Review Article

Effects of Systemic or Local Administration of Zoledronate on Implant Osseointegration: A Preclinical Meta-Analysis

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Objective. This study aims to investigate the effect of systemically administrated zoledronate on bone-implant fixation in animal models. *Methods.* We searched MEDLINE, Embase, and EBSCO for studies that explore the role of systemic or local zoledronate delivery in implant osseointegration in animal models. The Review Manager software was used to analyze selected studies by using the weighted mean difference random-effects model. Analytical data are mainly about bone ingrowth, such as bone-to-implant contact (BIC), bone volume/total volume (BV/TV), and bone area. *Results.* Twenty studies were selected from 182 publications. The mean quality score was 18/20 for all of the 20 studies (κ = 0.9). Despite differences in protocols, these studies showed consistent improvement of implant osseointegration with zoledronate administration. In addition, the osteoporotic animal model, systemic or local administration of zoledronate could improve the osseointegration of orthopedic implant in animal models. Results of this meta-analysis should be interpreted cautiously because of the inherent differences between preclinical and clinical subjects. For the local administration, there is a similar trend as well, but the results need to be confirmed and complemented with further analyses.

1. Introduction

Bone ingrowth into a prosthetic implant is crucial for the longevity of uncemented total hip arthroplasty (THA). Rapid and sound bone ingrowth can increase implant stability and improve long-term bone-implant fixation. In addition, adequate tissue ingrowth may protect the boneimplant interface against wear particle-induced osteolysis, which further decreases the risk of aseptic loosening [1]. Therefore, how to improve bone-implant fixation is always a topic of great interest for joint surgeons.

Zoledronate (ZOL) is a new-generation intravenous bisphosphonate (BP) with the greatest affinity and longest retention for bone mineral, and it has been largely utilized in the treatment of osteoporosis and metastatic bone disease. It has a well-documented profile of possible side effects, such as initial influenza-like illness which has been documented with the first infusion of BPs. Renal failure has been noted in patients with cancer after repetitive high-dose infusions, and an association between BPs and osteonecrosis of the jaw after tooth extraction has been recorded as well [2, 3]. ZOL is traditionally believed to be an antiresorptive agent; however, recent animal studies suggested that it could stimulate bone formation and improve implant mechanical fixation [4, 5]. Meanwhile, there are clinical studies revealing that ZOL is associated with decreased early implant migration and reduced peri-implant bone loss [6, 7]. Nevertheless, apart from these findings, because of small sample sizes and diverse study protocols of previous researches, the exact cellular and molecular mechanisms governing the improved bone content, structure, and strength, induced by the systemically or locally administered ZOL on implant osseointegration, have no consensus. Thus, a preclinical meta-analysis was performed to investigate the effect of systemically or local administrated ZOL on bone-implant fixation in animal models and to guide the design of evidence-based, largescale preclinical or clinical trials.

2. Materials and Methods

2.1. Inclusion and Exclusion Criteria. Inclusion criteria for the literature were as follows: (1) original animal studies, (2) studies that aimed to explore the role of ZOL delivery in implant osseointegration, (3) studies that included a control group, which received placebo or no drug, and (4) studies with the outcomes that included pertinent information regarding bone ingrowth, such as bone-to-implant contact (BIC), bone volume/total volume (BV/TV), and bone area. Exclusion criteria were as follows: (1) clinical studies, (2) implants embedded in the mandible or maxilla, and (3) presence of other confounding factors, such as wear debris in a local environment.

2.2. Search Strategy. Literature, which was published before September 30, 2018, was searched by using the electronic databases MEDLINE, Embase, and EBSCO. No language restriction was applied. The adopted search keywords were "implant AND (bisphosphonates OR zoledronate) AND osseointegration AND arthroplasty." Titles and abstracts of studies that fulfill the eligibility criteria were screened by the authors and checked for agreement. Finally, the reference lists of all full-text papers, which were identified pertinent to the study, were reviewed for any unidentified studies.

2.3. Study Selection. Two authors (Yao He and Zhengyun Li) independently applied the search strategy to select references from the aforementioned databases. The titles and abstracts were reviewed independently. When in doubt, the full-text articles were retrieved for further examination. These two authors independently assessed each full report to evaluate fulfillment of the inclusion criteria, and corresponding authors were contacted for more information and clarification regarding their data, if necessary. Any disagreement was discussed with the senior author, and when consensus could not be established, that study was excluded.

2.4. Quality Assessment. The methodological quality of included studies was assessed independently by two authors (Yao He and Xiang-Dong Wu) according to the ARRIVE guidelines which included title, abstract, background, objectives, ethical statement, study design, experimental procedures, experimental animals, housing and husbandry, sample size, allocation of animals to experimental groups, experimental outcomes, statistical methods, baseline data, numbers analyzed, outcomes and estimation, adverse events, interpretation/scientific implications, generalizability/translation, and funding. Each study was given a quality score out of a possible total of 20 points. Any disagreement was resolved by the senior author (Yao He).

2.5. Data Extraction. A data extraction form was designed and agreed by the authors, and a pilot test of five articles was performed to ensure their consistency. Initially, two authors (Yao He and Xiang-Dong Wu) independently extracted the data, which were later reviewed jointly to produce the agreed accurate data. Disagreements were resolved by consensus or consultation with the senior author. The extracted data included study design, animal species, implantation site, implant characteristics (material, shape, and coating), ZOL route and dosage, follow-up time, and outcome measurements (BIC, BV/TV, and bone area). In all studies, BIC was calculated as the length percentage of the direct bone-implant interface to the total implant surface, BV/TV was defined as the percentage of mineralized bone volume to total bone tissue volume in the peri-implant region, and bone area was evaluated as the percentage of bone tissue area to the total area of the bone and implant.

2.6. Statistical Analysis. Review Manager (RevMan version 5.0, The Cochrane Collaboration in 2008) was used to analyze the included studies. The primary outcome was the BIC between treatment and control groups. From a clinical point of view, the authors (Yao He and Xiang-Dong Wu) performed subgroup analyses according to the animal model (osteoporotic or normal), animal species, and drug dosage and frequency, as well as follow-up time. In case of multiple treatment groups next to a control group within one trial, the animal number in the control group was divided equally by the number of treatment groups. For each arm in a particular study, continuous data were expressed as means and standard deviations (SDs), and dichotomous data were expressed as proportions or risks. For continuous outcomes, we calculated the mean differences (MDs) with 95% confidence interval (CI). For dichotomous outcomes, we estimated the relative risks' 95% CI. Statistical heterogeneity was assessed by using the value of I^2 and the result of the chisquare test. An I^2 value >50% suggests statistical heterogeneity, which prompts a random-effects modeling estimate. Otherwise, a fixed-effects approach was used. A P value <0.05 was determined as statistically significant.

3. Results

3.1. Included Studies. A total of 182 articles were searched from multiple electronic databases. After screening their titles and available abstracts, 20 satisfied the eligibility criteria and were included in the meta-analysis [8–27] (Figure 1).

3.2. Characteristics of Enrolled Studies. The sample size ranged from 10 to 64. In twelve studies, rats were used as animal models; in other seven studies, rabbits were utilized as animal models; and in the last one study, dogs were used as animal models. The follow-up time ranged from 10 days to 1 year. The tail vertebra was used as the implantation site in one study; in another four studies, the femoral condyle was operated on; the remaining 15 studies all selected the proximal tibia as the surgical site. Out of 20 studies, one used



FIGURE 1: Flow of studies through the review.

the tantalum as the implant, another used the calcium phosphate bone cement, and the remaining eighteen studies used the titanium implant (Table 1).

3.3. Methodological Quality. The mean quality score was 18/ 20 for all of the 20 studies (Table 2) (κ =0.9). Among all publications, eighteen (90%) reported the animal breeding condition, fourteen (70%) reported random allocation, no study reported blinded surgical implantation, and only one (5%) reported blinded outcome assessment. One study (5%) reported sample size calculation.

4. Systemic Administration

4.1. Bone-to-Implant Contact. Ten articles [8–17] with 193 animals reported BIC measurements with substantial heterogeneity between studies ($I^2 = 98\%$, P < 0.001); thus, the random-effects model was used to evaluate the results. A significant difference of BIC could be found between the treatment and control groups (MD, 13.44; 95% CI, 7.34–19.55; P < 0.0001) (Figure 2).

Subgroup analyses were performed according to the animal model (osteoporotic *vs.* normal), animal species (rats *vs.* rabbits), ZOL dosage (>0.1 mg/kg *vs.* <0.1 mg/kg), administration frequency (single *vs.* multiple), and follow-up time (>8 weeks *vs.* <8 weeks). As shown in Figure 3, ZOL could significantly increase BIC in osteoporotic animals (mean difference, 16.52; 95% CI, 8.07–24.98; P = 0.0001); however, this effect was not obvious in normal animals (MD,

6.70; 95% CI, -1.75 to 15.15; P = 0.12). Based on animal species, similar effects of ZOL on BIC were observed for both rats (MD, 12.61; 95% CI, 5.50–19.72; P = 0.0005) and rabbits (MD, 13.89; 95% CI, 5.80–21.98; P = 0.0008).

Regarding the drug dosage, when ZOL dosage exceeded 0.1 mg/kg, BIC could be significantly improved compared with the control group (MD, 14.86; 95% CI, 8.20–21.51; P < 0.001) (Figure 4). On the contrary, administration frequency had no significant impact on BIC (test for subgroup difference: P = 0.86). Finally, although positive influence of ZOL on BIC has been demonstrated in short follow-up time studies (<8 weeks; MD, 3.92; 95% CI, 0.69–7.14; P = 0.02), this effect was much more significant in case of longer follow-up time (>8 weeks; MD, 17.65; 95% CI, 9.30–26.06; P < 0.001).

4.2. Bone Volume/Total Volume. The pooled analysis of five experiments [11, 13, 15, 18, 19] showed a significant difference of BV/TV between ZOL-treated and control groups (MD, 26.28; 95% CI, 7.58–44.99; P = 0.006) with heterogeneity ($I^2 = 99\%$, P < 0.001) (Figure 5).

Animal model (osteoporotic vs. normal) was initially used for subgroup analysis. The results showed that more effects of ZOL were seen improving BV/TV in osteoporotic animals (MD, 22.28; 95% CI, 11.98–32.58; P = 0.0004). With regard to animal species, ZOL significantly increased BV/TV in rats (MD, 18.60; 95% CI, 4.59–32.60; P = 0.009) and rabbits (MD, 8.00; 95% CI, 3.47–12.53; P = 0.0005). Finally, similar effects of ZOL on

Studies	Animals	Group		ZOL dosage and route, or		Implant		FU
Studies	71111111111	Number	Size	concentration	Site	Туре	Coating	time
Ayan et al. [8]	Male rabbits	1: control 2: ZOL	9	0.1 mg/kg IV Single dose	Tibia	Titanium screw	No	4 weeks
Cardemil et al. [9]	Female rats	1: control (OVX) 2: ZOL	28	0.1 mg/kg IV Single dose	Tibia	Titanium screw	No	4 weeks
Carvas et al. [10]	Male rabbits	1: control (GC) 2: ZOL	6	0.1 mg/kg IV Single dose	Tibia	Titanium screw	No	18 weeks
Chen et al. [11]	Female rats	1: control (OVX) 2: ALN 3: SR 4: ZOL	12	0.1 mg/kg IV Single dose	Tibia	Titanium rod	НА	12 weeks
Dikicier et al. [12]	Female rats	1: control (OVX) 2: ZOL	12	0.04 mg/kg IV 6 doses	Tibia	Titanium screw	No	8 weeks
Li et al. [13]	Female rabbits	1: control (OVX) 2: ZOL	6	0.1 mg/kg SC 3 doses	Tibia	Titanium screw	HA	8 weeks
de Oliveira et al. [14]	Male rats	1: control 2: ZOL	3	0.0075 mg/kg SC 3 doses	Tibia	Titanium	No	4 weeks
Qi et al. [15]	Female rabbits	1: control (OVX) 2: local ZOL 3: systemic ZOL 4: local and systemic ZOL	8	0.1 mg/kg SC 4 doses	Tibia	Titanium screw	НА	12 weeks
Yaman et al. [16]	Male rats	1: control 2: ZOL	7	0.1 mg/kg IV 3 doses	Tibia	Titanium screw	HA	12 weeks
Yildiz et al. [17]	Female rabbits	1: control (OVX) 2: ZOL	12	0.1 mg/kg IV Single dose	Tibia	Titanium screw	Resorbable blast media	8 weeks
Blazsek et al. [18]	Female rats	1: control 2: ZOL	5	0.6 mg/kg IP 3 doses	Tail vertebra	Titanium screw	No	6 weeks
von Knoch et al. [19]	Rabbits	1: control 2: ZOL	12	0.015 mg/kg IV Single dose	Femur	Titanium cylinder	Fiber metal mesh	12 weeks
Arnoldi et al. [20]	Rabbits	1: control 2: ZOL	6	1μg/ml	Femur	Titanium screw	No	10 days
Miettinen et al. [21]	Male rats	1: control 2: ZOL	7	20 µg/ml	Femur	Titanium	No	4 weeks
Gao et al. [22]	Female rats	1: control (OVX) 2: ZOL	10	1 mg/ml	Tibia	Titanium	HA	12 weeks
Stadlinger et al. [23]	Female rats	1: control (OVX) 2: ZOL	8	8.5 µg/implant	Tibia	Titanium	No	4 weeks
Ying et al. [24]	Female rats	1: control (OVX) 2: ZOL	10	$30\mu g/implant$	Tibia	Titanium cylinder	No	12 weeks
Pyo et al. [25]	Female rats	1: control (OVX) 2: ZOL	10	8 μg/ml 80 μg/ml 800 μg/ml	Tibia	Titanium screw	No	8 weeks
Sörensen et al. [26]	Rats	1: control 2: ZOL	10	50 µg/implant	Tibia	Calcium phosphate bone cement	No	3 weeks
Bobyn et al. [27]	Dogs	1: control 2: ZOL	5	0.2 mg/ml	Femur	Tantalum	HA	1 year

TABLE 1: Characteristics of enrolled studies.

ZOL: zoledronate; FU: follow-up; GC: glucocorticoid; HA: hydroxyapatite; ALN: alendronate; SR: strontium ranelate; SC: subcutaneous; IP: intraperitoneal.

bone area were observed for follow-up time <8 weeks (MD, 13.19; 95% CI, 1.22–25.17; P = 0.03) and >8 weeks (MD, 20.12; 95% CI, 6.81–33.43; P = 0.003). ZOL has a significant effect on BV/TV if ZOL dosage was more than 0.1 mg/kg (MD, 32.85; 95% CI, 20.15 to 45.55; P < 0.001) or if ZOL was given in multiple doses (MD, 28.52; 95% CI, 14.66 to 42.39; P < 0.001).

Animal model (osteoporotic *vs.* normal) was initially used for subgroup analysis. The results showed that ZOL could only improve BV/TV in osteoporotic animals (MD, 30.43; 95% CI, 16.05–44.80; P < 0.001), but not in normal ones (MD, 20.00; 95% CI, -22.18 to 62.19; P = 0.35). With regard to animal species, ZOL significantly increased BV/TV in rats (MD, 44.75; 95% CI, 37.48–52.01; P < 0.001);

TABLE 2: Quality assessment score of enrolled studies.

Studies	Quality score
Ayan et al. [8]	18
Cardemil et al. [9]	17
Carvas et al. [10]	16
Chen et al. [11]	19
Dikicier et al. [12]	17
Li et al. [13]	17
Marcio et al. [14]	19
Qi et al. [15]	17
Yaman et al. [16]	19
Yildiz et al. [17]	18
Blazsek et al. [18]	17
von Knoch et al. [19]	18
Arnoldi et al. [20]	18
Miettinen et al. [21]	18
Gao et al. [22]	19
Stadlinger et al. [23]	19
Ying et al. [24]	19
Pyo et al. [25]	19
Sörensen et al. [26]	18
Bobyn et al. [27]	17

however, this effect was not evidenced in rabbits (MD, 15.29; 95% CI, -7.01 to 37.59; P = 0.18). Finally, ZOL has an effect on BV/TV if ZOL dosage was more than 0.1 mg/kg (MD, 32.85; 95% CI, 20.15–45.55; P < 0.001) and if ZOL was given in multiple doses (MD, 28.52; 95% CI, 14.66–42.39; P < 0.001).

4.3. Bone Area. Only two studies [9, 10] with 68 animals reported bone area, and no heterogeneity was observed $(I^2 = 0, P = 0.87)$ (Figure 6).

5. Local Administration

5.1. Bone-to-Implant Contact. Seven articles [20–25] with 117 animals reported BIC measurements with substantial heterogeneity between studies ($I^2 = 95\%$, P < 0.001); thus, the random-effects model was used to evaluate the results. A significant difference of BIC could be found between the treatment and control groups (MD, 13.54; 95% CI, 4.43–22.65; P = 0.004) (Figure 2).

Subgroup analyses were performed according to the animal model (osteoporotic *vs.* normal), animal species (rats *vs.* rabbits *vs.* dogs), and follow-up time (>8 weeks *vs.* <8 weeks). As shown in Figure 7, ZOL with local administration could significantly increase BIC in osteoporotic animals (MD, 18.96; 95% CI, 10.27–27.65; P < 0.0001); but in normal animals, there was no effect at all (MD, –0.9; 95% CI, –11.22 to 9.42; P = 0.86).

Based on animal species, similar effects of ZOL on BIC were observed for rats (MD, 15.16; 95% CI, 0.66–29.66; P = 0.04) and rabbits (MD, 10.46; 95% CI, 2.21–18.70; P = 0.01). On the contrary, only the longer follow-up time has a increased effect on BIC (MD, 17.05; 95% CI, 7.21–26.89; P < 0.0007).

5.2. Bone Volume/Total Volume. Only three studies [15, 23, 25] with 51 animals reported bone volume/total

volume with heterogeneity between studies ($I^2 = 49\%$, P = 0.14) (Figure 4).

5.3. Bone Area. A total of seven articles [21–24, 26, 27] with 125 animals reported bone area showing a significant difference between ZOL-treated and no ZOL groups (MD, 16.02; 95% CI, 5.98–26.05; P = 0.002) with heterogeneity ($I^2 = 97\%$, P < 0.001) (Figure 6).

6. Discussion

THA is an effective technique owing to its ability to reduce pain, correct deformity, and improve function. However, its longevity is always an unsolved issue. According to previous reports, the most common reason of implant failure is aseptic loosening, which is caused by implant micromotion, prosthesis-related stress shielding, disuse osteoporosis, and wear-debris-induced osteolysis [6]. Thus, a simple, low-cost, and readily available method for improving implant fixation and decreasing periprosthetic bone loss is considerably important. Over the recent years, some animal study data proposed that ZOL might increase peri-implant bone stock and improve biological implant fixation, whereas other studies have denied this effect [14]. By pooling the currently available animal study data, the present meta-analysis provides evidence-based information about the positive effects of ZOL on implant osseointegration. In addition, our results indicated that the animal model, drug dosage, and follow-up time might influence study outcomes, which suggests possible reasons for the diversity of previous studies and gives insights into the design of future research.

Although bisphosphonates are well-known osteoclast inhibitors, they could reduce bone resorption by inhibiting and promoting apoptosis of osteoclasts [28, 29]. Several in vitro studies have demonstrated that they could also stimulate osteoblast function [30, 31]. However, according to our results, ZOL could only significantly improve implant osseointegration in osteoporotic animals but not in normal ones, which indicates that this effect was mediated mainly by decreasing the abnormal bone turnover rate rather than directly stimulating bone formation. Meanwhile, caution should be taken when interpreting this result. In the experimental animals, osteoporosis was acquired mainly by ovariectomy and was the only systemic condition. Nevertheless, in clinical settings, osteoporotic patients are usually old aged and sometimes diabetic. It has already been reported that aging and chronic hyperglycemia would lead to accumulation of advanced glycation end products (AGEs), which could negatively influence bone metabolism, and thus, the effect of ZOL may be less promising in clinical settings than in laboratories [32, 33].

In view of the widespread use of bisphosphonates and the increase in bisphosphonate-related cases of osteonecrosis of the jaw, some studies have shown that osteonecrosis with dental implants may be a side effect of treatment with BP. The incidence of bisphosphonate-related osteonecrosis of the jaws is accelerated at the end of or during BP treatment. Serra et al. [34] suggested the avoidance of

Study or subgroup		ZOL		1	No ZO	L	Waight (0/)	Mean difference	Mean difference
Study of subgroup	Mean	SD	Total	Mean	SD	Total	weight (%)	IV, random, 95% CI	IV, random, 95% CI
Systemic									
Avan et al. [8]	27.38	2.22	9	23.19	4.95	9	10.1	4.19 [0.65, 7.73]	+
Cardemil et al. [9]	72.89	6.27	28	66.27	6.33	28	10.1	6.62 [3.32, 9.92]	+
Carvas et al. [10]	38.95	4.24	6	25.98	2.35	6	10.0	12.97 [9.09, 16.85]	-
Chen et al. [11]	81	9.32	12	37.6	6.73	4	8.7	43.40 [34.96, 51.84]	
Dikicier et al. [12]	31.1	1.7	12	27.39	0.96	12	10.4	3.71 [2.61, 4.81]	
Li et al. [13]	43.11	2.99	6	23.12	2.32	6	10.2	19.99 [16.96, 23.02]	
de Oliveira et al. [14]	48.37	2.31	3	47.17	1.08	3	10.2	1.20 [-1.69, 4.09]	
Qi et al. [15] systemic	53.61	2.74	8	28.81	1.97	8	10.3	24.80 [22.46, 27.14]	
Yaman et al. [16]	45.8	3.77	7	31.14	1.04	7	10.2	14.66 [11.76, 17.56]	-
Yildiz et al. [17]	43.08	5.37	12	36.03	5.81	12	9.9	7.05 [2.57, 11.53]	-
Subtotal (95% CI)			103			95	100.0	13.44 [7.34, 19.55]	•
Heterogeneity: $tau^2 = 93$	3.09; chi	$^{2} = 433$	3.87; df	= 9 (P < 0)	0.0000	1); $I^2 = 9$	8%		
Test for overall effect: Z	= 4.31 (P < 0.0	0001)			,,			
Local									
Arnoldi et al. [20]	15	8	6	10	8	6	13.4	5.00 [-4.05, 14.05]	
Gao et al. [22]	51.74	10.07	10	29.09	9.33	10	13.6	22.65 [14.14, 31.16]	
Miettinen et al. [21]	69.7	4.5	7	75.3	6.3	7	14.6	-5.60 [-11.34, 0.14]	
Pvo et al. [25]	48.38	8.07	10	44.46	6.36	10	14.4	3.92 [-2.45, 10.29]	4 - -
Oi et al. [15] local	42.51	3.55	8	28.81	1.97	8	15.2	13.70 [10.89, 16.51]	-
Stadlinger et al. [23]	47.8	10.4	7	20.6	2.4	8	13.8	27.20 [19.32, 35.08]	
Ying et al. [24]	49.52	5.11	10	21.56	4.32	10	15.0	27.96 [23.81, 32.11]	
Subtotal (95% CI)			58			59	100.0	13.54 [4.43, 22.65]	•
Heterogeneity: $tau^2 = 1^2$	39.76: cł	$ni^2 = 11$	6.29: di	f = 6 (P <	0.0000	(1): $I^2 =$	95%		-
Test for overall effect: Z	= 2.91 (P = 0.0)04)	0 (1)	0.0000		5070		
Test for subgroup differ	ences: cl	$hi^2 = 0.$.00; <i>df</i> =	1 (P = 0)	.99); I ²	= 0%			ravours no ZOL ravours ZOL

FIGURE 2: Forest plot of comparison for bone-to-implant contact between control and treatment groups.

Cturday on out on our		ZOL		No ZOL			Mainht (0/)	Mean difference	Mean difference			
Study or subgroup	Mean	SD	Total	Mean	SD	Total	weight (%)	IV, random, 95% CI		IV, 1	random, 95	% CI
Normal												
Ayan et al. [8]	27.38	2.22	9	23.19	4.95	9	32.9	4.19 [0.65, 7.73]			-	
Marcio et al. [14]	48.37	2.31	3	47.17	1.08	3	33.6	1.20 [-1.69, 4.09]			_	
Yaman et al. [16]	45.8	3.77	7	31.14	1.04	7	33.5	14.66 [11.76, 17.56]			-	
Subtotal (95% CI)			19			19	100.0	6.70 [-1.75, 15.15]				
Heterogeneity: $tau^2 =$	53.24; cl	$hi^2 = 4$	4.79; df =	= 2 (<i>P</i> < 0	0.0000	1); $I^2 = 9$	96%					
Test for overall effect:	Z = 1.55	P = 0).12)									
OST												
Cardemil et al. [9]	72.89	6.27	28	66.27	6.33	28	14.5	6.62 [3.32, 9.92]			-	
Carvas et al. [10]	38.95	4.24	6	25.98	2.35	6	14.4	12.97 [9.09, 16.85]				
Chen et al. [11]	81	9.32	12	37.6	6.73	4	12.9	43.40 [34.96, 51.84]				
Dikicier et al. [12]	31.1	1.7	12	27.39	0.96	12	14.8	3.71 [2.61, 4.81]				
Li et al. [13]	43.11	2.99	6	23.12	2.32	6	14.5	19.99 [16.96, 23.02]			-	
Qi et al. [15] system	ic 53.61	2.74	8	28.81	1.97	8	14.7	24.80 [22.46, 27.14]				F .
Yildiz et al. [17]	43.08	5.37	12	36.03	5.81	12	14.2	7.05 [2.57, 11.53]				
Subtotal (95% CI)			84			76	100.0	16.52 [8.07, 24.98]			-	
Heterogeneity: tau ² =	125.60;	$chi^2 = 1$	384.47; 6	df = 6 (P	< 0.00	$001); I^2$	= 98%					
Test for overall effect:	Z = 3.83	P = 0	0.0001)									
									-50	_25	0 2	5 5
									Eavouro	-25 No 701	U Z	
Test for subgroup diff	erences	chi ² –	2 59. df	- 1 (P -	0 11).	$I^2 - 61$	1%		ravours	INU ZOL	Fav	ours ZC

FIGURE 3: Subgroup analysis with regard to the animal model (osteoporotic or normal) in the systemic group. Bone-to-implant contact was significantly improved in osteoporotic animals.

dental implant procedures in patients that have been receiving intravenous BPs. A recent review indicates that one hundred percent of the studies related to combined use of BPs have shown cases of osteonecrosis [35]. Others authors [36–41] suggested that bisphosphonate exposure and implant placement do not affect implant success and do not

Charles an and announ		ZOL		1	No ZO	L	\mathbf{M}_{1}	Mean difference	Mean diffe	rence	
Study or subgroup	Mean	SD	Total	Mean	SD	Total	weight (%)	IV, random, 95% CI	IV, random,	95% CI	
<0.1 mg/kg											
Marcio et al. [14]	48.37	2.31	3	47.17	1.08	3	100.0	1.20 [-1.69, 4.09]			
Subtotal (95% CI)			3			3	100.0	1.20 [-1.69, 4.09]	•		
Heterogeneity: not appl	icable										
Test for overall effect: Z	5 = 0.82	(P=0.	42)								
>0.1 mg/kg											
Ayan et al. [8]	27.38	2.22	9	23.19	4.95	9	11.2	4.19 [0.65, 7.73]			
Cardemil et al. [9]	72.89	6.27	28	66.27	6.33	28	11.3	6.62 [3.32, 9.92]	-	-	
Carvas et al. [10]	38.95	4.24	6	25.98	2.35	6	11.1	12.97 [9.09, 16.85]			
Chen et al. [11]	81	9.32	12	37.6	6.73	4	9.8	43.40 [34.96, 51.84]		_	_ →
Dikicier et al. [12]	31.1	1.7	12	27.39	0.96	12	11.5	3.71 [2.61, 4.81]	-		
Li et al. [13]	43.11	2.99	6	23.12	2.32	6	11.3	19.99 [16.96, 23.02]			
Qi et al. [15] systemic	53.61	2.74	8	28.81	1.97	8	11.4	24.80 [22.46, 27.14]		-	
Yaman et al. [16]	45.8	3.77	7	31.14	1.04	7	11.3	14.66 [11.76, 17.56]			
Yildiz et al. [17]	43.08	5.37	12	36.03	5.81	12	11.0	7.05 [2.57, 11.53]		_	
Subtotal (95% CI)			100			92	100.0	14.86 [8.20, 21.51]		•	
Heterogeneity: $tau^2 = 9^{\circ}$	9.46: ch	$i^2 = 40$	5.74: df	= 8 (P <	0.0000	1): $I^2 = 0$	98%				
Test for overall effect: Z	= 4.38	(P < 0.	0001)	- (-//-	/ -				
								-50	-25 0	25	50
								50	Favours No ZOI	Eavours 701	50
Test for subgroup differ	ences: c	$hi^2 = 1$	3.62; df	r = 1 (P = 1)	0.0002	2); $I^2 = 9$	2.7%		1 avouis 140 LOL	1 avouis 202	

FIGURE 4: Subgroup analysis with regard to the animal model (<0.1 mg/kg or >0.1 mg/kg) in the systemic group.

Study or subgroup	ZOL			1	No ZO	L	Maight (04)	Mean difference	Mean difference
Study of subgroup	Mean	SD	Total	Mean SD 7		Total	weight (%)	IV, random, 95% CI	IV, random, 95% CI
Systemic									
Blazsek et al. [18]	63.8	12.87	5	21.67	9.58	5	18.5	42.13 [28.07, 56.19]	
Chen et al. [11]	68.9	7.63	12	23.2	7.46	4	19.8	45.70 [37.21, 54.19]	
von Knoch et al. [19]	1.74	0.69	12	2.67	0.88	12	20.7	-0.93 [-1.56, -0.30]	
Li et al. [13]	36.98	4.03	6	21.81	2.6	6	20.5	15.17 [11.33, 19.01]	-
Qi et al. [15] systemic	53.49	3.07	8	21.77	2.65	8	20.6	31.72 [28.91, 34.53]	+
Subtotal (95% CI)			43			35	100.0	26.28 [7.58, 44.99]	
Heterogeneity: $tau^2 = 4$ Test for overall effect: 2	40.82; Z = 2.75	$chi^2 = 6$ 5 (P = 0	585.44; a .006)	lf = 4 (P)	< 0.00	001); I ² :	= 99%		
Local									
Pyo et al. [25]	54.94	14.19	10	30.74	17.9	10	15.5	24.20 [10.04, 38.36]	
Qi et al. [15] Local	35.37	4	8	23.19	2.96	8	55.1	12.18 [8.73, 15.63]	
Stadlinger et al. [23]	23	11.3	7	4.7	3.5	8	29.4	18.30 [9.58, 27.02]	
Subtotal (95% CI)			25			26	100.0	15.84 [9.47, 22.21]	
Heterogeneity: $tau^2 = 1$ Test for overall effect: 2	l 6.06; cl Z = 4.88	$hi^2 = 3.8$ P < 0.8	92; <i>df</i> = .00001)	2(P=0.	14); <i>I</i> ²	= 49%			•
									-50 -25 0 25 50
									Favours No ZOL Favours ZOL
Test for subgroup diffe	rences:	$chi^2 = 1$	1.07; <i>df</i> =	= 1 (<i>P</i> =	0.30);	$I^2 = 6.89$	6		

FIGURE 5: Forest plot of comparison for bone volume/total volume between control and treatment groups.

result in osteonecrosis. However, the duration of their follow-up was short. Najeeb et al. [42] believe that these results should be confirmed by more in-depth research before the dental implant can be used in the clinic. This is also the reason that one of the exclusion criteria is the implants embedded in the mandible or maxilla in our study.

Rats are the most commonly used animal model for osteoporosis studies because the ovariectomized rat exhibits most of the characteristics of human postmenopausal osteoporosis. However, the lack of intracortical remodeling process in this animal compromises the physiologic investigation of the cortical bone. By contrast, rabbits do have some inherent advantages as the osteoporosis animal model. For example, they achieve skeletal maturity shortly after reaching complete sexual development and show significant intracortical remodeling [43]. Thus, some researchers prefer rabbits as their ideal model. With respect to our results, studies with rats or rabbits have achieved similar yet slightly different outcomes, indicating that different animal models

Study or subgroup		ZOL		1	No ZO	L	Weight (%)	Mean difference	Mean difference
Study of subgroup	Mean	SD	Total	Mean	SD	Total	weight (%)	IV, random, 95% CI	IV, random, 95% CI
Systemic									
Cardemil et al. [7]	66.35	6.4	28	60	5.79	28	21.2	6.35 [3.15, 9.55]	-
Carvas et al. [8]	30.44	1.42	6	24.4	1.51	6	78.8	6.04 [4.38, 7.70]	
Subtotal (95% CI)			34			34	100.0	6.11 [4.63, 7.58]	•
Heterogeneity: tau ² =	0.00; c	$hi^2 = 0.$.03; <i>df</i> =	= 1 (P =	0.87); 1	$^{2} = 0\%$			
Test for overall effect	: Z = 8.1	13 (P <	0.0000	1)					
Local									_
Arnoldi et al. [18]	36	4	6	28	4	6	14.8	8.00 [3.47, 12.53]	
Bobyn et al. [25]	25.4	12.7	5	14.6	7.3	5	12.2	10.80 [-2.04, 23.64]	
Gao et al. [20]	52.6	7.36	10	19.79	8.75	10	14.2	32.81 [25.72, 39.90]	
Miettinen et al. [19]	14.3	1.1	14	13.8	1.5	14	15.3	0.50 [-0.47, 1.47]	
Sörensen et al. [24]	72.27	6.38	10	46.26	6.84	10	14.5	26.01 [20.21, 31.81]	
Stadlinger et al. [22]	23.8	8.6	7	4.5	4.2	8	14.2	19.30 [12.30, 26.30]	-
Ying et al. [22]	36.6	5.72	10	21.34	6.19	10	14.7	15.26 [10.04, 20.48]	
Subtotal (95% CI)			62			63	100.0	16.02 [5.98, 26.05]	•
Heterogeneity: $tau^2 =$ Test for overall effect	= 171.35 : <i>Z</i> = 3.1	; chi ² = 13 (<i>P</i> =	201.60	; $df = 6$	(<i>P</i> < 0.	00001);	$I^2 = 97\%$		
								-	
									-50 -25 0 25 50
									Favours No ZOL Favours ZOL
Test for subgroup diff	ferences	s: chi ² =	= 3.67; 0	lf = 1 (P)	= 0.06): $I^2 = 7$	2.7%		

FIGURE 6: Forest plot of comparison for bone area between control and treatment groups.

Study or subgroup		ZOL		No ZOL			Maight (0/)	Mean difference	Mean difference			
Study of Subgroup	Mean	SD	Total	Mean	SD	Total	weight (%)	IV, random, 95% CI		IV, ran	dom, 95% CI	
Normal												
Arnoldi et al. [20]	15	8	6	10	8	6	44.3	5.00 [-4.05, 14.05]				
Miettinen et al. [21]	69.7	4.5	7	75.3	6.3	7	55.7	-5.60 [-11.34, 0.14]		-		
Subtotal (95% CI)			13			13	100.0	-0.90 [-11.22, 9.42]		•		
Heterogeneity: $tau^2 =$ Test for overall effect:	41.23; cl Z = 0.17	$hi^2 = 3.$ 7 ($P = 0$.76; <i>df</i> = 0.086)	= 1 (P = 0	0.05); 1	² = 73%						
Cap at al [22]	51.74	10.07	10	20.00	0 33	10	18.3	22.65 [14.14.31.16]				
Gao et al. $[22]$	51./4 10.20	10.07 8.07	10	29.09	9.55	10	10.5	22.03 [14.14, 31.10]				
Pyo et al. [25]	40.50	0.07	10	44.40	1.07	10	19.9	3.92 [-2.43, 10.29]				
Qi et al. [15] local	42.51	3.35	0	28.81	1.97	0	21.8	15.70 [10.89, 16.51]				_
Stadlinger et al. [23]	4/.8	10.4	/	20.6	2.4	8	18.8	27.20 [19.32, 35.08]				
Ying et al. [24]	49.52	5.11	10	21.56	4.62	10	21.2	27.96 [23.81, 32.11]				
Subtotal (95% CI)			45			46	100.0	18.96 [10.27,27.65]				
Heterogeneity: tau ² = Test for overall effect:	88.29; cl Z = 4.28	$hi^2 = 50$ 3 (P < 0	6.13; <i>df</i> .0001)	= 4 (<i>P</i> <	0.000	()1); $I^2 =$	93%					
								-	-50	_25	0 25	50
									Favours	No ZOL	Favo	urs ZOI
Test for subgroup diffe	oren ceci	chi ² –	8 33. <i>d</i>	- 1 (P -	- 0.004). $I^2 - 8$	8.0%		1 470413	110 201	1 400	

FIGURE 7: Subgroup analysis with regard to the animal model (osteoporotic or normal) in the local group.

may influence implant osseointegration characteristics. However, because of the limited number of included studies, drawing the final conclusion now is too early.

The dosage and frequency of ZOL delivery varied among the included studies, and the best medication administration protocol remains unclear. According to the present study, administration frequency does not exert much influence as long as drug dosage exceeds 0.1 mg/kg. This information is quite important because concerns about the safety of long-term bisphosphonate usage are always present. If single and multiple administrations have similar osseointegrationimproving effects, long-term usage would be unnecessary, thus avoiding the risk of complications, such as osteonecrosis of the jaw or stress fracture [44].

This meta-analysis has several limitations. Firstly, because of the small number of included studies and the limited animal sample sizes, conclusions from this metaanalysis should be interpreted cautiously and should be substantiated by larger studies. Secondly, because of the diverse study characteristics, animal populations, and treatment protocols, significant heterogeneity existed among the included studies. Nevertheless, because the main focus of preclinical meta-analysis is to generate hypotheses, the existence of heterogeneity is quite rational and could provide insight into the design of future clinical trials [45].

7. Conclusions

In conclusion, current animal studies demonstrate that both systemic and local administration of ZOL could improve the osseointegration of the orthopedic implant in animal models. An appropriate animal model (osteoporotic), sufficient drug dosage (exceeding 0.1 mg/kg, only in the method of systemic administration), and enough follow-up time (more than eight weeks) are crucial influencing factors, which should be given particular attention in future animal or clinical studies. Nonetheless, caution should be taken when interpreting the results of this meta-analysis because of inherent differences between preclinical and clinical subjects.

Additional Points

Strengths and Limitations of This Study. (1) Both systemic and local administration of ZOL could improve the osseointegration of the orthopedic implant in animal models. (2) The osteoporotic model could be more effective to improve the osseointegration of the orthopedic implant than the normal model. (3) Sufficient drug dosage (exceeding 0.1 mg/kg, only the method of systemic administration) and enough follow-up time (more than eight weeks) are crucial influencing factors. (4) This metaanalysis should be considered cautiously and should be substantiated by larger studies. (5) Results of this metaanalysis should be interpreted cautiously because of the inherent differences between preclinical and clinical subjects.

Conflicts of Interest

The authors declare that they have no conflicts of interest.

Authors' Contributions

Yao He and Zhengyun Li designed the study. Yao He and Zhengyun Li collected the data. Yao He, Xiang-Dong Wu, Hong Chen, and Wei Huang analyzed the data. Yao He, Wei Bao, Xiang-Dong Wu, Hong Chen, Wei Huang, and Zhengyun Li interpreted the data. Yao He, Wei Bao, Xiang-Dong Wu, Hong Chen, Wei Huang, and Zhengyun Li drafted the manuscript. Hong Chen and Wei Huang revised the manuscript content. Yao He, Wei Bao, Xiang-Dong Wu, Hong Chen, Wei Huang, and Zhengyun Li approved the final version of the manuscript.

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