

The contribution of residential greenness to mortality among men with prostate cancer: a registry-based cohort study of Black and White men

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Background: Black men with prostate cancer (CaP) experience excess mortality compared with White men. Residential greenness, a health promoting contextual factor, could explain racial disparities in mortality among men with CaP.

Methods: We identified Pennsylvania Cancer Registry cases diagnosed between January 2000 and December 2015. Totally, 128,568 participants were followed until death or 1 January 2018, whichever occurred first. Residential exposure at diagnosis was characterized using the Normalized Difference Vegetation Index (NDVI) with 250 m resolution. We estimated hazard ratios (HRs) using Cox models, adjusting for area-level socioeconomic status, geographic healthcare access, and segregation. To determine whether increasing residential greenness could reduce racial disparities, we compared standardized 10-year mortality Black-White risk differences under a hypothetical intervention fixing NDVI to the 75th percentile of NDVI experienced by White men.

Results: We observed 29,978 deaths over 916,590 person-years. Comparing men in the highest to lowest NDVI quintile, all-cause (adjusted HR [aHR]: 0.88, 95% confidence interval [CI]: 0.84, 0.92, $P_{\text{trend}} < 0.0001$), prostate-specific (aHR: 0.88, 95% CI: 0.80, 0.99, $P_{\text{trend}} = 0.0021$), and cardiovascular-specific (aHR: 0.82, 95% CI: 0.74, 0.90, $P_{\text{trend}} < 0.0001$) mortality were lower. Inverse associations between an interquartile range increase in NDVI and cardiovascular-specific mortality were observed in White (aHR: 0.90, 95% CI: 0.86, 0.93) but not Black men (aHR: 0.97, 95% CI: 0.89, 1.06; $P_{\text{het}} = 0.067$). Hypothetical interventions to increase NDVI led to nonsignificant reductions in all-cause (−5.3%) and prostate-specific (−23.2%), but not cardiovascular-specific mortality disparities (+50.5%).

Discussion: Residential greenness was associated with lower mortality among men with CaP, but findings suggest that increasing residential greenness would have limited impact on racial disparities in mortality.

Keywords: Environmental epidemiology, Greenness, Mediation analysis, Prostate cancer, Racial disparities, Vegetation

Introduction

Cancer of the prostate (CaP) is the most commonly diagnosed cancer and the second leading cause of cancer-related death among men in the United States, accounting for one out of

every 10 cancer deaths in men.¹ In the US, Black men experience more than double the mortality from CaP compared with White men.² Although racial gaps in access to CaP care have narrowed over time, disparities in mortality rates among men with CaP have persisted for as long as reliable registry data have been available.^{3–5} Causes of racial disparities in mortality among men with CaP are multifactorial, requiring a multilevel framework that considers genetic and lifestyle risk factors along with historic policies, environments, social attitudes, and community norms that differently shape experiences of Black and White men.^{6–11} Recent advances in epidemiologic methods have provided investigators with analytic tools to quantify the impact of social and environmental policy changes on racial disparities.^{12–15} Although most research to date on cancer disparities has focused on biological and social factors, few studies have

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We are unable to share data because it contains patient identifiers. Please contact the corresponding author to request code used for this analysis.

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What this study adds

Racial disparities in prostate cancer are among the largest for any major cancer and could be mediated by neighborhood context. We conducted a cohort study in the Pennsylvania Cancer Registry to estimate associations between residential greenness and cause-specific mortality among Black and White men with prostate cancer. We observed statistically significant inverse associations between residential greenness and all-cause, prostate, and cardiovascular mortality. Though residential greenness does not appear to mediate disparities in all-cause and prostate cancer mortality, residential greenness could mediate cardiovascular mortality disparities due to differences in how Black and White men with prostate cancer interact with green spaces.

investigated the mediating role of the built and contextual environment on racial disparities in cancer.¹⁶

A growing literature describes numerous health benefits of neighborhood greenness, defined as the extent of green, natural vegetation within a given area. More comprehensive than “green space,” the term neighborhood “greenness” includes all vegetation in a given area, regardless of type (e.g., parks, forests, gardens, and street trees). Neighborhood greenness is hypothesized to confer health benefits through promotion of healthy lifestyles and social cohesion, and reduction of harmful environmental exposures and biopsychosocial stressors.^{17–20} Cohort studies have reported inverse associations between neighborhood greenness and several diseases, including all-cause mortality, cardiovascular disease (CVD), and depression.^{21–24} In the US, neighborhoods with higher proportion of Black residents have lower levels of neighborhood greenness,²⁵ suggesting that neighborhood greenness could mediate racial disparities in mortality among men with prostate cancer.

We studied the association between residential greenness and mortality in a cohort of Black and White men with CaP in Pennsylvania. Because earlier studies reported stronger associations between neighborhood greenness and specific causes of death,^{21,22,26,27} we assessed the magnitude of the association between residential greenness and all-cause mortality, prostate-specific mortality, and CVD mortality. We further evaluated whether the mortality disparity between Black and White men with CaP could be mediated by residential greenness.

Methods

Study design and participants

We used data from the population-based Pennsylvania Cancer Registry. We included 145,399 Black and White men with CaP diagnosed from 2000 to 2015 and followed them until death, 10 years postdiagnosis, or 1 January 2018, whichever came first. Participant addresses were geocoded using ArcGIS software version 10.2. We excluded cases who were diagnosed with *in situ* cancers ($n = 69$), missing address at diagnosis ($n = 85$), or missing stage or grade ($n = 16,677$). A total of 128,568 (88%) men with CaP were included in the study. The Institutional Review Board of Harvard T. H. Chan School of Public Health approved this study protocol. Because existing data sources were used, no written consent was required for participation in the study.

Mortality assessment

CaP diagnoses were staged according to the 2000 Surveillance, Epidemiology, and End Results (SEER) summary staging guidelines.²⁸ Race was extracted from facility medical records and included in data provided by state health providers to the Pennsylvania Cancer Registry. Each year, the Pennsylvania Cancer Registry conducts a Death Clearance in which reportable cause of death information from Pennsylvania Death Certificates is linked with Pennsylvania Cancer Registry files. If deaths occur out of state, linkage is done through data exchanges. Causes of death were categorized based on ICD-09 and ICD-10 codes. For CaP-specific mortality, we included deaths coded as 185 (ICD-09) and C61 (ICD-10). For CVD mortality, we included deaths coded as 390–459 and I00–I99.

Exposure assessment

To estimate exposure to residential greenness for CaP cases at time of diagnosis, we used the normalized difference vegetation index (NDVI), a satellite-derived spatial measure of neighborhood greenness.²⁹ NDVI values range from -1 to 1 and quantify the amount of infrared light absorbed vs reflected by plant life. NDVI values approaching 1 correspond to lush forests, close to 0 reflect

barren areas, and below 1 indicate bodies of water. In this study, to focus specifically on associations related to green vegetation, we set values below 0 to missing. In sensitivity analysis, results were unchanged after applying this procedure. Moderate Resolution Imaging Spectroradiometer data capturing NDVI at a 250-m resolution were obtained using Google Earth Engine. We used Google Earth Engine’s cloud cover algorithm to extract the least cloudy image in January, April, July, and September for every year from 2000 to 2015, representing seasonal variation in residential greenness. Exposure was modeled using NDVI averaged over seasonal measures during calendar year of diagnosis (baseline), as well as cumulative updated average NDVI measurements over each participant’s entire follow-up period as a sensitivity analysis. Participants were assigned the baseline NDVI value for the 250 m² pixel containing their residential address. Because we did not have information on participant mobility, in the case of cumulative updated average, we assigned the seasonal average NDVI over all years of follow-up at their residential address. We chose to model baseline NDVI as the primary exposure because we did not have time-varying information for any other variables in the analysis and wished to limit threat of collider stratification bias.^{30,31}

We used Krieger’s ecosocial theory⁶ to develop a conceptual framework that integrates socioeconomic position, geographic barriers to access, along with demographic and clinical risk factors into our analytic framework. Area-level socioeconomic data at census Block Group level in 2000 were obtained from the National Historical Geographic Information System Database³² and spatially joined to cohort participant addresses using the R statistical package. We chose to link data at the census Block Group level because this is the smallest geographic unit for which the US census publishes data. In the 2010 census, Pennsylvania reported 9,740 Block Groups, which contained a median of 1,160 people (interquartile range [IQR]: 866–1,574). When Block Group data were not available ($n = 101$), we used census tract-level data.

Statistical analysis

We estimated adjusted hazard ratios (aHRs) and 95% confidence intervals (CIs) for the association between NDVI and each of the mortality outcomes (all-cause, prostate, CVD) using multiple Cox proportional hazards models in SAS. NDVI was modeled using quintiles and as a continuous exposure using restricted cubic splines with three knots to test for nonlinearity. We chose to present results from models using quintiles alongside continuous NDVI because quintiles facilitate comparison of HRs between extremes of the NDVI distribution, mitigate the influence of outliers, and allow investigation of possible thresholds that could be used to guide a policy recommendation. When no evidence of nonlinearity was observed, we modeled continuous NDVI using a linear term scaled in units of IQR for the study population (0.14 U). We estimated *P* values for linear trend in categorical models using the median for each quintile. Because NDVI and CaP rates vary by urbanicity, we stratified our analyses by population density ($\geq 1,000$ people/mi² vs. $< 1,000$ people/mi²). This threshold was chosen to differentiate more rural settings from suburban and urban settings.

Because NDVI is tied to location and therefore socioeconomic status, we chose to control for possible confounding variables guided by the literature on socioeconomic position and health,³³ racial disparities in prostate cancer,⁵ and neighborhood greenness and mortality.³⁴ We considered advanced stage and marital status assessed at diagnosis as mediators because they could be interpreted as consequences of residential greenness exposure. We then sequentially adjusted for sets of confounding variables. In model 1, we stratified by age in 10-year categories and diagnosis year (categorical), and adjusted for race (Black vs. White); census Block Group median income (continuous: US\$); median home value (continuous: US\$); percent poverty (continuous);

percent adults age 25 years and older with less than high school education (continuous); the joint race and income Index of Concentration at Extremes (ICE), a measure of inequality based on income- or race-based privilege in a given geographic area (quintiles)³⁵; four indicator variables for receipt of care at a currently NCI-designated cancer center (The University of Pennsylvania, University of Pittsburgh Medical Center, Fox Chase Cancer Center, or Thomas Jefferson University Hospital); population density (continuous), distance between each participant's geocoded address to the closest cancer center using road network distances (continuous, minutes, calculated using ArcMap 10.2). Although estimating road network distances makes strong assumptions (car-based transport, constant speed at road speed limit), this measure does allow rank ordering participants based on the proximity to cancer care. In model 2, we additionally adjusted for stage (categorical: localized, regional, distant) and grade (categorical: I–IV). In model 3, we further adjusted for marital status, using logistic regression models and Monte Carlo imputation with 10 repetitions to impute missing marital status ($n = 46,519$) conditional on the covariates used in model 1. We assessed whether primary associations varied by race (binary: Black vs. White), stage (binary: localized vs. regional/distant), and population density ($\geq 1,000$ people/mi², $< 1,000$ people/mi²). Though census Block Group socioeconomic variables were correlated (range [absolute value]: 0.43–0.84), no pair-wise correlation exhibited perfect collinearity (eTable 1; <http://links.lww.com/EE/A71>). Tests for effect modification were performed by fitting interaction terms between these modifiers and NDVI (continuous and as quintiles).

To evaluate the role of residential greenness as a potential mediator of racial disparities in cause-specific mortality among men with CaP, we estimated racial disparities among men with CaP after hypothetical interventions that fix NDVI for all participants to a specific value using previously described statistical methods.^{13–15,36,37} This approach assumes no unmeasured confounding of race and cause-specific mortality, no unmeasured confounding between residential greenness and cause-specific mortality, and correct model specification. Technical details are provided in eMethods 1; <http://links.lww.com/EE/A71>.

First, we fit the outcome model described above (Cox model 1) for each mortality outcome, omitting NDVI. Resulting model parameters were used to estimate 10-year mortality among Black and White men, standardized to covariates described above. The difference in standardized 10-year mortality for Black and White men with CaP was defined as the racial disparity. Because most Black men with CaP in our study lived in high population density areas, we repeated this procedure separately among men living in high and low population density areas. Next, we estimated the racial disparity that would remain after hypothetical interventions to fix NDVI to target values for all study participants. We again estimated expected racial disparities between Black and White men using our outcome model, with two additional parameters (continuous NDVI and an NDVI-race interaction). Bootstrapping with 500 repetitions was used to estimate 95% CIs.

Three levels of NDVI were chosen to set bounds on estimated changes in mortality among men with prostate cancer that could result from a policy change (1); the observed racial disparity with no change in NDVI, (2) the 25th percentile of NDVI among Black men with CaP (lower bound), and (3) the 75th percentile of NDVI among White men with CaP (upper bound). We then estimated the proportion of racial disparity that could be eliminated by implementing policy change (3).^{12,38} Details regarding sensitivity analyses for competing risks, and estimation of bounds for bias due to unmeasured confounding using E-values³⁹ are provided in eMethods 1; <http://links.lww.com/EE/A71>.

Results

After exclusions, we observed 29,978 deaths over 916,590 person-years of postdiagnosis follow-up. Study population characteristics are presented in Table 1 overall and by NDVI in the year before diagnosis. Median age at diagnosis was 66 and did not vary by quintile of NDVI. Black men made up 11% of the study population and were less likely than Whites to reside in neighborhoods in the highest quintile of NDVI (NDVI Q1: 33% vs. Q5: 3%). Most participants were diagnosed with localized disease (85%). Participants in greener neighborhoods (Q5) had lower population density, higher census Block Group income and median home value than participants in less green neighborhoods (Q1). Study participants were concentrated in the Southeast and Western parts of Pennsylvania, corresponding to the Pittsburgh and Philadelphia metropolitan areas where NDVI was relatively lower than in other regions of the state (Figs. 1 and 2). CVD was leading cause of death ($n = 7,677$), followed by CaP ($n = 6,515$).

In adjusted analysis, we observed statistically significant inverse associations between NDVI in the year of diagnosis and rates of mortality using quintiles and continuous exposure parameterizations (Table 2). Tests for splines were not significant, so we assumed linear dose response between continuous NDVI and mortality. When considering confounding factors (model 1), there was a 12% lower rate of all-cause mortality comparing participants with NDVI Quintile 5 to 1 (Q5 to 1) (aHR: 0.88, 95% CI: 0.84, 0.92, $P_{\text{trend}} < 0.0001$). This association was similar for prostate-specific mortality (aHR: 0.88, 95% CI: 0.80, 0.98, $P_{\text{trend}} = 0.0021$), but the relative stability of HR estimates across NDVI quintiles suggests that this result should be interpreted with caution. For CVD mortality, there was an 18% lower rate comparing NDVI Q5 to 1 (aHR: 0.82, 95% CI: 0.74, 0.90, $P_{\text{trend}} < 0.0001$).

The associations with prostate-specific mortality were attenuated in models additionally adjusting for stage and grade (aHR NDVI Q5 to 1: 0.93, 95% CI: 0.84, 1.03) and then marital status (aHR NDVI Q5 to 1: 0.96, 95% CI: 0.86, 1.06). Adjusting for stage and grade did not result in major changes in inference with respect to all-cause mortality or CVD mortality. However, adjusting for marital status resulted in modest attenuation of the association with all-cause mortality (aHR NDVI Q5 to 1: 0.92, 95% CI: 0.87, 0.96), but not CVD mortality.

In stratified analyses, we found no evidence of effect modification by race, stage, or population density with respect to all-cause mortality (Table 3). The inverse association between an IQR increase in continuous NDVI and prostate-specific mortality was stronger among participants with localized (aHR: 0.92, 95% CI: 0.87, 0.97) compared with distant CaP (aHR: 0.98, 95% CI: 0.93, 1.03, $P_{\text{het}} = 0.032$). In addition, the inverse association was stronger among participants in high (aHR: 0.88, 95% CI: 0.83, 0.93) compared with low (aHR: 0.96, 95% CI: 0.91, 1.01) population density areas ($P_{\text{het}} = 0.028$). There was no association between continuous NDVI and CVD mortality among Black men with CaP (aHR: 0.97, 95% CI: 0.89, 1.06), but there was an inverse association among White men with CaP (aHR: 0.90, 95% CI: 0.86, 0.93, $P_{\text{het}} = 0.067$), suggesting increasing levels of NDVI could increase disparities by preferentially benefiting White but not Black men with CaP.

Racial disparities in 10-year mortality without adjustment for NDVI were greatest for all-cause mortality, resulting in 29.3 excess deaths per 1,000 (95% CI: 22.1, 36.5) among Black men with CaP, and least for CVD mortality (11.5, 95% CI: 6.4, 16.7 excess deaths per 1,000). Disparities were greater in low (all-cause: 33.9, 95% CI: 20.9, 47.8; prostate: 22.1, 95% CI: 13.0, 31.2; CVD: 16.8, 95% CI: 7.3, 26.3 per 1,000) compared with high population density areas (all-cause: 25.1, 95% CI: 15.2, 35.0; prostate: 15.1, 95% CI: 8.1, 22.1; CVD: 8.5, 95% CI: 1.6, 15.3 per 1,000). There were no statistically significant differences in racial disparities among men with CaP after hypothetical interventions fixing residential NDVI to the 25th percentile (Black), observed

Table 1. Descriptive characteristics of Pennsylvania Cancer Registry Cohort stratified by quintile of baseline NDVI quintile, from 2000 to 2015.

	Quintile of NDVI					Total
	Q1	Q2	Q3	Q4	Q5	
Total population	25,708	25,723	25,719	25,717	25,701	128,568
NDVI baseline ^a	0.36 (0.30, 0.41)	0.50 (0.48, 0.52)	0.56 (0.55, 0.58)	0.62 (0.60, 0.63)	0.67 (0.65, 0.68)	0.56 (0.48, 0.63)
Age at diagnosis, years ^a	66 (59, 73)	67 (60, 74)	67 (60, 74)	66 (60, 73)	66 (60, 73)	66 (60, 73)
Black reported race [n (%)]	8,599 (33)	2,512 (10)	1,371 (5)	1,028 (4)	654 (3)	14,164 (11)
Married [n (%)]						
Yes	9,868 (38)	11,887 (46)	13,142 (51)	13,956 (54)	14,515 (56)	63,368 (49)
No ^b	5,910 (23)	3,752 (15)	3,342 (13)	3,001 (12)	2,676 (10)	18,681 (15)
(Missing)	9,930 (39)	10,084 (39)	9,235 (36)	8,760 (34)	8,510 (33)	46,519 (36)
Stage at diagnosis [n (%)]						
Localized	21,665 (84)	22,007 (86)	22,060 (86)	22,097 (86)	21,990 (86)	109,819 (85)
Regional	2,721 (11)	2,679 (10)	2,672 (10)	2,804 (11)	2,902 (11)	13,778 (11)
Distant	1,322 (5)	1,037 (4)	987 (4)	816 (3)	809 (3)	4,971 (4)
Grade [n (%)]						
I	1,329 (5)	1,387 (5)	1,244 (5)	1,348 (5)	991 (4)	6,299 (5)
II	13,227 (51)	13,505 (53)	13,489 (52)	13,425 (52)	13,307 (52)	66,953 (52)
III	10,993 (43)	10,698 (42)	10,863 (42)	10,821 (42)	11,284 (44)	54,659 (43)
IV	159 (1)	133 (1)	123 (0)	123 (0)	119 (0)	657 (1)
Year of diagnosis [n (%)]						
2000–2004	8,791 (34)	9,082 (35)	8,309 (32)	7,884 (31)	7,985 (31)	42,051 (33)
2005–2009	8,059 (31)	7,406 (29)	8,884 (35)	9,842 (38)	10,014 (39)	44,205 (34)
2010–2015	8,858 (34)	9,235 (36)	8,526 (33)	7,991 (31)	7,702 (30)	42,312 (33)
Contextual factors						
Population Density (100 people/mi ²) ^a	24.5 (9.9, 59.3)	6.1 (1.2, 14.1)	2.4 (0.5, 8.6)	1.9 (0.5, 6.3)	1.3 (0.5, 4.3)	4.0 (0.8, 12.3)
Census block group						
Census block income (1,000 US\$) ^a	31.1 (23.6, 39.4)	40.6 (32.2, 51.0)	43.7 (35.5, 55.0)	47.6 (38.0, 60.6)	52.9 (40.4, 72.0)	42,453 (33.2, 55.6)
Census block % poverty ^a	11 (5, 21)	5 (2, 9)	4 (2, 7)	3 (1, 6)	3 (1, 5)	4 (2, 9)
% over 25 with less than high school education ^a	16 (11, 21)	11 (7, 15)	10 (6, 14)	9 (6, 13)	8 (4, 13)	11 (6, 15)
Median home value (1,000 US\$) ^a	63.2 (42.3, 87.3)	94.7 (72.6, 123.6)	104.6 (80.3, 135.8)	116.3 (88.5, 148.6)	134.1 (96.3, 188.0)	99.5 (72.4, 139.0)
Index concentration at the extremes						
Income ^a	-0.09 (-0.14, -0.03)	-0.02 (-0.07, 0.04)	0.00 (-0.05, 0.06)	0.02 (-0.04, 0.09)	0.05 (-0.02, 0.13)	-0.01 (-0.07, 0.06)
Joint race/income ^a	0.00 (-0.18, 0.02)	0.02 (0.00, 0.06)	0.03 (0.01, 0.07)	0.04 (0.02, 0.10)	0.07 (0.03, 0.17)	0.03 (0.00, 0.08)
Race ^a	0.74 (-0.31, 0.94)	0.95 (0.85, 0.98)	0.96 (0.90, 0.98)	0.96 (0.92, 0.98)	0.96 (0.92, 0.98)	0.95 (0.86, 0.98)
Quintiles joint race-income ICE [n (%)]						
Q1	12,886 (50)	4,216 (16)	2,250 (9)	1,243 (5)	573 (2)	21,168 (16)
Q2	6,290 (24)	7,302 (28)	6,742 (26)	5,632 (22)	4,295 (17)	30,261 (24)
Q3	3,256 (13)	5,663 (22)	6,192 (24)	5,818 (23)	4,791 (19)	25,720 (20)
Q4	2,152 (8)	5,072 (20)	5,969 (23)	6,695 (26)	5,835 (23)	25,723 (20)
Q5	1,124 (4)	3,470 (13)	4,566 (18)	6,329 (25)	10,207 (40)	25,696 (20)
Geographic access to oncology services						
Accredited cancer center where patient was diagnosed [n (%)]						
University of Pennsylvania	1,131 (4)	275 (1)	253 (1)	250 (1)	537 (2)	2,446 (2)
University of Pittsburgh Medical Center	1,579 (6)	2,533 (10)	2,818 (11)	2,857 (11)	2,514 (10)	12,301 (10)
Fox Chase	829 (3)	487 (2)	443 (2)	535 (2)	798 (3)	3,092 (2)
Jefferson Health	834 (3)	159 (1)	142 (1)	126 (0)	212 (1)	1,473 (1)
Network distance to closest facility (Minutes) ^{a,b}	3.12 (1.97, 5.01)	5.73 (3.22, 11.39)	7.69 (4.34, 15.61)	8.9 (4.89, 17.32)	9.15 (5.38, 16.35)	6.45 (3.45, 13.65)
Mortality and Follow-up [n (%)]						
Follow-up (months) ^a	89.4 (49.1, 120.0)	91.6 (52.2, 120.0)	95.7 (54.0, 120.0)	101.2 (56.2, 120.0)	106.3 (63.6, 120.0)	96.4 (54.9, 120.0)
All deaths over 10 years	6,971 (27)	6,150 (24)	6,029 (23)	5,624 (22)	5,204 (20)	29,978 (23)
Of deaths, prostate cancer	1,633 (23)	1,305 (21)	1,257 (21)	1,159 (21)	1,161 (22)	6,515 (22)
Of deaths, CVD	1,831 (26)	1,647 (27)	1,505 (25)	1,428 (25)	1,266 (24)	7,677 (26)

^aMedian (IQR).^bSingle, divorced, and widowed.

values of NDVI, or the 75th percentile among (White) (Table 4). Fixing NDVI to the 75th percentile (White) resulted in the lowest cause-specific mortality, and fixing NDVI to the 25th percentile (Black) resulted in the highest cause-specific mortality in all scenarios except for CVD mortality among Black men in low population density areas (Fig. 3).

Estimated proportions of racial disparity in mortality that would be eliminated by fixing residential greenness to the 75th percentile of NDVI (White) were modest for all-cause (5.3%) and prostate-specific (23.2%) mortality. However, for CVD mortality, we estimated a relative 50.5% increase in the racial disparity after this hypothetical intervention (Table 4). These findings are consistent with results from our race-stratified models, in which NDVI was associated with lower CVD mortality among White but not Black men with CaP. Stratification by population density preserved these patterns, though estimated

proportions of racial disparities eliminated for all-cause and prostate mortality were greater in low compared with high population density areas (Table 4).

In sensitivity analysis for competing risks, results for stratified associations between NDVI and prostate- and CVD-specific mortality resulted in slightly weaker estimates compared with primary results and no change to inference so we did not use competing risk models for our main analysis (eTable 2; <http://links.lww.com/EE/A71>). E-values summarizing bounds of bias due to unmeasured confounding for our primary effect estimates are provided in eTable 3; <http://links.lww.com/EE/A71>. Effect sizes for the association between socioeconomic status, a likely confounding variable, and mortality among men with CaP from previous registry-based studies range from 1.14 to 1.52.⁴⁰ E-values for all-cause mortality and prostate-specific mortality lie within this range, meaning that if an unmeasured factor exhibited

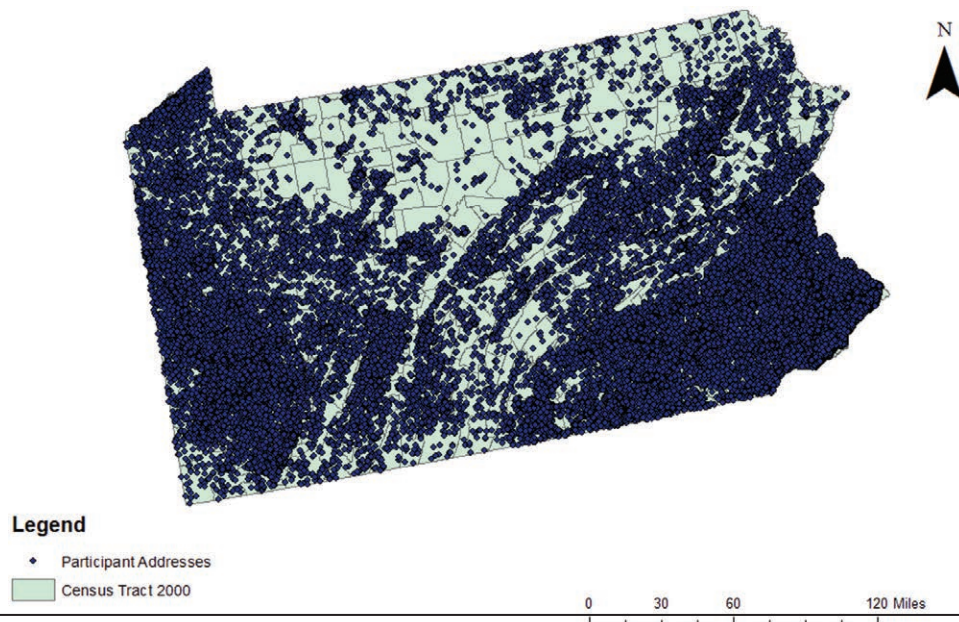


Figure 1. Participant residential address locations in Pennsylvania Cancer Registry prostate cancer cohort study from 2000 to 2015.

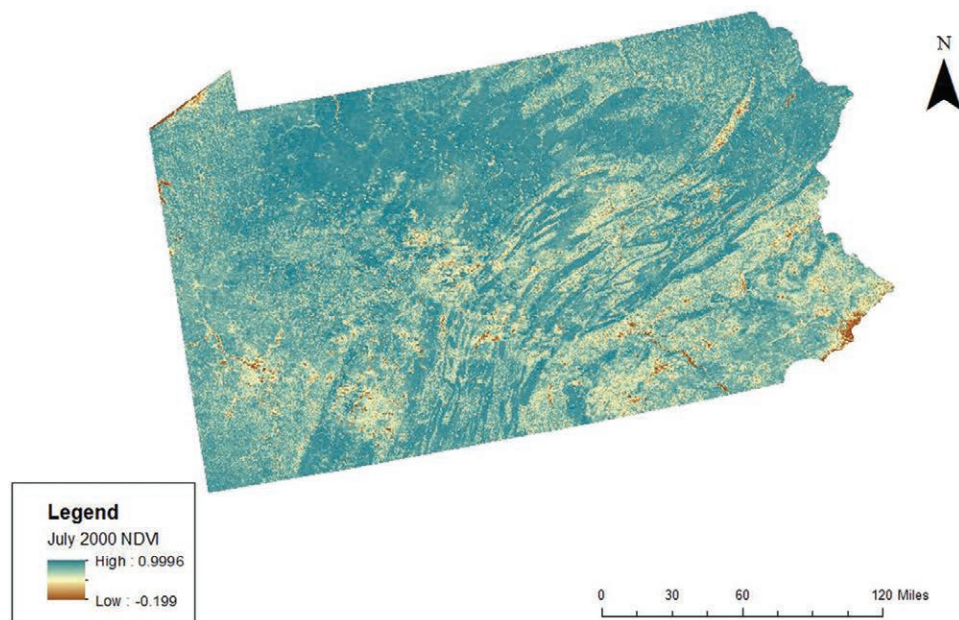


Figure 2. Normalized difference vegetation index July 2000 values at Pennsylvania Cancer Registry participants' residential address locations from 2000 to 2015.

patterns of association with NDVI and mortality similar to that of socioeconomic status, adjusting for that factor could explain away these results. However, this unmeasured factor would need to be sufficiently correlated with NDVI and mortality even after adjusting for the demographic, socioeconomic, and geographic access variables already included in our analysis. Our strongest e-values corresponding to the HR for CVD-mortality comparing men in Q5 to Q1 are 1.75 for point estimate, and 1.47 for CI, suggesting that these results are unlikely to be explained by unmeasured confounding bias. Associations between cumulative updated average NDVI and mortality exhibited non-linear dose response, with increased all-cause mortality and prostate-specific mortality in the lowest and highest quintiles of NDVI (eTable 4; <http://links.lww.com/EE/A71>).

Discussion

In this cohort of Black and White men with CaP, we observed inverse associations between NDVI and lower all-cause, prostate-, and CVD-specific mortality after adjusting for demographics, neighborhood socioeconomic context, and geographic healthcare access. Our results suggest that increasing levels of residential greenness could result in modest, nonsignificant decreases in racial disparities in all-cause and prostate-specific mortality. However, we estimated increases in racial disparities in CVD mortality among men with CaP after hypothetical interventions to increase residential greenness. In our sensitivity analysis using cumulative updated average NDVI, we observed different dose-response patterns compared with analyses using

Table 2. Cox proportional hazards models for association between NDVI at diagnosis and cause-specific mortality among Pennsylvania prostate cancer patients diagnosed between 2000 and 2015.

	Cases/person-years	Full aHR (95% CI)	+ Stage/grade aHR (95% CI)	+ Marital status ^a aHR (95% CI)
All-cause mortality	29,978/916,590			
Linear (per interquartile range)		0.94 (0.92, 0.96)	0.95 (0.93, 0.97)	0.96 (0.94, 0.98)
Q1	6,972/175,795	Ref	Ref	Ref
Q2	6,149/179,406	0.94 (0.90, 0.98)	0.95 (0.92, 0.99)	0.96 (0.93, 1.00)
Q3	6,026/182,719	0.94 (0.90, 0.98)	0.95 (0.91, 0.99)	0.96 (0.93, 1.01)
Q4	5,624/185,936	0.89 (0.85, 0.93)	0.92 (0.88, 0.96)	0.94 (0.90, 0.98)
Q5	5,207/192,733	0.88 (0.84, 0.92)	0.89 (0.85, 0.94)	0.92 (0.87, 0.96)
<i>P</i> _{trend}		<0.0001	<0.0001	0.0004
Prostate-specific mortality	6,515/916,590			
Linear (per interquartile range)		0.93 (0.89, 0.98)	0.95 (0.91, 1.00)	0.96 (0.92, 1.01)
Q1	1,633/175,795	Ref	Ref	Ref
Q2	1,305/179,406	0.89 (0.82, 0.97)	0.93 (0.86, 1.01)	0.94 (0.87, 1.02)
Q3	1,256/182,719	0.88 (0.81, 0.96)	0.92 (0.84, 1.01)	0.94 (0.86, 1.02)
Q4	1,157/185,936	0.83 (0.76, 0.92)	0.92 (0.84, 1.01)	0.94 (0.85, 1.03)
Q5	1,164/192,733	0.88 (0.80, 0.98)	0.93 (0.84, 1.03)	0.96 (0.86, 1.06)
<i>P</i> _{trend}		0.0021	0.10	0.28
Cardiovascular mortality	7,677/916,590			
Linear (per interquartile range)		0.90 (0.87, 0.94)	0.90 (0.87, 0.94)	0.91 (0.88, 0.95)
Q1	1,832/175,795	Ref	Ref	Ref
Q2	1,646/179,406	0.95 (0.88, 1.03)	0.95 (0.88, 1.03)	0.96 (0.89, 1.04)
Q3	1,505/182,719	0.88 (0.81, 0.96)	0.88 (0.81, 0.96)	0.90 (0.83, 0.98)
Q4	1,428/185,936	0.86 (0.78, 0.93)	0.86 (0.79, 0.94)	0.88 (0.81, 0.96)
Q5	1,266/192,733	0.82 (0.74, 0.90)	0.82 (0.74, 0.90)	0.84 (0.76, 0.92)
<i>P</i> _{trend}		<0.0001	<0.0001	<0.0001

Models adjusted for age (deciles), diagnosis year, race, census block group socioeconomic status (% poverty, median income, median home value, % 25 and older with less than high school education, joint race-income index concentration at extremes (quintiles), site at diagnosis (University of Pennsylvania Medical Center, University of Pittsburgh Medical Center, Fox Chase, Jefferson Health), network distance to closest cancer facility (minutes), population density.

^aMissing values for marital status were obtained using multiple imputation.

NDVI at time of diagnosis. Increased all-cause and prostate specific mortality observed with increasing cumulative updated average NDVI could be attributable to reverse causation, resulting from tree planting and greening interventions such as Philadelphia’s “Green Works” program, implemented from 2009 to present. These interventions were targeted precisely at those urban areas which were most deprived and experienced worse outcomes during the study period.⁴¹

Although few studies have reported associations between neighborhood greenness and mortality among men with CaP, our findings are consistent with results from earlier prospective population-based and occupational cohort studies in the US, Canada, and Europe, which have also reported inverse associations between neighborhood greenness and all-cause mortality.^{21,23,26,27} Most men in our study were diagnosed with localized CaP. The 10-year survival is relatively high among these men, and deaths from prostate cancer are few relative to deaths from other causes like CVD.^{42,43} This suggests mortality risks for these men could be similar to the general population. Cohort studies in Canada, Europe, and the US have also reported inverse associations between neighborhood greenness and cardiovascular mortality.^{22,23,26,27,44} Though we did not have data to evaluate lifestyle risk factors, prior research shows that physical activity is associated with lower mortality risk,⁴⁵⁻⁴⁷ and obesity is associated with higher risk^{48,49} among men with CaP. In our study, CVD-specific mortality was the leading cause of death among men with CaP, so inverse associations between residential greenness and mortality reported here could be due to reduced CVD-specific mortality, possibly through pathways related to physical activity and obesity.^{17,18,20} Empirical estimates of the proportion of inverse association between NDVI and mortality mediated by physical activity are few, with a single large prospective cohort of female nurses reporting 2.1% mediated by

physical activity (based on questionnaire responses) using a 1,250 m buffer for NDVI.²¹ Future investigations using more precise estimates of physical activity, for example, through accelerometry, may yield stronger relationships.⁵⁰

The second question we sought to answer was whether increasing residential greenness could reduce racial disparities in mortality among men with CaP. No differences in the association between NDVI and either all-cause or prostate-specific mortality comparing Black to White men with CaP were observed. However, for CVD mortality, we observed an inverse association with NDVI in White but not Black men. Wide confidence intervals for the cause-specific racial disparities from our simulation-based approach limited our ability to statistically evaluate differences in disparities under hypothetical interventions to fix NDVI to different thresholds. However, estimates of the proportion of disparity eliminated suggest that increasing residential greenness could lead to modest reductions in disparities in all-cause mortality. Estimated reductions in racial disparities for CaP mortality were offset by increases in disparities for CVD mortality. Cohort studies in the general population from Canada and Europe looking at all-cause and CVD mortality have also reported stronger inverse associations among high income or privileged racial groups.^{26,27} Better understanding of how contextual environment and CaP outcomes vary by race in different US and global settings will be essential to informing policy interventions.

Although we lacked data to explain racial differences in the association between NDVI and CVD mortality, the literature on differing patterns of park use between Black and White men and women offers some guidance. Parks are a major contributor to urban neighborhood greenness. The ways in which Black and White men experience neighborhood greenness could be different, which in turn could have consequences for potential health

Table 3.**Cox proportional hazards models for association between NDVI at diagnosis and cause-specific mortality among Pennsylvania prostate cancer patients diagnosed between 2000 and 2015, stratified by race, stage and population density**

	All-cause mortality aHR (95% CI)	Prostate-specific mortality aHR (95% CI)	Cardiovascular mortality aHR (95% CI)
Race^a			
White			
Linear (per interquartile range)	0.94 (0.92, 0.96)	0.94 (0.90, 0.98)	0.90 (0.86, 0.93)
Q1	Ref	Ref	Ref
Q2	0.94 (0.90, 0.98)	0.90 (0.83, 0.99)	0.94 (0.87, 1.02)
Q3	0.93 (0.89, 0.98)	0.88 (0.80, 0.96)	0.88 (0.80, 0.95)
Q4	0.89 (0.85, 0.93)	0.84 (0.76, 0.92)	0.84 (0.76, 0.91)
Q5	0.88 (0.84, 0.92)	0.89 (0.80, 0.99)	0.81 (0.74, 0.90)
P_{trend}	<0.0001	0.0046	<0.0001
Black			
Linear (per interquartile range)	0.94 (0.89, 0.98)	0.90 (0.82, 0.98)	0.97 (0.89, 1.06)
Q1	Ref	Ref	Ref
Q2	0.94 (0.86, 1.04)	0.82 (0.68, 0.99)	1.01 (0.84, 1.21)
Q3	0.96 (0.84, 1.08)	0.99 (0.78, 1.26)	0.88 (0.68, 1.14)
Q4	0.95 (0.82, 1.10)	0.83 (0.62, 1.12)	1.23 (0.94, 1.60)
Q5	0.75 (0.62, 0.92)	0.69 (0.46, 1.03)	0.70 (0.46, 1.06)
P_{trend}	0.022	0.048	0.66
P_{het} (quintiles)	0.47	0.41	0.076
P_{het} (linear)	0.84	0.33	0.067
Stage^b			
Localized			
Linear (per interquartile range)	0.94 (0.92, 0.96)	0.92 (0.87, 0.97)	0.90 (0.87, 0.94)
Q1	Ref	Ref	Ref
Q2	0.95 (0.91, 0.99)	0.91 (0.81, 1.02)	0.94 (0.87, 1.01)
Q3	0.95 (0.90, 0.99)	0.88 (0.78, 0.99)	0.88 (0.81, 0.96)
Q4	0.91 (0.87, 0.95)	0.85 (0.75, 0.96)	0.85 (0.77, 0.93)
Q5	0.88 (0.83, 0.92)	0.84 (0.74, 0.96)	0.82 (0.74, 0.90)
P_{trend}	<0.0001	0.0038	<0.0001
Regional/distant			
Linear (per interquartile range)	0.95 (0.92, 0.98)	0.98 (0.93, 1.03)	0.91 (0.83, 0.99)
Q1	Ref	Ref	Ref
Q2	0.93 (0.86, 1.01)	0.94 (0.84, 1.04)	1.07 (0.86, 1.32)
Q3	0.94 (0.87, 1.02)	0.95 (0.85, 1.06)	0.93 (0.74, 1.16)
Q4	0.89 (0.82, 0.96)	0.90 (0.80, 1.01)	0.94 (0.74, 1.18)
Q5	0.94 (0.86, 1.02)	0.98 (0.87, 1.11)	0.83 (0.65, 1.06)
P_{trend}	0.024	0.42	0.12
P_{het} (quintiles)	0.34	0.33	0.79
P_{het} (linear)	0.55	0.032	0.89
Population density^c			
Low			
Linear (per interquartile range)	0.94 (0.91, 0.96)	0.96 (0.91, 1.01)	0.89 (0.85, 0.94)
Q1	Ref	Ref	Ref
Q2	0.95 (0.89, 1.00)	0.92 (0.82, 1.04)	0.95 (0.85, 1.06)
Q3	0.94 (0.89, 1.00)	0.91 (0.81, 1.03)	0.88 (0.79, 0.98)
Q4	0.90 (0.85, 0.96)	0.86 (0.76, 0.97)	0.86 (0.77, 0.97)
Q5	0.89 (0.83, 0.94)	0.91 (0.80, 1.04)	0.82 (0.73, 0.92)
P_{trend}	<0.0001	0.090	0.0002
High			
Linear (per interquartile range)	0.94 (0.92, 0.97)	0.88 (0.83, 0.93)	0.93 (0.88, 0.98)
Q1	Ref	Ref	Ref
Q2	0.93 (0.88, 0.98)	0.85 (0.76, 0.95)	0.96 (0.87, 1.06)
Q3	0.93 (0.88, 0.99)	0.83 (0.72, 0.95)	0.90 (0.79, 1.02)
Q4	0.86 (0.79, 0.93)	0.78 (0.66, 0.93)	0.82 (0.70, 0.95)
Q5	0.85 (0.76, 0.96)	0.74 (0.56, 0.96)	0.77 (0.60, 0.98)
P_{trend}	<0.0001	<0.0001	0.0028
P_{het} (quintiles)	0.87	0.58	0.91
P_{het} (linear)	0.84	0.028	0.24

Models adjusted for age (deciles), diagnosis year, race, census block group socioeconomic status (% poverty, median income, median home value, % 25 and older with less than high school education, joint race-income index concentration at extremes (quintiles), site at diagnosis (University of Pennsylvania Medical Center, University of Pittsburgh Medical Center, Fox Chase, Jefferson Health), network distance to closest cancer facility (minutes), population density.

^aModel includes race and NDVI interaction (1 degree of freedom).

^bModel includes stage and NDVI interaction (1 degree of freedom).

^cModel includes population density (high: $\geq 1,000$ people/mi², low: $< 1,000$ people/mi²) and NDVI interaction (1 degree of freedom).

Table 4.

Cause-specific 10-year mortality risks^a, disparities^b, and 95% CIs under three levels of NDVI at diagnosis among Black and White men with prostate cancer in Pennsylvania, 2000 to 2015.

10-year mortality risk ^a	All-cause mortality			Prostate-specific mortality			Cardiovascular-specific mortality		
	Black/100	White/100	Disparity/1,000	Black/100	White/100	Disparity/1,000	Black/100	White/100	Disparity/1,000
Total population									
No NDVI intervention	21.1 (20.4, 21.8)	18.2 (18.0, 18.3)	29.3 (22.1, 36.5)	6.4 (5.8, 6.9)	4.6 (4.5, 4.8)	17.3 (11.6, 22.9)	6.5 (6.0, 7.0)	5.4 (5.2, 5.5)	11.5 (6.4, 16.7)
NDVI threshold ^c									
25th percentile (Black)	22.8 (21.7, 23.9)	19.6 (19.1, 20.1)	31.8 (21.7, 41.9)	7.2 (6.3, 8.1)	5.1 (4.7, 5.4)	21.3 (13.2, 29.4)	7.1 (6.4, 7.9)	6.3 (5.9, 6.6)	8.8 (1.8, 15.8)
Observed	21.1 (20.2, 21.9)	18.2 (18.0, 18.3)	29.0 (20.1, 37.8)	6.2 (5.6, 6.8)	4.6 (4.5, 4.7)	15.6 (9.0, 22.2)	6.9 (6.2, 7.5)	5.4 (5.2, 5.5)	14.9 (8.4, 21.5)
75th percentile (White)	20.3 (19.1, 21.5)	17.6 (17.3, 17.8)	27.8 (15.7, 39.9)	5.8 (4.9, 6.6)	4.4 (4.3, 4.6)	13.3 (4.5, 22.0)	6.7 (5.9, 7.6)	5.0 (4.8, 5.2)	17.4 (8.4, 26.3)
Proportion eliminated ^d									
75th percentile (White)			-0.053			-0.232			+0.505
Urban (≥1,000 people/mi ²)									
No NDVI intervention	22.1 (21.3, 22.9)	19.6 (19.2, 19.9)	25.1 (15.2, 35.0)	6.5 (5.9, 7.1)	5.0 (4.7, 5.3)	15.1 (8.1, 22.1)	6.8 (6.2, 7.4)	5.9 (5.7, 6.2)	8.5 (1.6, 15.3)
NDVI threshold ^c									
25th percentile (Black)	23.0 (21.9, 24.1)	20.5 (19.8, 21.2)	25.6 (13.0, 38.2)	7.1 (6.2, 7.9)	5.5 (5.0, 5.9)	15.8 (6.5, 25.1)	7.1 (6.3, 7.9)	6.5 (6.0, 7.0)	6.3 (-2.4, 15.1)
Observed	22.0 (21.1, 22.9)	19.6 (19.2, 20.0)	24.0 (13.6, 34.5)	6.5 (5.8, 7.1)	5.0 (4.7, 5.3)	14.6 (7.4, 21.8)	6.9 (6.2, 7.5)	6.0 (5.7, 6.2)	9.1 (1.8, 16.5)
75th percentile (White)	21.2 (19.8, 22.6)	18.9 (18.4, 19.5)	22.8 (7.8, 37.8)	6.0 (5.1, 6.9)	4.6 (4.3, 5.0)	13.6 (3.7, 23.5)	6.7 (5.7, 7.7)	5.6 (5.2, 5.9)	11.1 (0.7, 21.6)
Proportion eliminated ^d									
75th percentile (White)			-0.092			-0.098			+0.314
Rural (<1,000 people/mi ²)									
No NDVI intervention	21.0 (19.7, 22.3)	17.6 (17.4, 17.8)	33.9 (20.9, 47.8)	6.7 (5.8, 7.6)	4.5 (4.4, 4.6)	22.1 (13.0, 31.2)	6.8 (5.9, 7.7)	5.1 (5.0, 5.3)	16.8 (7.3, 26.3)
NDVI threshold ^c									
25th percentile (Black)	21.9 (20.4, 23.4)	18.3 (17.9, 18.7)	36.0 (20.6, 51.4)	7.4 (6.3, 8.5)	4.6 (4.4, 4.9)	27.7 (16.8, 38.6)	6.9 (5.8, 8.0)	5.5 (5.3, 5.8)	14.3 (3.0, 25.5)
Observed	20.7 (19.4, 22.0)	17.6 (17.4, 17.8)	31.2 (17.8, 44.5)	6.3 (5.4, 7.2)	4.5 (4.4, 4.6)	17.9 (8.4, 27.4)	7.0 (6.0, 7.9)	5.1 (5.0, 5.3)	18.6 (9.0, 28.2)
75th percentile (White)	20.0 (18.3, 21.8)	17.2 (16.9, 17.4)	28.5 (11.1, 46.0)	5.7 (4.5, 6.8)	4.4 (4.2, 4.6)	12.4 (0.7, 24.1)	7.0 (5.8, 8.2)	4.9 (4.7, 5.1)	21.2 (8.7, 33.7)
Proportion eliminated ^d									
75th percentile (White)			-0.163			-0.439			+0.265

^aTen-year risks among Black and White men estimated using Cox models with continuous NDVI (interquartile range increase), adjusted for age (deciles), diagnosis year, race, census block group socioeconomic status (% poverty, median income, median home value, %25 and older with less than high school education, joint race-income index concentration at extremes (quintiles), site at diagnosis (University of Pennsylvania Medical Center, University of Pittsburgh Medical Center, Fox Chase, Jefferson Health), network distance to closest cancer facility (minutes), population density (total only), and interaction term between NDVI and race, standardized to distribution of confounders in total, urban, and rural populations.

^bDisparity denotes racial disparity, estimated by taking difference in 10-year mortality risks in Black and White men standardized to covariates. Confidence intervals were estimated using bootstrapping with 500 repetitions.

^cNDVI thresholds were set to (Black, 25th percentile) 0.31, 0.29, and 0.46 and (White, 75th percentile) 0.63, 0.56, and 0.64 corresponding to levels experienced by each group in the total, urban, and rural populations.

^dProportion disparity eliminated by increasing NDVI estimated by taking the difference of racial disparity under No NDVI intervention and 75th percentile of NDVI (White), divided by disparity under No NDVI intervention. (-) indicates disparity would be reduced and (+) indicates disparity would be increased.

benefits of exposure to high levels of greenness. Parks in predominantly Black neighborhoods may be used less frequently due to fewer resources for security and maintenance.⁵¹ Black men in the US may use parks differently because they were historically excluded from public parks through segregation.^{52,53} Residents' perception of higher crime rates, lower levels of walkability, and lack of upkeep could make parks less welcoming for physical activity and socializing, particularly for older community members.⁵⁴⁻⁵⁶ Surveys of park users in the US have found that Black community members often cite greater obstacles to using parks compared with White users, including feeling unwelcome, inconvenient schedules, and financial barriers.^{53,57} These findings suggest that merely introducing neighborhood green spaces in these communities, without any attempts to ensure that the space fits the needs and social mores of that community, could fail to produce any health benefits.

Results from our study should be interpreted in light of its limitations. This study was conducted in the state of Pennsylvania, which has a unique history, geography, and racial composition. Thus, our results may not be generalizable to dissimilar populations. NDVI is a popular objective measure of neighborhood and residential greenness, but does not capture quality or accessibility of green spaces, which may be necessary to inform appropriate interventions. It is possible that by focusing on residential greenness at diagnosis, we fail to capture other possibly meaningful sources of greenness exposure that occur in work or recreational settings, or at different residences during follow-up, leading to measurement error with respect to total greenness exposure. We did not have measurements of screening, health insurance, diet, and lifestyle factors (including physical activity) at diagnosis, which could influence neighborhood selection and mortality, leading to unmeasured confounding

(a threat to all observational studies). However, assuming that socioeconomic status lies upstream of these proximal confounding variables, controlling for socioeconomic status should mitigate this bias. Our sensitivity analysis using E-values suggests that unmeasured variables would be unlikely to completely explain our effect estimates, particularly for CVD mortality. If lifestyle factors lie on the causal path between neighborhood greenness and mortality, we would not adjust even if those data were available. The same argument applies to treatment and quality of life measures post-diagnosis-although they may be important mediators, we lacked data to evaluate these pathways. Measurements of residential air pollution, which has been proposed as a confounder of the greenness-mortality association, were unavailable. However, other large cohort studies reveal that further adjusting for air pollution after adjustment for demographic and socioeconomic variables does not lead to major changes in the estimate of inverse association between NDVI and mortality.^{22,26} Strengths of our study include a cohort design with long follow-up, a large, racially diverse population, adjustment for major sociodemographic, clinical, and contextual environmental confounders, and analysis of the contribution of environment to racial disparities.

In conclusion, we report an inverse association between residential greenness and rate of all-cause, prostate- and CVD-specific mortality among men with CaP in Pennsylvania. Although interventions to set thresholds of residential greenness could have limited impact on reducing racial disparities, increases in greenness were associated with reduced all-cause and prostate-specific mortality rates among both Black and White men with CaP. Enhanced understanding of differences in how Black and White men interact with green spaces could inform targeted nature-based interventions to allow all men with CaP to experience those benefits.

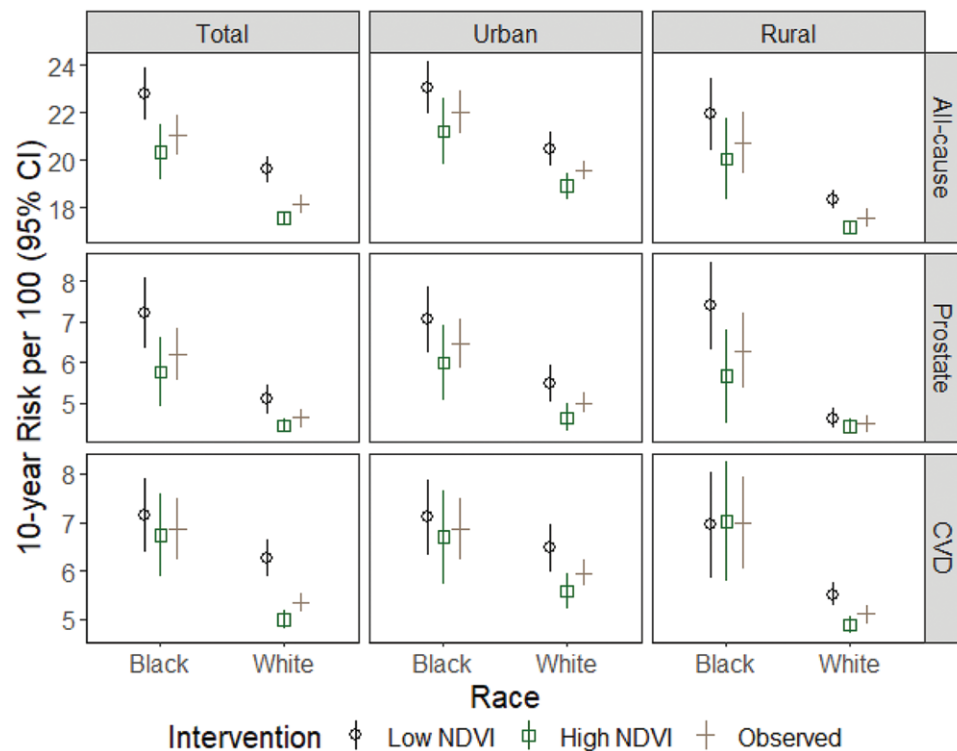


Figure 3. Standardized 10-year cause-specific (all-cause, prostate-, and cardiovascular-specific) mortality risk and 95% CIs under three scenarios^a of NDVI at diagnosis among Black and White men with prostate cancer in Pennsylvania. ^aScenario 1: observed NDVI, Scenario 2: low NDVI (25th percentile of Black men with CaP), Scenario 3: high NDVI (75th percentile of White men with CaP), CVD, cardiovascular-specific mortality, (Urban: $\geq 1,000$ people/mi², Rural: < 1000 people/mi²).

Conflict of interest statement

M.D.H. declares receipt of aspirin for trial NCT 02927249 from Bayer AG, consulting for Arla Foods, United States Social Security Administration, VISIONS Inc., and service on Cambridge Savings Bank's advisory board. The other authors have no conflicts to report.

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References

1. Siegel RL, Miller KD, Jemal A. Cancer statistics, 2019. *CA Cancer J Clin*. 2019;69:7–34.
2. Hsing AW, Tsao L, Devesa SS. International trends and patterns of prostate cancer incidence and mortality. *Int J Cancer*. 2000;85:60–67.
3. Kelly SP, Rosenberg PS, Anderson WF, et al. Trends in the incidence of fatal prostate cancer in the United States by race. *Eur Urol*. 2017;71:195–201.
4. Chornokur G, Dalton K, Borysova ME, Kumar NB. Disparities at presentation, diagnosis, treatment, and survival in African American men, affected by prostate cancer. *Prostate*. 2011;71:985–997.
5. Rebbeck TR. Prostate cancer disparities by race and ethnicity: from nucleotide to neighborhood. *Cold Spring Harbor Perspect Med*. 2018;8:a030387.
6. Krieger N. Epidemiology and the web of causation: has anyone seen the spider? *Soc Sci Med*. 1994;39:887–903.
7. Lynch SM, Rebbeck TR. Bridging the gap between biologic, individual, and macroenvironmental factors in cancer: a multilevel approach. *Cancer Epidemiol Biomarkers Prev*. 2013;22:485–495.
8. Bailey ZD, Krieger N, Agénor M, Graves J, Linos N, Bassett MT. Structural racism and health inequities in the USA: evidence and interventions. *Lancet*. 2017;389:1453–1463.
9. Pager D, Shepherd H. The sociology of discrimination: racial discrimination in employment, housing, credit, and consumer markets. *Annu Rev Sociol*. 2008;34:181–209.
10. Morello-Frosch R, Lopez R. The riskscape and the color line: examining the role of segregation in environmental health disparities. *Environ Res*. 2006;102:181–196.
11. Massey DS, Denton NA. The dimensions of residential segregation. *Soc Forces* 1988;67:281–315.
12. VanderWeele T, ed. *Explanation in Causal Inference: Methods for Mediation and Interaction*. New York: Oxford University Press; 2015.
13. Valeri L, Chen JT, Garcia-Albeniz X, Krieger N, VanderWeele TJ, Coull BA. The role of stage at diagnosis in colorectal cancer black-white survival disparities: a counterfactual causal inference approach. *Cancer Epidemiol Biomarkers Prev*. 2016;25:83–89.
14. VanderWeele TJ, Robinson WR. On the causal interpretation of race in regressions adjusting for confounding and mediating variables. *Epidemiology*. 2014;25:473–484.
15. Jackson JW, VanderWeele TJ. Decomposition analysis to identify intervention targets for reducing disparities. *Epidemiology*. 2018;29:825–835.
16. Gomez SL, Shariff-Marco S, DeRouen M, et al. The impact of neighborhood social and built environmental factors across the cancer continuum: current research, methodological considerations, and future directions. *Cancer*. 2015;121:2314–2330.
17. James P, Banay RF, Hart JE, Laden F. A review of the health benefits of greenness. *Curr Epidemiol Rep*. 2015;2:131–142.
18. Markevych I, Schoierer J, Hartig T, et al. Exploring pathways linking greenspace to health: theoretical and methodological guidance. *Environ Res*. 2017;158:301–317.
19. Frumkin H, Bratman GN, Breslow SJ, et al. Nature contact and human health: a research agenda. *Environ Health Perspect*. 2017;125:075001.
20. Fong KC, Hart JE, James P. A review of epidemiologic studies on greenness and health: updated literature through 2017. *Curr Environ Health Rep*. 2018;5:77–87.

21. James P, Hart JE, Banay RF, Laden F. Exposure to greenness and mortality in a nationwide prospective cohort study of women. *Environ Health Perspect*. 2016;124:1344–1352.
22. Wilker EH, Wu CD, McNeely E, et al. Green space and mortality following ischemic stroke. *Environ Res*. 2014;133:42–48.
23. Twohig-Bennett C, Jones A. The health benefits of the great outdoors: a systematic review and meta-analysis of greenspace exposure and health outcomes. *Environ Res*. 2018;166:628–637.
24. South EC, Hohl BC, Kondo MC, MacDonald JM, Branas CC. Effect of greening vacant land on mental health of community-dwelling adults: a cluster randomized trial. *JAMA Netw Open*. 2018;1:e180298.
25. Casey JA, James P, Cushing L, Jesdale BM, Morello-Frosch R. Race, ethnicity, income concentration and 10-year change in urban greenness in the United States. *Int J Environ Res Public Health*. 2017;14:E1546.
26. Vienneau D, de Hoogh K, Faeh D, Kaufmann M, Wunderli JM, Rössli M; SNC Study Group. More than clean air and tranquillity: residential green is independently associated with decreasing mortality. *Environ Int*. 2017;108:176–184.
27. Crouse DL, Pinault L, Balram A, et al. Urban greenness and mortality in Canada's largest cities: a national cohort study. *Lancet Planet Health*. 2017;1:e289–e297.
28. Young JL Jr, Roffers SD, Ries LAG, Fritz AG, Hurlbut AA, eds. SEER Summary Staging Manual – 2000: Codes and Coding Instructions. Bethesda: National Cancer Institute; 2001.
29. Kriegler FJ, Malila WA, Nalepka, RF, Richardson W. Preprocessing transformations and their effects on multispectral recognition. In: *Proceedings of the Sixth International Symposium on Remote Sensing of Environment*. Ann Arbor, MI: Institute of Science and Technology of the University of Michigan; 1969:97–131.
30. Greenland S, Pearl J, Robins JM. Causal diagrams for epidemiologic research. *Epidemiology*. 1999;10:37–48.
31. Hernán MA, Hernández-Díaz S, Robins JM. A structural approach to selection bias. *Epidemiology*. 2004;15:615–625.
32. Manson S, Schroeder J, Van Riper D, Ruggles S. IPUMS National Historical Geographic Information System: Version 13.0 [Database]. Minneapolis: University of Minnesota; 2018. Available at: [https://data2.nhgis.org/main]. Accessed 18 July 2019.
33. Krieger N, Williams DR, Moss NE. Measuring social class in US public health research: concepts, methodologies, and guidelines. *Annu Rev Public Health*. 1997;18:341–378.
34. Rojas-Rueda D, Nieuwenhuijsen MJ, Gascon M, Perez-Leon D, Mude P. Green spaces and mortality: a systematic review and meta-analysis of cohort studies. *Lancet Planet Health*. 2019;3:e469–e477.
35. Krieger N, Waterman PD, Spasojevic J, Li W, Maduro G, Van Wye G. Public health monitoring of privilege and deprivation with the index of concentration at the extremes. *Am J Public Health*. 2016;106:256–263.
36. Naimi AI, Schnitzer ME, Moodie EE, Bodnar LM. Mediation analysis for health disparities research. *Am J Epidemiol*. 2016;184:315–324.
37. Hernán MA, Robins JM. *Causal Inference*. Boca Raton: Chapman & Hall/CRC, forthcoming. Available at: <https://www.hsph.harvard.edu/miguel-hernan/causal-inference-book/>. Accessed 5 November 2019.
38. Valeri L, Vanderweele TJ. Mediation analysis allowing for exposure-mediator interactions and causal interpretation: theoretical assumptions and implementation with SAS and SPSS macros. *Psychol Methods*. 2013;18:137–150.
39. VanderWeele TJ, Ding P. Sensitivity analysis in observational research: introducing the E-Value. *Ann Intern Med*. 2017;167:268–274.
40. Klein J, von dem Knesebeck O. Socioeconomic inequalities in prostate cancer survival: a review of the evidence and explanatory factors. *Soc Sci Med*. 2015;142:9–18.
41. Greenworks Philadelphia. City of Philadelphia, Mayor's Office of Sustainability. Available at: <https://www.phila.gov/media/20160419140515/2009-greenworks-vision.pdf>. Accessed 18 July 2019.
42. Albertsen PC, Hanley JA, Gleason DF, Barry MJ. Competing risk analysis of men aged 55 to 74 years at diagnosis managed conservatively for clinically localized prostate cancer. *JAMA*. 1998;280:975–980.
43. Lu-Yao GL, Albertsen PC, Moore DF, Lin Y, DiPaola RS, Yao SL. Fifteen-year outcomes following conservative management among men aged 65 years or older with localized prostate cancer. *Eur Urol*. 2015;68:805–811.
44. Yitshak-Sade M, James P, Kloog I, et al. Neighborhood greenness attenuates the adverse effect of PM_{2.5} on cardiovascular mortality in neighborhoods of lower socioeconomic status. *Int J Environ Res Public Health*. 2019;16:e814.
45. Wekesa A, Harrison M, Watson RW. Physical activity and its mechanistic effects on prostate cancer. *Prostate Cancer Prostatic Dis*. 2015;18:197–207.
46. Kenfield SA, Stampfer MJ, Giovannucci E, Chan JM. Physical activity and survival after prostate cancer diagnosis in the health professionals follow-up study. *J Clin Oncol*. 2011;29:726–732.
47. Wang Y, Jacobs EJ, Gapstur SM, et al. Recreational physical activity in relation to prostate cancer-specific mortality among men with nonmetastatic prostate cancer. *Eur Urol*. 2017;72:931–939.
48. Allott EH, Masko EM, Freedland SJ. Obesity and prostate cancer: weighing the evidence. *Eur Urol*. 2013;63:800–809.
49. Zhang X, Zhou G, Sun B, et al. Impact of obesity upon prostate cancer-associated mortality: a meta-analysis of 17 cohort studies. *Oncol Lett*. 2015;9:1307–1312.
50. James P, Jankowska M, Marx C, et al. “Spatial Energetics”: integrating data from GPS, accelerometry, and GIS to address obesity and inactivity. *Am J Prev Med*. 2016;51:792–800.
51. Bruton CM, Floyd MF. Disparities in built and natural features of urban parks: comparisons by neighborhood level race/ethnicity and income. *J Urban Health*. 2014;91:894–907.
52. Byrne J, Wolch J. Nature, race, and parks: past research and future directions for geographic research. *Prog Hum Geogr*. 2009;33:743–765.
53. Wolch JR, Byrne J, Newell JP. Urban green space, public health, and environmental justice: the challenge of making cities “just green enough”. *Landsc Urban Plan*. 2014;125:234–244.
54. Cohen DA, McKenzie TL, Sehgal A, Williamson S, Golinelli D, Lurie N. Contribution of public parks to physical activity. *Am J Public Health*. 2007;97:509–514.
55. Kaczynski AT, Wilhelm Stanis SA, Hastmann TJ, Besenyi GM. Variations in observed park physical activity intensity level by gender, race, and age: individual and joint effects. *J Phy Act Health*. 2011;8(suppl 2):S151–S160.
56. Gobster PH. Managing Urban parks for a racially and ethnically diverse clientele. *Leis Sci*. 2002;24: 143–159.
57. Das KV, Fan Y, French SA. Park-use behavior and perceptions by race, hispanic origin, and immigrant status in minneapolis, MN: implications on park strategies for addressing health disparities. *J Immigr Minor Health*. 2017;19:318–327.
58. VanderWeele TJ. Policy-relevant proportions for direct effects. *Epidemiology*. 2013;24:175–176.
59. Breslow NE. Discussion of the paper by D. R. Cox. *J R Statist Soc B*. 1972;34: 216–217.