

Dietary acrylamide intake and the risk of endometrial or ovarian cancers in Japanese women

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Funding information

Food Safety Commission, Cabinet Office, Government of Japan, Grant/Award Number: 1503; National Cancer Center Research and Development Fund; Grant-in-Aid for Cancer Research from the Ministry of Health, Labour and Welfare of Japan

A meta-analysis published in 2015 noted a marginally increased risk of endometrial and ovarian cancers in non-smoking women with dietary acrylamide intake, but only a few studies were included, and they were limited to Western countries. The aim of this study was to investigate the association between dietary acrylamide intake and endometrial or ovarian cancer risk in the Japan Public Health Center-based Prospective Study (JPHC Study). In this prospective cohort study, 47 185 participants aged 45–74 years at the follow-up starting point in the JPHC Study were enrolled. Dietary acrylamide intake was assessed using a validated food frequency questionnaire. Cox proportional hazards regression models were used to estimate hazard ratios (HR) and 95% confidence intervals (95%CI). In participants with endometrial and ovarian cancer, the average follow-up periods were 15.5 and 15.6 years, respectively, and 161 and 122 cases of endometrial and ovarian cancer were diagnosed, respectively. Energy-adjusted dietary acrylamide intake was negatively associated with endometrial cancer, but the association disappeared after adjusting for coffee consumption with an adjusted HR for the highest vs lowest tertile of 0.85 (95%CI: 0.54–1.33). No association was observed, however, for ovarian cancer (adjusted HR, 0.77; 95%CI: 0.49–1.23). Furthermore, after stratifying by smoking status, coffee consumption, alcohol consumption, body mass index, and menopause status, no association was observed. Dietary acrylamide intake was not associated with the risk of endometrial or ovarian cancer in Japanese women with a relatively lower dietary intake of acrylamide compared with Western populations.

KEYWORDS

Asia, dietary acrylamide, endometrial cancer, epidemiology, ovarian cancer

Abbreviations: CI, confidence interval; DCO, death certificate only; EPIC, European Prospective Investigation into Cancer and Nutrition; FFQ, food frequency questionnaire; HR, hazard ratio; ICD-O-3, International Classification of Diseases for Oncology, Third Edition; JPHC Study, Japan Public Health Center-based Prospective Study; NHS, Nurses' Health Study; NLCS, Netherlands Cohort Study on Diet and Cancer; SMC, Swedish Mammography Cohort.

1 | INTRODUCTION

The International Agency for Research on Cancer classified acrylamide as a probable human carcinogen (group 2A) in 1994 given the evidence for the carcinogenicity of acrylamide in animal studies.¹ In the general population, smoking was thought to be the main

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source of acrylamide exposure,² but in 2002, Swedish researchers also found acrylamide in starchy foods cooked at high temperatures (such as fried potato).³ Therefore, in Western countries, epidemiological studies have been conducted to clarify the risk of dietary acrylamide intake with regard to the incidence of cancers.

A meta-analysis published in 2015 showed no increased risk for most cancers but, due to the marginally increased risk of endometrial and ovarian cancers identified in non-smoking women, there is a need for further studies.⁴ One possible mechanism by which dietary acrylamide exerts its carcinogenic effect is thought to be via the genotoxic pathway of glycidamide, which is an acrylamide metabolite.⁵⁻⁷ Furthermore, a hormone-related pathway has also been debated.⁷⁻⁹ Thus, acrylamide is more likely to cause a non-negligible increase in the risk of endometrial and ovarian cancers through synergistic genotoxicity effects and hormone changes than with other cancers. The small number of studies included in the meta-analysis (four studies for each cancer), however, is also considered one of the reasons for this finding. Therefore, accumulation of evidence from further studies is needed.

In the Japan Public Health Center-based Prospective Study (JPHC Study), the main sources of dietary acrylamide intake based on the dietary records were coffee and green tea, followed by confectionery, vegetables, and potatoes.¹⁰ In contrast, the main sources in Western countries were potato-based foods, wheat-based products, and coffee.¹¹⁻¹³ These differences might influence the effect of dietary acrylamide on the risk of cancer in Japan. This is because coffee is also known to be a preventive factor for endometrial cancer.¹⁴ Thus, in the case of endometrial cancer, it is expected that the carcinogenic effect of acrylamide may be attenuated by the protective effect of coffee. Therefore, in order to evaluate the safety of dietary acrylamide, it is important to examine its influence on cancers in various countries with different dietary sources. Only one study in Asia, however, has assessed the influence of dietary acrylamide intake on the incidence of cancers.¹⁵

Therefore, the aim of this study was to investigate the association between dietary acrylamide intake and the incidence of endometrial or ovarian cancers in Japanese women.

2 | MATERIALS AND METHODS

2.1 | Study participants

The JPHC Study was launched in the 1990s to investigate the associations between lifestyle-related diseases in two cohorts as a population-based prospective cohort study. Cohort I areas included Iwate, Akita, Nagano, Okinawa (Chubu), and Tokyo, while cohort II areas included Ibaraki, Niigata, Kochi, Nagasaki, Okinawa (Miyako), and Osaka. The study protocol has been described elsewhere.^{16,17} Participants in the JPHC Study aged 40-69 years in these 11 areas consisted of 140 420 inhabitants (68 722 men and 71 698 women). Dietary surveys were conducted using a self-administered food frequency questionnaire (FFQ) at baseline and at the 5-year follow-up survey. The number of food items and the number of food items

with the option of portion size in the FFQ at baseline were limited. In contrast, the FFQ at 5-year follow-up contained more detailed dietary information. Thus, we treated the 5-year follow-up survey as the starting point of the follow up, and calculated dietary acrylamide intake as the exposure variable using the 5-year follow-up FFQ. The study protocol was approved by the institutional review boards of the National Cancer Center, Tokyo, Japan; Osaka University; and Azabu University. All study participants provided informed consent prior to inclusion in the study.

Participants in the Tokyo area were not included in the present study because the incidence data for cancer cases were not available. After excluding the participants who were disqualified, had died, moved out of the study area, or were lost to follow up before the starting point, 62 750 women were eligible for inclusion in this study. Of these, respondents to the 5-year follow-up survey consisted of 52 483 women (response rate, 83.6%). Furthermore, the participants who had a past history of endometrial (n = 654) or

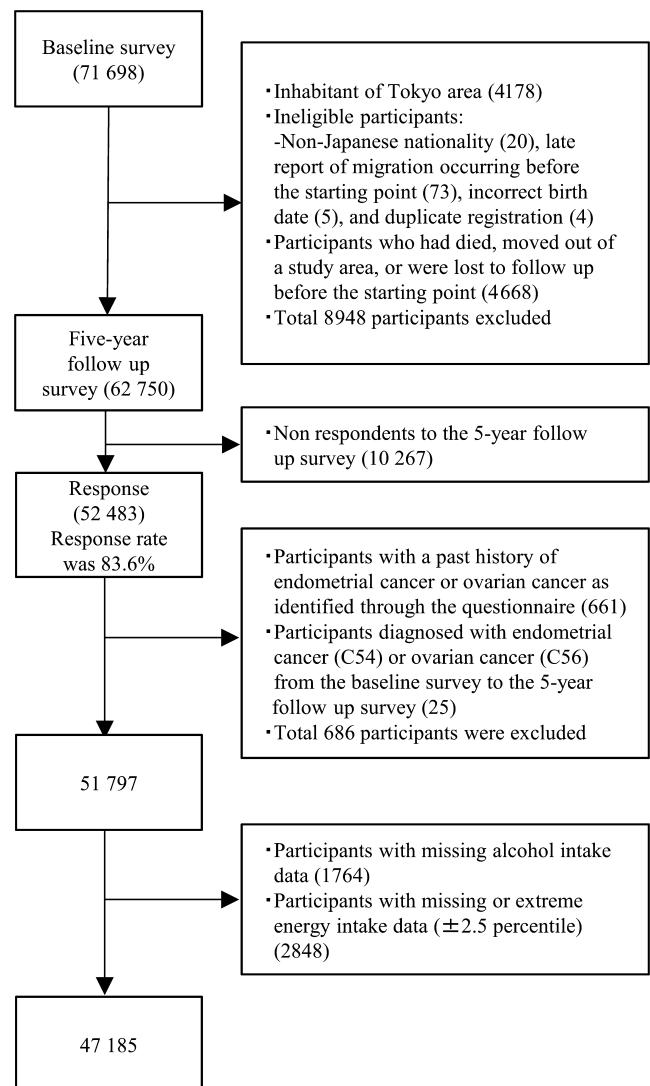


FIGURE 1 Flowchart of participant selection; Japan Public Health Center-based Prospective Study

ovarian ($n = 7$) cancers identified through the questionnaire and diagnosed during the baseline survey up to the 5-year follow-up survey were excluded (endometrial cancer, $n = 14$; ovarian cancer, $n = 11$). Participants with missing data or extreme energy intake data (upper and lower 2.5 percentile) were also excluded ($n = 4612$). Therefore, 47 185 women were included in the present analysis of endometrial and ovarian cancers (Figure 1).

2.2 | Assessment of energy and acrylamide intake from FFQ

Overall, 138 food and beverage items were included in the 5-year follow-up FFQ. The options for each food item were grouped into 9 categories according to the frequency of eating (never; 1-3 times/month; 1-2 times/week; 3-4 times/week; 5-6 times/week; once/day; 2-3 times/day; 4-6 times/day; or ≥ 7 times/day) and 3 categories according to portion size (less than half the standard portion size; standard portion size; or >1.5 -fold the standard portion size). The options for each drink were grouped into 9 categories according to the frequency of drinking (<1 cup/week; 1-2 cups/week; 3-4 cups/week; 5-6 cups/week; 1 cup/day; 1-3 cups/day; 4-6 cups/day; 7-9 cups/day; or ≥ 10 cups/day).

Energy content in each food item was referenced from the fifth revised and enlarged edition of the Standard Tables of Food Composition in Japan.¹⁸ The estimated energy intake was calculated as the sum of the product of the eating frequency, portion size, and energy content of each food. From validation studies, on comparison of 28-day dietary record (DR) in subsamples of two cohorts, the correlation coefficients of energy intake in women were 0.41 and 0.24 in cohort I ($n = 113$) and cohort II ($n = 176$), respectively.¹⁹⁻²¹

Acrylamide intake was also estimated from the amount of food and beverage intake and the acrylamide content database. The database was developed from published reports of measurements of common Japanese foods.²²⁻³¹ In addition to heated processed foods such as bread, biscuit and cookies, or coffee, we considered the influence of home cooking such as stir-fried vegetables, toast, or fried batter, to estimate the dietary acrylamide more accurately.²³ The de-attenuated correlation coefficients of energy-adjusted dietary acrylamide intake in women were 0.48 and 0.37 in cohort I and cohort II, respectively.¹⁰

2.3 | Follow up and identification of endometrial and ovarian cancers

The follow-up period for all participants was from the starting point of the 5-year follow-up survey until 31 December 2013 (until 31 December 2012, only in the Osaka area). Residential status was confirmed annually through the residential registry. During the follow-up period, 5741 participants (12.2%) died, 3191 (6.8%) moved from the study area, and 31 (0.1%) were lost to follow up.

The cancer incidence was identified through the following data sources: active patient notification from major local hospitals in the

study area and data linkage with population-based cancer registries. Death certificates were used as a supplementary information source. The International Classification of Diseases for Oncology, Third Edition (ICD-O-3) was used for coding endometrial (C54) and ovarian (C56) cancer cases. The proportion of cases determined using death certificate only (DCO) was 1.9% and 7.4% for endometrial and ovarian cancers, respectively. Given that these percentages were $<10\%$, they were considered satisfactory for the present study.³² A total of 161 endometrial and 122 ovarian cancers were newly diagnosed by 31 December 2013.

2.4 | Statistical analysis

Person-years of follow up were determined from the starting point until the date of diagnosis of endometrial or ovarian cancer, the date of a participant's death, the date of relocation from the study area, or the end of the study period (31 December 2012 for Osaka area and 31 December 2013 for all other areas), whichever occurred first. For participants lost to follow up, the last confirmed date of their presence in the study area was used as the date of censor. The mean follow-up period was 15.5 and 15.6 years for endometrial and for ovarian cancer analysis, respectively.

Cox proportional hazard modeling was used to estimate the hazard ratio (HR) and 95% confidence intervals (95%CI) to determine the association between tertiles of energy-adjusted dietary acrylamide intake and endometrial or ovarian cancers. For energy adjustment, the residual method was used. The trend of HR was also assessed using the ordinal values of the tertiles of energy-adjusted dietary acrylamide intake. HR were adjusted for the following potential confounders in model 1 for endometrial cancer analysis: age (years), area (10 public health center areas), body mass index (BMI; <25 , ≥ 25 kg/m², or missing), age at menarche (≤ 13 , 14, 15, ≥ 16 years, or missing), age at first delivery (<26 , ≥ 26 years, or missing), number of deliveries (0, 1-2, 3, ≥ 4 , or missing), menopause status and age at menopause (premenopause, postmenopause [age <49 , 50-54, ≥ 55 years], or missing), use of exogenous female hormones (yes, no, or missing), smoking status (current, ever, never, or missing), and alcohol intake (<150 or ≥ 150 g/week). Furthermore, in the multivariate-adjusted model 2, energy-adjusted coffee intake (continuous) was adjusted for in addition to the variables in model 1.

For ovarian cancer analysis, similar confounding factors as those for endometrial cancer were also adjusted for. These variables were identified from the FFQ and are known or suspected risk factors of endometrial or ovarian cancers.^{33,34} For sensitivity analysis in model 2, we repeated the same analysis for each cancer but excluded cancer cases diagnosed at ≤ 3 years of follow-up. Further, we conducted a stratified analysis using smoking status, frequency of coffee consumption (<1 cup/week, ≥ 1 cup/week), alcohol consumption, BMI, and menopause status at starting point (pre- or postmenopause). All *P*-values were two-sided, and statistical significance level was set at $P < .05$ using SAS 9.3 (SAS Institute, Cary, NC, USA).

TABLE 1 Subject characteristics for endometrial and ovarian cancer analysis

	Tertiles of acrylamide intake			P-value ^a
	Lowest T1 Mean ± SD or %	Middle T2 Mean ± SD or %	Highest T3 Mean ± SD or %	
No. participants	15 728	15 729	15 728	
Acrylamide intake				
Median (µg/day) ^b	3.9	6.3	10.2	
Range (µg/day) ^b	0.0-5.1	5.1-7.9	7.9-59.0	
Mean (µg/day) ^b	3.7 ± 1.0	6.4 ± 0.8	11.1 ± 3.3	
Mean (µg/kg bodyweight/day) ^b	0.07 ± 0.05	0.12 ± 0.11	0.22 ± 0.15	
Age at 5-year follow-up survey (years)	58 ± 8	57 ± 8	55 ± 8	<.001
Body mass index (kg/m ²)	24 ± 3	23 ± 3	23 ± 3	<.001
Smoking status				
Current	4.1	4.5	7.2	<.001
Past	0.9	1.0	1.2	
Never	89.6	89.3	86.7	
Missing	5.5	5.3	5.0	
Menarche age				
≤13 years	18.8	24.1	27.7	<.001
14 years	18.9	21.7	22.2	
15 years	19.1	18.7	17.1	
≥16 years	28.3	23.4	19.6	
Missing	14.9	12.2	13.4	
Age at first delivery				
<26 years	50.7	50.9	48.5	<.001
≥26 years	25.9	28.8	30.0	
Missing	23.4	20.3	21.5	
No. deliveries				
None	4.8	5.5	5.7	<.001
1-2	33.4	36.5	37.0	
3	23.7	24.8	23.7	
≥4	19.4	17.9	17.5	
Missing	18.8	15.4	16.1	
Menopause status				
Premenopause	15.7	21.7	29.0	<.001
Postmenopause from unknown age	2.0	1.5	1.5	
Postmenopause from age <49 years	36.9	34.6	32.6	
Postmenopause from age 50-54 years	36.2	36.2	31.8	
Postmenopause from age ≥55 years	4.6	3.7	3.2	
Missing	4.6	2.3	2.0	
Exogenous hormone use				
Yes	2.7	2.5	2.8	<.001
No	89.9	93.1	93.5	
Missing	7.4	4.4	3.7	
Dietary intake				
Energy (kcal/day)	1845 ± 570	1874 ± 553	1858 ± 560	<.001
Alcohol intake (g/week)	16 ± 80	13 ± 58	14 ± 55	<.001
Coffee (g/day) ^b	37 ± 51	92 ± 90	232 ± 238	<.001

(Continues)

TABLE 1 (Continued)

	Tertiles of acrylamide intake			P-value ^a
	Lowest T1 Mean ± SD or %	Middle T2 Mean ± SD or %	Highest T3 Mean ± SD or %	
Green tea (g/day) ^b	343 ± 352	543 ± 429	791 ± 698	<.001
Biscuit and cookies (g/day) ^b	1 ± 1	2 ± 3	6 ± 8	<.001
Potato (g/day) ^b	12 ± 10	19 ± 15	23 ± 24	<.001
Vegetables (g/day) ^b	204 ± 123	231 ± 124	235 ± 135	<.001

^aKruskal-Wallis test for continuous variables and chi-squared test for categorical variables.

^bEnergy-adjusted intake.

3 | RESULTS

3.1 | Participant characteristics

Participant characteristics according to acrylamide intake are listed in Table 1. The mean ± SD dietary acrylamide intake was 3.7 ± 1.0, 6.4 ± 0.8, and 11.1 ± 3.3 µg/day in the lowest, middle, and highest tertiles of dietary acrylamide intake, respectively. Overall, the median dietary acrylamide intake was 6.3 µg/day (IQR, 4.5-8.8 µg/day), and the mean ± SD dietary acrylamide intake was 7.1 ± 3.7 µg/day and 0.14 ± 0.13 µg/kg bodyweight/day in all participants. Foods that highly contributed to the total acrylamide intake were coffee (24%), green tea (22%), biscuit and cookies (13%), potatoes (12%), and vegetables (11%). When the percentages of contributing food were compared between tertiles, coffee (from 14% in the lowest to 29% in the highest), green tea (from 20% in the lowest to 24% in the highest), and biscuit and cookies (from 8% in the lowest to 15% in the highest) increased linearly. The percentages of potatoes (from 16% in the lowest to 10% in the highest) and vegetables (from 17% in the lowest to 8% in the highest), however, decreased linearly.

The highest acrylamide consumption group were younger; had lower BMI; had a larger proportion of current smokers, higher proportion of younger menarche, lower proportion of older first delivery, lower proportion of none or few deliveries, higher proportion of premenopause status, higher proportion of exogenous female hormone non-users; and consumed a higher energy diet, less alcohol, more coffee, more green tea, more biscuit and cookies, more potatoes, and more vegetables than the lowest acrylamide consumption group.

3.2 | Dietary acrylamide intake and endometrial or ovarian cancers

The association between dietary acrylamide intake and endometrial cancer risk is shown in Table 2. In model 1, dietary acrylamide intake significantly decreased the risk of endometrial cancer. When cancer cases occurring ≤3 years of the starting point were excluded, the risk did not change. In addition to the covariates in model 1, however, we added coffee consumption in model 2; the association attenuated and no significant association was observed. This did not change after excluding cases that occurred ≤3 years of the starting

point. Furthermore, although similar associations were observed when stratified according to the confounding factors, there was no significant association in model 2.

The associations between dietary acrylamide intake and ovarian cancer risk are given in Table 3. In contrast to the endometrial cancer analysis, no significant association was observed with ovarian cancer. Furthermore, no significant associations were seen on stratification in any of the strata.

Given the wide range of dietary acrylamide intake in the highest group (7.9-59 µg/day), we divided all participants into 9 quantiles and conducted further analysis to clarify the risk of extremely high consumption (Figure 2). Mean dietary acrylamide intake increased by approximately 1 µg/day between quantiles. No significant association was observed when the highest quantile was compared with the lowest quantile: HR was 0.55 (95%CI: 0.23-1.33) and *P* for trend was 0.18 for endometrial cancer risk; with 0.66 (95%CI: 0.25-1.73) and *P* for trend = 0.32 for ovarian cancer risk.

4 | DISCUSSION

This study identified no associations between dietary acrylamide intake and endometrial or ovarian cancer risks in Japanese women. Specifically, energy-adjusted dietary acrylamide intake was inversely associated with endometrial cancer in model 1, but the significant association disappeared after adjustment for coffee consumption. Furthermore, no associations were observed in either cancers after smoking status, coffee consumption, alcohol consumption, BMI, and menopause status stratifications.

In addition to the null association in all women, we did not detect any significant associations between dietary acrylamide intake and endometrial or ovarian cancers in non-smoking Japanese women. Furthermore, the point estimates showed no increase. In the previous meta-analysis, a non-negligible association was observed in non-smoking women with endometrial cancer (HR, 1.23; 95%CI: 1.00-1.51) and with ovarian cancer (HR, 1.39; 95%CI: 0.97-2.00).⁴ Moreover, the Netherlands Cohort Study on Diet and Cancer (NLCS), the Nurses' Health Study (NHS), and the European Prospective Investigation into Cancer and Nutrition (EPIC) noted increased HR for endometrial and ovarian cancers.³⁵⁻³⁷ In contrast, the Swedish Mammography Cohort (SMC) and the Italian Case-Control

TABLE 2 Acrylamide intake and the risk of endometrial cancer

	Total	Tertiles of acrylamide intake			P for trend
		Lowest (T1) HR (95%CI)	Middle (T2) HR (95%CI)	Highest (T3) HR (95%CI)	
All women					
No. participants	47 185	15 728	15 729	15 728	
No. cases	161	67	51	43	
Person-years	733 067	246 682	244 634	241 751	
Age- and area-adjusted		1.00 (Reference)	0.77 (0.53-1.12)	0.64 (0.43-0.96)	.03
Multivariate-adjusted ^a		1.00 (Reference)	0.76 (0.53-1.10)	0.65 (0.44-0.97)	.03
Multivariate-adjusted (excluding cases ≤ 3 years) ^a		1.00 (Reference)	0.79 (0.53-1.18)	0.68 (0.44-1.05)	.08
Multivariate-adjusted ^b		1.00 (Reference)	0.83 (0.57-1.22)	0.85 (0.54-1.33)	.43
Multivariate-adjusted (excluding cases ≤ 3 years) ^b		1.00 (Reference)	0.85 (0.56-1.28)	0.85 (0.52-1.38)	.46
By smoking status					
Current or past smoker					
No. cases	5	1	0	4	
Multivariate-adjusted ^a		1.00 (Reference)	–	1.68 (0.12-22.83)	.49
Never smoker					
No. cases	149	64	48	37	
Multivariate-adjusted ^a		1.00 (Reference)	0.77 (0.52-1.12)	0.62 (0.40-0.94)	.02
Multivariate-adjusted ^b		1.00 (Reference)	0.85 (0.57-1.25)	0.82 (0.51-1.31)	.37
By coffee consumption					
<1 cup/week					
No. cases	47	32	7	8	
Multivariate-adjusted ^a		1.00 (Reference)	0.51 (0.22-1.18)	1.14 (0.52-2.53)	.77
≥1 cup/week					
No. cases	114	35	44	35	
Multivariate-adjusted ^a		1.00 (Reference)	0.83 (0.53-1.31)	0.59 (0.36-0.96)	.03
Multivariate-adjusted ^b		1.00 (Reference)	0.91 (0.58-1.44)	0.79 (0.46-1.36)	.40
By alcohol consumption					
<150 g/week					
No. cases	157	64	51	42	
Multivariate-adjusted ^a		1.00 (Reference)	0.80 (0.55-1.16)	0.67 (0.44-1.00)	.05
Multivariate-adjusted ^b		1.00 (Reference)	0.88 (0.60-1.29)	0.89 (0.57-1.40)	.58
≥150 g/week					
No. cases	4	3	0	1	
Multivariate-adjusted ^a		1.00 (Reference)	–	–	–
By body mass index					
<25 kg/m ²					
No. cases	103	40	35	28	
Multivariate-adjusted ^a		1.00 (Reference)	0.87 (0.55-1.39)	0.72 (0.43-1.19)	.20
≥25 kg/m ²					
No. cases	55	25	16	14	
Multivariate-adjusted ^a		1.00 (Reference)	0.67 (0.35-1.26)	0.57 (0.29-1.13)	.10
By menopause status					
Premenopause					
No. cases	49	17	14	18	
Multivariate-adjusted ^a		1.00 (Reference)	0.62 (0.30-1.27)	0.67 (0.33-1.35)	.29

(Continues)

TABLE 2 (Continued)

	Total	Tertiles of acrylamide intake			P for trend
		Lowest (T1) HR (95%CI)	Middle (T2) HR (95%CI)	Highest (T3) HR (95%CI)	
Postmenopause					
No. cases	107	45	37	25	
Multivariate-adjusted ^a		1.00 (Reference)	0.90 (0.58-1.40)	0.69 (0.41-1.14)	.15

^aMultivariate-adjusted model 1, adjusted for age (years), area (10 public health center areas), body mass index (<25, ≥25 kg/m², or missing), age at menarche (≤13, 14, 15, ≥16 years, or missing), age at first delivery (<26, ≥26 years, or missing), number of deliveries (0, 1-2, 3, ≥4, or missing), menopause status and age at menopause (premenopause, postmenopause [<49, 50-54, ≥55 years], or missing), use of exogenous female hormones (yes, no, or missing), smoking status (current or ever, never, or missing), and alcohol intake (<150 or ≥150 g/week).

^bMultivariate-adjusted model 2 was further adjusted for energy-adjusted coffee intake (continuous) in addition to the variables in model 1.

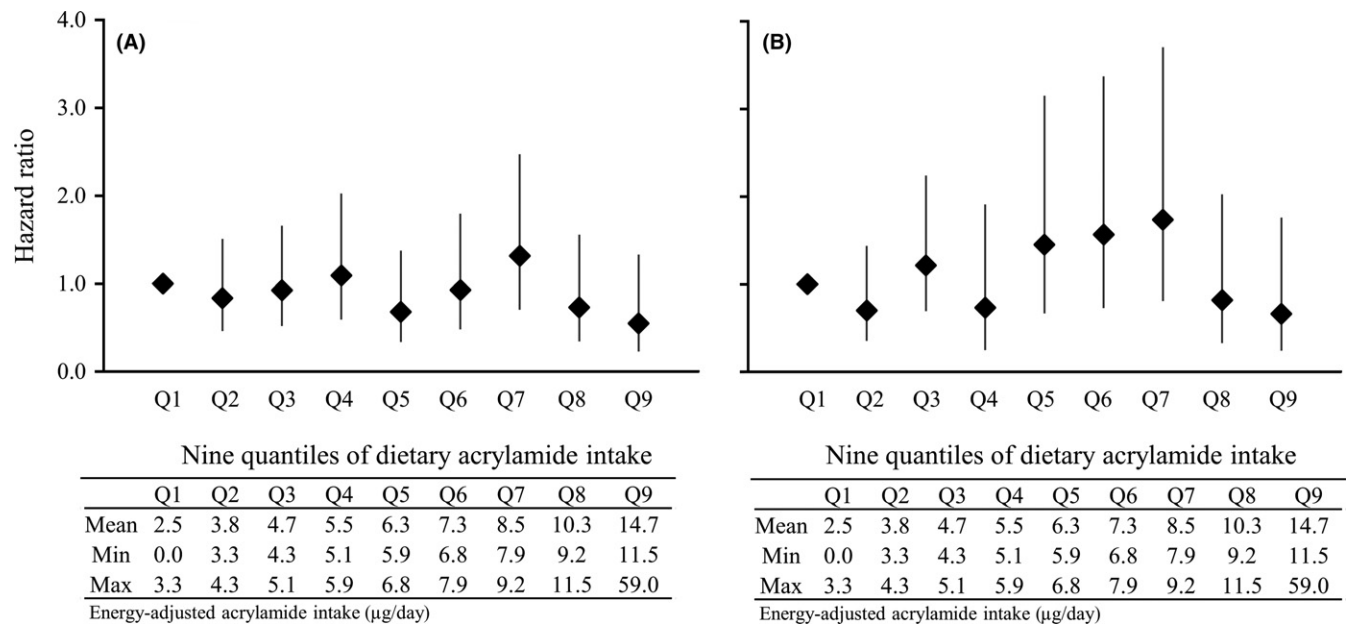


FIGURE 2 Hazard ratio (HR) for A, endometrial cancer risk and B, ovarian cancer risk vs 9 quantiles of dietary acrylamide intake. The reference group was the lowest ninth quantile of energy-adjusted dietary acrylamide intake. A, HR and 95%CI were adjusted for age (years), area (10 public health center areas), body mass index (<25, ≥25 kg/m², or missing), age at menarche (≤13, 14, 15, ≥16 years, or missing), age at first delivery (<26, ≥26 years, or missing), number of deliveries (0, 1-2, 3, ≥4, or missing), menopause status and age at menopause (premenopause, postmenopause [<49, 50-54, ≥55 years], or missing), use of exogenous female hormones (yes, no, or missing), smoking status (current, ever, never, or missing), alcohol intake (<150 or ≥150 g/week), and energy-adjusted coffee intake (continuous). Number of cases from the lowest to the highest of the nine quantiles was 25, 20, 22, 21, 13, 17, 23, 12, and 8, respectively. B, HR and 95%CI were adjusted for age (years), area (10 public health center areas), body mass index (<25, ≥25 kg/m², or missing), age at menarche (≤14, ≥15 years, or missing), age at first delivery (<26, ≥26 years, or missing), number of deliveries (0, 1-2, ≥3, or missing), menopause status (premenopause, postmenopause, or missing), use of exogenous female hormones (yes, no, or missing), smoking status (current or ever, never, or missing), and alcohol intake (<150 or ≥150 g/week). Number of cases from the lowest to the highest of the nine quantiles was 16, 11, 19, 8, 16, 17, 19, 9, and 7, respectively

Studies found null associations for endometrial and ovarian cancers.³⁸⁻⁴¹ Differences in the dietary acrylamide intake in these cohorts may be one of the reasons for the lack of association. The average acrylamide intakes in the lowest and highest quintiles were 9.5 and 36.8 µg/day in the NLCS, and 9 and 26 µg/day in the NHS, respectively.^{35,36} Likewise, the EPIC cohort, which included 10 European countries, also had a wide intake range: the lowest was 8.8 µg/day in Italy and the highest was 35.5 µg/day in Denmark.³⁷ In contrast to these cohorts, the range of intake was narrower in the studies that showed no association.³⁸⁻⁴¹ In the Japanese participants, intake range

of quartiles, as well as the average intake, was considerably smaller than in the other studies that showed a significant association.

Differences in the contributing foods also may have had an impact on the results. In this study, specifically, decreased risk was observed with endometrial cancer in model 1, but the association was attenuated after the adjustment for coffee consumption. Coffee intake was reported as a probable preventive factor for endometrial cancer in the World Cancer Research Fund International.¹⁴ This was consistent with the decreased risk of endometrial cancer due to coffee consumption in the JPHC Study.³⁴

TABLE 3 Acrylamide intake and the risk of ovarian cancer

	Total	Tertiles of acrylamide intake			P for trend
		Lowest (T1) HR (95%CI)	Middle (T2) HR (95%CI)	Highest (T3) HR (95%CI)	
All women					
No. participants	47 185	15 728	15 729	15 728	
No. cases	122	46	41	35	
Person-years	733 572	246 889	244 758	241 925	
Age- and area-adjusted		1.00 (Reference)	0.90 (0.59-1.38)	0.76 (0.48-1.21)	.26
Multivariate-adjusted ^a		1.00 (Reference)	0.90 (0.59-1.38)	0.77 (0.49-1.23)	.28
Multivariate-adjusted (excluding cases ≤ 3 y) ^a		1.00 (Reference)	0.83 (0.52-1.33)	0.69 (0.41-1.16)	.16
By smoking status					
Current or past smoker					
No. cases	4	2	1	1	
Multivariate-adjusted ^a		1.00 (Reference)	0.42 (0.03-5.14)	0.23 (0.02-3.39)	.27
Never smoker					
No. cases	111	41	38	32	
Multivariate-adjusted ^a		1.00 (Reference)	0.94 (0.60-1.48)	0.82 (0.50-1.33)	.43
By coffee consumption					
<1 cup/week					
No. cases	39	25	10	4	
Multivariate-adjusted ^a		1.00 (Reference)	0.90 (0.43-1.90)	0.62 (0.21-1.82)	.40
≥1 cup/week					
No. cases	83	21	31	31	
Multivariate-adjusted ^a		1.00 (Reference)	1.02 (0.58-1.79)	0.95 (0.53-1.71)	.86
By alcohol consumption					
<150 g/week					
No. cases	117	42	40	35	
Multivariate-adjusted ^a		1.00 (Reference)	0.95 (0.61-1.48)	0.84 (0.52-1.35)	.47
≥150 g/week					
No. cases	5	4	1	0	
Multivariate-adjusted ^a		1.00 (Reference)	0.74 (0.05-12.14)	–	.35
By body mass index					
<25 kg/m ²					
No. cases	89	31	33	25	
Multivariate-adjusted ^a		1.00 (Reference)	1.09 (0.66-1.80)	0.83 (0.48-1.45)	.53
≥25 kg/m ²					
No. cases	31	14	8	9	
Multivariate-adjusted ^a		1.00 (Reference)	0.54 (0.22-1.32)	0.64 (0.26-1.55)	.29
By menopause status					
Premenopause					
No. cases	25	8	9	8	
Multivariate-adjusted ^a		1.00 (Reference)	0.87 (0.33-2.27)	0.70 (0.25-1.92)	.48
Postmenopause					
No. cases	94	35	32	27	
Multivariate-adjusted ^a		1.00 (Reference)	0.96 (0.59-1.57)	0.86 (0.51-1.46)	.58

^aMultivariable Cox proportional hazard models were adjusted for age (years), area (10 public health center areas), body mass index (<25, ≥25 kg/m², or missing), age at menarche (≤14, ≥15 y, or missing), age at first delivery (<26, ≥26 y, or missing), number of deliveries (0, 1-2, ≥3, or missing), menopausal status (premenopause, postmenopause, or missing), use of exogenous female hormones (yes, no, or missing), smoking status (current or ever, never, or missing), and alcohol intake (<150 or ≥150 g/week).

In the present study, as dietary acrylamide intake increased, the proportion of acrylamide intake from coffee was increased. It was thought that the significantly decreased risk observed in model 1 might be due to the preventive effect of coffee intake. Therefore, the beneficial effect of coffee intake may be greater than the influence of acrylamide intake accompanying coffee consumption, given that coffee was the highest contributing food to total acrylamide intake in this study. In the NHS, coffee was also one of the contributing foods, but, the risk of endometrial cancer increased.³⁶ This difference may be due to the high acrylamide intake and the difference in contributing foods other than coffee.

This study has several limitations. First, the associations might have been attenuated by FFQ measurement errors. Second, the results may be affected by residual confounding factors such as passive smoking. Given, however, that there were no questions regarding passive smoking in the 5-year follow-up FFQ, we could not include passive smoking in the adjusted model in the present study. To eliminate these effects, therefore, further studies using biomarkers are needed. Also, of the risk assessments on endometrial cancer, the FFQ result is inconsistent with the results using biomarkers such as hemoglobin adduct concentration in blood.^{6,37} Third, the assessment of acrylamide intake was done once. Therefore, we could not consider individual dietary changes during the follow-up period. There may be differences in the contributing food groups by generation, given that the proportion of acrylamide intake from beverages and from vegetables was lower and higher, respectively, in the Japanese 2012 national dietary survey estimations compared with the present results in the 1990s.²² Individual dietary habits, however, might not have changed dramatically, because the present participants were aged 45-74 years and their dietary habits were considered to be well established. Fourth, the low cancer incidence might affect the statistical power. The number of endometrial and of ovarian cancer cases was low in this cohort, reflecting the low incidence rate in Japan: age-standardized rates per 100 000 population in 2012 in Japan were 10.6 for endometrial cancer and 8.4 for ovarian cancer.⁴²

The strengths of this study were that the JPHC Study is one of the largest prospective cohort studies on lifestyle diseases; and recall bias on the exposure was prevented because the data were collected before cancer diagnosis. Participants were selected from the general population and the survey response rate was >80% while the loss to follow-up rate was considerably small. The proportion of DCO was <10% each for endometrial and ovarian cancers. Therefore, the follow-up data and the cancer registry in this study population were of sufficient quality.

This is the first study to assess the effect of dietary acrylamide intake on endometrial or ovarian cancer risks in Asian countries. We found no association between dietary acrylamide intake and endometrial or ovarian cancer risks regardless of smoking status, coffee consumption, alcohol consumption, body size, or menopause status in this large prospective cohort study of Japanese women with a relatively lower dietary intake of acrylamide compared with Western populations.

ACKNOWLEDGMENTS

The members of the JPHC study are listed at the following site (as of April 2017): <https://epi.ncc.go.jp/en/jphc/781/7951.html>. This study was supported by a grant from the Food Safety Commission, Cabinet Office, Government of Japan (Research Program for Risk Assessment Study on Food Safety, No. 1503, the principal investigator is T.S.), the National Cancer Center Research and Development Fund (since 2011, the principal investigator is S.T.), and a Grant-in-Aid for Cancer Research from the Ministry of Health, Labour and Welfare of Japan (from 1989 to 2010; principal investigator from 1997 to 2010 is S.T.).

CONFLICT OF INTEREST

The authors declare no conflict of interest.

AUTHOR CONTRIBUTIONS

J.I. and T.S. designed the research; S.T., T.S., J.I., N.S., and M.I. conducted the research; A.K. contributed to the calculation of dietary acrylamide intake; A.K., L.Z., and R.L. performed the statistical analysis; A.K. interpreted the results and wrote the paper; and J.I. was primarily responsible for the final content. All authors reviewed the manuscript and contributed to the discussion.

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How to cite this article: Kotemori A, Ishihara J, Zha L, et al. Dietary acrylamide intake and the risk of endometrial or ovarian cancers in Japanese women. *Cancer Sci.* 2018;109:3316-3325. <https://doi.org/10.1111/cas.13757>