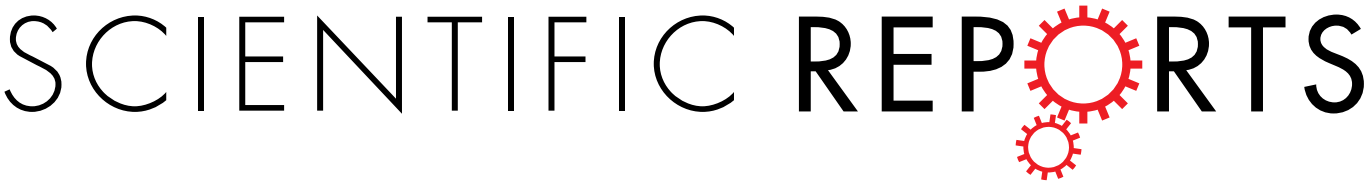

OPEN

Methylation-to-Expression Feature Models of Breast Cancer Accurately Predict Overall Survival, Distant-Recurrence Free Survival, and Pathologic Complete Response in Multiple Cohorts

Received: 13 December 2017
Accepted: 13 March 2018
Published online: 26 March 2018

Jeffrey A. Thompson[1], Brock C. Christensen[2] & Carmen J. Marsit[3]

Prognostic biomarkers serve a variety of purposes in cancer treatment and research, such as prediction of cancer progression, and treatment eligibility. Despite growing interest in multi-omic data integration for defining prognostic biomarkers, validated methods have been slow to emerge. Given that breast cancer has been the focus of intense research, it is amenable to studying the benefits of multi-omic prognostic models due to the availability of datasets. Thus, we examined the efficacy of our methylation-to-expression feature model (M2EFM) approach to combining molecular and clinical predictors to create risk scores for overall survival, distant metastasis, and chemosensitivity in breast cancer. Gene expression, DNA methylation, and clinical variables were integrated via M2EFM to build models of overall survival using 1028 breast tumor samples and applied to validation cohorts of 61 and 327 samples. Models of distant recurrence-free survival and pathologic complete response were built using 306 samples and validated on 182 samples. Despite different populations and assays, M2EFM models validated with good accuracy (C-index or AUC $\geq$ 0.7) for all outcomes and had the most consistent performance compared to other methods. Finally, we demonstrated that M2EFM identifies functionally relevant genes, which could be useful in translating an M2EFM biomarker to the clinic.

The extreme heterogeneity of breast cancer makes one-size-fits all treatments difficult to achieve and although a number of targeted therapies are available[1] personalized therapies work only for a limited number of patients. Better methods are needed to rapidly identify patients that can benefit from specific treatments, reveal new targets for historically difficult to treat patients, and provide accurate prognoses to patients to enable better life planning.

Large scale collection of cancer genomic data, such as The Cancer Genome Atlas (TCGA)[2] and the International Cancer Genome Consortium[3], and expression repositories such as NCBI's Gene Expression Omnibus[4] have enabled developments in prognostic and predictive cancer models. In addition to improved understanding of how cancer disrupts normal cellular function, such data have led to the development of biomarker panels, such as the PAM50 risk of recurrence score (rorS)[5] and the 21 gene Oncotype DX[6], which are both now used for clinical decision making in the treatment of breast cancer[7]. Furthermore, to the extent that such tests are able to predict the occurrence of distant metastases, they may enable early prevention strategies to be put in place.

Despite such promising developments, the goals of precision medicine remain elusive due to specific limitations in molecular biomarker studies. These include limited understanding of the optimal genomic features to

[1]Department of Biostatistics, University of Kansas Medical Center, Kansas City, USA. [2]Department of Epidemiology, Geisel School of Medicine at Dartmouth College, Hanover, USA. [3]Department of Environmental Health, Rollins School of Public Health at Emory University, Atlanta, USA. Correspondence and requests for materials should be addressed to J.A.T. (email: jthompson21@kumc.edu)

utilize, lack of validation on external data, and inconsistent use of best practices for creating biomarkers with the best chances of entering clinical practice[8,9].

There has been growing interest in the integration of molecular and traditional clinicopathological markers to achieve higher levels of prognostic power and utility and small but significant gains in prognostic accuracy have been observed through this approach[5,10]. To the extent that these sources are independently prognostic, they should result in higher accuracy than either source alone and better understanding of prognostic features by revealing holistic relationships in the data. As yet, few prognostic or predictive models are designed to take advantage of multiple data types. In this work, we apply our own previously developed data integration approach for combining gene expression and DNA-methylation data to model the impact of regulatory disruption on breast cancer prognosis and prediction while also incorporating prognostic clinical variables[11]. This study overcomes prior limitations by creating a model that can identify biologically plausible targets, using best practices in assessing the model[8,9], and by validating the actual model (not just the gene signature) in multiple, external datasets, representing different demographics and disease distributions.

Methods

M2EFM. Previously, we introduced a data-integrated modeling approach we call Methylation-to-Expression Feature Model (M2EFM)[11] and applied it in the case of clear cell renal cell carcinoma (ccRCC). The basis of this approach is to pre-filter loci to those that are differentially methylated between matched tumor and adjacent, pathologically normal tissue and to associate the methylation status of those loci with significant differences in gene expression in the tumor samples. The differentially methylated loci are then called m2eQTL (for methylation-to-expression QTL) and the genes are called m2eGenes. An initial model is built from m2eQTL and m2eGenes with the largest effects to generate risk scores for all samples using Cox-Ridge regression (using the exponential function of the weighted sum of the model features). A final model is built from the risk scores in conjunction with clinical variables and a final risk score is generated from the weighted sum of the features of this model. Here we build M2EFM models of overall survival (OS), distant recurrence-free survival (DRFS), and pathologic complete response (pCR) in breast cancer, where OS is defined as time from initial diagnosis to death from any cause, DRFS is defined as time from diagnosis to distant recurrence, death from cancer, or death from unknown cause, and pCR is defined as a patient being recorded as having a pathologic complete response or a residual cancer burden class of RCB 0/I as in Hatzis, *et al.*[12].

Data. A full description of data pre-processing appears in the Supplementary Methods. For this work, we supplied a list of probes to M2EFM from a study of CpG loci involved in epigenetic field effects in breast cancer based on data available in GEO (GSE69914). The loci were delineated CpG sites exhibiting subtle differences in methylation in samples from breast tissue from non-diseased healthy subjects compared to samples from tumor-adjacent normal breast cancer cases, that are also shown to exhibit differential methylation in tumor tissue itself[13]. Therefore, they represent regions that may be involved in early loss of regulation in breast cancer.

Aside from the probe list, five different breast cancer cohorts were used in this work. The first dataset is referred to as **TCGA**, was used to train and test an OS model for breast cancer, and consisted of gene expression and DNA methylation profiles created by TCGA, using the Illumina HiSeq 2000 and Illumina Infinium HumanMethylation 450 platforms, respectively. Methylation data were functionally normalized[14] using the *minfi* package[15] in R. Expression data were TDM normalized, which makes the distributions between microarray and RNA-seq datasets similar, making them more comparable to the validation datasets[16]. The samples are described in Table S1. After pre-processing, we retained 1028 samples.

The second dataset is referred to as **Terunuma**, is used for external validation of the OS model, and contains 61 tumor samples with clinical and survival data, downloaded from ArrayExpress (E-GEOD-39004)[17]. These data were assayed on the Affymetrix GeneChip Human Gene 1.0 ST Array. They were normalized and then batch corrected in conjunction with the following dataset. A description of the samples is in Table S2.

The third dataset is referred to as **Kao**, is used as a second external validation dataset for the OS model, and contains 327 tumor samples, downloaded from ArrayExpress (E-GEOD-20685)[18]. These data were assayed on the Affymetrix GeneChip Human Genome U133 Plus 2.0. A description of the samples appears in Table S2.

The fourth dataset is referred to as **Hatzis1**, was used to train and test a model of DRFS and another model of pCR, and includes HER2-negative breast cancer cases with neo-adjuvant treatment by taxane-anthracycline (followed by endocrine therapy for ER-positive cases)[12]. These data were assayed on the Affymetrix Human Genome U133A Array. They were downloaded from GEO (GSE25055), and were normalized and then batch corrected in conjunction with the following dataset. A description of the samples appears in Table S3. After pre-processing we retained 306 samples.

The final dataset is referred to as **Hatzis2**, was used for external validation of both the DRFS and pCR models of (HER2-negative) breast cancer, and includes HER2-negative breast cancer cases with neo-adjuvant treatment by taxane-anthracycline (followed by endocrine therapy for ER-positive cases)[12]. These data were assayed on the Affymetrix Human Genome U133A Array. They were downloaded from GEO (GSE25065). A description of the samples appears in Table S3. After pre-processing we retained 182 samples.

In order to reduce categories with very few samples and the number of variables in the model, all AJCC tumor stage classifications were reduced to simple Stage I, II, III, or IV.

Experimental Design. Three different types of models were built: (1) a model of overall survival (OS), (2) a model of distant recurrence-free survival (DRFS), (3) a model of pathologic complete response (pCR) to neoadjuvant taxane/anthracycline treatment. For all model types, we built 100 different models using different random splits of the data, divided into 70% training and 30% testing data sets, to assess how consistently a good model could be found given different training data. We assessed the performance of the models built with and

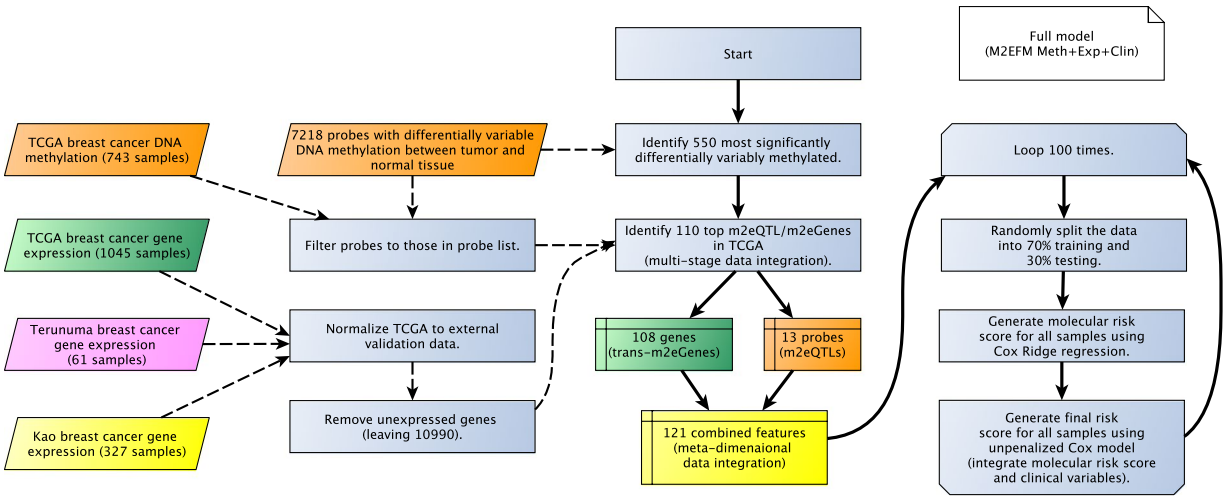

Figure 1. Overview of the M2EFM OS model construction.

without clinical variables, with clinical variables alone, and separately to include external validation data. The external validation data could only be tested on models without DNA methylation values, because methylation data were not available (an overview is shown in Fig. 1). For external validation, a final model using all training data was also built.

M2EFM OS and DRFS models were compared with Cox-Ridge regression, using the *glmnet* package[19] for R, without pre-selection of genes or probes (although probes were pre-filtered by median absolute deviation to reduce them to a number of the most variable probes roughly equal to the number of genes). The M2EFM pCR model was compared to a logistic-Ridge regression model, again using *glmnet*. Furthermore, the M2EFM OS and DRFS models were compared to a model incorporating the PAM50 risk of recurrence score (rorS)[5] with clinical variables. Finally, the M2EFM model was compared to a model built from a gene signature obtained by the network clinical association (NCA) algorithm[20], which is also based on data integration.

Performance of the OS and DRFS models was evaluated using concordance or C-index, a measure of how likely it is, in any given pair of individuals, that the individual with the higher risk score has the event first. For the pCR model, the performance was measured using area under the receiver operating characteristic curve (AUC). Significance tests for comparisons used the two-tailed Wilcoxon signed-rank test.

Clinical variables were chosen using the least absolute shrinkage and selection operator (LASSO) for each outcome. Candidate variables included AJCC stage, patient age at diagnosis, ER status, PR status, HER2 status, and PAM50 subtype (calculated using the *genefu* package for R)[21]. Selected variables include AJCC stage and patient age at diagnosis for all outcomes, although stage was reduced to simply presence or absence of stage III cancer for DRFS and pCR outcomes due to the absence of stage IV cases and small number of stage I cases for these outcomes.

The m2eGenes were used to analyze the functional relevance of our approach using the online tool WEB-based GEne SeT AnaLysis Toolkit (WebGestalt 2017)[22]. Using WebGestalt, we performed protein network topology-based analysis and enrichment analysis for GO biological process terms. Furthermore, we tested the enrichment of m2eGenes in a database of genes with causal cancer connections, the Catalog of Somatic Mutations in Cancer (COSMIC) Cancer Gene Census[23], using Fisher's exact test. Finally, we tested the enrichment of genes in our signature in genes known to be targeted by breast cancer drugs listed in the DrugBank database[24], again using Fisher's exact test.

Data availability. The datasets generated and/or analyzed during the current study are available in the following repositories:

- TCGA: UCSC Cancer Genomics Browser
- Terunuma: ArrayExpress (E-GEOD-39004)
- Kao: ArrayExpress (E-GEOD-20685)
- Hatzis1: GEO (GSE25055)
- Hatzis2: GEO (GSE25065)
- Cancer Genes: COSMIC Cancer Gene Census
- Cancer Drug Targets: DrugBank

M2EFM is available as a package for R at https://github.com/jeffreyat/M2EFM.

Results

We applied M2EFM to the TCGA dataset to select dysregulated probes and genes for our models. The lists of probes and genes selected by M2EFM are shown in Table S4.

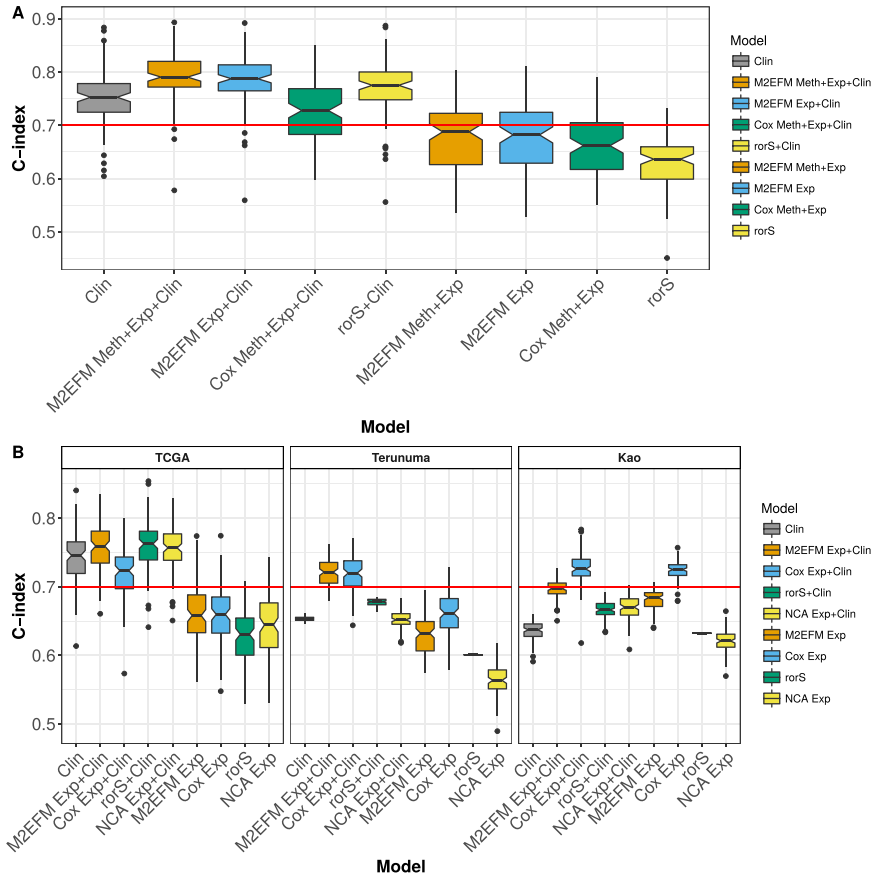

Figure 2. C-index for full overall survival models across 100 random splits of TCGA data into training and testing data for the various approaches. M2EFM models are the most consistent performers, achieving a median C-index of at least 0.7 in both internal and external validation data in all cases. (**A**) The M2EFM model with gene expression, DNA methylation, and clinical data (M2EFM Meth + Exp + Clin) was statistically significantly stronger than any other model. (**B**) The models with external validation data show that only the M2EFM and Cox-Ridge models consistently validate at a high median C-index.

Overall Survival – Meta-dimensional Models Without External Validation Data. Results of the meta-dimensional double-integrated model, (Fig. 1 and fully described in the Supplemental Methods) are shown in Fig. 2A and Table S5, allowing a comparison to how frequently equivalent strength results could be obtained from random gene sets. Across the 100 random splits of the training and testing data, the double integrated model containing methylation values from m2eQTL, expression values from m2eGenes, and clinical covariates (M2EFM Meth + Exp + Clin) had the highest median C-index at 0.790, followed by the model incorporating PAM50 rorS values for the samples and clinical variables (rorS + Clin) at 0.775, then the clinical variables only model (Clin) at 0.753, and finally the Cox-Ridge regression with an additional regression integrating unpenalized clinical variables (Cox Meth + Exp + Clin) at 0.728. The C-index scores across the splits were statistically significantly greater for the M2EFM Meth + Exp + Clin model than rorS + Clin, Clin, or Cox Meth + Exp + Clin ($p = 1.84e\text{-}08$, 1.26e-12, and 3.23e-16 respectively). For models built from only the molecular risk score, M2EFM Meth + Exp had the highest median C-index although scores tended to be lower than models incorporating clinical data. M2EFM Meth + Exp results were significantly greater than rorS and Cox Meth + Exp ($p = 3.44e\text{-}11$, and 1.13e-02 respectively).

Overall Survival - Expression Only Models with External Validation Data. To allow external validation of our models using other datasets we next focused on breast cancer OS, building models using only expression data, with the m2eQTL discovered previously. In such cases, we substitute the *cis*-m2eGenes (genes proximal to the differentially methylated loci) for the associated methylation values. We also added a model built from a gene signature discovered using NCA (except for two genes: *HES5* and *NKIRAS1*, which were not in our profiles).

The median C-index for each model on each dataset is shown in Table S6, and the distribution of C-indices is shown in Fig. 2B. On TCGA, the median C-index for rorS + Clin was the highest (0.762), followed closely by M2EFM Exp + Clin (0.759) and NCA Exp + Clin (0.757), with the other models further behind. The differences were not significant for rorS + Clin compared to M2EFM Exp + Clin ($p = 6.76e\text{-}02$) but were for NCA Exp + Clin, Clin, and Cox Exp + Clin ($p = 3.99e\text{-}02$, 3.26e-11, and 5.80e-13 respectively). However, the M2EFM model was the most consistent performer over all the validation data. On the Terunuma external validation data,

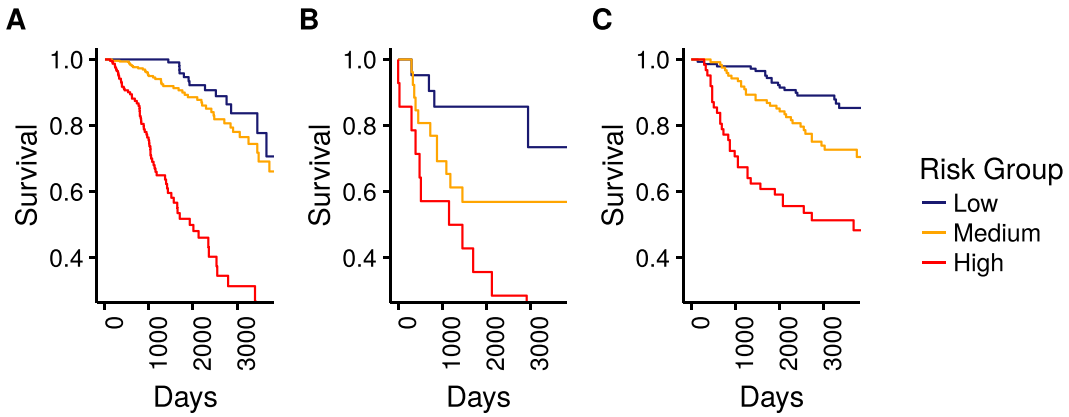

Figure 3. Kaplan-Meier plots of survival in the overall survival datasets shown out to 10 years. In each dataset, the samples were split into three risk groups based on their predicted M2EFM risk score: (**A**) TCGA, (**B**) Terunuma, (**C**) Kao. The three risk groups (denoted Low, Medium, and High), appear to have good discrimination on all datasets. In each case, the Low risk group had the best survival and the High risk group the worst.

	HR	95% CI	p-value
TCGA Medium vs. Low	1.77	[1.00, 3.14]	4.90e-02
TCGA High vs. Low	10.20	[5.86, 17.77]	2.20e-16
Terunuma Medium Vs. Low	2.96	[0.93, 9.43]	6.57e-02
Terunuma High vs. Low	6.20	[1.96, 19.66]	1.90e-03
Kao Medium vs. Low	2.05	[1.16, 3.63]	1.40e-02
Kao High vs. Low	5.59	[3.17, 9.85]	2.60e-09

Table 1. Hazard Ratios for Overall Survival Model Risk Groups.

M2EFM Exp + Clin validated better than any other model, with a median C-index of 0.721, followed by Cox Exp + Clin at 0.720, while the other models trailed behind. The M2EFM Exp + Clin C-index scores were not significantly greater than Cox Exp + Clin ($p = 8.17e-01$), but were significantly greater than rorS + Clin ($p < 2.20e-16$) and NCA Exp + Clin ($p < 2.20e-16$). In the Kao data, the best performer was Cox Exp (0.726), followed by Cox Exp + Clin (0.725), then M2EFM Exp + Clin (0.697), followed by the other models. The C-index scores for Cox Exp were significantly stronger than M2EFM Exp + Clin ($p = 3.14e-16$), but M2EFM Exp + Clin scores were still significantly stronger than either rorS + Clin ($p < 2.20e-16$) or NCA Exp + Clin ($p < 2.20e-16$).

Full Training Data Model for Overall Survival. We used the full training data set (TCGA), to build a final overall survival model and validated it on Terunuma and Kao. For Terunuma, the C-index was 0.724, 95% CI [0.499, 0.874]. For Kao, the C-index was 0.703, 95% CI [0.582, 0.800]. These scores were very similar to their median C-index scores. The hazard ratios for this model are given in Table S7. We further divided the samples into three groups based on their risk score, with low risk below the 25th percentile, medium risk from the 25th to the 75th percentile, and high risk above the 75th percentile, and created Kaplan-Meier plots of the survival of the risk groups in each dataset (Fig. 3). Hazard ratios comparing risk groups for the risk score are shown in Table 1.

We generated calibration curves for the proportional hazards models based on the full training data model (Fig. S1), using the *pec* package for R[25], as well as the integrated Brier score (IBS)[26], a quantitative measure of calibration or prediction error. For M2EFM Exp + Clin, the IBS on Kao was 0.060 and on Terunuma was 0.166. For Cox Exp + Clin, the IBS on Kao was 0.055 and on Terunuma was 0.161, indicating that the Cox Exp + Clin models were better calibrated over all time points.

ER Status. Most prognostic markers for breast cancer have limited utility for patients with ER-negative tumors[27]. Unfortunately, ER status was not annotated in the Kao data, and there were too few samples in Terunuma to effectively test the prognostic utility of our marker on ER-negative patients. Therefore, we inferred the ER, PR, and HER2 status of samples in Kao by clustering samples using expression probes associated with each receptor status using a Gaussian mixture model to capture each bimodal distribution[28]. We then assessed the relative prognostic power of the M2EFM models in samples that varied by ER or triple negative status (Table 2). The overall C-index was 0.703, but for ER-positive patients it was 0.665, for ER-negative patients it was 0.727, and for triple negative patients it was 0.682.

Subtypes. Given that the covariates in our model are unpenalized, we needed to select a limited number of clinical covariates to avoid overfitting the model. This resulted in models without PAM50 intrinsic subtypes, ER,

	C-index	95% CI
Overall	0.703	[0.582, 0.800]
ER-Positive	0.665	[0.489, 0.806]
ER-Negative	0.727	[0.558, 0.849]
Triple Negative	0.682	[0.388, 0.879]

Table 2. Performance of M2EFM on Kao Samples by ER and Triple Negative Status.

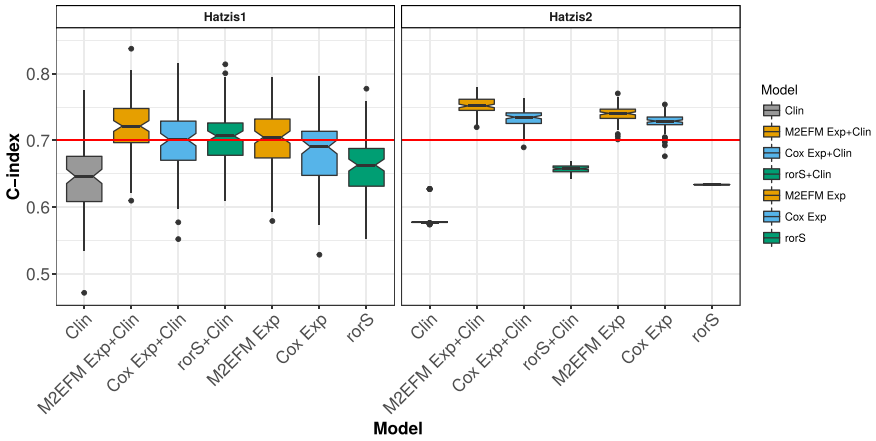

Figure 4. C-index for expression-only DRFS models across 100 random splits of Hatzis1 data into training and testing data for the various approaches, which were then validated on Hatzis2. The M2EFM models were the best performing on both testing and external validation data.

PR, or HER2 status, which are all known to be associated with breast cancer prognosis. Therefore, we regressed the M2EFM molecular risk score (calculated for all TCGA samples) on these variables and calculated the variance inflation factor (VIF), a commonly used measure of the multicollinearity between variables. The VIF was approximately 1.20. A VIF between 1 and 5 is generally regarded to imply only moderate multicollinearity. Therefore, the molecular risk score from M2EFM contains prognostic information independent of breast cancer subtype.

Distant Recurrence-Free Survival (DRFS) Models. On both the testing data and the external validation data, M2EFM was the most prognostic model of DRFS (Fig. 4 and Table S8). On the testing data, the M2EFM Exp + Clin had a median C-index score of 0.722, followed by rorS + Clin (0.707), then M2EFM Exp (0.705), and Cox Exp + Clin (0.701), with the other models below 0.7. In the external validation data, M2EFM Exp + Clin and M2EFM Exp were the most prognostic at 0.752 and 0.741 respectively, although Cox Exp + Clin and Cox Exp were strong at 0.735 and 0.729 respectively.

Finally, an M2EFM Exp + Clin model trained on the full training data (Hatzis1) achieved a C-index of 0.766, 95% CI [0.604, 0.875] on the external validation data (Hatzis2). The hazard ratios for this model are given in Table S9. Calibration curves at 3 years are shown in Fig. S2. The IBS score for M2EFM Exp + Clin on Hatzis2 was 0.079. For Cox Exp + Clin it was 0.083, demonstrating that the M2EFM Exp + Clin model is better calibrated.

Negative Predictive Value for Node-Negative Cases. One interest in genomic cancer biomarkers is for determining cases with a low risk of recurrence who may not benefit from adjuvant chemotherapy[29,30]. For these cases, it is important that the marker has a high negative predictive value (NPV). Therefore, we assessed the NPV of the M2EFM model in node-negative samples using the *timeROC* package for R, which uses a time-dependent AUC and can handle censored survival times. A cutoff for the risk score was determined in Hatzis1 that equalized the sensitivity and specificity in predicting DRFS for node-negative cases at 3 years (the median follow-up in these data). This cutoff was applied to node-negative cases in Hatzis2 to determine the positive predictive value (PPV) and NPV. At the 3-year follow up point, the PPV was 0.49 and the NPV was 0.95. We predicted 39 node-negative cases would not have a distant metastasis by 3 years and we were correct for 36. We predicted 12 node-negative cases would have a distant metastasis and were correct for 7.

Pathologic Complete Response (pCR) Models. The final set of models are for pCR following neoadjuvant treatment with Taxane/Anthracycline in HER2-negative breast cancer. Results are shown in Fig. 5 and Table S10. For these data, the clinical models were not very informative. The reference logistic-Ridge regression model performed best on the testing data, with Logistic Exp + Clin scoring a median AUC of 0.767 and Logistic Exp at 0.763 compared to M2EFM Exp + Clin with a median AUC of 0.739 and M2EFM Exp at 0.733. Nevertheless, M2EFM performed best on the external validation data, with M2EFM Exp + Clin at a median AUC of 0.716 and M2EFM Exp at 0.727 compared to Logistic Exp + Clin at a median AUC of 0.707 and Logistic Exp at 0.709.

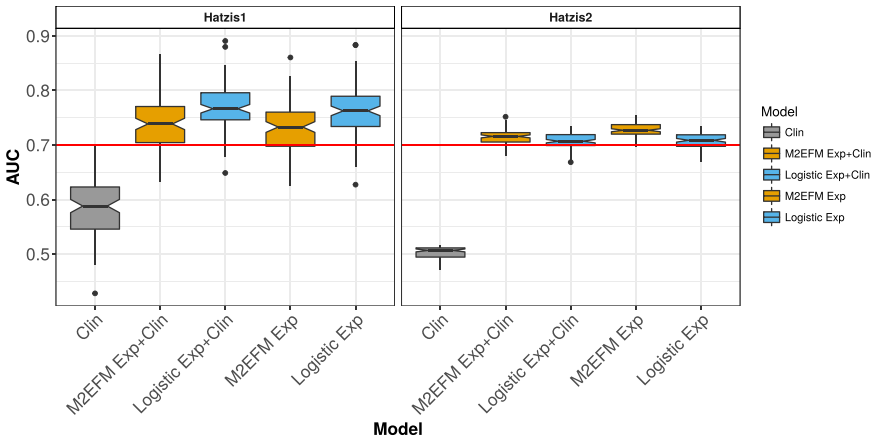

Figure 5. AUC for expression-only pCR models across 100 random splits of Hatzis1 data into training and testing data for the various approaches, which were then validated on Hatzis2. Although Logistic-Ridge models performed better on the internal validation models, M2EFM models performed better on external validation data.

Finally, an M2EFM Exp + Clin model trained on the full training data (Hatzis1) achieved an AUC of 0.720, 95% CI [0.639, 0.802] on the external validation data (Hatzis2). The odds ratios for this model are given in Table S11. Calibration curves appear in Fig. S3. The mean Brier score for M2EFM Exp + Clin on Hatzis2 was 0.184 and for Cox Exp + Clin it was 0.202, showing that M2EFM Exp + Clin is better calibrated.

Functional Analysis. *Protein Network Topology-Based Analysis.* Of the 115 m2eGenes, 55 were found to interact in a sub-network of protein-protein interactions using WebGestalt. The sub-network included 65 genes in all and was enriched for 184 biological process terms, demonstrating that many of the genes in our signature are functionally related. The top result was for lymphocyte activation ($p = 2.50e-12$), which included 23 genes from the sub-network and is about 8.96 times more overlap than would be expected by chance. The sub-network is shown in Fig. S4, visualized using Cytoscape[31].

Biological Process, Disease, and Drug Enrichment. Using Gene Ontology enrichment testing, our gene set shows a clear enrichment for immune functions (Fig. 6). The results further demonstrate that expression tends to be repressed in these pathways in the samples with the highest M2EFM risk scores and more moderately repressed in samples with medium M2EFM risk scores.

Many of the genes in our gene signature are already known to be associated with breast cancer (such as *ESR1*[32] and *EGFR*[33]). We calculated the enrichment of genes in our signature in the Catalog of Somatic Mutations in Cancer (COSMIC) Cancer Gene Census[23], downloaded on 9/13/2016. Enrichment would suggest a causal link between genes in our signature and the progression of the disease. There were 601 genes in the catalog, 515 of which were in our gene expression profiles. The enrichment for causally implicated cancer genes in our gene signature was about 4 times greater than expected by chance (OR: 3.87, $p = 6.09e-06$). The following 18 genes used in our M2EFM models appear in the Cancer Gene Census: *AFF3, BCL11A, CD79A, CD79B, EGFR, ERBB4, ESR1, FOXA1, GATA3, IRF4, ITK, LCK, MYB, POU2AF1, PRF1, PTPRC, RET, TNFRSF17*.

We also examined the known targets of FDA approved breast cancer drugs using the DrugBank database version 5.0[24]. There were 43 targets among the genes in our expression profiles. Of these, 4 genes were found in our signature (*EGFR, ESR1, MAPT*, and *PRKCB*). Thus, our signature was enriched for known targets of breast cancer treatments 10 times more than would be expected by chance (OR: 10.00, $p = 1.03e-03$). These drugs included paclitaxel (*MAPT*), raloxifene (*ESR1*), toremifene (*ESR1*), fulvestrant (*ESR1*), trastuzumab (*EGFR*), tamoxifen (*ESR1*, *PRKCB*), docetaxel (*MAPT*), and lapatinib (*EGFR*).

Discussion

In this work, we have tried to address common shortcomings in molecular biomarker studies, have applied the REMARK guidelines in our assessment[34], and validated a data-integrated approach to creating prognostic and predictive cancer models. The M2EFM approach shows remarkable promise for developing prognostic and predictive cancer biomarkers using both clinical and molecular data. For overall survival, distant recurrence free survival, and pathologic complete response in breast cancer it achieved a median C-index or AUC of at least 0.7 in all internal and external validation datasets (including 5 different cohorts). The fact that our model can reliably attain this level of discrimination across platforms and populations is promising indeed.

Unsurprisingly, one of our comparison methods, a modified Cox-Ridge approach, performed about as well as M2EFM in terms of discriminative power. M2EFM depends on Cox-Ridge itself, and the Cox-Ridge models contain a superset of the predictors in M2EFM. Nevertheless, M2EFM models in many cases outperform the Cox-Ridge models, and more importantly, M2EFM models depend on only 115 genes, as opposed to 10990 in the Cox-Ridge models. Therefore, clinical use for Cox-Ridge would require a transcriptome-wide gene expression assay, and the situation would be even more extreme if the model with methylation values was used. Additionally,

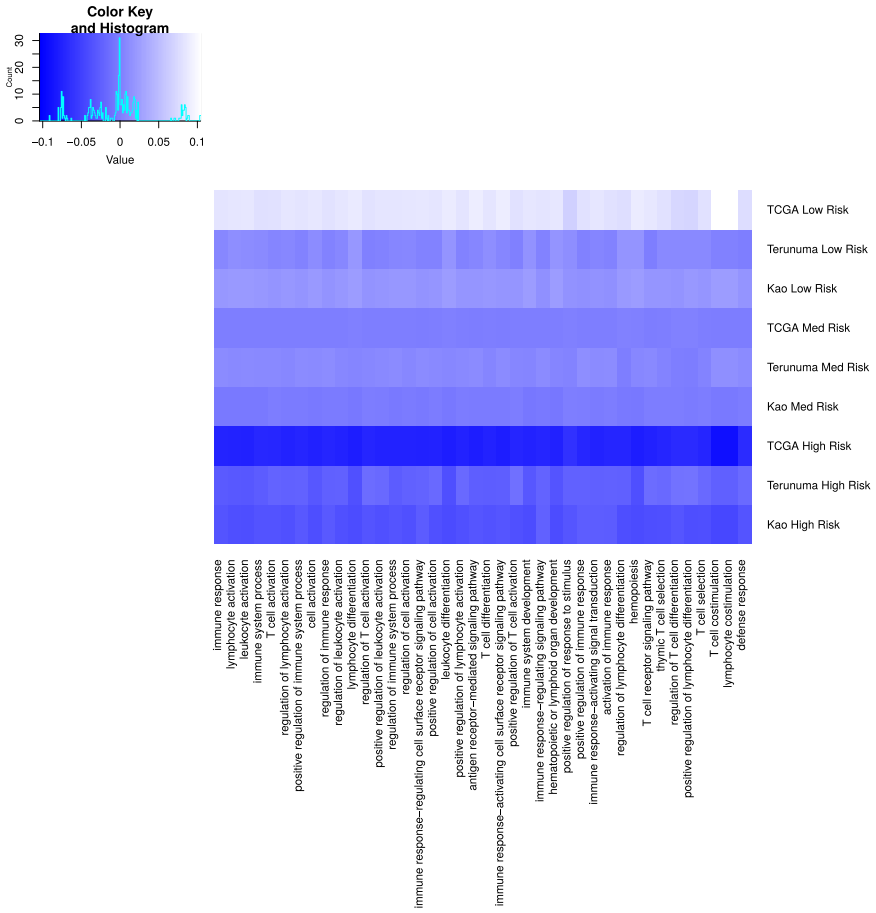

Figure 6. Mean relative expression of biological process pathways enriched for genes in the M2EFM gene signature. The pathways are composed primarily of immune-related functions. The heat map shows the mean relative expression of genes from our signature in each pathway for each overall survival dataset. The data are subset into samples predicted by M2EFM to be at low, medium, or high risk. There is a trend toward repression of the pathways as the risk score goes from low to high.

for DRFS and pCR, M2EFM was better calibrated than Cox-Ridge, suggesting that Cox-Ridge would not be as useful clinically. We also demonstrated that M2EFM has a high NPV for node-negative cases, while maintaining a meaningful PPV, suggesting M2EFM might be useful for identifying patients who could avoid adjuvant chemotherapy. Finally, M2EFM has functional relevance, increasing the possibility that it could serve as the basis for clinical tests that determine eligibility for treatment.

Thus, using three different outcomes and five different data sets, we maintained a consistent level of accuracy. In conjunction with our prior study of M2EFM applied to ccRCC[11], these results consistently show that integrating both molecular and clinical data results in more prognostic models than either data type alone. Our results strongly suggest that a data-integrated approach that models genes affected by regulatory changes in cancer is more reliable than many extant approaches. Furthermore, our results suggest our method builds a model that is prognostic in the case of ER-negative samples and possibly triple negative samples (although unfortunately few samples were available to test this hypothesis). Given the lack of approved prognostic indicators for ER-negative patients, partly due to its greater proliferative rate and the reliance of many tests on proliferative genes[27], and the greater difficulty in treating triple negative disease, a marker, such as ours, that is largely independent of proliferation is desirable[27,35,36].

Our results also show that our approach does more than build a reliable prognostic model: it identifies biologically relevant pathways and, importantly, possible therapeutic targets and genes involved in cancer progression. Our gene signature contains genes that are functionally involved in cancer and that are targeted by current breast cancer treatments. Perhaps more interestingly, given recent advances in immunotherapy[37,38], we note the repression of immune related genes in our signature for patients predicted to be at high risk. There has been growing interest in finding signatures that are independent of proliferation and ER status that might also be predictive of possible immune therapies[27].

Although we provide evidence of the efficacy of M2EFM, there remains one limitation before attempting to translate these results to the clinic: calibration of the model to a single platform. To obtain external validation data, we performed cross-platform normalization. Having established the utility of the M2EFM approach, it remains to train the model on a high-quality dataset generated on the platform that will be used for any clinical assays.

Conclusions

M2EFM can predict risk based on data integrated from multiple sources, taking advantage of tried-and-true prognostic factors while incorporating data on relevant relationships between molecular data types at the individual level. For breast cancer, this resulted in models based on several genes with known functional associations to cancer progression and which are targeted by currently available drugs, and therefore, may result in a useful tool for clinical decision support. This also suggests that other genes in the model should be examined as possible targets, because many of the genes are functionally connected. Given the strength of our findings, we believe that further development of this technique and application to other cancers is warranted.

References

1. Cho, S. H., Jeon, J. & Kim, S. I. Personalized Medicine in Breast Cancer: A Systematic Review. *J Breast Cancer* **15**, 265–272, https://doi.org/10.4048/jbc.2012.15.3.265 (2012).
2. Weinstein, J. N. *et al.* The Cancer Genome Atlas Pan-Cancer analysis project. *Nat Genet* **45**, 1113–1120, https://doi.org/10.1038/ng.2764 (2013).
3. Zhang, J. J. *et al.* International Cancer Genome Consortium Data Portal-a one-stop shop for cancer genomics data. *Database-Oxford* https://doi.org/10.1093/database/bar026 (2011).
4. Edgar, R., Domrachev, M. & Lash, A. E. Gene Expression Omnibus: NCBI gene expression and hybridization array data repository. *Nucleic Acids Res* **30**, 207–210, https://doi.org/10.1093/nar/30.1.207 (2002).
5. Parker, J. S. *et al.* Supervised risk predictor of breast cancer based on intrinsic subtypes. *J Clin Oncol* **27**, 1160–1167, https://doi.org/10.1200/JCO.2008.18.1370 (2009).
6. Paik, S. *et al.* A multigene assay to predict recurrence of tamoxifen-treated, node-negative breast cancer. *N Engl J Med* **351**, 2817–2826, https://doi.org/10.1056/NEJMoa041588 (2004).
7. Coates, A. S. *et al.* Tailoring therapies-improving the management of early breast cancer: St Gallen International Expert Consensus on the Primary Therapy of Early Breast Cancer 2015. *Ann Oncol* **26**, 1533–1546, https://doi.org/10.1093/annonc/mdv221 (2015).
8. Kern, S. E. Why Your New Cancer Biomarker May Never Work: Recurrent Patterns and Remarkable Diversity in Biomarker Failures. *Cancer Res* **72**, 6097–6101, https://doi.org/10.1158/0008-5472.Can-12-3232 (2012).
9. Yuan, Y. *et al.* Assessing the clinical utility of cancer genomic and proteomic data across tumor types. *Nat Biotechnol* **32**, 644−+, https://doi.org/10.1038/nbt.2940 (2014).
10. Ritchie, M. D., Holzinger, E. R., Li, R. W., Pendergrass, S. A. & Kim, D. Methods of integrating data to uncover genotype-phenotype interactions. *Nat Rev Genet* **16**, 85–97, https://doi.org/10.1038/nrg3868 (2015).
11. Thompson, J. A. & Marsit, C. J. A Methylation-to-Expression Feature Model for Generating Accurate Prognostic Risk Scores and Identifying Disease Targets in Clear Cell Kidney Cancer. *Pac Symp Biocomput* **22**, 509–520, https://doi.org/10.1142/9789813207813_0047 (2017).
12. Hatzis, C. *et al.* A Genomic Predictor of Response and Survival Following Taxane-Anthracycline Chemotherapy for InvasiveBreast Cancer. *Jama-J Am Med Assoc* **305**, 1873–1881, https://doi.org/10.1001/jama.2011.593 (2011).
13. Teschendorff, A. E. *et al.* DNA methylation outliers in normal breast tissue identify field defects that are enriched in cancer. *Nat Commun* **7** https://doi.org/10.1038/ncomms10478 (2016).
14. Fortin, J. P. *et al.* Functional normalization of 450k methylation array data improves replication in large cancer studies. *Genome Biol* **15** https://doi.org/10.1186/s13059-014-0503-2 (2014).
15. Aryee, M. J. *et al.* Minfi: a flexible and comprehensive Bioconductor package for the analysis of Infinium DNA methylation microarrays. *Bioinformatics* **30**, 1363–1369, https://doi.org/10.1093/bioinformatics/btu049 (2014).
16. Thompson, J. A., Tan, J. & Greene, C. S. Cross-platform normalization of microarray and RNA-seq data for machine learning applications. *Peerj* **4**, https://doi.org/10.7717/peerj.1621 (2016).
17. Terunuma, A. *et al.* MYC-driven accumulation of 2-hydroxyglutarate is associated with breast cancer prognosis. *J Clin Invest* **124**, 398–412, https://doi.org/10.1172/Jci71180 (2014).
18. Kao, K. J., Chang, K. M., Hsu, H. C. & Huang, A. T. Correlation of microarray-based breast cancer molecular subtypes and clinical outcomes: implications for treatment optimization. *Bmc Cancer* **11**, https://doi.org/10.1186/1471-2407-11-143 (2011).
19. Friedman, J., Hastie, T. & Tibshirani, R. Regularization Paths for Generalized Linear Models via Coordinate Descent. *J Stat Softw* **33**, 1–22 (2010).
20. Martinez-Ledesma, E., Verhaak, R. G. W. & Trevino, V. Identification of a multi-cancer gene expression biomarker for cancer clinical outcomes using a network-based algorithm. *Sci Rep-Uk* **5**, https://doi.org/10.1038/srep11966 (2015).
21. Gendoo, D. M. A. *et al.* Genefu: an R/Bioconductor package for computation of gene expression-based signatures in breast cancer. *Bioinformatics* **32**, 1097–1099, https://doi.org/10.1093/bioinformatics/btv693 (2016).
22. Wang, J., Duncan, D., Shi, Z. & Zhang, B. WEB-based GEne SeT AnaLysis Toolkit (WebGestalt): update 2013. *Nucleic Acids Res* **41**, W77–W83, https://doi.org/10.1093/nar/gkt439 (2013).
23. Forbes, S. A. *et al.* COSMIC: exploring the world's knowledge of somatic mutations in human cancer. *Nucleic Acids Res* **43**, D805–D811, https://doi.org/10.1093/nar/gku1075 (2015).
24. Wishart, D. S. *et al.* DrugBank: a comprehensive resource for *in silico* drug discovery and exploration. *Nucleic Acids Res* **34**, D668–D672, https://doi.org/10.1093/nar/gkj067 (2006).
25. Mogensen, U. B., Ishwaran, H. & Gerds, T. A. Evaluating Random Forests for Survival Analysis Using Prediction Error Curves. *J Stat Softw* **50**, 1–23 (2012).
26. Graf, E., Schmoor, C., Sauerbrei, W. & Schumacher, M. Assessment and comparison of prognostic classification schemes for survival data. *Stat Med* **18**, 2529–2545 (1999).
27. Gyorffy, B. *et al.* 3 Multigene prognostic tests in breast cancer: past, present, future. *Breast Cancer Research* **17** https://doi.org/10.1186/s13058-015-0514-2 (2015).
28. Lehmann, B. D. *et al.* Identification of human triple-negative breast cancer subtypes and preclinical models for selection of targeted therapies. *J Clin Invest* **121**, 2750–2767, https://doi.org/10.1172/Jci45014 (2011).
29. Duffy, M. J. *et al.* Clinical use of biomarkers in breast cancer: Updated guidelines from the European Group on Tumor Markers (EGTM). *Eur J Cancer* **75**, 284–298, https://doi.org/10.1016/j.ejca.2017.01.017 (2017).
30. Loncaster, J. *et al.* Impact of Oncotype DX breast Recurrence Score testing on adjuvant chemotherapy use in early breast cancer: Real world experience in Greater Manchester, UK. *Ejso-Eur J Surg Onc* **43**, 931–937, https://doi.org/10.1016/j.ejso.2016.12.010 (2017).
31. Shannon, P. *et al.* Cytoscape: a software environment for integrated models of biomolecular interaction networks. *Genome Res* **13**, 2498–2504, https://doi.org/10.1101/gr.1239303 (2003).
32. Masuda, H. *et al.* Role of epidermal growth factor receptor in breast cancer. *Breast Cancer Res Treat* **136**, 331–345, https://doi.org/10.1007/s10549-012-2289-9 (2012).
33. Lauring, J. & Wolff, A. C. Evolving Role of the Estrogen Receptor as a Predictive Biomarker: ESR1 Mutational Status and Endocrine Resistance in Breast Cancer. *J Clin Oncol* **34**, 2950-+, https://doi.org/10.1200/Jco.2016.68.4720 (2016).
34. McShane, L. M. *et al.* REporting recommendations for tumor MARKer prognostic studies (REMARK). *Breast Cancer Res Tr* **100**, 229–235, https://doi.org/10.1007/s10549-006-9242-8 (2006).

35. Shao, F., Sun, H. & Deng, C. X. Potential therapeutic targets of triple-negative breast cancer based on its intrinsic subtype. *Oncotarget* **8**, 73329–73344, https://doi.org/10.18632/oncotarget.20274 (2017).
36. Collignon, J., Lousberg, L., Schroeder, H. & Jerusalem, G. Triple-negative breast cancer: treatment challenges and solutions. *Breast Cancer (Dove Med Press)* **8**, 93–107, https://doi.org/10.2147/BCTT.S69488 (2016).
37. Liu, Z., Li, M., Jiang, Z. & Wang, X. A Comprehensive Immunologic Portrait of Triple-Negative Breast Cancer. *Transl Oncol* **11**, 311–329, https://doi.org/10.1016/j.tranon.2018.01.011 (2018).
38. Tolba, M. F. & Omar, H. A. Immunotherapy, an evolving approach for the management of triple negative breast cancer: Converting non-responders to responders. *Crit Rev Oncol Hematol*. https://doi.org/10.1016/j.critrevonc.2018.01.005 (2018).

Acknowledgements

The analyses described in this report were supported by NIH grants R01ES022223, P30CA138292, P30ES019776, and R01DE022772.

Author Contributions

J.T. and C.M. conceived of and designed the study. J.T. collected the data, performed the analyses, and drafted the article. J.T., C.M. and B.C. performed critical revision. All authors read and approved the final manuscript.

Additional Information

Supplementary information accompanies this paper at https://doi.org/10.1038/s41598-018-23494-0.

Competing Interests: The authors declare no competing interests.

Publisher's note: Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.