

Short Term Synaptic Depression Imposes a Frequency Dependent Filter on Synaptic Information Transfer

Robert Rosenbaum^{1,2*}, Jonathan Rubin^{1,2}, Brent Doiron^{1,2}

1 Mathematics, University of Pittsburgh, Pittsburgh, Pennsylvania, United States of America, **2** Center for the Neural Basis of Cognition, Pittsburgh, Pennsylvania, United States of America

Abstract

Depletion of synaptic neurotransmitter vesicles induces a form of short term depression in synapses throughout the nervous system. This plasticity affects how synapses filter presynaptic spike trains. The filtering properties of short term depression are often studied using a deterministic synapse model that predicts the mean synaptic response to a presynaptic spike train, but ignores variability introduced by the probabilistic nature of vesicle release and stochasticity in synaptic recovery time. We show that this additional variability has important consequences for the synaptic filtering of presynaptic information. In particular, a synapse model with stochastic vesicle dynamics suppresses information encoded at lower frequencies more than information encoded at higher frequencies, while a model that ignores this stochasticity transfers information encoded at any frequency equally well. This distinction between the two models persists even when large numbers of synaptic contacts are considered. Our study provides strong evidence that the stochastic nature of neurotransmitter vesicle dynamics must be considered when analyzing the information flow across a synapse.

Citation: Rosenbaum R, Rubin J, Doiron B (2012) Short Term Synaptic Depression Imposes a Frequency Dependent Filter on Synaptic Information Transfer. *PLoS Comput Biol* 8(6): e1002557. doi:10.1371/journal.pcbi.1002557

Editor: Olaf Sporns, Indiana University, United States of America

Received: February 11, 2012; **Accepted:** April 25, 2012; **Published:** June 21, 2012

Copyright: © 2012 Rosenbaum et al. This is an open-access article distributed under the terms of the Creative Commons Attribution License, which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Funding: This work was supported by NIH-1R01NS070865-01A1, NSF-DMS-1021701, and NSF-DMS-1121784; and B Doiron is a Sloan Research Fellow. The funders had no role in study design, data collection and analysis, decision to publish, or preparation of the manuscript.

Competing Interests: The authors have declared that no competing interests exist.

* E-mail: robertr@pitt.edu

Introduction

Synapses act as information gates in neuronal networks. Presynaptic action potentials are communicated to postsynaptic neurons by causing synaptic neurotransmitter vesicles to release their contents, which then bind to receptors on a postsynaptic neuron's membrane, evoking a transient change in membrane conductance. After a vesicle is released, it typically takes several hundred milliseconds for it to be replaced at a synaptic contact (see Fig. 1 for a schematic of synaptic release and recovery). This refractoriness induces a form of short term synaptic depression that alters the filtering properties of synapses [1]. An accurate description of synaptic vesicle dynamics and their impact on information transfer is necessary for a thorough understanding of coding in neuronal networks.

A widely used model of synaptic depression treats vesicle release and recovery as deterministic processes [2–6]. While this deterministic model accurately describes the trial-averaged synaptic response to a presynaptic spike train presented repeatedly to a cell [7–11], it fails to capture the variability introduced at each trial by the probabilistic nature of vesicle release and recovery [12]. Regardless, the model has been used in studies for which neural variability and information transfer are central themes [13–18]. The aim of our paper is to determine the impact (if any) of stochastic vesicle dynamics on the filtering properties of depressing synapses.

Past studies have begun to address this aim by considering how variability from stochastic vesicle release and recovery affects the amount of information transmitted through a synapse as well as the firing rate of a postsynaptic cell [12,19,20], but a thorough investigation of the impact of stochastic vesicle dynamics on

synaptic filtering has not been performed. We derive a compact description of the filters imposed by short term synaptic depression when stochastic vesicle dynamics are taken into account and when they are ignored. We find that variability introduced by stochastic vesicle dynamics plays a fundamental role in shaping the way in which depressing synapses filter presynaptic information. In particular, a model that ignores this variability transmits presynaptic information encoded at any frequency with the same fidelity [16,17]. In contrast, a model that captures this variability reduces overall information transmission, and transmits quickly varying signals with higher fidelity than slowly varying signals. Differences between the two models persist over a broad range of physiologically motivated parameter values, even when a large number of synaptic contacts is considered and even at the population level. Our results suggest important implications for how signals encoded at different timescales are propagated through the nervous system and show that synaptic variability must be taken into account to accurately address such questions.

Results

We study the synaptic filter induced by short term depression with both a stochastic model and a deterministic model of synaptic vesicle dynamics (see Fig. 2A–D for an illustration and Methods for a detailed discussion). For both models, we consider a presynaptic spike train, $I(t)$, with rate ν that induces a postsynaptic conductance,

$$g(t) = \sum_j w_j \alpha(t - t_j).$$

Author Summary

Neurons communicate through electro-chemical connections called synapses. Action potentials in a presynaptic neuron cause neurotransmitter vesicles to release their contents which then bind to nearby receptors on a postsynaptic neuron's membrane, transiently altering its conductance. After it is released, the replacement of a neurotransmitter vesicle takes time and the depletion of vesicles can prevent subsequent action potentials from eliciting a postsynaptic response, an effect that represents a form of short term synaptic depression. When a vesicle is available for release, an action potential elicits its release probabilistically and depleted vesicles are replenished randomly in time, making the transmission of presynaptic signals inherently unreliable. We analyze a mathematical model of vesicle release and recovery to understand how signals encoded in sequences of presynaptic action potentials are reflected in the fluctuations of a postsynaptic neuron's conductance. We find that slow modulations in the rate of presynaptic action potentials are more difficult for a postsynaptic neuron to detect than faster modulations. This phenomenon is only observed when randomness in vesicle release and replacement is taken into account. Thus, by including stochasticity in the workings of synaptic dynamics we give new qualitative understanding to how information is transferred in the nervous system.

Here, t_j is the time of the j th presynaptic spike, w_j is the number of vesicles released by the j th presynaptic spike, and $\alpha(t)$ represents the time course of conductance induced by the release of a single synaptic vesicle. The presynaptic cell makes M contacts with the postsynaptic cell. We make a simplifying assumption that each contact contains only one release site, so that a single presynaptic action potential can release at most one vesicle per contact [21], hence $0 \leq w_j \leq M$. Alternately, to model biological settings where this single vesicle hypothesis is violated [22,23], M can be interpreted as the total number of release sites across all contacts (see Discussion). We rescale conductance units so that $\int_0^\infty \alpha(t) dt = 1$. This rescaling causes $g(t)$ to have dimension time^{-1} but simplifies the exposition.

In the stochastic model of vesicle dynamics [12,19,24,25], a presynaptic spike releases each available vesicle at each contact independently with probability p_r . After a contact releases its vesicle, it is unavailable to release again until the vesicle is replaced, a process known as recovery. The waiting time until the vesicle is replaced follows an exponential distribution with mean τ_u (Fig. 2B,C). For the deterministic model of vesicle dynamics [2], the number of available vesicles is treated as a continuous variable where a proportion p_r of the total available vesicles are released by each presynaptic spike and the number of available vesicles increases exponentially towards M with timescale τ_u between releases (Fig. 2D). Stochasticity in the conductance, $g(t)$, produced by the deterministic model is introduced solely by the stochasticity in the input, $I(t)$. Several presentations of the same realization of

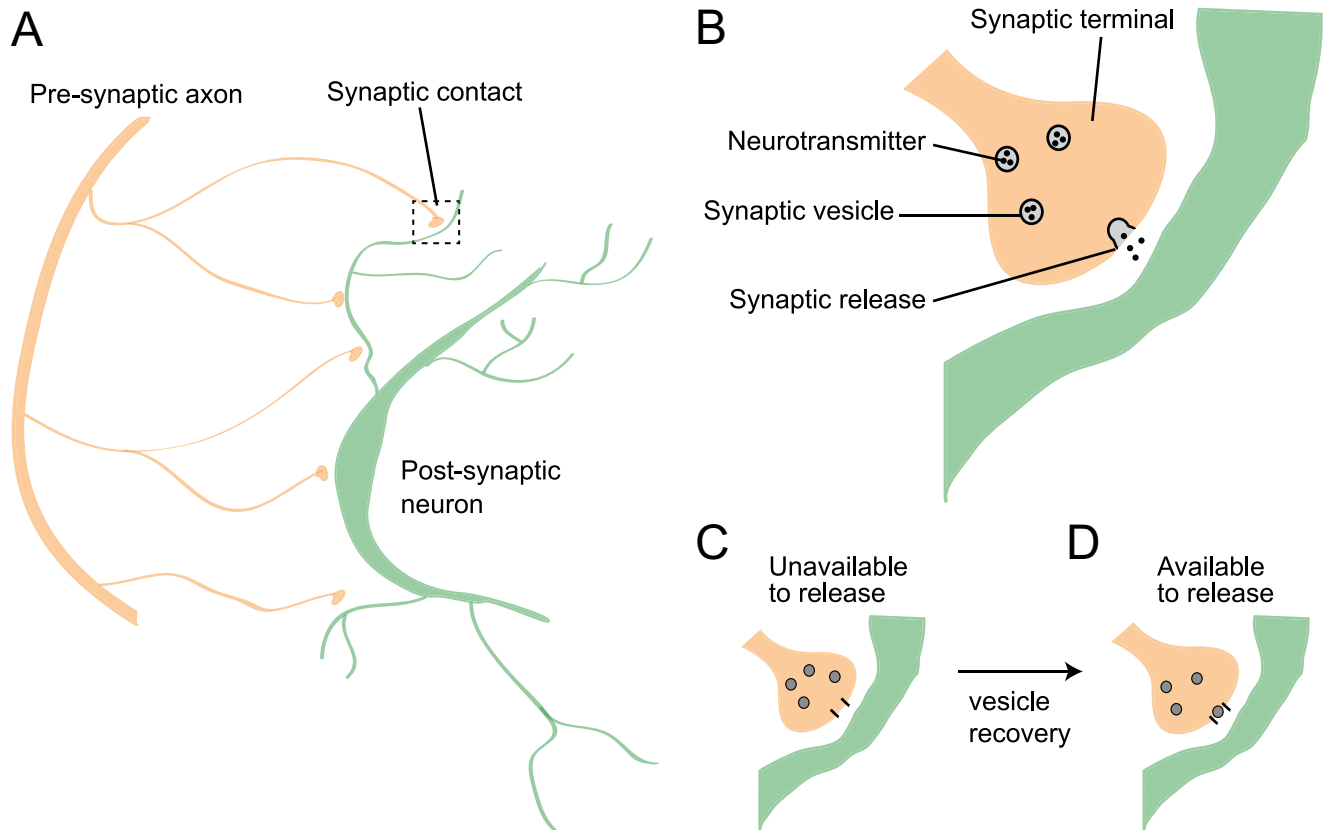


Figure 1. Synaptic vesicle dynamics. (A) The axon of a presynaptic neuron (orange) makes $M = 5$ synaptic contacts onto a postsynaptic neuron (green). (B) Synaptic vesicles in the synaptic terminal of the presynaptic neuron contain neurotransmitter molecules. A presynaptic action potential releases these neurotransmitter molecules with some probability, p . Once released, these molecules bind to the postsynaptic neuron's membrane and cause a transient change in membrane conductance. (C,D) After a vesicle is released, the synapse enters a refractory state where it is unavailable to release additional neurotransmitter until it recovers by replacing the released vesicle. doi:10.1371/journal.pcbi.1002557.g001

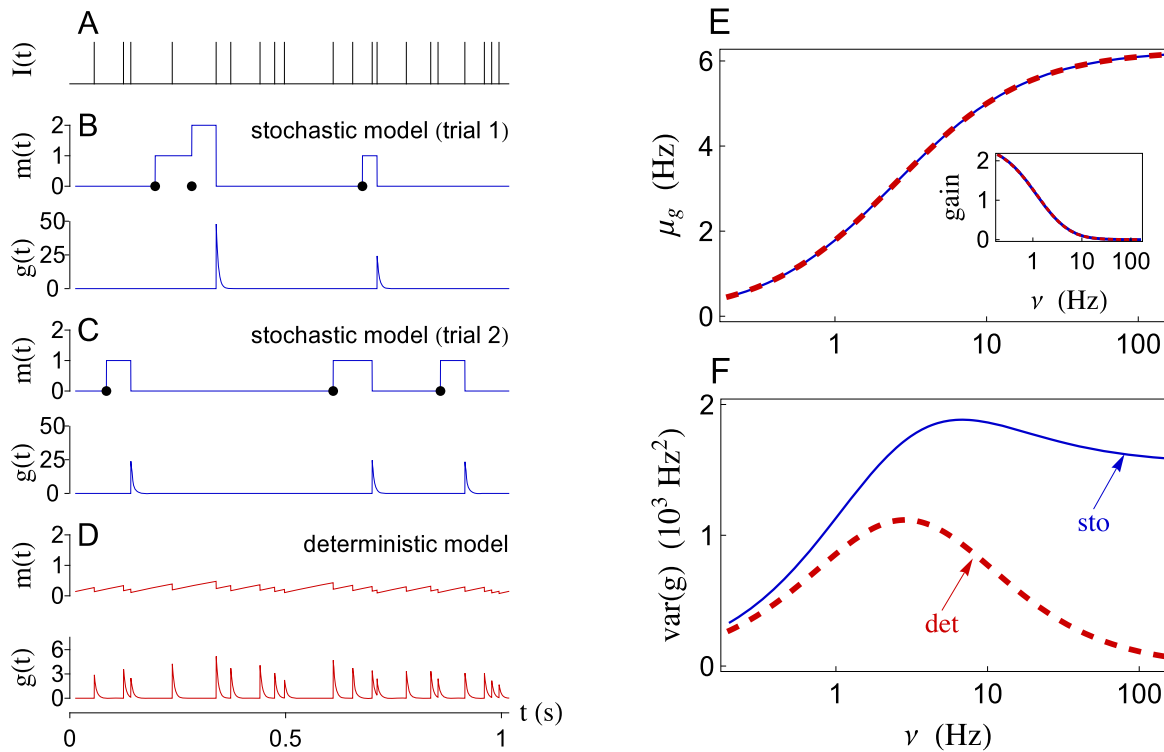


Figure 2. Stochastic versus deterministic models of short term depression. (A) An example presynaptic spike train, $I(t)$. Each vertical bar represents an action potential. (B) The number of synaptic vesicles, $m(t)$, available for release and the conductance, $g(t)$, induced in the postsynaptic cell for one realization of the stochastic model. Filled circles in (B) represent vesicle recovery events. (C) A second realization of the stochastic model with the same input. Observe in (B) and (C) that the number of vesicles released by the stochastic model during one second is primarily determined by the number of recovery events during that second and does not reflect the number of presynaptic spikes. (D) The number of synaptic vesicles and the conductance induced by the deterministic model with the input from (A). Parameters in (A–D) were chosen for illustrative purposes as $M=2$, $\tau_u=650\text{ms}$, $p_r=0.5$, and $\tau_x=5\text{ms}$. (E) The steady state mean conductance, μ_g , as a function of the presynaptic firing rate, ν . The inset shows the gain, $d\mu_g/d\nu$. (F) The steady state variance of $g(t)$ as a function of ν for the deterministic (solid blue) and stochastic (dashed red) models of vesicle dynamics with Poisson inputs. Variability in the deterministic model is introduced only by variability in the input, $I(t)$. Synaptic parameters for (E–F) and for all subsequent figures are given in Table 1. doi:10.1371/journal.pcbi.1002557.g002

$I(t)$ produce the same $g(t)$ for the deterministic model, but not for the stochastic model (Fig. 2A–D).

The conductance produced by the deterministic model represents the quantity that would be obtained by presenting the same realization of $I(t)$ to the stochastic model over several trials, then computing the trial-averaged conductance. Despite the agreement of their trial-averages, though, individual realizations of the two models differ substantially. The deterministic model responds to every presynaptic input, but releases a fractional number of vesicles at each response (Fig. 2D). In contrast, the stochastic model responds to only a few inputs, but releases a larger, quantal number of vesicles at each response (Fig. 2B,C).

The steady state mean conductance induced by a presynaptic spike train $I(t)$ with rate $\nu=\langle I(t) \rangle$ is given by $\mu_g = \lim_{t \rightarrow \infty} \langle g(t) \rangle = p_r \nu M / (1 + p_r \tau_u \nu)$ for both the stochastic and deterministic models of vesicle dynamics (Fig. 2E and Eq. (25)). The degree to which a small shift of the presynaptic rate is reflected in a shift of the steady state mean conductance is measured by the gain,

$$\frac{d\mu_g}{d\nu} = \frac{p_r M}{(1 + \nu \tau_u p_r)^2}, \quad (1)$$

which is a decreasing function that decays to zero as ν increases, a well-known effect that is due to the saturation of the mean

conductance for large presynaptic firing rates (see Fig. 2E, inset and [2,3,26]). However, the gain only measures changes in the steady state mean of $g(t)$ after a sustained shift in the mean of $I(t)$, whereas the signal processing properties of a synapse also depend on the temporal response of $g(t)$ to transient fluctuations in $I(t)$ [3,10,27,28]. Below, we use a cross-spectral measure to quantify the temporal response properties of $g(t)$.

The information processing capabilities of a synapse depend not only on the response of $g(t)$ to temporal fluctuations in $I(t)$, but also on the temporal and trial-to-trial variability of $g(t)$. Noise introduced by stochastic vesicle release and recovery leads to larger variability in $g(t)$, as measured by its variance (Fig. 2F). However, the variance alone does not capture the timescale over which this variability occurs. Below, we use a power-spectral measure to describe the variability of $g(t)$ over different timescales.

Synaptic filtering of a Poisson presynaptic spike train

To gain an intuition for the signal processing properties of depressing synapses, we first study the case of a single Poisson presynaptic spike train, $I(t)$, with constant rate ν . Since a homogeneous Poisson process has equal power at every frequency, this approach allows us to investigate synaptic filtering at all frequencies simultaneously. Later, we will consider the response to an inhomogeneous Poisson process whose rate encodes a signal.

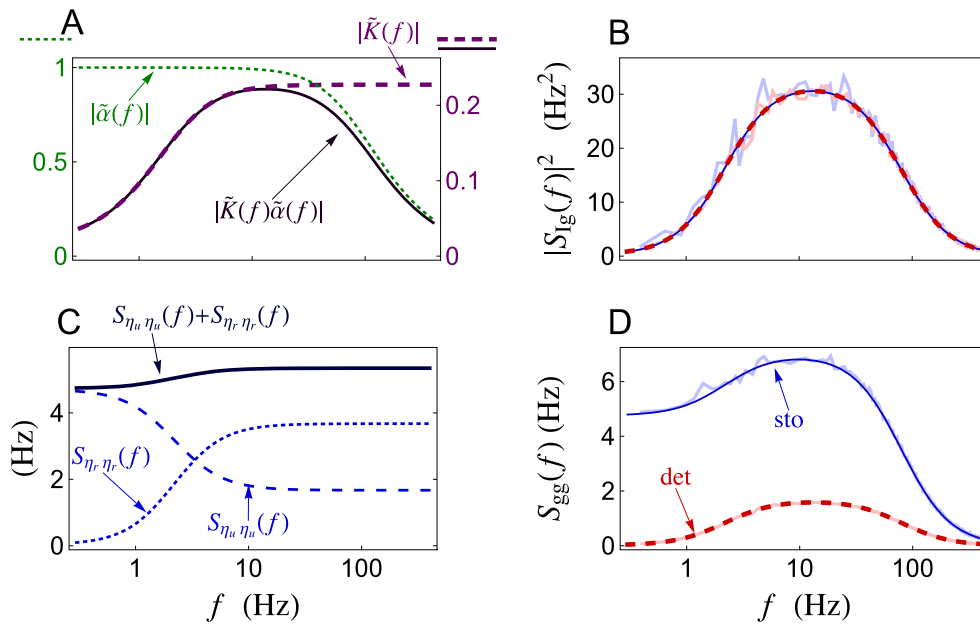


Figure 3. Synaptic filtering of a single Poisson presynaptic spike train. (A)–(B) The low-pass filter, $\tilde{\alpha}(f)$, and the high-pass filter, $\tilde{K}(f)$, are multiplied with the presynaptic rate (cf. Eq. (2)) to determine the band-pass cross-spectrum, $S_{I_g}(f)$, between a Poisson presynaptic spike train, $I(t)$, and postsynaptic conductance, $g(t)$. The cross-spectrum is identical for the stochastic (solid blue) and deterministic (dashed red) models. (C)–(D) The power spectrum, $S_{g_g}(f)$, of the conductance is larger for the stochastic model than the deterministic model due to the additive terms, $S_{\eta_r, \eta_r}(f)$ and $S_{\eta_u, \eta_u}(f)$, that quantify the increase in variability due to stochastic vesicle release and recovery (see Eq. (3)). For this and all subsequent figures, solid blue lines and dashed red lines show plots obtained from closed form expressions for the stochastic and deterministic models, respectively. Light blue and light red lines indicate simulations of the stochastic and deterministic models, respectively. doi:10.1371/journal.pcbi.1002557.g003

The magnitude of the response of the conductance, $g(t)$, at frequency f to fluctuations in the input, $I(t)$, is quantified by the cross-spectrum, $S_{I_g}(f)$, between these quantities (see Methods). For both the deterministic and stochastic models of vesicle dynamics, the cross-spectrum is given by (see Eq. (25) in Methods)

$$S_{I_g}(f) = \tilde{\alpha}(f)\tilde{K}(f)v, \quad (2)$$

where $\tilde{u}(f) = \int u(t)e^{-2\pi ift} dt$ denotes the Fourier transform and $\tilde{K}(f)$ is a kernel that captures the filtering properties of synaptic depression (see Eq. (20) in Methods and Fig. 3A). The fact that $S_{I_g}(f)$ is identical for the stochastic and deterministic models can be understood intuitively by noting that stochasticity in vesicle dynamics is uncorrelated from $I(t)$ and therefore does not contribute to the covariability of $I(t)$ and $g(t)$. It should be noted that, while Eq. (2) is exact for the deterministic model, it is an approximation for the stochastic model (see Methods), which is validated by simulations (Fig. 3B).

The shape of $S_{I_g}(f)$ can be understood by its components in Eq. (2). The low-pass filter, $\tilde{\alpha}(f)$, which captures postsynaptic channel dynamics, suppresses power at frequencies higher than $1/(2\pi\tau_x)$ (see Fig. 3A and [29]). The high-pass filter $\tilde{K}(f)$, which captures the deterministic dynamics of short term depression, suppresses power at frequencies lower than $1/(2\pi\tau_0) = (1 + p_r v \tau_u)/(2\pi\tau_u)$ (see Fig. 3A, Methods and [17]). Their product, which determines $S_{I_g}(f)$ through Eq. (2), is then band-pass with most of its power at frequencies between $1/(2\pi\tau_0)$ and $1/(2\pi\tau_x)$ (Fig. 3B). Thus, only fluctuations in the presynaptic input within this frequency band are reflected faithfully by fluctuations in the postsynaptic conductance.

The low-frequency limit of $S_{I_g}(f)$ is nearly zero for the parameter values chosen in Table 1 (Fig. 3B). This can be

explained by noting that the zero-frequency cross-spectrum is related to the gain by [30]

$$S_{I_g}(0) = v \frac{d\mu_g}{dv}.$$

For large v , the mean conductance saturates and the gain decays to zero like v^{-2} (see Eq. (1) and Fig. 2E). Thus, $S_{I_g}(0) \sim v^{-1}$ which decays to zero for large v (Fig. 4Ai). More specifically, $S_{I_g}(0) \approx 0$

Table 1. Table of synaptic parameters.

Name	Definition	Default value
τ_u	timescale of vesicle recovery	800ms
M	number of contacts between a pre- and postsynaptic cell	5
p_r	probability of release when vesicle is available	0.5
v	presynaptic rate	25 Hz
$\alpha(t)$	synaptic activation kernel	$\Theta(t)\tau_x^{-1}e^{-t/\tau_x}$
τ_x	time constant of postsynaptic channels	2ms
σ_s	bandwidth of rate-coded signal	0.1 Hz
D_s	peak power of rate-coded signal	20 Hz
c	noise correlation between presynaptic spike trains	0.1

Parameters for synapses and presynaptic spike trains. These parameter values are used in all figures unless otherwise indicated. Here, $\Theta(t)$ represents the Heaviside step function.

doi:10.1371/journal.pcbi.1002557.t001

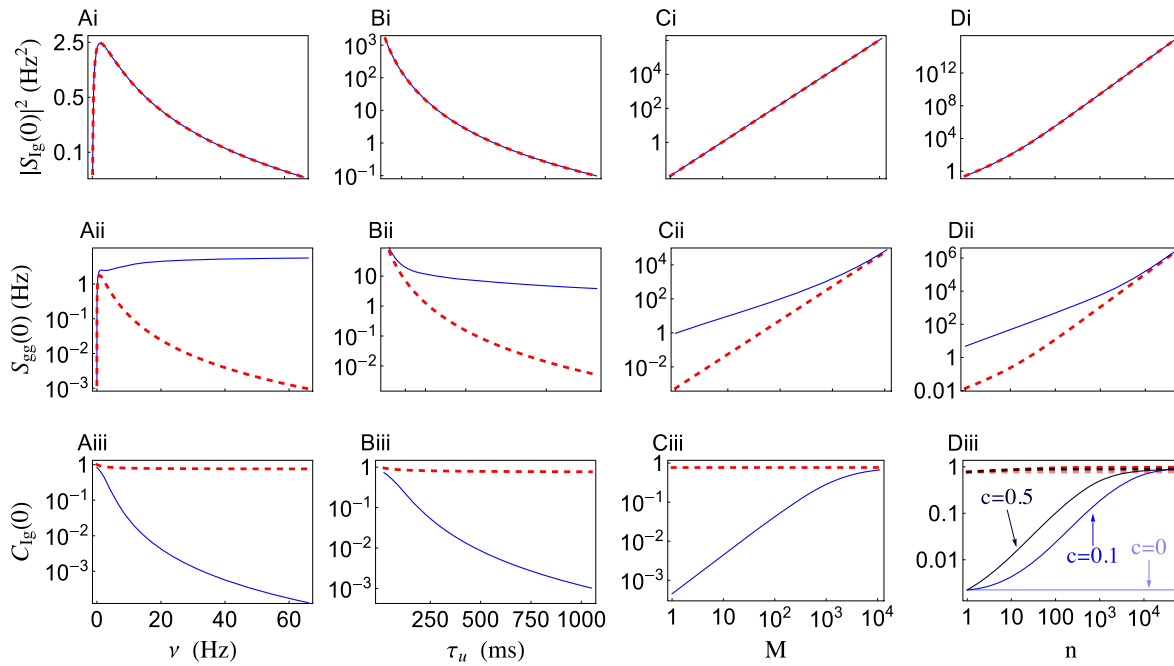


Figure 4. Low frequency signal transfer in a variety of parameter regimes. Low frequency cross-spectrum ($S_{I_g}(0)$), auto-spectrum ($S_{g_g}(0)$), and coherence ($C_{I_g}(0)$) between a Poisson presynaptic spike train, $I(t)$, and postsynaptic conductance, $g(t)$, plotted as a function of the presynaptic rate, ν (**Ai–iii**), the vesicle recovery timescale, τ_u (**Bi–iii**), the number of synaptic contacts, M (**Ci–iii**), and presynaptic population size, n (**Di–iii**). Columns A–C are for a single presynaptic spike train ($n=1$). The zero-frequency coherence in Diii is shown for three values of the presynaptic correlation coefficient: $c=0, 0.1$, and 0.5 . The power spectrum and coherence predicted by the stochastic model (solid blue) and the deterministic model (dashed red) disagree by orders of magnitude unless ν is small, M is large, τ_u is small, or n is large with $c>0$. doi:10.1371/journal.pcbi.1002557.g004

when vesicles become depleted, which occurs when release is faster than recovery, i.e., $p_r \nu \gg M/\tau_u$. Note, though, that $S_{I_g}(f)$ is larger for higher frequencies, meaning that faster fluctuations in $I(t)$ cause larger transient fluctuations in $g(t)$ when compared to changes in the steady state mean conductance, μ_g , caused by static changes in ν [3,10,27,28].

The trial-to-trial and temporal variability of the conductance at frequency f is quantified by its power spectrum, $S_{g_g}(f)$, which is given by (see Eq. (25) in Methods)

$$S_{g_g} = (1 + D_0) |\tilde{\mathbf{K}} \tilde{\mathbf{a}}|^2 \nu + |\tilde{\mathbf{a}}|^2 (S_{\eta_u \eta_u} + S_{\eta_r \eta_r}). \quad (3)$$

Here D_0 is a constant that represents variability introduced by the interaction of Poisson input with deterministic vesicle dynamics, $S_{\eta_u \eta_u}(f)$ captures variability introduced by stochastic recovery, and $S_{\eta_r \eta_r}(f)$ captures variability introduced by probabilistic vesicle release. For the deterministic model, $S_{\eta_r \eta_r}(f) = S_{\eta_u \eta_u}(f) = 0$, but $S_{\eta_r \eta_r}(f)$ and $S_{\eta_u \eta_u}(f)$ are positive for the stochastic model (see Methods and Fig. 3C). As a result, the stochastic model predicts a larger power spectrum than the deterministic model (Fig. 3D). The decay of $S_{g_g}(f)$ at high frequencies is due to the low-pass nature of the synaptic conductance kernel, $\tilde{\alpha}(f)$ (see Fig. 3A and [29]).

The power spectrum predicted by the two models differs most significantly at low frequencies, where it is nearly zero for the deterministic model but much larger for the stochastic model (Fig. 3D). This can be understood by noting that [30]

$$S_{g_g}(0) = \lim_{T \rightarrow \infty} \text{var}(N_x(T))/T$$

where $N_x(T)$ is the number of vesicles released in a window of

length T . For the parameter values in Table 1, $p_r \nu \gg M/\tau_u$ so that vesicles are mostly depleted and therefore the number of vesicles released in a large time window is determined largely by the number of recovery events during that window (Fig. 2A–D). For the stochastic model, recovery events at each contact occur as a Poisson process with rate $1/\tau_u$. Since there are M contacts and a Poisson process has power equal to its rate, $S_{g_g}(0) \approx M/\tau_u$ when ν is large. This intuition is confirmed by noting that $S_{g_g}(0) = M/\tau_u + O(\nu^{-1})$ for the stochastic model. In contrast, for the deterministic model, recovery is deterministic and therefore the amount of neurotransmitter taken up, and hence released, over a large time window has a small variance. This is confirmed by noting that $S_{g_g}(0) \sim \nu^{-3}$ for the deterministic model and therefore approaches zero for large ν . For the synaptic parameters in Table 1, the power spectra produced by the stochastic and deterministic models disagree for ν larger than a few Hz (Fig. 4Aii).

The fidelity with which fluctuations in the postsynaptic conductance, $g(t)$, reflect fluctuations of the input, $I(t)$, at frequency f is quantified by their coherence

$$C_{I_g}(f) = \frac{|S_{I_g}(f)|^2}{S_{II}(f) S_{g_g}(f)}$$

where $S_{II}(f) = \nu$ is the power spectrum of the Poisson input. Since $S_{I_g}(f)$ is identical for the two models, but $S_{g_g}(f)$ is larger for the stochastic model (Fig. 3B,D), it follows that $C_{I_g}(f)$ is smaller for the stochastic model (Fig. 5). We now investigate the differences between the coherences produced by the two models in more depth.

Since $S_{\eta_u \eta_u}(f) = S_{\eta_r \eta_r}(f) = 0$ for the deterministic model, the cross-spectrum, $S_{I_g}(f)$, and power spectrum, $S_{g_g}(f)$, are proportional to one another (see Eqs. (2) and (3)) so that dividing them

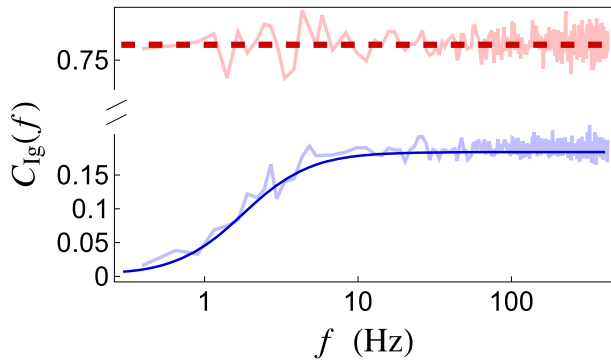


Figure 5. Coherence between a single presynaptic spike train and the postsynaptic conductance it induces. The coherence, $C_{I_g}(f)$, between a Poisson presynaptic spike train, $I(t)$, and the resulting postsynaptic conductance, $g(t)$. The stochastic model (solid blue) yields a high pass coherence that is dramatically smaller than the flat coherence predicted by the deterministic model (dashed red). doi:10.1371/journal.pcbi.1002557.g005

gives a flat coherence (i.e., a coherence that does not depend on f , Fig. 5 and [16,17]),

$$C_{I_g}^{\text{det}}(f) = (1 + D_0)^{-1}.$$

Here and in subsequent expressions, a *det* (*sto*) superscript indicates identities for the deterministic (stochastic) model. Synaptic variability in the stochastic model increases the power spectrum, giving a frequency-dependent coherence

$$C_{I_g}^{\text{sto}}(f) = \left(1 + D_0 + \frac{S_{\eta_r \eta_r}(f) + S_{\eta_u \eta_u}(f)}{|\bar{K}(f)|^2 v} \right)^{-1},$$

which is high-pass (Fig. 5). Thus, stochastic vesicle dynamics introduce high-pass frequency dependence into the fidelity of a synaptic filter.

In addition to introducing frequency dependence, stochastic vesicle dynamics also decrease the coherence substantially, especially at lower frequencies where the coherence is nearly zero for the stochastic model (Fig. 5). The fact that coherence is small at low frequencies for the stochastic model can be understood intuitively through the following relation [30],

$$C_{I_g}(0) = \lim_{T \rightarrow \infty} \text{corr}(N_I(T), N_X(T))^2,$$

where $\text{corr}(N_x(T), N_I(T))$ is the Pearson correlation coefficient between the number of presynaptic spikes, $N_I(T)$, and the number of vesicles released, $N_X(T)$, in a window of length T . When $p_r v \gg M/\tau_u$, synapses are mostly depleted in the steady state. As a result, the number of vesicles released during a long time interval is determined primarily by the number of recovery events in that time window and hence mostly independent of the number of presynaptic spikes (Fig. 2A–C and [31]). Therefore, for the stochastic model, the number of vesicles released over a long time window is uncorrelated from the number of presynaptic spikes and, as a result, $C_{I_g}(0)$ is small.

These intuitions are confirmed by appealing to the asymptotic expressions derived for the cross-spectrum and power spectrum above. For the stochastic model, $S_{I_g}(0) \sim v^{-1}$ and $S_{g_g}(0) \sim M/\tau_u + v^{-1}$ when $p_r v \gg M/\tau_u$. Since $S_{II}(f) = v$ for Poisson input, it is then clear that

$$C_{I_g}(0) = \frac{|S_{I_g}(0)|^2}{S_{II}(0)S_{g_g}(0)} \sim v^{-3}$$

for the stochastic model when $p_r v \gg M/\tau_u$. For the deterministic model, however, $S_{I_g}(0) \sim v^{-1}$, $S_{II}(0) = v$, and $S_{g_g}(0) \sim v^{-3}$ so that $C_{I_g}(0) = |S_{I_g}(0)|^2 / (S_{II}(0)S_{g_g}(0))$ approaches a positive constant for v sufficiently larger than $M/(\tau_u p_r)$. For the parameter values in Table 1, the coherences for the stochastic and deterministic models disagree substantially when v is more than a few Hz (Fig. 4Aiii).

The disagreement between the stochastic and deterministic models is most dramatic when $p_r v \gg M/\tau_u$ since the postsynaptic response is determined primarily by vesicle recovery dynamics in this regime, as discussed above. In the figures considered so far, we have used $\tau_u = 800\text{ms}$, motivated by measurements of pyramidal-to-pyramidal synapses in rodent neocortex [2,19]. However, both shorter and longer time constants have also been reported in cortex [5,7,8,32,33]. When other parameters are set to the values from Table 1, the two models disagree substantially when $\tau_u > 100\text{ms}$ (see Fig. 4Bi–iii).

A proposed justification for using a deterministic model of vesicle dynamics is that stochasticity introduced at each contact averages out when a presynaptic cell makes several contacts [17]. The number, M , of contacts a presynaptic cell makes with a single postsynaptic cell varies greatly across cell subtypes and brain regions. Rodent and cat pyramidal cells in the hippocampus and neocortex typically make $M = 1\text{--}12$ contacts onto other pyramidal cells or onto interneurons. Interneurons in the same regions make $M = 1\text{--}17$ contacts onto pyramidal cells. On the other hand, the Calyx of Held synapse can make more than $M = 700$ contacts onto a single postsynaptic target in the rodent auditory brainstem and Purkinje cells can receive over $M = 500$ contacts from single presynaptic cells in the rodent cerebellum (see [34] for values of M measured in various animals and synapses). When other parameters are set to the values from Table 1, the stochastic and deterministic models disagree substantially for $M < 1000$ (see Fig. 4Ci–iii).

In summary, over a broad range of synaptic parameters, stochastic vesicle dynamics both attenuate and impart a high-pass nature to the coherence between a pre-synaptic spike train and the post-synaptic conductance response. We next explore the implications of these effects on the transfer of rate-coded information.

Synaptic filtering of a rate-coded signal

Time-varying stimuli are often encoded in fluctuations of the firing rate of neuronal populations [35]. To address the question of how information about a rate-coded signal is filtered by vesicle dynamics, we use a model from [16] and [17] in which a time-varying signal is encoded in the firing rate of a presynaptic spike train to yield a doubly stochastic Poisson process, $I(t)$ (see Methods).

In this model, the instantaneous presynaptic rate conditioned on a signal, $s(t)$, is given by $\langle I(t)|s(t) \rangle = v + s(t)$ and, without conditioning on $s(t)$, is given by $\langle I(t) \rangle = v$. The power spectrum of the presynaptic spike train is given by

$$S_{II}(f) = v + S_{ss}(f), \quad (4)$$

where $S_{ss}(f)$ is the power spectrum of $s(t)$. Eq. (4) can be interpreted as follows: v represents the power of Poisson noise and $S_{ss}(f)$ represents the power of the signal. Unless $s(t)$ is identically zero, $I(t)$ inherits non-Poisson statistics from $s(t)$, which violates the Poisson assumptions used to derive the spectral properties given above. In the Methods, we derive a linear approximation (valid when

$S_{ss}(f) \ll v$) to the synaptic filter induced by the deterministic and stochastic models of vesicle dynamics and use it to obtain approximations to the cross-spectrum, $S_{sg}(f)$, between the signal and conductance as well as the power spectrum, $S_{gg}(f)$, of the conductance for this model (see Eqs. (27) and (28) in the Methods). These approximations allow an investigation of the information transfer of the signal across the synapse in various frequency bands.

We model $s(t)$ as a Gaussian process with Gaussian-shaped power spectrum (Fig. 6A,B),

$$S_{ss}(f) = D_s e^{-\frac{(f_s - f)^2}{2\sigma_s^2}}, \quad f \geq 0 \quad (5)$$

where σ_s is the bandwidth, f_s the central frequency, and D_s the peak power of the signal. We use a narrow-band signal (σ_s small) to more clearly illustrate the dependence of synaptic fidelity on signal frequency. Since $s(t)$ is Gaussian, there is a positive probability that $s(t) + v < 0$ so that the instantaneous firing rate of the presynaptic cells becomes negative. However, when $D_s \sigma_s \ll v^2$, this occurs rarely and can be disregarded by considering negative rates as zero [17]. The coherence, $C_{sg}(f) = |S_{sg}(f)|^2 / (S_{ss}(f) S_{gg}(f))$, between the signal and the conductance quantifies the fidelity with which the signal, $s(t)$, is represented in the postsynaptic conductance, $g(t)$. For the deterministic model of vesicle dynamics, the coherence is given by (from Eqs. (27))

$$C_{sg}^{\text{det}}(f) = \frac{S_{ss}(f)}{(1 + D_0)(v + S_{ss}(f))}$$

so that changing f_s merely shifts $C_{sg}^{\text{det}}(f)$, but does not change its amplitude (Fig. 6C,D dashed red line). Thus, a signal coded within any frequency band is transmitted with the same fidelity, consistent with the conclusions reached above using the Poisson model and also consistent with previous studies [16,17]. For the stochastic model, however,

$$C_{sg}^{\text{sto}}(f) = \frac{|\tilde{K}(f)|^2 S_{ss}(f)}{|\tilde{K}(f)|^2 (1 + D_0)(v + S_{ss}(f)) + S_{\eta_u \eta_u}(f) + S_{\eta_r \eta_r}(f)}$$

Since $\tilde{K}(f)$ is high pass (Fig. 3A) and $S_{\eta_u \eta_u}(f) + S_{\eta_r \eta_r}(f)$ is mostly flat (Fig. 3B), $C_{sg}^{\text{sto}}(f)$ is larger when $S_{ss}(f)$ concentrates its power in higher frequencies. For example, the amplitude of the coherence is larger when $f_s = 10\text{Hz}$ than when $f_s = 1\text{Hz}$ for the stochastic model, but independent of f_s for the deterministic model (Fig. 6C,D).

The rate of linear information transferred from the signal to the conductance is given by [36,37]

$$I_L(g; s) = - \int_0^\infty \log_2(1 - C_{sg}(f)) df.$$

In particular, $I_L(g; s)$ represents the total information per unit time that a linear decoder can obtain about the signal, $s(t)$, by observing the conductance, $g(t)$, and also represents a lower bound on the Shannon information [36,37]. The stochastic model predicts a dramatically lower linear information rate than the deterministic model (Fig. 7A). Since, for the deterministic model, the amplitude of $C_{sg}(f)$ is independent of the central signal frequency, f_s , the linear information rate is also independent of the central frequency (Fig. 7A). The stochastic model, however, transmits quickly varying signals with more fidelity than slowly varying signals (Fig. 7A). Hence, stochastic vesicle dynamics introduce frequency dependence into the transfer of linear information across a synapse.

In summary, our results show that the high pass nature of synaptic depression combined with low frequency synaptic noise limits the transfer of low frequency information through a synapse, while higher frequency information is transmitted more reliably. We next investigate these conclusions in a population setting.

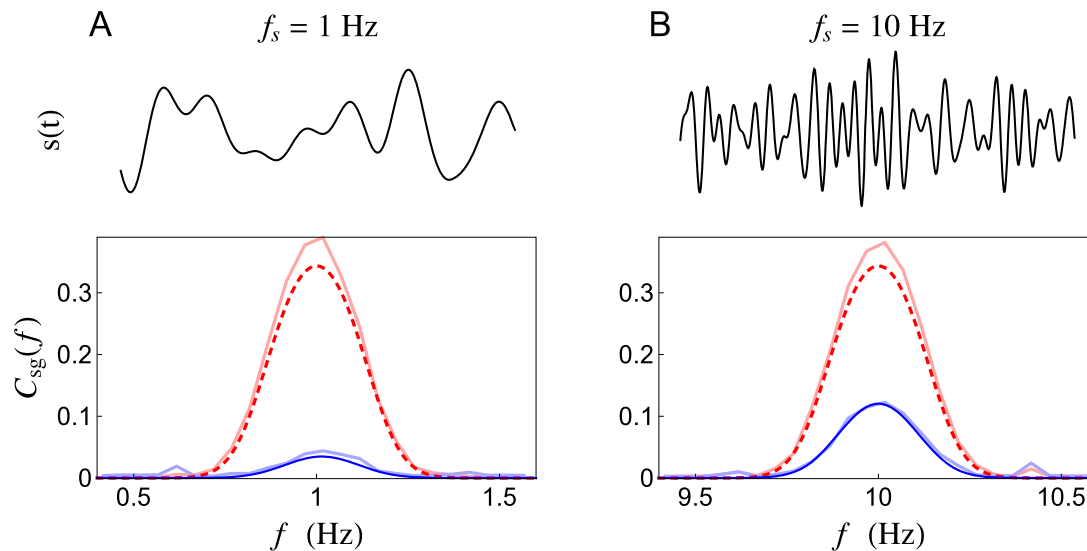


Figure 6. Signal transfer at high and low frequencies. The firing rate of a single presynaptic spike train ($n=1$) is modulated by the signal, $s(t)$, producing a postsynaptic conductance, $g(t)$. The coherence between the signal and conductance for (A) a slowly varying signal with peak frequency $f_s = 1\text{Hz}$ and (B) a quickly varying signal with $f_s = 10\text{Hz}$. The stochastic model (solid blue) transmits the higher frequency signal more reliably than the lower frequency signal. The deterministic model (dashed red) transmits the signal with equal fidelity in both cases. doi:10.1371/journal.pcbi.1002557.g006

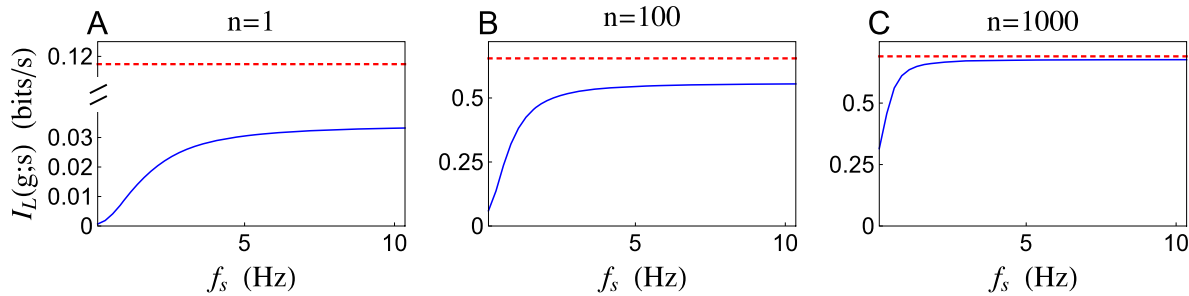


Figure 7. Linear information transfer rate as a function of signal frequency. The linear mutual information rate, $I_L(g; s)$, between a rate-coded signal, $s(t)$, and the total conductance, $g(t)$, produced by (A) $n = 1$, (B) $n = 100$, and (C) $n = 1000$ presynaptic spike trains, each encoding $s(t)$. The information rate is plotted as a function of the central frequency, f_s , at which $s(t)$ is encoded. The stochastic model (solid blue) transmits quickly varying signals more reliable than slowly varying signals. The deterministic model (dashed red) transmits information encoded at any frequency equally well.

doi:10.1371/journal.pcbi.1002557.g007

Synaptic filtering at the population level

So far, we have studied the conductance induced by a single presynaptic spike train that makes several contacts onto a postsynaptic cell. However, information about a stimulus is often encoded by populations of several presynaptic cells. We now consider a population model in which a collection, $\{I_k(t)\}_{k=1}^n$, of n presynaptic spike trains all encode the same signal, $s(t)$, as described for the single-cell model above. These inputs induce individual synaptic conductances, $\{g_k(t)\}_{k=1}^n$, in a single postsynaptic cell. Define the total presynaptic input, $I(t) = \sum_k I_k(t)$, and the conductance induced by this input, $g(t) = \sum_k g_k(t)$. For simplicity, we assume that all synapses have the same synaptic parameters p_r , τ_u , M , and $\alpha(t)$.

The signal, $s(t)$, introduces variability that is shared between the presynaptic spike trains. Such shared variability is commonly referred to as *signal correlation* since it is informative of the signal. Populations of presynaptic neurons that code for the same stimulus also share non-informative variability, known as *noise correlation* [38,39]. As a simple model of presynaptic noise correlation, we assume that each pair of spike trains, $I_j(t)$ and $I_k(t)$ with $j \neq k$, share a proportion c of their spike times. The pairwise cross-spectra are then given by

$$S_{I_j I_k}(f) = cv + S_{ss}(f), \quad j \neq k$$

where cv represents the contribution of noise correlations and $S_{ss}(f)$ represents the contribution of signal correlations.

As we have done for the single input model above, we gain an intuition for the population-level filter imposed by short term depression by first considering purely Poisson spike trains, which is achieved by setting $s(t) = 0$ so that $S_{ss}(f) = 0$. Even though the cross-spectrum, $S_{I_g}(f)$, is identical for the stochastic and deterministic models, the power spectrum, $S_{gg}(f)$, is larger for the stochastic model due to noise introduced by synaptic variability (see Fig. 8A,B and Eq. (29) in Methods). Therefore the coherence, $C_{I_g}(f)$, between the total presynaptic signal and the total conductance is smaller for the stochastic model. Moreover, the deterministic model predicts a flat coherence, while the stochastic model predicts a high-pass coherence (Fig. 8C). These conclusions are identical to those reached for a single input above, but the disparity between the two models is reduced at the population level (compare Figs. 3 and 5 with Fig. 8).

Notice also that the power spectrum, $S_{gg}(f)$, is peaked within the beta frequency band even though the inputs are Poisson and

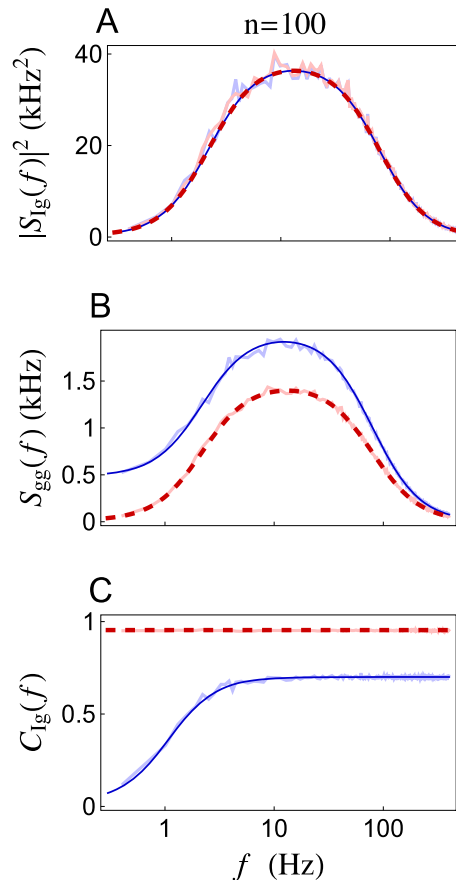


Figure 8. Synaptic filtering at the population level. A population, $\{I_k(t)\}$, of $n = 100$ Poisson presynaptic spike trains with pairwise correlation $c = 0.1$ drive a postsynaptic neuron to produce postsynaptic conductances, $\{g_k(t)\}$. (A) The cross-spectrum between the total presynaptic input and the total conductance. (B) The power spectrum of the total conductance has maximal power within the beta frequency band for both the deterministic (dashed red) and stochastic (solid blue) models. (C) The coherence between the total presynaptic input and the total conductance. Stochastic vesicle dynamics increase the power spectrum and therefore decrease the coherence, especially at low frequencies. All three plots are obtained in the absence of a rate-coded signal ($s(t) = 0$).

doi:10.1371/journal.pcbi.1002557.g008

therefore have a flat power spectrum. This effect could exaggerate beta frequencies in recorded data. We return to this topic in the Discussion.

A potential justification for using a deterministic model of vesicle dynamics is that, since stochastic release and recovery events are uncorrelated across all synapses, the extra variability introduced by synaptic noise averages out at the population level. So far, we have compared the two models for a population size of $n=100$. For the parameter values in Table 1, the low frequency cross-spectrum is identical for the two models, but the coherence and power spectrum disagree considerably until $n \approx 5000$ (Fig. 4Di–iii). The value of n at which the models begin to agree depends on the pairwise correlation, c , between the presynaptic inputs. Notably, in the absence of correlations ($c=0$ and $s(t)=0$), the population-level coherence is identical to the individual coherences, $C_{I_g}(f) = C_{I_{i,g_k}}(f)$, so that the coherence predicted by the stochastic and deterministic models disagree by the same amount for any value of n (Fig. 4Diii, lightest lines). As c increases, the two models agree at smaller population sizes (Fig. 4Diii, darker lines). Hence, presynaptic correlations must be present and n must be large if the deterministic model is to be used in place of the stochastic model for large populations.

We now study the transfer of rate coded information at the population level by allowing $s(t) \neq 0$. In particular, we are interested in how information about a rate-coded signal, $s(t)$, is transferred to the population conductance, $g(t)$. As above, we use a signal with Gaussian shaped power spectrum given by Eq. (5). A linear approximation to the cross-spectrum, $S_{sg}(f)$, for this model is calculated in the Methods (see Eq. (29)), which allows us to calculate the coherence, $C_{sg}(f)$, between the signal and the postsynaptic response and the linear information rate, $I_L(g; s)$, which depends on the central frequency at which the signal is coded in a qualitatively similar manner as for a single presynaptic spike train (compare Figs. 7A to 7B,C). In particular, low frequency information transfer is reduced for the stochastic model of synaptic depression. Moreover, the stochastic model transfers information in a frequency dependent manner and the deterministic model transfers information at all frequencies equally (Fig. 7). The disparity between the models is substantial when $n=100$, but reduced considerably when $n=1000$ (compare panels B and C in Fig. 7). We remind the reader that n represents the number of presynaptic neurons that encode the shared signal, $s(t)$, which could be much smaller than the total number of presynaptic inputs a cell receives. This suggests that, due to the stochastic nature of vesicle release and recovery, large presynaptic populations must be used to encode slowly varying signals.

Discussion

We derived a concise mathematical description of the synaptic filter induced by short term depression arising from neurotransmitter vesicle depletion. We found that stochasticity in vesicle release and recovery plays an important role in shaping this filter and determining the information processing capabilities of depressing synapses. For example, ignoring the stochasticity introduced by stochastic vesicle dynamics gives rise to a filter that transmits rate-coded signals encoded at all frequencies equally well [16,17], but taking this stochasticity into account reduces information transfer and causes slowly varying signals to be transferred with higher fidelity than slowly varying signals.

The deterministic model of short term depression provides a usable approximation to the stochastic model when considering large populations of correlated presynaptic spike trains (Figs. 4Di–iii and 7C). While a postsynaptic neuron typically receives

thousands of inputs, only a fraction of these inputs might be devoted to encoding a single stimulus. Our results show that a slowly varying stimulus must be encoded by large presynaptic populations, but quickly varying stimuli can be encoded by smaller populations. This conclusion is not true the deterministic model of synaptic depression, which ignores the inherent randomness of vesicle dynamics.

Since the two models predict the same mean conductance, the deterministic model is valid for studies that focus on mean postsynaptic activity and for which noise is not a concern. For example, the deterministic model has been used to describe the effects of depression on gain and temporal changes in postsynaptic firing rate [3,10,26,27]. Using the deterministic model in these cases is justified only if changes in postsynaptic firing rate result primarily from changes in the mean conductance and the variability of the conductance is inconsequential. When spiking is fluctuation driven, the postsynaptic firing rate is underestimated by the deterministic model [12].

A number of experimental studies have successfully fit parameters for the deterministic model to recorded neural data. This is achieved by first repeating the same presynaptic stimulus to a cell, then averaging the cell's response and fitting the averaged response to the response predicted by the deterministic model [2,5,7,8,18,32,33]. Since the stochastic model discussed here uses the same parameters as the deterministic model, the parameters obtained through this procedure can also be used to constrain the stochastic model.

Spectral analysis of synaptic depression

There is an extensive experimental and theoretical literature addressing how synapses that exhibit short term depression transmit different patterns of presynaptic spikes [3,26,27,40,41]. One recurring observation in these studies is that the steady state mean conductance (equivalently, the mean rate of vesicle release) saturates with the presynaptic firing rate, which causes the gain, $d\mu_g/dv$, to approach zero for large presynaptic rates (Fig. 2E). However, the gain only captures the sensitivity of the steady-state mean, μ_g , to static changes in v . Previous studies show that temporal changes in v are reflected more reliably in the transient mean of $g(t)$ than static changes of v are reflected in the steady-state mean of $g(t)$ [3,10,27,28]. This observation can be understood through our analysis by noting that higher frequency components of $S_{I_g}(f)$ are larger than the low-frequency components (Fig. 3B). Note that the decay of $S_{I_g}(f)$ at very high frequencies is due to the low-pass properties of the post-synaptic conductance kernel, $\tilde{\alpha}(f)$, (Fig. 3A and [29]) and not to synaptic depression. The filtering effects of depression are captured by the kernel $\tilde{K}(f)$, which is high-pass (Fig. 3A).

A second shortcoming of the gain as a descriptive quantity is that it does not capture the trial-to-trial variability in the conductance, which is a vital component of information transfer. We quantify this trial-to-trial variability as a function of frequency using the power spectrum, $S_{gg}(f)$. We show that the frequency-independence of information transfer through a deterministic synapse model depends on the precise shape of $S_{gg}(f)$ [16,17], and the high-pass frequency-dependence of information transfer through a stochastic synapse model likewise depends on the shape of $S_{gg}(f)$. Furthermore, we show that stochastic vesicle dynamics cause an overall decrease in information transfer by increasing $S_{gg}(f)$. Thus, trial-to-trial variability in $g(t)$ must be considered to obtain an accurate description of information transfer through a synapse.

While other studies of synaptic depression have investigated the transfer of rate-coded signals at various frequencies, we are not aware of a study that derives an explicit approximation to the filter induced by a depressing synapse. Such an approximation is derived in the Methods, giving

$$\tilde{g} = \left((1 + \sqrt{D_0}) \tilde{K} \tilde{I} + \tilde{\eta}_u + \tilde{\eta}_r \right) \tilde{a}$$

where $\tilde{I}(f)$ and $\tilde{g}(f)$ are the Fourier transforms of the presynaptic spike train and postsynaptic conductance respectively (see Methods for definitions of other terms). This expression can be used to predict the spectral properties of the postsynaptic response to a presynaptic input with a given power spectrum. A generalization of this expression that can be used in the case of a population of correlated presynaptic spike trains is given by Eq. (26).

Synaptic depression and neural rhythms

For the parameters in Table 1, the power spectrum is peaked within the beta frequency band (13–30 Hz) for both the stochastic and deterministic models (Fig. 8B). We emphasize that the presynaptic spike trains in this case are Poisson processes with flat power spectra and cross-spectra. Thus, the peaked power spectrum of the conductance is due completely to synaptic filtering: Frequencies below $1/(2\pi\tau_0) = (1 + p_r v \tau_u)/(2\pi\tau_u)$ Hz are suppressed by synaptic depression and frequencies above $1/(2\pi\tau_x)$ Hz are suppressed by post-synaptic channel dynamics. The conductance power spectrum is peaked between these two frequencies. This effect could potentially cause an exaggeration of beta or other frequencies in recordings such as local field potentials that reflect large pools of synaptic currents. Parameters can be chosen within a physiologically realistic range to produce a more exaggerated peak than that shown in Fig. 8B or to produce a peak within another frequency band (not shown). Further work is needed to determine the role that synaptic filtering plays in generating or exaggerating rhythms within beta or other frequency bands in functioning neural circuits.

Possible extensions

We used a simplified model of neurotransmitter release and recovery. In particular, we assumed that each contact contains only one release site. However, individual contacts can have multiple release sites and recent results show that multiple vesicles can be released by a single contact in response to a single presynaptic action potential [22,23]. Such situations can be modeled in our framework by interpreting M as the total number of release sites at all contacts. However, this interpretation is only valid if the release of vesicles is statistically independent between release sites that share a contact. If the probability of release at one site depends on release at another site – for instance if a contact has several release sites but can only release one vesicle per presynaptic spike [12,42] – then our model would need to be adjusted to account for this dependency. To the authors' knowledge, the precise structure of such dependencies are a subject of current research and not presently understood. In the depleted state ($U v \tau_u \gg 1$), a contact with several release sites will rarely have more than one vesicle available for release at any point in time and our single-vesicle model should provide an accurate approximation regardless of dependencies between release sites, as long as the recovery time constant is properly adjusted [12].

We modeled stochasticity introduced by probabilistic vesicle release and random recovery times, but did not model stochasticity introduced by randomness in the amount of neurotransmitter

contained in each vesicle [43,44]. In addition we did not model variability at the postsynaptic site (e.g., randomness in the number of bound receptors, the number of open channels, or the availability of messenger molecules), which could introduce variability in the amplitude of the postsynaptic conductance elicited by each vesicle released. Assuming statistical independence of these sources of variability between release events, they can be captured by multiplying each response amplitude, w_k , by a random number. This would simply scale the power spectrum of the conductance linearly and would not alter our central conclusions.

The cross-spectrum between presynaptic input and postsynaptic conductance decays to zero at high frequencies, but the coherence between the two does not (Figs. 3A and 5). This is due to the fact that the power spectrum also decays at high frequencies and cancels perfectly with the cross-spectrum. However, any additional high frequency noise would destroy this balance. For example, if one were to instead compute the coherence between the presynaptic input and the *current* across the postsynaptic membrane, high frequency channel noise [45] could increase the power spectrum without increasing the cross-spectrum and therefore cause the coherence to decay at high frequencies. Thus, information transfer from presynaptic input to postsynaptic current is effectively bandpass. Similar observations were discussed in [17] for the deterministic model of vesicle dynamics with additive noise.

We used a linear approximation to predict the spectral properties of the postsynaptic conductance induced by non-Poisson presynaptic spike trains. However, the approximation is only assured to be accurate when inputs are approximately Poisson, i.e., have a nearly flat power spectrum. This restriction is implicit in our assumption that $S_{ss}(f) \ll v$ (see Eq. (4) and the surrounding discussion). Presynaptic spike trains that exhibit highly non-Poisson properties, such as bursts or a high degree of regularity, can interact with synaptic depression in a fundamentally different manner than Poisson spike trains [12,46]. Further work is needed to extend our results to highly non-Poisson presynaptic spiking statistics.

We focused on short term depression caused by the depletion of synaptic neurotransmitter vesicles. However, other sources of short term depression as well as several forms of short term facilitation affect the filtering properties of synapses [1,40]. Our mathematical methods could be extended to take these additional forms of plasticity into account.

Synaptic transmission of Shannon information

To quantify information transfer through a synapse, we used an information metric that only captures the amount of information available to a linear decoder observing the conductance. The Shannon information measures the maximum amount of information available to any decoder [47]. Interestingly, for our choice of $\alpha(t)$, the deterministic model of vesicle dynamics transmits Shannon information perfectly because every presynaptic spike elicits a postsynaptic response (Fig. 2D) and hence each spike time can be resolved by detecting jumps in $g(t)$ [17,19]. In contrast, the stochastic model of vesicle dynamics exhibits failures due both to probabilistic release and to vesicle depletion (Fig. 2C,E). Due to the presence of synaptic failure, the stochastic model reduces Shannon information since some presynaptic spikes have no effect on the postsynaptic conductance.

A few studies have investigated the reduction of Shannon information through synapses with synaptic failure [20,46,48] but focus on the impact of probabilistic release and ignore stochasticity in vesicle recovery dynamics. In contrast, we studied the reduction

of linear information induced by both probabilistic release and stochastic recovery. The qualitative differences we observed between stochastic and deterministic models depend on the stochasticity of vesicle recovery since it introduces low frequency variability into the conductance (Fig. 3C,D). To our knowledge, only one study [19] has investigated information transmission in a model with both probabilistic release and stochastic recovery. Using simulations, they found that stochastic vesicle dynamics reduce Shannon information by orders of magnitude, consistent with our results for linear information. These previous studies of information transmission do not quantify the dependence of information transfer on the frequency band in which presynaptic information is encoded. Furthermore, care must be taken when drawing conclusions about neural coding from studies of Shannon information. Shannon information quantifies the maximal information that can be extracted by a decoder, but it is not always clear whether a neural decoder can achieve optimal or even near-optimal decoding.

Methods

Definition of the models and derivation of first moments

Consider a single presynaptic neuron that fires action potentials at times $\{t_j\}$ and define the presynaptic spike train as a point process,

$$I(t) = \sum_j \delta(t - t_j),$$

where $\delta(t)$ is the Dirac delta function. The number of presynaptic spikes in $[0, t]$ is then given by $N_I(t) = \int_0^t I(s) ds$. Define M to be the number of functional contacts that the presynaptic neuron makes onto a postsynaptic cell [48] and, for simplicity, assume that each contact can have at most one vesicle available for release at any point in time. Let $0 \leq m(t) \leq M$ be the total number of vesicles available for release at time t . Let w_j be the number of vesicles released by the j th presynaptic spike, with $0 \leq w_j \leq m(t_j)$. The total number of vesicles released up to time t is given by $N_x(t) = \sum_{t_j < t} w_j$ and the effective synaptic input is a marked point process defined by

$$x(t) = \frac{dN_x(t)}{dt} = \sum_j w_j \delta(t - t_j). \quad (6)$$

We first consider a model of synaptic vesicle dynamics that treats vesicle release and recovery stochastically [12,19,24,25]. At each presynaptic spike time, t_j , each contact at which a vesicle is available releases this vesicle independently with probability p_r . After a synaptic contact releases its vesicle, vesicle recovery occurs as a Poisson process with rate $1/\tau_u$. That is, the waiting time from vesicle release until recovery at a single contact is exponentially distributed with mean τ_u and independent from the state of other contacts, so that the probability of a recovery event during the interval $[t, t + dt]$ is $dt(M - m(t))/\tau_u + \mathcal{O}(dt^2)$. This model can be described by the equation

$$dm(t) = -dN_x(t) + dN_u(t) \quad (7)$$

where $dN_u(t) = u(t)dt$ is the increment of an inhomogeneous Poisson process with instantaneous rate that depends on $m(t)$ through $\langle dN_u(t) | m(t) \rangle = dt(M - m(t))/\tau_u$ (here, $\langle \cdot | \cdot \rangle$ denotes

conditional expectation) and $dN_x(t)$ is given by Eq. (6) where each w_j is a binomial random variable with mean $p_r m(t_j)$. Since each trial with a fixed input, $I(t)$, yields a different, random realization of the response, $x(t)$, we hereafter refer to this model as the “stochastic model” of vesicle dynamics.

A popular simplification of the stochastic model replaces the random increments, $dN_x(t)$ and $dN_u(t)$, in Eq. (7) with their expected values conditioned on $m(t)$ and $dN_I(t)$ [2,3,5,6]. Since $\langle dN_x(t) | m(t_j), dN_I(t) \rangle = p_r m(t) dN_I(t)$ and $\langle dN_u(t) | m(t), dN_I(t) \rangle = dt(M - m(t))/\tau_u$, this gives

$$\begin{aligned} dm(t) &= -dN_x(t) + \frac{M - m(t)}{\tau_u} dt \\ dN_x(t) &= p_r m(t) dN_I(t). \end{aligned} \quad (8)$$

This model treats $m(t)$ as a continuous variable where a proportion p_r of the available vesicles are released at each input and recovery occurs exponentially with time constant τ_u . We hereafter refer to the model described by Eq. (8) as the “deterministic model” of vesicle dynamics since the response, $x(t)$, is determined completely by the presynaptic input, $I(t)$. Stochasticity in this model is only introduced by randomness in $I(t)$.

When $I(t)$ is a homogeneous Poisson process, the deterministic model is analytically tractable: the first two moments of $x(t)$ and $m(t)$ can be derived exactly, as we show below. We also show that the first moments agree for two models. The second moments for the stochastic model are difficult to derive analytically, but we derive a more tractable diffusion approximation below. Furthermore, when $I(t)$ is not a homogeneous Poisson processes, closed form approximations can be obtained for both the deterministic and stochastic models.

Assume that $I(t)$ is a homogeneous Poisson process with rate v . Then the increment, $dN_I(t)$, is independent from the current value of $m(t)$ so that, by taking expectations in Eq. (8), $\langle dN_x(t) \rangle = \langle p_r m(t) dN_I(t) \rangle = p_r v \langle m(t) \rangle$ for the deterministic model. Similarly, $\langle dm(t) \rangle = -\langle dN_x(t) \rangle + dt(M - \langle m(t) \rangle)/\tau_u$. Combining these gives

$$\begin{aligned} \frac{d\langle m(t) \rangle}{dt} &= \frac{M}{\tau_u} - \left(\frac{1 + p_r v \tau_u}{\tau_u} \right) \langle m(t) \rangle \\ \frac{d\langle N_x(t) \rangle}{dt} &= p_r v \langle m(t) \rangle. \end{aligned} \quad (9)$$

Eq. (9) is also obtained by taking expectations in Eq. (7), which implies that the deterministic model and the stochastic model yield the same means when $I(t)$ is a homogeneous Poisson process. The following equation for $\langle x(t) \rangle$ can be obtained using Eq. (9) and the fact that $x(t) = dN_x(t)/dt$,

$$\frac{d\langle x(t) \rangle}{dt} = \frac{p_r v M}{\tau_u} - p_r v \left(\frac{1 + p_r v \tau_u}{\tau_u} \right) \langle m(t) \rangle. \quad (10)$$

The stationary mean of $m(t)$ is given by the unique steady state solution to Eq. (9) [4],

$$\mu_m := \lim_{t \rightarrow \infty} \langle m(t) \rangle = \frac{M}{1 + p_r v \tau_u}. \quad (11)$$

Furthermore, after a perturbation of $m(t)$ or starting from an initial condition $\langle m(0) \rangle \neq \mu_m$, $\langle m(t) \rangle$ decays exponentially back to μ_m with time constant

$$\tau_0 = \frac{\tau_u}{1 + p_r v \tau_u}.$$

The stationary mean number of vesicles released by each presynaptic spike is given by $\lim_{j \rightarrow \infty} \langle w_j \rangle = p_r \mu_m$ and the stationary mean of the postsynaptic signal is $\lim_{t \rightarrow \infty} \langle x(t) \rangle = p_r v \mu_m$, which represents the steady state rate of vesicle release. Furthermore, $\langle x(t) \rangle$ approaches its steady state exponentially with the same time constant, τ_0 , as $\langle m(t) \rangle$.

The calculations of first moments above depend on the fact that $m(t)$ and $dN_I(t)$ are independent for any t . This can only be assumed to hold when Eq. (8) is interpreted in the Itô sense (so that $m(t)$ is updated directly after a spike) and $I(t)$ is a homogeneous Poisson process. If $I(t)$ is not a homogeneous Poisson process, then the equations for the first moments are not valid and the first moments may not agree for the two models.

A diffusion approximation of the stochastic model

Second moments for the stochastic model are difficult to derive analytically, so we obtain approximations by considering a diffusion approximation

$$\begin{aligned} dm(t) &= \frac{M - m(t)}{\tau_u} dt - dN_x(t) + \sqrt{D_u} dW_u(t) \\ dN_x(t) &= p_r m(t) dN_I(t) + \sqrt{D_r} Z_r(t) dN_I(t) \end{aligned} \quad (12)$$

where $W_u(t)$ is a standard Wiener process that models stochasticity in vesicle recovery. Stochasticity in vesicle release is captured by the stationary process, $Z_r(t)$, with moments given by $\langle Z_r(t) \rangle = 0$, $\langle Z_r(t)^2 \rangle = 1$, and $\langle Z_r(t) Z_r(s) \rangle = 0$ for $s \neq t$. We assume that $Z_r(t)$, $W_u(t)$, and $I(t)$ are mutually independent. These equations should be interpreted in the Itô sense, so that the increments $dm(t) = m(t + dt) - m(t)$ and $dN_x(t) = N_x(t + dt) - N_x(t)$ are independent from the history of the noise terms, $\{W_u(s), Z_r(s), N_I(s)\}_{s \leq t}$, for any time t [49]. Since $\langle Z_r(t) \rangle = \langle dW_u(t) \rangle = 0$, it is clear that the diffusion approximation defined by Eq. (12) has first moments that satisfy Eq. (9).

The noise coefficients, D_r and D_u , quantify the degree of randomness introduced by stochastic release and recovery respectively. To find appropriate values for these coefficients, we compute the infinitesimal variance of $dm(t)$ and $dN_x(t)$ conditioned on the drift terms that appear in their respective equations in Eq. (12) [50]. Since vesicle recovery events are Poissonian, the variance of its increment is equal to its rate, giving the conditional variance

$$\text{var}(dm|m, dN_x) = \frac{M - m}{\tau_u} dt.$$

Note that the $dN_x(t)$ term that appears on the right hand side of Eq. (12) does not contribute to this conditional variance since $\text{var}(dN_x|m, dN_x) = 0$. Conditioned on $m(t)$ and the occurrence of a presynaptic spike, the number of vesicles released has a binomial distribution with mean $\langle dN_x|m, dN_I = 1 \rangle = p_r m$ and therefore has conditional variance given by

$$\text{var}(dN_x|m, dN_I = 1) = p_r(1 - p_r)m.$$

Optimally, we would set $D_u = \text{var}(dm|m, dN_x)/dt$ and $D_r = \text{var}(dN_x|m, dN_I > 0)/dt$, but doing so would give rise to nonlinear multiplicative noise in Eq. (12), which is difficult to treat

mathematically. Instead, we obtain an approximation by replacing $m(t)$ with its stationary mean, μ_m , to obtain

$$D_u = \frac{M - \mu_m}{\tau_u} \text{ and } D_r = p_r(1 - p_r)\mu_m. \quad (13)$$

All calculations for the stochastic model are carried out using the diffusion approximation from Eq. (12) with the noise coefficients from Eq. (13), and therefore expressions obtained are approximations to the full stochastic model described above. However, in all figures, simulations are performed using the full stochastic model from Eq. (7) (light blue lines) and show excellent agreement with the closed form approximations (dark blue lines).

Note that the deterministic model can be recovered by taking $D_u = D_r = 0$ in Eq. (12). Thus, we can proceed in our analysis by considering Eq. (12) without instantiating D_u or D_r to obtain results that apply to both the deterministic and stochastic models.

Derivation of the auto-covariance and power spectrum of $x(t)$

We quantify temporal and trial-to-trial variability between two stationary processes, $v(t)$ and $y(t)$, using the cross-covariance function,

$$R_{vy}(\tau) = \text{cov}(v(t), y(t + \tau)),$$

and its Fourier transform, the cross-spectrum,

$$S_{vy}(f) = \int_{-\infty}^{\infty} R_{vy}(\tau) e^{-2\pi i f \tau} d\tau.$$

The cross-covariance (cross-spectrum) between a process and itself is called an auto-covariance (power spectrum). To quantify the variability of the postsynaptic response, we now derive the auto-covariance, $R_{xx}(\tau)$, and the power spectrum, $S_{xx}(f)$, for the synapse model in Eq. (12).

From Eqs. (9) and (10) it is apparent that, for $\tau > 0$, the expectations $\langle m(t + \tau) \rangle$ and $\langle x(t + \tau) \rangle$ decay exponentially to their steady state, given any initial distribution, P_0 , imposed on $m(t)$ and $x(t)$. From this fact, it is apparent that $\langle x(t + \tau)x(t) \rangle = \int x \langle x(t + \tau) | m(t) = m, x(t) = x \rangle dP_0(m, x)$ should inherit this exponential shape and therefore that $R_{xx}(\tau)$ should have an exponential shape with time constant τ_0 .

We now make this argument more precise using a regression theorem from [49]. Define the bivariate Markov process,

$$Y(t) = \begin{pmatrix} m(t) - \mu_m \\ x(t) - p_r v \mu_m \end{pmatrix}.$$

Then Eqs. (9) and (10) show that

$$\frac{d \langle Y(t + \tau) | Y(t) \rangle}{dt} = -A \langle Y(t + \tau) | Y(t) \rangle$$

for $\tau > 0$ where

$$A = \begin{pmatrix} 1/\tau_0 & 0 \\ p_r v/\tau_0 & 0 \end{pmatrix}.$$

In Sec. 3.7.4 of [49], it is shown that this implies

$$\frac{d\langle Y(t+\tau)Y(t)^T \rangle}{dt} = -A\langle Y(t+\tau)Y(t)^T \rangle$$

for $\tau > 0$. Solving this linear differential equation gives

$$\langle (x(t+\tau) - p_r v \mu_m)(x(t) - p_r v \mu_m) \rangle = v p_r \langle x(t) m(t) \rangle e^{-\tau/\tau_0}$$

for $\tau > 0$. Thus, due to stationarity,

$$\begin{aligned} R_{xx}(\tau) &= \lim_{t \rightarrow \infty} \langle (x(t+\tau) - p_r v \mu_m)(x(t) - p_r v \mu_m) \rangle \\ &= B e^{-\tau/\tau_0} \end{aligned}$$

for $\tau > 0$ and where B is a constant. By symmetry, we have $R_{xx}(-\tau) = R_{xx}(\tau)$. Note also that, since $x(t)$ is a marked point process, there is a Dirac delta function that contributes to $R_{xx}(\tau)$ at $\tau = 0$ [51]. Finally, we may conclude that the auto-covariance of $x(t)$ has the form

$$R_{xx}(\tau) = A \delta(\tau) + B e^{-|\tau|/\tau_0} \tag{14}$$

for some constants A and B .

To calculate the coefficients A and B in Eq. (14), we must first calculate a few infinitesimal moments using stochastic calculus techniques [52]. In our calculations, we ignore terms of order $dt^2 := (dt)^2$ and higher, but must include terms of the form order dm^2 and dN_x^2 because their expectation is of the order dt [50].

The second moment of dN_x conditioned on m is given by

$$\begin{aligned} \langle dN_x^2 | m \rangle &= \langle (p_r m dN_I + \sqrt{D_r} dN_I Z_r)^2 | m \rangle \\ &= p_r^2 m^2 \langle dN_I^2 \rangle + D_r \langle dN_I^2 \rangle \end{aligned} \tag{15}$$

$$= (p_r^2 m^2 + D_r) v dt \tag{16}$$

where (15) follows from the fact that $Z_r(t)$ and $dN_I(t)$ are independent from each other and from $m(t)$, that $\langle Z_r(t) \rangle = 0$, and that $\langle Z_r(t)^2 \rangle = 1$; and (16) follows from the fact that $\langle dN_I^2 \rangle = v dt$. The calculation of the conditional mixed moment, $\langle m dN_x | m \rangle$, is similar and gives

$$\begin{aligned} \langle m dN_x | m \rangle &= \langle m (p_r m dN_I + \sqrt{D_r} dN_I Z_r) | m \rangle \\ &= p_r m^2 v dt. \end{aligned}$$

To calculate the stationary second moment, $\lim_{t \rightarrow \infty} \langle m(t)^2 \rangle$, we modify a strategy from Sec. 4.4.7c of [49] to derive a linear differential equation for the time dependent second moment and find its steady state. First note that

$$d\langle m^2 \rangle = \langle d(m^2) \rangle = 2\langle m dm \rangle + \langle dm^2 \rangle.$$

The first term in this sum is given by

$$\begin{aligned} \langle m dm \rangle &= \langle m (dt(M-m)/\tau_u - dN_x + \sqrt{D_u} dW_u) \rangle \\ &= \frac{\langle m \rangle M}{\tau_u} dt - \frac{\langle m^2 \rangle}{\tau_u} dt - \langle m dN_x \rangle \\ &= \frac{\langle m \rangle M}{\tau_u} dt - \langle m^2 \rangle \left(\frac{1}{\tau_u} + p_r v \right) dt \end{aligned}$$

where we used the fact that $m(t)$ and $dW_u(t)$ are independent (see above) and the last line follows from the equation for $\langle m dN_x | m \rangle$ derived above. Now calculate

$$\begin{aligned} \langle dm^2 \rangle &= \left\langle \left(dt(M-m)/\tau_u - dN_x + \sqrt{D_u} dW_u \right)^2 \right\rangle \\ &= \langle dN_x^2 \rangle + D_u dt \\ &= (p_r^2 \langle m^2 \rangle + D_r) v dt + D_u dt \end{aligned}$$

where we have eliminated terms of order dt^2 and used the fact that dW_u is independent from all other terms; and the last line follows from the equation for $\langle dN_x^2 | m \rangle$ above. Combining these expressions gives a differential equation for the time course of the second moment of m ,

$$\begin{aligned} \frac{d\langle m(t)^2 \rangle}{dt} &= -\langle m(t)^2 \rangle \left(\frac{2}{\tau_u} + (2p_r - p_r^2) v \right) \\ &+ \frac{2\langle m(t) \rangle M}{\tau_u} + D_r v + D_u \end{aligned}$$

where $\langle m(t) \rangle$ is given by the solution of Eq. (9) above. The stable fixed point of this linear differential equation is the stationary second moment of $m(t)$,

$$\begin{aligned} \langle m^2 \rangle &:= \lim_{t \rightarrow \infty} \langle m(t)^2 \rangle \\ &= \frac{2\mu_m M + D_r v \tau_u + D_u \tau_u}{2 + (2p_r - p_r^2) v \tau_u} \end{aligned} \tag{17}$$

where μ_m is the stationary mean of $m(t)$, given in Eq. (11). The delta function in $R_{xx}(\tau)$ has area given by

$$A = \frac{\langle dN_x^2 \rangle}{dt} = (p_r^2 \langle m^2 \rangle + D_r) v \tag{18}$$

where we used Eq. (16) above and where $\langle m^2 \rangle$ is given by Eq. (17).

To calculate the one-sided limit, $B = \lim_{\tau \rightarrow 0^+} R_{xx}(\tau)$, first calculate

$$\begin{aligned} &\lim_{\tau \rightarrow 0^+} \langle x(t)x(t+\tau) \rangle \\ &= \lim_{\tau \rightarrow 0^+} \left\langle \left(p_r m(t) I(t) + \sqrt{D_r} I(t) Z_r(t) \right) \times \right. \\ &\quad \left. \left(p_r m(t+\tau) I(t+\tau) + \sqrt{D_r} I(t+\tau) Z_r(t+\tau) \right) \right\rangle \\ &= v p_r \lim_{\tau \rightarrow 0^+} \left\langle \left(p_r m(t) I(t) + \sqrt{D_r} I(t) Z_r(t) \right) m(t+\tau) \right\rangle \\ &= v p_r^2 \lim_{\tau \rightarrow 0^+} \left\langle I(t) m(t) m(t+\tau) \right\rangle + \sqrt{D_r} \langle I(t) Z_r(t) m(t+\tau) \rangle \end{aligned}$$

where we have used the fact that $Z_r(t+\tau)$ and $I(t+\tau)$ are independent of all of the other terms when $\tau > 0$. Each of the terms in the sum above can be calculated by conditioning on a spike at time t and on the value of $m(t)$,

$$\begin{aligned} & \lim_{\tau \rightarrow 0} \langle I(t)m(t)m(t+\tau) \rangle \\ &= \lim_{\tau \rightarrow 0} v \langle m(t) \langle m(t+\tau) | m(t), dN_I(t) > 0 \rangle \rangle_m \\ &= v \langle m(t) (m(t) - p_r m(t) - \sqrt{D_r} Z_r(t)) \rangle_m \\ &= v(1 - p_r) \langle m^2 \rangle \end{aligned}$$

where $\langle \cdot \rangle_m$ is expectation over the variable $m(t)$. Similarly,

$$\begin{aligned} & \lim_{\tau \rightarrow 0} \langle I(t)Z_r(t)m(t+\tau) \rangle \\ &= \lim_{\tau \rightarrow 0} v \langle Z_r(t) \langle m(t+\tau) | Z_r(t), dN_I(t) > 0 \rangle \rangle_z \\ &= v \langle Z_r(t) (m(t) - p_r m(t) - \sqrt{D_r} Z_r(t)) \rangle_z \\ &= -v \sqrt{D_r} \end{aligned}$$

where $\langle \cdot \rangle_z$ is expectation over $Z_r(t)$. Combining the expressions above gives

$$\lim_{\tau \rightarrow 0} \langle x(t)x(t+\tau) \rangle = v^2 p_r (p_r (1 - p_r) \langle m^2 \rangle - D_r)$$

Finally, since $\langle x(t) \rangle = \langle dN_x \rangle / dt = p_r v \mu_m$ from above, we have

$$\begin{aligned} B &= \lim_{\tau \rightarrow 0} R_{xx}(\tau) = \lim_{\tau \rightarrow 0} \langle x(t)x(t+\tau) \rangle - \langle x(t) \rangle^2 \\ &= v^2 p_r (p_r (1 - p_r) \langle m^2 \rangle - D_r - p_r \mu_m^2) \end{aligned} \tag{19}$$

where μ_m and $\langle m^2 \rangle$ are the stationary first and second moments of $m(t)$, given in Eqs. (11) and (17). The auto-covariance of $x(t)$ is then given by Eq. (14) with A and B given by Eqs. (18) and (19).

The power spectrum is obtained from the auto-covariance through a Fourier transform,

$$S_{xx}(f) = (1 + D_0) |\tilde{K}(f)|^2 v + S_{\eta_u \eta_u}(f) + S_{\eta_r \eta_r}(f)$$

where

$$\tilde{K}(f) = p_r \mu_m - v p_r^2 \mu_m \frac{i \tau_0}{i - 2\pi \tau_0 f} \tag{20}$$

is a deterministic linear kernel,

$$D_0 = \frac{v \tau_u p_r^2}{v \tau_u (2 - p_r) p_r + 2}$$

is the noise intensity introduced by the interaction between the stochastic input and deterministic vesicle dynamics,

$$S_{\eta_u \eta_u}(f) = D_u D_0 \left(1 + v(1 - p_r) \frac{2\tau_0}{4\pi^2 \tau_0^2 f^2 + 1} \right)$$

is the noise introduced by stochasticity in vesicle recovery, and

$$S_{\eta_r \eta_r}(f) = D_r D_0 \left(\frac{2}{p_r^2 \tau_0} - v \left(\frac{\tau_0 + \tau_u}{p_r \tau_0 \tau_u} \right) \frac{2\tau_0}{4\pi^2 \tau_0^2 f^2 + 1} \right)$$

is the noise introduced by stochasticity in vesicle release. Note that

$S_{\eta_u \eta_u}(f) = S_{\eta_r \eta_r}(f) = 0$ for the deterministic model since $D_u = D_r = 0$.

Derivation of the cross-covariance and cross-spectrum between $I(t)$ and $x(t)$

To measure the covariability between the presynaptic spike trains and the postsynaptic response, we now derive the cross-covariance between the input, $I(t)$, and the response $x(t)$. By a similar argument to the one made above for $R_{xx}(\tau)$, we may conclude that $R_{Ix}(\tau)$ is the sum of a delta function and an exponential, except that the exponential is one-sided since $R_{Ix}(\tau) = \text{cov}(I(t), x(t+\tau)) = 0$ for $\tau < 0$. For $\tau > 0$, we can find the peak of the exponential by first conditioning on a spike at time t , then conditioning on a spike at time $t + \tau$,

$$\begin{aligned} \lim_{\tau \rightarrow 0^+} \langle I(t)x(t+\tau) \rangle &= v \lim_{\tau \rightarrow 0^+} \langle x(t+\tau) | dN_I(t) > 0 \rangle \\ &= v^2 \lim_{\tau \rightarrow 0^+} \langle p_r m(t+\tau) + \sqrt{D_r} Z_r(t+\tau) | dN_I(t) > 0, dN_I(t+\tau) > 0 \rangle \\ &= v^2 \langle p_r (m(t+\tau) - p_r m(t+\tau) - \sqrt{D_r} Z_r(t+\tau)) + \sqrt{D_r} Z_r(t+\tau) \rangle \\ &= v^2 p_r (1 - p_r) \mu_m \end{aligned}$$

since $\langle Z_r(t+\tau) \rangle = 0$. Thus,

$$\begin{aligned} \lim_{\tau \rightarrow 0^+} R_{Ix}(\tau) &= \lim_{\tau \rightarrow 0^+} \langle x(t+\tau)I(t) \rangle - \langle x(t) \rangle \langle I(t) \rangle \\ &= v^2 p_r (1 - p_r) \mu_m - v^2 p_r \mu_m \\ &= -v^2 p_r^2 \mu_m. \end{aligned}$$

The area of the delta function in R_{Ix} is given by

$$\begin{aligned} \frac{\langle dN_x(t)dN_I(t) \rangle}{dt} &= v \langle dN_x | dN_I > 0 \rangle \\ &= v p_r \mu_m \end{aligned}$$

since $\langle dN_x | dN_I > 0 \rangle = \mu_w = p_r \mu_m$. Thus, we have

$$R_{Ix}(\tau) = v p_r \mu_m \delta(\tau) - \Theta(\tau) v^2 p_r^2 \mu_m e^{-\tau/\tau_0}$$

where Θ is the Heaviside step function. Taking the Fourier transform gives the cross-spectrum

$$S_{Ix}(f) = v \tilde{K}(f)$$

where $\tilde{K}(f)$ is defined in Eq. (20) above.

Postsynaptic response to several correlated presynaptic spike trains

The statistics of the postsynaptic response to a population, $\{I_k(t)\}_{k=1}^n$, of *uncorrelated* presynaptic spike trains can be easily calculated from the statistics of individual responses, which are calculated above. However, neurons that contact a shared postsynaptic cell often exhibit correlations between their spiking activity [39,53]. To determine the postsynaptic response to a population of correlated presynaptic spike trains, we must first calculate the pairwise cross-spectra of the conductances induced by these inputs. Assume that each spike train, $I_k(t)$, in the presynaptic population is a Poisson process with rate v . Introduce correlations by assuming that each pair, $I_j(t)$ and $I_k(t)$, of spike

trains share a proportion c of their spike times so that $S_{I_j I_k}(f) = cv$ [54]. We use subscripts to denote quantities associated with each spike train and double subscripts as necessary. For simplicity, assume that the synaptic parameters M , p_r , and τ_u are identical for all synapses. The asymmetric case can be treated identically, but the expressions obtained are more cumbersome. The power spectrum, $S_{x_k x_k}(f)$, and the cross-spectrum, $S_{I_k x_k}(f)$, are given above (where they are written as $S_{xx}(f)$ and $S_{Ix}(f)$). Below, we derive expressions for $S_{x_j x_k}(f)$ and $S_{I_j x_k}(f)$ for $j \neq k$.

First, following the same arguments used above to derive the moments of $N_x(t)$ and $m(t)$ in the case of a single presynaptic spike train, we obtain the bivariate moments

$$\begin{aligned} \langle dN_{x_j} dN_{x_k} | m_j, m_k \rangle &= \left\langle \left(p_r m_j dN_{I_j} + \sqrt{D_r} dN_{I_j} Z_{r,j} \right) \times \right. \\ &\quad \left. \left(p_r m_k dN_{I_k} + \sqrt{D_r} dN_{I_k} Z_{r,k} \right) | m_j, m_k \right\rangle \\ &= p_r^2 m_j m_k \langle dN_{I_j} dN_{I_k} \rangle \\ &= p_r^2 m_j m_k cv dt. \end{aligned}$$

Similarly,

$$\begin{aligned} \langle dN_{x_j} m_k | m_j, m_k \rangle &= \langle (p_r m_j dN_{I_j} + \sqrt{D_r} dN_{I_j} Z_{r,j}) m_k \rangle \\ &= p_r v m_j m_k dt \end{aligned}$$

and, equivalently,

$$\langle dN_{x_k} m_j | m_j, m_k \rangle = p_r v m_j m_k dt.$$

We now derive a differential equation for $\langle m_j(t) m_k(t) \rangle$ to get the stationary second moment. First note that $d(m_j m_k) = m_j dm_k + m_k dm_j + dm_j dm_k$ so that

$$d \langle m_j m_k \rangle = \langle m_j dm_k \rangle + \langle m_k dm_j \rangle + \langle dm_j dm_k \rangle. \quad (21)$$

By symmetry, the first and second terms in Eq. (21) are the same and they can be derived from Eq. (12) as

$$\begin{aligned} \frac{\langle m_j dm_k \rangle}{dt} &= \frac{\langle m_k dm_j \rangle}{dt} \\ &= \langle m_j ((M - m_k) / \tau_u - dN_{x_k} / dt + \sqrt{D_u} dW_{u,k} / dt) \rangle \\ &= \frac{M \mu_m}{\tau_u} - \frac{\langle m_j m_k \rangle}{\tau_u} - \langle m_j dN_{x_k} \rangle \\ &= \frac{M \mu_m}{\tau_u} - \frac{\langle m_j m_k \rangle}{\tau_u} - p_r v \langle m_j m_k \rangle. \end{aligned}$$

The last term in Eq. (21) is given by

$$\begin{aligned} \langle dm_j dm_k \rangle &= \left\langle \left(dt(M - m_j) / \tau_u - dN_{x_j} + \sqrt{D_u} dW_{u,j} \right) \times \right. \\ &\quad \left. \left(dt(M - m_k) / \tau_u - dN_{x_k} + \sqrt{D_u} dW_{u,k} \right) \right\rangle \\ &= \langle dN_{x_j} dN_{x_k} \rangle \\ &= p_r^2 \langle m_j m_k \rangle cv dt. \end{aligned}$$

Combining these gives

$$\begin{aligned} \frac{d \langle m_j m_k \rangle}{dt} &= 2 \frac{M \mu_m}{\tau_u} - 2 \frac{\langle m_j m_k \rangle}{\tau_u} \\ &\quad - 2 p_r v \langle m_j m_k \rangle + p_r^2 \langle m_j m_k \rangle cv \end{aligned}$$

which has a fixed point at

$$\begin{aligned} \langle m_j m_k \rangle &: = \lim_{t \rightarrow \infty} \langle m_j(t) m_k(t) \rangle \\ &= \frac{2 M \mu_m}{2 + 2 v \tau_u p_r - c v \tau_u p_r^2}. \end{aligned} \quad (22)$$

We now calculate the cross-covariance between $x_j(t)$ and $x_k(t)$. By a similar argument to that used to derive Eq. (14) above, the cross-covariance between $x_j(t)$ and $x_k(t)$ has the form

$$R_{x_j x_k}(\tau) = A_2 \delta(\tau) + B_2 e^{-|\tau|/\tau_0} \quad (23)$$

where we have used the symmetry of $x_j(t)$ and $x_k(t)$, inherited from the symmetry in parameters, to conclude that $R_{x_j x_k}(\tau) = R_{x_j x_k}(-\tau)$. The area of the delta function is given by

$$A_2 = \frac{\langle dN_{x_j} dN_{x_k} \rangle}{dt} = p_r^2 cv \langle m_j m_k \rangle$$

where $\langle m_j m_k \rangle$ is given in Eq. (22). To find B_2 , we first calculate

$$\begin{aligned} \lim_{\tau \rightarrow 0^+} \langle x_j(t) x_k(t + \tau) \rangle &= \\ \lim_{\tau \rightarrow 0^+} \left\langle \left(p_r m_j(t) I_j(t) + \sqrt{D_r} I_j(t) Z_{r,j}(t) \right) \right. \\ &\quad \left. \left(p_r m_k(t + \tau) I_k(t + \tau) + \sqrt{D_r} I_k(t + \tau) Z_{r,k}(t + \tau) \right) \right\rangle \\ &= p_r^2 \lim_{\tau \rightarrow 0^+} \langle m_j(t) I_j(t) m_k(t + \tau) I_k(t + \tau) \rangle \\ &= v^2 p_r^2 \lim_{\tau \rightarrow 0^+} \langle m_j(t) m_k(t + \tau) | dN_{I_j}(t), dN_{I_k}(t + \tau) > 0 \rangle \\ &= v^2 p_r^2 \lim_{\tau \rightarrow 0^+} \langle m_j(t) m_k(t + \tau) | dN_{I_j}(t) > 0 \rangle \\ &= v^2 p_r^2 \lim_{\tau \rightarrow 0^+} (1 - c) \langle m_j(t) m_k(t + \tau) | dN_{I_j}(t) > 0, dN_{I_k}(t) = 0 \rangle \\ &\quad + c \langle m_j(t) m_k(t + \tau) | dN_{I_j}(t), dN_{I_k}(t) > 0 \rangle \\ &= v^2 p_r^2 ((1 - c) \langle m_j(t) m_k(t) \rangle + c \langle m_j(t) (1 - p_r) m_k(t) \rangle) \\ &= v^2 p_r^2 (1 - c p_r) \langle m_j m_k \rangle \end{aligned}$$

so that

$$\begin{aligned} B_2 &= \lim_{\tau \rightarrow 0^+} R_{x_j x_k}(\tau) = \lim_{\tau \rightarrow 0^+} \langle x_j(t) x_k(t + \tau) \rangle - \langle x_j \rangle \langle x_k \rangle \\ &= v^2 p_r^2 (\langle m_j m_k \rangle - \mu_m^2) - c v^2 p_r^3 \langle m_j m_k \rangle \end{aligned}$$

which gives $R_{x_j x_k}(\tau)$ through Eqs. (22) and (23).

Finally, we will derive $R_{I_j x_k}(\tau)$ and $R_{I_k x_j}(\tau)$. Once again, by linearity, each of these is the sum of a delta function and an exponential. The area of the delta function is given by

$$\langle dN_{I_j}(t)dN_{x_k}(t) \rangle / dt = cv \langle dN_{x_k} | dN_{I_j}, dN_{I_k} > 0 \rangle = cv p_r \mu_m.$$

We also have

$$\begin{aligned} & \lim_{\tau \rightarrow 0^+} \langle I_j(t)x_k(t+\tau) \rangle \\ &= v^2 p_r \lim_{\tau \rightarrow 0^+} \langle m_k(t+\tau) | dN_{I_j}(t), dN_{I_k}(t+\tau) > 0 \rangle \\ &= v^2 p_r \lim_{\tau \rightarrow 0^+} (1-c) \langle m_k(t+\tau) | dN_{I_k}(t) = 0 \rangle \\ &\quad + c \langle m_k(t+\tau) | dN_{I_k}(t) > 0 \rangle \\ &= v^2 p_r ((1-c)\mu_m + c(1-p_r)\mu_m) \\ &= v^2 p_r (1 - c p_r) \mu_m \end{aligned}$$

Thus,

$$\begin{aligned} \lim_{\tau \rightarrow 0^+} R_{I_j x_k}(\tau) &= \lim_{\tau \rightarrow 0^+} \langle x_k(t+\tau) I_j(t) \rangle - \langle x_k(t) \rangle \langle I_j(t) \rangle \\ &= v^2 p_r (1 - c p_r) \mu_m - v p_r \mu_m v \\ &= -c v^2 p_r^2 \mu_m \end{aligned}$$

and therefore

$$R_{I_j x_k}(\tau) = c v p_r \mu_m \delta(\tau) - \Theta(\tau) c v^2 p_r^2 \mu_m e^{-\tau/\tau_0}.$$

By symmetry, $R_{I_k x_j}(\tau) = R_{I_j x_k}(\tau)$.

Finally, the cross-spectra can now be found through a Fourier transform to obtain

$$S_{I_j x_k}(f) = \tilde{K} c v \text{ and } S_{x_j x_k}(f) = (1 + D_0 c_0) |\tilde{K}|^2 c v$$

where

$$c_0 = \frac{v \tau_u (2 - p_r) p_r + 2}{v \tau_u (2 - c p_r) p_r + 2} c. \tag{24}$$

Statistics of the postsynaptic conductance

So far we have described the statistics of the processes, $x_k(t)$, which quantify the release of vesicles released over time. The postsynaptic conductance induced by vesicle release is then defined as $g_k = \alpha * x_k$ where $*$ denotes convolution and $\alpha(t)$ represents the time course of conductance induced by the release of a single vesicle (with $\alpha(t) = 0$ for $t < 0$). The statistics of $g_k(t)$ can easily be derived from those of $x_k(t)$ using standard signal processing identities [29] to give

$$\begin{aligned} \mu_g &= \frac{p_r v M}{1 + p_r \tau_u v} \int_0^\infty \alpha(t) dt \\ S_{I_k g_k} &= \tilde{a} \tilde{K} v \\ S_{g_k g_k} &= (1 + D_0) |\tilde{a} \tilde{K}|^2 v + |\tilde{a}|^2 (S_{\eta_u \eta_u} + S_{\eta_r \eta_r}) \\ S_{I_j g_k} &= \tilde{a} \tilde{K} c v \\ S_{g_j g_k} &= (1 + D_0 c_0) |\tilde{a} \tilde{K}|^2 c v \end{aligned} \tag{25}$$

for $j \neq k$ and the steady state variance of $g(t)$ is given by $\lim_{t \rightarrow \infty} \text{var}(g_k(t)) = 2 \int_0^\infty S_{g_k g_k}(f) df$.

Synaptic filtering of presynaptic spike trains with rate coded signals

So far, we have discussed statistics of the conductance induced by a population of homogeneous Poisson presynaptic spike trains, but spike trains measured *in vivo* do not always exhibit homogeneous Poisson statistics [55]. For example, time-varying stimuli can induce fluctuations in the firing rate of presynaptic neurons. As a simple model of rate-coded signals, we assume that a shared, time-varying signal, $s(t)$, is encoded in the firing rates of a presynaptic population, $\{I_k(t)\}_{k=1}^n$.

In this model, each presynaptic spike train is a doubly stochastic Poisson process [51]. The instantaneous firing rate of each presynaptic neuron, conditioned on $s(t)$, is given by $\langle I_k(t) | s(t) \rangle = v + s(t)$. Without loss of generality, we assume that the signal has zero bias, $\langle s(t) \rangle = 0$, so that the unconditioned firing rates are $\langle I_k(t) \rangle = v$. Signal correlations are introduced in this model by the shared signal, $s(t)$. We include noise correlations, i.e., correlations that are not due to shared signal [38,39], by assuming each pair of presynaptic spike trains share a proportion c of their spike times.

To compute the auto- and cross-covariance functions we first note that, for $\tau \neq 0$,

$$\begin{aligned} & \langle I_j(t) I_k(t+\tau) \rangle \\ &= \int \langle I_j(t) I_k(t+\tau) | s(t) = s_1, s(t+\tau) = s_2 \rangle dP_\tau(s_1, s_2) \\ &= \int s_1 s_2 dP_\tau(s_1, s_2) \\ &= \langle s(t) s(t+\tau) \rangle \end{aligned}$$

where $P_\tau(s_1, s_2)$ is distribution of $(s(t), s(t+\tau))$ in the steady state ($t \rightarrow \infty$). In addition, $R_{I_j I_k}(\tau)$ has a Dirac delta function at $\tau = 0$ with mass equal to the rate of synchronous spikes, cv . Thus, $R_{I_j I_k}(\tau) = cv \delta(\tau) + R_{ss}(\tau)$ for $j \neq k$. The auto-covariance ($j = k$) can be obtained by taking $c = 1$. The cross-covariance function between $s(t)$ and $I_k(t)$ is computed similarly to obtain $R_{sI}(\tau) = R_{ss}(\tau)$. Taking Fourier transforms gives the spectra,

$$\begin{aligned} S_{I_k I_k}(f) &= v + S_{ss}(f) \\ S_{I_j I_k}(f) &= cv + S_{ss}(f), \quad j \neq k \\ S_{sI_k}(f) &= S_{ss}(f) \end{aligned}$$

where $S_{ss}(f)$ is the power spectrum of the signal.

Exact expressions for the statistics of the postsynaptic conductance are difficult to obtain for this inhomogeneous Poisson model because $I_k(t)$ is correlated with $I_k(t+\tau)$ and with $I_j(t+\tau)$, which invalidates the methods used in the derivations for the homogeneous Poisson model above. However, when $S_{ss}(f) \ll v$, the firing rate inhomogeneities are weak compared to the background firing rate and temporal correlations are weak as a result (analogously, $S_{I_k I_k}(f) \approx v$). In this case, a linear approximation to the synaptic response can be obtained. To obtain this approximation, we find a linear filter that maps presynaptic spike trains to conductances and that is consistent with Eqs. (25) when inputs are Poisson. The following filter satisfies this requirement

$$\tilde{g}_k = \left((1 + \sqrt{D_0} \tilde{w}_{0,k}) \tilde{K} \tilde{I}_k + \tilde{\eta}_{u,k} + \tilde{\eta}_{r,k} \right) \tilde{a}. \quad (26)$$

Here, $w_{0,k}(t)$ is standard Gaussian white noise, $\eta_{u,k}(t)$ is unbiased stationary noise with power spectrum $S_{\eta_u \eta_u}(f)$ that accounts for stochasticity in vesicle recovery, and similarly for $\eta_{r,k}(t)$, which accounts for stochastic vesicle release. The noise terms $\eta_{u,k}(t)$ and $\eta_{r,k}(t)$ are zero for the deterministic model. All noise terms here are independent except that $w_{0,j}(t)$ and $w_{0,k}(t)$ are correlated with cross-spectrum

$$S_{w_{0,j} w_{0,k}}(f) = c_0, \quad j \neq k$$

where c_0 is given by Eq. (24).

The spectra predicted by Eq. (26) can be easily calculated using the fact that $S_{uv}(f) = \langle \tilde{u}^{0*} \tilde{v}^0 \rangle$ for stationary processes, $u(t)$ and $v(t)$, where $u^0(t) = u(t) - \langle u \rangle$ and $*$ denotes complex conjugation [56]. Thus,

$$\begin{aligned} S_{I_k g_k} &= \langle \tilde{I}_k^{0*} g_k^0 \rangle \\ &= \left\langle \tilde{I}_k^{0*} \left((1 + \sqrt{D_0} \tilde{w}_{0,k}) \tilde{K} \tilde{I}_k^0 + \tilde{\eta}_{u,k} + \tilde{\eta}_{r,k} \right) \tilde{a} \right\rangle \\ &= \tilde{K} \tilde{a} \langle \tilde{I}_k^{0*} \tilde{I}_k^0 \rangle \\ &= \tilde{K} \tilde{a} S_{I_k I_k} \end{aligned}$$

where we used the independence of the noise sources to eliminate several terms. Other spectra can be derived in a similar manner to obtain the following generalizations of Eqs. (25)

$$\begin{aligned} S_{g_k g_k} &= (1 + D_0) |\tilde{a} \tilde{K}|^2 S_{I_k I_k} + |\tilde{a}|^2 (S_{\eta_u \eta_u} + S_{\eta_r \eta_r}) \\ S_{I_j g_k} &= \tilde{a} \tilde{K} S_{I_j I_k} \\ S_{g_j g_k} &= (1 + D_0 c_0) |\tilde{K} \tilde{a}|^2 S_{I_j I_k} \end{aligned} \quad (27)$$

for $j \neq k$. These expressions agree with Eqs. (25) when inputs are Poisson, i.e., when $s(t) = 0$, because $S_{I_k I_k}(f) = v$ and $S_{I_j I_k}(f) = cv$ in this case. When $s(t) \neq 0$ and $S_{ss}(f) \ll v$, these expressions give a linear approximation which is verified using simulations in several figures below. The fidelity with which the signal, $s(t)$, is represented in the conductances, $g_k(t)$, depends on the cross-spectrum which can be calculated in analogous manner to $S_{I_k g_k}(f)$ above to obtain

$$\begin{aligned} S_{s g_k} &= \langle g_k^0 \tilde{s}^* \rangle \\ &= K a \langle \tilde{I}_k^0 \tilde{s}^* \rangle \\ &= K a S_{s I_k} = K a S_{ss}. \end{aligned} \quad (28)$$

References

- Zucker R, Regehr W (2002) Short-term synaptic plasticity. Annual Rev of Phys 64: 355–405.
- Tsodyks MV, Markram H (1997) The neural code between neocortical pyramidal neurons depends on neurotransmitter release probability. Proc Natl Acad Sci U S A 94: 719–723.
- Abbott LF, Varela JA, Sen K, Nelson SB (1997) Synaptic depression and cortical gain control. Science 275: 220–224.
- Tsodyks M, Pawelzik K, Markram H (1998) Neural networks with dynamic synapses. Neural Comput 10: 821–835.
- Markram H, Wang Y, Tsodyks M (1998) Differential signaling via the same axon of neocortical pyramidal neurons. Proc Natl Acad Sci U S A 95: 5323.
- Senn W, Markram H, Tsodyks M (2001) An algorithm for modifying neurotransmitter release probability based on pre- and postsynaptic spike timing. Neural Comput 13: 35–67.

We are especially interested in the population spectra, $S_{I_g}(f)$, $S_{s g}(f)$ and $S_{g g}(f)$, where $I(t) = \sum_k I_k(t)$ is the total presynaptic input and $g(t) = \sum_{k=1}^n g_k(t)$ is the total conductance induced by $I(t)$. These are given by using the bilinearity of covariances to obtain

$$\begin{aligned} S_{I_g}(f) &= n(n-1) S_{I_j g_k}(f) + n S_{I_k g_k}(f) \\ S_{s g}(f) &= n S_{s g_k}(f) \\ S_{g g}(f) &= n(n-1) S_{g_j g_k}(f) + n S_{g_k g_k}(f). \end{aligned} \quad (29)$$

A similar inhomogeneous Poisson input model was used in [17] to investigate the transfer of rate-coded signals for the deterministic model of synaptic depression. Their model is analogous to our deterministic model with $M = 1$ (since their response amplitudes are normalized) and $\tilde{\alpha}(f) = 1$ (since they consider the postsynaptic response, $x(t)$, before convolution with a conductance kernel). Under these substitutions, our expression for $S_{s g_k}(f)$ agrees with their expression for $S_{R_x}^M(f)$ exactly (where we use an “ M ” superscript to indicate expressions from [17]). However, our expression for $S_{g_k g_k}(f)$ for the deterministic model only agrees with their expression for $S_{xx}^M(f)$ when $s(t) = 0$ (i.e., when the input is a homogeneous Poisson process). Our expression has an additional term that accounts for power introduced by the signal $s(t)$. In particular, $S_{g_k g_k}(f) = S_{xx}^M(f)(1 + S_{ss}(f))$ for the deterministic model when $M = 1$ and $\tilde{\alpha}(f) = 1$.

Parameters used for figures

Theoretical results are obtained for arbitrary parameter values, but for all figures we use the parameters from Table 1, which are chosen to represent values from experimental studies. The values used for τ_u and p_r have been deemed “typical” for pyramidal-to-pyramidal synapses in the rodent neocortex [2,19] and the value of M is typical for several cortical areas [34]. The form of $\alpha(t)$ is chosen to model AMPA dynamics and its units are rescaled so that $\int_0^\infty \alpha(t) dt = 1$. This rescaling simplifies the exposition in the Results.

Acknowledgments

We would like to thank Anne-Marie Oswald for comments on an early draft of this manuscript. We would also like to thank the reviewers for several helpful suggestions that improved the quality and clarity of this manuscript.

Author Contributions

Wrote the paper: RR JR BD. Conceived the study: RR JR BD Performed calculations and simulations: RR.

7. Varela JA, Sen K, Gibson J, Fost J, Abbott LF, et al. (1997) A quantitative description of short-term plasticity at excitatory synapses in layer 2/3 of rat primary visual cortex. *J Neurosci* 17: 7926–7940.
8. Hanson JE, Jaeger D (2002) Short-term plasticity shapes the response to simulated normal and parkinsonian input patterns in the globus pallidus. *J Neurosci* 22: 5164–5172.
9. Chance F, Nelson S, Abbott L (1998) Synaptic depression and the temporal response characteristics of v1 cells. *J Neurosci* 18: 4785.
10. Cook DL, Schwandt PC, Grande LA, Spain WJ (2003) Synaptic depression in the localization of sound. *Nature* 421: 66–70.
11. Heiss J, Katz Y, Ganmor E, Lampl I (2008) Shift in the balance between excitation and inhibition during sensory adaptation of s1 neurons. *J Neurosci* 28: 13320.
12. de la Rocha J, Parga N (2005) Short-term synaptic depression causes a non-monotonic response to correlated stimuli. *J Neurosci* 25: 8416–8431.
13. Tsodyks M, Uziel A, Markram H, et al. (2000) Synchrony generation in recurrent networks with frequency-dependent synapses. *J Neurosci* 20: 825–835.
14. Vladimirovskii B, Tabak J, O'Donovan M, Rinzel J (2008) Episodic activity in a heterogeneous excitatory network, from spiking neurons to mean field. *J Comput Neurosci* 25: 39–63.
15. Barbieri F, Brunel N (2008) Can attractor network models account for the statistics of firing during persistent activity in prefrontal cortex? *Front Neurosci* 2: 114.
16. Lindner B, Gangloff D, Longtin A, Lewis JE (2009) Broadband Coding with Dynamic Synapses. *J Neurosci* 29: 2076–2087.
17. Merkel M, Lindner B (2010) Synaptic filtering of rate-coded information. *Phys Rev E* 81: 041921.
18. Oswald A, Urban N (2012) Interactions between behaviorally relevant rhythms and synaptic plasticity alter coding in the piriform cortex. *J Neurosci* 32: 6092–104.
19. Fuhrmann G, Segev I, Markram H, Tsodyks M (2002) Coding of temporal information by activity-dependent synapses. *J Neurophys* 87: 140.
20. Goldman M (2004) Enhancement of information transmission efficiency by synaptic failures. *Neural Comput* 16: 1137–1162.
21. Korn H, Triller A, Mallet A, Faber D (1981) Fluctuating responses at a central synapse: n of binomial fit predicts number of stained presynaptic boutons. *Science* 213: 898–901.
22. Loebel A, Silberberg G, Helbig D, Markram H, Tsodyks M, et al. (2009) Multiquantal release underlies the distribution of synaptic efficacies in the neocortex. *Front Comput Neurosci* 3:27.
23. Huang C, Bao J, Sakaba T (2010) Multivesicular release differentiates the reliability of synaptic transmission between the visual cortex and the somatosensory cortex. *J Neurosci* 30: 11994–12004.
24. Vere-Jones D (1966) Simple stochastic models for the release of quanta of transmitter from a nerve terminal. *Aust N Z J Stat* 8: 53–63.
25. Wang XJ (1999) Fast burst firing and short-term synaptic plasticity: a model of neocortical chattering neurons. *J Neurosci* 19: 347–362.
26. Rothman JS, Cathala L, Steuber V, Silver RA (2009) Synaptic depression enables neuronal gain control. *Nature* 457: 1015–1018.
27. Grande LA, Spain WJ (2005) Synaptic depression as a timing device. *J Physiol* 20: 201–210.
28. de la Rocha J, Parga N (2008) Thalamocortical transformations of periodic stimuli: the effect of stimulus velocity and synaptic short-term depression in the vibrissa-barrel system. *J Comput Neurosci* 25: 122–140.
29. Tetzlaff T, Rotter S, Stark E, Abeles M, Aertsen A, et al. (2008) Dependence of neuronal correlations on filter characteristics and marginal spike train statistics. *Neural Comput* 20: 2133–2184.
30. Stratonovich R (1963) Topics in the theory of random noise. London: Gordon & Breach Science Publishers.
31. de la Rocha J, Moreno R, Parga N (2004) Correlations modulate the non-monotonic response of a neuron with short-term plasticity. *Neurocomputing* 58: 313–319.
32. Galarreta M, Hestrin S (1998) Frequency-dependent synaptic depression and the balance of excitation and inhibition in the neocortex. *Nature Neurosci* 1: 587–594.
33. Rav-Acha M, Sagiv N, Segev I, Bergman H, Yarom Y (2005) Dynamic and spatial features of the inhibitory pallidal GABAergic synapses. *J Neurosci* 135: 791–802.
34. Branco T, Staras K (2009) The probability of neurotransmitter release: variability and feedback control at single synapses. *Nat Rev Neurosci* 10: 373–383.
35. Dayan P, Abbott L (2001) Theoretical Neuroscience: Computational and mathematical modeling of neural systems. Taylor & Francis.
36. Gabbiani F, Koch C, eds. (1998) Principles of spike train analysis. Methods in Neuronal Modeling Cambridge, MA: MIT Press. pp 313–360.
37. Rieke F (1999) Spikes: exploring the neural code. The MIT Press.
38. Averbeck B, Latham P, Pouget A (2006) Neural correlations, population coding and computation. *Nat Rev Neurosci* 7: 358–366.
39. Cohen M, Kohn A (2011) Measuring and interpreting neuronal correlations. *Nat Neurosci* 14: 811–819.
40. Abbott LF, Regehr WG (2004) Synaptic computation. *Nature* 431: 796–803.
41. Silver RA (2010) Neuronal arithmetic. *Nat Rev Neurosci* 11: 474–489.
42. Dobrunz L (2002) Release probability is regulated by the size of the readily releasable vesicle pool at excitatory synapses in hippocampus. *Int J Developmental Neurosci* 20: 225–236.
43. Bekkers J, Richerson G, Stevens C (1990) Origin of variability in quantal size in cultured hippocampal neurons and hippocampal slices. *Proc Natl Acad Sci U S A* 87: 5359.
44. McAllister A, Stevens C (2000) Nonsaturation of ampa and nmda receptors at hippocampal synapses. *Proc Natl Acad Sci U S A* 97: 6173.
45. White J, Rubinstein J, Kay A (2000) Channel noise in neurons. *Trends Neurosci* 16: 3219–3235.
46. Rotman Z, Deng PY, Klyachko VA (2011) Short-Term Plasticity Optimizes Synaptic Information Transmission. *J Neurosci* 31: 14800–14809.
47. Cover T, Thomas J, Wiley J (1991) Elements of information theory. Volume 6. Wiley Online Library.
48. Zador A (1998) Impact of synaptic unreliability on the information transmitted by spiking neurons. *J Neurophys* 79: 1219.
49. Gardiner C (1985) Handbook of stochastic methods. Berlin: Springer.
50. Karlin S, Taylor H (1981) A second course in stochastic processes. Gulf Professional Publishing.
51. Cox D, Isham V (1980) Point processes. London: Chapman and Hall.
52. Øksendal B (2003) Stochastic differential equations: an introduction with applications. Springer Verlag.
53. Zohary E, Shadlen M, Newsome W (1994) Correlated neuronal discharge rate and its implications for psychophysical performance. *Nature* 370: 140–3.
54. Shea-Brown E, Josić K, de la Rocha J, Doiron B (2008) Correlation and synchrony transfer in integrate-and-fire neurons: basic properties and consequences for coding. *Phys Rev L* 100: 108102.
55. Maimon G, Assad J (2009) Beyond poisson: increased spike-time regularity across primate parietal cortex. *Neuron* 62: 426–440.
56. Yaglom A (2004) An introduction to the theory of stationary random functions. Dover Publications.