

1 **FRONT MATTER**

2
3 **Title**

4 Lifespan Normative Models of White Matter Fractional Anisotropy: Applications to Early
5 Psychosis

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24 **Abstract**

25 This study presents large-scale normative models of white matter (WM) organization
26 across the lifespan, using diffusion MRI data from over 25,000 healthy individuals aged 0-
27 100 years. These models capture lifespan trajectories and inter-individual variation in
28 fractional anisotropy (FA), a marker of white matter integrity. By addressing non-
29 Gaussian data distributions, race, and site effects, the models offer reference baselines
30 across diverse ages, ethnicities, and scanning conditions. We applied these FA models to
31 the HCP Early Psychosis cohort and performed a multivariate analysis to map symptoms
32 onto deviations from multimodal normative models using multi-view sparse canonical
33 correlation analysis (msCCA). Our results reveal extensive white matter heterogeneity in
34 psychosis, which is not captured by group-level analyses, with key regions identified,
35 including the right uncinate fasciculus and thalami. These normative models offer
36 valuable tools for individualized WM deviation identification, improving precision in
37 psychiatric assessments. All models are publicly available for community use.

38 **Teaser**

39 Lifespan models of white matter offer insights into brain health, providing tools for
40 tracking individual deviations across ages.
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46 **MAIN TEXT**

47
48 **Introduction**

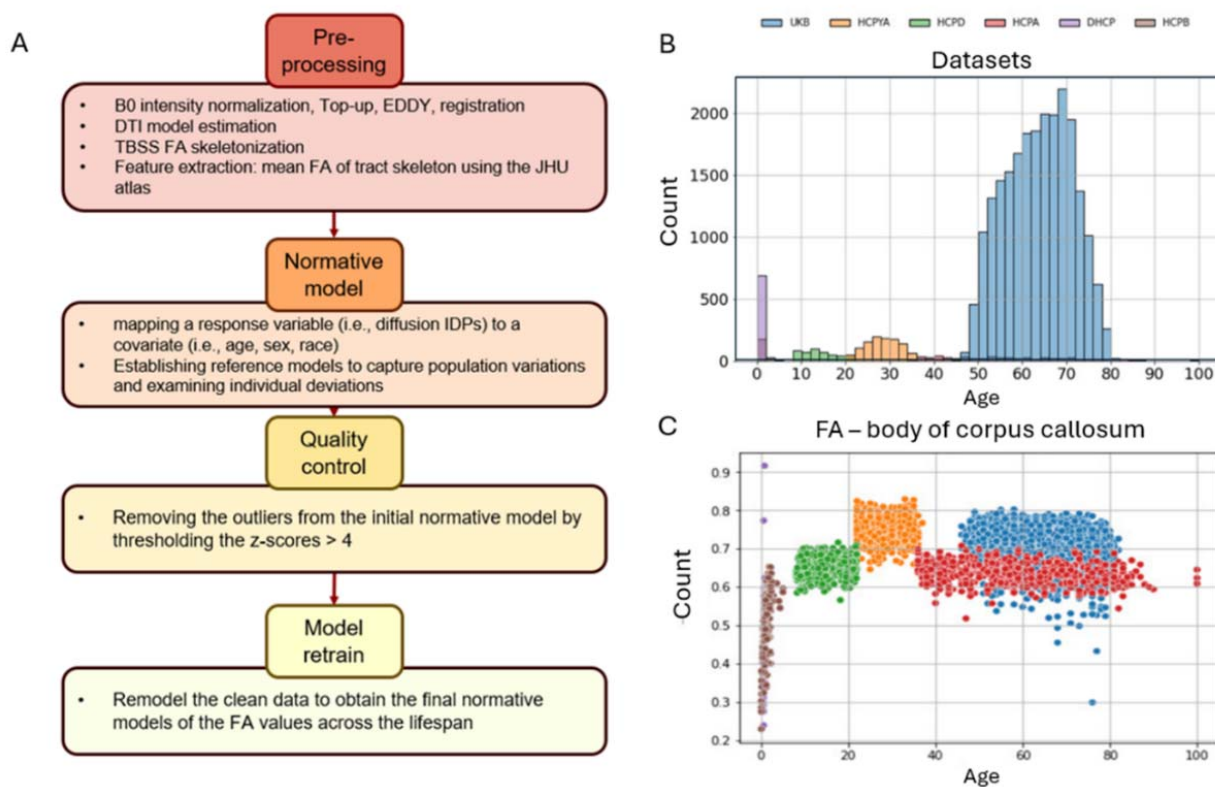
49 Over the past century, normative growth charts have become integral to paediatric
50 practice, providing essential benchmarks for comparing individual growth patterns
51 (height, weight, head circumference) with established population standards. These charts
52 have facilitated a better understanding of typical developmental trajectories and have been
53 crucial in identifying deviations from expected growth patterns which are used in clinical
54 practice to determine if additional medical workup or treatment is required [1]. This
55 concept has recently been extended to the field of neuroimaging, where it allows for
56 detailed, individual-level insights into lifespan trajectories of brain measures. By
57 comparing individual neuroimaging data against large, normative reference datasets,
58 researchers and clinicians can gain a deeper understanding of both typical and atypical
59 brain development and aging [2], [3], [4], [5].

60 In psychiatric disorders, traditional case-control studies have been valuable for
61 detecting abnormalities in structural, microstructural, functional and neurometabolic brain
62 signatures in patient groups compared to control groups. However, group comparisons are
63 not designed to capture inter-individual heterogeneity which is prominent at the
64 phenotypic and biological levels in virtually all psychiatric disorders. This significant
65 translational gap hampers identification of specific biological markers that explain clinical
66 heterogeneity in these disorders such as disease risk, severity, and progression, as well as
67 responsiveness to pharmacological and non-pharmacological treatments and overall
68 clinical outcomes. Normative modelling provides a precision framework that has emerged
69 as a promising tool in this endeavour [6], [7], [8]. By comparing brain imaging data
70 against large reference cohorts, this method allows us to quantify deviations from
71 expected norms at the individual level. It is now possible to capture deviation profiles in a
72 single patient, which offers a more nuanced understanding of biological variations in
73 psychiatric disorders. Even more importantly, it also has promise for bridging this
74 translational gap by providing a foundational framework for developing tailored tools that
75 capture disease risk and progression, as well as precision treatments tailored to individual
76 brain pathology. For instance, normative models capture inter-individual biological
77 variations that provided important insights into heterogeneity in schizophrenia, major
78 depressive disorder, bipolar disorder, ADHD and autism spectrum disorders [6], [9], [10].
79 Moreover, we have demonstrated that normative measures frequently outperform raw
80 measures (e.g. cortical thickness in mm) in group difference testing, disease classification
81 [11] and treatment response prediction [12].

82 We and others have created large-scale normative models that leveraged >50,000
83 healthy volunteer imaging datasets for structural [4], [5], [13] and functional MRI [11],
84 [14]. To our knowledge, no comprehensive normative models for diffusion weighted
85 imaging measures at comparable scale exist at this time. There are several reasons why
86 this is the case. First, diffusion imaging was developed more recently than structural and
87 functional MRI, and a broader adoption in neuroscience research did not happen until the
88 early 2010s. Second, processing of diffusion data is more computationally demanding
89 compared to structural and functional MRI and the gold standard for diffusion data quality
90 control remains visual inspection; both of these factors have been substantial limitations to
91 scaling efforts. Third, diffusion imaging measures are very sensitive to differences
92 between vendors, individual scanners (e.g. signal intensity variations, eddy currents), and
93 sequence acquisition parameters (e.g. b-values), making it difficult to integrate different
94 datasets necessary to develop lifespan normative models. To date, only two preliminary

95 studies have fit normative models to diffusion weighted data [15],[16]. In [15], the authors
96 used approximately 1,300 single-shell DTI datasets collected at eight different sites using
97 the same vendor to test performance of different statistical methods within the normative
98 framework and in [16] the authors focus principally on generating reference curves for
99 data harmonisation.

100 The aims of this study are to: (i) develop normative models of Fractional
101 Anisotropy (FA), the most widely used diffusion metric in neuroimaging [17], across
102 major white matter tracts using a large dataset of over 25,000 healthy individuals across a
103 broad age range. By using high-quality diffusion MRI data from the UK Biobank and the
104 Human Connectome Project, we seek to establish robust models that capture lifespan
105 trajectories of white matter organization; (ii) investigate white matter FA in early
106 psychosis, a prototypical psychiatric disorder that is known to be highly heterogeneous in
107 disease severity and course, as well as clinical symptom expression and clinical outcomes.
108 Using the HCP Early Psychosis (HCP-EP) dataset [18], we aim to map both group level
109 differences and individual deviations from the normative model in order to better
110 understand individual variability in white matter integrity; (iii) we aim to illustrate the
111 value of normative models for multi-modal data fusion, by combining FA deviations with
112 cortical thickness and subcortical brain volume deviations with the goal to identify multi-
113 modal biological signatures and specific white matter pathways in psychosis associated
114 with different psychosis symptom domains. Finally, (iv) we release all models freely to
115 the community via our existing open-source software platforms [19].



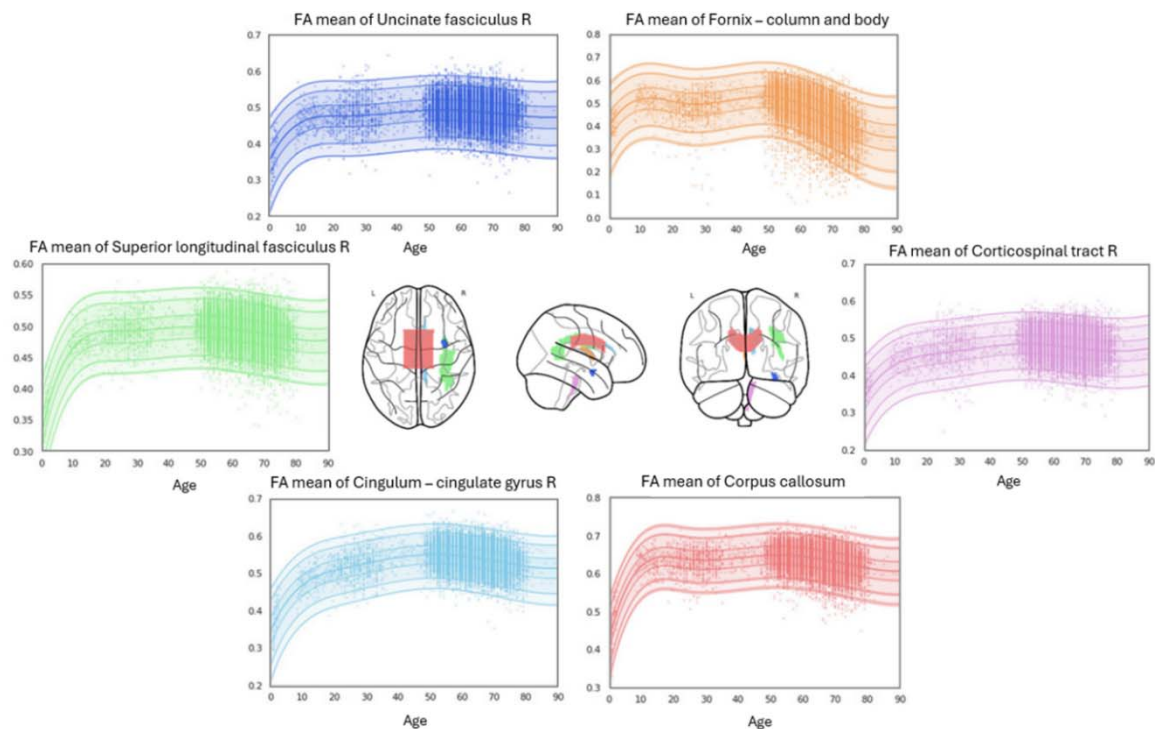
116 Figure 1. A) Flow chart of the main diffusion image processing steps B) Histogram plot of
117 the data used for normative modeling, showing the population density at each age and
118 highlighting the different datasets used C) Scatterplot exemplifying the quality control
119 process using normative modeling and outlier exclusion based on Z-score thresholding. In
120 this plot, site effects are clearly evident, which are accommodated by the normative
121 models (see Figure 2).
122

123 Results

124 Normative modelling

125 First, we assembled high-quality multi-shell diffusion data from five cohorts
126 having closely matched acquisition and processing pipelines, namely Human Connectome
127 Project (HCP) Baby [20], HCP Development [21], HCP Young Adult [22], HCP Aging
128 [23] datasets, and UK Biobank [24]. Total $N=24,915$, ($N=12,457$ for training and
129 $N=12,457$ for test, stratified for sex, self-reported race, dataset and site). A summary of the
130 sample and processing is provided in Figure 1 with further details in the methods. In short,
131 the datasets were processed using harmonised FSL-based pipelines, involving pre-
132 processing (intensity normalisation, distortion and movement corrections), DTI modelling
133 to extract fractional anisotropy (FA) values, Tract-Based Spatial Statistics (TBSS) for
134 skeletonised FA images, and segmentation with the Johns Hopkins University (JHU) atlas
135 to compute mean FA values across 48 white matter tracts. We then fit lifespan normative
136 models to these data on the basis of age, sex, site and race using warped Bayesian linear
137 regression (BLR) and a non-linear basis expansion over age, in line with our prior work
138 [4], [25]. We assessed the quality of the normative modeling fit using three key out-of-
139 sample metrics, namely explained variance (EV), evaluating the fit of the median
140 regression line, in addition to skewness and kurtosis, which evaluate the shape of the
141 distribution used to model the centiles. These metrics offer insight into how well the
142 models capture the underlying distribution of the data across 48 white matter tracts. The
143 mean (standard deviation) EV was 0.37 (0.10), indicating good fit across different models.
144 Skewness, and kurtosis were respectively -0.09 (0.12) and 0.42 (0.27), which together
145 indicate that the shape was also appropriate for the data. Supplementary figure 1 shows a
146 histogram of the EV, skew and kurtosis of the models.

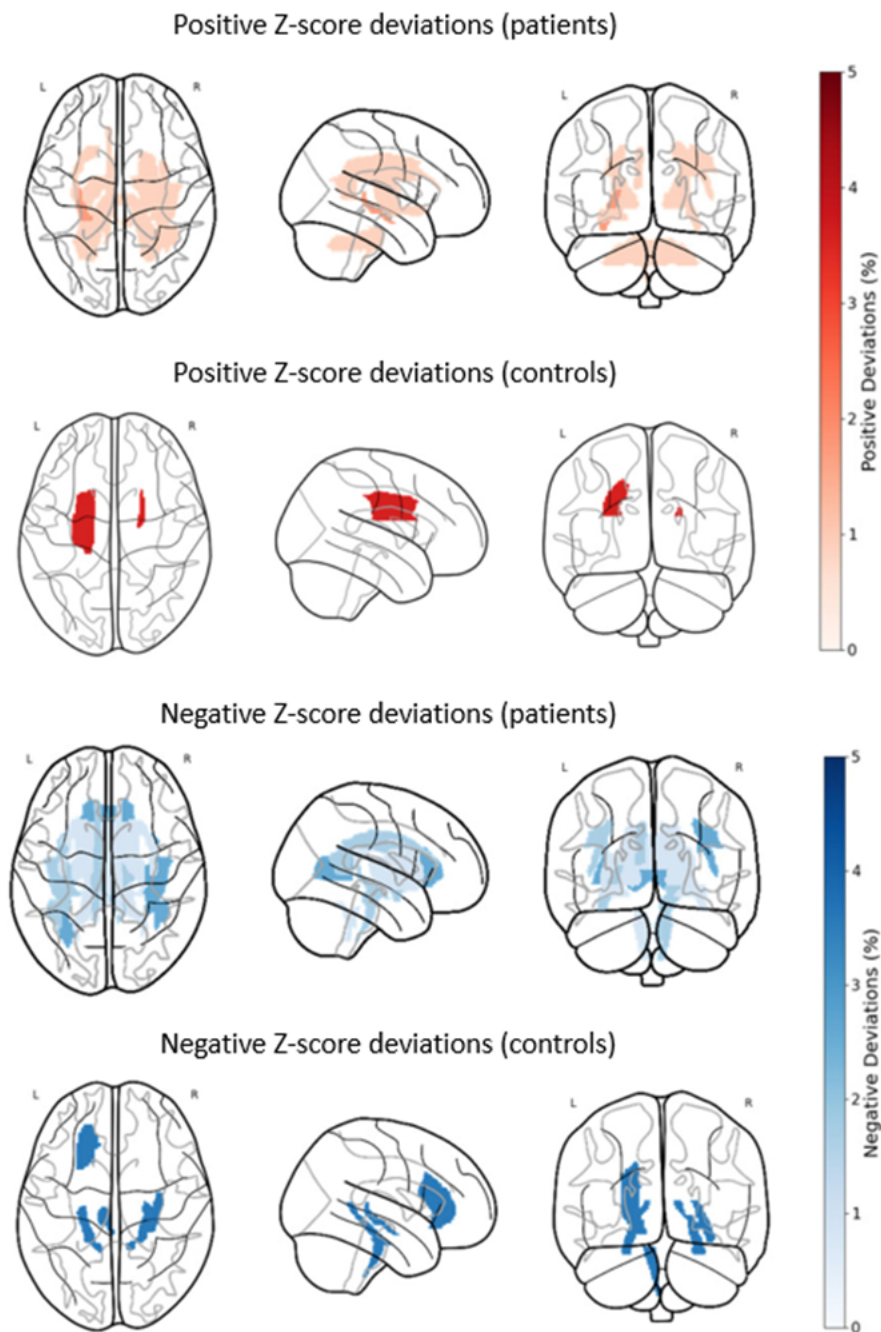
147 We illustrate the trajectory and fitted centiles for a selection of white matter tracts
148 across the lifespan in Figure 2. The complete set can be found in the supplementary figure
149 2. In addition, we also show the results of models that do not include race in the
150 supplementary figure 3.



151 Figure 2. A selection of six white matter tracts and their corresponding normative
152 modelling centile plots highlighting the similarity in white matter formation and
153 degeneration along the lifespan as well as tract specific differences in terms of shapes and
154 variance of the FA values. For visualization purposes, data from different sites are aligned
155 to a common reference (e.g. the mean centiles or the centiles for an arbitrary chosen site)
156 by computing the z-scores separately for each site using the site-specific means and
157 standard deviations, then inverting the z-scores using the mean and standard deviation
158 derived from the common reference.

159 **Application to a clinical dataset**

160 Next, we used these models to understand heterogeneity in white matter FA in
161 psychosis. To achieve this, we applied these reference models to the HCP early psychosis
162 (HCP-EP) dataset (N=173 with diffusion data - see supplementary table 2 for
163 demographic information) in order to derive z-scores for each individual and tract. We
164 evaluated the mean differences in normative deviations between patients and controls for
165 each tract using a t-test, applying false discovery rate (FDR) correction [26] to account for
166 multiple comparisons. There were no significant differences in the mean deviations
167 between individuals with psychosis and healthy controls that survived false discovery rate
168 (FDR) multiple comparison correction, although we did find nominally significant effects
169 in the fornix (column and body and the stria terminalis bilaterally). However, we did find
170 evidence for significantly more heterogeneity in individuals with psychosis relative to
171 controls in terms of the proportion of extreme deviations. More specifically, individuals
172 with schizophrenia had a greater proportion of extreme positive (Mann-Whitney
173 $U=1403.0$, $p=0.0036$) and extreme negative ($U=1517.0$, $p=0.0016$) Z-scores relative to
174 controls, indicating substantial differences between groups that were highly variable
175 across individuals. Notably, while both positive and negative deviations were present, the
176 prominence of negative outliers (subjects with Z-scores exceeding ± 2.6 was particularly
177 pronounced, highlighting a consistent trend where patients exhibited a greater number of
178 extreme Z-scores across white matter tracts. We show the percentage overlap of extreme
179 deviations across all tracts in Figure 3, which reveals that extreme positive and negative
180 deviations were observed in some individuals with psychosis in nearly all tracts. In
181 contrast, the extreme deviations in controls were more focused, confined to only several
182 white matter tracts, as illustrated in the bar plots in the supplementary figure 4, with an
183 alternative representation highlighting overlap in individual tracts.



185 Figure 3: Glass brain representations illustrating the overlap between extreme
186 positive and negative Z-score deviations for patients and controls, with thresholds set at -
187 2.6 and 2.6 which correspond to a p-value of 0.01. This stringent threshold enhances the
188 detection of significant deviations while controlling for false positives. The top two panels
189 depict positive Z-score deviations for patients and controls, while the bottom two panels
190 show negative Z-score deviations for patients and controls. The legend indicates the
191 percentage of subjects having extreme deviations in each tract.

192

193 Next, we sought to demonstrate the utility of these normative models in identifying
194 multivariate brain-behaviour associations within a clinical cohort. To achieve this, we
195 combined FA deviations with cortical thickness and subcortical volume measures from

196 our previously published models [4] in a multimodal analysis. Symptom severity was
197 quantified using the Positive and Negative Syndrome Scale (PANSS) [27], with domain
198 scores for positive, negative, and cognitive symptoms, as well as the total score,
199 summarised using a standard factor model, the ‘Marder’ factors, which were estimated
200 and released by the HCP-EP consortium. More specifically, we included the positive
201 symptom factor, negative symptom factor, cognitive/disorganised symptom factor in
202 addition to the total PANSS score. To determine the multivariate association with
203 symptoms, we used an approach we have employed in prior work [28], based on a multi-
204 view sparse canonical correlation analysis (msCCA) and stability selection [29] (see
205 methods for details). Briefly, aimed to learn the association between three ‘views’ of the
206 data, namely symptom domains, FA deviations and structural deviations (i.e. deviations
207 from normative models of cortical thickness and subcortical volume). Next, we randomly
208 split the data 1000 times into training (70%) and test (30%) sets, then fit an msCCA model
209 and report the mean canonical correlation on the test set. This analysis yielded a
210 significant mean test canonical correlation of $r=0.25$ for the leading component ($p=0.003$
211 under permutation testing, see Methods for details). This model showed good predictive
212 performance for both the associations between symptoms and FA deviations and
213 symptoms and structural (cortical and subcortical) deviations, but not between diffusion
214 and structural deviations (Figure 4 A). This is expected because we deliberately do not
215 optimise directly for this to prevent the model learning the trivial correlation between
216 different types of brain features (see methods). The second and third components
217 achieved test canonical correlations of $r=0.04$ ($p=0.11$) and $r=0.02$ ($p=0.03$) respectively.
218 However, considering the limited clinical relevance of associations of this magnitude and
219 their marginal significance of the third component, we focus principally on the first.

220 The symptom loadings derived from the msCCA analysis show that the association
221 was principally driven by the cognitive factor and total PANSS scores (Figure 4 B). We
222 used stability selection to determine the most informative features driving the association
223 by counting the number of times each feature was selected under the 1000 random splits
224 described above and considered samples having a selection probability greater than 0.8 as
225 informative. Note that this threshold is theoretically justified in order to control the type 1
226 error rate [29]. Under this threshold, FA in the right uncinate fasciculus and volume of the
227 thalamus bilaterally were predictive of PANSS symptoms Figure 4 C-F.

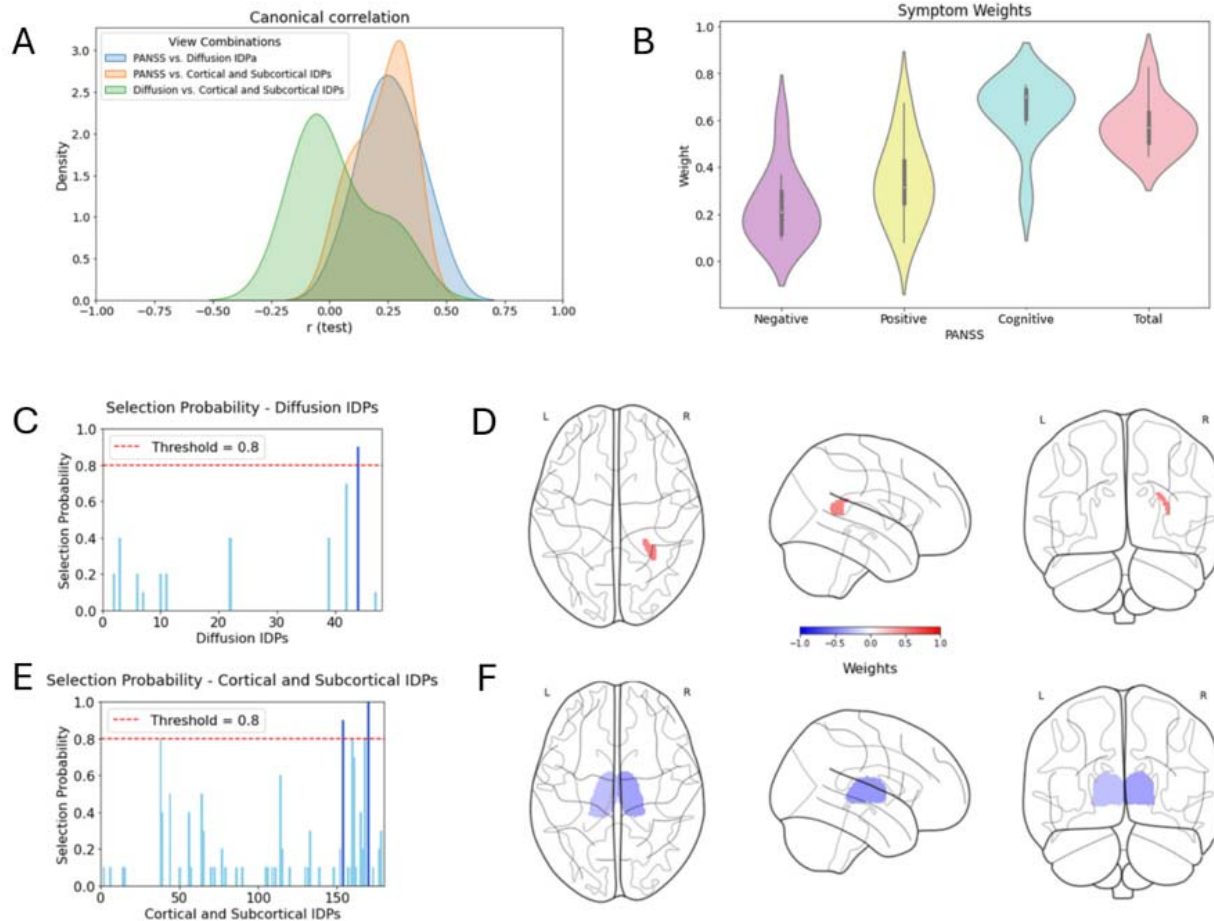


Figure 4: (A) Density plot of the multiple sparse Canonical Correlation Analysis (msCCA) main components, highlighting the distribution of test canonical correlations separately for each pair of views. (B) Violin plots representing the weights of PANSS symptom scores across the four symptom categories, namely negative symptoms, positive symptoms, cognitive symptoms, and total symptoms. (C) and (D) Selection probabilities for diffusion white matter tracts and Cortical Thickness white matter tracts, respectively, with a red threshold line indicating the chosen selection threshold of 80%. (E) and (F) Glass brain representations of the significantly selected white matter tracts and subcortical regions of interest, respectively. Note that no cortical thickness ROIs survived the selection threshold. The highlighted regions include the uncinate fasciculus (right) for diffusion and the cortical thickness regions: Left-Thalamus and Right-Thalamus

Discussion

This study presents a set of large-scale normative models for FA across major white matter tracts, estimated from a dataset of over 25,000 individuals spanning infancy to old age. Leveraging high-quality multi-shell diffusion MRI data, these models map the trajectory of white matter development and degeneration over the lifespan whilst also quantifying variance across the population. We showcase the clinical utility of these models by mapping inter-individual variation in cohorts of individuals with early psychosis. We show a high degree of inter-individual heterogeneity in these individuals, evidenced by relative increases in both positive- and negative deviations from the normative model in individuals with psychosis relative to controls. These differences were

251 evident despite an absence of case control effects, indicating that the differences were
252 highly individualized. Finally, we show that normative deviations of FA, cortical
253 thickness and subcortical brain volume were accurate multi-modal predictors of
254 symptomatology. Taken together, our findings provide a step toward advancing the
255 understanding of the heterogeneity of white matter alterations in early psychosis.

256 Our normative models show region-specific developmental trajectories in white
257 matter organization that align well with foundational findings on lifespan changes in FA
258 [30], [31]. However, we also show that inter-individual variability is considerably higher
259 than the magnitude of lifespan-related changes in FA, underscoring the importance of
260 using approaches such as this to characterize this at the individual level. Studies suggest
261 that increased FA during development relates to synaptic pruning and myelination, while
262 declines in old age are linked to axonal degradation and reduced fiber coherence [32],
263 [33]. Our models robustly capture these patterns, underscoring their relevance as a
264 normative reference sample and utility for studies examining brain aging and clinical
265 conditions.

266 In the HCP-EP cohort, we show a high degree of inter-individual variability in
267 white matter organization in psychosis, consistent with the variability that has already
268 been described in brain structure [9], [10], and in other psychiatric conditions [6], [7],
269 [34], which speaks to the potential for normative models as a basis for stratifying cohorts
270 [3]. Note that this variability was evident in an absence of case-control effects, which
271 indicates that inter-individual variability masks group level effects, which we also have
272 observed in gray matter in autism [7]. We also show a multivariate correspondence
273 between brain connectivity deviations, structural deviations and clinical symptoms driven
274 by decreased volume in the thalamus and FA in the right uncinate fasciculus. The left and
275 right thalamus showed negative weights, indicating that reductions in subcortical volume
276 may be linked to greater symptom severity. The uncinate fasciculus exhibited a positive
277 weight in relation to clinical symptoms, suggesting a possible compensatory role, although
278 we cannot rule out that this finding may also reflect other factors (such as crossing fibres).
279 In line with this interpretation, alterations in the uncinate fasciculus have been previously
280 reported in psychotic disorders, suggesting its involvement in the pathophysiology of
281 these conditions [14], [35]. Studies utilizing DTI have reported abnormalities in the
282 uncinate fasciculus among individuals with schizophrenia and affective psychosis. For
283 instance, Kawashima et al. [36] found reduced FA in the uncinate fasciculus of patients
284 with recent-onset schizophrenia, indicating compromised white matter organisation in this
285 tract [36].

286 One of the benefits of this study is that we focus on acquiring a high-quality
287 diffusion sample with closely harmonized protocols. This maximizes the ability to
288 attribute detected variations to biological differences, rather than artefacts such as data
289 quality or residual site effects. In this study, we prioritised modelling FA as it is the most
290 commonly used diffusion metric in the field due to its sensitivity to microstructural
291 integrity factors like axonal density, fibre coherence, and myelination, making it a
292 valuable and accessible measure for understanding white matter architecture. Additionally,
293 FA is less affected by CSF contamination compared to metrics like mean diffusivity
294 (MD), allowing for more accurate assessments in regions prone to such contamination,
295 such as the fornix [37]. However, this is only the first step, we intend to augment these
296 models with further models, including other tensor-based metrics, such as mean diffusivity
297 MD and non-tensor models (e.g. neurite orientation dispersion and density imaging;
298 NODDI [38], [39], to take full advantage of the multi-shell diffusion data and provide an
299 even more comprehensive resource for white matter analysis. Finally, we provide these

300 models to the field via our established no-code software platform [19] and via open-source
301 software tools (https://github.com/ramonacirstian/fa_normative_modeling), so that others
302 in the field can easily apply these models to their own data.

303 Finally, we acknowledge some limitations to the current study. The age
304 distribution in our dataset is skewed, with fewer data points at the extremes of the lifespan
305 particularly in young children between the ages of 5 and 8 and adults over 85 years old.
306 Although great care was taken to ensure that this did not bias the analysis (e.g. by ensuring
307 smoothness for the interpolating centile curves), this gap should be considered as limiting
308 the generalizability of the models for younger and older populations at this time. We
309 intend to augment our dataset with additional samples to increase data density in these
310 regions as future high-quality datasets come online. Additionally, while our models
311 effectively account for site-specific differences, variability due to demographic factors like
312 socioeconomic background was not fully explored and should be considered in future
313 normative modeling efforts. A strength of our analysis is that we specifically account for
314 ethnicity in our models, by including self-reported race using fixed effects in the analysis,
315 following our prior work [40]. We consider this important to reduce the risk of racial bias,
316 but it should be remembered that the datasets on which these models were trained on are
317 not representative of the wider population and are themselves biased towards ‘Western
318 Educated, Industrialised, Rich and Democratic’ (WEIRD) populations [40]. Self-reported
319 race is also an imperfect proxy for ethnicity, and it is likely that using more flexible
320 modelling approaches may be needed to properly account for these effects [41]. For these
321 reasons we also release the models that do not include race so that each researcher using
322 these models can decide for themselves which model is more appropriate for their needs.

323
324 In summary, this study provides comprehensive normative reference models for FA across
325 the lifespan, using an extensive dataset that spans infancy to old age. By integrating high-
326 quality diffusion MRI data and using robust modeling techniques, we captured the typical
327 trajectory of white matter development and decline, aligning with prior research and
328 enhancing the field’s understanding of brain aging. Our application of these models to a
329 clinical early psychosis cohort underscores their potential utility in identifying atypical
330 white matter patterns in psychiatric conditions. These models not only serve as a
331 benchmark for individual-level assessments but also offer valuable insights for precision
332 medicine, facilitating more personalized interventions. This study highlights the relevance
333 of normative modeling in neuroimaging, paving the way for its integration into clinical
334 and research settings focused on individual variability in brain structure and pathology.

335 **Materials and Methods**

336 **Data acquisition and processing**

337
338 The construction of the lifespan dataset involved integrating data from five cohorts having
339 high-quality multi-shell diffusion data, i.e.: the HCP Baby [20], HCP Development [21],
340 HCP Young Adult [22], HCP Aging [23] datasets, and the UK Biobank [24]. The
341 demographic information is available in supplementary table 1.

342 The processing of these datasets followed harmonized FSL-based pipelines, summarized
343 in Figure 1A. Initially, pre-processing was performed: B0 intensity normalization,
344 correction for EPI distortions, eddy-current-induced and movement corrections. These
345 corrections were executed using the HCP-pipeline [42] for the HCP datasets while the
346 UKB dataset was already processed according to the UKB documentation [43].
347 Subsequently, we estimated the DTI model using DTIfit on the lowest shell value in order
348 to extract the fractional anisotropy (FA) values. Following this, we ran Tract-Based

Spatial Statistics (TBSS) [44] on the FA images which included registration to a standard space (FMRIB58_FA), projection of each individual's FA image to the standard space skeletonized image (threshold at 0.2) to generate skeletonized FA images for each individual in the same space. Finally, segmentation was conducted using the Johns Hopkins University (JHU) atlas [45]. This process delineated 48 white matter (WM) tracts (listed the supplementary figure 4), for which we computed the mean FA values along the skeleton of each tract.

Normative modeling

To prepare for the modelling stage, we began by splitting the dataset of subjects (N=24,915) into two equal groups: a test set (N=12,457) and a training set (N=12,457), stratified to ensure an even distribution of sex, race, dataset and site. A normative model was then fit to the training set for each white matter tract. The model incorporated several covariates, including sex, age, and dummy coded race, and site. To address potential non-linear effects and non-Gaussian distributions, we employed a warped Bayesian linear regression (BLR) model and used in previous research [4], [25]. This approach involved applying a third-order polynomial B-spline basis expansion over age, with five evenly spaced knots, combined with a SinhArcsinh warping function.

Next, we estimated deviation scores for each subject and white matter tract. In line with our prior work [46] we refit the models after excluding gross outliers having deviations larger than 5 standard deviations from the mean (Figure 1C). Once the models were refit with the cleaned data, we calculated the fit statistics, including explained variance, skew, and kurtosis. The extent of deviation for each subject was visualized by plotting individual z-scores against the mean and centiles of variation predicted by the model. All statistical analyses were conducted using Python version 3.8, with the Predictive Clinical Neuroscience PCN toolkit (GitHub, PCNtoolkit).

Application to a clinical dataset

Next, we applied the model to the Human Connectome Project Early Psychosis (HCP-EP) dataset [18] (Jacobs et al., 2024), which includes multi-shell diffusion data and T1-weighted structural MRI derived from participants diagnosed with early psychosis (n=118) and control participants (n=55). The dataset's demographic distribution comprises 37% females and 63% males, with a racial composition of 58% White, 28% Black, 9% Asian, 1% Mixed, and 3% Other. Participants with early psychosis were diagnosed using the Structured Clinical Interview for DSM-5 (SCID-5) (First et al., 2015) and symptoms assessed with the Positive and Negative Syndrome Scale (PANSS) [27], including negative symptoms (e.g., social withdrawal), positive symptoms (e.g., hallucinations), disorganisation, and general psychopathology. The item-level data were subsequently summarized by the HCP-EP consortium using a standard factor model [47] and the positive, negative and cognitive symptom domain scores were used in addition to the PANSS total score to quantify symptomatology across multiple domains [47]. Medication status was also documented, including antipsychotic type and dosage converted to chlorpromazine equivalents.

The diffusion data were processed with the same pipeline as described above (Figure 1A), and structural data were processed using Freesurfer version 6.0 following similar procedures as we have described previously [4]. Next, we divided this dataset into a training set, consisting of half of the control participants, and combined it with the larger training set described above to retrain the normative models for each white matter tract. Using transfer learning, as in our previous work, we can efficiently adapt the models with only a small amount of calibration data to account for site-specific effect. We then

397 computed z-scores for the patients and remaining controls for the FA data and computed
398 the deviations for cortical thickness and subcortical volumes derived from models we have
399 previously brought online [4]. Note that the splits for this analysis were matched so that
400 the same participants were in the training and test sets for diffusion and structural
401 measures at each iteration.

402 We next assessed the mean difference of the deviations between patients and controls for
403 each tract using a t-test with false discovery rate (FDR) correction for multiple testing
404 [26]. We then tested whether the proportion of extreme deviations differ between groups
405 for each tract. To achieve this, we calculated the percentage of participants falling below
406 and above the threshold in each of the 48 tracts. To achieve this, we set a z-score threshold
407 between -2.6 and 2.6, which correspond to a p-value of 0.01 as in prior work to identify
408 extreme deviations then employed a non-parametric Mann-Whitney U test [48], again
409 followed by FDR correction for multiple comparisons. This stringent threshold enhances
410 the detection of significant deviations while controlling for false positives

411 Next, we combined the FA deviations with structural deviations in a multimodal analysis
412 aiming to predict the four symptom domains described above. To achieve this, we
413 conducted a multi-view sparse canonical correlation analysis (msCCA), using an approach
414 we have described previously [28]. To identify relationships between multiple datasets,
415 msCCA maximises the cross-correlation between weighted sums of variables from each
416 dataset (Equation 1).

$$417 \quad \max_{\mathbf{w}_1, \dots, \mathbf{w}_2} \mathbf{w}_1^T X_1^T \sum_{i=2}^m X_i \mathbf{w}_i \quad (\text{Equation 1})$$

418 where X_1 represents psychiatric symptoms and X_2 to X_m represent neuroimaging
419 measures from m different modalities. The weights ($\mathbf{w}_1, \mathbf{w}_2, \dots, \mathbf{w}_m$) are subject to
420 constraints: $\|\mathbf{w}_1\|_2 = 1, \|\mathbf{w}_i\|_2 = 1, \|\mathbf{w}_1\|_1 \leq c_1, \|\mathbf{w}_i\|_1 \leq c_i$, ensuring sparsity and
421 interpretability. The regularization parameters for each view v are assumed to be set such
422 that $0 < c_v < \sqrt{p_v}$ where p_v is the number of feature in view v , which is helpful to bound
423 the (approximate) number of selected variables [28]. Importantly, this approach avoids
424 optimising correlations between different neuroimaging modalities directly, focusing
425 instead on shared variance between psychiatric symptoms and neuroimaging measures.

426 This involved creating three views of the data (i.e. symptoms, structural deviations
427 and diffusion deviations) and then fitting an msCCA model to maximise the association
428 between symptom domains and each of the diffusion and structural deviations but –
429 crucially– not the imaging views with one another [28]. This requires setting (L1-norm)
430 sparsity parameters for each of the data views c_v . These were fixed throughout such that
431 approximately 90% of the PANSS features were selected and 20% of the FA and
432 structural image features. This corresponds respectively to light regularization for the
433 symptoms, and moderately high regularization for the FA and structural measures. Note
434 that we deliberately choose fixed parameters rather than optimizing them via nested cross-
435 validation given the moderate sample size for the clinical data. Instead, we employed
436 stability selection to assess the generalizability of the coefficients, which is theoretically
437 guaranteed to provide tight type-I family-wise error control [29].

438 In more detail, we performed 1000 random splits of the dataset into a training
439 (70%) and test set (30%) and selected the most stable features, i.e. features that were
440 selected in more than 80% of the splits. This threshold is justified as it is sufficiently high
441 that the theoretical guarantees on controlling the type 1 error rate become operative. In
442 order to assess generalizability, we then ran an additional 1000 permutations, where
443 within each permutation, we computed the test canonical correlation averaged across 10
444 random splits of the data, both before and after randomly permuting the order of the
445 PANSS data view to destroy the relationship between the symptom scores and imaging
446 data. We did this for the first three canonical components, which were derived by

447 successively applying projection deflation to the data matrices [28], [49]. In order to
448 compute significance, we then counted the number of times the true mean test canonical
449 correlation exceeded the permuted value and divided by the number of permutations.

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602

603 **Competing interests:**

604 CB is director and shareholder for SBGneuro
605

606 **Data and materials availability:**

607 The data used in the present study is part of the UK Biobank dataset which is available to
608 be downloaded upon completing an access application. More information can be found on
609 the dedicated webpage (UK Biobank, n.d.). The code used to process the data and train the
610 normative models is also available online on GitHub
611 (https://github.com/ramonacirstian/fa_normative_modeling n.d.)