#### RESEARCH



# Are sleep disorders associated with the risk of gastrointestinal cancer?—A case-control study

 $Sven \ Loosen^1 \cdot Sarah \ Krieg^1 \cdot Andreas \ Krieg^2 \cdot Catherine \ Leyh^1 \cdot Tom \ Luedde^1 \cdot Céline \ Vetter^3 \cdot Karel \ Kostev^3 \cdot Christoph \ Roderburg^1$ 

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#### Abstract

**Purpose** Sleep disorders are among the most common health problems worldwide and are linked to a variety of physical and mental health problems. Recently, there has been increasing evidence of an association between sleep disorders and cancer risk. We aimed to investigate this association specifically for cancers of the gastrointestinal (GI) tract.

**Methods** Using the DA database (IQVIA), adult patients diagnosed with GI cancer between January 2010 and December 2022 were retrospectively compared to a 1:1 propensity score-matched cohort of patients without cancer. The outcome of the study was the association between sleep disorders and subsequent diagnosis of GI cancer. To determine whether sleep disorders were more common in patients with GI cancer than in patients without GI cancer, logistic regression models were used to estimate odds ratios (ORs) with 95% confidence intervals (95% CI).

**Results** After matching, 37,161 cases with GI cancer and 37,161 controls without cancer were available for analysis. No association with cancer was found for sleep disorders in the overall history before the index date (OR 1.04; 95% CI 0.96–1.12), but considering sleep disorders documented within 1 year before the index date showed a positive association with GI cancer overall (OR 1.20; 95% CI 1.08–1.34). Stratified analyses by cancer site revealed higher odds of sleep disorders prior to diagnosis of gastric, pancreatic, and colorectal cancer.

**Conclusion** Our findings suggest that sleep disorders might be indicative of short-term health outcomes, including GI cancer, suggesting a role for sleep disorder screening in the context of cancer prevention efforts.

Keywords Sleep · Sleep disorders · Cancer · Gastrointestinal cancer · Epidemiology · Prevention

Sven Loosen and Sarah Krieg authors share first authorship.

Karel Kostev and Christoph Roderburg authors share last authorship.

Christoph Roderburg christoph.roderburg@med.uni-duesseldorf.de

- <sup>1</sup> Department of Gastroenterology, Hepatology and Infectious Diseases, Medical Faculty of Heinrich Heine University Duesseldorf, University Hospital Duesseldorf, Moorenstraße 5, 40225 Duesseldorf, Germany
- <sup>2</sup> Department of Surgery (A), Medical Faculty of Heinrich Heine University Duesseldorf, University Hospital Duesseldorf, 40225 Duesseldorf, Germany
- <sup>3</sup> Epidemiology, IQVIA, 60549 Frankfurt, Germany

## Introduction

Sleep disorders are one of the most common health problems in the general population and result in significant socioeconomic and healthcare costs (Lateef et al. 2011; Ohayon 2011; Grandner 2020). Of these, insomnia and obstructive sleep apnea (OSA) have an estimated prevalence of 6–20% and 9–38%, respectively, in the general population (Riemann et al. 2017; Senaratna et al. 2017). Insomnia is characterized by delayed sleep onset, impaired sleep maintenance, reduced total sleep time, and/or early morning awakening (Sutton 2021). Obstructive sleep apnea (OSA), on the other hand, is defined by episodes of complete or partial airway collapse associated with decreased oxygen saturation or arousal from sleep (Slowik et al. 2023).

Sleep is known to be important for human health (Schlack et al. 2013; Grandner 2020). Indeed, sleep disorders have been reported to be associated with a wide range of physical

and mental health problems (Schlack et al. 2013; Grandner 2020). As such, sleep disorders have been discussed as an independent risk factor for myocardial infarction, heart failure and hypertension (Laugsand et al. 2011, 2014; Palagini et al. 2013; Lin et al. 2017; Li et al. 2021), as well as stroke (Bassetti 2005). Additionally, sleep disorders have been associated with weight gain and an increased risk of developing diabetes or metabolic syndrome (Patel et al. 2008; Anothaisintawee et al. 2016). A meta-analysis by Baglioni et al. found that disrupted sleep was also a risk factor for subsequent depression, with twice the risk compared with undisturbed sleep (Baglioni et al. 2011). Other studies have linked sleep disorders to the risk of infectious and inflammatory diseases and all-cause mortality (Kripke et al. 2002; Dew et al. 2003; Irwin et al. 2004).

In recent years, there has even been growing evidence of an association between sleep and cancer (Büttner-Teleagă et al. 2021; Mogavero et al. 2021). In this context, several studies have suggested an increased risk of several cancer types including breast cancer (Davis et al. 2001), liver cancer (Hu et al. 2013), prostate cancer (Fang et al. 2015), and colorectal cancer (Stevens et al. 2011; Lin et al. 2019). However, the relationship between different sleep disorders and specific cancer site is not very clear and appears to be bidirectional (Mogavero et al. 2021). Given the limited data available to date, this study aims to investigate this association specifically for gastrointestinal (GI) cancer using a large real-world cohort from Germany based on the IQVIA Disease Analyzer (DA) database.

## Methods

#### Database

This study used data from German primary care practices from the DA database (IQVIA). Details of the methodology have been published previously (Rathmann et al. 2018). In brief, the DA database contains data on demographic variables, diagnoses and prescriptions obtained in general and specialized practices in Germany. Practices included in the database are selected according to the yearly statistics of the German Medical Association, which include information on physician's age, specialty group, community size category, and German federal state. The database covers approximately 3–6% of all private practices in Germany. It has previously been shown that the panel of practices included in the DA database is representative of general and specialized practices in Germany (Rathmann et al. 2018).

#### **Study population**

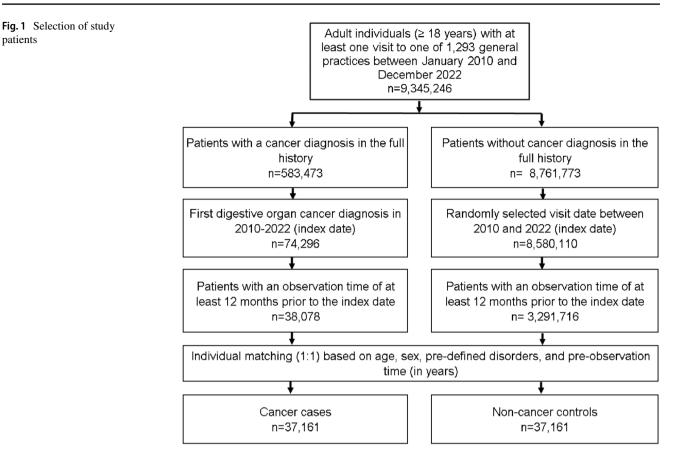
The study population included all patients aged  $\geq 18$  years with a cancer diagnosis of digestive tract (ICD-10 code: C15-C26) between January 2010 and December 2022 (index date) who had at least one year of observation prior to the index date. Controls were patients without cancer who were matched (1:1) by greedy propensity scores based on age, sex, their pre-diagnostic observation time in years, and predefined diagnoses which may be associated with increased risk of GI cancer. These diagnoses documented prior to the index date included diabetes mellitus (ICD-10: E10-E14), obesity (ICD-10: E66), chronic bronchitis and COPD (as possible proxy for smoking, ICD-10: J42-J44), diseases of esophagus, stomach, and duodenum (ICD-10: K20-K31), inflammatory bowel diseases (ICD-10: K50, K51) and diseases of liver (ICD-10: B18, K70-K77). For individuals without cancer, the index date was a randomly selected visit date between January 2010 and December 2022. The flow diagram of study participants is shown in Fig. 1.

#### Study outcome

Outcome of the study was the association between sleep disorders documented prior to the index date and subsequent GI cancer diagnosis. Sleep disorders (ICD-10: G47) included insomnia, hypersomnia, circadian rhythm sleep disorders, sleep apnea, narcolepsy and cataplexy, parasomnia, sleep related movement disorders, and unspecified sleep disorders. However, general practitioner usually documented unspecified sleep disorders (ICD-10: G47.9). Other ICD-10 codes were very rarely documented. Based on the frequency of sleep disorder detail documentation, analyses were conducted for sleep disorders in total only.

#### Statistical analyses

Demographic and clinical characteristics of cases and controls after 1:1 propensity-score matching were evaluated using the Wilcoxon signed-rank test for continuous variables, the McNemar test for categorical variables with two categories, and the Stuart-Maxwell test for categorical variables with more than two categories. To examine whether a history of sleep disorders was more frequent in patients with GI cancer, as compared to those with no GI cancer diagnosis, we used logistic regression models and estimated odds ratios (ORs) with 95% confidence intervals (95%CI) for sleep disorders across both patient groups. A first model analyzed the association between GI cancer and a history of sleep disorders (no matter when sleep disorder was documented in the history). In the second model, sleep disorders



were classified based on the duration of disease which were calculated as number of years since sleep disorder diagnosis. Specifically, if a given sleep disorder was documented within one year prior to the GI cancer index date, time since diagnosis was set to one year. Both models were also stratified by age groups ( $\leq 50$ , 51–60, 61–70, 71–80, > 80 years), and sex. In addition to all GI cancer cases, analyses were stratified by cancer site, colorectal (ICD-10: C18-C20), gastric (ICD-10: C16), esophageal (ICD-10: C15), liver (ICD-10: C22), and pancreas cancer (ICD-10: C25). To account for multiple comparisons, we adjusted the threshold for statistical significance to p-values < 0.01. All analyses were performed using SAS 9.4 (SAS Institute, Cary, US).

## Results

#### **Baseline characteristics**

After 1:1 matching, 37,161 GI cancer cases and 37,161 noncancer controls were available for analyses. Mean (standard deviation, SD) age at the index date was 69.4 (12.7) years and 45.1% were female. On average, both cases and controls had 8.7 (SD 6.1) years of pre-observation time prior to index date. Case and controls had the same proportions of predefined diagnoses (Table 1).

## Association between sleep disorders and cancer by cancer site

Sleep disorders in the whole history prior to the index date were not associated with cancer (OR: 1.04; 95% CI 0.96–1.12), however sleep disorders documented within one year prior to the index date was positively associated with overall GI cancer (OR: 1.20; 95% CI 1.08–1.34). Stratified analyses by cancer site showed higher odds of sleep disorders prior to gastric, pancreas, and colorectal cancer diagnosis, but these associations did not reach statistical significance (Table 2).

## Association between sleep disorders and cancer by age and sex

Overall, stratified analyses by age and sex demonstrated that positive associations between cancer and prior sleep disorder diagnoses were only observed in the age group 61–70 years (OR: 1.48; 95% CI 1.18–1.86), and in men (OR: 1.29; 95% CI 1.11–1.49) (Table 3).

Table 1 Characteristics of study patients after 1.1 matching

| Variable  | GI cancer ( $n = 37, 161$ ) | No cancer $(n = 37, 161)$ | P-value |
|---|-----------------------------|---------------------------|---------|
| Age (in years)  |                             |                           |         |
| Mean (SD)   | 69.4 (12.7)                 | 69.4 (12.7)               | 1.000   |
| ≤50   | 16.5                        | 16.5                      | 1.000   |
| 51-60   | 24.3                        | 24.3                      |         |
| 61–70   | 31.4                        | 31.4                      |         |
| 71–80   | 31.4                        | 31.4                      |         |
| > 80  | 20.4                        | 20.4                      |         |
| Sex   |                             |                           |         |
| Female  | 45.1                        | 45.2                      | 0.802   |
| Male  | 54.9                        | 54.8                      |         |
| Observation time prior to the index date (years), mean (SD) | 8.7 (6.1)                   | 8.7 (6.1)                 | 0.778   |
| Conditions documented prior to or at the index              | ate*                        |                           |         |
| Diabetes mellitus   | 31.6                        | 31.6                      | 1.000   |
| Obesity   | 12.5                        | 12.5                      | 1.000   |
| Chronic bronchitis / COPD                                   | 13.9                        | 13.9                      | 1.000   |
| Diseases of esophagus, gastric and duodenum                 | 35.0                        | 35.0                      | 1.000   |
| Inflammatory bowel diseases                                 | 0.9                         | 0.8                       | 0.259   |
| Liver diseases  | 17.6                        | 17.6                      | 1.000   |

Data are absolute samples and percentages unless otherwise specified

SD standard deviation

## Discussion

Using the IQVIA DA database, this retrospective case-control study examined the association between sleep disorders and a subsequent diagnosis of GI cancer in a real-world cohort in Germany. Adult patients diagnosed with GI cancer were compared with a control group of patients without cancer, matched 1:1 for age, sex, and diagnoses associated with an increased risk of GI cancer. Overall, there was no association between sleep disorders and cancer when the whole history before the index date was considered. Interestingly, however, when looking at the year prior to the index date, sleep disorders were positively associated with all GI cancers. These findings suggest that sleep disorders are not a risk factor for cancer, but rather an early symptom or indication of cancer. Stratified analyses by cancer site provided evidence of higher odds of sleep disorders before diagnosis of gastric, pancreatic, and colorectal cancer.

The results of our study add to an existing body of evidence linking sleep disorders to GI cancers. A retrospective case-control study by Lin et al. (Lin et al. 2019) estimated the risk of colorectal cancer in patients with sleep disorders (Lin et al. 2019) using multivariate logistic regression analysis on data from 7355 participants with colorectal cancer from the National Health Insurance Research Database of Taiwan (Lin et al. 2019). The authors found that sleep disorders were significantly associated with an increased risk of colorectal cancer. If a patient suffered from both sleep

disorders and depression, the incidence of colorectal cancer was five times higher than in the control group (Lin et al. 2019).

In particular, OSA has been implicated in the development and progression of cancer; however, the evidence for this association is conflicting (Phillips et al. 2012; López et al. 2016). While some studies have found an association between OSA and cancer incidence (Lowery-Allison et al. 2018; Walker et al. 2019), other studies have failed to find this association (Mogavero et al. 2021). A retrospective nationwide analysis from Chicago examined a cohort of about 5.6 million individuals based on a health insurance database for the incidence of OSA and 12 different cancer types between 2003 and 2012. As a result, OSA was associated with only a limited number of different cancer types, including kidney cancer, pancreatic cancer, and melanoma (Gozal et al. 2016). Another study of 34,402 subjects found a differential tumor risk in OSA patients depending on the anatomic site, with a significantly higher incidence of breast, kidney, uterine, and melanoma cancers and, remarkably, a lower incidence of lung and colorectal cancers compared to the general population (Sillah et al. 2018). In contrast, a case-control study by Lee et al. examining the prevalence of colorectal neoplasia in patients with and without OSA who underwent screening colonoscopy showed that patients with OSA were approximately three times more likely to develop advanced colorectal neoplasia than controls matched for age, sex, BMI, and smoking. Consequently, the authors

| Table 2 Association between sleep disorders and subsequent GI cancer by cancer site and sleep disorder duration in patients followed in general |
|---|
| practices in Germany  |

| Variables / groups                            | Proportion (%) in patients with GI cancer | Proportion (%) in patients without cancer | OR (95% CI)      | <i>P</i> -value |
|---|---|---|------------------|-----------------|
| Cancer in total $(n = 37, 161 \text{ pairs})$ |   |   |                  |                 |
| Any past sleep disorder diagnosis             | 4.34                                      | 4.19                                      | 1.04 (0.96–1.12) | 0.326           |
| Within $\leq 1$ year                          | 2.22                                      | 1.85                                      | 1.20 (1.08–1.34) | < 0.001         |
| Within $> 1-3$ years                          | 0.93                                      | 0.96                                      | 0.97 (0.83-1.14) | 0.725           |
| Within $> 3-5$ years                          | 0.39                                      | 0.42                                      | 0.94 (0.74–1.20) | 0.635           |
| >5 years ago                                  | 0.79                                      | 0.96                                      | 0.83 (0.70-0.98) | 0.025           |
| Colorectal cancer ( $n = 17,730$ pairs)       |   |   |                  |                 |
| Any diagnosis in the history                  | 3.93                                      | 4.00                                      | 0.98 (0.88-1.10) | 0.743           |
| $\leq 1$ year                                 | 2.05                                      | 1.75                                      | 1.17 (1.00–1.37) | 0.047           |
| > 1-3 years                                   | 0.83                                      | 0.89                                      | 0.93 (0.74-1.18) | 0.564           |
| >3-5 years                                    | 0.34                                      | 0.42                                      | 0.81 (0.57-1.15) | 0.232           |
| >5 years                                      | 0.70                                      | 0.94                                      | 0.75 (0.59-0.95) | 0.016           |
| Gastric cancer ( $n = 3,566$ pairs)           |   |   |                  |                 |
| Any diagnosis in the history                  | 4.63                                      | 4.21                                      | 1.10 (0.87–1.39) | 0.409           |
| $\leq 1$ year                                 | 2.27                                      | 1.60                                      | 1.42 (1.00-2.03) | 0.051           |
| > 1-3 years                                   | 0.93                                      | 1.32                                      | 0.70 (0.45-1.11) | 0.134           |
| >3-5 years                                    | 0.36                                      | 0.41                                      | 0.90 (042-1.94)  | 0.781           |
| >5 years                                      | 1.07                                      | 0.88                                      | 1.22 (0.75–1.99) | 0.434           |
| Esophagus cancer $(n = 2,175 pairs)$          |   |   |                  |                 |
| Any diagnosis in the history                  | 4.32                                      | 4.42                                      | 0.98 (0.72-1.32) | 0.878           |
| $\leq 1$ year                                 | 2.21                                      | 2.00                                      | 1.10 (0.72–1.69) | 0.662           |
| > 1-3 years                                   | 0.97                                      | 0.72                                      | 1.34 (0.68–2.64) | 0.398           |
| >3-5 years                                    | 0.18                                      | 0.41                                      | 045 (0.13-1.49)  | 0.189           |
| >5 years                                      | 0.97                                      | 1.28                                      | 0.75 (0,42–1.35) | 0.336           |
| Liver cancer $(n = 2,271 \text{ pairs})$      |   |   |                  |                 |
| Any diagnosis in the history                  | 5.15                                      | 4.35                                      | 1.19 (0.90–1.59) | 0.219           |
| $\leq 1$ year                                 | 2.51                                      | 2.47                                      | 1.02 (0.70-1.51) | 0.904           |
| > 1-3 years                                   | 1.23                                      | 0.94                                      | 1.32 (0.74–2.38) | 0.348           |
| > 3-5 years                                   | 0.44                                      | 0.20                                      | 2.24 (0.70-7.17) | 0.172           |
| >5 years                                      | 0.97                                      | 0.74                                      | 1.32 (0.68–2.55) | 0.412           |
| Pancreas cancer ( $n = 4,483$ pairs)          |   |   |                  |                 |
| Any diagnosis in the history                  | 5.60                                      | 4.42                                      | 1.28 (1.05-1.56) | 0.014           |
| $\leq 1$ year                                 | 2.70                                      | 1.99                                      | 1.37 (1.03–1.83) | 0.032           |
| > 1-3 years                                   | 1.23                                      | 0.96                                      | 1.29 (0.85–1.96) | 0.224           |
| > 3–5 years                                   | 0.67                                      | 0.45                                      | 1.49 (0.83-2.69) | 0.182           |
| >5 years                                      | 1.00                                      | 1.01                                      | 1.01 (0.66-1.54) | 0.978           |

OR odds ratio, CI confidence interval

recommended that physicians explain the need for colonoscopy to patients with OSA (Lee et al. 2017). Additionally, a large retrospective multicenter cohort study of clinical and health administrative data from more than 30,000 individuals from Canada with suspected obstructive sleep apnea (OSA) found that OSA severity was associated with cancer incidence independent of known cancer risk factors, with the highest risk in individuals with severe hypoxemia (Kendzerska et al. 2021). Furthermore, several studies have been published that specifically address the relationship between cancer and insomnia, but again, the results are inconsistent (Shi et al. 2020). In this context, a recent review by Shi et al. evaluated the available studies of the last 10 years after a rigorous screening and included a total of 8 cohort studies with a predominantly prospective design with 578,809 participants meeting diagnostic criteria for insomnia and 7451 cancer events (Senaratna et al. 2017). The results showed a Table 3Association betweensleep disorders and subsequentGI cancer by cancer site andsleep disorder duration inpatients followed in generalpractices in Germany

| Variables/groups                     | Proportion (%) in<br>patients with GI<br>cancer | Proportion (%) in<br>patients without<br>cancer | OR (95% CI)      | <i>P</i> -value |
|--------------------------------------|---|---|------------------|-----------------|
| $Age \le 50 \ (n=2568 \ pairs)$      |   |   |                  |                 |
| Any diagnosis in the history         | 2.69  | 2.05  | 1.32 (0.92–1.90) | 0.133           |
| ≤1 year                              | 1.75  | 1.24  | 1.43 (0.90-2.25) | 0.128           |
| >1-3 years                           | 0.62  | 0.46  | 1.35 (0.64–2.86) | 0.431           |
| >3–5 years                           | 0.12  | 0.19  | 0.61 (0.15-2.55) | 0.497           |
| >5 years                             | 0.19  | 0.15  | 1.27 (0,34–4.73) | 0.724           |
| Age 51–60 (n=5564 pairs)             |   |   |                  |                 |
| Any diagnosis in the history         | 3.45  | 3.31  | 1.05 (0.85-1.28) | 0.676           |
| ≤1 year                              | 2.07  | 1.91  | 1.09 (0.83–1.42) | 0.551           |
| >1-3 years                           | 0.74  | 0.78  | 0.94 (0.62–1.45) | 0.794           |
| >3–5 years                           | 0.27  | 0.24  | 1.14 (0.54–2.40) | 0.725           |
| >5 years                             | 0.38  | 0.38  | 0.99 (0.54–1.82) | 0.975           |
| Age 61–70 (n=8119 pairs)             |   |   |                  |                 |
| Any diagnosis in the history         | 4.00  | 3.48  | 1.16 (0.98–1.36) | 0.079           |
| ≤1 year                              | 2.32  | 1.57  | 1.48 (1.18–1.86) | < 0.001         |
| >1-3 years                           | 0.79  | 0.95  | 0.84 (0.60–1.17) | 0.296           |
| >3–5 years                           | 027   | 0.37  | 0.73 (0.42–1.26) | 0.260           |
| >5 years                             | 0.63  | 0.59  | 1.08 (0.73–1.61) | 0.710           |
| Age 71-80 $(n=10,522 \text{ pairs})$ |   |   |                  |                 |
| Any diagnosis in the history         | 4.50  | 4.55  | 0.99 (0.87–1.13) | 0.876           |
| ≤1 year                              | 2.16  | 1.82  | 1.19 (0.98–1.44) | 0.085           |
| >1-3 years                           | 0.99  | 0.97  | 1.02 (0.78–1.35) | 0.873           |
| >3–5 years                           | 0.47  | 0.45  | 1.03 (0.69–1.53) | 0.903           |
| >5 years                             | 0.89  | 1.31  | 0.68 (0.52–0.89) | 0.004           |
| Age >80 (n= 6692 pairs)              |   |   |                  |                 |
| Any diagnosis in the history         | 5.86  | 6.00  | 0.97 (0.84–1.12) | 0.722           |
| ≤1 year                              | 2.51  | 2.42  | 1.04 (0.83–1.29) | 0.740           |
| >1-3 years                           | 1.30  | 1.30  | 1.00 (0.74–1.34) | 0.975           |
| >3–5 years                           | 0.63  | 0.64  | 0.98 (0.64–1.51) | 0.938           |
| >5 years                             | 1.42  | 1.65  | 0.86 (0.65–1.14) | 0.291           |
| Women (n=15,023 pairs)               |   |   |                  |                 |
| Any diagnosis in the history         | 4.89  | 4.99  | 0.98 (0.88-1.09) | 0.696           |
| ≤1 year                              | 2.30  | 2.05  | 1.12 (0.96–1.31) | 0.160           |
| >1-3 years                           | 1.10  | 1.09  | 1.01 (0.81–1.25) | 0.950           |
| >3–5 years                           | 0.47  | 0.50  | 0.94 (0.67–1.30) | 0.689           |
| >5 years                             | 1.03  | 1.35  | 0.76 (0.62–0.94) | 0.012           |
| Men (n=18,442 pairs)                 |   |   | ,                |                 |
| Any diagnosis in the history         | 3.89  | 3.53  | 1.11 (0.99–1.23) | 0.072           |
| ≤1 year                              | 2.16  | 1.68  | 1.29 (1.11–1.49) | < 0.001         |
| >1-3 years                           | 0.80  | 0.86  | 0.94 (0.75–1.17) | 0.559           |
| >3-5 years                           | 0.33  | 0.35  | 0.95 (0.67–1.35) | 0.782           |
| >5 years                             | 0.60  | 0.64  | 0.94 (0.72–1.22) | 0.636           |

OR odds ratio, CI confidence interval

modest 24% increased risk of cancer in people with insomnia compared with people without insomnia. Subgroup analyses revealed a significantly higher risk of cancer in studies conducted in women, but not in men (Senaratna et al. 2017). Similarly, for specific cancers, the pooled hazard ratio was significantly increased only for thyroid cancer, but not for other cancers. Despite the limited number of included studies and potential bias, the authors suggested, consistent with our conclusions, that insomnia may be an early warning sign of cancer development and may provide an opportunity for early detection and intervention (Senaratna et al. 2017).

Although the pathophysiological mechanisms underlying the association between sleep disorders and cancer are not fully understood, possible mechanisms have been discussed in the literature (Mogavero et al. 2021). It has been hypothesized that cancer may induce sleep disturbances, which in turn appear to be involved in the development and progression of cancer (Walker et al. 2019; Mogavero et al. 2021). In particular, one suggested role for the association of sleep disorders in cancer patients is the activation of an inflammatory response. Tumors produce large amounts of proinflammatory cytokines such as interleukin-1 beta (IL-1 beta), interleukin-6 (IL-6), and tumor necrosis factor-alpha (TNF-alpha), which in turn affect numerous neurotransmitters involved in sleep, including adenosine, prostaglandins, nitric oxide, and GABA (Walker et al. 2019; Mogavero et al. 2021). In this context, interestingly, sleep deprivation, short sleep duration, and other sleep disorders are associated with elevated levels of inflammatory parameters (Sillah et al. 2018). Specifically in OSA, sleep fragmentation and/or systemic intermittent hypoxia are thought to be mechanisms of carcinogenesis, inducing oxidative stress, systemic inflammation, endothelial dysfunction, and alterations in angiogenesis, sympathetic tone, immune function, and transcription factor expression that influence tumor progression (Hunyor et al. 2018; Martinez et al. 2019).

Melatonin is known to play a central role in circadian rhythm and sleep. The mechanisms of action of this hormone in cancer are not fully understood, but in addition to its role in regulating the sleep-wake cycle, melatonin is thought to be involved in inflammatory, immunological, metabolic, and neoplastic processes (Mogavero et al. 2021). Melatonin is mainly produced and secreted by the pineal gland in response to the light-dark cycle, but interestingly, it is also found in other organs such as the skin, bone marrow, and gastrointestinal tract (Acuña-Castroviejo et al. 2014). The antitumor effects of melatonin are attributed primarily to its involvement in DNA repair (Reiter 2004; Lin et al. 2019), enhancement of mitochondrial respiratory chain function, and inhibition of telomerase activity (Talib 2018). In addition, melatonin is thought to have antioxidant effects and act as a scavenger of reactive oxygen species (Talib 2018).

Sleep is hypothesized to influence the two primary effector systems, the hypothalamic-pituitary-adrenocortical axis and the sympathetic nervous system, which in turn control adaptive and innate immune responses (Irwin 2015). It has also been suggested that insufficient or disturbed sleep may decrease the release of immune-stimulating hormones such as growth hormones, prolactin, and dopamine, which could affect the natural state of the immune system (Lange et al. 2010; Lin et al. 2019). Regarding gender differences, a review by Irwin et al. described that women appear to be more susceptible to the effects of sleep disorders on inflammatory dynamics, whereas men appear to have a higher risk of cardiovascular disease associated with sleep disorders, especially short sleep duration, and, consistent with our findings, cancer (Irwin 2015). However, future studies, especially of a prospective nature, are needed to further investigate the question of a gender-specific relationship.

In fact, only a few studies have examined the association between different types of cancer and the prevalence of sleep disorders, mainly single case reports and subjective assessments, and conducted while cancer therapy was ongoing, leading to potential bias (Mogavero et al. 2021). In particular, insomnia seems to be a very common sleep disorder in patients with malignant tumors, reported in the literature with a highly variable prevalence ranging from 19 to 63%, often associated with anxiety and depression, pain and fatigue, and surprisingly more common in men than in women (Mogavero et al. 2021). Early and adequate treatment of sleep disorders is considered important to improve the therapeutic response of cancer patients in terms of survival, quality of life, and reduction of comorbidities (Mogavero et al. 2021).

Several potential limitations of this study should be noted, which are primarily due to the study design and are unavoidable. First, the data used in the DA database (IQVIA) are based on the ICD-10 coding system for specific sleep disorders, which means that misclassification and undercoding of certain diagnoses cannot be excluded. The ICD coding for sleep disorders (ICD-10: G47) basically includes the diagnoses of insomnia, hypersomnia, circadian rhythm disorders, sleep apnea, narcolepsy and cataplexy, parasomnias, sleep-related movement disorders, and unspecified sleep disorders. It should be emphasized that primary care physicians generally coded only unspecified sleep disorders (ICD-10: G47.9), whereas specific sleep disorders were very rarely documented. Therefore, analyses were performed only for sleep disorders as a whole, and no evaluations can be made for individual subgroups of sleep disorders. Similarly, data on sleep duration, sleep onset latency, or frequency and duration of nighttime awakenings were not available. In addition, the extent of the disorders cannot be graded in our study. However, the inclusion of these parameters would be useful for follow-up studies with a different study design. It should also be considered that the ICD-10 coding of sleep disorders is mostly based on a subjective description of the patient's complaints. Furthermore, the selection of our sample may have introduced some bias, as we only used data from general practitioners in the database. Other specialties, such as gastroenterology, which are thought to have a higher detection rate for gastrointestinal cancers,

were not included. Additionaly, acute or severe cancers are likely to be treated by specialists or in hospitals not included in the database. Another limitation of the study is that the DA database (IQVIA) does not include information on potential mediating effects, such as lifestyle factors like smoking, alcohol, diet, and physical activity, or information on socioeconomic status (e.g., education and income), which could potentially bias the association between sleep disorders and cancer risk. The DA database also does not provide larger panels of laboratory values or histologic parameters that would have allowed us to perform further analyses. Finally, no prognostic conclusions about cancer progression, metastasis, or survival can be drawn from our study. Thus, further studies are needed to investigate sleep problems as a modifiable risk factor associated with cancer progression. Finally, it should be stated that our study cannot establish causal relationships, only associations.

Nevertheless, it should be highlighted the large sample size as a strength of the study. Finally, to our knowledge, our study is the first population-based longitudinal study to investigate the association between sleep disorders and GI cancer in a German cohort.

## Conclusion

In conclusion, our data did not show a significant association between sleep disorders and GI cancers in the overall history. Interestingly, there was an association between sleep disorders and GI cancers when only the year of diagnosis was considered. These results suggest that sleep disorders can be interpreted as an early symptom and an indication of tumor disease. Individuals with sleep disorders may be encouraged to undergo tumor screening for early detection and treatment to positively influence disease progression and improve survival. However, large-scale prospective studies with more precise stratification of individual sleep disorders and detailed measurement of sleep duration are needed in the future to clarify the cause-effect relationship between sleep disorders and cancer.

Author contributions SHL, KK, and CR designed the study, KK performed statistical analyses and generated figures and tables, SHL, SK, and CR wrote the manuscript, AK, CL, TL, and CV provided intellectual input; all authors agreed to the final version of the manuscript.

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## Declarations

**Competing interests** The authors have no relevant financial or non-financial interests to disclose.

Ethics approval The database used includes only anonymized data in compliance with the regulations of the applicable data protection laws. German law allows the use of anonymous electronic medical records for research purposes under certain conditions. According to this legislation, it is not necessary to obtain informed consent from patients or approval from a medical ethics committee for this type of observational study that contains no directly identifiable data. Because patients were only queried as aggregates and no protected health information was available for queries, no Institutional Review Board approval was required for the use of this database or the completion of this study.

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