

Micronutrient Requirements for Growth and Hydrocarbon Production in the Oil Producing Green Alga *Botryococcus braunii* (Chlorophyta)

Liang Song¹, Jian G. Qin^{1*}, Shengqi Su², Jianhe Xu³, Stephen Clarke⁴, Yichu Shan⁵

1 School of Biological Sciences, Flinders University, Adelaide, Australia, **2** School of Animal Science and Technology, Southwest University, Chongqing, P. R. China, **3** Key Laboratory of Marine Biotechnology of Jiangsu Province, Huaihai Institute of Technology, Lianyungang, P. R. China, **4** School of Chemistry, Physics and Earth Sciences, Flinders University, Adelaide, Australia, **5** CAS Key Laboratory of Separation Science for Analytical Chemistry, National Chromatographic Research and Analysis Centre, Dalian Institute of Chemical Physics, Chinese Academy of Sciences, Dalian, P. R. China

Abstract

The requirements of micronutrients for biomass and hydrocarbon production in *Botryococcus braunii* UTEX 572 were studied using response surface methodology. The concentrations of four micronutrients (iron, manganese, molybdenum, and nickel) were manipulated to achieve the best performance of *B. braunii* in laboratory conditions. The responses of algal biomass and hydrocarbon to the concentration variations of the four micronutrients were estimated by a second order quadratic regression model. Genetic algorithm calculations showed that the optimal level of micronutrients for algal biomass were 0.266 μM iron, 0.707 μM manganese, 0.624 μM molybdenum and 3.38 μM nickel. The maximum hydrocarbon content could be achieved when the culture media contained 10.43 μM iron, 6.53 μM manganese, 0.012 μM molybdenum and 1.73 μM nickel. The validation through an independent test in a photobioreactor suggests that the modified media with optimised concentrations of trace elements can increase algal biomass by 34.5% and hydrocarbon by 27.4%. This study indicates that micronutrients play significant roles in regulating algal growth and hydrocarbon production, and the response surface methodology can be used to optimise the composition of culture medium in algal culture.

Citation: Song L, Qin JG, Su S, Xu J, Clarke S, et al. (2012) Micronutrient Requirements for Growth and Hydrocarbon Production in the Oil Producing Green Alga *Botryococcus braunii* (Chlorophyta). PLoS ONE 7(7): e41459. doi:10.1371/journal.pone.0041459

Editor: Terence Evens, US Dept. of Agriculture – Agricultural Research Service (USDA-ARS), United States of America

Received: January 28, 2012; **Accepted:** June 25, 2012; **Published:** July 25, 2012

Copyright: © 2012 Song et al. This is an open-access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Funding: These authors have no support or funding to report.

Competing Interests: The authors have declared that no competing interests exist.

* E-mail: jian.qin@flinders.edu.au

Introduction

Microalgae have recently been receiving much attention in an attempt to explore their use as a potential feedstock for biofuel production [1,2]. *Botryococcus braunii* is a green colonial microalga found in freshwater lakes, reservoirs, and ponds [3,4] and is classified into A, B and L races depending on the type of hydrocarbons synthesized [5]. Race A produces C₂₃–C₃₃ odd numbered *n*-alkadienes, mono-, tri-, tetra-, and pentaenes and race B produces C₃₀–C₃₇ triperpenes while race L produces C₄₀ tetraerpenes [5]. This species is characterised by a conspicuous ability to synthesise and accumulate a variety of hydrocarbons [6,7,8]. These hexane-soluble hydrocarbons have the potential to be converted into biofuels by catalytic cracking [9]. However, the great variation of hydrocarbon content in *B. braunii* (0.1–86% of dry weight) provides an opportunity to explore the optimal growing conditions to maximise hydrocarbon production for a given *B. braunii* strain [10,11,12]. Therefore, it is necessary to identify the most efficient growing conditions for sustainable mass and hydrocarbon production in *B. braunii*.

The requirements for macronutrients by *B. braunii* have been intensively studied in the past a few decades. Largeau *et al.* [13] pointed out that the phosphorus (0.46 mM) in the Chu 13 medium was not limiting through the stationary growth phase in *B. braunii*, while the nitrogen concentration of 0.5 mM NO₃⁻ is only adequate

to sustain the growth of *B. braunii* for 10 days and the initial concentration of 8 mM NO₃⁻ is required to maintain the growth of *B. braunii* for 35 days. Ammonia can inhibit botryococcene biosynthesis in the *B. braunii* race B [14], but the replacement of nitrite nitrogen for nitrate nitrogen benefits the growth of race A *B. braunii* [15]. Air enriched with 1% CO₂ can enhance algal growth by doubling algal biomass and achieving 5-fold hydrocarbon production compared to aeration without CO₂ enrichment [16]. Dayanada *et al.* [17] reported that the N: P ratio played a significant role in both biomass and hydrocarbon production in *B. braunii* and the N: P ratio of 1:4 by weight favoured hydrocarbon production while the N:P ratio of 1:0.5 by weight increased the yield of algal biomass.

Given the depth of understanding in the growth requirement for macronutrients in *B. braunii*, it is surprising that the requirements for trace elements are little known. Trace elements such as iron, molybdenum and manganese can play critical roles in a variety of metabolic pathways involving utilization of light, nitrogen, phosphorus, and CO₂ [18,19]. Among trace elements, iron is essential for photosynthetic electron transport, respiratory electron transport, nitrate and nitrite reduction, and detoxification of reactive oxygen species [20,21,22]. Mojaat *et al.* [23] demonstrated that the addition of iron to the *Dunaliella salina* culture medium stimulated β -carotene production. The iron enrichment in the *Chlorella vulgaris* culture could increase algal growth and lipid

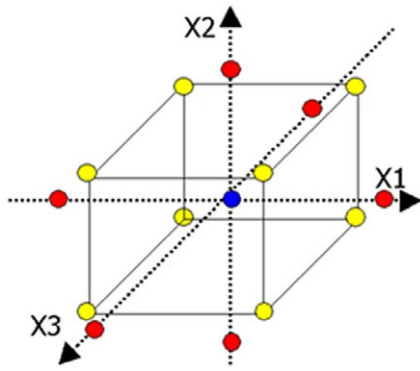


Figure 1. Illustration of the central composite design (only 3 out of the 4 dimensions are shown).
doi:10.1371/journal.pone.0041459.g001

accumulation [24], where the total lipid content of algae grown in the medium supplemented with 1.2×10^{-5} M FeCl_3 reached 56.6% of the dry biomass, which was a 3–7 fold increase compared to the medium without iron enrichment. Manganese is another important component in algal photosynthesis and also presents in enzymes to remove toxic superoxide radicals to sustain algal growth [25]. Chernikova *et al.* [26] reported that manganese (MnCl_2) enhanced the capacity to accumulate inorganic minerals and catalysed protein synthesis in *Spirulina platensis*. Molybdenum is coupled with iron in the enzymes for nitrate reduction, and its deficiency diminishes the nitrate uptake mechanism and interferes with lipid synthesis [27]. Nickel can facilitate nitrogen uptake to enhance the growth of *Thalassiosira weissflogii* when urea is the nitrogen source, suggesting the positive role of Ni in enhancing algal growth [28]. Berges *et al.* [29] also reported that the addition of nickel and molybdenum to the algal culture medium increased the overall primary productivity. Coincidentally, in a field survey, Wake and Hillen [3] found that wherever the *B. braunii* bloom occurred in the Darwin River reservoir, the nickel concentration in the environment was always higher than that in adjacent water bodies where no *B. braunii* bloomed, suggesting this trace element may trigger the occurrence of *B. braunii*. However, no laboratory testing has been conducted so far to test the need of nickel to enhance the growth of *B. braunii* in the laboratory since the early field survey work of Wake and Hillen’s in the 1980’s.

Optimization of micronutrient requirements is an important undertaking prior to the establishment of sustainable production of *B. braunii* on a large scale. The conventional method to optimise the level of multiple nutrients in algal culture has been focussed on one-factor-at-a-time approach, studying the effect of one nutrient on the response of algae by keeping the other nutrients constant. However, this approach is time consuming and does not take into account interactions between nutrients, which usually results in poor optimization results [30,31].

Techniques in experimental design are critical to identify key nutrients required for algal growth. In this study we used the response surface methodology (RSM) [32] to explore the requirement of micronutrients in the culture of *B. braunii* because the RSM approach can optimise the nutrient requirement with low input of time and resources [33,34,35]. This approach has been widely used in optimization of plant nutrients [36,37], bacterial medium composition [38], enzymatic hydrolysis [39,40], synthesis of polymers [41], food processing [42,43] and operation conditions for photobioreactors [44]. The RSM approach has also been used for medium optimisation in algal culture. Azma *et al.* [45] optimised the culture medium for *Tetraselmis suecica* by RSM

Table 1. Coded and actual values of experimental variables used in the central composite experimental design.

Independent variables	Symbols	Levels				
		-1.72*	-1	0	1	1.72*
Fe (μM)	x_1	0.03	2.39	5.35	8.31	10.44
Mn (μM)	x_2	0.02	2.67	6.36	10.05	12.70
Mo (μM)	x_3	0	0.13	0.31	0.50	0.62
Ni (μM)	x_4	0	0.71	1.69	2.68	3.39

*Alpha values used for the axial points in this study.
doi:10.1371/journal.pone.0041459.t001

Table 2. Central composite design matrix and the responses of biomass and hydrocarbon production to Fe (x_1), Mn (x_2), Mo (x_3) and Ni (x_4).

Runs	Independent variables				Responses	
	Coded levels				Biomass (g/L)	Hydrocarbon (% w/w)
	x_1	x_2	x_3	x_4		
1	1	1	1	1	0.246	14.82
2	-1	-1	1	1	0.292	14.31
3	1	-1	-1	1	0.251	15.45
4	-1	1	-1	1	0.296	14.56
5	1	-1	1	-1	0.124	13.99
6	-1	1	1	-1	0.120	13.42
7	1	1	-1	-1	0.136	14.83
8	-1	-1	-1	-1	0.125	13.86
9	1	-1	1	1	0.257	13.96
10	-1	1	1	1	0.320	14.12
11	1	1	-1	1	0.248	14.19
12	-1	-1	-1	1	0.306	14.00
13	1	1	1	-1	0.116	13.96
14	-1	-1	1	-1	0.121	15.26
15	1	-1	-1	-1	0.105	14.68
16	-1	1	-1	-1	0.126	13.96
17	1.72	0	0	0	0.215	20.23
18	-1.72	0	0	0	0.231	19.24
19	0	1.72	0	0	0.123	12.25
20	0	-1.72	0	0	0.121	11.59
21	0	0	1.72	0	0.118	18.57
22	0	0	-1.72	0	0.124	20.18
23	0	0	0	1.72	0.289	12.54
24	0	0	0	-1.72	0.094	11.90
25*	0	0	0	0	0.124	19.31
26*	0	0	0	0	0.120	18.46
27*	0	0	0	0	0.123	19.17
28*	0	0	0	0	0.127	20.13
29*	0	0	0	0	0.122	19.74
30*	0	0	0	0	0.126	18.45

*Central point values contributing to the degree of freedom for pure error calculation.

doi:10.1371/journal.pone.0041459.t002

Table 3. Analysis of variance (ANOVA) for the fitted quadratic polynomial regression model for optimization of the algal biomass production.

Source	Sum of squares	df	Mean square	F-value	Probability P-value
Model	0.162049	14	0.011575	31.64	<0.001
Residual	0.005488	15	0.000366		
Lack of fit	0.005354	10	0.000535	20.08	0.002
Pure error	0.000133	5	0.000027		
Cor. total	0.167537	29			
$R^2 = 0.967$					
Adj. $R^2 = 0.937$ Pred. $R^2 = 0.824$					

doi:10.1371/journal.pone.0041459.t003

and increased algal production by two times. Similarly, by using RSM, Isleten-Hosoglu *et al.* [46] optimised the carbon and nitrogen concentrations for *Chlorella saccharophila* and improved biomass production by 7.7 fold.

The objectives of this study were to (1) estimate the roles of the four micronutrients iron, manganese, molybdenum, and nickel in regulating the responses of algal biomass and hydrocarbon, and (2) identify the optimum requirements of micronutrients for the cultivation of *B. braunii* to maximise hydrocarbon production.

Methods

Materials and Procedures

Botryococcus braunii UTEX 572 was obtained from the University of Texas Culture Collection, USA. The basic macronutrients for algal growth were adapted from the Bold 3N medium, which also contains micronutrients including 5.35 μM Fe, 6.36 μM Mn, and 0.31 μM Mo [47]. All chemicals were of analytical reagent grade. To avoid the effect of other unknown trace elements, soil residuals were not added into the medium in this study. The experiment for model construction was conducted at $24 \pm 1^\circ\text{C}$ with illumination provided by fluorescent lights at $150 \mu\text{mol}/\text{m}^2/\text{s}$ at 12 h light and 12 h dark. The algal growth experiments lasted 3 weeks.

The dry weight of algal cells was measured by vacuum filtration onto pre-weighed Whatman[®] GF/C filters [48]. The filters with algal cells were freeze-dried, weighed, and expressed as algal biomass (g/L). Hydrocarbons in dry biomass were extracted on glass filters using η -hexane [48]. Solvents were removed from the extracts by a rotary evaporator and the residues were rinsed with η -hexane. Hydrocarbon fractions were purified by passing the samples through an alumina gel plug and eluting with η -hexane.

Solvents were evaporated under a stream of nitrogen to dry, and the pure hydrocarbon fractions were measured gravimetrically and expressed as hydrocarbon content (% w/w).

Experimental Design

Central composite design (CCD) is one type of RSM approach [49] which allows estimating the polynomial regression between independent variables and dependant variables [50]. In this study, a 2^4 CCD with 24 runs and six replications of the centre points were used to determine the optimal concentrations of iron, manganese, molybdenum, and nickel on the yield of algal biomass and hydrocarbon production (Fig. 1). The coded and corresponding actual values are given in Table 1. The corresponding central composite experimental design and their values are shown in Table 2. All the design points except the centre point (0, 0, 0, 0) were run in three replications. Due to the restriction of modeling protocol, only one mean value of the three replicates for each

Table 5. Concentration of micronutrients in different algal culture media.

Culture media	Micronutrients (μM)			
	Fe	Mn	Mo	Ni
Original Bold 3N	2.150	1.240	0.099	0.00
Modified Bold 3N-1	0.276	0.707	0.624	3.38
Modified Bold 3N-2	10.430	6.530	0.012	1.73

doi:10.1371/journal.pone.0041459.t005

Table 4. Analysis of variance (ANOVA) for the fitted quadratic polynomial regression model for optimization of the hydrocarbon production.

Source	Sum of squares	df	Mean square	F-value	Probability P value
Model	218.69	14	15.621	36.58	<0.001
Residual	6.406	15	0.427		
Lack of fit	4.127	10	0.413	0.91	0.584
Pure error	2.279	5	0.456		
Cor. total	225.096	29			
$R^2 = 0.972$					
Adj. $R^2 = 0.945$ Pred. $R^2 = 0.875$					

doi:10.1371/journal.pone.0041459.t004

Table 6. Results of regression analysis of the full second-order polynomial model for optimization of algal biomass production with Fe (x_1), Mn (x_2), Mo (x_3) and Ni (x_4).

Model term	Coefficients estimated	P-value	t-Statistic
intercept	0.2196	<0.001	5.04
x_1	-0.0433	<0.001	-5.93
x_2	-0.0036	0.547	0.55
x_3	-0.1471	0.249	-1.20
x_4	-0.0058	0.795	-0.26
x_1x_2	-0.0001	0.999	-0.00
x_1x_3	0.0021	0.813	0.24
x_1x_4	-0.0043	0.019	-2.62
x_2x_3	0.0001	0.992	0.01
x_2x_4	-0.0003	0.798	-0.26
x_3x_4	0.0149	0.581	0.56
x_1^2	0.0044	<0.001	8.64
x_2^2	0.0004	0.290	1.10
x_3^2	0.1703	0.269	1.15
x_4^2	0.0294	<0.001	6.22

doi:10.1371/journal.pone.0041459.t006

dependent variable was allowed to enter the model. Therefore, the degree of freedom of the triplicate for each non-centrepoint could not be used for pure error calculation. Experiments were repeated six times at the central point to provide an estimate of pure error [51,52,53,54] thus providing adequate degree of freedom ($df=5$) for pure error calculation (Tables 3 and 4).

Data from the CCD experiment were analysed by RSM. A mathematical model with a second-order polynomial regression

Table 7. Results of regression analysis of the full second-order polynomial regression model for optimization of hydrocarbon production with Fe (x_1), Mn (x_2), Mo (x_3) and Ni (x_4).

Model term	Coefficients estimated	P-value	t-Statistic
intercept	4.4600	<0.001	3.00
x_1	-0.0082	0.974	-0.03
x_2	2.3089	<0.001	11.46
x_3	1.7040	0.690	0.41
x_4	8.4303	<0.001	11.17
x_1x_2	0.0062	0.683	0.42
x_1x_3	-0.3608	0.245	-1.21
x_1x_4	0.0100	0.861	0.18
x_2x_3	-0.0720	0.768	-0.30
x_2x_4	0.0273	0.554	0.61
x_3x_4	-0.1031	0.911	-0.11
x_1^2	0.0120	0.497	0.70
x_2^2	-0.1865	<0.001	-16.22
x_3^2	-0.3860	0.940	-0.08
x_4^2	-2.5090	<0.001	-15.54

doi:10.1371/journal.pone.0041459.t007

was developed to describe the relationships between the predicted response variables (biomass or hydrocarbon) and the independent variables (Fe, Mn, Mo and Ni). The regression equation was described as follows (Eq. 1):

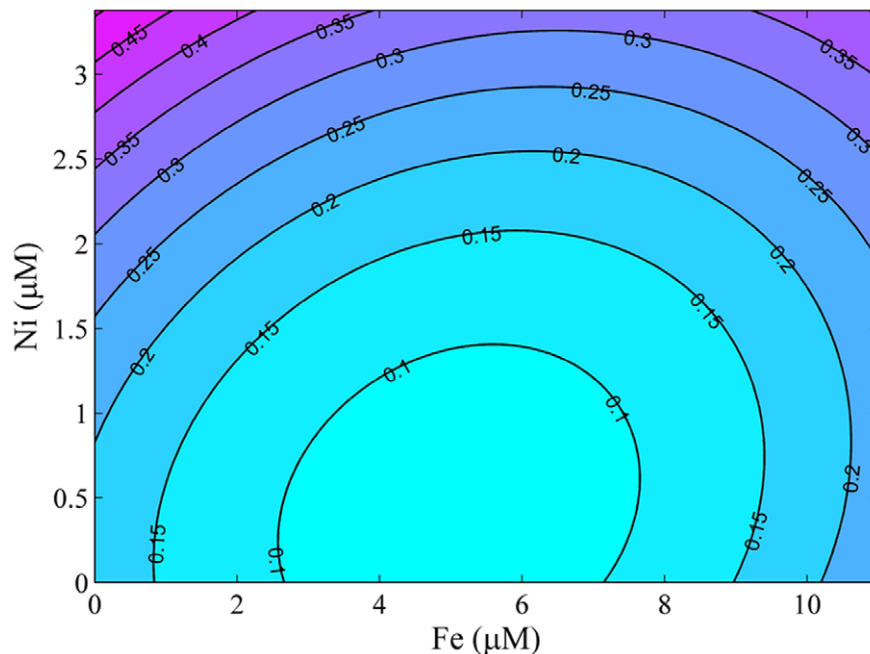


Figure 2. Contour plot showing biomass prediction from Fe (x_1) Ni (x_4) with other independent variables Mn (x_2) and Mo (x_3) being constant.

doi:10.1371/journal.pone.0041459.g002

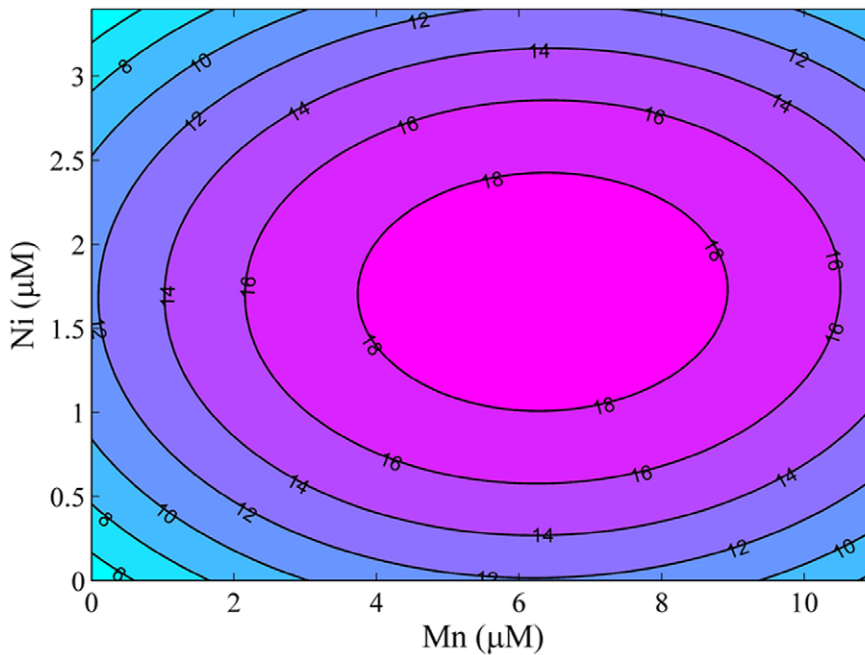


Figure 3. Contour plot showing hydrocarbon prediction from Mn (x_2) and Ni (x_4) with other independent variables Fe (x_1) and Mo (x_3) being constant.

doi:10.1371/journal.pone.0041459.g003

$$y = \beta_0 + \sum_{i=1}^4 \beta_i x_i + \sum_{i=1}^4 \beta_{ii} x_i^2 + \sum_{i,j=1}^4 \beta_{ij} x_i x_j \quad (1)$$

where y is the predicted response variables (biomass or hydrocarbon production); β_0 is a constant, β_i is the linear coefficient, β_{ii} is the quadratic coefficients, β_{ij} is the interaction coefficients of the model, respectively; x_i and x_j ($i = 1, 4; j = 1, 4; i \neq j$) represent the non-coded independent variables (micronutrient concentrations).

Model Validation

The predicted models on algal biomass and hydrocarbon production of *B. braunii* were validated in an independent experiment using optimized micronutrient concentrations from the genetic algorithms calculations [55]. A flat plate photobioreactor (3.2 L) was used as the culture vessel under a light intensity of $300 \mu\text{mol}/\text{m}^2/\text{s}$ and a mixing rate of 1.10 L/L/min. The *B. braunii* cells were separately inoculated into the original Bold 3N medium, the modified Bold 3N-1 medium for producing algal biomass, and the modified Bold 3N-2 for producing hydrocarbon with different micronutrient compositions (Table 5). The experimental protocols in the validation study were the same as those in the model construction. Algal biomass and hydrocarbon content were separately measured at 3-day intervals over 12 days to assess the response of algal performance to modified media. The productivities of algal biomass and hydrocarbon during the experimental period were also calculated and expressed as g/L/day. All data points in the figures were the mean of three replicates to provide a better estimate of the response of each dependent variable.

Statistical Analysis

The data analyses for model construction were performed with MINITAB 16, based on the response surface methodology. The F -

test for the analysis of variance (ANOVA) was performed on experimental data to evaluate the statistical significance of the model. The significance of regression coefficients was evaluated using t -test. The contour plots described by the regression model were drawn using MATLAB 7 to illustrate the effects of the independent variables and interactive effects of each independent variable on the response variables.

Optimisation of nutrient composition in the medium was determined by the procedure of genetic algorithms (MATLAB 7), which is a computer simulation program based on the best fit theory of natural selection to generate optimal solutions to problems [55]. In simulations, the program selected the best-fit concentration of each nutrient to maximise the algal response such as biomass and hydrocarbon production. In the validation experiment, data from the original 3N medium and modified medium were analysed by quadratic regression to compare the significant differences of curves. The probability level for significant difference was set at $P < 0.05$.

Results and Discussion

Model Fitting

The application of RSM yielded the following regression equations for biomass (Eq. 2) and hydrocarbon production (Eq. 3). A central composite design (CCD) with five coded levels for all the four factors: iron, manganese, molybdenum, and nickel were used for model simulations. The range of variables, experimental designs and results for biomass and hydrocarbon production are presented in Table 2. The second order polynomial regression equations were used to fit the dependent variables (\hat{Y}_{biomass} and $\hat{Y}_{\text{hydrocarbon}}$) to the independent variables x_1 (iron), x_2 (manganese), x_3 (molybdenum) and x_4 (nickel).

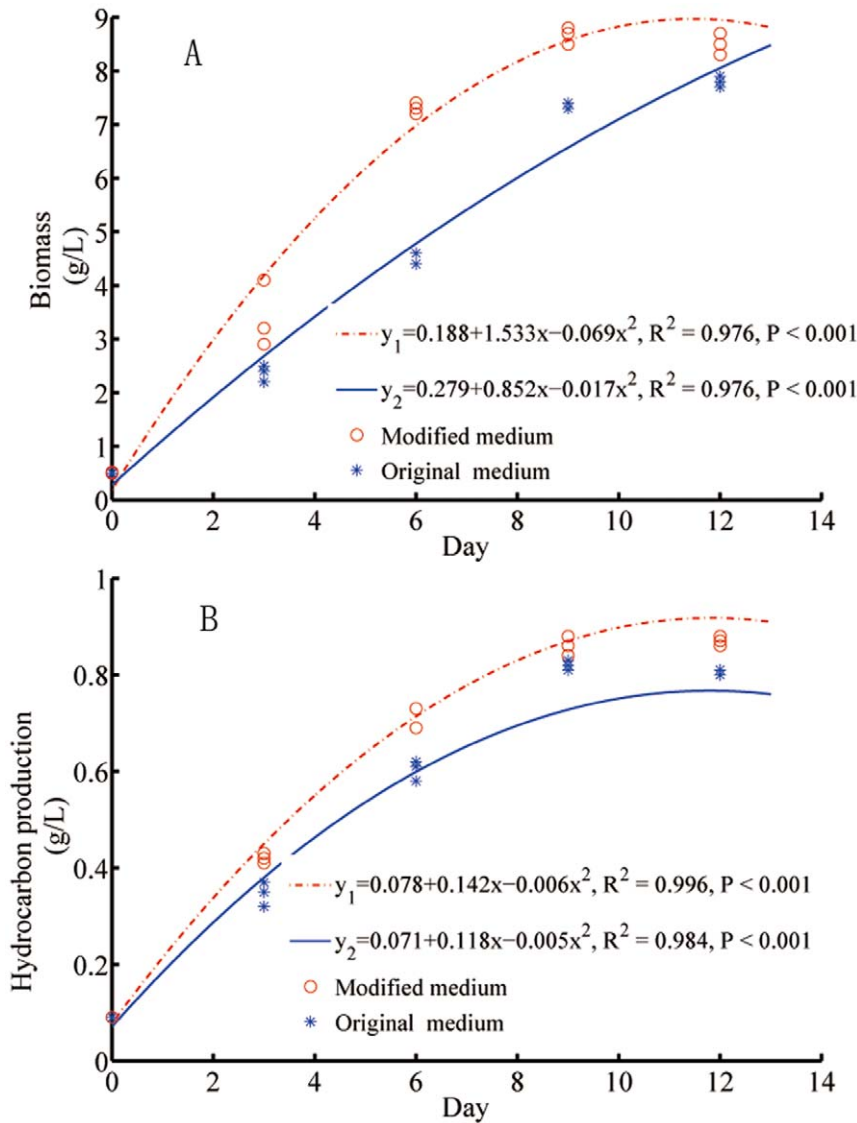


Figure 4. Regression plots of biomass (A) and hydrocarbon (B) productions in the modified and original Bold 3N media.
doi:10.1371/journal.pone.0041459.g004

$$\begin{aligned}
 Y_{biomass} = & 0.2196 - 0.0433x_1 - 0.0036x_2 - 0.1470x_3 \\
 & - 0.0058x_4 - 0.0001x_1x_2 + 0.0021x_1x_3 - 0.0043x_1x_4 \\
 & + 0.0001x_2x_3 - 0.0003x_2x_4 + 0.0150x_3x_4 + 0.0044x_1^2 \\
 & + 0.0004x_2^2 + 0.1703x_3^2 + 0.0294x_4^2
 \end{aligned} \tag{2}$$

$$\begin{aligned}
 Y_{hydrocarbon} = & 4.4600 - 0.0082x_1 + 2.3089x_2 + 1.7040x_3 \\
 & + 8.4303x_4 + 0.0062x_1x_2 - 0.3608x_1x_3 + 0.0100x_1x_4 \\
 & - 0.0720x_2x_3 + 0.0273x_2x_4 - 0.1031x_3x_4 + 0.0120x_1^2 \\
 & - 0.1865x_2^2 - 0.3860x_3^2 - 2.5090x_4^2
 \end{aligned} \tag{3}$$

The significance and adequacy of the regression model were tested using ANOVA. These two regression models could

significantly predict algal biomass ($P < 0.001$) and hydrocarbon production ($P < 0.001$) from the four micronutrients (Tables 3 and 4). The predicted R^2 (0.824 for Eq. 2 and 0.875 for Eq. 3) agreed well with the adjusted model R^2 (0.937 for Eq. 2 and 0.945 for Eq. 3), suggesting a close correlation between the observed values and the predicted values. Therefore, we can use the regression models to predict algal biomass and hydrocarbon production from the amount of micronutrients in the culture medium.

Effect of Micronutrients on Algal Biomass

The regression coefficients of the model for biomass prediction are presented in Table 6. The linear effect of x_j and the quadratic effect of x_j^2 and x_i^2 had significant effects ($P < 0.001$) on $Y_{bioamss}$ followed by the interaction effect of x_1x_4 ($P = 0.019$). Other terms of the model had no significant effect on $Y_{bioamss}$. Negative coefficients of x_j and interaction term x_1x_4 decreased $Y_{bioamss}$. However, the quadratic terms of x_j^2 and x_i^2 had positive effects on $Y_{bioamss}$.

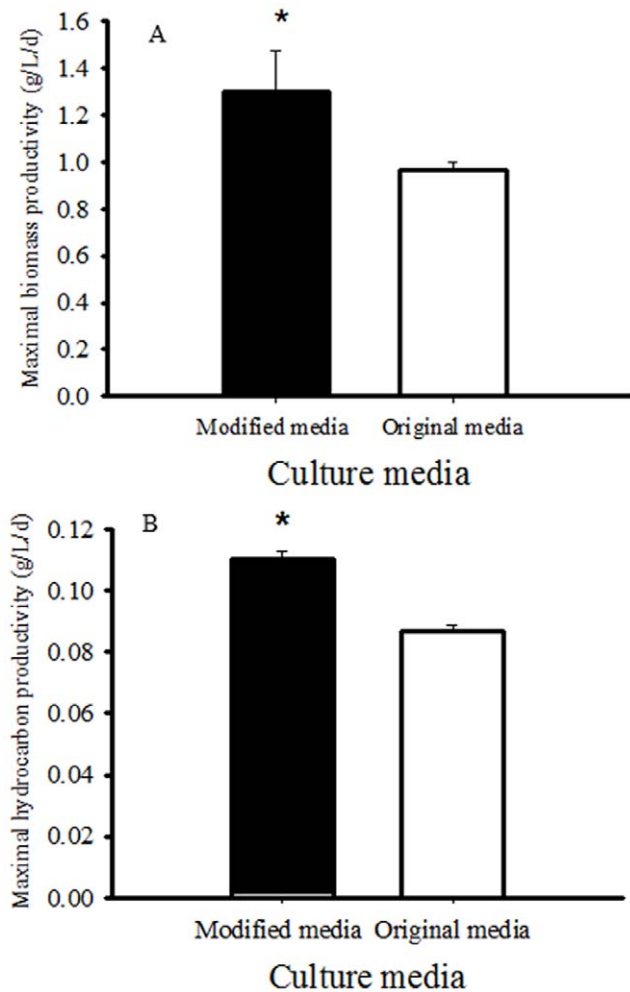


Figure 5. Comparison of maximal biomass (A) and hydrocarbon (B) productivities in the modified and original Bold-3N media.

doi:10.1371/journal.pone.0041459.g005

The interaction between two independent variables (Fe and Ni) and the response variable (biomass) was shown by the contour plots generated by keeping the independent variables (Mn and Mo) as constants (Fig. 2). The algal biomass was sensitive to the change of Fe and Ni concentrations. As the concentration of Ni increased, algal biomass increased progressively. The Fe in the medium at either low or high concentrations increased algal biomass when Ni concentrations were high.

In this study, the positive relationship between algal biomass and Ni concentrations corroborates an early report by Wake and Hillen [3] that the *B. braunii* bloom occurred in waters with the nickel concentration of 0.1 mg/L. In other studies, however, nickel accumulation in cells has been shown to cause a detrimental effect on algal growth as nickel is toxic to some physiological processes [56]. Wong *et al.* [57] reported that both *Chlorella vulgaris* and *Chlorella miniata* were capable of cell division after being treated with wastewater containing nickel for 24 h, but the growth rate was reduced in proportion to the concentrations of nickel in the wastewater. Despite this inhibition effect of nickel on other algal species, the present study does suggest that the use of nickel stimulated the growth of *B. braunii*.

Effect of Micronutrients on Hydrocarbon Production

The regression coefficients of the model for hydrocarbon production are presented in Table 7. The linear effect of x_2 and x_4 , and the quadric effect of x_2^2 and x_4^2 had significant effects ($P < 0.001$) on $Y_{\text{hydrocarbon}}$. Other terms of the model had no significant effect on $Y_{\text{hydrocarbon}}$. Positive coefficient of x_2 and x_4 indicated their role to enhance $Y_{\text{hydrocarbon}}$. However, the quadratic terms of x_2^2 and x_4^2 had negative effects on $Y_{\text{hydrocarbon}}$.

The interaction effects of two independent variables (Mn and Ni) on the response variable (hydrocarbon) are shown by the contour plots generated by keeping the independent variables (Fe and Mo) as constants (Fig. 3). Hydrocarbon production was more sensitive to the change of Mn and Ni concentrations. An increase in hydrocarbon production was observed with the increase of Mn concentrations. But this trend was reversed when the Mn concentration was above 9 μM . The effect of Ni on $Y_{\text{hydrocarbon}}$ followed the similar trend. With the increase of Ni concentration, $Y_{\text{hydrocarbon}}$ firstly increased and then decreased as a result of excessive Ni concentration. The circular profile of the contour plots indicated that the interaction between the Mn and Ni concentrations on hydrocarbon was negligible (Fig. 3).

The composition of the culture medium affects not only algal productivity, but also secondary metabolites [58]. This finding was consistent with result of Wang *et al.* [59] who found that the increase of Fe and Mn concentrations stimulated the growth of blue green algae, while a further increase in their concentrations inhibited algal growth. Cloëz *et al.* [60] found that lipid synthesis increased by three times after adding manganese, copper and nickel at 2 mM. On the other hand, Mohammady and Fathy [61] reported that the total lipid content in *Dunaliella salina* cultivated in nickel supplemented media (0.5 mg/L NiCl_2) has reduced in comparison to the control. In another study, Rousch and Sommerfeld [62] found that manganese had stronger impact on the growth of a green alga (*Ulothrix* sp.) than nickel. However, in this study, both nickel and manganese regulated the production of hydrocarbon, though the algal biomass was only affected by nickel.

Optimisation of Micronutrients

The concentrations of these four micronutrients for producing algal biomass were optimized using the genetic algorithm calculation. The optimal medium for biomass consisted of 0.266 μM Fe, 0.707 μM Mn, 0.624 μM Mo and 3.38 μM Ni. By running the optimization simulation within the experimental range, the optimal medium for hydrocarbon production is recommended to contain 10.43 μM Fe, 6.53 μM Mn, 0.012 μM Mo and 1.73 μM Ni. It is worth noting that the optimal composition of these four micronutrients for algal biomass was different from that for hydrocarbon production. This difference highlights the importance of selecting culture medium to achieve different objectives in algal culture since the nutrient requirement differs for algae cell division and accumulation of secondary metabolites [63].

Validation of Algal Growth and Hydrocarbon Production

The reliability of nutrient requirement generated from the predicted models and the genetic algorithm calculations for biomass and hydrocarbon production in *B. braunii* were validated in an independent photobioreactor study. From day 3 to day 12, the algal biomass produced in the Bold 3N medium supplemented with 0.266 μM Fe, 0.707 μM Mn, 0.624 μM Mo, 3.38 μM Ni was significantly higher than that produced in the original Bold 3N medium ($P < 0.05$, Fig. 4A). The maximal algal biomass productivity (1.300 ± 0.176 g/L/day) in dry weight with modified media

was significantly higher than that (0.967 ± 0.033 g/L/day) in the original media ($P < 0.05$, Fig. 5A).

The hydrocarbon production of algae in the Bold 3N medium supplemented with $10.43 \mu\text{M}$ Fe, $6.53 \mu\text{M}$ Mn, $0.012 \mu\text{M}$ Mo and $1.73 \mu\text{M}$ Ni was significantly higher than that in the original medium from day 3 to day 12 ($P < 0.05$, Fig. 4B). The maximal hydrocarbon productivity (0.110 ± 0.003 g/L/day) in the modified media was significantly higher than that (0.087 ± 0.002 g/L/day) in the original media ($P < 0.05$, Fig. 5B).

The biomass and hydrocarbon productivity are key parameters affecting the economic feasibility of producing bioproducts from algae. The micronutrient concentrations optimised by modelling were validated in a photobioreactor, and the accuracy and reliability of the model in predicting nutrient requirements for producing algal biomass and hydrocarbon have been confirmed.

Conclusion

The application of response surface methodology (RSM) is a reliable approach to model and optimize the requirements for iron, manganese, molybdenum, and nickel in producing algal biomass and hydrocarbon in *B. braunii*. Nickel and iron played significant roles but manganese and molybdenum had a trivial role in algal biomass production. In contrast, nickel and manganese

were more important than molybdenum and iron in regulating algal hydrocarbon production. The production of algal biomass and production of hydrocarbon require different micronutrients in the culture medium. The recommended levels of micronutrients in the Bold 3N medium are $0.266 \mu\text{M}$ iron, $0.707 \mu\text{M}$ manganese, $0.624 \mu\text{M}$ molybdenum and $3.38 \mu\text{M}$ nickel for *B. braunii* biomass and $10.43 \mu\text{M}$ iron, $6.53 \mu\text{M}$ manganese, $0.012 \mu\text{M}$ and $1.73 \mu\text{M}$ nickel for hydrocarbon production. The model validation showed that by using modified algal culture media, algal biomass productivity increased 1.345 fold and hydrocarbon productivity increased 1.274 fold compared with the original Bold 3N medium without addition of the trace elements.

Acknowledgments

The authors would like to thank Dr. Daniel Jardine for his advice on chemical analysis and Dr. David Kehoe for commenting on the early draft manuscript.

Author Contributions

Conceived and designed the experiments: LS JGQ SC. Performed the experiments: LS. Analyzed the data: SS JX YS. Wrote the paper: LS JGQ JX SC.

References

- Chisti Y (2007) Biodiesel from microalgae. *Biotechnol Adv* 25: 294–306.
- Qin JG (2010) Hydrocarbons from algae. In: Timmis KN editor. *Microbiology of hydrocarbons, oils, lipids, and derived compounds*, vol 4: Consequences of microbial interactions with hydrocarbons, oils and lipids. Berlin: Springer, 2817–2816.
- Wake LV, Hillen LW (1980) Study of a “bloom” of the oil-rich alga *Botryococcus braunii* in the Darwin River reservoir. *Biotechnol Bioeng* 22: 1637–1656.
- Wake LV, Hillen LW (1991) Nature and hydrocarbon content of blooms of the alga *Botryococcus braunii* occurring in Australian freshwater lakes. *Aust J Mar Freshwater Res* 32: 353–367.
- Metzger P, Largeau C (2005) *Botryococcus braunii*: a rich source for hydrocarbons and related ether lipids. *Appl Microbiol Biot* 66: 486–496.
- Brown AC, Knights BA, Conway E (1969) Hydrocarbon content and its relationship to physiological state in green alga *Botryococcus braunii*. *Phytochemistry* 8: 543–547.
- Knights BA, Brown AC, Conway E, Middleditch BS (1970) Hydrocarbons from the green form of the freshwater alga *Botryococcus braunii*. *Phytochemistry* 9: 1317–1324.
- Li Y, Qin JG (2005) Comparison of growth and lipid content in three *Botryococcus braunii* strains. *J Appl Phycol* 17: 551–556.
- Hillen L, Pollard G, Wake LV, White N (1982) Hydrocracking of the oils of the alga *Botryococcus braunii* to transport fuels. *Biotechnol Bioeng* 24: 193–205.
- Qin JG (2005) Bio-hydrocarbons from algae: impacts of temperature, light and salinity on algae growth. *Development* 5: 1–26.
- Qin JG, Li Y (2006) Optimization of the growth environment of *Botryococcus braunii* strain CHN 357. *J Freshwater Ecol* 21: 169–176.
- Metzger P, Largeau C (1999) Chemicals of *Botryococcus braunii*. In Cohen Z editor. *Chemicals from microalgae*. London: Taylor & Francis. 205–260.
- Largeau C, Casadevall E, Berkaloff C, Dhamelincoourt P (1980) Sites of accumulation and composition of hydrocarbons in *Botryococcus braunii*. *Phytochemistry* 19: 1043–1051.
- Ohmori M, Wolf FR, Bassham JA (1984) *Botryococcus braunii*: carbon/nitrogen metabolism as affected by ammonia addition. *Arch Microbiol* 140: 101–106.
- Yang SL, Wang J, Cong W, Cai ZL, Fan OY (2004) Utilization of nitrite as a nitrogen source by *Botryococcus braunii*. *Biotechnol Lett* 26: 239–243.
- Chirac C, Casadevall E, Largeau C, Metzger P (1985) Bacterial influence upon growth and hydrocarbon production of the green alga *Botryococcus braunii*. *J Phycol* 21: 380–387.
- Dayananda C, Sarada R, Bhattacharya S, Ravishankar GA (2005) Effect of media and culture conditions on growth and hydrocarbon production by *Botryococcus braunii*. *Process Biochem* 40: 3125–3131.
- Raven JA (1988) The iron and molybdenum use efficiencies of plant growth with different energy, carbon, and nitrogen sources. *New Phytol* 109: 279–287.
- Raven JA (1990) Predictions of Mn and Fe use efficiencies of phototrophic growth as a function of light availability for growth and C assimilation pathway. *New Phytol* 116.
- Maldonado MT, Price NM (1996) Influence of N substrate on Fe requirements of marine centric diatoms. *Mar Ecol Prog Ser* 141: 161–172.
- Sunda WG, Huntsman SA (1997) Interrelated influence of iron, light, and cell size on growth of marine phytoplankton. *Nature* 390: 389–92.
- Sunda WG, Huntsman SA (2004) Relationships among photoperiod, carbon fixation, growth, chlorophyll a, and cellular iron and zinc in a coastal diatom. *Limnol Oceanogr* 49: 1742–1753.
- Mojaat M, Pruvost J, Foucault A, Legrand J (2008) Effect of organic carbon sources and Fe^{2+} ions on growth and β -carotene accumulation by *Dunaliella salina*. *Biochem Eng J* 39: 177–184.
- Liu ZY, Wang GG, Zhou BC (2008) Effect of iron on growth and lipid accumulation in *Chlorella vulgaris*. *Bioresource Technol* 99: 4717–4722.
- Peers GS, Price NM (2004) A role for manganese in superoxide dismutases and the growth of iron-deficient diatoms. *Limnol Oceanogr* 49: 1174–1783.
- Chernikova AA, Tsoglin LN, Markelova AG, Zorin SN, Mazo VK, et al. (2006) Capacity of *Spirulina platensis* to accumulate manganese and its distribution in dell. *Russ J Plant Physl* 53: 800–806.
- Carvalho AP, Pontes I, Gaspar H, Malcata FX (2006) Metabolic relationships between macro- and micronutrients, and the cicosapentaenoic acid and docosahexaenoic acid contents of *Pavlova lutheri*. *Enzyme Microb Tech* 38: 358–366.
- Price NM, Morel FMM (1991) Co-limitation of phytoplankton growth by nickel and nitrogen. *Limnol Oceanogr* 36: 1071–1077.
- Berges JA, Franklin DJ, Harrison PJ (2001) Evolution of an artificial seawater medium: improvements in enriched seawater, artificial water over the last two decades. *J Phycol* 37: 1138–1145.
- Hernay-Ramirez J, Lampinen M, Vicente MA, Costa CA, Madeira LM (2008) Experimental design to optimize the oxidation of orange II dye solution using a clay-based Fenton-like catalyst. *Ind Eng Chem Res* 47: 284–294.
- Oliveira R, Almeida MF, Santos L, Madeira LM (2006) Experimental design of 2, 4-dichlorophenol oxidation by Fenton's reaction. *Ind Eng Chem Res* 45: 1266–1276.
- Myers RH, Montgomery DC (2002) Response surface methodology: process and product optimization using designed experiments, second (ed). New York: John Wiley & Sons. 704 p.
- Ren J, Lin WT, Shen YJ, Wang JF, Lou XC, et al. (2008) Optimization of fermentation media for nitrite oxidizing bacteria using sequential statistical design. *Bioresour Technol* 99: 7923–7927.
- Kammoun R, Naili B, Bejar S (2008) Application of a statistical design to the optimization of parameters and culture media for α -amylase production by *Aspergillus oryzae* CBS 819.72 grown on gruel (wheat grinding by-product). *Bioresour Technol* 99: 602–609.
- Pan CM, Fan YF, Xing Y, Hou HW, Zhang ML (2008). Statistical optimization of process parameters on biohydrogen production from glucose by *Clostridium* sp. *Fanp2. Bioresour Technol* 99: 3146–3154.
- De Rijk G, Schrevens E (1998) Multifactorial optimisation of the nutrient solution for hydroponically grown chichory plants. *Scientia Horti* 76: 149–159.
- Niedz RP, Hyndman SE, Evens TJ (2007) Using a gestalt to measure the quality of in vitro responses. *Sci Horti-Amsterdam* 112: 349–359.
- Rao KJ, Kim CH, Rhee SK (2000) Statistical optimization of medium for the production of recombinant hirudin from *Saccharomyces cerevisiae* using response surface methodology. *Process Biochem* 35: 639–647.

39. Kunamneni A, Singh S (2005) Response surface optimization of enzymatic hydrolysis of maize starch for higher glucose production. *Biochem Eng J* 27: 179–190.
40. Nilsang S, Lertsiri S, Suphantharika M, Assavanic A (2005) Optimization of enzymatic hydrolysis of fish soluble concentrate by commercial proteases. *J Food Eng* 70: 571–578.
41. Shieh CJ, Lai YF (2000) Application of response surface methodology to the study of methyl glucoside polyester synthesis parameters in a solvent-free system. *J Agr Food Chem* 48: 1124–1128.
42. Castro IA, Tirapegui J, Silva RSSF (2000) Protein mixtures and their nutritional properties optimized by response surface methodology. *Nutr Res* 20: 1341–1353.
43. Ozer EA, Herken EN, Guzel S, Ainsworth S, Ibanoglu S (2006) Effect of extrusion process on the antioxidant activity and total phenolics in a nutritious snack food. *Int J Food Sci Tech* 41: 289–293.
44. Jacob-lobes E, Lacerda LMCF, Franco TT (2008) Biomass production and carbon dioxide fixation by *Aphanothece microscopica nageli* in a bubble column photobioreactor. *Biochem Eng J* 40: 27–34.
45. Azma M, Mohamed MS, Mohamad R, Rahim RA, Ariff AB (2011) Improvement of medium composition for heterotrophic cultivation of green microalgae, *Tetraselmis suecica*, using response surface methodology. *Biochem Eng J* 53: 187–195.
46. Isleten-Hosoglu M, Gultepe L, Elibol M (2012) Optimization of carbon and nitrogen sources for biomass and lipid production by *Chlorella saccharophila* under heterotrophic conditions and development of Nile red fluorescence based method for quantification of its neutral lipid content. *Biochem Eng J* 61: 11–19.
47. Provasoli L, McLaughlin JJA, Droop MR (1957) The development of artificial media for marine algae. *Arch Mikrobiol* 25: 392–428.
48. Okada S, Devarenne TP, Murakami M, Abe H, Chappell J (2004) Characterization of botryococcene synthase enzyme activity, a squalene synthase-like activity from the green microalga *Botryococcus braunii*, race B. *Arch Biochem Biophys* 422: 110–118.
49. Wang JP, Chen YZ, Ge XW, Yu HQ (2007) Optimization of coagulation–flocculation process for a paper-recycling wastewater treatment using response surface methodology. *Colloids Surf. Afri Physicochem Eng* 302: 204–210.
50. Zheng ZM, Hu QI, Hao J, Xu F, Guo NN, et al. (2008) Statistical optimization of culture conditions for 1,3-propanediol by *Klebsiella pneumoniae* AC15 via central composite design. *Bioresour Technol* 99: 1052–1056.
51. Ghadge SV, Raheman H (2006) Process optimization for biodiesel production from mahua (*Madhuca indica*) oil using response surface methodology. *Bioresour Technol* 97: 379–384.
52. Cui FJ, Li Y, Xu ZH, Xu HY, Sun K, et al. (2006) Optimization of the medium composition for production of mycelial biomass and exo-polymer by *Grifola frondosa* GF9801 using response surface methodology. *Bioresour Technol* 97: 1209–1216.
53. Gu XB, Zheng ZM, Yu HQ, Wang J, Liang FL, et al. (2005) Optimization of medium constituents for a novel lipopeptide production by *Bacillus subtilis* MO-01 by a response surface method. *Process Biochem* 40: 3196–3201.
54. Kaushik R, Saran S, Isar J, Saxena RK (2006) Statistical optimization of medium components and growth conditions by response surface methodology to enhance lipase production by *Aspergillus carneus*. *J Mol Catal B: Enzyme* 40: 121–126.
55. Goldberg ED (1989) Genetic algorithms in search optimization and machine learning. Boston: Addison-Wesley Longman Publishing. 412 p.
56. Jin X, Nalewajko C, Kushner DJ (1996) Comparative study of nickel toxicity to growth and photosynthesis in nickel-resistant and -sensitive strains of *Scenedesmus acutus* f. *alternans* (Chlorophyceae). *Microbial Ecol* 31: 103–114.
57. Wong JPK, Wong YS, Tam NFY (2000) Nickel biosorption by two chlorella species, *C. vulgaris* (a commercial species) and *C. miniata* (a local isolate). *Biores Technol* 73: 133–137.
58. Shay LK, Hunt HR, Wegner GH (1987) Highproductivity fermentation process for cultivation industrial microorganisms. *J Ind Microbiol* 2: 79–85.
59. Wang Z, Chen S, Cao X (2010) Micro-nutrients effects on algae colony: growth rate and biomass response to various micro-nutrients and competitive inhibitions among multi-microelements. *Symposium of 4th Internat Con Bioinformat Biomed Eng* 1–8.
60. Cloëz I, Dumont O, Piciotti M, Bourre JM (1987) Alterations of lipid synthesis in the normal and dysmyelinating trembler mouse sciatic nerve by heavy metals (Hg, Pb, Mn, Cu, Ni). *Toxicology* 46: 65–71.
61. Mohammady NGE, Fathy AA (2007) Humic acid mitigates viability reduction, lipids and fatty acids of *Dunaliella salina* and *Nannochloropsis salina* grown under nickel stress. *Internat J Bot* 3: 64–70.
62. Rousch JM, Sommerfeld MR (1999) Effect of manganese and nickel on growth of selected algae in pH buffered medium. *Water Res* 33: 2448–2454.
63. Lee YK, Ding SY (1994) Cell cycle and accumulation of astaxanthin in *Haematococcus lacustris* (Chlorophyta). *J Phycol* 30: 445–449.