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## **OPEN** Prognostic role of Gli1 expression in solid malignancies: a meta-analysis

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Gli1 is a downstream transcriptional factor of Sonic hedgehog pathway in mammalians, and has been recognized as a proliferative indicator of carcinogenesis. However, its actual role in prognosis among solid malignancies remains unclear. Therefore we performed this meta-analysis aiming to discover the correlation between Gli1 positivity and clinical prognosis in patients suffering from diverse carcinomas. A total of 39 studies containing 4496 cases were selected into our quantitative analysis via electronic database search. Original data of 3-year, 5-year, 10-year overall survival and disease-free survival were extracted and calculated using odds ratio and Mantel-Haenszel model. Subgroup analysis was also conducted to clarify the possible confounding factors. P < 0.05 was considered significant in statistics. Gli1 redundancy was associated with worse 3-year, 5-year, 10-year overall survival and diseasefree survival in solid malignancies. Different source regions, sample-size, mean-age and detection approaches had no impact on the negative prognostic effect of Gli1 over-expression. Nevertheless, stratified by cancer type and subcellular localization, cytoplasmic Gli1 expression and Gli1 positivity in intracranial tumors was not correlated to poorer 3-year and 5-year prognosis. The over-expression of Gli1 is a credible indicator of poorer prognosis in most of solid malignancies, irrespective of intracranial tumors.

As a life-threatening disorder and costly burden of healthcare system, cancer has drawn tremendous academic attention to the mechanisms towards its formation and uncontrollable spread. To date, accumulating evidence has disclosed the close kinship between normal embryonic development processes and malignant proliferation, both of which share similar cellular mechanisms and morphological alterations<sup>1</sup>. Including the conserved Sonic hedgehog (Shh) signaling, aberrant activation of pivotal developmental pathways generally facilitates carcinogenesis through a direct signal transduction in a variety of tumors, implicating a potential therapeutic targeted strategy for future medication<sup>2,3</sup>.

It is well documented that the Sonic hedgehog signaling initially emerged as an essential pathway involved in tissue growth and patterning during the fetal development<sup>4</sup>. The pathway has been strictly inhibited in mature tissues to ensure the physiological function and prevent abnormal proliferation. Following studies subsequently confirm its out-of-control plays a crucial role within the tumor origination and metastatic dissemination, identified as a tumor-stimulating pathway in diversified malignancies<sup>5</sup>. The transduction of this relatively conserved pathway is mediated by interactions between Shh, Patched and Smoothened proteins. As long as ligand Shh binds to trans-membrane protein Patched, the normally activity-inhibited Smoothened protein is then released to activate the translocation of cytoplasmic Gli1 into nucleus. Gli1 acts as a transcriptional factor downstream of the Shh pathway, whose nuclear translocation serves as a hallmark indicator of Shh pathway activation, culminating in up-regulation of target genes especially certain proliferative oncogenes<sup>6</sup>. Given the tumorigenic feature of Gli1, a specific targeted therapy may be prone to benefit patients with better survival quality and longer lifespan. Nevertheless, the prognostic impact of Gli1 presence in solid malignancies remains in controversy, despite an overwhelming majority of evidence has explored a negative prognostic value of Gli1 over-expression across miscellaneous neoplasms. Failed to draw a similar conclusion, Pizem et al.<sup>7</sup> presented with a better survival status of patients under stronger nuclear Gli1 expression, suggesting an astonishing positive prognostic significance in circumstance of Gli1 redundancy. Therefore, in the present study through gathering available evidence, we carried out an exhaustive meta-analysis as well as subgroups analysis to verify the prognostic influence of Gli1 positivity across solid malignancies, aiming to provide more theoretical supports for targeted regimens.

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Figure 1. Selection flow chart of the meta-analysis.

### Results

**Eligible studies.** According to the selection criteria, most of the preliminarily included entries were eliminated on account of duplicated data, inappropriate article type or inadequate original information. Eventually, a total of 39 observational studies consisting of 4496 cases were retained for subsequent pooling calculation. Figure 1 concisely displayed the selection workflow of all eligible studies in our meta-analysis.

**Demographic characteristics of included studies.** As for the source regions of included studies, the majority were carried out in China (n = 21), followed by USA (n = 6), Japan (n = 5) and other sporadic nations. None of the eligible entries scored less than six by Newcastle-Ottawa Scale, revealing a high methodological quality across all studies. Studies concerning breast cancer occupied the largest proportion of cancer type among all primary literatures (n = 6), followed by hepatocellular carcinoma (n = 5), esophageal malignancy (n = 4) and remaining types of solid neoplasm. The sample-size ranged from 19 to 339, with a median of 90 patients. A total of 38 studies described the correlation of overall survival and Gli1 expression, while 11 trials reported relationship between disease-free survival and Gli1 presence. Other detailed features were recorded and summarized in Table 1.

**Correlation of Gli1 expression with 3-year overall survival and its subgroup analysis.** 38 observational trials offered original data on 3-year overall survival in terms of different Gli1 expressions. It demonstrated that higher Gli1 activity referred to unfavorable 3-year overall survival (OR: 2.42, 95% CI: [1.91, 3.07], P < 0.00001). A moderate heterogeneity was observed ( $I^2 = 40\%$ ) so that a random-effects model was applied for statistical adjustment. In order to explore the possible sources of heterogeneity across studies, we stratified the original articles for subgroup analysis, according to various confounding factors (Fig. 2).

In the subgroup analysis by cancer type, a worse 3-year overall survival of Gli1 positivity was observed in breast cancer (n = 6, OR: 1.63, 95% CI: [1.09, 2.45], P = 0.02, I<sup>2</sup> = 1%), cancer of digestive tract (n = 8, OR: 3.51, 95% CI: [1.97, 6.24], P < 0.0001, I<sup>2</sup> = 44%), liver cancer (n = 5, OR: 3.54, 95% CI: [2.23, 5.63], P < 0.00001, I<sup>2</sup> = 0%), pancreatic cancer (n = 3, OR: 2.47, 95% CI: [1.24, 4.92], P = 0.01, I<sup>2</sup> = 7%) and ovarian cancer (n = 3, OR: 3.80, 95% CI: [1.50, 9.61], P = 0.005, I<sup>2</sup> = 0%) respectively. Nevertheless, over-expression of Gli1 in intracranial tumors was irrelevant with poorer 3-year prognosis, along with a significant heterogeneity observed (n = 5, OR: 0.99, 95% CI: [0.20, 4.98], P = 0.99, I<sup>2</sup> = 84%) (Forest plot was displayed in Supplementary Figure S1).

As for different TNM clinical stages of solid malignancies, a worse 3-year survival was strongly linked to Gli1 positivity in advanced TNM stages (TNM II-IV) (n = 5, OR: 4.51, 95% CI: [1.85, 11.00], P = 0.0009,  $I^2 = 35\%$ ) as well as pre-terminal stages (TNM I-III) (n = 5, OR: 4.28, 95\% CI: [2.35, 7.79], P < 0.00001,  $I^2 = 0\%$ ). Moreover, in

Reference	Country	Cancer type	No.	Mean age	Male/Female	TNM stage	Follow-up (range) months	Gli1(-/+) No.	3-year OS(-/+)%	5-year OS(-/+)%	10-year OS(-/+)%	NOS score
Li et al. <sup>12</sup>	China	Gallbladder carcinoma	93	NA	39/54	I–IV	32(5-66)	27/66	48.1/24.2	14.8/7.6	NA	7
Xie et al.9	China	Gallbladder carcinoma	32	$67.0\pm12.0$	11/21	I–IV	20	16/16	18.8/12.5	NA	NA	7
Hong et al.13	China	Lung cancer	55	$61.9\pm8.1$	34/21	I–IV	NA	31/24	74.2/50.0	64.5/20.8	54.8/12.5	8
Ishikawa <i>et al</i> . <sup>14</sup>	Japan	Lung cancer	102	$64.8\pm9.8$	68/34	II–IV	NA	87/15	59.8/40.0	41.4/20.0	NA	7
Che et al. <sup>15</sup>	China	Liver cancer	46	$51.7\pm11.2$	40/6	NA	30(1-83)	21/25	57.1/24.0	47.6/24.0	NA	7
Zheng et al.16	USA	Liver cancer	139	NA	NA	NA	NA	65/74	63.1/35.1	44.6/13.5	27.7/13.5	6
Tang et al. <sup>17</sup>	China	Liver cancer	108	NA	77/31	I–IV	17(2-82)	63/45	39.7/13.3	20.6/11.1	NA	6
Xu et al. <sup>18</sup>	China	Liver cancer	63	NA	49/14	I–IV	NA	6/57	83.3/35.1	NA	NA	7
Zhang et al.19	China	Liver cancer	58	51.7	43/15	I–IV	23(3-36)	31/27	41.9/22.2	NA	NA	6
Chaudary et al.20	Canada	Cervical cancer	85	47.0	All female	I–IV	72(9-127)	43/42	NA	NA	NA	7
Pressey et al.21	USA	Rhabdomyosarcoma	68	NA	46/22	I–IV	70(0-123)	26/42	84.6/92.9	80.8/92.9	80.8/92.9	7
Yan et al. <sup>22</sup>	China	Glioma	65	NA	18/47	II–IV	>3 years	18/47	77.8/42.6	61.1/42.6	NA	7
Ding et al.23	China	Colon cancer	96	NA	60/36	I–IV	38(6-60)	20/76	85.0/48.7	25.0/6.6	NA	7
Xu et al. <sup>24</sup>	China	Colon cancer	228	NA	108/120	I–III	52(5-109)	56/172	96.4/70.9	91.1/65.1	NA	7
Liao et al. <sup>25</sup>	China	Ovarian cancer	44	NA	All female	I–IV	64(5-111)	33/11	75.8/45.5	69.7/18.2	66.7/0.0	7
McCann et al.26	USA	Ovarian cancer	19	$61.0\pm13.6$	All female	III–IV	NA	10/9	40.0/11.1	40.0/0.0	20.0/0.0	8
Ciucci et al.27	Italy	Ovarian cancer	56	54.0	All female	III–IV	35(9-127)	46/10	69.6/40.0	63.0/30.0	52.2/30.0	8
He et al. <sup>28</sup>	China	Bladder cancer	118	56.0	105/13	I–III	NA	23/95	78.3/61.1	NA	NA	7
Sverrisson et al.29	China	Bladder cancer	261	NA	194/67	I–IV	33	60/201	40.0/17.9	33.3/12.4	NA	7
Haaf et al. <sup>30</sup>	Germany	Breast cancer	186	56.0	All female	NA	78(0-148)	66/120	86.4/83.3	80.3/66.7	77.3/57.5	8
Xu et al. <sup>31</sup>	USA	Breast cancer	139	NA	All female	I–IV	94	94/45	81.9/80.0	76.6/68.9	64.9/60.0	7
O'Toole et al. <sup>32</sup>	USA	Breast cancer	267	55.0	All female	NA	NA	184/83	90.8/85.5	89.7/71.1	79.9/65.1	7
Li et al. <sup>33</sup>	China	Breast cancer	284	52.0	All female	I–III	62(3-83)	48/236	95.8/81.8	95.8/77.1	NA	7
Ramaswamy et al.34	USA	Breast cancer	289	54.5	All female	NA	96(1-139)	89/200	95.5/88.0	89.9/78.0	71.9/51.0	8
He et al.35	China	Breast cancer	290	NA	All female	NA	NA	131/159	96.9/96.6	92.4/82.4	NA	6
Souzaki <i>et al.</i> <sup>36</sup>	Japan	Neuroblastoma	92	NA	56/36	I–IV	NA	30/62	80.0/95.2	70.0/93.5	NA	7
Mori et al. <sup>8</sup>	Japan	Esophageal cancer	104	63.0	92/12	I–IV	NA	52/52	71.2/48.1	71.2/34.6	NA	6
Yoshikawa <i>et al.</i> <sup>37</sup>	Japan	Esophageal cancer	69	60.7	58/11	II–IV	NA	12/57	50.0/0.0	NA	NA	7
Zhu et al. <sup>38</sup>	China	Esophageal cancer	100	NA	85/15	I–IV	23(3-83)	28/72	71.4/51.4	50.0/30.6	NA	7
Wei et al. <sup>39</sup>	China	Esophageal cancer	35	60.0	29/6	I–IV	NA	10/25	60.0/28.0	60.0/8.0	NA	6
Pizem et al.40	Slovenia	Medulloblastoma	65	12.4	51/14	NA	71(2-234)	25/40	60.0/100.0	52.0/90.0	52.0/80.0	8
Buczkowicz et al.41	Canada	Medulloblastoma	90	6.6	NA	NA	80	70/20	68.6/40.0	65.7/30.0	60.0/25.0	6
Cordeiro et al.42	Brazil	Medulloblastoma	40	NA	26/14	NA	NA	28/12	78.6/50.0	71.4/41.7	64.3/33.3	7
Chung et al.43	USA	Head and neck squamous cell carcinoma	339	NA	263/76	NA	96	192/147	43.2/29.9	33.9/17.7	NA	7
Saze et al.44	Japan	Gastric cancer	41	64.8	30/11	I–IV	NA	19/22	57.9/50.0	47.4/36.4	NA	8
Wang et al.45	China	Gastric cancer	121	63.0	92/29	I–IV	30	25/96	68.8/47.9	56.0/26.0	NA	8
Jiang et al.46	China	Pancreatic cancer	90	62.0	57/33	I–IV	31(0-87)	35/55	57.1/45.5	57.1/43.6	NA	7
Sheng et al.47	China	Pancreatic cancer	57	NA	38/19	I–III	NA	28/29	32.1/10.3	NA	NA	6
Liu et al.48	China	Pancreatic cancer	62	NA	43/19	I–III	NA	29/33	31.0/9.1	NA	NA	7

Table 1. Demographic information of included studies. NOS: Newcastle-Ottawa Scale; NA: not available.

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studies with coverage to all stages (TNM I-IV), aberrant Gli1 activation likewise suggested a poorer 3-year overall outcome (n = 18, OR: 2.16, 95% CI: [1.55, 3.01], P < 0.00001,  $I^2 = 38\%$ ) (Supplementary Figure S2).

There were totally two stratified subgroups in terms of mean age among studies. In both groups of mean age >60-year (n = 10, OR: 2.33, 95% CI: [1.52, 3.57], P < 0.0001, I<sup>2</sup> = 18%) and mean age <60-year (n = 10, OR: 2.23, 95% CI: [1.35, 3.69], P = 0.002, I<sup>2</sup> = 47%), negative Gli1 expression was verified to significantly associate with more favorable 3-year outcome (Supplementary Figure S3).

Subgroups analysis by Gli1 detection methods explored that high Gli1 expression status was identified as a worse prognostic marker of 3-year outcome in mRNA-examined group (n = 8, OR: 2.49, 95% CI: [1.61, 3.84], P < 0.0001,  $I^2 = 7\%$ ). Similarly, obvious Gli1 positivity analyzed by immunohistochemistry (IHC) as well predicted an unfavorable 3-year overall survival in solid cancer (n = 30, OR: 2.43, 95% CI: [1.85, 3.21], P < 0.00001,  $I^2 = 47\%$ ) (Supplementary Figure S4).

Included studies were divided into three subgroups according to the subcellular localization of Gli1 expression by IHC. Trials based on nuclear staining of Gli1 significantly correlated its over-reactivity with a worse 3-year survival outcome (n = 13, OR: 3.04, 95% CI: [1.82, 5.07], P < 0.0001, I<sup>2</sup> = 56%), which made a similar conclusion as group of unspecific expression did (n = 14, OR: 2.27, 95% CI: [1.74, 2.96], P < 0.00001, I<sup>2</sup> = 8%). However,

	Gli1	-	Gli1-	F.		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
Buczkowicz 2011	48	70	8	20	3.1%	3.27 [1.17, 9.14]	
Che 2012	12	21	6	25	2.4%	4.22 [1.20, 14.90]	
Chung 2011	83	192	44	147	5.8%	1.78 [1.13, 2.81]	
Ciucci 2013	32	46	4	10	2.0%	3.43 [0.84, 14.08]	
Cordeiro 2014	22	28	6	12	2.0%	3.67 [0.86, 15.59]	
Ding 2012	17	20	37	76	2.3%	5.97 [1.62, 22.08]	—
Haaf 2009	57	66	100	120	3.8%	1.27 [0.54, 2.97]	
He 2012	18	23	58	95	2.9%	2.30 [0.79, 6.72]	<del></del>
He 2015	127	131	154	159	2.2%	1.03 [0.27, 3.92]	
Hong 2014	23	31	12	24	2.7%	2.88 [0.92, 8.94]	
Ishikawa 2014	52	87	6	15	2.8%	2.23 [0.73, 6.82]	
Jiang 2014	20	35	25	55	3.8%	1.60 [0.68, 3.76]	
Li 2012-B	46	48	193	236	1.9%	5.12 [1.20, 21.93]	
Li 2012-G	13	27	16	66	3.4%	2.90 [1.13, 7.44]	
Liao 2009	25	33	5	11	2.0%	3.75 [0.90, 15.66]	
Liu 2015	9	29	3	33	2.0%	4.50 [1.08, 18.69]	
McCann 2011	4	10	1	9	0.8%	5.33 [0.47, 60.80]	
Mori 2006	37	52	25	52	4.0%	2.66 [1.19, 5.99]	
OToole 2011	167	184	71	83	4.0%	1.66 [0.75, 3.66]	+
Pizem 2011	15	25	40	40	0.6%	0.02 [0.00, 0.33]	
Pressey 2011	22	26	39	42	1.7%	0.42 [0.09, 2.07]	
Ramaswamy 2012	85	89	176	200	2.9%	2.90 [0.97, 8.62]	
Saze 2012	11	19	11	22	2.4%	1.38 [0.40, 4.73]	
Sheng 2014	9	28	3	29	2.0%	4.11 [0.98, 17.23]	
Souzaki 2010	24	30	59	62	1.9%	0.20 [0.05, 0.88]	
Sverrisson 2014	24	60	36	201	4.8%	3.06 [1.63, 5.74]	
Tang 2013	25	63	6	45	3.2%	4.28 [1.58, 11.59]	
Wang 2014	17	25	46	96	3.4%	2.31 [0.91, 5.86]	<b>—</b> •—
Wei 2011	6	10	7	25	1.8%	3.86 [0.83, 17.94]	
Xie 2014	2	16	3	16	1.2%	0.62 [0.09, 4.32]	
Xu 2010	77	94	36	45	3.6%	1.13 [0.46, 2.78]	
Xu 2012	54	56	122	172	2.0%	11.07 [2.60, 47.14]	
Xu 2014	5	6	20	57	1.0%	9.25 [1.01, 84.73]	
Yan 2014	14	18	20	47	2.4%	4.72 [1.35, 16.54]	
Yoshikawa 2008	6	12	0	57	0.6%	115.00 [5.79, 2283.65]	→
Zhang 2014	13	31	6	27	2.7%	2.53 [0.80, 8.02]	
Zheng 2012	41	65	26	74	4.5%	3.15 [1.58, 6.31]	_ <del>_</del> +
Zhu 2011	20	28	37	72	3.4%	2.36 [0.92, 6.06]	
Total (95% CI)		1834		2577	100.0%	2.42 [1.91, 3.07]	•
Total events	1282		1467				
Heterogeneity: Tau <sup>2</sup> =	0.20; Chi	²= 61.	97, df = 3	7 (P = 0	0.006); l <sup>e</sup> =	= 40%	
Test for overall effect:	Z = 7.33 (	(P < 0.0	00001)				Favours [Gli1+] Favours [Gli1-]

Figure 2. The correlation between Gli1 expression and 3-year overall survival in solid malignancies.

the abnormal cytoplasmic existence of Gli1 did not stand for a detrimental tendency on 3-year overall survival through our pooling analysis (n = 3, OR: 1.11, 95% CI: [0.25, 4.83], P = 0.89, I<sup>2</sup> = 76%) (Supplementary Figure S5). Among the subgroups determined by sample size of the studies, elevated Gli1 expression was confirmed to play a worse prognostic role in terms of 3-year overall survival in solid malignancies, regardless of greater amount (>90) (n = 21, OR: 2.27, 95% CI: [1.77, 2.91], P < 0.00001, I<sup>2</sup> = 35%) or smaller magnitude (<90) of participants

(n = 17, OR: 2.75, 95% CI: [1.63, 4.66], P = 0.0002, I<sup>2</sup> = 49%) (Supplementary Figure S6). Stratified by source regions of the included studies, there was an analogical trend of trials implemented by Asian (n = 26, OR: 2.79, 95% CI: [2.12, 3.69], P < 0.00001, I<sup>2</sup> = 30%) or non-Asian investigators (n = 12, OR: 1.83, DSW (CL [1.20, 2.78], P = 0.005, I<sup>2</sup> = 50%) bet Cl: 1 even preservitive was identified as a preserve 2 were preserved.

Asian (n = 26, OR: 2.79, 95% CI: [2.12, 3.69], P < 0.00001,  $I^2$  = 30%) or non-Asian investigators (n = 12, OR: 1.83, 95% CI: [1.20, 2.78], P = 0.005,  $I^2$  = 50%) that Gli1 over-reactivity was identified as a poorer 3-year prognostic marker in solid tumors (Supplementary Figure S7).

**Correlation of Gli1 expression with 5-year overall survival and its subgroup analysis.** Concerning the 5-year overall survival in solid malignancies, the pooling analysis revealed a negative impact of Gli1 over-expression on clinical prognosis of patients, along with a moderate heterogeneity of undefined source

	Gli1	-	Gli1-	F		Odds Ratio		Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H	, Random, 95% Cl	
Buczkowicz 2011	46	70	6	20	3.3%	4.47 [1.52, 13.12]			
Che 2012	10	21	6	25	2.8%	2.88 [0.82, 10.10]			
Chung 2011	65	192	26	147	5.0%	2.38 [1.42, 4.00]			
Ciucci 2013	29	46	3	10	2.4%	3.98 [0.91, 17.47]			
Cordeiro 2014	20	28	5	12	2.5%	3.50 [0.85, 14.34]		<u> </u>	
Ding 2012	5	20	5	76	2.6%	4.73 [1.22, 18.42]			
Haaf 2009	53	66	80	120	4.4%	2.04 [1.00, 4.17]			
He 2015	121	131	131	159	4.3%	2.59 [1.21, 5.55]			
Hong 2014	20	31	5	24	2.9%	6.91 [2.02, 23.63]			
Ishikawa 2014	36	87	3	15	2.7%	2.82 [0.74, 10.73]			
Jiang 2014	20	35	24	55	4.0%	1.72 [0.73, 4.05]			
Li 2012-B	46	48	182	236	2.4%	6.82 [1.60, 29.03]		<u> </u>	
Li 2012-G	4	27	5	66	2.5%	2.12 [0.52, 8.60]			
Liao 2009	23	33	2	11	2.0%	10.35 [1.89, 56.80]			
McCann 2011	4	10	0	9	0.8%	13.15 [0.60, 288.33]			
Mori 2006	37	52	18	52	4.0%	4.66 [2.03, 10.67]			
OToole 2011	165	184	59	83	4.6%	3.53 [1.81, 6.91]			
Pizem 2011	13	25	36	40	2.7%	0.12 [0.03, 0.44]		—	
Pressey 2011	21	26	39	42	2.3%	0.32 [0.07, 1.49]			
Ramaswamy 2012	80	89	156	200	4.2%	2.51 [1.17, 5.39]			
Saze 2012	9	19	8	22	2.8%	1.57 [0.45, 5.50]			
Souzaki 2010	21	30	58	62	2.8%	0.16 [0.04, 0.58]			
Sverrisson 2014	20	60	25	201	4.5%	3.52 [1.78, 6.95]			
Tang 2013	13	63	5	45	3.2%	2.08 [0.68, 6.32]		+	
Wang 2014	14	25	25	96	3.8%	3.61 [1.45, 9.00]			
Wei 2011	6	10	2	25	1.7%	17.25 [2.53, 117.72]			
Xu 2010	72	94	31	45	4.2%	1.48 [0.67, 3.26]		<b></b>	
Xu 2012	51	56	112	172	3.6%	5.46 [2.07, 14.42]			
Yan 2014	11	18	20	47	3.2%	2.12 [0.70, 6.44]		+•	
Zheng 2012	29	65	10	74	4.0%	5.16 [2.26, 11.78]			
Zhu 2011	14	28	22	72	3.8%	2.27 [0.93, 5.56]			
Total (95% Cl)		1689		2263	100.0%	2,58 [1,93, 3.46]		•	
Total events	1078		1109			,,,			
Heterogeneity: Tau <sup>2</sup> =	0.37: Ch	i <sup>z</sup> = 73.	97. df = 3	0 (P < 1	0.0001) <sup>,</sup> P	²= 59%		<u> </u>	
Test for overall effect:	Z = 6.38	(P < ∩ f	10001)				0.001 0.1	1 10	1000
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Figure 3. The correlation between Gli1 expression and 5-year overall survival in solid malignancies.

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 $(n = 31, OR: 2.58, 95\% CI: [1.93, 3.46], P < 0.00001, I^2 = 59\%)$ . Thus we performed the following subgroup analysis aiming to clarify the possible confounding elements (Fig. 3).

According to subgroups of different cancer types, Gli1 over-expression in breast cancer (n = 6, OR: 2.52, 95% CI: [1.83, 3.47], P < 0.00001, I<sup>2</sup> = 0%), cancer of digestive tract (n = 7, OR: 3.78, 95% CI: [2.50, 5.70], P < 0.00001, I<sup>2</sup> = 7%), liver cancer (n = 3, OR: 3.53, 95% CI: [1.96, 6.34], P < 0.0001, I<sup>2</sup> = 0%) and ovarian cancer (n = 3, OR: 6.57, 95% CI: [2.30, 18.79], P = 0.0004, I<sup>2</sup> = 0%) contributed to a significantly worse 5-year prognosis, except for intracranial tumor, in which Gli1 expression was not associated with the outcome prediction as usual (n = 5, OR: 0.93, 95% CI: [0.21, 4.19], P = 0.92, I<sup>2</sup> = 87%) (Supplementary Figure S8).

With respect to subgroups by different TNM stages, our quantitative results suggested that Gli1 abundance in cancer tissues closely correlated to unfavorable 5-year overall survival in all stratified groups, including advanced stages (n = 4, OR: 2.95, 95% CI: [1.44, 6.06], P = 0.003, I<sup>2</sup> = 0%), pre-terminal stages (n = 2, OR: 5.85, 95% CI: [2.61, 13.11], P < 0.0001, I<sup>2</sup> = 0%) and unselected stages (n = 15, OR: 2.37, 95% CI: [1.46, 3.85], P = 0.0005, I<sup>2</sup> = 66%) (Supplementary Figure S9).

Gli1 over-expression in different mean-age subgroups similarly imposed a negative effect on 5-year overall survival, especially in the group with older mean-age (Mean-age >60-year: n = 8, OR: 3.54, 95% CI: [2.19, 5.72], P < 0.00001,  $I^2 = 25\%$ ) (Mean-age <60-year: n = 8, OR: 2.32, 95% CI: [1.17, 4.58], P = 0.02,  $I^2 = 73\%$ ) (Supplementary Figure S10).

Different detection approaches seemed not influence the negative prognostic role of Gli1 redundancy on 5-year overall survival in solid cancer, including PCR-analyzing group (n = 8, OR: 2.85, 95% CI: [1.48, 5.49], P = 0.002,  $I^2 = 48\%$ ) as well as immunohistochemistry staining group (n = 23, OR: 2.63, 95% CI: [1.88, 3.69], P < 0.00001,  $I^2 = 63\%$ ) (Supplementary Figure S11).

	Gli1-		Gli1+		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
Buczkowicz 2011	42	70	5	20	8.0%	4.50 [1.47, 13.78]	
Ciucci 2013	24	46	3	10	6.0%	2.55 [0.58, 11.08]	
Cordeiro 2014	18	28	4	12	6.2%	3.60 [0.86, 15.01]	
Haaf 2009	51	66	69	120	11.1%	2.51 [1.27, 4.96]	
Hong 2014	17	31	3	24	6.3%	8.50 [2.09, 34.52]	
Liao 2009	22	33	0	11	2.2%	45.00 [2.43, 833.69]	
McCann 2011	2	10	0	9	1.9%	5.59 [0.23, 133.61]	
OToole 2011	147	184	54	83	11.9%	2.13 [1.20, 3.80]	
Pizem 2011	13	25	32	40	8.1%	0.27 [0.09, 0.82]	
Pressey 2011	21	26	39	42	5.7%	0.32 [0.07, 1.49]	
Ramaswamy 2012	64	89	102	200	12.1%	2.46 [1.43, 4.22]	
Xu 2010	61	94	27	45	10.7%	1.23 [0.59, 2.56]	
Zheng 2012	18	65	10	74	9.8%	2.45 [1.04, 5.79]	
Total (95% Cl)		767		690	100.0%	2.11 [1.32, 3.39]	<b></b>
Total events	500		348				
Heterogeneity: Tau <sup>2</sup> =	0.40; Ch	i <sup>z</sup> = 32.1					
Test for overall effect:	Z = 3.11 (	(P = 0.0	Eavoure [Gli1+] Eavoure [Gli1_]				

Figure 4. The correlation between Gli1 expression and 10-year overall survival in solid malignancies.

In terms of subcellular localization of Gli1 staining, the excessive expression of Gli1 distributed in nucleus (n = 11, OR: 3.00, 95% CI: [1.61, 5.58], P = 0.0006, I<sup>2</sup> = 70%) or unspecific locations (n = 9, OR: 2.61, 95% CI: [2.03, 3.36], P < 0.00001, I<sup>2</sup> = 0%) referred to a disappointing 5-year overall survival in a significant way. Nevertheless, a cytoplasmic positivity of Gli1 expression failed to make a similar conclusion via the pooling analysis, displaying a dubious connection with 5-year prognosis (n = 3, OR: 1.34, 95% CI: [0.17, 10.48], P = 0.78, I<sup>2</sup> = 88%) (Supplementary Figure S12).

Among the subgroups divided by different amount of sample sizes, no matter with a greater (>90) (n = 20, OR: 2.71, 95% CI: [2.07, 3.53], P < 0.00001, I<sup>2</sup> = 44%) or smaller volume (<90) (n = 11, OR: 2.49, 95% CI: [1.04, 5.95], P = 0.04, I<sup>2</sup> = 74%) of sample-size, the merged analysis consistently provided an unfavorable effect on 5-year overall survival owing to Gli1 aberrant activation in solid malignancies (Supplementary Figure S13).

Included studies were stratified into Asian group (n = 19, OR: 2.94, 95% CI: [2.06, 4.20], P < 0.00001,  $I^2 = 50\%$ ) and non-Asian group (n = 12, OR: 2.10, 95% CI: [1.27, 3.48], P = 0.004,  $I^2 = 70\%$ ) on the basis of different source regions. There was a significant association between Gli1 over-expression and poorer 5-year overall survival through our pooling analysis, regardless of the multiple nationality among the studies (Supplementary Figure S14).

**Correlation of Gli1 expression with 10-year overall survival and its subgroup analysis.** In our pooling analysis, Gli1 positivity was verified to have a strong connection with a worse 10-year prognosis in solid carcinomas, albeit a moderate heterogeneity was observed across studies (n = 13, OR: 2.11, 95% CI: [1.32, 3.39], P = 0.002,  $I^2 = 63\%$ ). A subgroup analysis is therefore conducted (Fig. 4).

Studies were stratified according to different types of cancer. Gli1 positivity in both breast cancer (n = 4, OR: 2.10, 95% CI: [1.54, 2.86], P < 0.00001, I<sup>2</sup> = 0%) and ovarian cancer (n = 3, OR: 6.28, 95% CI: [1.07, 36.69], P = 0.04, I<sup>2</sup> = 38%) acted as an indicator of a worse 10-year overall survival, while its over-expression in medul-loblastoma (n = 3, OR: 1.60, 95% CI: [0.25, 10.40], P = 0.62, I<sup>2</sup> = 86%) was unable to offer a prognostic prediction as above (Supplementary Figure S15).

**Correlation of Gli1 expression with 3-year and 5-year disease-free survival.** Our quantitative analysis confirmed that high Gli1 expression was shown to be an unfavorable prognostic marker of 3-year (n = 11, OR: 2.62, 95% CI: [1.79, 3.85], P < 0.00001, I<sup>2</sup> = 25%) and 5-year disease-free survival (n = 8, OR: 3.82, 95% CI: [2.19, 6.68], P < 0.00001, I<sup>2</sup> = 55%) among solid malignancies, along with a moderate heterogeneity existed (Fig. 5).

**Sensitivity analysis.** Removal of studies featuring intracranial tumors had no substantial impact on the outcomes of 3-year (n = 33, OR: 2.52, 95% CI: [2.12, 3.00], P < 0.00001, I<sup>2</sup> = 17%), 5-year (n = 26, OR: 2.95, 95% CI: [2.46, 3.54], P < 0.00001, I<sup>2</sup> = 21%) and 10-year overall survival (n = 10, OR: 2.26, 95% CI: [1.47, 3.46], P = 0.0002, I<sup>2</sup> = 47%), however, a huge decline on heterogeneity was observed respectively.

Exclusion of studies with mesenchymal origin obtained similar results of 3-year (n = 37, OR: 2.49, 95% CI: [1.97, 3.14], P < 0.00001, I<sup>2</sup> = 37%), 5-year (n = 30, OR: 2.71, 95% CI: [2.04, 3.59], P < 0.00001, I<sup>2</sup> = 56%) and 10-year overall survival (n = 12, OR: 2.34, 95% CI: [1.49, 3.67], P = 0.0002, I<sup>2</sup> = 59%) respectively.

Elimination of studies scoring 6 in NOS assessment was unable to alter the negative prognostic effect of Gli1 positivity in terms of 3-year (n = 30, OR: 2.30, 95% CI: [1.71, 3.08], P < 0.00001,  $I^2 = 49\%$ ), 5-year (n = 25, OR: 2.29, 95% CI: [1.63, 3.22], P < 0.00001,  $I^2 = 62\%$ ) and 10-year overall survival among solid malignancies (n = 11, OR: 1.95, 95% CI: [1.12, 3.38], P = 0.02,  $I^2 = 67\%$ ).

	Gli 1	-	Gli1+	F		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.7.1 3-year DFS							
Chaudary 2012	32	43	26	42	6.6%	1.79 [0.71, 4.52]	+
Che 2012	11	21	6	25	4.6%	3.48 [0.99, 12.22]	
Ciucci 2013	15	46	1	10	2.0%	4.35 [0.50, 37.61]	
He 2012	16	23	55	95	6.3%	1.66 [0.63, 4.42]	
He 2015	118	131	128	159	8.7%	2.20 [1.10, 4.40]	
Ishikawa 2014	24	87	2	15	3.3%	2.48 [0.52, 11.80]	-+
Ramaswamy 2012	78	89	160	200	8.5%	1.77 [0.86, 3.64]	+
Tang 2013	9	63	4	45	4.6%	1.71 [0.49, 5.94]	_ <b>+-</b> _
Xu 2012	52	56	108	172	5.7%	7.70 [2.66, 22.30]	
Yoshikawa 2008	5	12	0	57	1.1%	84.33 [4.23, 1682.72]	
Zhu 2011	18	28	9	27	5.4%	3.60 [1.18, 10.95]	
Subtotal (95% CI)		599		847	<b>56.6</b> %	2.62 [1.79, 3.85]	•
Total events	378		499				
Heterogeneity: Tau <sup>2</sup> =	0.10; Chi	i <sup>z</sup> = 13.3	30, df = 1	0 (P = 0	0.21); I <sup>z</sup> =	25%	
Test for overall effect:	Z = 4.92 (	(P < 0.0	00001)				
1.7.2 5-year DFS							
Chaudary 2012	31	43	25	42	6.8%	1.76 [0.71, 4.35]	
Che 2012	9	21	3	25	3.6%	5.50 [1.25, 24,26]	<b>_</b>
Ciucci 2013	13	46	1	10	1.9%	3.55 [0.41, 30.85]	
He 2015	88	131	59	159	11.0%	3.47 [2.13, 5.64]	
Ishikawa 2014	18	87	1	15	2.1%	3.65 [0.45, 29.65]	
Ramaswamy 2012	76	89	150	200	9.0%	1.95 [1.00, 3.81]	<b>⊢</b> •−
Xu 2012	47	56	64	172	7.9%	8.81 [4.05, 19.17]	
Zhu 2011	10	28	0	72	1.2%	82.30 [4.61, 1469.89]	
Subtotal (95% CI)		501		695	43.4%	3.82 [2.19, 6.68]	•
Total events	292		303				
Heterogeneity: Tau <sup>2</sup> =	0.30; Chi	i <sup>z</sup> = 15.1	73, df = 7	(P = 0.	03); <b>P</b> = 5	5%	
Test for overall effect:	Z= 4.71 (	(P < 0.0	00001)	,			
Total (95% Cl)		1100		1542	100.0%	3.10 [2.24, 4.28]	•
Total events	670		802				-
Heterogeneity: Tau <sup>2</sup> =	0.19: Chi	i <sup>z</sup> = 31.1	12.df=1	8 (P = 1	1.03): P=	42%	
Test for overall effect:	7 = 6.85 (	Έ≺በበ	0.001 0.1 1 10 1000				
Test for subaroup diff	erences:	Favours [Gli1+] Favours [Gli1-]					

Figure 5. The correlation between Gli1 expression and disease-free survival in solid malignancies.

**Publication bias.** The funnel plots, Egger test and Begg test jointly demonstrated that there was no publication bias concerning 3-year (Egger: P = 0.505; Begg: P = 0.138), 5-year (Egger: P = 0.996; Begg: P = 0.415) and 10-year (Egger: P = 0.653; Begg: P = 0.583) overall survival in our meta-analysis (Fig. 6).

### Discussion

As the laboratorial evidence has revealed, inactivation of Sonic hedgehog pathway is generally marked by translocation of Gli1 transcriptional factor into the nucleus. The frequently over-expressed Gli1 strongly triggers the carcinogenesis and dissemination in various cancer models. Its oncogenic role takes effects by up-regulation of the downstream target genes, especially including certain detrimental oncoproteins<sup>8</sup>. The pivotal action of Gli1 inside the molecular carcinogenesis network emphasizes its potential value for targeted treatment modalities. However, from the clinical perspective, a persuasive support of Gli1's clinical significance is still unavailable, partially due to the uncertainty of the association between Gli1 positivity and prognosis implication. The majority of investigations established potent evidence suggesting an unfavorable impact of Gli1 abnormality on clinical prognosis in a wide spectrum of carcinomas. Nevertheless, on the contrary, several researchers recently highlighted that an obvious advantage on survival duration was obtained in Gli1 over-expression cases, with mechanisms not fully elucidated<sup>7</sup>. A comprehensive study is therefore in urgent demand.

To our knowledge, the present study is the first and most full-scale meta-analysis systemically exploring the possible prognostic role of Gli1 up-regulation in solid malignancies. On the whole, our quantitative results strongly supported the current mainstream viewpoint that an undesirable impact of Gli1 redundancy was correlated with the 3-year, 5-year, 10-year overall survival and disease-free survival, taking no account of subgroup confounding factors. Additionally, this negative prognostic role was confirmed to be independent of mean-age, source countries, detection measures, sample-size and TNM stages. With respect to subcellular localization, nuclear emergence of Gli1 was identified to tightly associate with awful prognosis while cytoplasmic expression culminated in an absolutely contrary conclusion that there was no obvious connection between Gli1 cytoplasmic



**Figure 6.** The funnel plots of this meta-analysis. (A) 3-year overall survival; (B) 5-year overall survival; (C) 10-year overall survival.

expression and worse prognosis in solid malignancies. This inconsistency is probably attributed to the academic consensus that translocation of Gli1 into the nucleus is the hallmark of inactivation of Sonic hedgehog pathway, exclusive of the cytoplasmic positioning pattern<sup>9</sup>. Moreover, Gli1 lost its prediction efficacy in intracranial tumors by subgroup analysis, displaying an indistinctive prognostic effect on long-term survival. The possible explanation of this exception may be owing to the less possibility of metastasis and cerebrospinal fluid dissemination in medulloblastomas with Gli1 nuclear staining, although the molecular mechanism remains unclear<sup>10</sup>. Furthermore, Gershon *et al.* reported a proliferation-inhibitory effect of Gli1 over-expression on neuroblastoma

plausible reasons, it warrants more investigations to make further explanations to these controversies. Apart from the inspiring outcomes, limitations still lay in this quantitative meta-analysis. First of all, despite the usage of random-effects model and subgroup analysis, the heterogeneity across studies failed to be eliminated completely, which could result in bias of the outcome in certain extent. Secondly, on account of the lacking of effective data, we merely analyzed the correlation between Gli1 redundancy and prognosis in terms of certain clinical elements. Other parameters that may partially contribute to the heterogeneity were not explored, such as pathological grade and body mass index.

cell lines, which triggered the malignancy towards mature differentiation instead<sup>11</sup>. In spite of aforementioned

In spite of the limitations mentioned above, there are still numerous valuable implications of this comprehensive mete-analysis. Firstly, the formerly recognized oncoprotein Gli1 has been identified as a biomarker for poor prognosis of 3-year, 5-year, 10-year overall survival and disease-free survival in solid malignancies for the first time, irrespective of several exceptions including intracranial tumors. Secondly, nuclear localization instead of cytoplasmic positivity serves as a prognostic predictor in most cancers, which provides a more specific subcellular positioning to guide the clinical evaluations for diagnosis and prognosis.

#### Methods

**Search strategy.** We performed a thorough search for available literatures in electronic databases of PubMed and Web of Science until September 2015. The search terms "Gli1 AND (cancer OR neoplasm OR carcinoma OR malignancy)" was applied and we initially identified 904 entries for further examination. Both abstracts and full-texts were elaborately screened to exclude irrelevant articles. Citation lists of retrieved articles were additionally reviewed to guarantee the sensitivity of the search process. Two authors independently carried out this procedure and any discrepancy was resolved by mutual discussion.

**Selection criteria.** Studies that met the following requirements were considered eligible and selected into our quantitative meta-analysis: 1. English-written articles published from 2000–2015; 2. Studies exploring the correlation between Gli1 expression and prognosis in human solid malignancies; 3. A minimal follow-up duration of 3 years.

Studies were eliminated on the basis of the following criteria: 1. Duplicated or overlapped studies; 2. Studies with a sample-size less than 10 participants; 3. Lack of enough statistical data for further quantification calculation; 4. Review articles or case reports.

All evaluations were separately undertaken by two individuals to warrant the precision of selection process.

**Data extraction.** By aid of predefined standardized extraction forms, two investigators independently extracted data from each qualified studies, in terms of general information, 3-year survival, 5-year survival, 10-year survival and disease-free survival respectively. The original survival data were obtained from the text, tables or Kaplan-Meier curves for both comparative groups. GetData Graph Digitizer 2.2 helped us to digitize and extract survival information from the Kaplan-Meier curves. A joint decision was offered in the case of any disagreement.

**Methodological quality assessment.** Given that all of the included studies were observational studies, a Newcastle-Ottawa Scale (NOS) was adopted to assist the quality assessment of each eligible article. The scale was

revised with certain adaptive modifications to match the practical needs of the pooled analysis. There were generally three categories contained in the scale, including selection, comparability and outcome, with a maximum score of nine. Studies graded with more than six scores were classified as high quality literatures in methodology. Two reviewers independently conducted the assessment process and a consensus on NOS score of each study had achieved.

**Statistical analysis.** Review Manager 5.3 was adopted for the quantitative calculation in this meta-analysis. The original data was inserted as dichotomous variables, therefore odds ratio (OR) along with 95% confidential interval (CI) was applied to measure the correlation between Gli1 presence and long-term survival, including the general survival analysis and sub-groups comparison. I<sup>2</sup> was designated as an indicator of heterogeneity across studies, with its value below 25% defined as low heterogeneity. In the absence of significant heterogeneity (I<sup>2</sup> < 25%), a fixed-effects model was appropriately employed, while a random-effects model took its place in the remaining situations. P < 0.05 certainly signified a statistical significance within the comparisons. Additionally, we implemented a sensitivity analysis examining the consistency of the pooled outcomes. Funnel plots as well as Egger's test and Begg's test was used to investigate the internal publication bias across the included studies, based on the calculation by Stata 12.0 software.

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#### **Author Contributions**

J.C. contributed to the design of the study and manuscript writing. J.C., J.G. and P.Y. contributed to the data extraction and analysis process of the study. P.Y. and K.T. contributed to the financial support and revision of the manuscript.

#### Additional Information

Supplementary information accompanies this paper at http://www.nature.com/srep

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