

1 **Using sero-epidemiology to monitor disparities in vaccination and infection with**
2 **SARS-CoV-2**

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18
19 **Abstract**

20 *Background:* As COVID-19 vaccines continue to be rolled-out, the “double burden” of health
21 disparities in both exposure to infection and vaccination coverage intersect to determine the
22 current and future patterns of infection, immunity, and mortality. Serology provides a unique
23 opportunity to measure biomarkers of infection and vaccination simultaneously, and to relate
24 these metrics to demographic and geographic factors.

25
26 *Methods:* Leveraging algorithmically selected residual serum samples from two hospital
27 networks in San Francisco, we sampled 1014 individuals during February 2021, capturing
28 transmission during the first 11 months of the epidemic and the early roll out of vaccination.
29 These samples were tested using two serologic assays: one detecting antibodies elicited by
30 infection, and not by vaccines, and one detecting antibodies elicited by both infection and
31 vaccination. We used Bayesian statistical models to estimate the proportion of the
32 population that was naturally infected and the proportion protected due to vaccination.

33
34 *Findings:* We estimated that the risk of prior infection of Latinx residents was 5.3 (95% CI:
35 3.2 - 10.3) times greater than the risk of white residents aged 18-64 and that white San
36 Francisco residents over the age of 65 were twice as likely (2.0, 95% CI: 1.1 - 4.6) to be
37 vaccinated as Black residents. We also found socioeconomically deprived zipcodes in the
38 city had high probabilities of natural infections and lower vaccination coverage than wealthier
39 zipcodes.

40
41 *Interpretation:* Using a platform we created for SARS-CoV-2 serologic data collection in San
42 Francisco, we characterized and quantified the stark disparities in infection rates and
43 vaccine coverage by demographic groups over the first year of the pandemic. While the
44 arrival of the SARS-CoV-2 vaccine has created a ‘light at the end of the tunnel’ for this
45 pandemic, ongoing challenges in achieving and maintaining equity must also be considered.

46
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48 the Chan Zuckerberg Biohub.

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51 **Background**

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53 During the initial waves of the COVID-19 pandemic, disparities in disease burden were
54 largely driven by differences in infection rates, as a result of factors including occupation,
55 ability to shelter in place or to take sick leave, access to testing, housing status and
56 crowding, and neighborhood exposure. In addition to driving disparities in infection rates with
57 this novel virus, existing structural inequalities are associated with disparities in the risk of
58 comorbidities such as diabetes and heart disease (as a result of factors such as, but not
59 limited to, nutrition, access to exercise and increased stress), which increase the likelihood
60 of hospitalization and death from COVID-19, and with disparities in access to healthcare
61 both in managing comorbidities and in accessing care for COVID-19.

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63 As vaccine roll-outs advance in the United States and globally, there are disparities in both
64 vaccine access and uptake. These disparities are multifactorial and complex, including
65 reduced technology access and literacy¹, barriers in access to healthcare, concern about the
66 safety of the vaccines², mistrust as a result of a history of medical racism and discrimination,
67 and poor access to reliable information about the vaccine. In the age of vaccination,
68 policymakers must understand the way in which societal structures affect disparities in both
69 infection and vaccination. These disparities may interact to affect both population level
70 immunity and the burden of COVID-19 in different communities. This is relevant both in the
71 present and in the future, as policymakers consider the continued roll-out of vaccines in the
72 context of new variants, as well as preparing for and responding to other diseases.

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75 Given the high levels of disease under-ascertainment, serology (i.e., the measurement of
76 antibodies) has been particularly useful for understanding SARS-CoV-2 infection levels in
77 the population. When there is variability in testing rates and healthcare seeking behavior,
78 serology is an even more useful tool. Serology provides a unique opportunity to measure
79 biomarkers of infection and vaccination simultaneously, and to relate these metrics to
80 demographic and geographic factors. In settings where vaccines based on the SARS-CoV-2
81 spike protein (e.g., currently available mRNA or adenovirus vector vaccines) are used,
82 measuring long-lived antibody responses to both spike and non-spike proteins can be used
83 to disentangle immune responses elicited by infection from vaccination. While structural
84 inequalities are by no means limited to the United States, here we focus on a case example
85 leveraging serology to understand inequalities from a domestic perspective.

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88 Even in San Francisco, a city which has had a relatively successful early and sustained
89 COVID-19 response, and has achieved high vaccination coverage over the past few months,
90 reported case counts of COVID-19 and hospitalization rates have been higher in
91 socioeconomically deprived areas, amongst homeless individuals, and within the city's
92 Latinx and Black communities^{3,4}. Disparities in vaccination coverage have also been
93 reported, particularly in the early months of vaccine roll-out, creating a double burden for
94 some vulnerable communities. In San Francisco, whilst some disparities have now been
95 addressed, vaccination remains much lower in homeless individuals and in Black/African
96 American individuals⁵.

97

98 To measure disparities in both infection rates and vaccination, we leveraged a SARS-CoV-2
99 serosurveillance platform launched in March 2020 that utilizes residual blood samples taken
100 from two hospital networks in San Francisco. Estimates derived from this platform during the
101 first wave of the pandemic showed seroprevalence in Latinx individuals to be nearly two
102 times higher than in white individuals, and nearly two times higher in homeless individuals
103 than the population average⁶. We collected samples from 1,014 individuals undergoing

104 routine blood draws between February 4 and February 17, 2021, capturing transmission
105 during the first 11 months of the epidemic and the early roll-out of vaccination for those over
106 65 years old. These samples were tested using two serologic assays: one detecting
107 antibodies to SARS-CoV-2 elicited by infection and not by vaccines currently used in the US,
108 and one detecting antibodies to SARS-CoV-2 elicited by both infection and vaccination. We
109 used Bayesian statistical models to estimate the proportion of the population that was
110 seropositive due to natural infection and the proportion seropositive due to vaccination,
111 stratified by age, race and ZIP code of residence.

112 **Methods**

113 Data

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117 As part of an existing serological survey⁶, residual serum samples from routine blood draws
118 from the University of California, San Francisco (UCSF) and San Francisco Department of
119 Public Health (SFDPH) inpatient and outpatient healthcare systems were sampled between
120 February 4 and February 17, 2021. A total of 1,091 samples were collected, of which 77
121 were excluded due to participation in a separate COVID research study, and a further 15
122 were later excluded from further analyses as they could not be linked to antibody test results.
123 The characteristics of the samples are illustrated in Table 1. The full inclusion and exclusion
124 criteria and sampling algorithm are described previously⁶.

125 Laboratory analysis

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128 Each sample (N = 1014) was tested on two commercial SARS-CoV-2 serologic platforms.
129 The Ortho Clinical Diagnostics VITROS Anti-SARS-CoV-2 Total assay measures the total Ig
130 antibody response to the S1 subunit of the SARS-CoV-2 spike (S) protein and therefore is
131 expected to yield a positive results after natural infection or vaccination⁷. The Roche Elecsys
132 Anti-SARS-CoV-2 assay measures the total Ig antibody response to the SARS-CoV-2
133 nucleocapsid (N) protein⁸ and therefore is expected to yield a positive results after natural
134 infection but not after vaccination with vaccines based on the spike protein. In a previous
135 analysis assessing the test performance characteristics of many SARS-CoV-2 serological
136 assays, we found both assays to exhibit high sensitivity over time following infection⁹.

137 Univariate data analysis (by assay)

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140 Seropositivity on the Vitros assay indicates whether or not an individual has had any prior
141 immune response to SARS-CoV-2, either through natural infection and/or vaccination. The
142 SARS-CoV-2 mRNA and adenovirus vector vaccines elicit immune responses to only the S
143 protein of the virus. Therefore, in contexts where these vaccines are used exclusively (such
144 as the United States), seropositivity on the Roche assay indicates whether an individual has
145 had a prior immune response to SARS-CoV-2 via infection. Assuming perfect test
146 performance characteristics, the difference between the proportion seropositive on Vitros
147 and the proportion seropositive on Roche indicates the proportion of the population that has
148 been vaccinated and has not been infected.

149
150 We used Binomial models in a Bayesian framework to first estimate seropositivity separately
151 by assay. We adjusted for the manufacturer-reported specificity of each assay (100% for
152 Vitros and 99.80% for Roche) and for in-house estimates of the sensitivity of each assay at 2
153 months post symptom onset among non-hospitalized individuals based on a longitudinal
154 post-infection study⁹, corresponding to 83.8% for Vitros and 90.0% for Roche. We note that
155 sensitivity was particularly consistent over time following infection for these assays, so our
156 results are robust to the choice of exact time point used. We computed 95% credible
157 intervals (Crl) to quantify uncertainty in posterior estimates. For these univariate analyses,

158 age was stratified into 4 groups (0-17 y, 18-34 y, 35-64 y, and 65+ y) and race/ethnicity was
159 stratified into 5 groups (Asian, Black, Latinx, White, and Other).

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161 Bivariate data analysis (infection vs. vaccination)

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163 We then conducted bivariate analyses using the results of each sample on both assays. We
164 used the data set of individuals who had a test result on both the Vitros and Roche assays,
165 and removed the 81 samples that had a result on only one assay. In addition, we removed
166 the 3 samples that tested negative on Vitros and positive on Roche, which likely reflects a
167 false negative result on the Vitros assay and/or a false positive result on the Roche assay
168 (Supplementary Table 1).

169

170 *For demographic analyses:* Age was stratified into 2 groups (18-64 y and 65+ y), and
171 race/ethnicity was stratified into 4 groups (Asian, Black, Latinx, and White). We omitted
172 individuals aged 0-17 y from this portion of the analysis, as individuals in that age group
173 were not eligible for vaccination during this time frame in San Francisco (February 2021).
174 The decision to choose 65 y as a cutoff was due to the age-based roll-out of SARS-CoV-2
175 vaccination and the expected resultant differences in vaccine coverage by age. We also
176 omitted individuals with race/ethnicity of "Other" due to lack of data on vaccine doses in this
177 demographic group that are used as a prior for estimation (see below). For each
178 combination of age group and race/ethnicity j , we estimated the marginal probabilities of
179 natural infection, $\text{Pr}(\text{inf})$, and of vaccination, $\text{Pr}(\text{vacc})$, separately as follows.

180

$$181 n_{\text{Vitros +ve \& Roche +ve}, j} \sim \text{Binomial}(N_j, \text{Pr}(\text{inf}))$$

$$182 n_{\text{Vitros +ve \& Roche -ve}, j} \sim \text{Binomial}(N_j, \text{Pr}(\text{vacc}|\text{uninf}) * [1 - \text{Pr}(\text{inf})])$$

$$183 n_{\text{Vitros -ve \& Roche -ve}, j} \sim \text{Binomial}(N_j, [1 - \text{Pr}(\text{vacc}|\text{uninf})] * [1 - \text{Pr}(\text{inf})])$$

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$$185 \text{Pr}(\text{vacc}) \sim \text{Beta}(\mu_j, \kappa_j)$$

$$186 \text{Pr}(\text{vacc}) = \text{Pr}(\text{vacc}|\text{inf}) * \text{Pr}(\text{inf}) + \text{Pr}(\text{vacc}|\text{uninf}) * [1 - \text{Pr}(\text{inf})]$$

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[INSERT FIGURE 1 HERE]

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201 N_j represents the number of individuals in our data set in group j who were included, and $n_{x,j}$
202 represents the number of individuals in group j with antibody results x . For this analysis, we
203 assumed perfect test performance characteristics. As $\text{Pr}(\text{vacc}|\text{inf})$ is not identified by our
204 data, and we do not assume that $\text{Pr}(\text{vacc}|\text{inf}) = \text{Pr}(\text{vacc}|\text{uninf})$, we set up a process to
205 estimate the hyper-priors μ_j (mean) and κ_j (precision) for each group j based on reported
206 vaccination coverage data. To do this, we first obtained the total population size M_j of group j
207 in San Francisco as well as the reported number of individuals in that group m_j who had
208 been vaccinated up to January 20, 2021 (i.e., 3 weeks before the weighted mean date of
209 sample collection, allowing for time to sero-conversion after vaccination). We then used a
210 hypergeometric distribution to sample N_j individuals without replacement from a population in
211 which m_j individuals had been vaccinated and $M_j - m_j$ had not. We then calculated the
212 empirical proportion of vaccinated individuals among the N_j in that simulation, and repeated

213 this procedure 10,000 times to obtain a prior distribution of $\text{Pr}(\text{vacc})$. This distribution was
214 used to estimate the hyper-priors of the Beta distribution. This procedure was performed
215 separately for each group j .

216
217 For geographic analyses, ZIP codes with fewer than 10 individuals were excluded. As data
218 on vaccine doses distributed by ZIP code during this time-frame was not available, we
219 modified the model above by assuming that vaccination was independent of infection, and
220 estimated a single $\text{Pr}(\text{inf})$ and $\text{Pr}(\text{vacc})$ for each ZIP code. To measure disparities in both
221 infection rates and vaccination, we leveraged a SARS-CoV-2 serosurveillance platform
222 launched in March 2020 that utilizes residual blood samples taken from two hospital
223 networks in San Francisco. Estimates derived from this platform during the first wave of the
224 pandemic showed seroprevalence in Latinx individuals to be nearly two times higher than in
225 white individuals, and nearly two times higher in homeless individuals than the population
226 average⁶. We collected samples from 1,014 individuals undergoing routine blood draws
227 between February 4 and February 17, 2021, capturing transmission during the first 11
228 months of the epidemic and the early roll-out of vaccination. These samples were tested
229 using two serologic assays: one detecting antibodies to SARS-CoV-2 elicited by infection
230 and not by vaccines currently used in the US, and one detecting antibodies to SARS-CoV-2
231 elicited by both infection and vaccination. We used Bayesian statistical models to estimate
232 the proportion of the population that was seropositive due to natural infection and the
233 proportion seropositive due to vaccination, stratified by age, race and ZIP code of residence.

234 Role of funding source

235 The authors confirm that the funding sources for this research had no role in the study
236 design, collection, analysis, interpretation of data, writing of the article or decision to submit
237 for publication.
238

239 **Results**

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241
242 Between February 4, 2021, and February 17, 2021, we collected samples, from 1014
243 individual patients, from UCSF Health ($n = 698$ patients) and the San Francisco
244 Department of Public Health ($n = 316$ patients) networks. By design, the geographic
245 distribution of residents matched the proportion of the San Francisco population living in
246 each zip code (**Fig. 2**). Our sample was equally distributed by sex, however over-
247 represented the 65+ age range and underrepresented the 0-34 age range relative to the San
248 Francisco population (**Table 1**). Our results were relatively representative of the San
249 Francisco population by race and ethnicity, although our sample overrepresented those who
250 identified as Black/African American and slightly underrepresented those who identified as
251 Asian.

252
253 Following testing samples on both Vitros and Roche assays, of the sampled population
254 where assay results were complete ($N = 915$), we found that while 28.4% ($N = 260$) tested
255 positive on the Vitros assay and therefore antibodies to SARS-CoV-2 were detected, only
256 8.6% ($N = 79$) tested positive on the Roche assay, detecting antibodies elicited by prior
257 natural infection (Figure 1). Of the 999 samples where assay results were returned, $N = 81$
258 samples were excluded from the analyses due to missing data in at least one assay, and 3
259 additional samples were excluded due to positive results on the Roche assay despite
260 negative results on the Vitros assay.

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[INSERT FIGURE 2 HERE]

267 **Table 1: Sample Characteristics.** Table showing the sample size and distribution of the
 268 sample population by demography and hospital system, compared to the San Francisco
 269 Population as determined by the American Community Survey 2019.

| | N | % | SF Population (ACS 2019) |
|---------------------------|----------|----------|---------------------------------|
| Age | | | |
| 0-17 | 21 | 2.10% | 13.40% |
| 18-34 | 157 | 15.50% | 30.60% |
| 35-64 | 442 | 43.60% | 40.60% |
| 65+ | 393 | 38.80% | 15.40% |
| Unknown | 1 | 0.10% | N/A |
| Sex | | | |
| Female | 509 | 50.20% | 49.30% |
| Male | 504 | 49.70% | 50.70% |
| Unknown | 1 | 0.10% | N/A |
| Hospital | | | |
| UCSF | 698 | 68.80% | N/A |
| ZSFG | 316 | 31.20% | N/A |
| Race/Ethnicity | | | |
| Asian | 282 | 27.80% | 34.60% |
| Black or African American | 115 | 11.30% | 5.20% |
| Hispanic or Latino | 175 | 17.30% | 15.20% |
| Other | 75 | 7.40% | 5.20% |
| White | 367 | 36.20% | 39.80% |

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279 Our estimated probabilities of vaccination and prior infection stratified by age, race/ethnicity
280 and ZIP code showed striking differences in prior infection rates and vaccination rates
281 across the city (**Figure 3, Supplementary Table 2**). Across all age and demographic
282 groups, ZIP codes in the southeastern region of the city, comprising medically underserved
283 neighborhoods, had demonstrably higher rates of prior infection and lower rates of
284 vaccination. This pattern is not evident in estimated seroprevalence by the vitros assay
285 which captures antibody responses acquired through natural infection and/or vaccination.
286 For example, within the 94124 zipcode, Bayview-Hunter's Point, one of the city's most
287 deprived zipcodes, the mean probability of prior infection was 0.155 (95% CI: 0.077 - 0.254)
288 and vaccination was 0.079 (0.019-0.163), whereas 94115, Pacific Heights, one of the
289 wealthier zipcodes in San Francisco with a median household income almost double that of
290 Bayview-Hunter's Point at 123,037¹⁰, the probability of infection was just 0.023 (0.001,
291 0.080) and vaccination was 0.359 (0.258 - 0.467).

292
293 We found the highest seroprevalence as a result of prior infection in younger age groups
294 (using the Roche assay, **Figure 4a, Supplementary Table 3**). We estimated
295 seroprevalence derived from both vaccination and natural infection using the Vitros assay
296 showed much higher seroprevalence in those aged over 65 (**Figure 4b, Supplementary**
297 **Table 3**), consistent with the eligibility criteria for vaccination in the weeks before the
298 sampling period.

299
300 We identified differences in prior infection rates by race/ethnicity (**Figure 4c-d**): we
301 estimated that the risk of prior infection of Latinx residents was 5.3 (95% CI: 3.2 - 10.3) times
302 greater than the risk of white residents aged 18-64 (**Figure 5a, Supplementary Table 3**).
303 These trends were echoed in older individuals (aged 65+) (**Figure 5a, Supplementary**
304 **Table 3**). We also identified disparities in vaccination coverage among the 65+ year old
305 population, who were eligible to receive the vaccine during this time period. We estimated
306 that White San Francisco residents over the age of 65 were twice as likely (2.0, 95% CI: 1.1
307 - 4.6) to be vaccinated as Black residents.

308
309 Taken together, these findings imply that there is an imbalance between the risk of infection
310 and the rate of vaccination in certain populations. Among the 65+ year old population, we
311 found greatly increased ratios of vaccination compared to infection risk among Asian and
312 white individuals, while this was much lower among Black and Latinx individuals (**Figure 5b,**
313 **Supplementary Table 4**). For every naturally infected Asian resident of this age group,
314 there were 12.2 vaccinated Asian residents (95%CI: 4.2 - 55.5), whereas for every naturally
315 infected Latinx resident of this age group, there were only 1.6 vaccinated Latinx residents
316 (95% CI: 0.7 – 3.7). For both Latinx and Black individuals over 65 years old, the risk of
317 having immunity acquired through vaccination, relative to natural infection, was up to four
318 times lower than for white individuals.

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[INSERT FIGURES 3-5 HERE]

327 Discussion

328 Using serological data, we quantified disparities in both vaccination coverage and infection
329 rates across different demographic groups and geographies and show that, during early
330 vaccine roll-out, vaccination coverage was much higher in Asian and White populations,
331 despite experiencing lower risk of infection by SARS-CoV-2 than Black and Hispanic/Latinx
332 populations.

333 The “double burden” we observed in San Francisco during the early vaccine roll-out echoes
334 broader patterns that have been observed in San Francisco and elsewhere. Even though
335 San Francisco was hailed as the first major US city to reach the milestone of 80%
336 vaccination coverage in adults¹¹, recent increases in infection have been found to be
337 concentrated in the neighborhoods which were hardest hit by initial infections and where we
338 found vaccination-related immunity was lowest¹². A report from the University of Texas found
339 striking geographic and racial stratification of cases of COVID-19 and vaccination rates in
340 Austin, Texas, which also closely mapped with indices of deprivation and social vulnerability
341 over ZIP codes¹³. Like in San Francisco, the neighborhoods which were predominantly
342 Latinx communities and had higher indices of deprivation also had higher incidence of
343 SARS-CoV-2 infection and lower vaccination coverage. Disparities in SARS-CoV-2
344 vaccination coverage among socially vulnerable populations have been documented across
345 the United States¹⁴ and in other parts of the world¹⁵.

346 There are several caveats and limitations to the approach introduced here. The Roche assay
347 can only differentiate antibody responses resulting from natural infections in settings where
348 Spike- based vaccines are used (that do not generate antibody responses against the
349 nucleocapsid), so in geographies where other vaccines are used this approach won't be
350 suitable. Although we used samples obtained through DPH health network, meaning we
351 included un-insured and under-insured individuals, we still are only able to capture those
352 seeking healthcare or reached by the SF DPH.

353
354 While inequalities revealed during COVID-19 are not new, the pandemic has highlighted the
355 ways in which even a city such as San Francisco which invests deeply in public health and
356 social safety nets still has deep structural inequalities, through a combination of higher
357 infection rates, incompatibility of living or work conditions with risk reduction, and lower or
358 delayed access to vaccines as they were rolled out.

359
360 Various initiatives are underway around the country to pinpoint geographic and other
361 disparities in the context of COVID-19 (e.g.,^{16,17}). However, as well as highlighting
362 disparities, it is important to consider what successful testing and vaccination initiatives may
363 look like and which may be learnt from in future measures. For example, within San
364 Francisco, robust community-academic partnerships have been key for effectively
365 responding to the pandemic in vulnerable communities¹⁸ and for narrowing gaps in
366 vaccination coverage, such as low-barrier neighborhood vaccination sites¹⁹. Prospectively,
367 as we gain a better understanding of waning immunity and the potential need for vaccine
368 booster doses, considerations of equity will remain an important consideration for allocating
369 resources. In the context of the United States, where vaccination and infection elicit different
370 immune responses, serology provides a powerful lens through which we can quantify these
371 disparities directly. In addition, it is important to consider the desired metric before
372 conducting a serosurvey, as assays measure different pathways to immunity and any
373 disparities in infection rates may be masked by using assays that measure overall antibody
374 prevalence.

375
376
377 Since the early days of the pandemic, many policy recommendations have been made for
378 ways to reduce health disparities in infection²⁰ and vaccination²¹. Policymakers must invest

379 in addressing both the upstream, structural drivers of health disparities, such as providing
380 workers with a living wage, affordable housing, and access to quality healthcare and also
381 downstream drivers such as improved community engagement, targeted testing and
382 vaccination provision, and assistance with common barriers to accessing healthcare such as
383 technology access/literacy, transport, and providing accessible health information in multiple
384 languages. While the arrival of the SARS-CoV-2 vaccine has created a 'light at the end of
385 the tunnel' for this pandemic, ongoing challenges that long predate COVID-19 in achieving
386 and maintaining equity must also be considered.

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467 **Figure Legends**

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469 **Figure 1: Schematic of parameters to be estimated using serosurveillance platform**
470 **(shown in red, blue and gold).** Red represents the probability of vaccination given prior
471 infection, $\Pr(\text{vacc}|\text{inf})$, blue represents the probability of prior infection, $\Pr(\text{inf})$, and gold
472 represents the probability of vaccination given no prior history of infection, $\Pr(\text{vacc}|\text{uninf})$.

473

474 **Figure 2: Sample characteristics. (A)** Age distribution by hospital week of sample
475 collection within the University of California, San Francisco (UCSF) and San Francisco
476 Department of Public Health (ZSFG) hospital networks. Each point represents a sample and
477 colors correspond to age bins used for analysis. **(B)** Proportion of samples from a given San
478 Francisco zipcode plotted against the proportion of the San Francisco population within that
479 zipcode. Colors show the percentage of residents below the poverty line within that zipcode,
480 as determined by the American Community Survey 2019.

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482

483 **Figure 3: Maps showing geographic disparities in SARS-CoV-2 within San Francisco.**

484 Maps show (a) estimated probability of prior infection and (b) probability of vaccination by
485 ZIP code in San Francisco, as of February 2021.

486

487 **Figure 4: Stratified seroprevalence by assay and by demographic group. (A)** Univariate
488 Roche seropositivity estimates by age (elicited by prior infection). **(B)** Univariate Vitros
489 estimates by age (elicited by either prior infection or vaccination). **(C)** Univariate Roche
490 estimates by race/ethnicity. **(D)** Univariate Vitros estimates by race/ethnicity.

491

492 **Figure 5: Relationship between probability of vaccination and probability of prior**
493 **infection by race/ethnicity. a)** Probability of infection vs. probability of vaccination by age
494 and race/ethnicity. **b)** Infographic showing the number of estimated people vaccinated for
495 every one person previously naturally infected in San Francisco within each
496 racial/demographic group.

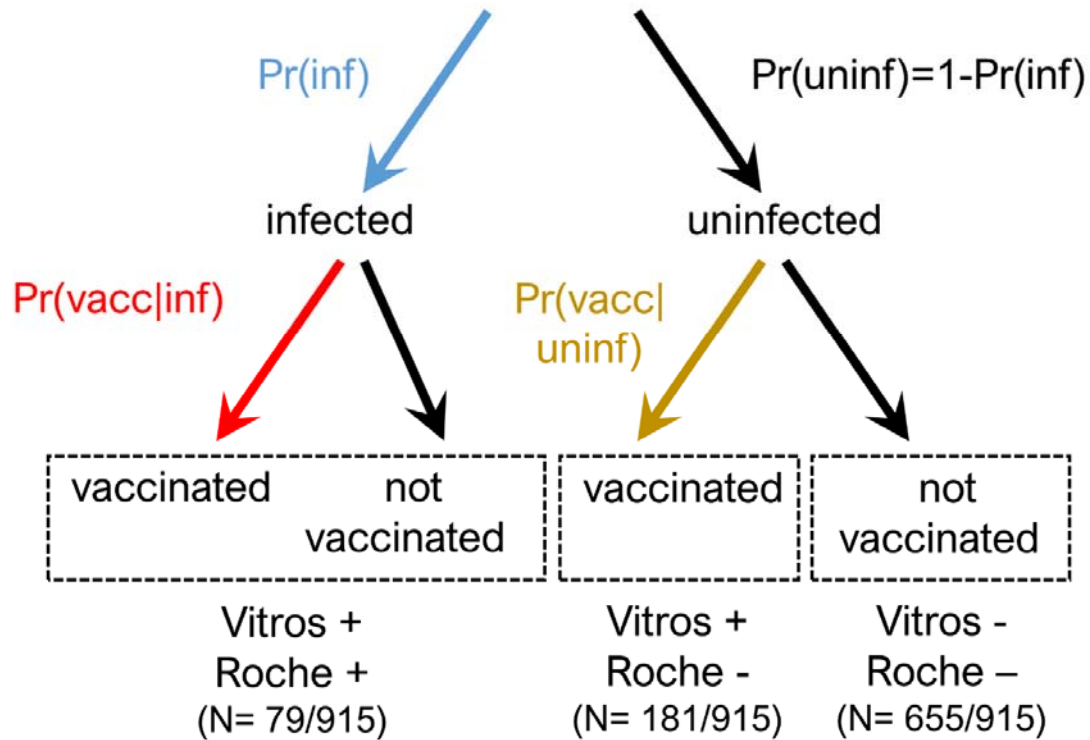
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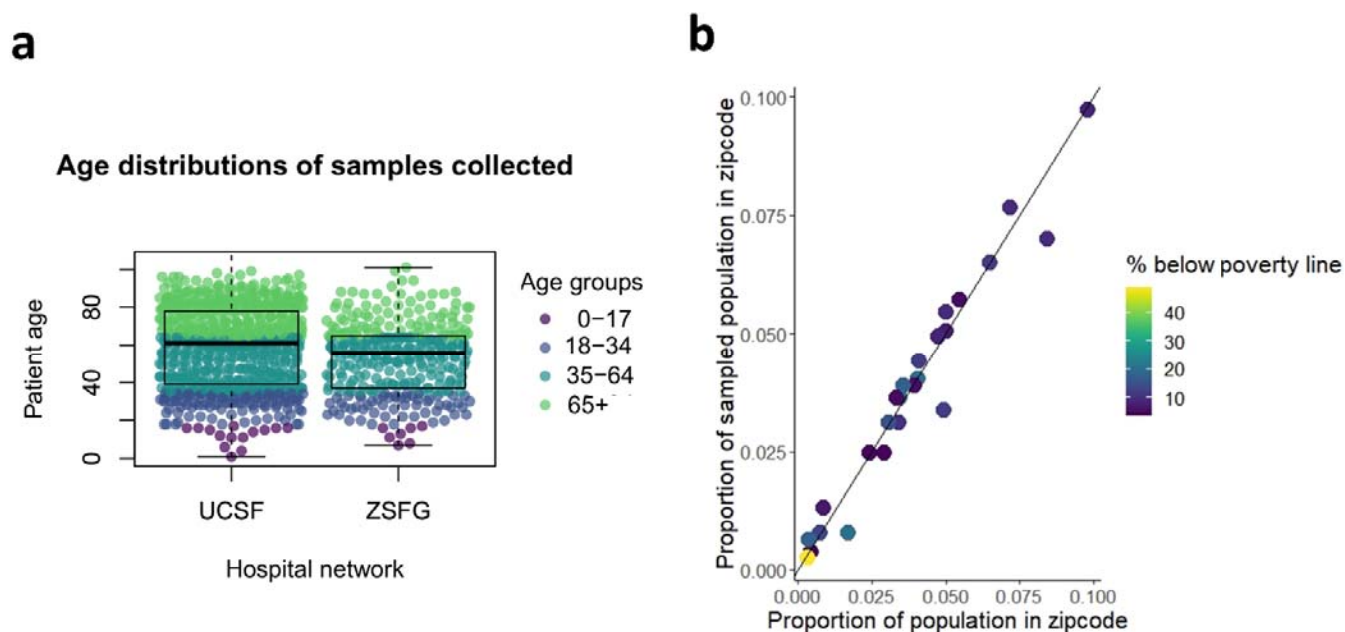
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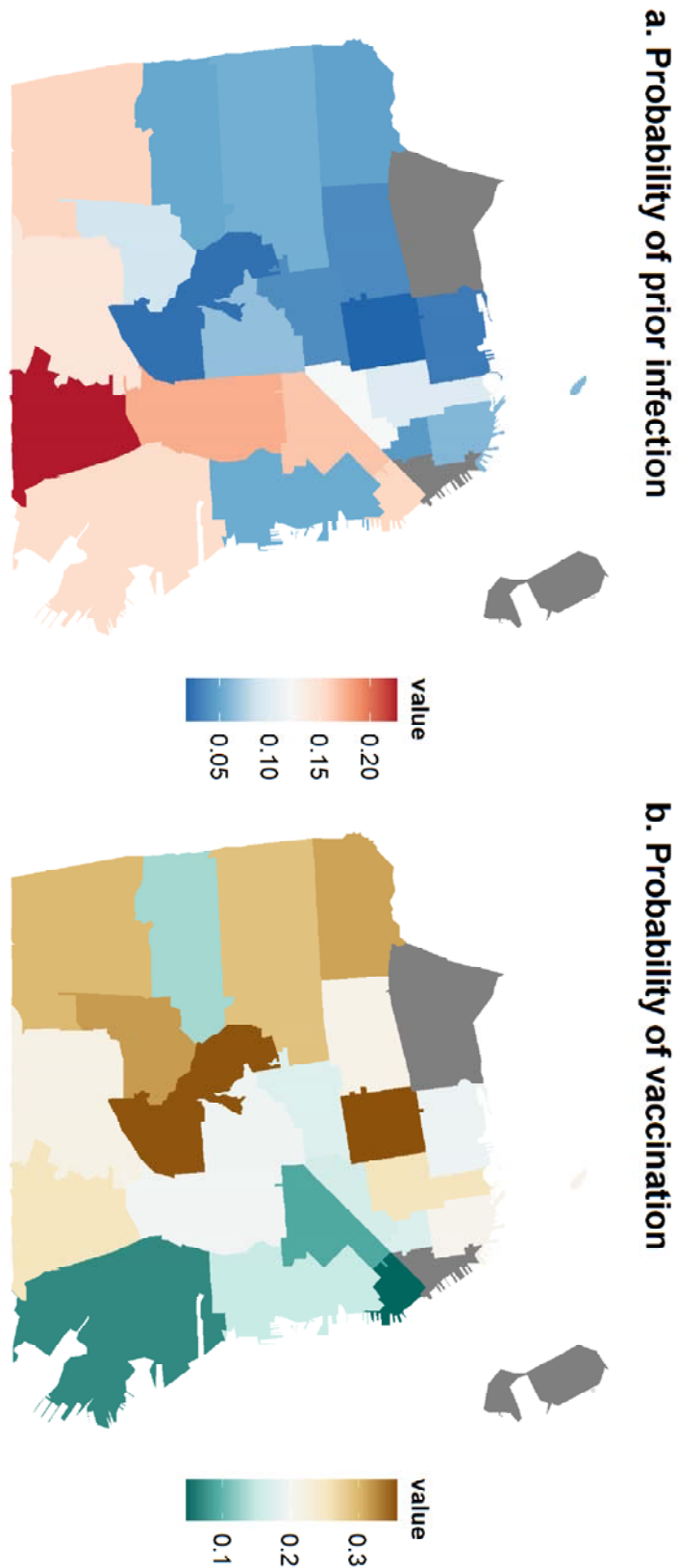


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516 Francisco zipcode plotted against the proportion of the San Francisco population within that
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518 as determined by the American Community Survey 2019.
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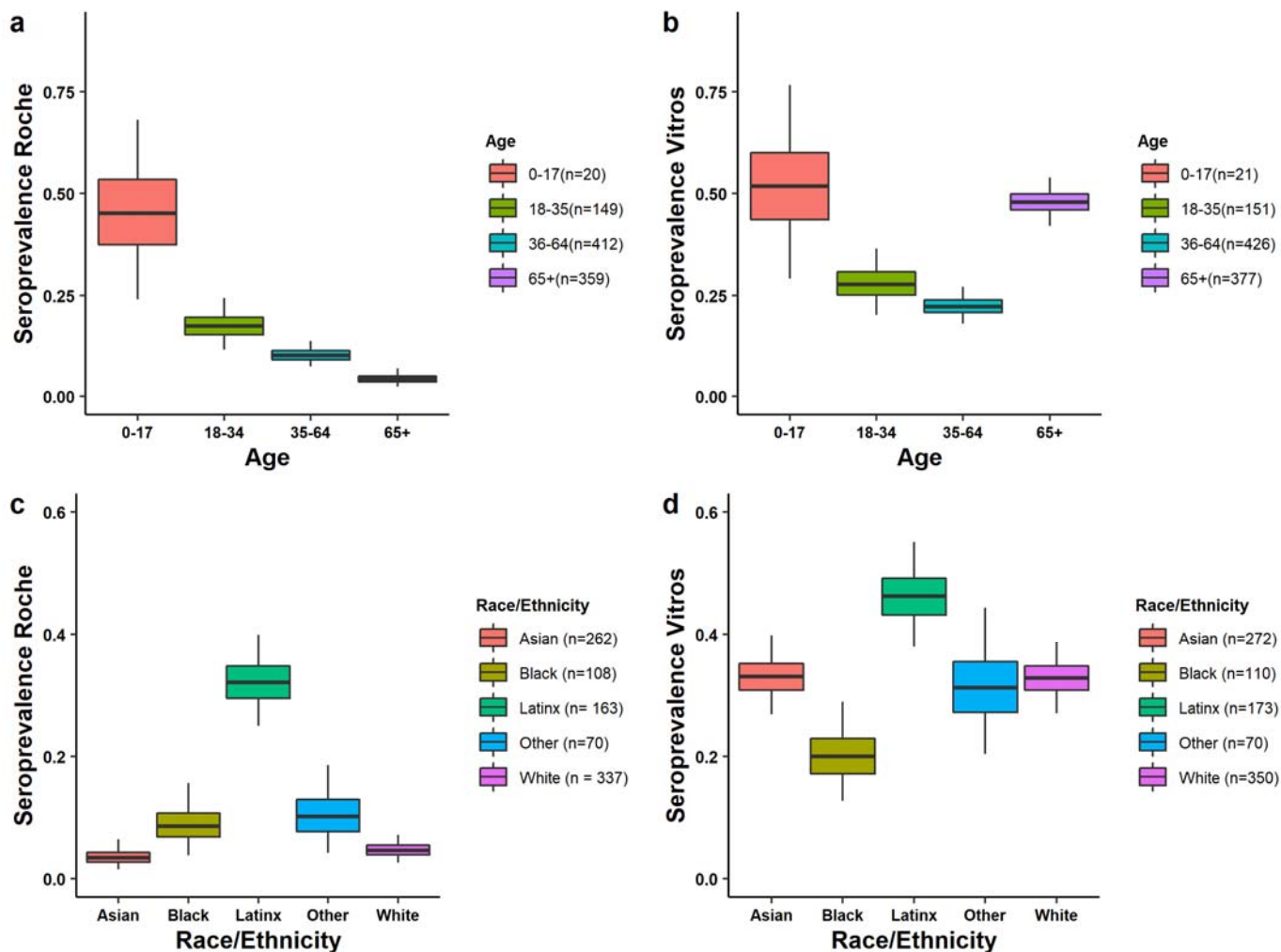


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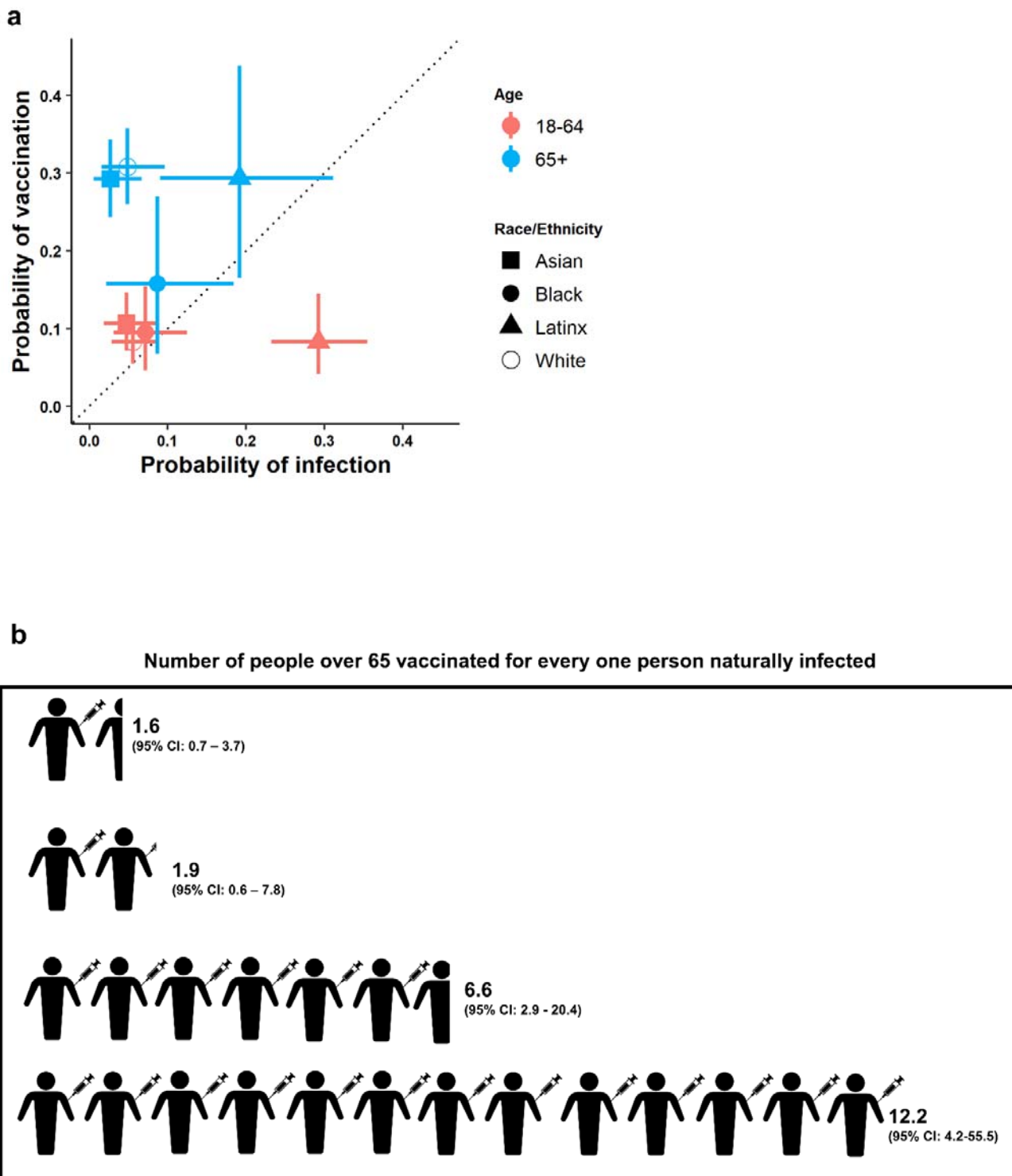


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527 **Figure 4: Stratified seroprevalence by assay and by demographic group. (A)** Univariate
528 Roche seropositivity estimates by age (elicited by prior infection). **(B)** Univariate Vitros
529 estimates by age (elicited by either prior infection or vaccination). **(C)** Univariate Roche
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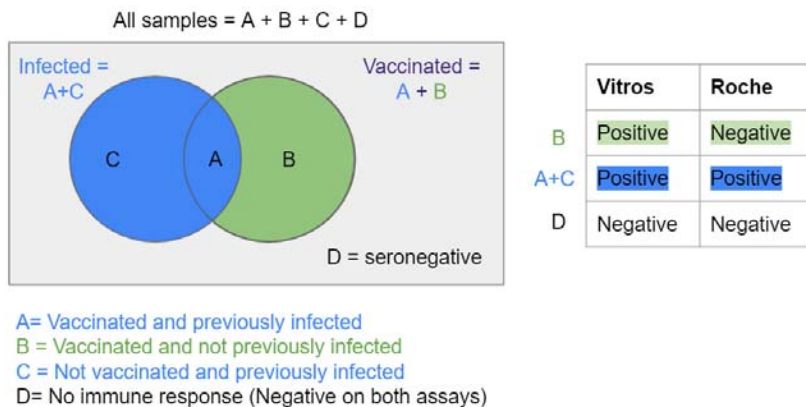


532 **Figure 5: Relationship between probability of vaccination and probability of prior**
533 **infection by race/ethnicity. a) Probability of infection vs. probability of vaccination by age**
534 **and race/ethnicity. b) Infographic showing the number of estimated people vaccinated for**
535 **every one person previously naturally infected in San Francisco within each**
536 **racial/demographic group.**
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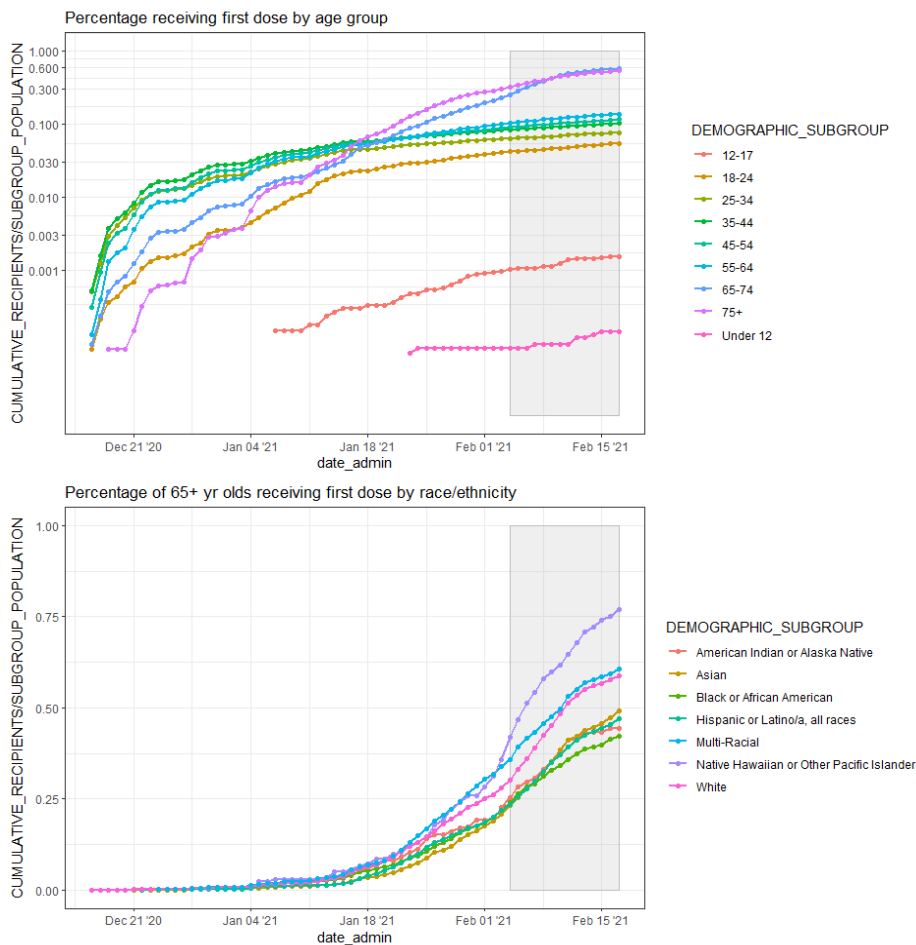
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Supplementary Figure 2: Pathways to immunity via infection and vaccination.



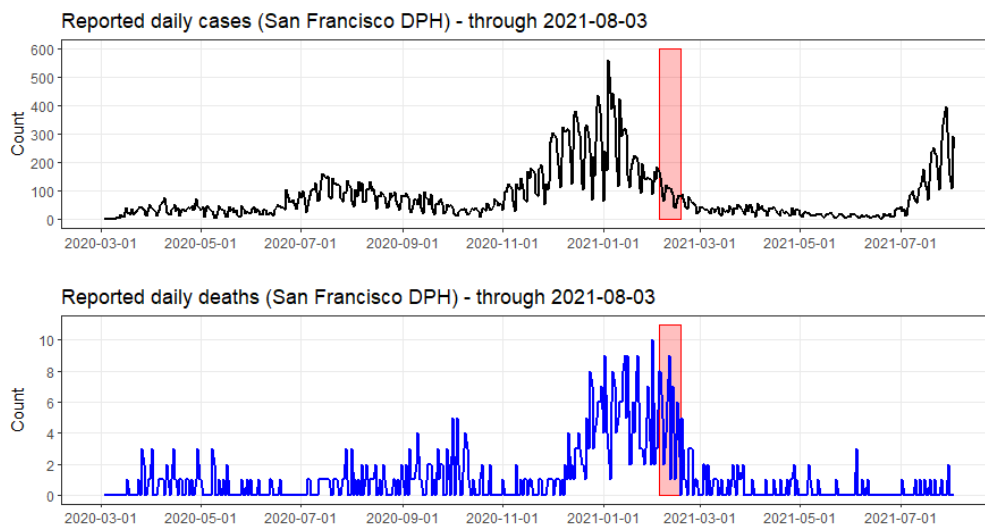
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Supplementary Figure 3: Cumulative proportion of individuals who received the first dose of SARS-CoV-2 vaccine by date, age group, and race/ethnicity. Sampling period of this study shown in grey box. Data downloaded from ²².



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563 **Supplementary Figure 4: Time series of reported SARS-CoV-2 cases and deaths in**
564 **San Francisco.** Sampling period of this study shown in red box. Data downloaded from ²³.
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601 **Supplementary Table 1 : Interpretation of bivariate antibody results on the Vitros and**
602 **Roche assays.**
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| Vitros result | Roche result | Interpretation for bivariate analyses (n) |
|----------------------|---------------------|--|
| Positive | Positive | Infected and (either vaccinated or unvaccinated) (n=79) |
| Positive | Negative | Uninfected and vaccinated (n=181) |
| Negative | Negative | Uninfected and unvaccinated (n=655) |
| Negative | Positive | False negative on Vitros and/or false positive on Roche (n=3; removed) |
| Positive | No result | (n=15; removed) |
| Negative | No result | (n=43; removed) |
| No result | Positive | (n=2; removed) |
| No result | Negative | (n=21; removed) |

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Supplementary Table 2: Bivariate posteriors by ZIP code.

| zcta | Mean, P(vaccination) | Mean, P(infection) | 2.5%, P(vaccination) | 2.5%, P(infection) | 50%, P(vaccination) | 50%, P(infection) | 97.5%, P(vaccination) | 97.5%, P(infection) | n_samples |
|-------|----------------------|--------------------|----------------------|--------------------|---------------------|-------------------|-----------------------|---------------------|-----------|
| 94102 | 0.166 | 0.124 | 0.073 | 0.048 | 0.164 | 0.121 | 0.273 | 0.224 | 36 |
| 94103 | 0.097 | 0.169 | 0.025 | 0.079 | 0.091 | 0.166 | 0.203 | 0.278 | 31 |
| 94105 | 0.060 | 0.166 | 0.002 | 0.044 | 0.046 | 0.158 | 0.193 | 0.333 | 13 |
| 94107 | 0.156 | 0.059 | 0.066 | 0.008 | 0.152 | 0.053 | 0.266 | 0.147 | 28 |
| 94108 | 0.180 | 0.057 | 0.056 | 0.001 | 0.171 | 0.044 | 0.349 | 0.186 | 12 |
| 94109 | 0.257 | 0.101 | 0.178 | 0.046 | 0.256 | 0.099 | 0.341 | 0.170 | 65 |
| 94110 | 0.198 | 0.178 | 0.123 | 0.108 | 0.198 | 0.176 | 0.279 | 0.256 | 75 |
| 94112 | 0.219 | 0.145 | 0.145 | 0.087 | 0.218 | 0.143 | 0.298 | 0.212 | 86 |
| 94114 | 0.198 | 0.070 | 0.105 | 0.017 | 0.195 | 0.065 | 0.307 | 0.153 | 38 |
| 94115 | 0.359 | 0.023 | 0.258 | 0.001 | 0.357 | 0.016 | 0.467 | 0.080 | 39 |
| 94116 | 0.135 | 0.056 | 0.066 | 0.014 | 0.132 | 0.051 | 0.219 | 0.120 | 48 |
| 94117 | 0.177 | 0.042 | 0.097 | 0.006 | 0.175 | 0.037 | 0.273 | 0.108 | 41 |
| 94118 | 0.217 | 0.041 | 0.131 | 0.005 | 0.215 | 0.036 | 0.313 | 0.107 | 43 |
| 94121 | 0.315 | 0.052 | 0.229 | 0.012 | 0.313 | 0.048 | 0.407 | 0.115 | 53 |
| 94122 | 0.292 | 0.058 | 0.210 | 0.017 | 0.291 | 0.055 | 0.382 | 0.118 | 65 |
| 94123 | 0.197 | 0.038 | 0.086 | 0.001 | 0.192 | 0.028 | 0.331 | 0.126 | 21 |
| 94124 | 0.075 | 0.155 | 0.019 | 0.077 | 0.070 | 0.153 | 0.163 | 0.254 | 40 |
| 94127 | 0.322 | 0.097 | 0.166 | 0.015 | 0.317 | 0.088 | 0.494 | 0.235 | 17 |
| 94131 | 0.357 | 0.031 | 0.232 | 0.001 | 0.355 | 0.022 | 0.491 | 0.111 | 28 |
| 94132 | 0.299 | 0.162 | 0.185 | 0.073 | 0.298 | 0.159 | 0.421 | 0.266 | 40 |
| 94133 | 0.215 | 0.067 | 0.106 | 0.011 | 0.211 | 0.060 | 0.348 | 0.161 | 25 |
| 94134 | 0.257 | 0.231 | 0.136 | 0.130 | 0.255 | 0.227 | 0.392 | 0.347 | 36 |

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613 **Supplementary Table 3: Univariate posteriors by demography.**

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| Assay | Demographic Group | Median | CI.2.5% | CI.97.5% |
|--------------|---------------------------|---------------|----------------|-----------------|
| Roche | 0-17 | 0.451 | 0.240 | 0.681 |
| Vitros | 0-17 | 0.518 | 0.291 | 0.767 |
| Roche | 18-34 | 0.173 | 0.116 | 0.243 |
| Vitros | 18-34 | 0.277 | 0.202 | 0.364 |
| Roche | 35-64 | 0.101 | 0.074 | 0.136 |
| Vitros | 35-64 | 0.223 | 0.180 | 0.271 |
| Roche | 65+ | 0.043 | 0.024 | 0.070 |
| Vitros | 65+ | 0.479 | 0.420 | 0.539 |
| Roche | Asian | 0.035 | 0.015 | 0.065 |
| Vitros | Asian | 0.331 | 0.269 | 0.399 |
| Roche | Black or African American | 0.086 | 0.038 | 0.157 |
| Vitros | Black or African American | 0.201 | 0.127 | 0.289 |
| Roche | Hispanic or Latino | 0.321 | 0.250 | 0.399 |
| Vitros | Hispanic or Latino | 0.463 | 0.380 | 0.551 |
| Roche | Other | 0.102 | 0.042 | 0.186 |
| Vitros | Other | 0.313 | 0.204 | 0.444 |
| Roche | White | 0.046 | 0.026 | 0.072 |
| Vitros | White | 0.328 | 0.270 | 0.387 |

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618 **Supplementary Table 4: Bivariate antibody results by age and race/ethnicity.**
 619 Probability of vaccination given infected, probability of vaccination given uninfected,
 620 probability of infection, probability of vaccination, probability of vaccination divided by
 621 probability of infection (“risk ratio”), and relative risk ratio with a reference group of White
 622 race/ethnicity and the corresponding age.
 623

| Age | Race/Ethnicity | parameter | median | lb_2.5 | ub_97.5 |
|-------|----------------|--------------------------|--------|--------|---------|
| 18-64 | Asian | p_vacc_given_inf | 0.216 | 0.008 | 0.875 |
| 18-64 | Black | p_vacc_given_inf | 0.303 | 0.012 | 0.932 |
| 18-64 | Latinx | p_vacc_given_inf | 0.053 | 0.002 | 0.261 |
| 18-64 | White | p_vacc_given_inf | 0.125 | 0.005 | 0.620 |
| 65+ | Asian | p_vacc_given_inf | 0.233 | 0.007 | 0.891 |
| 65+ | Black | p_vacc_given_inf | 0.387 | 0.015 | 0.958 |
| 65+ | Latinx | p_vacc_given_inf | 0.270 | 0.010 | 0.914 |
| 65+ | White | p_vacc_given_inf | 0.105 | 0.003 | 0.688 |
| 18-64 | Asian | p_vacc_given_uninf | 0.099 | 0.065 | 0.137 |
| 18-64 | Black | p_vacc_given_uninf | 0.074 | 0.036 | 0.124 |
| 18-64 | Latinx | p_vacc_given_uninf | 0.086 | 0.045 | 0.141 |
| 18-64 | White | p_vacc_given_uninf | 0.078 | 0.052 | 0.108 |
| 65+ | Asian | p_vacc_given_uninf | 0.294 | 0.245 | 0.344 |
| 65+ | Black | p_vacc_given_uninf | 0.130 | 0.054 | 0.241 |
| 65+ | Latinx | p_vacc_given_uninf | 0.285 | 0.168 | 0.420 |
| 65+ | White | p_vacc_given_uninf | 0.316 | 0.269 | 0.364 |
| 18-64 | Asian | p_inf | 0.046 | 0.019 | 0.086 |
| 18-64 | Black | p_inf | 0.069 | 0.031 | 0.124 |
| 18-64 | Latinx | p_inf | 0.291 | 0.232 | 0.355 |
| 18-64 | White | p_inf | 0.054 | 0.028 | 0.088 |
| 65+ | Asian | p_inf | 0.024 | 0.005 | 0.067 |
| 65+ | Black | p_inf | 0.081 | 0.022 | 0.184 |
| 65+ | Latinx | p_inf | 0.188 | 0.090 | 0.311 |
| 65+ | White | p_inf | 0.046 | 0.016 | 0.096 |
| 18-64 | Asian | p_vacc | 0.106 | 0.072 | 0.147 |
| 18-64 | Black | p_vacc | 0.093 | 0.046 | 0.154 |
| 18-64 | Latinx | p_vacc | 0.080 | 0.042 | 0.145 |
| 18-64 | White | p_vacc | 0.083 | 0.056 | 0.116 |
| 65+ | Asian | p_vacc | 0.292 | 0.244 | 0.343 |
| 65+ | Black | p_vacc | 0.155 | 0.068 | 0.270 |
| 65+ | Latinx | p_vacc | 0.290 | 0.165 | 0.438 |
| 65+ | White | p_vacc | 0.308 | 0.260 | 0.358 |
| 18-64 | Asian | ratio_p_vacc_p_inf | 2.314 | 1.016 | 6.304 |
| 18-64 | Black | ratio_p_vacc_p_inf | 1.368 | 0.510 | 3.388 |
| 18-64 | Latinx | ratio_p_vacc_p_inf | 0.275 | 0.134 | 0.526 |
| 18-64 | White | ratio_p_vacc_p_inf | 1.515 | 0.754 | 3.350 |
| 65+ | Asian | ratio_p_vacc_p_inf | 12.220 | 4.188 | 55.534 |
| 65+ | Black | ratio_p_vacc_p_inf | 1.912 | 0.559 | 7.817 |
| 65+ | Latinx | ratio_p_vacc_p_inf | 1.572 | 0.655 | 3.705 |
| 65+ | White | ratio_p_vacc_p_inf | 6.676 | 2.930 | 20.444 |
| 18-64 | Asian | RRR_vs_white_of_that_age | 1.571 | 0.493 | 4.832 |
| 18-64 | Black | RRR_vs_white_of_that_age | 0.875 | 0.265 | 2.883 |

| | | | | | |
|-------|--------|--------------------------|-------|-------|-------|
| | | e | | | |
| 18-64 | Latinx | RRR_vs_white_of_that_age | 0.184 | 0.064 | 0.495 |
| 18-64 | White | RRR_vs_white_of_that_age | 1.000 | 1.000 | 1.000 |
| 65+ | Asian | RRR_vs_white_of_that_age | 1.831 | 0.405 | 9.568 |
| 65+ | Black | RRR_vs_white_of_that_age | 0.291 | 0.060 | 1.489 |
| 65+ | Latinx | RRR_vs_white_of_that_age | 0.231 | 0.058 | 0.788 |
| 65+ | White | RRR_vs_white_of_that_age | 1.000 | 1.000 | 1.000 |

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626 **Data sharing statement**

627

628 All analysis was conducted using the R statistical software and the Stan programming
629 language. All code to reproduce these results are available at:

630 <https://github.com/EPPICenter/scale-it-2>.

631

632

633 **Declaration of Interests**

634 The authors have no conflicts of interest to declare.

635

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650 **Research in context**

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652 **Evidence before this study**

653 We searched PubMed on October 13, 2021, using the following search terms: (inequalit* OR
654 disparit*) AND (vaccination OR vaccine) AND (infect* OR exposure) AND (COVID* OR
655 SARS-CoV-2 OR SARS-CoV), published in 2021 (i.e. after the roll-out of vaccination) with
656 no language restrictions. We identified 145 publications, including original research,
657 commentaries, systematic reviews and meta-analyses and reviews. Although many studies
658 discussed or quantified disparities in either vaccination or infection risk, we did not find any
659 studies which measured the combined effect or interaction between disparities in vaccination
660 and infection risk, or quantified the effect of these disparities on population-level immunity.

661

662 **Added value of this study**

663 To our knowledge, this is the first study to jointly measure disparities in vaccination and
664 infection using serological data. Serosurveillance systems leveraging demographically-
665 stratified residual serum samples, like the one we implemented, are affordable and flexible,
666 and could be used by Public Health systems to monitor health disparities as they avoid some
667 of the biases inherent in metrics such as case counts and test positivity rates. We also show
668 how vaccination coverage may be monitored, which can be challenging when integrating
669 data from multiple healthcare systems and when denominator populations may be
670 inaccurate for calculating vaccination coverage.

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