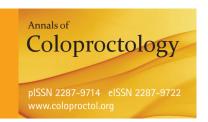
# **Original Article**

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# The Impact of Primary Tumor Resection on Survival in Asymptomatic Colorectal Cancer Patients With Unresectable Metastases

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Purpose: This study was conducted to evaluate the effectiveness of primary tumor resection (PTR) in asymptomatic colorectal cancer (CRC) patients with unresectable metastases using the inverse probability of treatment weighting (IPTW) method to minimize selection bias.

Methods: We selected 146 patients diagnosed with stage IV CRC with unresectable metastasis between 2001 and 2018 from our institutional database. In a multivariate logistic regression model using the patients' baseline covariates associated with PTR, we applied the IPTW method based on a propensity score and performed a weighted Cox proportional regression analysis to estimate survival according to PTR.

Results: Upfront PTR was performed in 98 patients, and no significant differences in baseline factors were detected. The upweighted median survival of the PTR group was 18 months and that of the non-PTR group was 15 months (P = 0.15). After applying the IPTW, the PTR was still insignificant in the univariate Cox regression (hazard ratio [HR], 0.26; 95% confidence interval [CI], 0.5–1.21). However, in the multivariate weighted Cox regression with adjustment for other covariates, the PTR showed a significantly decreased risk of cancer-related death (HR, 0.61; 95% CI, 0.40–0.94).

Conclusion: In this study, we showed that asymptomatic CRC patients with unresectable metastases could gain a survival benefit from upfront PTR by analysis with the IPTW method. However, randomized controlled trials are mandatory.

Keywords: Colonic neoplasms; Surgery; Neoplasm metastasis; Prognosis; Probability; Selection bias

### **INTRODUCTION**

According to reports of the 2018 International Agency for Research on Cancer, colorectal cancer (CRC) is the third most common cancer worldwide. In this report, the Korean incidence of CRC was the highest worldwide, although there has been a recent decrease in the annual domestic statistics [1, 2]. When the disease is diagnosed without metastasis, the 5-year survival rate is over

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70%. However, more than 20% of CRC cases are found with distant metastases, and overall survival (OS) is only about 14% to 20% [3]. Even for stage IV CRC, radical resection, including metastatic lesions, is mandatory to achieve long-term survival. However, in 65% of patients with stage IV CRC, radical surgery is not possible [4]. Upfront active chemotherapy, including various biological agents, is recommended in the current treatment guidelines [5]. However, primary tumor resection (PTR) or other palliative surgical treatments such as stoma formation or bypass surgeries are inevitable in cases of complications such as intestinal obstruction, hemorrhage, or perforation for symptomatic relief and further medical treatment. In asymptomatic patients at the time of diagnosis, many clinicians have performed upfront PTR before chemotherapy [6] because of the concern of future complications from the primary lesion during chemotherapy despite the debate about the usefulness of PTR in those situations [4, 6-15].

The treatment for these cases should be provided by a multidisciplinary team approach consisting of surgical, medical, and radi-

ation oncologists, along with pathologists and radiologists [16, 17]. By considering the resectability, location, and size of the primary/metastatic tumor, the treatment strategies are decided in consideration with the patient's preference, socioeconomic status, symptoms, and general condition. Therefore, various combinations of medical or nonmedical pretreatment factors can be involved in treatment decisions. To date, the reason why the debate about the advantage of PTR for asymptomatic CRC patients with unresectable metastasis continues is because of the lack of confirmative results of a well-controlled randomized trial [18, 19]. For studies supporting PTR, most points of criticism have been about their selection bias because of the various pretreatment factors mentioned earlier, which cannot be easily controlled [20].

In this study, we evaluated the effectiveness of PTR using the IPTW method, a recently emerging statistical tool to minimize selection bias by adjusting the baseline confounders.

#### **METHODS**

#### Patients and treatments

We selected 245 patients who were diagnosed with stage IV CRC with unresectable metastasis between 2001 and 2018 from our institutional CRC database. The staging workups consisted of computed tomography (CT), including the chest, abdomen, and pelvis, and positron emission tomography fused with CT after tissue confirmation by endoscopic biopsy. At the clinician's discretion, pelvic or liver magnetic resonance imaging was performed if necessary. The resectability of the primary and the metastatic lesions was assessed by reviewing the radiological findings. Whether the patient was symptomatic or not was determined by reviewing the medical records and radiologic results. Symptoms of the primary lesion were defined as follows: obstruction was determined when gas distention existed at a bowel site proximal to the primary lesion; defecation difficulty was defined as a large amount of fecal impaction without gas distension; perforation/abscess was defined as free air in the abdominal cavity or abnormal fluid collection around the primary tumor with any systemic inflammatory sign or symptoms; bleeding was defined as hemorrhage from the primary lesion that required blood transfusion to maintain stable blood pressure. After reviewing the medical records and radiologic results, we excluded 91 symptomatic patients and 8 patients with unresectable primary lesions. Our Institutional Review Board approved this study, and the requirement for informed consent was waived because of the retrospective nature of the study (No. 2020-05-005).

Surgical treatment consisted of PTR, stoma formation, and bypass surgery. PTR was performed in the same manner as surgery, indicating nonmetastatic CRC with a possible level of lymphadenectomy. For the intention-to-treat analysis, the patients who underwent PTR during the course of chemotherapy were classified as the non-PTR group.

Chemotherapies were classified as monotherapy (fluorouracil or

capecitabine only), doublet chemotherapy (monotherapy coupled with irinotecan or oxaliplatin), and doublet chemotherapy with a biologic agent (bevacizumab or cetuximab).

#### Statistical analysis

For the description of the study population, the chi-square or Fisher exact test was used to compare the categorical variables. For continuous variables, independent or paired t-test and the Wilcoxon rank-sum or signed-rank tests were used appropriately as required.

To identify the possible factors that influence the decision of performing PTR or non-PTR, we performed a multivariate binary logistic regression and calculated the propensity score (PS). Then, the inverse probability treatment weighting (IPTW) method was applied to minimize the selection bias. The variables for estimating the PS were as follows; sex, age, pretreatment Charlson comorbidity score, level of carcinoembryonic antigen, primary tumor location, organs involved in metastasis, multiple organ metastases, and chemotherapy. To check the goodness of fit and the discriminability of the model employed for calculating the PS, we used the Hosmer and Lemeshow test and the receiver-operating curve (ROC) analysis. The weights designated to each patient were estimated using the following formula: for the PTR group, the weights were allocated as 1; and for the non-PTR group, the weights were calculated as PS/(1 – PS) [21].

To check the imbalance of covariates between the PTR and non-PTR groups, the standardized mean difference (SMD) was calculated before and after weighting, and an absolute value of SMD less than 0.2 was considered balanced [22].

As the endpoint of this study, the OS was defined as the time from the date of diagnosis of stage IV CRC to the date of cancerrelated death. For survival analysis and comparison, the Kaplan-Meier and log-rank tests were used. Cox proportional hazard regression analysis with/without weighting was performed for the univariate and multivariate analyses. Variables with a P-value of < 0.2 in the weighted univariate analysis were selected for weighted multivariate analysis. For the multivariate analysis, a backward stepwise elimination method was used. All tests were 2-sided, and a P-value of < 0.05 was considered significant. R software ver. 4.0.0 (R Foundation for Statistical Computing, Vienna, Austria) was used for the statistical analyses.

#### RESULTS

#### Patients' characteristics

The median age of the 146 patients was 68.5 years (interquartile range [IQR], 59–76 years). The number of male patients was 78 (53.4%). The median Charlson comorbidity index (CCI) score was 3.0 (IQR, 2.0–4.0). The most common metastatic organ was the liver (70.5%), and peritoneal metastasis was observed in 29 patients (19.9%). Chemotherapy was administered to 121 patients (82.9%). The number of patients who received doublet chemo-

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therapy or targeted therapy was 106 (72.6%). Surgical treatment was performed in 140 patients (95.9%). In addition, 110 patients (75.3%) underwent PTR. Among them, 12 patients underwent upfront chemotherapy before PTR, and they were classified into the non-PTR group for intention-to-treat analysis. Twenty-seven patients with non-PTR (n=48) underwent upfront chemotherapy. Among them, 13 patients (48.1%) received surgical treatment because of complications of the primary lesion (PTR, 4; stoma/bypass, 9). Among the 98 upfront PTR patients, 1 patient developed anastomotic leakage after low anterior resection, 2 patients underwent wound evisceration after surgery, and there was no opera-

tion-related mortality. After upfront PTR, 80 patients (81.6%) received chemotherapy. The median chemotherapy initiation time was 29 days (IQR, 21–36 days) after upfront PTR (Table 1).

# Balance between the PTR and non-PTR groups

Although there were no significant differences in the baseline characteristics between the groups in the upweighted descriptive analysis (Table 1), there were potentially important imbalances in the 3 covariates with SMD > 0.2 (Fig. 1); the levels of carcinoembryonic antigen (SMD = 0.41), the presence of distant lymph node metastasis (SMD = 0.36), and the primary tumor location in the

Table 1. Descriptive analyses between the primary tumor resection (PTR) vs. non-PTR group

Variable	PTR	Non-PTR	Total	P-value
No. of patients	98	48	146	
Male sex	49 (50.0)	29 (60.4)	78 (53.4)	0.313
Age (yr)	69.0 (58.0–77.0)	66.5 (62.0-75.0)	68.5 (59.0–76.0)	0.881
Charlson comorbidity index	3.0 (2.0-4.0)	3.0 (2.0-4.0)	3.0 (2.0-4.0)	0.541
Level of CEA (ng/dL)	24.2 (4.7–123.5)	16.9 (4.8–105.5)	20.4 (4.7–120.2)	0.452
Location of primary lesion				0.262
Right colon	29 (29.6)	10 (20.8)	39 (26.7)	
Left colon	39 (39.8)	17 (35.4)	56 (38.4)	
Rectum	30 (30.6)	21 (43.8)	51 (34.9)	
Multiple organ metastases	43 (43.9)	21 (43.8)	64 (43.8)	1.000
Site of metastasis				
Liver	72 (73.5)	31 (64.6)	103 (70.5)	0.361
Lung	37 (37.8)	17 (35.4)	54 (37.0)	0.926
Peritoneum	19 (19.4)	10 (20.8)	29 (19.9)	1.000
Distant lymph node	17 (17.3)	15 (31.2)	32 (21.9)	0.09
Bone	6 (6.1)	1 (2.1)	7 (4.8)	0.509
Ovary	7 (7.1)	1 (2.1)	8 (5.5)	0.382
Others	2 (2.0)	1 (2.1)	3 (2.1)	1.000
Surgical treatment				< 0.0001*
None	0 (0)	6 (12.5)	6 (4.1)	
PTR	98 (100)	12 (25.0)	110 (75.3)	
Stoma/bypass	0 (0)	30 (62.5)	30 (20.5)	
Chemotherapy, yes	80 (81.6)	41 (85.4)	121 (82.9)	0.737
First-line regimen				0.644
None	18 (18.4)	7 (14.6)	25 (17.1)	
Monotherapy	9 (9.2)	6 (12.5)	15 (10.3)	
Doublet	57 (58.2)	25 (52.1)	82 (56.2)	
Biologic agent	14 (14.3)	10 (20.8)	24 (16.4)	
Upfront chemotherapy, yes	0 (0)	27 (56.2)	27 (18.5)	< 0.0001*

Values are presented as number only, number (%), or median (interquartile range).

CEA, carcinoembryonic antigen.

<sup>\*</sup>Statistically significant (P < 0.05).

rectum (SMD = 0.32). As these could be a major source of bias for further comparison, we constructed a PS to predict probabilities of treatment allocation with baseline covariates by a multivariable logistic regression and calculated the weights for applying the

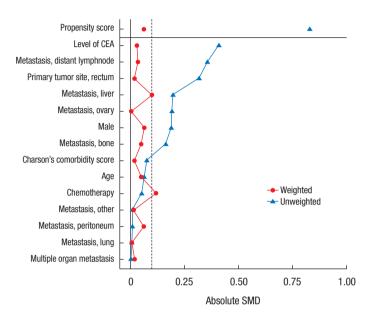


Fig. 1. Standardized mean difference (SMD) of baseline covariates before and after inverse probability treatment weighting. CEA, carcinoembryonic antigen.

IPTW as described in the method section. This treatment allocation model showed acceptable adequacy with a Hosmer and Lemeshow goodness of fit test ( $\chi^2$  = 6.7, degree of freedom = 8, P=0.57) and discrimination with an area under the ROC of 0.72. After weighing all 14 covariates, SMDs were less than 0.2 and were considered to be well balanced (Fig. 1).

#### Survival analysis

During the median 18 months (IQR, 11-29 months), 132 of the 146 patients died because of disease, and the median survival time was 18 months. Comparing the survival according to whether the PTR was performed in the upweighted population, the median survival months of the PTR group was 18 months, and that of the non-PTR group was 15 months (P=0.152).

In the univariate Cox regression analysis without weights, the age (hazard ratio [HR], 1.03; 95% confidence interval [CI], 1.01–1.05], CCI score (HR, 1.26; 95% CI, 1.10–1.44), and chemotherapy (HR, 0.5; 95% CI, 0.36–0.9) were found to be significant coprognostic factors affecting cancer-specific survival. The PTR showed a decreased risk of death but was not significant (HR, 0.77; 95% CI, 0.53–1.13). In the weighted univariate Cox regression, the significant prognostic factors were age (HR, 1.03; 95% CI, 1.01–1.05), CCI score (HR, 1.25; 95% CI, 1.04–1.5), the organ of metastasis in spleen, brain (indicated as others) (HR, 2.06; 95% CI, 1.01–4.21), multiple organ metastasis (HR, 1.53; 95% CI, 1.03–2.29), and chemotherapy (HR, 0.49; 95% CI, 0.26–0.94). The PTR was still an insignificant factor in this univariate weighted

Table 2. Univariate Cox analysis of the factors associated with cancer-related survival

Variable	Unweighted		Weighte	Weighted	
	HR (95% CI)	P-value	HR (95% CI)	P-value	
Primary tumor resection, yes	0.772 (0.529–1.126)	0.179	0.773 (0.495–1.208)	0.258	
Male sex	0.912 (0.644-1.292)	0.604	0.818 (0.542-1.234)	0.338	
Age	1.030 (1.012–1.047)	0.001	1.031 (1.011–1.051)	0.002*	
Charlson comorbidity index	1.256 (1.095–1.442)	0.001	1.254 (1.043–1.507)	0.016*	
Level of CEA	1 (1–1.001)	0.095	1 (1–1)	0.165	
Site of metastasis					
Liver	1.480 (1.001–2.188)	0.050	1.099 (0.665-1.816)	0.713	
Lung	1.017 (0.708-1.462)	0.927	1.071 (0.721-1.590)	0.733	
Distant lymph node	0.921 (0.598-1.419)	0.709	0.889 (0.562-1.406)	0.615	
Peritoneum	0.871 (0.558-1.360)	0.544	1.116 (0.630-1.975)	0.707	
Bone	1.664 (0.773–3.583)	0.193	2.337 (0.705–7.749)	0.165	
Ovary	0.672 (0.313-1.446)	0.310	1.163 (0.390-3.469)	0.787	
Others	1.952 (0.616–6.188)	0.256	2.059 (1.008-4.208)	0.048*	
Multiple organ metastases, yes	1.362 (0.952–1.951)	0.091	1.529 (1.026–2.281)	0.037*	
Chemotherapy, yes	0.549 (0.343-0.876)	0.012	0.494 (0.259-0.943)	0.032*	

HR, hazard ratio; CI, confidence interval; CEA, carcinoembryonic antigen.

<sup>\*</sup>Statistically significant (P < 0.05).

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Table 3. Multivariate Cox analysis of the factors associated with cancer-related survival

Variable	HR (95% CI)	P-value
Primary tumor resection, yes	0.614 (0.402-0.938)	0.024*
Age	1.030 (1.007-1.054)	0.012*
Metastasis, bone	2.592 (0.84-7.992)	0.097
Multiple organ metastases, yes	1.736 (1.121–2.688)	0.013*
Chemotherapy, yes	0.570 (0.292-1.116)	0.101

HR, hazard ratio; CI, confidence interval.

Cox regression (HR, 0.26; 95% CI, 0.5–1.21) (Table 2).

However, in the multivariate weighted Cox regression with the factors whose P-values were less than 0.2, including the PTR at the weighed univariate Cox regression, the PTR showed a significantly decreased risk of cancer-related death (HR, 0.61; 95% CI, 0.40–0.94) (Table 3, Fig. 2).

#### **DISCUSSION**

In this study, we showed that asymptomatic CRC patients with unresectable metastases could achieve a survival gain by upfront PTR independently from the patients' likelihood of undergoing PTR based on their baseline characteristics by analysis with the IPTW method.

There are significant conflicts in the study of the impact of PTR on survival [6]. According to a study, PTR is performed in over 50% of patients with stage IV CRC, and many doctors still prefer routine use of this treatment. This may be because many current studies suggest that PTR can provide significant survival benefits [7, 10, 12, 15, 23-25]. In a Canadian study of stage IV CRC patients who underwent chemotherapy between and 2006 to 2010, 199 patients underwent upfront PTR, and 127 patients did not [24]. Ninety-one percent and 67% of patients received doublet chemotherapy and biologic agents, respectively. The median OS was significantly higher in the PTR group (27 months vs. 14 months). However, their results were not free from selection bias because there was a significant imbalance in the distribution of confounders, a higher rate of multiple organ metastasis (M1b), and worse performance status in the non-PTR group. In a post hoc analysis of 294 patients with non-resectable CRC metastases who were enrolled in a phase III randomized controlled trial (RCT), patients who underwent PTR before enrollment showed higher rates of OS than non-PTR patients [23]. However, the majority of PTR patients had single organ metastases (M1a) in the liver. Similarly, in a meta-analysis using survival data from 21 retrospective studies with a total of 44,226 patients with metastatic CRC, the mortality rate was lower with PTR than with chemotherapy alone. However, all but 2 of the studies in the analysis were single-center and the other 2 studies used administrative da-

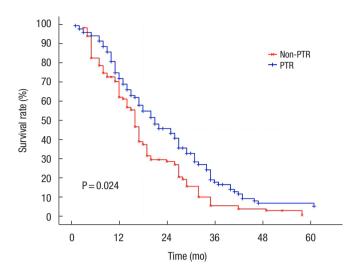


Fig. 2. Estimated survival of primary tumor resection (PTR) and non-PTR group by weighted multivariate Cox regression.

tabases with limited information about how patients were selected for the PTR [7].

These continued conflicts are because of the lack of firm results of the level I study. To answer this question, some clinical trials are ongoing [18, 19, 26, 27]. Kanemitsu et al. [28] reported their interim analysis results of an RCT at the 2020 American Society of Clinical Oncology meeting after early termination due to 50% of the expected event. They enrolled 160 patients, which was about half of the target enrollment number (n = 280). The upfront PTR showed no survival benefit over the chemotherapy alone; the PTR group's median progression-free survival was 10.4 months (8.6-13.4 months) and that of chemotherapy alone was 12.1 months (9.4-13.2 months; HR, 1.08; 95% CI, 0.77-1.50). Although the overall results of the remaining RCTs are expected, Kanemitsu's assertions imply that they would not be easy to see. Because all RCTs on this issue carry numerous intrinsic difficulties caused by strict eligibility criteria, patient preferences, and clinician bias lead to poor acceptance of the randomized assignment and difficulties in patient accrual.

In this retrospective study, we could see the survival benefit of upfront PTR. However, as mentioned earlier, this study has a fundamental limitation that cannot be free in this kind of study. The metastatic tumor burden and personal response to chemotherapy may be different for each patient. However, because of the lack of information on retrospective medical record review, we could not obtain detailed data. Although a sincere effort to reduce the selection bias, the patients selected from our database are those who visited our department of surgical oncology. Therefore, doctors in this department might have a strong tendency to prefer PTR. Additionally, because of the long period of study, most of the earlier decisions were not made with multidisciplinary team discussions. Inevitably, the patients enrolled in this study have a strong likeli-

<sup>\*</sup>Statistically significant (P < 0.05).

hood of getting upfront PTR. To overcome this issue, we used the weight from the average treatment effect on the treated (ATT). ATT is the one of the generally used treatment effects typically considered in causal inference. The other is the average treatment effect (ATE), which is the average effect of the treatment across the entire eligible sample, including the control sample's treatment effect. In a well-constructed RCT with perfect compliance to treatment assignment, both values must be the same. However, in the nonrandomized, observational studies like ours, the assignment of treatment can be influenced by various confounders such as patients' willingness or surgeon's preference for the PTR, which is not easily measurable. Endogenous selection bias is inevitable when using the ATE, referring to the entire study population, including those who did not receive the PTR. However, by using the ATT weight focusing only among those who received the PTR, we can estimate the effect of PTR in patients who received the real PTR by comparing in an assumed situation where the PTR patients might not receive the PTR [21, 29].

In conclusion, after minimizing the fundamental limitations of this retrospective study, this study demonstrated the beneficial impact of PTR on unresectable stage IV CRC patients. However, the debate about this issue will continue until the confirmative results of ongoing RCTs.

#### **CONFLICT OF INTEREST**

No potential conflict of interest relevant to this article was reported.

## **ACKNOWLEDGMENTS**

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