# Updated association of tea consumption and bone mineral density 

A meta-analysis

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#### Abstract

Background: Current studies evaluating the association of tea consumption and bone mineral density (BMD) have yielded inconsistent findings. Therefore, we conducted a meta-analysis to assess the relationship between tea consumption and BMD. Methods: The PubMed, Embase, and Cochrane Library databases were comprehensively searched, and a meta-analysis performed of all observational studies assessing the association of tea consumption and BMD. Forest plots were used to illustrate the results graphically. The Q-test and $I^{2}$ statistic were employed to evaluate between-study heterogeneity. Potential publication bias was assessed by the funnel plot. Results: Four cohort, 1 case-control, and 8 cross-sectional studies including a total of 12,635 cases were included. Tea consumption was shown to prevent bone loss [odds ratio (OR): 0.66 ; $95 \%$ confidence interval (Cl), $0.47-0.94 ; P=0.02$ ], yielding higher mineral densities in several bones, including the lumbar spine [standardized mean difference (SMD): 0.19; 95\% CI, 0.08-0.31; $P=0.001$ ], hip (SMD: $0.19 ; 95 \% \mathrm{Cl}, 0.05-0.34 ; P=0.01$ ), femoral neck [mean difference (MD): 0.01; 95\% CI, 0.00-0.02; $P=0.04$ ], Ward triangle (MD: $0.02 ; 95 \% \mathrm{Cl}, 0.01-0.04 ; P=0.001$ ), and greater trochanter (MD: $0.03 ; 95 \% \mathrm{Cl}, 0.02-0.04 ; P<0.00001$ ), than the non-tea consumption group. Conclusion: This meta-analysis provided a potential trend that tea consumption might be beneficial for BMD, especially in the lumbar spine, hip, femoral neck, Ward triangle, and greater trochanter, which might help prevent bone loss. Abbreviations: $\mathrm{AHRQ}=$ Agency for Healthcare Research and Quality, BMD = bone mineral density, $\mathrm{Cl}=$ confidence interval, EGCG = epigallocatechin gallate, $\mathrm{MD}=$ mean difference, $\mathrm{NOS}=$ Newcastle-Ottawa scale, $\mathrm{OR}=$ odds ratio, $\mathrm{SMD}=$ standardized mean difference.


Keywords: bone mineral density, meta-analysis, osteoporosis, tea consumption

## 1. Introduction

Osteoporosis, a serious hazard to human health, ${ }^{[1]}$ is a systemic skeletal disease caused by decreased bone mass and microarchitectural degradation of the bone tissue; it produces increasing physical illness and pain. ${ }^{[1-5]}$ Bone mineral density (BMD) is a major indicator of osteoporosis, and determines its severity.

Tea is the most popular beverage in the world; the health advantages of this beverage have been reported in cardiovascular diseases, rheumatoid arthritis, influenza, and cancer. ${ }^{[6-10]}$ The association of tea consumption with BMD has been investigated since the 1990s. ${ }^{[11]}$ However, previous studies have yielded

[^0]inconsistent conclusions. Indeed, some researchers claim a positive relationship between tea consumption and BMD, ${ }^{[12-14]}$ while others support an inverse relationship between them. ${ }^{[11]}$ Meanwhile, no correlation was found in some studies. ${ }^{[15-17]}$

To clarify the relationship of tea consumption and BMD more exactly and systematically, a meta-analysis of all available studies was performed.

## 2. Methods

### 2.1. Literature search

Electronic databases, including Embase, PubMed, and Cochrane Library, were searched comprehensively for all relevant literature, without restriction to regions, languages, or publication types. The following MeSH-free words and their combinations were searched in all fields: (("Tea" [MeSH]) or ("Tea" [Free words])) AND ((("Osteoporosis" [MeSH]) or ("Osteoporosis" [Free words])) or (("Bone Density" [MeSH]) or ("Bone Density" [Free words]))). In addition, the reference lists of the selected literature were also used to expand the search. Two investigators independently searched for articles, and reviewed all the retrieved studies. Disagreement between the 2 investigators was resolved by consensus, involving a third reviewer.

### 2.2. Inclusion and exclusion criteria

Inclusion criteria were as follows: observational studies with cohort, case-control, and cross-sectional designs, respectively (there was no relevant double-blind, placebo controlled trials); tea
consumption as exposure; BMD or osteoporosis as outcome; results including means and standard deviations or dichotomous data, or providing sufficient information to derive the latter. Exclusion criteria were reviews, repeated, or overlapped publications; herbal tea; no measurement locations of BMD; animal studies; and studies with unavailable data.

### 2.3. Data extraction

Relevant information from each eligible study was extracted independently by 2 reviewers (Zhang and Yang) using a standardized form. Information included study name (first author), publication year, study location, study design, sample size, tea type, measurement locations of BMD, and study results.

### 2.4. Quality assessment

The Newcastle-Ottawa scale (NOS), with scores of 0 to 9 (allocated as stars), was used to assess the methodological quality of cohort and case-control studies, ${ }^{[18,19]}$ and high quality was considered for articles with 6 or more stars. In addition, crosssectional studies were assessed by an 11-item checklist recommended by Agency for Healthcare Research and Quality (AHRQ). ${ }^{[20]}$ An "UNCLEAR" or "NO" answer conferred a score of " 0 " to the item, and "YES" a score of " 1 ." The quality of the article was assessed as follows: low quality, 0 to 3 ; medium quality, 4 to 7 ; high quality, 8 to 11 . ${ }^{[20]}$

### 2.5. Ethical statement

All results and analyses were from previous published studies; thus, no ethical approval and patient consent are required.

### 2.6. Statistical analysis

All statistical analyses were performed with Review Manager 5.3 (The Cochrane Collaboration, Copenhagen, Denmark). In the pooled results, mean difference (MD) and odds ratio (OR) were used to compare continuous and dichotomous variables, respectively; standardized MD (SMD) were used for different units. All results were expressed with $95 \%$ confidence interval (CI). Heterogeneity was measured by the Q-test with $\mathrm{I}^{2}$ statistics, with $P<0.10$ and $\mathrm{I}^{2}>50 \%$ indicating high heterogeneity. ${ }^{[21,22]}$ The random-effects model was used, when there was significant heterogeneity between-studies; otherwise, the fixed-effects model was employed. ${ }^{[23]}$ Sensitivity analyses were used to evaluate the impact of each study on the pooled results by removing each study in turn, ${ }^{[10,24]}$ assessing whether the quality of articles affected the overall results. ${ }^{[25]}$ Funnel plots were used to assess potential publication bias.

## 3. Results

### 3.1. Study selection

A total of 364 relevant articles were obtained in the initial literature search. Then, 82 articles were excluded because of duplication; 248 others which were obviously irrelevant, reviews, or animal studies were also excluded by screening titles or abstracts. Studies with unavailable data or assessing herbal tea were also removed after full text reading. Finally, 13 articles were included in this study, evaluating 12,635 cases (6059 and 6576 individuals in the tea- and non-tea consumption groups) (Fig. 1).


Figure 1. Flow diagram of screened, included, and excluded studies.

Table 1
Characteristics of included studies (continuous data).

| Study ID | Study location | Study design | Tea types | Measurement locations (BMD) | Non-tea consumption (means $\pm$ SDs, except: ${ }^{\dagger,+}$ ) | Tea consumption (means $\pm$ SDs,except: ${ }^{\dagger,+ \text { ) }}$ ) | Quality score |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Wang et al ${ }^{[32]}$ | China | Cohort | Oolong tea | ( $\mathrm{g} / \mathrm{cm}^{2}$ ) | $\mathrm{n}=556$ | $n=124$ | 7 |
|  |  |  |  | L2-4 | $0.745 \pm 0.05$ | $0.767 \pm 0.010$ |  |
|  |  |  |  | Femoral neck | $0.796 \pm 0.121$ | $0.796 \pm 0.108$ |  |
|  |  |  |  | Greater trochanteric | $0.759 \pm 0.116$ | $0.793 \pm 0.119$ |  |
|  |  |  |  | Ward triangle | $0.637 \pm 0.135$ | $0.668 \pm 0.133$ |  |
| Hsiao et al ${ }^{[33]}$ | Taiwan | Cross-sectional | Tea | Hip | $\begin{gathered} n=55 \\ -1.35 \pm 0.94^{*} \end{gathered}$ | $\begin{gathered} n=45 \\ -0.89 \pm 0.93^{*} \end{gathered}$ | 3 |
| Muraki et al ${ }^{[35]}$ | Japan | Cross-sectional | Green tea | $\left(\mathrm{g} / \mathrm{cm}^{2}\right)$ | n=52 | $n=580$ | 7 |
|  |  |  |  | Lumbar spine | $0.733 \pm 0.182$ | $0.807 \pm 0.187$ |  |
| Hossein et al ${ }^{[34]}$ | Iran | Cross-sectional | Tea | ( $\mathrm{g} / \mathrm{cm}^{2}$ ) | $\mathrm{n}=705$ | $n=125$ | 9 |
|  |  |  |  | Spine ( $L_{2-4}$ ) men | $1.2 \pm 0.2(\mathrm{n}=290)$ | $1.2 \pm 0.2(n=35)$ |  |
|  |  |  |  | women | $1.1 \pm 0.2(\mathrm{n}=415)$ | $1.15 \pm 0.0(\mathrm{n}=90)$ |  |
|  |  |  |  | Hip men | $1.01 \pm 0.1(n=290)$ | $1.01 \pm 0.1(n=35)$ |  |
|  |  |  |  | women | $0.95 \pm 0.1(n=415)$ | $0.99 \pm 0.11$ ( $\mathrm{n}=90$ ) |  |
| Devine et al ${ }^{[31]}$ | Australian | Cross-sectional | Tea | $\left(\mathrm{mg} / \mathrm{cm}^{2}\right)$ | $\mathrm{n}=172$ | $n=855$ | 10 |
|  |  |  |  | Hip | $782(764,800)^{\dagger}$ | $806(799,814)^{\dagger}$ |  |
|  |  |  |  | Femoral neck | $665(650,680)^{\dagger}$ | $681(674,688){ }^{\dagger}$ |  |
|  |  |  |  | Trochanter | $603(587,620)^{\dagger}$ | $630(623,637)^{\dagger}$ |  |
|  |  |  |  | Intertrochanter | $927(905,949){ }^{\dagger}$ | $950(941,960)^{\dagger}$ |  |
| Hamdi et al ${ }^{[30]}$ | Turkey | Cross-sectional | Tea | Phalanges | $\begin{gathered} n=60 \\ -1.51 \pm 1.68^{*} \end{gathered}$ | $\begin{gathered} n=393 \\ -1.09 \pm 1.66 \end{gathered}{ }^{*} \text { * }$ | 7 |
| Chen et al ${ }^{[17]}$ | United States | Cohort | Tea | $\left(\mathrm{g} / \mathrm{cm}^{2}\right)$ | <1 cup/day ( $\mathrm{n}=3683$ ) | 1 cup/day ( $\mathrm{n}=566$ ) | 8 |
|  |  |  |  | Total body | 1.021 (1.020, 1.022) ${ }^{\dagger}$ | 1.019 (1.016,1.022) ${ }^{\dagger}$ |  |
|  |  |  |  | Total hip | 0.846 (0.845, 0.847) ${ }^{\dagger}$ | 0.847 (0.844,0.850) ${ }^{\dagger}$ |  |
|  |  |  |  | Lumbar spine | $0.989(0.988,0.991)^{\dagger}$ | $0.991(0.986,0.995)^{\dagger}$ |  |
|  |  |  |  | Total bodyTotal hip |  | 2-3 cups/day ( $n=588$ ) |  |
|  |  |  |  |  |  | $1.024(1.021,1.027)^{\dagger}$ |  |
|  |  |  |  | Lumbar spine |  | $0.848(0.845,0.851)^{\dagger}$ |  |
|  |  |  |  | Total body |  | 0.995 (0.991, 0.999) ${ }^{\dagger}$ |  |
|  |  |  |  | Total hip |  | $\geq 4$ cups/day ( $\mathrm{n}=142$ ) |  |
|  |  |  |  | Lumbar spine |  | 1.029 (1.023,1.036) ${ }^{\dagger}$ |  |
|  |  |  |  |  |  | 0.846 (0.840,0.850) ${ }^{\dagger}$ |  |
|  |  |  |  |  |  | 0.990 (0.981,0.998) ${ }^{\dagger}$ |  |
| Wu et al ${ }^{[14]}$ | Taiwan | Cohort | Tea | ( $\mathrm{g} / \mathrm{cm}^{2}$ ) | $\mathrm{n}=535$ | $1-5$ y ( $\mathrm{n}=226$ ) | 6 |
|  |  |  |  | Total body | $1.150 \pm 0.007^{\text {\# }}$ | $1.155 \pm 0.008^{\ddagger}$ |  |
|  |  |  |  | Spine ( $\mathrm{L}_{1-4}$ ) | $1.114 \pm 0.014^{\ddagger}$ | $1.127 \pm 0.015^{\text {* }}$ |  |
|  |  |  |  | Hip neck | $0.851 \pm 0.011^{\text { }}$ | $0.851 \pm 0.012^{\text { }}$ |  |
|  |  |  |  | Ward triangle | $0.741 \pm 0.012^{\ddagger}$ | $0.748 \pm 0.014^{\text {* }}$ |  |
|  |  |  |  | Total body |  | $6-10$ y ( $\mathrm{n}=152$ ) |  |
|  |  |  |  | Spine ( $\mathrm{L}_{1-4}$ ) |  | $1.158 \pm 0.009$ * |  |
|  |  |  |  | Hip neck |  | $1.150 \pm 0.017^{*}$ |  |
|  |  |  |  | Ward triangle |  | $0.865 \pm 0.013^{\text {\# }}$ |  |
|  |  |  |  | Total body |  | $0.758 \pm 0.015^{\ddagger}$ |  |
|  |  |  |  | Spine ( $\mathrm{L}_{1-4}$ ) |  | $>10$ y ( $\mathrm{n}=124$ ) |  |
|  |  |  |  | Hip neck |  | $1.174 \pm 0.009^{\text {F }}$ |  |
|  |  |  |  | Ward triangle |  | $1.162 \pm 0.018^{\ddagger}$ |  |
|  |  |  |  |  |  | $0.891 \pm 0.014^{\text {* }}$ |  |
|  |  |  |  |  |  | $0.787 \pm 0.016^{\text { }}$ |  |
| Hegarty et al ${ }^{[13]}$ | United Kingdom | Cohort | Tea | $\left(\mathrm{g} / \mathrm{cm}^{2}\right)$ | $\mathrm{n}=122$ | $\mathrm{n}=1134$ | 6 |
|  |  |  |  | Lumbar spine | $0.889 \pm 0.16$ | $0.917 \pm 0.17$ |  |
|  |  |  |  | Femoral neck | $0.660 \pm 0.10$ | $0.669 \pm 0.11$ |  |
|  |  |  |  | Greater trochanter | $0.588 \pm 0.09$ | $0.612 \pm 0.15$ |  |
|  |  |  |  | Ward triangle | $0.454 \pm 0.12$ | $0.476 \pm 0.13$ |  |

$\mathrm{BMD}=$ bone mineral density, $\mathrm{L}=$ lumbar, $\mathrm{SDs}=$ standard deviations

* T score.
${ }^{\dagger}$ Mean, $95 \%$ confidence interval in parentheses.
${ }^{\text {Mean, standard error. }}$


### 3.2. Characteristics of the included studies

Ten studies were from PubMed, 3 from Embase, and none from the Cochrane Library. All included articles had available full texts. There were 5 studies with count data ( 3 and 2 studies with ordinal and dichotomous data, respectively) ${ }^{[26-30]} ; 9$ studies had continuous data, ${ }^{[13,14,17,30-35]}$ with, 1 presenting both dichotomous and continuous data. ${ }^{[30]}$ Besides, for continuous data,

BMD values of the femoral neck, Ward triangle, greater trochanter, total body, intertrochanteric hip, lumbar spine, hip, and phalanges were presented in the 9 studies. No further evaluation study was found in the reference lists of these studies. In addition, the reference lists of relevant articles did not produce further studies for evaluation. The characteristics of all studies are presented in Tables 1 and 2.

Table 2
Characteristics of included studies (dichotomous data).

| Study ID | Study location | Study design | Tea types | Measurement locations | Degree | Non-tea consumption ( n ) | Tea consumption ( n ) | Quality score |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Keskin et al ${ }^{[28]}$ | Turkey | Cross-sectional | Tea | Phalanges | Normal | 226 | 306 | 8 |
|  |  |  |  |  | Osteoporosis | 52 | 36 |  |
| Chan et al ${ }^{[29]}$ | Malaysia | Cross-sectional | Tea | Hip | Normal | 23 | 11 | 7 |
|  |  |  |  |  | Osteopenia | 31 | 11 |  |
| Alquaiz et al ${ }^{[27]}$ | Saudi Arabia | Cross-sectional | Green tea | Spine ( $L_{1-4}$ ) and dual femur | Normal | 10 | 140 | 6 |
|  |  |  |  |  | Osteopenia | 27 | 185 |  |
| Wang et al ${ }^{[26]}$ | China | Case-control | Tea | Left femur | Normal | 81 | 29 | 7 |
|  |  |  |  |  | Osteopenia | 97 | 40 |  |
|  |  |  |  |  | Osteoporosis | 59 | 18 |  |
| Hamdi et al ${ }^{[30]}$ | Turkey | Cross-sectional | Tea | Phalanges | Normal | 36 | 265 | 8 |
|  |  |  |  |  | Osteopenia | 20 | 169 |  |
|  |  |  |  |  | Osteoporosis | 34 | 178 |  |

$\mathrm{L}=$ lumbar.

### 3.3. Meta-analysis results

Data from 5 studies with dichotomous variables (1 study both dichotomous and continuous data), including 1398 and 696 in the tea- and non-tea consumption groups, were assessed in this meta-analysis. Two of the studies claimed tea consumption to have a beneficial impact on BMD compared with non-tea consumption ${ }^{[28,29]}$; no correlation between them was found in the remaining 3 studies. ${ }^{[26,27,30]}$ The pooled results showed that tea consumption could reduce the occurrence of low bone mass ( $\mathrm{OR}=0.66,95 \% \mathrm{CI}=0.47-0.94, P=0.02$ ) (Fig. 2).

We also performed a meta-analysis for continuous variables according to BMD at different locations in the 9 studies. ${ }^{[13,14,17,30-35]}$ The pooled results showed that BMD values of the femoral neck (MD: $0.01,95 \% \mathrm{CI}, 0.00-0.02 ; P=$ 0.04 ), Ward triangle (MD: $0.02,95 \% \mathrm{CI}, 0.01-0.04, P=0.001$ ), greater trochanter (MD: $0.03,95 \% \mathrm{CI}, 0.02-0.04, P<0.00001$ ), lumbar spine (SMD: $0.19,95 \% \mathrm{CI}, 0.08-0.31, P=0.001$ ), and hip (SMD: $0.19,95 \% \mathrm{CI}, 0.05-0.34, P=0.01$ ) were higher in the tea consumption group than in the non-tea consumption group (Figs. 3 and 4). However, no statistically significant differences in BMD values of the total body (MD: $0.00,95 \% \mathrm{CI},-0.00$ to 0.00 , $P=0.06$ ), intertrochanteric hip (MD: 0.02, $95 \% \mathrm{CI},-0.01$ to $0.05, P=0.06$ ), and phalanges (SMD: 0.25, $95 \% \mathrm{CI},-0.02$ to $0.52, P=0.07$ ) were obtained (Figs. 3 and 4).

### 3.4. Sensitivity and publication bias

Sensitivity analysis was performed by sequentially excluding each study in order to examine the influence of individual studies on
the overall assessment. Interestingly, none of the studies fairly affected the overall findings. In addition, sensitivity analysis was conducted to access high-quality studies, and the results did not obviously change. Figure 5 shows a funnel plot for continuous variables, with slight asymmetry.

## 4. Discussion

Tea is the most frequently consumed beverage in daily life, and its protective or harmful impacts on human health are a major public issue. This was the first meta-analysis to address this subject. Our analysis of 4 cohort, 1 case-control, and 8 crosssectional studies with 12,635 cases indicated a potential trend that tea consumption might result in higher BMD at the femoral neck, Ward triangle, greater trochanter, lumbar spine, and hip than non-tea consumption, and might prevent bone loss.

These findings corroborate many other studies. ${ }^{[12,36,37]}$ For instance, a study including 2016 cases reported that tea consumption has protective effects on the femoral neck and lumbar spine with T scores $>-0.75 .{ }^{[36]}$ In addition, Hoover et al ${ }^{[12]}$ conducted a study of 62 postmenopausal women, and found that tea consumption yields $10 \%$ and $14 \%$ higher BMD values in the lumbar spine and femoral neck, respectively, than the non-tea consumption group. Similar results were found in animal experiments. Indeed, studies showed that tea polyphenols increase femoral BMD of female rats. ${ }^{[38-41]}$ However, others reported that tea consumption has no effect on BMD. ${ }^{[11,42]}$ Hernández et al ${ }^{[11]}$ found no association of tea consumption with BMD, and even an inverse correlation. In addition, a study


Figure 2. Forest plot and meta-analysis of dichotomous variables. $\mathrm{Cl}=$ confidence interval, $\mathrm{M}-\mathrm{H}=$ Mantel-Haenszel method.


Figure 3. Forest plot and meta-analysis of BMD in the femoral neck, Ward triangle, greater trochanteric, total body, and intertrochanteric hip. $\mathrm{Cl}=\mathrm{confidence}$ interval, IV = inverse variance method, SD = standard deviation.
evaluating 201 cases with osteoporosis or osteopenia in Turkey reported no association of tea consumption with BMD of the lumbar spine, and a negligible correlation with the femoral neck. ${ }^{[42]}$ The above 2 studies, which were not adjusted, were inconsistent with our series. Besides, a prospective randomized controlled study also found no correlation between tea and BMD, ${ }^{[43]}$ the main reason likely being the small sample size.

Tea polyphenols contain abundant epigallocatechin gallate (EGCG) as the main component of tea. ${ }^{[44,45]}$ EGCG has been widely studied. Its beneficial effects on bone formation are mainly through increasing alkaline phosphatase activity at both the protein and gene expression levels in osteoblastic-like cells, including SaOS-2 cells ${ }^{[46]}$ and MC3T3-E1 cells, and increased formation of mineralized bone. ${ }^{[47]}$ Moreover, tea polyphenols significantly promote osteoblastic survival and decrease osteoblastic apoptosis, ${ }^{[46]}$ thereby leading to elevated cell proliferation and differentiation. ${ }^{[48]}$ This was the theoretical basis of our research.

In this meta-analysis, no correlation was found between tea consumption and BMD of the whole body, intertrochanteric hip, and phalanges. This finding might be unreliable because only 1 or 2 studies assessing these entities were included. Meanwhile, 5 studies of dichotomous variables were included in this metaanalysis; the pooled results indicated that tea consumption could prevent bone loss. However, we did not analyze the effects of tea
consumption on BMD at each location due to insufficient data provided in the above 5 studies.

Means and standard deviations were not provided for continuous data in some included articles; means and standard deviations were derived on the basis of the Cochrane Handbook. Furthermore, other studies provided ordinal data, such as normal, osteopenia, and osteoporosis; osteopenia and osteoporosis were combined into low bone mass to generate dichotomous data, as they all belong to bone mass reduction; the pooled results were similar to those of continuous data. In addition, considering that T values reflect the true levels of BMD, we included 2 studies that provided only T values. ${ }^{[30,33]}$ We combined T values with BMD values using SMD, and only BMD values using MD.

Next, we performed sensitivity analysis, including only highquality studies in order to evaluate any impact of study quality on the effect estimates. The results did not change obviously. Consequently, the low-quality studies included had no significant effects on the overall results. Moreover, between-study heterogeneity was not significant for dichotomous and continuous variables, but significant for the lumbar spine and hip. The main reason for the heterogeneity might be from different locations, populations, tea types, or study designs. These studies used the random-effects model to reduce the effect of heterogeneity; however, it was not eliminated.

This meta-analysis had the following limitations that must be taken into account. First and foremost, it included a limited


Figure 4. Forest plot and meta-analysis of BMD in the lumbar spine, hip, and phalanges. ; $\mathrm{Cl}=$ confidence interval, $\mathrm{IV}=$ inverse variance method, $\mathrm{SD}=$ standard deviation.


Figure 5. Funnel plot. $\mathrm{SE}=$ standard error, $\mathrm{SMD}=$ standardized mean difference.
number of observational trials with some being low and mediumquality studies, which might result in selective and performance bias owing to the absence of random allocation, allocation concealment, and blinding. Secondly, subgroup analysis of age, tea types, and tea in cups was not performed because of few included studies.

In summary, these findings provide a potential trend of beneficial effects of tea consumption on BMD, especially in the lumbar spine, hip, femoral neck, Ward triangle, and greater trochanter, which might help prevent bone loss.

## References

[1] WHO Scientific Group on the Assessment of Osteoporosis at Primary Care Level. Summary Meeting Report. May 5-7, 2004; Brussels, Belgium.
[2] G. Office of the Surgeon, Reports of the Surgeon General. Bone Health and Osteoporosis: A Report of the Surgeon General. Office of the Surgeon General (US), Rockville (MD), 2004.
[3] Brown JP, Josse RG. Scientific Advisory Council of the Osteoporosis Society of Canada. 2002 Clinical practice guidelines for the diagnosis and management of osteoporosis in Canada. CMAJ 2002;156:S1-34.
[4] Yuen CK, Kendler D, Khan AA, et al. Osteoporosis: Canadian consensus on menopause and osteoporosis. JOGC 2002;24:35-45.
[5] Johnell O, Kanis JA. An estimate of the worldwide prevalence and disability associated with osteoporotic fractures. Osteoporos Int 2006; 17:1726-33.
[6] Trevisanato SI, Kim YI. Tea and health. Nutr Rev 2000;58:1-0.
[7] Mikuls TR, Cerhan JR, Criswell LA, et al. Coffee, tea, and caffeine consumption and risk of rheumatoid arthritis: results from the Iowa Women's Health Study. Arthritis Rheum 2002;46:83-91.
[8] Sueoka N, Suganuma M, Sueoka E, et al. A new function of green tea: prevention of lifestyle-related diseases. Ann N Y Acad Sci 2001;928: 274-80.
[9] Kazuki I, Hiroshi Y, Yohei K. Effect of gargling with tea and ingredients of tea on the prevention of influenza infection: a meta-analysis. BMC Public Health 2016;16:396.
[10] Shihao W, Fei L, Xiao H, et al. The association of tea consumption with bladder cancer risk: a meta-analysis. Asia Pac J Clin Nutr 2013;22: 128-37.
[11] Hernández AM, Stampfer MJ, Ravnikar VA, et al. Caffeine and other predictors of bone density among pre- and perimenopausal women. Epidemiology 1993;4:128-34.
[12] Hoover PA, Webber CE, Beaumont LF, et al. Postmenopausal bone mineral density: relationship to calcium intake, calcium absorption, residual estrogen, body composition, and physical activity. Can J Physiol Pharmacol 1996;74:911-7.
[13] Hegarty VM, May HM, Khaw KT. Tea drinking and bone mineral density in older women. Am J Clin Nutr 2000;71:1003-7.
[14] Wu CH, Yang YC, Yao WJ, et al. Epidemiological evidence of increased bone mineral density in habitual tea drinkers. Arch Intern Med 2002;162:1001-6.
[15] Kyriazopoulos P, Trovas G, Charopoulos J, et al. Lifestyle factors and forearm bone density in young Greek men. Clin Endocrinol 2006; 65:234-8.
[16] Saitoglu M, Ardicoglu O, Ozgocmen S, et al. Osteoporosis risk factors and association with somatotypes in males. Arch Med Res 2007;38: 746-51.
[17] Chen X, Pettinger MB, Ritenbaugh C, et al. Habitual tea consumption and risk of osteoporosis: a prospective study in the women's health initiative observational cohort. Am J Epidemiol 2003;158:772-81.
[18] Taggart DP, Amico R, Altman DG. Effect of arterial revascularization on survival: a systematic review of studies comparing bilateral and single internal mammary arteries. Lancet 2001;358:870-5.
[19] Wells GA, Shea B, O'Connell D, et al. The Newcastle-Ottawa Scale (NOS) for Assessing the Quality of Non-Randomized Studies in MetaAnalysis. Appl Eng Agric 2000;18:727-34.
[20] Jiande H, Yonghai D, Xiaodan C, et al. Prevalence of suicide attempts among Chinese adolescents: a meta-analysis of cross-sectional studies. Compr Psychiatry 2015;61:78-89.
[21] Higgins JP, Thompson SG, Deeks JJ, et al. Measuring inconsistency in meta-analyses. BMJ 2003;327:557-60.
[22] Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. Stat Med 2002;21:1539-58.
[23] Higgins J, Green S. Cochrane Handbook for Systematic Reviews of Interventions. Cochrane Collaboration, New York, NY:2008.
[24] Li Z, Yu J, Miao Q, et al. The association of fish consumption with bladder cancer risk: a meta analysis. World J Surg Oncol 2011;9:107.
[25] Xinxiang F, Tianxin L, Kewei X, et al. Laparoendoscopic single-site nephrectomy compared with conventional laparoscopic nephrectomy: a systematic review and meta-analysis of comparative studies. Eur Urol 2012;62:601-12.
[26] Wang S, Lin S, Zhou Y, et al. Social and behavior factors related to aged Chinese women with osteoporosis 2008;24:538-45.
[27] Alquaiz AM, Kazi A, Tayel S, et al. Prevalence and factors associated with low bone mineral density in Saudi women: a community based survey. BMC Musculoskelet Disord 2014;15:5.
[28] Keskin Y, Çekin MD, Gündüz H, et al. The prevalence of osteoporosis in the Thrace region of Turkey: a community-based study. Turk Fiz Tip Rehab D 2014;60:335-40.
[29] Chan PJ, Nurul ZZ, Chuah JS, et al. Association between risk factors of osteoporosis and bone mineral density in women of different ethnic groups in a Malaysian hospital. Int J Osteoporosis Metab Disord 2014;7:1-1.
[30] Hamdi KI, Aydin S, Gemalmaz A, et al. Habitual tea drinking and bone mineral density in postmenopausal Turkish women: investigation of prevalence of postmenopausal osteoporosis in Turkey. Int J Vitam Nutr Res 2007;77:389-97.
[31] Devine A, Hodgson JM, Dick IM, et al. Tea drinking is associated with benefits on bone density in older women. Am J Clin Nutr 2007;86: 1243-7.
[32] Wang G, Liu G, Zhao H, et al. Oolong tea drinking could help prevent bone loss in postmenopausal Han Chinese women. Cell Biochem Biophys 2014;70:1289-93.
[33] Hsiao MC, Liu CY, Wang CJ. Factors associated with low bone density among women with major depressive disorder. Int J Geriatr Psych 2012;44:77-90.
[34] Hossein N, Maghbooli ZH, Shafaie AR, et al. Relationship between tea drinking and bone mineral density in Iranian population. Iranian J Publ Health 2007;13:57-62.
[35] Muraki S, Yamamoto S, Ishibashi H, et al. Diet and lifestyle associated with increased bone mineral density: cross-sectional study of Japanese elderly women at an osteoporosis outpatient clinic. J Orthop Sci 2007;12:317-20.
[36] Vestergaard P, Hermann AP, Gram J, et al. Evaluation of methods for prediction of bone mineral density by clinical and biochemical variables in perimenopausal women. Maturitas 2001;40:211-20.
[37] Keramat A, Patwardhan B, Larijani B, et al. The assessment of osteoporosis risk factors in Iranian women compared with Indian women. BMC Musculoskelet Disord 2008;27:28.
[38] Shen CL, Wang P, Guerrieri J, et al. Protective effect of green tea polyphenols on bone loss in middle-aged female rats. Osteoporosis Int 2008;19:979-90.
[39] Shen CL, Yeh JK, Stoecker BJ, et al. Green tea polyphenols mitigate deterioration of bone microarchitecture in middle-aged female rats. Bone 2009;44:684-90.
[40] Kanis J, Johnell O, Gullberg B, et al. Risk factors for hip fracture in men from southern Europe: the MEDOS Study. Osteoporos Int 1999; 9:45-54.
[41] Shen CL, Yeh JK, Stoecker BJ, et al. Green tea polyphenols protects bone microarchitecture in female rats with chronic inflammation-induced bone loss. JBMR 2008;23:458.
[42] Beyazal MS, Çapkin E, Karkucak M, et al. The relationship of osteoporosis risk factors with bone mineral density in patients admitted our outpatient clinic in Trabzon. Turk Osteoporoz Dergisi 2016;22: 17-23.
[43] Dostal AM, Arikawa A, Espejo L. Long-term supplementation of green tea extract does not modify adiposity or bone mineral density in a randomized trial of overweight and obese postmenopausal women. Brit J Nutr 2016;146:256-64.
[44] Chwan LS, James K, Yeh JC, et al. Green tea and bone metabolism. Nutr Res 2009;29:437-56.
[45] Shen CL, Chyu MC. Tea flavonoids for bone health: from animals to humans. J Investig Med 2016;64:1151-7.
[46] Choi EM, Hwang JK. Effects of (+)-catechin on the function of osteoblastic cells. Biol Pharm Bull 2003;26:523-6.
[47] Vali B, Rao LG, El-Sohemy A. Epigallocatechin-3-gallate increases the formation of mineralized bone nodules by human osteoblast-like cells. J Nutr Biochem 2007;18:341-7.
[48] Mizutani A, Sugiyama I, Kuno E, et al. Expression of matrix metalloproteinases during ascorbate-induced differentiation of osteoblastic MC3T3-E1 cells. J Bone Miner Res 2001;16:2043-9.


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