SYSTEMATIC REVIEW



The Effect of Nitazoxanide on the Clinical Outcomes in Patients with COVID-19: A Systematic Review and Meta-Analysis of Randomized Controlled Trials

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Abstract

Background and Objective Nitazoxanide, a US Food and Drug Administration-approved antiparasitic agent, was reported to be effective in treating coronavirus disease 2019 (COVID-19). The lack of effective and precise treatments for COVID-19 infection earlier in the pandemic forced us to depend on symptomatic, empirical, and supportive therapy, which overburdened intensive care units and exhausted hospital resources. Therefore, the aim of this systematic review and meta-analysis was to assess the efficacy and safety of nitazoxanide for COVID-19 treatment.

Methods A systematic review and meta-analysis synthesizing relevant randomized controlled trials from six databases (MedRxiv, WOS, SCOPUS, EMBASE, PubMed, and CENTRAL) until 17 May 2022 was conducted. Risk ratio (RR) for dichotomous outcomes was used and data with a 95% confidence interval (CI) are presented. The protocol was registered in PROSPERO with ID: CRD42022334658.

Results Six randomized controlled trials with 1412 patients were included in the analysis. Nitazoxanide was effective in accelerating viral clearance compared with placebo (RR: 1.30 with 95% CI 1.08, 1.56, p = 0.006) and reducing oxygen requirements (RR: 0.48 with 95% CI 0.39, 0.59, p = 0.00001), but we found no difference between nitazoxanide and placebo in improving clinical resolution (RR: 1.01 with 95% CI 0.94, 1.08, p = 0.88), reducing the mortality rate (RR: 0.88 with 95% CI 0.4, 1.91, p = 0.74), and intensive care unit admission (RR: 0.69 with 95% CI 0.43, 1.13, p = 0.14). Moreover, nitazoxanide was as safe as placebo (RR: 0.9 with 95% CI 0.72, 1.12, p = 0.34).

Conclusions Compared with placebo, nitazoxanide was effective in expediting viral clearance and decreasing oxygen requirements. However, there was no difference between nitazoxanide and placebo regarding clinical response, all-cause mortality, and intensive care unit admission. Therefore, more large-scale studies are still needed to ascertain the clinical applicability of nitazoxanide in COVID-19.

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Key Points

Nitazoxanide is potentially effective in accelerating coronavirus disease 2019 (COVID-19) viral clearance and reducing oxygen requirements compared with placebo.

It showed no efficacy in improving clinical symptoms or reducing all-cause mortality and intensive care unit admission.

1 Introduction

The coronavirus disease 2019 (COVID-19) pandemic is the worst global health crisis since the influenza pandemic of 1918. It crippled numerous healthcare systems around the world and tremendously downturned the global economy. Severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) infection results in a broad spectrum of clinical presentations. Most cases are asymptomatic or have mild-to-moderate symptoms; however, 5–14% develop a severe, potentially life-threatening disease [1-3]. The disease can progress to different clinical phenotypes ending with severe pneumonia, acute respiratory distress syndrome, cytokine storm, disseminated intravascular coagulopathy, multi-organ failure, shock, and eventually death [1, 4-7]. We can relate this progression to different reasons, including age, comorbidities, viral genotype, and viral load [8-12].

Earlier in the pandemic, the lack of effective and specific treatments for COVID-19 infection left us with symptomatic, empirical, and supportive therapies that overworked the intensive care units (ICUs) and depleted hospitals' resources. Coupling vaccination programs, social distancing, and these therapies played a pivotal role in controlling the pandemic. Based on disease severity, treatment options included pre-existed antiviral therapies, anti-SARS-CoV-2 neutralizing antibody products, immunomodulatory agents, ventilation, and oxygen therapies [5]. With the focus on treating and developing new drugs for severe and complicated cases of COVID-19 infections, therapies for mild and moderate infections in outpatient settings are limited. Paxlovid and molnupiravir, two newly developed and US Food and Drug Administrationapproved oral medications for COVID-19, have shown a significant reduction in hospitalization and death in mild to moderate infections [13, 14]. With barriers to worldwide access to these recently developed medications, the need to repurpose existing anti-microbial agents has been accepted as an alternative treatment option [15–17].

Nitazoxanide (NTZ) is a US Food and Drug Administration-approved antiparasitic drug with an excellent safety profile. It has been suggested as one of the alternative therapies for COVID-19 infection for different reasons. Hong et al. found that NTZ decreased the plasma level of interleukin (IL)-6 markedly when administered in mice [18]. Shou et al. also suggested that tizoxanide, the main active metabolite of NTZ, wielded anti-inflammatory effects in vivo [19]. This advocated for the possible beneficiary effects of NTZ in controlling cytokine storms where large amounts of IL-6 and other pro-inflammatory cytokines are released. Treatment with NTZ also showed wide antiviral activities with different mechanisms against various viral infections, including influenza, Middle East respiratory syndrome coronavirus, and other coronaviruses [20–25]. Jasenosky et al. also found that NTZ amplified the host's innate immune response to viruses and inhibited Ebola virus replication [26].

To date, a total of 31 clinical trials have been registered on ClinicalTrials.gov to investigate the effect of NTZ on COVID-19 infections. Nitazoxanide was administered alone or combined with other drugs compared to a placebo. The results ranged from accelerated symptom resolution, a shorter time to hospital discharge, a decreasing viral load, and a well-tolerated safety profile to no difference between the placebo and NTZ groups [27–32]. Therefore, we performed this systematic review and meta-analysis to synthesize evidence from the published randomized controlled trials (RCTs) on the efficacy and safety of NTZ in patients with COVID-19 infection.

2 Methods

2.1 Protocol Registration

For this systematic review and meta-analysis, we rigorously adhered to the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) statement [33] and the *Cochrane Handbook for Systematic Reviews and Interventions* [34]. The review protocol was registered in PROSPERO with ID: CRD42022334658.

2.2 Data Sources and Search Strategy

Until 17 May 2022, two reviewers (B.A. and M.A.) conducted a systematic search of the following electronic databases: MedRxiv, Web of Science, SCOPUS, EMBASE, PubMed (MEDLINE), and Cochrane Central Register of Controlled Trials (CENTRAL). There were no search filters applied. The search approach and results are outlined in Table S1 of the Electronic Supplementary Material (ESM).

2.3 Eligibility Criteria

We included RCTs with the following PICO criteria: population (P): patients with COVID-19 symptoms and confirmed by either chest computed tomography suggestive of viral pneumonia or a positive nasopharyngeal swab test for SARS-CoV-2 (reverse-transcriptase polymerase chain reaction [RT-PCR]); intervention (I): nitazoxanide regardless of dosage, route, and duration of administration; control I: placebo; outcomes (O): primary outcome: confirmed viral clearance by negative RT-PCR irrespective of the time of assessment. Our secondary outcomes are clinical resolution, all-cause mortality, oxygen supplementation, ICU admission or mechanical ventilation, and incidence of adverse events (diarrhea, nausea, vomiting, abdominal pain, pruritis, and headache). Animal studies, pilot studies, observational studies (cohort, case-control, cross-sectional, case series, and case reports), single-arm clinical trials, in vitro investigations (tissue and culture studies), book chapters, editorials, press articles, and conference abstracts were excluded.

2.4 Study Selection

After duplicates were deleted by Covidence online software [35], two reviewers (R.A. and F.L.) independently screened the titles and abstracts of the included records. The full texts of the relevant records were then screened for the preceding eligibility criteria. Any disagreements were resolved by inviting a third reviewer (M.A).

2.5 Data Extraction

Four reviewers (B.K., F.L., R.A., and A.A.) independently extracted the following data from the included trials using a pre-tested extraction sheet: study characteristics (first author name, year of publication, country, study design, total participants, the dose, route of administration, and duration of administration; time of viral eradication assessment of NTZ; and follow-up duration); baseline information (age, sex, viral load, race, basal metabolic index, and comorbidities); efficacy outcomes data (negative RT-PCR, all-cause mortality, oxygen supplementation, and ICU admission); and safety outcomes data (diarrhea, nausea, vomiting, abdominal pain, pruritis, and headache). Dissension was used to resolve conflicts.

2.6 Risk of Bias and Quality Assessment

Using The Cochrane Collaboration's tool for assessing the risk of bias in randomized trials, four reviewers (M.A., F.L., R.A., and B.K.) independently assessed the included studies for risk of bias [36]. Random sequence generation (selection bias), allocation concealment (selection bias), blinding of participants and personnel (performance bias), blinding of outcome assessment (detection bias), incomplete outcome data (attrition bias), selective reporting (reporting bias), and other potential sources of bias were considered. Conflicts were resolved by discussion. Two reviewers (M.A. and F.L.) used the Grading of Recommendations Assessment, Development and Evaluation (GRADE) Working Group recommendation [37, 38] for the quality of evidence assessment. Inconsistency, imprecision, indirectness, publication bias, and bias risk were all considered. Our findings on the quality of evidence were justified, documented, and included in each outcome's reporting. Any disputes were handled by the third reviewer (B.A.).

2.7 Statistical Analysis

The statistical analysis was carried out with RevMan version 5.3 software [39]. We pooled dichotomous outcomes using the risk ratio (RR) presented with the corresponding 95% confidence interval (CI). We used the I² and Chi-square tests to examine heterogeneity; the Chi-square test determines if there is substantial heterogeneity, while the I² determines the magnitude of heterogeneity. A substantial heterogeneity (for the Chi-square test) is defined as an alpha level below 0.1, according to the *Cochrane Handbook for Systematic Reviews and Interventions (Chapter Nine)* [34], while the I² test is interpreted as follows: (0–40%: not significant; 30–60%: moderate heterogeneity; 50–90%: considerable heterogeneity). We utilized the fixed-effects model.

We conducted a sensitivity analysis in the case of considerable heterogeneity by deleting one study at a time and reconducting the analysis to see how each study affected the total effect size of the outcomes. We also conducted a subgroup analysis depending on the time of the viral clearance assessment. Because we only included fewer than ten studies in each outcome, we did not offer funnel plots to reveal publication bias, as advised by Egger et al. [40].

3 Results

3.1 Search Results and Study Selection

A total of 777 articles were collected by searching six databases: PubMed (110), Cochrane (45), Web of Science (106), Scopus (379), Embase (96), and MedRxiv (41), respectively. Two hundred and sixty-five duplicates were initially excluded. After title and abstract screening, 487 records were excluded leaving 25 articles for full-text screening. Sixteen articles were excluded after full-text screening (Table S2 of the ESM). Only six articles met our inclusion criteria. Figure 1 shows the selection process in a PRISMA flow diagram.

3.2 Characteristics of Included Studies

Our study included six RCTs: three conducted in Brazil [27, 29, 30], one in Egypt [28], one in Argentina [32], and another in the USA and Puerto Rico [31]. Our included studies had a total of 1412 participants who were randomized to receive either NTZ (n = 705) in the oral form or placebo (n = 707). Participants had a mean age of 48.1 years with a predominant white race (n = 714), then the black race (n = 108). Time of administration differed between one study and another, twice [27, 31], three [29, 30], or four [28, 32] times a day with a mean treatment duration of 8 days and a mean

Fig. 1 Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) flow chart of the screening process



follow-up duration of 2 months. The method of COVID-19 assessment was RT-PCR in all our studies, with a mean viral eradication time assessment of 14 days. Further description of the summary and baseline characteristics of included trials can be found in Tables 1 and 2.

3.3 Risk of Bias and Quality of Evidence

We assessed the quality of the included studies according to the Cochrane risk of bias tool, as shown in Fig. 2. All studies had a low risk of bias regarding the "random sequence generation". All studies had a low risk of bias regarding "allocation concealment" except Blum et al. [27] had an unclear risk of bias. All studies had a low risk of bias regarding "performance bias" except Medhat et al., an open-label study [28], with a high risk of bias. All studies had a low risk of bias regarding the "detection bias" except Medhat et al. [28] and Silva et al. [32], with a high risk of bias owing to a lack of outcome assessor blinding. Regarding the "attrition bias", three studies had a low risk of bias [28, 30, 31]. However, Blum et al. [27], Rocco et al. [29], and Silva et al. [32] had a high risk of bias because of a significant loss of follow-up. All studies had a low risk of bias regarding "reporting bias" except Medhat et al. [28], which had a high risk of bias due to not reporting clinical response data. All studies had a high risk of bias regarding the "other bias" owing to the presence of a funding source, except the Silva et al. [32] study, which had a low risk of bias because of the absence of a funding source.

Using the GRADE system, all the included outcomes yielded very low-quality evidence. Details and explanations are clarified in Table 3.

Table 1 Summa	rry characteristics	of the included st	udies								
Study	Study design	Study site	Sample size	Disease sever- ity	Nitazoxanide re,	gimen			Method of assessment	Time of viral eradication assessment	Follow-up dura- tion
					Dose (mg)	Formulation	Times of administra- tion/day	Dura- tion (days)			
Blum et al. (2021) [<mark>27</mark>]	Pilot RCT	Six centers in Brazil	50	Moderate	600 mg	Immediate- release tablets with food	2	7	RT-PCR	On day 21	5 months
Medhat et al. (2022) [28]	Open-label placebo- controlled randomized trial	Three centers in Egypt	150	Mild and mod- erate	500 mg	Immediate- release tablets with food	4	14	RT-PCR	On day 14	4 months
Rocco et al. (2021) [30]	Double- blinded, placebo- controlled randomized trial	Seven centers in Brazil	392	Mild	500 mg	Oral solution	ς	Ś	RT-PCR	On day 5	3 months
Rocco et al. (2022) [29]	Double- blinded, placebo- controlled randomized trial	19 centers in Brazil	405	Hospitalized with COVID- 19 pneumo- nia	500 mg	Oral solution	ς	Ś	RT-PCR	On day 7	14 days
Rossignol et al. (2022) [31]	Double- blinded, placebo- controlled randomized trial	36 centers in USA and Puerto Rico	379	Mild and mod- erate	600 mg	Extended- release tablets with food	0	Ś	RT-PCR	On days 4 and 10	28 days
Silva et al. (2021) [32]	Pilot RCT	One center in Argentina	36	Mild and mod- erate	Started with 1 g every 8 h, then changed to 500 mg every 6 h	Immediate- release tablets with food	4	14	RT-PCR	On day 7	14 days
COVID-19 coro	navirus disease 20	119, RCT randomi	ized controlled	l trial, RT-PCR rev	erse-transcriptase	polymerase chair	n reaction, h hou	IIS			

é		Age (years) Mean (SD)	_	Sex (male N (%)		BMI, mea	an (SD)	Race, N (?	(9)									Viral load. Mean (SD	, log 10	Comorbio	dities, N (%)		
								White		Black		Asian		Hispanic	or Latino	Mixed				DM		HTN	
on- NTZ	ZTN		Control	NTZ	Control	ZTN	Control	ZIN	Control	ZTN	Con- trol	ZIN	Con- trol	ZIN	Control	ZIN	Con- trol	ZIN	Control	ZTN	Control	ZTN	Control
5 64 (17)	64 (17)		64 (21)	7 (28)	8 (32)	89.2 (17)	76.18 (13.04)	21 (42)	16 (32)	1 (2)	2 (4)	N/A	N/A	N/A	N/A	3 (6)	7 (14)	29.03 (4.84)	28.88 (4.89)	N/A	N/A	N/A	N/A
3 44.32 (8.2	44.32 (8.4	13)	45.34 (5.59)	27 (35.1)	41 (56.2)	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	12 (15.6)	17 (23.3)	14 (18.2)	18 (24.7)
98 N/A	N/A		N/A	101 (52)	83 (42)	N/A	N/A	131 (68)	138 (70)	31 (16)	32 (16)	5 (3)	2(1)	N/A	N/A	27 (14)	24 (12)	6.99 (1.76)	7.32 (1.62)	N/A	N/A	N/A	N/A
33 56 (3.	56 (3.	83)	56 (3.5)	117 (58)	131 (65)	N/A	N/A	87 (43)	88 (43)	17 (8)	17 (8)	7 (4)	3 (2)	N/A	N/A	90 (45)	93 (46)	4.87 (2.17)	4.93 (2.02)	42 (21)	48 (24)	73 (36)	73 (36)
35 38 (1	38 (1	1.83)	42 (11.33)	83 (45.1)	82 (42.1)	29.16 (6.12)	29.46 (5.9)	117 (63.6)	116 (59.5)	4 (2.2)	4 (2.1)	2 (1.1)	4 (2.1)	59 (32.1)	71 (36.4)	N/A	N/A	5.93 (2.47)	6.05 (2.24)	N/A	N/A	V/N	V/N
3 43.5(43.5(12)	50.5 (11.64)	17 (73.9)	9 (69.3)	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A

 Table 2
 Baseline characteristics of the included studies

	0											
Certainty assessn	nent						Nº of patients		Effect		Certainty	Importance
No. of studies	Study design	Risk of bias	Inconsistency	Indirectness	Imprecision	Other con- siderations	ZTN	Placebo	Relative (95% CI)	Absolute (95% CI)		
Viral clearance 4	Randomized trials	Very serious ^a	Serious ^b	Not serious	Serious ^c	None	167/471 (35.5%)	127/472 (26.9%)	RR 1.30 (1.08, 1.56)	81 more per 1000 (from 22 more to 151 more)	⊕⊖⊖⊖ Very low	Critical
All-cause mortalı 4	ity Randomized trials	Very serious ^d	Not serious	Not serious	Very serious ^e	None	11/722 (1.5%)	12/704 (1.7%)	RR 0.88 (0.40, 1.91)	2 fewer per 1000 (from 10 fewer to 16 more)	⊕⊖⊖⊖ Very low	Critical
Clinical resolutio 2	n Randomized trials	Very serious ^f	Serious ^b	Serious ^e	Very serious ^e	None	311/396 (78.5%)	313/401 (78.1%)	RR 1.01 (0.94, 1.08)	8 more per 1000 (from 47 fewer to 62 more)	⊕⊖⊖⊖ Very low	Important
ICU admission 3	Randomized trials	Very serious ^h	Not serious	Not serious	Very serious ^e	None	24/425 (5.6%)	35/426 (8.2%)	RR 0.69 (0.43, 1.13)	25 fewer per 1000 (from 47 fewer to 11	⊕⊖⊖⊖ Very low	Important
Oxygen requirem 2	tent Randomized trials	Very serious ⁱ	Not serious	Not serious	Serious ^j	None	101/681 (14.8%)	212/684 (31.0%)	RR 0.48 (0.39, 0.59)	more) 161 fewer per 1000 (from 189 fewer to 127 fewer to	⊕⊖⊖⊖ Very low	Important
Patients with at I. 4	east one adverse e Randomized trials	vent Very serious ^k	Not serious	Not serious	Very serious ^e	None	201/893 (22.5%)	218/889 (24.5%)	OR 0.90 (0.72, 1.12)	19 fewer per 1000 (from 56 fewer to 22 more)	⊕⊖⊖⊖ Very low	Important
<i>Cl</i> confidence ^a Blum et al. al ^b f ² test more ti ^c The CI does 1 ^c The CI does 1 ^f Rocco et al. 2 ^f Rocco et al. 2 ^s The criteria c ^h Blum et al. ar ⁱ Blum et al. an ⁱ Blum et al. an ⁱ Blum et al. an	interval, <i>ICU</i> ¹ ind Medhat et a han 50% not exclude the silva et al., and not exclude the of clinical resol ind Rocco et al d Rocco et al d Rocco et al	intensive care ur I. are of low qua RR of 1.25 Medhat et al. ar RR of 1.25 and quality, while Ro ution may not ac 2022 are of low dence interval, o	iit, <i>NTZ</i> nitazox: lity, while Rossi e of low quality 0.75 (appreciab occo et al. 2021 j occurately represe v quality, while J quality or analysis inclu	anide <i>OR</i> odds ignol et al. and , while Rossig he benefit/harr is of moderate ant the effectiv Rocco et al. 20 aded less than	ratio, <i>RR</i> risk Rocco et al. 21 nol et al. is of r (n) quality eness 21 is of moder and Rocco et a	natio 021 are of mc noderate qual ate quality ach arm	derate quality ity					



Fig. 2 Summary of risk of bias. A Review authors' judgments about each risk of bias item for each included study and **B** review authors' judgments about each risk of bias item presented as percentages across all included studies

3.4 Primary Outcome: Confirmed Viral Clearance by Negative RT-PCR

The pooled RR significantly favored NTZ over placebo (RR: 1.30 with 95% CI 1.08, 1.56, p = 0.006) [very-low quality evidence] (Fig. 3A, Table 3). Pooled studies were heterogenous (p = 0.03, $I^2 = 66\%$). To resolve heterogeneity, we conducted a sensitivity analysis. However, heterogeneity was not resolved by a sensitivity analysis. Furthermore, pooled RR showed no difference between

NTZ and placebo after excluding Medhat et al. [28] and Rocco et al. [30] [(RR: 1.21 with 95% CI 1.00, 1.47, p =0.06) and (RR: 1.16 with 95% CI 0.94, 1.44, p = 0.16), respectively] (Table S3 of the ESM). We conducted a subgroup analysis based on the time of assessment; pooled RR favored NTZ over placebo from 1 to 7 days (RR: 1.49 with 95% CI 1.07, 2.08, p = 0.02); however, we found no difference either from 8 to 14 days (RR: 1.16 with 95% CI 0.87, 1.54, p = 0.31) or from 15 to 21 days (RR: 1.26 with 95% CI 0.99, 1.61, p = 0.06) (Fig. 3B).

3.5 Secondary Outcomes

3.5.1 Clinical Resolution

The pooled RR showed no difference between NTZ and placebo (RR: 1.01 with 95% CI 0.94, 1.08, p = 0.88) [very-low quality evidence] (Fig. 4A, Table 3). Pooled studies were heterogenous (p = 0.11, $I^2 = 60\%$).

3.5.2 All-Cause Mortality

The pooled RR showed no difference between NTZ and placebo (RR: 0.88 with 95% CI 0.4, 1.91, p = 0.74) [very-low quality evidence] (Fig. 4B, Table 3). Pooled studies were homogenous (p = 0.36, $I^2 = 7\%$).

3.5.3 ICU Admission

The pooled RR showed no difference between NTZ and placebo (RR: 0.69 with 95% CI 0.43, 1.13, p = 0.14) [very-low quality evidence] (Fig. 4C, Table 3). Pooled studies were homogenous (p = 0.28, $I^2 = 21\%$).

3.5.4 Oxygen Requirement

The pooled RR significantly favored NTZ over placebo (RR: 0.48 with 95% CI 0.39, 0.59, p = 0.00001) [very-low quality evidence] (Fig. 4D, Table 3). Pooled studies were homogenous (p = 0.14, $I^2 = 39\%$). We conducted a subgroup analysis based on the day of assessment; pooled RR significantly favored NTZ over placebo on day 4 or 5 (RR: 0.4 with 95% CI 0.3, 0.52, p = 0.00001) and on day 7 (RR: 0.54 with 95% CI 0.36, 0.81, p = 0.003); however, we found no difference between NTZ and placebo on day 14 (RR: 0.67 with 95% CI 0.41, 1.08, p = 0.1) (Fig. 4D).

3.5.5 Safety (Incidence of Adverse Events)

The pooled RR showed no difference between NTZ and placebo in patients with at least one adverse event (RR: 0.9 with 95% CI 0.72, 1.12, p = 0.34) [very-low quality evidence]

Fig. 3 Forest plot of the primary outcome. A Viral clearance and **B** viral clearance subgroubed by the time of assessment. *CI* confidence interval, *M-H Mantel-Haenszel method*

A Viral Clearance

	Nitazoxa	anide	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% CI
Blum et al. 2021	23	23	15	19	13.2%	1.26 [0.99, 1.61]	
Medhat et al. 2022	27	77	12	73	9.6%	2.13 [1.17, 3.89]	_
Rocco et al. 2021	58	194	36	198	27.8%	1.64 [1.14, 2.37]	
Rossignol et al. 2022	59	177	64	182	49.3%	0.95 [0.71, 1.26]	+
Total (95% CI)		471		472	100.0%	1.30 [1.08, 1.56]	◆
Total events	167		127				
Heterogeneity: Chi ² = 8.	90, df = 3 ((P = 0.0	3); l² = 66	i%			
Test for overall effect: Z	= 2.76 (P =	= 0.006)					Favours [Placebo] Favours [Nitazoxanide]

В

Subgroub analysis based on time of assessment.

	Nitazoxa	anide	Place	bo		Risk Ratio		Risk	Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI		M-H, Fixe	ed, 95% CI	
1.2.1 From 1 to 7 days										
Rocco et al. 2021	58	194	36	198	27.8%	1.64 [1.14, 2.37]				
Rossignol et al. 2022	11	177	11	178	8.6%	1.01 [0.45, 2.26]			<u> </u>	
Subtotal (95% CI)		371		376	36.4%	1.49 [1.07, 2.08]			◆	
Total events	69		47							
Heterogeneity: Chi ² = 1.	18, df = 1	(P = 0.2)	B); I² = 15	%						
Test for overall effect: Z:	= 2.37 (P =	= 0.02)								
1.2.2 From 8 to 14 days	;									
Medhat et al. 2022	27	77	12	73	9.6%	2.13 [1.17, 3.89]				
Rossignol et al. 2022	48	177	53	182	40.8%	0.93 [0.67, 1.30]		-	-	
Subtotal (95% CI)		254		255	50.4%	1.16 [0.87, 1.54]			◆	
Total events	75		65							
Heterogeneity: Chi ² = 5.	65, df = 1	(P = 0.0)	2); I² = 82	%						
Test for overall effect: Z	= 1.02 (P =	= 0.31)								
1.2.3 From 15 to 21 day	s									
Blum et al. 2021	23	23	15	19	13.2%	1.26 (0.99, 1.61)			-	
Subtotal (95% CI)		23		19	13.2%	1.26 [0.99, 1.61]			◆	
Total events	23		15							
Heterogeneity: Not appli	icable									
Test for overall effect: Z	= 1.88 (P =	= 0.06)								
Total (95% CI)		648		650	100.0%	1.30 [1.07, 1.57]			•	
Total events	167		127							
Heterogeneity: Chi ² = 8.	52, df = 4	(P = 0.0	7); I² = 53	%				01	1 10	100
Test for overall effect: Z	= 2.66 (P =	= 0.008)					0.01	Eavours (Placebo)	Favours (Nitazox)	our anidel
Test for subgroup differ	ences: Ch	i² = 1.29), df = 2 (F	P = 0.52	2), I ² = 0%	5		i aroaro (i lacebo)	1 310310 [141132076	11100]

(Fig. 5A, Table 3). Pooled studies were homogenous (p = 0.45, $I^2 = 0\%$). Moreover, there was no difference between NTZ and placebo regarding the incidence of diarrhea, headache, nausea, abdominal pain, pruritis, and urticaria; however, NTZ was significantly associated with vomiting (Fig. 5B, Table S4 of the ESM).

4 Discussion

With the lack of a definitive antiviral therapy for SARS-CoV-2 infection, trials of repurposing existing medications started to trend. Nitazoxanide was considered a potential treatment for COVID-19 based on the existing evidence of its various antiviral and immunomodulatory properties either in vivo or in vitro [18–21, 23, 25, 26, 41, 42]. Our recent meta-analysis involving six RCTs demonstrated that NTZ is effective in increasing the viral clearance rate and decreasing oxygen requirements; however, we detected no difference between NTZ and placebo in reducing mortality, ICU admission, and improving clinical resolution. Additionally, NTZ

was safe, well tolerated, and with similar rates of adverse events compared to placebo, except for vomiting.

Regarding viral clearance, NTZ was effective compared to placebo up to 7 days after initiating treatment but not effective afterward up to 14 and 21 days. On the one hand, Blum et al. [27] and Rossignol et al. [31] did not support NTZ. In Blum et al. [27], viral clearance was assessed on day 21 after treatment; hence, this difference can be attributed to the long duration of assessment after treatment, because in most patients, a longer duration would result in decreased viral load regardless of therapy [30]. This supports the findings of our subgroup analysis that NTZ was effective for only up to 7 days; however, only one to two studies were included in each subgroup, which can undermine our findings. In contrast, Medhat et al. [28] supported NTZ after 14 days, which can be attributed to some different methodological aspects, including using standard treatment along with NTZ, a longer treatment duration (14 days), and more frequent NTZ administration (four times a day). Moreover, Rossignol et al. [31] attributed this difference to the procedures used to collect, process, and quantify viral

Fig. 4 Forest plot of the secondary outcomes. A Clinical resolution, B all-cause mortality, C intensive care unit admission, and D oxygen requirement. CI confidence interval, *M-H Mantel-Haenszel method*

A Clinical Resolution.

	Nitazoxa	anide	Place	bo		Risk Ratio		Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI		M-H, Fixed, 95% Cl	
Rocco et al. 2021	135	194	146	198	46.5%	0.94 [0.83, 1.07]		•	
Rocco et al. 2022	176	202	167	203	53.5%	1.06 [0.97, 1.15]		•	
Total (95% CI)		396		401	100.0%	1.01 [0.94, 1.08]			
Total events	311		313						
Heterogeneity: Chi ² =	2.50, df=	1 (P = 0	.11); I² =	60%			0.01		100
Test for overall effect:	Z=0.15 (P = 0.88	i)				0.01	Favours [Placebo] Favours [Nitazoxanide]	100

B All-Cause Mortality

	Nitazoxa	anide	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% Cl
Blum et al. 2021	2	25	6	25	47.0%	0.33 [0.07, 1.50]	
Rocco et al. 2022	6	202	5	203	39.1%	1.21 [0.37, 3.89]	_
Rossignol et al. 2022	2	472	0	463	4.0%	4.90 [0.24, 101.89]	
Silva et al. 2021	1	23	1	13	10.0%	0.57 [0.04, 8.30]	
Total (95% CI)		722		704	100.0%	0.88 [0.40, 1.91]	-
Total events	11		12				
Heterogeneity: Chi ² = 3.3	22, df = 3 ((P = 0.3)	6); I ² = 79	6			
Test for overall effect: Z	= 0.33 (P =	= 0.74)					Favours [Nitazoxanide] Favours [Placebo]

C ICU admission.

	Nitazoxa	anide	Place	bo		Risk Ratio	Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% Cl	
Blum et al. 2021	2	25	6	25	16.9%	0.33 [0.07, 1.50]		
Rocco et al. 2021	2	198	0	198	1.4%	5.00 [0.24, 103.49]		
Rocco et al. 2022	20	202	29	203	81.7%	0.69 [0.41, 1.18]		
Total (95% CI)		425		426	100.0%	0.69 [0.43, 1.13]	•	
Total events	24		35					
Heterogeneity: Chi ² =	2.55, df =	2 (P = 0	.28); l² =	21%				100
Test for overall effect:	Z=1.48 (P = 0.14)				Favours [Nitazoxanide] Favours [Placebo]	100

D Oxygen Requirement.

	Nitazoxa	anide	Place	bo		Risk Ratio	Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% CI	
2.4.1 On day 4 or 5								_
Blum et al. 2021	8	25	10	25	4.7%	0.80 [0.38, 1.69]		
Rocco et al. 2022	40	202	112	203	52.8%	0.36 [0.26, 0.49]	+	
Subtotal (95% CI)		227		228	57.5%	0.40 [0.30, 0.52]	◆	
Total events	48		122					
Heterogeneity: Chi ² =	3.82, df =	1 (P = 0	.05); I² =	74%				
Test for overall effect: .	Z = 6.50 (F	P < 0.00	001)					
0.4.0.0								
2.4.2 On day /	_		-					
Blum et al. 2021	5	25	6	25	2.8%	0.83 [0.29, 2.38]		
Rocco et al. 2022	24	202	48	203	22.6%	0.50 [0.32, 0.79]		
Total quanta	20	221	54	220	25.5%	0.54 [0.50, 0.61]	•	
Hotorogonoity: Chiž -	29 076 df-	1 /0 - 0	04 20\+IZ=	n 04				
Tect for overall effect:	0.70, ui = 7 - 2 0 / /i	- 0 00	.30), IT = 1 2)	U 70				
restion overall cheet.	2 - 2.34 ()	- 0.00	5)					
2.4.3 On day 14								
Blum et al. 2021	2	25	3	25	1.4%	0.67 [0.12, 3.65]		
Rocco et al. 2022	22	202	33	203	15.6%	0.67 [0.41, 1.11]		
Subtotal (95% CI)		227		228	17.0%	0.67 [0.41, 1.08]	◆	
Total events	24		36					
Heterogeneity: Chi ² =	0.00, df=	1 (P = 1	.00); I ^z = I	0%				
Test for overall effect:	Z = 1.63 (F	^o = 0.10)					
Total (95% CI)		681		684	100.0%	0.48 [0.39, 0.59]	•	
Total events	101		212					
Heterogeneity: Chi ² =	8.24, df=	5 (P = 0	.14); I² =	39%				1
Test for overall effect:	Z = 6.95 (F	P < 0.00	001)				Favours [Nitazoxanide] Favours [Placebo]	-
Test for subaroup diffe	erences: (>hi⁼ = 3.	97. df = 2	! (P = 0.	.14), I ² = 4	19.6%	. , , ,	

loads from nasopharyngeal swabs not being validated to predict viral load, inflammation, lung symptoms, or clinical outcomes at the patient or trial level. It is also still questionable if RT-PCR adequately detects infectious viruses because viral RNA can survive even in the absence of the replication-competent virus for a long duration [31]. On the other hand, Rocco et al. 2021 [30] and Medhat et al. [28] supported NTZ as they only included patients with mild disease with no mortality recorded and with only two patients admitted to the ICU in Rocco et al. [30].

Fig. 5 Forest plot of the safety outcomes. A Patients with at least one adverse event and **B** adverse events. *CI* confidence interval, *M-H Mantel-Haenszel method*

A Patients with at least one adverse event.

	Nitazoxa	anide	Place	bo		Odds Ratio	Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% CI	
Blum et al. 2021	8	25	13	25	5.5%	0.43 [0.14, 1.37]		
Rocco et al. 2021	60	194	60	198	25.5%	1.03 [0.67, 1.58]		
Rocco et al. 2022	70	202	70	203	28.3%	1.01 [0.67, 1.52]	-+-	
Rossignol et al. 2022	63	472	75	463	40.7%	0.80 [0.55, 1.15]		
Total (95% CI)		893		889	100.0%	0.90 [0.72, 1.12]	•	
Total events	201		218					
Heterogeneity: Chi ² = 2.6	64, df = 3 (P = 0.4	5); I² = 09	6				400
Test for overall effect: Z =	= 0.96 (P =	= 0.34)					Favours [Nitazoxanide] Favours [Placebo]	100

B Adverse events.

	Nitazoxa	anide	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% CI	M-H, Fixed, 95% Cl
2.6.1 Diarrhea							
Blum et al. 2021	1	25	10	25	0.2%	3.00 [0.13, 70.30]	
Roccoletial, 2021 Roccoletial, 2022	5/	194	49	198	21.3%	1.19 [0.86, 1.65]	
Ruccu et al. 2022	16	472	10	462	1.970	1.23 [0.06, 2.22]	
Subtotal (95% CI)	10	893	10	889	33.8%	1.26 [0.96, 1.65]	◆
Total events	96		77				
Heterogeneity: Chi ² = 0.	73, df = 3	(P = 0.8)	7); I² = 0%	5			
Test for overall effect: Z:	= 1.68 (P :	= 0.09)					
2.6.2 Hoodoobo							
Z.0.2 neauache	24	104	22	100	12.000	1 00 0 70 1 601	
Rocco et al. 2021	33	202	45	203	10,7%	0.74 [0.49 1.10]	
Subtotal (95% CI)		396	45	401	33.6%	0.88 [0.65, 1.18]	
Total events	67		77				
Heterogeneity: Chi ² = 1.	60, df = 1	(P = 0.2	1); I² = 38	%			
Test for overall effect: Z	= 0.84 (P :	= 0.40)					
2.6.2 Nouses							
Z.0.3 Mausea	20	104	20	100	12606	0.00.00.61.1.601	
Rocco et al. 2021	20	202	29	203	5.7%	1 47 [0 75 2 89]	
Subtotal (95% CI)	10	396	10	401	18.3%	1.14 [0.77, 1.68]	
Total events	47		42				
Heterogeneity: Chi ² = 0.4	89, df = 1	(P = 0.3	5); I² = 0%	5			
Test for overall effect: Z	= 0.64 (P :	= 0.52)					
2.6.4 Abdominal pain							
2.0.4 Abuominai pain Dium et el. 2021		25	1	25	0.70	0 33 (0 04 7 04)	
Diumietal. 2021 Rocco et al. 2021	10	104	5	100	2.2%	2.04 [0.71, 6.96]	
Rocco et al. 2021	11	202	14	203	6.1%	0.79 [0.37 1.70]	
Subtotal (95% CI)		421	14	426	8.9%	1.06 [0.59, 1.91]	•
Total events	21		20				
Heterogeneity: Chi ² = 2.	57, df = 2	(P = 0.2)	B); I² = 22	%			
Test for overall effect: Z	= 0.19 (P :	= 0.85)					
2.6.5.Vomiting							
Posso of al. 2021	0	104	2	100	1 206	206100411141	
Rocco et al. 2021	3	202	0	203	0.2%	7.03 [0.37, 135.32]	
Subtotal (95% CI)		396		401	1.5%	3.63 [1.12, 11.76]	
Total events	12		3				
Heterogeneity: Chi ² = 0.1	26, df = 1	(P = 0.6	1); I² = 0%	6			
Test for overall effect: Z :	= 2.15 (P :	= 0.03)					
2.6.6 Druritis							
Rocco et al. 2021	4	194	1	198	0.4%	4 08 00 46 36 201	
Rocco et al. 2022	2	202	1	203	0.4%	2.01 [0.18, 21.99]	
Subtotal (95% CI)		396		401	0.9%	3.04 [0.62, 14.97]	
Total events	6		2				
Heterogeneity: Chi ² = 0.	19, df = 1	(P = 0.6	7); I² = 0%	5			
Test for overall effect: Z:	= 1.37 (P =	= 0.17)					
2.6.7 Urticaria							
Rocco et al. 2021	1	194	3	198	1.3%	0.34 (0.04, 3.24)	
Rocco et al. 2022	4	202	4	203	1.7%	1.00 [0.25, 3.96]	
Subtotal (95% CI)		396		401	3.1%	0.72 [0.23, 2.25]	
Total events	5		7				
Heterogeneity: Chi ² = 0.1	65, df = 1	(P = 0.4)	2); I² = 0%				
lest for overall effect: Z:	= U.56 (P :	= 0.57)					
Total (95% CI)		3294		3320	100.0%	1.13 [0.96, 1.33]	▲
Total events	254	0204	228	2020			ľ
Heterogeneity: Chi ² = 15	5.45, df = 1	6 (P = 0	1.49); I ² =	0%			
Test for overall effect: Z	= 1.42 (P =	= 0.15)					U.UT U.1 1 10 100 Eavours [Nitazoxanide] Eavours [Placebo]
Test for subgroup different	ences: Ch	i ² = 9.21	, df = 6 (F	P = 0.18	5), I ² = 34.	9%	i avorio (ivitazovanine) i avorio (i iaceno)

Nitazoxanide may have an antiviral effect in more than one stage of the COVID-19 replication cycle; it suppresses viral RNA and DNA replication as well as direct viral protein production in a variety of viruses [20, 21, 43]. To clarify, NTZ has been reported to be effective against Middle-East respiratory syndrome severe acute respiratory syndrome-1 (SARS-CoV) [21, 44]. Given that the genomic similarity between COVID-19 and Middle-East respiratory syndrome is about 50% and between COVID-19 and SARS-Cov is about 79% [45], the effective therapeutic approaches against Middle-East respiratory syndrome and SARS-CoV, NTZ in our case, can be effective against COVID-19 [46].

Furthermore, it interferes with the host's cellular metabolism by modulating interferon (IFN) surge [20, 21]. Nitazoxanide prevents COVID-19-induced IFN surge, subsequently preventing the development of a cytokine storm [47]. To clarify, the entry of the COVID-19 virus into the alveolar type II pneumocyte cells uses angiotensin-converting enzyme 2 receptors [48], which leads to cellular pyroptosis and a damage-associated molecular pattern release [49]. This is detected by alveolar macrophages, leading to the secretion of the pro-inflammatory cytokines, including tumor necrosis factor-alpha (TNF- α), IL-6, IL-8, and macrophage inflammatory protein-1 alpha [50]. This cycle is inhibited and controlled by IFN-1, leading to diminished viral replication and decreased cellular damage [51]; however, IFN-1 is downregulated by COVID-19, leading to immunological escape and cytokine over-secretion leading to a cytokine storm [52]. This effect is inhibited by NTZ, preventing immunological escape and the subsequent cytokine storm [46]. Additionally, it interferes with host-regulated mechanisms responsible for viral replication of mammalian target of rapamycin complex 1 signaling [53]. Finally, NTZ has been reported to enhance autophagic cell death [46], which was reported to be beneficial in controlling COVID-19 [54]. The autophagy of necrotic cells, which is considered a pro-inflammatory trigger, can ameliorate the inflammatory process decreasing the amount of secreting cytokines [55].

Regarding inflammatory markers, acute inflammatory markers decreased significantly with NTZ in COVID-19. To clarify, C-reactive protein, which is associated with a worse prognosis in COVID-19 [56, 57], was decreased with NTZ [27]. Furthermore, IL-6 was reduced with NTZ; IL-6 is a key pro-inflammatory mediator involved in the development of the acute phase response, which results in a variety of local and systemic responses such as fever, leucocyte recruitment, activation, and hemodynamic effects [27]. It also predicts a higher risk of disease deterioration [58].

Furthermore, NTZ was also associated with TNF- α reduction [27]. Tumor necrosis factor-alpha is one of the main cytokines responsible for the immunological responses to COVID-19 [59, 60]. Hence, anti-TNF- α drugs can reduce COVID-19 respiratory insufficiency and mortality by lowering inflammatory-driven capillary leak [61]. Moreover, IL-8, a powerful pro-inflammatory cytokine that plays a key function in the inflammatory recruitment and activation of neutrophils, was decreased significantly with NTZ [27]. It is also possible that IL-8 has a role in the frequent neutrophilia seen in patients with COVID-19 [42]. Regarding cellular

immunity, CD4⁺ HLA-DR⁺ T-cell lymphocytes were also significantly decreased with NTZ [27]. To conclude, NTZ can be beneficial in COVID-19 via its antiviral and immunomodulatory effects, decreasing IL-6, IL-8, TNF- α , and CD4 T cells [62].

Despite the previous effects, to effectively treat a viral infection resembling COVID-19, antiviral agents that act at different steps of the viral replication cycle must be used together [27, 29]. This would reduce the virus's genetic variation and reduce the likelihood of the fast evolution of resistant strains [27]. Supporting this hypothesis, higher transmission rates and viral loads with COVID-19 can lead to new mutant strains, such as the Brazilian P.1 strain [63]. In this line, multiple trials have evaluated NTZ in combination with other drugs. To clarify, COVID-19 clearance from the nasopharynx was substantially faster with NTZ in conjunction with ribavirin, ivermectin, and zinc supplements compared with symptomatic treatment [53]. Another study assessed the efficacy of NTZ against COVID-19 in combination with azithromycin [53]. Furthermore, several in-vitro investigations revealed synergistic effects when NTZ is combined with other agents [64-66]. However, more research is still warranted in this regard.

Regarding the clinical resolution, we found no difference between NTZ and placebo; however, only two trials [29, 30] were included in our analysis. This can be attributed to symptoms in mild COVID-19 that can resolve spontaneously regardless of antiviral therapy [30], i.e., the median time from symptom onset to resolution was reported to be 8 (6.25–11.5) days [67]. Therefore, we can speculate that using clinical improvement as a marker of the efficacy of NTZ is inaccurate, especially when used in conjunction with symptomatic and supportive treatment [28]. However, Rocco et al. [29] reported a 1-day faster symptom resolution with NTZ compared with placebo, which can be attributed to the previous anti-inflammatory effects [29].

Accordingly, the magnitude of the effect of NTZ in managing COVID-19 is affected by the population being studied. To clarify, five out of six trials included only mild to moderate cases [27, 28, 30–32], with only Rocco et al. [29] including hospitalized patients with pneumonia. Additionally, despite the significant effect of viral load, the effects on substantial clinical data, such as mortality and ICU admission, were insignificant. Thus, our findings should be interpreted with caution, especially in the management of severe cases of pneumonia.

Regarding the all-cause mortality rates, we detected no difference between NTZ and placebo. In the RCT conducted by Rossignol et al. [31], two participants in the NTZ arm died. One because of severe COVID-19 infection and the other, who tested negative for SARS-CoV-2, because of secondary aspiration 19 days after completing the treatment. Both events were not tracked back to the study medication [31]. Rocco et al. [29] reported no difference by day 14 in the number of deaths between the NTZ group (six deaths) and the placebo group (five deaths). Blum et al. [27] reported a total of eight deaths; two in the NTZ group and six in the placebo group, all because of acute respiratory distress syndrome. The difference between the two groups was not statistically significant; however, they argued that this difference is clinically relevant, and a difference might be detected with a larger sample size [27]. Silva et al. [32] reported two deaths, one in each group. Both were aged older than 65 years and had other comorbidities [32]. The other two RCTs did not report any mortalities [28, 30] as they only included mild cases, as previously clarified.

Regarding the ICU admission, we detected no difference between NTZ and placebo. Rocco et al. [29] detected no difference between the NTZ and placebo group regarding ICU admission. However, he also found that participants who presented with oxygen saturation >90% on day 1 and were treated with NTZ had a lower odds of ICU admission compared with placebo [29]. Similarly, adding corticosteroids to NTZ decreased the odds of ICU admission compared with corticosteroids alone in the placebo group [29]. No ICU admission was reported in the rest of the included studies [28, 31, 32].

Oxygen requirements for treating the NTZ group were less than the placebo group. However, this effect showed a decreasing pattern with longer follow-ups, according to our subgroup analysis. Supporting our findings, Rocco et al. [29] found that NTZ reduced oxygen requirements of any type compared with placebo only from day 3 to day 7 [29]. He also found that the time on supplemental oxygen was reduced by a median of 2 days compared with placebo [29]. Blum et al. [27] reported a lower time to withdraw from oxygen supplementation in the NTZ group compared with the placebo group (3 vs 8 days, respectively). This effect is important because reducing the necessity of supplementary oxygen subsequently reduces the load on the healthcare system and perhaps enhances hospital capacity [29].

Regarding safety, there was no difference between the NTZ group and the placebo group in the incidence of at least one adverse event. Among the reported adverse events, only the incidence of vomiting was significantly associated with the NTZ compared with placebo. No severe adverse events associated with NTZ were reported in any of the included studies. The US Food and Drug Administration-approved dose of NTZ in treating parasite infection is 500 mg twice daily (BID). Different doses of NTZ have been suggested for their efficacy and safety against SARS-CoV-2 infection [68, 69]. Moreover, NTZ has a short half-life, thus its main action is achieved through its active metabolite, tizoxanide, which has a relatively long half-life. Maximum serum concentration (C_{max}) and the time to reach C_{max} determine the bioavailability (area under the curve) of NTZ and tizoxanide. Maximum serum concentration and time to reach the

 $C_{\rm max}$ are affected by the formulation of the NTZ where the suspension form is 41% less bioavailable than tablets [70]. Furthermore, food affects the absorption and bioavailability of NTZ; when NTZ is taken with food its C_{max} , T time to reach the C_{max} , and area under the curve increase. The dosing interval also determines the area under the curve and concentration needed to inhibit 90% (IC90) of SARS-CoV-2 [70]. To clarify, with food, a plasma concentration of more than IC90 was expected most of the time and IC50 almost all the time by using NTZ 500-mg tablets every 6 h. However, with fasting, the same dose achieved only IC50. A less frequent dose of 500-mg tablets every 8 hours also achieved plasma concentrations more than IC50 with food [70]. Therefore, the variability noticed among the included studies regarding the formulation and associated food administration can affect our findings.

Rajoli et al. [69] reported a physiologically based pharmacokinetic model about the optimal doses of NTZ that provide plasma and lung concentrations above its reported in vitro 90% effective concentration against SARS-CoV-2 $(4.64 \ \mu M \text{ or } 1.43 \ \mu g/mL)$ [69]. Ninety percent effective concentration was achieved when given in the fasting state with the doses of 1200 mg four times a day (QID), 1600 mg three times daily (TID), or 2900 mg BID. While with food, the needed doses were 700 mg OID, 900 mg TID, or 1400 mg BID [69]. Nitazoxanide was also reported to be safe up to 4 g/day [68]; however, most of the RCTs investigating its efficacy and safety in COVID-19 infection have been using doses of less than 2 g daily [27–32]. Haffizulla et al. [25] investigated a higher dose of 600 mg BID to achieve antiviral activities against influenza without reporting safety issues [25]. In the same line, Blum et al. [27] and Rossignol et al. [31] investigated the same dose of 600 mg BID in cases of mild and moderate COVID-19 infection without safety issues as well [27, 31]. Rocco et al. used a higher dose of 500 mg TID in both mild and hospitalized cases of COVID-19 infection with no safety issues [29, 30]. Silva et al. [32] used a higher dose of 1 g TID in mild and moderate COVID-19 infection; however, they changed it sooner to 500 mg QID as the first two participants did not tolerate the first dose (3 g/day), but the latter dose (2 g/day) was well tolerated [32]. Medhat et al. [28] used a similar dose of 500 mg QID in mild and moderate COVID-19 cases without safety issues as well [28]. In the AGILE trial [71], a higher dose of NTZ (1500 mg BID) for 7 days was investigated in healthy volunteers to determine the optimal dose, safety, and efficacy of NTZ in preventing COVID-19 infection [71]. Only self-limited to moderate gastrointestinal disturbance, urine, and scleral discoloration were reported. No severe adverse events were reported [71]. Therefore, with the above evidence about the least effective doses of NTZ being 500 mg orally TID to achieve at least IC50 against SARS-CoV2, the optimal dosing regimen and formulation are still being investigated.

4.1 Strengths

To the best of our knowledge, this is the first systematic review and meta-analysis of the effectiveness of NTZ in the treatment of COVID-19 infection, constituting the most robust evidence in this regard. Moreover, we strictly followed the PRISMA statement [33] and the *Cochrane Handbook for Systematic Reviews and Interventions* [34] and prospectively registered and published our protocol. Moreover, the quality of evidence was assessed using the GRADE recommendations.

4.2 Limitations

Our review has a few limitations: first, we only included six RCTs with limited demographic distribution; three studies in south America [29, 30, 32], one in North Africa [28], and another in the USA [31]. Second, we could not control multiple confounding variables, including baseline viral load and comorbidities. Third, all of the included RCTs recruited patients with mild to moderate disease except Rocco et al. [29], who included patients with COVID-19 pneumonia requiring hospitalization. Fourth, the NTZ treatment regimen, including dosage, formulation, administration times, and duration of treatment varied among the included RCTs. Fifth, none of the included studies assessed the effect of NTZ against the different variants of COVID-19. Sixth, all of the included trials have a high risk of bias in different domains, as we previously clarified. Seventh, we could not conduct a dose-response meta-analysis based on the included data as we only included six RCTs with three different dosing regimens. Finally, we detected significant heterogeneity regarding the viral clearance, and the GRADE assessment yielded very low-quality evidence for all the included outcomes; thus, the generalizability of our findings is limited.

4.3 Implications for Future Research

Despite the globally available vaccination protocols, widespread vaccination will require a long period to adequately prevent further infection transmission. Therefore, a safe, well-tolerated, and easy-to-administer antiviral agent is required for mild to moderate COVID-19 treatment [31]. Nitazoxanide looks promising in this regard; however, further research is still required to ascertain the following: first, the most effective dosage regimen is still to be investigated with various regimens used in the previous trials. In this regard, we support Blum et al. [27] that given the lack of information about the most clinically applicable dosage, conducting a pharmacokinetic study is important. Second, more work is needed to evaluate the effect of NTZ in combination with other antiviral agents to prevent the evolution of NTZ-resistant strains on the wide implementation of a single NTZ treatment regimen. Additionally, the NTZ viral evasion in a high viral load is yet to be evaluated. Third, although multiple immunological effects of NTZ have been clarified, more work is still required to evaluate the effect of NTZ on monocytes and interferons (IFN- α and IFN- β), given their important role in COVID-19 pathogenesis [72–74]. Finally, more phase III, multi-center, large-scale clinical trials are still required to ascertain the effects of NTZ in COVID-19.

5 Conclusions

Despite the efficacy of NTZ in accelerating viral clearance compared with placebo, evidence regarding the efficacy of NTZ in improving clinical resolution, reducing all-cause mortality, reducing ICU admission, and oxygen requirements is uncertain. This warrants more large-scale clinical trials to yield more generalizable and clinically applicable findings.

Supplementary Information The online version contains supplementary material available at https://doi.org/10.1007/s40261-022-01213-y.

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Consent to Participate Not applicable.

Consent for Publication Not applicable.

Availability of Data and Material The data are available upon request from the corresponding author.

Code Availability Not applicable.

Authors' Contributions MA conceived the idea. BA and MA designed the research workflow. BA and MA searched the databases. FL, RF, and BK screened the retrieved records, and MA resolved the conflicts. AK, FL, RF, and BK extracted relevant data, assessed the quality of evidence, and MA resolved the conflicts. MA and BA performed the analysis. MA and AG wrote the final manuscript. All authors have read and agreed to the final version of the manuscript.

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