



Effect of Sodium Tanshinone IIA Sulfonate Injection on Blood Lipid in Patients With Coronary Heart Disease: A Systematic Review and Meta-Analysis of Randomized Clinical Trials

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Front. Cardiovasc. Med. 8:770746. doi: 10.3389/fcvm.2021.770746 Hufang Zhou^{1†}, Ying Zhao^{2†}, Wenhua Peng^{1†}, Wenbo Han³, Zichen Wang¹, Xiaoxia Ren¹, Dayang Wang¹, Guozhong Pan^{1*}, Qian Lin^{4*} and Xian Wang^{1*}

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Background: Lipid-lowering therapy is very important in secondary prevention of coronary heart disease (CHD). In many clinical trials, it has been found that Sodium Tanshinone IIA Sulfonate Injection (STS) have a lipid-lowering effect while reducing major cardiovascular events in patients with CHD. However, up to now, there is no system review on the effectiveness and safety of STS affecting blood lipids.

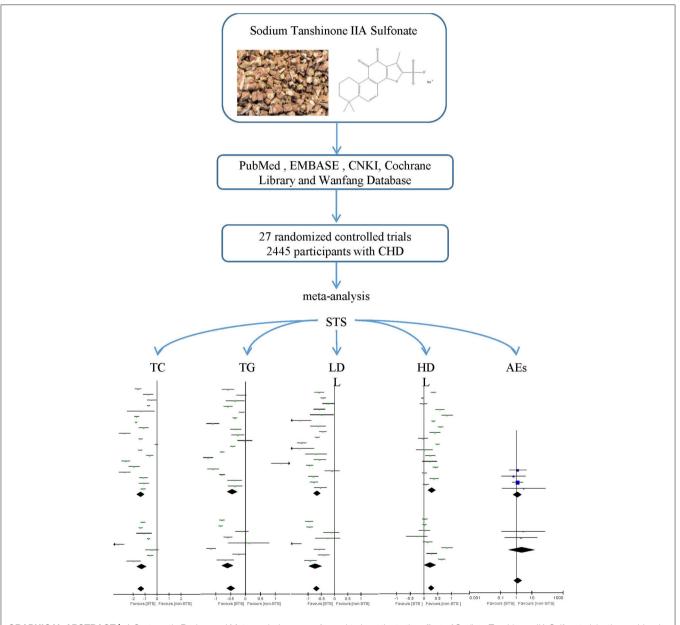
Purpose: The aim of this review is to systematically assess the effects of STS on blood lipid levels in patients with CHD.

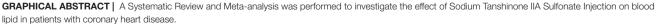
Methods: Until Mar 2021, five databases (PubMed, EMBASE, Cochrane Library, China National Knowledge Infrastructure, and Wanfang Database) were searched for randomized controlled trials (RCTs) about STS treating patients with CHD. Risk bias was assessed for included studies according to Cochrane handbook. The primary outcome was total cholesterol (TC). The secondary outcomes were triglycerides (TG), low-density lipoprotein cholesterol (HDL-c), high-density lipoprotein cholesterol (HDL-c), and adverse events (AEs).

Results: A total of 27 trials including 2,445 CHD patients met the eligibility criteria. Most trials had high risks in random sequence generation, allocation concealment, blinding of patients and personal, blinding of outcome assessment. Meta-analysis showed that STS significantly reduced plasma TC levels [MD = -1.34 mmol/l 95% CI (-1.59, -1.09), p < 0.00001, $l^2 = 98\%$], TG levels [MD = -0.49 mmol/l 95% CI (-0.62, -0.35), p < 0.00001, $l^2 = 97\%$], LDL-c levels [MD = -0.68 mmol/l (-0.80, -0.57), p < 0.00001, $l^2 = 96\%$], increased HDL-c levels [MD = 0.26 mmol/l (0.15, 0.37), p < 0.00001, $l^2 = 97\%$], without increasing the incidence of AEs [RR = 1.27 95% CI (0.72, 2.27), p = 0.94, $l^2 = 0\%$] in patients with CHD.

Conclusion: STS can safely and effectively reduce plasma TC, TG and LDL-c levels in patients with CHD, and improve plasma HDL-c levels. However, these findings require careful recommendation due to the low overall quality of RCTs at present. More multi-center, randomized, double-blind, placebo-controlled trials which are designed follow the CONSORT 2010 guideline are needed.

Keywords: Sodium Tanshinone IIA Sulfonate, blood lipid, coronary heart disease, randomized controlled trials, systematic review, meta-analysis





INTRODUCTION

Lipid-lowering therapy, such as Statins, is very important in secondary prevention of CHD. Statins play an irreplaceable role as first-line lipid-lowering drugs for secondary prevention of CHD. Clinical practice guidelines recommend that all non-ST-elevation acute coronary syndromes (NSTE-ACS) patients should be treated with high-intensity statins as soon as possible after admission (1). Statins use should also be considered in patients with confirmed stable coronary heart disease (SCAD) or chronic coronary syndrome (CCS) regardless of LDL-c levels (2). However, the use of statin lipid-lowering drugs is susceptible to clinical side effects, including statin-related muscle symptoms (SAMS), as well as the risks of liver disease, diabetes, cognitive impairment and hemorrhagic stroke (3, 4). Guidelines also recommend that dietary supplements, including phytosterols, should be considered for high-risk CHD patients with statin intolerance.

Traditional Chinese medicine, which has a history of thousands of years, provides us with a large number of complementary and alternative drugs. In recent years, red yeast rice extract (5), berberine (6), ryegrass extract (7), policosanol (8), astaxanthin (9) and other drugs were found to have significant lipid-lowering effects in clinical and basic research, and some of them had been used to assist in clinical lipidlowering therapy. Salvia miltiorrhiza Bunge is one of the most commonly used herbs to accelerate blood circulation in China for treating cardiovascular diseases (CVDs) and other circulatory disturbance-related diseases (10). Tanshinone IIA is one of the most pharmacologically active ingredients extracted from Salvia miltiorrhiza Bunge and is considered as a promising natural cardioprotective agent which can reduce the levels of many inflammatory factors related to atherosclerosis progression, such as C-reactive protein (CRP), interleukin-6 (IL-6), tumor necrosis factor-alpha (TNF-α), vascular cell adhesion molecule-1 (VCAM-1), CD40, monocyte chemotactic protein-1 (MCP-1) and matrix metalloproteinase-9 (MMP-9) (10-13). Because of low oral bioavailability of tanshinone IIA Sulfonate, intravenous sodium tanshinone IIA sulfate (STS) has been developed. Tanshinone IIA preparation is one of the most widely used medical treatment for patients with CHD in China. In many clinical trials in China, STS has been found to have a lipidlowering effect while reducing major cardiovascular events in patients with CHD (14). However, up to now, there has been no systematic review on the efficacy and safety of STS affecting blood lipids in patients with CHD. Therefore, the purpose of this systematic review and meta-analysis is to systematically assess the efficacy and safety of STS on blood lipid levels in patients with CHD.

METHODS

The review protocol was registered at PROSPERO (No: CRD42018109290, https://www.crd.york.ac.Uk/prospero/). This study was conducted according to Cochrane Handbook for Systematic Reviews of Interventions (15) and was reported according to Preferred Reporting Items for Systematic reviews and Meta-Analyses (PRISMA) (16).

Literature Search

The literature searches were conducted in the following five databases: PubMed (1966 to 2021), EMBASE (1974 to 2021), China National Knowledge Infrastructure (CNKI) (1979 to 2021), Cochrane Library (2000 to 2021), and Wanfang Database (1985 to 2021). In order to check the whole, we searched all the literatures concerning STS. The main search terms was "Sodium Tanshinone IIA Sulfonate." There are no restrictions on language, date of publication or status of publication. All searches ended on March 4, 2021.

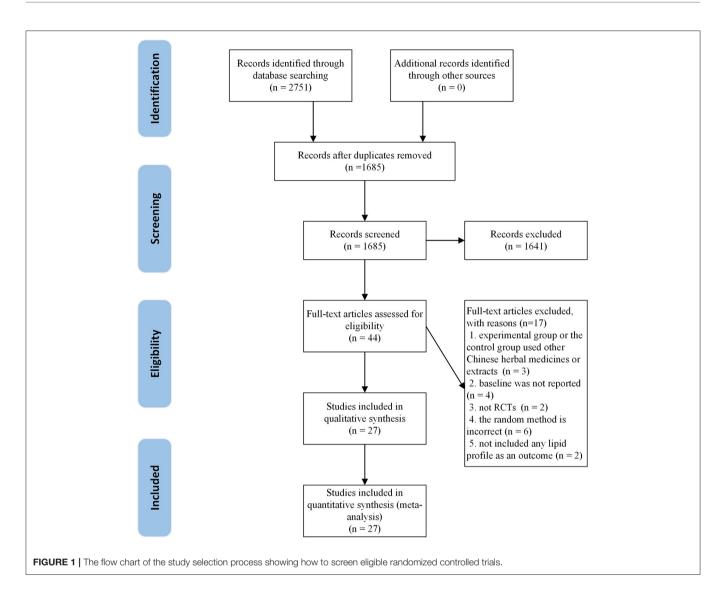
Eligibility Criteria

Original literatures were included if they met the following inclusion criteria: (1) Types of studies (S): RCTs with or without blinding; (2) Types of participants (P): Patient with CHD, including patients with CCS and acute coronary syndrome (ACS); (3) Types of interventions (I): STS with or without lipid-lowering drugs (LLDs); (4) Types of comparators (C): placebo, no drug or LLDs; intervention and control groups that used routine therapy (RT) simultaneously; (5) Types of outcome measures (O): any of the lipid profile parameters (including TC, LDL-c, HDL-c, TG and AEs). Exclusion criteria: (1) Duplicated publications; (2) Trials whose methods of allocation use date of birth, date of admission, hospital numbers, or alternation; (3) Use of other herbal medicine or extracts either in intervention or control group; (4) Lack of baseline or related lipid information.

Data Extraction

Data extraction was performed independently by two reviewers (Zhou HF, Lin Q). The extracted data included: authors, title of study, year of publication, sample size, PICOS details and treatment duration. We defined TC as the primary outcome and LDL-c, HDL-c, TG and AEs as secondary outcomes. AEs were defined as a composite of events including headache, dizziness, gastrointestinal reaction, liver function damage, allergy, renal function damage and bleeding.

Abbreviations: ACS, Acute Coronary Syndrome; AEs, Adverse Events; AMI, Acute Myocardial Infarction; CCS, Chronic Coronary Syndrome; CHD, Coronary Heart Disease; CI, Confidence Interval; CNKI, China National Knowledge Infrastructure; CONSORT, Consolidated Standards For Reporting Trials; CRP, C-reactive Protein; CVDs, Cardiovascular Diseases; GRADE, grading of recommendations assessment, development, and evaluation; HDLc, High-Density Lipoprotein Cholestero; IL-6, Interleukin-6; LDL-c, Lowdensity Lipoprotein Cholesterol; LLDs, Lipid-lowering Drugs; MCP-1, Monocyte Chemotactic Protein-1; MD, Mean Difference; MMP-9, Matrix Metalloproteinase-9; NSTE-ACS, Non-ST-elevation Acute Coronary Syndromes; PRISMA, Preferred Reporting Items For Systematic Reviews and Meta-analyses; RCTs, Randomized Controlled Trials; RIS, required information size; RR, Risk Ratio; RT, Routine Therapy; SAMS, Statin-Related Muscle Symptoms; SCAD, Stable Coronary Heart Disease; STS, Sodium Tanshinone IIA Sulfate; TC, Total Cholesterol; TG, Triglycerides; TNF-α, Tumor Necrosis Factor-Alpha; TSA, Trial Sequential Analysis; UAP, Unstable Angina Pectoris.



Risk of Bias Assessment

Two reviewers (Zhou HF, Lin Q) independently assessed the risk of bias for each included trial according to the Cochrane Handbook for Systematic Reviewers of Interventions version 5.1.0 (15). The items included random sequence generation (selection bias), allocation concealment (selection bias), blinding of participants and personnel (performance bias), blinding of outcome assessment (detection bias), incomplete outcome data (attrition bias), selective reporting (reporting bias), and baseline data comparability (other bias). Each item was categorized as low/unclear/high risk of bias. Disagreements were resolved by discussion, with involvement of a third review author (Wang X) when necessary. In addition, we used the grading of recommendations assessment, development, and evaluation (GRADE) approach to evaluate the quality of included evidences.

Data Analysis

If the index to be analyzed were continuous variables, we selected mean difference (MD) as the effect scale. If it were

binary variables, we selected risk ratio (RR). All results were presented with 95% confidence intervals (CI). Heterogeneity among trials was assessed by Cochrane's Q test and I-squared statistic. According to Cochrane Handbook for Systematic Reviews of Interventions, the scale of I^2 had a range of 0-100% and values on the order of 0-40%, 30-60%, 50-90%, and 75-100% were considered might not be important, may represent moderate heterogeneity, may represent substantial heterogeneity and considerable heterogeneity, respectively (15). For heterogeneous studies, we adopted use a random-effects model to estimate the overall effect instead of a fixed-effect model, because random-effects models assess the outcomes of the study according to within-trial as well as between-trial variance (17), thus providing more conservative results. Furthermore, meta-regression analysis was used to explore the sources of heterogeneity. The sensitivity analysis was also performed by removing each study one at a time to evaluate the stability of the results. The Trial Sequential Analysis (TSA) was used to determine the robustness of on primary outcome and calculate

TABLE 1 | Characteristics of the included RCTs and the detail of PICOS.

Study ID	Participants	No. of participants (I/C)	Age (years)	Male (%)	Interventions	Comparators	Duration
Fang et al., (63)	CHD	33/33	l: 58.6 ± 4.8	l: 17(51.52%)	STS 80 mg qd; statin; RT	statin; RT	14D
			C: 57.9 \pm 5.4	C: 16(48.48%)			
Yan et al., (58)	CHD	36/36	I: 57.84 ± 7.63	l: 23 (63.89%)	STS 60 mg qd; Simvastatin 10 mg qn; RT	Simvastatin 10 mg qn; RT	1M
			C: 57.76 \pm 7.82	C: 21 (58.33%)			
Zhang (45)	CHD with hyperuricemia	50/50	I: 61.70 ± 6.89	l: 35 (70.00%)	STS 60 mg qd; Simvastatin 10 mg qn; RT	Simvastatin 10 mg qn; RT	14D
			$C: 61.41 \pm 6.85$	C: 33 (66.00%)			
He et al., (54)	CHD	29/29	I: 56.6 \pm 16.1	I: 21 (72.41%)	STS 60 mg qd; Atorvastatin; RT	Atorvastatin; RT	14D
			C: 53.9 ± 14.5	C: 17 (58.62%)			
Li et al., (55)	CHD	50/50	I: 66 ± 5	I: 29 (58.00%)	STS 80 mg qd; statin; RT	Atorvastatin 10mg qn; RT	1M
			C: 64 ± 5	C: 28(56.00%)			
Yang (60)	CHD	52/52	I: 67.64 ± 9.46	I: 33 (63.46%)	STS 50 mg qd; statin; RT	statin; RT	14D
			C: 68.42 ± 9.32	C: 32 (61.54%)			
Cao et al., (61)	CHD	49/49	I: 61.49 ± 9.14	I: 33 (67.35%)	STS 60 mg qd; statin; RT	statin; RT	28D
			C: 62.59 ± 8.21	C: 30 (61.22%)			
Zhu et al., (62)	CHD	48/48	I: 65.3 ± 8.8	l: 28 (58.33%)	STS 60 mg qd; statin; RT	statin; RT	28D
			C: 64.7 ± 8.2	C: 26 (54.17%)			
Lin et al., (39)	ACS	49/49	l: 55.98 ± 5.49	I: 30 (61.22%)	STS 50 mg qd; statin; RT	statin; RT	14D
			C: 56.44 ± 5.86	C: 32 (65.31%)			
Guo (56)	CHD	50/50	I: 60.53 ± 9.42	l: 33 (66.00%)	STS 60 mg qd; Atorvastatin 20 mg qn; RT	Atorvastatin 20 mg qn; RT	14D
			C: 60.14 ± 8.43	C: 31 (62.00%)			
Jin (44)	UAP	50/50	I: 66.24 ± 8.3	I: 35 (70.00%)	STS 60 mg qd; Atorvastatin 20 mg qn; RT	Atorvastatin 20 mg qn; RT	14D
			C: 64.35 ± 9.4	C: 38 (76.00%)			
He et al., (57)	CHD	49/49	I: 64.19 ± 5.82	l: 28 (57.14%)	STS 60 mg qd; Simvastatin 10 mg qn; RT	Simvastatin 10 mg qn; RT	1M

(Continued)

Tanshinone IIA and Lipid-Lowing

TABLE 1 | Continued

Study ID	Participants	No. of participants (I/C)	Age (years)	Male (%)	Interventions	Comparators	Duration
			C: 62.31 ± 6.37	C: 30 (61.22%)			
Gao et al., (52)	CHD	46/46	I: 58.6 ± 8.9	l: 29 (63.04%)	STS 60 mg qd; Atorvastatin 20 mg qn; RT	Atorvastatin 20 mg qn; RT	14D
			C: 57.2 ± 8.1	C: 27 (58.70%)			
Xun et al., (53)	CHD	80/80	I: 60.12 ± 9.47	l: 37 (46.25%)	STS 80 mg qd; Atorvastatin 10 mg qn; RT	Atorvastatin 10 mg qn; RT	30D
			C: 58.75 ± 9.25	C: 38 (47.50%)			
Li et al., (38)	ACS	30/30	unclear	unclear	STS 80 mg qd; statin; RT	statin; RT	7D
Wang and Zheng (42)	AMI	46/46	unclear	unclear	STS 60 mg qd; Atorvastatin; RT	Atorvastatin; RT	7D
Huang (59)	CHD	30/30	l: 65.77 ± 8.90	l: 14 (46.67%)	STS 40 mg qd; statin; RT	statin; RT	14D
			C: 65.53 ± 7.78	C: 10 (33.33%)			
Zhang et al., (37)	ACS	32/32	l: 60 ± 7	I: 24 (75.00%)	STS 60 mg qd; statin; RT	statin; RT	10D
			C: 58 ± 9	C: 22 (68.75%)			
Ge et al., (43)	UAP	30/30	l: 57 ± 7	I: 19 (63.33%)	STS 60mg qd; statin; RT	statin; RT	15D
			C: 58 \pm 7	C: 20 (66.67%)			
Zhao et al., (46)	CHD	39/39	l: 47.45 ± 2.13	I: 21 (53.85%)	STS 80mg qd; Aspirin 50 bid	no drug; Aspirin 50 bid	14D
			l: 51.23 \pm 2.72	I: 19 (48.72%)			
Fan and Yang (41)	AMI	47/42	I: 59.7 \pm 5.6	I: 26 (55.32%)	STS 60mg qd; RT	no drug; RT	7D
			C: 58.9 ± 4.4	C: 23 (54.76%)			
Wei and Shen (50)	CHD	60/60	I: 49.18 ± 12.50	I: 38 (63.33%)	STS 60mg qd; RT	no drug; RT	14D
			I: 52.36 \pm 10.18	I: 36 (60.00%)			
Zhang et al., (40)	AMI	48/48	unclear	unclear	STS 40-80mg qd; RT	no drug; RT	7D
Gao et al., (51)	CHD	30/30	unclear	unclear	STS 60mg qd; RT	no drug; RT	14D
Cheng and Zhang (49)	CHD	60/60	I: 67.5 ± 6.1	I: 36 (60.00%)	STS 60mg qd; RT	no drug; RT	14D
			C: 66.2±4.5	C: 38 (63.33%)			
Guo and Sun (48)	CHD	65/62	unclear	I: 35 (53.85%)	STS 50mg qd; RT	no drug; RT	14D
				C: 34 (54.84%)			
Tao and Jia (47)	CHD	40/37	l: 72 ± 10	I: 27 (67.50%)	STS 80mg qd; RT	no drug; RT	14D
			C: 74 \pm 8.5	C: 25 (67.57%)			

I, Intervention; C, control; RT, routine therapy; D, days; M, months.

the required information size (RIS) in the meta-analysis (18). Subgroup analysis was performed according to various types of interventions (with or without LLDs). The publication bias was detected by the funnel plot, the Begger's test and the Egger's test (19).

RESULTS

Study Selection

A total of 2,751 articles were retrieved from the initial search. After deleting the duplicate literature, 1,685 articles remained. After reading the title and abstract, 1,641 articles were excluded and 44 articles were screened in detail. By reading the full text of the remaining 44 articles, 17 articles were excluded, which did not meet our inclusion criteria: two were not RCTs (20, 21); three were combined with other herbs or extracts (22–24); two did not include any blood lipids profile in the outcome (25, 26); four did not report baseline (27–30) and six used nonstandard randomization methods (31–36). Finally, 27 trials (37–63) were qualified and included in the meta-analysis. The flow chart of literature screening is as follows (**Figure 1**).

Study Characteristics

The 27 RCTs were all from China and single-center trials. The 2,445 patients recruited (1,228 in the trial group and 1,217 in the control group) were all hospitalized patients. The average age ranged from 47 to 72 years. Most studies recruited more male patients. Three studies (37–39) included patients with ACS; Three studies (40–42) included patients with acute myocardial infarction (AMI); Two studies (43, 44) included patients with unstable angina pectoris (UAP) and one study included CHD patients with hyperuricemia (45). The remaining studies only showed that patients with CHD or myocardial ischemia are included.

Except for one study (46), which was based on oral aspirin, the remaining studies were based on secondary prevention of CHD. Eight studies (40, 41, 46-51) did not use statins in combination. Seven studies (42, 44, 52-56) clearly stated that atorvastatin was used in combination; Three studies (45, 57, 58) stated that simvastatin was used in combination; Nine studies (37-39, 43, 59-63) only stated the use of statins, but did not specify which statins were used; Other studies stated that no statins were used in combination. The dosage of STS ranged from 40 to 80 mg per day. The intervention duration ranged from 1 week to 1 month. Most studies reported in detail the results of blood lipids. All 27 included RCTs provided data for TC, 26 RCTs provided data for TG and LDL-c and 25 RCTs provided data for HDL-c. Of the 27 RCTs, 15 studies (37-42, 44, 48, 50-52, 54-56, 60) did not report information about AEs. Six trials (43, 46, 53, 57, 59, 61) reported no AEs and the remaining six trials (45, 47, 49, 58, 62, 63) reported a total of 39 AEs. Characteristics of the trials and the detail of PICOS are shown in Table 1.

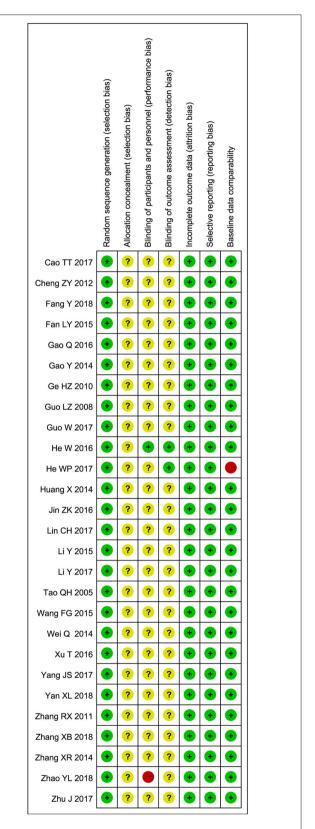


FIGURE 2 | Risk of bias summary.

STS					n-STS			Mean Difference		Mean Difference		
tudy or Subgroup			Total	Mean	SD	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% Cl		
1.1 STS with LLD v	s. LLD a	alone										
e HZ 2010	4.74	0.36	30	6.32	0.6	30	3.9%	-1.58 [-1.83, -1.33]	2010	-		
hang RX 2011	3.02	0.71	32	3.76	0.95	32	3.6%	-0.74 [-1.15, -0.33]	2011			
uang X 2014	3.44	0.83	30	3.88	0.66	30	3.7%	-0.44 [-0.82, -0.06]	2014			
ang FG 2015	2.31	0.22	46	3.02	0.28	46	4.0%	-0.71 [-0.81, -0.61]	2015	-		
Y 2015	2.38	1.2	30	3.73	2.91	30	2.2%	-1.35 [-2.48, -0.22]	2015			
u T 2016	3.16	0.45	80	4.86	0.73	80	3.9%	-1.70 [-1.89, -1.51]	2016	-		
n ZK 2016	2.46	0.18	50	4.2	0.19	50	4.0%	-1.74 [-1.81, -1.67]	2016	• ·		
e W 2016	4.28	0.79	49	5.39	1.14	49	3.7%	-1.11 [-1.50, -0.72]	2016			
ao Q 2016	4.1	1.1	46	6.8	1.2	46	3.5%	-2.70 [-3.17, -2.23]	2016 —	-		
ang JS 2017	4.75	0.38	52	6.12	0.41	52	4.0%	-1.37 [-1.52, -1.22]	2017	÷		
n CH 2017	3.26	0.28	49	3.31	0.3	49	4.0%	-0.05 [-0.16, 0.06]		+		
Y 2017	3.3	0.3	50	4.7	0.5	50	4.0%	-1.40 [-1.56, -1.24]		-		
nu J 2017	3.74	0.63	48	4.35	0.71	48	3.9%	-0.61 [-0.88, -0.34]				
ao TT 2017	4.29	0.71	49	6.89	1.02	49	3.7%	-2.60 [-2.95, -2.25]				
e WP 2017	4.2		29	6.1	0.8	29	3.6%	-1.90 [-2.34, -1.46]				
uo W 2017	4.09	1.03	50	6.79	1.36	50	3.5%	-2.70 [-3.17, -2.23]	2017 —	-		
nang XB 2018	4.29	0.68	50	5.42	0.95	50	3.8%	-1.13 [-1.45, -0.81]				
an XL 2018	4.32	0.66	36	5.41	1.07	36	3.6%	-1.09 [-1.50, -0.68]				
ang Y 2018	4.94	0.33	33	6.11	0.44	33	3.9%	-1.17 [-1.36, -0.98]	2018			
ubtotal (95% CI)			839			839	70.7%	-1.36 [-1.68, -1.04]		◆		
eterogeneity: Tau ² =	0.48; Cł	ni² = 9'	19.44. c	f = 18 (P < 0.0	00001);	l² = 98%					
est for overall effect:	Z = 8.28	3 (P < (0.00001	D)								
				,								
1.2 STS vs. no drug	1											
ao QH 2005	4.95	0.35	40	6.12	0.44	37	4.0%	-1.17 [-1.35, -0.99]	2005	-		
uo LZ 2008	4.76	0.37	65	6.12	0.44	62	4.0%	-1.36 [-1.50, -1.22]	2008	÷		
heng ZY 2012	4.95	1.61	60	6.03	1.83	60	3.2%	-1.08 [-1.70, -0.46]	2012			
nang XR 2014	2.31	0.22	48	3.02	0.28	48	4.0%	-0.71 [-0.81, -0.61]	2014	-		
ei Q 2014	3.24	1.05	60	6.48	1.6	60	3.5%	-3.24 [-3.72, -2.76]	2014	-		
ao Y 2014	3.48	0.85	30		1.18	30	3.4%	-0.42 [-0.94, 0.10]				
an LY 2015	3.22	0.23	47	3.79	0.75	42	3.9%	-0.57 [-0.81, -0.33]				
nao YL 2018		1.28	39		1.33	39	3.3%	-1.95 [-2.53, -1.37]				
			389			378	29.3%	-1.28 [-1.68, -0.89]		◆		
ubtotal (95% CI)	0.29; Cł	ni² = 17	70.09, c	df = 7 (P	< 0.0	0001);	² = 96%	-				
ubtotal (95% CI) eterogeneity: Tau ² =		(P < (0.00001)								
eterogeneity: Tau ² =	Z = 6.34							4 24 5 4 50 4 001				
eterogeneity: Tau ² = est for overall effect:	Z = 6.34		1228			1217	100.0%	-1.34 [-1.59, -1.09]				
eterogeneity: Tau ² = est for overall effect: otal (95% CI) eterogeneity: Tau ² =	0.40; Cł		122.93,		(P < 0							
eterogeneity: Tau ² = est for overall effect: otal (95% CI) eterogeneity: Tau ² = est for overall effect:	0.40; Cł Z = 10.5	51 (P <	122.93, 0.0000	01)		0.00001); I² = 98%			-2 -1 0 1 2 Favours [STS] Favours [non-STS]		
eterogeneity: Tau ² = est for overall effect: otal (95% CI) eterogeneity: Tau ² =	0.40; Cł Z = 10.5	51 (P <	122.93, 0.0000	01)		0.00001); I² = 98%			-2 -1 0 1 2 Favours [STS] Favours [non-STS]		

Risk of Bias in Individual Studies

Randomization had been mentioned in all studies. Fifteen of them explicitly described methods of randomization, such as random number tables or dichroic spheres. None of the studies mentioned allocation concealment. Except for two studies, none of the other studies mentioned the blind method. One study made it clear that no blind method was used, and another made it clear that blind method was used. All studies have a lowrisk selective reporting bias. All included studies were baseline comparable. The Risk of bias summary is shown in **Figure 2**. Supplement one (**Supplementary Tables S1, S2**) summarized the level of evidence for the studies included and indicated that the overall quality of the evidence was very low.

Overall Results of Meta-Analysis

In addition to AEs, the remaining meta-analysis showed a high degree of heterogeneity among studies. Therefore, the meta-

analysis of AEs adopted a fixed-effect model, while the remaining meta-analysis used a random-effect model.

TC: As shown in **Figure 3**, meta-analysis of 27 studies showed that STS significantly reduced plasma TC levels in patients with CHD [MD = -1.34 mmol/l, 95% CI (-1.59, -1.09), P < 0.00001, $I^2 = 98\%$]. Subgroup analysis revealed that STS significantly lowered TC levels when used alone compared to blank control [MD = -1.28 mmol/l, 95% CI (-1.68, -0.89), P < 0.00001, $I^2 = 96\%$], or used in combination with LLDs compared to LLDs [MD = -1.36mmol/l, 95% CI (-1.68, -1.04), P < 0.00001, $I^2 = 98\%$].

TG: As shown in **Figure 4**, pooled results of 26 studies revealed that STS significantly reduced plasma TG levels in patients with CHD [MD = -0.49 mmol/l, 95% CI (-0.62, -0.35), P < 0.00001, $I^2 = 97\%$). Subgroup analysis revealed that STS significantly lowered TG levels when used alone compared to blank control [MD = -0.60 mmol/l, 95% CI (-0.79, -0.41), P < 0.00001, $I^2 = 97\%$], or used in combination with LLDs compared

	ST	rs		non-ST	3		Mean Difference	Mean Difference
Study or Subgroup						Weight	IV, Random, 95% CI Yea	
2.1.1 STS with LLD vs								
Ge HZ 2010	1.51 0).16	30 2	1 0.53	30	3.9%	-0.59 [-0.79, -0.39] 2010	0
Zhang RX 2011	1.31 0	0.33	32 1.4	5 0.31	32	4.0%	-0.14 [-0.30, 0.02] 201	
Huang X 2014	1.12 0).44	30 1.4	7 0.75	30	3.4%	-0.35 [-0.66, -0.04] 2014	
Wang FG 2015	2.19 0	0.25	46 2.7	8 0.42	46	4.0%	-0.59 [-0.73, -0.45] 201	
He W 2016	1.7 0	0.37	49 2.0	3 0.52	49	3.9%	-0.33 [-0.51, -0.15] 2010	6
Gao Q 2016	1.4	0.4	46 2	5 0.6	46	3.8%	-1.10 [-1.31, -0.89] 2010	6
Xu T 2016	1.21 0).18	80 1.5	0.25	80	4.2%	-0.36 [-0.43, -0.29] 2010	6 -
Jin ZK 2016	0.89 0	0.15	50 1.0	2 0.27	50	4.2%	-0.13 [-0.22, -0.04] 2010	
Yang JS 2017	1.44 0	0.12	52 2.2	3 0.22	52	4.2%	-0.79 [-0.86, -0.72] 201	7 -
Li Y 2017	1.06 0).17	50 1.4	9 0.23	50	4.2%	-0.43 [-0.51, -0.35] 201	7 -
Lin CH 2017	5.68 0).62	49 5	7 0.59	49	3.7%	-0.02 [-0.26, 0.22] 201	7
Zhu J 2017	1.18 0).47	48 1.4	4 0.56	48	3.8%	-0.26 [-0.47, -0.05] 201	7
Guo W 2017	1.39 0).37	50 2.4	8 0.57	50	3.9%	-1.09 [-1.28, -0.90] 201	7
He WP 2017	3.1	0.8	29 1	9 0.4	29	3.4%	1.20 [0.87, 1.53] 201	7
Cao TT 2017	1.43	0.3	49 2.6	9 0.48	49	4.0%	-1.26 [-1.42, -1.10] 201	7
Yan XL 2018	1.67 0).41	36 2.0	2 0.6	36	3.7%	-0.35 [-0.59, -0.11] 2018	8
Fang Y 2018	1.49 0).12	33 2.2	8 0.21	33	4.2%	-0.79 [-0.87, -0.71] 2018	8
Zhang XB 2018	2.01 0).48	50 2.3	8 0.7	50	3.7%	-0.37 [-0.61, -0.13] 2018	8
Subtotal (95% CI)		8	809		809	70.3%	-0.45 [-0.61, -0.28]	◆
Heterogeneity: Tau ² = (Test for overall effect: 2 2.1.2 STS vs. no drug				/ (i = 0	00001)	1 - 57 %		
Tao QH 2005	1.5 0	11	40 2.2	9 0.2	37	4.2%	-0.79 [-0.87, -0.71] 200	5
Guo LZ 2008	1.45 0			9 0.21	62	4.2%	-0.84 [-0.90, -0.78] 200	
Cheng ZY 2012	2.19 0			0.65	60	3.8%	-0.12 [-0.34, 0.10] 2012	
Gao Y 2014	1.98 1			7 1.09	30	2.0%	0.11 [-0.57, 0.79] 2014	
Wei Q 2014	1.43 0			9 0.46	60	4.0%	-1.16 [-1.31, -1.01] 2014	
Zhang XR 2014	2.19 0			8 0.42	48	4.1%	-0.59 [-0.73, -0.45] 2014	
Fan LY 2015	1.31 0			3 0.36	42	3.8%	-0.22 [-0.43, -0.01] 201	
Zhao YL 2018	0.82 0			5 0.83	39	3.6%	-0.68 [-0.96, -0.40] 2018	
Subtotal (95% CI)			89		378	29.7%	-0.60 [-0.79, -0.41]	•
Heterogeneity: Tau ² = (Test for overall effect: 2				(P < 0.0	0001);	² = 93%		
Total (95% CI)		11	98		1187	100.0%	-0.49 [-0.62, -0.35]	•
Heterogeneity: $Tau^2 = ($) 11· Chi2			5 (P < 0			5140 [010£, -0100]	
Test for overall effect: 2 Test for subaroup differ	Z = 7.15 (F	P < 0.00	0001)		,			-1 -0.5 0 0.5 1 Favours [STS] Favours [non-STS]
GURE 4 Forest plot of c	ompariso	n: STS v	vs. non-8	TS in T(à.			

to LLDs [MD = -0.45 mmol/l, 95% CI (-0.61, -0.28), P < 0.00001, $I^2 = 97\%$].

LDL-c: Meta-analysis results of 25 studies demonstrated that STS also significantly reduced LDL-c levels in patients with CHD [MD = -0.68 mmol/l, 95% CI (-0.80, -0.57), P < 0.00001, $I^2 =$ 96%). Subgroup analysis revealed that STS significantly lowered LDL-c levels when used alone compared to blank control [MD = -0.73 mmol/l, 95% CI (-0.97, -0.49), P < 0.00001, $I^2 = 91\%$], or used in combination with LLDs compared to LLDs [MD= -0.66 mmol/l, 95% CI (-0.79, -0.53), P < 0.00001, $I^2 = 89\%$]. These results are shown in **Figure 5**.

HDL-c: Pooled results of 25 RCTs showed that STS significantly increased HDL-c levels [MD = 0.26 mmol/l, 95% CI (0.15, 0.37), P < 0.00001, $I^2 = 97\%$]. Subgroup analysis revealed that STS significantly increased HDL-c levels when used alone compared to blank control [MD = 0.22 mmol/l, 95% CI (0.01, 0.44), P < 0.00001, $I^2 = 95\%$], or used in combination with LLDs compared to LLDs [MD = 0.28 mmol/l, 95% CI (0.15, 0.42), P < 0.00001, $I^2 = 97\%$]. These results are shown in **Figure 6**.

Adverse events: Meta-analysis of 12 studies showed that no statistically significant difference was found between groups in terms of the incidence of adverse events [RR = 1.27, 95% CI (0.72, 2.27), P = 0.94, $I^2 = 0\%$]. Subgroup analysis revealed that no statistically significant difference was found between groups in terms of the incidence of adverse events when used alone compared to blank control [RR = -2.27, 95% CI (0.34, 15.09), P = 0.94, $I^2 = 0\%$], or used in combination with LLDs compared to LLDs [RR = 1.18, 95% CI (0.64, 2.17), P < 0.00001, $I^2 = 0\%$] (Figure 7).

Meta-Regression

The results above showed that there was considerable heterogeneity (75%-100%) among 27 studies. We conducted a meta-regression analysis which showed that the total dose of STS and the type of statins were important sources of heterogeneity in the included studies (**Supplementary Material 2**, **Supplementary Table S3**).

		STS		no	n-STS	6		Mean Difference	Mean Difference
			Total	Mean	SD	Total	Weight	IV, Random, 95% CI Year	IV, Random, 95% CI
3.1.1 STS with LLD v	s. LLD a	alone							
Ge HZ 2010	2.52	0.27	30	3.1	0.41	30	4.2%	-0.58 [-0.76, -0.40] 2010	
Zhang RX 2011	1.82	0.18	32	2.25	0.21	32	4.6%	-0.43 [-0.53, -0.33] 2011	-
Huang X 2014	1.91	0.38	30	2.15	0.46	30	4.0%	-0.24 [-0.45, -0.03] 2014	
Wang FG 2015	2.58	0.42	46	3.15	0.59	46	4.1%	-0.57 [-0.78, -0.36] 2015	
Li Y 2015	1.62	0.94	30	2.1	0.78	30	2.8%	-0.48 [-0.92, -0.04] 2015	
Gao Q 2016	2.8	0.9	46	4.1	1.1	46	2.9%	-1.30 [-1.71, -0.89] 2016	←
He W 2016	2.39	0.68	49	3.01	0.74	49	3.7%	-0.62 [-0.90, -0.34] 2016	
Xu T 2016	2.57	0.39	80	3.01	0.44	80	4.4%	-0.44 [-0.57, -0.31] 2016	
Jin ZK 2016	1.7	0.16	50	2.35	0.23	50	4.6%	-0.65 [-0.73, -0.57] 2016	
Li Y 2017	2.1	0.3	50	2.9	0.6	50	4.2%	-0.80 [-0.99, -0.61] 2017	
Lin CH 2017	2.73	0.75	49	2.82	0.63	49	3.7%	-0.09 [-0.36, 0.18] 2017	
Yang JS 2017	2.9	0.24	52	3.84	0.39	52	4.5%	-0.94 [-1.06, -0.82] 2017	-
Zhu J 2017	2.05	0.51	48	2.61	0.62	48	4.0%	-0.56 [-0.79, -0.33] 2017	
He WP 2017	2.6	0.6	29	3.3	0.9	29	3.0%	-0.70 [-1.09, -0.31] 2017	
Guo W 2017	2.71	1.29	50	4.03	1.44	50	2.3%	-1.32 [-1.86, -0.78] 2017	←
Cao TT 2017	2.03	0.29	49	3.27	0.5	49	4.3%	-1.24 [-1.40, -1.08] 2017	
Zhang XB 2018	2.19	0.43	50	2.72	0.7	50	4.0%	-0.53 [-0.76, -0.30] 2018	
Fang Y 2018	2.86	0.25	33	3.64	0.35	33	4.4%	-0.78 [-0.93, -0.63] 2018	
Subtotal (95% CI)			803			803	69.6%	-0.66 [-0.79, -0.53]	•
Heterogeneity: Tau ² =	0.06; Ch	ni² = 15	57.10, d	lf = 17 (P < 0.	00001);	l² = 89%		
Test for overall effect:	Z = 10.1	4 (P <	0.0000	1)					
3.1.2 STS vs. no drug	1								
Tao QH 2005	2.89	0.27	40	3.85	0.38	37	4.4%	-0.96 [-1.11, -0.81] 2005	
Guo LZ 2008	2.91	0.26	65	3.85	0.38	62	4.5%	-0.94 [-1.05, -0.83] 2008	-
Cheng ZY 2012	3.8	0.87	60	3.92	0.75	60	3.6%	-0.12 [-0.41, 0.17] 2012	
Wei Q 2014	2.33	0.59	60	3.96	1.37	60	3.1%	-1.63 [-2.01, -1.25] 2014	←
Gao Y 2014	2.01	0.82	30	2.27	0.99	30	2.7%	-0.26 [-0.72, 0.20] 2014	
Zhang XR 2014	2.58	0.42	48	3.15	0.59	48	4.1%	-0.57 [-0.77, -0.37] 2014	
Fan LY 2015	1.51	0.42	47	1.92	0.57	42	4.1%	-0.41 [-0.62, -0.20] 2015	
Zhao YL 2018	1.36	0.36	39	2.27	0.62	39	4.0%	-0.91 [-1.14, -0.68] 2018	
Subtotal (95% CI)			389			378	30.4%	-0.73 [-0.97, -0.49]	•
Heterogeneity: Tau ² =	0.10; Cł	ni² = 74	4.16, df	= 7 (P <	< 0.00	001); l²	= 91%		
	Z = 5.90) (P < (0.00001)					
Test for overall effect:			1192			1181	100.0%	-0.68 [-0.80, -0.57]	◆
Test for overall effect: . Total (95% CI)			1192						
	0.07; Ch	1i² = 24		lf = 25 (P < 0.	00001):	l ² = 90%		
Total (95% CI)			17.55, d		P < 0.0	00001);	l² = 90%		-1 -0.5 0 0.5 1
Total (95% CI) Heterogeneity: Tau ² =	Z = 11.8	81 (P <	47.55, d 0.0000)1)		,			-1 -0.5 0 0.5 1 Favours [STS] Favours [non-STS]
Total (95% CI) Heterogeneity: Tau ² = Test for overall effect:	Z = 11.8	81 (P <	47.55, d 0.0000)1)		,			

Sensitivity Analysis

Sensitivity analysis was conducted by excluding the studies from the analysis one by one to understand the impact of this on the results. The results showed that the pooled effects of STS on TC did not change substantially if a single studies were omitted (**Supplementary Material 2, Supplementary Figure S1**).

Trial Sequential Analysis

TSA was performed at the level of an overall 5% risk of type I error and a power of 20%. Regarding the results of TSA, shown in the **Supplementary Figure S2** (**Supplementary Material 2**). When the first study was included, the sample size had passed the traditional boundary value and TSA boundary value. After the third study was included, the sample size reached RIS value (RIS value = 220). Although it had reached the statistical threshold, the risk of methodological bias in the original study was high, which may affect the results. It may indicate that the evidence needed to reach a conclusion was sufficient and no further trials were needed.

Publication Bias

Figure 8 is a funnel diagram of the impact of STS on TC in patients with CHD, showing asymmetry, indicating possible publication bias. Begge's test and egger's test, respectively obtained z = -0.980 (p = 0.327) and t = -0.702 (p = 0.489), indicating that there was no publication bias in statistics. Although the probability of publication bias was statistically small, we still believed that there was a greater possibility of publication bias, because the published studies were all Chinese literature, and the positive results were easier to publish.

DISCUSSION

In this systematic review, twenty-seven RCTs with a total of 2,445 participants were included to assess the effect of STS on blood lipids in patients with CHD. It showed that STS could significantly reduce the levels of plasma TC, LDL-c and TG, and increase the levels of HDL-c. Through the subgroup analysis we

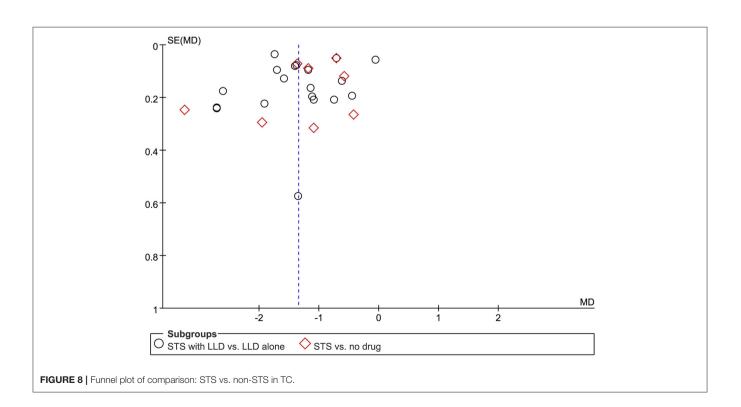
		STS			n-STS			Mean Difference		Mean Difference
tudy or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% Cl
.1.1 STS with LLD ve	s. LLD a	lone								
e HZ 2010	1.36	0.25	30	1.02	0.07	30	4.2%	0.34 [0.25, 0.43]	2010	
hang RX 2011		0.07	32	1.13	0.06	32	4.4%	-0.06 [-0.09, -0.03]	2011	-
uang X 2014	1.31	0.22	30	1.33	0.26	30	4.1%	-0.02 [-0.14, 0.10]		-
Vang FG 2015	3.12	0.29	46	2.28	0.62	46	3.8%	0.84 [0.64, 1.04]	2015	
Y 2015	1.8	0.27	30	1.22	0.33	30	4.0%	0.58 [0.43, 0.73]	2015	
(u T 2016	1.98	0.1	80	1.64	0.25	80	4.3%	0.34 [0.28, 0.40]	2016	-
n ZK 2016	2.32	0.17	50	1.72	0.27	50	4.3%	0.60 [0.51, 0.69]	2016	
ao Q 2016	1.9	0.4	46	1.5	0.5	46	3.9%	0.40 [0.21, 0.59]	2016	
uo W 2017	1.93	0.38	50	1.7	0.51	50	3.9%	0.23 [0.05, 0.41]	2017	
le WP 2017	1.9	0.5	29	1.7	0.5	29	3.5%	0.20 [-0.06, 0.46]	2017	
ao TT 2017	1.74	0.26	49	1.32	0.2	49	4.2%	0.42 [0.33, 0.51]	2017	-
ang JS 2017	1.12	0.22	52	1.11	0.21	52	4.3%	0.01 [-0.07, 0.09]	2017	+
hu J 2017	1.08	0.43	48	1.12	0.41	48	4.0%	-0.04 [-0.21, 0.13]	2017	-
i Y 2017	1.97	0.19	50	1.48	0.16	50	4.3%	0.49 [0.42, 0.56]	2017	-
in CH 2017	1.69	0.72	49	1.68	0.75	49	3.3%	0.01 [-0.28, 0.30]	2017	
hang XB 2018	1.69	0.47	50	1.32	0.35	50	4.0%	0.37 [0.21, 0.53]	2018	
ang Y 2018	1.17	0.14	33	1.1	0.25	33	4.2%	0.07 [-0.03, 0.17]	2018	—
bubtotal (95% CI) leterogeneity: Tau ² =				f = 16 (P < 0.0	754 00001);	68.8% I² = 97%	0.28 [0.15, 0.42]		•
eterogeneity: Tau ² = est for overall effect: 2	Z = 4.08		65.00, d	f = 16 (P < 0.(0.28 [0.15, 0.42]		•
eterogeneity: Tau ² = est for overall effect: 1.2 STS vs. no drug	Z = 4.08		65.00, d		P < 0.0			0.28 [0.15, 0.42] 0.03 [-0.06, 0.12]	2005	-
eterogeneity: Tau ² = est for overall effect: 1.2 STS vs. no drug ao QH 2005	Z = 4.08 I 1.13	8 (P < (65.00, d 0.0001)	1.1		00001);	l² = 97%			•
eterogeneity: Tau ² = est for overall effect: 2 .1.2 STS vs. no drug ao QH 2005 suo LZ 2008	Z = 4.08 I 1.13 1.12	6 (P < 0 0.16	65.00, d 0.0001) 40	1.1	0.23 0.23)0001); 37	l² = 97% 4.3%	0.03 [-0.06, 0.12]	2008	
eterogeneity: Tau ² = est for overall effect: 1.2 STS vs. no drug ao QH 2005 uo LZ 2008 heng ZY 2012	Z = 4.08 1.13 1.12 1.12	0.16 0.15	65.00, d 0.0001) 40 65	1.1 1.1	0.23 0.23 0.64	37 62	l ² = 97% 4.3% 4.3%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09]	2008 2012	
eterogeneity: Tau ² = est for overall effect: ; 1.2 STS vs. no drug ao QH 2005 uo LZ 2008 heng ZY 2012 ao Y 2014	Z = 4.08 1.13 1.12 1.12 1.22	0.16 0.15 0.45	65.00, d 0.0001) 40 65 60	1.1 1.1 1.11	0.23 0.23 0.64 1.02	37 62 60	² = 97% 4.3% 4.3% 3.8%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21]	2008 2012 2014	
eterogeneity: Tau ² = est for overall effect: 3 1.2 STS vs. no drug ao QH 2005 uo LZ 2008 heng ZY 2012 ao Y 2014 /ei Q 2014	Z = 4.08 1.13 1.12 1.12 1.22 1.06	0.16 0.15 0.45 0.24	65.00, d 0.0001) 40 65 60 30	1.1 1.1 1.11 1.48	0.23 0.23 0.64 1.02 0.08	37 62 60 30	² = 97% 4.3% 4.3% 3.8% 2.9%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21] -0.26 [-0.63, 0.11]	2008 2012 2014 2014	
	Z = 4.08 1.13 1.12 1.12 1.22 1.06 3.12	0.16 0.15 0.45 0.24 0.64	65.00, d 0.0001) 40 65 60 30 60	1.1 1.1 1.11 1.48 0.93 2.28	0.23 0.23 0.64 1.02 0.08	00001); 37 62 60 30 60	² = 97% 4.3% 4.3% 3.8% 2.9% 4.0%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21] -0.26 [-0.63, 0.11] 0.13 [-0.03, 0.29]	2008 2012 2014 2014 2014	
eterogeneity: Tau ² = est for overall effect: : 1.2 STS vs. no drug ao QH 2005 uo LZ 2008 heng ZY 2012 ao Y 2014 /ei Q 2014 hang XR 2014 an LY 2015	Z = 4.08 1.13 1.12 1.12 1.22 1.06 3.12 1.83	0.16 0.15 0.45 0.24 0.64 0.29	65.00, d 0.0001) 40 65 60 30 60 48	1.1 1.1 1.48 0.93 2.28 1.55	0.23 0.23 0.64 1.02 0.08 0.62	37 62 60 30 60 48	 I² = 97% 4.3% 4.3% 3.8% 2.9% 4.0% 3.8% 	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21] -0.26 [-0.63, 0.11] 0.13 [-0.03, 0.29] 0.84 [0.65, 1.03]	2008 2012 2014 2014 2014 2014 2015	
eterogeneity: Tau ² = est for overall effect: 2 ao QH 2005 uo LZ 2008 heng ZY 2012 ao Y 2014 fei Q 2014 hang XR 2014 an LY 2015 hao YL 2018 ubtotal (95% CI)	Z = 4.08 1.13 1.12 1.12 1.22 1.06 3.12 1.83 2.03	0.16 0.15 0.45 0.24 0.64 0.29 0.53 0.21	65.00, d 0.0001) 40 65 60 30 60 48 47 39 389	1.1 1.1 1.48 0.93 2.28 1.55 1.39	0.23 0.23 0.64 1.02 0.08 0.62 0.32 0.33	37 62 60 30 60 48 42 39 378	² = 97% 4.3% 4.3% 3.8% 2.9% 4.0% 3.8% 3.9% 4.1% 31.2%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21] -0.26 [-0.63, 0.11] 0.13 [-0.03, 0.29] 0.84 [0.65, 1.03] 0.28 [0.10, 0.46]	2008 2012 2014 2014 2014 2014 2015	
eterogeneity: Tau ² = est for overall effect: 2 ao QH 2005 uo LZ 2008 heng ZY 2012 ao Y 2014 /ei Q 2014 hang XR 2014 an LY 2015 hao YL 2018 ubtotal (95% CI) eterogeneity: Tau ² =	Z = 4.08 1.13 1.12 1.12 1.22 1.06 3.12 1.83 2.03	0.16 0.15 0.45 0.24 0.64 0.29 0.53 0.21	65.00, d 0.0001) 40 65 60 30 60 48 47 39 389 39.47, d	1.1 1.1 1.48 0.93 2.28 1.55 1.39	0.23 0.23 0.64 1.02 0.08 0.62 0.32 0.33	37 62 60 30 60 48 42 39 378	² = 97% 4.3% 4.3% 3.8% 2.9% 4.0% 3.8% 3.9% 4.1% 31.2%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21] -0.26 [-0.63, 0.11] 0.13 [-0.03, 0.29] 0.84 [0.65, 1.03] 0.28 [0.10, 0.46] 0.64 [0.52, 0.76]	2008 2012 2014 2014 2014 2014 2015	
eterogeneity: Tau ² = est for overall effect: 2 ao QH 2005 uo LZ 2008 heng ZY 2012 ao Y 2014 /ei Q 2014 hang XR 2014 an LY 2015 hao YL 2018 ubtotal (95% CI) eterogeneity: Tau ² =	Z = 4.08 1.13 1.12 1.12 1.22 1.06 3.12 1.83 2.03	0.16 0.15 0.45 0.24 0.64 0.29 0.53 0.21	65.00, d 0.0001) 40 65 60 30 60 48 47 39 389 39.47, d	1.1 1.1 1.48 0.93 2.28 1.55 1.39	0.23 0.23 0.64 1.02 0.08 0.62 0.32 0.33	37 62 60 30 60 48 42 39 378	² = 97% 4.3% 4.3% 3.8% 2.9% 4.0% 3.8% 3.9% 4.1% 31.2%	0.03 [-0.06, 0.12] 0.02 [-0.05, 0.09] 0.01 [-0.19, 0.21] -0.26 [-0.63, 0.11] 0.13 [-0.03, 0.29] 0.84 [0.65, 1.03] 0.28 [0.10, 0.46] 0.64 [0.52, 0.76]	2008 2012 2014 2014 2014 2014 2015	
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found that the effect of STS was significantly better than control no matter when STS as monotherapy or as adjunctive therapy. At the same time, STS has add-on effect when used in combination with conventional lipid-lowering drugs. In terms of AEs, there was no significant difference between STS group and non-STS group. STS treatment may be safe and effective to reduce the blood lipid level of patients with CHD and has great potential as an assistant drug to reduce blood lipid. However, due to the low overall quality of RCTs at present, the above conclusions need to be verified in larger samples of high-quality research.

Previous studies have mostly focused on the effects of STS on the incidence of major cardiovascular events and inflammatory factors in patients with CHD (64, 65). This systematic review focused for the first time on the effect of STS on blood lipid in patients with CHD and evaluated its safety. At present, there are few studies on the mechanism of STS improving blood lipid level. Researches have shown that STS can effectively reduce LDL-c oxidation in serum and LDL-c level in plasma, which may be related to its peroxy radical scavenging and LDL- c binding activity (66). Tanshinone IIA could increase LDLR levels and LDL uptake by cells through the downregulation of PCSK9 expression in HepG2 cells (67). Tanshinone IIA regulated the SREBP2/PCSK9 pathway, increased the levels of HDL, and decreased lipid deposition in the livers of hyperlipidemic rats (68). On the other hand, tanshinone IIA could lower blood lipid level by downregulating the expression of lipogenic enzymes, including FASN, ACC1, and SCD1, through suppression of SREBP1 expression and LXRα-mediated transcriptional activation, resulting in reducing lipogenesis and lipid accumulation in the HepG2 cells (69). Furthermore, STS can not only affect the changes of AMPK/ACC/CPT1 signal involved in lipid metabolism to inhibit fatty acid beta oxidation during ischemia, but also maintain circulating lipid levels, including TC, TG and free fatty acids (FFA) (70). Moreover, tanshinone IIA is considered to be an inhibitor of apoB secretion and TG secretion in hepatocytes (71) and a natural antagonist of peroxisome proliferator-activated receptor- γ (PPAR γ), which is a member of the nuclear receptor superfamily of ligand-activated

	STS		non-S	TS		Risk Ratio		Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	Year	M-H, Fixed, 95% Cl
5.1.1 STS with LLD v	s. LLD alo	ne						
Ge HZ 2010	0	30	0	30		Not estimable	2010	
luang X 2014	0	30	0	30		Not estimable	2014	
Ku T 2016	0	80	0	80		Not estimable	2016	
He W 2016	0	49	0	49		Not estimable	2016	
Cao TT 2017	0	49	0	49		Not estimable	2017	
Zhu J 2017	5	48	4	48	22.2%	1.25 [0.36, 4.37]	2017	_
Zhang XB 2018	11	50	9	50	49.9%	1.22 [0.56, 2.69]	2018	
ang Y 2018	1	33	0	33	2.8%	3.00 [0.13, 71.07]	2018	
Yan XL 2018	2	36	3	36	16.6%	0.67 [0.12, 3.75]	2018	
Subtotal (95% CI)		405		405	91.6%	1.18 [0.64, 2.17]		•
Total events	19		16					
Heterogeneity: Chi ² =	0.77, df = 3	3 (P = 0	0.86); l ² =	0%				
Test for overall effect:	Z = 0.54 (F	P = 0.5	9)					
4.0.070								
5.1.2 STS vs. no dru	•							
Гао QH 2005	1	40	0	37	2.9%	2.78 [0.12, 66.20]		
Cheng ZY 2012	2	60	1	60	5.5%	2.00 [0.19, 21.47]		-
Zhao YL 2018	0	39	0	39	0.40/	Not estimable	2018	
Subtotal (95% CI)		139		136	8.4%	2.27 [0.34, 15.09]		
Fotal events	3		1					
Heterogeneity: Chi ² =		•		0%				
0,	Z = 0.85 (F	P = 0.4	0)					
Test for overall effect:						1.27 [0.72, 2.27]		•
0,		544		541	100.0%	1.21 [0.12, 2.21]		•
Test for overall effect:	22	544	17	541	100.0%	1.27 [0.72, 2.27]		
Fest for overall effect: Fotal (95% CI)					100.0%	1.27 [0.72, 2.27]		
Fest for overall effect: Fotal (95% CI) Fotal events	1.20, df = 5	5 (P = (0.94); l ² =		100.0%	1.27 [0.72, 2.27]		0.001 0.1 1 10 100 Favours [STS] Favours [non-STS]

FIGURE 7 | Forest plot of comparison: STS vs. non-STS in AEs.



transcription factors and major transcription factor that regulates the gene expression of adipocyte differentiation, lipogenesis, and glucose metabolism (72). Tanshinone IIA can reduce fat mass and weight, improve glucose tolerance, reduce the ratio of LDL-c to HDL-c without changing food intake (73). In addition, STS can improve the PPAR-alpha down-regulation (74) caused by myocardial infarction and restore insulin signaling pathway (75), thereby improving the hyperlipidemia accumulation of ischemic heart (76).

Comparison With Previous Studies

Only one systematic review and meta-analysis on the efficacy and safety of Traditional Chinese Medicine injection for blood lipid in patients with unstable angina pectoris (77) was conducted before, in which STS was analyzed, but the results were not described in detail. Our meta-analysis has some strength in comparison with the previous systematic review. First, several newly-published and well-conducted trials were included in our meta-analysis. Second, in the meta-analysis, we performed subgroup analysis from different aspects to make the results more stable.

The results of our meta-analysis are similar to those of the previous review regarding TC and HDL-c when STS was used in combination with LLDs compared to LLDs alone, but they differ with respect to TG and LDL-c. Our results indicated that, compared with LLDs alone, the STS combined with LLDs significantly lowered LDL-c and TG levels. Additionally, we found that STS as monotherapy significantly reduced TG, TC, and LDL-c levels and increased HDL-c levels compared to the controls. Our study carried out a more comprehensive, up-to-date, and PRISMA-compliant systematic review. Meta-analysis provides more reliable clinical evidence.

Limitation

Although the systematic evaluation shows that STS has a good effect on blood lipid level in patients with CHD, the results are not convincing. Firstly, high-quality, multi-center, large-sample, double-blind RCTs were lacked in the 27 trials included. The 27 trials were not registered in advance, and the relevant test schemes were not published. Secondly, most of the trials did not specify the specific details of random mode, allocation concealment, blind method and so on. This greatly weakens the credibility of the evidence. Finally, the greater heterogeneity and publication bias of the results require us to interpret the final results carefully.

Implications for Research

The systematic review provides a small amount of evidence for STS in improving the prognosis of patients with CHD. For hospitalized patients with CHD, especially those with statin intolerance, clinicians may consider STS as a complementary alternative therapy. However, the toxicological and pharmacological mechanism of the optimal dose and duration of STS needs further study. Especially for traditional Chinese medicine extracts, we should pay more attention to their adverse reactions to patients. In addition, STS is currently only intravenous preparations, which is obviously not suitable for long-term use of patients. Finding optimum route of administration for STS is another problem we will face in the future. Finally, the design of RCTs and the reporting of clinical study results should be carried out in strict accordance with the requirements of the CONSORT 2010 statement to ensure the scientific quality and rigor of studies (78).

CONCLUSION

STS can safely and effectively reduce plasma TC, TG and LDL-c levels in patients with CHD, and improve plasma HDL-c levels. However, these findings require careful recommendation due to the low overall quality of RCTs at present. More multi-center, randomized, double-blind, placebo-controlled trials which are designed follow the CONSORT 2010 guideline are needed.

DATA AVAILABILITY STATEMENT

The original contributions presented in the study are included in the article/**Supplementary Material**, further inquiries can be directed to the corresponding authors.

AUTHOR CONTRIBUTIONS

QL, HZ, and YZ conceived and drafted this systematic review and registered the protocol at PROSPERO. HZ and QL developed the search strategy and conducted the literature research, study selection, data extraction and risk of rias assessment, and contributed to manuscript drafting. WH, DW, GP, WP, ZW, and XR interpreted the evidence from methodological and clinical perspective. XW oversaw the conduct of the study. All authors have read, critically reviewed, and approved the final manuscript.

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SUPPLEMENTARY MATERIAL

The Supplementary Material for this article can be found online at: https://www.frontiersin.org/articles/10.3389/fcvm. 2021.770746/full#supplementary-material

REFERENCES

- Roffi M, Patrono C, Collet JP, Mueller C, Valgimigli M, Andreotti F, et al. 2015 ESC guidelines for the management of acute coronary syndromes in patients presenting without persistent ST-segment elevation: task force for the management of acute coronary syndromes in patients presenting without persistent ST-segment elevation of the European society of cardiology (ESC). *Eur Heart J.* (2016) 37:267–315. doi: 10.1093/eurheartj/ehv320
- Knuuti J, Wijns W, Saraste A, Capodanno D, Barbato E, Funck BC, et al. 2019 ESC guidelines for the diagnosis and management of chronic coronary syndromes. *Eur Heart J.* (2020) 41:407–77. doi: 10.1093/Eurheartj/Ehz425
- Collins R, Reith C, Emberson J, Armitage J, Baigent C, Blackwell L, et al. Interpretation of the evidence for the efficacy and safety of statin therapy. *Lancet.* (2016) 388:2532–61. doi: 10.1016/S0140-6736(16)31357-5
- Adhyaru BB, Jacobson TA. Safety and efficacy of statin therapy. Nat Rev Cardiol. (2018) 15:757–69. doi: 10.1038/s41569-018-0098-5
- Gordon RY, Becker DJ. The role of red yeast rice for the physician. Curr Atheroscler Rep. (2011) 13:73–80. doi: 10.1007/s11883-010-0145-0
- Ju J, Li J, Lin Q, Xu H. Efficacy and safety of berberine for dyslipidaemias: a systematic review and meta-analysis of randomized clinical trials. *Phytomedicine*. (2018) 50:25–34. doi: 10.1016/j.phymed.2018.09.212
- Ruscica M, Gomaraschi M, Mombelli G, Macchi C, Bosisio R, Pazzucconi F, et al. Nutraceutical approach to moderate cardiometabolic risk: results of a randomized, double-blind and crossover study with armolipid plus. *J Clin Lipidol.* (2014) 8:61–8. doi: 10.1016/j.jacl.2013.11.003
- Chen JT, Wesley R, Shamburek RD, Pucino F, Csako G. Metaanalysis of natural therapies for hyperlipidemia: plant sterols and stanols versus policosanol. *Pharmacotherapy.* (2005) 25:171– 83. doi: 10.1592/phco.25.2.171.56942
- Yoshida H, Yanai H, Ito K, Tomono Y, Koikeda T, Tsukahara H, et al. Administration of natural astaxanthin increases serum HDL-cholesterol and adiponectin in subjects with mild hyperlipidemia. *Atherosclerosis.* (2010) 209:520–3. doi: 10.1016/j.atherosclerosis.2009.10.012
- Shang QH, Xu H, Huang Li. Tanshinone IIA: a promising natural cardioprotective agent. *Evid Based Complement Alternat Med.* (2012) 2012: 716459. doi: 10.1155/2012/716459
- Wang QL, Deng XJ, Li XQ, Chen XT. Effect of sodium tanshinone injection on CRP and D-dimer level in patients with unstable angina. *J Chin Med.* (2007) 39:16–7. doi: 10.13457 /j.Cnki.Jncm.2007.07.011
- Gao S, Liu Z, Li H, Little PJ, Liu P, Xu S. Cardiovascular actions and therapeutic potential of tanshinone IIA. *Atherosclerosis.* (2012) 220:3– 10. doi: 10.1016/j.atherosclerosis.2011.06.041
- Li S, Jiao Y, Wang H, Shang Q, Lu F, Huang L, et al. Sodium tanshinone IIA sulfate adjunct therapy reduces high-sensitivity C-reactive protein level in coronary artery disease patients: a randomized controlled trial. *Sci Rep.* (2017) 7:17451. doi: 10.1038/s41598-017-16980-4
- Yu ML, Li SM, Gao X, Li JG, Xu H, Chen KJ. Sodium tanshinone II A sulfonate for coronary heart disease: a systematic review of randomized controlled trials. *Chin J Integr Med.* (2020) 26:219–26. doi: 10.1007/s11655-018-2556-7
- Higgins JPT, Green S. Cochrane Handbook for Systematic Reviews of Interventions Version 5.1.0. New Jersey: Wiley Blackwell (2011). p. 649.
- Liberati A, Altman DG, Tetzlaff J, Mulrow C, Gøtzsche PC, Ioannidis JP, et al. The PRISMA statement for reporting systematic reviews and meta-analyses of studies that evaluate healthcare interventions: explanation and elaboration. *BMJ*. (2009) 339:B2700. doi: 10.1136/bmj.b2700
- Laird NM, Mosteller F. Some statistical methods for combining experimental results. Int J Technol Assess Health Care. (1990) 6:5–30. doi: 10.1017/S0266462300008916
- Thorlund K, Engstrom J, Wetterslev J, Brok J, Imberger G, Gluud C. User Manual for Trial Sequential Analysis (TSA). Copenhagen: Copenhagen Trial Unit (2011). p. 115.
- Egger M, Davey SG, Schneider M, Minder C. Bias in metaanalysis detected by a simple graphical test. *BMJ*. (1997) 315:629–34. doi: 10.1136/bmj.315.7109.629
- Lin JC. To analyze the efficacy and safety of sodium tanshinone IIA sulfonate injection in the treatment of angina pectoris of coronary heart disease. *For all Health.* (2014) 8:158–9.

- Liang CH, Xu CL, Zhang YQ, Yan JJ, Zhou BJ, Zhu T, et al. Clinical significance of change of plasma CRP, UA, lipids and bilirubin and role of tanshinone II A sulfonate in patients with angina. J Clin Med. (2012) 16:25–7.
- Li SD. Clinical Analysis of 39 Unstable Angina Cases. China Tropical Medicine (2008). p. 2276–7.
- Ren S, Li BX. The effect of tan on indicators and blood lipid levels in patients with coronary heart disease. *Journal of LiaoNing University of TCM*. (2015) 17:211–2.
- Wei XX. Therapeutic effect of sodium tanshinone IIA sulfonate injection on patients with coronary heart disease. *Chinese Journal of Clinical Rational Drug Use.* (2018) 11:33–4. doi: 10.15887/j.Cnki.13-1389/r.2018.35.016
- Cheng JD. Therapeutic effect of sodium tanshinone IIA sulfonate combined with low molecular weight heparin calcium on unstable angina pectoris. *Lingnan Journal of Emergency Medicine*. (2007) 12:225–6.
- Wang G, Tang F, Zhang P, Yang LN, Li F. Observation of therapeutic effects of tanshinone IIA on ischemic heart disease in diabetic patients receiving maintenance hemodialysis. *Medical Journal of National Defending Forces in Southwest China*. (2011) 21:256–8. doi: 10.3969/j.Issn.1004–0188.2011.03.008
- Yu HP, Zhao CJ, Wang T, Jia RY, Pu YH. Clinical study of cinepazide maleate combined with tanshinone in the treatment of unstable angina pectoris. *Journal of Hebei Medical University.* (2010) 31:202–4. doi: 10.3969/J.Issn.1007-3205.2010.02.029
- Cheng J. Observation on the efficacy of sodium tanshinone IIA sulfonate injection in the treatment of stable angina pectoris of coronary heart disease. *Clinical Medicine*. (2014) 34:48–9.
- Lv XL. Therapeutic effect of "sodium tanshinone II A sulfonate + simvastatin" on angina pectoris in patients with coronary heart disease. J Front Med. (2017) 7:154.
- Wang HT. Clinical study of tanshinone IIA sulfonate sodium combined with atorvastatin in the treatment of coronary heart disease. *Clinical Medicine*. (2017) 37:123–24. doi: 10.19528/j.Issn.1003-3548.2017.06.056
- Zhang GZ. Clinical observation of tanshinone IIA sodium sulfonate injection in the treatment of angina pectoris of coronary heart disease (blood stasis syndrome). *Modern Diagnosis and Treatment*. (2015) 26:2680–1.
- 32. Xia SJ, Yin YG, Fang Y. Atorvastatin combined with danshen ketone IIA sulfonic acid sodium injection in the treatment of angina pectoris of coronary heart disease. *Hebei Medicine*. (2016) 22:1067–70. doi: 10.3969/j.Issn.1006-6233.2016.07.05
- Chen LQ, Yan WJ, Yao YQ. Sodium tanshinone IIA sulfonic acid in patients with coronary heart disease, myocardial ischemia of TC, TG, LDL-C Level. *Contemporary Medicine*. (2017) 23:66–9. doi: 10.3969/j.Issn.1009-4393.2017.35.026
- Peng J. Effect of nuoxinkang on inflammatory factors in patients with acute coronary syndrome. *Smart Healthcare*. (2018) 04:25–7. doi: 10.19335/j.Cnki.2096-1219.2018.31.011
- Wei WJ. Clinical observation of tanshinone IIA sodium sulfonate injection combined with tirofiban in the treatment of acute coronary syndrome. *Chinese Journal of Integrative Medicine on Cardio-Cerebrovascular Disease.* (2019) 17:1693–5. doi: 10.12102/j.Issn.1672-1349.2019.11.026
- Zeng W, Wei Y. Clinical efficacy of sodium tanshinone IIA sulfonate combined with nicorandil in the treatment of stable angina pectoris of coronary heart disease. *Capital Food Medicine*. (2019) 26:73.
- Zhang RX, Wang MW, Wang SW. Effects of tanshinone on the level of adiponectin in patients with acute coronary syndrome. *China Medical Engineering*. (2011) 19:39–41.
- Li Y, Fan XZ, Han YS, Shen L. Clinical study of tanshinone IIA sodium sulfonate on blood lipid and inflammatory factor in patients with acute coronary syndrome. *Journal of Perscription Drug.* (2015) 13:75–6.
- Lin CH, Liu YJ, Gu X. Influence of sodium tanshinone IIA sulfonate on oxidative stress level in patients with acute coronary syndrome. *Chin J Evid Based Cardiovasc Med.* (2017) 01:63–5. doi: 10.3969/j.Issn.1674-4055.2017.01.17
- Zhang XR, Wang LJ, Liu HJ, Dong W, Guo YB. The impact of tan treatment on haemodynamics and blood lipid of acute myocardial infarction patients. *Journal of Guiyang Medical College*. (2014) 39:379– 82. doi: 10.19367/j.Cnki.1000-2707.2014.03.027

- 41. Fan LY, Yang J. Impact of tanshinone IIA sulfonic acid natrium on blood lipid metabolism, serum inflammatory cytokines levels and cardiac function in patients with acute myocardial infarction. *Practical Journal of Cardiac Cerebral Pneumal and Vascular Disease*. (2015) 23:14– 7. doi: 10.3969/j.Issn.1008-5971.2015.09.005
- Wang FG, Zheng XD. Effect of tanshinone injection on hemorheology and platelet aggregation in patients with acute myocardial infarction. *Chin Med.* (2015) 10:149–50.
- Ge HZ, Cheng YN, Qin HY. Clinical observation of tanshinone II A sulfonate sodium in the treatment of unstable angina pectoris. *Chin J Diffic and Compl Cas.* (2010) 09:201–2. doi: 10.3969/j.Issn.1671-6450.2010.03.021
- Jin ZK. Observation on the efficacy of sodium tanshinone IIA sulfonate in the treatment of 100 cases of unstable angina pectoris. *Guide of China Medicine*. (2016) 14:177–8. doi: 10.15912/j.Cnki.Gocm.2016.23.146
- 45. Zhang XB. Effects of sodium tanshinone IIA sulfonate injection on serum uric acid, CRP and hemorheology in patients with coronary heart disease complicated with hyperuricemia. *Modern Journal of Integrated Traditional Chinese and Western Medicine*. (2018) 27:1530–2.
- Zhao YL, Hou W, Yan YX, Wang HX, Shang R. Clinical trial of tanshinone IIA sulfonate injection combined with aspirin tablets in the treatment of myocardial ischemia. *Chinese J Pharmacol.* (2018) 34:1276– 8. doi: 10.13699/j.Cnki.1001-6821.2018.11.004
- Tao XH, Jia LW. Therapeutic action of tanshinone IIA sulfonic acid natrium injection on myocardial ischemia of coronary artery disease. *Modern Journal* of Integrated Traditional Chinese and Western Medicine. (2005) 14:2507–9.
- Guo LZ, Sun F. Effect of sodium tanshinone IIA sulfonate on angina pectoris in patients with coronary heart disease. *China & Foreign Medical Treatment*. (2008) 26:62–3. doi: 10.16662/j.Cnki.1674-0742.2008. 26.021
- Cheng ZY, Zhang R. Observation of the effect of sulfotanshinone IIA sodium sulfonate injection on the treatment of coronary heart disease and angina pector. *Chin J Pharmacoepidemiol.* (2012) 21:313–5.
- Wei Q, Shen XJ. Effects of tanshinone IIA sulfonate on clinical symptoms and vascular aging factor in angina pectoris of coronary heart disease patients. *Chinese Journal of Basic Medicine in Traditional Chinese Medicine*. (2014) 20:783–5.
- Gao Y, Liu GB, Wang J. The effect of tanshinone combined with nitrate for unstable angina pectoris. *Proceeding of Clinical Medicine*. (2014) 23:258–60.
- Gao Q, Bo DN, Liu YY. Curative effect of atorvastatin combining tanshinone IIA sodium injection on angina pectoris of coronary heart disease. *Chin J Evid Based Cardiovasc Med.* (2016) 08:177–9. doi: 10.3969/j.Issn.1674-4055.2016.02.14
- 53. Xu T, Zhang Z, Wang JP, Tian YF, Liu X. Improvement of blood lipid, blood pressure and hemodynamics in patients with coronary heart disease by sodium tanshinone IIA sulfonate combined with atorvastatin. *Hebei Medical Journal.* (2016) 38:351–3. doi: 10.3969/j.Issn.1002-7386.2016.03.008
- He WP, Wang YJ, Wang HH. Effect of tanshinone II A sodium sulfonate injection on coronary atherosclerotic heart disease. *Pratical Clinical Medicine*. (2017) 18:20–1. doi: 10.13764/j.Cnki.Lcsy.2017.07.008
- 55. Li Y, Chen ZF, Zhang KS, Li C. Effect of tanshinone IIA sodium sulfonate combined with atorvastatin on hemodynamics in patients with coronary heart disease. *Clinical Pharmacy.* (2017) 12:89–91. doi: 10.12010/j.Issn.1673-5846.2017.05.031
- Guo W. Clinical observation of atorvastatin combined with tanshinone IIA sodium sulfonate injection in the treatment of coronary heart disease angina pectoris. *Medical Information*. (2017) 30:123–4.
- He W, Shi YL, Zheng CG. Clinical study of tanshinone IIA sodium sulfonate combined with simvastatin in the treatment of patients with angina of coronary heart disease. *Chin J Clin Pharmacol.* (2016) 32:393– 5. doi: 10.13699/j.Cnki.1001-6821.2016.05.003
- Yan XL. Effect of sodium tanshinone IIA sulfonate on blood lipid and cardiac function in patients with coronary heart disease and angina pectoris. *Henan Medical Research.* (2018) 27:1311–2. doi: 10.3969/j.Issn.1004-437X.2018.07.095
- Huang X. Effect of tanshinone IIA sodium sulfonate injection on clinical efficacy and blood lipid level of chronic stable angina pectoris. *Chinese Journal* of *Traditional Medical Science and Technology*. (2014) 21:71.

- Yang JS. Observation of curative effect and clinical index of sodium tanshinone IIA sulfonate in treating angina pectoris of coronary heart disease. *The Journal of Medical Theory and Practice*. (2017) 30: 659– 60. doi: 10.19381/j.Issn.1001-7585.2017.05.015
- Cao TT, Xu HL, He Y. Therapeutic effect of tanshinone IIA-sulfonic sodium on CHD patients and its influence on hemorheology, cytokines and blood lipid levels. *Chin J Cardiovasc Rehabil Med.* (2017) 27:104– 7. doi: 10.3969/j.Issn.1008-0074.2017.01.28
- 62. Zhu J, Dong W, Lin JH. Clinical effect of tanshinone IIA sulfonic acid natrium combined with nicorandil in treating coronary heart disease patients with stable angina pectoris and the impact on blood lipids and serum homocysteine level. *Practical Journal of Cardiac Cerebral Pneumal and Vascular Disease*. (2017) 25:107–10. doi: 10.3969/j.Issn.1008-5971.2017.02.028
- 63. Fang Y, Yang KP, Jin D. Effects of sodium tanshinone IIA sulfonate on serum metabolic factor, vascular endothelial growth factor, serum brain natriuretic peptide precursor and soluble signaling factor 4D in patients with coronary heart disease. *Modern Journal of Integrated Traditional Chinese and Western Medicine*. (2018) 27:3964–7. doi: 10.3969/j.Issn.1008-8849.2018.35.025
- Wen JX, Chang YM, Huo SS, Li WY, Huang HL, Gao YH, et al. Tanshinone IIA attenuates atherosclerosis *via* inhibiting NLRP3 inflammasome activation. *Aging*. (2021) 13:910–32. doi: 10.18632/aging.202202
- 65. Zeng H, Wang LL, Zhang JW, Pan T, Yu YH, Lu JX, et al. Activated PKB/GSK-3 β synergizes with PKC-δ signaling in attenuating myocardial ischemia/reperfusion injury via potentiation of NRF2 activity: therapeutic efficacy of dihydrotanshinone-I. Acta Pharm Sin B. (2021) 11:71–88. doi: 10.1016/j.apsb.2020.09.006
- Niu XL, Ichimori K, Yang X, Hirota Y, Hoshiai K, Li M, et al. Tanshinone IIA inhibits low density lipoprotein oxidation *in vitro. Free Radic Res.* (2000) 33:305–12. doi: 10.1080/10715760000301471
- Chen HC, Chen PY, Wu MJ, Tai MH, Yen JH. Tanshinone IIA modulates low density lipoprotein uptake via down-regulation of pcsk9 gene expression in HepG2 cells. *PLoS ONE*. (2016) 11:E0162414. doi: 10.1371/journal.pone.0162414
- 68. Jia LQ, Song N, Yang GL, Ma YX, Li XT, Lu R, et al. Effects of tanshinone IIA on the modulation of miR-33a and the SREBP-2/Pcsk9 signaling pathway in hyperlipidemic rats. *Mol Med Rep.* (2016) 13:4627– 35. doi: 10.3892/mmr.2016.5133
- 69. Gao WY, Chen PY, Hsu HJ, Lin CY, Wu MJ, Yen JH. Tanshinone IIA downregul-ates lipogenic gene expression and attenuates lipid accumulation through the modulation of LXRα/SREBP1 pathway in HepG2 cells. *Biomedicines*. (2021) 9:326. doi: 10.3390/biomedicines9030326
- Wei B, You MG, Ling JJ, Wei LL, Wang K, Li WW, et al. Regulation of antioxidant system, lipids and fatty acid β-oxidation contributes to the cardioprotective effect of sodium tanshinone IIA sulphonate in isoproterenolinduced myocardial infarction in rats. *Atherosclerosis.* (2013) 230:148– 56. doi: 10.1016/j.atherosclerosis.2013.07.005
- Kang YJ, Jin UH, Chang HW, Son JK, Lee SH, Son KH, et al. Inhibition of microsomal triglyceride transfer protein expression and atherogenic risk factor apolipoprotein B100 secretion by tanshinone IIA in HepG2 Cells. *Phytotherapy Research*. (2008) 22:1640–5. doi: 10.1002/ptr.2542
- 72. Evans RM, Barish GD, Wang YX. PPARs and the complex journey to obesity. *Nat Med.* (2004) 10:355–61. doi: 10.1038/nm1025
- 73. Gong Z, Huang C, Sheng X, Zhang Y, Li Q, Wang MW, et al. The role of tanshinone IIA in the treatment of obesity through peroxisome proliferator-activated receptor gamma antagonism. *Endocrinology*. (2009) 150:104–13. doi: 10.1210/en.2008-0322
- 74. Gilde AJ, van DLKA, Willemsen PH, Chinetti G, van DLFR, van DVGJ, et al. Peroxisome proliferator-activated receptor (PPAR) alpha and PPARbeta/Delta, but not PPARgamma, modulate the expression of genes involved in cardiac lipid metabolism. *Circ Res.* (2003) 92:518– 24. doi: 10.1161/01.RES.0000060700.55247.7C
- Hu QH, Zhang X, Pan Y, Li YC, Kong LD. Allopurinol, quercetin and rutin ameliorate renal NLRP3 inflammasome activation and lipid accumulation in fructose-fed rats. *Biochem Pharmacol.* (2012) 84:113– 25. doi: 10.1016/j.bcp.2012.03.005
- 76. Hu Q, Wei B, Wei L, Hua K, Yu X, Li H, et al. Sodium tanshinone IIA sulfonate ameliorates ischemia-induced myocardial inflammation and lipid

accumulation in beagle dogs through NLRP3 inflammasome. *Int J Cardiol.* (2015) 196:183–92. doi: 10.1016/j.ijcard.2015.05.152

- Yin C, Cheng J, Sun H, Li J. Effect of traditional chinese medicine injection for unstable angina pectoris on blood lipid level: systematic review and network meta-analysis. *Journal of Liaoning University of TCM*. (2021) 23:179–89.
- Moher D, Liberati A, Tetzlaff J, Altman DG, PRISMA Group. Preferred reporting items for systematic reviews and meta-analyses: the PRISMA statement. *BMJ*. (2009) 339:B2535. doi: 10.1136/bm j.b2535

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