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Author manuscript *Pediatr Res.* Author manuscript; available in PMC 2014 August 01.

Published in final edited form as:

Pediatr Res. 2014 February ; 75(2): 358-366. doi:10.1038/pr.2013.209.

# TRAJECTORY AND CORRELATES OF GROWTH OF EXTREMELY LOW BIRTH WEIGHT ADOLESCENTS

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### Abstract

**Background**—Catch-up growth may predispose to obesity and metabolic sequelae. We sought to examine the trajectory and correlates of growth and catch-up among extremely low birth weight (ELBW, <1 kg) adolescents.

**Methods**—Cohort study of 148 neurologically normal ELBW children and 115 normal birth weight (NBW) controls born 1992 through 1995. Longitudinal measures of gender-specific growth of ELBW children from birth, and growth and measures of obesity of ELBW and NBW children at 14 years.

**Results**—Following neonatal growth failure, ELBW children had accelerated growth, but at 8 years they still had lower weight and height z scores than NBW children. By 14 years ELBW boys had caught up in growth to their NBW controls but ELBW girls remained significantly smaller. ELBW children however did not differ from their controls in measures of obesity. In hierarchical multiple regression analyses only maternal BMI and weight gain during infancy and childhood predicted the ELBW children's 14-year weight z scores, BMI z scores and abdominal circumference. Perinatal risk factors including intrauterine growth only predicted growth up to 20 months.

**Conclusion**—Maternal BMI and rate of growth, rather than perinatal factors, predict 14-year obesity among neurologically normal ELBW adolescents.

Preterm infants have traditionally suffered from neonatal growth failure due to inadequate nutrition and chronic complications of prematurity. The majority catch up in growth, although their growth attainment may be less than that of NBW children (1,2). The implications of this catch-up growth for long term cardiovascular and metabolic health have

Category of Study: Population Study

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been the subject of interest for many years (3,4), but have become more critical since childhood obesity became epidemic (5).

We recently reported on the increase in rates of obesity between ages 8 and 14 years among ELBW children born 1992–1995 (6). In the current report, we sought to examine the children's gender specific trajectory and correlates of growth and catch-up from birth. We hypothesized that by age 14 years, the ELBW children would catch up to NBW controls in weight, height, and clinical measures of obesity, and that the predictors of growth would include socioenvironmental, perinatal and neonatal risk factors.

### RESULTS

### **Descriptors of ELBW and NBW Children**

ELBW boys and girls did not differ significantly from their respective NBW peers in maternal sociodemographic factors (SES) with the exception that mothers of ELBW girls were older and had a higher mean family income than mothers of NBW girls. ELBW girls reported less physical activity than NBW girls. Age of puberty was similar between groups (Table 1). Within the ELBW population, bronchopulmonary dysplasia (BPD, oxygen dependence at 36 weeks) and sepsis were the most common neonatal complications. ELBW boys had higher rates of BPD and postnatal steroid therapy than girls.

### Longitudinal Changes in Growth

The mean weight z scores, i.e. standard deviation scores (ZWT) of ELBW children at birth was -0.72 for boys and -0.96 for girls. Due to neonatal growth failure, these decreased by 40 weeks to -1.97 among boys and -2.02 among girls, and then increased to +0.07 and +0.14 respectively by age 14 years. Mean birth length or height z scores similarly decreased by 40 weeks and then increased to 14 years (Figure 1 and Supplemental Table S1 (online)). Among the ELBW children, catch-up growth (0.67 SD) occurred between all periods of study. Between ages 8 and 14 years, the increases in growth as measured by an increase in mean weight and height z scores per month or by catch-up growth, were significantly greater for ELBW boys than for NBW boys, whereas this was not evident among girls (Table 2). The growth of the children with measures from their biologic mothers was similar to that of the children of mothers who did not have growth measures (data not shown).

### Eight- and Fourteen-Year Growth of ELBW and NBW Children

Among both ELBW and NBW children the height z score was lower than ZWT at both 8 and 14 years (Table 3). At 8 years, ELBW boys and girls had significantly lower mean weight and height z scores than their NBW peers. At 14 years, although ELBW boys still had a lower weight and height than NBW boys, the differences were not significant. ELBW girls, however, remained significantly smaller than their peers. Nine of 15 girls who were overweight at 8 years became obese by 14 years with rates of obesity increasing from 12% to 21% (p=0.049) compared to an increase of 18% to 19% among the NBW children. The mean body mass index (BMI, wt/ht<sup>2</sup>) z scores and rates of obesity did not differ significantly between ELBW and NBW boys or girls at 8 or 14 years nor did the mean 14-year abdominal circumference.

### Longitudinal Correlates of Growth

The univariate perinatal and neonatal correlates of the weight and BMI z scores were significant up to 8 years, the majority only up to 20 months. They included preeclampsia, birth weight z score, small for gestational age (SGA) status, gestational age, total number of neonatal complications, and duration of hyperalimentation (Supplemental Tables S2 and S3 (online)). The predictors of the height z scores were similar (data not shown).

In the hierarchical multiple regression analyses (Table 4), birth weight z score was positively, and duration of hyperalimentation was negatively associated with ZWT at 40 weeks. Birth weight z score, gestational age, and change in ZWT from 40 weeks to 8 months' conceptual age (CA), along with maternal BMI and the interaction of maternal BMI and gender, were associated with ZWT at 8 months. Maternal BMI was positively associated with 8-month ZWT in girls ( $\beta = 0.048$  (95% CI, 0.018 to 0.078), p = 0.002), but not in boys  $(\beta = 0.004 (95\% \text{ CI}, -0.024 \text{ to } 0.031), \text{ p} = 0.80)$ . This relationship also held at other ages as illustrated in a plot of ZWT vs. maternal BMI by gender at age 14 years (Figure 2). Results of the regressions of ZWT at 20 months were similar to those at 8 months, except that gestational age was no longer a significant predictor. In the regression of ZWT at 8 years, maternal BMI along with changes in the child's ZWT from 40 weeks to 8 months, and from 8 months to 8 years were all significant predictors. In addition, changes in ZWT during all three prior periods were significantly associated with ZWT at 14 years (all p's <0.001). After adjusting for variables included in the final model, neither age of puberty, nor physical activity were significant predictors (data not shown). Hierarchical multiple regressions of the adolescent BMI z score and of abdominal circumference at 14 years were very similar to results for the 14-year ZWT (Table 5). Only 13 ELBW children of biologic mothers were obese at 14 years, precluding multivariable logistic regression modeling of obesity.

The findings were similar when SGA, defined either as weight < -2SD or <10<sup>th</sup> percentile for gestational age, was used in place of ZWT at birth, with the exception that gestational age was no longer a significant predictor (Supplemental Tables S4–S7 (online)).

### DISCUSSION

The results of this longitudinal study reveal that neurologically normal ELBW children demonstrated accelerated growth following neonatal growth failure. Although their weight and height z scores were still lower than those of NBW controls at age 8 years, by 14 years these did not differ significantly among boys, whereas ELBW girls continued to have significantly lower weight and height z scores than their NBW peers. Both ELBW and NBW children had higher weight than height z scores. The rates of obesity, mean BMI z scores and 14-year abdominal circumference of the ELBW children, although lower, did not differ significantly from those of their NBW peers. Multivariable analyses revealed that neither intrauterine growth, as measured by birth weight z score or SGA status, nor neonatal risk factors were associated with the child's gain in weight after 20 months' CA. The only factors that predicted 14-year measures of obesity were the child's weight gain during each of the periods studied and maternal BMI which affected girls more than boys.

The only study of the adolescent growth of preterm children born in the 1990s pertains to 11-year old <26-week gestation Swedish children who remained smaller than controls, but similar to our findings, did not differ in BMI (7). It is interesting that the obesity epidemic in the 1990s has been associated with an increase in the rates of obesity of both preterm and NBW children as compared to their rates of obesity reported in the 1980s (2,5,8).

Studies of predictors of growth among preterm children have pertained mainly to infancy and childhood. Correlates of reduced growth have included male gender, lower birth weight and gestation, SGA birth weight, and various neonatal complications (9–14). Similar to our results, Ehrenkranz found the duration of hyperalimentation to be inversely associated with growth, as sicker infants need hyperalimentation for longer periods (14). Smoking during pregnancy and postnatal steroid therapy were not significant predictors of growth in our univariate analyses possibly due to the fact that the effects of postnatal steroids on growth may be transient (9), and that maternal smoking affects fetal growth only in term born children (15). These factors were thus not included in our multivariable models. SES and ethnicity, known correlates of growth, were not significant predictors in our population. Self-reported physical activity was also not predictive, possibly related to the lesser physical activity of ELBW children (6).

Significant associations between parent and child growth have been reported in normal (16,17) and preterm populations (1,2,7,18,19), and similar to our findings, to be greater in girls than boys (17). Multiple factors contribute to effect of parental growth including genetic, hormonal, and shared environmental and psychosocial characteristics.

Our finding that weight gain throughout infancy and childhood is associated with 14-year measures of obesity is in agreement with the literature that there is no specific critical period of child growth that predicts later obesity (20) or its cardiovascular and metabolic sequelae (21–23). Although the ELBW children attained similar rates of obesity to their peers, their catch-up growth is of concern especially among girls whose rates of obesity increased significantly between 8 and 14 years, a finding also reported by Saigal (2). Among preterm children, rapid growth is associated with insulin sensitivity and may be a risk factor for type 2 diabetes and cardiovascular risk (24,25).

This is the first report of the sequential correlates of growth of preterm children born in the 1990s. Strengths of the study include its longitudinal design, relatively good follow-up rate, and the many risk factors considered. The rates of obesity of our NBW children and their mothers were representative of national data (5,26). We also acknowledge several limitations. Our results may have been influenced by the lower follow-up rate of ELBW boys and that participant boys had less BPD, thus lower neonatal risk, than non-participant boys. Difficulty in stretching immature sick infants may have influenced the initial length z scores, which were much lower than those of weight at birth and 40 weeks. Had we included the neurologically abnormal children, neonatal risk factors such as postnatal steroid therapy and periventricular hemorrhage, which predispose to the development of cerebral palsy, may have been predictive of the growth outcomes. Multiple births were included in the study as they did not influence longitudinal growth which was similar to that of singletons (data not presented). A further consideration is that early postnatal nutrition was less than currently

advised (27), and that we lack detailed information on neonatal nutrition including breast milk. However, current modes of neonatal nutrition have not eliminated childhood growth failure, and although beneficial for brain growth (28), may contribute to future metabolic risk (29). Maternal pre-pregnancy weight and paternal weight were also lacking, although the mother's current weight should reflect her pre-pregnancy weight and maternal effects on obesity are greater than paternal effects (17). The CDC norms are not representative of our urban, predominantly minority population, but we had a sociodemographically similar control population for comparison. We lacked measures of body composition and metabolic markers, but BMI, the measure we used, is associated with elevated body fat (30), and abdominal circumference is a proxy measure of abdominal fat mass (31).

The accelerated catch-up growth of the ELBW children and its potential associated cardiovascular and metabolic risk (32) is concerning as it may add to their high rates of chronic problems and further increase health care utilization (6). Possible intervention strategies include attempts to decrease maternal obesity and optimize diet to promote catch-up in height without promoting overweight. The latter may be very difficult as the lower height than weight z scores that we have documented reflect the notion that growth in weight may end in overshooting and obesity, whereas growth in height may be limited by a "self-stat" mechanism (33). Physical activities should also be encouraged despite the respiratory and subtle neurologic difficulties of ELBW children.

### **METHODS**

### **Extremely Low Birth Weight Group**

The birth cohort of 161 boys and 183 girls was admitted to Rainbow Babies and Children's Hospital, Cleveland, Ohio between 1992 and 1995. Thirteen children were excluded because of congenital conditions (6). Of the remaining children, 101 (65%) boys and 137 (78%) girls survived, of whom 70 (69%) boys and 111 (81%) girls were followed to age 14 years. Sixteen boys and 15 girls with cerebral palsy were excluded because of the known poor growth of neurologically abnormal children and 2 boys did not have 14-year growth measures. The study population thus included 52 boys and 96 girls who had 14-year growth measures. They did not differ from the non-participant birth cohort with the exception that fewer boys participated (51% boys vs 70% girls, p<0.01) and that the participant boys had lower rates of BPD and a shorter neonatal hospitalization.

### Normal Birth Weight Children

Sixty-five NBW boys and 111 NBW girls, of the same sex, race, school and age within 3 months, were recruited at 8 years, of whom 42 boys (65%) and 73 girls (65%) were followed to 14 years, all of whom had growth measures. They did not differ in SES from the non-participants with the exception that more mothers of boys were married.

### **Biologic Mothers**

Biologic mothers represented 116 (78%) caregivers of ELBW children and 101 (88%) caregivers of NBW children of whom 105 (91%) and 96 (83%) of mothers respectively had growth measures. Mothers of the ELBW children with growth measures were significantly

younger than those who did not have growth measures, but did not differ in SES, perinatal, or their children's neonatal risk factors. The mothers of ELBW children did not differ in weight, height or BMI from those of NBW children (Supplemental Table S8 (online)) and were also representative of national data for women age 40–59 (26).

### **Neonatal Care and Measures of Outcome**

Neonatal care was according to practice during the 1990s. The majority of infants received parenteral nutrition (hyperalimentation) of <3 mg/kg of protein per day. Sociodemographic, perinatal, and neonatal data were documented at neonatal hospital discharge. Weight and length were measured at birth and then at 40 weeks (term date as estimated from the last menstrual period and pregnancy ultrasound, when available); at 8 and 20 months' CA; and at 8 and 14 years' postnatal age. The children were measured according to standard procedures. The children were weighed unclothed but lightly clothed at 14 years. To correct for this clothing, we subtracted 1.0 kg for boys and 0.5 kg for girls. Length was measured supine with a tape measure at birth; with an infantometer at 40 weeks, 8 and 20 months' CA; and with a stadiometer after removing shoes at 8 and 20 years (Harpenden, Holtain, Crymych, UK). Maternal weight and height were similarly measured. The children's abdominal circumference, a proxy for visceral fat, was measured at 14 years according to the NHANES procedure (34).

Weight z scores were computed at birth and 40 weeks using standards which exclude infants delivered for maternal and fetal indications, many complicated with intrauterine growth failure (35). Length z scores at birth were computed according to Usher (36). At 8 and 20 months' CA and at 8 and 14 years, weight and height z scores were computed from the Center for Disease Control and Prevention (CDC) growth data (37). The CDC BMI norms (z scores) are only available from age 24 months. We could thus only calculate the BMI z scores of the ELBW cohort at 8 and 14 years. BMI was thus computed at 8 and 14 years and obesity defined as BMI 95th percentile. Catch-up growth was defined as an increase in weight or height z score (SD) of >0.67 (i.e., crossing of percentiles) (38). Additional 14-year measures included the adolescent self-report of physical activity during the last 4 weeks (39) and pubertal development (40).

The study was approved by the Institutional Review Board of University Hospitals Case Medical Center, Cleveland, Ohio. Written consent was obtained from parents and assent from children.

### Data analysis

Within the ELBW cohort, we examined gender specific growth parameters at each age, changes in z scores between each age studied, and rates of catch-up growth. The 8- and 14- year growth measures of the ELBW and NBW children were compared using two sample t-tests after adjusting for race and z-SES.

Correlates of growth were considered only for children of biologic mothers with growth measures (16). Pearson correlation coefficients calculated at each age studied included maternal education, race, and z-SES, defined as a composite of the sample z score for maternal education and family income (6). Perinatal data included a history of smoking

during pregnancy, preeclampsia, antenatal steroid therapy, birth weight z score, gestational age, SGA, considered both as birth weight < -2SD and as <the 10<sup>th</sup> percentile for gestational age, and multiple birth. Neonatal risk factors included the rates of BPD, sepsis (positive blood culture), severe cerebral ultrasound abnormality, necrotizing enterocolitis, the total number of these neonatal complications, the duration of parenteral nutrition (hyperalimentation), duration of hospitalization, and postnatal steroid therapy. Maternal growth correlates considered included weight, height, and BMI.

A hierarchical multiple regression approach was used to examine risk factors related to ZWT at 40 weeks, 8 and 20 months' corrected age and 8 and 14 years post natal age. ZWT, rather than BMI, was used for these longitudinal analyses as the CDC norms for BMI are only available from age 24 months (37). In the first and all stages, sociodemographic factors (z-SES, race, gender) and maternal BMI were forced into the models. In the first stage, interactions of gender with the other factors were tested. In the second stage, factors forced in or found to be statistically significant (p<0.05) in the first stage were included and perinatal factors associated with intrauterine and/or postnatal growth and the change in weight z score from birth to 40 weeks were then examined using stepwise regression, retaining those factors significant at p < 0.05. Interactions of birth and perinatal factors with gender were also examined in stage 2 and included if found significant (p<0.05). The third stage, carried out only when examining ZWT at 8 and 14 years, included terms retained in stages 1 and 2, and examined changes in ZWT from birth to 40 weeks, 40 weeks to 8 months, 8 months to 8 years, and 8 to 14 years (when examining ZWT at 14 years) using stepwise regression. Interactions of predictors found significant in stage 3 with gender were also examined. The maternal BMI\*gender interaction was significant in modeling ZWT at 8 and 20 months and 14 years, and bordered on significance in modeling these scores at 8 years (p=0.07); hence, this term was also included in the final 8-year model. Age of puberty and physical activity were each examined by testing whether they added significantly to the final model. A similar approach was used at 14 years to examine predictors of the child's BMI z score, rates of obesity and abdominal circumference. In addition, in separate analyses we examined the effect of SGA birth rather than ZWT on the longitudinal growth. All the analyses included only subjects with no missing covariates and growth measured at all-time points (n=94 ZWT, n=95 for BMI z score).

### Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

### ACKNOWLEDGEMENTS

Dr. Hack supervised the study, had full access to all of the data in the study, and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Drs. Hack, Taylor, and Schluchter developed the study concept and design.

Drs. Hack and Andreias acquired the data.

Drs. Hack, Schluchter, Taylor, Cuttler, Andreias, and Ms. Margevicius participated in the analysis, including statistical analysis, and interpretation of the data.

Dr. Hack drafted the first version of the manuscript and all the co-authors participated in the critical revision of the manuscript and approve its submission for publication.

We thank Kathy Winter who coordinated the project and participated in the interview of the parents; Ellen Durand MA and Heather Marcinick MA, research assistants, who tested the children, administered the questionnaires and measured the subjects; Bonnie Tarantino BA who provided clerical assistance; and Bonnie Siner RN who provided editorial assistance.

### Statement of Financial Support:

This study was supported by grants R01 HD 39756, M01 RR000, and ULI RR024989 from the U.S. National Institutes of Health (Bethesda, MD).

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a.



b.



### Figure 1.

Mean weight and height z scores  $\pm$  SD of extremely low birth weight (ELBW) boys and girls at birth, 40 weeks (term), 8 and 20 months, and 8 and 14 years, and mean weight and height z scores  $\pm$ SD of normal birth weight (NBW) boys and girls at ages 8 and 14 years. Weight measures are indicated in Panel a and height measures in Panel b. Measures of ELBW boys are indicated by dots and solid lines and measures of ELBW girls are indicated by squares and dashed lines. Measures of NBW boys are indicated by diamonds and solid lines and measures of NBW girls are indicated by triangles and dashed lines.

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### Figure 2.

Plots of weight z scores of ELBW boys and girls at 14 years vs. maternal BMI by gender illustrating the interactive effect of gender and maternal BMI on the weight z score. Estimated regression coefficients of maternal BMI on weight z score were 0.1068 (95 CI, 0.0633 to 0.1503) for girls and 0.0202 (-0.0279 to 0.0682) for boys.

### Table 1

### MATERNAL DEMOGRAPHIC FACTORS, PERINATAL DATA AND 14-YEAR OUTCOMES

	E	loys	6	Firls
	ELBW	NBW	ELBW	NBW
	n=52	n=42	n=96	n=73
Caregiver Demographic Data <sup>a</sup>		-		-
Age (years $\pm$ SD) <sup>b</sup>	41.3 ± 6	$40.3\pm7$	$43.1 \pm 6^{*}$	$40.6\pm6^{*}$
Married	25 (48%)	26 (62%)	36 (38%)	37 (51%)
Education				
<high school<="" td=""><td>5 (10%)</td><td>3 (7%)</td><td>6 (6%)</td><td>9 (12%)</td></high>	5 (10%)	3 (7%)	6 (6%)	9 (12%)
High school <sup>C</sup>	16 (31%)	9 (21%)	28 (29%)	13 (18%)
>High school	31 (60%)	30 (71%)	62 (65%)	51 (70%)
Race				
White	22 (42%)	15 (36%)	33 (34%)	24 (33%)
Black	30 (58%)	27 (64%)	63 (66%)	49 (67%)
Family income (mean dollars) $^d$	\$43,485	\$43,861	\$44,728	\$38,056*
Perinatal and Birth Data				I
Preeclampsia	8 (15%)	NA	17 (18%)	NA
Smoking	2 (4%)	NA	4 (7%)	NA
Antenatal steroid therapy	15 (42%)	NA	21 (30%)	NA
Birth weight (gm $\pm$ SD)	$825\pm119$	$3323\pm597^{\dagger}$	$813\pm124$	$3238\pm411^{\dagger}$
Gestational age (wk ± SD)	$26.5\pm2$	37	$26.5\pm2$	37
Small for gestational age				
<-2SD	7(13%)	NA	18 (19%)	NA
<10 <sup>th</sup> percentile	16 (31%)	NA	37 (39%)	NA
Multiple birth	7 (13%)	0	19 (20%)	0
Neonatal Risk Factors				
Necrotizing enterocolitis	2 (4%)	NA	5 (5%)	NA
Septicemia <sup>e</sup>	24 (46%)	NA	42 (44%)	NA
Cerebral abnormality <sup>f</sup>	11 (21%)	NA	15 (16%)	NA
Bronchopulmonary dysplasia	26 (50%)	NA	30 (31%)	NA
Total complications (mean $\pm$ SD)	$1.12 \pm 1$	NA	$0.96 \pm 1$	NA
Postnatal steroid therapy	33 (63%)	NA	44 (46%)	NA
Hyperalimenation (days)	32.8 ± 25	NA	$28.8 \pm 18$	NA
14-Year Outcomes				
Age at study (years $\pm$ SD)	$14.6\pm0.6$	$14.6\pm0.7$	$14.8\pm0.6$	$14.9\pm0.7$
Age of puberty (years $\pm$ SD)	12.0 ± 1.3	11.8 ± 1.3	11.8 ± 1.4	$12.0\pm1.2$
Physical activity $(\text{mean} \pm \text{SD})^g$	3.0 ± 1.0	$3.3\pm0.9$	$2.5 \pm 1.0$	$2.9 \pm 0.9^{**}$

ELBW, extremely low birth weight; NBW, normal birth weight; NA, not available or applicable

\* p<0.05;

\*\* p<0.01;

<sup>†</sup>p<0.001.

<sup>a</sup>Unless otherwise stated, refers to primary caregiver

 $^{b}$ Biologic mothers only: 40 ELBW and 36 NBW mothers of boys and 76 ELBW and 65 NBW mothers of girls.

<sup>C</sup>Includes General Education Diploma (GED)

 $^{d}$ Mean of median family income per US \$1000, according to the 2000 census tract neighborhood in which the families lived.

<sup>e</sup>Positive blood culture

<sup>f</sup>Cerebral ultrasound grade III-IV hemorrhage, periventricular leukomalacia and/or ventricular dilatation at discharge

 $^{g}$ Physical activity, subdomain mean score<sup>13</sup>

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# CHANGES IN WEIGHT AND HEIGHT Z SCORES BETWEEN TIME PERIODS OF STUDY

		EXTREME	LY LOW BIRT	H WEIGHT		NORMAL BIRTH WEIGHT
CHANGES IN Z SCORES	Birth to 40 weeks <sup>a</sup>	40 weeks to 8 months $^{b}$	8 to 20 months <sup>c</sup>	20 months to 8 years <sup>d</sup>	8 to 14 years <sup>e</sup>	8 to 14 years
BOYS, n=52						n=41
Z-WEIGHT						
0.67	1 (2%)	21 (47%)	7(16%)	29 (60%)	$15(29\%)^{*}$	5(12%)*
Per month	$-0.390\pm0.339$	$0.030 \pm 0.224$	$0.015\pm0.065$	$0.015\pm0.065$	$0.006\pm0.009^{*}$	$0.002\pm0.011^{*}$
Z-LENGTH/HEI	GHT					
0.67	6 (13%)	23 (56%)	15 (37%)	16 (33%)	18 (35%) <sup>**</sup>	2 (5%)**
Per month	$-0.336\pm0.576$	$0.132 \pm 0.179$	$0.054 \pm 0.124$	$0.005 \pm 0.013$	$0.005\pm0.008^{**}$	$**800.0\pm000.0$
GIRLS, n=96						n=72
Z-WEIGHT						
0.67	7 (7%)	44 (49%)	18 (20%)	55 (60%)	32 (33%)	21 (29%)
Per month	$-0.333\pm0.356$	$0.076\pm0.129$	$0.005 \pm 0.069$	$0.013 \pm 0.015$	$0.004\pm0.011$	$0.004 \pm 0.012$
Z-LENGTH/HEM	GHT					
0.67	8 (10%)	65 (83%)	36 (43%)	22 (25%)	8 (8%)	7 (10%)
Per month	$-0.391{\pm}0.520$	$0.195\pm0.129$	$0.048 \pm 0.067$	$0.001{\pm}0.012$	$-0.004\pm0.011$	$-0.002\pm0.011$
* p<0.05						

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\*\* p<0.01 Numbers of missing children: Weight,  $n = a_7$ , b9, c4, d1, e0. Length,  $n = a_{12}$ , b18, c7, d1.

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			Boys				Girls		
8 YEARS	ELBW	NBV	N M	ean Difference <sup>a</sup> or	ELBW	NBW	Mean	Difference <sup>a</sup>	
	n=51	n=4	1	Odds Ratio (95% $CI)^b$	96=u	n=73		dds Ratio 5% CI) <sup>b</sup>	
Weight									
$Mean \; (kg \pm SD)$	$28.1 \pm 1 0.$	8 33.1 ±	10.4	-4.8(-9.4,-0.3)*	$29.3 \pm 9.0$	$35.2 \pm 1$	1.2 -5.7(-	$-2.6, -8.7)^{\ddagger}$	
$Z \ score \pm SD$	$-0.37 \pm 1.$	4 0.37 ±	1.26 _0	.74(-1.31, -0.17)*	$-0.18 \pm 1.4$	$3  0.37 \pm 1.$	22 -0.52(-	0.10, -0.93)**	
Less than –2 SD	5 (10%)	1 (29	()	4.5(0.5,41.2)	8 (8%)	0 (0%)		NA	
Height									
Mean $(cm \pm SD)$	126.6 ± 8.	8 133.4 ±	- 8.0	$6.8(-3.2, -10.4)^{\dagger}$	$128.7 \pm 8.3$	3 136.2 ± 3	3.3 –7.3(·	$-4.7, -9.9)^{\dagger}$	
$Z \ score \pm SD$	$-0.75 \pm 1.$	[8 −0.03 ±	: 088 –0.	$72(-1.17, -0.27)^{**}$	$-0.46\pm1.12$	$0.18 \pm 1.$	04 -0.62(-	-0.30, -0.96) <sup>†</sup>	
Less than -2 SD	7 (14%)	1 (29	()	6.0(0.7,51.4)	7 ( <i>7</i> %)	0 (0%)		NA	
Measures of Obesity									
BMI, mean $(kg/m^2 \pm SD)$	$17.1 \pm 4.2$	e 18.3 ±	4.2	-1.1(-2.9,0.7)	$17.4 \pm 4.0$	$18.6 \pm 4$	.2 –1.2	(-2.4,0.08)	
Z score $\pm$ SD	$0.07 \pm 1.1$	0.4 ±	1.2 -	-0.36(-0.86, 0.13)	$0.05 \pm 1.4$	$0.42 \pm 1$	.1 -0.36	(-0.03,0.75)	
Underweight (<5 <sup>th</sup> %)	2 (4%)	2 (5%	( 9	0.8(0.1, 6.2)	12 (13%)	2 (3%)	4.8(	1.0,22.5)*	
Normal weight (5 <sup>th</sup> –84 <sup>th</sup> %)	40 (78%)	25 (61	(%)	2.4(0.9,6.0)	57 (60%)	45 (62%	.0 (9	)(0.5,1.7)	
Overweight (85 <sup>th</sup> –94 <sup>th</sup> %)	3 (6%)	8 (20	(%)	0.3(0.1, 1.1)	15 (16%)	13 (18%	.0 (9	9(0.4,2.0)	
Obese ( 95 <sup>th</sup> %)	6 (12%)	6 (15	(%)	0.7(0.2,2.5)	11 (12%)	13 (18%	0.0	5(0.5,1.7)	
			В	OYS			GIRI	S	
14 YEARS		ELBW	NBW	Mean Differen	ce <sup>a</sup> or	ELBW	NBW	Mean Differer	ce <sup>a</sup> or
		n=52	n=42	Odds Ratio (95	% CI)b	n=96	n=73	Odds Ratio (95	% CI) <sup>b</sup>
Weight									
$Mean \; (kg \pm SD)$		$8.5 \pm 21.0$	$63.0 \pm 18$	.0 -4.6(-12.7).	3.6) 2	$66.6 \pm 15.1$	$63.2 \pm 19.3$	-6.2(-11.4, -	*(6.0
$Z \text{ score } \pm SD$	)	).07 ± 1.38	$0.49 \pm 13$	9 -0.43(-1.01)	).15) (	$1.14 \pm 1.21$	$0.64\pm1.05$	-0.47(-0.12, -0	.83)**
Less than -2 SD		2 (4%)	2 (5%)	0.80(0.1,6	()	3 (3%)	(%0) 0	NA	

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		BOY	S		GIRI	S
14 YEARS	ELBW	NBW	Mean Difference <sup>a</sup> or	ELBW	NBW	Mean Difference <sup>d</sup> or
	n=52	n=42	Odds Ratio (95% CI)b	96=u	n=73	Odds Ratio (95% CI) <sup>b</sup>
Height						
Mean (cm $\pm$ SD)	$164.0 \pm 9.4$	167.1±7.6	-3.1(-6.7, 0.5)	$156.8\pm6.2$	$161.9 \pm 6.1$	$-5.1(-3.2,-7.0)^{\ddagger}$
$Z \text{ score} \pm SD$	$-0.36 \pm 1.06$	$0.02 \pm 0.82$	-0.38(-0.78,0.02)	$-0.73 \pm 0.97$	$0.05 \pm 0.9$	$-0.78(-1.08, -0.48)\dot{7}$
Less than –2 SD	3 (6%)	1 (2%)	2.7(0.3,29.0)	6 (%)	0 (0%)	NA
Measures of Obesity						
BMI, mean (kg/m <sup>2</sup> $\pm$ SD)	$21.4 \pm 6.1$	$22.4 \pm 5.5$	-0.9(-3.4,1.5)	$22.9 \pm 5.6$	$24.0\pm6.8$	-0.9(-2.8,1.0)
Z score $\pm$ SD	$0.17 \pm 1.1$	$0.40 \pm 1.4$	-0.24(-0.76, 0.28)	$0.48 \pm 1.1$	$0.65 \pm 1.0$	-0.15(-0.47, 0.18)
Underweight (<5 <sup>th</sup> %)	2 (4%)	2 (5%)	0.8(0.1, 6.0)	3 (3%)	1 (1%)	2.4(0.2,24.3)
Normal weight (5 <sup>th</sup> -84 <sup>th</sup> %)	39 (75%)	26 (62%)	1.8(0.7, 4.4)	(%02) (20%)	46 (63%)	1.3(0.7,2.5)
Overweight (85 <sup>th</sup> -94 <sup>th</sup> %)	3 (6%)	5 (12%)	0.5(0.1,2.1)	9 (%) 9	12 (16%)	$0.3(0.1,0.96)^{*}$
Obese ( 95 <sup>th</sup> %)	8 (15%)	9 (21%)	0.7(0.2, 1.9)	20 (21%)	14 (19%)	1.2(0.5, 2.6)
Abdominal Circumference (cm±SD)	$74.7 \pm 15.1$	$75.0 \pm 12.7$	-0.45 (-6.3,5.4)	$73.5\pm13.6$	$75.3\pm14.9$	-1.5(-5.9,2.9)
ELBW, extremely low birth weight; NBW	/, normal birth w	eight				

Postnatal age of NBW children at 8 years was  $9.0 \pm 0.9$  for boys and  $9.4 \pm 0.8$  for girls and at 14 years was  $14.6 \pm 0.7$  for boys and  $14.9 \pm 0.7$  for girls.

 $^{a}$ Extremely low birth weight minus normal birth weight

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 $^{b}$  Adjusted difference in means or adjusted Odds Ratio, when adjusting for race and socioeconomic status.

\* <0.05; \*\* <0.01;

 $t^{+}_{< 0.001}$ 

NA, not applicable as no NBW children had measures less than -2 SD

# Table 4

# MULTIPLE REGRESSION ANALYSIS OF PREDICTORS OF Z-WEIGHT FROM 40 WEEKS TO 14 YEARS (N=94)<sup>a</sup>

	40 Weeks	8 Months	20 Months	8 Years	14 Years
	Beta (95% CI)	Beta (95% CI)	Beta (95% CI)	Beta (95% CI)	Beta (95% CI)
Demographics, Maternal BMI					
Race (Black)	0.004 (-0.406,0.414)	-0.180(-0.626, 0.265)	-0.421 (-0.981,0.139)	0.014 (-0.377,0.406)	0.006 (-0.367,0.379)
Gender (male)	-0.084 (-0.449,0.282)	$-0.558 (-1.051, -0.066)^{*}$	-0.448 (-1.074,0.179)	-0.208 (-0.554,0.139)	-0.136 (-0.470,0.20)
z-SES	-0.027 (-0.253,0.199)	0.032 (-0.172,0.236)	0.046 (-0.218,0.310)	0.013 (-0.207,0.234)	0.033 (-0.177,0.244)
Maternal $BMI^b$	0.011 (-0.012,0.033)	$0.048(0.018,0.078)^{**}$	$0.060 \left( 0.021, 0.010  ight)^{**}$	$0.056~(0.023,0.088)^{\dagger}$	$0.065~(0.034,0.096)^{\dagger}$
Gender <sup>at</sup> Maternal BMI	1	$-0.044 (-0.085, -0.004)^{*}$	-0.050 (-0.103,0.003)	-0.040 (-0.083,0.004)	$-0.049 (-0.091, -0.008)^{*}$
Gender <sup>41</sup> Race	1	0.429 (-0.238,1.096)	0.594 (-0.245,1.433)		1
Birth and Perinatal Factors					
Z-birth weight	$0.388~(0.240,0.536)^{\dagger\prime}$	$0.501~(0.325,0.676)^{\dot{T}}$	$0.310\ (0.137, 0.483)^{*}$		
Gestational age		$0.120\ (0.009, 0.232)^{*}$			
Preeclampsia					
Hyperalimentation <sup>C</sup>	$-0.016(-0.023,-0.008)^{\dagger}$				
Neonatal risk factors <sup>d</sup>	1				
Change in ZWT from birth to 40 weeks	-			-	
Intermediate Growth (change in weight z	t score)				
40 weeks to 8 months		$0.595~(0.474,0.717)^{\dot{T}}$	$0.655 \left(0.497, 0.813\right)^{*}$	$0.565(0.436,0.694)^{\dagger}$	$0.523~(0.397,0.649)^{\dagger}$
8 months to 8 years				$0.732~(0.608, 0.857)^{\dagger}$	$0.622~(0.484,0.760)^{\dagger\prime}$
8 years to 14 years					$0.652~(0.430,0.874)^{\dagger}$
R-square	0.3911	0.6799	0.5647	0.7390	0.6857

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 $^{a}$ Biologic mothers and their ELBW children;

p < 0.05;p < 0.01;r = p < 0.001; Author Manuscript

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 $d_{\mathrm{Total}}$  number of neonatal risk factors.

 $^{c}$ Duration (days);

### TABLE 5

### PREDICTORS OF BMI Z SCORE AND WAIST CIRCUMFERENCE AT 14 YEARS (N=95)<sup>a</sup>

Independent Variables	BMI Z SCORE Beta (95% CI)	Waist Circumference Beta (95% CI)			
Demographics, Maternal BMI					
Race (Black)	-0.024 (-0.332, 0.283)	-3.525 (-8.045, 0.995)			
Gender (male)	-0.395 (-0.674, -0.115)**	-2.266 (-6.368, 1.837)			
ZSES	0.013 (-0.159, 0.185)	-1.217 (-3.748, 1.314)			
Maternal <sup>b</sup>	$0.055~(0.028,~0.081)^{\dagger}$	$0.754~(0.368,1.140)^{\dagger}$			
Gender <sup>a</sup> Maternal BMI	-0.045 (-0.078, -0.011)**	-0.775 (-1.264, -0.286)**			
Birth and Perinatal Factors	-				
Z-birth weight					
Gestational age					
Preeclampsia					
Hyperalimentation					
Total neonate risk factors					
Change Z-WT from birth to 40 weeks					
Intermediate Growth (Change in WI	Intermediate Growth (Change in WT z score)				
40 weeks to 8 months	0.411 (0.305, 0.517) <sup>†</sup>	4.245 (2.692, 5.798) <sup>†</sup>			
8 months to 8 years	$0.537~(0.424,~0.649)^{\dagger}$	$5.934~(4.284,~7.585)^{\dagger}$			
8 years to 14 years	0.616 (0.433, 0.799) <sup>†</sup>	$6.286~(3.597, 8.974)^{\dagger}$			
R-square	0.6931	0.5692			

\*\* p<0.01

 $^{\dagger}{}_{p<0.001}$ 

 $^{a}$ Biologic mothers and their ELBW children

 $^{b}{\rm Maternal~BMI}$  was centered at its overall mean