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A prognostic index model for predicting overall survival in patients with metastatic castration-resistant prostate cancer treated with abiraterone acetate after docetaxel

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Background: Few prognostic models for overall survival (OS) are available for patients with metastatic castrationresistant prostate cancer (mCRPC) treated with recently approved agents. We developed a prognostic index model using readily available clinical and laboratory factors from a phase III trial of abiraterone acetate (hereafter abiraterone) in combination with prednisone in post-docetaxel mCRPC.

Patients and methods: Baseline data were available from 762 patients treated with abiraterone–prednisone. Factors were assessed for association with OS through a univariate Cox model and used in a multivariate Cox model with a stepwise procedure to identify those of significance. Data were validated using an independent, external, population-based cohort.

Results: Six risk factors individually associated with poor prognosis were included in the final model: lactate dehydrogenase > upper limit of normal (ULN) [hazard ratio (HR) = 2.31], Eastern Cooperative Oncology Group performance status of 2 (HR = 2.19), presence of liver metastases (HR = 2.00), albumin ≤ 4 g/dl (HR = 1.54), alkaline phosphatase > ULN

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(HR = 1.38) and time from start of initial androgen-deprivation therapy to start of treatment \leq 36 months (HR = 1.30). Patients were categorized into good (*n* = 369, 46%), intermediate (*n* = 321, 40%) and poor (*n* = 107, 13%) prognosis groups based on the number of risk factors and relative HRs. The *C*-index was 0.70 ± 0.014. The model was validated by the external dataset (*n* = 286).

Conclusion: This analysis identified six factors used to model survival in mCRPC and categorized patients into three distinct risk groups. Prognostic stratification with this model could assist clinical practice decisions for follow-up and monitoring, and may aid in clinical trial design.

Trial registration numbers: NCT00638690.

Key words: castration-resistant prostate cancer, abiraterone acetate, prognostic, risk, survival

introduction

Prostate cancer is the second most common cancer and the fifth leading cause of cancer death in men worldwide [1]. Virtually, all patients with advanced prostate cancer respond to medical or surgical castration, but these therapies are not curative and the disease eventually progresses to metastatic castration-resistant prostate cancer (mCRPC) [2, 3]. Several available therapies improve outcomes for patients with mCRPC [4–14].

Abiraterone acetate (hereafter abiraterone) is a prodrug of abiraterone, an inhibitor of CYP17 that blocks androgen biosynthesis [15]. COU-AA-301 (clinicaltrials.gov identifier: NCT00638690) was a multinational, randomized, double-blind, phase III trial comparing abiraterone plus prednisone/prednisolone (hereafter prednisone) with placebo plus prednisone/prednisolone (hereafter prednisone alone) in patients with mCRPC progressing after docetaxel. Abiraterone–prednisone significantly prolonged overall survival (OS) compared with prednisone alone [6, 7]. Abiraterone in combination with prednisone is indicated for the treatment of patients with mCRPC and is considered standard of care [16].

The wide selection of life-prolonging agents for the treatment of mCRPC complicates clinical decision-making [4–13]. Prognostic models estimate risk for clinically significant diseaserelated morbidity or mortality [17], and can be important for stratification and patient selection in clinical trials. Many prognostic models for prostate cancer have focused on clinically localized disease, and those for mCRPC were developed before the introduction of newer therapies, or focus on markers that can be costly and more difficult to obtain, making them impractical for daily clinical practice [18, 19].

Recognizing the need for a prognostic tool that reflects outcomes from currently available treatments, we explored factors associated with OS in the abiraterone–prednisone arm of the COU-AA-301 study. We report on a tripartite model that defines a meaningful range of risk, based on factors that can be obtained rapidly in routine patient management.

methods

patient population

The study design and efficacy results of COU-AA-301 are published [6, 7]. Nineteen routinely available and readily assessable clinical and baseline laboratory factors were identified and included in the analysis. Data from 762 of 797 patients in the abiraterone–prednisone arm formed the basis for the modeling. Thirty-five patients with missing relevant baseline data were not included in the modeling. The model was applied to the 398 patients in the prednisone alone arm as a validation to test its discriminative ability in a cohort of patients not treated with abiraterone–prednisone. The external validation set consisted of data from 286 sequentially treated mCRPC patients from 11 centers in Canada who received abiraterone–prednisone as standard therapy after docetaxel. Baseline demographics and disease characteristics for the abiraterone–prednisone arm and the prednisone alone arm were well balanced. The Canadian validation cohort was generally similar, although this cohort had a greater percentage of patients with Eastern Cooperative Oncology Group performance status (ECOG PS) of 2 (36% versus 10%–11%) and other apparent differences that reflect a non-clinical trial selected population (summarized in supplementary Table S1, available at *Annals of Oncology* online).

The review boards at all participating institutions approved the study, which was conducted according to the principles set forth in the Declaration of Helsinki and the Good Clinical Practice guidelines of the International Conference on Harmonisation. All patients provided written informed consent to participate.

statistical analyses

Distributions of OS were estimated by the Kaplan–Meier product limit method [6, 7]. Laboratory factors were dichotomized into high or low risk according to the lower and upper limits of normal (LLN and ULN, respectively) or median values. For non-laboratory parameters, the median values were used due to the skewed distribution observed. The factors were dichotomized for ease of interpretation of the results derived from the Cox regression model.

Selected clinically relevant baseline factors previously associated with prognosis were assessed for significant association with OS using a univariate Cox regression model. $P \leq 0.05$ was required for inclusion in the subsequent stepwise selection procedure. A multivariate Cox regression model was then used with a stepwise procedure to identify the prognostic factors for OS with a significance level of 0.05 for entry into the model and 0.01 for removal of each factor from the model. The final model was determined based on the Akaike information criterion (AIC) and the model χ^2 score. The combination of the significance levels for entry and removal and the AIC/ χ^2 score were used to derive a model that is limited to the factors that contribute most to the model. Finally, patients were categorized into risk groups based on the number of baseline risk factors, and the median OS was calculated for each group.

The final model was subjected to several validation steps (supplementary Figure S1, available at *Annals of Oncology* online). Internal validation of the predictive performance of the final model was assessed by a bootstrap resampling procedure [20]. Five hundred samples were generated randomly, with replacement from the original data (n = 762). Stepwise Cox regression was used in each sample, with the same selection criteria as the original model. The frequency with which each factor was selected in the resulting model was tabulated. Consistency between the most frequently selected factors and

those in the final model was assessed; the model was deemed internally consistent if the factors were common between the two models.

The parameter estimates for the final model were also validated by randomly generating 500 bootstrap samples from the original population for the final model. For each sample, a Cox regression model was employed to obtain the parameter estimates using the same factors selected in the final model. Summary statistics were computed and compared with the final model. After the final model was established, patients were assessed for the number of risk factors and categorized into good, intermediate or poor prognosis groups, according to the number of risk factors and the relative hazard ratio (HR). The concordance index (C-index) [21] was computed for the final model. A *C*-index ($0.5 \le C$ -index ≤ 1) of 0.5 suggests no predictive discrimination power, while an index of 1.0 indicates perfect discriminatory power. The final model was applied to patients who received prednisone alone (n = 398) and to an independent, external dataset as validation. Statistical analyses were carried out using SAS® Version 9.2 (SAS Institute Inc., Cary, NC) and the receiver operating characteristic analysis was carried out using R Version 2.15.3 (Comprehensive R Archive Network).

results

univariate and multivariate analyses

Fifteen of 19 baseline clinical and laboratory factors were found to be significantly associated with OS ($P \le 0.05$) through a univariate Cox model and were advanced forward (Table 1). A multivariate Cox regression model with a stepwise procedure identified the following 6 of 15 adverse prognostic factors to be the strongest independent predictors of OS: lactate dehydrogenase (LDH), ECOG PS, presence of liver metastases, albumin, alkaline phosphatase (ALP) and time from start of initial androgen-deprivation therapy to start of treatment (Table 2). They were included in the final model. The *C*-index was 0.70 ± 0.014.

model checking and bootstrap validation

To avoid overfitting, the independent factors were limited to those that contributed most to the model based on the AIC and the model χ^2 score. Results from the best subset selection indicated that including additional risk factors was unlikely to improve the model's predictability (supplementary Table S2 and Figure S2, available at *Annals of Oncology* online). The six risk factors were selected most frequently via application of the stepwise Cox regression procedure to the 500 bootstrap samples, suggesting robust internal consistency (supplementary Table S3, available at *Annals of Oncology* online).

risk grouping

Patients were categorized into three risk groups based on the number of baseline risk factors significantly associated with OS and similar HRs (relative to the good prognosis group). Different OS rates were found for each group (Table 3). Patients with zero to one risk factor were in the good prognosis category (n = 369, 46%); the median OS was 21.3 months. Patients with two to three risk factors were in the intermediate prognosis category (n = 321, 40%), the median OS was 13.9 months [HR = 2.3; 95% confidence interval (CI) 1.9–2.8]. Patients with four to six risk factors were in the poor prognosis category (n = 107, 13%), the median OS was 6.1 months (HR = 6.2; 95% CI

Table 1. Baseline clinical and laboratory factors assessed forinclusion in the model, and results from univariate analysis

Baseline risk factor	P-value	HR (95% CI)
LDH [N III N (250 III/l) wareus < III N]	<0.0001	3 01 (2 51 3 60)
ECOC PS (2 variable 0, 1)	<0.0001	2.55(1.08, 2.28)
ECOG PS (2 versus 0-1)	<0.0001	2.55(1.98-5.28)
Liver metastases (present versus absent)	<0.0001	2.55 (1.98-5.24)
ALP [> ULN (160 IU/I) versus \leq ULN]	<0.0001	2.02 (1.69–2.41)
Hemoglobin [≤ LLN (12.5 g/dl) versus > LLN]	<0.0001	1.76 (1.44–2.16)
Albumin (≤4 versus >4 g/dl)	< 0.0001	1.71 (1.43-2.04)
Presence of pain (BPI-SF item $3 \ge 4$ versus	< 0.0001	1.64 (1.38-1.96)
<4)		· · · ·
PSA (>131.4 versus ≤131.4 ng/ml)	< 0.0001	1.59 (1.33–1.90)
Visceral metastases (present versus absent)	< 0.0001	1.46 (1.21–1.75)
Start of androgen-deprivation therapy to	< 0.0001	1.46 (1.21–1.76)
initiation of abiraterone-prednisone		
$(\leq 36 \text{ versus} > 36 \text{ months})$		
Prior radiation therapy (yes versus no)	0.0014	1.40 (1.14–1.72)
End of chemotherapy to initiation	0.0012	1.34 (1.12–1.61)
of abiraterone-prednisone		
$(\leq 3 \text{ versus } > 3 \text{ months})$		
Radiographic progression (with or without	0.0061	1.31 (1.08–1.59)
PSA progression)		
Start of chemotherapy to initiation	0.0035	1.30 (1.09-1.56)
of abiraterone-prednisone		· · ·
$(\leq 12 \text{ versus } > 12 \text{ months})$		
Prior duration of docetaxel treatment	0.0101	1.27 (1.06–1.52)
(≤6 versus >6 months)		
Gleason score (≥8 versus <8) ^a	0.1653	1.14 (0.95–1.38)
Bone and soft-tissue metastases	0.0661	1.19 (0.99–1.43)
(present versus absent) ^a		
Age (≥69 versus <69 years) ^a	0.2722	0.91 (0.76-1.08)
Prior prostatectomy (yes versus no) ^a	0.1576	0.88 (0.74-1.05)

ALP, alkaline phosphatase; BPI-SF, Brief Pain Inventory-Short Form; CI, confidence interval; ECOG PS, Eastern Cooperative Oncology Group performance status; HR, hazard ratio; LDH, lactate dehydrogenase; LLN, lower limit of normal; PSA, prostate-specific antigen; ULN, upper limit of normal.

^aFactors with *P*-values >0.05 were excluded from modeling.

4.8–8.0) (Figure 1A and B). The 2-year survival probabilities were 42%, 14% and 4% for the good, intermediate and poor prognosis categories, respectively.

validation cohorts

The model was evaluated for its discriminative ability in patients in the COU-AA-301 study treated with prednisone alone and an external validation cohort. Of the 398 patients treated with prednisone alone, 193 (48%) were in the good prognosis category, the median OS was 19.7 months; 149 (37%) patients were in the intermediate prognosis category, the median OS was 8.7 months (HR = 3.1; 95% CI 2.4–4.1); and 56 (14%) patients were in the poor prognosis category (four to six risk factors), the median OS was 5.3 months (HR = 5.9; 95% CI 4.1–8.4) (Figure 1C and D). The *C*-index using the prednisone alone treatment arm was 0.72 ± 0.019 .

Baseline risk factor	Factor estimate (standard error)	P-value	HR
LDH [> ULN (250 IU/l) versus ≤ ULN]	0.84 (0.10)	<0.0001	2.31
ECOG PS (2 versus 0–1)	0.78 (0.14)	< 0.0001	2.19
Liver metastases	0.69 (0.13)	< 0.0001	2.00
(present versus absent)			
Albumin (≤4 versus >4 g/dl)	0.43 (0.09)	< 0.0001	1.54
ALP [> ULN (160 IU/l) versus \leq ULN]	0.32 (0.10)	0.0016	1.38
Start of androgen-deprivation	0.26 (0.10)	0.0078	1.30
therapy to initiation of abiraterone-			
prednisone (\leq 36 versus >36 months)			

ALP, alkaline phosphatase; ECOG PS, Eastern Cooperative Oncology Group performance status; HR, hazard ratio; LDH, lactate dehydrogenase; ULN, upper limit of normal.

Table 3. Definition of risk groups by pooling groups with similarrelative HRs					
Risk group	No. of baseline risk factors	No. of patients (%)	HR relative to patients with 0 risk factors		
Good prognosis ($n = 369$)	0 1	152 (19.07) 217 (27.23)	- 1.61		
Intermediate prognosis	2	192 (24.09)	2.73		
(<i>n</i> = 321)	3	129 (16.19)	3.79		
Poor prognosis ($n = 107$)	4	85 (10.66)	7.04		
	5	20 (2.51)	24.64		
	6	2 (0.25)	53.14		
HR, hazard ratio.					

The external validation cohort was an independent cohort of 286 patients with mCRPC who were sequentially treated in a routine clinical care setting with abiraterone–prednisone after docetaxel at 11 centers in Canada. Sixty-three patients (22%) were categorized as having a good prognosis, 146 patients (51%) an intermediate prognosis and 77 patients (27%) a poor prognosis. Analysis of the validation set confirmed the ability of the model to prognosticate for OS: for the good prognosis group, the median OS was 23.9, and for the intermediate and poor risk groups, the median OS was 16.2 months (relative HR = 1.9; 95% CI 1.3–2.9) and 8.2 months (relative HR = 4.1; 95% CI 2.7–6.4), respectively (Figure 1E and F). The *C*-index using the 286 patients was 0.69 ± 0.023 .

discussion

A prognostic index model was developed for post-docetaxel mCRPC using six factors that are highly associated with OS. The factors are available during routine patient treatment and enable patients to be categorized into three distinct risk groups

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(poor, intermediate and good prognosis). The model was validated with an external cohort of 286 patients treated with abiraterone–prednisone outside of a clinical trial and, notably, the model enabled similar stratification of the prednisone alone group. The diversity of the patient populations studied suggests that our findings are generalizable. Ravi et al. [22] have provided further external validation of the model by applying the prognostic index to an independent cohort of 94 patients treated with abiraterone–prednisone after docetaxel [*C*-index = 0.71 (95% CI 0.60–0.80)].

Applications of the model include the homogenization of the risk of death of patients to be enrolled in clinical trials and may be a useful tool in addition to treatment guidelines to help physicians determine appropriate follow-up and monitoring of patients with poor prognosis. The St Gallen conference guidelines recommend baseline examinations that measure the readily available factors that were included in our analysis. Some factors, such as duration of initial androgen-deprivation therapy and established prognostic factors, like hemoglobin, LDH and ALP, could potentially aid in prognostic stratification of patients with mCRPC, but their value in guiding treatment decisions is not well established [23]. Prospective studies are needed to determine the utility of the model for selecting therapies for patients with mCRPC and poor prognostic factors.

We purposefully developed this model using continuous variables dichotomously and weighting each risk factor equally in order to facilitate ease of use in the clinic. Prognostic models employing an index design similar to that developed here have seen widespread application in patients with metastatic renal cell cancer to classify patients into prognostic risk groups and in patients with non-Hodgkin's lymphoma to classify risk of death, complete response and relapse risk [24, 25]. These prognostic indices have evolved over time to incorporate changes in treatment and contemporary outcomes [26, 27]. Armstrong et al. [28] and Halabi et al. [29, 30] developed prognostic nomograms in the post-docetaxel and first- and second-line chemotherapy mCRPC settings, and identified several similar pretreatment clinical prognostic factors that are associated with survival. More recent studies of docetaxel for mCRPC report greater survival than previous studies; however, these do not account for the introduction of recent life-prolonging therapies. Other prognostic factors warranting further study include bone-associated biomarkers [31], bone-specific ALP and urinary N-telopeptide, although these factors are not usually obtained in routine practice [28, 32-34]. Enumeration of circulating tumor cells with the Veridex system [19] and of androgen levels as determined by ultrasensitive assays have also been identified as independently prognostic in patients with mCRPC treated with abirateroneprednisone [18]; these tests were not included in the development of this model because they are not routinely carried out or widely available.

conclusion

We have developed a contemporary prognostic index model, composed of six routinely available and readily assessable factors, that categorizes patients with mCRPC treated with abiraterone-prednisone into distinct risk groups. Further external



Figure 1. Kaplan-Meier curves for overall survival and relative hazard ratio by risk group, as estimated by the stepwise final model for patients treated with abiraterone-prednisone (A and B) and prednisone alone (C and D) and for the independent validation dataset (E and F). CI, confidence interval; OS, overall

validation of our model is required. This model could be useful in clinical practice to aid in the determination of patient prognosis so that follow-up and monitoring may be planned accordingly, and may aid in patient stratification in clinical trials.

survival. ^aVersus patients with good prognosis.

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disclosure

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Association between PSA kinetics and cancer-specific mortality in patients with localised prostate cancer: analysis of the placebo arm of the SPCG-6 study

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Background: The prognostic value of prostate-specific antigen (PSA) kinetics in untreated prostate cancer (PCa) patients is debatable. We investigated the association between PSA doubling time (PSAdt), PSA velocity (PSAvel) and PSAvel risk count (PSAvRC) and PCa mortality in a cohort of patients with localised PCa managed on watchful waiting.

Patients and methods: Patients with clinically localised PCa managed observationally, who were randomised to and remained on placebo for minimum 18 months in the SPCG-6 study, were included. All patients survived at least 2 years and had a minimum of three PSA determinations available. The prognostic value of PSA kinetics was analysed and patients were stratified according to their PSA at consent: \leq 10, 10.1–25, and >25 ng/ml. Cumulative incidences of PCa-specific mortality were estimated with the Aalen-Johansen method.

Results: Two hundred and sixty-three patients were included of which 116, 76 and 71 had a PSA at consent \leq 10, 10.1–25, and >25 ng/ml, respectively. Median follow-up was 13.6 years. For patients with PSA at consent between 10.1 and 25 ng/ml, the 13-year risks of PCa mortality were associated with PSA kinetics: PSAdt \leq 3 years: 62.0% versus

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