# Sex-specific association between coffee consumption and incident chronic kidney disease: a population-based analysis of 359,906 participants from the UK Biobank 

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#### Abstract

Background: The risk for chronic kidney disease (CKD) is influenced by genetic predisposition, sex, and lifestyle. Previous research indicates that coffee is a potentially protective factor in CKD. The current study aims to investigate whether sex disparity exists in the coffee-CKD association, and whether genetic risk of CKD or genetic polymorphisms of caffeine metabolism affect this association. Methods: A total of 359,906 participants from the UK Biobank who were enrolled between 2006 and 2010 were included in this prospective cohort study, which aimed to estimate the hazard ratios for coffee intake and incident CKD using a Cox proportional hazard model. Allele scores of CKD and caffeine metabolism were additionally adjusted for in a subsample with qualified genetic data ( $n=255,343$ ). Analyses stratified by genetic predisposition, comorbidities, and sex hormones were performed. Tests based on Bayesian model averaging were conducted to ascertain the robustness of the results. Results: Coffee was inversely associated with CKD in a dose-dependent manner. The effects of coffee did not differ across different strata of genetic risk for CKD, but were more evident among slower genetically predicted caffeine metabolizers. Significant sex disparity was observed ( $P$ value for interaction $=0.013$ ), in that coffee drinking was only associated with the risk reduction of CKD in females. Subgroup analysis revealed that testosterone and sex hormone-binding globulin (SHBG), but not estradiol, modified the coffee-CKD association. Conclusions: In addition to the overall inverse coffee-CKD association that was observed in the general population, we could also establish that a sex disparity existed, in that females were more likely to experience the benefit of the association. Testosterone and SHBG may partly account for the sex disparity.


Keywords: Coffee; Chronic kidney diseases; Genotype; Sex

## Introduction

Chronic kidney disease (CKD) is a major public health problem with substantial comorbidities and disease burden. The statistics from the Global Burden of Disease Study 2017 reveal that approximately one-tenth of the world's population was affected by CKD, and that it ranked as the 12 th leading cause of death globally, causing 35.8 million disability-adjusted life years in 2017. ${ }^{[1,2]}$

Coffee is one of the most commonly consumed beverages worldwide, and is reported to be related to risk reduction of all-cause mortality, ${ }^{[3]}$ as well as multiple health

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outcomes, such as obesity, metabolic syndrome, type 2 diabetes, ${ }^{[4]}$ and cardiovascular disease (CVD). ${ }^{[5]}$ Studies examining the overall associations between coffee consumption and CKD have yielded mixed results. Several cohort studies and meta-analyses reported that coffee is associated with decreased CKD risk, ${ }^{[6-10]}$ while others found no significant association. ${ }^{[11,12]}$

Despite the accumulating evidence supporting the renoprotective effect of coffee, considering that CKD exhibits sex disparities in its incidence and progression, ${ }^{[13]}$ the question still remains on whether both sexes could benefit alike from coffee. Coffee consumption has actually been

[^0]reported to affect the risk of some diseases sex-specifically. For instance, Hsu et al ${ }^{[14]}$ found that coffee consumption significantly increased the high-density lipoprotein cholesterol level only in females but not males. Similarly, Lee et al ${ }^{[15]}$ reported that the protective effect of habitual coffee drinking on incident stroke presented with sex disparity.

The genetic predisposition also contributes to the development of CKD. ${ }^{[16]}$ Meanwhile, the genetic polymorphisms affecting caffeine metabolism are also associated with increased risk of several health impairments, including hypertension, impaired fasting glucose, and myocardial infarction. ${ }^{[17-19]}$ Therefore, it is worth examining whether the effect of coffee on incident CKD is independent of genetic factors.

Integrating individual phenotype and genotype data from the UK Biobank, we conducted a comprehensive prospective cohort study to investigate the association between coffee and CKD, considering the impact of sex, genetic risk of CKD, and caffeine metabolism polymorphisms.

## Methods

## Study design

The UK Biobank, the source of the data used in the present study, is a large-scale population-based cohort with indepth genetic and health information of more than 500,000 participants. The UK Biobank data used in this study were derived from the details of participants recruited from 22 assessment centers across the United Kingdom during 2006 to $2010 .{ }^{[20]}$ With the consent of participants, health-related outcomes were obtained periodically from external health care providers. ${ }^{[21]}$ Hospital inpatient records were linked to Hospital Episode Statistics (HES) for England, Scottish Morbidity Record for Scotland, and Patient Episode Database for Wales. Data on mortality were available from National Health Service (NHS) Digital in England and Wales, and NHS Central Register in Scotland. Primary care data were linked with these records by general health care practitioners. Participants were genotyped using UK BiLEVE and UK Biobank Axiom array, which share $95 \%$ common markers, and variants were imputed using Haplotype Reference Consortium, as well as merged UK10K and 1000 Genomes phase 3 reference panels. ${ }^{[22]}$ The researches had applied to access the UK Biobank database with the application approval number of 54803 .

## Study population

For the primary analysis, we excluded participants based on the following criteria: (1) lost to follow-up for any reason ( $n$ $=1346$ ); (2) without complete information on coffee consumption ( $n=2248$ ); (3) without results of testosterone or sex hormone-binding globulin (SHBG) ( $n=112,349$ ); and (4) with any congenital or acquired CKD preceding or within 3 months of recruitment, where the diagnostic criteria used to infer CKD were estimated glomerular filtration rate $(e G F R)<60 \mathrm{~mL} \cdot \mathrm{~min}^{-1} \cdot 1.73 \mathrm{~m}^{-2}$ and urine albumin to creatinine ratio $(\mathrm{uACR}) \geq 30 \mathrm{mg} / \mathrm{g}(n=26,658)$ [Figure 1]. Supplementary Table 1 [http://links.lww.com/

CM9/B116] presents a comparison of the baseline characteristics of participants enrolled in the primary analysis cohort, those lost to follow-up, and those with missing information on coffee intake, testosterone, and SHBG. In order to explore the influence of genetic predisposition on the studied association, we derived a genetic analysis cohort by further making exclusions based on the following criteria: (1) non-Caucasian ( $n=59,248$ ); (2) did not pass the quality control of genetic data (ie, with inconsistent selfreported and genetic sex, or high rate of genotype missingness and heterozygosity; $n=1037$ ); and (3) with first or second level of relatedness (ie, with the kinship coefficient $>0.0884 ; n=44,278$ ) [Figure 1]. ${ }^{[23]}$

## Assessment of coffee consumption

At the recruitment assessment center, participants completed a food frequency questionnaire (FFQ) that included 29 questions on diet. For habitual coffee consumption, participants were first asked "How many cups of coffee do you drink each day? (Include decaffeinated coffee)." We defined coffee intake as follows: none, $\leq 1,2-3,4-5$, and $\geq 6$ cups/day. Coffee drinkers would be further asked "What type of coffee do you usually drink?," and they could choose from "decaffeinated" "instant" "ground" "other type" "do not know" and "prefer not to answer".

The reproducibility of the FFQ in a subsample of around 20,000 participants who repeated the visit 4 years after recruitment and its concordance with post-recruitment online 24-h recall have been described elsewhere. ${ }^{[24]}$ The weighted kappa of reported coffee intake in FFQs 4 years apart was 0.83 and the ability of FFQ to discriminate between high and low intakes was confirmed by $24-\mathrm{h}$ recall.

## Assessment of outcomes

CKD outcomes were identified as follows: (1) first diagnosis of incident chronic renal failure, initiation of renal replacement therapy, development of renal complications of hypertension or diabetes, glomerular diseases, or other renal structural abnormality $>3$ months after recruitment; (2) presenting with decreased kidney function, defined by eGFR $<60 \mathrm{~mL} \cdot \mathrm{~min}^{-1} \cdot 1.73 \mathrm{~m}^{-2}$ based on $\mathrm{eGFR}_{\text {creat }}$ or eGFR creat-cys as appropriate, ${ }^{[25]}$ or uACR $\geq$ $30 \mathrm{mg} / \mathrm{g}$ in the follow-up assessment during 2012 to 2013. The time to event was determined by the first diagnosis records for incident CKD, or, December 31, 2012, which was set as the timing for ascertaining incident CKD by abnormal laboratory tests. Cases were obtained from records linked to inpatients, death register, and primary care, and classified using the International Classification of Diseases (ICD), 10th Revision (ICD-10) codes, and Office of Population Censuses and Surveys Classification of Interventions and Procedures [Supplementary Table 2, http://links.lww.com/CM9/B116].

## Assessment of covariates

Sociodemographic factors (ie, age, sex, race, Townsend deprivation index, assessment center, and the highest education level) and lifestyles (ie, smoking, alcohol


Figure 1: Flow chart of participants' enrollment. CKD: Chronic kidney disease; QC: Quality control; SHBG: Sex hormone-binding globulin.
consumption, and milk and tea intake) were collected at recruitment using questionnaires. Body mass index (BMI), as the anthropometric measurement, was calculated from participants' height and weight. Participants' history of hypertension, diabetes, CVD, and cancer was obtained from health-related data. The detailed description of covariates is available in Supplementary Methods, http://links.lww.com/CM9/B116. Serum sex hormones and SHBG were collected during 2006 to 2010 at the recruitment assessment centers and measured by Beckman Coulter Unicel Dxl 800 (Beckman Coulter, London, UK).

## Polygenetic risk score of CKD

The polygenic risk scores (PRS) of CKD were derived from the summary statistics of a recent genome-wide association study of eGFR. ${ }^{[26]}$ Independent single nucleotide polymorphisms (SNPs) were identified using the "clumping" method. ${ }^{[27]}$ Aggregating the numbers of risk allele at each locus, weighted by the corresponding beta coefficient, we constructed PRS and then $z$-standardized them. To optimize the capability of risk prediction, multiple $P$ value thresholds $\left(5 \times 10^{-8}, 5 \times 10^{-6}, 5 \times 10^{-4}, 0.05\right.$, $0.01,0.1,0.5$ ) were examined. PRS calculated from SNPs with $P$ value $<0.01$ finally came out to be the best score carried forward, explaining $4.53 \%$ of the variance of eGFR in the UK Biobank [Supplementary Table 3, http:// links.lww.com/CM9/B116]. A higher PRS represented a better genetically predicted kidney function. The hazard ratios (HRs) of CKD were, respectively, 1.33 ( $95 \%$ confidence interval [CI] 1.23-1.43, $P$ value $<0.001$ ) and 1.79 ( $95 \%$ CI $1.67-1.92, P$ value $<0.001$ ) for the intermediate (ie, second tertile of PRS) and high-risk (ie, first tertile of PRS) groups compared with the low-risk group (ie, third tertile of PRS), adjusting for age, sex, third
degree of relatedness, first ten principle components, and genotyping arrays.

## Genetic polymorphisms of caffeine metabolism

We chose SNPs near AHR, CYP1A2, and CYP2A6 reported in the genome-wide association study of caffeine metabolites conducted by Cornelis et al ${ }^{[28]}$ to construct the allele score [Supplementary Table 4, http://links.lww.com/ CM9/B116]. The score correlated positively with caffeine metabolizing rate. Coffee consumption reduced by 0.039 ( $95 \%$ CI $0.031-0.047, P$ value $<0.001$ ) cups/day for every 1 -standard-deviation (SD) increase in the allele score. Faster caffeine metabolizers (ie, higher half of the score) were $6 \%$ (odds ratio $1.06,95 \%$ CI $1.04-1.08, P$ value $<0.001$ ) more likely to become heavy coffee drinkers (ie, $\geq 4$ cups/day) compared with slower ones (ie, lower half of the score).

## Statistical analysis

The distributions of baseline characteristics were presented across eGFR categories. Continuous variables were shown as means SDs if normally distributed and medians (interquartile ranges) if skewed. Categorical variables were displayed as count (\%). We compared continuous variables using analysis of variance or Kruskal-Wallis test, as appropriate, and categorical variables using chisquared test.

The end of follow-up was recorded as the date of CKD incidence, the date of death, or the end of data collection of the attended assessment center (i., February 28, 2018 for centers in England and Wales; December 31, 2016 for centers in Scotland), whichever came first. A Cox proportional hazards model was applied using the "survival" package in R (The R Foundation for Statistical Computing, Vienna, Austria) to calculate HRs and $95 \%$ CIs of coffee consumption and CKD, stratified by 5 -year age groups, sex, and assessment centers. The proportional hazards assumption for the Cox model was checked using Schoenfeld residuals, and no violation was found.

Coffee consumption, measured by the number of consumed coffee cups with non-drinkers as the reference, was first introduced into the model as a multi-categorical variable. Then, a test for linearity was performed by modeling coffee as a continuous variable. To investigate the extent of confounding, we adjusted for sociodemographic factors (Townsend deprivation index [in quartiles] and highest education level), lifestyle (smoking [never, past, $<1,1-9,10-14,15-19$, and $\geq 20$ cigarettes/ day], alcohol consumption [never, past, $<1,1-7,8-15$, $16-29$, and $\geq 30 \mathrm{~g} /$ day], milk intake [none, $<150,150-$ 299 , and $\geq 300 \mathrm{~mL} /$ day], and tea intake [none, $\leq 1,2-3$, and $\geq 3$ cups/day]), anthropometric measurement (BMI $\left[<18.5, \quad 18.5-24.9, \quad 25.0-29.9\right.$, and $\left.\geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}\right]$ ), comorbidities (history or comorbidities of hypertension, diabetes, CVD, and cancer), and sex hormones (logtransformed SHBG and testosterone). In the genetic analysis cohort, genetic risk of CKD (low, intermediate, and high risk defined by PRS) and caffeine metabolizing
rate (fast and slow metabolizers defined by allele score) was further adjusted. Participants' missing variables were grouped into a single category and the proportion of missing observations was $<1 \%$ for all covariates [Supplementary Table 5, http://links.lww.com/CM9/ B116].

We evaluated the associations between coffee intake and CKD for various coffee types, and non-coffee drinkers were treated as the reference group; further, coffee drinkers preferring a certain type of coffee were included in each subgroup analysis. To find other potential modifiers on the coffee-CKD association, we also performed stratified analyses by age, sex, Townsend deprivation index, smoking status, alcohol consumption, BMI, prevalent hypertension, diabetes, CVD and cancer, genetic risk of CKD, caffeine metabolizing rate, sex hormones, and SHBG. Since estradiol was not routinely measured in the UK Biobank, the subgroup analysis was restricted to 44,921 females with available assay results. The $P$ value of heterogeneity corresponds to the likeli-hood-ratio test comparing the models with and without the interaction terms.

## Sensitivity analysis

We used the following sensitivity analyses to test the robustness of the results: (1) performing Bayesian model averaging (BMA) using the " $B A S$ " package in R to verify the sex-specific coffee-CKD association, which is based on specific priors, to generate posterior distributions of candidate effect sizes of variables under each of the models selected ${ }^{[29]}$; (2) excluding participants with unavailable baseline eGFR and uACR since we could not rule out the possibility that they had prevalent CKD; and (3) excluding incident CKD cases within the first 2 and 3 years to reduce reverse causation.

Analyses were done using R version 4.0.2 (The R Foundation for Statistical Computing, Vienna, Austria). A two-tailed $P$ value $<0.05$ was interpreted as statistically significant.

## Ethics approval and consent to participate

All the UK Biobank participants gave written informed consent before data collection. The UK Biobank has full ethical approval from the NHS National Research Ethics Service (16/NW/0274), and this study was approved by the biomedical research ethics committee of West China Hospital (2019-1171). The study conformed to the Declaration of Helsinki.

## Results

## Coffee consumption and incident CKD

Of the 359,906 UK Biobank participants enrolled in the current study, the median age was 57 years, $90.8 \%$ were Caucasians, $49.8 \%$ were females, and $78.4 \%$ were coffee drinkers. Participants' characteristics by sex and coffee intake are presented in Table 1. Heavy coffee drinkers (ie, $\geq 4$ cups/day) were more likely to be males, fast caffeine
metabolizers, obese, current smokers, and alcohol drinkers. Over a median follow-up period of 8.8 years, $3454(1.9 \%)$ cases of incident CKD in females and 3800 $(2.1 \%)$ in males were observed. Regular coffee consumption was associated with a $6 \%$ to $15 \%$ reduced risk of CKD. The adjusted HRs across coffee intake varied in a dose-dependent manner ( $P$ value $<0.001$ for trend; Table 2). The coffee-CKD association did not differ by coffee types [Supplementary Table 6, http://links.lww. com/CM9/B116]. In the subset of 255,343 participants with qualified genetic data, after the additional adjustment for the genetic risk of CKD and caffeine metabolizing rate, the results remain the same [Table 2].

## Subgroup analyses and sex-specific coffee-CKD association

Habitual coffee consumption could offset the genetic risk of CKD. Compared with non-drinkers, coffee consumption reduced the risk of CKD by $6 \%$ to $17 \%$ [Supplementary Table 7, http://links.lww.com/CM9/B116]. The inverse coffee-CKD association seemed stronger among slower caffeine metabolizers than faster ones. However, the formal test of interaction did not reach statistical significance ( $P$ value for interaction $=0.14$; Supplementary Table 7, http://links.lww.com/CM9/B116). The coffee-CKD association did not significantly differ by age, Townsend deprivation index, smoking status, alcohol consumption, BMI, prevalent hypertension, diabetes, CVD, and cancer [Supplementary Tables 7 and 8, http://links.lww.com/CM9/B116].

Stratified by sex, the inverse coffee-CKD relationship existed in females, but not males ( $P$ value for interaction $=$ 0.013 ; Table 3). Observing the sex-specific association between coffee consumption and CKD, we further explored the possible modification effect of sex hormones and SHBG, and found that the coffee-CKD association was more obvious in participants with lower testosterone and higher SHBG concentrations.

To be specific, in the general population, coffee intake brought about a $12 \%$ to $30 \%$ decreased CKD risk in the lowest tertile of testosterone concentration. However, such an inverse coffee-CKD association became less evident as testosterone increased, and eventually disappeared in the highest tertile ( $P$ value for interaction $=$ 0.031; Figure 2). Similarly, results diverged in different strata of plasma SHBG concentration. The coffee-CKD association was greatest in the highest tertile, while could not be noticed as SHBG fell down ( $P$ value for interaction $=0.057$; Figure 3). However, the available assay results indicate that estradiol did not significantly modify the reno-protective effect of coffee in females with available assay results ( $P$ value for interaction $=0.96$; Supplementary Figure 1, http://links.lww.com/CM9/B116). Risk patterns across testosterone and SHBG subgroups were generally similar in both sexes, although no estimate could be derived from the lowest and the highest tertiles of testosterone, in males and females, respectively, due to inadequate sample sizes [Figures 2 and 3]. It is notable that a weak tendency of inverse coffee-CKD association was found in males with the highest SHBG level. In this subgroup, compared with non-drinkers, drinking $\geq 4$ cups/
Table 1: Baseline characteristics of study participants stratified by sex and coffee intake in the UK Biobank.

| Characteristics | Coffee intake in the female (cups/day) |  |  |  |  |  | Coffee intake in the male (cups/day) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | None | $\leq 1$ | 2-3 | 4-5 | $\geq 6$ | $P$ value | None | $\leq 1$ | 2-3 | 4-5 | $\geq 6$ | $P$ value |
| No. of participants | 41,455 | 50,470 | 55,288 | 22,328 | 9586 |  | 36,315 | 46,761 | 57,514 | 27,101 | 13,088 |  |
| Age (years) | 54 [48, 61] | 57 [50, 63] | $58[50,63]$ | 57 [50, 63] | 56 [49, 62] | $<0.001$ | 57 [49, 63] | $59[51,64]$ | $59[51,64]$ | 57 [49, 63] | $56[48,62]$ | <0.001 |
| Townsend deprivation index |  |  |  |  |  |  |  |  |  |  |  |  |
| Q1 | 9029 (21.8) | 12,784 (25.3) | 14,847 (26.9) | 5779 (25.9) | 2046 (21.3) | $<0.001$ | 7640 (21.0) | 12,003 (25.7) | 15,533 (27.0) | 7197 (26.6) | 3024 (23.1) | $<0.001$ |
| Q2 | 9747 (23.5) | 12,795 (25.4) | 14,425 (26.1) | 5850 (26.2) | 2323 (24.2) |  | 8086 (22.3) | 11,604 (24.8) | 14,799 (25.7) | 7133 (26.3) | 3133 (23.9) |  |
| Q3 | 10,535 (25.4) | 12,777 (25.3) | 13,838 (25.0) | 5648 (25.3) | 2507 (26.2) |  | 9016 (24.8) | 11,707 (25.0) | 14,160 (24.6) | 6463 (23.8) | 3171 (24.2) |  |
| Q4 | 12,093 (29.2) | 12,052 (23.9) | 12,114 (21.9) | 5028 (22.5) | 2705 (28.2) |  | 11,522 (31.7) | 11,383 (24.3) | 12,947 (22.5) | 6275 (23.2) | 3747 (28.6) |  |
| Race |  |  |  |  |  |  |  |  |  |  |  |  |
| White | 37,487 (90.4) | 47,429 (94.0) | 53,428 (96.6) | 21,871 (98.0) | 9415 (98.2) | $<0.001$ | 32,695 (90.0) | 43,707 (93.5) | 55,285 (96.1) | 26,379 (97.3) | 12,785 (97.7) | <0.001 |
| Mixed | 345 (0.8) | 377 (0.7) | 331 (0.6) | 101 (0.5) | 45 (0.5) |  | 203 (0.6) | 257 (0.5) | 259 (0.5) | 110 (0.4) | 56 (0.4) |  |
| Asian | 1355 (3.3) | 830 (1.6) | 390 (0.7) | 83 (0.4) | 28 (0.3) |  | 1713 (4.7) | 1244 (2.7) | 685 (1.2) | 159 (0.6) | 53 (0.4) |  |
| Black | 1303 (3.1) | 934 (1.9) | 492 (0.9) | 104 (0.5) | 35 (0.4) |  | 969 (2.7) | 795 (1.7) | 553 (1.0) | 173 (0.6) | 54 (0.4) |  |
| Chinese | 285 (0.7) | 260 (0.5) | 119 (0.2) | 15 (0.1) | 6 (0.1) |  | 157 (0.4) | 149 (0.3) | 104 (0.2) | 30 (0.1) | 7 (0.1) |  |
| Other | 542 (1.3) | 517 (1.0) | 395 (0.7) | 92 (0.4) | 42 (0.4) |  | 418 (1.2) | 417 (0.9) | 415 (0.7) | 133 (0.5) | 77 (0.6) |  |
| Highest education level |  |  |  |  |  |  |  |  |  |  |  |  |
| None | 7664 (18.5) | 7640 (15.1) | 7785 (14.1) | 3504 (15.7) | 1809 (18.9) | $<0.001$ | 8245 (22.7) | 7483 (16.0) | 7880 (13.7) | 3925 (14.5) | 2350 (18.0) | $<0.001$ |
| College/university | 11,541 (27.8) | 17,069 (33.8) | 19,317 (34.9) | 7035 (31.5) | 2455 (25.6) |  | 9698 (26.7) | 16,451 (35.2) | 22,295 (38.8) | 9845 (36.3) | 3823 (29.2) |  |
| A/AS | 3568 (8.6) | 4878 (9.7) | 5424 (9.8) | 2116 (9.5) | 861 (9.0) |  | 2573 (7.1) | 3670 (7.8) | 4528 (7.9) | 1955 (7.2) | 934 (7.1) |  |
| O/GCSEs | 7800 (18.8) | 9569 (19.0) | 10,713 (19.4) | 4330 (19.4) | 1854 (19.3) |  | 4844 (13.3) | 6146 (13.1) | 7348 (12.8) | 3477 (12.8) | 1748 (13.4) |  |
| CSEs | 2244 (5.4) | 1952 (3.9) | 2217 (4.0) | 1036 (4.6) | 509 (5.3) |  | 1819 (5.0) | 1600 (3.4) | 2010 (3.5) | 1104 (4.1) | 687 (5.2) |  |
| NVQ/HND/HNC | 5943 (14.3) | 6035 (12.0) | 6253 (11.3) | 2774 (12.4) | 1465 (15.3) |  | 7006 (19.3) | 8804 (18.8) | 10,301 (17.9) | 5423 (20.0) | 2854 (21.8) |  |
| Other | 2157 (5.2) | 2867 (5.7) | 3135 (5.7) | 1365 (6.1) | 557 (5.8) |  | 1614 (4.4) | 2098 (4.5) | 2622 (4.6) | 1127 (4.2) | 568 (4.3) |  |
| Smoking |  |  |  |  |  |  |  |  |  |  |  |  |
| None | 26,667 (64.3) | 32,007 (63.4) | 32,889 (59.5) | 11,987 (53.7) | 4172 (43.5) | <0.001 | 19,281 (53.1) | 24,491 (52.4) | 29,028 (50.5) | 12,415 (45.8) | 4760 (36.4) | $<0.001$ |
| Previous | 12,042 (29.0) | 16,278 (32.3) | 19,154 (34.6) | 79,76 (35.7) | 3354 (35.0) |  | 13,843 (38.1) | 19,528 (41.8) | 24,443 (42.5) | 11,633 (42.9) | 5373 (41.1) |  |
| <10 cigarettes/day | 616 (1.5) | 653 (1.3) | 1009 (1.8) | 534 (2.4) | 244 (2.5) |  | 452 (1.2) | 533 (1.1) | 760 (1.3) | 415 (1.5) | 220 (1.7) |  |
| 10-14 cigarettes/day | 668 (1.6) | 544 (1.1) | 807 (1.5) | 566 (2.5) | 445 (4.6) |  | 595 (1.6) | 568 (1.2) | 842 (1.5) | 596 (2.2) | 438 (3.3) |  |
| 15-19 cigarettes/day | 598 (1.4) | 436 (0.9) | 649 (1.2) | 576 (2.6) | 522 (5.4) |  | 712 (2.0) | 649 (1.4) | 965 (1.7) | 745 (2.7) | 650 (5.0) |  |
| $\geq 20$ cigarettes/day | 857 (2.1) | 537 (1.1) | 761 (1.4) | 677 (3.0) | 843 (8.8) |  | 1422 (3.9) | 981 (2.1) | 1458 (2.5) | 1289 (4.8) | 1643 (12.6) |  |
| Alcohol consumption |  |  |  |  |  |  |  |  |  |  |  |  |
| None | 5725 (13.8) | 3551 (7.0) | 2981 (5.4) | 1299 (5.8) | 740 (7.7) | $<0.001$ | 2705 (7.4) | 1549 (3.3) | 1350 (2.3) | 578 (2.1) | 377 (2.9) | <0.001 |
| Previous | 2170 (5.2) | 1297 (2.6) | 1261 (2.3) | 677 (3.0) | 499 (5.2) |  | 2028 (5.6) | 1194 (2.6) | 1372 (2.4) | 882 (3.3) | 670 (5.1) |  |
| $<1$ grams/day | 6378 (15.4) | 5880 (11.7) | 5251 (9.5) | 2344 (10.5) | 1383 (14.4) |  | 2791 (7.7) | 2469 (5.3) | 2403 (4.2) | 1219 (4.5) | 911 (7.0) |  |
| $1-7$ grams/day | 10,684 (25.8) | 14,085 (27.9) | 13,606 (24.6) | 5341 (23.9) | 2326 (24.3) |  | 6147 (16.9) | 7271 (15.5) | 8039 (14.0) | 3770 (13.9) | 2028 (15.5) |  |
| 8-15 grams/day | 8235 (19.9) | 13,206 (26.2) | 15,690 (28.4) | 5924 (26.5) | 2041 (21.3) |  | 5454 (15.0) | 9101 (19.5) | 10,999 (19.1) | 4798 (17.7) | 2037 (15.6) |  |
| 16-30 grams/day | 5531 (13.3) | 8918 (17.7) | 11,783 (21.3) | 4664 (20.9) | 1743 (18.2) |  | 7765 (21.4) | 12,402 (26.5) | 16,812 (29.2) | 7625 (28.1) | 3185 (24.3) |  |
| $\geq 30$ grams/day | 2674 (6.5) | 3497 (6.9) | 4678 (8.5) | 2056 (9.2) | 838 (8.7) |  | 9357 (25.8) | 12,723 (27.2) | 16,495 (28.7) | 8202 (30.3) | 3856 (29.5) |  |
| Tea intake |  |  |  |  |  |  |  |  |  |  |  |  |
| None | 4758 (11.5) | 2755 (5.5) | 7791 (14.1) | 7431 (33.3) | 4845 (50.5) | $<0.001$ | 3352 (9.2) | 2053 (4.4) | 7026 (12.2) | 7434 (27.4) | 5500 (42.0) | <0.001 |
| $\leq 1$ cups/day | 2225 (5.4) | 4319 (8.6) | 7411 (13.4) | 4352 (19.5) | 1740 (18.2) |  | 1841 (5.1) | 4386 (9.4) | 8149 (14.2) | 5277 (19.5) | 2380 (18.2) |  |
| $2-3$ cups/day | 9107 (22.0) | 15,281 (30.3) | 21,083 (38.1) | 5715 (25.6) | 1345 (14.0) |  | 7890 (21.7) | 13,810 (29.5) | 21,800 (37.9) | 7380 (27.2) | 2399 (18.3) |  |
| $\geq 3$ cups/day | 25,263 (60.9) | 28,033 (55.5) | 18,951 (34.3) | 4809 (21.5) | 1637 (17.1) |  | 23,150 (63.7) | 26,463 (56.6) | 20,491 (35.6) | 6987 (25.8) | 2781 (21.2) |  |
| BMI |  |  |  |  |  |  |  |  |  |  |  |  |
| $<18.5 \mathrm{~kg} / \mathrm{m}^{2}$ | 276 (0.7) | 343 (0.7) | 351 (0.6) | 117 (0.5) | 71 (0.7) | <0.001 | 119 (0.3) | 95 (0.2) | 93 (0.2) | 49 (0.2) | 42 (0.3) | $<0.001$ |
| $18.5-24.9 \mathrm{~kg} / \mathrm{m}^{2}$ | 15,630 (37.7) | 21,244 (42.1) | 21,807 (39.4) | 7543 (33.8) | 2996 (31.3) |  | 9350 (25.7) | 12,961 (27.7) | 14,858 (25.8) | 5908 (21.8) | 2781 (21.2) |  |
| 25.0-29.9 | 14,860 (35.8) | 18,241 (36.1) | 20,817 (37.7) | 8754 (39.2) | 3582 (37.4) |  | 17,584 (48.4) | 23,301 (49.8) | 29,169 (50.7) | 13,670 (50.4) | 6323 (48.3) |  |
| $\geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$ | 10,504 (25.3) | 10,484 (20.8) | 12,168 (22.0) | 5870 (26.3) | 2905 (30.3) |  | 9033 (24.9) | 10,207 (21.8) | 13,199 (22.9) | 7393 (27.3) | 3896 (29.8) |  |
| History of hypertension |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 23,920 (57.7) | 28,562 (56.6) | 31,642 (57.2) | 12,677 (56.8) | 5666 (59.1) | $<0.001$ | 15,542 (42.8) | 20,211 (43.2) | 24,945 (43.4) | 11,896 (43.9) | 5929 (45.3) | $<0.001$ |
| Yes | 17,535 (42.3) | 21,908 (43.4) | 23,646 (42.8) | 9651 (43.2) | 3920 (40.9) |  | 20,773 (57.2) | 26,550 (56.8) | 32,569 (56.6) | 15,205 (56.1) | 7159 (54.7) |  |
| History of diabetes |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 39,639 (95.6) | 48,670 (96.4) | 53,409 (96.6) | 21,508 (96.3) | 9192 (95.9) | $<0.001$ | 33,577 (92.5) | 43,546 (93.1) | 53,907 (93.7) | 25,241 (93.1) | 12,097 (92.4) | <0.001 |
| Yes | 1816 (4.4) | 1800 (3.6) | 1879 (3.4) | 820 (3.7) | 394 (4.1) |  | 2738 (7.5) | 3215 (6.9) | 3607 (6.3) | 1860 (6.9) | 991 (7.6) |  |


| Characteristics | Coffee intake in the female (cups/day) |  |  |  |  |  | Coffee intake in the male (cups/day) |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | None | $\leq 1$ | 2-3 | 4-5 | $\geq 6$ | $P$ value | None | $\leq 1$ | 2-3 | 4-5 | $\geq 6$ | $P$ value |
| History of CVD |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 39,562 (95.4) | 48,498 (96.1) | 53,230 (96.3) | 21,434 (96.0) | 9092 (94.8) | <0.001 | 32,224 (88.7) | 42,027 (89.9) | 52,239 (90.8) | 24,757 (91.4) | 11,781 (90.0) | <0.001 |
| Yes | 1893 (4.6) | 1972 (3.9) | 2058 (3.7) | 894 (4.0) | 494 (5.2) |  | 4091 (11.3) | 4734 (10.1) | 5275 (9.2) | 2344 (8.6) | 1307 (10.0) |  |
| History of cancer |  |  |  |  |  |  |  |  |  |  |  |  |
| No | 37,073 (89.4) | 44,639 (88.4) | 48,841 (88.3) | 19,884 (89.1) | 8497 (88.6) | <0.001 | 33,446 (92.1) | 42,658 (91.2) | 52,620 (91.5) | 25,075 (92.5) | 12,138 (92.7) | <0.001 |
| Yes | 4382 (10.6) | 5831 (11.6) | 6447 (11.7) | 2444 (10.9) | 1089 (11.4) |  | 2869 (7.9) | 4103 (8.8) | 4894 (8.5) | 2026 (7.5) | 950 (7.3) |  |
| Testosterone ( $\mathrm{nmol} / \mathrm{L}$ ) | $\begin{gathered} 1.02 \\ {[0.72,1.38]} \end{gathered}$ | $\begin{gathered} 1.00 \\ {[0.71,1.36]} \end{gathered}$ | $\begin{gathered} 1.02 \\ {[0.72,1.37]} \end{gathered}$ | $\begin{gathered} 1.03 \\ {[0.74,1.38]} \end{gathered}$ | $\begin{gathered} 1.04 \\ {[0.75,1.41]} \end{gathered}$ | <0.001 | $\begin{gathered} 11.61 \\ {[9.40,14.15]} \end{gathered}$ | $\begin{gathered} 11.61 \\ {[9.47,14.11]} \end{gathered}$ | 11.70 [9.53, 14.18] | $\begin{gathered} 11.70 \\ {[9.50,14.18]} \end{gathered}$ | 11.89 [9.60, 14.46] | $<0.001$ |
| Sex hormone-binding globulin ( $\mathrm{nmol} / \mathrm{L}$ ) | $\begin{gathered} 55.97 \\ {[39.26,77.38]} \end{gathered}$ | $\begin{gathered} 56.98 \\ {[40.68,77.20]} \end{gathered}$ | $\begin{gathered} 56.85 \\ {[41.04,76.91]} \end{gathered}$ | $\begin{gathered} 56.57 \\ {[40.60,76.63]} \end{gathered}$ | $\begin{gathered} 58.20 \\ {[41.17,79.76]} \end{gathered}$ | <0.001 | $\begin{gathered} 36.23 \\ {[27.21,47.55]} \end{gathered}$ | $\begin{gathered} 37.06 \\ {[28.18,48.12]} \end{gathered}$ | $\begin{gathered} 37.40 \\ {[28.39,48.47]} \end{gathered}$ | $\begin{gathered} 37.12 \\ {[28.07,48.28]} \end{gathered}$ | $\begin{gathered} 37.55 \\ {[28.27,49.11]} \end{gathered}$ | $<0.001$ |
| Genetic risk of CKD* |  |  |  |  |  |  |  |  |  |  |  |  |
| Low | 8836 (32.5) | 11,527 (33.2) | 13,163 (33.4) | 5403 (32.9) | 2291 (32.8) | 0.007 | 8092 (33.1) | 11,191 (33.4) | 14,039 (33.0) | 6633 (32.6) | 3088 (31.5) | 0.016 |
| Intermediate | 9212 (33.9) | 11,742 (33.9) | 13,503 (34.3) | 5547 (33.8) | 2462 (35.3) |  | 8168 (33.4) | 11,405 (34.0) | 14,451 (34.0) | 6957 (34.2) | 3370 (34.4) |  |
| High | 9160 (33.7) | 11,412 (32.9) | 12,719 (32.3) | 5451 (33.2) | 2231 (31.9) |  | 8184 (33.5) | 10,948 (32.6) | 14,042 (33.0) | 6766 (33.2) | 3350 (34.2) |  |
| Allele score of caffeine metabolism | $-3.14 \pm 2.23$ | $-3.13 \pm 2.23$ | $-3.08 \pm 2.22$ | $-3.03 \pm 2.21$ | $-2.94 \pm 2.17$ | <0.001 | $-3.11 \pm 2.25$ | $-3.12 \pm 2.24$ | $-3.08 \pm 2.24$ | $-3.04 \pm 2.23$ | $-3.01 \pm 2.19$ | <0.001 |


 Diploma; HTN: Hypertension; IQR: Interquartile range; NVQ: National Vocational Qualification; SD: Standard deviation.
day of coffee reduced the risk of CKD by $15 \%$ (HR 0.85 , 95\% CI 0.64-1.13).

In the genetic analysis cohort, after additionally adjusting for allele scores, results of subgroup analysis by sex hormones and SHBG remain similar [Supplementary Figures 2-4, http://links.lww.com/CM9/B116].

## Sensitivity analyses

In BMA, coffee drinking was inversely related to CKD in the general population ( $\beta=-0.074, \mathrm{SD}=0.067$ ) and had a $58.3 \%$ posterior probability [Supplementary Table 9 and Supplementary Figure 5, http://links.lww.com/CM9/ B116]. Posterior probabilities of regular coffee intake were $0 \%$ and $99.4 \%$ for males and females, respectively [Supplementary Tables 10 and 11 and Supplementary Figures 7 and 8, http://links.lww.com/CM9/B116]. Further, as indicated in Supplementary Tables 12-14, http:// links.lww.com/CM9/B116, no major change could be obtained in the result pursuant to performing the same analysis in the advanced-age group, restricting it to participants with available baseline kidney biomarker results, or excluding CKD cases diagnosed within two years and three years after recruitment.

## Discussion

Among more than 350,000 participants in the UK Biobank, coffee consumption reduced the risk of CKD regardless of the genetic risk of CKD, but possibly depending partly on the caffeine metabolizing rate. The effect was sex-specific, and was modified by testosterone and SHBG.

Our finding adds to the growing evidence on the possible effect of coffee consumption on CKD. The results of the current work were in line with previous observational studies and meta-analyses indicating significant association between coffee consumption and the reduced risk of CKD in the general population, ${ }^{[6-8,10]}$ as well as reports about the causal effect of coffee on kidney function based on Mendelian randomization. ${ }^{[9]}$ The hypothesis that bioactive components of coffee, such as caffeine and chlorogenic acids, have a beneficial impact on health outcomes through multiple interconnected pathways, including insulin sensitivity improvement, sex hormone production, and inflammation reduction, can be mentioned as a plausible causative factor for the inverse coffee-CKD association observed in the present study and elsewhere in the literature. ${ }^{[30]}$ In addition, our study further extends the existing findings by addressing a paramount question: that of ascertaining the populations to which this association could be generalized.

First, it merits attention that, in the current study, only females could benefit from coffee drinking. Other investigators have also assessed the potential difference in the coffee-CKD association stratified by sex, yielding conflicting results. Hu et al ${ }^{[6]}$ reported a null finding of the coffeeCKD association in males, but failed to confirm the interaction of coffee and sex. However, Lew et al ${ }^{[7]}$ found that coffee consumption could only reduce the risk of endstage renal disease in males. Compared with previous ones,

Table 2: Hazard ratios of coffee intake and incident CKD in the UK Biobank.
Coffee intake (cups/day)

| Items | None | $\leq 1$ | 2-3 | 4-5 | $\geq 6$ | $P$ value for trend |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Primary analysis cohort ( $n=359,906$ ) |  |  |  |  |  |  |
| No. of cases | 1719 | 2010 | 2155 | 925 | 445 |  |
| Person years at risk | 668,466 | 834,613 | 968,851 | 425,409 | 195,232 |  |
| Cases per 1000 person-years | 2.57 | 2.41 | 2.22 | 2.18 | 2.28 |  |
| Model $1^{*}$ | 1 (ref) | $0.82(0.77-0.88){ }^{\text {II }}$ | $0.75(0.71-0.8)^{\text {II }}$ | 0.77 (0.71-0.83) ${ }^{\text {II }}$ | 0.84 (0.76-0.93) ${ }^{\text {II }}$ | $<0.001$ |
| Model $2^{\dagger}$ | 1 (ref) | 0.94 (0.88-1) | 0.89 (0.83-0.95) ${ }^{\text {\|\| }}$ | 0.86 (0.79-0.94) ${ }^{\text {II }}$ | 0.84 (0.75-0.95) ${ }^{\text {I }}$ | <0.001 |
| Model $3^{\ddagger}$ | 1 (ref) | 0.94 (0.88-1) | 0.89 (0.83-0.95) ${ }^{\text {II }}$ | 0.86 (0.79-0.94) ${ }^{\text {\|\| }}$ | 0.85 (0.75-0.95) ${ }^{\text {I }}$ | $<0.001$ |
| Genetic analysis cohort ( $n=255,343$ ) |  |  |  |  |  |  |
| No. of cases | 1133 | 1461 | 1579 | 713 | 316 |  |
| Person years at risk | 455,622 | 586,948 | 704,140 | 315,849 | 144,412 |  |
| Cases per 1000 person-years | 2.54 | 2.49 | 2.24 | 2.26 | 2.19 |  |
| Model $1^{*}$ | 1 (ref) | 0.86 (0.8-0.93) ${ }^{\text {II }}$ | $0.78(0.72-0.84)^{\text {II }}$ | $0.82(0.75-0.9)^{\text {II }}$ | 0.83 (0.73-0.94) ${ }^{\text {\| }}$ | <0.001 |
| Model $2^{\dagger}$ | 1 (ref) | 0.97 (0.9-1.05) | 0.89 (0.82-0.96) ${ }^{\text {II }}$ | 0.88 (0.79-0.97) ${ }^{\text {§ }}$ | 0.79 (0.69-0.92) ${ }^{\text {I }}$ | <0.001 |
| Model $3^{\ddagger}$ | 1 (ref) | 0.97 (0.9-1.05) | $0.89(0.82-0.96)^{\text {II }}$ | 0.88 (0.79-0.98) ${ }^{\text {8 }}$ | 0.80 (0.69-0.92) ${ }^{\text {I }}$ | $<0.001$ |

[^1]one strength of our study lies in the prospective design and large scale of the UK Biobank, providing us with a unique opportunity to investigate the sex-specific effect of coffee on CKD with greater statistical power and less bias. Especially, in addition to the Cox proportional hazard model with comprehensive adjustment for confounders, the sex disparity was also verified by BMA, which generally performs better than traditional statistical methods in variable selection as it evaluates all potential combinations of the candidate variables and avoids uncertainty. ${ }^{[29]}$ In BMA, coffee was inversely associated with renal deficiency in the general population and had a posterior probability of around $95 \%$, indicating positive evidence for its renoprotective effect. ${ }^{[31]}$ Performing BMA by sex, we found extremely strong evidence, measured by posterior probability $>99 \%$, for the reno-protective effect of coffee in females, in contrast to that of $2.5 \%$ in males.

Second, with the in-depth genetic information provided by the UK Biobank, the genetic predisposition of CKD and the genetically predicted caffeine metabolizing rate could be measured as allele scores at the individual level; thus, we can evaluate the coffee-CKD association in subgroups of varied genetic predisposition. While genetic risk for CKD seemingly did not modify the beneficial effects of coffee consumption on incident CKD, the caffeine metabolizing rate might partly influence the coffeeCKD association. Previous research reports a modification effect of the CYP1A2 genotype, which is responsible primarily for metabolizing caffeine, on some coffee-health outcome associations, such that slow metabolizers could
not benefit from habitual coffee consumption. ${ }^{[17-19]}$ Nonetheless, constructing a comprehensive allele score representing caffeine metabolism, we surprisingly found a stronger inverse association in slow metabolizers, but failed to confirm the gene-diet interaction.

In the current study, we observed sex-specific association between coffee and CKD. Sex disparity in the pathogenesis of CKD is well acknowledged. Both animal and epidemiological studies have revealed that sex hormones largely contribute to the phenomenon. Previous Mendelian randomization analyses have also reported the casual role of sex hormones in the incidence and progression of CKD, especially in males. ${ }^{[32,33]}$ Therefore, with available individual-level data in the UK Biobank, we performed a series of analyses to explore whether sex hormones and SHBG may be involved this gender difference. We observed their potential modification roles in the cof-fee-CKD association.

For instance, in females, the reno-protective effect of coffee was more robust in those with higher SHBG and lower testosterone concentrations. SHBG per se also relates to metabolic syndrome, including dyslipidemia, hypertension, dysregulated glucose homeostasis, and obesity, ${ }^{[34]}$ which are all confirmed risk factors of CKD. ${ }^{[35]}$ Acting as a transporting protein, SHBG binds sex hormones with certain affinity and acts as a modulator of their bioactivity. Previous studies have revealed that the alteration of SHBG brings about a more drastic fluctuation in testosterone than estradiol, and low SHBG is often associated with hyper-
Table 3: Sex-specific association of coffee and incident CKD in the whole sample and genetic analysis cohort in the UK Biobank.

| Variables | Coffee intake (cups/day) |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | None | $\leq 1$ | 2-3 | 4-5 | $\geq 6$ | $P$-value for trend | $P$-value for interaction ${ }^{\text {8 }}$ |
| Primary analysis cohort |  |  |  |  |  |  | 0.013 |
| Female ( $n=179,127$ ) |  |  |  |  |  |  |  |
| No. of cases | 903 | 983 | 989 | 392 | 187 |  |  |
| Person years at risk | 358,473 | 435,777 | 477,462 | 193,234 | 83,088 |  |  |
| Cases per 1000 person-years | 2.52 | $2.26$ | $2.07$ | $2.03$ | 2.25 |  |  |
| $\operatorname{HR}(95 \% \mathrm{CI})^{*}$ | 1 (ref) | $0.88(0.8-0.97)^{\dagger}$ | $0.81(0.73-0.89)^{\text {* }}$ | 0.73 (0.64-0.84) ${ }^{\text {\# }}$ | $0.72(0.6-0.86)^{\text {* }}$ | $<0.001$ |  |
| Male ( $n=180,779$ ) |  |  |  |  |  |  |  |
| No. of cases | 816 | 1027 | 1166 | 533 | 258 |  |  |
| Person years at risk | 309,993 | 398,836 | 491,390 | 231,581 | 112,144 |  |  |
| Cases per 1000 person-years | 2.63 | 2.57 | 2.37 | 2.30 | 2.30 |  |  |
| $\operatorname{HR}(95 \% \mathrm{CI})^{*}$ | 1 (ref) | 1 (0.9-1.1) | 0.97 (0.88-1.07) | 1 (0.88-1.13) | 0.98 (0.83-1.15) | 0.730 |  |
| Genetic analysis cohort |  |  |  |  |  |  | 0.018 |
| Female ( $n=124,659$ ) |  |  |  |  |  |  |  |
| No. of cases | 584 | 690 | 717 | 288 | 136 |  |  |
| Person years at risk | 236,293 | 300,248 | 340,293 | 142,000 | 60,457 |  |  |
| Cases per 1000 person-years | 2.47 | 2.30 | 2.11 | 2.03 | 2.25 |  |  |
| $\operatorname{HR}(95 \% \mathrm{CI})^{*}$ | 1 (ref) | 0.9 (0.81-1.01) | $0.82(0.73-0.92)^{\dagger}$ | $0.72(0.61-0.85)^{\text {* }}$ | 0.71 (0.57-0.80) ${ }^{\text {* }}$ | $<0.001$ |  |
| Male $(n=130,684)$ |  |  |  |  |  |  |  |
| No. of cases | 549 | 771 | 862 | 425 | 180 |  |  |
| Person years at risk | 209,236 | 286,700 | 363,847 | 173,850 | 83,955 |  |  |
| Cases per 1000 person-years | 2.62 | 2.69 | 2.37 | 2.44 | 2.14 |  |  |
| $\operatorname{HR}(95 \% \mathrm{CI})^{*}$ | 1 (ref) | 1.05 (0.94-1.17) | 0.96 (0.86-1.08) | 1.04 (0.9-1.2) | 0.89 (0.73-1.08) | 0.350 |  |






 day $\times$ sex are respectively $0.02,0.004,0.002$ and 0.11 in the primary analysis cohort and $0.019,0.026,0.001$ and 0.55 in the genetic analysis cohort.


Figure 2: Subgroup analysis of coffee intake and incident chronic kidney disease by testosterone level in the primary analysis cohort in the UK Biobank. *Values in parentheses represent the incidence rates of CKD expressed in cases per 1,000 person-years. ${ }^{\dagger} \mathrm{P}$ values of HRs are within the range of 0.01 to 0.05 . ${ }^{*} \mathrm{P}$ values of HRs are within the range of 0.001 to 0.01 . ${ }^{\S} P$ values of HRs are within the range of $<0.001$. "Model is stratified for 5 -year age groups, and 22 assessment centers and adjusted for sociodemographic factors (race [in the primary analysis cohort only], Townsend deprivation index [in quartiles], and highest education level [college or university degree, A levels/AS levels or equivalent, 0 levels/GCSEs or equivalent, CSEs or equivalent, NVQ or HND or HNC equivalent, or other professional qualifications]), lifestyle (smoking [never, past, $<1,1-10,10-14,15-19$, and $\geq 20$ cigarettes per day], alcohol consumption [never, past, $<1,1-7,8-15,16-29$ and $\geq 30$ grams/day], milk intake [none, $<150,150-299$ and $\geq 300 \mathrm{ml} /$ day], and tea intake [none, $1,2-3$ and $\geq 3 \mathrm{cups} /$ day]], anthropometric measurement (BMI [<18.5, 18.5-24.9, 25.0-29.9, and $\geq 30 \mathrm{~kg} / \mathrm{m} 2]$ ), comorbidities (history of hypertension, diabetes, cardiovascular disease, and cancer), and sex hormones (log-transformed SHBG and testosterone). "Coffee intake and testosterone are included in the interaction term as binary variables (ie, non-drinker vs. drinker and low level vs. intermediate + high level). Pvalue for interaction equals to 0.031 for coffee and testosterone. The analysis is not performed in the lowest and the highest tertile of testosterone, in males and females, respectively, due to inadequate sample sizes.


Figure 3: Subgroup analysis of coffee intake and incident CKD by SHBG level in the primary analysis cohort in the UK Biobank. *Values in parentheses represent the incidence rates of CKD expressed in cases per 1,000 person-years. ${ }^{\dagger} P$ values of HRs are within the range of 0.01 to 0.05 . ${ }^{\ddagger} P$ values of HRs are within the range of 0.001 to $0.01 .{ }^{\S} P$ values of HRs are within the range of $<0.001$. "Model is stratified for 5 -year age groups, and 22 assessment centers and adjusted for sociodemographic factors (race [in the primary analysis cohort only], Townsend deprivation index [in quartiles], and highest education level [college or university degree, A levels/AS levels or equivalent, 0 levels/GCSEs or equivalent, CSEs or equivalent, NVQ or HND or HNC equivalent, or other professional qualifications], lifestyle (smoking [never, past, $<1,1-10,10-14,15-19$, and $\geq 20$ cigarettes per day], alcohol consumption [never, past, $<1,1-7,8-15,16-29$ and $\geq 30$ grams/ day], milk intake [none, $<150,150-299$ and $\geq 300 \mathrm{ml} /$ day], and tea intake [none, $\leq 1,2-3$ and $\geq 3$ cups/day]), anthropometric measurement (BMI [ $<18.5,18.5-24.9,25.0-29.9$, and $\geq 30 \mathrm{~kg} /$ $\mathrm{m}^{2} \mathrm{~J}$ ), comorbidities (history of hypertension, diabetes, cardiovascular disease, and cancer), and sex hormones (log-transformed SHBG and testosterone). ${ }^{\text {T }}$ Coffee intake and SHBG are included in the interaction term as binary variables (ie, non-drinker vs. drinker and low + intermediate level vs. high level). Pvalue for interaction equals to 0.057 for coffee and SHBG. AS: Advanced subsidiary; CI: Confidence interval; CKD: Chronic kidney disease; CSE: Certificate of Secondary Education; GCSE: General Certificate of Secondary Education; HR: Hazard ratio; HNC: Higher National Certificate; HND: Higher National Diploma; NVQ: National Vocational Qualification; SHBG: Sex hormone-binding globulin.
androgenism in females. ${ }^{[36]}$ So we presume that not only lower SHBG but also higher testosterone, partly induced by the decrease in SHBG, hinders the beneficial effects of coffee in female drinkers.

For males, only a small fraction of heavy habitual coffee drinkers with the highest SHBG presented with a tendency of risk reduction in CKD. Thus, we hypothesize, one possible explanation underlying the sex-specific renoprotective effect of coffee may be that daily coffee intake, as an isolated aspect of a person's lifestyle, is insufficient to
adequately overcome the CKD susceptibility brought about by the naturally high-testosterone and low-SHBG concentrations in males.

Despite the reported effect of estradiol in offering protection against incident CKD, ${ }^{[37,38]}$ we failed to confirm its role as a modifier of CKD susceptibility by virtue of the coffee-CKD association.

To the best of our knowledge, we are among the first, using large prospective cohort, to comprehensively
explore the sex disparity and the potential modification effect of sex hormones and SHBG in the coffee-CKD association. While it has several strengths, some limitations need to be pointed out. First, the study is mainly constituted of Caucasians with a median age of 57 years, making it difficult to extrapolate our results to other ethnic backgrounds or age groups. Besides, a "healthy volunteer" selection bias regarding UK Biobank has been identified. ${ }^{[39]}$ Second, since the exposure was ascertained only by self-report at baseline assessment, reporting bias of coffee consumption was inevitable, and we were not aware of the changes in participants' dietary habits. Third, biomarker concentrations were based on one single measurement and random measurement error existed, leading to the misclassification of sex hormones and SHBG subgroups. Fourth, when conducting the subgroup analysis, smaller sample size may lead to inadequate statistical power and increase the likelihood of false negatives (ie, type II error), ${ }^{[40]}$ especially for the analysis stratified by estradiol, since only a small proportion of participants underwent the measurement of estradiol.

In conclusion, despite the overall observed reno-protective effect of coffee in the general population, sex disparity existed, with the result that females are more likely to experience the benefit. Sex hormones and SHBG may partly account for the sex disparity. Further studies investigating the full mechanisms sex-specifically linking coffee to CKD risk reduction are warranted.

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## Conflicts of interest

None.

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[^1]:    *Model 1 is stratified for 5 -year age groups, sex, and 22 assessment centers. ${ }^{\dagger}$ Model 2 is stratified for 5 -year age groups, sex, and 22 assessment centers and adjusted for sociodemographic factors (race [in the primary analysis cohort only], Townsend deprivation index [in quartiles], and highest education level [college or university degree, A levels/AS levels or equivalent, O levels/GCSEs or equivalent, CSEs or equivalent, NVQ, HND, or HNC equivalent, or other professional qualifications]), lifestyle (smoking [never, past, $<1,1-10,10-14,15-19$, and $\geq 20$ cigarettes/day], alcohol consumption [never, past, $<1,1-7,8-15,16-29$, and $\geq 30 \mathrm{~g} /$ day], milk intake [none, $<150,150-299$, and $\geq 300 \mathrm{~mL} /$ day], and tea intake [none, $\leq 1,2-$ 3 , and $\geq 3$ cups/day]), anthropometric measurement (body mass index $\left[<18.5,18.5-24.9,25-29.9\right.$, and $\left.\geq 30 \mathrm{~kg} / \mathrm{m}^{2}\right]$ ), comorbidities (history of hypertension, diabetes, CVD, and cancer), and genetic factors (in the genetic analysis cohort only, polygenetic risk score of CKD [low risk, intermediate risk, and high risk] and genetic polymorphisms of caffeine metabolizing rate [fast and slow]). ${ }^{\ddagger}$ Model 3 is adjusted for model $2+$ log-transformed SHBG and testosterone. ${ }^{8} P$ values of HRs are within the range of 0.01 to 0.05 . ${ }^{\|} P$ values of HRs are within the range of 0.001 to 0.01 . $P$ values of HRs are within the range of $<0.001$. AS: Advanced subsidiary; CI: Confidence interval; CKD: Chronic kidney disease; CSE: Certificate of Secondary Education; CVD: Cardiovascular disease; GCSE: General Certificate of Secondary Education; HR: Hazard ratio; HNC: Higher National Certificate; HND: Higher National Diploma; HTN: Hypertension; NVQ: National Vocational Qualification; SHBG: Sex hormone-binding globulin.

