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ORIGINAL ARTICLE

Association of polymorphisms of A260G and A386G in *DAZL* gene with male infertility: a meta-analysis and systemic review

Ping Chen, Xiao Wang, Chang Xu, He Xiao, Wen-Hao Zhang, Xing-Huan Wang, Xin-Hua Zhang

To investigate the association of single nucleotide polymorphism 260 and 386 (SNP260 and SNP386) gene with male infertility, an electronic search was performed to identify case-control studies evaluating the relationship of SNP260 or SNP386 of deleted in azoospermia-like (*DAZL*) and male infertility. Review Manager 5 was used to process the meta-analysis and other statistical analysis. A total of 139 records were retrieved, of which 13 case-control studies with total 2715 patients and 1835 normozoospermic men were included. SNP260 was found not to play a functional role in male oligo/azoospermia either for Caucasians or for Asians. But for SNP386, models of allele (A/G), dominant (AA/AG + GG), co-dominant (AA/AG) and super-dominant (AA + GG/AG) had a strong correlation to spermatogenic failure with related odds ratio being 0.15 (95% confidence interval [95% CI] 0.07 to 0.34, P < 0.00001), 0.16 (95% CI 0.07 to 0.35, P < 0.00001), 0.15 (95% CI 0.06 to 0.33, P < 0.00001) and 0.15 (95% CI 0.06 to 0.33, P < 0.00001), respectively. Moreover, this correlation was only found in the Chinese Han population (decreasing around 85% risk of oligo/azoospermia infertility) and not found in India, Japan, and Caucasian countries. Our analysis demonstrated that SNP260 of *DAZL* did not contribute to oligo/azoospermia while SNP386 was correlated to male infertility. However, this correlation was only found in China with a country-specific and ethnicity-specific manner.

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Keywords: deleted in azoospermia-like; male infertility; meta-analysis; polymorphism; single nucleotide polymorphisms

INTRODUCTION

The prevalence of infertility may be as high as 15% with about 50% of cases attributed to male factors. Male infertility is commonly believed to result from an exogenous insult after birth or in utero, such as an exposure to gonadotoxins, trauma, or infection, but there is no effective treatment for patients with nonobstructive azoospermia, in which there is an absence of mature sperm in the testes. This suggests the exogenous insult after birth or in utero may not play a role for nonobstructive azoospermia patients.1 But from an evidence-based medicine perspective, rigorous studies demonstrating direct cause and effect relationship in human infertility are frequently lacking. Changes in sperm motility parameters are the predominant factor in most cases of male infertility, however, in 30%-45%, the cause of the abnormal sperm parameters is not identified (idiopathic male infertility).² In recent years, there has been remarkable progress in understanding the regulation of spermatogenesis and the causes of male infertility, largely as a result of the molecular and biochemical insights. It is estimated that in about 30% of cases, male infertility is due to a genetic disorder such as aneuploidy, structural chromosomal abnormalities, DNA damage, and gene mutations including a variety of newly discovered genes.3 The most frequent causes are rare variants causing the spermatogenic dysfunction.

The incidence of chromosomal abnormalities is 5.8% for infertile men.⁴ Of these, sex chromosome abnormalities account for 4.2% and

Department of Urology, Zhongnan Hospital of Wuhan University, Wuhan 430071, China. Correspondence: Dr. XH Zhang (zxhmd2000@yahoo.com) Received: 27 August 2014; Revised: 19 October 2014; Accepted: 2 February 2015

autosomal abnormalities for 1.5%. Among the azoospermia/severe oligozoospermia patients, microdeletions of the Y chromosome, including deleted in azoospermia (DAZ) gene family, are the major, well-characterized genetic causes. DAZ-like (DAZL), an autosomal homolog of DAZ, localized on chromosome 3 and expressed in germ cells, is essential for germ cell lineage development in several species.^{5,6} Mutation analysis of DAZL in infertile men identified the first two nonsynonymous single nucleotide polymorphism (SNP) at nucleotide position 260 (exon 2) and 386 (exon 3), resulting in the amino acid exchange T12A and T54A, respectively.7 In recent years, a number of studies investigated the possible association of azoospermia or oligozoospermia with A260G or A386G sites polymorphism but controversies exist. Tüttelmann et al.8 summarized all polymorphisms and male infertility that have been investigated in single case-control studies in their meta-analysis in 2007 and revealed significant associations between polymorphism and spermatogenic failure only for AZF gr/gr deletions and MTHFR 677C \rightarrow T but not for POLG, DAZL, USP26, or FSHR. However, for DAZL polymorphism, just A260G gene loci were analyzed with only six included studies in their review. Thus, no complete meta-analysis has been conducted to verify this correlation. In this review, we apply the methods of evidence-based medicine to evaluate and analyze the documented studies of DAZL polymorphism and male infertility so as to provide a more systematic and comprehensive assessment of their associations.

MATERIALS AND METHODS

Inclusion criteria

Case-control studies investigated the association of SNP260 or SNP386 of *DAZL* with male infertility should provide sufficient data that could be extracted and used to calculate odds ratios (ORs) and 95% confidence intervals (95% CI). Patients were diagnosed as male infertility with sperm count $<20 \times 10^6$ ml⁻¹ or no sperm in at least two seminal fluid examinations after 3/4 days of sexual abstinence, which was accordant with 2012 European Association of Urology⁹ or 2010 WHO Laboratory Manual for the Examination and Processing of Human Semen.¹⁰ There was no limitation to the methods of assays for detecting SNP of *DAZL*.

Exclusion criteria

Repeat publications and grey literature that were unpublished and reported superficially studies, such as in the form of an abstract and the studies in which sample size <10 or main data could not be obtained would be excluded.

Literature search strategy

We performed an electronic search of PubMed (1966 to September 2013), ScienceDirect Online (1995 to September 2013), Wiley Online Library (2010 to September 2013), China National Knowledge Infrastructure (CNKI) (1999 to September 2013), VIP Database (2000 to September 2013) and Chinese Dissertation Database (CDDB) (1998 to September 2013) for case-control studies evaluating the relationship of SNP260 or SNP386 of DAZL and male idiopathic oligozoospermia (sperm number <20 × 106 spermatozoa per ml) or azoospermia. The search keywords were used with different combinations with both medical subject headings terms and text words: "male infertility" or "male sterility" or "male infecundity" or "male sterile" or "male dysgenesis" or "azoospermia" or "oligozoospermia" or "azoospermic" or "oligozoospermic" or "OAT" plus "DAZL" or "deleted in azoospermia-like" or "spgyla" or "Daz-like autosomal" or "SNP260" or "SNP386" or "T12A" or "T54A" or "SNP." Publication date was not restricted in our search. Reference lists of the included studies and supplemental materials were checked manually to further identify related studies. Meanwhile, published genome-wide association studies (GWASs) about male infertility were also examined.

Selection of studies

Two reviewers (CX and HX) independently screened the title, abstract, and keywords of each article retrieved. Full-text papers were screened for further assessment if the information given suggested that the study fulfilled the inclusion criteria and did not meet the exclusion criteria. Discrepancies were settled by discussion and consensus with all the authors.

Data extraction

The following information was independently extracted from the identified studies by three reviewers (CX, HX and WHZ) using a standard form with first author's surname, year of publication, country, ethnicity, allele frequencies, and genotype distribution in cases and controls, method of genotype test and Hardy–Weinberg equilibrium (HWE) test. Ethnicities were stratified as Asians, Caucasians, and others. The authors of original studies were consulted for missing information where necessary. Discrepancies were resolved by open discussion.

HWE test

HWE test was performed, and the HWE significance of the control groups was calculated with StataSE 12.0 (StataCorp LP, College Station, TX, USA) when the original information was not provided.

Data synthesis and analysis

The significance for five genetic models (allele, dominant, recessive, co-dominant, and super-dominant genetic models) was evaluated for each study separately. All the associations were indicated as ORs with the corresponding 95% CI. Based on the individual ORs, a pooled OR was estimated. Subgroup analysis was also performed by stratifying country and ethnicity. Fixed-effects model¹¹ or the random-effects model¹² of meta-analysis was chosen according to the results of heterogeneity tests among individual studies by Review Manager 5 (The Cochrane Library). The significance of the pooled OR was determined using the *Z* test and *P* < 0.05 was considered statistically significant.

Heterogeneity assumption was assessed by Cochran's Q statistic¹³ and *I*² statistic. The heterogeneity was considered statistically significant if *P* < 0.10. The random-effects model (if *P* < 0.10 and *I*² >50%) or the fixed-effects model (if *P* ≥ 0.10 and *I*² <50%) was used to pool the ORs. Egger's test was used to evaluate the publication bias,which was considered when *P* < 0.05.

RESULTS

Characteristics of included studies

Using the database search strategy, a total of 139 records were retrieved from PubMed, Science Direct, Wiley Online Library, CNKI, VIP and CDDB, of which 13 case-control studies finally met full inclusion criteria for this review.^{7,14-25} No dataset of genotype frequencies of SNP was acquired from GWASs of male infertility.^{26,27} **Supplementary Figure 1** depicts the flowchart of the search process. **Supplementary Table 1** describes the characteristics of the included trials. **Supplementary Tables 2** and **3** describe the distribution of genotypes and allele frequencies of SNP260 and SNP386 in infertile men and control group, respectively.

Meta-analysis of single nucleotide polymorphism 260

A total of ten studies were included for the analysis of SNP260. Figure 1 showed the meta-analysis for allele model (A/G), dominant model (AA/AG+GG), recessive model (AA+AG/GG), co-dominant model (AA/AG), co-dominant model (AA/GG) and super-dominant model (AA + GG/AG), respectively. I^2 standing for the heterogeneity among studies for all models was 0%; thus, fixed-effects models were applied. Related ORs of the six models were 0.88 (95% CI 0.73 to 1.07, P = 0.21), 0.85 (95% CI 0.69 to 1.05, P = 0.12), 0.70 (95% CI 0.31 to 1.58, P = 0.39), 0.83 (95% CI 0.67 to 1.02, P = 0.08), 1.39 (95% CI 0.61 to 3.17, P = 0.43) and 0.82 (95% CI 0.66 to 1.02, P = 0.08), respectively, which indicated that SNP260 was not significantly associated with male infertility for all models. Sensitivity analysis was performed by eliminating each of the included studies, and the statistical significance was not changed, suggesting that the results above were reliable. Egger's test revealed that there was no any obvious evidence of publication bias (P = 0.737).

 Table 1 shows the summary of the subgroup analysis stratified

 by different ethnicity. No significant difference was detected between

 each subgroup, either.

Meta-analysis of single nucleotide polymorphism 386

Single nucleotide polymorphism 386 of *DAZL* was not found in populations from Germany, Italy, Japan, Northern China, and Western India. It was only detected in populations from Western China, Taiwan of China, Northern, and Eastern India. **Figure 2** showed the meta-analysis for the allele model (A/G), dominant model (AA/AG + GG), recessive model (AA + AG/GG), co-dominant model (AA/AG), co-dominant model (AA/GG), and super-dominant model (AA + GG/AG), in



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which *I*² standing for the heterogeneity among the studies was 19%, 28%, 0%, 38%, 0%, and 38%, respectively. Thus fixed-effects models were applied. Related OR was 0.15 (95% CI 0.07 to 0.34, *P* < 0.00001), 0.16 (95% CI 0.07 to 0.35, *P* < 0.00001), 2.11 (95% CI 0.24 to 18.67, *P* = 0.50), 0.15 (95% CI 0.06 to 0.33, *P* < 0.00001), 0.45 (95% CI 0.05 to 4.04, *P* = 0.48) and 0.15 (95% CI 0.06 to 0.33, *P* < 0.00001), indicating that model of A/G, AA/AG + GG, AA + GG/AG AA/AG of SNP386 was significantly associated with spermatogenic failure, while for model AA + AG/GG and AA/GG, the difference was not significant. Sensitivity analysis was performed by eliminating each of the included studies, and the statistical significance was not changed, suggesting that the results above were reliable. Egger's test revealed that there was obvious evidence of publication bias (*P* = 0.009).

Table 1 shows subgroup analysis stratified by different countries. For allele model (A/G), dominant model (AA/AG + GG), super-dominant model (AA + GG/AG) and co-dominant model (AA/AG) model, we

detected significant association of SNP386 with male infertility. But the significant difference was kept consistent only in China. In co-dominant model (AA/GG) and recessive model (AA + AG/GG), the genotype of GG was not detected in India and meta-analysis of this subgroup was not applicable.

DISCUSSION

This meta-analysis and systemic review investigated the association of *DAZL* polymorphisms and male infertility. Our novel data demonstrated that SNP260 of *DAZL* did not contribute to oligozoospermia or azoospermia, while SNP386 of *DAZL* was correlated to male infertility only in the Chinese Han.

Polymorphisms or genetic variants in genes involved in spermatogenesis are considered potential risk factors which may contribute to the severity of spermatogenic failure. Several polymorphic variants have been described in association with oligo/

а	Ca	se	Con	trol			d	C	ase	Co	ntrol				
Study	Events	Total	Events	Total	ORs [95%CI]	M-H, Fixed	Study	Event	s Total	Event	s Total	ORs [95%CI]	M-H,Fi	xed	
Becherini L ²⁴ 2004	259	306	400	458	0.80 [0.53,1.21]	-8-	Becherini L ²⁴ 2004	108	151	176	224	0.68 [0.43,1.10]	-8-		
Chen P ²² 2010	291	314	100	114	1.77 [0.88,3.57]		Chen P22 2010	135	156	44	56	1.75 [0.80,3.85]	+	-	
Kishlay K ²³ 2011	179	200	183	200	0.79 [0.40,1.55]		Kishlay K ²³ 2011	79	100	83	100	0.77 [0.38,1.57]			
Teng YN ⁷ 2002	274	284	226	232	0.73 [0.26,2.03]		Teng YN ⁷ 2002	132	142	110	116	0.72 [0.25,2.04]		-	
Teng YN ¹⁴ 2006	443	462	370	382	0.76 [0.36,1.58]		Teng YN14 2006	212	231	179	191	0.75 [0.35,1.58]			
Thangaraj K ¹⁸ 2006	1264	1320	671	700	0.98 [0.62,1.54]	+	Thangaraj K ¹⁸ 2006	606	658	325	346	0.75 [0.45,1.27]			
Tschanter P16 2004	349	404	290	330	0.88 [0.57,1.35]	+	Tschanter P16 2004	148	201	125	165	0.89 [0.56,1.44]	-		
Wang H ²⁵ 2009	380	392	78	80	0.81 [0.18,3.70]	<u> </u>	Wang H ²⁵ 2009	190	190	39	39	Not estimable			
Wen XH ¹⁹ 2007	97	100	105	106	0.31 [0.03,3.01]		Wen XH ¹⁹ 2007	47	50	52	53	0.30 [0.03,3.00]		_	
Yang XJ ¹⁷ 2005	430	468	244	262	0.83 [0.47,1.49]	-	Yang XJ ¹⁷ 2005	198	232	113	131	0.93 [0.50,1.72]	-+	-	
Total (95%CI)		4250		2864	0.88 [0.73,1.07]	•	Total (95%CI)		2111		1421	0.83 [0.67,1.02]	•		
Total events	3966		2667				Total events	1855		1246					
Hotorogonoity: Chi2	- 5 46 4	df - 0 /	P - 0 70). R = 0	× –		Heterogeneity: Chi2=	= 5.38.	df = 8 (<i>F</i>	e 0.72): <i>f</i> ² = 0%	, H			
Test for overall effect	- 3.40, t ct: Z = 1	.25 (P	= 0.21)), 7 – 0	0.01	0.1 1 10 100	Test for overall effec	t: Z = 1	.75 (P=	0.08)		0.01	0.1 1	10	100
b		(.				[A] [G]	Α			_			[AA]	[AG]	
	Ca	se	Cont	trol	OB+ (05% CI)	M LL Floor d	Shudu	- 0	ase	Cor	ntrol	OB+ (05%/ 01)	MUE		
Study	Events	lotal	Events	lotal	ORS [95%CI]	- IVI-H, Fixed	Boohorini L 24 2004	Even	IS IOTAI	Events	s lotal	ORS [95%CI]	IVI-FI,FI	xea -	
Becherini L ^{ar} 2004	108	153	176	229	0.72 [0.45,1.15]		Chop B22 2010	108	110	176	181	1.53 [0.29,8.05]			
Chen P** 2010	135	107	44	57	1.81 [0.84,3.90]		Kichlay K ²³ 2011	135	130	44	45	3.07 [0.19,50.08]			_
Kishiay K ²⁰ 2011	79	100	83	100	0.77 [0.38,1.57]		Topa XN/ 2002	19	79	83	83	Not estimable			
Teng YN 2002	132	142	110	116	0.72 [0.25,2.04]		Teng VNI4 2002	132	132	110	110	Not estimable			
Teng YN* 2006	212	231	179	191	0.75 [0.35,1.58]	I	Thangarai K18 2006	212	212	1/9	1/9	Not estimable			
Thangaraj K ¹⁰ 2006	606	660	325	350	0.86 [0.53,1.41]	1	Techanter P ¹⁶ 2004	140	140	325	329	3.73 [0.68,20.47]		· ·	
Ischanter P ¹⁰ 2004	148	202	125	165	0.88 [0.55,1.41]		Mana H25 2009	148	149	125	125	0.38 [0.02,9.77]			
Wang H ²³ 2009	190	196	39	40	0.81 [0.10,6.93]		Wang TH 2009	190	190	39	40	0.81 [0.10,6.93]	1		
Vien XH ¹² 2007	4/	50	52	53	0.30 [0.03,3.00]		Vien X117 2007	4/	4/	52	52	Not estimable			
rang XJ ¹¹ 2005	198	234	113	131	0.88 [0.48, 1.61]	1	Tang Xu 2000	190	200	113	113	0.35 [0.02,7.35]	-		
Total (95%Cl)		2125		1432	0.85 (0.69.1.05)	•	Total (95%CI)		1869		1257	1 39 [0 61 3 17]			
Total events	1855	2120	1246	1452	0.00 [0.00, 1.00]		Total events	1855		1246	1207	1.55 [5.61,5.17]			
	1000		1210		⊢		Heterogeneity: Chi2=	= 3.23,	df = 5 (<i>F</i>	P = 0.66); <i>I</i> ² = 0%	. ⊢			
Heterogeneity: Chi ²	= 5.33,	df = 9 ((P = 0.80)); <i>I</i> ² = 0	0.01	0.1 1 10 100	Test for overall effec	t: Z = 0	.79 (P =	0.43)		0.01	0.1 1	10	100
Test for overall effe	ect: Z =	1.54 (P	= 0.12)			[AA] [AG+GG]	f						[AA]	[GG]	
C	Cas	se	Cont	rol				Ca	se	Con	trol				
Study	Events	Total	Events	Total	ORs [95%CI]		Study	E∨ent	s Total	Events	Total	ORs [95%CI]			
Becherini L ²⁴ 2004	151	153	224	229	1.69 [0.32,8.80]		Becherini L ²⁴ 2004	110	153	181	229	0.68 [0.42,1.09]			
Chen P ²² 2010	156	157	56	57	2.79 [0.17,45.29]		Chen P ²² 2010	136	157	45	57	1.73 [0.79,3.79]	+	-	
Kishlay K ²³ 2011	100	100	100	100	Not estimable		Kishlay K ²³ 2011	79	100	83	100	0.77 [0.38,1.57]			
Teng YN ⁷ 2002	142	142	116	116	Not estimable		Teng YN/ 2002	132	142	110	116	0.72 [0.25,2.04]		-	
Teng YN ¹⁴ 2006	231	231	191	191	Not estimable		Teng YN ¹⁴ 2006	212	231	179	191	0.75 [0.35,1.58]			
Thangaraj K ¹⁸ 2006	658	660	346	350	0.80 [0.69,20.87]		Thangaraj K ¹⁸ 2006	608	660	329	350	0.75 [0.44,1.26]			
Tschanter P18 2004	201	202	165	165	0.41 [0.02,10.03] -		Tschanter P ¹⁶ 2004	149	202	125	165	0.90 [0.56,1.45]	-		
Wang H ²⁵ 2009	190	196	39	40	0.81 [0.10,6.93]		Wang H ²⁵ 2009	196	196	40	40	Not estimable			
Wen XH ¹⁹ 2007	50	50	53	53	Not estimable		Wen XH ¹⁹ 2007	47	50	52	53	0.30 [0.03,3.00]		_	
Yang XJ ¹⁷ 2005	232	234	131	131	0.35 [0.02,7.42] -		Yang XJ ¹⁷ 2005	200	234	113	131	0.94 [0.51,1.74]	1	-	
Tatal (059(OI)		0405		1 4 2 2	4 42 10 62 2 251		Total (05% CI)		2425			0.0010.004.003			
Total (95%CI)		2125	1.001	1432	1.43 [U.03,3.25]		Total (95%CI)	1055	2125	40.10	1432	0.82 [0.66,1.02]	•		
i utai events	2111		1421				Heterogeneity: Chi2:	1805	df = 8 (4	1246 2 = 0.72): <i>P</i> = 0%	6 H			
Heterogeneity: Chi2	= 3.19, (at: 7 = 0	dt = 5 (.	P = 0.67); I= 0	% 0.01	0.1 1 10 100	Test for overall effect	st: Z = 1	.77 (P=	0.08)	,, , = 07	0.01	0.1 1	10	100
rest for overall effer	u. ∠ - U	.00 (P	- 0.59)		[A	A+AG] [GG]							[AA+GG]	[AG]	

Figure 1: The association of single nucleotide polymorphism 260 with male infertility. A forest plot in (**a**) allele model (A/G), (**b**) dominant model (AA/AG + GG), (**c**) recessive model (AA + AG/GG), (**d**) co-dominant model (AA/AG), (**e**) co-dominant model (AA/GG), (**f**) super-dominant model (AA + GG/AG). The association was indicated as odds ratio (OR) estimate with the corresponding 95% confidence interval (CI). The OR estimate of each study is marked with a solid black square. The size of the square represents the weight that the corresponding study exerts in the meta-analysis. The CIs of pooled estimates are displayed as a horizontal line through the diamond. OR <1 indicates increased risk of male infertility. A: wild type gene, G: mutant gene.

azoospermia,²⁸⁻³⁵ among which two *DAZL* polymorphisms (SNP260 and SNP386) were the first description of SNPs of autosomal genes influencing human spermatogenesis. Analysis of the distribution and frequency of these SNPs among fertile and infertile men has demonstrated a strong association of the heterozygous genotype for SNP386 with spermatogenic failure for Asians while it was not detected in Caucasians. But no such association could be found for SNP260.

In order to provide a complete assessment, we performed the present meta-analysis of 13 independent case-control studies with total number of 2715 patients and 1835 normozoospermic men, of which there were 2500 oligo/azoospermic men. We found that SNP260 of *DAZL* had no correlation with male oligozoospermia or azoospermia which was consistent with Tüttelmann's meta-analysis⁸ including 6 studies with 1600 patients and 1100 controls of different ethnic origin. In this previous review,⁸ the authors did not perform meta-analysis on the relationship between SNP386 and male infertility. Instead, they did a literature review and drew a conclusion

that SNP386 was never found in non-Chinese populations. However, SNP386 was later detected in India though there was no association with male infertility.

We further conducted a subgroup analysis by ethnicity and country in the current meta-analysis. In the stratified analysis according to ethnicity, we found that SNP260 had no correlation with male oligozoospermia or azoospermia either in Caucasian or Asian populations. For SNP386, it was only detected in Asians and not in Caucasians. We also further performed the stratified analysis by country and demonstrated that models of SNP386 of allele (A/G), dominant (AA/ AG + GG), co-dominant (AA/AG), and super-dominant (AA + GG/AG) had a strong association with spermatogenic failure in China. Although these variants were detected in northern and eastern India, the difference did not reach statistical significance. Therefore, functional DAZL A/G polymorphism may play a penetrance role in male oligo/azoospermia susceptibility in ethnicity-specific and country-specific manners.

Discrepancies between results of association studies are a rather frequent phenomenon and may be related to many different factors.³⁶

Fable 1: Subgroup analysis stratifie	d by different	population and	countries
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Gene	Genetic model	Population	P O test	l² (%)	P Z test	OR (95% CI)
SNP260		Overall	0.79	0	0.21	0.88 (0.73_1.07)
5111 200		Caucasian	0.75	0	0.24	0.83 (0.62-1.13)
		Asian	0.65	0	0.52	0.92 (0.71_1.19)
	Dominant (AA/AG+GG)	Overall	0.8	0	0.12	0.85 (0.69–1.05)
	Dominant (NVNG) (G)	Caucasian	0.57	0	0.12	0.80 (0.57–1.11)
		Asian	0.69	0	0.37	0.88 (0.67–1.16)
	Recessive (GG/AA+AG)	Overall	0.67	0	0.86	1.43 (0.63–3.25)
		Caucasian	0.44	0	0.28	1.22 (0.31-4.87)
		Asian	0.48	0	0.88	1.57 (0.58-4.29)
	Co-dominant (AA/AG)	Overall	0.72	0	0.08	0.83 (0.67–1.02)
		Caucasian	0.44	0	0.16	0.78 (0.56-1.10)
		Asian	0.6	0	0.27	0.85 (0.65–1.13)
	Co-dominant (AA/GG)	Overall	0.66	0	0.43	1.39 (0.60–3.17)
		Caucasian	0.46	0	0.87	1.12 (0.28-4.52)
		Asian	0.47	0	0.38	1.57 (0.58–4.30)
	Super-dominant (AA+GG/AG)	Overall	0.72	0	0.08	0.82 (0.66–1.02)
		Caucasian	0.41	0	0.15	0.78 (0.56–1.09)
		Asian	0.61	0	0.26	0.85 (0.65–1.13)
SNP386	Allele (A/G)	Overall	0.29	19	< 0.01	0.15 (0.07–0.34)
		India	0.4	0	0.79	0.80 (0.16-4.09)
		China	0.79	0	< 0.01	0.11 (0.04–0.28)
	Dominant (AA/AG+GG)	Overall	0.23	28	< 0.01	0.16 (0.07–0.35)
		India	0.4	0	0.79	0.80 (0.16-4.09)
		China	0.53	0	< 0.01	0.11 (0.04–0.28)
	Recessive (GG/AA+AG)	Overall	0.89	0	0.5	2.11 (0.24–18.67)
		India	-	-	-	-
		China	0.89	0	0.5	2.11 (0.24–18.67)
	Co-dominant (AA/AG)	Overall	0.19	38	< 0.01	0.15 (0.06–0.33)
		India	0.4	0	0.79	0.80 (0.16-4.09)
		China	0.88	0	< 0.01	0.09 (0.03–0.26)
	Co-dominant (AA/GG)	Overall	0.86	0	0.48	0.45 (0.05–4.04)
		India	-	-	-	-
		China	0.86	0	0.48	0.45 (0.05–4.04)
	Super-dominant (AA+GG/AG)	Overall	0.19	38	< 0.01	0.15 (0.06–0.33)
		India	0.4	0	0.79	0.80 (0.16-4.09)
		China	0.89	0	< 0.01	0.09 (0.03–0.26)

OR: odds ratio; CI: confidence interval; SNP: single nucleotide polymorphisms

An important feature of these studies is that the discrepancy between Caucasian analysis and the Chinese studies is not related to sampling biases (small sample size, inadequate control group or population substructuring), but it is related to the complete absence of the "at risk" SNP in Caucasians, indicating that in Western countries, this polymorphism is probably absent or rare. This remarkable difference is paradigmatic as it shows how ethnic background is important for polymorphisms involved in spermatogenesis, thereby underscoring that different genetic risk factors may be present in different populations. A similar phenomenon is observed for *cystic fibrosis transmembrane conductance regulator* (*CFTR*) gene mutations causing the azoospermia because of congenital bilateral absence of vas deferens. The frequency of a particular *CFTR* mutation is also influenced by the ethnic composition of the population analyzed.³⁷ Our study, therefore, contributes to a better definition of clinically relevant tests, specifically

based on the ethnic origin of the infertile patients.

The low heterogeneity in our analysis contributes to highly reliable results through meta-analysis. There is no publication bias in the studies about *DAZL* SNP260. But there exists publication bias about SNP386 of *DAZL*. The sources of bias might be from the small sample sizes or incorporated literatures of low quality. Bias may be also resulted from more chance of positive findings published than the negative ones.

The results of our meta-analysis should be interpreted after taking into consideration several inevitable study limitations. First, although we set a comprehensive search strategy for the retrieval of eligible studies, we cannot eliminate that some studies might have been lost. Second, there is language limitation as current meta-analysis only contains English and Chinese literature with some articles possibly published in other languages and not accessible to the international journals. Finally, our results are based on unadjusted ORs while a

а	Ca	se	Con	trol					d	Cas	se	Con	trol			
Study	Events	Total	Events	Total	ORs [95%CI]	M-H,	Fixed		Study	Events	Total	Events	Total	ORs [95%CI]	M-H,F	ixed
Bartoloni L ¹⁵ 2004	190	190	126	126	Not estimable				Bartoloni L15 2004	95	95	126	126	Not estimable		
Becherini L ²⁴ 2004	306	306	458	458	Not estimable				Becherini L ²⁴ 2004	153	153	458	458	Not estimable		
Poongothai J ²⁰ 200	8 188	188	280	280	Not estimable				Poongothai J ²⁰ 2008	94	94	280	280	Not estimable		
Singh K ²¹ 2009	330	330	399	400	2.48 [0.10,61.13]				Singh K ²¹ 2009	165	165	399	400	2.49 [0.10,61.50]		
Teng XN ⁷ 2002	263	284	230	232	0.11 [0.03,0.47]	_			Tena XN ⁷ 2002	121	142	230	232	0.10 [0.02,0.44]		
Teng XN ¹⁴ 2006	435	462	380	382	0.08 [0.02,0.36]	_			Tena XN ¹⁴ 2006	205	230	380	382	0.09 [0.02,0.37]	_	
Thangaraj K18 2006	1316	1320	699	700	0.47 [0.05,4.22]	-			Thangarai K ¹⁸ 2006	656	660	699	700	0.47 [0.05.4.22]	-	
Tschanter P16 2004	404	404	330	330	Not estimable				Tschanter P ¹⁶ 2004	202	202	330	330	Not estimable		
Wang H ²⁵ 2009	384	392	78	78	0.29 [0.02.5.04]				Wang H ²⁵ 2009	192	192	78	78	Not estimable		
Wen XH ¹⁹ 2007	100	100	106	106	Not estimable				Men XH19 2007	50	50	106	106	Not estimable		
Yang XJ ¹⁷ 2005	468	468	262	262	Not estimable				Yang XJ ¹⁷ 2005	234	234	262	262	Not estimable		
Total (95%CI)		4444		3354	0.15 [0.07.0.34]				Total (95%CI)		2217		3354	0 15 [0 06 0 33]		
Total events	4384		3348			•			Total events	2167		3348			-	
Heterogeneity: Chi2	= 4.95	df = 4	(P = 0.2)	29)· 8=	19%		L .		Heterogeneity: Chi ² =	4.82.0	df = 3 (P = 0.19	$f^2 =$	38%		
Test for overall effe	ct: 7 =	4 63 ()	P<0.00	001)	0	.005 0.1	1 10	200	Test for overall effect	: Z = 4	56 (P	< 0.0000	01)	0	005 0.1	1 10 20
b				,		[A]	[G]		е	_					[AA]	[AG]
كالمنار	Ca	ise	Co	ntrol						Ca:	se	_ Cont	rol		• • • • •	
Study	Event	s Tota	l Event	s Tota	ORs [95%CI]	M-H,	Fixed		Study	Events	Total	Events	Total	URs [95%CI]	M-H,F	-ixed
Bartoloni L ¹⁵ 2004	95	95	63	63	Not estimable				Bartoloni L ¹⁵ 2004	95	95	63	63	Not estimable		
Becherini L ²⁴ 2004	153	153	229	229	Not estimable				Becherini L ²⁴ 2004	153	153	229	229	Not estimable		
Poongothai J ²⁰ 2008	8 94	94	140	140	Not estimable				Poongothai J ²⁰ 2008	94	94	140	140	Not estimable		
Singh K ²¹ 2009	165	165	199	200	2.49 [0.10,61.50]		+ • • • •		Singh K ²¹ 2009	165	165	199	199	Not estimable		
Teng XN ⁷ 2002	121	142	114	116	0.10 [0.02,0.44]				Teng XN ⁷ 2002	121	121	114	114	Not estimable		
Teng XN ¹⁴ 2006	205	231	189	191	0.08 [0.02,0.36]	_			Teng XN14 2006	205	206	189	189	0.36 [0.01,8.93]		
Thangaraj K18 2006	656	660	349	350	0.47 [0.05,4.22]		<u> </u>		Thangaraj K ¹⁸ 2006	656	656	349	349	Not estimable		
Tschanter P16 2004	202	202	165	165	Not estimable				Tschanter P16 2004	202	202	165	165	Not estimable		
Wang H ²⁵ 2009	192	196	39	39	0.54 [0.03,10.26]	· · · ·	<u> </u>		Wang H ²⁵ 2009	192	196	39	39	0.54 [0.03,10.26]		
Wen XH ¹⁹ 2007	50	50	53	53	Not estimable				Wen XH ¹⁹ 2007	50	50	53	53	Not estimable		
Yang XJ ¹⁷ 2005	234	234	131	131	Not estimable				Yang XJ ¹⁷ 2005	234	234	131	131	Not estimable		
Total (95%CI)		2222		1677	0.16 [0.07,0.35]	•			Total (95%CI)		2172		1671	0.45 [0.05,4.04]		
Total events	2167		1671	1					Total events	2167		1671				
Heterogeneity: Chi2	= 5.57,	df = 4	(P = 0.2)	23); <i>f</i> ²=	28%	+	+ +		Heterogeneity: Chi ² =	0.03, (df = 1 (P = 0.29	9); f ² =	0% H	1	
Test for overall effe	ct: Z =	4.58 (<i>l</i>	P < 0.00	001)	0	005 0.1	1 10	200	Test for overall effect	: Z = 0	.71 (P	= 0.48)		0	.005 0.1	1 10 200
С						[AA]	[AG+G	G]	f	Car		Cont	rol		[AA]	[GG]
Study	Event			ha Tata		М-Н	Fived		Study	Evente	Total	Evente	Total	ORe [95%CI]	М-Н І	Fixed
Bartoloni I 15 2004	Lveni	3 1018 0=		62	Not estimable				Bartoloni I 15 2004	95	95	63	63	Not estimable	,	
Becherini L 24 2004	0	150	0	200	Not octimete				Becherini 24 2004	152	152	220	220	Not estimable		
Boongothai 120 2004	• •	153	U	229	Not estimable				Boongothai 120 2004	04	04	140	140	Not estimable		
Floongothar J=200 Singh K ²¹ 2000	υ 0 ~	94	0	140	Not estimable				Singh K ²¹ 2009	94 16F	94 16F	140	200	1 10 L ESUINADIE		
Tong VN/ 2009	0	165	0	200	Not estimable				Tong VN/ 2009	100	140	199	200	2.48 [0.10,01.50]	_	
Teng XNI 2002	0	142	0	116	Not estimable		_		Tong XNI4 2002	121	142	114	101	0.10[0.02,0.44]		
There are 1 K18 0000	1	231	0	191	2.49 [0.10,61.53]				There are in K18 0000	206	231	189	191	0.09 [0.02,0.37]		
i nangaraj K.º 2006	0	660	0	350	Not estimable				Trangaraj K ¹⁰ 2006	656	660	349	350	0.47 [U.U5,4.22]		
Ischanter P ¹⁰ 2004	0	202	0	165	Not estimable		L_		Ischanter P ¹⁰ 2004	202	202	165	165	Not estimable		
VVang H ²⁵ 2009	4	196	0	39	1.85 [0.10,34.99]			_	Vvang H ²⁰ 2009	196	196	39	39	Not estimable		
Wen XH19 2007	0	50	0	53	Not estimable				vven XH ¹⁹ 2007	50	50	53	53	Not estimable		
Yang XJ ¹⁷ 2005	0	234	0	131	Not estimable				Yang XJ ¹⁷ 2005	234	234	131	131	Not estimable		
Total (95%CI)		2222		1677	2.11 [0.24,18.67]				Total (95%CI)		2222		1677	0.15 [0.06,0.33]	•	
Total events	5		Ω						Total events	2172		1671			-	
Heterogeneity: Chi2	= 0 02	df = 1	(P=0.9	39)· /2=	0%		+ +		Heterogeneity: Chi2=	4.81	-ff = 3 /	P = 0.20	a): P=	38%	1	
Test for overall effe	ct: Z =	0.67 (F	P = 0.50)	C	.005 0.1	1 10	200	Test for overall effect	:: Z = 4	.56 (P	< 0.0000),, , = 01)	<i>/</i>	.005 0.1 [AA+GG]	1 10 200 [AG1

Figure 2: The association of single nucleotide polymorphism 386 with male infertility. A forest plot in (a) allele model (A/G), (b) dominant model (AA/AG + GG), (c) recessive model (AA + AG/GG), (d) co-dominant model (AA/AG), (e) co-dominant model (AA/GG), (f) super-dominant model (AA + GG/AG). The association was indicated as odds ratio (OR) with the corresponding 95% confidence interval (CI). The OR estimate of each study is marked with a solid black square. The size of the square represents the weight that the corresponding study exerts in the meta-analysis. The CIs of pooled estimates are displayed as a horizontal line through the diamond. OR <1 indicates increased risk of male infertility. A: wild type gene and G: mutant gene.

more precise estimation should take into account the effect of multiple confounders such as age and disease severity on the association. Therefore, multicenter, large sample sizes, high-quality case-control or cohort trials are further required to study the exact relationship between *DAZL* polymorphisms and male oligozoospermia or azoospermia.

CONCLUSIONS

Our analysis demonstrated that SNP260 of *DAZL* did not contribute to oligozoospermia or azoospermia in any ethnic subgroups. Instead, SNP386 was correlated to male infertility though this correlation was only found in China with a country-specific and ethnicity-specific manner. However, high-quality trials are further required to study the exact relationship between *DAZL* polymorphisms and male infertility.

AUTHOR CONTRIBUTIONS

All authors have fulfilled all conditions required for authorship. XHZ and PC conceived and designed the experiments. CX, HX, and WHZ performed the electronic search, selected studies, extracted data and performed quality assessment. PC and XW analyzed data and conducted a meta-analysis. XHZ and XHW supervised the research, edited and drafted revisions to the article.

COMPETING INTERESTS

The authors declare no competing interests.

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