (5) testing for HBsAg, HBeAg/HBeAb at baseline, 48 weeks, 96 weeks, and 144 weeks; (6) rigorous monitoring of HBV-DNA at baseline, 12 weeks, 24 weeks, 36 weeks, 48 weeks, 72 weeks, 96 weeks, 120 weeks, and 144 weeks; (7) the primary study endpoint was the occurrence of HCC and liver-related deaths during the 144-week follow-up period, with HCC diagnosis confirmed through liver biopsy.

At the conclusion of the study at week 144, a total of 12 patients had developed HCC. For research purposes, 24 patients who did not develop HCC were randomly matched as controls. These 36 patients were subsequently included in our prospective validation cohort. The AUC was used to assess the diagnostic efficacy of the nomograms at the time of HCC diagnosis, as well as at weeks -24 and -48 within the prospective validation cohort.

## Statistical Analysis

The categorical variable was described as numbers and percentages of patients and compared using the  $\chi^2$  test or Fisher's exact test. For the normal distribution of continuous variables, we expressed as mean  $\pm$  SD and compared it using the unpaired, 2-tailed *t*-test. Meanwhile, for the non-normal distribution of continuous variables, we expressed as median (interquartile range) and compared using the Mann–Whitney *U*-test. In all analyses, *P* < 0.05 was considered to indicate statistical significance. All analyses were performed using SPSS 26.0 for Window software (SPSS, Chicago, IL) and R version 4.0.2 for windows.

## Results

#### **Patient Characteristics**

The study flowchart was shown in <u>Figure S1</u>. Patients diagnosed at EHBH from January 2015 to December 2017 formed a training cohort for developing the nomogram. Those diagnosed at EHBH between January 2018 and December 2018 were designated as an internal independent validation cohort. Eligible patients from four external medical centers (CZ, KM, FJ, and YY) from January 2016 to December 2019 constituted an external multicenter validation cohort. Following the training and internal validation phase, 204 patients from EHBH were enrolled in a longitudinal follow-up validation cohort to assess the long-term prognostic prediction efficacy of the models. Furthermore, a prospective cohort from JA, comprising 12 HCC patients and 24 matched BLDs, was included for the prospective validation of the models. The clinical characteristics of the BLDs and HCC patients are summarized in Tables 1, S1 and S2.

## Development of HCC-Predicting Nomograms

In the training cohort, variables with the *P* values <0.05 in the univariable logistic regression analysis were entered into the multivariable logistic regression analysis. The results of univariate and multivariate logistic regression analysis are shown in Table 2. After multivariate logistic regression, seven significant associated factors were employed to build HCC diagnostic nomograms. The  $\beta$  coefficients and odds ratios (95% CI) for the seven variables in Table 2 including log(AFP) at 4.399 (3.658–5.290), AFP-L3 at 1.065 (1.040–1.090), log(DCP) at 30.906 (23.279–41.033), sex at 5.582 (4.515–6.901), age at 1.073 (1.063–1.082), log(TBIL) at 0.049 (0.032–0.074), and ALB at 1.168 (1.144–1.192). Subsequently, a diagnostic nomogram named LAD (Figure 1A) was created, including AFP-L3 and log-transformed <u>AFP</u> and <u>DCP</u> values. Likewise, nomograms named C-GALAD (the initial "C" indicating the Chinese population differ from published

VariableTraining Cohort (n=7729)BLDs (n=1731)HCC (n=5998)		hort (n=7729)	P-value	Internal Validatio	Internal Validation Cohort (n=2317)		
			BLDs (n=253)	HCC (n=2064)			
Age, y	49.74±12.81	54.18±10.71	<0.001	53.44±11.83	55.32±10.70	0.002	
Sex, %						<0.001	
Male	887 (51.24)	5087 (84.81)		180 (71.15)	1773 (85.90)		
Female	844 (48.76)	911 (15.19)		73 (28.86)	291 (14.10)		
HBsAg, %			<0.001			<0.001	
Positive	1200 (69.32)	5039 (84.01)		144 (56.92)	1689 (81.83)		
Negative	503 (29.06)	890 (14.84)		103 (40.71)	354 (17.15)		
WBCs, 10 <sup>9</sup> /L	4.71 (3.34–5.95)	5.08 (4.12-6.20)	<0.001	3.27 (2.26-4.72)	5.03 (4.04-6.14)	<0.001	
PLT, 10 <sup>9</sup> /L	160.50 (70.81-222.29)	153.00 (112.15–198.17)	0.110	69.89 (45.28–109.10)	150.79 (109.61–193.57)	<0.001	
TBIL, μmol/L	14.55 (10.35–24.14)	13.80 (10.64–17.97)	<0.001	22.40 (14.72–49.88)	13.71 (10.51–18.98)	<0.001	
ALB, g/L	41.05 (34.58-44.26)	42.02 (39.47-44.28)	<0.001	33.30 (29.85–38.95)	42.40 (39.76–44.90)	<0.001	
AFP, log µg/L	0.48 (0.30-0.71)	1.76 (0.74–3.02)	<0.001	0.55 (0.33–0.94)	1.64 (0.67–2.91)	<0.001	
AFP-L3, %	0.10 (0.00-0.70)	4.79 (0.04–22.82)	<0.001	0.37 (0.00-1.49)	3.64 (0.12-24.47)	<0.001	
DCP, log mAU/mL	1.30 (1.20–1.41)	2.59 (1.80–3.46)	<0.001	1.32 (1.19–1.49)	2.56 (1.77–3.50)	<0.001	

**Notes**: Categorical variables presented as number (percentage) and continuous data presented as means ± SDs or median (interquartile range); categorical variables were compared using the chi-square test or Fisher's exact test. Continuous variables were compared using the independent *t*-test or Mann–Whitney *U*-test. **Abbreviations**: BLD, benign liver disease; HCC, hepatocellular carcinoma; HBsAg, hepatitis B surface antigen; WBC, leukocyte count; PLT, platelet; TBIL, total bilirubin;

Abbreviations: BLD, benign liver disease; HCC, hepatocellular carcinoma; HBsAg, hepatitis B surface antigen; WBC, leukocyte count; PLT, platelet; TBIL, total bilirubin; ALB, albumin; AFP, alpha fetoprotein; AFP-L3, Lens culinaris agglutinin-reactive fraction of AFP; DCP, des-gamma-carboxy prothrombin.

GALAD) (Figure 1B)<sup>12</sup> and TAGALAD (the initial TA indicating <u>TBIL</u> and <u>ALB</u>, respectively) (Figure 1C) were also established. As shown in the nomogram C-GALAD (Figure 1B),  $\log(DCP)$ ,  $\log(AFP)$ , and AFP-L3 still made the most considerable contribution to this nomogram, followed by sex and age at diagnosis. While in the nomogram TAGALAD (Figure 1C), TBIL and  $\log(DCP)$  made the most significant contribution, followed by  $\log(AFP)$ , AFP-L3, age at diagnosis, sex, and ALB.

#### Cross-Sectional Validation of the Nomograms

The three above diagnostic nomograms all demonstrated higher accuracy in distinguishing HCC from BLDs (AUC of LAD: 0.943, 95% CI: 0.938–0.948; C-GALAD: 0.952, 95% CI: 0.947–0.956; TAGALAD: 0.969, 95% CI: 0.965–0.973) compared with AFP (AUC: 0.828, 95% CI: 0.819–0.836, P<0.001), DCP (AUC: 0.919, 95% CI: 0.913–0.925, P<0.001) and AFP-L3 (AUC: 0.764, 95% CI: 0.755–0.774, P<0.001), in the training cohort (Table 3 and Figure 2A). As the published GALAD score has been established and validated for early diagnosis of HCC, we also compared the performances of our nomograms and the formerly published GALAD score in distinguishing HCC. The results suggested that our nomograms also demonstrated better diagnostic efficacy than the GALAD score (AUC: 0.925, 95% CI: 0.919–0.931, P<0.001) (Table 3 and Figure 2A).

In subgroup analysis, our three nomograms consistently achieved AUCs above 0.900 for distinguishing HCC from LC in the training cohort. Specifically, the AUCs were 0.910 for LAD (95% CI: 0.903–0.917), 0.912 for C-GALAD (95% CI: 0.905–0.919), and 0.972 for TAGALAD (95% CI: 0.967–0.975). These results significantly surpassed those of AFP (AUC: 0.746, 95% CI: 0.735–0.757, P<0.001), DCP (AUC: 0.894, 95% CI: 0.886–0.901), AFP-L3 (AUC: 0.732, 95% CI: 0.721–0.743, P<0.001), and GALAD (AUC: 0.860, 95% CI: 0.851–0.917) (Table 3 and Figure 2B).

In the HBV-related case subgroup analysis, our models (LAD, C-GALAD, and TAGALAD) also demonstrated superior diagnostic performance for HBV-HCC compared to GALAD (AUC: 0.859, 95% CI: 0.844–0.875), with all AUCs above 0.900. LAD achieved an AUC of 0.928 (95% CI: 0.918–0.938), C-GALAD reached 0.927 (95% CI: 0.917–0.938), and TAGALAD showed the best performance with an AUC of 0.966 (95% CI: 0.961–0.971) (Table 3 and Figure 2C).

In individuals without HBV infection, TAGALAD (AUC: 0.973, 95% CI: 0.965–0.980) and C-GALAD (AUC: 0.962, 95% CI: 0.954–0.969) maintained superior HCC diagnostic performance compared to GALAD (AUC: 0.958, 95% CI:

Variable	Univariate Logistic R	legression	Multivariate Logistic Regression			
	OR (95% CI)	P-value	β	OR (95% CI)	P-value	
Sex, Male vs Female	5.313 (4.723–5.977)	<0.001	1.719	5.582 (4.515–6.901)	<0.001	
Age, year	1.035 (1.030-1.040)	<0.001	0.070	1.073 (1.063-1.082)	<0.001	
AFP, log ug/L	4.803 (4.327–5.333)	<0.001	1.481	4.399 (3.658-5.290)	<0.001	
DCP, log mAU/mL	49.890 (39.164–63.553)	<0.001	3.431	30.906 (23.279-41.033)	<0.001	
AFP-L3, %	1.242 (1.215–1.269)	<0.001	0.063	1.065 (1.040-1.090)	<0.001	
WBC, 10 <sup>9</sup> /L	1.135 (1.101–1.170)	<0.001				
RBC, 10 <sup>12</sup> /L	3.653 (3.307-4.035)	<0.001				
PLT, 10 <sup>9</sup> /L	1.001 (1.000-1.001)	0.115				
RDW, %	0.655 (0.629–0.681)	<0.001				
PDW, %	1.008 (0.989-1.028)	0.413				
MCV, fL	1.025 (1.016–1.034)	<0.001				
MPV, fL	1.028 (0.985-1.074)	0.206				
PCT, %	1.212 (0.926-1.589)	0.163				
TBIL, log μmol/L	0.189 (0.150-0.237)	<0.001	-3.026	0.049 (0.032-0.074)	<0.001	
ALB, g/L	1.112 (1.100–1.125)	<0.001	0.155	1.168 (1.144–1.192)	<0.001	
AST, U/L	1.000 (1.000–1.001)	0.376				
ALT, U/L	1.002 (1.001–1.004)	0.005				
GGT, U/L	1.003 (1.002-1.003)	<0.001				
ALP, U/L	0.999 (0.998-1.000)	0.111				

**Table 2** The Univariate and Multivariate Logistic Regression Analysis for Distinguishing HCC from BLDs Based

 on Clinical Laboratory Index

**Abbreviations**: OR, odds ratio; AFP, alpha fetoprotein; DCP, des-gamma-carboxy prothrombin; AFP-L3, Lens culinaris agglutinin-reactive fraction of AFP; WBC, leukocyte count; RBC, erythrocyte blood cell; PLT, platelet; RDW, red blood cell distribution width; PDW, platelet distribution width; MCV, mean corpuscular volume; MPV, mean platelet volume; PCT, plateletcrit; TBIL, total bilirubin; ALB, albumin; AST, aspartate aminotransferase; ALT, alanine aminotransferase; GGT,  $\gamma$ -glutamyl transpeptidase; ALP, alkaline phosphatase.

0.950–0.965), with significantly higher accuracy (Table 3 and Figure 2D). Table 3 also provides comprehensive data on sensitivity, specificity, and AUCs for all cohorts studied.

In the internal validation cohort, the nomograms from the training cohort showed consistent performance discriminating HCC from matched controls with AUCs of 0.905 (95% CI: 0.892–0.917) for LAD, 0.908 (95% CI: 0.895–0.919) for C-GALAD, and 0.964 (95% CI: 0.956–0.972) for TAGALAD (<u>Table S3</u>, Figure 2E). Subgroup analysis of the internal cohort as shown in <u>Table S3</u> revealed that the TAGALAD model excelled in diagnosing HBV-HCC with an AUC of 0.957 (95% CI: 0.947–0.966) and HCC without HBV infection with an AUC of 0.980 (95% CI: 0.962–0.991), surpassing the diagnostic accuracy of GALAD (AUC: 0.845 for HBV cases, 95% CI: 0.813–0.877; 0.911 for non-HBV cases, 95% CI: 0.883–0.939), LAD (AUC: 0.900 for HBV cases, 95% CI: 0.873–0.926; 0.913 for non-HBV cases, 95% CI: 0.885– 0.942), and C-GALAD (AUC: 0.901 for HBV cases, 95% CI: 0.875–0.927; 0.917 for non-HBV cases, 95% CI: 0.889– 0.945) (Figure 2F and G).

In external validation cohort, TAGALAD again showed superior performance in differentiating HCC from BLDs with an AUC of 0.931 (95% CI: 0.895–0.958), exceeding AFP (AUC: 0.769, 95% CI: 0.713–0.825), DCP (AUC: 0.840, 95% CI: 0.792–0.887), AFP-L3 (AUC: 0.752, 95% CI: 0.698–0.806), and GALAD (AUC: 0.911, 95% CI: 0.878–0.944) (Table S3 and Figure 2H). AUCs, sensitivity, specificity, PPV, and NPV are summarized in Table S3.

## Longitudinal Follow-Up Validation of the Nomograms

In the prognostic assessment of HCC patients using the LAD model, those with scores above the 58.17 cutoff exhibited a median OS and RFS of  $27.83\pm14.05$  and  $27.25\pm14.02$  months, respectively, which is approximately 5 months shorter than those below the cutoff, who experienced OS and RFS of  $33.38\pm11.85$  and  $32.83\pm12.12$  months, respectively (P<0.001, Figure 3A and B).

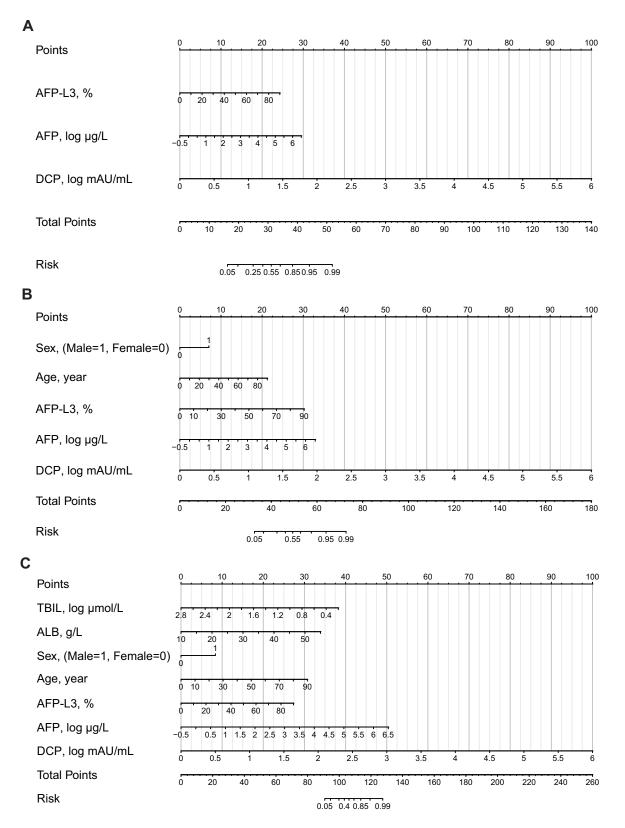


Figure I The three nomogram algorithms (LAD, C-GALAD and TAGALAD) to estimate the risk of HCC. Univariate and multivariate analyses identified AFP, AFP-L3, DCP, TBIL, and ALB as significant risk factors independently associated with HCC. Utilizing these factors, we developed HCC risk estimation nomograms: LAD (A), C-GALAD (B), and TAGALAD (C). To apply a nomogram, locate each indicator on its axis, assign points, sum these points, and then determine the HCC probability at the bottom of the nomogram based on the total points.

Table 3 Comparison Among the Four	Nomograms a	and the Individual	HCC Biomar	rker in the -	Training Cohort
	1		1		

Variable		AUC (95% CI)	P-value	Cut-Off	SEN (%, 95% CI)	SPE (%, 95% CI)	PPV (%, 95% CI)	NPV (%, 95% CI)
				value				
Training cohort: HCC vs BLDs	AFP, ng/mL	0.828 (0.819–0.836)	<0.001	7.01	70.32 (69.10–71.50)	82.26 (80.40-84.00)	93.20 (92.50–93.80)	44.40 (43.30–45.60)
	DCP, mAU/mL	0.919 (0.913-0.925)	<0.001	35.00	84.04 (83.10-85.00)	88.04 (86.40-89.50)	96.10 (95.50-96.50)	61.40 (60.00-62.90)
	AFP-L3, %	0.764 (0.755–0.774)	<0.001	10.00	41.46 (40.20-42.70)	98.27 (97.50-98.80)	98.80 (98.30-99.20)	32.60 (32.20-33.10)
	GALAD	0.925 (0.919-0.931)	<0.001	0.70	83.39 (82.40-84.30)	85.33 (83.60-87.00)	95.20 (94.60–95.70)	59.70 (58.30-61.20)
	LAD	0.943 (0.938-0.948)	<0.001	31.80	89.38 (88.60-90.10)	86.83 (85.10-88.40)	95.90 (95.40-96.40)	70.20 (68.60–71.80)
	C-GALAD	0.952 (0.947-0.956)	<0.001	53.30	85.93 (85.00-86.80)	90.58 (89.10–91.90	96.90 (96.50–97.30)	65.00 (63.50–66.50)
	TAGALAD	0.969 (0.965–0.973)	n.d.	110.93	88.86 (88.00-89.60)	94.45 (93.30–95.50)	98.20 (97.90–98.50)	71.00 (68.50–72.50)
Training cohort: HCC vs LC	AFP, ng/mL	0.746 (0.735–0.757)	<0.001	24.20	57.34 (56.10–58.60)	81.82 (78.70–84.70)	96.60 (96.10–97.10)	17.40 (16.80–18.10)
	DCP, mAU/mL	0.894 (0.886-0.901)	<0.001	36.00	83.59 (82.60-84.50)	77.88 (74.50-81.00)	97.20 (96.70–97.50)	34.30 (32.70-35.90)
	AFP-L3, %	0.732 (0.721–0.743)	<0.001	10.00	41.46 (40.20-42.70)	96.52 (94.80-97.80)	99.10 (98.60-99.40)	15.40 (15.00–15.70)
	GALAD	0.860 (0.851-0.917)	<0.001	1.97	70.86 (69.70–72.00)	84.55 (81.60-87.20)	97.70 (97.20-98.00)	24.20 (23.30–25.10)
	LAD	0.910 (0.903-0.917)	<0.001	38.23	79.53 (78.50-80.50)	86.82 (84.00-89.30)	98.20 (97.80-98.50)	31.80 (30.60-33.10)
	C-GALAD	0.912 (0.905-0.919)	<0.001	54.74	83.36 (82.40-84.30)	82.88 (79.80-85.70)	97.80 (97.40-98.10)	35.40 (33.90-36.90)
	TAGALAD	0.972 (0.967–0.975)	n.d.	110.77	89.03 (88.20-89.80)	93.77 (91.60–95.50)	99.20 (99.00–99.40)	48.40 (46.50–50.30)
Training cohort: HBV-HCC vs CHB	AFP, ng/mL	0.732 (0.712–0.752)	<0.001	34.26	55.96 (54.60–57.30)	79.32 (75.50–82.80)	96.40 (95.80–97.00)	15.20 (14.60–16.00)
	DCP, mAU/mL	0.927 (0.916-0.937)	<0.001	32.00	85.77 (84.80-86.70)	85.69 (82.30-88.60)	98.40 (98.00–98.70)	37.50 (35.80–39.40)
	AFP-L3, %	0.726 (0.711–0.741)	<0.001	10.00	42.21 (40.80-43.60)	96.22 (94.20-97.70)	99.10 (98.60–99.40)	14.30 (13.90–14.60)
	GALAD	0.859 (0.844–0.875)	<0.001	0.81	82.06 (81.00-83.10)	71.97 (67.80–75.90)	96.70 (96.20-97.10)	28.60 (27.00-30.30)
	LAD	0.928 (0.918-0.938)	<0.001	38.23	79.92 (78.80-81.00)	91.05 (88.20-93.40)	98.90 (98.50-99.20)	31.20 (29.90-32.50)
	C-GALAD	0.927 (0.917–0.938)	<0.001	54.09	84.72 (83.70-85.70)	86.68 (83.40-89.50)	98.50 (98.10-98.80)	36.20 (34.50-37.90)
	TAGALAD	0.966 (0.961–0.971)	n.d.	110.98	88.65 (87.70–89.50)	93.64 (91.10–95.60)	99.30 (99.00–99.50)	45.20 (43.20-47.20)
Training cohort: HCC vs BLDs	AFP, ng/mL	0.825 (0.806-0.844)	<0.001	5.60	65.17 (61.90–68.30)	88.92 (87.00–90.60)	81.30 (78.70–83.80)	77.50 (75.80–79.10)
(without HBV infection)	DCP, mAU/mL	0.918 (0.905-0.931)	<0.001	35.00	83.93 (81.40-86.30)	88.67 (86.70-90.40)	84.60 (82.40-86.60)	88.20 (86.50-89.60)
	AFP-L3, %	0.736 (0.719–0.753)	<0.001	10.00	36.40 (33.20-39.70)	99.08 (98.40–99.50)	96.70 (94.20–98.20)	67.70 (66.60–68.80)
	GALAD	0.958 (0.950-0.965)	<0.001	0.10	91.01 (88.90–92.80)	86.67 (84.60-88.50)	83.50 (81.40-85.40)	92.90 (91.30–94.10)
	LAD	0.938 (0.927-0.950)	<0.001	32.77	85.28 (82.80-87.50)	92.00 (90.30-93.50)	88.80 (86.70-90.60)	89.40 (87.80–90.80)
	C-GALAD	0.962 (0.954-0.969)	<0.001	51.27	90.45 (88.30–92.30)	90.00 (88.20-91.60)	87.00 (85.00-88.80)	92.70 (91.20-94.00)
	TAGALAD	0.973 (0.965–0.980)	n.d.	112.60	87.42 (85.10-89.50)	97.17 (96.10–98.00)	95.80 (94.30–97.00)	91.20 (89.70–92.50)

Abbreviations: AUC, area under the receiver operating characteristic curve; SEN, sensitivity; SPE, specificity; PPV, positive predict value; NPV, negative predict value; AFP, alpha fetoprotein; AFP-L3, the lens culinaris agglutinin-reactive fraction of AFP; DCP, des-gamma-carboxy prothrombin; HCC, hepatocellular carcinoma; BLD, benign liver diseases; LC, liver cirrhosis; CHB, chronic hepatitis B.

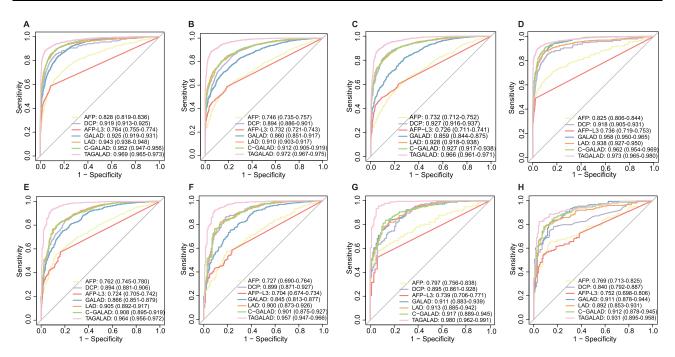


Figure 2 The ROC curve analyses of GALAD, LAD, C-GALAD, TAGALAD and individual tumor marker in training and validation cohorts. (**A–D**) Training cohort: The AUC analysis showed the diagnostic efficacy of the three nomograms for HCC and BLDs (**A**) (LAD: AUC: 0.943, 95% CI: 0.938–0.948; C-GALAD: AUC: 0.952, 95% CI: 0.947–0.956; TAGALAD: AUC:0.969, 95% CI: 0.965–0.973) better than GALAD (AUC: 0.925, 95% CI: 0.919–0.931) and individual tumor markers (AFP: AUC: 0.828, 95% CI: 0.819–0.836; DCP: AUC: 0.919, 95% CI: 0.913–0.925; AFP-L3: 0.764, 95% CI: 0.755–0.774). In the subgroup analysis of HCC vs LC (**B**), HBV-HCC vs CHB (**C**) and HCC vs BLDs (both without HBV infection) (**D**), the diagnostic algorithm still achieved excellent diagnostic performance, with AUCs exceeding 0.900. (**E–G**) Internal validation cohort: The AUC revealed that the nomograms (LAD: AUC: 0.905, 95% CI: 0.892–0.917; C-GALAD: AUC: 0.908, 95% CI: 0.895–0.919; TAGALAD: AUC: 0.964, 95% CI: 0.956–0.972) showed the best diagnostic power for HCC vs BLDs (**E**) in the internal validation cohort (AFP: AUC: 0.763, 95% CI: 0.745–0.780; DCP: 0.894, 95% CI: 0.881–0.906; AFP-L3: AUC: 0.724, 95% CI: 0.705–0.742; GALAD: AUC: 0.866, 95% CI: 0.879–0.379). Similarly, in the subgroup analyses of HBV-HCC vs CHB (**F**) and HCC vs BLDs (both without HBV infection) (**G**) in the internal validation cohort, the diagnostic algorithms continued to show excellent performance, with all AUCs exceeding 0.900. (**H**) External validation cohort: For the external validation cohort, the nomograms (LAD: AUC: 0.892, 95% CI: 0.892–0.979). Similarly, in the subgroup analyses of HBV-HCC vs CHB (**F**) and HCC vs BLDs (both without HBV infection) (**G**) in the internal validation cohort, the diagnostic algorithms continued to show excellent performance, with all AUCs exceeding 0.900. (**H**) External validation cohort: For the external validation cohort, the nomograms (LAD: AUC: 0.892, 95% CI: 0.853–0.931; C-GALAD: AUC: 0.912, 95% CI: 0.878–0.945; TAGALAD: AUC: 0.931, 95% CI: 0.895–0.958) again

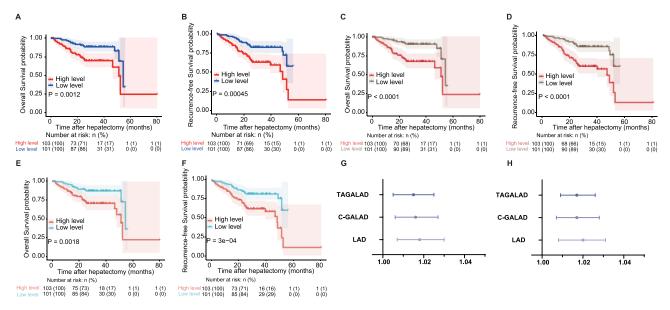


Figure 3 Performances of three nomograms in predicting OS and RFS in the follow-up cohort. (A and B) LAD algorithm: In HCC prognostic evaluation using the LAD model, patients with scores over 58.17 showed median OS (A) and RFS (B) of 27.83±14.05 and 27.25±14.02 months, respectively, which were about 5 months less than those with lower scores (OS: 33.38±11.85, RFS: 32.83±12.12 months). (C and D) C-GALAD algorithm: In HCC prognosis using C-GALAD, high-risk patients (scores > 78.10) had OS (C) and RFS (D) of 27.32±14.16 and 26.74±14.11 months, respectively. The low-risk group showed longer OS (33.90±11.44 months) and RFS (33.36±11.75 months). (E and F) TAGALAD algorithm: Using 139.36 as the median cutoff in the TAGALAD model, the high-risk group's OS (E) and RFS (F) were 28.31±14.18 and 27.73 ±14.17 months, respectively, about 4.58 and 4.62 months less than the low-risk group's OS (32.89±11.91) and RFS (32.35±12.15). (G and H) Cox regression forest plot: Cox regression revealed LAD, C-GALAD, and TAGALAD were HCC independent predictors of survival (LAD HR: 1.018; C-GALAD HR: 1.016; TAGALAD HR: 1.015) (G) and recurrence (LAD HR: 1.020; C-GALAD HR: 1.017; TAGALAD HR: 1.017) (all P-values <0.05) (H).

Using the C-GALAD model for HCC prognostic assessment, patients with a score above the median cutoff of 78.10 in the high-risk group showed OS and RFS of  $27.32\pm14.16$ ,  $26.74\pm14.11$  months, respectively, compared to OS 33.90  $\pm11.44$  and RFS  $33.36\pm11.75$  months in the low-risk group, indicating a significant difference over 6.28 months (P<0.001, Figure 3C and D).

In the TAGALAD model, using the median value 139.36 as the cutoff, OS and RFS in the high-risk group were 28.31  $\pm$ 14.18 and 27.73 $\pm$ 14.17 months, respectively, approximately 4.58 and 4.62 months shorter than in the low-risk group (OS 32.89 $\pm$ 11.91, RFS 32.35 $\pm$ 12.15 months), respectively (P<0.001, Figure 3E and F).

Furthermore, the Cox regression analysis indicated that LAD, C-GALAD, and TAGALAD were independent risk factors for both survival (HR of LAD: 1.018, 95% CI: 1.007–1.030; C-GALAD: 1.016, 95% CI: 1.006–1.027; TAGALAD: 1.015, 95% CI: 1.005–1.025; P<0.05; Figure 3G) and recurrence (HR of LAD: 1.020, 95% CI: 1.008–1.031; C-GALAD: 1.017, 95% CI: 1.007–1.028; TAGALAD: 1.017, 95% CI: 1.009–1.026; P<0.05; Figure 3H) in HCC patients.

The results indicated that LAD, C-GALAD, and TAGALAD algorithms could stratify HCC patients into different prognostic categories (high-risk and low-risk). Kaplan–Meier survival curves confirmed that higher nomogram scores correlated statistically with poorer OS and RFS (P < 0.001, Figure 3).

#### Longitudinal Prospective Validation of the Nomograms

To determine the predictive time frame prior to HCC diagnosis, we evaluated the nomogram points of patients with very early-stage HCC (n=12) and matched CHB controls (n=24) in prospective validation cohort from JA. At -48, -24, and 0 weeks, TAGALAD continued to demonstrate its leading performance among the four HCC diagnostic algorithms, with AUCs of 0.984 (95% CI: 0.954–1.000), 0.900 (95% CI: 0.740–1.000), and 0.996 (95% CI: 0.900–1.000), respectively (Table 4, Figure 4).

At week -48, TAGALAD, LAD, C-GALAD, and GALAD had AUCs of 0.984 (95% CI: 0.954–1.000), 0.964 (95% CI: 0.912–1.000), 0.937 (95% CI: 0.848–1.000), and 0.845 (95% CI: 0.708–0.982), respectively, ranking in descending order (Figure 4A, Table 4). ROC curve analysis determined cutoff values for these models, with 103.54 for TAGALAD, 25.90 for LAD, 46.15 for C-GALAD, and -2.23 for GALAD based on the Youden index. TAGALAD and LAD both achieved sensitivity of 100.00% (95% CI: 66.40%–100.00%), surpassing both GALAD and C-GALAD with 88.89% sensitivity (95% CI: 51.80%–99.70%) (Figure 4A, Table 4).

At week -24, the AUC for TAGALAD was 0.900 (95% CI: 0.740-1.000) for LAD 0.889 (95% CI: 0.711-1.000), C-GALAD 0.871 (95% CI: 0.693-1.000), GALAD 0.830 (95% CI: 0.646-1.000) (Figure 4B, Table 4). These results highlighted TAGALAD as the most effective diagnostic model among the four algorithms. At an optimal cutoff value of 102.15, determined by the maximum Youden index, TAGALAD achieved sensitivity and specificity for diagnosing very early HCC at 90.91% (95% CI: 58.70-99.80%) and 90.32% (95% CI: 74.20-98.00%), respectively (Table 4).

At week 0, the AUCs of the four algorithms were all above 0.900, with TAGALAD (AUC: 0.996, 95% CI: 0.900– 1.000) and LAD (AUC: 0.992, 95% CI: 0.893–1.000) leading closely, followed by C-GALAD (AUC: 0.985, 95% CI: 0.880–1.000) and GALAD (AUC: 0.939, 95% CI: 0.810–0.991) (Figure 4C, Table 4). Table 4 summarizes the sensitivity, specificity, PPV, and NPV of these models at -48W, -24w, and week 0.

## Discussion

In this real-world multi-center study, three diagnostic nomogram algorithms based on AFP, AFP-L3, DCP (LAD), Age, Gender (C-GALAD), TBIL, and ALB (TAGALAD) were developed and validated to predict the risk and the prognoses of HCC in Chinese populations when different parameters were available. Especially, the TAGALAD nomogram demonstrated strong diagnostic accuracy to predict early-stage HCC and monitor the prognosis. As such, this nomogram algorithm could be used as a valuable tool to screen high-risk populations, detect HCC at the resectable stage, and monitor the prognosis of HCC after hepatectomy.

Despite the progress in the diagnosis and treatment of HCC, the morbidity and mortality rate of HCC is still rising rapidly demonstrating the urgent need to strengthen the research in developing high-sensitivity and non-invasive biomarkers for HCC early diagnosis.<sup>19</sup> The AFP is a widely used, yet imperfect, biomarker for early HCC diagnosis.<sup>7</sup> New strategies are therefore urgently needed.<sup>20</sup> The AFP-L3 has been proved by the US Food and Drug Administration

Variable		AUC (95% CI)	P-value	Cut-Off value	SEN (%, 95% CI)	SPE (%, 95% CI)	PPV (%, 95% CI)	NPV (%, 95% CI)
HCC-48w AFP, ng/mL		0.696 (0.457–0.936)	0.0174	5.90	55.56 (21.20-86.30)	92.86 (76.50–99.10)	71.40 (36.80–91.50)	86.70 (75.70–93.10)
	DCP, mAU/mL	0.960 (0.903-1.000)	0.2647	21.00	100.00 (66.40-100.00)	85.71 (67.30-96.00)	69.20 (47.60-84.80)	100.00
	AFP-L3, %	0.726 (0.534–0.919)	0.0074	4.40	55.56 (21.20-86.30)	89.29 (71.80–97.70)	62.50 (33.00-84.90)	86.20 (74.90-92.90)
	GALAD	0.845 (0.708-0.982)	0.0319	-2.23	88.89 (51.80-99.70)	67.86 (47.60-84.10)	47.10 (33.10–65.10)	95.00 (74.60-99.20)
	LAD	0.964 (0.912-1.000)	0.3882	25.90	100.00 (66.40-100.00)	82.14 (63.10–93.90)	64.30 (44.90–79.90)	100.00
	C-GALAD	0.937 (0.848–1.000)	0.1568	46.15	88.89 (51.80-99.70)	92.86 (76.50–99.10)	80.00 (50.80-93.90)	96.30 (80.30–99.40)
	TAGALAD	0.984 (0.954–1.000)	n.d.	103.54	100.00 (66.40–100.00)	92.86 (76.50–99.10)	81.80 (54.20–94.50)	100.00
HCC-24w	AFP, ng/mL	0.834 (0.642–1.000)	0.6068	5.90	63.64 (30.80-89.10)	93.55 (78.60–99.20)	77.80 (46.00–93.50)	87.90 (76.70–94.10)
	DCP, mAU/mL	0.878 (0.701-1.000)	0.1623	24.00	90.91 (58.70-99.80)	87.10 (70.20–96.40)	71.40 (49.60-86.40)	96.40 (80.60–99.40)
	AFP-L3, %	0.789 (0.632-0.945)	0.2008	5.00	54.55 (23.40-83.30)	93.55 (78.60–99.20)	75.00 (41.40–92.70)	85.30 (75.10–91.80)
	GALAD	0.830 (0.646–1.000)	0.0704	-0.09	72.73 (39.00–94.00)	100.00 (88.80-100.00)	100.00	91.20 (79.70–96.40)
	LAD	0.889 (0.711–1.000)	0.4322	28.49	90.91 (58.70-99.80)	90.32 (74.20–98.00)	76.90 (52.80–90.90)	96.60 (81.20-99.50)
	C-GALAD	0.871 (0.693-1.000)	0.2379	53.65	72.73 (39.00–94.00)	100.00 (88.80-100.00)	100.00	91.20 (79.70–96.40)
	TAGALAD	0.900 (0.740-1.000)	n.d.	102.15	90.91 (58.70–99.80)	90.32 (74.20–98.00)	76.90 (52.80–90.90)	96.60 (81.20-99.50)
HCC-0w	AFP, ng/mL	1.000 (0.907-1.000)	0.4795	5.90	100.00 (66.40-100.00)	100.00 (88.10-100.00)	100.00	100.00
	DCP, mAU/mL	0.979 (0.870-1.000)	0.2385	27.00	100.00 (66.40-100.00)	89.66 (72.60–97.80)	75.00 (50.70–89.80)	100.00
	AFP-L3, %	0.833 (0.677–0.934)	0.0459	4.40	33.33 (7.50-70.10)	100.00 (88.10-100.00)	100.00	82.90 (75.30-88.50)
	GALAD	0.939 (0.810-0.991)	0.2044	-1.47	77.78 (40.00–97.20)	100.00 (88.10-100.00)	100.00	93.50 (81.00-98.00)
	LAD	0.992 (0.893-1.000)	0.4795	29.57	100.00 (66.40-100.00)	96.55 (82.20-99.90)	90.00 (56.70–98.40)	100.00
	C-GALAD	0.985 (0.880-1.000)	0.3822	52.59	88.89 (51.80–99.70)	100.00 (88.10-100.00)	100.00	96.70 (82.00–99.50)
	TAGALAD	0.996 (0.900-1.000)	n.d.	104.48	100.00 (66.40-100.00)	96.55 (82.20–99.90)	90.00 (56.70–98.40)	100.00

Table 4 Comparison Among the Four Nomograms and the Individual HCC Biomarker in the Prospective Study Cohort

Abbreviations: AUC, area under the receiver operating characteristic curve; SEN, sensitivity; SPE, specificity; PPV, positive predict value; NPV, negative predict value; AFP, alpha fetoprotein; AFP-L3, the lens culinaris agglutinin-reactive fraction of AFP; DCP, des-gamma-carboxy prothrombin; HCC, hepatocellular carcinoma.

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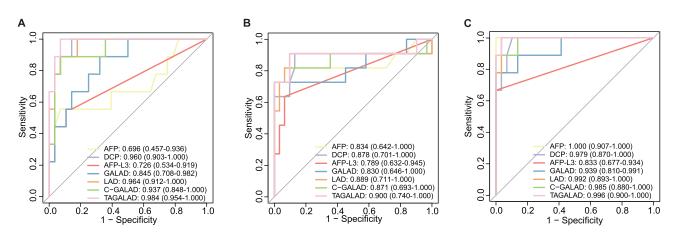


Figure 4 The diagnostic efficacies of AFP, AFP-L3, DCP, GALAD, LAD, C-GALAD and TAGALAD in distinguishing the very early-stage HCC from the matched controls at corresponding time points. (A) At week –48, the AUCs of TAGALAD, LAD, C-GALAD, and GALAD were 0.984, 0.964, 0.937, and 0.845 respectively, in descending order. (B) At week –24, AUCs were 0.900 for TAGALAD, 0.889 for LAD, 0.871 for C-GALAD, and 0.830 for GALAD. TAGALAD was the most effective one among these four diagnostic models. (C) At week 0, the best AUC was shown in TAGALAD (0.996), then was the LAD (0.992), the C-GALAD (0.985) and the GALAD (0.939).

(FDA) to be used as a serum marker for high-risk populations screening,  $^{21-23}$  as a prognostic biomarker after treatment,  $^{24-26}$  and as a malignant biologic factor of HCC.  $^{27,28}$  The DCP is an abnormal protein, and several reports have indicated that DCP is more specific than AFP in the diagnosis of HCC.  $^{29,30}$ 

Due to the heterogeneity of tumors, single biomarker could not perform well in early detection.<sup>31</sup> Recent diagnostic algorithms combining multiple biomarkers, such as GALAD and HES, are more sensitive for early diagnosis, with the sensitivity all over 70%.<sup>6,10–14</sup> However, the previous studies were mostly limited to defined cross-sectional exploration, which may overestimate the diagnostic performance of biomarkers.<sup>32</sup> The International Liver Cancer Association (ILCA) recently stressed the critical importance of retrospective longitudinal study to determine the performance of biomarkers in diagnosis and prognosis assessment.<sup>33</sup>

Of all the available machine learning algorithms, the nomogram is an easy-to-use visualized alternative LR scoring method and can facilitate management-related decision-making.<sup>34–36</sup> Currently, in this study, compared with the three individual tumor biomarkers (AFP, AFP-L3, and DCP), the LAD nomogram, which digitally combined the AFP, AFP-L3, and DCP, could improve the diagnostic accuracy to detect HCC (LAD: AUC: 0.943, 95% CI: 0.938–0.948; AFP: AUC: 0.828, 95% CI: 0.819–0.836; AFP-L3: AUC: 0.764, 95% CI: 0.755–0.774; DCP: AUC: 0.919, 95% CI: 0.913–0.925) (Figures 1A, 2A and Table 3).

The GALAD algorithm incorporates the three biomarkers and has shown excellent diagnostic efficacy in several study cohorts, higher than each individual HCC biomarker.<sup>12–14</sup> It is not surprising that the diagnostic efficacy of the panel is better than a single biomarker.<sup>12–14</sup> For example, for the initial single-center study in the UK, the GALAD achieved a good diagnostic efficacy for detecting HCC with an AUC of 0.97.<sup>12</sup> Similarly, the GALAD showed an encouraging diagnostic performance, with the sensitivity exceeding 80% for early HCC detection in a multi-national study that enrolled over 6,500 participants from Germany, UK, Hong Kong, and Japan.<sup>11</sup> GALAD also achieved satisfactory diagnostic accuracy in the multi-center cohort with the nonalcoholic steatohepatitis (NASH)-related HCC patients.<sup>13</sup> However, our previous study indicated that the diagnostic performance of DCP was higher than that of GALAD in HBV-positive patients in Chinese populations (AUC of DCP: 0.904, 95% CI: 0.865–0.934; GALAD: 0.813, 95% CI: 0.765–0.855; P<0.05).<sup>14</sup> This finding indicated that the etiology as well as the ethnicity were among the important factors influencing the diagnostic accuracy and sensitivity.<sup>14</sup> Therefore, the original GALAD score and cutoff value might not be perfectly suitable for Chinese HCC patients with HBV infection background.

In the present study, the  $\beta$  coefficients of each variable (gender, age, AFP, AFP-L3, and DCP) from the GALAD score were retrained and used to assign new nomogram points named Chinese-GALAD, C-GALAD (Figure 1B). We retrospectively compared the diagnostic performances of our C-GALAD and the GALAD score, the results demonstrated that the C-GALAD nomogram was better than the GALAD in Chinese populations (Training cohort: C-GALAD: 0.952, 95%)

CI: 0.947–0.956; GALAD: 0.925, 95% CI: 0.919–0.931; Internal validation cohort: C-GALAD: 0.908, 95% CI: 0.895–0.919; GALAD: 0.866, 95% CI: 0.851–0.879) (Figure 2A and E, Tables 3 and S3).

However, the most suitable combination of serum biomarkers for detecting HCC in China has not been thoroughly investigated. In our newly constructed HCC risk estimation nomogram (TAGALAD), the TBIL and ALB were found to be among risk factors for HCC in this study (Figure 1C). TBIL and ALB are important indicators for liver metabolic and synthetic functions. They are frequently used for liver function assessment and are among the important parameters in Child-Pugh grading.<sup>37</sup> Although the variables in our TAGALAD nomograms partially overlapped with the BALAD-2<sup>11</sup> score, using the TBIL and ALB together with the five indicators of C-GALAD/GALAD, we obtained the best diagnostic efficacy among these diagnostic algorithms in both the training, internal and external validation cohort (AUCs of Training cohort: 0.969, 95% CI: 0.965–0.973; Internal validation cohort: 0.964, 95% CI: 0.956–0.972; External validation cohort: 0.931, 95% CI: 0.895–0.958) (Figure 2A, E, and H; Tables 3 and <u>S3</u>).

The associated risk factors of the long-term clinical outcome of HCC are scarcely understood.<sup>14</sup> In our current study, we successfully validated the OS and RFS prediction values of our newly developed nomograms in the longitudinal follow-up validation cohort. Our three algorithm models (LAD, C-GALAD, and TAGALAD) all could stratify HCC populations into two distinct subgroups with high or low risk of OS and RFS. As shown in Figure 3, most patients in the high-risk subgroup suffered a poor prognosis after resection. Conversely, the low-risk populations have a consistently high overall survival rate and recurrence-free survival rate. Thus, if identified earlier, the high-risk subjects could receive active surveillance and intensified regimens to achieve a prolonged survival period.<sup>36</sup>

In addition to retrospective analysis, our newly developed nomogram was applied to a longitudinal (144 weeks) prospective cohort of chronic hepatitis B patients who might progress into HCC. The AUC results from TAGALAD, C-GALAD, and LAD demonstrated their capability to predict HCC up to 48 weeks prior to confirmed diagnosis. Importantly, at -48, -24, and 0 weeks, TAGALAD consistently showed the highest performance among the four HCC diagnostic algorithms, with AUCs of 0.984 (95% CI: 0.954–1.000), 0.900 (95% CI: 0.740–1.000), and 0.996 (95% CI: 0.900–1.000), respectively.

Evidences from both cross-sectional and longitudinal real-world data support the feasibility and effectiveness of our newly developed nomograms in the early detection and monitoring of HCC. However, this study has some limitations. Despite a considerable sample size, most samples came from case–control studies. Moreover, detection methods were limited to one defined system from the same manufacturer. Future validation of clinical feasibility necessitates larger cohorts in longitudinal and prospective studies using various detection systems.

#### Conclusion

In conclusion, the proposed nomograms generated from real-world Chinese populations are easy-to-use for risk estimation and prognostic prediction for HCC. The three nomograms (LAD, C-GALAD, and TAGALAD) not only could significantly outperform the GALAD score for the detection of HCC, but also were closely related to the HCC prognosis independently. While further validation in larger-scale longitudinal prospective cohorts assessing the benefit-to-harm ratio for blood-based surveillance is necessary, our results show the promise that blood-based tests could improve outcomes substantially for at-risk patients undergoing HCC surveillance.

#### **Data Sharing Statement**

Data sources and handling of the publicly available datasets used in this study are described in the Materials and Methods. Further details and other data that support the findings of this study are available from the corresponding authors upon request.

#### **Statement of Ethics**

The data obtained for this study complied with relevant data protection and privacy legislation. This study complied with the Declaration of Helsinki. This study has been approved by the Ethics Committee of Shanghai Oriental Hepatobiliary Surgery Hospital (EHBHKY2018-1-010). All enrolled patients have signed informed consent forms.

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# Disclosure

The authors report no conflicts of interest in this work.

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