Does My Patient With Shoulder Pain Have a Rotator Cuff Tear?

A Predictive Model From the ROW Cohort

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Background: Rotator cuff tears are the leading cause of shoulder pain and disability. However, the diagnosis of a rotator cuff tear based on patient characteristics, symptoms, and physical examination findings remains a challenge because of a lack of data. Moreover, data on the predictive ability of a combination of these characteristics and tests are not available from a large cohort of patients. Consequently, clinicians rely on expensive imaging, such as magnetic resonance imaging (MRI), to make a diagnosis.

Purpose: To model patient characteristics, symptoms, and physical examination findings that predict a rotator cuff tear. We present a nomogram based on our predictive model that can be used in patients with shoulder pain to determine the probability of the diagnosis of a rotator cuff tear without the need for imaging.

Study Design: Cohort study (diagnosis); Level of evidence, 2.

Methods: We recruited patients from outpatient clinics who were \geq 45 years of age and who had shoulder pain of at least 4 weeks' duration. A rotator cuff tear was diagnosed based on expert clinical impression and the presence/absence of a tear on a blinded review of MRI. Ultimately, 301 patients were included in the analysis.

Results: A total of 123 patients (41%) had rotator cuff tears, and 178 patients (59%) did not. The predictors of the diagnosis of a rotator cuff tear included external rotation strength ratio of the affected versus unaffected shoulder (odds ratio [OR], 1.20 [95% CI, 1.08-1.34]), male sex (OR, 1.98 [95% CI, 1.10-3.56]), positive lift-off test result (OR, 4.33 [95% CI, 1.46-12.86]), and positive Jobe test result (OR, 9.19 [95% CI, 4.69-17.99]). A nomogram based on these predictor variables was plotted.

Conclusion: Presented is a model that can accurately predict the diagnosis of a rotator cuff tear with satisfactory discrimination and calibration based on 4 variables: sex, lift-off test, Jobe test, and external rotation strength ratio. Data from this study can be used to aid in the diagnosis of a rotator cuff tear in day-to-day clinical practice in outpatient settings without the need for expensive imaging such as MRI.

Keywords: rotator cuff tears; diagnostic accuracy; predictive modeling

In the United States, shoulder symptoms accounted for an estimated 10.7 million ambulatory care visits to physician offices in 2013.¹⁰ The prevalence of shoulder pain in the population ranges from 14% to 34%.^{2,8,46,50,55,56} Rotator cuff disorders are the underlying issue in 65% to 70% of patients with shoulder pain,^{11,57} and they accounted for 272,148 surgeries in the United States in 2006.^{12,27} Recent reports have highlighted the importance of patient symptoms and shoulder pain.⁷ However, the diagnosis of a rotator cuff tear based on patient characteristics, symptoms, and physical examination findings remains a challenge.³⁰ The lack of data on this issue is well documented, and a need for the

synthesis of such data has been expressed by leading experts.^{22,42} As such, there is a heavy reliance on expensive imaging such as magnetic resonance imaging (MRI) to make a diagnosis of a rotator cuff tear.

There are several "special" shoulder physical examinations described for the diagnosis of rotator cuff tears. Data on the diagnostic accuracy of individual special shoulder tests for rotator cuff tears are variable²⁸ and are mostly available from single-site/single-provider studies or retrospective chart reviews.^{4,25,31,38,39,47,63,64} Several of these studies assessed patient populations undergoing rotator cuff surgery,^{6,16,17,40,49} which creates a spectrum bias, thus limiting the inferences that can be made to the larger population of patients with rotator cuff tears, most of whom do not undergo surgery. To the best of our knowledge, data from patient characteristics, symptoms, and physical examinations have not been used to build a robust,

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diagnostic predictive model in a large cohort for use in clinical practice.

In a large cohort of patients with shoulder pain, we modeled patient characteristics, symptoms, and physical examination findings that predict a symptomatic rotator cuff tear. We present a nomogram based on our predictive model that can be used in patients with shoulder pain to determine the probability of the diagnosis of a rotator cuff tear without the need for imaging.

METHODS

Patient Population

We recruited a cohort of 390 patients with shoulder pain and/or limitation in range of motion lasting at least 4 weeks and who completed baseline history questionnaires. These patients were recruited from 4 sports medicine/shoulder clinics (3 academic and 1 community) between February 2011 and June 2015. Other inclusion criteria were patients aged \geq 45 years because a diagnostic dilemma of shoulder pain is most common in this age group. Patients provided informed consent, and the study was approved by our institutional review board. Exclusion criteria were a current shoulder fracture of the affected shoulder, history of index shoulder surgery, and active cervical radiculopathy. Patients were asked if they had neck pain radiating to the shoulder to elicit cervical radiculopathy. Patients missing information on the degree of clinical certainty of a rotator cuff tear (n = 4), those found to be ineligible after enrollment (n = 1), and those missing physical examination findings (n = 19) were excluded from the analysis.

History and Shoulder Pain/Function Assessment

The structured shoulder and general health questionnaire assessed patient demographics; comorbidities; symptoms; smoking habits; prior treatments, including physical therapy and corticosteroid injections; and occupational history. Patients were asked about their medication use, including acetaminophen, nonsteroidal anti-inflammatory drugs, glucosamine, and narcotics for shoulder pain in the past 2 weeks. Patients also completed the Shoulder Pain and Disability Index (SPADI),⁵¹ a standardized 13-item questionnaire. The SPADI has a pain scale (5 items) and a disability scale (8 items) that elicit information on pain and functional limitations with activities of daily living. The composite SPADI score ranges from 0 to 100, with higher scores reflecting worse pain and function.

Physical Examination Protocol

Special Tests. Our standardized physical examination protocol has been previously published and was based on the original descriptions of the special shoulder tests.^{28,29} Briefly, shoulder special tests were performed by a physician/orthopaedic physician assistant trained in our research protocol. The research protocol was developed before the recruitment of patients. The 13 shoulder special tests that were performed included the lift-off test,¹⁷ passive lift-off test,¹⁶ belly-press test,¹⁶ belly-off sign,⁵³ and bear hug test⁶ for the subscapularis; the external rotation lag sign at 0°,²³ external rotation lag sign at 90°,²³ Jobe test (empty can test),^{32,33} full can test,³⁶ Neer sign,⁴⁸ Hawkins-Kennedy sign,²⁰ and drop arm test⁶⁰ for the infraspinatus and supraspinatus; and the hornblower sign⁵⁸ for the teres minor/infraspinatus. Good interrater and intrarater reliability have been reported for the performance of many of these tests.^{9,15,34,43}

Strength Testing. Strength testing was performed using a handheld dynamometer in abduction, external rotation, and internal rotation by trained research assistants. Both the affected and contralateral shoulders were assessed, and a mean of 2 consecutive measurements that were at least 10 seconds apart was used in our analysis. Our detailed protocol for standardized strength testing has been previously described.^{29,45} Strength testing using a dynamometer has good intrarater and interrater reliability.²¹ We used a ratio of the affected shoulder strength versus the contralateral shoulder strength in our analysis.

Diagnostic Imaging

Shoulder MRI scans were read in a blinded fashion by consensus of 2 shoulder experts (L.D.H. and N.B.J. or J.E.K. and N.B.J.). Although this was a pragmatic cohort with no specific requirements for MRI, most patients underwent MRI on a 1.5-T or 3.0-T magnet with a dedicated shoulder coil. The following sequences were usually obtained: coronal oblique fast spin echo (FSE) proton density-weighted images, coronal oblique FSE short tau inversion recovery (STIR) images, sagittal FSE proton density- or T2weighted images with fat suppression, sagittal FSE T1weighted images, axial T1-weighted images, and axial T2-weighted gradient echo images. Scans were read at least 2 months after the clinical encounter. Our previous work has shown good interrater and intrarater reliability for these MRI readings as compared with those of a musculoskeletal radiologist. Kappa values ranged from 0.75 to

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0.90 for tear presence, tear size, and tear thickness.²⁶ MRI features that were assessed in a standardized manner included tear thickness, tear size in the longitudinal and transverse planes, fatty infiltration of the rotator cuff muscles, and rotator cuff muscle atrophy. Criteria for each of these measurements have been previously described²⁶ and follow standard radiology methodology.

Diagnosis of Rotator Cuff Tears (Reference Standard)

Prior literature has shown structural evidence for rotator cuff tears even in asymptomatic patients.^{44,54,61,62} Moreover, a rotator cuff tear is a clinical syndrome, and a case definition simply based on imaging findings is incomplete and clinically inaccurate. Our study was designed to account for this issue because an accurate case definition is paramount for our results to be valid. In addition to our standardized protocol for assessing the structural presence of a rotator cuff tear on MRI, a fellowship-trained shoulder or sports medicine attending physician performed an independent clinical assessment (by history and physical examination). The attending physician then rated the degree of certainty that the patient's symptoms were clinically attributable to a rotator cuff tear on a scale ranging from 0 (certainly not a tear) to 100 (certainly a tear). If the degree of certainty was marked as \geq 50 and a tear was documented on the blinded MRI review, the patient was considered to have a clinical diagnosis of a rotator cuff tear. If the degree of certainty was marked as <50, the patient was considered not to have a clinical diagnosis of a rotator cuff tear.

In our analysis, we excluded 65 patients who had a degree of certainty of \geq 50 but (1) who did not undergo MRI because it was not clinically indicated/ordered (n = 31) or (2) whose MRI did not show a structural deficit indicating a rotator cuff tear (n = 34). Thus, 301 patients were included in our final analysis for this study, termed the Rotator Cuff Outcomes Workgroup (ROW). Patients with a negative MRI finding but a degree of certainty of \geq 50 were excluded because the sensitivity of MRI in partial-thickness rotator cuff tears is 64%.¹³ Thus, it was possible that a structural deficit in the case of a partial-thickness tear was not detected by MRI, which would lead to the misclassification of such patients. This strict definition of a rotator cuff tear ensured that a patient diagnosed as having a tear in our study indeed had a tear.

Statistical Analysis

We followed the methodology described by Harrell¹⁹ for creating our predictive model. Among 115 individual and composite variables, 41 variables were selected a priori by clinical judgment as potential predictors. With 123 patients classified as having a rotator cuff tear in our cohort, our goal was to reduce our predictor set to include no more than 12 (123/10) degrees of freedom. A hierarchical cluster analysis based on the Spearman correlation of our initial 41 variables was conducted for an independent review by 2 shoulder-trained clinicians (N.B.J. and J.E.K.) to choose, among highly correlated variables, a relatively independent variable set that would have significant predictive power. Among the 9 variables agreed upon by the 2 independent reviewers, a separate hierarchical cluster analysis and redundancy analysis were conducted. The highest pairwise correlation was <0.1, and the greatest R^2 value with which each variable could be predicted from all the others was 0.184. Nonlinearity of the covariates overall was assessed, and the *P* value to reject the hypothesis of linearity was not rejected (P = .413). Consequently, our full model was based on 9 variables with 10 degrees of freedom selected without using our primary response.

We used logistic regression to model the probability of a rotator cuff tear with our predictor set. The effect size from this model was the odds ratio, which estimated the increased odds of a rotator cuff tear diagnosis for a unit increase in a given covariate. To account for bias due to missing data, we used predictive mean matching, including variables previously deleted by clinical judgment but correlated to the predictor set, to create 20 inputed data sets for analysis. For each of these 20 data sets, a predictive model was generated, resulting in 20 sets of parameter values and variability estimates. These parameter sets were pooled using Rubin's⁵² rules. Model validation and calibration were conducted using 500 bootstrap samples.

Seeking the most parsimonious predictive model (termed the final model), we used ordinary least squares to fit the full model–predicted values with all 9 variables, with backward elimination for each of the 20 imputation data sets. For each of these 20 models, variables were eliminated if they were not in the set of variables, ordered from largest to smallest contribution, with a cumulative R^2 value <0.95. Variables that presented in at least 10 models were retained in the final model for predicting a rotator cuff tear. Model validation of the subsequent 4-variable logistic regression was conducted as described above. A nomogram was also developed.

RESULTS

Of the 301 patients in our cohort who were included in this analysis, 123 (41%) were diagnosed with a rotator cuff tear, and 178 patients (59%) did not have a tear (Table 1). Patients without tears had glenohumeral osteoarthritis, adhesive capsulitis, biceps lesions, impingement syndrome (without a tear), labral tears, and acromioclavicular osteoarthritis as their diagnosis. The mean age for those with a tear was 61.6 ± 8.7 years compared with 59.2 ± 9.0 years for those without a tear. The dominant shoulder was affected in a majority of patients with (68%) and without (54%) a tear. The median SPADI score was lower in patients without a tear versus those with a tear (40 vs 50, respectively).

Our final model was derived using bootstrapped backward elimination of the full model variable fitted to the predicted values of the full model. The subsequent final model variable set was refit to the observed outcome. External rotation strength ratio, sex, lift-off test, and Jobe test constituted the final set of significant predictors of the diagnosis of a rotator cuff tear (Table 2 and Figure 1).

	$\begin{array}{l} Rotator\\ Cuff Tear\\ (n=123) \end{array}$	No Rotator Cuff Tear (n = 178)	
Sex			
Female	49 (40)	95 (53)	
Male	74 (60)	83(47)	
Age, y			
25th/50th (median)/75th percentiles	54.0/63.0/66.9	51.3/58.6/66.6	
$Mean \pm SD$	61.6 ± 8.7	59.2 ± 9.0	
Dominant shoulder affected?			
Yes	84 (68)	97 (54)	
No	35(28)	70 (39)	
Missing	4 (3)	11 (6)	
Daily shoulder use at work?			
Heavy/moderate manual work	35(28)	29 (16)	
Light/no manual labor	83 (67)	139 (78)	
Missing	5(4)	10 (6)	
Lift-off test			
Positive	18(15)	8 (4)	
Negative	76 (62)	152(85)	
Unable to perform	6 (5)	5(3)	
Missing	23 (19)	13 (7)	
Jobe test			
Positive	91 (74)	57(32)	
Negative	13 (11)	116 (65)	
Missing	19 (15)	5 (3)	
Use of medication for shoulder			
pain/corticosteroid injection			
Yes	106 (86)	125(70)	
No	17 (14)	53 (30)	
External rotation strength ratio ^b			
25th/50th (median)/75th percentiles	0.44/0.62/0.86	0.67/0.86/0.99	
$Mean \pm SD$	0.63 ± 0.30	0.85 ± 0.26	
SPADI score			
25th/50th (median)/75th percentiles	34/50/69	24/40/58	
Mean \pm SD	51 ± 22	42 ± 23	

TABLE 1

Baseline Characteristics of the ROW Cohort

^aData are shown as n (%) unless otherwise indicated. ROW, Rotator Cuff Outcomes Workgroup; SPADI, Shoulder Pain and Disability Index.

^bRatio of the affected versus contralateral shoulder.

A patient with a positive Jobe test result was 9.19~(95%CI, 4.69-17.99) times more likely to have a rotator cuff tear versus a patient with a negative test result. The lift-off test had an odds ratio of 4.33 (95% CI, 1.46-12.86) for a positive versus negative test finding. The full model with additional variables (age, SPADI score, affected shoulder, shoulder use at work, and use of medications/corticosteroid injection in addition to variables in the final model) is presented in the Appendix (Table A1 and Figures A1 and A3).

For ease of use in clinical settings, we plotted a nomogram with each of the significant variables in our final model (Figure 2A). The nomogram can be used to calculate the probability of a rotator cuff tear (Figure 2B). For instance, a male patient (20 points) with a positive lift-off

TABLE 2 Predictors of the Diagnosis of a Rotator Cuff Tear (Final Model)

	Odds Ratio (95% CI)
External rotation strength ratio (affected vs contralateral shoulder)	1.20 (1.08-1.34)
Sex (male vs female) Lift-off test (positive vs negative) Lift-off test (unable to perform vs negative) Jobe test (positive vs negative)	$\begin{array}{c} 1.98 \ (1.10\text{-}3.56) \\ 4.33 \ (1.46\text{-}12.86) \\ 2.01 \ (0.46\text{-}8.75) \\ 9.19 \ (4.69\text{-}17.99) \end{array}$

test result (40 points), positive Jobe test result (~ 60 points), and external rotation strength ratio of 1 (~ 50 points) has a cumulative score of approximately 170 points and has a probability exceeding 80% for the diagnosis of a rotator cuff tear. We generated a table with various possible clinical scenarios and the corresponding probability of a rotator cuff tear (Table A2). A similar nomogram is also available for our full model, which includes additional variables (Figure A2).

A calibration curve for our final model showed almost perfect overlap of predicted versus actual values (slope = 0.95) for the diagnosis of a rotator cuff tear (Figure 3). The corrected c-index for this model was 0.83. The van Houwelingen-Le Cessie heuristic shrinkage estimate was 0.956, suggesting that our final model would validate an independent data set only approximately 4.4% worse than our data set.

DISCUSSION

The diagnostic dilemma of a rotator cuff tear in patients with shoulder pain is recognized by experts in the recent literature.^{22,42} In a large cohort of patients with shoulder pain, our results show that external rotation strength ratio, sex, lift-off test, and Jobe test were significant predictors of the diagnosis of a rotator cuff tear. We constructed a nomogram that can be used in clinical settings to predict the diagnosis of a rotator cuff tear in patients aged >45 years presenting with shoulder pain. Thus, data from our study fill a void and the unmet need of clinicians managing patients with shoulder pain in outpatient settings.

Several studies, including data from our cohort, have reported on the sensitivity and specificity of special tests in the diagnosis of rotator cuff tears.^{4-6,25,28,38,39,47,63} The limitations of prior studies are well described in the literature³⁰ and include data from retrospective chart reviews, recruitment from a single institution/provider, and the use of either MRI or surgical findings as the "gold standard" for the diagnosis of a rotator cuff tear.^{4-6,25,38,39,47,63,64} The sole use of MRI or arthroscopic visualization for a rotator cuff tear diagnosis, without the inclusion of the clinical impression of an expert, is a major drawback. This methodology assumes that the diagnosis of a rotator cuff tear can be based only on structural findings. Asymptomatic rotator cuff tears documented on imaging are prevalent in approximately 40% of people aged >50 years, 54% of those aged >60 years, and 65% of those aged >70 years.^{44,54} Thus, the



Figure 1. Odds ratios with CIs for the probability of a rotator cuff tear (final model). The bars represent 90% (dark blue), 95% (light blue), and 99% (gray) CIs for the odds ratios.



Figure 2. Nomogram to predict the probability of a rotator cuff tear (final model). (A) The values of a predictor variable for a patient are compared via a perpendicular line with the points scale. This is repeated for all variable values for a patient, and the total sum of points is calculated. (B) The total points value is then located on the total points scale, and the estimated probability of a rotator cuff tear for that patient is found by dropping a vertical line down to the probability line. SPADI, Shoulder Pain and Disability Index.

combination of imaging/surgical deficits and clinical findings is essential in making a diagnosis of the clinical syndrome of rotator cuff tear. Our study addresses this issue, although the clinical diagnosis of a rotator cuff tear, as in our study, is subjective. Prior studies that only recruited patients undergoing surgery create a disease spectrum bias because it is likely that patients with more severe disease undergo surgery. This patient population may not represent the majority of patients with rotator cuff tears, especially in primary care settings.

Our study found that in addition to a patient's sex, the performance of simple physical maneuvers such as the



Figure 3. Calibration curve for the fitted final model. Small vertical lines at the top of the graph denote the frequency of predicted probabilities for the patient set. The apparent line depicts the relationship between predicted and observed probabilities of a rotator cuff tear. The bias-corrected line compares predicted probabilities from a bootstrap-corrected model with observed probabilities. These lines are constructed using locally weighted smoothing. The ideal line of perfect agreement originates at the origin (0,0) with a slope of 1.

lift-off test and Jobe test as well as the external rotation strength ratio can assist in the diagnosis of a rotator cuff tear. Strength measurements using a handheld dynamometer are easily performed in an office setting (although many practices including shoulder practices may not currently use them on a routine basis). External rotation predominantly measures force exerted by the infraspinatus. We used the ratio of strength of the affected versus contralateral shoulder to account for the variability in shoulder strength among patients by body weight, muscle mass, and sex.

The Jobe test³² or empty can test is performed by first assessing the deltoid with the arm in 90° of abduction and neutral rotation. The shoulder is then internally rotated and angled forward 30° so that the thumb is pointing to the floor. The test result is positive when manual muscle strength is weaker for the second maneuver as compared with the first. Prior studies have reported a sensitivity of 19% to 99%, specificity of 39% to 100%, and likelihood ratio of 0.6 to 2.7 for the Jobe test. $^{4,25,37-39,47,64}$ Jain et al²⁸ reported a sensitivity of 88% (95% CI, 80%-96%), specificity of 62% (95% CI, 53%-71%), and likelihood ratio of 2.30 (95% CI, 1.79-2.95) for the Jobe test. The Jobe test assesses the supraspinatus tendon, which functions as an abductor of the arm and is also the most commonly torn tendon of the rotator cuff.²⁴ However, there is debate whether the Jobe test assesses exclusively the supraspinatus or multiple muscles. Most of the patients in our cohort had a supraspinatus tear, as would be the case in most clinical settings. Hence, our findings of a large effect size of the Jobe test are plausible. A recent meta-analysis also reported that the

Jobe test had the highest pooled diagnostic odds ratio among physical examinations for diagnosing fullthickness rotator cuff tears.¹⁸

The lift-off test for subscapularis tears is based on the tendon's function as an internal rotator. With the clinician's assistance, the patient touches his or her lower back with the arm fully extended and internally rotated. If the test result is positive, the patient is unable to lift the dorsum of his or her hand off the back. In previous studies, the lift-off test had a sensitivity ranging from 6% to 79% and specificity ranging from 23% to 100% for the detection of subscapularis tears.^{5,25,38,39,64} A few patients in our cohort were unable to perform the lift-off test (likely because of pain/inability). Hence, we have reported positive lift-off test findings versus negative test findings and also results for those unable to perform the test versus those in whom the test results were negative. It should be noted that the lift-off test may be difficult to perform if the patient is limited in internal rotation movement due to pain.

Our study was performed in a specialty clinic setting. This study design was ideal and was chosen a priori because an accurate clinical diagnosis of a rotator cuff tear was an essential component of the "gold standard." Reliance on a diagnosis by trained shoulder/sports physicians was the most accurate way to achieve this goal. An expert clinician's impression has been used in prior studies of disease classification in which no gold standard exists, such as rheumatoid arthritis, spondyloarthropathy, fibromyalgia, spinal stenosis, and osteoarthritis.^{1,3,14,35,59} This broad clinical concept extends beyond musculoskeletal disorders and is embedded in the classification of acute coronary heart disease and congestive heart failure.⁴¹ However, clinical diagnosis is subjective and has a potential for bias.

We did not assess the role of a local anesthetic injection in the subacromial space to diagnose rotator cuff tears or impingement syndrome. Another limitation is the potential variability in the performance of physical examination maneuvers. However, good interrater and intrarater reliability have been reported for strength testing²¹ and special tests,^{9,15,34,43} and our protocols were standardized. Our study did not use ultrasound for the diagnosis of rotator cuff tears. MRI offers at least equally reliable information for the structural diagnosis of a rotator cuff tear. In some instances, especially in patients needing surgical intervention, MRI may be required, despite our algorithm showing a high probability of a rotator cuff tear. Our model also does not predict which patients may need surgery. We did not have the sample size to stratify our predictive models by tear thickness and tendon(s) involved. Few patients in our cohort had clinical suspicion of a tear with negative MRI findings or may have had a tear on MRI but were negative based on the clinical impression. This may have introduced a bias.

CONCLUSION

We present a model that can predict the diagnosis of a rotator cuff tear based on 4 variables—sex, lift-off test, Jobe test, and external rotation strength ratio—without the need for expensive imaging such as MRI. There may be situations, such as considerations for surgery, profound loss of strength, and major traumatic events leading to the current presentation, that may warrant MRI. Clinicians should use their judgment in these situations. Our model predicts the diagnosis of a rotator cuff tear accurately with satisfactory discrimination and calibration. This predictive model can be used by clinicians in the management of patients with shoulder pain when making the diagnosis of a rotator cuff tear and may possibly reduce the reliance on MRI.

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Figure A1. Odds ratios with CIs for the probability of a rotator cuff tear (full model). The bars represent 90% (dark blue), 95% (light blue), and 99% (gray) CIs for the odds ratios. SPADI, Shoulder Pain and Disability Index.

APPENDIX



Figure A2. Nomogram to predict the probability of a rotator cuff tear (full model). (A) The values of a predictor variable for a patient are compared via a perpendicular line with the points scale. This is repeated for all variable values for a patient, and the total sum of points is calculated. (B) The total points value is then located on the total points scale, and the estimated probability of a rotator cuff tear for that patient is found by dropping a vertical line down to the probability line. SPADI, Shoulder Pain and Disability Index.



Figure A3. Calibration curve for the fitted full model. Small vertical lines at the top of the graph denote the frequency of predicted probabilities for the patient set. The apparent line depicts the relationship between predicted and observed probabilities of a rotator cuff tear. The bias-corrected line compares predicted probabilities from a bootstrap-corrected model with observed probabilities. These lines are constructed using locally weighted smoothing. The ideal line of perfect agreement originates at the origin (0,0) with a slope of 1.

TABLE A1					
Predictors of the Diagnosis of a Rotator Cuff Tear (Full Model) ^{a}					

	Odds Ratio (95% CI)	
Age	1.02 (0.99-1.06)	
External rotation strength ratio (affected vs contralateral shoulder)	1.21 (1.08-1.35)	
SPADI score	1.00 (0.98-1.01)	
Sex (male vs female)	1.98 (1.09-3.61)	
Dominant shoulder affected? (yes vs no)	1.46(0.77-2.76)	
Daily shoulder use at work? (heavy/moderate manual labor vs light/no manual labor)	1.45(0.70-3.01)	
Use of medication for shoulder pain/corticosteroid injection (yes vs no)	1.76 (0.81-3.80)	
Lift-off test (positive vs negative)	3.92(1.27-12.10)	
Lift-off test (unable to perform vs negative)	1.81 (0.40-8.33)	
Jobe test (positive vs negative)	8.06 (4.06-15.99)	

^aSPADI, Shoulder Pain and Disability Index.

Frobability of a Rotator Cull Tear for Selected values of Predictive variables							
	Probability						
	5.0%	9.4%	18.4%	28.6%	32.4%	96.9%	
External rotation strength ratio ^a Male sex Positive lift-off test result Positive Jobe test result	1.1 (-) (-) (-)	1.1 (+) (-) (-)	1.1 (-) (+) (-)	0.0 (-) (-) (-)	1.1 (-) (-) (+)	0.0 (+) (+) (+)	

 TABLE A2

 Probability of a Rotator Cuff Tear for Selected Values of Predictive Variables

 a Low effect value for external rotation strength ratio set at 1.1 to elicit a probability of a rotator cuff tear at 5.0%.