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Preliminary estimation of the basic reproduction number of novel coronavirus (2019-nCoV) in China, from 2019 to 2020: A data-driven analysis in the early phase of the outbreak



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ABSTRACT

Backgrounds: An ongoing outbreak of a novel coronavirus (2019-nCoV) pneumonia hit a major city in China, Wuhan, December 2019 and subsequently reached other provinces/regions of China and other countries. We present estimates of the basic reproduction number, R_0 , of 2019-nCoV in the early phase of the outbreak.

Methods: Accounting for the impact of the variations in disease reporting rate, we modelled the epidemic curve of 2019-nCoV cases time series, in mainland China from January 10 to January 24, 2020, through the exponential growth. With the estimated intrinsic growth rate (γ), we estimated R_0 by using the serial intervals (SI) of two other well-known coronavirus diseases, MERS and SARS, as approximations for the true unknown SI.

Findings: The early outbreak data largely follows the exponential growth. We estimated that the mean R_0 ranges from 2.24 (95%CI: 1.96–2.55) to 3.58 (95%CI: 2.89–4.39) associated with 8-fold to 2-fold increase in the reporting rate. We demonstrated that changes in reporting rate substantially affect estimates of R_0 . *Conclusion*: The mean estimate of R_0 for the 2019-nCoV ranges from 2.24 to 3.58, and is significantly larger than 1. Our findings indicate the potential of 2019-nCoV to cause outbreaks.

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Introduction

The atypical pneumonia case, caused by a novel coronavirus (2019-nCoV), was first reported and confirmed in Wuhan, China on December 31, 2019 (World Health Organization, 2020a). As of January

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To the best of our knowledge, there is no existing peer-reviewed literature quantifying the transmissibility of 2019-nCoV as of January 22, 2020. In this study, we estimated the transmissibility of 2019-nCoV via the basic reproduction number, R_0 , based on the limited data in the early phase of the outbreak.

Methods

We obtained the number of 2019-nCoV cases time series data in mainland China released by Wuhan Municipal Health Commission, China and National Health Commission of China from January 10 to January 24, 2020 from (Wuhan Municipal Health Commission, China, 2020). All cases were laboratory confirmed following the case definition by the National Health Commission of China (National Health Commission of the People's Republic of China, 2020). Although the date of submission of this study is January 26, we choose to use data up to January 24. Note that the data of the most recent few days contain a number of infections that were infected outside Wuhan due to travel, and thus this part of the infections is excluded from the analysis.

Although there were cases confirmed on or before January 16, the official diagnostic protocol was released by WHO on January 17 (World Health Organization, 2020b). To adjust the impact of this event, we considered a time-varying reporting rate that follows a linear increasing trend, motivated by the previous study (Wu et al., 2010). We assumed that the reporting rate, r(t), started increasing on January 17, and stopped at the maximal level on January 21. The reporting rate increase corresponds to accounts for the announcement on improving the 2019-nCoV surveillance of the Hubei provincial government (Hubei provincial government, 2020). The length of the reporting increasing part roughly equals the average of the incubation periods of two other well-known coronavirus diseases, i.e., the Middle East Respiratory Syndrome (MERS) and the Severe Acute Respiratory Syndrome (SARS), i.e., 5 days (Bauch et al., 2005; Lin et al., 2018; Donnelly et al., 2003). Denoting the daily reported number of new cases by *c*(*t*) for the *t*-th day, then the adjusted cumulative number of cases, C(t), is $C(t) = \sum_{\tau=0}^{t} c(\tau)/r(\tau)$. Instead of finding the exact value of r(t), we calculated the fold change in r(t) that is defined by the ratio of r on January 10 over that on January 24 minus 1. We illustrated six scenarios with 0- (no change), 0.5-, 1-, 2-, 4- and 8-fold increase in reporting rate, see Figure 1(a), (c), (e), (g), (i) and (k).

Following previous studies (Zhao et al., 2019; de Silva et al., 2009), we modelled the epidemic curve obeying the exponential growth. The nonlinear least square (NLS) framework is adopted for data fitting and parameter estimation. The intrinsic growth rate (γ) of the exponential growth was estimated, and the basic reproduction number could be obtained by $R_0 = 1/M(-\gamma)$ with 100% susceptibility for 2019-nCoV at this early stage. The function $M(\cdot)$ is the Laplace transform, i.e., the moment generating function, of the probability distribution for the serial interval (SI) of the disease (Zhao et al., 2019; Wallinga and Lipsitch, 2007), denoted by h(k) and k is the mean SI. Since the transmission chain of 2019nCoV remains unclear, we adopted the SI information from SARS and MERS, which share a similar pathogen as 2019-nCoV. We modelled h(k) as Gamma distributions with a mean of 7.6 days and standard deviation (SD) of 3.4 days for MERS (Assiri et al., 2013), and mean of 8.4 days and SD of 3.8 days for SARS (Lipsitch et al., 2003) as well as their average, see the row heads in Table 1 for each scenario

Results and discussion

The exponential growth fitting results are shown in Figure 1(b), (d), (f), (h), (j) and (l). The coefficient of determination, R-squared, ranges from 0.91 to 0.92 for all reporting rate changing scenarios, which implies that the early outbreak data were largely following the exponential growth. In Table 1, we estimated that the R_0 ranges from 2.24 (95%CI: 1.96-2.55) to 5.71 (95%CI: 4.24-7.54) associated with an 8-fold to 0-fold increase in the reporting rate. All R_0 estimates are significantly larger than 1, which indicates the potential of 2019-nCoV to cause outbreaks. Since the official diagnostic protocol was released by WHO on January 17 (World Health Organization, 2020b), an increase in the diagnosis and reporting of 2019-nCoV infections probably occurred. Thereafter, the daily number of newly reported cases started increasing around January 17, see Figure 1, which implies that more infections were likely being diagnosed and recorded. We suggested that



Figure 1. The scenarios of the change in the reporting rate (top panels) and the exponential growth fitting (bottom panels). The top panels, i.e., (a), (c), (e), (g), (i) and (k), show the assumed change in the reporting rate. The bottom panels, i.e., (b), (d), (f), (h), (j) and (l), show the reported (or observed, green circles), adjusted (blue dots) and fitted (blue curve) number of 2019-nCoV infections, and the blue dashed lines are the 95%CI. The vertical grey line represents the date of January 16, 2020, after which the official diagnostic protocol was released by the WHO (World Health Organization, 2020). Panels (a) and (b) show the scenarios that the reporting rate increased by 0.5-fold. Panels (e) and (f) show the scenarios that the reporting rate increased by 1-fold. Panels (g) and (h) show the scenarios that the reporting rate increased by 2-fold. Panels (i) and (j) show the scenarios that the reporting rate increased by 2-fold. Panels (i) and (j) show the scenarios that the reporting rate increased by 8-fold.

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The summary table of the estimated basic reproduction number, *R*₀, under different scenarios. The estimated *R*₀ is shown as in the 'median (95%Cl)' format. The 'reporting rate increased' indicates the number of fold increase in the reporting rate from January 17, when WHO released the official diagnostic protocol (World Health Organization, 2020b), to January 20, 2020.

Reporting rate increased	Estimated R ₀		
	Same as MERS SI 7.6 \pm 3.4	SI in average 8.0 \pm 3.6	Same as SARS SI 8.4 \pm 3.8
(unchanged)	5.31 (3.99-6.96)	5.71 (4.24-7.54)	6.11 (4.51-8.16)
0.5-fold	4.52 (3.49-5.76)	4.82 (3.69-6.20)	5.14 (3.90-6.67)
1-fold	4.01 (3.17-5.02)	4.26 (3.34-5.38)	4.53 (3.51-5.76)
2-fold	3.38 (2.75-4.12)	3.58 (2.89-4.39)	3.77 (3.02-4.67)
4-fold	2.73 (2.31-3.22)	2.86 (2.40-3.39)	3.00 (2.50-3.58)
8-fold	2.16 (1.90-2.45)	2.24 (1.96-2.55)	2.32 (2.02-2.66)

Note: 'SI' is serial interval. 'MERS' is Middle East Respiratory Syndrome, and 'SARS' is the Severe Acute Respiratory Syndrome.

changes in reporting might exist, and thus it should be considered in the estimation, i.e., 8-, 4- and 2-fold changes are more likely than no change in the reporting efforts. Although six scenarios about the reporting rate were explored in this study, the real situation is difficult to determine given limited data and (almost) equivalent model fitting performance in terms of R-squared. However, with increasing reporting rate, we found the mean R_0 is likely to be between 2 and 3.

Our analysis and estimation of R_0 rely on the accuracy of the SI of 2019-nCoV, which remains unknown as of January 25. In this work, we employed the SIs of SARS and MERS as approximations to that of 2019-nCoV. The determination of SI requires knowledge of the chain of disease transmission that needs a sufficient number of patient samples and periods of time for follow-up (Cowling et al., 2009), and thus this is unlikely to be achieved shortly. However, using SIs of SARS and MERS as approximation could provide an insight to the transmission potential of 2019-nCoV at the early stage of the outbreak. We reported that the mean R_0 of 2019-nCoV is likely to be from 2.24 (8-fold) to 3.58 (2-fold), and it is largely in the range of those of SARS, i.e., 2-5 (Bauch et al., 2005; Lipsitch et al., 2003; Wallinga and Teunis, 2004), and MERS, i.e., 2.7-3.9 (Lin et al., 2018).

We note that WHO reported the basic reproduction number for the human-to-human (direct) transmission ranged from 1.4 to 2.5 (World Health Organization, 2020b), which is marginally lower than ours. However, many of the existing online preprints estimate the mean R_0 ranging from 2 to 5 (Imai et al., 2020; Riou and Althaus, 2020; Read et al., 2020; Shen et al., 2020), which is largely consistent with our results.

Conclusion

We estimated the mean R_0 of 2019-nCoV ranging from 2.24 (95%CI: 1.96-2.55) to 3.58 (95%CI: 2.89-4.39) if the reporting effort has been increased by a factor of between 8- and 2-fold, respectively, after the diagnostic protocol released on January 17, 2020 and many medical supplies reached Wuhan.

Ethics approval and consent to participate

The ethical approval or individual consent was not applicable.

Availability of data and materials

All data and materials used in this work were publicly available.

Consent for publication

Not applicable.

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Disclaimer

The funding agencies had no role in the design and conduct of the study; collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; or decision to submit the manuscript for publication.

Conflict of interests

The authors declare no competing interests.

Authors contributions

All authors conceived the study, carried out the analysis, discussed the results, drafted the first manuscript, critically read and revised the manuscript, and gave final approval for publication.

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