

RESEARCH ARTICLE

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# Transcapillary escape rate of $^{125}\text{I}$ -albumin in relation to timing of blood sampling: the need for standardization

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## Abstract

**Background:** Increased vascular permeability is an early sign of vascular damage and can be measured with the transcapillary escape rate of albumin ( $\text{TER}_{\text{alb}}$ ). Although  $\text{TER}_{\text{alb}}$  has a multi-exponential kinetic model, most published  $\text{TER}_{\text{alb}}$  data are based on mono-exponential kinetic models with variation in blood sampling schemes. Aim of this posthoc study was to evaluate the influence of variation in blood sampling schemes and the impact of mono- or bi-exponential analyses on the calculation of  $\text{TER}_{\text{alb}}$ . Study participants were part of a cross-over intervention study protocol, investigating effects of sodium loading on blood pressure, endothelial surface layer and microcirculation. Multiple blood samples were drawn between 3 and 60 min after injection of radioactive iodide labeled human serum albumin (rHSA).

**Results:** In total 27 male participants with 54 measurements were included. For all participants the maximum serum radioactivity was reached within 20 min, while 85% of the participants had their maximum serum activity within 10 min. The  $\text{TER}_{\text{alb}}$  calculated with the subsequently chosen  $T_{20-60 \text{ min}}$  reference scheme ( $6.19 \pm 0.49\%/\text{h}$ ) was significantly lower compared to the  $\text{TER}_{\text{alb}}$  of the  $T_{3-60 \text{ min}}$ ,  $T_{5-60 \text{ min}}$ , and  $T_{\text{max}-60 \text{ min}}$  schemes. There was no significant difference between the  $T_{20-60 \text{ min}}$  reference scheme and the  $T_{10-60 \text{ min}}$  and  $T_{15-60 \text{ min}}$  schemes. Bi-exponential kinetic modeling did not result in significant different observations compared to the mono-exponential kinetic analysis.

**Conclusions:** As there is variation in the timing of the maximum serum radioactivity of rHSA, blood sampling schemes starting before 10 min after administration of rHSA will result in a significant overestimation of  $\text{TER}_{\text{alb}}$ . In addition, variation in kinetic modeling did not result in significant changes in  $\text{TER}_{\text{alb}}$ . Therefore, we emphasize the need to standardize  $\text{TER}_{\text{alb}}$  and for practical and logistical reasons advocate the use of a mono-exponential model with blood sampling starting 20 min after rHSA administration.

**Keywords:** Radioactive iodide labeled human serum albumin, Transcapillary escape rate of albumin,  $\text{TER}_{\text{alb}}$ , Vascular permeability

## Background

Diabetes mellitus and hypertension are characterized by an increased risk of vascular complications. An early sign of vascular damage is increased vascular permeability, which can be determined by the transcapillary escape rate of albumin ( $TER_{alb}$ ) (Broekhuizen et al., 2010).

$TER_{alb}$  is the rate in which intravenous albumin escapes from the intravascular to the extravascular volume in the first hour after injection of radioactive iodide labeled human serum albumin (rHSA) (Parving, 1975). The pharmacokinetics of rHSA could be described as the sum of three exponential components, with respective half-life's of 6.8 h, 1.29 days and 19.4 days (Berson et al., 1953; Bauman et al., 1955; CIS bio International, 2010). The disappearance of rHSA, in the first hour after injection, could be described as a bi-exponential decay curve with a inflection point after approximately 10 min (Margaron & Soni, 2002).

Despite the fact that rHSA has a multi-compartment kinetic model, all published  $TER_{alb}$  data analyses are based on a mono-exponential kinetic model. This mono-exponential  $TER_{alb}$  model has four assumptions: the rHSA behaves like endogenous albumin; the albumin metabolism is in steady state during the  $TER_{alb}$  test; rHSA has a mono-exponential blood pool elimination during the first hour after injection, with a rate constant equal to that of time zero; the initial blood pool elimination reflects extravasation and is not influenced by the rHSA metabolism rate (Parving & Gyntelberg, 1973).

The original protocol of Parving et al. describes that a small amount of I-125 or I-131 labeled rHSA is injected in an arm vein, and eight venous blood samples were drawn from the contralateral arm at 10, 15, 20, 30, 40, 50, 55, and 60 min after the injection. The radioactivity of the rHSA in each blood sample was measured in duplicate. The  $TER_{alb}$  was calculated and expressed as the percentage decline of radioactivity during the first hour (%/h) (Parving & Gyntelberg, 1973).

However most studies using  $TER_{alb}$  values show variation in sampling schemes ranging from 3 to 13 blood samples (Margaron & Soni, 2002; Jensen et al., 1992). Some of the schemes started already 1 min after the injection of rHSA, while others started blood sampling 20 min after the administration of rHSA (Margaron & Soni, 2002; Norberg et al., 2015).

This variation in sampling schemes does impact the calculated  $TER_{alb}$ . Sampling schemes that started 5 min after administration found  $TER_{alb}$  in the range of 6.9–9.1%/h (Dell'omo et al., 2006; Dell'Omo et al., 2000; Haskell et al., 1997; Pedrinelli et al., 2000; Pedrinelli et al., 1999; van Eijk et al., 2005; Rorije et al., 2018). While studies which started sampling 10 min after injection found a lower  $TER_{alb}$  of approximately 5.5%/h (Parving, 1975; Jensen et al., 1992; Nannipieri et al., 1997; Jensen, 1995; Nannipieri et al., 1995; Staberg et al., 1982; Zietse et al., 1995). These differences in  $TER_{alb}$  were not related to differences in patient population, but are in line with the multi-exponential kinetics of rHSA.

As the use of  $TER_{alb}$  for clinical research seems to gain in popularity, standardization of the technique is essential: i.e. reducing variation in performing the test and thereby reducing variation in the test result. As we observed large variations between different publications in  $TER_{alb}$  sampling schemes and most likely thereby variation in  $TER_{alb}$  results, we therefore aimed to study the influence of

different sampling schemes and the use of a mono- or bi-exponential analysis on the calculation of  $TER_{alb}$ .

## Methods

### Study population and study design

Selected participants of this post hoc study were part of a cross-over intervention study protocol investigating whether an acute intravenous sodium load, as compared to a chronic dietary sodium load, differs in its effects on blood pressure, the endothelial surface layer and microcirculation (Rorije et al., 2018). Participants included healthy men, and both male type 1 diabetes mellitus and hereditary multiple exostosis patients (i.e., patients with, respectively, acquired and genetically determined glycoalyx changes) (Mooij et al., 2014). Exclusion criteria were hypertension ( $\geq 140/90$  mmHg), obesity (body mass index (BMI)  $\geq 30$  kg/m<sup>2</sup>), history of primary hyperlipoproteinemia, coagulation disorders, and renal or cardiovascular diseases. All participants were randomized to a low sodium diet (LSD,  $< 50$  mmol Na<sup>+</sup> daily) or to a high sodium diet (HSD,  $> 200$  mmol Na<sup>+</sup> daily) for 8 days, separated by a crossover period of at least 1 week. The study was performed at the Amsterdam UMC, location AMC, Amsterdam, The Netherlands. All participants provided written informed consent and approval was obtained from the local ethics committee. The trial is registered in the Netherlands Trial Register (NTR4095 and NTR4788).

### Transcapillary escape rate of rHSA

An intravenous (IV) bolus of saline solution with rHSA labeled with 100 kBq I-125 was administered in a cubital vein. Blood samples were drawn from the contralateral arm at baseline and between 3 and 60 min after injection of rHSA. Radioactivity in plasma was measured in duplicate with a Wizard2 2480 automatic gamma counter (PerkinElmer, Waltham, Massachusetts, USA) with a coefficient of variation of  $< 3\%$ . The routine quality controls of the gamma counter were performed according to the standard GLP features of PerkinElmer, including detector energy resolution, background, absolute - and relative detector efficiency, detector stability probability and calibration.

The  $TER_{alb}$  was calculated with PKSolver, a free Microsoft Excel add-in for pharmacokinetic (PK) and pharmacodynamic (PD) data analysis (Zhang et al., 2010). PKSolver has been validated and has been used in different PK/PD studies (Zhang et al., 2010; Kulo et al., 2017; de Velde et al., 2016; Nezcic et al., 2014; Balakumar et al., 2013; Wenstedt et al., 2020).

$TER_{alb}$  was expressed as percentage decline in plasma radioactivity per hour (%/h). The  $TER_{alb}$  calculation with PKSolver was performed for an IV bolus administration. The formula used for the calculation of  $TER_{alb}$  was:

$$TER_{alb} = (A_{0 \text{ min}} - A_{60 \text{ min}}) / A_{0 \text{ min}}$$

The predicted activity of rHSA at  $T_{0 \text{ min}}$  ( $A_{0 \text{ min}}$ ) and at  $T_{60 \text{ min}}$  ( $A_{60 \text{ min}}$ ) were calculated by PKSolver (Microsoft Excel 2016) based on a mono- and bi-exponential kinetic model. This program also calculated the correlation coefficient (R) between the observed and predicted data.

### Sampling schemes

After acquiring the PK curves of rHSA, we calculated the  $TER_{alb}$  according the following simulated blood sampling schemes:

- $T_{3-60 \text{ min}}$ : 3, 4, 5, 10, 15, 20, 30, 45, and 60 min
- $T_{5-60 \text{ min}}$ : 5, 10, 15, 20, 30, 45, and 60 min
- $T_{10-60 \text{ in}}$ : 10, 15, 20, 30, 45, and 60 min
- $T_{15-60 \text{ min}}$ : 15, 20, 30, 45, and 60 min
- $T_{20-60 \text{ min}}$ : 20, 30, 45, and 60 min
- $T_{\max - 60 \text{ min}}$ : from individual  $A_{\max}$  till 60 min

All blood samples before  $A_{\max}$  of the PK curves were excluded for the calculation of  $TER_{alb}$ , irrespective of the sampling scheme.

### Statistics

The effect of different blood sampling schemes on the  $TER_{alb}$  values were analyzed by fitting a mixed model as implemented in IBM SPSS Statistics (version 26, IBM, USA). This mixed model uses a compound symmetry covariance matrix and is fitted using maximum likelihood. In the absence of missing values, this method results in the same  $p$  values as multiple comparisons tests (e.g. repeated measures ANOVA) that are less able to deal with missing values. Therefore, in the presence of missing values, the results can be interpreted like repeated measures ANOVA (Harrison et al., 2018). We used Bonferroni correction as post hoc test and  $p$  values  $< 0.05$  were considered statistically significant. Results were reported as mean  $\pm$  standard error of the mean (SEM). Bland-Altman plots were used to evaluate the level of agreement between two different blood sample schemes.

## Results

### Patient demographics

In total 27 men were included resulting in 54 PK curves (27 linked to the LSD and 27 linked to the HSD), based on 486 (54\*9 samples) blood sample analyses. The study population consisted of 12 healthy volunteers, 8 diabetes mellitus type I patients, and 7 patients with hereditary multiple exostoses. All volunteers were between 18 and 38 years old with a median age of 24 (range 18–38 years). Other characteristics of study participants are displayed in Table 1.

The blood serum disappearance of rHSA of the study participants is shown in Fig. 1. The pharmacokinetic graphic shows a bi-exponential slope of decay curve with inflexion point at 15–20 min.

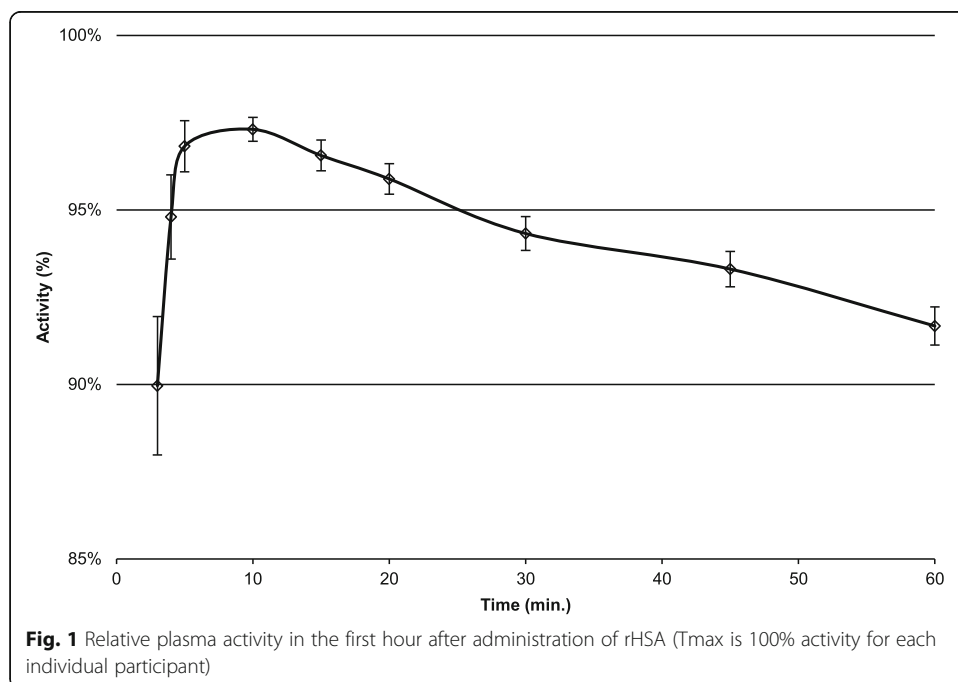
### $T_{\max}$ after rHSA administration

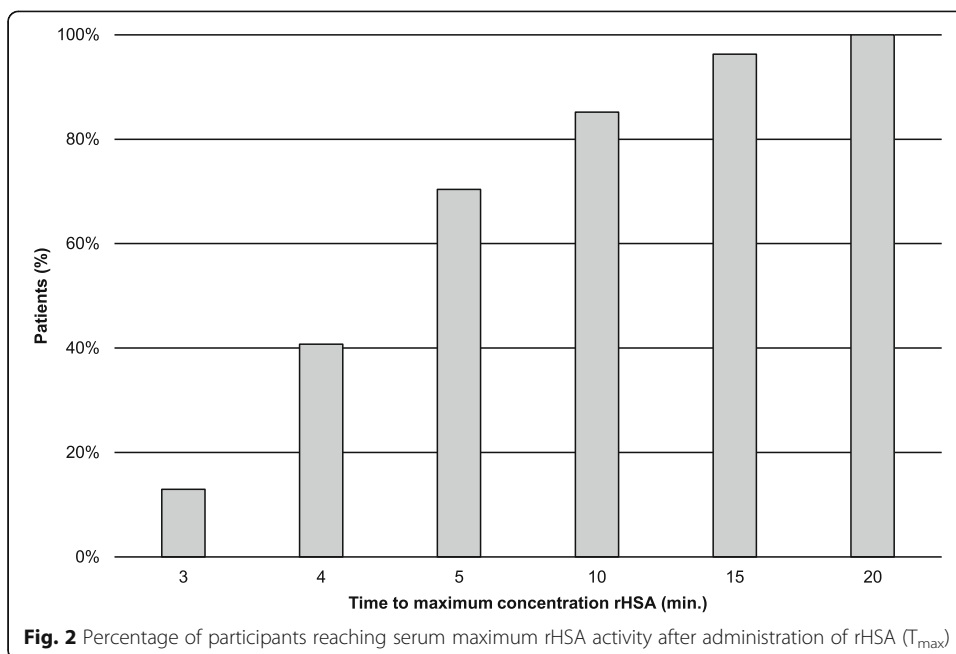
The  $T_{\max}$  after rHSA administration showed a large inter-individual variability (Fig. 2). The mean  $T_{\max}$  was  $6.9 \pm 0.6$  min. In 85% of the participants  $A_{\max}$  of rHSA was reached within 10 min, while  $T_{\max}$  was reached at 20 min after administration for all participants without an effect of subject category (HME, DM type 1 or healthy volunteer) or the diet followed (LSD vs HSD). Therefore  $T_{20-60 \text{ min}}$  was used as the reference

**Table 1** Characteristics of study participants

Characteristics of participants	Result
Health status (n)	27
Healthy	12
DM type I	8
HME <sup>a</sup>	7
Age (years median, range)	24 (18–38)
Healthy	21 (18–31)
DM type I	26 (21–37)
HME	21 (19–38)
Length (cm ± SD <sup>b</sup> )	183.7 (5.8)
Healthy	185.6 (6.3)
DM type I	184.3 (5.0)
HME	179.9 (4.1)
Weight (kg ± SD)	77.0 (7.6)
Healthy	75.7 (6.8)
DM type I	77.4 (9.4)
HME	78.8 (7.3)
BMI (kg/m <sup>2</sup> ± SD)	22.9 (2.5)
Healthy	22.0 (2.2)
DM type I	22.8 (2.5)
HME	24.4 (2.9)
eGFR <sup>c</sup> (ml/min ± SD)	118.1 (10.3)
Healthy	114.7 (12.1)
DM type I	120.3 (9.6)
HME	121.6 (6.3)

<sup>a</sup>HME (hereditary multiple exostoses), <sup>b</sup>SD (standard deviation), <sup>c</sup>eGFR based on CKD-EPI equation

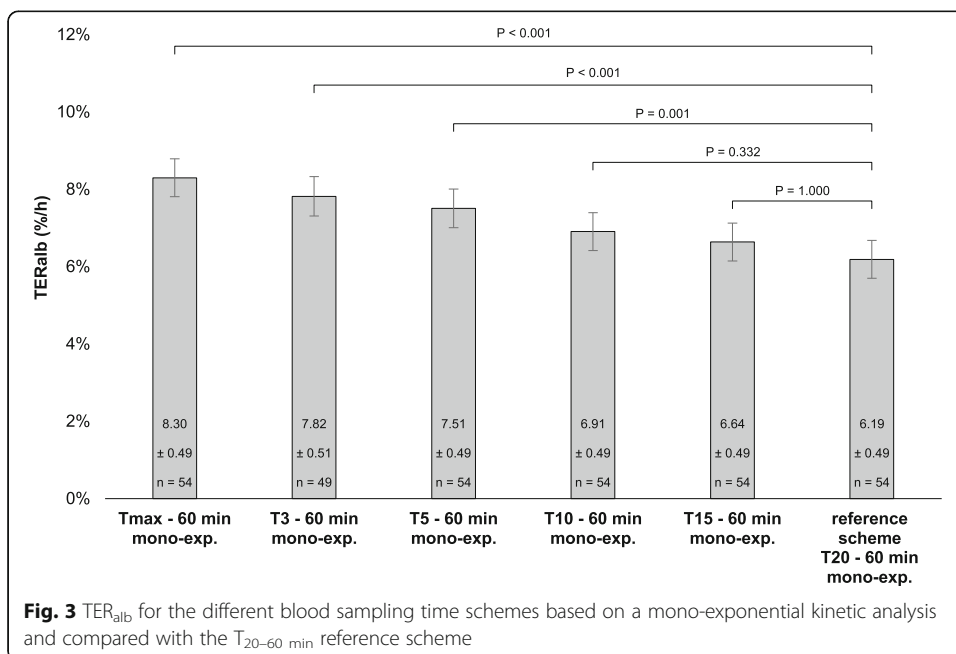




scheme. The mean  $TER_{alb}$  values of the other time schemes were compared to the reference scheme  $T_{20-60 \text{ min}}$  based on mono-exponential kinetic analysis.

**$TER_{alb}$  based on mono-exponential kinetic analysis**

The reference  $T_{20-60 \text{ min}}$  scheme included 54 of the 54 PK curves. The mean  $TER_{alb}$  of the  $T_{max} - 60 \text{ min}$  scheme resulted in the highest calculated  $TER_{alb}$ :  $8.30 \pm 0.49\%/h$

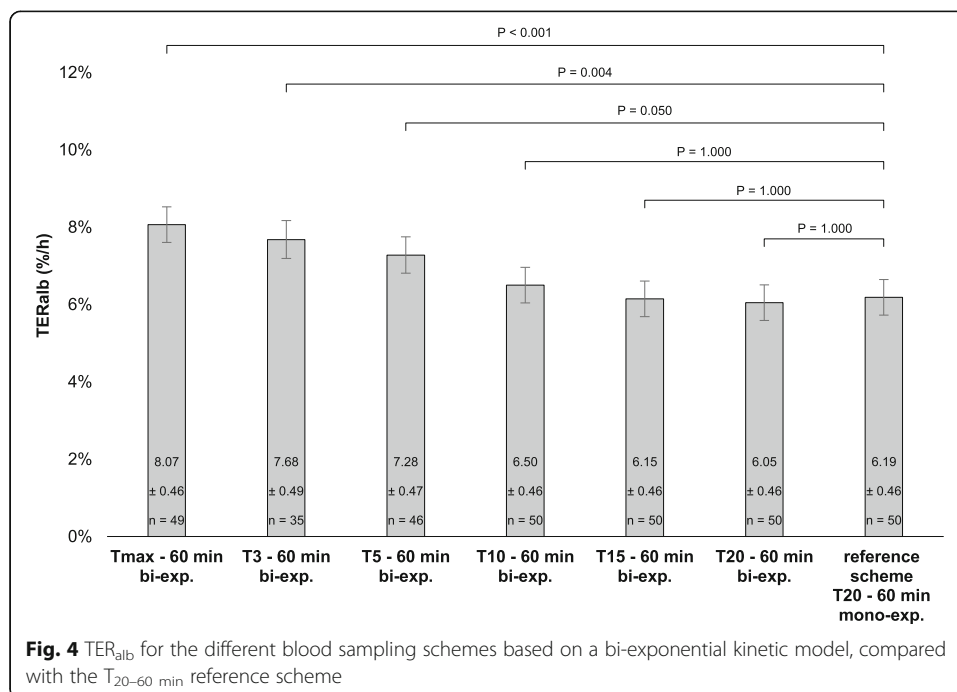


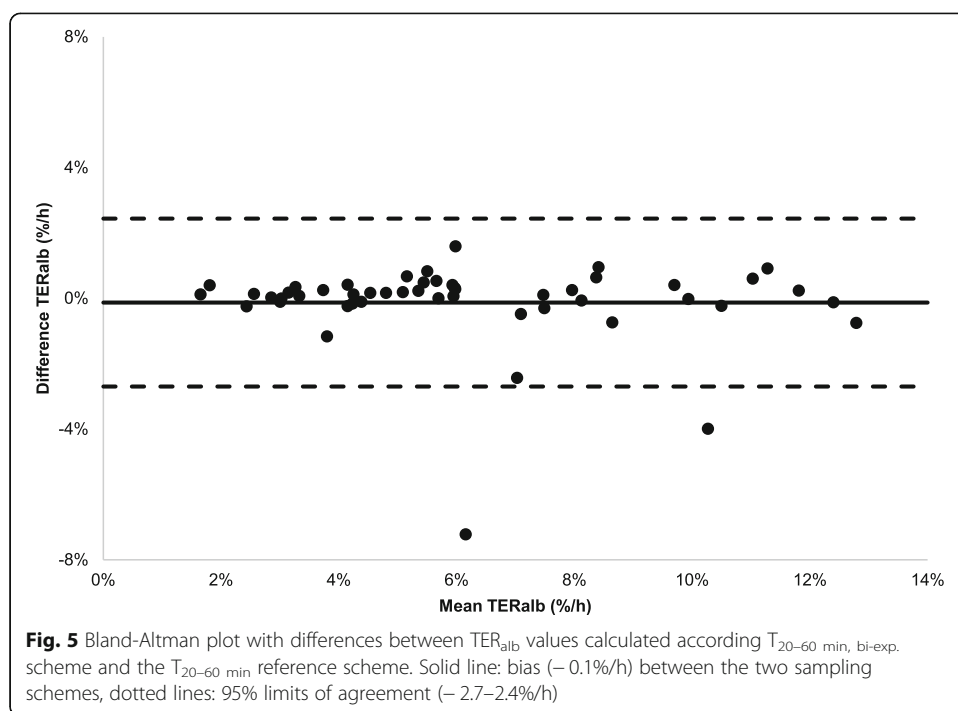
(Fig. 3). The  $TER_{alb}$  calculated with the  $T_{20-60 \text{ min}}$  reference scheme ( $6.19 \pm 0.49\%/h$ ) was significantly lower compared to the  $TER_{alb}$  of the  $T_{3-60 \text{ min}}$  (mean difference =  $-1.63\%/h$ ,  $CI = -2.65 - -0.61\%/h$ ,  $p < 0.001$ ),  $T_{5-60 \text{ min}}$  (mean difference =  $-1.32\%/h$ ,  $CI = -2.27 - -0.36\%/h$ ,  $p = 0.001$ ), and  $T_{\max - 60 \text{ min}}$  (mean difference =  $-2.11\%/h$ ,  $CI = -3.04 - -0.48\%/h$ ,  $p < 0.001$ ) schemes. There were no significant difference between the mean  $TER_{alb}$  of the  $T_{20-60 \text{ min}}$  reference scheme and the  $T_{10-60 \text{ min}}$  and  $T_{15-60 \text{ min}}$  scheme.

#### $TER_{alb}$ based on bi-exponential kinetic analysis

Using a bi-exponential analysis according to the  $T_{20-60 \text{ min}}$  scheme did not result in significant different  $TER_{alb}$  values when compared to the mono-exponential analysis based  $T_{20-60 \text{ min}}$  reference scheme (respectively  $6.19 \pm 0.46\%/h$  vs.  $6.05 \pm 0.46\%/h$ ,  $p = 1.000$ ). The mean  $TER_{alb}$  of the reference  $T_{20-60 \text{ min}}$  scheme was significantly lower compared to the mean  $TER_{alb}$  of the bi-exponential kinetic analysis of the  $T_{3-60 \text{ min}}$  (mean difference =  $-1.49\%/h$ ,  $CI = -2.68 - -0.29\%/h$ ,  $p = 0.004$ ),  $T_{5-60 \text{ min}}$  (mean difference =  $-1.09\%/h$ ,  $CI = -2.18 - -0.01\%/h$ ,  $p = 0.050$ ), and  $T_{\max - 60 \text{ min}}$  (mean difference =  $-1.87\%/h$ ,  $CI = -2.95 - -0.80\%/h$ ,  $p < 0.001$ ) schemes. There were no significant difference between the mean  $TER_{alb}$  of the  $T_{20-60 \text{ min}}$  reference scheme and the  $T_{10-60 \text{ min}}$  and  $T_{15-60 \text{ min}}$  schemes based on bi-exponential kinetic analysis (Fig. 4).

Figure 5 shows the Bland-Altman plot with agreement between the bi-exponential analysis based on  $T_{20-60 \text{ min}}$  scheme and  $T_{20-60 \text{ min}}$  reference scheme. The  $TER_{alb}$  showed a bias of  $-0.1\%/h$  between the different time schemes without a significant trend over the data range ( $1.7-12.8\%/h$ ) and with a consistent variability over the data range.





## Discussion

To our knowledge, this study is the first to examine the influence of different blood sampling schemes and the impact of mono- or bi-exponential analyses on the calculation of  $TER_{alb}$ . Our findings emphasize the necessity to standardize  $TER_{alb}$  calculations.

We found that the  $TER_{alb}$  became lower when blood sample collection started later. This phenomenon has been reported previously (Margaron & Soni, 2002). In this context it is remarkable that the majority of published studies used a fixed time sampling scheme with the first blood sampling within 10 min (Dell'omo et al., 2006; Dell'Omo et al., 2000; Haskell et al., 1997; Pedrinelli et al., 2000; Pedrinelli et al., 1999; van Eijk et al., 2005; Rorije et al., 2018). This practice will have caused an overestimation of the reported  $TER_{alb}$ . In addition, this makes the reported findings based on  $TER_{alb}$  difficult to reproduce and troublesome to extrapolate. Especially when  $TER_{alb}$  values of different sampling schemes are compared with each other.

Although the blood serum disappearance of rHSA should be described as a bi-exponential kinetic model, as shown in Fig. 1, the mean  $TER_{alb}$  values between mono- and bi-exponential analysis were not significantly different. Therefore, we concluded that the mono-exponential kinetic analysis, which is commonly used for  $TER_{alb}$  analysis, is a robust and easy-to-use approach to calculate the  $TER_{alb}$  in daily practice.

Our data showed that the biodistribution of rHSA seems to be complete after 15–20 min. Apparently, rHSA may need up to 20 min to reach an equilibrium. This inter-individual variation may be explained by the rate of lymphatic return or redistribution into the hepatic and splenic interstitium (Margaron & Soni, 2002; Henriksen & Schlichting, 1981). To minimize the number of blood samples, we advocate the use of a mono-exponential model with blood sampling starting 20 min after rHSA administration for daily practice. For scientific purposes, we suggest to use the  $T_{max}$  scheme to correct for inter- and intra-individual variability. It should be



noted that these TER<sub>alb</sub> values are significant higher compared to the daily practice scheme.

This study has several limitations that need to be addressed. First, we pooled data, because no differences between healthy participants, type 1 diabetes mellitus and hereditary multiple exostosis patients were detected. One would have expected higher TER<sub>alb</sub> values in type 1 diabetes mellitus patients. However, our study included young male with uncomplicated type 1 diabetes without albuminuria. Likely, at this stage of the disease, TER<sub>alb</sub> is still unaffected. Secondly, we did not collect any blood samples after T<sub>60 min</sub>. Blood sampling for longer time periods after administration, for example up to 24 h after rHSA injection, could have helped in better understanding the kinetics of rHSA blood clearance.

## Conclusions

To our knowledge, this study examined for the first time whether different blood sampling schemes impact TER<sub>alb</sub> values. We found significant differences between the blood sampling schemes which will cause bias in reporting TER<sub>alb</sub> and makes it difficult to reproduce and extrapolate outcomes of TER<sub>alb</sub>.

As there is a large variation in the timing of the maximum serum radioactivity of rHSA, blood sampling schemes starting before 10 min after administration of rHSA will result in a significant overestimation of TER<sub>alb</sub>. In addition, variation in mono- or bi-exponential kinetic modeling did not result in significant changes in TER<sub>alb</sub>. Therefore, we emphasize the need to standardize TER<sub>alb</sub> and for practical and logistical reasons advocate the use of a mono-exponential model with blood sampling starting 20 min after rHSA administration.

## Abbreviations

rHSA: radioactive iodide labeled human serum albumin; TER<sub>alb</sub>: Transcapillary escape rate of albumin; T<sub>max</sub>: Time to peak drug concentration; C<sub>max</sub>: Peak drug concentration; I-125: Iodine-125; I-131: Iodine-131; BMI: Body mass index; LSD: Low sodium diet; HSD: High sodium diet; GLP: Good Laboratory Practice

## Acknowledgements

Not applicable.

## Authors' contributions

YC and HV analyzed and interpreted the data. YC, SB and HV were major contributors in writing the manuscript. NR, LV and RM have substantively revised the manuscript. All authors read and approved the final manuscript.

## Authors' information

Not applicable.

## Funding

Not applicable.

## Availability of data and materials

The datasets used and/or analysed during the current study are available from the corresponding author on reasonable request.

## Ethics approval and consent to participate

All participants provided written informed consent and approval was obtained from the local ethics committee. The trial is registered in the Netherlands Trial Register (NTR4095 and NTR4788).

## Consent for publication

Not applicable.

## Competing interests

The authors declare that they have no competing interests.

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Received: 29 October 2020 Accepted: 3 February 2021

Published online: 16 February 2021

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