

British Journal of Cancer (2016) 115, 592–598 | doi: 10.1038/bjc.2016.232

Keywords: breast cancer; statins; adjuvant therapy; survival; epidemiology

De novo post-diagnosis statin use, breast cancer-specific and overall mortality in women with stage I-III breast cancer

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Background: Prior evidence suggests a role for statins in the management of cancer. However, the benefit of statin use in the adjuvant setting remains uncertain. This study investigates associations between statin use initiated after a breast cancer diagnosis and mortality.

Methods: Women with stage I–III breast cancer were identified from the National Cancer Registry of Ireland (N = 4243). Post-diagnostic statin initiators were identified from pharmacy claims data (N = 837). Multivariate models were used to estimate hazard ratios (HRs) and 95% confidence intervals (Cls) for associations between *de novo* statin use and mortality.

Results: The median duration of statin use was 6.7 years. No association was found between post-diagnostic statin use and breast cancer-specific (HR 0.88, 95% CI 0.66, 1.17) or all-cause mortality (HR 1.00, 95% CI 0.82, 1.21).

Conclusions: The results from our study suggest that initiating statin use after a diagnosis of stage I–III breast cancer is not associated with a reduction in breast cancer-specific mortality.

Statins, or 3-hydroxy-3-methylglutaryl coenzyme-A reductase (HMGCR) inhibitors, are prescribed for cholesterol reduction and cardiovascular disease prevention (Holmes and Chen, 2012); however, some epidemiological evidence suggests a role in breast cancer management (Kwan et al, 2008; Ahern et al, 2011; Chae et al, 2011; Nielsen et al, 2012; Nickels et al, 2013; Boudreau et al, 2014; Murtola et al, 2014; Cardwell et al, 2015; Desai et al, 2015). Uncertainty over the benefits of statins in the adjuvant breast cancer setting remain, as significant effects may be limited to reductions in locoregional recurrence, rather than distant recurrence (Ahern et al, 2011), and to date, no studies of statin use have reported reductions in breast cancer-specific mortality (Nickels et al, 2013; Cardwell et al, 2015; Desai et al, 2015). Previous studies have included women who initiated statin use prior to their breast cancer diagnosis, limiting their

utility in clinical decision making in the adjuvant setting (Ahern *et al*, 2011; Chae *et al*, 2011; Nickels *et al*, 2013; Boudreau *et al*, 2014; Cardwell *et al*, 2015; Desai *et al*, 2015). This study aimed to: (a) measure associations between statin use initiated after a breast cancer diagnosis (*de novo*), and breast cancer-specific and all-cause mortality, and (b) investigate whether these associations are modified by statin solubility or tumour characteristics.

MATERIALS AND METHODS

This study used patient records from the National Cancer Registry Ireland (NCRI), linked to individual-level prescription dispensing data from Ireland's Primary Care Reimbursement Services (PCRS),

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Received 12 April 2016; revised 1 July 2016; accepted 6 July 2016; published online 2 August 2016

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as described previously (Barron *et al*, 2014). The study included women diagnosed with stage I–III invasive breast cancer (ICD-10 C50) between 1 January 2001 and 31 December 2011, aged between 50–80 years at diagnosis, with GMS eligibility from at least 1 year prior to diagnosis and no history of invasive cancer, other than non-melanoma skin cancer. Women receiving statin therapy in the year prior to breast cancer diagnosis were excluded.

De novo post-diagnostic statin exposure was identified from prescriptions dispensed between breast cancer diagnosis and end of follow-up (death or 31 December 2012, whichever occurred first). The number of days' supply on each prescription was extracted and the statin dosing intensity was calculated on the basis of the number of days' statin supply in the prior year (Peterson et al, 2007). These exposure histories were used to define the following time varying exposure categories: (i) exposed (yes/no) from the date of their first statin prescription following diagnosis; (ii) within statin users, women were identified as having high-intensity exposure from the date they had received a statin at an intensity of ≥80%, for at least 1 year (e.g., at least 292 out of 365 days is considered high intensity). Once allocated to an exposure category, women remained in this category to the end of follow-up.

The following data were obtained from the NCRI database: age (years) at diagnosis, smoking status at diagnosis (never, past, current and unspecified), tumour stage (I, IIa, IIb, IIIa and IIIb-c), histologic tumour grade (low, intermediate, high and unspecified), oestrogen (ER), progesterone and human epidermal growth factor-2 (HER2) receptor status (positive, negative and unspecified), and chemotherapy (yes, no) in the year after diagnosis. The PCRS database was used to identify anti-oestrogen therapy in the year after breast cancer diagnosis (yes, no) and potentially confounding medication use in the year prior to diagnosis (exposed, unexposed); aspirin (Holmes et al, 2010), anti-diabetics (Holmes et al, 2010), non-steroidal anti-inflammatory drugs (Marshall et al, 2005) and bisphosphonates (Coleman et al, 2013). The number of drug classes (fourth level WHO-ATC classification) dispensed in the year before diagnosis was used as a proxy measure of comorbidity (Schneeweiss et al, 2001). Death certificates provided the date and cause of death (all-cause or breast cancer-specific). Breast cancerspecific deaths were identified using SEER definitions (Supplementary Table S1; Howlader et al, 2010).

Analyses were performed using SAS v9.3 (SAS Institute Inc, Cary, NC, USA). The proportion of post-diagnostic statin users was tabulated and differences in the rates of statin initiation across covariates were compared using Poisson regression (significance at a two-sided α -level of 0.05). Kaplan–Meier analysis was used to estimate the median duration of statin use from initiation to the last exposure (censored at the date of death or end of follow-up). The overall statin exposure intensity was calculated as the number of days' supply as a proportion of the number of days from initiation to last exposure.

For survival analyses, person time was calculated from the date of breast cancer diagnosis to the end of follow-up. Multivariate Cox proportional hazards models were used to estimate adjusted hazard ratios (HRs) and 95% confidence intervals (CIs) for associations between post-diagnosis statin use, and breast cancer-specific and all-cause mortality. Statin exposure was lagged by 2 years to reduce the possibility that changes in breast cancer prognosis or treatment (i.e., cancer recurrence or approaching death) influenced statin initiation or continuation (Tevaarwerk *et al*, 2013; Chubak *et al*, 2013).

Subgroup analyses included stratification by: (a) high-/low-exposure intensity as a measure of drug adherence, (b) statin solubility: lipophilic (atorvastatin, fluvastatin and simvastatin), hydrophilic (pravastatin and rosuvastatin), or both, and (c) ER status (positive, negative and unspecified). An interaction term was included in the multivariable model to assess effect modification. In sensitivity analyses, we defined high-intensity statin exposure as

≥80% intensity for longer than two consecutive years, extended the time without pre-diagnostic statin exposure from 1 to 3 years, varied the lag time from 0 to 4 years and stratified lipophilic/hydrophilic statin use by high-/low-exposure intensity.

RESULTS

Cohort and exposure characteristics. For the 4243 eligible women, the median post-diagnostic follow-up was 4.9 years and their characteristics are described in Table 1. A study flow diagram is shown in Supplementary Figure S1. Within this cohort, 837 (19.7%) women initiated statin use after their breast cancer diagnosis. Rates of initiation were significantly higher in women with a history of diabetes, lower tumour stage at diagnosis and positive ER status. The median time from diagnosis to statin initiation was 2.1 years, the median duration of statin use was 6.7 years and the mean on-treatment exposure intensity was 86.3% (Table 2). Person time attributed to *de novo* statin users and non-users was 2426 and 12 369 years, respectively.

De novo statin use and mortality. No significant association was found between *de novo* statin initiation, and breast cancer-specific (HR 0.88, 95% CI 0.66, 1.17) or all-cause mortality (HR 1.00, 95% CI 0.87, 1.18) (Table 2). Subgroup analyses in women taking statins at an intensity of \geq 80% for longer than 12 consecutive months also yielded null associations with breast cancer-specific mortality (HR 1.04, 95% CI 0.71, 1.51). The median length of time to statin initiation in this high-intensity exposure group was 2.0 years, the median duration of statin use was 8.5 years and the mean ontreatment exposure intensity was 89.2%. Our results were unchanged in sensitivity analyses (Table 3).

We found no statistically significant associations between hydrophilic or lipophilic statin use and breast cancer-specific mortality in subgroup analyses (Table 2). There was no evidence of effect modification by ER status ($P_{\rm interaction} = 0.69$).

DISCUSSION

This study sought to address the clinically relevant question of whether there is a benefit associated with statin initiation for women following a breast cancer diagnosis. We observed no significant association between *de novo* post-diagnostic statin exposure and breast cancer-specific mortality in a cancer registry-based cohort of 4243 women newly diagnosed with stage I–III breast cancer. Within statin initiators, we observed long treatment durations and high treatment intensity, suggesting that our results are unlikely to be due to inadequate statin exposure. A statistically significant association with reduced all-cause and breast cancer-specific mortality was observed in the low-intensity lipophilic statin subgroup. However, this finding is very unlikely to be causal, as the median duration of exposure in this subgroup was only 6 months and high-intensity lipophilic statin use was not associated with a reduction in breast cancer-specific mortality.

Several studies have examined post-diagnostic statin use in women who initiated statin treatment prior to their breast cancer diagnosis (Ahern *et al*, 2011; Chae *et al*, 2011; Nickels *et al*, 2013; Boudreau *et al*, 2014; Murtola *et al*, 2014; Cardwell *et al*, 2015; Desai *et al*, 2015), with some reporting large reductions in breast cancer recurrence, in particular for lipophilic statin users (Ahern *et al*, 2011; Murtola *et al*, 2014). However, these findings may be at least partly attributable to residual confounding due to statin-prescribing patterns and healthy user effects. There is evidence that statins are preferentially prescribed for, and taken by, patients who make better healthcare choices, engage in healthier behaviours and have superior health outcomes (Evans *et al*, 1995; Haley and Dietschy, 2000;

Table 1. Characteristics of	f women included in the stud	dy cohort, by post-dia	ignosis statin exposure, wi	th statin initiation rate
		De novo statin use pos	st breast cancer diagnosis ^{a,b}	
Characteristic	Non-user (<i>N</i> = 2759)	User (N = 837)	Initiation rate (per 1 associate	1000 person years), d <i>P</i> -value
Age in years				
Median (IQR)	66 (58, 73)	65 (58, 72)		_
Comorbidity score ^c				
Median (IQR)	6 (3, 11)	7 (3, 11)		_
Smoking (%)				
Current	583 (21.1)	171 (20.4)	41.3	0.53
Past	306 (11.1)	106 (12.7)	47.5	
Never	1324 (48.0)	422 (50.4)	43.8	
Unspecified	546 (19.8)	138 (16.5)	38.8	
Aspirin (%) ^c				
Yes	432 (15.7)	153 (18.3)	49.2	0.06
No	2327 (84.3)	684 (81.7)	41.6	
NSAID (%) ^c				
Yes	1178 (42.7)	384 (45.9)	44.8	0.22
No	1581 (57.3)	453 (54.1)	41.2	
Anti-diabetic (%) ^{c,d}				
Yes	60 (2.2)	38 (4.5)	74.7	0.001
No	2699 (97.8)	799 (95.5)	41.9	
Bisphosphonate (%) ^c				
Yes	198 (7.2)	46 (5.5)	39.4	0.40
No	2561 (92.8)	791 (94.5)	43.0	
Tumour stage (%) ^{d,e}				
l lla	917 (33.2)	297 (35.5)	44.1 47.5	0.02
IIb	843 (30.6) 610 (22.1)	297 (35.5) 162 (19.4)	38.0	
Illa	166 (6.0)	40 (4.8)	39.6	
IIIb-c	223 (8.1)	41 (4.9)	31.7	
Tumour grade (%)				
Low	301 (10.9)	101 (12.1)	44.8	0.18
Intermediate	1357 (49.2)	416 (49.7)	43.9	
High	866 (31.4)	254 (30.4)	42.4	
Unspecified	235 (8.5)	66 (7.9)	35.8	
ER (%) ^d				
Negative	471 (17.1)	110 (13.1)	35.3	0.01
Positive Unspecified	2028 (73.5) 260 (9.4)	610 (72.9) 117 (14.0)	43.7 47.5	
<u>'</u>	200 (7.4)	117 (14.0)	47.5	
PR (%)	717 (27.0)	170 (21 4)	20.2	0.22
Negative Positive	717 (26.0) 1393 (50.5)	179 (21.4) 415 (49.6)	39.2 44.7	0.22
Unspecified	649 (23.5)	243 (29.0)	42.7	
HER2 (%)		·		
Negative	1679 (60.9)	419 (50.1)	40.8	0.06
Positive	339 (12.3)	99 (11.8)	44.7	5.50
Unspecified	741 (26.9)	319 (38.1)	45.1	
Chemotherapy (%) ^f				
Yes	1123 (40.7)	344 (41.1)	43.2	0.78
No	1636 (59.3)	493 (58.9)	42.5	
Anti-oestrogen (%) ^f				
Yes	2065 (74.9)	642 (76.7)	43.8	0.25
No	694 (25.1)	195 (23.3)	39.9	

Abbreviations: ER = oestrogen receptor; HER2 = human epidermal growth factor receptor 2; IQR = interquartile range; NSAID = non-steroidal anti-inflammatory drug; PR = progesterone receptor.

^aNo statin use in the year prior to diagnosis and at least one statin prescription received between diagnosis and the end of follow-up, 31 December 2011.

^bPatients identified as statin users/non-users after lagging exposure by 2 years.

 $^{^{\}mathbf{c}}$ In the year prior to breast cancer diagnosis.

 $[\]mathbf{d}$ Difference in statin initiation rate P < 0.05 (Poisson regression).

eAJCC Cancer Staging Manual 6th Edition. Springer, 2002.

 $[{]f f}$ In the year post breast cancer diagnosis.

Years to treatment initiation (median)									
Years to treatment ing intensity					All-cause mortality	tality	Breas	Breast cancer-specific mortality	ic mortality
	Years on treatment (median)	On-treatment exposure intensity (mean %)	Follow-up (person years)	Deaths (rate) ^a	Univariate HR (95% CI)	Multivariate HR (95% CI) ^b	Deaths (rate) ^a	Univariate HR (95% CI)	Univariate HR Multivariate HR (95% CI) (95% CI)
	6.7	86.3	12369 2426	692 (55.9) 128 (52.8) C	Ref 1.93 (0.77, 1.14)	Ref Ref 0.93 (0.77, 1.14) 1.00 (0.82, 1.21)	398 (32.2) 56 (23.1)	98 (32.2) Ref Ref 55 (23.1) 0.79 (0.59, 1.06) 0.88 (0.66, 1.17)	Ref 0.88 (0.66, 1.17)
	1 ;	L	12369	692 (55.9)	Ref	Ref	398 (32.2)		Ref
	0.7	82.1 89.2	1165 1261	54 (46.4) C	0.82 (0.62, 1.08) 1.05 (0.82, 1.35)	0.88 (0.67, 1.17) 1.11 (0.86, 1.43)	24 (20.6) 32 (25.4)		0.68 (0.45, 1.02) 0.76 (0.50, 1.15) 0.92 (0.63, 1.34) 1.03 (0.71, 1.50)
	1	1	12369	692 (55.9)	Ref	Ref	398 (32.1)	Ref	Ref
	5.0	88.9	610	41 (67.2) 1	1.18 (0.68, 1.63)	1.43 (1.04, 1.97) [‡]	21 (34.4)	1.16 (0.74, 1.81)	1.35 (0.86, 2.11)
	5.8	88.2 71.6	1579 236	74 (46.9) C	0.83 (0.65, 1.06)	0.83 (0.65, 1.06)	31 (19.6)		0.67 (0.46, 0.97) 0.72 (0.49, 1.04) 0.62 (0.23, 1.66) 0.77 (0.28, 2.08)
103									
103	ı	I	12369	692 (55.9)	Ref	Ref	398 (32.1)	Ref	Ref
212	0.7	85.5 91.9	290	22 (75.9) 1	1.33 (0.87, 2.03) 1.03 (0.65, 1.61)	1.60 (1.05, 2.46) ^f 1.23 (0.78, 1.92)	13 (44.8) 8 (25.0)		1.44 (0.83, 2.51) 1.68 (0.96, 2.94) 0.92 (0.47, 1.80) 1.07 (0.55, 2.10)
y ^d 292	0.5	85.2 90.4	805 774	28 (34.8) C 46 (59.4) 1	0.62 (0.42, 0.90)	0.63 (0.43, 0.92) ^f 1.06 (0.79, 1.44)	9 (11.2) 22 (28.4)		0.37 (0.19, 0.72) 0.39 (0.20, 0.76) [‡] 0.95 (0.61, 1.48) 1.05 (0.67, 1.63)
Both 107 2.3	7.9	71.6	236	13 (55.0)	1.96 (0.48, 1.93)	0.96 (0.48, 1.93) 1.23 (0.61, 2.48)	4 (16.9)		0.72 (0.23, 2.26) 0.91 (0.29, 2.86)

Abbreviations: CI = confidence interval; HR = hazard ratio; Ref = referent group.

*Deaths per 1000 person years.

*Deaths person

Univariate HR Multivariate HR 1.18) 0.86 (0.69, 1.07) 0.94 (0.74, 1.19) 0.96 (0.68, 1.34) 1.46) 1.21) (95% CI)^b Breast cancer-specific mortality 0.84 (0.60, 1.02 (0.63, (0.62, (0.67, Ref Ref Ref Ref Ref Ref 0.95 0.90 Ref 0.76 (0.54, 1.06) 0.88 (0.55, 1.42) 1.35) 1.10) 0.78 (0.63, 0.97) 0.85 (0.67, 1.08) 0.87 (0.62, 1.22) (95% CI) Ref 0.82 (0.61, 0.88 (0.57, Ref Ref Ref Ref 398 (32.2) 37 (22.9) (1 19 (23.4) 562 (30.7) 107 (23.9) 308 (31.5) 40 (23.7) 221 (29.3) 25 (22.4) 392 (32.4) 55 (23.8) 482 (31.5) 85 (25.3) Deaths (rate)^a Table 3. Sensitivity analyses – univariate and multivariate hazard ratios for association between de novo post-diagnostic statin use and mortality 0.96 (0.76, 1.21) Multivariate HR (95% CI)^b 1.31) 1.25) (0.87, 1.18) 1.33) 1.25) Ref (0.84, 0.99 (0.74, 1.06 (0.84, 1.06 (0.89, Ref Ref Ref Ref Ref 1.01 1.03 All-cause mortality Univariate HR 0.99 (0.84, 1.17) 0.91 (0.72, 1.14) 1.17) 0.94 (0.81, 1.09) 0.99 (0.79, 1.25) 0.96 (0.73, 1.27) (95% CI) Ref 0.96 (0.78, ' Ref Ref Ref Ref Ref 692 (55.9) 83 (51.5) (45 (55.3) 909 (49.6) 804 (52.6) 183 (54.6) 564 (57.7) 93 (55.2) 427 (56.6) 59 (52.8) 677 (56.0) 124 (53.8) Deaths (rate)^a Follow-up (person years) 18 339 4496 9776 1686 12369 1613 813 12 096 2307 15 291 3354 7540 exposure intensity On-treatment (mean %) 82.8 91.0 35.6 36.0 35.9 85.7 86.1 >24 consecutive months^c treatment (median) Years on 1.6 6.1 1.9 5.7 6.7 6.7 Sensitivity analysis: no statin exposure in 3 years prior to diagnosis 4 years by 0, 1, 3 and treatment initiation (median) exposure≥80% for 1.6: 1.8 2.3 1 [: Sensitivity analysis: yes/no exposure lagged 2759 480 357 2425 2046 2670 3038 3058 Z Statin exposure – yes/no (lag 0 years) Statin exposure – yes/no (lag 3 years) Statin exposure – yes/no (lag 4 years) Sensitivity analysis: high-intensity Statin exposure – yes/no (lag 1 year) De novo post-diagnostic statin Statin exposure – yes/no Statin user - high intensity Statin user - low intensity exposure definitions Statin user Statin user Statin user Statin user Non-user Non-user Non-user Non-user Non-user

Abbreviations: CI=confidence interval; HR=hazard ratio; Ref=referent group

Deaths per 1000 person years

exposure lagged by 2 years in analysis

and distributed for age at diagnosis (years); smoking status (never, past, current and unspecified); comorbidity score, turnour stage (I, IIa, IIb, IIIa and IIIb—c); turnour grade (low, intermediate, high and unspecified); ER, PR and HER2 receptor status (positive, negative, and no). unspecified); chemotherapy in year post diagnosis (yes, no); anti-oestrogen therapy in year post diagnosis (yes, no); aspirin, bisphosphonate, NSAID and anti-diabetic medication use (yes,

Brookhart *et al*, 2007; Dormuth *et al*, 2009) and have a better breast cancer prognoses (Snyder *et al*, 2009a, b). If unaccounted for in analyses, this residual confounding can lead to an overestimation of any beneficial effect of statins (Glynn *et al*, 2001, 2006). Moreover, these studies included women who initiated statin use prior to their breast cancer diagnosis, limiting the relevance of their findings to clinical decision making in the adjuvant setting.

Although our study is larger and more methodologically robust, our results are consistent with those from the small number of studies that have specifically examined de novo post-diagnostic statin use and breast cancer-specific mortality (Kwan et al, 2008; Cardwell et al, 2015). In these studies, statin use initiated after diagnosis was not associated with an improvement in breast cancer outcomes. In a study by Murtola et al (Murtola et al, 2014) investigating statin use and breast cancer survival, a sensitivity analysis was carried out that limited their analysis to de novo statin users. A large reduction in breast cancer mortality was observed (HR 0.31, 95% CI 0.22, 0.44), however, this association lacked a clear dose response. In addition, this study did not employ a lagged statin exposure, thereby, increasing the risk of reverse causation bias (Chubak et al, 2013). Although we observed no overall association between de novo statin use and breast cancer-specific mortality in an unselected population, experimental studies suggest there may be specific subgroups of patients for whom statin treatment could be beneficial (Garwood et al, 2010; Bjarnadottir et al, 2013, 2015). In a study by Bjarnadottir et al (Bjarnadottir et al, 2013, 2015), in which women received atorvastatin (80 mg per day) for 2 weeks between diagnosis and surgical resection of their breast tumour, statin treatment was associated with a statistically significant reduction in Ki67 proliferation index among women with tumours expressing HMGCR. It would be worthwhile to evaluate tumour expression of HMGCR as a predictor of response to statin treatment in future studies.

Study strengths include the use of prospectively collected outcome and statin exposure data, whereas limitations include the potential for (a) residual confounding owing to a lack of information on lifestyle factors that could influence disease progression (i.e., obesity) and (b) misclassification bias owing to non-adherence (although the risk is small, as women are unlikely to continue filling a prescription they are no longer taking). A limitation of this study is the unavailability of reliable cancer recurrence data. Finally, the generalisability of study findings is limited by the use of the GMS-eligible population, which is constrained by age and socioeconomic status.

In conclusion, the results from our study suggest that initiating statin use after a diagnosis of stage I–III breast cancer is not significantly associated with a reduction in breast cancer-specific mortality. We observed no evidence of effect modification by statin solubility or hormone receptor characteristics.

ACKNOWLEDGEMENTS

We would like to thank the NCRI and the Irish Health Services Executive PCRS for providing access to the data upon which this study was based. In particular, we are grateful to the Data Team at the NCRI for linking the data sets, and Dr Sandra Deady and Mr Christopher Brown for preparing these for analysis. The interpretation and reporting of these data are the responsibility of the authors and should in no way be seen as the official policy or interpretation of the NCRI or the Irish Health Services Executive PCRS. This work was supported by the Irish Cancer Society Collaborative Cancer Research Centre BREAST-PREDICT (CCRC13GAL to WMG, KB, DOC and TIB) and the Health Research Board Ireland (ICE20119 to KB and LS). LM and AS are funded by the Irish Cancer Society Collaborative Cancer Research

Centre BREAST-PREDICT (CCRC13GAL). TIB was funded by the Health Research Board Ireland (ICE20119). The Health Research Board Ireland and the Irish Cancer Society had no role in the study design; collection, analysis, and interpretation of data; writing of the report; or the decision to submit for publication.

CONFLICT OF INTEREST

LS reports receiving commercial research support from Sanofi-Aventis for a project on treatment and outcomes in breast cancer; 2011–2012. WMG holds a part-time role as Chief Scientific Officer in OncoMark Limited, and was a co-founder and current shareholder of the same. The remaining authors declare no conflict of interest.

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Supplementary Information accompanies this paper on British Journal of Cancer website (http://www.nature.com/bjc)