

The effect of anaesthetic dose on response and remission in electroconvulsive therapy for major depressive disorder: nationwide register-based cohort study

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Background

Electroconvulsive therapy (ECT) is a safe and effective treatment for major depressive disorder (MDD). ECT treatment effect relies on induced generalised seizures. Most anaesthetics raise the seizure threshold and shorten seizure duration. There are no conclusive studies on the effect of anaesthetic dose on response and remission rates with ECT for MDD.

Aims

We aimed to examine the effect of different dose intervals of anaesthetics on response and remission after ECT for MDD.

Method

We conducted a nationwide cohort study, using data from Swedish registers. Low-, medium- and high-dose intervals, adjusted for age and gender, were constructed for each anaesthetic drug. Response and remission were measured with the Clinical Global Impression – Severity and Improvement scales (CGI-I and CGI-S), and a self-rated version of the Montgomery–Åsberg Depression Rating Scale (MADRS-S). Logistic regression models were used to calculate adjusted odds ratios for response and remission rates.

Results

The study included 7917 patients who received ECT for MDD during 2012–2018. Patients were given either thiopental (64.1%) or propofol (35.9%). Low-dose intervals of anaesthetics were

associated with increased rates of response (CGI-I: odds ratio 1.22, 95% CI 1.07–1.40, $P = 0.004$; MADRS-S: odds ratio 1.31, 95% CI 1.09–1.56, $P = 0.004$) and remission (CGI-S: odds ratio 1.37, 95% CI 1.17–1.60, $P \leq 0.001$; MADRS-S: odds ratio 1.31, 95% CI 1.10–1.54, $P = 0.002$).

Conclusions

We found improved treatment outcomes with low- compared with high-dose anaesthetic during ECT for MDD. To enhance treatment effect, deep anaesthesia during ECT for MDD should be avoided.

Keywords

Depressive disorders; electroconvulsive therapy; epidemiology; antidepressants; outcome studies.

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Electroconvulsive therapy (ECT) is a safe and effective treatment for major depressive disorder (MDD).^{1,2} Early studies on the mechanisms of action of ECT demonstrated that the improvement in depression primarily was accomplished with the seizure and not the electrical stimulus itself.³ Thus, inducing a generalised seizure is crucial to achieve adequate MDD treatment results with ECT.⁴ In its initial form, ECT was used without anaesthesia and muscle relaxation, which today is referred to as ‘unmodified ECT’. However, because of the risk of injury such as vertebrae fractures and joint dislocations during treatment, muscle relaxants and anaesthesia were introduced to the treatment, termed ‘modified ECT’.⁵ Modified ECT is currently the recommended practice, but unmodified ECT is still used in some parts of the world.⁶ The seizure threshold elevating effect of several anaesthetic agents is well known. Thiopental and propofol are commonly used to treat status epilepticus.⁷ Other anaesthetic agents, such as ketamine and etomidate, have less effect on seizure quality during ECT.⁸ The effects of thiopental and propofol on seizure threshold seem to be dose-dependent.⁹ Several studies have examined the effects of different anaesthetic agents and dose on seizure threshold and seizure lengths.^{10–13} It has been suggested that a light anaesthesia is of importance to increase remission rates.¹⁴ However, until now, no report has addressed the relationship

between anaesthetic dose during ECT and MDD treatment outcomes.¹⁵ Recent studies suggest that light anaesthesia during surgical procedures does not come with increased short- or long-term risks.¹⁶ The purpose of this study is to examine the effect of different dose intervals of anaesthetic agents on response and remission after ECT for MDD. We hypothesise that low doses of anaesthetic agents are associated with increased likelihood of MDD treatment effect.

Method

Study design

This was a nationwide observational study. Data on ECT procedures and outcomes were gathered from the Swedish National Quality Register for ECT (Q-ECT). Data for this register is gathered prospectively. The Q-ECT had an inclusion rate of 90% during 2018.¹⁷ Data on comorbidities and pharmacotherapy were gathered from the Swedish National Patient Register and the Swedish Prescribed Drug Register. This study was approved by the regional ethical review board in Uppsala, Sweden (Ethical permit 2014/174). No informed consent was required for this register-based study using anonymised data.

Study population

We included patients aged >18 years with MDD, who received treatment during 2012–2018 and were included in the Q-ECT. Only index series of ECT were included; maintenance and booster sessions were excluded. Only the first registered ECT series for each patient was included in the study. An additional inclusion criterion was complete data on anaesthetic medication. With our initial inclusion criteria, we found 10 815 patients. Of these, 2500 patients were excluded because of incomplete data on anaesthetics. The vast majority of patients were anaesthetised with either thiopental or propofol. Because of their infrequent use in this material, we excluded 42 patients with ketamine listed as their primary anaesthetic agent and 46 patients with remifentanyl listed as their primary anaesthetic agent; 310 patients who received a combination of at least two anaesthetic medications were also excluded. After these exclusions, 7917 patients remained. Among these patients, 5078 received treatment with thiopental and 2839 received treatment with propofol. After calculating age- and gender-adjusted dose intervals there were 3455 patients in the low-dose interval group, 2430 patients in the medium-dose interval and 2032 patients in the high-dose interval (Fig. 1).

Exposure

The exposure of this study was anaesthetic dose given during ECT, and the dose was defined as a low, medium and high interval. The anaesthetic dose for each drug was registered in the Q-ECT for the last treatment in the series. If several anaesthetic agents were used, then all of them were reported and listed as primary, secondary, third, etc. The continuous dose variable was adjusted for age and gender. For each drug, we calculated dose intervals for men and women separately, for the following age categories: 18–34, 35–54, 55–74 and ≥ 75 years. The low-, medium- and high-dose intervals for each of these categories were calculated with the 33rd and 67th percentiles. Doses up to the 33rd percentile were included in the low-dose interval, doses up to the 67th percentile were included in the medium-dose interval and doses above the 67th percentile were included in the high-dose interval. As an example, a 76-year-old woman receiving 275 mg of thiopental would be considered to be in the high-dose interval, whereas a 30-year-old man receiving the same dose would be considered to be in the low-dose interval. All dose intervals for each category can be seen in Supplementary Appendix 1 available at <https://doi.org/10.1192/bjo.2021.31>.

Primary outcome and measurements

The primary outcome of our study was response to treatment based on the Clinical Global Impression – Improvement scale (CGI-I), and was defined as a score of 1 or 2.¹⁸ The scale comprises the following seven ratings: 1, ‘Very much improved’; 2, ‘Much improved’; 3, ‘Minimally improved’; 4, ‘No change’; 5, ‘Minimally worse’; 6, ‘Much worse’ and 7, ‘Very much worse’. These measurements are recorded by the evaluating doctor in the ECT evaluation form and then reported to the Q-ECT. Evaluation is routinely done the day after the last ECT treatment.

In addition to our primary outcome, we also included the following secondary outcomes: distinct response and remission according to the CGI, response and remission according to the Montgomery–Åsberg Depression Rating Scale (MADRS-S), and subjective memory worsening. Distinct response was defined as a score of 1 with the CGI-I. Remission based on clinical rating was defined as a score of 1 on the Clinical Global Impression – Severity scale (CGI-S).¹⁸ The CGI-S is a seven-point rating scale that rates severity of the patient’s illness as follows: 1, ‘Normal, not at all ill’; 2, ‘Borderline mentally ill’; 3, ‘Mildly ill’; 4,

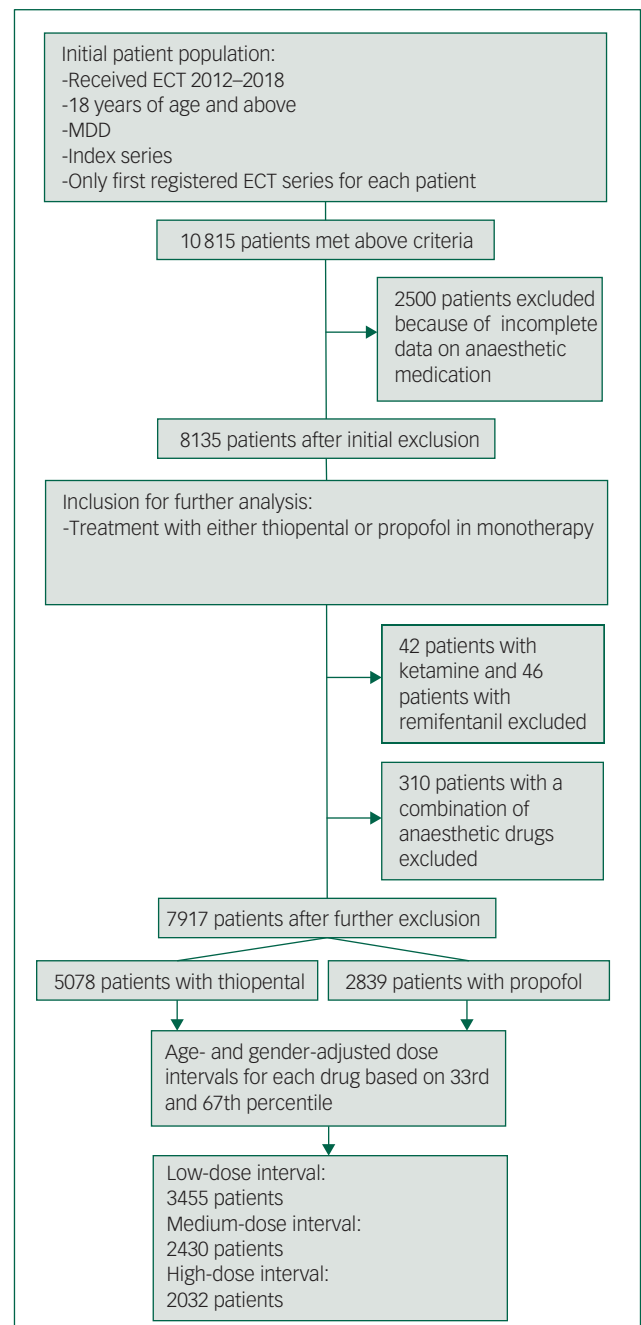


Fig. 1 Flow chart of study. Illustrates each step of inclusion and exclusion for the study population and the size of the study population for each step. ECT, electroconvulsive therapy; MDD, major depressive disorder.

‘Moderately ill’; 5, ‘Markedly ill’; 6, ‘Severely ill’ and 7, ‘Among the most extremely ill patients’.¹⁹ The MADRS-S was used for self-rating depressive symptoms, which measures nine different items for depression, each rating 0–6 points.²⁰ Remission based on MADRS-S was defined as a score of <10 and response as $\geq 50\%$ reduction on the MADRS-S scale.²¹

Subjective memory complaints were measured with a seven-point scale based on the ‘failing memory’ item of the Comprehensive Psychopathological Rating Scale (CPRS-M).²² The scale ranges from 0 (‘Memory as usual’) to 6 (‘Complaints of complete inability to remember’). Subjective memory worsening was defined as a two-point increase on this scale after ECT.²³

Comorbidities and pharmacotherapy

Data on the following categories of somatic and psychiatric comorbidities were gathered from the Swedish National Patient Register: diabetes mellitus type 2, heart disease, chronic obstructive pulmonary disease, kidney failure, anxiety, substance misuse and personality disorder. A patient was considered to have a comorbidity based on ICD-10 codes (see Supplementary Appendix 2 for all codes for each category). A comorbidity was considered to be present if the code was registered within 3 years before the first date of ECT. Considering that ICD-10 codes are recorded in the register at the end of a hospital stay, ICD-10 codes registered within a week after concluded treatment or at discharge after the current ECT series were also included.

Data on pharmacotherapy was gathered from the Swedish Prescribed Drug Register. Data is recorded based on the Anatomical Therapeutic Chemical Classification System (ATC). Data on the following prescription categories was gathered based on ATC codes (see Supplementary Appendix 3): anti-epileptics, antipsychotics, antidepressants, benzodiazepines and lithium. A patient was considered to be on treatment if the patient had filled a prescription within 100 days before the date of the first ECT in the series.

Statistical analysis

In addition to using age- and gender-adjusted dose interval, we applied crude and adjusted logistic regression models, yielding odds ratios and 95% confidence intervals, to associate anaesthetic dose intervals with our outcomes. We adjusted for the following potential confounders in our final regression models: age, gender, number of sessions in the series, psychiatric comorbidity (anxiety disorders, substance misuse and personality disorder) and psychiatric pharmacotherapies (antidepressants, anti-epileptics, antipsychotics, benzodiazepines and lithium). The choice of covariates in our model was based on a variable being a known confounder, such as age,^{24–27} or a potential confounder based on clinical reasoning.

Initially, we included choice of anaesthetic drug in our logistic regression model to rule out potential effects of this variable. There was no statistically significant difference between CGI-I score for thiopental and propofol (odds ratio 1.06, 95% CI 0.95–1.19, $P = 0.298$). We also included an interaction term for dose interval and choice of anaesthetic drug, which showed no statistical significance. With our calculations of age- and gender-adjusted dose interval combined with the above results, choice of drug was left out of further analysis owing to it being an unlikely confounder between dose interval and our outcomes.

All analyses were conducted with SPSS version 26.0 for Windows 10 (IBM Corp, Armonk, USA).

Missing data

In our initial material, 23.1% of patients had missing data on the exposure. The patients with missing data on anaesthetic agent and dose were, on average, older than those with complete data (55.9 v. 53.1 years, $P = 0.001$). There was a larger proportion of women in the group with missing data compared with the group with available data on the exposure (61.0% v. 57.8%, $P = 0.006$). For our covariates (age, gender, psychiatric pharmacotherapy, psychiatric comorbidity, number of sessions) used in our logistic regression model, the data was fully available for all patients in our study. For the outcome, a total of 7211 (91%) patients had available data on CGI-I score and 6800 (86%) had data on CGI-S score. For self-rated scales, MADRS-S score before treatment was available for 5252 patients (66%), and after treatment for 4042 patients (51%). Data for CPRS-M to calculate subjective memory worsening was available for 4497 patients (57%). There was no significant

difference in terms of missing data regarding our primary outcome (CGI-I) for age and gender, or for the CGI-S and CPRS-M. However, patients with missing data on the self-rated scale, MADRS-S, were on average younger than patients with available data on the self-rated scale (52.5 v. 53.7 years, $P = 0.019$).

Results

Sociodemographic, clinical and treatment characteristics for each of the dose-interval groups can be seen in Table 1. On average, patients who received high-dose intervals required more sessions in a series compared with patients who received low-dose intervals (7.5, 7.8 and 8.1 sessions, on average, for low-, medium- and high-dose intervals, respectively; $P \leq 0.001$). Patients who received low-dose intervals had longer seizures, on average, than patients who received high-dose intervals (low-dose interval, 38.7 s; medium-dose interval, 37.1 s; high-dose interval, 36.0 s; $P \leq 0.001$) (Table 2). For our primary outcome, response measured with CGI-I, we found that low-dose intervals were associated with increased chance of response compared with high-dose intervals (odds ratio 1.22, 95% CI 1.07–1.40, $P = 0.004$) (Table 3). For our secondary outcomes, we found that low-dose intervals were associated with increased likelihood of response (based on MADRS-S score reduction) and remission (based on CGI-S and MADRS-S scores) compared with high-dose intervals (response based on MADRS-S score: odds ratio 1.31, 95% CI 1.09–1.56, $P = 0.004$; remission based on CGI-S score: odds ratio 1.37, 95% CI 1.17–1.60, $P \leq 0.001$; remission based on MADRS-S score: odds ratio 1.31, 95% CI 1.10–1.54, $P = 0.002$). For distinct response (score of 1 on CGI-I), we found that both low- and medium-dose intervals were associated with increased chance of distinct response compared with high-dose interval (low-dose interval: odds ratio 1.51, 95% CI 1.31–1.73,

Table 1 Sociodemographic, clinical and treatment characteristics for each of the dose interval groups

	Low-dose interval, $n = 3455$	Medium-dose interval, $n = 2430$	High-dose interval, $n = 2032$	
	Mean (s.d.) or n (%)	Mean (s.d.) or n (%)	Mean (s.d.) or n (%)	P -value
Age, years	53.7 (19.3)	53.7 (18.4)	51.5 (18.1)	<0.001
Gender				
Female	2026 (58.6%)	1511 (62.2%)	1054 (51.9%)	<0.001
Male	1429 (41.4%)	919 (37.8%)	978 (48.1%)	
Anxiety disorder	859 (24.9%)	619 (25.5%)	569 (28.0%)	0.033
Personality Disorder	223 (6.5%)	170 (7.0%)	144 (7.1%)	0.588
Substance misuse	320 (9.3%)	274 (11.3%)	282 (13.9%)	<0.001
COPD	50 (1.4%)	28 (1.2%)	22 (1.1%)	0.426
Diabetes type 2	129 (3.7%)	107 (4.4%)	72 (3.5%)	0.274
Heart disease	99 (2.9%)	65 (2.7%)	39 (1.9%)	0.093
Kidney failure	26 (0.8%)	14 (0.6%)	8 (0.4%)	0.248
Antidepressants	2665 (77.1%)	1973 (81.2%)	1675 (82.4%)	<0.001
Anti-epileptics	231 (6.7%)	223 (9.2%)	197 (9.7%)	<0.001
Antipsychotics	1174 (34.0%)	887 (36.5%)	763 (37.5%)	0.017
Benzodiazepines	1513 (43.8%)	1135 (46.7%)	964 (47.4%)	0.014
Lithium	173 (5.0%)	133 (5.5%)	132 (6.5%)	0.066
Severity				
CGI-S before	5.3 (0.8)	5.2 (0.9)	5.0 (0.9)	<0.001
MADRS-S before	33.7 (9.3)	34 (9.0)	33.7 (8.6)	0.447

P -values were calculated using one-way ANOVA for mean and Pearson chi-squared for proportions. COPD, chronic obstructive pulmonary disease; CGI-S, Clinical Global Impression – Severity Scale, MADRS-S, Montgomery–Åsberg Depression Rating Scale, self-rated version.

Table 2 ECT treatment characteristics

	Low-dose interval, <i>n</i> = 3455	Medium-dose interval, <i>n</i> = 2430	High-dose interval, <i>n</i> = 2032	<i>P</i> -value
	Mean (s.d.) or <i>n</i> (%)	Mean (s.d.) or <i>n</i> (%)	Mean (s.d.) or <i>n</i> (%)	
Number of sessions in series	7.5 (3.5)	7.8 (3.4)	8.1 (3.3)	<0.001
Charge (mC)	401.7 (186.8)	407.8 (186.6)	407.7 (181.1)	0.353
Seizure duration(s)	38.7 (16.1)	37.1 (16)	36.0 (33.2)	<0.001
Electrode placement				0.863
Unilateral	3069 (89.5%)	2160 (89.6%)	1800 (89.1%)	
Bitemporal	279 (8.1%)	199 (8.3%)	178 (8.8%)	
Bifrontal	82 (2.4%)	52 (2.2%)	43 (2.1%)	

The table shows treatment characteristics for the ECT series of each group, including the total number of sessions in the series, charge given for the last treatment in each series, seizure duration for the last treatment in the series and electrode placement. *P*-values were calculated using one-way ANOVA for mean and Pearson chi-square for proportions. ECT, electroconvulsive therapy; mC, millicoulomb.

Table 3 Proportion and odds ratios of response, remission and subjective memory worsening

	<i>N</i> included	Response/remission/memory worsening, %	Crude odds ratio (95% CI)	<i>P</i> -value	Adjusted odds ratio (95% CI)	<i>P</i> -value
Response, CGI-I						
Low-dose interval	3140	75.7%	1.27 (1.11–1.45)	<0.001	1.22 (1.07–1.40)	0.004
Medium-dose interval	2266	74.8%	1.21 (1.05–1.39)	0.008	1.15 (1.00–1.33)	0.056
High-dose interval	1805	71.1%	Reference	Reference	Reference	Reference
Distinct response, CGI-I						
Low-dose interval	3140	32.7%	1.60 (1.40–1.83)	<0.001	1.51 (1.32–1.73)	<0.001
Medium-dose interval	2266	29.4%	1.37 (1.19–1.58)	<0.001	1.32 (1.14–1.53)	<0.001
High-dose interval	1805	23.3%	Reference	Reference	Reference	Reference
Remission, CGI-S						
Low-dose interval	2950	23.0%	1.48 (1.27–1.72)	<0.001	1.37 (1.17–1.60)	<0.001
Medium-dose interval	2134	21.6%	1.36 (1.16–1.60)	<0.001	1.30 (1.10–1.54)	0.002
High-dose interval	1716	16.8%	Reference	Reference	Reference	Reference
Response, MADRS-S						
Low-dose interval	1470	65.4%	1.39 (1.17–1.65)	<0.001	1.31 (1.09–1.56)	0.004
Medium-dose interval	1026	60.7%	1.13 (0.94–1.36)	0.180	1.06 (0.87–1.28)	0.556
High-dose interval	889	57.7%	Reference	Reference	Reference	Reference
Remission, MADRS-S						
Low-dose interval	1801	43.9%	1.45 (1.24–1.70)	<0.001	1.31 (1.11–1.55)	0.002
Medium-dose interval	1210	39.7%	1.22 (1.03–1.45)	0.023	1.13 (0.94–1.36)	0.18
High-dose interval	1031	35.0%	Reference	Reference	Reference	Reference
Subjective memory worsening						
Low-dose interval	1923	21.6%	1.34 (1.11–1.62)	0.002	1.32 (1.09–1.60)	0.004
Medium-dose interval	1398	17.0%	1.00 (0.81–1.22)	0.971	0.99 (0.80–1.22)	0.893
High-dose interval	1176	17.0%	Reference	..	Reference	..

For each outcome, the table shows the number of included patients (if patients had missing data on the outcome they were excluded from the logistic regression models). The proportion of each outcome. Crude odds ratios and their 95% confidence intervals were calculated by correlating the age and gender-adjusted dose intervals and the outcomes without any further variables. A regression model adjusted for age, gender, number of treatments, psychiatric comorbidity and psychiatric pharmacotherapy was used to calculate adjusted odds ratios and their 95% confidence intervals. *P*-values are shown for both crude and adjusted odds ratios separately. CGI-S, Clinical Global Impression – Severity Scale; CGI-I, Clinical Global Impression – Improvement Scale; MADRS-S, Montgomery–Åsberg Depression Rating Scale, self-rated version.

$P \leq 0.001$; medium-dose interval: odds ratio 1.32, 95% CI 1.14–1.52, $P \leq 0.001$). Furthermore, for distinct response we also investigated a possible dose-dependent relationship by calculating odds ratios with medium-dose interval as a reference, which showed an association with higher chance for distinct response with low-dose intervals compared with medium-dose intervals, and a lower chance for distinct response with high-dose intervals compared with medium-dose intervals (low-dose interval: odds ratio 1.14, 95% CI 1.01–1.29, $P \leq 0.028$; high-dose interval: odds ratio 0.76, 95% CI 0.75–0.87, $P \leq 0.001$). Low-dose intervals were associated with higher risk of subjective memory worsening compared with high-dose intervals (odds ratio 1.32, 95% CI 1.09–1.60, $P = 0.004$).

Discussion

Main findings

In this study, patients with MDD who received low-dose anaesthetic during ECT had higher response and remission rates compared

with patients who received high-dose anaesthetic. All of our outcomes pointed in the same direction, with superior MDD treatment outcomes, measured with both clinical assessments and self-rated scales, for low- compared with high-dose intervals. For one of our secondary outcomes, distinct response, there was a dose-dependent relationship. Furthermore, treatment series were shorter and seizure durations were longer for patients who received low-dose intervals. We found an increased risk of subjective memory worsening for patients who received low-dose intervals compared with patients who received high-dose intervals. The measurements of memory disturbance are routinely done the day after the last ECT session. With this in mind, it is more likely that the seizure activity itself was associated with the greater effect on memory the day after treatment rather than a direct effect of the anaesthetics. Previous studies have shown that objective negative effects on memory typically disappear within 15 days, and after this, the cognitive function improves.²⁸ The prolonged epileptic seizures in the low-dose interval group could explain the increased subjective memory worsening the day after the last ECT. Postictal agitation/post-ECT confusion is

more frequent in unmodified ECT, which is an additional possible explanation to increased memory disturbances in the low-dose interval group.²⁹ On the other hand, improved response and remission rates with lower doses of anaesthesia are likely to be associated with cognitive improvement in a longer perspective. However, a long-term follow-up of this finding is needed.

Comparison with previous findings

To our knowledge, this is the first study to investigate the effects of anaesthetic dose on response and remission rates after ECT for MDD. Our results are in line with previous findings of a dose-dependent relationship between anaesthetic dose and seizure thresholds and durations.⁹


Strengths and limitations

This study has several strengths. It is the first study to investigate the association between dose of anaesthetic agents and MDD treatment outcome with ECT. We have access to high-resolution outcome data from the Q-ECT. Cross-matching this unique database with validated Swedish nationwide registers allowed us to adjust for comorbidities and pharmacotherapy. This study also has several limitations. Data on the anaesthetic agent was missing in 2500 patients in our initial cohort, and these were excluded from further analysis. The patients with missing data on anaesthetic agents were, on average, older than the patients who had complete data (patients with available data had a mean age of 53.1 years, patients with missing data had a mean age of 55.9 years; $P \leq 0.001$). We believe that it is unlikely that age itself is the cause of the missing data. The mechanism of missing data on the exposure is most likely underreporting of this data to the register rather than unknown or unrecorded data for certain age groups. The reporting to the register is not carried out by the anaesthesiologist and thus, it is unlikely there would be a systematic bias between age and missing data on anaesthetic dose. However, older age is an established predictor of ECT response.^{24,25} Considering that older patients typically receive lower doses of anaesthetics, it could be difficult to disentangle the effect of age and anaesthetic on ECT response.^{26,27} To remove potential confounding from age, we used both age- and gender-adjusted dose intervals, and included a continuous age variable in our logistic regression models. An additional limitation was lack of data on weight, and thus, we cannot present administered anaesthetic dose per kilogram. Ideally, this information could have helped to construct our adjusted dose intervals. However, by adjusting dose intervals for both age and gender, we believe that weight is indirectly considered. We also lacked information on time from induction of anaesthesia to administered ECT, which could affect anaesthetic depth during ECT. As a consequence, we cannot rule out any systematic differences in time from induction to administered ECT being dependent on anaesthetic dose. Moreover, our three dose groups were different in size. This was because of the distribution of anaesthetic dose, with accumulation of the same doses in lower-dose intervals. Furthermore, because of the observational nature of this study, there could be residual confounding in our analyses.

In our study, we lack measurements of potential awareness during ECT. However, the induced seizure causes an immediate loss of consciousness, minimising the risk of awareness during the procedure. Furthermore, the general risks with light anaesthesia are low.¹⁶ Nevertheless, further studies need to address if the risk of awareness and postictal agitation during the procedure increase with low-dose intervals of anaesthetics.

In conclusion, our study shows more favourable outcomes in terms of response and remission rates with low-dose intervals of anaesthetics for MDD after ECT, compared with high-dose

intervals. For anaesthesia during ECT, anaesthetic depth is likely the main issue regarding MDD treatment effect rather than the actual dose. In our study anaesthetic dose likely corresponds to anaesthetic depth, although other factors also contribute. The clinical implications of our results would be that a light anaesthesia increases the likelihood of response and remission with ECT for MDD. This should be weighed against the risk of awareness during anaesthesia and the potential risk of postictal agitation. In conclusion, deep anaesthesia should be avoided during ECT, to improve treatment outcomes for MDD. Further studies need to explore if similar effects are achieved by increasing the time from induction to stimulation, as by lowering the anaesthetic dose. In addition, further studies on long-term effects of anaesthetic dose during ECT for MDD are needed.

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First received 25 Nov 2020, final revision 24 Feb 2021, accepted 25 Feb 2021

Supplementary material

Supplementary material is available online at <https://doi.org/10.1192/bjpo.2021.31>.

Data availability

Requests for the data that supports the findings of this study can be sent to the corresponding author, A.K.. The data are not publicly available due to the privacy of research participants.

Author contributions

M.T., A.K., A.N., M.B. and E.M.-R. participated in study design. A.K. and R.A. carried out statistical analysis. A.K., M.T. and A.N. carried out initial preparations of the manuscript. A.K. had full access to the data in the study, and takes responsibility for the integrity of the data and the accuracy of data analysis. In addition, M.T. and R.A. verified the underlying data and the results. All authors reviewed and provided comments on the manuscript. All authors have read and approved the final version of the manuscript before publication.

Funding

M.T. has received funding from the Söderström-König Foundation. A.K. has received funding from Stiftelsen Professor Bror Gadelius Minnesfond. No funder had any role in study design, data collection, data analysis, data interpretation or writing of the manuscript.

Declaration of interest

None.

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