

was absent, however, for most of these relationships.

The present study, though hampered by a small sample size, identified selected variables as being candidate confounders and mediators in studies of sperm concentration and disease risk.

ACKNOWLEDGMENTS

GDB is supported by the UK Medical Research Council (MR/P023444/1) and the US National Institute on Aging (1R56AG052519-01; 1R01AG052519-01A1), and Martin Shipley by the British Heart Foundation.

G. David Batty,^a

Laust H. Mortensen,^b and

Martin J. Shipley^a

From the ^aDepartment of Epidemiology and Public Health, University College London, London, United Kingdom; and

^bStatistics Denmark and Department of Public Health, University of Copenhagen, Denmark.

REFERENCES

- Jensen TK, Jacobsen R, Christensen K, Nielsen NC, Bostofte E. Good semen quality and life expectancy: a cohort study of 43,277 men. *Am J Epidemiol.* 2009;170:559–565.
- Latif T, Kold Jensen T, Mehlsen J, et al. Semen quality as a predictor of subsequent morbidity: a Danish cohort study of 4,712 men with long-term follow-up. *Am J Epidemiol.* 2017;186:910–917.
- Latif T, Lindahl-Jacobsen R, Mehlsen J, et al. Semen quality associated with subsequent hospitalizations - can the effect be explained by socio-economic status and lifestyle factors? *Andrology.* 2018;6:428–435.
- Sharma R, Harlev A, Agarwal A, Esteves SC. Cigarette smoking and semen quality: a new meta-analysis examining the effect of the 2010 world health organization laboratory methods for the examination of human semen. *Eur Urol.* 2016;70:635–645.
- Ricci E, Al Beitawi S, Cipriani S, et al. Semen quality and alcohol intake: a systematic review and meta-analysis. *Reprod Biomed Online.* 2017;34:38–47.
- Sermondade N, Faure C, Fezeu L, et al. BMI in relation to sperm count: an updated systematic review and collaborative meta-analysis. *Hum Reprod Update.* 2013;19:221–231.
- The Centers for Disease Control Vietnam Experience Study. Health status of Vietnam veterans. II. Physical Health. *JAMA* 1988;259(18):2708–2714.
- Batty GD, Shipley MJ, Mortensen LH, et al. IQ in late adolescence/early adulthood, risk factors in middle age and later all-cause mortality in men: the Vietnam experience study. *J Epidemiol Community Health.* 2008;62:522–531.

OPEN Controlled Mediation as a Generalization of Interventional Mediation

To the Editor:

Moreno-Betancur and Carlin contrast interventional with natural mediation excellently.¹ They did not discuss the policy-relevant controlled mediation.² I illustrate how controlled is a generalization of interventional mediation, at least in simple settings, using Moreno-Betancur and Carlin's directed acyclic graph (DAG) in part A of their Figure that contains an outcome (Y), an exposure (A), and a mediator (M). Assumptions are no uncontrolled confounding of exposure and outcome, mediator and outcome, and exposure and mediator.³ With a binary exposure, the interventional effects are estimated from a hypothetical three-arm trial with exposure, control, and exposure then mediator intervention arms. The interventional effects arising are defined as potential outcomes weighted by the distribution of the mediator. The total effect is $\sum_m E(Y_{1m}) * P(M_1 = m) - \sum_m E(Y_{0m}) * P(M_0 = m)$; the direct effect is $\sum_m E(Y_{1m}) * P(M_0 = m) - \sum_m E(Y_{0m}) * P(M_0 = m)$; and the indirect effect is $\sum_m E(Y_{1m}) * P(M_1 = m) - \sum_m E(Y_{1m}) * P(M_0 = m)$. The mediator intervention in the third arm changes the distribution to the control group's.¹ Although the controlled direct

effect is defined at the individual level,⁴ the group-level interventional effect has been defined as a controlled direct effect under a stochastic intervention.³ Generalizing to include mediator change in the control group means adding a fourth arm to the trial, control then mediator intervention. The total effect is the same, but now the direct effect is $\sum_m E(Y_{1m}) * P(M_1 = m) - \sum_m E(Y_{0m}) * P(M_1 = m)$, and the indirect effects are $\sum_m E(Y_{1m}) * P(M_1 = m) - \sum_m E(Y_{1m}) * P(M_0 = m)$ and $\sum_m E(Y_{0m}) * P(M_0 = m) - \sum_m E(Y_{0m}) * P(M_1 = m)$ where M_i is the mediator distribution of the mediator intervention. There are indirect effects for the exposed and control groups. Their difference is the overall indirect effect which sums with the direct effect to the total effect. The direct effect measures the effect of A on Y not through M , whereas the overall indirect effect measures the effect of A on Y through M after mediation intervention. The two group-specific indirect effects compare the impact of intervening on M within each group. In contrast, the controlled indirect effect is usually not estimated, as its sum with the direct effect does not equal the total effect. Interaction of exposure with mediator on the outcome means the controlled direct effect can differ from the total effect, even when the exposure has no effect on the mediator.^{4,5} In other words, the difference is owing to interaction and not mediation. From an interventional perspective, this is an important result as mediator interventions may benefit health even when there is no relationship between exposure and mediator.⁶

In social epidemiology, for example, there is debate about the relative impact of population and targeted health behavior (smoking in the example below) interventions on socioeconomic health inequalities.⁷ Adapting a toy dataset,⁸ three mediation intervention scenarios are shown in the Table. The first changes the smoking rate of those disadvantaged to that of those advantaged, the next decreases the rate in both groups to 20%, and the third eliminates smoking. The outcome is 10-year probability of death. The direct effects show that eliminating differences in smoking do not fully

The author's post is funded by grants from the Medical Research Council, UK (MC_UU_12017/13), and the Chief Scientist Office of the Scottish Government Health Directorates (SPHSU 13).

The authors report no conflicts of interest. Reproducibility: Data is in the paper's table.

Copyright © 2019 The Author(s). Published by Wolters Kluwer Health, Inc. This is an open-access article distributed under the terms of the Creative Commons Attribution-Non Commercial-No Derivatives License 4.0 (CCBY-NC-ND), where it is permissible to download and share the work provided it is properly cited. The work cannot be changed in any way or used commercially without permission from the journal.

ISSN: 1044-3983/2019/303-e21
DOI: 10.1097/EDE.0000000000000980

TABLE. Differing Interventions on Smoking and Direct and Indirect Effects on Socioeconomic Inequalities

Trial arm	To That of Control Group	Probability of smoking	Lowered to Same Level	Probability of smoking	Smoke Free	Probability of smoking
	Probability of death		Probability of death		Probability of death	
1) Advantage, no intervention on smoking	0.48	0.4	0.48	0.4	0.48	0.4
2) Disadvantage, no intervention on smoking	0.736	0.6	0.736	0.6	0.736	0.6
3) Advantage, intervention on smoking	0.48	0.4	0.34	0.2	0.2	0
4) Disadvantage, intervention on smoking	0.624	0.4	0.512	0.2	0.4	0
Effects (probability of death)						
Total effect (=Arms 2–1)		0.256		0.256		0.256
Direct effect (=Arms 4–3)		0.144		0.172		0.2
Indirect effect, advantage (=Arms 3–1)		0		0.14		0.28
Indirect effect, disadvantage (=Arms 4–2)		0.112		0.224		0.336
Overall indirect effect (=IE, A1–IE, A0)		0.112		0.084		0.056
Not real data.						

explain away the socioeconomic differences and that the more drastic cuts in smoking have less impact on inequality (at least on the difference scale), while

decreasing the death risk the most. Why is this? Although intervention on smoking has a greater impact on the disadvantaged's death rate, the advantaged also

see reductions as their smoking rate falls and so the disadvantaged's relative gain is lessened.

Frank Popham

MRC/CSO Social and Public Health
Sciences Unit
University of Glasgow
Glasgow, United Kingdom
frank.popham@glasgow.ac.uk

REFERENCES

1. Moreno-Betancur M, Carlin JB. Understanding interventional effects: a more natural approach to mediation analysis? *Epidemiology*. 2018;29:614–617.
2. Naimi AI, Kaufman JS, MacLehose RF. Mediation misgivings: ambiguous clinical and public health interpretations of natural direct and indirect effects. *Int J Epidemiol*. 2014;43:1656–1661.
3. Vansteelandt S, Daniel RM. Interventional effects for mediation analysis with multiple mediators. *Epidemiology*. 2017;28:258–265.
4. VanderWeele TJ. Controlled direct and mediated effects: definition, identification and bounds. *Scand Stat Theory Appl*. 2011;38:551–563.
5. VanderWeele TJ. A unification of mediation and interaction: a 4-way decomposition. *Epidemiology*. 2014;25:749–761.
6. Kivimäki M, Shipley MJ, Ferrie JE, et al. Best-practice interventions to reduce socioeconomic inequalities of coronary heart disease mortality in UK: a prospective occupational cohort study. *Lancet*. 2008;372:1648–1654.
7. Capewell S, Graham H. Will cardiovascular disease prevention widen health inequalities? *PLoS Med*. 2010;7:e1000320.
8. Pang M, Kaufman JS, Platt RW. Studying non-collapsibility of the odds ratio with marginal structural and logistic regression models. *Stat Methods Med Res*. 2016;25:1925–1937.