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Can Vitamin D Supplementation in Addition to Asthma Controllers Improve Clinical Outcomes in Patients With Asthma?

A Meta-Analysis

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Abstract: Effects of vitamin D on acute exacerbation, lung function, and fraction of exhaled nitric oxide (FeNO) in patients with asthma are controversial. We aim to further evaluate the roles of vitamin D supplementation in addition to asthma controllers in asthmatics.

From 1946 to July 2015, we searched the PubMed, Embase, Medline, Cochrane Central Register of Controlled Trials, and ISI Web of Science using "Vitamin D," "Vit D," or "VitD" and "asthma," and manually reviewed the references listed in the identified articles. Randomized controlled trials which reported rate of asthma exacerbations and adverse events, forced expiratory volume in 1 s (FEV₁, % of predicted value), FeNO, asthma control test (ACT), and serum 25hydroxyvitamin D levels were eligible. We conducted the heterogeneities test and sensitivity analysis of the enrolled studies, and randomeffects or fixed-effects model was applied to calculate risk ratio (RR) and mean difference for dichotomous and continuous data, respectively. Cochrane systematic review software Review Manager (RevMan) was used to test the hypothesis by Mann–Whitney U test, which were displayed in Forest plots.

Seven trials with a total of 903 patients with asthma were pooled in our final studies. Except for asthma exacerbations (I2 = 81%, $\chi 2 = 10.28$, P = 0.006), we did not find statistical heterogeneity in outcome measures. The pooled RR of asthma exacerbation was 0.66 (95% confidence interval: 0.32–1.37), but without significant difference (z = 1.12, P = 0.26), neither was in FEV₁ (z = 0.30, P = 0.77), FeNO (z = 0.28, P = 0.78), or ACT (z = 0.92, P = 0.36), although serum 25hydroxyvitamin D was significantly increased (z = 6.16, P < 0.001).

Vitamin D supplementation in addition to asthma controllers cannot decrease asthma exacerbation and FeNO, nor improve lung function and asthma symptoms, although it can be safely applied to increase serum 25-hydroxyvitamin D levels.

JL and DL have contributed equally to this work.

JL and DL designed the study, conducted the literature searching and data analysis, and prepared the manuscript; C-TL reviewed the manuscript and made the decision to submit the report for publication.

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Abbreviations: ACTa = sthma control test, ASM = airway smooth muscle, BMI = body mass index, CENTRAL = Cochrane Central Register of Controlled Trials, CI = confidence interval, FeNO = fraction of exhaled nitric oxide, FEV₁ = forced expiratory volume in 1 s, FEV₁/FVC = ratio of FEV₁ to forced vital capacity, GINA = the Global Initiative for Asthma, HR = hazard ratio, ILs = interleukins, INF- γ = interferon- γ , MAP = mitogen-activated protein, MKP-1 = mitogen-activated protein kinase 1, NO = nitric oxide, NOS = nitric oxide synthase, OR = odds ratio, RCT = randomized controlled trial, RR = risk ratio, SD = standard derivation, Th lymphocytes = T helper lymphocytes, TNF- α = tumor necrosis factor alpha, VDD = vitamin D deficiency, VDR = vitamin D receptor.

INTRODUCTION

A sthma is a chronic inflammatory disorder of the airways in which many cells and cellular elements involved, such as eosinophils, mast cells, and T helper (Th) lymphocytes.¹ It is estimated an annual worldwide deaths from asthma is as high as 250,000, and acute exacerbations are major causes of morbidity and mortality.² A variety of factors may cause asthma exacerbations, including allergens, pollutants, and poor inherence to the controllers, while viral upper respiratory infections are predominant reasons. In order to improve asthma clinical outcomes, it is essential to seek strategies to decrease asthma exacerbations.

Vitamin D were reported to prevent adverse influence of corticosteroids, which are the core controllers for asthma due to the potent anti-inflammation effects, such as bone mass density decrease and fractures.^{3,4} Observational studies have demonstrated that lower serum 25-hydroxyvitamin D levels were associated with worse lung function, worse corticosteroid responsiveness, greater exacerbation frequency, as well as more susceptibility to upper respiratory infections in patients with asthma.^{5–7} Arshi et al⁸ enrolled 130 patients with moderate persistent asthma in an open-label randomized controlled trial (RCT) and they found that vitamin D supplementation significantly improved forced expiratory volume in 1 s (FEV₁) and ratio of FEV₁ to forced vital capacity (FEV₁/FVC). However, Castro et al⁹ recently illustrated that treatment with vitamin D had no effect on lung function (FEV₁: -0.07 [95% confidence interval (CI): -0.14 to 0.01] vs -0.04 [95% CI: -0.11 to 0.03], P = 0.64), and did not reduce the rate of first asthma exacerbation (13% [95% CI: 8-18%] vs 19% [95% CI: 13-24%], P = 0.21). Therefore, the effects of vitamin D supplementation in addition to asthma controllers on acute exacerbation and lung

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function in patients with asthma are still controversial thus necessitating further evaluations.

In this study, we conducted a meta-analysis of all published RCTs to further clarify the efficacy and safety of vitamin D in patients with asthma.

METHODS

The Institutional Ethical Committee for Clinical and Biomedical Research of West China Hospital (Sichuan, China) approved our study protocol. Each enrolled trial included approved protocol by respective institutional review board, and all participants provided written informed consent.

Search Strategies

We conducted a comprehensive computer search, from 1946 to July 2015, in PubMed, Embase, Medline, Cochrane Central Register of Controlled Trials (CENTRAL), and ISI Web of Science using "Vitamin D," "Vit D," or "VitD" and "asthma." Publication type of RCTs was limited. A review of references listed in the identified articles and a manual search of the related articles were performed to identify all relevant and eligible studies and minimize publication bias.

Inclusion and Exclusion Criteria

Eligible clinical trials were defined based on the following criteria: study design was RCT; asthma was diagnosed by physicians according to the Global Initiative for Asthma (GINA) or relevant guidelines before randomization; intervention treatment was vitamin D, regardless of the drug names, doses, and administration routines; outcome measures included rate of asthma exacerbations and adverse events, lung functions such as FEV₁ (% of predicted value), fraction of exhaled nitric oxide (FeNO), asthma symptom scores such as asthma control test (ACT), and serum 25-hydroxyvitamin D levels. We did not include trials that were nonrandomized controlled or observational, neither were cohort or case–control studies.

Study Selection

Two investigators conducted study selection independently in 2 phases. First, they discarded duplicated and nonrandomized controlled studies by screening titles and abstracts. Second, eligible studies were extracted by reviewing full texts according to the study inclusion criteria. Any disagreement was solved by mutual consensus in the presence of a third investigator.

Data Extraction

Independently, 2 data collectors extracted and recorded desirable information of each enrolled study in a standard form recommended by Cochrane,¹⁰ which included authors, publication year, country, study design, participants and population, demographic characteristics (age, gender, body mass index [BMI], etc.), baseline characteristics (FEV₁, FeNO, serum 25-hydroxyvitamin D levels, etc.), details of intervention treatment (drug name, dose, administration routine, and duration), follow-up period, and outcome measures and study results. For any missing data information, corresponding authors were contacted via email to request the full original data. Different opinions between the 2 collectors were determined by mutual consensus in the presence of a third investigator.

Quality Assessment

We used the Cochrane risk of bias tool to assess the quality of methods conducted and outcomes reported by each eligible study.¹⁰ The studies were considered as good quality if they meet the following criteria: random sequence generation (low selection bias); allocation concealment (low selection bias); blinding of participants and personnel (low performance bias); blinding of related outcomes assessment (low detection bias); incomplete outcome data (low attrition bias); selective reporting (low reporting bias); and without other bias. Two independent assessors performed the quality assessment, and any divergence was resolved by reaching a consensus or consulting a third investigator.

Statistical Analysis

An independent statistician performed the statistical analysis using Cochrane systematic review software Review Manager (RevMan; Version 5.3.5., The Nordic Cochrane Centre, The Cochrane Collaboration, Copenhagen, 2014). Mann–Whitney U test was used for hypothesis test with statistical significance rendered as z-value and P value < 0.05, 0.05, and the results were displayed in Forest plots.

Dichotomous variables were reported as frequency and proportion, while continuous variables were showed as mean and standard derivation (SD). An initial test for clinical, methodological, and statistical heterogeneities was conducted, and we used the χ^2 test with P < 0.1 and $I^2 > 50\%$ indicating significance. We also performed the sensitivity analysis to substitute alternative decisions or ranges of values for decisions that were arbitrary or unclear.

Random-effects model was applied in the presence of statistical heterogeneity; otherwise fixed-effects model was used. For dichotomous data we calculated risk ratio (RR) and 95% CI, while for continuous data we calculated mean difference and SD. Furthermore, in terms of FEV₁, FeNO, ACT, and serum 25-hydroxyvitamin D levels, we separately conducted sub-analysis at different follow-up time points.

RESULTS

Primordially we identified 90 records in the electronic databases and extracted another 2 records from the reference lists. After screening the titles and abstracts, 83 studies were discarded, of which 25 studies were duplicated, 36 studies were not RCTs, 17 studies did not enroll asthmatic patients, and 5 studies did not prescribe vitamin D as intervention treatment. The remaining 9 studies were searched for full-text articles, and eventually we enrolled 7 trials^{3,4,9,11-14} in the final analysis, because the 2 studies discarded did not report our predefined outcome measures (Fig. 1).

Study Description

Seven enrolled trials were conducted separately in different countries, that is, Germany, Poland, Turkey, Hershey, India, The Netherlands, and The United Kingdom. Three studies were conducted in children,^{4,11,12} while 4 studies were in adults.^{3,9,13,14} Patients with clinically stable asthma, persistent asthma, and IgE-dependent asthma were, respectively, enrolled in 1 study,^{3,4,11} while the other 4 studies did not report the clinical features of the enrolled patients. Two studies applied identical dose (1000 IU) of vitamin D,^{3,4} and 2 studies^{4,11} and 4 studies,^{9,12–14} respectively, administered vitamin D subcutaneously and orally. Treatment duration were reported in 3 studies,^{9,12,14} and all studies reported the follow-up period, 1 in 9 weeks,¹³ 1 in 28 weeks,⁹ 2 in 6 months,^{3,12} and 3 in 12 months.^{4,11,14} In terms of the outcome measurements, 3 studies

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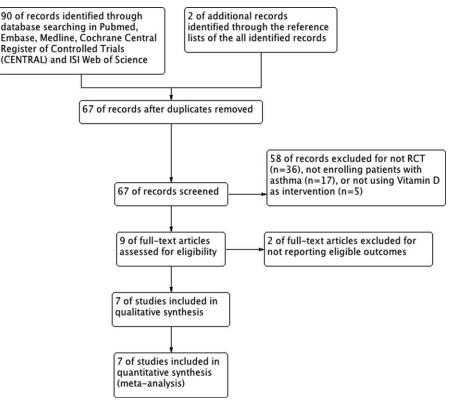


FIGURE 1. Study flow diagram. RCT = randomized controlled trial.

reported asthma exacerbation rate, 9,12,14 4 studies reported change of FEV₁ (% of predicted value), $^{4,11,13,14}_{4,11,13,14}$ 2 studies reported change of FeNO, $^{13,14}_{4,11,13,14}$ 2 studies provided data on change of ACT, $^{9,14}_{4,14}$ 4 studies provided change levels of 25hydroxyvitamin D, $^{3,4,11,14}_{3,4,11,14}$ and 3 studies provided rate of adverse events. 11,13,14 Details of patients' characteristics, intervention strategies, and outcomes were summarized in Table 1.

A total of 903 patients with asthma were pooled in our final studies, of which, 453 (50.2%) received vitamin D supplementation in addition to common asthma controllers. Baseline characteristics of the patients in the 7 enrolled studies were summarized in Table 2, all studies showed the data of sex proportion, but not for other demographic measures, such as age, BMI, FEV₁, FeNO, as well as ACT. Baseline serum 25hydroxyvitamin D levels in patients receiving vitamin D supplementation ranged from 19.0 to 21.0 ng/mL (49.8–60.0 nmol/L), which were also included in Table 2.

In patients with administration of vitamin D, 3 trials^{11,13,14} reported serious adverse events such as malignant cancer, acute appendicitis, internal fixation and upper gastrointestinal bleeding, and nonserious adverse events; for example, local urticarial plaques at the injection sites, acute respiratory infection, allergy symptoms, and psychiatric events. However, hypercalcemia was not observed in any of the patients.

Quality assessment of the 7 enrolled studies showed no bias in selection, attribution, or reporting, although 2 studies^{3,11} did not blind participants and personnel as well as outcome assessment (Figs. 2 and 3). We did sensitivity analysis, which resulted in no studies excluding for low quality or dubious decisions.

Heterogeneity

For all outcome measures, except for asthma exacerbations (I2 = 81%, $\chi 2$ = 10.28, P = 0.006), we did not find statistical heterogeneity (Figs. 4–9), thus we used fixed-effects model to analyze the effects of vitamin D on FEV₁, FeNO, ACT, and serum 25-hydroxyvitamin D levels.

Outcomes

The pooled RR of asthma exacerbation and adverse events rate was 0.66 (95% CI: 0.32–1.37) and 1.16 (95% CI: 0.74– 1.81), but both data analysis did not show any significant differences (z = 1.12, P = 0.26 and z = 0.64, P = 0.52, respectively) (Figs. 4 and 9). Figure 5 showed that the pooled SD of FEV₁ change at the 2-month and 12-month follow-up point from baseline were -0.03 (95% CI: -0.26 to 0.20) and 0.01 (95% CI: -0.21 to 0.23), but they did not show significant differences (z = 0.23, P = 0.82 and z = 0.07, P = 0.94, respectively), neither did the overall effect (z = 0.30, P = 0.77). Figures 6 and 7, respectively, presented the pooled effects of vitamin D on FeNO and ACT; however, we did not find significant differences at different follow-up points or in the overall effects.

Four studies reported the effects of vitamin D on serum 25hydroxyvitamin D, of which 2 studies reported the effects at 12-month after enrollment (Fig. 8). The pooled SD of serum 25hydroxyvitamin D change at 12-month follow-up point from baseline and the overall effect was 0.67 (95% CI: 0.43–0.91) and 0.51 (95% CI: 0.35–0.68), and both analyses show vitamin D supplementation could significantly increase serum 25-

				IJ	Intervention					
Author, Year	Country	Participants	No	Drug	Dose	Routine	Control	Duration	Duration Follow-Up	Outcomes
Worth, 1994	Germany	Clinically stable asthmatics without hypoxemia who received more than 10 mg prednisone or an equivalent dose of another corticosteroid for	33	Vit D + EHDP + calcium + ICS	Vit D: 1000 IU	WZ	Calcium + ICS	WN	6 mo	Ś
Majak, 2009	Poland	more than 9 mo IgE-dependent asthma with regular symptoms requiring long- term treatment with ICS and disease duration of	36	Vit D + ICS (prednisone)	Vit D3: 1000 IU	SC	ICS (prednisone)	WN	12 mo	, 2
Baris, 2014	Turkey	at least 2 y Mild to moderate	32	Vit D + SCIT	Vit D: 650 IU	SC	SCIT	MN	12 mo	2, 5, 6
Castro, 2014	Hershey	persistent asthma Asthma with a serum 25- hydroxyvitamin D level of <30 ng/	408	Vit D + ICS (ciclesonide) + levalbuterol	Vit D: 100,000 IU, followed by 4000 IU	Ю	Placebo + ICS (ciclesonide) + levalbuterol	28 wk	28 wk	1, 4
Yadav, 2014	India	Moderate to severe	100	Vit $D + asthma$	Vit D: 60,000 IU	Ю	Placebo + asthma	6 mo	6 mo	1
De Groot, 2015	The Netherlands	asunna Asthma	44	controllers Vita D + asthma controllers	Vit D: 40,000 IU	Ю	controllers Placebo + asthma controllers	MN	9 wk	2, 3, 6
Martineau, 2015	The United Kingdom	The United Kingdom Asthma treated with ICS	250	Vit $D + ICS$	VitD: 120,000 IU	Ю	Placebo + ICS	12 mo	12 mo	1, 2, 3, 4, 5, 6

Author, Year	Participants	No*	$\substack{ \mathbf{Age} \\ (\mathbf{y}, \mathbf{SD})^{*} }$	Sex (Male, %) [*] (Sex BMI (Male, %)* (kg/m ² , SD)*	dicted Value, SD)*	FeNO (ppb, SD)*	ACT (Score, SD)*	Sputum Eosinophil Count (%, SD)*	Concentration (ng/mL, SD)*
Worth, 1994	Clinically stable asthmatics without hypoxemia who received more than 10 mg prednisone or an equivalent dose of another corticosteroid for	f 1	55 (11)	3 (21%)	WZ	WN	WN	WN	WN	21.0 (2.0)
Majak, 2009	IgE-dependent asthma with regular symptoms requiring long- term treatment with ICS and disease duration of	f 18	WZ	10 (55%)	WX	93% (95% CI, 92– 95)	MN	WN	MN	ЖX
Baris, 2014	at least $\angle y$ Mild to moderate	17	9.2 (2.6)	5 (29%)	MN	NM	MN	NM	NM	19 (9)
Castro, 2014	persistent asthma Asthma with a serum 207 25- hydroxyvitamin D level of <30 ng/ mI	с 207 г	39.9 (13.1)	66 (31.9%)	66 (31.9%) 32.00 (8.19)	80.7% (13.8%)	MM	19.0 (12.0–26.0)	19.0 (12.0–26.0) 0.3 (0.00–1.30)	19.9 (14.5–25.0)
Yadav, 2014	Moderate to severe	50	9.15 (2.4)	24 (48%)	NM	NM	MN	MM	NM	NM
De Groot, 2015 Asthma Martineau, 2015 Asthma ICS	De Groot, 2015 Asthma 22 Martineau, 2015 Asthma treated with 125 ICS	22 1 125	59.0 (9.7) 49.4 (14.8)	11 (50%) 55 (44%)	26.6 (4.2) NM	99.1% (15.7%) 82.0% (18.7%)	24 (19–36) 38.1 (29.1)	NM 19.2 (3.9)	3.2 (0.3–13.3) NM	60 nmol/L (47–78) 49.8 nmol/L (25.2)

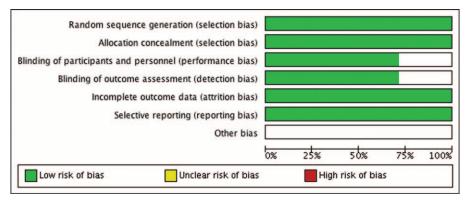


FIGURE 2. Risk of bias graph.

hydroxyvitamin D levels (z = 5.50, P < 0.001 and z = 6.16, P < 0.001).

DISCUSSION

This meta-analysis included 7 studies, and the data analysis showed that vitamin D supplementation in addition to asthma controllers could not decrease asthma exacerbation and FeNO,

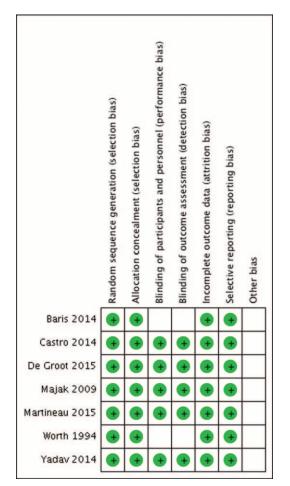
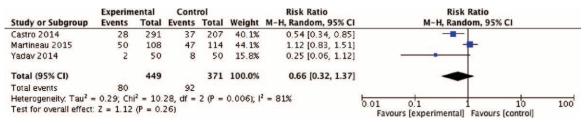


FIGURE 3. Risk of bias summary.

nor improve lung function (FEV₁, % of predicted value) and asthma symptoms (ACT), although it could significantly increase serum 25-hydroxyvitamin D without inducing higher adverse events frequency.

Recently, Kerley et al^{15,16} reviewed the potential mechanisms underlying the influence of vitamin \hat{D} on asthmatic disease, and summarized the current human literature involving the possible consequences of vitamin D deficiency (VDD) for asthma. The interaction between vitamin D and the pathogenesis of asthma is complex, but the known mechanisms raise the potential for vitamin D repletion as adjunct therapy which include: effects on lung development (structural effects), VDD could alter lung structure and create deficits in lung function thus leading to permanent susceptibility to poorer respiratory outcomes; anti-inflammatory effects, vitamin D could inhibit the production of pro-inflammatory cytokines and vitamin D receptor (VDR) could suppress NF-kB activation and signaling; immunomodulation, addition of vitamin D to human monocytes could inhibits the expression of the toll-like receptors 2 and 4 resulting in reduced production of tumor necrosis factor alpha (TNF- α), but the role of vitamin D on Th2 cells is not consistent; airway smooth muscle (ASM) modulation, ASM cells possess the enzymatic machinery to form 1,25-dihydroxyvitamin D from 25-hydroxyvitamin D and 1,25dihydroxyvitamin D could in turn suppress ASM proliferation and decrease expression of inflammatory chemokines; altering the effect of anti-asthmatic therapy, VDD may lead to downregulation of corticosteroids pathways and vitamin D could improve corticosteroids response by increasing mitogen-activated protein kinase 1 (MKP-1), which is a protein involved in directing cellular responses to a diverse array of stimuli. However, the consequences from human studies were inconsistent with some demonstrating benefit from increased vitamin D intake, while others suggesting inferior asthmatic outcomes associated with higher vitamin D, thus activating our interests in further evaluating the effects of vitamin D supplementation in patients with asthma.

Poorly controlled airway inflammation and hyperresponsiveness are pathological and clinical hallmarks underlying asthma exacerbations, which can be induced and enhanced by interleukins (ILs) released by Th lymphocytes via activation of arginase activity.^{17,18} Zhang et al found that the production of IL-6 and TNF- α could be suppressed by vitamin D via the inhibition of p38 mitogen-activated protein (MAP) kinase in monocytes,¹⁹ and several studies have also shown that higher levels of IL-17A and IL-22 were





	Expe	erimen	tal	C	ontrol		1	Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
2.1.1 Effect on FEV1	at 2-m	onth							
De Groot 2015	-0.11	1.42	22	-0.2	1.41	22	5.1%	0.06 [-0.53, 0.65]	
Martineau 2015 Subtotal (95% CI)	-0.03	1.4	125 147	0.03	1.43	125 147	29.1% 34.2%	-0.04 [-0.29, 0.21] -0.03 [-0.26, 0.20]	*
Heterogeneity. Chi ² =	0.10, di	f = 1(P = 0.7	'5); 1 ² =	0%				
Test for overall effect	Z = 0.2	3 (P =	0.82)						
2.1.2 Effect on FEV1	at 6-m	onth							
Martineau 2015 Subtotal (95% CI)	-0.07	1.43	125 125	0	1.43	125 125	29.1% 29.1%	-0.05 [-0.30, 0.20] -0.05 [-0.30, 0.20]	‡
Heterogeneity. Not ap	plicable								121 0
Test for overall effect	Z = 0.3	9 (P =	0.70)						
2.1.3 Effect on FEV1	at 12-n	nonth							
Baris 2014	0.4	1.43	17	0.08	1.41	15	3.7%	0.22 [-0.48, 0.92]	
Majak 2009	0.15	1.42	17	0.49	1.43	17	3.9%	-0.23 [-0.91, 0.44]	
Martineau 2015 Subtotal (95% CI)	-0.02	1.43	125 159	-0.04	1.43	125 157	29.1% 36.7%	0.01 [-0.23, 0.26] 0.01 [-0.21, 0.23]	‡
Heterogeneity. $Chi^2 =$	0.85. dt	f = 2(P = 0.6	61; 1 ² =	0%				
Test for overall effect									
Total (95% CI)			431			429	100.0%	-0.02 [-0.15, 0.11]	•
Heterogeneity. Chi ² =	1.07. dt	f = 5(P = 0.9	(6); I ² =	0%				<u> </u>
Test for overall effect									-4 -2 0 2 4
Test for subgroup dif				df = 2	(P = 0)	94), 12	= 0%		Favours [control] Favours [experimental]

FIGURE 5. Effect of vitamin D versus placebo on FEV_1 (% of predicted value). $CI = confidence interval, FEV_1 = forced expiratory volume in 1 s, SD = standard derivation.$

	Expe	erimen	tal	C	ontrol		1	Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
3.1.1 Effect on FeNO	at 2-m	onth							
De Groot 2015	-0.11	1.41	22	-0.11	1.42	22	5.5%	0.00 [-0.59, 0.59]	
Martineau 2015 Subtotal (95% CI)	-0.07	1.43	125 147	-0.02	1.43	125 147	31.5% 37.0%	-0.03 [-0.28, 0.21] -0.03 [-0.26, 0.20]	↓
Heterogeneity. Chi ² =	0.01, d	f = 1()	P = 0.9	(2); $I^2 =$	0%				
Test for overall effect:	Z = 0.2	5 (P =	0.80)	1					
3.1.2 Effect on FeNO	at 6-m	onth							
Martineau 2015 Subtotal (95% CI)	-0.05	1.43	125 125	-0.08	1.43	125 125	31.5% 31.5%	0.02 [-0.23, 0.27] 0.02 [-0.23, 0.27]	‡
Heterogeneity. Not ap	plicable								(2 TA)
Test for overall effect:	Z = 0.1	.7 (P =	0.87)						
3.1.3 Effect on FeNO	at 12-1	nonth							
Martineau 2015 Subtotal (95% CI)	-0.02	1.43	125 125		1.4	125 125	31.5% 31.5%	-0.05 [-0.30, 0.20] -0.05 [-0.30, 0.20]	‡
Heterogeneity. Not ap	plicable								
Test for overall effect:	Z = 0.3	9 (P =	0.70)						
Total (95% CI)			397			397	100.0%	-0.02 [-0.16, 0.12]	•
Heterogeneity. Chi ² =	0.18, d	f = 3(0)	P = 0.9	(8); 1 ² =	0%				4 5 8 3
Test for overall effect:	Z = 0.2	8 (P =	0.78)						Favours [experimental] Favours [control]
Test for subgroup diff	ferences	Chi ² =	0.17,	df = 2	(P = 0	92), 12	= 0%		ravours (experimental) ravours (control)

FIGURE 6. Effect of vitamin D versus placebo on FeNO. CI = confidence interval, FeNO = fraction of exhaled nitric oxide, SD = standard derivation.

	Expe	Experimental			Control			Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
4.1.1 Effect on ACT	at 2-mo	onth							
Martineau 2015	0.23	1.43	125	0.14	1.4	125	21.6%	0.06 [-0.18, 0.31]	+
Subtotal (95% CI)			125			125	21.6%	0.06 [-0.18, 0.31]	*
Heterogeneity. Not ap	plicable								
Test for overall effect:	Z = 0.5	50 (P =	0.62)						
4.1.2 Effect on ACT	at 6-mo	onth							
Castro 2014	0.5	4.7	201	1.1	3.99	207	35.2%	-0.14 [-0.33, 0.06]	-
Martineau 2015	0.38	1.43	125	0.44	1.43	125	21.6%	-0.04 [-0.29, 0.21]	+
Subtotal (95% CI)			326			332	56.8%	-0.10 [-0.25, 0.05]	•
Heterogeneity. $Chi^2 =$	0.35, d	f = 1 (P = 0.5	(5); 1 ² =	0%				
Test for overall effect:	: Z = 1.3	30 (P =	0.20)	30 L					
4.1.3 Effect on ACT	at 12-m	onth							
Martineau 2015	0.3	1.43	125	0.37	1.43	125	21.6%	-0.05 [-0.30, 0.20]	+
Subtotal (95% CI)			125			125	21.6%	-0.05 [-0.30, 0.20]	•
Heterogeneity. Not ap	plicable								
Test for overall effect:	Z = 0.3	89 (P =	0.70)						
Total (95% CI)			576			582	100.0%	-0.05 [-0.17, 0.06]	
Heterogeneity. Chi ² =	1.58, d	f = 3(P = 0.6	56); 1 ² =	0%				
Test for overall effect:									-4 -2 0 2 4
Test for subgroup diff					(P = 0)	0.541, 1 ²	= 0%		Favours [experimental] Favours [control]

FIGURE 7. Effect of vitamin D versus placebo on ACT. ACT = asthma control test. CI = confidence interval, SD = standard derivation.

synthesized in asthmatic patients which could not be inhibited by glucocorticoids but vitamin D could significantly reduce both.^{20,21} However, our pooled data analysis showed that treatment of vitamin D did not significantly decrease asthma exacerbations, which we supposed to be caused by the multiple factors associated with development of exacerbations, such as indoor and outdoor allergens, pollutants, specific drugs, poorly adherent to asthma controllers, as well as viral infections. Martineau et al¹⁴ conducted a trial to evaluate the effect of vitamin D on upper respiratory infections in 250 patients with asthma, but the results showed no influence (adjusted hazard ratio [HR] 0.87, 95% CI: 0.64– 1.16, P = 0.34). Moreover, asthma is considered mainly as a Th2-mediated disease, but the reported effects of vitamin D on Th2 cells were not identical.¹⁵ Kuo et al²² and Mathieu et al²³ suggested a direct signaling effect of vitamin D on naïve CD4+ T cells toward Th2 differentiation or maintenance, which was further demonstrated in murine experiment that vitamin D could shift the Th1–Th2 cytokine balance toward Th2, thus potentially increasing risk of asthma. While a recent animal model study reported that perinatal VDD in mice leaded to Th2 skewing and reduced IL-10-secreting regulatory T cells, and vitamin D supplementation was associated with reduced serum IgE levels, pulmonary eosinophilia and peri-bronchiolar collagen deposition.²⁴ On the other hand, interpretation of our result should be cautious,

	Exp	erimen	tal	C	ontrol		S	td. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
5.1.1 Effect on 25(O	H)D at 2	-mon	th						
Martineau 2015 Subtotal (95% CI)	0.48	1.43	125 125	-0.04	1.4	125 125	42.7% 42.7%	0.37 [0.12, 0.62] 0.37 [0.12, 0.62]	-
Heterogeneity. Not ap	plicable								
Test for overall effect	: Z = 2.8	87 (P =	0.004	1					
5.1.2 Effect on 25(O	H)D at 3	-mon	th						
Majak 2009	0.24	1.42		-0.31	1.43	17	5.6%	0.38 [-0.31, 1.07]	
Subtotal (95% CI)			16			17	5.6%	0.38 [-0.31, 1.07]	-
Heterogeneity. Not ap	plicable	1							
Test for overall effect	Z = 1.0	07 (P =	0.28)						
5.1.3 Effect on 25(O	H)D at 6	-mon	th						
Worth 1994	0.8	1.48	14	0.11	1.53	19	5.5%	0.45 [-0.25, 1.15]	
Subtotal (95% CI)			14			19	5.5%	0.45 [-0.25, 1.15]	-
Heterogeneity: Not ap	plicable								
Test for overall effect	: Z = 1.2	25 (P =	0.21)						
5.1.4 Effect on 25(O	H)D at 1	2-mo	nth						
Baris 2014	1.25	1.57	17	0	1.42	15	5.1%	0.81 [0.09, 1.54]	
Martineau 2015	0.84	1.48		-0.12	1.43	125	41.2%	0.66 [0.40, 0.91]	+
Subtotal (95% CI)			142			140	46.2%	0.67 [0.43, 0.91]	•
Heterogeneity. Chi ² =	0.15, 0	if = 1 (P = 0.7	70); $I^2 =$	0%				
Test for overall effect	: Z = 5.5	50 (P <	0.000	01)					
Total (95% CI)			297			301	100.0%	0.51 [0.35, 0.68]	•
Heterogeneity. Chi ² =	3.40, c	if = 4 (P = 0.4	49); l ² =	0%			N .	-4 -2 0 2 4
Test for overall effect	: Z = 6.:	16 (P <	0.000	01)					Favours [control] Favours [experimental]
Test for subgroup dif	ferences	: Chi ² =	= 3.24,	df = 3	(P = 0)	.36), I ²	= 7.5%		ravours (control) ravours (experimental)

FIGURE 8. Effect of vitamin D versus placebo on serum 25(OH)D levels. 25(OH)D = 25-hydroxyvitamin D, CI = confidence interval, SD = standard derivation.

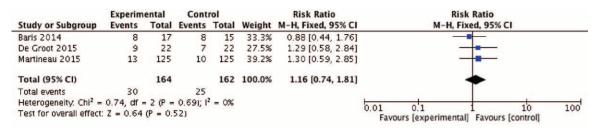


FIGURE 9. Effect of vitamin D versus placebo on adverse events. CI = confidence interval, M-H = Mantel-Haenszel.

because significant statistical heterogeneity existed in the pooled data of enrolled studies, and we did not analyze the serum ILs levels as well as upper respiratory infection rate due to the incomplete data report.

Observational studies have associated lower vitamin D status with decreased lung function, which was further demonstrated by Arshi and Thuesen, ^{5,8,25,26} but a recent cohort study did not find such an association (adjusted odds ratio [OR] 0.996, 95% CI: 0.990–1.002, P = 0.171),²⁷ thus making the conclusions of relation between vitamin D and lung function illusive. Moreover, exacerbations are characterized by decreases in expiratory airflow that can be quantified by measurement of FEV_1 , and the degree of symptoms is a sensitive measure of the onset of an exacerbation.^{28,29} Our data analysis showed that treatment of vitamin D could not improve FEV1 and ACT no matter in short or long terms, although serum 25-hydroxyvitamin D was significantly increased, which we think may result from the variable factors involving vitamin D status. As explained by Kerley's review, exposure to ultraviolet-B. not through diet, is the major source of vitamin D: therefore. vitamin D status varies markedly depending on local weather conditions and individual sun behaviors.¹⁶ Due to the variable and changeable human behavior and season, both asthma and vitamin D outcomes are often influenced. In addition, although our study showed a significant elevation in serum 25-hydroxyvitamin D after administration of vitamin D, but not all patients reached the sufficiency level of 75 to 250 nmol/L released by the Endocrine Society,³⁰ thus may offset the possible effect of vitamin D on improvement in lung function and symptoms.

Nitric oxide (NO) in respiratory tract can be synthesized from L-arginine by NO synthase (NOS) in a wide variety of cell types including inflammatory cells, epithelial cells, and airway nerves, and large amount of NO generated by inducible NOS in asthma can be detected in exhaled breath, that is, FeNO, which is regarded as a potential noninvasive method to reflect the degree of airway inflammation and diagnose asthma.31,32 In theory, vitamin D could decrease FeNO levels in asthma through the mechanism of antiinflammation by reducing ILs mentioned above; however, de Groot et al¹³ did not find such an effect of vitamin D, and what is more, they further demonstrated that vitamin D did not decrease sputum eosinophils and neutrophils. Our pooled data analysis showed that treatment with vitamin D could not significantly decrease FeNO levels in patients with asthma, which was in line with the findings reported with de Groot, but we should interpret it cautiously because of the relatively limited studies pooled in the final analysis and the absence in the comparisons of biomarkers associated with FeNO such as sputum eosinophils, ILs, TNF- α , and interferon- γ (INF- γ).

Limitations of our meta-analysis are as follows: First, the name, dose, administration routine, and duration of the intervention drugs were not identical in the enrolled studies, which may result in performance biases. Second, the baseline characteristics of the patients were not completely provided, and we included all trials conducted in adults and children, which may lead to selection biases. Third, the total number of studies and patients enrolled was relatively small. And finally, these trials relied on a single measurement of vitamin D status or intake with asthmatic outcomes assessed up to 31 years later; therefore, it is impossible to discount variables other than vitamin D contributing to the development of asthma.

CONCLUSIONS

Vitamin D supplementation in addition to asthma controllers cannot decrease asthma exacerbation and FeNO, nor improve lung function and asthma symptoms, although it can be safely applied to increase serum 25-hydroxyvitamin D levels. More large double-blind RCTs are needed, particularly in identical medication doses and administration routines and durations, to further determine the role of vitamin D, including the cost-effectiveness, in patients with asthma.

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