

A new model to predict acute kidney injury after cardiac surgery in patients with renal insufficiency

Xijian Wang^{a*}, Naifeng Guo^{a*}, Ying Chen^b and Houyong Dai^a 

^aDepartment of Nephrology, Affiliated Hospital of Nantong University, Nantong Jiangsu, China; ^bDepartment of Epidemiology and Medical Statistics, Nantong University School of Public Health, Nantong Jiangsu, China

ABSTRACT

Objective: To establish a simple model for predicting postoperative acute kidney injury (AKI) requiring renal replacement therapy (RRT) in patients with renal insufficiency (CKD stages 3–4) who underwent cardiac surgery.

Methods: A total of 330 patients were enrolled. Among them, 226 were randomly selected for the development group and the remaining 104 for the validation group. The primary outcome was AKI requiring RRT. A nomogram was constructed based on the multivariate analysis with variables selected by the application of the least absolute shrinkage and selection operator. Meanwhile, the discrimination, calibration, and clinical power of the new model were assessed and compared with those of the Cleveland Clinic score and Simplified Renal Index (SRI) score in the validation group. Results: The rate of RRT in the development group was 10.6% ($n = 24$), while the rate in the validation group was 14.4% ($n = 15$). The new model included four variables such as postoperative creatinine, aortic cross-clamping time, emergency, and preoperative cystatin C, with a C-index of 0.851 (95% CI, 0.779–0.924). In the validation group, the areas under the receiver operating characteristic curves for the new model, SRI score, and Cleveland Clinic score were 0.813, 0.791, and 0.786, respectively. Furthermore, the new model demonstrated greater clinical net benefits compared with the Cleveland Clinic score or SRI score.

Conclusions: We developed and validated a powerful predictive model for predicting severe AKI after cardiac surgery in patients with renal insufficiency, which would be helpful to assess the risk for severe AKI requiring RRT.

ARTICLE HISTORY

Received 5 August 2021

Revised 17 April 2022

Accepted 17 April 2022

KEYWORDS



Acute kidney injury; renal replacement therapy; renal insufficiency; cardiac surgery; risk model

Introduction

With the rapid growth of the aging population and the increasing number of chronic diseases, such as hypertension, arteriosclerosis, and diabetes, more and more patients with renal insufficiency will be referred to cardiac surgery [1–4]. Whereas, the reduction of renal function reserve as a result of underlying chronic kidney disease (CKD) makes the kidney more vulnerable to this operation. The incidence of acute kidney injury (AKI) after cardiac surgery varies from 7% to 40%, and about 1% of them need renal replacement therapy (RRT) [5–7]. Obviously, the prevalence will be more serious in renal insufficiency patients, especially in patients with CKD stages 3–4 due to the lack of renal function reserve. The development of AKI requiring RRT is associated with worse postoperative outcomes and higher costs in patients undergoing cardiac surgery [8].

Therefore, given the importance to patients and cost implications of RRT, there is an urgent need to predict the occurrence of AKI, especially in patients with renal insufficiency.

Several models for cardiac surgery patients to predict severe AKI requiring RRT have been established, including the Cleveland Clinic score and the Simplified Renal Index (SRI) score [9,10]. However, there are some limitations to their adoption, including the measurement of RRT risk, time consumption and complexity of variables required, and poor discrimination among patients at the highest risk. The pathophysiology of cardiac surgery-associated AKI appears to be complex because of various risk factors, resulting in significant heterogeneity among different results [11]. Patients with renal insufficiency often have different comorbidities and risk factors for AKI than general patients [1].

CONTACT Houyong Dai  daihy520@126.com  Department of Nephrology, Affiliated Hospital of Nantong University, Nantong Jiangsu, 226001, China
†Xijian Wang and Naifeng Guo contributed equally to the paper.

© 2022 The Author(s). Published by Informa UK Limited, trading as Taylor & Francis Group.

This is an Open Access article distributed under the terms of the Creative Commons Attribution License (<http://creativecommons.org/licenses/by/4.0/>), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Ideally, as a risk prediction model for clinical application, it should have a better clinical utility and reliability, and incorporate as few variables as possible to make it rapidly accessible for the clinician. To our knowledge, there are no previous retrospective studies that develop the RRT risk model after cardiac surgery in patients with renal insufficiency. In this study, the preoperative, intraoperative, and immediate postoperative data were collected, which could be helpful to improve the model discrimination while reducing the number of variables.

The aim of the present study was to establish a simple risk model to enable bedside risk stratification with sufficient accuracy, which was focused on intensive monitoring and therapies for high-risk patients. In addition, the new model was compared with the SRI score and Cleveland Clinic score, with regard to their discrimination, calibration, and clinical usefulness for predicting RRT risk.

Materials and methods

Patients

Renal insufficiency patients (estimated glomerular filtration rate [eGFR] based on the Chronic Kidney Disease-Epidemiology Collaboration formula [12], less than 60 mL/min/1.73 m²) older than 18 years old who underwent cardiac surgery at the Affiliated Hospital of Nantong University between January 2011 and January 2019 were collected and retrospectively analyzed. The exclusion criteria were: (1) severe preoperative renal insufficiency (preoperative RRT dependence or eGFR less than 15 mL/min/1.73 m²), (2) death during or within

24 h after surgery, (3) incomplete clinical data. For some patients who had more than one cardiac surgery procedure performed during the study period, only the first surgical episode was considered. A total of 393 patients were registered. According to the application of the inclusion and exclusion criteria, 63 patients were excluded. The remaining 330 patients were enrolled in the current study. These patients were randomly assigned to the development group and validation group at a ratio of 7:3. A detailed flow chart of patient selection was shown in Figure 1.

This study was approved by the Ethics Committee of the Affiliated Hospital of Nantong University (approval number: 2020-K096-01) and conducted according to the Declaration of Helsinki. This study was registered in the Chinese Clinical Trial Registry (ChiCTR2100043161).

Outcomes

The primary endpoint was AKI which required dialysis during the postoperative period. The indications for dialysis based on the clinical judgment by senior nephrologists, included uremia, volume overload, and biochemical abnormalities.

Data collection

Attempts to improve clinical outcomes of patients with AKI have centered on early diagnosis and customized treatment. Considering that arriving in the intensive care unit (ICU) after surgery is a better time point for intervention, the risk factors of three periods were included in our study: pre-operation, intra-operation,

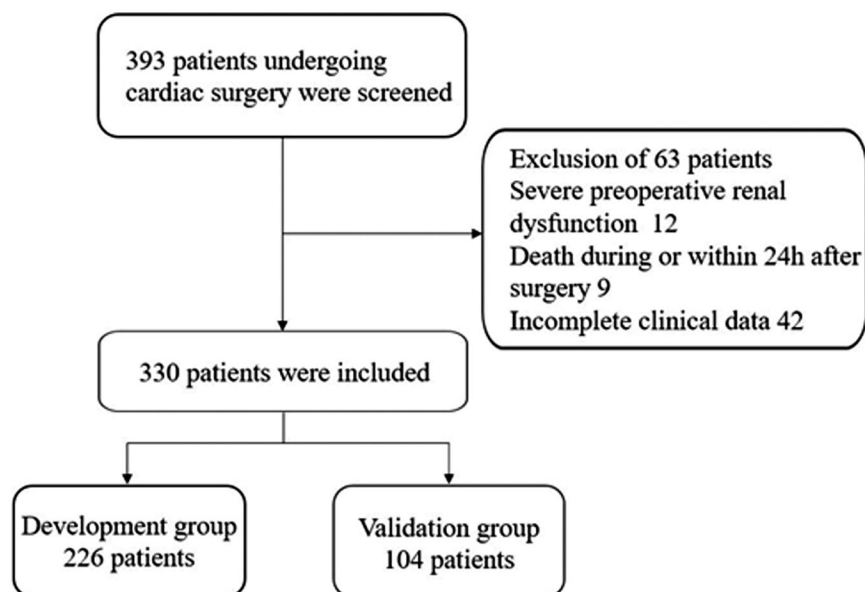


Figure 1. Flowchart of participant selection.

and post-operation immediately (up to 4 h from the ICU admission time). The clinical variables from the electronic medical records of patients were extracted. Demographic characteristics included gender, age, and Body Mass Index (BMI). Preoperative data: diabetes, hypertension, cerebrovascular disease, recent contrast agent exposure (within 7 days before surgery), recent myocardial infarction (occurred within 1 month before surgery), previous cardiac surgery, intra-aortic balloon pump (IABP), chronic obstructive pulmonary disease, baseline eGFR, emergency surgery, laboratory parameters at admission (hemoglobin, leucocyte, platelet, albumin, alanine aminotransferase, aspartate aminotransferase, blood electrolyte levels, serum creatinine, serum uric acid, serum cystatin C, blood urea nitrogen, B-type natriuretic peptide, high sensitivity C-reactive protein), New York Heart Association (NYHA) functional class was assessed and echocardiography was performed before cardiac surgery, left ventricular ejection fraction, Interventricular septum thickness (IVST) and left ventricular end diastolic diameter (LVEDD) were measured as recommended, the usages of medications (nonsteroidal anti-inflammatory drugs [NSAID], vasoactive drugs, antibiotics, angiotensin-converting enzyme inhibitors or angiotensin receptor inhibitors [ACEI/ARB], diuretics) before surgery were recorded. Intraoperative data: operative procedure, aortic cross-clamping time, cardiopulmonary bypass time, IABP, erythrocyte transfusion. Immediate postoperative data: serum creatinine, serum uric acid, serum Cystatin C, blood urea nitrogen, eGFR, medications (NSAID, vasoactive drug, antibiotics, ACEI/ARB, diuretics), IABP.

Statistical analysis

The Kolmogorov-Smirnov test was used to perform the normality test for continuous variables, and a comparison of the differences between the two groups was performed by using the *t* test, chi-square test, or Mann-Whitney *U* test. The results were presented as mean \pm standard deviation ($\bar{x} \pm SD$), frequencies (percentages) or median (interquartile range). The least absolute shrinkage and selection operator (LASSO) regression analysis was used for the best predictors of RRT after cardiac surgery. This approach was able to avoid issues of multicollinearity and overfitting, even with a high number of potential predictors and a small sample size. Five-fold cross-validation and the 1-SE rule were performed to control overfitting [13]. All predictors were included in the logistic regression, and a new multivariable regression model was developed. A nomogram was constructed based on the multivariate analysis; the

weighted point was calculated by the beta coefficient of each variable in the model. The variable with the highest beta coefficient was scored on a 100 points scale, and the remaining variables were scored according to their individual weighted effect. Finally, the total number of points was calculated [14]. The nomogram was subjected to 1,000 bootstraps resamples for internal validation of the development set. At the same time, the performance of the new model was also analyzed in the verification set and compared with that of the SRI score [10] and Cleveland Clinic score [9]. The area under the receiver operating characteristic curve (AUC) was used to assess discrimination and compared by the DeLong method [15]. A calibration curve was plotted to evaluate the calibration and was accompanied by the Hosmer–Lemeshow test. Finally, a decision curve was carried out to calculate the net benefits of individuals in different threshold probabilities and confirmed the clinical efficacy of the predictive model [16]. All data analysis was performed by SPSS 22.0 software (IBM Corporation, 2013, USA) and R software version 4.0.2 (<http://www.r-project.org>). The difference was considered to be statistically significant at $p < 0.05$.

Results

Characteristics of patients undergoing cardiac surgery in development and validation cohorts

The demographic and baseline clinical data of 330 patients were listed in Table 1; 226 and 104 patients were divided into the development and validation groups, respectively. The rate of RRT in the derivation group was 10.6% ($n = 24$), while the rate in the validation group was 14.4% ($n = 15$). There was no significant difference in baseline characteristics and intraoperative variables between the development and validation groups.

Potential variables used in the model for predicting the risk of RRT

The results of the LASSO regression analysis were presented in Figure 2. A total of 55 potential variables were used in LASSO regression analyses, which were collected from preoperative, intraoperative, and immediate postoperative data. Among of the 55 variables, eight predictors such as postoperative creatinine, aortic cross-clamping time, emergency, preoperative cystatin C, high sensitivity C-reactive protein, leucocyte, preoperative eGFR, and erythrocyte transfusion were selected to construct a new model for predicting the risk of RRT through logistic regression. In multivariate

Table 1. Baseline characteristics of development and validation group.

Variables	Development group (n = 226)	Validation group (n = 104)	P value
Demographics			
Age (years)	63 ± 10	64 ± 10	0.609
Gender (male)*	133 (58.8)	57 (54.8)	0.490
Body mass index (kg/m ²)	24.01 ± 3.16	24.41 ± 3.68	0.310
Preoperative			
Hypertension*	110 (48.7)	46 (44.2)	0.453
Diabetes mellitus*	28 (12.4)	18 (17.3)	0.231
Contrast media exposure*	100 (44.2)	39 (37.5)	0.249
Previous cardiac surgery*	7 (3.1)	4 (3.8)	0.725
Recent myocardial infarction*	51 (22.6)	18 (17.3)	0.275
Cerebrovascular disease*	13 (5.8)	6 (5.8)	0.995
Chronic obstructive pulmonary disease*	3 (1.3)	1 (1.0)	0.778
IABP*	14 (6.2)	4 (3.8)	0.383
NYHA classification III or IV*	110 (48.7)	48 (46.2)	0.670
IVST (mm)	10.00 (9.00, 11.00)	10.00 (9.00, 11.00)	0.060
LVEDD (mm)	49.93 ± 8.78	49.66 ± 8.07	0.790
LVEF (%)	59 (55, 64)	58 (53, 64)	0.379
Drugs use*			
Antibiotic	71 (31.4)	44 (42.3)	0.054
ACEI/ARB	74 (32.7)	34 (32.7)	0.993
NSAID	5 (2.2)	0 (0)	0.126
Diuretic	170 (75.2)	79 (76.0)	0.885
Vasoactive drug	8 (3.5)	4 (3.8)	0.890
Hemoglobin (g/L)	127.89 ± 16.70	130.66 ± 19.46	0.211
Leucocyte (× 10 ⁹ /L)	5.91 ± 1.62	6.08 ± 1.95	0.424
Platelet (× 10 ⁹ /L)	186.78 ± 56.46	198.41 ± 66.23	0.101
Albumin (g/L)	38.38 ± 3.66	38.42 ± 3.45	0.926
Alanine aminotransferase (U/L)	23.50 (15.00, 39.00)	26.00 (17.00, 35.75)	0.587
Aspartate aminotransferase (U/L)	25.00 (21.00, 33.25)	25.00 (21.00, 29.75)	0.561
Natremia (mmol/L)	141.00 (140.00, 143.00)	141.00 (139.00, 143.00)	0.681
Potassium (mmol/L)	4.00 (3.80, 4.30)	4.00 (3.80, 4.20)	0.291
Magnesium (mmol/L)	0.94 (0.89, 1.00)	0.95 (0.88, 1.00)	0.995
Calcium (mmol/L)	2.23 ± 0.11	2.23 ± 0.11	0.777
Phosphorus (mmol/L)	1.18 ± 0.20	1.22 ± 0.20	0.055
Creatinine (μmol/L)	116.34 ± 19.97	117.31 ± 17.70	0.672
Uric acid (μmol/L)	346.50 (287.50, 414.25)	365.50 (306.25, 444.00)	0.213
Cystatin C (μg/L)	1117.30 ± 321.61	1142.00 ± 302.54	0.509
Blood urea nitrogen (mmol/L)	8.61 ± 2.49	8.31 ± 1.99	0.287
eGFR (ml/min/1.73 m ²)	52.23 ± 7.42	50.45 ± 8.15	0.051
B-type natriuretic peptide (pg/mL)	200.70 (83.93, 530.63)	207.35 (83.53, 477.20)	0.828
High sensitivity C-reactive protein (mg/L)	3.70 (2.19, 5.65)	3.44 (2.24, 6.65)	0.733
Emergency*	18 (8.0)	9 (8.7)	0.832
Intraoperative			
Operative procedure*			
CABG	64 (28.3)	27 (26.0)	0.432
Valve	153 (67.7)	73 (70.2)	
Valve and CABG	5 (2.2)	4 (3.8)	
Others [#]	4 (1.8)	0 (0)	
Aortic cross-clamping time (min)	67 (0, 79)	67 (52, 85)	0.385
CPB time (min)	106 (0, 122)	107 (88, 122)	0.417
IABP*	17 (7.5)	7 (6.7)	0.797
Erythrocyte transfusion (U)	3 (3, 3)	3 (3, 3)	0.172
Immediate postoperative^{&}			
Creatinine (μmol/L)	153.87 ± 37.83	158.22 ± 40.50	0.343
Uric acid (μmol/L)	371.50 (310.50, 451.25)	392.00 (318.00, 478.50)	0.239
Cystatin C (μg/L)	1304.20 ± 313.04	1349.40 ± 291.42	0.213
Blood urea nitrogen (mmol/L)	11.24 ± 3.51	11.28 ± 3.72	0.914
eGFR (ml/min/1.73 m ²)	38.67 ± 9.67	36.86 ± 10.02	0.118
Drugs use*			
NSAID	22 (9.7)	3 (2.9)	0.029
Vasoactive drug	221 (97.8)	104 (100)	0.126
Antibiotic	224 (99.1)	104 (100)	0.336
ACEI/ARB	82 (36.3)	40 (38.5)	0.703
Diuretic	225 (99.6)	102 (98.1)	0.188
IABP*	19 (8.4)	11 (10.6)	0.524
Acute kidney injury*	154 (68.1)	77 (74.0)	0.277
Renal replacement therapy*	24 (10.6)	15 (14.4)	0.320
Inhospital mortality*	4 (1.8)	2 (1.9)	0.923

IABP, intra-aortic balloon pump; NYHA, New York Heart Association; IVST, interventricular septum thickness; LVEDD, left ventricular end-diastolic dimension; LVEF, left ventricular ejection fraction; ACEI/ARB, angiotensin-converting enzyme inhibitor or angiotensin receptor inhibitors; NSAID, nonsteroidal anti-inflammatory drugs; eGFR, estimated glomerular filtration rate; CABG, coronary artery bypass grafting; CPB, cardiopulmonary bypass.

*Categorical variables were expressed in frequency (percentage).

[#]Other surgery includes replacement of the ascending aorta and surgery of atrial septal defects.

[&]Up to 4 h from the ICU admission time.

The difference was considered to be statistically significant at $p < 0.05$.

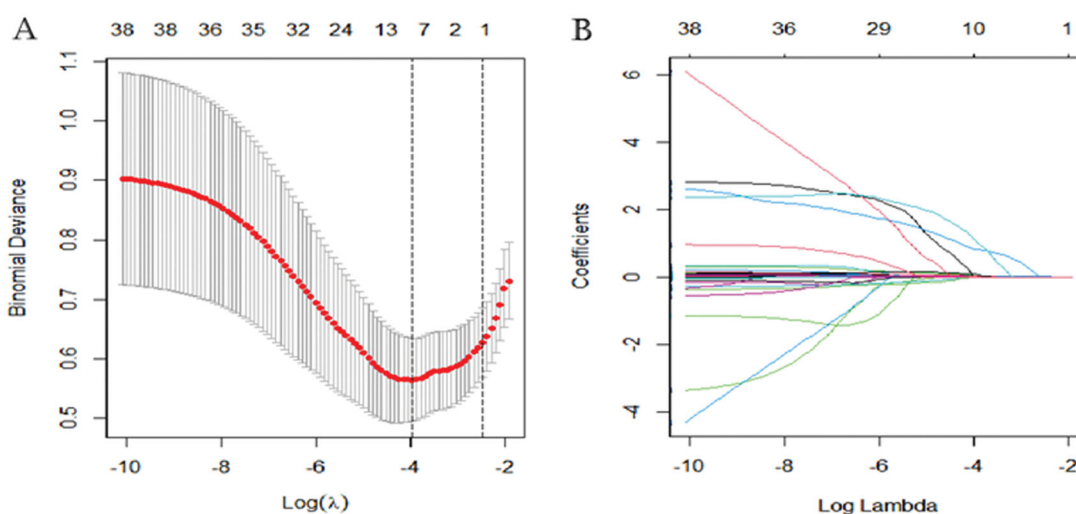


Figure 2. Least absolute shrinkage and selection operator (LASSO) binary logistic regression analysis in prediction of RRT. (A) The optimal parameter (λ) of Lasso is selected by the minimum criterion for five times cross-validation. The dotted vertical lines were plotted at the optimal values using the minimum criteria and the one standard error of the minimum criteria (the 1 – SE criteria). Finally, the λ value of 0.0191 was selected. (B) The distribution of the lasso coefficient of fifty-five variables. A coefficient profile plot was produced against the log (λ) sequence. Predictors were selected based on the minimum criteria, where the best λ produced fourteen predictors with non-zero coefficients.

Table 2. Multivariate logistic regression analyses of variables selected with LASSO.

Variables	OR	95%CI	P
Postoperative creatinine level up to 4 h from ICU admission ($\mu\text{mol/L}$)	1.026	1.013–1.038	<0.001
Aortic cross-clamping time (min)	1.026	1.008–1.044	0.004
Emergency (yes/no)	14.779	3.206–68.136	0.001
Preoperative cystatin C ($\mu\text{g/L}$)	1.001	1.000–1.003	0.049
HsCRP (mg/L)	1.094	0.979–1.223	0.113
Leucocyte ($\times 10^9/\text{L}$)	1.218	0.964–1.537	0.098
Preoperative eGFR ($\text{ml/min}/1.73 \text{ m}^2$)	0.984	0.928–1.043	0.588
Erythrocyte transfusion (U)	1.318	0.738–2.354	0.352

HsCRP, High sensitivity C-reactive protein, eGFR, estimated glomerular filtration rate; OR, odds ratio; CI, confidence interval.

analysis, postoperative creatinine ($p < 0.001$; odds ratio [OR]: 1.026), aortic cross-clamping time ($p = 0.004$; OR: 1.026), emergency ($p = 0.001$; OR: 14.779), and preoperative cystatin C ($p = 0.049$; OR: 1.001) were considered to be independent risk factors for RRT (Table 2).

A nomogram model for prediction of RRT risk

A nomogram was also drawn according to the logistic regression results (Figure 3). By creating an intuitive graph of a statistical predictive model, a nomogram showed the numerical probability of each clinical event. For example, the application of this model to renal insufficiency patients undergoing cardiac surgery would show the following results: postoperative creatinine $271 \mu\text{mol/L}$, aortic cross-clamping time 106 min, non-emergency surgery, preoperative cystatin C $800 \mu\text{g/L}$, the total score of 100, and the risk of RRT was about 68%. The calibration curves of the new model showed a good agreement (Figure 4), the C-index of the new model was 0.851 (95% CI, 0.779–0.924) in the validation

group. Compared with the SRI score and Cleveland Clinic score in the validation group, the AUC for our model, SRI score and Cleveland Clinic score were 0.813, 0.791, and 0.786, respectively. Our new model versus the SRI score or Cleveland Clinic score demonstrated a non-significant difference ($p = 0.809$ and 0.746 , respectively, Figure 5).

Decision curve analysis for the nomogram in the validation group

The decision curve analysis displayed that the net benefit of the prediction model was higher within almost all ranges of prediction thresholds, this new model obtained greater clinical net benefits in comparison with the Cleveland Clinic score or SRI score (Figure 6).

Discussion

AKI is one of the most common serious complications after cardiac surgery [17]. The pathogenesis of

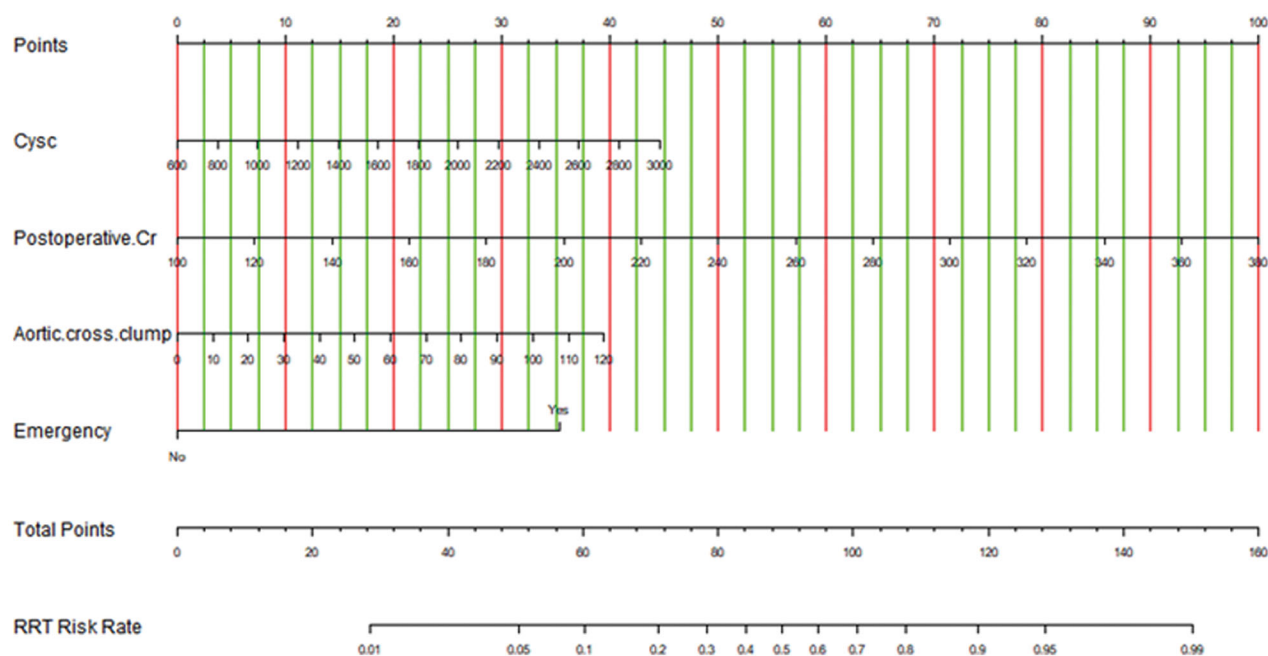


Figure 3. Prediction of RRT in renal insufficiency patients after cardiac surgery by nomogram model. In order to get every factor's position on the corresponding axis, lines were drawn on the point axis to represent the number of points. Added all points, find the position of the total score to determine the RRT probability of that line in the nomogram. Cys C, preoperative cystatin C ($\mu\text{g/L}$); Cr, creatinine ($\mu\text{mol/L}$); RRT, renal replacement therapy.

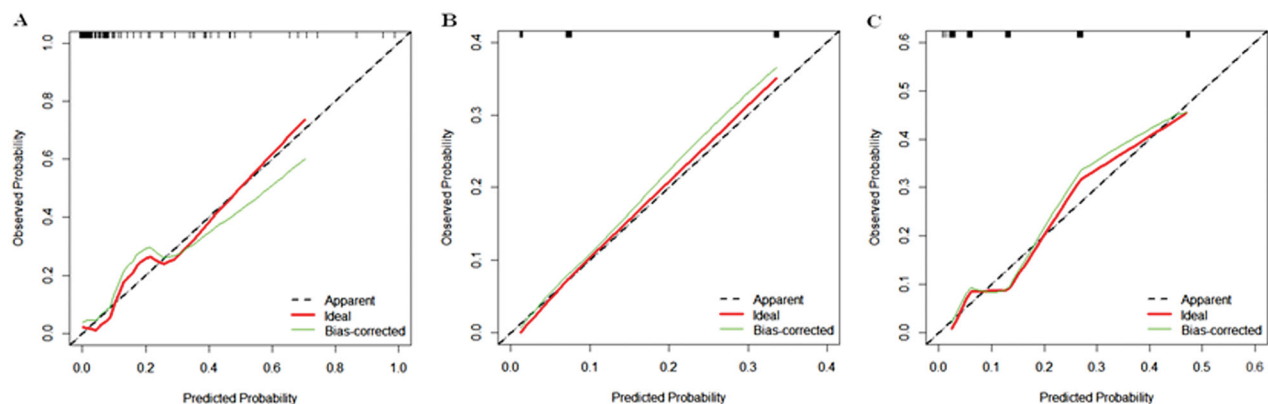


Figure 4. Calibration curves in the validation group for the new model (A), SRI score (B), and Cleveland score (C), respectively. The predicted RRT was plotted on the X-axis, and the actual RRT occurrence was plotted on the Y-axis. A plot along the 45° line would indicate a perfect calibration model in which the predicted RRT is identical to the actual RRT. The dotted line has a close fit to the solid line, which indicated better predictive accuracy of the model.

postoperative AKI is multifactorial and there are various potential causative mechanisms, such as inflammatory reactions, ischemic-reperfusion injury, hemolysis, nephrotoxic agents, and oxidative stress [18]. Patients with renal dysfunction before surgery are more likely to experience AKI and are associated with increased mortality, morbidity, and medical costs [19,20]. In recent years, researchers have been committed to the early diagnosis and risk stratification of AKI after cardiac surgery, and thereby providing patients with appropriate individualized prevention and treatment strategies. Unlike previous studies, our study is dedicated to

developing and validating a novel predictive tool to identify the risk of RRT in renal insufficiency patients after cardiac surgery. The reason for choosing RRT as the endpoint was based on clinical relevance, and more importantly, considering that AKI requiring hemodialysis after cardiac surgery is probably associated with worse postoperative outcomes [21]. In a study of 13,847 patients, the 1-year survival rate was reported to be only 10% among patients requiring hemodialysis in the postoperative period [22].

Nomograms have been widely accepted as reliable tools to quantify AKI risk by incorporating and

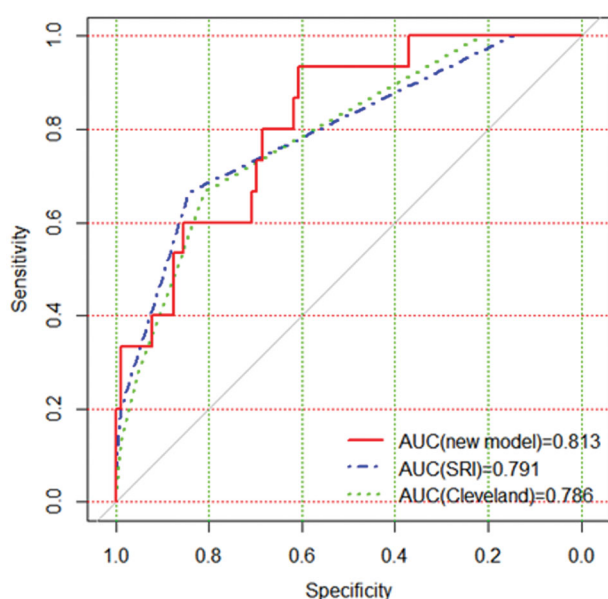


Figure 5. The AUC for models in the validation group. Comparison of AUC among models for RRT in renal inadequacy patients after cardiac surgery. New model AUC: 0.813; SRI score AUC: 0.791; Cleveland Clinic score AUC: 0.786. The new model versus SRI score, $P=0.809$; new model versus Cleveland score, $P=0.746$.

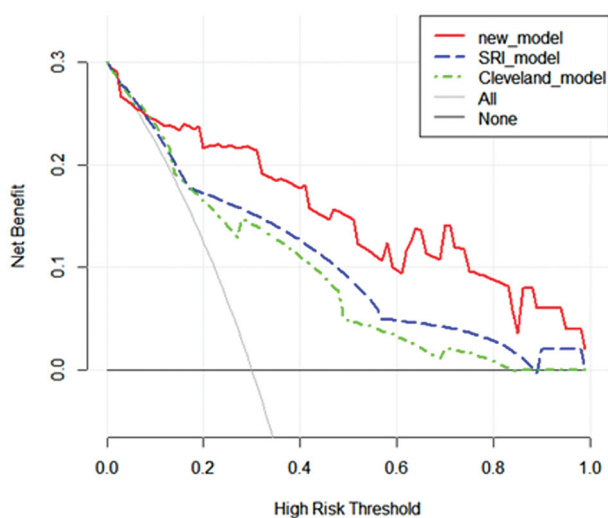


Figure 6. Decision curve analyses for prediction models. The x-axis shows the threshold probability. The y-axis shows the net benefit. The black solid lines hypothesized that all patients were RRT positive or negative, respectively. Across the range of decision thresholds, the new model was positive and had a larger net benefit than the SRI and Cleveland scores.

illustrating important factors for AKI [23–25]. In our study, we built a simple and clear nomogram model, the RRT risks were calculated by adding each point according to different parameters, which was more simple and more practical for clinicians compared to the Cleveland Clinic score and SRI score. To date, several clinical risk models for predicting severe AKI after

cardiac surgery have been developed, including the SRI score and Cleveland Clinic score, which have been widely validated in patients undergoing cardiac surgery and have been shown to have a better clinical value. However, so far as we know, most existing models have focused on patients undergoing cardiac surgery with normal preoperative renal function. In the general population, the incidence of RRT after cardiac surgery is low and occurs in the later stage in clinical practice, which limits the application of these models. Furthermore, with the development of medical technology, cardiac surgery is booming in patients with renal insufficiency in developing countries, such as China. Lastly, patients with renal insufficiency often have different comorbidities and risk factors for AKI than general patients. Therefore, there is a specific need for an appropriate model to guide clinical management, which is capable of predicting AKI in patients with renal insufficiency undergoing cardiac surgery. As shown in Figure 3, the four top risk factors were included in the new model, including postoperative creatinine, aortic cross-clamping time, emergency, and preoperative cystatin C. Nomograms have been widely used to quantify the combined contribution of several risk factors and provide a predicted probability of the event of interest. Their prediction applies to an individual instead of situating this individual within a risk group [26–29].

One of the risk factors for RRT is preoperative baseline renal function. The results of our study showed that a preoperative increase in cystatin C was associated with a higher risk of RRT after cardiac surgery. Serum cystatin C has recently been proposed as an alternative marker to serum creatinine for estimating renal clearance, due to its serum levels unaffected by variables other than kidney function [30]. Some equations using serum cystatin C levels to estimate the GFR with or without anthropometric data have recently been reported [31–33]. In addition, Yuan et al. and Saydam et al. reported that serum cystatin C is a reliable biomarker in the early detection and follow-up of AKI after cardiac surgery in a cohort of children and adults, respectively [34,35]. Recently, another study also indicated that a predictive nomogram incorporating cystatin C was helpful to evaluate the possibilities of AKI in patients with traumatic brain injury [36].

Although a remarkable increase in serum creatinine concentration needs a longer time to be detected, it is still a robust variable for predicting the outcome in patients after cardiac surgery [37]. Serum creatinine is conversely and exemplarily associated with poor renal perfusion [38]. Furthermore, it is indicated that a little increase in serum creatinine without evolution to an

acutely uremic condition could serve as a predictor of poor prognosis [37]. It appears that microcirculation disturbances leading to local ischemia and reperfusion phenomena determine very early with the initiation of a systemic inflammatory response. For renal insufficiency patients, the early rise in serum creatinine concentration after cardiac surgery may be more significant, at least in our research.

The cardiopulmonary bypass (CPB) time and emergency have been identified as risk factors for AKI after cardiac surgery in previous studies. The CPB provokes a complex systemic inflammatory response, mainly triggered by contact activation of blood by artificial surfaces [39]. Furthermore, the extracorporeal circuit provides nonpulsatile blood flow that may dysregulate the balance between cortical and medullary perfusion in the patient's kidney [40]. Again, aortic cross-clamping further exacerbates ischemia and induces inflammation [41]. So, it is reasonable that aortic cross-clamping time is a variable factor that contributes to the development of AKI. CPB time was not included in the model as it was highly correlated with aortic cross-clamping time [42].

Validation of the nomogram is essential to avoid overfitting the model and determine generalizability [43]. In our study, calibration plots showed optimal agreement between prediction and actual observation, which guaranteed the repeatability and reliability of the established nomogram. Furthermore, compared with that the SRI score and Cleveland Clinic score, the AUC of the new model was the best, and the decision curve analysis displayed that it obtained greater clinical net benefits. The results demonstrated that the new model was simple and clinically useful for predicting RRT immediately postoperatively. The model could be used for clinical risk stratification and to allow for risk stratification for future clinical trials.

The study has several limitations. First, samples were recruited from a single-center, which perhaps introduced bias to the result. In the future, we will apply the model to patients in other centers, hoping to prove the clinical utility of our proposed method in more patients. Second, given its retrospective design, information on some important clinical risk factors of postoperative AKI was not prospectively collected. Therefore, the magnitude of effects of certain risk factors may have been under- or overestimated. Third, the incorporation of novel biomarkers into prediction algorithms may provide additional opportunities to identify patients at high risks, such as heart-type fatty acid-binding protein, midkine, and soluble tumor necrosis factor 1 or 2 [44]. New tools based on novel biomarkers and prospective

clinical risk factors need to be developed in future studies with a larger multi-center sample to assist clinicians in identifying cardiac surgery patients at risk of AKI and guide patient management.

Taken together, we developed and validated a powerful predictive model for predicting severe AKI after cardiac surgery in renal insufficiency patients. Furthermore, the new model exhibited greater clinical net benefits than the SRI score and Cleveland Clinic score. It will be used in the perioperative patient management and decision-making about targeted therapies to improve prognosis.

Disclosure statement

The authors report no conflicts of interest. The authors alone are responsible for the content and writing of the paper.

Funding

This work was supported by the Nantong Science and Technology Project under Grant number JCZ19045.

ORCID

Houyong Dai  <http://orcid.org/0000-0001-7736-7108>

References

- [1] Balkan B, Magin H. What are the factors affecting the progression of kidney failure, mortality and morbidity after cardiac surgery in patients with chronic kidney disease? *J Card Surg.* 2021;36(6):2021–2028.
- [2] Nicolini F, Agostinelli A, Vezzani A, et al. The evolution of cardiovascular surgery in elderly patient: a review of current options and outcomes. *Biomed Res Int.* 2014;2014:736298.
- [3] Parikh DS, Swaminathan M, Archer LE, et al. Perioperative outcomes among patients with end-stage renal disease following coronary artery bypass surgery in the USA. *Nephrol Dial Transplant.* 2010; 25(7):2275–2283.
- [4] Wiedemann D, Bernhard D, Laufer G, et al. The elderly patient and cardiac surgery – a mini-review. *Gerontology.* 2010;56(3):241–249.
- [5] Kurth MJ, McBride WT, McLean G, et al. Acute kidney injury risk in orthopaedic trauma patients pre and post surgery using a biomarker algorithm and clinical risk score. *Sci Rep.* 2020;10(1):20005.
- [6] McBride WT, Prasad PS, Armstrong M, et al. Cytokine phenotype, genotype, and renal outcomes at cardiac surgery. *Cytokine.* 2013;61(1):275–284.
- [7] Ortega-Loubon C, Fernández-Molina M, Carrascal-Hinojal Y, et al. Cardiac surgery-associated acute kidney injury. *Ann Card Anaesth.* 2016;19(4):687–698.
- [8] Leacche M, Rawn JD, Mihaljevic T, et al. Outcomes in patients with normal serum creatinine and with artificial renal support for acute renal failure developing

- after coronary artery bypass grafting. *Am J Cardiol.* 2004;93(3):353–356.
- [9] Thakar CV, Arrigain S, Worley S, et al. A clinical score to predict acute renal failure after cardiac surgery. *J Am Soc Nephrol.* 2005;16(1):162–168.
- [10] Wijeyesundera DN, Karkouti K, Dupuis JY, et al. Derivation and validation of a simplified predictive index for renal replacement therapy after cardiac surgery. *JAMA.* 2007;297(16):1801–1809.
- [11] Husain-Syed F, Quattrone MG, Ferrari F, et al. Clinical and operative determinants of acute kidney injury after cardiac surgery. *Cardiorenal Med.* 2020;10(5):340–352.
- [12] Levey AS, Stevens LA, Schmid CH, et al. A new equation to estimate glomerular filtration rate. *Ann Intern Med.* 2009;150(9):604–612.
- [13] Tibshirani R. The lasso method for variable selection in the cox model. *Statist. Med.* 1997;16(4):385–395.
- [14] Xie HY, Zhu Y, Yao XD, et al. Development of a nomogram to predict non-organ-confined bladder urothelial cancer before radical cystectomy. *Int Urol Nephrol.* 2012;44(6):1711–1719.
- [15] DeLong ER, DeLong DM, Clarke-Pearson DL. Comparing the areas under two or more correlated receiver operating characteristic curves: a nonparametric approach. *Biometrics.* 1988;44(3):837–845.
- [16] Vickers AJ, Cronin AM, Elkin EB, et al. Extensions to decision curve analysis, a novel method for evaluating diagnostic tests, prediction models and molecular markers. *BMC Med Inf Decis Making.* 2008;8:53.
- [17] Kellum JA, Prowle JR. Paradigms of acute kidney injury in the intensive care setting. *Nat Rev Nephrol.* 2018;14(4):217–230.
- [18] Bruetto RG, Rodrigues FB, Torres US, et al. Renal function at hospital admission and mortality due to acute kidney injury after myocardial infarction. *PLOS One.* 2012;7(4):e35496.
- [19] Howell NJ, Keogh BE, Bonser RS, et al. Mild renal dysfunction predicts in-hospital mortality and post-discharge survival following cardiac surgery. *Eur J Cardiothorac Surg.* 2008;34(2):390–395; discussion 5.
- [20] Ibáñez J, Riera M, Saez de Ibarra JI, et al. Effect of pre-operative mild renal dysfunction on mortality and morbidity following valve cardiac surgery. *Interact Cardiovasc Thorac Surg.* 2007;6(6):748–752.
- [21] Lysak N, Bihorac A, Hobson C. Mortality and cost of acute and chronic kidney disease after cardiac surgery. *Curr Opin Anaesthesiol.* 2017;30(1):113–117.
- [22] Koyner JL, Garg AX, Coca SG, et al. Biomarkers predict progression of acute kidney injury after cardiac surgery. *JASN.* 2012;23(5):905–914.
- [23] Hu P, Chen Y, Wu Y, et al. Development and validation of a model for predicting acute kidney injury after cardiac surgery in patients of advanced age. *J Card Surg.* 2021;36(3):806–814.
- [24] Lin H, Hou J, Tang H, et al. A novel nomogram to predict perioperative acute kidney injury following isolated coronary artery bypass grafting surgery with impaired left ventricular ejection fraction. *BMC Cardiovasc Disord.* 2020;20(1):517.
- [25] Hu P, Mo Z, Chen Y, et al. Derivation and validation of a model to predict acute kidney injury following cardiac surgery in patients with normal renal function. *Ren Fail.* 2021;43(1):1205–1213.
- [26] Djavan B, Remzi M, Zlotta A, et al. Novel artificial neural network for early detection of prostate cancer. *J Clin Oncol.* 2002;20(4):921–929.
- [27] Eastham JA, May R, Robertson JL, et al. Development of a nomogram that predicts the probability of a positive prostate biopsy in men with an abnormal digital rectal examination and a prostate-specific antigen between 0 and 4 ng/mL. *Urology.* 1999;54(4):709–713.
- [28] Garzotto M, Hudson RG, Peters L, et al. Predictive modeling for the presence of prostate carcinoma using clinical, laboratory, and ultrasound parameters in patients with prostate specific antigen levels $<=$ 10 ng/mL. *Cancer.* 2003;98(7):1417–1422.
- [29] Lopez-Corona E, Ohori M, Scardino PT, et al. A nomogram for predicting a positive repeat prostate biopsy in patients with a previous negative biopsy session. *J Urol.* 2003;170(4 Part 1):1184–1188. discussion 8.
- [30] Tangri N, Stevens LA, Schmid CH, et al. Changes in dietary protein intake has no effect on serum cystatin C levels independent of the glomerular filtration rate. *Kidney Int.* 2011;79(4):471–477.
- [31] Grubb A, Björk J, Lindström V, et al. A cystatin C-based formula without anthropometric variables estimates glomerular filtration rate better than creatinine clearance using the Cockcroft-Gault formula. *Scand J Clin Lab Invest.* 2005;65(2):153–162.
- [32] Larsson A, Malm J, Grubb A, et al. Calculation of glomerular filtration rate expressed in mL/min from plasma cystatin C values in mg/L. *Scand J Clin Lab Invest.* 2004;64(1):25–30.
- [33] Sjöström P, Tidman M, Jones I. Determination of the production rate and non-renal clearance of cystatin C and estimation of the glomerular filtration rate from the serum concentration of cystatin C in humans. *Scand J Clin Lab Invest.* 2005;65(2):111–124.
- [34] Saydam O, Türkmen E, Portakal O, et al. Emerging biomarker for predicting acute kidney injury after cardiac surgery: cystatin C. *Turk J Med Sci.* 2018;48(6):1096–1103.
- [35] Yuan SM. Acute kidney injury after pediatric cardiac surgery. *Pediatr Neonatol.* 2019;60(1):3–11.
- [36] Wang RR, He M, Gui X, et al. A nomogram based on serum cystatin C for predicting acute kidney injury in patients with traumatic brain injury. *Ren Fail.* 2021;43(1):206–215.
- [37] Lassnigg A, Schmid ER, Hiesmayr M, et al. Impact of minimal increases in serum creatinine on outcome in patients after cardiothoracic surgery: do we have to revise current definitions of acute renal failure?. *Crit Care Med.* 2008;36(4):1129–1137.
- [38] Chertow GM, Burdick E, Honour M, et al. Acute kidney injury, mortality, length of stay, and costs in hospitalized patients. *J Am Soc Nephrol.* 2005;16(11):3365–3370.
- [39] Day JR, Taylor KM. The systemic inflammatory response syndrome and cardiopulmonary bypass. *Int J Surg.* 2005;3(2):129–140.

- [40] O'Neal JB, Shaw AD, Billings FT. Acute kidney injury following cardiac surgery: current understanding and future directions. *Crit Care*. 2016;20(1):187.
- [41] Granata A, Insalaco M, Di Pietro F, et al. Atheroembolism renal disease: diagnosis and etiologic factors. *Clin Ter*. 2012;163(4):313–322.
- [42] Swinkels BM, Ten Berg JM, Kelder JC, et al. Effect of aortic cross-clamp time on late survival after isolated aortic valve replacement. *Interact Cardiovasc Thorac Surg*. 2021;32(2):222–228.
- [43] Iasonos A, Schrag D, Raj GV, et al. How to build and interpret a nomogram for cancer prognosis. *J Clin Oncol*. 2008;26(8):1364–1370.
- [44] McBride WT, Kurth MJ, McLean G, et al. Stratifying risk of acute kidney injury in pre and post cardiac surgery patients using a novel biomarker-based algorithm and clinical risk score. *Sci Rep*. 2019;9(1):16963.