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Time-related trends in variability of cIMT changes in statin trials



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ABSTRACT

This brief article provides complementary data supporting the results reported in "Changing Characteristics of Statin-related cIMT Trials from 1988 to 2006" [1]. That article described time-related trends in baseline factors and study characteristics that may have influenced the variability of carotid intima media thickness (cIMT) endpoints (mean of mean and maximum common carotid artery [CCA]/cIMT) in published statin trials. In this brief report, additional details for the studies included in the analysis, and further supporting data, including mean of the maximum CCA/cIMT changes and subgroup data (mean and maximum CCA/cIMT) are provided. For the analysis, study-level data was extracted from 17 statin cIMT trials conducted during 1988-2006, selected on the basis of having at least one statin monotherapy arm in the absence of mixed therapy, and baseline- and study-end values for mean mean and mean maximum CCA/cIMT endpoints. The baseline mean CCA/cIMT, maximum mean CCA/cIMT and LDL-C levels, and annualized cIMT changes were estimated for the overall studies, those conducted before/after 2000, and in risk-based subgroups. Interestingly, all 8 studies conducted before 2000 were significant for cIMT change in which patients did not receive prior LLT; whereas after 2000, the results were more variable and in 4 of 6 trials that did not show a significant cIMT change, patients had received prior treatment. Baseline mean maximum cIMT and LDL-C levels, and annualized changes in studies conducted before 2000 were higher than those

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conducted after 2000, similar to the results reported in the original article for the mean mean cIMT endpoint. These findings were consistent across study populations of patients with CHD risk versus those without, and in studies with greater LDL-C reductions and with thickened baseline cIMT at study entry for both mean and maximum cIMT changes. Taken together, these results are consistent with trends in recent years toward greater use of lipidlowering therapy and control of LDL-C that may have impacted the variability in the results of cIMT studies.

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Specifications table

| Subject area | Vascular Biology |
|----------------------------|--|
| More specific subject area | Carotid artery imaging |
| Type of data | Tables and figures |
| How data was acquired | Study-level data from published statin cIMT trials |
| Data format | Aggregate study-level data |
| Experimental factors | Brief description of any pretreatment of samples |
| Experimental features | Very brief experimental description |
| Data source location | Published cIMT trials as referenced |
| Data accessibility | Data are supplied with this article |

Value of the data

- The relationship between baseline and study characteristics and the variability of cIMT change in statin cIMT trials over time has not been well-studied.
- These data show time-related trends in baseline and study characteristics that may influence the variability of mean mean and mean max cIMT results in statin clinical trials.
- Statin cIMT trials conducted before 2000 had greater cIMT baseline values, LDL-C reduction and annualized cIMT changes than studies after 2000.
- These trends are consistent with increased statin treatment and management of LDL-C lowering over recent years, bringing into question the utility of cIMT as a surrogate marker in clinical trials given the high standard of background statin therapy.
- These results highlight the need to consider such factors in the design of future cIMT trials.

1. Data

The predictive power of cIMT as a surrogate marker for the assessment of CVD risk reduction in lipid-lowering trials may be limited in some settings due to differences in study design, cIMT methods, and patient characteristics [2–4]. With increased use of lipid-lowering therapy over recent years, it is possible that patient characteristics and study designs of cIMT trials have also changed in a manner which could influence the utility of cIMT in assessing LLT treatment effects. This analysis was undertaken to better understand how baseline factors and study characteristics may influence the variability of cIMT changes in intervention trials.

A summary of the cIMT trials included in the analyses reported in the original article [1] and in this brief report are summarized in Table 1. All 8 of the studies that were conducted before year 2000 were significantly positive for cIMT change, and the patients in those trials did not receive prior lipid-lowering therapy. However, the 9 studies conducted after 2000 were more variable for cIMT changes

Table 1Summary of the included trials.

| Study Treatment (Reference) (mg)*** | | Age (yr) | Female (%) | Population | Prior treatment | Study length (vr) | Study (yr) | Baseline IMT (mm) | thickness | Baselin | e mg/dl (r | nmol/l) | |
|--|--------------------------|----------------|------------------|------------|--|--|---------------|----------------------|--|--|---------------|----------------|---------------|
| | Arn | n n | | | | | (91) | | Mean CCA/cIMT | Mean max CCA/cIMT | LDL-C | HDL-C | TG |
| Statin studie | s | | | | | | | | | | | | |
| ACAPS** [9] | L20-40 Pbo | 23 23 | l 62) | 49 | Asymptomatic CAD; LDL-C 60–90th percentiles; IMT \ge 1.5 - < 3.5 mm | No prior LLT for 1 yr | 3 | 1989– 1990 | $L \pm W = 1.15$ Pbo = 1.14 | $L \pm W = 1.33$ Pbo = 1.32 | 155 (4.01) | 45.8 (1.19) | 140 (1.58) |
| ARBITER I** [20] | P40 A80 | 70 68 | 60 | 29 | Adults (> 18 yr), ATPII lipid-lowering criteria | No prior LLT | 1 | 1999– 2001 | P = 0.615 A = 0.625 | <i>P</i> =0.808 <i>A</i> =0.935 | 152 (3.94) | 49 (1.27) | 207 (2.34) |
| ASAP** [19] | A80 S40 | 160 163 |) 48 3 | 61 | hFH; mean CCA \geq 0.7 mm and/or \geq 0.9 mm in bulb | Untreated or treated ≤ 1 yr, LDL-C > 173 mg/dL (4.48 mmol/l) | 2 | 1997– 1998 | A=0.86 S=0.87 | NA | 315 (8.16) | 46 (1.19) | 165 (1.86) |
| BCAPS*** [10] | Met F40 Pbo | 20 20 20 | 0 61.8 0 0 | 54.5 | Asymptomatic plaque (focal IMT > 1.2 mm) in right coronary artery | No beta blocker or statin | 3 | 1991– 1994 | Met $=$ 0.92 F $=$ 0.886 Pbo $=$ 0.898 | Bulb: Met = 1.936 F = 1.893 Pbo = 1.875 | 158 (4.09) | 53.2 (1.38) | 104 (1.18) |
| CAIUS** [15] | P40 Pbo | 151 154 | 55 1 | 47 | Moderate HC, LDL-C 150– 250 mg/dL (3.88– 6.47 mmol/l), ≥ 1 cIMT | No prior LLT | \leq 3 yr | 1991– 1995 | NA | P = 1.06 Pbo = 1.04 | 181 (4.69) | 52.5 (1.36) | 138 (1.56) |
| CAPTIVATE* | Pact 100 | 44 | 3 55 | 39 | hesion 1.3–3.5 mm hFH, LDL-C $> 100 \text{ mg/dL}$ ($> 2.6 \text{ mmol/l}$); max | Statin Rx, $\langle or \rangle$ 24 mon | 1 | 2004– 2005 | Pac=0.785 Pbo=0.775 | Pac=0.937 Pbo=0.927 | 140 (3.63) | 52 (1.35) | 130 (1.47) |
| CASHMERE* | PD0 A80 | 433 | s 2 57 | 100 | CIMT > 0.7 - \leq 2.5 mm Postmenopause women \leq 70 yr; LDL-C 130- | No statins > 3 mon w/ in last yr, or no LLT w/ | 1 | 2003– 2006 | A=0.699 | NA | 159 (4.12) | 66.6 (1.72) | 120 (1.36) |
| [17] | Pbo | 20 | 5 | | 190 mg/dL (3.36– 4 86 mmol/l) | in last 6 wks | | | Pbo=0.683 | | | | |
| KAPS** | P40 | 212 | 2 57 | 0 | LDL-C 164 mg/dL (\geq 4.25 mmol/l), BMI | No prior LLT | 3 | 1989– 1990 | 1.35 (overall) | NA | 189 (4.90) | 46 (1.19) | 150 (1.70) |
| [18] | Pbo | 212 | 2 | | \leq 32 kg/m ² | | | | | | | | |

| LIPID** [13] | P40 | 273 249 | 61 | 12 | MI or hospitalization for unstable angina [†] | 2 mon run-in diet, no prior LLT | 4 | 1990– 1992 | P=0.804 Pbo=0.786 | NA | 154 (3.99) | 34.7 (0.90) | 151 (1.71) |
|---------------------|--|--------------------------------|-----------|----|--|--|---|---------------|--------------------------|----------------------------------|---------------|----------------|---------------|
| METEOR** [7] | R40 | 702 | 61 | 12 | Asympotmatic, moderately elevated cholesterol with | No LLT 12 mon prior | 2 | 2002– 2004 | R = 0.76 Pbo=0.76 | R = 1.15 Pbo = 1.17 | 165 (4.27) | 50 (1.30) | 130 (1.47) |
| | PDO | 282 | C1 | 10 | low CVD risk per AIPIII* | No. and an LUT | 2 | 1000 | D 101 | D 100 | 100 | 41 | 171 |
| [6] | P20→40 ⁻ Pbo | 75 76 | 01 | 12 | centile, ≥ 1 IMT lesion ≥ 1.3 mm | NO PHOT LLI | 3 | 1988– 1990 | P = 1.01 Pbo = 1.01 | P = 1.32 Pbo = 1.32 | (4.30) | (1.06) | (1.93) |
| RADIANCE I* | A10, 20, 40, 80 | 454 | 61 | 12 | hFH | Aspirin 30%; beta blocker 20%; ACE inh | 2 | 2003– 2004 | A=0.72 A+T=0.71 | A = 1.15 A + T = 1.09 | 139 (3.60) | 53 (1.37) | 97 (1.10) |
| RADIANCE II* | A10, 20, 40, 80 | 430 375 | 57 | 37 | Mixed HL eligible for statin Rx by ATPIII; max | Aspirin 55%, beta blocker 24–29%; ACE | 2 | 2003– 2006 | A=0.83 A+T=0.83 | A = 1.3 A + T = 1.32 | 101 (2.62) | 48 (1.24) | 167 (1.89) |
| [5] | A+T60 | 377 | | | $IMT = 1.2 - 3.5 \text{ mm HDL-C}$ $\leq 1.6 \text{ mmol/l}^{\$}$ | inh or ARB 37–39%; no LLT ≥ 4 wks | | | | | | | |
| REGRESS** | P40 | 131 | 56 | 0 | CAD, symptomatic, cor- onary angiogram $\geq 50\%$ | No LLT ≤ 6 wks (≤ 12 wks for fibrates) | 2 | 1989– 1991 | P=0.87 Pbo=0.86 | P = 1.08 Pbo = 1.07 | 168 (4.35) | 38 (0.98) | 163 (1.84) |
| [8] | Pbo | 124 | | | reduction in ≥ 1 major coronary artery [¶] | | | | | | | | |
| Studies with | ı statin arms | | | | | | | | | | | | |
| ENHANCE* [12] | S80 S80+E10 | 342 338 | 46 | 51 | hFH | Untreated, LDL-C > 210 mg/dL 5.44 mmol/l), 81% prior statin use | 2 | 2002– 2006 | S = 0.68 S + E = 0.68 | S = 0.8 S + E = 0.8 | 319 (8.26) | 47 (1.22) | 159 (1.80) |
| VYCTOR** [14] | P40 S40 S20+E10 ^b | ≥ 18 ≥ 18 > 18 | 58 | 57 | 10 yr absolute risk for CHD or MI \geq 20 per ATPIII | Low dose statins, none received E previously | 1 | 2005 | NA | P = 1.33 S = 1.3 S + E = 1.23 | 130 (3.37) | 45.3 (1.17) | 192 (2.17) |
| ARBITER II* [21] | ERN 1000 Pbo | 87 80 | 67 | 9 | Known CVD, LDL-C < 130 mg/dL (< 3.37 mmol/l) and HDL- C < 45 mg/dL (< 1.17 mmol/l) | All patients on statins (93% S) | 1 | 2001– 2003 | ERN=0.89 Pbo=0.87 | NA | 89 (2.31) | 40 (1.04) | 163 (1.84) |

A=atorvastatin; ACE=angiotensin-converting enzyme; ARB=angiotensin-receptor blocker; CAD=coronary artery disease; CHD=coronary heart disease; CVD=cardiovascular disease; E=ezetimibe; hFH= heterozygous familial hypercholesterolemia L=lovastatin; Met=metropolo; P=pravastatin; Pac=pactimibe; Pbo=placebo; S=simvastatin; R=rosuvastatin; T=torcetrapib, W= warfarin.

*** Bold text denotes treatment arm in analysis.

** Study was statistically significant (p < 0.05) for primary endpoint.

* Study was non-significant (p > 0.05) for primary endpoint.

[†] MI or hospitalization for unstable angina 3 mon-5 yr before enrollment and a total serum cholesterol 4–7 mmol/L (155–271 mg/dL) after 2 mon on low-fat diet; 88% men, 75% infarction, 5% diabetics.

 $^{\pm}$ LDL-C 120–190 mg/dL (3.10–4.92 mmol/l) w/0–1 CHD risk factors or 120–160 mg/dL (3.10–4.14 mmol/l) w/ \geq 2 CHD risk factors and 10 yr risk < 10%, HDL-C \leq 60 mg/dL (1.55 mmol/l); max IMT = 1.2–3.5 mm.

 $\frac{100}{100}$ LDL-C \leq 160 - > 220 mg/dL (4.1 mmol/L) w/10 yr CHD risk < 10%; LDL-C \geq 130 - < 190mg/dL (3.37-4.92 mmol/L) w/10 yr CHD risk \geq 10% and \leq 20%, LDL-C \geq 115 - < 190 mg/dL (2.98-4.92 mmol/L) w/10 yr CHD risk > 20%, and TG \geq 150 - \leq 500 mg/dL (1.70-5.65 mmol/L).

[¶] Cholesterol 155–310 mg/dl (4.01–8.03 mmol/l).

^a P20 initially, change to 40 mg based on LDL-C.

^b S20+E10 initially, change to S40+E based on LDL-C.

Table 2

Pooled baseline mean maximum CCA/cIMT and annualized rates of change.

| | Studies | Patients | Baseline mean max CCA/cIMT $[mm]^{\dagger}$ | Annualized change mean max CCA/cIMT [mm/yr] $^{\dagger \ddagger}$ | |
|-------------------------------------|---------|----------|---|---|--------------------|
| | # | Ν | Mean (SE) | Mean (SE) | 95% CI |
| Overall Study year [§] | 12 | 5051 | 1.0866 (0.0032) | -0.0023 (0.0014) | (-0.0050, -0.0004) |
| Before 2000 | 5 | 1955 | 1.2550 (0.0059) | -0.0134 (0.0020) | (-0.0173, -0.0096) |
| After 2000 | 7 | 3096 | 1.0181 (0.0038) | 0.0083 (0.0019) | (-0.0046, 0.0121) |
| CHD risk | | | | | (|
| Yes | 4 | 1242 | 1.3022 (0.0071) | 0.0021 (0.0030) | (-0.0038, 0.0081) |
| No | 8 | 3809 | 1.0317 (0.0036) | -0.0034 (0.0015) | (-0.0065, -0.0004) |
| hFH | | | | | |
| Yes | 3 | 1677 | 0.9130 (0.0050) | 0.0053 (0.0024) | (0.0005, 0.0100) |
| No | 9 | 3374 | 1.2032 (0.0041) | -0.0059(0.0017) | (-0.0091, -0.0026) |
| Thickened baseline IMT [¶] | | | | | |
| Yes | 3 | 1964 | 1.1703 (0.0052) | -0.0135 (0.0020) | (-0.0175, -0.0095) |
| No | 9 | 3087 | 1.0375 (0.0040) | 0.0074 (0.0019) | (-0.0037, 0.0111) |
| LDL-C reduction | | | | | |
| < -27.6% | 6 | 3259 | 1.1335 (0.0042) | -0.0036 (0.0016) | (-0.0066, -0.0005) |
| $\geq -27.6\%$ | 6 | 1792 | 1.0200 (0.0050) | 0.0025 (0.0030) | (-0.0034, -0.0083) |
| Mean age (yr) | | | | | |
| < 57 | 5 | 2237 | 0.9454 (0.0044) | -0.0045 (0.0017) | (-0.0077, -0.0012) |
| ≥ 57 | 7 | 2814 | 1.2423 (0.0046) | 0.0023 (0.0024) | (-0.0024, 0.0071) |
| Female (%) | | | | | |
| < 40 | 5 | 1800 | 0.9853 (0.0058) | 0.0204 (0.0043) | (0.0119, 0.0289) |
| \geq 40 | 7 | 3251 | 1.1299 (0.0038) | -0.0048 (0.0015) | (-0.0077, -0.0020) |

hFH=heterozygous familial hypercholesterolemia.

* Studies included: ACAPS, ARBITER I, BCAPS, CAIUS, CAPTIVATE, ENHANCE simvastatin, REGRESS, METEOR, PLAC II, VYCTOR, and RADIANCE I & II (atorvastatin arms). The PLAC II study was excluded from the baseline analysis because of missing baseline standard error information.

[†] Pooled estimates for subgroups are weighted based on the inverse of the square of the standard error for the individual studies.

[‡] SE and 95%CI are based on study-end values for the treatment difference (placebo-controlled) or change from baseline (monotherapy arms) in cIMT.

[§] Studies conducted before 2000: ACAPS, BCAPS, CAIUS, PLAC II, REGRESS. Studies conducted after 2000: ARBITER I, CAPTIVATE, ENHANCE, METEOR, VYCTOR, and RADIANCE I & II (atorvastatinarms)...

[¶] Entry criteria.

with 6 of these not showing significant cIMT change, the majority of which (4 of 6 trials) included patients that had received prior lipid-lowering therapy. The baseline mean mean and mean maximum cIMT and LDL-C levels were also generally greater for those studies conducted before 2000 than after 2000. The durations of the trials were also longer (≥ 2 years) in those conducted before 2000.

Table 2 summarizes the mean maximum baseline CCA/cIMT level and annualized changes for the overall studies (1988–2006) and those conducted before and after 2000, as well as for various study populations. For the overall studies combined (1988–2006), the baseline mean maximum CCA/cIMT level was 1.0866 mm and the annualized rate of change for the mean maximum CCA/cIMT from baseline was -0.0023 mm/yr (95% CI: -0.0050, -0004), indicating a statistically significant annualized change in cIMT. Baseline mean maximum cIMT levels were higher in populations of studies conducted before 2000 (1.2550 mm) than after 2000 (1.0181 mm), as well as in study populations of patients with CHD risk versus those without, and in those studies with greater LDL-C reductions, and those with thickened baseline cIMT. Annualized rates of change in maximum CCA/ cIMT levels were also larger for the combined study populations conducted before year 2000 (-0.0134 mm/yr) than after 2000 (-0.0083 mm/yr), and for those studies in which patients had thickened baseline cIMT at study entry and also those which showed greater LDL-C reductions.

The baseline mean CCA/cIMT and annualized changes observed in pooled studies conducted before and after 2000 for various prespecified subgroups of interest are displayed in Table 3. Consistent with the findings in the overall analysis cohort, baseline cIMT levels and annualized change for mean CCA/ cIMT were generally greater in all subgroups assessed for the combined studies conducted before 2000 than after 2000. Similarly, mean max CCA/cIMT and annualized change in studies were also generally greater in all subgroups assessed in studies conducted before 2000 than after 2000 (Table 4).

2. Experimental design, materials and methods

2.1. Study design

This was an exploratory analysis of study-level data from statin treatment arms of published cIMT imaging trials, as previously described in the original study article [1]. Fig. 1 displays the method of trial selection for the analysis. Following a detailed review of cIMT trials in the literature, 24 statin trials were identified that had both baseline- and study-end measurements for mean of the mean CCA/cIMT (mean of all mean measurements on CCA or a single mean cIMT value when not available), and/or mean of the maximum CCA/cIMT (mean of all maximum mean measurements on CCA, or single maximum mean or bulb values when not available). Of these, 7 trials were excluded due to insufficient data or mixed lipid-lowering therapy in the statin arm, and 17 studies conducted during 1988–2006 were selected on the basis of having at least one statin monotherapy arm in the absence of mixed therapy, baseline- and study-end values for mean of the mean and mean of the maximum CCA/cIMT. The selected studies included a mix of placebo (ACAPS, ARBITER II, BCAPS, CAIUS, CAPTI-VATE, CASHMERE, KAPS, LIPID, METEOR, PLAC II, REGRESS) and active-controlled (ARBITER I, ASAP, RADIANCE I, RADIANCE II, ENHANCE, VYCTOR) studies (Table 1 and Fig. 1) [5–21]. There were 13 trials with mean mean CCA/cIMT as the primary endpoint and 12 with mean maximum CCA/cIMT end-points; these were assessed in 2 separate, but similar analyses.

2.2. Data extraction and analysis

Only aggregate data in the published literature were collected from the studies. Demographic information (e.g., age and gender), study characteristics (study dates, CHD risk, heterozygous familial hypercholesterolemia [FH], baseline LDL-C, HDL-C and TG levels), and baseline- and study-end measurements for mean mean CCA/cIMT and mean max CCA/cIMT, and variability estimates were extracted from the publications. Baseline means of the mean CCA/cIMT and maximum CCA/cIMT levels were summarized for the overall population of studies, those conducted before and after 2000 (the median of enrollment dates for all studies), and studies that enrolled patients based on CHD risk,

Table 3 Pooled baseline mean mean CCA/cIMT and annualized rates of change in studies before and after 2000 for subgroups.

| | Studies I | before 200 |)0* | | | Studies after 2000** | | | | | | |
|----------------|------------|-------------------|---|--|--------------------|----------------------|----------|--------------------------------|---|--------------------|--|--|
| | Studies | Patients | Baseline mean CCA/cIMT [mm] [†] | Annualized change mean CCA/cIMT Str [mm/yr] ^{†‡} | | Studies | Patients | Baseline mean CCA/cIMT [mm] | Annualized chang [mm/yr] ^{†‡} | e mean CCA/cIMT | | |
| | # | Ν | Mean (SE) | Mean (SE) | 95% CI | # | Ν | Mean (SE) | Mean (SE) | 95% CI | | |
| CHD risk | | | | | | | | | | | | |
| Yes | 3 | 1201 | 0.8008 (0.0064) | -0.0184 (0.0051) | (-0.0283, -0.0084) | 1 | 375 | 0.8300 (0.0072) | 0.0080 (0.0020) | (0.0041, 0.0119) | | |
| No | 2 | 1064 | 0.8898 (0.0054) | -0.0098 (0.0029) | (-0.0155, -0.0042) | 7 | 3131 | 0.7367 (0.0024) | -0.0033 (0.0009) | (-0.0051, -0.0015) | | |
| hFH | | | | | | | | | | | | |
| Yes | 1 | 281 | 0.8645 (0.0094) | -0.0138 (0.0069) | (-0.0273, -0.0003) | 3 | 1677 | 0.7435 (0.0036) | -0.0005 (0.0012) | (-0.0113, -0.0057) | | |
| No | 4 | 1984 | 0.8492 (0.0046) | -0.0117 (0.0027) | (-0.0170, -0.0064) | 5 | 1829 | 0.7473 (0.0029) | -0.0021 (0.0012) | (0.0003, 0.0043) | | |
| Thickened ba | seline cIN | 1T <mark>§</mark> | | | | | | | | | | |
| Yes | 1 | 783 | 0.9026 (0.0067) | -0.0090 (0.0032) | (-0.0153, -0.0027) | 4 | 1781 | 0.7782 (0.0036) | -0.0085 (0.0015) | (-0.0113, -0.0057) | | |
| No | 4 | 1482 | 0.8207 (0.0053) | -0.0167(0.0041) | (-0.0247, -0.0087) | 4 | 1725 | 0.7242 (0.0029) | 0.0023 (0.0010) | (0.0003, 0.0043) | | |
| LDL-C reducti | ion (medi | an) | | | | | | | | | | |
| <-27.6% | 2 | 1305 | 0.8484 (0.0047) | -0.0096 (0.0030) | (-0.0156, -0.0036) | 4 | 1781 | 0.7782 (0.0036) | 0.0021 (0.0011) | (0.0000, 0.0042) | | |
| $\geq -27.6\%$ | 3 | 960 | 0.8645 (0.0086) | -0.0170 (0.0045) | (-0.0258, -0.0082) | 4 | 1725 | 0.7242 (0.0029) | -0.0066 (0.0013) | (-0.0092, -0.0040) | | |
| Mean age (yr | .) | | | | | | | | | | | |
| < 57 | 2 | 536 | 0.8645 (0.0086) | -0.0148 (0.0066) | (-0.0277, -0.0019) | 3 | 1677 | 0.7435 (0.0036) | -0.0005 (0.0012) | (-0.0028, 0.0019) | | |
| ≥ 57 | 3 | 1729 | 0.8484 (0.0047) | -0.0115 (0.0027) | (-0.0168, -0.0061) | 5 | 1829 | 0.7473 (0.0029) | -0.0021 (0.0012) | (-0.0043, 0.0002) | | |
| Female (%) | | | | | | | | | | | | |
| < 40 | 3 | 1201 | 0.8008 (0.0064) | -0.0184 (0.0051) | (-0.0283, -0.0084) | 4 | 1465 | 0.7835 (0.0040) | 0.0094 (0.0019) | (0.0057, 0.0130) | | |
| \geq 40 | 2 | 1064 | 0.8898 (0.0054) | -0.0098 (0.0029) | (-0.0155, -0.0042) | 4 | 2041 | 0.7277 (0.0027) | -0.0040 (0.0009) | (-0.0058, -0.0021) | | |
| | | | | | | | | | | | | |

hFH=heterozygous familial hypercholesterolemia; CHD=coronary heart disease; IMT=statins.

* Studies before 2000 included REGRESS, ASAP, LIPID, and BCAPS, KAPS study was excluded from the baseline analysis due to missing baseline standard error information.

** Studies after 2000 included ARBITER I, METEOR, CASHMERE, CAPTIVATE, ENHANCE (simvastatin arm), ARBITER II (placebo arm), and Radiance I & II (atorvastatin arm).

[†] Pooled estimates for subgroups are weighted based on the inverse of the square of the standard error for the individual studies.

¹ SE and 95%CI are based on study-end values for the treatment difference (placebo-controlled) or change from baseline (monotherapy arms) in cIMT.

[§] Entry criteria.

Table 4

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Pooled baseline mean maximum CCA/cIMT and annualized rates of change in studies before and after 2000[°] for subgroups.

| StudiesPatientsBaseline max CCA/ $CIMT [mm]^{\uparrow}$ Annualized change max CCA/cIMT $[mm/yr]^{\uparrow\uparrow}$ StudiesPatientsBaseline max CCA/ $CIMT [mm]$ Annualized change max CCA/cIMT $[mm/yr]^{\uparrow\uparrow}$ Mean (SE)95% CIMean (SE)Mean (SE)95% CICHD risk Yes38671.3028 (0.0080) $-0.0140 (0.0038) (-0.0215, -0.0066)$ 13751.3000 (0.0150)0.0300 (0.0050) (0.0202, 0.0398)No210881.1975 (0.0088) $-0.0132 (0.0023) (-0.0177, -0.0087)$ 627210.9989 (0.0039)0.0046 (0.0021) (0.0005, 0.0087) | |
|--|------|
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| 111 11 | |
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| No 5 1955 1.2550 (0.0059) -0.0134 (0.0020) (-0.0173, -0.0096) 4 1419 1.1547 (0.0057) 0.0135 (0.0032) (0.0074, 0.0197) | |
| Thickened baseline cIMT [§] | |
| Yes 2 1088 1.1975 (0.0088) -0.0132 (0.0023) (-0.0177, -0.0087) 1 876 1.1553 (0.0065) -0.0145 (0.0042) (-0.0228, -0.0 | 062) |
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| LDL-C reduction (median) | |
| < -27.6% 3 1549 1.2642 (0.0061) -0.0136 (0.0020) (0.0176, -0.0097) 3 1710 1.0180 (0.0057) 0.0112 (0.0024) (0.0064, 0.0160) | |
| $\geq -27.6\% \ 2 \qquad 406 \qquad 1.0747 \ (0.0269) \qquad -0.0088 \ (0.0095) \ (-0.0275, \ 0.0099) \ 4 \qquad 1386 \qquad 1.0181 \ (0.0050) \qquad 0.0037 \ (0.0031) \ (-0.0025, \ 0.0098) \ (-0.0025, \ 0.008) \ (-0.002$ | 8) |
| Mean age (yr) | |
| < 57 2 560 1.0515 (0.0091) -0.0132 (0.0023) (-0.0177, -0.0087) 3 1677 0.9130 (0.0050) 0.0053 (0.0024) (0.0005, 0.0100) | |
| $\geq 57 \qquad 3 \qquad 1395 \qquad 1.4045 \ (0.0078) \qquad -0.0139 \ (0.0038) \ (-0.0214, -0.0065) \ 4 \qquad 1419 \qquad 1.1547 \ (0.0057) \qquad 0.0135 \ (0.0031) \ (0.0074, \ 0.0197) \ 1.4045 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0031) \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ (0.0074, \ 0.0197) \ 0.0135 \ ($ | |
| Female (%) | |
| < 40 2 406 1.0747 (0.0269) -0.0088 (0.0095) (-0.0275, 0.0099) 3 1394 0.9809 (0.0060) 0.0280 (0.0049) (0.0185, 0.0375) | |
| $\geq 40 \qquad 3 \qquad 1549 \qquad 1.2642 \ (0.0061) \qquad -0.0136 \ (0.0020) \ (-0.0176, -0.0097) \ 4 \qquad 1702 \qquad 1.0430 \ (0.0049) \qquad 0.0047 \ (0.0021) \ (0.0006, 0.0088) \ -0.0136 \ (0.0021) \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ (0.006, 0.0088) \ -0.0136 \ -0$ | i - |

hFH=heterozygous familial hypercholesterolemia; CHD=coronary heart disease; cIMT=carotid intima media thickness.

* Studies conducted before 2000: ACAPS, BCAPS, CAIUS, PLAC II, REGRESS. ** Studies conducted after 2000: ARBITER I, CAPTIVATE, ENHANCE, METEOR, VYCTOR, and RADIANCE I & II (atorvastatin arms).

[†] Pooled estimates for subgroups are weighted based on the inverse of the square of the standard error for the individual studies.

[‡] SE and 95%CI are based on study-end values for the treatment difference (placebo-controlled) or change from baseline (monotherapy arms) in cIMT.

§ Entry criteria.



Fig. 1. Study selection flow chart.

heterozygous FH, and thickened baseline cIMT. Subgroups of patients categorized by LDL-C reduction, age and gender were also assessed. Dichotomized variables for the subgroups were prespecified based on median values for the combined studies.

Studies were combined using a meta-analytic approach. The 13 trials included in the mean mean CCA/cIMT analysis were: ARBITER I, REGRESS, ASAP, KAPS, LIPID, METEOR, BCAPS, CASHMERE, CAP-TIVATE, ENHANCE (simvastatin), ARBITER II (placebo), and Radiance I & II (atorvastatin). The 12 studies included in the analysis of mean maximum CCA/cIMT were: ACAPS, ARBITER I, BCAPS, CAIUS, CAPTIVATE, ENHANCE (simvastatin), REGRESS, METEOR, PLAC II, VYCTOR, and RADIANCE I & II (atorvastatin).

Study-end values were reported as a change from baseline in primary variables (mean mean or mean max CCA/cIMT) indicating regression and/or progression of cIMT. Annualized rates of change (regression and/or progression) were extracted as treatment differences for placebo-controlled studies and as change from baseline in each individual treatment arm for active-control studies. For active-controlled studies with 2 or more monotherapy arms, each treatment arm was included in the analysis as a separate data point. For trials in which annualized rates of change was not provided, it was calculated by dividing the reported change in cIMT by the duration of the study in years.

Estimates of variance for the baseline and study-end values for mean mean and mean max cIMT were extracted from the publications. In most cases the standard deviation was reported for baseline values and the standard error (SE) was reported for study-end values. In some instances the SE was not reported, and it was calculated from the reported 95% CIs or *p*-values for the treatment difference (placebo-controlled) or change from baseline (monotherapy arms) in cIMT. The rate of cIMT change was standardized (annualized) for this analysis; however, the SE reported in the publication was used regardless of study duration. An overall pooled estimate for the baseline CCA/ cIMT value and annualized rate of cIMT change was calculated by weighting each individual study by the inverse of the square of its standard error which was extracted from the publications. This resulted in studies with small variability having more weight in the pooled values than studies with large variability. Overall pooled estimates for the baseline CCA/cIMT value and annualized rate of max cIMT, were calculated using a fixed-effects model, by weighting each individual study by the inverse of the square of max cIMT, were calculated using a fixed-effects with small variability having with small variability having each individual study by the inverse of the square of its standard error; thus, studies with small variability having each individual study by the inverse of the square of its standard error; thus, studies with small variability having each individual study by the inverse of the square of its standard error; thus, studies with small variability had more weight in the pooled values than studies with large variability.

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References

- M.H. Davidson, J.E. Tomassini, E. Jensen, D. Neff, A.B. Polis, A.M. Tershakovec, Changing characteristics of statin-related cIMT trials from 1988 to 2006, Atherosclerosis (2015), http://dx.doi.org/10.1016/j.atherosclerosis.2015.11.023.
- [2] A.L. Gould, J. Koglin, R.P. Bain, et al., Effects of sources of variability on sample sizes required for RCTs, applied to trials of lipid-altering therapies on carotid artery intima-media thickness, Clin. Trials 6 (2009) 305–319.
- [3] J.J. Kastelein, E. de Groot, R. Sankatsing, Atherosclerosis measured by B-mode ultrasonography: effect of statin therapy on disease progression, Am. J. Med. 116 (Suppl. 6A) (2004) S31–S36.
- [4] D.H. O'Leary, J.F. Polak, Intima-media thickness: a tool for atherosclerosis imaging and event prediction, Am. J. Cardiol. 90 (2002) 18L-21L.
- [5] M.L. Bots, F.L. Visseren, G.W. Evans, et al., Torcetrapib and carotid intima-media thickness in mixed dyslipidaemia (RADI-ANCE 2 study): a randomised, double-blind trial, Lancet 370 (2007) 153–160.
- [6] J.R. Crouse III, R.P. Byington, M.G. Bond, et al., Pravastatin, lipids, and atherosclerosis in the carotid arteries (PLAC-II), Am. J. Cardiol. 75 (1995) 455–459.
- [7] J.R. Crouse III, J.S. Raichlen, W.A. Riley, et al., Effect of rosuvastatin on progression of carotid intima-media thickness in lowrisk individuals with subclinical atherosclerosis: the METEOR trial, JAMA 297 (2007) 1344–1353.
- [8] E. de Groot, J.W. Jukema, A.D. Montauban van Swijndregt, et al., B-mode ultrasound assessment of pravastatin treatment effect on carotid and femoral artery walls and its correlations with coronary arteriographic findings: a report of the regression growth evaluation statin study (REGRESS), J. Am. Coll. Cardiol. 31 (1998) 1561–1567.
- [9] C.D. Furberg, H.P. Adams Jr., W.B. Applegate, et al., Effect of lovastatin on early carotid atherosclerosis and cardiovascular events. Asymptomatic Carotid Artery Progression Study (ACAPS) Research Group, Circulation 90 (1994) 1679–1687.
- [10] B. Hedblad, J. Wikstrand, L. Janzon, H. Wedel, G. Berglund, Low-dose metoprolol CR/XL and fluvastatin slow progression of carotid intima-media thickness: main results from the beta-blocker cholesterol-lowering asymptomatic plaque study (BCAPS), Circulation 103 (2001) 1721–1726.
- [11] J.J. Kastelein, S.I. van Leuven, L. Burgess, et al., Effect of torcetrapib on carotid atherosclerosis in familial hypercholesterolemia, N. Engl. J. Med. 356 (2007) 1620–1630.
- [12] J.J. Kastelein, F. Akdim, E.S. Stroes, et al., Simvastatin with or without ezetimibe in familial hypercholesterolemia, N. Engl. J. Med. 358 (2008) 1431–1443.
- [13] S. MacMahon, N. Sharpe, G. Gamble, et al., Effects of lowering average of below-average cholesterol levels on the progression of carotid atherosclerosis: results of the LIPID atherosclerosis substudy. LIPID Trial Research Group, Circulation 97 (1998) 1784–1790.
- [14] A. Meaney, G. Ceballos, J. Asbun, et al., The VYtorin on carotid intima-media thickness and overall arterial rigidity (VYCTOR) study, J. Clin. Pharmacol. 49 (2009) 838–847.
- [15] M. Mercuri, M.G. Bond, C.R. Sirtori, et al., Pravastatin reduces carotid intima-media thickness progression in an asymptomatic hypercholesterolemic mediterranean population: the Carotid Atherosclerosis Italian Ultrasound Study, Am. J. Med. 101 (1996) 627–634.
- [16] M.C. Meuwese, E. de Groot, R. Duivenvoorden, et al., ACAT inhibition and progression of carotid atherosclerosis in patients with familial hypercholesterolemia: the CAPTIVATE randomized trial, JAMA 301 (2009) 1131–1139.
- [17] Pfizer Inc. PhRMA web synopsis, Protocol A2581051, Carotid Atorvastatin Study in Hyperlipidemic Post-Menopausal Women: A Randomised Evaluation of Atorvastatin Versus Placebo (CASHMERE). 10-29-2007.
- [18] R. Salonen, K. Nyyssonen, E. Porkkala, et al., Kuopio Atherosclerosis Prevention Study (KAPS). A population-based primary preventive trial of the effect of LDL lowering on atherosclerotic progression in carotid and femoral arteries, Circulation 92 (1995) 1758–1764.
- [19] T.J. Smilde, S. van Wissen, H. Wollersheim, M.D. Trip, J.J. Kastelein, A.F. Stalenhoef, Effect of aggressive versus conventional lipid lowering on atherosclerosis progression in familial hypercholesterolaemia (ASAP): a prospective, randomised, double-blind trial, Lancet 357 (2001) 577–581.

- [20] A.J. Taylor, S.M. Kent, P.J. Flaherty, L.C. Coyle, T.T. Markwood, M.N. Vernalis, ARBITER: Arterial Biology for the Investigation of the Treatment Effects of Reducing Cholesterol: a randomized trial comparing the effects of atorvastatin and pravastatin on carotid intima medial thickness, Circulation 106 (2002) 2055–2060.
 [21] A.J. Taylor, L.E. Sullenberger, H.J. Lee, J.K. Lee, K.A. Grace, Arterial Biology for the Investigation of the Treatment Effects of
- [21] A.J. Taylor, L.E. Sullenberger, H.J. Lee, J.K. Lee, K.A. Grace, Arterial Biology for the Investigation of the Treatment Effects of Reducing Cholesterol (ARBITER) 2: a double-blind, placebo-controlled study of extended-release niacin on atherosclerosis progression in secondary prevention patients treated with statins, Circulation 110 (2004) 3512–3517.