JBMR PLUS

ASBMR

Are the Relationships of Lean Mass and Fat Mass With Bone Microarchitecture Causal or Due to Familial Confounders? A Novel Study of Adult Female Twin Pairs

Minh Bui,¹ D Roger Zebaze,² Shuai Li,^{1,3} John L Hopper,¹ and Åshild Bjørnerem^{4,5}

¹Centre for Epidemiology and Biostatistics, School of Population and Global Health, University of Melbourne, Melbourne, Victoria, Australia ²Department of Medicine, School of Clinical Sciences, Monash University, Melbourne, Victoria, Australia

³Centre for Cancer Genetic Epidemiology, Department of Public Health and Primary Care, University of Cambridge, Cambridge, UK

⁴Department of Clinical Medicine, UiT – The Arctic University of Norway, Tromsø, Norway

⁵Department of Obstetrics and Gynecology, University Hospital of North Norway, Tromsø, Norway

ABSTRACT

It is not known whether the relationships of lean mass (LM) and fat mass (FM) with bone microarchitecture and geometry are causal and/or are because of confounders, including familial confounders arising from genetic and environment effects shared by relatives. We tested the hypotheses that: (i) LM is associated with cortical bone traits, (ii) FM is associated with trabecular bone traits, and (iii) these relationships of LM and FM with bone microarchitecture and geometry have a causal component. Total body composition was quantified for 98 monozygotic (MZ) and 54 dizygotic (DZ) white female twin pairs aged 31 to 77 years. Microarchitecture at the distal tibia and distal radius was quantified using HRpQCT and StrAx software. We applied the Inference about Causation through Examination of FAmiliaL CONfounding (ICE FALCON) method. Within-individuals, distal tibia total bone area, cortical area, cortical thickness, and trabecular number were positively associated with LM (standardized regression coefficient (β) = 0.13 to 0.43; all p < 0.05); porosity of the inner transitional zone (ITZ) was negatively associated with LM ($\beta = -0.22$; p < 0.01). Trabecular number was positively associated with FM (β = 0.40; p < 0.001), and trabecular thickness was negatively associated with FM (β = -0.27; p < 0.001). For porosity of ITZ and trabecular number, the cross-pair cross-trait association with LM was significant before and after adjustment for the within-individual association with LM (all ps < 0.05). For trabecular number, the cross-pair cross-trait association with FM was significant before and after adjustment for the within-individual association with FM (p < 0.01). There were no significant changes in these cross-pair cross-trait associations after adjustment for the within-individual association (p = 0.06 to 0.99). Similar results were found for distal radius measures. We conclude that there was no evidence that the relationships of LM and FM with bone microarchitecture and geometry are causal; they must in part due to by familial confounders affecting both bone architecture and body composition. © 2020 The Authors. JBMR Plus published by Wiley Periodicals LLC. on behalf of American Society for Bone and Mineral Research.

KEY WORDS: BONE MICROARCHITECTURE; CAUSATION; FAT MASS; GENETIC FACTORS; LEAN MASS; TWIN PAIRS

Introduction

W ith advancing age, the loss of muscle (lean mass [LM]) is associated with a reduced bone mass, more falls, and increased fracture risk.⁽¹⁻⁸⁾ The relationship of fat mass (FM) with fracture risk is complex; low BMI is associated with low bone mass and fractures at some sites, whereas obesity is associated with fractures at other sites.⁽⁹⁻¹¹⁾

When studying the relationship of bone architecture with body composition, both LM and FM must be considered

independently and together.^(12,13) When considered together, it has been found that cortical, but not trabecular, microarchitecture is associated with LM, whereas trabecular, not cortical, microarchitecture is mainly associated with FM, at least for post-menopausal women.⁽¹²⁾ Cortical porosity does not appear to be associated with LM or FM.⁽¹²⁾ Visceral adipose tissue is suggested to be a negative predictor and muscle mass a positive predictor of microarchitecture in obese men.⁽¹⁴⁾

A causal association of LM on bone microarchitecture and geometry has been proposed.^(12,15) The mechanostatic

JBMR[®] Plus (WOA), Vol. 4, No. 9, September 2020, e10386.

DOI: 10.1002/jbm4.10386

© 2020 The Authors. JBMR Plus published by Wiley Periodicals LLC. on behalf of American Society for Bone and Mineral Research.

This is an open access article under the terms of the Creative Commons Attribution License, which permits use, distribution and reproduction in any medium, provided the original work is properly cited.

Received in original form January 27, 2020; revised form June 12, 2020; accepted June 24, 2020. Accepted manuscript online June 24, 2020.

Address correspondence to: Åshild Bjørnerem, MD, PhD, Department of Clinical Medicine, UiT The Arctic University of Norway, N-9037 Tromsø, Norway. E-mail: ashild.bjornerem@uit.no

hypothesis proposes that muscle contractions apply forces to bones that cause deformations or strains within the bone tissue.⁽¹⁵⁾ These forces are sensed by osteocytes, which increase bone formation through the bone-remodeling process, resulting in increased cortical area and thickness, and increased bone strength. It has been argued that the associations of bone size with both muscle size (forearm and lower leg cross-sectional area of muscle) and grip strength support the mechanostat hypothesis and a role of the muscle-bone unit.^(4,12) However, as the association of muscle size with bone structure is stronger than the association of muscle strength, other mechanisms could be involved, such as genetic, developmental, or hormonal factors.^(4,12) Dietary factors and physical activity have also been proposed to play a role in explaining the association between bone traits and muscle.^(1,7,12)

A causal association of FM with bone traits has been proposed through the action of estrogen.⁽¹²⁾ A beneficial association of estrogen on bone is well-established,⁽¹⁶⁻¹⁸⁾ as shown by bone loss after the drop in serum estrogen levels across the menopausal transition.⁽¹⁹⁻²⁰⁾ Estrogen produced by adipocytes is an important source of estrogen for postmenopasual women.⁽²¹⁾

Twin studies have predicted that 42%-92% of the variance in bone mass,^(1,22-23) 50% to 80% of the variance in bone microarchitecture,⁽²⁴⁾ and 52% to 84% of the variance in LM,^(2,22) and 65% of the variance in LM⁽²²⁾ are due to genetic factors. During the last decade, genome-wide association studies have identified single-nucleotide polymorphisms in many loci associated with BMD,⁽²⁵⁻²⁷⁾ cortical porosity and trabecular bone fraction,⁽²⁸⁾ LM⁽²⁹⁾ and BMI in adults,⁽³⁰⁾ Mendelian randomization uses genetic variants to infer whether risk factors have a causal influence on health outcomes under strong assumptions, and found evidence that adiposity-related traits have a causal effect on BMD at the heel for children.⁽³¹⁾ In addition, deletion of genome-wide association study-identified genes has been shown to result in increased cortical porosity and decreased bone strength of KO mice.⁽²⁷⁾ We had previously investigated if genetic factors explained the associations between LM and bone density using a twin study.⁽¹⁾ In that study we assumed that the LM measures did not have a causal effect on bone density.

To the best of our knowledge, it is not known whether the relationships of LM and FM with bone microarchitecture are causal, and/or due to genetic or environmental confounders. We therefore tested the hypotheses that: (i) LM is associated with cortical bone traits, (ii) FM is associated with trabecular bone traits, and (iii) these relationships of LM and FM with bone micro-architecture and other bone traits are causal. We did this applying a novel approach to the analysis of a twin study, Causation through Examination of FAmiliaL CONfounding (ICE FALCON), which allows inference on causation to be made from examination of familial confounding and changes in pairs of regression coefficients.⁽³²⁾

Materials and Methods

Study sample

This twin study included 324 female twin pairs, 199 monozygotic (MZ) and 125 dizygotic (DZ), aged 31 to 77 years at baseline and was conducted in Melbourne, Australia from 2008 to 2011.^(19,24,32-34) At follow-up in 2011 to 2013, participants had a total body scan for assessment of body composition. After excluding 39 women who had been treated with hormone replacement therapy or whose bone scans had movement artifacts, 388 women

had valid measurements of distal tibia at the follow-up visit. Of these, we excluded 24 women with missing total body scans and 60 for whom we did not have measurements for their cotwin. This left 152 complete pairs (54 DZ and 98 MZ) for the analysis of distal tibia traits. After a similar exclusion process, there were 124 complete pairs (45 DZ and 79 MZ) for the analysis of distal radius traits. All women had given written informed consent. The study was approved by the Austin Health Human Research Ethics Committee.

Bone microstructure and other measurements

Three-dimensional HRpQCT (isotropic resolution of 82 µm; XtremeCT; Scanco Medical AG, Brüttisellen, Switzerland) was used to obtain images at 60 kVp using 900 µA at the nondominant distal tibia and distal radius.^(35,36) The region of interest consisted of 110 CT slices obtained at 22.5 and 9.5 mm from a reference line at the endplate of the distal tibia and distal radius, respectively. The 49 most proximal slices were chosen because the relatively thicker cortex allows accurate assessment of porosity. Porosity within the total cortex and its compartments (compact cortex, outer and inner transitional zones [TZs]), matrix mineralization density, trabecular number, thickness, separation, and total, cortical, and trabecular cross-sectional area (CSA), total, cortical, and trabecular volumetric bone mineral density (vBMD) were quantified using StrAx software (Straxcorp, Melbourne, Australia), a nonthreshold-based method that automatically segments bone from background and into its compartments. The precision was 0.5% to 3.0%. (37,38) Daily quality control was carried out by scanning a phantom containing rods of hydroxyapatite (QRM, Moehrendorf, Germany). Cortical and trabecular microstructure was derived based on the photon attenuation by mineralized bone. Porosity is the proportion of voxels within the cortical compartment that contains void. Once deposited, osteoid is mineralized reaching \geq 80% of full mineralization (1,200 mg HA/cm³) within days. Matrix mineralization is quantified as the mean density of voxels with attenuation between 80% to 100% of fully mineralized bone. These voxels are unlikely to contain a pore because a pore results in voxel attenuation <80% of the maximum. So, variation in attenuation within 80% to 100% of full mineralization reflects heterogeneity in mineralization. Voxels with attenuating <80% are used to calculate porosity. Total body LM and FM were quantified using DXA (Lunar, Madison, WI, USA).

Statistical methods

Summary statistics were presented as mean and SD. Within-pair correlations were estimated for MZ and DZ pairs, and Fisher's *z*-transform was used to test for differences in correlations between these two groups. To test whether the MZ twins bone traits were similar to the DZ twins, we compared trait means between these two groups, adjusted for age and height, using the generalized estimating equation method, which takes into account correlation within twin pairs.

This method was also used to apply the Inference about Causation through Examination of FAmiliaL CONfounding (ICE FAL-CON) models so as to investigate whether there was evidence consistent with an association being due to causal effects or to familial confounding.^(32,39,40)

In brief, three models were fitted to the data in regression analysis using the generalized estimating equation method to allow for the outcomes being correlated within twin pairs. The first model estimated within individual-association of LM or FM with each bone trait, giving β_{self} . The second model estimated the cross-pair crosstrait association between LM or FM for one twin and a bone trait for the other twin, giving β_{cotwin} . The third model estimated the within-individual and cross-pair cross-trait associations concurrently, in effect adjusting each predictor for the other, giving β_{ceff}^{adj} and β_{cotwin}^{adj} . If a predictor has a causal effect on the outcome, the cross-pair cross-trait association will be attenuated towards zero ($\beta_{cotwin} > \beta_{cotwin}^{adj}$), but the within-individual association will be unchanged ($\beta_{self} = \beta_{self}^{adj}$). Hypothesis testing of the changes in regression coefficients from before and after adjustment, β_{self} $-\beta_{self}^{adj}$ and $\beta_{cotwin} - \beta_{cotwin}^{adj}$, was conducted by using the method proposed by Yan and colleagues,⁽⁴¹⁾ implemented in the R package "geepack,"⁽⁴²⁾ to estimate the standard errors of the changes. Because of small sample size for DZ twins (54 pairs), which does not have power to detect significant cross-pair cross-trait association, we therefore conducted ICE FALCON analyses for the combined MZ and DZ twin pairs. All models were adjusted for age and height. Outcome and predictor variables were all standardized to have a mean of zero and SD of 1; all p values were two-sided. Furthermore, we also conducted simulation studies to assess the performance of the changes in coefficients under causal effect and familial confounding (see Supplemental Information). The ICE FALCON and all other analyses were conducted using our own programs written in R language (R Foundation for Statistical Computing, Vienna, Austria; https://www.r-project.org/). Following convention, we have defined statistical significance as p < 0.05and not adjusted for multiple comparisons.

Results

The mean (SD) of age was 50.8 years (8.1). The MZ and DZ twins did not differ in mean age, LM, FM, and distal tibial bone traits, but MZ twins were shorter than DZ twins (161.9 versus 164.1 cm; p = 0.019; Table 1). Within-pair correlations, adjusted for age and height, were higher for MZ pairs [r(MZ)] than for DZ pairs [r(DZ)] for LM (0.79 versus 0.54; p = 0.003) and for FM (0.79 versus 0.40, p < 0.001).

Univariably, greater LM was associated with greater total bone CSA (standardized regression coefficient $\beta = 0.13$; p < 0.05), and both greater LM and greater FM were associated higher total BMC (β = 0.36; p < 0.01 and β = 0.317; p < 0.001, respectively), larger cortical CSA, thicker cortices and lower cortical porosity, and increased trabecular number of thinner trabeculae (absolute β ranges from 0.11 to 0.43; all *p*s < 0.05; see Table 2). When fitted together, greater LM was associated with larger total bone CSA, higher total BMC, larger cortical CSA, thicker cortices, lower porosity of the inner TZ, and increased trabecular number (absolute β ranges from 0.18 to 0.47; all *p*s < 0.05), but the strength of association for the latter was attenuated. Greater FM was no longer associated with the cortical bone traits after adjustment for LM, but remained associated with increased trabecular number and thinner trabeculae (absolute β ranges from 0.22 to 0.34; all ps < 0.001).

We conducted ICE FALCON analyses of the bone traits association with LM or FM based on the significant associations found from the analyses in Table 2 (see Table 3). The cross-pair crosstrait association of LM with the distal tibia total BMC and cortical

Table 1. Characteristics of and Comparison Between Dizygotic (DZ) and Monozygotic (MZ) Twins

	DZ (<i>n</i> = 108)		MZ (n	= 196))	
	Mean	SD	Mean	SD	p	
Age (years)	50.3	6.26	51.0	8.95	0.585	
Height (cm)	164.1	6.18	161.9	5.77	0.019	
Weight (kg)	71.0	15.0	69.1	14.9	0.413	
Total body lean mass (kg)	40.2	4.66	39.7	5.11	0.544	
Total body fat mass (kg)	27.4	11.2	26.6	11.1	0.646	
Distal tibia bone traits						
Total bone CSA (mm ²)	632	104	608	93.8	0.116	
Total vBMD (mg HA/cm ³)	307	59.6	311	52.8	0.596	
Total BMC (mg HA)	763	126	750	119	0.482	
Cortical CSA (mm ²)	210	20.9	207	21.0	0.432	
Cortical CSA/Total CSA	33.8	4.86	34.6	4.58	0.271	
Cortical thickness (mm)	2.40	0.25	2.43	0.25	0.516	
Total cortex porosity (%)	62.2	6.29	61.2	5.24	0.241	
Compact cortex porosity (%)	43.9	7.61	42.7	6.26	0.295	
Outer TZ porosity (%)	44.8	6.92	43.8	5.44	0.312	
Inner TZ porosity (%)	85.6	3.14	85.8	2.54	0.775	
Cortical vBMD (mg HA/cm ³)	640	80.5	653	66.4	0.253	
Cortical BMC (mg HA)	538	82.8	544	79.3	0.644	
Matrix mineralization density (%)	64.2	1.49	64.4	1.30	0.489	
Medullary CSA (mm ²)	422	95.3	401	84.5	0.122	
Medullary CSA/Total CSA	66.2	4.86	65.4	4.58	0.271	
Trabecular number (1/mm)	2.32	0.52	2.29	0.50	0.688	
Trabecular thickness (mm)	0.20	0.01	0.20	0.01	0.822	
Trabecular separation (mm)	1.44	0.28	1.50	0.28	0.195	
Trabecular vBMD (mg HA/cm ³)	133	37.8	128	36.2	0.345	
Trabecular BMC (mg HA)	225	83.0	206	76.0	0.114	

CSA = Cross-sectional area; HA = hydroxyapatite; TZ = transitional zone; vBMD = volumetric bone mineral density.

Table 2. Within-Individual Associations (Regression Coefficient b and Standard Error) of Lean Mass and F	at Mass (Predictors) With Distal
Tibia Bone Traits Adjusted for Age and Height (Outcomes)	

	Univariab	le models	Mutually adjusted models		
Distal tibia bone traits	Lean mass (kg) b \pm SE	Fat mass (kg) b \pm SE	Lean mass (kg) b \pm SE	Fat mass (kg) b \pm SE	
Total bone CSA (mm ²)	$\textbf{0.134} \pm \textbf{0.068}^{*}$	-0.002 ± 0.048	0.181 ± 0.069**	-0.071 ± 0.154	
Total vBMD (mg HA/cm ³)	$0.184 \pm 0.062^{**}$	$0.133 \pm 0.043^{**}$	$\textbf{0.127} \pm \textbf{0.071}$	$\textbf{0.086} \pm \textbf{0.050}$	
Total BMC (mg HA)	$0.364 \pm 0.069^{**}$	$0.167 \pm 0.043^{***}$	$0.332 \pm 0.074^{***}$	$\textbf{0.048} \pm \textbf{0.049}$	
Cortical CSA (mm ²)	$0.425 \pm 0.073^{***}$	$0.110 \pm 0.050 *$	$0.473 \pm 0.081^{***}$	-0.071 ± 0.061	
Cortical CSA/Total CSA	$0.131 \pm 0.064^{*}$	$\textbf{0.071} \pm \textbf{0.049}$	$\textbf{0.112} \pm \textbf{0.067}$	$\textbf{0.028} \pm \textbf{0.053}$	
Cortical thickness (%)	$0.322 \pm 0.074^{***}$	$0.112 \pm 0.054^{*}$	$0.331 \pm 0.078^{***}$	-0.014 ± 0.060	
Total cortex porosity (%)	$-0.146 \pm 0.059^{*}$	$-0.110 \pm 0.035^{**}$	-0.097 ± 0.066	-0.074 ± 0.040	
Compact cortex porosity (%)	-0.033 ± 0.058	-0.006 ± 0.039	-0.039 ± 0.061	$\textbf{0.009} \pm \textbf{0.040}$	
Outer TZ porosity (%)	-0.027 ± 0.055	$\textbf{0.011} \pm \textbf{0.029}$	-0.045 ± 0.064	$\textbf{0.028} \pm \textbf{0.034}$	
Inner TZ porosity (%)	$-0.215 \pm 0.072^{**}$	-0.098 ± 0.051	$-0.200 \pm 0.086^{*}$	-0.022 ± 0.063	
Cortical vBMD (mg HA/cm ³)	$0.145 \pm 0.059^{*}$	$0.112 \pm 0.035^{**}$	$\textbf{0.094} \pm \textbf{0.067}$	$\textbf{0.077} \pm \textbf{0.040}$	
Cortical BMC (mg HA)	$0.398 \pm 0.069^{***}$	$0.161 \pm 0.044^{***}$	$0.387 \pm 0.074^{***}$	$\textbf{0.017} \pm \textbf{0.051}$	
Matrix mineralization density (%)	$\textbf{0.040} \pm \textbf{0.057}$	$\textbf{0.004} \pm \textbf{0.035}$	$\textbf{0.050} \pm \textbf{0.066}$	-0.015 ± 0.041	
Medullary CSA (mm ²)	$\textbf{0.050} \pm \textbf{0.068}$	-0.027 ± 0.049	$\textbf{0.091} \pm \textbf{0.068}$	-0.061 ± 0.050	
Medullary CSA/Total CSA	$-0.131 \pm 0.064*$	-0.071 ± 0.049	-0.112 ± 0.067	-0.028 ± 0.053	
Trabecular number (1/mm)	$0.390 \pm 0.067^{***}$	$0.401 \pm 0.055^{***}$	$0.177 \pm 0.066^{**}$	$0.335 \pm 0.060^{***}$	
Trabecular thickness (mm)	$-0.243 \pm 0.071^{***}$	$-0.272 \pm 0.047^{***}$	-0.082 ± 0.078	$-0.240 \pm 0.054^{***}$	
Trabecular separation (mm)	$-0.247 \pm 0.062^{***}$	$-0.262 \pm 0.050^{***}$	-0.103 ± 0.070	$-0.224 \pm 0.056^{***}$	
Trabecular vBMD (mg HA/cm ³)	$\textbf{0.149} \pm \textbf{0.070}^{*}$	$0.121 \pm 0.050^{*}$	$\textbf{0.093} \pm \textbf{0.082}$	$\textbf{0.087} \pm \textbf{0.058}$	
Trabecular BMC (mg HA)	$\textbf{0.134} \pm \textbf{0.068}$	$\textbf{0.087} \pm \textbf{0.050}$	$\textbf{0.102} \pm \textbf{0.071}$	$\textbf{0.050} \pm \textbf{0.051}$	

CSA = cross-sectional area; HA = hydroxyapatite; TZ = transitional zone; vBMD = volumetric bone mineral density. Outcome variables and predictors were standardised to have mean zero and standard deviation of 1.

****p* < 0.001.

BMC, porosity of the inner TZ, and trabecular number remained significant after adjustment for the respective within-individual association. The cross-pair cross-trait association of FM with distal tibia trabecular number, separation, and vBMD remained after

adjustment for the respective within-individual association. None of the changes in the cross-pair cross-trait associations, i.e., $\beta_{cotwin} - \beta_{cotwin}^{adj}$, was significant (*p* ranging from 0.149 to 0.998), except marginally for cortical CSA with LM (*p* = 0.062),

	Univa cross-1	Univariable cross-pair cross-trait association		Cross-pair cross-trait association allowing for within-individual association			Absolute cl cross-pair c associa	Absolute change in cross-pair cross-trait association	
Distal tibia bone traits (outcomes)	β_{cotwin}	SE	р	$\beta_{\text{cotwin}}^{\text{adj}}$	SE	р	Change	p	
Lean mass (predictor)									
Total bone CSA (mm ²)	0.028	0.053	0.609	0.018	0.056	0.741	-0.009	0.494	
Total BMC (mg HA)	0.183	0.055	0.001	0.143	0.050	0.004	-0.040	0.220	
Cortical CSA (mm ²)	0.150	0.068	0.028	0.072	0.059	0.223	-0.079	0.062	
Cortical thickness (mm)	0.113	0.070	0.110	0.078	0.061	0.197	-0.034	0.249	
Inner TZ porosity (%)	-0.150	0.059	0.011	-0.122	0.058	0.035	0.028	0.169	
Cortical BMC (mg HA)	0.144	0.061	0.018	0.108	0.049	0.028	-0.036	0.312	
Trabecular number (1/mm)	0.159	0.057	0.006	0.112	0.050	0.024	-0.047	0.149	
Fat mass (predictor)									
Trabecular number (1/mm)	0.162	0.052	0.002	0.162	0.041	0.001	0.0001	0.998	
Trabecular separation (mm)	-0.133	0.045	0.003	-0.150	0.044	0.001	-0.017	0.533	
Trabecular vBMD (mg HA/cm ³)	0.137	0.042	0.001	0.152	0.045	0.001	0.015	0.298	

Table 3. ICE FALCON Analyses for the Associations (Regression Coefficients b and Standard Error) of Distal Tibia Bone Traits Adjusted for

 Age and Height (Outcomes) With Lean Mass and Fat Mass (Predictors)

Outcome and predictor variables were standardized to have mean of zero and SD of 1, p were for two-sided.

CSA = Cross-sectional area; HA = hydroxyapatite; ICE FALCON = Inference about Causation through Examination of FAmiliaL CONfounding; TZ = transitional zone; vBMD = volumetric bone mineral density.

^{*}p < 0.05.

^{**}*p* < 0.01.

as well as the changes in the within-individual association, $\beta_{self} - \beta_{self}^{adj}$ (all ps > 0.100, results not shown). The results in Table 3 suggest that the associations between bone trait and LM and FM were confounded by familial factors, and this was confirmed by simulation results (Supplementary Table S1), where estimated biases and mean square errors were smaller for the model simulated under familial confounding than for the model in which LM or FM causes the bone trait (again except for cortical CSA). For the distal radius, the associations were weaker, but they followed similar patterns as for the distal tibia (Supplementary Table S2), and because there was no significant cross-pair cross-trait association, no further analysis was performed (Supplementary Table S3).

Discussion

For adult women, we found that within an individual the cortical bone traits were mainly associated with LM, not FM, and the trabecular bone traits were mainly associated with FM. For each of these body compositions-bone trait associations-we found that the body composition trait of a twin was associated with the bone trait of the cotwin. But these cross-pair cross-trait associations did not change after adjustment for the individual's body composition. Following the logic of the ICE FALCON approach, we have found no evidence that the body composition measure of an individual had a causal effect on their bone trait, except for cortical CSA and LM. Given the cross-pair crosstrait associations with LM and FM were significant after adjustment for the respective within-individual associations with LM and FM, the cross-pair cross-trait associations must be attributable, at least in part, to familial factors shared by twins in the same pair, and these could include both genetic and environmental factors.

We confirmed that greater LM was associated with a larger bone size and cortical area, as well as thicker cortices independent of FM.⁽¹²⁾ In addition, greater LM was associated with a lower cortical porosity of the inner transitional zone-a novel finding. Measurement of porosity in the transitional zone was made possible by an accurate segmentation of bone using StrAx software. As previously reported, bone loss due to unbalanced remodeling upon intracortical canal surfaces starts and is more pronounced in the portion of the cortex adjacent to the marrow space, which corresponds to the inner transitional zone.⁽⁴³⁾ This might explain why the greater LM was associated with lower cortical porosity of the inner transitional zone only. We also confirmed that greater FM was associated with an increased trabecular number,⁽¹²⁾ but we also found that greater LM was associated with that outcome even after adjusting for FM. Having a larger sample size of 304 women in the current study versus 167 women in the previous study⁽¹²⁾ could be one reason why we found that both FM and LM are associated with a higher trabecular number.

The developmental origins hypothesis proposes that the body composition and bone traits may be correlated because both are associated with early life environment.⁽¹²⁾ However, twin studies have found that the variances in bone mass, LM, and FM appear to be largely determined by genetic factors.^(1,2,2,2,3) The heritabilities for the cortical and trabecular traits were estimated to be from 67% to 88% based on the twin baseline data, under the assumptions of the classic twin model.⁽²⁴⁾ Here we have found evidence consistent with those high heritabilities for LM and FM under the same assumptions, in keeping with previous

reports.^(2,22) Whether the relationships of bone traits with LM or FM also are determined solely by genetic factors is not known.⁽¹⁾

This study has several limitations. We cannot exclude causal roles of LM and FM as they might not be detectable with our sample size. Larger sample sizes are needed to examine the hypotheses with more power and to provide more precision on the amount of cross-pair cross-trait correlations that could be explained by a causal relationship of bone geometry and microarchitecture with LM and FM. Our approach is also novel, so replication studies using a twin or sister-pair design for the prospective evaluation of the relationships of changes in bone traits with LM and FM will be important in trying to validate our findings and approach. The clinical utility is to not delude adult women of this age into thinking there is evidence that changing their FM or LM will have a causal effect on the bone measures.

In conclusion, noninvasive assessment of bone morphology is feasible and increasingly available. Use of this methodology to study the skeletons of twin pairs permits insights into the pathophysiology of bone fragility. Here, we report that greater LM was associated with larger bone size and improved cortical and trabecular microarchitecture, and FM was associated with improved trabecular microarchitecture, but these associations were not causal for adult women, or not strong enough to be detectable by this study. This issue needs to be addressed separately for younger women, including children with growing bones, because any evidence for causation would have important implications for prevention of fracture. Given the familial nature of the cross-pair cross-trait associations, there must be familial factors that predispose to both the body composition and the bone traits. These familial factors could have genetic or environmental origins, or both.

Disclosures

RZ is director and shareholder of StrAx Corp Pty Ltd and has received research grants, served on the advisory boards, and/or received honoraria from Amgen, MSD, Servier, and Sanofi Genzyme. All authors state that they have no other conflicts of interest.

Acknowledgments

This research was facilitated through access to the Australian Twin Registry, a national resource supported by an Enabling Grant (ID 628911) from the National Health & Medical Research Council (NHMRC) of Australia. This study was funded by NHMRC (Project Grant ID: 1004938) and by the Research Council of Norway (RCN) Grant (ID 178588/V50). JLH is a NHMRC Senior Principal Research Fellow. We thank research staff at the University of Melbourne who managed the data and the images.

Authors' roles: Study design and conduct: MB, JLH, ÅB. Data collection: ÅB. StrAx analyses: RZ. Statistical analyses: MB, SL. Drafting manuscript, data interpretation, and approving final version of manuscript: all authors. MB takes responsibility for the integrity of the data analysis.

Peer Review

The peer review history for this article is available at https://publons.com/publon/10.1002/jbm4.10386.

References

- Seeman E, Hopper JL, Young NR, Formica C, Goss P, Tsalamandris C. Do genetic factors explain associations between muscle strength, lean mass, and bone density? A twin study. Am J Physiol. 1996;270: E320–7.
- Arden NK, Spector TD. Genetic influences on muscle strength, lean body mass, and bone mineral density: a twin study. J Bone Miner Res. 1997;12:2076–81.
- Park JH, Song YM, Sung J, et al. The association between fat and lean mass and bone mineral density: the Healthy Twin Study. Bone. 2012; 50:1006–11.
- Edwards MH, Gregson CL, Patel HP, et al. Muscle size, strength, and physical performance and their associations with bone structure in the Hertfordshire Cohort Study. J Bone Miner Res. 2013;28:2295–304.
- Leslie WD, Orwoll ES, Nielson CM, et al. Estimated lean mass and fat mass differentially affect femoral bone density and strength index but are not FRAX independent risk factors for fracture. J Bone Miner Res. 2014;29:2511–9.
- Scott D, Hayes A, Sanders KM, Aitken D, Ebeling PR, Jones G. Operational definitions of sarcopenia and their associations with 5-year changes in falls risk in community-dwelling middle-aged and older adults. Osteoporos Int. 2014;25:187–93.
- 7. Scott D, Daly RM, Sanders KM, Ebeling PR. Fall and fracture risk in sarcopenia and dynapenia with and without obesity: the role of lifestyle interventions. Curr Osteoporos Rep. 2015;13:235–44.
- Wagner P, Chapurlat R, Ecochard R, Szulc P. Low muscle strength and mass is associated with the accelerated decline of bone microarchitecture at the distal radius in older men: the prospective STRAMBO study. J Bone Miner Res. 2018;33:1630–40.
- 9. De Laet C, Kanis JA, Oden A, et al. Body mass index as a predictor of fracture risk: a meta-analysis. Osteoporos Int. 2005;16:1330–8.
- Johansson H, Kanis JA, Oden A, et al. A meta-analysis of the association of fracture risk and body mass index in women. J Bone Miner Res. 2014;29:223–33.
- 11. Compston JE, Flahive J, Hosmer DW, et al. Relationship of weight, height, and body mass index with fracture risk at different sites in postmenopausal women: the Global Longitudinal study of Osteoporosis in Women (GLOW). J Bone Miner Res. 2014;29:487–93.
- Edwards MH, Ward KA, Ntani G, et al. Lean mass and fat mass have differing associations with bone microarchitecture assessed by high resolution peripheral quantitative computed tomography in men and women from the Hertfordshire Cohort Study. Bone. 2015;81:145–51.
- Sornay-Rendu E, Boutroy S, Vilayphiou N, Claustrat B, Chapurlat RD. In obese postmenopausal women, bone microarchitecture and strength are not commensurate to greater body weight: the Os des Femmes de Lyon (OFELY) study. J Bone Miner Res. 2013;28:1679–87.
- 14. Bredella MA, Lin E, Gerweck AV, et al. Determinants of bone microarchitecture and mechanical properties in obese men. J Clin Endocrinol Metab. 2012;97:4115–22.
- 15. Frost HM. Bone's mechanostat: a 2003 update. Anat Rec A Discov Mol Cell Evol Biol. 2003;275:1081–101.
- 16. Compston JE. Sex steroids and bone. Physiol Rev. 2001;81:419-47.
- Bjørnerem Å, Emaus N, Berntsen GKR, et al. Circulating estradiol and sex hormone-binding globulin predict bone loss in women and men. The Tromsø study. Calcif Tissue Int. 2007;81:65–72.
- Bjørnerem Å, Ahmed LA, Berntsen GKR, et al. A prospective study of sex steroids, sex hormone-binding globulin and non-vertebral fractures in women and men in the Tromsø study. Eur J Endocrinol. 2007;157:119–25.
- Bjørnerem Å, Wang X, Bui M, et al. Menopause-related appendicular bone loss is mainly cortical and results in increased cortical porosity. J Bone Miner Res. 2018;33:598–605.
- Bjørnerem Å, Straume B, Midtby M, et al. Endogenous sex hormones in relation to age, sex, lifestyle factors, and chronic diseases in a general population: the Tromsø study. J Clin Endocrinol Metab. 2004;89: 6039–47.
- 21. Simpson ER. Sources of estrogen and their importance. J Steroid Biochem Mol Biol. 2003;86:225–30.

- Nguyen TV, Howard GM, Kelly PJ, Eisman JA. Bone mass, lean mass, and fat mass: same genes or same environments? Am J Epidemiol. 1998;147:3–16.
- Pocock NA, Eisman JA, Hopper JL, Yeates MG, Sambrook PN, Eberl S. Genetic determinants of bone mass in adults. A twin study. J Clin Invest. 1987;80:706–10.
- 24. Bjørnerem Å, Bui M, Wang X, et al. Genetic and environmental variances of bone microarchitecture and bone remodeling markers: a twin study. J Bone Miner Res. 2015;30:519–27.
- 25. Estrada K, Styrkarsdottir U, Evangelou E, et al. Genome-wide metaanalysis identifies 56 bone mineral density loci and reveals 14 loci associated with risk of fracture. Nat Genet. 2012;44:491–501.
- Kemp JP, Morris JA, Medina-Gomez C, et al. Identification of 153 new loci associated with heel bone mineral density and functional involvement of GPC6 in osteoporosis. Nat Genet. 2017;49:1468–75.
- 27. Morris JA, Kemp JP, Youlten SE, et al. An atlas of genetic influences on osteoporosis in humans and mice. Nat Genet. 2019;51:258–66.
- 28. Paternoster L, Lorentzon M, Lehtimäki T, et al. Genetic determinants of trabecular and cortical volumetric bone mineral densities and bone microstructure. PLoS Genet. 2013;9:e1003247.
- 29. Karasik D, Zillikens MC, Hsu YH, et al. Disentangling the genetics of lean mass. Am J Clin Nutr. 2019;109:276–87.
- 30. Speliotes EK, Willer CJ, Berndt SI, et al. Association analysis of 249,796 individuals reveal 18 new loci associated with body mass index. Nat Genet. 2010;42:937–48.
- 31. Kemp JP, Sayers A, Smith GD, Tobias JH, Evans DM. Using Mendelian randomization to investigate a possible causal relationship between adiposity and increased bone mineral density at different skeletal sites in children. Int J Epidemiol. 2016;45:1560–72.
- Bui M, Bjørnerem Å, Ghasem-Zadeh A, Dite GS, Hopper JL, Seeman E. Architecture of cortical bone determines in part its remodelling and structural decay. Bone. 2013;55:353–8.
- 33. Bjørnerem Å, Ghasem-Zadeh A, Bui M, et al. Remodeling markers are associated with larger intracortical surface area but smaller trabecular surface area: a twin study. Bone. 2011;49:1125–30.
- Bjørnerem Å, Bui QM, Ghasem-Zadeh A, Hopper JL, Zebaze R, Seeman E. Fracture risk and height: an association partly accounted for by cortical porosity of relatively thinner cortices. J Bone Miner Res. 2013;28:2017–26.
- 35. Laib A, Hauselmann HJ, Ruegsegger P. In vivo high resolution 3D-QCT of the human forearm. Technol Health Care. 1998;6:329–37.
- Boutroy S, Bouxsein ML, Munoz F, Delmas PD. In vivo assessment of trabecular bone microarchitecture by high-resolution peripheral quantitative computed tomography. J Clin Endocrinol Metab. 2005; 90:6508–15.
- Zebaze R, Ghasem-Zadeh A, Mbala A, Seeman E. A new method of segmentation of compact-appearing, transitional and trabecular compartments and quantification of cortical porosity from high resolution peripheral quantitative computed tomographic images. Bone. 2013;54:8–20.
- Chiang CY, Zebaze R, Wang XF, Ghasem-Zadeh A, Zajac JD, Seeman E. Cortical matrix mineral density measured noninvasively in pre- and postmenopausal women and a woman with vitamin D– dependent rickets. J Bone Miner Res. 2018;33:1312–7.
- Li S, Wong EM, Bui M, Nguyen TL^{1,} Joo J-HE, Stone J, et al. Inference about causation between body mass index and DNA methylation in blood from a twin family study. Int J Obes (Lond). 2019;43:243–52.
- 40. Li S, Bui M, Hopper JL. Inference about causation from examination of familial confounding (ICE FALCON): a model for assessing causation analogous to Mendelian randomization. Int J Epidemiol. 2020 Jun 4;dyaa065. doi:10.1093/ije/dyaa065. Online ahead of print.
- 41. Yan J, Aseltine R, Harel O. Comparing regression coefficients between nested linear models for clustered data with generalized estimating equations. J Educ Behav Stat. 2013;38:172–89.
- 42. Hojsgaard S, Halekoh U, Yan J. The R package geepack for generalized estimating equations. J Stat Softw. 2006;15(2):1–11.
- 43. Zebaze RM, Ghasem-Zadeh A, Bohte A, et al. Intracortical remodelling and porosity in the distal radius and post-mortem femurs of women: a cross-sectional study. Lancet. 2010;375:1729–36.