# Ambulatory Care Fragmentation and Incident Stroke 

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BACKGROUND: More fragmented ambulatory care (ie, care spread across many providers without a dominant provider) has been associated with excess emergency department and inpatient care. We sought to determine whether more fragmented ambulatory care is associated with an increase in the hazard of incident stroke, overall and stratified by health status and by race.

METHODS AND RESULTS: We conducted a secondary analysis of data from the REGARDS (Reasons for Geographic and Racial Differences in Stroke) study (2003-2016), including participants aged $\geq 65$ years who had linked Medicare fee-for-service claims and no history of stroke ( $\mathrm{N}=12510$ ). We measured fragmentation of care with the reversed Bice-Boxerman index. We used Poisson models to determine the association between fragmentation and adjudicated incident stroke. The average age of participants was 70.5 years; $53 \%$ were women, $32 \%$ were Black participants, and $16 \%$ were participants with fair or poor health. Overall, the adjusted rate of incident stroke was similar for high versus low fragmentation (8.2 versus 8.1 per 1000 person-years, respectively; $P=0.89$ ). Among participants with fair or poor self-rated health, having high versus low fragmentation was associated with a trend toward a higher adjusted rate of incident strokes (14.8 versus 10.4 per 1000 person-years, respectively; $P=0.067$ ). Among Black participants with fair or poor self-rated health, having high versus low fragmentation was associated with a higher adjusted rate of strokes (19.3 versus 10.3 per 1000 person-years, respectively; $P=0.02$ ).

CONCLUSIONS: Highly fragmented ambulatory care is independently associated with incident stroke among Black individuals with fair or poor health.

Key Words: ambulatory care $\square$ medicare $\square$ stroke

Patients who are at risk for incident stroke often see multiple ambulatory providers to manage their chronic conditions, such as diabetes mellitus, hypertension, heart disease, and atrial fibrillation. ${ }^{1-3}$ For example, patients at risk for incident stroke may have a primary care provider, cardiologist, endocrinologist, and other subspecialists who they see regularly. Patients with either diabetes mellitus or heart disease see more ambulatory providers than the median Medicare beneficiary nationally. ${ }^{2}$ Indeed, the more chronic conditions a patient has, the more providers he or she tends to see. ${ }^{2}$

Unfortunately, providers do not consistently communicate with each other about their common patients. ${ }^{4}$ In a national survey of nearly 5000 providers, $30 \%$ of primary care providers reported that they do not "always or most of the time" send information about a patient's medical history or the reason for requesting a consultation at the time of referral. ${ }^{4}$ Similarly, 20\% of specialists reported that they do not "always or most of the time" send the results of the consultation back to the referring provider. ${ }^{4}$ Thus, clinically relevant information is frequently missing at the point of care. ${ }^{5}$

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## CLINICAL PERSPECTIVE

## What Is New?

- Highly fragmented ambulatory care was associated with an increased hazard of incident stroke among Black individuals with fair or poor health, even after adjusting for clinically detailed potential confounders.


## What Are the Clinical Implications?

- Excess strokes attributable to highly fragmented ambulatory care may be preventable.

| Nonstandard Abbreviations and Acronyms |  |
| :--- | :--- |
| NCQA | National Committee for Quality <br> REGARDS |
| Assurance  <br> Reasons for Geographic and Racial  <br> Differences in Stroke  <br> self-rated general health (the first  <br> question in the Short-Form-12  <br> [SF-12] survey)  <br> WHO World Health Organization |  |

As a result, more fragmented ambulatory care (ie, care spread across many providers without a dominant provider) has been associated with more drug-drug interactions, ${ }^{6}$ testing, ${ }^{7}$ procedures, ${ }^{8}$ emergency department visits, ${ }^{9-11}$ and hospitalizations, ${ }^{9,12}$ compared with less fragmented care. However, the relationship between fragmented ambulatory care and health outcomes is not well understood. ${ }^{13}$ To our knowledge, no study has explored whether fragmented ambulatory care is associated with more incident strokes.

We previously found that the association between fragmented ambulatory care and subsequent healthcare utilization varies with patients' health status., ${ }^{7,14}$ We hypothesized that the association between fragmented ambulatory care and stroke might also vary with health status. Healthy people might be able to see multiple providers without experiencing harm from gaps in communication, whereas sicker individuals might be more vulnerable to harm.

We sought to determine the association between fragmented ambulatory care and incident stroke using data from the nationwide prospective REGARDS (Reasons for Geographic and Racial Differences in Stroke) cohort study. As the name of the investigation conveys, the purpose of the REGARDS study is to explain geographic and racial disparities in stroke mortality. ${ }^{15}$ Thus, we sought to
measure the association between fragmentation and incident stroke overall and to determine whether any association varied with health status, geographic region, or race.

## METHODS

## Study Design, Population, and Data Sources

We conducted an ancillary study to the nationwide, prospective REGARDS cohort study, using data from 2003 to 2016. ${ }^{15}$ The institutional review boards of the participating institutions approved the protocol. All participants provided written informed consent. Because of the restrictions of our data use agreement with the Centers for Medicare and Medicaid Services, the data used for this study cannot be made publicly available to other researchers for the purpose of replicating the results. However, access to REGARDS Medicare-linked data can be requested first through the REGARDS study (www.uab.edu/ soph/regardsstudy/researchers) and then through Medicare's Research Data Assistance Center (www. resdac.org).

Between 2003 and 2007, 30239 communitydwelling Black and White adults aged $\geq 45$ years were enrolled in the REGARDS study, with oversampling of Black adults and individuals living in the Southeastern United States. ${ }^{15}$ Baseline data collection involved computer-assisted telephone interviews and in-home visits with a physical examination, blood test, urine test, ECG, and medication inventory. Participants or their proxies are contacted by telephone every 6 months to detect study end points. The report of a potential event triggers adjudication, which involves expert review of medical records, death certificates, proxy interviews, autopsy reports, Social Security Death Index, and National Death Index. ${ }^{16}$

For this study, we used REGARDS study baseline data, REGARDS-adjudicated events, and REGARDS Medicare-linked claims. ${ }^{17}$

## Variables <br> Exposure

We used Medicare claims to identify ambulatory visits, which were defined using a modified National Committee for Quality Assurance (NCQA) definition ${ }^{18}$ that was restricted to Clinical Procedure Terminology codes for in-person, evaluation-and-management visits for adults in an office setting. ${ }^{7}$ The NCQA definition of ambulatory visits does not include emergency department visits. We calculated fragmentation scores for each participant using the previously validated Bice-Boxerman index, which ranges from 0 to 1 and
captures both the spread of ambulatory visits across providers and the relative share of visits by each provider (Data S1). 9 , 12,19,20 We reverse-coded scores, calculating 1 minus Bice-Boxerman index, so that higher scores reflected more fragmentation. ${ }^{7,11,14}$

## Outcome

Stroke events were adjudicated by the REGARDS study according to study protocols. ${ }^{21}$ Briefly, stroke events were defined in part using the World Health Organization (WHO) definition as "rapidly developing clinical signs of focal, at times global, disturbance of cerebral function, lasting more than 24 hours or leading to death with no apparent cause other than that of vascular origin." ${ }^{21,22}$ Events not meeting this definition but characterized by symptoms lasting $<24$ hours and with neuroimaging consistent with acute ischemia or hemorrhage were also classified as "clinical strokes."21 We used strokes based on WHO or clinical classification, adjudicated through December 31, 2016.

## Baseline Covariates

Demographics included self-reported age, sex, race, marital status, educational attainment, and annual household income; as well as geographic region and rural/urban setting using Rural/Urban Commuting Area codes ${ }^{23}$ based on baseline contact information. Medical conditions included hypertension (self-reported use of antihypertensive medication, systolic blood pressure $\geq 140 \mathrm{~mm} \mathrm{Hg}$ or diastolic blood pressure $\geq 90 \mathrm{~mm} \mathrm{Hg}$ from the in-home visit), dyslipidemia (use of lipid-lowering medication, total cholesterol $\geq 240 \mathrm{mg} / \mathrm{dL}$, low-density lipoprotein cholesterol $\geq 160 \mathrm{mg} / \mathrm{dL}$, or high-density lipoprotein cholesterol $\leq 40 \mathrm{mg} / \mathrm{dL}$ from the in-home visit), diabetes mellitus (fasting glucose $\geq 126 \mathrm{mg} / \mathrm{dL}$ or nonfasting glucose $\geq 200 \mathrm{mg} / \mathrm{dL}$ from the in-home visit, or selfreported use of oral glucose-lowering medication or insulin), history of coronary heart disease (selfreported history of myocardial infarction, coronary artery bypass graft, angioplasty, or coronary stents; or evidence of myocardial infarction on ECG at baseline), and atrial fibrillation (self-reported or present on the study ECG at baseline). Medication variables included a validated measure of self-reported medication adherence ${ }^{24}$; the total number of medications taken in the past 2 weeks by medication inventory; self-reported use of antihypertensive medication, insulin, and aspirin from the computer-assisted telephone interview; and statin and warfarin use from the medication inventory. Health behaviors were selfreported and included smoking status, alcohol consumption, ${ }^{25}$ and exercise frequency. Psychosocial
variables included whether a participant reported being the primary caretaker for another individual and whether he or she saw any close friends or relatives in the past month. Depressive symptoms were defined as Center for Epidemiologic Studies Depression Scale scores $\geq 4 .{ }^{26}$ Physiological variables included body mass index, heart rate, systolic blood pressure, left ventricular hypertrophy (on ECG); total, low-density lipoprotein cholesterol, and high-density lipoprotein cholesterol; glucose; estimated glomerular filtration rate;; ${ }^{27}$ urinary albumin to creatinine ratio; ${ }^{28}$ and high-sensitivity C-reactive protein. Self-rated health was assessed using self-rated general health (SF-1) from the Short-Form-12 (SF-12) survey. ${ }^{29-32}$ Hospitalization during the first year of observation was measured in Medicare claims.

## Statistical Analysis

We included participants aged $\geq 65$ years whose REGARDS study data were linked to Medicare claims at any time during the study period, allowing for staggered entry as participants aged into Medicare. We excluded participants who: (1) were eligible for Medicare on the basis of end-stage renal disease; (2) had fewer than 12 months of continuous Medicare fee-for-service coverage or had Medicare managed care; (3) were lost to follow-up before entry into this ancillary study; (4) had $\leq 3$ ambulatory visits in the first year of observation, because calculating fragmentation based on $\leq 3$ visits can lead to unreliable estimates ${ }^{12}$; (5) had a history of stroke; or (6) had a missing value for self-rated health.

We used descriptive statistics to characterize the sample at baseline and summarize ambulatory care utilization in the first year of observation. We used the 75th percentile fragmentation score rounded to the nearest 0.05 from the first year of observation (0.85) to dichotomize fragmentation into high ( $\geq 0.85$ ) and low (<0.85). This approach was informed by our previous studies ${ }^{7,14}$ and by the goal of identifying individuals who have such high scores that they might benefit from future interventions if high fragmentation is found to be associated with increased stroke risk. To compare differences between participants with high versus low fragmentation, we used Wilcoxon rank sum, Student $t$, and chi-square tests. We conducted sensitivity analyses changing the cutoff for high fragmentation to $\geq 0.80$ and, separately, $\geq 0.90$. We also conducted a sensitivity analysis treating fragmentation as a 4-level variable.

Observation for each participant began after the in-home visit with the first 12-month period for which there was continuous Medicare fee-for-service coverage, starting with calendar year 2004. Because fragmentation can change over time, we treated
fragmentation as a time-dependent exposure. ${ }^{12,14,33}$ Because the effects of fragmentation are hypothesized to occur relatively quickly (eg, drug-drug interactions in the context of multiple prescribers), ${ }^{14,33}$ we measured outcomes in the 3 months following each 12-month exposure period. Specifically, for each participant, we calculated a fragmentation score for the first 12 months of his or her observation period and then determined whether a stroke occurred in the 3 months immediately following the exposure period (months 13-15). If no event occurred, we moved the window of observation forward by 3 months, measuring fragmentation in months 4 to 15 and any event in months 16 to 18, and so on (Data S2). If the number of visits in any 12-month exposure period fell below 4, we carried forward the last fragmentation score based on $\geq 4$ visits. ${ }^{12}$ Observation continued until an outcome or censoring occurred; censoring occurred when a participant died of nonstroke causes, no longer had Medicare fee-for-service coverage, dropped out of the REGARDS study, or reached December 31, 2016.

We used Poisson regression to estimate absolute unadjusted and adjusted rates of strokes. We used Cox proportional hazards models to estimate unadjusted and adjusted hazards ratios. ${ }^{34-36}$ Using Schoenfeld residuals, we determined that the assumption of proportional hazards was met. We included in the fully adjusted multivariate models covariates that had a bivariate association with fragmentation ( $P<0.10$ ), were used in the sampling frame, or were retained because of high clinical importance (sex, coronary heart disease, atrial fibrillation, statin, aspirin, warfarin, heart rate). If 2 variables had bivariate $P$ values $<0.10$, but 1 was embedded in the definition of another (eg, use of antihypertensive medication was embedded in the definition of hypertension), we adjusted only for the broader variable to avoid collinearity. To maximize power, given that some subgroups had relatively few stroke events, we also generated a more parsimonious model, which included covariates that had multivariate $P$ values $<0.10$ in the fully adjusted model, were used in the sampling frame, or were retained because of high clinical importance (dyslipidemia, atrial fibrillation, aspirin).

We conducted all model analyses using stochastic regression imputation to handle missing covariates. Stochastic regression is a method that adds a residual error term to each imputed predicted value, thereby preserving uncertainty while reducing bias. ${ }^{37}$ The most frequently missing variable was income (missing for $12 \%$ of the sample).

We used self-rated health as a measure of health status. ${ }^{29,30}$ We used 3 groupings of self-rated health (excellent, very good or good, fair or poor) to detect
any differences among subgroups, while maximizing statistical power. We used Wald tests to determine interactions between fragmentation and self-rated health in the overall sample, by race, and by geographic region. The Wald tests were significant for self-rated health in the overall sample ( $P=0.04$ ) and by race ( $P=0.02$ ) but not by geographic region ( $P=0.94$ ). Thus, we conducted analyses overall and stratified by selfrated health and by race.

Analyses were conducted with SAS version 9.4 (SAS Institute Inc) and Stata version 14 (StataCorp). Multivariate $P$ values $<0.05$ were considered statistically significant.

## RESULTS

## Study Sample

Of the 30239 REGARDS study participants, 20403 had linked Medicare claims and were aged $\geq 65$ years during the study period. Of those, 12510 met our inclusion criteria (Figure S1).

## Sample Characteristics

Participants in our sample were aged 70.5 years on average; $53 \%$ were women and $68 \%$ were of White race (Table 1). More than half ( $60 \%$ ) lived in the Southeastern United States. Chronic conditions were common, with $62 \%$ having hypertension, $62 \%$ having dyslipidemia, 22\% having diabetes mellitus, $21 \%$ having coronary heart disease, and $9 \%$ having atrial fibrillation. In terms of self-rated health, 17\% reported excellent health, $67 \%$ reported very good or good health, and $13 \%$ reported fair or poor health. Additional characteristics are shown in Table 1. Appendices show participant characteristics stratified by self-rated health (Table S1), by race (Table S2), by self-rated health among Black participants (Table S3), and by self-rated health among White participants (Table S4).

## Ambulatory Utilization

Participants with high fragmentation (fragmentation score $\geq 0.85$ ) in the first year of observation were younger, more likely to be of White race, and less likely to have hypertension, dyslipidemia, and diabetes mellitus than those with low fragmentation (fragmentation score $<0.85$ ). Additional differences are shown in Table 1.

Participants with high fragmentation had on average 11.1 ambulatory visits with 6.4 providers in the first year of observation, with the most frequently seen provider accounting for 29.1\% of visits (Table 2). By contrast, participants with low fragmentation had on average 9.2 ambulatory visits with 3.6 providers,

Table 1. Baseline Characteristics, Overall and Stratified by Fragmentation Score in the First Year of Observation*

| Characteristic | Overall $(\mathrm{N}=12510)$ | Low Fragmentation During First Year of Observation ( $\mathrm{n}=9301$ ) | High <br> Fragmentation During First Year of Observation ( $\mathrm{n}=3209$ ) | $P$ Value |
| :---: | :---: | :---: | :---: | :---: |
| Demographic characteristics |  |  |  |  |
| Age, mean (SD), y | 70.5 (6.0) | 70.7 (6.0) | 70.2 (5.9) | <0.001 |
| Women, n (\%) | 6623 (52.9) | 4896 (52.6) | 1727 (53.8) | 0.25 |
| White race, n (\%) | 8513 (68.0) | 6108 (65.7) | 2405 (74.9) | <0.001 |
| Marital status, married, n (\%) | 7487 (59.8) | 5475 (58.9) | 2012 (62.7) | <0.001 |
| Education, less than high school diploma, n (\%) | 1545 (12.4) | 1237 (13.3) | 308 (9.6) | <0.001 |
| Annual household income, <\$35 000, n (\%) | 5422 (49.5) | 4199 (51.5) | 1223 (43.7) | <0.001 |
| Geographic region, n (\%) |  |  |  |  |
| Stroke belt ${ }^{\dagger}$ | 4551 (36.4) | 3334 (35.8) | 1217 (37.9) | 0.010 |
| Stroke buckle ${ }^{\ddagger}$ | 2899 (23.2) | 2214 (23.8) | 685 (21.3) |  |
| Neither stroke belt nor stroke buckle | 5060 (40.4) | 3753 (40.4) | 1307 (40.7) |  |
| Residence in urban area, n (\%) | 8467 (75.3) | 6307 (75.2) | 2160 (75.4) | 0.89 |
| Medical conditions ${ }^{\S}$ |  |  |  |  |
| Hypertension, n (\%) | 7663 (61.5) | 5857 (63.2) | 1806 (56.4) | <0.001 |
| Dyslipidemia, n (\%) | 7538 (62.3) | 5668 (62.9) | 1870 (60.6) | 0.019 |
| Diabetes mellitus, n (\%) | 2623 (21.7) | 2016 (22.4) | 607 (19.6) | <0.001 |
| History of coronary heart disease, n (\%) | 2521 (20.5) | 1852 (20.2) | 669 (21.2) | 0.25 |
| Atrial fibrillation, n (\%) | 1149 (9.4) | 840 (9.2) | 309 (9.8) | 0.33 |
| Medications |  |  |  |  |
| Medication adherence, n (\%) | 8302 (71.1) | 6205 (71.4) | 2097 (70.3) | 0.23 |
| No. of medications, median (25th, 75 th percentiles) | 6.0 (3.0, 9.0) | 6.0 (3.0, 8.0) | 6.0 (3.0, 9.0) | <0.001 |
| Antihypertensive medication, n (\%) | 6693 (55.4) | 5126 (57.1) | 1567 (50.3) | <0.001 |
| Insulin use, n (\%) | 657 (5.5) | 483 (5.4) | 174 (5.7) | 0.62 |
| Statin use, n (\%) | 4439 (35.5) | 3297 (35.4) | 1142 (35.6) | 0.89 |
| Aspirin use, n (\%) | 6084 (48.6) | 4509 (48.5) | 1575 (49.1) | 0.55 |
| Warfarin use, n (\%) | 522 (4.2) | 402 (4.3) | 120 (3.7) | 0.16 |
| Health behaviors |  |  |  |  |
| Current smoker, n (\%) | 1344 (10.8) | 1031 (11.1) | 313 (9.8) | 0.035 |
| Alcohol use, n (\%) |  |  |  |  |
| Heavy | 464 (3.8) | 339 (3.7) | 125 (4.0) | <0.001 |
| Moderate | 4049 (33.0) | 2924 (32.0) | 1125 (35.7) |  |
| None | 7775 (63.3) | 5874 (64.3) | 1901 (60.3) |  |
| Exercise frequency, 0 times per week, n (\%) | 4192 (34.1) | 3121 (34.1) | 1071 (33.9) | 0.84 |
| Psychosocial variables |  |  |  |  |
| Cares for a family member with a chronic illness or disability, n (\%) | 1493 (11.9) | 1102 (11.9) | 391 (12.2) | 0.62 |
| Lack of social support, n (\%) | 485 (4.0) | 343 (3.8) | 142 (4.5) | 0.073 |
| Depressive symptoms, n (\%) | 710 (5.7) | 528 (5.7) | 182 (5.7) | 0.99 |
| Physiological variables |  |  |  |  |
| Body mass index, mean (SD), $\mathrm{kg} / \mathrm{m}^{2}$ | 28.9 (5.8) | 29.0 (5.8) | 28.6 (5.7) | $<0.001$ |
| Heart rate, beats per min, mean (SD) | 66.6 (24.8) | 66.5 (22.4) | 66.7 (30.8) | 0.76 |
| Systolic blood pressure, mean (SD), mm Hg | 128.6 (16.3) | 129.1 (16.3) | 127.2 (16.0) | <0.001 |
| Left ventricular hypertrophy, n (\%) | 1247 (10.1) | 959 (10.4) | 288 (9.1) | 0.032 |
| Total cholesterol, mean (SD), mg/dL | 189.9 (39.4) | 190.5 (39.7) | 188.4 (38.3) | 0.011 |
| Low-density lipoprotein cholesterol, mean (SD), mg/dL | 111.5 (33.9) | 112.1 (34.2) | 109.6 (32.8) | <0.001 |

(Continued)

Table 1. Continued

| Characteristic | $\begin{aligned} & \text { Overall } \\ & (\mathrm{N}=12 \text { 510) } \end{aligned}$ | Low Fragmentation During First Year of Observation ( $\mathrm{n}=9301$ ) | High <br> Fragmentation During First Year of Observation ( $\mathrm{n}=3209$ ) | $P$ Value |
| :---: | :---: | :---: | :---: | :---: |
| High-density lipoprotein cholesterol, mean (SD), mg/dL | 51.6 (16.3) | 51.4 (16.1) | 52.3 (16.6) | 0.009 |
| Glucose, mean (SD), mg/dL | 103.8 (33.7) | 104.3 (34.2) | 102.4 (32.2) | 0.008 |
| Estimated glomerular filtration rate, mean (SD), mL/min per $1.73 \mathrm{~m}^{2}$ | 80.6 (18.9) | 80.5 (19.1) | 80.8 (18.3) | 0.44 |
| Urinary albumin to creatinine ratio, median (25th, 75th percentiles), mg/g | 2.1 (0.9, 4.8) | 2.2 (1.0, 4.8) | 2.0 (0.9, 4.5) | 0.010 |
| C-reactive protein, median (25th, 75th percentiles), mg/L | 7.8 (4.9, 16.7) | 7.9 (4.9, 17.0) | 7.6 (4.8, 15.7) | 0.010 |
| Self-rated health |  |  |  |  |
| Self-rated general health (SF-1), n (\%) |  |  |  |  |
| Excellent | 2145 (17.1) | 1544 (16.6) | 601 (18.7) | 0.049 |
| Very good | 3940 (31.5) | 2926 (31.5) | 1014 (31.6) |  |
| Good | 4423 (35.4) | 3336 (35.9) | 1087 (33.9) |  |
| Fair | 1669 (13.3) | 1251 (13.5) | 418 (13.0) |  |
| Poor | 333 (2.7) | 244 (2.6) | 89 (2.8) |  |
| Mental component summary score, mean (SD) | 55.1 (7.6) | 55.1 (7.6) | 54.9 (7.6) | 0.24 |
| Physical component summary score, mean (SD) | 46.6 (10.3) | 46.6 (10.2) | 46.5 (10.5) | 0.51 |
| Healthcare utilization |  |  |  |  |
| Hospitalization in the first y of observation, n (\%) | 1245 (10.0) | 880 (9.5) | 365 (11.4) | 0.002 |

SF indicates short-form survey. ${ }^{19}$ Percentages may not sum to 100 because of rounding.
*Missing data: education ( $n=4$ ), income ( $n=1554$ ), residence in urban area ( $n=1261$ ), medication count ( $n=12$ ), smoking ( $n=48$ ), exercise ( $n=207$ ), caregiver status ( $n=12$ ), lack of social support ( $n=308$ ), depressive symptoms ( $n=72$ ), body mass index $(n=65)$, heart rate ( $n=111$ ), systolic blood pressure ( $n=34$ ), total cholesterol ( $n=454$ ), low-density lipoprotein cholesterol ( $n=680$ ), high-density lipoprotein cholesterol ( $n=529$ ), glucose ( $n=456$ ), estimated glomerular filtration rate ( $n=454$ ), urinary albumin to creatinine ratio ( $n=505$ ), C-reactive protein ( $n=724$ ), mental component summary score ( $n=591$ ), and physical component summary score ( $n=591$ ).
${ }^{\dagger}$ Stroke Belt consists of North Carolina, South Carolina, Georgia, Tennessee, Mississippi, Alabama, Louisiana, and Arkansas (except for 153 coastal counties that constitute the Stroke Buckle). ${ }^{15}$
${ }^{\ddagger}$ Stroke Buckle consists of 153 coastal counties in North Carolina, South Carolina, and Georgia. ${ }^{15,21}$
\$See Methods section for detailed definitions of variables.
with the most frequently seen provider accounting for $57.5 \%$ of visits ( $P<0.001$ for each pairwise comparison). The average fragmentation score for participants with high fragmentation was 0.91, compared with 0.63 for participants with low fragmentation ( $P<0.001$ ).

Participants with fair or poor self-rated health had more visits and more providers, compared with participants with very good/good self-rated health and compared with those with excellent self-rated health ( $P<0.001$ for each comparison). However, highly fragmented care (score $\geq 0.85$ ) was observed in all 3 subgroups by self-rated health, affecting 25\% to 28\% participants in each subgroup.

Black participants had on average more visits than White participants (10.0 visits versus 9.5 visits, $P<0.001$ ) but fewer providers (4.1 providers versus 4.4 providers, $P<0.001$ ). Highly fragmented care was less common among Black participants than White participants, affecting 20\% of Black participants and $28 \%$ of White participants.

## Rates of Incident Strokes

Participants were observed for a median of 5.6 years (range, 1 day to 11.8 years). We observed 611 incident strokes over 75106 person-years, for an adjusted rate of 8.1 strokes per 1000 person-years (Table 3). The majority of strokes observed were ischemic (85\%), rather than hemorrhagic (11\%); stroke type was unclassifiable for 4\%. Overall, the adjusted rates for incident strokes was similar for high versus low fragmentation (8.2 versus 8.1 per 1000 person-years, $P=0.89$ ). However, there were significant differences among subgroups. The subgroup with the highest adjusted rate of incident strokes was Black participants with high fragmentation and fair or poor self-rated health who had an adjusted rate of 19.3 strokes per 1000 person-years $(95 \% \mathrm{Cl}, 11.5$ per 1000 person-years; 27.2 per 1000 person-years) (Figure 1).

## Hazard of Incident Stroke

Overall, there was no association between high fragmentation and the hazard of incident stroke (adjusted
Table 2. Ambulatory Utilization During the First Year of Observation, Overall and Stratified By Fragmentation Score, Self-Rated Health, and Race*

|  | Overall |  |  | Self-Rated Health |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Excellent |  |  | Very Good or Good |  |  | Fair or Poor |  |  |
|  | Total | Low rBBI $(<0.85)$ | High rBBI $(\geq 0.85)$ | Total | Low rBBI $(<0.85)$ | High rBBI $(\geq 0.85)$ | Total | Low rBBI $(<0.85)$ | High rBBI ( $\geq 0.85$ ) | Total | Low rBBI (<0.85) | High rBBI $(\geq 0.85)$ |
| All participants |  |  |  |  |  |  |  |  |  |  |  |  |
| No. (\%) | 12510 (100) | 9301 (74.4) | 3209 (25.7) | 2145 (100) | 1544 (72.0) | 601 (28.0) | 8363 (100) | 6262 (74.9) | 2101 (25.1) | 2002 (100) | 1495 (74.9) | 507 (25.1) |
| Visits, mean (SD) | 9.68 (6.57) | 9.21 (6.28) | 11.05 (7.18) | 7.59 (4.56) | 7.36 (4.49) | 8.18 (4.70) | 9.56 (6.36) | 9.06 (6.00) | 11.06 (7.11) | 12.45 (8.13) | 11.78 (7.94) | 14.41 (8.36) |
| Providers, mean (SD) | 4.30 (2.30) | 3.58 (1.73) | 6.40 (2.46) | 3.88 (1.85) | 3.23 (1.39) | 5.56 (1.83) | 4.25 (2.26) | 3.54 (1.68) | 6.38 (2.44) | 4.94 (2.73) | 4.07 (2.11) | 7.48 (2.76) |
| Percent of visits with the most frequently seen provider, mean (SD) | 50.18 (19.96) | 57.45 (17.74) | 29.13 (6.74) | 49.96 (19.50) | 57.88 (16.92) | 29.62 (6.85) | 50.38 (19.89) | 57.45 (17.73) | 29.31 (6.73) | $\begin{aligned} & 49.59 \\ & (20.72) \end{aligned}$ | 56.98 (18.57) | 27.79 (6.46) |
| Fragmentation score, rBBI, mean (SD) | 0.70 (0.22) | 0.63 (0.22) | 0.91 (0.04) | 0.72 (0.22) | 0.64 (0.21) | 0.92 (0.05) | 0.70 (0.22) | 0.63 (0.22) | 0.91 (0.04) | 0.69 (0.23) | 0.62 (0.22) | 0.90 (0.04) |
| White participants |  |  |  |  |  |  |  |  |  |  |  |  |
| No. (\%) | 8513 (100) | 6108 (71.6) | 2405 (28.3) | 1760 (100) | 1230 (69.9) | 530 (30.1) | 5739 (100) | 4147 (72.3) | 1592 (27.7) | 1014 (100) | 731 (72.3) | 283 (27.7) |
| Visits, mean (SD) | 9.52 (6.57) | 9.07 (6.37) | 10.66 (6.93) | 7.50 (4.50) | 7.27 (4.44) | 8.05 (4.62) | 9.53 (6.44) | 9.03 (6.18) | 10.82 (6.90) | 13.00 (8.60) | 12.36 (8.57) | 14.65 (8.47) |
| Providers, mean (SD) | 4.38 (2.27) | 3.63 (1.72) | 6.28 (2.38) | 3.95 (1.82) | 3.28 (1.36) | 5.51 (1.79) | 4.37 (2.26) | 3.62 (1.70) | 6.32 (2.38) | 5.19 (2.73) | 4.28 (2.11) | 7.55 (2.74) |
| Percent of visits with the most frequently seen provider, mean, (SD) | 48.87 (19.25) | $\begin{gathered} 56.59 \\ (16.96) \end{gathered}$ | 29.28 (6.75) | 48.80 (18.70) | 57.03 (16.00) | 29.72 (6.76) | 49.09 (19.35) | 56.64 (17.18) | 29.42 (6.75) | 47.81 (19.62) | 55.61 (17.29) | 27.65 (6.53) |
| Fragmentation score, rBBI, mean (SD) | 0.72 (0.21) | 0.65 (0.20) | 0.91 (0.04) | 0.74 (0.20) | 0.66 (0.19) | 0.92 (0.05) | 0.72 (0.21) | 0.65 (0.21) | 0.91 (0.04) | 0.72 (0.20) | 0.64 (0.20) | 0.90 (0.04) |
| Black participants |  |  |  |  |  |  |  |  |  |  |  |  |
| No. (\%) | 3997 (100) | 3193 (79.9) | 804 (20.1) | 385 (100) | 314 (81.6) | 71 (18.4) | 2624 (100) | 2115 (80.6) | 509 (19.4) | 988 (100) | 764 (80.6) | 224 (19.4) |
| Visits, mean (SD) | 10.03 (6.55) | 9.48 (6.08) | 12.22 (7.77) | 7.97 (4.82) | 7.70 (4.71) | 9.15 (5.15) | 9.63 (6.19) | 9.11 (5.65) | 11.81 (7.69) | 11.88 (7.57) | 11.23 (7.25) | 14.10 (8.21) |
| Providers, mean (SD) | 4.13 (2.36) | 3.48 (1.75) | 6.75 (2.66) | 3.56 (1.94) | 3.02 (1.45) | 5.93 (2.07) | 4.01 (2.24) | 3.40 (1.63) | 6.57 (2.61) | 4.67 (2.71) | 3.88 (2.10) | 7.39 (2.79) |
| Percent of visits with the most frequently seen provider, mean, (SD) | 52.97 (21.14) | $\begin{gathered} 59.08 \\ (19.03) \end{gathered}$ | 28.69 (6.69) | 55.25 (22.08) | 61.21 (19.82) | 28.92 (7.51) | 53.21 (20.76) | 59.05 (18.67) | 28.98 (6.70) | $\begin{gathered} 51.42 \\ (21.66) \end{gathered}$ | 58.29 (19.65) | 27.97 (6.38) |
| Fragmentation score, rBBI, mean (SD) | 0.66 (0.24) | 0.60 (0.24) | 0.90 (0.04) | 0.64 (0.26) | 0.58 (0.25) | 0.92 (0.05) | 0.66 (0.24) | 0.61 (0.23) | 0.90 (0.04) | 0.67 (0.25) | 0.60 (0.24) | 0.90 (0.04) |

The reversed Bice-Boxerman index (rBBI) is equivalent to 1-Bice-Boxerman index, such that higher scores indicate more fragmentation. ${ }^{19}$
All comparisons of low vs high rBBI, overall and within self-rated health subgroup, are statistically significant ( $P<0.001$ ).
*To be included in the analysis, a participant had to have $\geq 4$ ambulatory visits in his or her first year of observation.

Table 3. Incident Rates of Stroke, Stratified by Self-Rated General Health, Fragmentation Status, and Race*

|  | Overall | Self-Rated General Health |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Excellent | Very Good or Good | Fair or Poor |
| Sample sizes and observation time |  |  |  |  |
| No. (\%) | 12510 | 2145 (17) | 8363 (67) | 2002 (16) |
| Range of observation time | 1 d to 11.8 y | 11 d to 11.8 y | 2 d to 11.8 y | 1 d to 11.8 y |
| Median observation time (IQR), y | 5.6 (2.9-9.1) | 6.4 (3.2-9.9) | 5.8 (3.0-9.4) | 4.1 (2.0-7.5) |
| Event counts and rates |  |  |  |  |
| All participants |  |  |  |  |
| Total |  |  |  |  |
| Observed no. of events (\% participants) | 611 (4.9) | 96 (4.5) | 399 (4.8) | 116 (5.8) |
| Observed no. of person-years | 75106 | 13836 | 51367 | 9903 |
| Unadjusted rate per 1000 person-years | 8.1 | 6.9 | 7.8 | 11.7 |
| Adjusted rate per 1000 person-years | 8.1 | 6.9 | 7.8 | 11.7 |
| Low fragmentation |  |  |  |  |
| Observed no. of events (\% participants) | 423 (5.2) | 73 (5.4) | 277 (5.1) | 73 (5.7) |
| Observed no. of person-years | 51053 | 9113 | 35127 | 6813 |
| Unadjusted rate per 1000 person-years | 8.3 | 8.0 | 7.9 | 10.7 |
| Adjusted rate per 1000 person-years | 8.1 | 7.8 | 7.9 | 10.4 |
| High fragmentation |  |  |  |  |
| Observed no. of events (\% participants) | 188 (4.2) | 23 (2.9) | 122 (4.2) | 43 (5.9) |
| Observed no. of person-years | 24052 | 4723 | 16239 | 3090 |
| Unadjusted rate per 1000 person-years | 7.8 | 4.9 | 7.5 | 13.9 |
| Adjusted rate per 1000 person-years | 8.2 | 5.1 | 7.5 | 14.8 |
| White participants |  |  |  |  |
| Total |  |  |  |  |
| Observed no. of events (\% participants) | 431 (5.1) | 84 (4.8) | 289 (5.0) | 58 (5.7) |
| Observed no. of person-years | 54516 | 11791 | 37303 | 5422 |
| Unadjusted rate per 1000 person-years | 7.9 | 7.1 | 7.7 | 10.7 |
| Adjusted rate per 1000 person-years | 7.9 | 7.1 | 7.7 | 10.8 |
| Low fragmentation |  |  |  |  |
| Observed no. of events (\% participantss) | 289 (5.5) | 63 (5.9) | 187 (5.3) | 39 (6.3) |
| Observed no. of person-years | 35641 | 7552 | 24511 | 3577 |
| Unadjusted rate per 1000 person-years | 8.1 | 8.3 | 7.6 | 10.9 |
| Adjusted rate per 1000 person-years | 7.9 | 8.0 | 7.5 | 10.6 |
| High fragmentation |  |  |  |  |
| Observed no. of events (\% participants) | 142 (4.3) | 21 (3.1) | 102 (4.6) | 19 (4.8) |
| Observed no. of person-years | 18875 | 4239 | 12792 | 1844 |
| Unadjusted rate per 1000 person-years | 7.5 | 5.0 | 8.0 | 10.3 |
| Adjusted rate per 1000 person-years | 7.9 | 5.3 | 8.3 | 11.1 |
| Black participants |  |  |  |  |
| Total |  |  |  |  |
| Observed no. of events (\% participants) | 180 (4.5) | 12 (3.1) | 110 (4.2) | 58 (5.9) |
| Observed no. of person-years | 20590 | 2045 | 14063 | 4481 |
| Unadjusted rate per 1000 person-years | 8.7 | 5.9 | 7.8 | 12.9 |
| Adjusted rate per 1000 person-years | 8.8 | 6.0 | 7.9 | 12.8 |
| Low fragmentation |  |  |  |  |
| Observed no. of events (\% participants) | 134 (4.7) | 10 (3.7) | 90 (4.6) | 34 (5.2) |
| Observed no. of person-years | 15413 | 1562 | 10616 | 3235 |

(Continued)

Table 3. Continued

|  |  | Self-Rated General Health |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Overall | Excellent | Very Good or Good | Fair or Poor |
| Unadjusted rate per 1000 person-years | 8.7 | 6.4 | 8.5 | 10.5 |
| Adjusted rate per 1000 person-years | 8.7 | 7.2 | 8.5 | 10.3 |
| High fragmentation |  |  |  |  |
| Observed no. of events (\% participants) | 46 (4.1) | 2 (1.8) | 20 (2.9) | 24 (7.2) |
| Observed no. of person-years | 5177 | 484 | 3447 | 1246 |
| Unadjusted rate per 1000 person-years | 8.9 | 4.1 | 5.8 | 19.3 |
| Adjusted rate per 1000 person-years | 8.9 | 3.4 | 6.0 | 19.3 |

*Adjusted rates were derived from Poisson models that adjusted for demographics (sex, age, race, income, region), medical conditions (hyperlipidemia, myocardial infarction, atrial fibrillation), medications (aspirin use, warfarin use), health behaviors (smoking), psychosocial variables (social support), physiological variables (body mass index, systolic blood pressure, albumin to creatinine ratio, C-reactive protein), self-rated health, and hospitalization. Fragmentation status is a time-varying exposure based on the reversed Bice-Boxerman index (low $<0.85$, high $\geq 0.85$ ). See Methods section for more details.
hazard ratio [HR], 1.01; 95\% CI, 0.85, 1.21) (Table 4). There was a trend suggesting that high fragmentation was associated with a higher hazard of stroke, compared with low fragmentation, among participants with fair or poor self-rated health (adjusted HR, 1.39; $95 \% \mathrm{Cl}, 0.95-2.04$ ). There was no association among participants with very good or good selfrated health (adjusted HR, 1.01; 95\% CI, 0.82-1.26)
or excellent self-rated health (adjusted HR, 0.66; 95\% CI, 0.41-1.07).

Among Black participants with fair or poor self-rated health, high fragmentation was associated with a significantly higher hazard of stroke, compared with low fragmentation (adjusted HR, 1.83; $95 \% \mathrm{Cl}, 1.06-3.14$ ). Among Black participants, there was no association between fragmentation and stroke among participants


Figure. Adjusted rates of incident strokes per 1000 person-years with $95 \%$ Cls, stratified by healthcare fragmentation, self-rated health, and race.*
*Fragmentation is defined using the reversed Bice-Boxerman index (high $\geq 0.85$, low $<0.85$ ) and modeled as a time-varying exposure. Adjusted rates were derived from multivariable Poisson models; see Table 3 for more details. The lower bound of the Cl for low fragmentation-Black race-excellent self-rated health was equal to -1.6

Table 4. Hazard of Incident Stroke, Comparing High With Low Fragmentation, Overall and Stratified by Self-Rated General Health and Race ( $\mathrm{N}=12$ 510)*

|  | Overall | Self-Rated General Health |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Excellent | Very Good or Good | Fair or Poor |
| HRs (95\% Cls) |  |  |  |  |
| All participants |  |  |  |  |
| Unadjusted | 0.92 (0.78-1.10) | 0.60 (0.37-0.96) | 0.93 (0.75-1.15) | 1.25 (0.86-1.82) |
| Model $1^{+}$ | 1.01 (0.84-1.21) | 0.67 (0.42-1.08) | 1.01 (0.82-1.26) | 1.41 (0.96-2.07) |
| Model $2^{\ddagger}$ | 1.01 (0.85-1.21) | 0.66 (0.41-1.07) | 1.01 (0.82-1.26) | 1.39 (0.95-2.04) |
| White participants |  |  |  |  |
| Unadjusted | 0.90 (0.73-1.10) | 0.58 (0.35-0.95) | 1.01 (0.80-1.29) | 0.89 (0.51-1.54) |
| Model $3^{\text {§ }}$ | 0.98 (0.80-1.21) | 0.68 (0.41-1.12) | 1.10 (0.86-1.41) | 0.99 (0.57-1.72) |
| Model $4^{\ddagger}$ | 0.99 (0.81-1.21) | 0.67 (0.40-1.10) | 1.11 (0.87-1.42) | 1.01 (0.58-1.76) |
| Black participants |  |  |  |  |
| Unadjusted | 1.04 (0.75-1.45) | ... | 0.69 (0.43-1.13) | 1.82 (1.08-3.07) |
| Model $3^{\text {§ }}$ | 1.06 (0.75-1.48) | $\ldots$ | 0.72 (0.44-1.17) | 1.87 (1.09-3.20) |
| Model $4^{\ddagger}$ | 1.05 (0.74-1.46) | $\ldots$ | 0.73 (0.45-1.19) | 1.83 (1.06-3.14) |

*HR indicates hazard ratio. Results were derived from Cox proportional hazards models that treat fragmentation as a time-varying exposure (considering 12 months of the exposure at a time and outcomes that occur in the 3 months immediately following the exposure period), with a total of up to 11.8 years of follow-up. High fragmentation is defined as a score of $\geq 0.85$ on a reversed Bice-Boxerman index. We were not able to calculate results for Black participants with excellent self-rated health because they had too few events. Wald tests for interaction were significant for self-rated health in the overall sample ( $P=0.04$ ) and by race ( $P=0.02$ ).
${ }^{\dagger}$ Model 1 is a fully adjusted model, which adjusts for 26 baseline covariates that were associated with fragmentation in the first year of observation in bivariate models ( $P<0.10$ ), were used in the sampling frame, or were retained because of high clinical importance: demographics (age, sex, race, marital status, education, income, region), medical conditions (dyslipidemia, diabetes mellitus, coronary artery disease, atrial fibrillation), medications (medication count, statin use, aspirin use, warfarin use), heath behaviors (smoking, alcohol use), psychosocial variables (social support), physiological variables (body mass index, heart rate, systolic blood pressure, left ventricular hypertrophy, albumin to creatinine ratio, and C-reactive protein), self-rated health, and hospitalization.
$\ddagger$ Model 2 is a more parsimonious model that includes 18 covariates that were significant ( $P<0.10$ ) in model 1, were used in the sampling frame, or were retained because of high clinical importance: demographics (sex, age, race, income, region), medical conditions (dyslipidemia, myocardial infarction, atrial fibrillation), medications (aspirin use, warfarin use), health behaviors (smoking), psychosocial variables (social support), physiological variables (body mass index, systolic blood pressure, albumin to creatinine ratio, C-reactive protein), self-rated health, and hospitalization.
§Model 3 is the same as model 1, except stratified by race, instead of including adjustment for race.
$\ddagger$ Model 4 is the same as model 2 , except stratified by race, instead of including adjustment for race.
with very good or good self-rated health (adjusted HR, 0.73 ; 95\% Cl, 0.45-1.19). There were too few stroke events among Black participants with excellent selfrated health to calculate a model for that subgroup. There was no association between fragmentation and stroke among White participants, overall or by selfrated health subgroup.

## Sensitivity Analyses

When we changed the cutoff for "high" fragmentation to $\geq 0.80$ instead of $\geq 0.85,46.7 \%$ of the cohort was classified as having high fragmentation (rather than $25.7 \%$ ). With this alternate cutoff of $\geq 0.80$, we did not find any association between high fragmentation and incident stroke (Table S5). When we changed the cutoff for "high" fragmentation to $\geq 0.90,13.7 \%$ of the cohort was classified as having high fragmentation. With this alternate cutoff of $\geq 0.90$, we found that the association between high fragmentation and incident stroke persisted among the subset of Black participants with fair or poor self-rated health (adjusted HR, 2.13; 95\% $\mathrm{Cl}, 1.15-3.95)$. When we modeled fragmentation as
a 4-level variable, Cls widened, but the association between fragmentation and incident stroke persisted among Black participants with fair of poor self-rated health and fragmentation scores $\geq 0.90$, compared with the reference category (adjusted HR, 2.18; 95\% CI, 1.11-4.28) (Table S6).

## DISCUSSION

In this nationwide, 11-year cohort study of 12510 Medicare beneficiaries, we found that the association between highly fragmented ambulatory care and incident stroke varied with self-rated health and race. Overall, the adjusted rate of incident stroke was similar between participants with high versus low fragmentation (8.2 versus 8.1 per 1000 person-years, $P=0.89$ ). Among participants with fair or poor self-rated health at baseline, having high versus low fragmentation was associated with a trend toward a higher adjusted rate of incident strokes (14.8 versus 10.4 per 1000 person-years, $P=0.067$ ). Among Black participants with fair or poor self-rated health, having high versus
low fragmentation was associated with a significantly higher adjusted rate of strokes ( 19.3 versus 10.3 per 1000 person-years, respectively; $P=0.02$ ), which was equivalent to a significantly increased adjusted hazard of stroke (HR, 1.83; $95 \% \mathrm{Cl}$, 1.06-3.14).

This study creates a novel bridge across several existing bodies of work. There is one body of work documenting physiologic risk factors for stroke, such as age, hypertension, diabetes mellitus, atrial fibrillation, coronary heart disease, and left ventricular hypertrophy, which our analysis includes. ${ }^{3}$ There is a second body of work linking self-rated health with stroke incidence ${ }^{38}$ and all-cause mortality. 29,30 There is a third body of work documenting racial disparities in stroke incidence, with Black individuals experiencing a disproportionate number of strokes. ${ }^{21}$ There is a fourth body of work documenting racial disparities in the delivery of health care. ${ }^{39}$ This study is a novel addition, contributing new data on how fragmented ambulatory care relates to incident strokes, by self-rated health and by race, after accounting for physiologic risk factors.

The exact mechanism by which fragmentation increases the hazard of stroke among Black individuals with fair or poor self-rated health cannot be fully elucidated by the data in this study. Previous work has shown that fragmented care can lead to gaps in communication across providers caring for the same patient. ${ }^{40}$ Such gaps in communication have been shown to be associated with an increased risk of drug-drug interactions, ${ }^{6}$ including drug-drug interactions that could inadvertently worsen blood pressure control or worsen anticoagulation control among individuals at risk for stroke. Additional factors to be considered include that Black adults may be cared for by less well-trained physicians than White adults, ${ }^{41}$ Black adults may have less access to specialist providers than White adults (even with similar Medicare insurance), ${ }^{42,43}$ Black adults have reason to distrust the healthcare system, ${ }^{44}$ and Black adults may face explicit and/or implicit bias in the care they receive. ${ }^{55}$ In this context, Black adults with fair or poor health may be particularly vulnerable to any adverse consequences of fragmented care.

If future interventions were to try to reduce fragmentation, by how much would care patterns need to change in order to reduce fragmentation scores? To illustrate this, we provide a nomogram (Figure S2), which shows all possible patterns of providers for individuals with 12 visits (the average number of visits for Black participants with fair or poor health). ${ }^{33}$ The nomogram shows all possible patterns of providers (1 to 12 providers) and all possible distributions of visits across those providers. The nomogram illustrates how fragmentation scores decrease as the proportion of visits with the most frequently seen provider increases. Although the nomogram illustrates only patterns of care for individuals with 12 visits, this negative correlation holds for any visit count. In our study, participants with high
fragmentation often had average fragmentation scores of $\approx 0.90$, whereas participants with low fragmentation often had fragmentation scores of $\approx 0.60$. Using the nomogram, we can see that, for individuals with 12 visits, changing a fragmentation score from 0.90 to 0.60 would involve increasing the proportion of visits with the most frequently seen provider from $\approx 25 \%$ to $58 \%$. The nomogram also shows how fragmentation scores vary with the distribution of visits to the other providers, not just the most frequently seen provider.

This study has several strengths, including the national sampling frame and large sample size, with oversampling of Black individuals to ensure robust subgroup analysis, the long follow-up period, the uniform insurance coverage through Medicare fee-for-service, the previously validated measure of fragmentation, the treatment of fragmentation as a time-varying exposure, adjudicated stroke events, and clinically detailed potential confounders that were collected through standardized protocols.

This study has several limitations. First, this is an observational study, so we cannot infer causation or rule out unmeasured confounding. There were differences in the characteristics for participants with high versus low fragmentation; the reasons for these differences are not known. However, the direction of many of these differences (such as lower rates of cardiovascular risk factors in the high fragmentation group) would bias our study toward the null. Second, we do not have data on communication among providers, so fragmentation cannot be interpreted as the definite absence of communication. Third, we cannot determine the clinical appropriateness of the ambulatory visit patterns that we observed. Fourth, we measured self-rated health and clinical covariates at baseline, and it is possible that they change over time. However, self-rated health at baseline was a significant predictor of incident stroke.

## CONCLUSIONS

High fragmentation of ambulatory care was associated with an increase in the hazard of incident stroke for Black individuals with fair or poor self-rated health, compared with low fragmentation (adjusted HR, 1.83; $95 \% \mathrm{Cl}, 1.06-3.14)$. Future research is needed to better understand how fragmentation affects Black adults with fair or poor health, in order to inform the design of future interventions to prevent unnecessary fragmentation and prevent excess strokes.

## ARTICLE INFORMATION

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## Supplementary Material

Data S1-S2
Supplemental Methods
Figures S1-S2
Tables S1-S6

## REFERENCES

1. Chambless LE, Heiss G, Shahar E, Earp MJ, Toole J. Prediction of ischemic stroke risk in the Atherosclerosis Risk in Communities Study. Am J Epidemiol. 2004;160:259-269. DOI: 10.1093/aje/kwh189.
2. Pham HH, Schrag D, O'Malley AS, Wu B, Bach PB. Care patterns in Medicare and their implications for pay for performance. N Engl J Med. 2007;356:1130-1139. DOI: 10.1056/NEJMsa063979.
3. Wolf PA, D’Agostino RB, Belanger AJ, Kannel WB. Probability of stroke: a risk profile from the Framingham Study. Stroke. 1991;22:312-318. DOI: 10.1161/01.STR.22.3.312.
4. O'Malley AS, Reschovsky JD. Referral and consultation communication between primary care and specialist physicians: finding common ground. Arch Intern Med. 2011;171:56-65. DOI: 10.1001/archintern med.2010.480.
5. Smith PC, Araya-Guerra R, Bublitz C, Parnes B, Dickinson LM, Van Vorst R, Westfall JM, Pace WD. Missing clinical information during primary care visits. JAMA. 2005;293:565-571. DOI: 10.1001/jama.293.5.565.
6. Guo JY, Chou YJ, Pu C. Effect of continuity of care on drug-drug interactions. Med Care. 2017;55:744-751. DOI: 10.1097/MLR.0000000000 000758.
7. Kern LM, Seirup JK, Casalino LP, Safford MM. Healthcare fragmentation and the frequency of radiology and other diagnostic tests: a crosssectional study. J Gen Intern Med. 2017;32:175-181. DOI: 10.1007/ s11606-016-3883-z.
8. Romano MJ, Segal JB, Pollack CE. The association between continuity of care and the overuse of medical procedures. JAMA Intern Med. 2015;175:1148-1154. DOI: 10.1001/jamainternmed.2015.1340.
9. Hussey PS, Schneider EC, Rudin RS, Fox DS, Lai J, Pollack CE. Continuity and the costs of care for chronic disease. JAMA Intern Med. 2014;174:742-748. DOI: 10.1001/jamainternmed.2014.245.
10. Katz DA, McCoy KD, Vaughan-Sarrazin MS. Does greater continuity of Veterans Administration primary care reduce emergency department visits and hospitalization in older Veterans? J Am Geriatr Soc. 2015;63:2510-2518. DOI: 10.1111/jgs.13841.
11. Liu CW, Einstadter D, Cebul RD. Care fragmentation and emergency department use among complex patients with diabetes. Am J Manag Care. 2010;16:413-420.
12. Nyweide DJ, Anthony DL, Bynum JP, Strawderman RL, Weeks WB, Casalino LP, Fisher ES. Continuity of care and the risk of preventable hospitalization in older adults. JAMA Intern Med. 2013;173:1879-1885. DOI: 10.1001/jamainternmed.2013.10059.
13. vanWalraven C, Oake N, Jennings A, ForsterAJ. The association between continuity of care and outcomes: a systematic and critical review. J Eval Clin Pract. 2010;16:947-956. DOI: 10.1111/j.1365-2753.2009.01235.x.
14. Kern LM, Seirup J, Rajan M, Jawahar R, Stuard SS. Fragmented ambulatory care and subsequent healthcare utilization among Medicare beneficiaries. Am J Manag Care. 2018;24:e278-e284.
15. Howard VJ, Cushman M, Pulley L, Gomez CR, Go RC, Prineas RJ, Graham A, Moy CS, Howard G. The reasons for geographic and racial differences in stroke study: objectives and design. Neuroepidemiology. 2005;25:135-143. DOI: 10.1159/000086678.
16. Safford MM, Brown TM, Muntner PM, Durant RW, Glasser S, Halanych JH, Shikany JM, Prineas RJ, Samdarshi T, Bittner VA, et al. Association of race and sex with risk of incident acute coronary heart disease events. JAMA. 2012;308:1768-1774. DOI: 10.1001/jama.2012.14306.
17. Xie F, Colantonio LD, Curtis JR, Safford MM, Levitan EB, Howard G, Muntner P. Linkage of a populaton-based cohort with primary data collection to Medicare claims: the REasons for Geographic and Racial Differences in Stroke (REGARDS) study. Am J Epidemiol. 2016;184:532544. DOI: 10.1093/aje/kww077.
18. National Committee for Quality Assurance. HEDIS volume 2: technical specifications. Available at: https://www.ncqa.org/hedis/measures/. Accessed January 25, 2021.
19. Bice TW, Boxerman SB. A quantitative measure of continuity of care. Med Care. 1977;15:347-349. DOI: 10.1097/00005650-19770 4000-00010.
20. Pollack CE, Hussey PS, Rudin RS, Fox DS, Lai J, Schneider EC. Measuring care continuity: a comparison of claims-based methods. Med Care. 2016;54:e30-e34. DOI: 10.1097/MLR. 0000000000 000018.
21. Howard VJ, Kleindorfer DO, Judd SE, McClure LA, Safford MM, Rhodes JD, Cushman M, Moy CS, Soliman EZ, Kissela BM, et al. Disparities in stroke incidence contributing to disparities in stroke mortality. Ann Neurol. 2011;69:619-627. DOI: 10.1002/ana.22385.
22. Stroke--1989. Recommendations on stroke prevention, diagnosis, and therapy. Report of the WHO Task force on stroke and other cerebrovascular disorders. Stroke 1989;20:1407-1431. DOI: 10.1161/01. str.20.10.1407.
23. U.S. Department of Agriculture. Rural-urban commuting area codes. Available at: https://www.ers.usda.gov/data-products/rural-urban -commuting-area-codes.aspx. Accessed January 25, 2021.
24. Morisky DE, Green LW, Levine DM. Concurrent and predictive validity of a self-reported measure of medication adherence. Med Care. 1986;24:67-74. DOI: 10.1097/00005650-198601000-00007.
25. Gunzerath L, Faden V, Zakhari S, Warren K. National Institute on Alcohol Abuse and Alcoholism report on moderate drinking. Alcohol Clin Exp Res. 2004;28:829-847. DOI: 10.1097/01.alc.0000128382.79375.b6.
26. Radloff LS. The CES-D scale: a self-report depression scale for research in the general population. Appl Psychol Meas. 1977;1:385-401.
27. Levey AS, Stevens LA, Schmid CH, Zhang YL, Castro AF 3rd, Feldman HI, Kusek JW, Eggers P, Van Lente F, Greene T, et al. A new equation to estimate glomerular filtration rate. Ann Intern Med. 2009;150:604-612. DOI: 10.7326/0003-4819-150-9-200905050-00006.
28. National Kidney Foundation. Albumin-to-creatinine ratio (ACR). Available at https://www.kidney.org/kidneydisease/siemens_hcp_acr. Accessed January 25, 2021.
29. DeSalvo KB, Bloser N, Reynolds K, He J, Muntner P. Mortality prediction with a single general self-rated health question: a meta-analysis. J Gen Intern Med. 2006;21:267-275. DOI: 10.1111/j.1525-1497.2005.00291.x.
30. DeSalvo KB, Fan VS, McDonell MB, Fihn SD. Predicting mortality and healthcare utilization with a single question. Health Serv Res. 2005;40:1234-1246. DOI: 10.1111/j.1475-6773.2005.00404.x.
31. DeSalvo KB, Fisher WP, Tran K, Bloser N, Merrill W, Peabody J. Assessing measurement properties of two single-item general health measures. Qual Life Res. 2006;15:191-201. DOI: 10.1007/s1113 6-005-0887-2.
32. Ware J Jr, Kosinski M, Keller SD. A 12-Item Short-Form Health Survey: construction of scales and preliminary tests of reliability and validity. Med Care. 1996;34:220-233. DOI: 10.1097/00005650-19960 3000-00003.
33. Kern LM, Seirup J, Rajan M, Jawahar R, Stuard SS. Fragmented ambulatory care and subsequent emergency department visits and hospital admissions among Medicaid beneficiaries. Am J Manag Care. 2019;25:107-112.
34. Chang HG, Lininger LL, Doyle JT, Maccubbin PA, Rothenberg RB Application of the Cox model as a predictor of relative risk of coronary heart disease in the Albany Study. Stat Med. 1990;9:287-292. DOI: 10.1002/sim. 4780090311.
35. Fisher LD, Lin DY. Time-dependent covariates in the Cox proportionalhazards regression model. Annu Rev Public Health. 1999;20:145-157. DOI: 10.1146/annurev.publhealth.20.1.145.
36. Fleming TR, Harrington DP. Counting Process and Survival Analysis. Hoboken, NJ: John Wiley \& Sons, Inc.; 2005.
37. Rubin DB, Schenker N. Multiple imputation in health-care databases: an overview and some applications. Stat Med. 1991;10:585-598. DOI: 10.1002/sim. 4780100410.
38. Mavaddat N, van der Linde R, Parker R, Savva G, Kimmonth AL, Brayne C, Mant J. Relationship of self-rated health to stroke incidence and mortality in older individuals with and without a history of stroke: a longitudinal study of the MRC Cognitive Function and Ageing (CFAS) Population. PLoS One. 2016;11:e0150178. DOI: 10.1371/journ al.pone.0150178.
39. Smedley BD, Stith AY, Nelson AR, eds. Unequal Treatment: Confronting Racial and Ethnic Disparities in Health Care. Washington, DC: National Academies Press; 2003.
40. Kern LM, Safford MM, Slavin MJ, Makovkina E, Fudl A, Carrillo JE, Abramson EL. Patients' and providers' views on the causes and consequences of healthcare fragmentation. J Gen Intern Med. 2019;34:899907. DOI: 10.1007/s11606-019-04859-1.
41. Bach PB, Pham HH, Schrag D, Tate RC, Hargraves JL. Primary care physicians who treat blacks and whites. N Engl J Med. 2004;351:575584. DOI: 10.1056/NEJMsa040609.
42. Blustein J, Weiss LJ. Visits to specialists under Medicare: socioeconomic advantage and access to care. J Health Care Poor Underserved. 1998;9:153-169. DOI: 10.1353/hpu.2010.0451.
43. Ferrer RL. Pursuing equity: contact with primary care and specialist clinicians by demographics, insurance, and health status. Ann Fam Med. 2007;5:492-502. DOI: 10.1370/afm.746.
44. Boulware LE, Cooper LA, Ratner LE, LaVeist TA, Powe NR. Race and trust in the health care system. Public Health Rep. 2003;118:358-365. DOI: 10.1016/S0033-3549(04)50262-5.
45. Cooper LA, Roter DL, Carson KA, Beach MC, Sabin JA, Greenwald AG, Inui TS. The associations of clinicians' implicit attitudes about race with medical visit communication and patient ratings of interpersonal care. Am J Public Health. 2012;102:979-987. DOI: 10.2105/ AJPH.2011.300558.

## SUPPLEMENTAL MATERIAL

Data S1. Supplemental Methods. Formula for the Bice-Boxerman Index (BBI)*19
$B B I=\frac{\left(\sum_{i=1}^{p} n_{i}^{2}\right)-n}{n(n-1)}$
where $n=$ total number of visits in the 12-month period
$n_{i}=$ number of visits to provider i
$p=$ total number of providers

* The BBI is a continuous index that ranges from 0 to 1 . We have reversed the direction of the Index, calculating 1 minus BBI , so that higher scores reflect more fragmentation.

Data S2. Supplemental Methods. Schematic diagram of analytic approach*


Time

REGARDS $=$ Reasons for Geographic and Racial Differences in Stroke.
*This schematic diagram illustrates a longitudinal design that treats the exposure (healthcare fragmentation) as time-varying. Observation continues until a stroke occurs, censoring occurs, or the end of the study period is reached. Note that the periods of time for the outcome do not overlap with each other, but they are contiguous (with the next one starting where the previous one left off).
+Observation began after the in-home visit with the first 12-month period for which there was continuous Medicare fee-for-service coverage, starting as early as January 1, 2004. The study continued (as shown in the schematic diagram) until the end of the study period, which was December 31, 2016.

Figure S1. Derivation of the study sample


Key: REGARDS = Reasons for Geographic and Racial Differences in Stroke

Table S1. Baseline characteristics, overall and stratified by self-rated health*

| Characteristic | Overall $(N=12,510)$ | Excellent $(N=2,145)$ | Very good or good $(\mathrm{N}=8,363)$ | $\begin{gathered} \text { Fair or } \\ \text { poor } \\ (\mathrm{N}=2,002) \\ \hline \end{gathered}$ | p -value |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Demographic characteristics |  |  |  |  |  |
| Age, years, mean (SD) | 70.5 (6.0) | 70.4 (5.9) | 70.6 (5.9) | 70.5 (6.2) | 0.38 |
| Gender, female, N (\%) | 6623 (52.9\%) | 994 (46.3\%) | 4425 (52.9\%) | 1204 (60.1\%) | <0.001 |
| Race, white, N (\%) | 8513 (68.0\%) | 1760 (82.1\%) | 5739 (68.6\%) | 1014 (50.6\%) | <0.001 |
| Marital status, married, N (\%) | 7487 (59.8\%) | 1432 (66.8\%) | 5049 (60.4\%) | 1006 (50.2\%) | <0.001 |
| Education, less than high school diploma, N (\%) | 1545 (12.4\%) | 127 (5.9\%) | 889 (10.6\%) | 529 (26.4\%) | <0.001 |
| Annual household income, <\$35,000, N (\%) | 5422 (49.5\%) | 652 (33.9\%) | 3568 (48.8\%) | 1202 (69.8\%) | <0.001 |
| Geographic region, N (\%) |  |  |  |  |  |
| Stroke Belt $\dagger$ | 4551 (36.4\%) | 783 (36.5\%) | 2980 (35.6\%) | 788 (39.4\%) | 0.015 |
| Stroke Buckle $\ddagger$ | 2899 (23.2\%) | 490 (22.8\%) | 1941 (23.2\%) | 468 (23.4\%) |  |
| Neither Stroke Belt nor Stroke Buckle | 5060 (40.4\%) | 872 (40.7\%) | 3442 (41.2\%) | 746 (37.3\%) |  |
| Residence in urban area, N (\%) | 8467 (75.3\%) | 1412 (73.5\%) | 5711 (75.8\%) | 1344 (75.0\%) | 0.003 |
| Medical conditions§ |  |  |  |  |  |
| Hypertension, N (\%) | 7663 (61.5\%) | 874 (40.8\%) | 5256 (63.1\%) | 1533 (76.8\%) | <0.001 |
| Dyslipidemia, N (\%) | 7538 (62.3\%) | 1109 (53.6\%) | 5082 (62.9\%) | 1347 (69.4\%) | <0.001 |
| Diabetes, N (\%) | 2623 (21.7\%) | 166 (8.1\%) | 1665 (20.6\%) | 792 (40.8\%) | <0.001 |
| History of coronary heart disease, N (\%) | 2521 (20.5\%) | 227 (10.7\%) | 1642 (20.0\%) | 652 (33.1\%) | <0.001 |
| Atrial fibrillation, N (\%) | 1149 (9.4\%) | 94 (4.4\%) | 723 (8.8\%) | 332 (17.0\%) | <0.001 |
| Medications |  |  |  |  |  |
| Medication adherence, N (\%) | 8302 (71.1\%) | 1437 (75.2\%) | 5591 (71.2\%) | 1274 (66.8\%) | <0.001 |
| Number of medications, median (IQR) | 6.0 (3.0, 9.0) | 4.0 (2.0, 7.0) | 6.0 (3.0, 8.0) | 8.0 (5.0, 11.0) | <0.001 |
| Anti-hypertensive medication, N (\%) | 6693 (55.4\%) | 671 (32.4\%) | 4602 (56.9\%) | 1420 (73.5\%) | <0.001 |
| Insulin use, N (\%) | 657 (5.5\%) | 23 (1.1\%) | 355 (4.4\%) | 279 (14.9\%) | <0.001 |
| Statin use, N (\%) | 4439 (35.5\%) | 547 (25.5\%) | 3074 (36.8\%) | 818 (40.9\%) | <0.001 |
| Aspirin use, N (\%) | 6084 (48.6\%) | 967 (45.1\%) | 4112 (49.2\%) | 1005 (50.2\%) | <0.001 |
| Warfarin use, N (\%) | 522 (4.2\%) | 40 (1.9\%) | 339 (4.1\%) | 143 (7.1\%) | <0.001 |
| Health behaviors |  |  |  |  |  |
| Current smoker, N (\%) | 1344 (10.8\%) | 144 (6.8\%) | 877 (10.5\%) | 323 (16.2\%) | <0.001 |
| Alcohol use, N (\%) |  |  |  |  |  |
| Heavy | 464 (3.8\%) | 113 (5.3\%) | 314 (3.8\%) | 37 (1.9\%) | <0.001 |
| Moderate | 4049 (33.0\%) | 851 (40.2\%) | 2791 (34.0\%) | 407 (20.7\%) |  |
| None | 7775 (63.3\%) | 1152 (54.4\%) | 5103 (62.2\%) | 1520 (77.4\%) |  |
| Exercise, 0 times per week, N (\%) | 4192 (34.1\%) | 419 (19.8\%) | 2740 (33.3\%) | 1033 (52.4\%) | <0.001 |
| Psychosocial variables |  |  |  |  |  |
| Cares for a family member with a chronic illness or disability, N (\%) | 1493 (11.9\%) | 242 (11.3\%) | 990 (11.9\%) | 261 (13.0\%) | 0.20 |
| Lack of social support, N (\%) | 485 (4.0\%) | 73 (3.5\%) | 320 (3.9\%) | 92 (4.7\%) | 0.13 |
| Depressive symptoms, N (\%) | 710 (5.7\%) | 54 (2.5\%) | 368 (4.4\%) | 288 (14.4\%) | <0.001 |
| Physiological variables |  |  |  |  |  |
| Body mass index, $\mathrm{kg} / \mathrm{m}^{2}$, mean (SD) | 28.9 (5.8) | 26.8 (4.5) | 28.9 (5.6) | 31.2 (6.8) | <0.001 |


| Heart rate, beats per min, mean (SD) | 66.6 (24.8) | 65.4 (36.6) | 66.1 (18.2) | 70.0 (31.9) | <0.001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Systolic blood pressure, mm Hg , mean (SD) | 128.6 (16.3) | 125.7 (15.0) | 128.8 (16.1) | 130.9 (17.6) | $<0.001$ |
| Left ventricular hypertrophy, N (\%) | 1247 (10.1\%) | 136 (6.4\%) | 846 (10.3\%) | 265 (13.4\%) | <0.001 |
| Total cholesterol, mg/dL, mean (SD) | 189.9 (39.4) | 193.3 (36.9) | 189.5 (39.4) | 188.0 (41.6) | <0.001 |
| Low-density lipoprotein cholesterol, mg/dL, mean (SD) | 111.5 (33.9) | 114.5 (31.8) | 111.2 (34.0) | 109.5 (35.4) | <0.001 |
| High-density lipoprotein cholesterol, $\mathrm{mg} / \mathrm{dL}$, mean (SD) | 51.6 (16.3) | 53.9 (17.0) | 51.5 (16.2) | 49.6 (15.6) | <0.001 |
| Glucose, mg/dL, mean (SD) | 103.8 (33.7) | 96.5 (22.6) | 103.2 (32.6) | 114.0 (44.3) | <0.001 |
| Estimated glomerular filtration rate, $\mathrm{mL} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$, mean (SD) | 80.6 (18.9) | 82.3 (15.3) | 80.8 (18.5) | 77.9 (23.2) | <0.001 |
| Urinary albumin-to-creatinine ratio, $\mathrm{mg} / \mathrm{g}$, median (IQR) | 2.1 (0.9, 4.8) | $1.4(0.7,3.2)$ | 2.1 (1.0, 4.6) | 3.4 (1.4, 7.7) | <0.001 |
| C-reactive protein, mg/L, median (IQR) | 7.8 (4.9, 16.7) | 6.4 (4.3, 11.3) | 7.9 (4.9, 16.8) | 10.2 (5.7, 30.2) | <0.001 |
| Self-rated health |  |  |  |  |  |
| Mental component summary score, mean (SD) | 55.1 (7.6) | 56.9 (5.5) | 55.6 (6.9) | 50.9 (10.4) | <0.001 |
| Physical component summary score, mean (SD) | 46.6 (10.3) | 54.1 (5.4) | 47.6 (8.5) | 34.0 (10.3) | <0.001 |
| Healthcare utilization |  |  |  |  |  |
| Hospitalization in the first year of observation, N (\%) | 1245 (10.0\%) | 153 (7.1\%) | 803 (9.6\%) | 289 (14.4\%) | <0.001 |

* Missing data: education ( $N=4$ ), income (1,554), residence in urban area ( $N=1,261$ ), medication count ( $N=12$ ), smoking ( $N=48$ ), exercise ( $N=207$ ), caregiver status ( $N=12$ ), lack of social support ( $N=308$ ), depressive symptoms ( $N=72$ ), body mass index ( $N=65$ ), heart rate (111), systolic blood pressure ( $N=34$ ), total cholesterol ( $N=454$ ), low density lipoprotein ( $N=680$ ), high-density lipoprotein ( $N=529$ ), glucose ( $N=456$ ), estimated glomerular filtration rate ( $\mathrm{N}=454$ ), urinary albumin-to-creatinine ratio ( $\mathrm{N}=505$ ), C-reactive protein ( $\mathrm{N}=724$ ), mental component summary score ( $\mathrm{N}=591$ ), and physical component summary score ( $\mathrm{N}=591$ ).
† Stroke Belt = North Carolina, South Carolina, Georgia, Tennessee, Mississippi, Alabama, Louisiana, and Arkansas (except for 153 coastal counties that constitute the Stroke Buckle). ${ }^{15}$
$\ddagger$ Stroke Buckle = 153 coastal counties in North Carolina, South Carolina, and Georgia. ${ }^{15,21}$
§ See methods section for detailed definitions of variables.
Key: $\mathrm{rBBI}=$ reversed Bice-Boxerman Index (BBI), equivalent to $1-\mathrm{BBI}$, such that higher scores indicate more fragmentation. ${ }^{19} \mathrm{IQR}=$ interquartile range. $\mathrm{SF}=$ short-form survey. ${ }^{32}$ Percentages may not sum to 100 due to rounding.

Table S2. Baseline characteristics, overall and stratified by race*

| Characteristic | Overall $(N=12,510)$ | $\begin{gathered} \text { Black } \\ (\mathrm{N}=3,997) \end{gathered}$ | White $(N=8,513)$ | p-value |
| :---: | :---: | :---: | :---: | :---: |
| Demographic characteristics |  |  |  |  |
| Age, years, mean (SD) | 70.5 (6.0) | 70.2 (5.8) | 70.7 (6.0) | <0.001 |
| Gender, female, N (\%) | 6623 (52.9\%) | 2531 (63.3\%) | 4092 (48.1\%) | <0.001 |
| Marital status, married, N (\%) | 7487 (59.8\%) | 1767 (44.2\%) | 5720 (67.2\%) | <0.001 |
| Education, less than high school diploma, N (\%) | 1545 (12.4\%) | 912 (22.8\%) | 633 (7.4\%) | <0.001 |
| Annual household income, <\$35,000, N (\%) | 5422 (49.5\%) | 2269 (64.6\%) | 3153 (42.4\%) | <0.001 |
| Geographic region, N (\%) |  |  |  |  |
| Stroke Belt $\dagger$ | 4551 (36.4\%) | 1345 (33.7\%) | 3206 (37.7\%) | <0.001 |
| Stroke Buckle $\ddagger$ | 2899 (23.2\%) | 809 (20.2\%) | 2090 (24.6\%) |  |
| Neither Stroke Belt nor Stroke Buckle | 5060 (40.4\%) | 1843 (46.1\%) | 3217 (37.8\%) |  |
| Residence in urban area, N (\%) | 8467 (75.3\%) | 3299 (89.5\%) | 5168 (68.3\%) | <0.001 |
| Medical conditions§ |  |  |  |  |
| Hypertension, N (\%) | 7663 (61.5\%) | 2990 (75.0\%) | 4673 (55.1\%) | <0.001 |
| Dyslipidemia, N (\%) | 7538 (62.3\%) | 2250 (58.9\%) | 5288 (63.9\%) | <0.001 |
| Diabetes, N (\%) | 2623 (21.7\%) | 1270 (33.1\%) | 1353 (16.4\%) | <0.001 |
| History of coronary heart disease, N (\%) | 2521 (20.5\%) | 693 (17.6\%) | 1828 (21.8\%) | <0.001 |
| Atrial fibrillation, N (\%) | 1149 (9.4\%) | 312 (8.0\%) | 837 (10.0\%) | <0.001 |
| Medications |  |  |  |  |
| Medication adherence, N (\%) | 8302 (71.1\%) | 2620 (70.7\%) | 5682 (71.3\%) | 0.53 |
| Number of medications, median (IQR) | 6.0 (3.0, 9.0) | 5.0 (3.0, 8.0) | 6.0 (3.0, 9.0) | <0.001 |
| Anti-hypertensive medication, N (\%) | 6693 (55.4\%) | 2699 (69.8\%) | 3994 (48.6\%) | <0.001 |
| Insulin use, N (\%) | 657 (5.5\%) | 374 (9.9\%) | 283 (3.5\%) | <0.001 |
| Statin use, N (\%) | 4439 (35.5\%) | 1304 (32.6\%) | 3135 (36.8\%) | <0.001 |
| Aspirin use, N (\%) | 6084 (48.6\%) | 1749 (43.8\%) | 4335 (50.9\%) |  |
| Warfarin use, N (\%) | 522 (4.2\%) | 101 (2.5\%) | 421 (5.0\%) | <0.001 |
| Health behaviors |  |  |  |  |
| Current smoker, N (\%) | 1344 (10.8\%) | 510 (12.8\%) | 834 (9.8\%) | <0.001 |
| Alcohol use, N (\%) |  |  |  | <0.001 |
| Heavy | 464 (3.8\%) | 79 (2.0\%) | 385 (4.6\%) |  |
| Moderate | 4049 (33.0\%) | 906 (23.2\%) | 3143 (37.5\%) |  |
| None | 7775 (63.3\%) | 2920 (74.8\%) | 4855 (57.9\%) |  |
| Exercise frequency, 0 times per week, N (\%) | 4192 (34.1\%) | 1515 (38.6\%) | 2677 (32.0\%) | <0.001 |
| Psychosocial variables |  |  |  |  |
| Cares for a family member with a chronic illness or disability, N (\%) | 1493 (11.9\%) | 531 (13.3\%) | 962 (11.3\%) | 0.001 |
| Lack of social support, N (\%) | 485 (4.0\%) | 154 (4.0\%) | 331 (4.0\%) | 0.99 |
| Depressive symptoms, N (\%) | 710 (5.7\%) | 306 (7.7\%) | 404 (4.8\%) | <0.001 |
| Physiological variables |  |  |  |  |
| Body mass index, $\mathrm{kg} / \mathrm{m}^{2}$, mean (SD) | 28.9 (5.8) | 30.6 (6.3) | 28.1 (5.4) | <0.001 |
| Heart rate, beats per minute, mean (SD) | 66.6 (24.8) | 68.0 (23.8) | 65.9 (25.3) | <0.001 |
| Systolic blood pressure, mm Hg , mean (SD) | 128.6 (16.3) | 131.8 (17.0) | 127.1 (15.7) | <0.001 |
| Left ventricular hypertrophy, N (\%) | 1247 (10.1\%) | 643 (16.3\%) | 604 (7.2\%) | $<0.001$ |
| Total cholesterol, mg/dL, mean (SD) | 189.9 (39.4) | 192.9 (40.9) | 188.5 (38.6) | <0.001 |
| Low-density lipoprotein cholesterol, mg/dL, mean (SD) | 111.5 (33.9) | 116.4 (36.1) | 109.2 (32.5) | <0.001 |


| High-density lipoprotein cholesterol, mg/dL, mean (SD) | 51.6 (16.3) | 53.9 (15.8) | 50.6 (16.4) | <0.001 |
| :---: | :---: | :---: | :---: | :---: |
| Glucose, mg/dL, mean (SD) | 103.8 (33.7) | 109.7 (41.9) | 101.1 (28.8) | <0.001 |
| Estimated glomerular filtration rate, $\mathrm{mL} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$, mean (SD) | 80.6 (18.9) | 83.9 (22.4) | 79.1 (16.8) | <0.001 |
| Urinary albumin-to-creatinine ratio, $\mathrm{mg} / \mathrm{g}$, median (IQR) | 2.1 (0.9, 4.8) | 2.9 (1.2, 6.4) | 1.9 (0.9, 4.1) | $<0.001$ |
| C-reactive protein, mg/L, median (IQR) | 7.8 (4.9, 16.7) | 8.3 (4.9, 20.2) | 7.6 (4.9, 15.4) | <0.001 |
| Self-rated health |  |  |  |  |
| Self-rated general health (SF-1), N (\%) |  |  |  |  |
| Excellent | 2145 (17.1\%) | 385 (9.6\%) | 1760 (20.7\%) | <0.001 |
| Very good | 3940 (31.5\%) | 943 (23.6\%) | 2997 (35.2\%) |  |
| Good | 4423 (35.4\%) | 1681 (42.1\%) | 2742 (32.2\%) |  |
| Fair | 1669 (13.3\%) | 843 (21.1\%) | 826 (9.7\%) |  |
| Poor | 333 (2.7\%) | 145 (3.6\%) | 188 (2.2\%) |  |
| Mental component summary score, mean (sd) | 55.1 (7.6) | 54.2 (8.4) | 55.5 (7.2) | <0.001 |
| Physical component summary score, mean (sd) | 46.6 (10.3) | 45.1 (10.4) | 47.2 (10.1) | <0.001 |
| Healthcare utilization |  |  |  |  |
| Hospitalization in the first year of observation, N (\%) | 1245 (10.0\%) | 454 (11.4\%) | 791 (9.3\%) | <0.001 |

Table S3. Baseline characteristics among black participants, overall and stratified by selfrated health

| Characteristic | Overall $(N=3,997)$ | Excellent $(\mathrm{N}=385)$ | Very good or good ( $\mathrm{N}=2624$ ) | $\begin{gathered} \text { Fair or } \\ \text { poor } \\ (\mathrm{N}=988) \\ \hline \end{gathered}$ | p-value |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Demographic characteristics |  |  |  |  |  |
| Age, years, mean (SD) | 70.2 (5.8) | 70.0 (5.5) | 70.2 (5.7) | 70.3 (6.2) | 0.7 |
| Gender, female, N (\%) | 2531 (63.3\%) | 207 (53.8\%) | 1644 (62.7\%) | 680 (68.8\%) | <0.001 |
| Marital status, married, N (\%) | 1767 (44.2\%) | 174 (45.2\%) | 1207 (46.0\%) | 386 (39.1\%) | <0.001 |
| Education, less than high school diploma, N (\%) | 912 (22.8\%) | 59 (15.3\%) | 499 (19.0\%) | 354 (35.8\%) | <0.001 |
| Annual household income, <\$35,000, N (\%) | 2269 (64.6\%) | 175 (51.0\%) | 1434 (61.7\%) | 660 (78.0\%) | <0.001 |
| Geographic region, N (\%) |  |  |  |  |  |
| Stroke Belt $\dagger$ | 1345 (33.7\%) | 125 (32.5\%) | 869 (33.1\%) | 351 (35.5\%) |  |
| Stroke Buckle $\ddagger$ | 809 (20.2\%) | 81 (21.0\%) | 522 (19.9\%) | 206 (20.9\%) |  |
| Neither Stroke Belt nor Stroke Buckle | 1843 (46.1\%) | 179 (46.5\%) | 1233 (47.0\%) | 431 (43.6\%) | 0.46 |
| Residence in urban area, N (\%) | 3299 (89.5\%) | 322 (90.7\%) | 2199 (90.8\%) | 778 (85.6\%) | <0.001 |
| Medical conditions§ |  |  |  |  |  |
| Hypertension, N (\%) | 2990 (75.0\%) | 216 (56.1\%) | 1937 (74.1\%) | 837 (84.8\%) | <0.001 |
| Dyslipidemia, N (\%) | 2250 (58.9\%) | 177 (48.6\%) | 1440 (57.6\%) | 633 (66.1\%) | <0.001 |
| Diabetes, N (\%) | 1270 (33.1\%) | 69 (18.8\%) | 759 (30.2\%) | 442 (46.1\%) | <0.001 |
| History of coronary heart disease, N (\%) | 693 (17.6\%) | 34 (9.1\%) | 384 (14.9\%) | 275 (28.3\%) | <0.001 |
| Atrial fibrillation, N (\%) | 312 (8.0\%) | 9 (2.4\%) | 162 (6.3\%) | 141 (14.8\%) | <0.001 |
| Medications |  |  |  |  |  |
| Medication adherence, N (\%) | 2620 (70.7\%) | 247 (75.1\%) | 1742 (71.4\%) | 631 (67.5\%) | 0.016 |
| Number of medications, median (IQR) | 5.0 (3.0, 8.0) | 3.0 (1.0, 5.0) | 5.0 (3.0, 7.0) | 7.0 (5.0, 10.0) | <0.001 |
| Anti-hypertensive medication, N (\%) | 2699 (69.8\%) | 173 (45.6\%) | 1743 (68.9\%) | 783 (81.6\%) | <0.001 |
| Insulin use, N (\%) | 374 (9.9\%) | 14 (3.8\%) | 196 (7.9\%) | 164 (17.7\%) | <0.001 |
| Statin use, N (\%) | 1304 (32.6\%) | 73 (19.0\%) | 842 (32.1\%) | 389 (39.4\%) | <0.001 |
| Aspirin use, N (\%) | 1749 (43.8\%) | 135 (35.1\%) | 1141 (43.5\%) | 473 (47.9\%) | <0.001 |
| Warfarin use, N (\%) | 101 (2.5\%) | 4 (1.0\%) | 54 (2.1\%) | 43 (4.4\%) | <0.001 |
| Health behaviors |  |  |  |  |  |
| Current smoker, N (\%) | 510 (12.8\%) | 46 (12.1\%) | 306 (11.7\%) | 158 (16.0\%) | 0.002 |
| Alcohol use, N (\%) |  |  |  |  |  |
| Heavy | 79 (2.0\%) | 10 (2.7\%) | 50 (2.0\%) | 19 (2.0\%) |  |
| Moderate | 906 (23.2\%) | 88 (23.4\%) | 667 (26.0\%) | 151 (15.6\%) |  |
| None | 2920 (74.8\%) | 278 (73.9\%) | 1845 (72.0\%) | 797 (82.4\%) | <0.001 |
| Exercise, 0 times per week, N (\%) | 1515 (38.6\%) | 89 (23.7\%) | 917 (35.6\%) | 509 (52.3\%) | <0.001 |
| Psychosocial variables |  |  |  |  |  |
| Cares for a family member with a chronic illness or disability, N (\%) | 531 (13.3\%) | 39 (10.1\%) | 362 (13.8\%) | 130 (13.2\%) | 0.14 |
| Lack of social support, N (\%) | 154 (4.0\%) | 12 (3.2\%) | 103 (4.1\%) | 39 (4.1\%) | 0.72 |
| Depressive symptoms, N (\%) | 306 (7.7\%) | 15 (3.9\%) | 145 (5.6\%) | 146 (14.8\%) | <0.001 |
| Physiological variables |  |  |  |  |  |
| Body mass index, $\mathrm{kg} / \mathrm{m}^{2}$, mean (SD) | 30.6 (6.3) | 28.5 (5.1) | 30.3 (6.0) | 32.3 (6.9) | <0.001 |


| Heart rate, beats per min, mean (SD) | 68.0 (23.8) | 68.5 (48.8) | 67.3 (21.4) | 69.5 (12.4) | 0.041 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Systolic blood pressure, mm Hg , mean (SD) | 131.8 (17.0) | 129.2 (16.1) | 131.7 (16.8) | 133.1 (17.8) | $<0.001$ |
| Left ventricular hypertrophy, N (\%) | 643 (16.3\%) | 50 (13.2\%) | 406 (15.7\%) | 187 (19.1\%) | 0.01 |
| Total cholesterol, mg/dL, mean (SD) | 192.9 (40.9) | 198.9 (39.2) | 192.9 (40.9) | 190.7 (41.4) | 0.005 |
| Low-density lipoprotein cholesterol, $\mathrm{mg} / \mathrm{dL}$, mean (SD) | 116.4 (36.1) | 122.3 (34.6) | 116.3 (36.1) | 114.5 (36.7) | 0.002 |
| High-density lipoprotein cholesterol, $\mathrm{mg} / \mathrm{dL}$, mean (SD) | 53.9 (15.8) | 55.4 (16.4) | 54.2 (15.8) | 52.4 (15.6) | 0.002 |
| Glucose, mg/dL, mean (SD) | 109.7 (41.9) | 102.0 (31.4) | 108.2 (39.8) | 116.6 (49.4) | <0.001 |
| Estimated glomerular filtration rate, $\mathrm{mL} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$, mean (SD) | 83.9 (22.4) | 88.0 (18.9) | 84.5 (21.6) | 80.6 (25.2) | <0.001 |
| Urinary albumin-to-creatinine ratio, $\mathrm{mg} / \mathrm{g}$, median (IQR) | 2.9 (1.2, 6.4) | 6.5 (4.3, 11.7) | 8.1 (4.8, 18.6) | 10.9 (5.5, 36.8) | <0.001 |
| C-reactive protein, mg/L, median (IQR) | 8.3 (4.9, 20.2) | 2.0 (0.8, 4.3) | $2.8(1.3,6.0)$ | 3.9 (1.5, 8.2) | <0.001 |
| Self-rated health |  |  |  |  |  |
| Mental component summary score, mean (SD) | 54.2 (8.4) | 56.7 (6.0) | 55.4 (7.2) | 50.2 (10.6) | <0.001 |
| Physical component summary score, mean (SD) | 45.1 (10.4) | 54.1 (5.4) | 47.3 (8.5) | 35.7 (10.2) | <0.001 |
| Healthcare utilization |  |  |  |  |  |
| Hospitalization in the first year of observation, N (\%) | 454 (11.4\%) | 39 (10.1\%) | 280 (10.7\%) | 135 (13.7\%) | 0.03 |

Table S4. Baseline characteristics among white participants, overall and stratified by self-rated health

| Characteristic | Overall $(N=8,513)$ | Excellent $(N=1,760)$ | Very good or good ( $\mathrm{N}=5,739$ ) | $\begin{gathered} \text { Fair or } \\ \text { poor } \\ (\mathrm{N}=1,014) \end{gathered}$ | $p$-value |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Demographic characteristics |  |  |  |  |  |
| Age, years, mean (SD) | 70.7 (6.0) | 70.5 (6.0) | 70.8 (6.0) | 70.7 (6.2) | 0.22 |
| Gender, female, N (\%) | 4092 (48.1\%) | 787 (44.7\%) | 2781 (48.5\%) | 524 (51.7\%) | 0.001 |
| Marital status, married, N (\%) | 5720 (67.2\%) | 1258 (71.5\%) | 3842 (66.9\%) | 620 (61.1\%) | <0.001 |
| Education, less than high school diploma, N (\%) | 633 (7.4\%) | 68 (3.9\%) | 390 (6.8\%) | 175 (17.3\%) | <0.001 |
| Annual household income, <\$35,000, N (\%) | 3153 (42.4\%) | 477 (30.2\%) | 2134 (42.8\%) | 542 (61.9\%) | <0.001 |
| Geographic region, N (\%) |  |  |  |  |  |
| Stroke Belt $\dagger$ | 3206 (37.7\%) | 658 (37.4\%) | 2111 (36.8\%) | 437 (43.1\%) | <0.001 |
| Stroke Buckle $\ddagger$ | 2090 (24.6\%) | 409 (23.2\%) | 1419 (24.7\%) | 262 (25.8\%) |  |
| Neither Stroke Belt nor Stroke Buckle | 3217 (37.8\%) | 693 (39.4\%) | 2209 (38.5\%) | 315 (31.1\%) |  |
| Residence in urban area, N (\%) | 5168 (68.3\%) | 1090 (69.6\%) | 3512 (68.7\%) | 566 (64.1\%) | 0.003 |
| Medical conditions§ |  |  |  |  |  |
| Hypertension, N (\%) | 4673 (55.1\%) | 658 (37.5\%) | 3319 (58.0\%) | 696 (69.0\%) | <0.001 |
| Dyslipidemia, N (\%) | 5288 (63.9\%) | 932 (54.7\%) | 3642 (65.2\%) | 714 (72.6\%) | <0.001 |
| Diabetes, N (\%) | 1353 (16.4\%) | 97 (5.7\%) | 906 (16.3\%) | 350 (35.6\%) | <0.001 |
| History of coronary heart disease, N (\%) | 1828 (21.8\%) | 193 (11.1\%) | 1258 (22.3\%) | 377 (37.9\%) | <0.001 |
| Atrial fibrillation, N (\%) | 837 (10.0\%) | 85 (4.9\%) | 561 (10.0\%) | 191 (19.2\%) | <0.001 |
| Medications |  |  |  |  |  |
| Ideal medication adherence | 5682 (71.3\%) | 1190 (75.2\%) | 3849 (71.1\%) | 643 (66.1\%) | <0.001 |
| Number of medications, median (IQR) | 6.0 (3.0, 9.0) | 4.0 (2.0, 7.0) | 6.0 (4.0, 9.0) | 9.0 (6.0, 12.0) | <0.001 |
| Anti-hypertensive medication, N (\%) | 3994 (48.6\%) | 498 (29.5\%) | 2859 (51.4\%) | 637 (65.5\%) | <0.001 |
| Insulin use, N (\%) | 283 (3.5\%) | 9 (0.5\%) | 159 (2.9\%) | 115 (12.2\%) | $<0.001$ |
| Statin use, N (\%) | 3135 (36.8\%) | 474 (26.9\%) | 2232 (38.9\%) | 429 (42.3\%) | <0.001 |
| Aspirin use, N (\%) | 4335 (50.9\%) | 832 (47.3\%) | 2971 (51.8\%) | 532 (52.5\%) | 0.002 |
| Warfarin use, N (\%) | 421 (5.0\%) | 36 (2.1\%) | 285 (5.0\%) | 100 (9.9\%) | <0.001 |
| Health behaviors |  |  |  |  |  |
| Current smoker, N (\%) | 834 (9.8\%) | 98 (5.6\%) | 571 (10.0\%) | 165 (16.3\%) | <0.001 |
| Alcohol use, N (\%) |  |  |  |  |  |
| Heavy | 385 (4.6\%) | 103 (5.9\%) | 264 (4.7\%) | 18 (1.8\%) | <0.001 |
| Moderate | 3143 (37.5\%) | 763 (43.9\%) | 2124 (37.6\%) | 256 (25.7\%) |  |
| None | 4855 (57.9\%) | 874 (50.2\%) | 3258 (57.7\%) | 723 (72.5\%) |  |
| Exercise, 0 times per week, N (\%) | 2677 (32.0\%) | 330 (19.0\%) | 1823 (32.3\%) | 524 (52.6\%) | <0.001 |
| Psychosocial variables |  |  |  |  |  |
| Cares for a family member with a chronic illness or disability, N (\%) | 962 (11.3\%) | 203 (11.5\%) | 628 (11.0\%) | 131 (12.9\%) | 0.18 |
| Lack of social support, N (\%) | 331 (4.0\%) | 61 (3.6\%) | 217 (3.9\%) | 53 (5.3\%) | 0.061 |
| Depressive symptoms, N (\%) | 404 (4.8\%) | 39 (2.2\%) | 223 (3.9\%) | 142 (14.1\%) | <0.001 |
| Physiological variables |  |  |  |  |  |
| Body mass index, $\mathrm{kg} / \mathrm{m}^{2}$, mean (SD) | 28.1 (5.4) | 26.4 (4.2) | 28.2 (5.3) | 30.1 (6.5) | <0.001 |


| Heart rate, beats per min, mean (SD) | 65.9 (25.3) | 64.7 (33.3) | 65.5 (16.5) | 70.4 (43.1) | <0.001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Systolic blood pressure, mm Hg , mean (SD) | 127.1 (15.7) | 124.9 (14.6) | 127.5 (15.7) | 128.8 (17.1) | <0.001 |
| Left ventricular hypertrophy, N (\%) | 604 (7.2\%) | 86 (4.9\%) | 440 (7.8\%) | 78 (7.8\%) | <0.001 |
| Total cholesterol, mg/dL, mean (SD) | 188.5 (38.6) | 192.1 (36.3) | 188.0 (38.6) | 185.4 (41.7) | <0.001 |
| Low-density lipoprotein cholesterol, $\mathrm{mg} / \mathrm{dL}$, mean (SD) | 109.2 (32.5) | 112.8 (30.9) | 108.9 (32.7) | 104.6 (33.3) | <0.001 |
| High-density lipoprotein cholesterol, $\mathrm{mg} / \mathrm{dL}$, mean (SD) | 50.6 (16.4) | 53.6 (17.1) | 50.3 (16.2) | 46.9 (15.0) | <0.001 |
| Glucose, mg/dL, mean (SD) | 101.1 (28.8) | 95.3 (20.0) | 101.0 (28.6) | 111.4 (38.7) | <0.001 |
| Estimated glomerular filtration rate, $\mathrm{mL} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$, mean (SD) | 79.1 (16.8) | 81.1 (14.1) | 79.2 (16.7) | 75.2 (20.7) | <0.001 |
| Urinary albumin-to-creatinine ratio, $\mathrm{mg} / \mathrm{g}$, median (IQR) | 1.9 (0.9, 4.1) | 6.3 (4.4, 11.1) | 7.8 (4.9, 16.2) | $9.9(5.8,25.0)$ | $<0.001$ |
| C-reactive protein, mg/L, median (IQR) | 7.6 (4.9, 15.4) | 1.3 (0.7, 3.0) | 1.9 (0.9, 4.1) | 3.0 (1.3, 7.0) | <0.001 |
| Self-rated health |  |  |  |  |  |
| Mental component summary score, mean (SD) | 55.5 (7.2) | 56.9 (5.4) | 55.7 (6.8) | 51.6 (10.1) | <0.001 |
| Physical component summary score, mean (SD) | 47.2 (10.1) | 54.0 (5.4) | 47.7 (8.4) | 32.4 (10.2) | <0.001 |
| Healthcare utilization |  |  |  |  |  |
| Hospitalization in the first year of observation, N (\%) | 791 (9.3\%) | 114 (6.5\%) | 523 (9.1\%) | 154 (15.2\%) | <0.001 |

Table S5. Fully adjusted hazard of incident stroke, comparing high to low fragmentation, overall and stratified by self-rated general health and race, with sensitivity analyses varying the cut-off for high fragmentation*


* Results were derived from Cox proportional hazards models that treat fragmentation as a time-varying exposure, consider 12 months of the exposure at a time and outcomes that occur in the 3 months immediately following the exposure period, and occur over up to 11.8 years of follow-up. Fragmentation is measured with the reversed BiceBoxerman Index. We were not able to calculate results for black participants with excellent self-rated health because they had too few events.

Table S6. Fully adjusted hazard of incident stroke, overall and stratified by self-rated general health and race, modeling fragmentation as a variable with four levels*

| Fragmentation score | N (\%) $\dagger$ | Number of strokes ${ }^{\dagger}$ | Overall | Self-rated general health |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Excellent | Very Good or Good | Fair or Poor |
|  |  |  | Adjusted hazard ratios (95\% confidence intervals) |  |  |  |
| ALL PARTICIPANTS | $\mathrm{N}=12,510$ |  |  |  |  |  |
| $\leq 0.80$ | 6672 (53.3\%) | 324 | 1.00 (ref) | 1.00 (ref) | 1.00 (ref) | 1.00 (ref) |
| 0.81-0.84 | 2629 (21.0\%) | 138 | 1.04 (0.84, 1.28) | 0.85 (0.51, 1.42) | 1.14 (0.89, 1.47) | 0.81 (0.47, 1.39) |
| 0.85-0.89 | 1496 (12.0\%) | 77 | 1.00 (0.79, 1.26) | 0.55 (0.27, 1.12) | 1.10 (0.82, 1.46) | 1.06 (0.63, 1.80) |
| $\geq 0.90$ | 1713 (13.7\%) | 72 | 1.06 (0.84, 1.34) | 0.70 (0.38, 1.27) | 1.01 (0.75, 1.37) | 1.63 (1.00, 2.67) |
| WHITE  <br> PARTICIPANTS $\mathrm{N}=8,513$ |  |  |  |  |  |  |
| $\leq 0.80$ | 4224 (49.6\%) | 218 | 1.00 (ref) | 1.00 (ref) | 1.00 (ref) | 1.00 (ref) |
| 0.81-0.84 | 1884 (22.1\%) | 108 | 1.05 (0.82, 1.34) | 0.92 (0.54, 1.58) | 1.17 (0.86, 1.58) | 0.76 (0.36, 1.62) |
| 0.85-0.89 | 1079 (12.7\%) | 51 | 1.02 (0.78, 1.33) | 0.61 (0.30, 1.26) | 1.24 (0.90, 1.71) | 0.77 (0.36, 1.64) |
| $\geq 0.90$ | 1326 (15.6\%) | 54 | 0.98 (0.74, 1.30) | 0.68 (0.35, 1.30) | 1.08 (0.77, 1.53) | 1.13 (0.54, 2.35) |
| BLACK  <br> PARTICIPANTS $\mathrm{N}=3,997$ |  |  |  |  |  |  |
| $\leq 0.80$ | 2448 (61.2\%) | 106 | 1.00 (ref) |  | 1.00 (ref) | 1.00 (ref) |
| 0.81-0.84 | 745 (18.6\%) | 30 | 0.98 (0.66, 1.45) | --- | 1.10 (0.69, 1.74) | 0.79 (0.35, 1.76) |
| 0.85-0.89 | 417 (10.4\%) | 26 | 0.88 (0.54, 1.42) | --- | 0.71 (0.37, 1.40) | 1.41 (0.67, 2.96) |
| $\geq 0.90$ | 387 (9.7\%) | 18 | 1.23 (0.79, 1.90) | --- | 0.78 (0.40, 1.54) | 2.18 (1.11, 4.28) |

* Results were derived from Cox proportional hazards models that treat fragmentation as a time-varying exposure, consider 12 months of the exposure at a time and outcomes that occur in the 3 months immediately following the exposure period, and occur over up to 11.8 years of follow-up. Fragmentation is measured with the reversed BiceBoxerman Index. We were not able to calculate results for black participants with excellent self-rated health because they had too few events.
$\dagger$ The sample size should be interpreted as the number of participants with the given level of fragmentation in the first year of observation. The number of strokes should be interpreted as the event count for the participants who had the given level of fragmentation in the first year of observation. The models, however, treat fragmentation as a time-varying exposure.

Figure S2. An illustrative nomogram showing how fragmentation scores (reversed BiceBoxerman scores*) vary with the proportion of visits with the most frequently seen provider, among those with 12 ambulatory visits ${ }^{\dagger}$

*Higher scores reflect more fragmentation of care.
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