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Longitudinal Trajectories of Digital Cognitive Biomarkers for Multiple Sclerosis

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ABSTRACT

Background: Cognitive impairment is one of the most common and debilitating symptoms of relapsing–remitting multiple sclerosis (RRMS). Digital cognitive biomarkers require less time and resources and are rapidly gaining popularity in clinical settings. We examined the longitudinal trajectory of the iPad-based Processing Speed Test (PST) and predictors of PST scores.

Methods: We prospectively enrolled RRMS patients between 2017 and 2021 across six Australian MS centres. Longitudinal data was analysed with mixed effect modelling and latent class mixed models. We then examined whether latent class group membership predicted confirmed decrease in correct PST responses.

Results: We recruited a total of 1093 participants, of which 724 had complete baseline data with a median follow up duration of 2 years. At a population level, PST trajectory was stable. A small practice effect was present up to the 4th visit. Age, baseline disability, T2 lesion volume, male sex and depression were associated with lower correct PST responses, whilst years of education and full/part-time employment were associated with more correct PST responses.

We identified four latent class trajectories of PST. The worst latent class was typified by low baseline PST and lack of a practice effect. Being in the worst latent class was associated with a greater hazard of time to sustained 5% decrease in PST (HR 2.84, 95% CI 1.16-6.94, p=0.02).

Conclusion: Worse baseline cognitive performance and lack of a practice effect predicted future cognitive decline in RRMS.

Helmut Butzkueven and Anneke van der Walt contributed equally as senior authors

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1 | Introduction

Multiple sclerosis (MS) is the most prevalent chronic inflammatory demyelinating neurological disease, affecting 2.8 million people worldwide [1]. Whilst the physical manifestations of MS have been well-studied, cognitive impairment is a common, debilitating, and under-recognised aspect of the disease. It affects up to 75% of people with MS (pwMS) and is independently associated with underemployment, poorer quality of life, driving impairment and social isolation [2–6]. As a result, the National MS society recommends baseline and annual screening for cognitive impairment in pwMS [7].

However, practical implementation of these recommendations is difficult. There is no standardised, easily administered screening tool appropriate for a routine clinical setting with limited appointment time and staff resources. The Symbol Digit Modalities Test (SDMT) is well-validated in MS and relatively quick to perform, thus suggested by Kalb and colleagues as the test of choice [7]. However, there are fees for its use and staff time for testing and scoring is required. As a result, it is rarely used in routine clinical practice.

The Processing Speed Test (PST), an iPad-based adaptation of the SDMT, has been introduced as an alternative. It has demonstrated excellent test-retest reliability, ecological and discriminative validity in cross-sectional studies [8–12]. However, due to the absence of longitudinal data we are unable to predict those at risk of future cognitive decline. This has significant treatment implications as early intervention may slow cognitive decline and increase cognitive reserve [13, 14]. Accordingly, we examined the trajectories of PST scores in a cohort of pwMS with mild-to-moderate disability and to identify predictors of PST scores.

2 | Methods

2.1 | Study Design

We recruited adult MS participants between December 2017 and March 2021 from six tertiary MS clinics in Australia. All participants were also enrolled in the MSBase registry, an international, observational cohort of pwMS [15]. Inclusion criteria were a confirmed diagnosis of relapsing-remitting MS (RRMS) or diagnosis of clinically isolated syndrome with evidence of MRI lesions that meet the Paty A or Paty B criteria, age > 18 years and an Expanded Disability Status Scale (EDSS) of less than 4 [16]. There were 28 minor deviations from the inclusion criteria where EDSS was four or above, mostly due to delays in baseline EDSS assessment because of restrictions in attending clinics in person due to COVID-19. These participants were allowed to continue in this study. For MSBase, all patients with MS who were attending a participating site and have provided informed consent were eligible. All assessments (EDSS, patient-reported outcomes (PROs), Multiple Sclerosis Performance Test (MSPT)) were performed during routine clinic visits (approximately six monthly). Clinical data were extracted from the MSBase registry. This study was approved by the Melbourne Health Human Research Ethics Committee and the University of Tasmania Human Research Ethics Committee. All participants provided informed consent prior to any data collection.

2.2 | Processing Speed Test

The PST is an iPad-based cognitive test, administered as part of the MSPT. The iPad displays a symbol key which contains nine symbols with corresponding numbers. After a practice test, participants were presented with rows of 15 symbols and instructed to select corresponding numbers. Once a row was completed a new row of symbols was presented. The total duration of the test is 2 min. The total number of correct responses in each test was used as the primary outcome for this assessment.

2.3 | Patient Reported Outcomes (PROs)

We utilised the Patient Health Questionnaire (PHQ-9) as a measure of depression. The PHQ-9 is a 9-item measure of depressive symptoms in the past 2weeks with a scale of 0–3 [17].

Higher scores in the PHQ-9 indicate a worse tendency to depression and has been validated in pwMS [18]. Electronic versions of these PROs were administered electronically using an iPad as previously described [19].

2.4 | Statistical Analysis

The baseline characteristic of the population was reported as the median (interquartile range) or mean (standard deviation) for continuous variables and number (percentile) for discrete variables, respectively.

We used a spaghetti plot with a fitted line to visualise the longitudinal trajectories of the PST. We examined mean change between visits to identify the magnitude and persistence of the practice effect. We then used multivariable linear mixed-effects models [20] with PST as the dependent variable and baseline covariates of age, disease duration (years), sex, EDSS, T2 lesion volume, T2 lesion number, 2-year relapse count, education (years), employment (unemployed, part-time, full-time), PHQ score and follow up duration. We then added interaction terms for all covariates and follow up duration in another model. A random intercept at the individual level was used to account for heterogeneity across patients.

We then examined the proportion of participants with 150 day sustained 5% sustained worsening from baseline assessment. This cut-off was used as in the original studies assessing change in SDMT for pwMS with relapses versus pwMS without relapse, the within group percentage change were as low as 2.2% [21]. We postulated that in our cohort of mild to moderate pwMS, primarily treated with high-efficacy disease modifying therapy (DMT), a lower cut-off could increase sensitivity for detection of clinically significant change [22]. We also conducted sensitivity analysis assessing a 4-point decrease in PST from baseline assessment [21, 23]. This cut-off has previously been validated as clinically meaningful for the SDMT [24].

A latent class mixed model (LCMM) was used to identify trajectories of digital cognitive biomarkers [25]. Initially, we explored the relationship between underlying latent processes and outcome. We fitted the LCMM of a single class with different linking

functions, including linear, beta cumulative distribution function or varying number of quadratic I-splines (3–5 and) with either equidistant knots or knots at quantiles. We selected the most appropriate link function by comparing the Bayesian Information Criterion (BIC) [26]. We next fitted models with varying numbers of latent classes (1–4) and the selected linking function appropriate for the data. To avoid selecting overfitted models based on BIC alone, we also considered clinically plausible trajectories and only included classes including at least 5% of the study cohort.

We examined the posterior probability of classification in a particular trajectory. A posterior probability of > 70% was considered acceptable, and proportions of individuals classified into each trajectory are reported [26].

A p < 0.05 was considered statistically significant. All statistical analyses were performed using R 4.1.2.

3 | Results

3.1 | Participant Characteristics

Participant characteristics are presented in Table 1. In total, 1093 participants were consented. Of these, 918 had baseline MSPT data, and 724 participants had complete questionnaire and MRI data for inclusion in the mixed models. The median number of study visits by participants was three, with a median follow-up duration of 728 days (IQR 264 to 1099; full table of visit numbers in Table A1).

3.2 | Longitudinal Analysis

Figure 1 shows the spaghetti plot of PST scores over time with a line of best fit. At a population level, PST scores stayed relatively stable over time. A practice effect was evident up to the 4th visit (Table 2 and Figure 1).

TABLE 1 | Participant characteristics at baseline (n = 724 unless otherwise specified).

	Median (IQR)
Age (years)	40.3 (33.4-47.9)
Disease duration (years)	7.5 (3.7–11.2)
Female (n, %)	549 (75.8)
EDSS	1.5 (1-2)
T2 lesion number	38 (19-74)
T2 lesion volume (mm³)	3074 (1187–7086)
2-year relapse count ^b	0 (0-1)
Years of education completed	14 (12–16)
PST at baseline ^a	55 (48-61)
PHQ-9 score (range 0–27)	5 (2-9)

 $Abbreviations: EDSS, Expanded\ Disability\ Status\ Scale; IQR, interquartile\ range; PHQ,\ Patient\ Health\ Questionnaire; PST,\ Processing\ Speed\ Test.$

PST trajectory over time

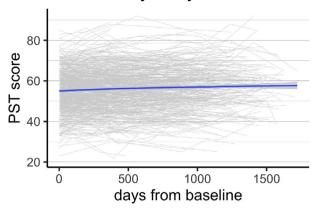


FIGURE 1 $\,\,$ Spaghetti plot of PST trajectory over time with line of best fit.

TABLE 2 | Mean change in PST up to the 6th visit.

Change between visits	Mean PST score (at preceding visit)	Mean PST change between visits	Total participant number
Visit 1 to 2	54.63	0.75	597
Visit 2 to 3	55.87	0.83	430
Visit 3 to 4	56.59	1.06	277
Visit 4 to 5	58.47	-0.35	154
Visit 5 to 6	58.69	-0.50	70

In the final linear mixed-effect model (Table 3), age, baseline EDSS, T2 lesion volume, male sex, and PHQ-9 score were negatively associated with PST, whilst years of education and full/part-time employment status were positively associated with PST. There was an increase in PST scores by 1 point every year, which we interpreted as the practice effect. None of the interaction terms between covariates and time were significant (data not shown).

3.3 | Sustained 4-Point Drop in PST

We then analysed those who had a confirmed 4-point drop in PST. Of the 430 participants with sufficient follow-up (minimum of three visits), 40 had a confirmed 4-point drop. Figure 2 shows the hazard of time to confirmed 4-point drop in PST.

3.4 | Latent Class Analysis

The linking function that provided the best fit to the longitudinal PST scores was a nonlinear linking function of three quadratic I-splines with nodes spaced in quantiles. Of the 430 people with sufficient follow-up, 77 (17.9%) were in class 1, 191 (44.4%) in class 2, 119 (27.67%) in class 3 and 43 (10.0%) in class 4. Figure 3 shows the scatterplot of the PST scores with overlying lines of mean SDMT scores per latent class. Class 1 started with the highest baseline PST score, and a slight improvement was noted over time. Class 2 and 3 had started with medium PST scores and were largely able

^aNumber of correct responses. ^bPrior to baseline.

TABLE 3 | Predictors of PST scores in linear mixed-effects model (n = 724).

Independent variables	Beta coefficient	95% CI	p
Age	-0.39	−0.46 to −0.33	< 0.01
Disease duration, years	-0.01	-0.13 to 0.10	0.84
EDSS	-0.66	−1.29 to −0.03	0.0
T2 lesion volume (per 1000 mm³)	-0.25	−0.36 to −0.14	< 0.01
T2 lesion number	0.00	-0.01 to 0.02	0.56
2-year relapse count ^b	-0.55	-1.39 to 0.28	0.20
Male sex	-3.54	-4.94 to -2.14	< 0.01
Education years	0.65	0.44 to 0.86	< 0.01
Part-time employment ^a	1.88	0.19 to 3.58	0.03
Full-time employment ^a	1.84	0.28 to 3.40	0.02
PHQ-9 score	-0.13	−0.25 to −0.02	0.03
Follow-up duration, years	0.58	0.40 to 0.77	< 0.01

Note: All covariates were baseline measures. Results with p < 0.05 are in boldface.

Abbreviations: CI, confidence interval; EDSS, Expanded Disability Status Scale; PHQ, Patient Health Questionnaire.

to maintain this over time—however, after 3 years a gradual decrease was noted in class 3. Finally, class 4 started with the lowest baseline PST scores, which remained relatively stable. Notably, a practice effect was lacking in this group. It should be noted that due to the small numbers in class 4, the increase towards the end of the follow-up period was driven by outliers.

3.5 | Latent Classes Predict Sustained Change in PST

We hypothesised that a sustained 4-point change would be more difficult to achieve if baseline PST score was low. For example, in class 4, the mean PST score was 40, whereas in class 1, mean PST score was 65. This represents a nearly 10% change for group 4 but only a 6% change for class 1. Therefore, we performed the primary analysis with sustained 5% change in PST scores from baseline instead of a 4-point change.

With this cut-off, 54 out of the 430 subjects with sufficient follow-up had a 5% sustained decrease in their PST. We found that class 4 had an increased hazard of time to 5% sustained PST decrease (HR 2.84, 95% CI 1.16 to 6.94, p = 0.02) (Table 4 and Figure 4). No other classes were associated.

3.6 | Sensitivity Analyses

3.6.1 | 4-Point Sustained PST Decrease

We then analysed whether latent class membership predicted a sustained 4-point decrease in PST. There was no significant relationship between latent class membership and 4-point confirmed PST change (Table 5 and Figure 5).

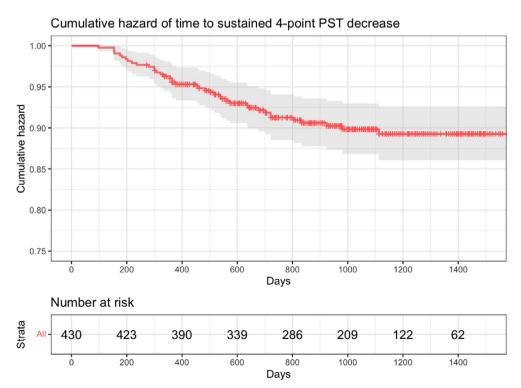


FIGURE 2 | Cumulative hazard of time to sustained 4-point PST decrease.

^aCompared to unemployed/retired.

^bPrior to baseline.

PST trajectory by latent class analysis

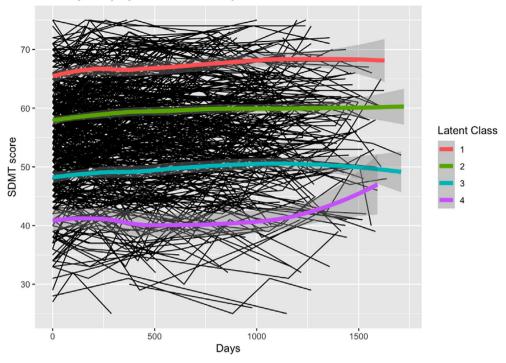


FIGURE 3 | PST trajectories defined by latent class analysis.

TABLE 4 | Hazard ratio and 95% CI for latent class trajectories and time to sustained 5% change in PST, compared to class 1 (reference).

Latent class group	Hazard ratio	95% CI	р
Class 2	0.87	0.39 to 2.00	0.75
Class 3	1.25	0.54 to 2.93	0.60
Class 4	2.84	1.16 to 6.94	0.02

Note: Results with p < 0.05 are in boldface.

4 | Discussion

In this large observational study investigating a digital cognitive biomarker in pwMS with mild to moderate disability, we observed that processing speed remained stable over 4 years' follow-up at a population level. We found that age, baseline EDSS, T2 lesion volume, male sex and PHQ-9 score were associated with lower PST scores, whilst years of education and full/ part-time employment were associated with higher PST scores. A small practice effect was noted up to the 4th visit despite the long intervals between tests. We identified four trajectories of PST over follow-up. The worst performing group was notable for having the worst baseline PST scores and lack of a practice effect. This group had greater hazard of time to sustained 5% decrease in PST. No other trajectories were associated with risks of sustained 5% decrease in PST, and no trajectories significantly predicted risk of sustained 4-point PST decrease. Our study illustrates that pwMS who have low baseline cognitive performance and no practice effect should be considered at greater risk of future cognitive decline, and early intervention should be considered to maximise cognitive reserve.

In this cohort of pwMS with mild to moderate disability, the worst latent class (class 4) predicted a sustained 5% PST change. This is a novel finding. Given the stability of latent classes over time, this signifies that baseline cognitive performance is an important predictor of subsequent cognitive decline. This class was also notable for a lack of practice effect. It may be that a practice effect demonstrates cognitive reserve that reduces the risk of subsequent cognitive decline, although this is difficult to delineate, as initial improvement with practice effect self-selects for a decreased likelihood of sustained decrease in PST. Overall, this represents a promising method to select pwMS at greater risk of deterioration for future progressive MS trials.

Our findings also show that for pwMS with minimal disability, a percentage-based measure instead of a uniform 4-point measure is the more sensitive change measure. In a prior longitudinal trajectory study on the digital MSReactor platform, a percentage-based cut-off was used as well [22]. In the original studies assessing change in SDMT for pwMS with relapses versus controls, the within group percentage change were as low as 2.2% [21]. In those with relapses with cognitive impairment, there was a mean 6.2% within group change. Furthermore, 6% change in SDMT has previously been shown to predict loss of employment over 3 years [23]. It should also be noted that most patients in the aforementioned studies were on low-efficacy DMT and had more severe disability. Our findings suggest that for pwMS with milder disability with access to high efficacy DMT use, a lower cut-off may increase sensitivity for detection of significant change. This will need to be clarified in future longitudinal studies to prove definitively that a 5% change in SDMT is clinically meaningful.

Cumulative hazard of time to 5% sustained PST decrease

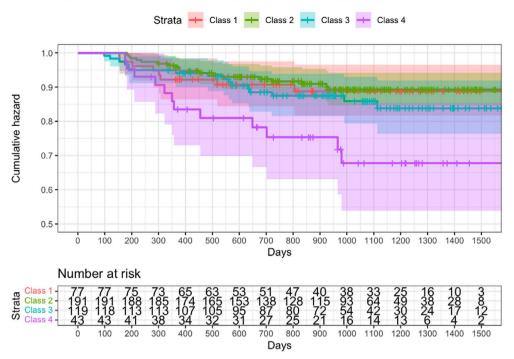


FIGURE 4 | Cumulative hazard of time to sustained 5% PST decrease. Kaplan–Meier curves were applied to show the cumulative hazard of time to sustained 5% PST decrease in latent classes 1–4.

TABLE 5 | Hazard ratio + 95% CI for latent class groupings and time to 4-point change in PST, compared to class 1 (reference).

Latent class group	Hazard ratio	95% CI	р
Class 2	0.67	0.28 to 1.59	0.36
Class 3	0.91	0.37 to 2.22	0.83
Class 4	1.32	0.46 to 3.80	0.61

At a population level, digital cognitive performance remained stable over time. Even within latent classes, deterioration was not seen until after the 4th visit. Even then, they did not drop below baseline PST values. This is similar to a recent study in a real-world cohort which showed improvement in the PST over time [27]. Prior studies with the SDMT have shown conflicting results. Some report deterioration over time [28], whilst a recent meta-analysis and a study over 11 years have reported similar findings with relative stability over time [29, 30]. It should be noted that given the lower level of disability in our cohort, significant cognitive decline would not be expected at a population level.

Our findings of age, baseline disability (as assessed by EDSS), male sex, education, employment and PHQ-9 as predictive of PST is similar to that found in prior studies [27, 31–35]. It should be noted that the association with employment could be bidirectional—whilst it is intuitive that declining cognition may lead to reduced employment, maintaining mental stimulation at work may protect against cognitive decline [36].

Our findings help to identify pwMS who are at greater risk of future cognitive deterioration. Early recognition and intervention should be considered in these patients. Confounding factors such as sleep disorders, pain fatigue, anxiety, depression, polypharmacy should be identified and addressed [37]. Whilst evidence is still lacking as to the benefit of a specific DMT compared to another, there is emerging evidence that early high-efficacy DMT use may slow cognitive decline [38]. Finally, exercise and cognitive rehabilitation should be encouraged as a way to slow cognitive decline [37].

Limitations of our study include the exclusion of progressive phenotypes and pwMS with severe disability. This limits the generalisability of our findings. Due to the COVID-19 pandemic, in-person clinic visits were significantly restricted, leading to longer than usual time between visits [39-41]. The number of follow up visits was relatively low, which affects the posterior probability of latent class trajectory membership. More frequent follow up would have allowed for better posterior probabilities. Furthermore, a practice effect was evident up to the 4th visit. Given the relative infrequency of the iPad testing, this meant that practice effects were evident up to the 2nd year of the study. Higher frequency of testing, particularly in the initial stages of the study, may alleviate this issue [42-44]. The Neuro-QOL, a patient-reported quality of life measure included in some MSPT versions, was not used in this study. Future studies of the MSPT may consider comparing the cognitive domains of Neuro-QOL against the PST.

5 | Conclusion

In this cohort of mild to moderate RRMS, PST scores remained stable over time at a population level. Latent class modelling identified a trajectory with low baseline PST and a lack of practice effect that predicted sustained 5% PST decrease. Our study

Cumulative hazard of time to 4-point PST decrease

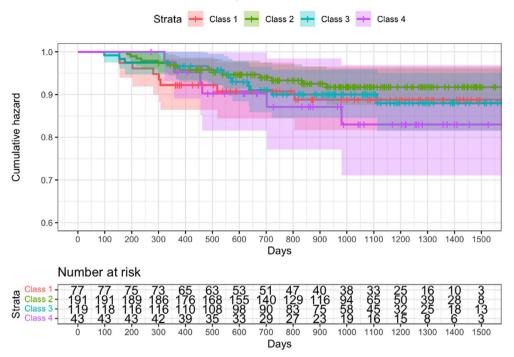


FIGURE 5 | Cumulative hazard of time to sustained 4-point PST decrease, stratified by latent classes. Kaplan–Meier curves were applied to show the cumulative hazard of time to sustained 4-point PST decrease in latent classes 1–4.

supports digital cognitive biomarkers and latent class modelling as a screening tool to identify those that may benefit from early intervention for maximising cognitive function. Future studies are needed to explore whether a 5% PST decrease predicts clinically meaningful cognitive decline.

Author Contributions

Y.F., H.B., A.v.d.W., D.M., M.G., C.Z., K.B., P.D., J.v.B., R.H., D.D., To.K., Tr.K., B.T., M.B. and J.L.S. conceived the study and methodology. J.L.S., M.B., B.T., To.K., Tr.K., D.D., H.B. and A.v.d.W contributed to data preparation. Y.F., M.G. and C.Z. performed statistical analysis and data extraction. Y.F. wrote the original draft of the manuscript. M.B. and C.W. contributed to the imaging data. M.G., D.M., C.Z., K.B., H.B. and A.v.d.W supervised the project. S.S.Y. provided domain-related feedback. H.B. and A.v.d.W were considered co-senior authors and contributed equally. All authors provided critical feedback and intellectual input throughout the study, revised the initial draft, and contributed to the final version, of the submitted manuscript.

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Conflicts of Interest

Yi Chao Foong received travel compensation and speaker's honoraria from Biogen. He also receives research funding support from National Health and Medical Research Council, Multiple Sclerosis Research Australia and Australian and New Zealand Association of Neurologists; Melissa Gresle is currently working on observational studies funded by Biogen and Roche; Daniel Merlo has received honoraria from Novartis; Katherine Buzzard has received honoraria for presentations and/or educational support from Biogen, Sanofi Genzyme, Merck, Roche, Alexion and Teva. She serves on medical advisory boards for Merck and Biogen; Jeannette Lechner-Scott received travel compensation from Novartis, Biogen, Roche and Merck. Her institution receives the honoraria for talks and advisory board commitment as well as research grants from Biogen, Merck, Roche, TEVA and Novartis; Michael Barnett served on scientific advisory boards for Biogen, Novartis and Genzyme and has received conference travel support from Biogen and Novartis. He serves on steering committees for trials conducted by Novartis. His institution has received research support from Biogen, Merck and Novartis; Chenyu Wang has no competing interests.; Trevor Kilpatrick receives support from Novartis in the form of consultancy fees, honoraria for giving lectures and funding for a pre-clinical Investigator-Initiated Study (IIS).; David Darby was a founder and shareholder in Cogstate Ltd., but has not been involved in this company since 2011. He is a consultant to UBrain, Brazil. His company CereScape Ltd. receives a stipend to maintain the msreactor.com website.;Tomas Kalincik served on scientific advisory boards for MS International Federation and World Health Organisation, BMS, Roche, Janssen, Sanofi Genzyme, Novartis, Merck and Biogen, steering committee for Brain Atrophy Initiative by Sanofi Genzyme, received conference travel support and/or speaker honoraria from WebMD Global, Eisai, Novartis, Biogen, Roche, Sanofi-Genzyme, Teva, BioCSL and Merck and received research or educational event support from Biogen, Novartis, Genzyme, Roche, Celgene and Merck; Bruce Taylor received funding for travel and speaker honoraria from Bayer Schering Pharma, CSL Australia, Biogen and Novartis, and has served on advisory boards for Biogen, Novartis, Roche and CSL Australia; Robert Hyde was an employee of Biogen and holds stock/stock options in Biogen; Johan van Beek is an employee of Biogen; Pamela Dobay is a full-time employee of Biogen and owns Biogen stock.; Steve Simpson-Yap has no competing interests; Anneke van der Walt served on advisory boards for Novartis, Biogen, Merck and Roche and NervGen. She received unrestricted research grants from Novartis, Biogen, Merck and Roche. She is currently a co-Principal investigator on a co-sponsored observational study with Roche, evaluating a Roche-developed

smartphone app, Floodlight-MS. She has received speaker's honoraria and travel support from Novartis, Roche, Biogen and Merck. She serves as the Chief operating Officer of the MSBase Foundation (not for profit). Her primary research support is from the National Health and Medical Research Council of Australia and MS Research Australia; Helmut Butzkueven's Institution has received compensation for advisory boards or lecture fees from Novartis, Biogen, Merck, UCB Pharma and Roche. His institutions receive research funding from Novartis, Biogen, Merck, Roche, The National Health and Medical Research Council of Australia, The Medical Research Future Fund (Australia), Monash Partners, the Trish MS Foundation, The Pennycook Foundation, and MS Australia. He receives personal compensation as the Managing Director of the MSBase Foundation and from the Oxford Health Policy Forum Brain Health Initiative.

Data Availability Statement

The data that support the findings of this study are available from the corresponding author upon reasonable request.

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Appendix A

TABLE A1 | Total number of visits by participant count and intervisit intervals.

Total number of visits	Participant count	Median days from prior visit
1	127	N/A
2	167	218
3	153	236
4	123	350
5	84	330
6	50	207
7	14	231
8	6	N/A