

# BMJ Open Detecting concurrent mood in daily contact networks: an online participatory cohort study with a diary approach

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## ABSTRACT

**Objectives** This paper examines how people express personal mood concurrently with those connected with them by one or two degrees of separation.

**Design** Participatory cohort study.

**Setting** Online contact diary.

**Participants** 133 participants kept online diaries for 7 months in 2014, which included 127 455 contacts with 12 070 persons.

**Main outcome measures** Diary keepers rated a contacted person's mood during each specific contact, as well as the strength of ties between any pairs of such contacted persons. Such rich information about ties and contacts enable us to construct a complete contact network for each diary keeper, along with the network members' mood and tie strength. We calculate one's overall mood by that person's average mood score during the study period and take the shortest path between any given pair of contacted persons as the degree of separation. We further assume that two connecting persons in a contact network have made contact with each other during the study period, which allows us to examine whether and how personal moods occur concurrently within these contact networks.

**Results** Using mixed-effects models while controlling for covariates at individual, tie and contact levels, we show that personal mood score positively and significantly correlates with the average mood among those directly tied to the person. The same effect remains positive and significant for those connected to the person by two degrees, although the effect size is reduced by about one-half. The mood of anyone separated by more than two degrees is statistically irrelevant.

**Conclusions** Applying network perspectives and rich data at both tie and contact levels to inquiries about subjective well-being, the current study sheds new light on how an improved diary approach can help explain the sophisticated ways in which individuals express their personal moods concurrently during social interactions in everyday life, contact by contact.

## INTRODUCTION

Experiments have demonstrated that emotions tend to spread to others during social interactions,<sup>1</sup> as in the case when

## Strengths and limitations of this study

- The observed online contact diaries allow us to construct 133 complete contact networks with >12 000 members, which help uncover how personal moods vary among network members.
- Some intertwining contact diaries further enable us to cross-examine personal mood and tie strength as rated by both parties of social interactions, solidifying our findings about how a bottom-up social network approach helps reveal concurrent mood in everyday life.
- Our unique approach of contact diaries yields complex network data and allows us to identify clear patterns of concurrent moods, which are particularly useful, because observational studies typically cannot provide sufficient empirical evidence as to how personal moods may spread to friends' friends through contacts.
- Without information about the exact timing of each contact among network members, we cannot infer the occurrence of contagion or diffusion of personal mood.
- With limited information about the contacted persons' personal background, our model cannot fully adjust for the effects of homophily.

diseases, behaviours and ideas transmit through social networking.<sup>2 3</sup> Analyses from other large-scale data have further revealed similar spreading patterns of both positive and negative emotions, such as happiness and depression.<sup>4 5</sup> Like the diffusion of behaviours and attitudes amid face-to-face social networking, emotive sharing and contagion may also cover large social circles through online social contacts by text or voice on Facebook, Twitter, Skype, Google Talk and other social media.<sup>6 7</sup>

The literature has identified such a common phenomenon, generally known as 'emotional contagion', through various study designs in different research settings.

One seminal work argued that emotion or mood can easily be transmitted within social networks, because people are inclined to synchronise others' facial expressions, voices, postures, movements and 'emotional behaviours' and get feedback from such 'mimicry'.<sup>8</sup> Diffusion of emotions and moods occurs among those who close to one another, and happens from moment to moment in everyday encounters with less known others. More recent experimental studies have shown similar contagion of happiness and fear through facial mimicry by measuring facial electromyography.<sup>9 10</sup> In addition to face-to-face contacts, some posts in social media also can trigger users' emotional contagion on a massive scale.<sup>11</sup>

Most existing studies have demonstrated that certain emotions tend to spread from direct contacts over a short period of time, but relatively few have examined whether personal mood may also spread from, or at least occur concurrently among indirect contacts within social networks over a longer period of time. Two recent studies, for example, showed that our mood can be influenced by those familiar to us, and by friends' friends whom we do not personally know.<sup>2 3</sup> More specifically, the spread of personal happiness reaches up to three degrees of separation along social networks, according to one such rare empirical study, the Framingham Heart Study, which analysed historical data over 20 years.<sup>4</sup> The longitudinal analyses indicated that an individual was 15.3% more likely to be happy if a directly connected network member (with one degree of separation) was happy; the effect decreased to 9.8% for those separated by two degrees.<sup>4</sup>

As such findings may have inspired studies of mood diffusion, most non-experimental studies are insufficient to examine the actual mechanism of 'diffusion', because the observational data on which they are based offer no such advantages. Thus, it would be particularly difficult to determine and infer from observational data how personal moods spread or diffuse.<sup>12</sup> Even without claiming the causal effects that are essential for explaining spread or diffusion, however, it would be revealing to examine whether and how similar patterns of concurrent mood may exist under other circumstances. While previous longitudinal studies have focused on how individuals' happiness or depression changes across waves of surveys, an alternative 'bottom-up' approach tends to uncover subtle patterns of concurrent mood by examining how one's mood fluctuates at the micro level, contact by contact.<sup>13-15</sup> Due to the methodological limitation, it may be inappropriate to pursue inquiries about 'mood diffusion or contagion' solely by using observational data, even if one could replicate the decades-long study by creating another huge number of personal networks with records of participants' relationships and emotions. As an alternative approach, our study, based on a different format, provides a new perspective that would further enrich studies on how personal moods can be linked to one another in contact networks.

By extending the bottom-up approach to social network studies, we aim to examine whether and how personal moods occur concurrently within a contact network, using data collected via an online platform, ClickDiary, over a 7-month period between 1 May 2014 and 30 November 2014. The ClickDiary program uses a web-based platform to collect data on participants' health behaviours and all one-on-one interpersonal contacts in everyday life.<sup>16</sup> The data retrieved for this study have a nested hierarchical structure, including detailed information about 133 participants (or 'diary keepers', who recorded details about their social interactions in daily life), 12 070 contacted persons (or 'network members', including 74 persons who also participated in the same diary keeping platform), as well as 127 455 contacts.

Two features in the ClickDiary platform are central to our research design. First, diary keepers reported their own mood for each one-on-one contact by selecting one of the following score categories that best matched their estimates: (1) poor, (2) good, (3) very good or (4) excellent; they then evaluated the mood of each network member during the contact. Second, diary keepers, to their best knowledge, rated how well a given pair of network members knew each other. Once a diary keeper confirmed all of their interpersonal ties, we used these two critical features as the backbone to construct a 'complete contact network' surrounding that diary keeper, the focal person. Within such a complete contact network, we linked the nodes that represent the network members to each other by interpersonal ties. The overall mood of each network member can be represented by taking the average of the person's mood scores recorded in the contact diary during the whole study period. Furthermore, we assumed that two connecting persons in a complete contact network have made contact with each other during the study period, which allows us to examine whether and how personal moods occur concurrently within these contact networks. That is, with the assumption, concurrent mood of the linked members could be explained partially as caused by personal contact. We applied a mixed-effects model to analyse overall mood scores of the members in the 133 complete contact networks to examine whether and how an individual's mood is associated with those within one or two degrees of separation, as well as other members in the complete contact networks.

In sum, our unique approach of online contact diaries is expected to facilitate sophisticated analyses of such phenomena as concurrent patterns of personal mood among network members. Unlike earlier studies that relied on surveys separated by years, our diary approach and analyses can better capture dynamic concurrent moods in everyday life, with complicated network data that help reveal how behaviours and emotions vary contact by contact along with personal ties embedded in different network structures.

## MATERIALS AND METHODS

### The ClickDiary program

The ClickDiary program (<http://cdiary.tw>) uses a web-based platform, written in Chinese, to collect data on participants' daily health behaviours and interpersonal one-on-one contacts.<sup>16</sup> One unique feature of ClickDiary is the friendly interface designed for clicking options on structured diary items via a website or mobile app, making it easier to record responses whenever it is convenient for participants (online supplementary figures S1, S2).

### Public involvement

Our research team recruited participants from various channels including university students, school teachers and administrative employees, volunteers at health-promotion centres, hospital patients and community college students, as well as other adults in the general population. When registering in the ClickDiary platform, all diary keepers are required to give online informed consent before starting to keep the diaries. On signing up for the program, moreover, participants provided sociodemographic information, including age, gender, place of residence, marital status and current job. The program also collects participants' Big-5 personality traits (openness, conscientiousness, extraversion, agreeableness, neuroticism),<sup>17</sup> height and weight, perceived health status and happiness, the number (and characteristics) of people contacted during the day, along with a baseline health module that borrowed items from the Taiwan Social Change Survey.<sup>18</sup> We provided an interactive web chart summarising the records in each participant's contact and health diaries. In addition, participants could gain insight from their overall contact patterns by checking their contact network tree we developed (online supplementary figure S3).

### Contact diary

When adding a person to the contact list for the first time, the diary keepers provided the person's background information, such as age and gender, and evaluated several aspects of their relationship with the person, including the duration of acquaintanceship, degree of familiarity, the most frequent mode used for contact (face-to-face, voice only or text only), contact frequency and the likelihood of discussing important matters. In this particular diary program, a 'contact' refers to one-on-one exchange that involves at least three verbal or written sentences, a definition somewhat narrower than most previous studies using the contact diary approach.<sup>19</sup>

The program also asked participants to evaluate the degree to which any two persons on the contact list were familiar with each other. Before starting to enter the contact details with any new person, a diary keeper had the chance to estimate, on a scale from 1 to 3, how well this particular person knew each of randomly selected five other persons already recorded in the diary. After this first step, another random sample of five different persons' names popped up, so that the diary keeper

could continue judging all the ties between any pair of persons within the contact network. The process continued, randomly adding five new names to the list at a time, until the diary keeper finished rating the strength of all alter pairs. The design allowed the diary keeper to evaluate all alter-alter ties when it was convenient to do so, thus achieving maximum flexibility and encouraging a higher completion rate. In addition, the diary entries also focused on 11 contact attributes, including when the contact took place, who initiated the contact, the major mode of the contact, the duration and content of the contact, where the diary keeper and the contacted person were during the contact, the extent to which the contact felt beneficial to the diary keeper, the mood of each party during the contact and whether the contacted person showed any cold symptoms.

### Data retrieved for the study

From 1 May 2014 to 30 November 2014, 133 residents in Taiwan each completed at least 30 days of contact diaries and recorded one-on-one interpersonal contacts with at least 30 persons. We retrieved the 7-month contact diaries from these 133 participants, with all personal identities removed, which consisted of 141 909 contacts with 16 139 contacted persons. The length of the contact lists varied substantially among the 133 participants, ranging from 30 to 1399, with a median of 76. The participants recorded an average of 12 contacts a day, with a minimum of 1 and a maximum of 56. The contacted persons varied greatly in terms of how often they appeared in the diaries, ranging from only once to daily (214 times during the 7-month period), while the average frequency was eight times.

To measure each person's mood, diary keepers selected one from the following score categories that best matched their estimates: (1) poor, (2) good, (3) very good or (4) excellent. As expected, diary keepers sometimes were unable to estimate a contacted person's mood and consequently answered 'do not know' for the item. The records during the study period showed such 'do not know' answers for the mood item appearing in 9042 contacts. We treated these contacts (with 845 contacted persons) as missing, which reduced the valid number of contacted persons to 15 294.

We further excluded the cases where the diary keepers did not know a contacted person (ie, the category that accounts for about 21.1% of the relationships) from the subsequent analysis for three reasons. First, interactions with strangers carry very different implications in studies of network diffusion. Even though they had actual contacts with the diary keepers, keeping total strangers in a personal network would have left too many unknown or uncertain links, because the diary keepers were unlikely to judge how well these strangers knew one another. Second, it would have been more difficult for the diary keepers to evaluate strangers' mood, which also tended to yield missing or less reliable mood rating. Third, probably due to such uncertainties, our different modelling efforts while retaining the ties with these strangers resulted in

unreasonable noise to the analysis. The final data for modelling mood spread included 12 070 contacted persons of 133 diary keepers who had made 127 455 contacts during the study period.

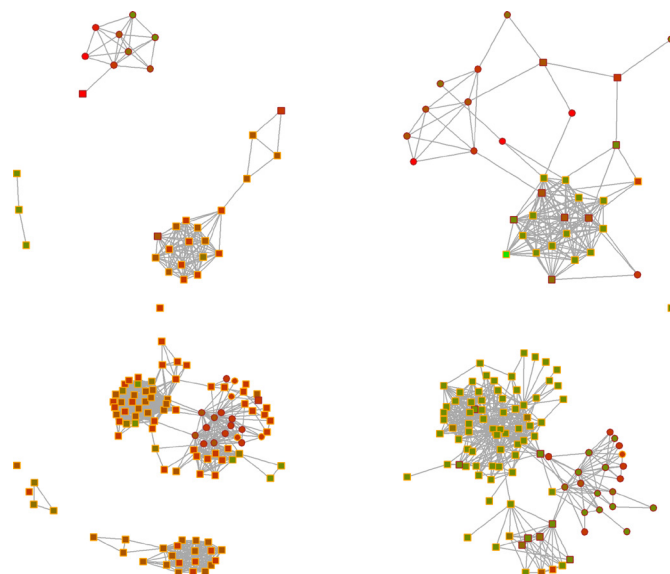
### Network construction

As a whole, diary keepers were able to confirm nearly all (99.97%) interpersonal ties in terms of the familiarity between any pair of contacted persons. Of these ties, averaged across the 133 diary keepers, 78.4% were considered absent (the pairs did not know each other), 10.6% were strong (knew each other well) and 11.1% were weak (knew each other, although not well). For each diary keeper, we constructed a complete contact network, in which each alter, or contacted person, represents a node, and two nodes (other than the focal person) are deemed to be connected when the reported alter-alter tie is either strong or weak, rather than absent. In contrast, nodes are not linked when the pair of network members are strangers to each other. Using the shortest path between any given pair of nodes in the network, we calculate 'distance' (or 'the degree of separation') between the network members pair by pair. We focus on the links among these network members only, while excluding any links leading to and from the unique node that represented the diary keeper (or 'ego-alter ties'). Excluding such ties linked by the diary keepers helps simplify the calculation of 'distance'. Otherwise, any two given nodes in the network would have at most two degrees of separation, because all the network members would be directly linked to the diary keepers. Finally, we defined a network member's overall mood by the average of the mood scores reported by diary keepers during the study period, ranging from 1 to 4.

To illustrate how alter-alter ties cluster into a subset of ego's complete contact network, we drew a figure, using R package igraph, to display the network patterns in each of four diary keepers' complete contact networks during the study period (figure 1). The figure shows that the clustering patterns may differ significantly among diary keepers. Within each of the four complete contact networks, those members who are closer to each other in terms of distance also tend to average similar scores in personal mood during all interpersonal contacts.

### Statistical analysis

To examine whether and how a network member's mood may be associated with those separated by one or two degrees in each of the 133 complete contact networks, we first calculated the average mood of those members who were directly connected to a particular member (ie, those connected by one degree of separation), and then we obtained the average mood of those at two degrees of separation and the average mood of all other members. The average mood score of the  $j^{\text{th}}$  member in the  $i^{\text{th}}$  complete contact network can be simply obtained by calculating  $Y_{ij} = \sum_{k=1}^{K_{ij}} O_{ijk} / K_{ij}$ , where  $O_{ijk}$  is the mood



**Figure 1** The clustering patterns of average personal mood in the complete contact networks of four diary keepers. Each node represents a person, whose relationship with the diary keeper is displayed with a circle for family members, relatives and good friends and a square for the others. The frame colour of the node reveals the strength of tie to the diary keeper (brown for 'know each other very well'; orange for 'know each other, but not well'). Node colour denotes average mood scores of the persons during the study period, with a colour gradient ranging from green, which indicates the worst mood, to red, which indicates the best mood. The figure does not include the diary keeper, who is linked to everyone in respective contact network.

score of their  $k^{\text{th}}$  contact given by the  $i^{\text{th}}$  diary keeper and  $K_{ij}$  is the number of contacts during the study period. Let  $D_{1ij}$  and  $D_{2ij}$  indicate the mean mood scores of those separated by one and two degrees from the  $j^{\text{th}}$  member in the  $i^{\text{th}}$  network, respectively. Finally,  $D_{3ij}$  measures the average mood scores for those beyond two degrees of separation from the  $j^{\text{th}}$  network member.

For this study, we applied a mixed-effects model to analyse the relationship between a person's mood score and the average mood scores of those surrounding the person in the network (or 'network neighbours'), while controlling for the effects of potential covariates on the person's mood. The model is given as:

$$Y_{ij} = (\alpha_0 + a_{0i}) + (\alpha_1 + a_{1i}) D_{1ij} + (\alpha_2 + a_{2i}) D_{2ij} + (\alpha_3 + a_{3i}) D_{3ij} + \sum_{h=1}^p \beta_h X_{hij} + \varepsilon_{ij} \text{ for } i = 1, 2, \dots, n \text{ and } j = 1, 2, \dots, n_i,$$

where  $n$  is the number of networks;  $n_i$  is the size of the  $i^{\text{th}}$  network; random components  $a_{li}$  are assumed to be normally distributed with a mean of 0 and variance  $\sigma_l^2$  for  $l = 0, 1, 2, 3$  and the error term is normally distributed with a mean of 0 and variance  $\tau^2$ . Our main interest is to estimate the fixed effects of  $\alpha_1$ ,  $\alpha_2$  and  $\alpha_3$ , which measure whether and how one's mood correlates with those separated by one degree, two degrees and the others, respectively, within the same complete contact network.

We applied exploratory data analysis tools to identify potential covariates  $X_{hij}$ ,  $h = 1, 2, \dots, p$ , for the  $j^{\text{th}}$  member in the  $i^{\text{th}}$  network. The final  $p=14$  covariates selected in the model are described in the following. In addition to controlling for the possible gender effect of the individual member, the other covariates included in the mixed-effects model measure a variety of features at both tie and contact levels. Two groups of covariates tap the relationships between the diary keepers and their network members. The first is a binary variable indicating that diary keepers knew a network member either well or 'not well' (the latter serves as the base category in the model). The second group of covariates distinguishes four types of the network members' relationships to the diary keepers: family members or relatives (13.2%); good friends (13.5%); coworkers or trade partners (12.7%) and schoolmates, teachers or students (18.4%). The base category for the group is 'others'.

The next group of covariates covers three major contact features: mode, purpose and duration. Because the main goal of the analysis is to examine how personal moods vary among network members, we sum up these contact features for each member. The modes of contact, for example, may play a key role in determining how well the diary keepers judged others' mood. To verify such an effect, we include two contact modes as covariates: face-to-face and 'voice only', leaving 'text only' out of the model. From all of the contacts between the diary keepers and each particular network member during the 7-month study period, we calculate the proportion (percentage) of each of the three modes for the network member. Because the percentages of all three modes for each network member add up to 1, we keep only the first two modes (face-to-face and voice) in the analysis. Suppose that a member-keeper pair had  $K$  times of contacts during the 7-month study period, and among which, the contact modes of 'face-to-face' and 'voice only' were  $k_1$  and  $k_2$  times, respectively. We defined the two covariates face-to-face as the proportion  $k_1/K$  and voice only as  $k_2/K$  for the analysis, and excluded the other mode of 'text only' from the model. The averaged percentages of face-to-face and voice-only contacts among the network members were about 79.8% and 8.2%, respectively.

The model also includes the percentage for each of the two kinds of contact purposes, 'work-related or school-related' and 'daily routine', which were about 30.4% and 12.4% on average, respectively, while excluding 'other purposes'. We further add contact duration into the model, using the percentages for contacts that last 5–59 min, 1–4 hours or >4 hours, excluding those lasting under 5 min. On average, 47% of the contact durations was 5–59 min, 20% was 1–4 hours and 6% was >4 hours.

Because diary keepers reported their own mood as well as the network members' mood in the same diaries, the two scores are expected to be highly correlated. To take such an effect into account, we further controlled for the most influential covariate that measures the average mood score of the  $i^{\text{th}}$  diary keeper when contacted with

the  $j^{\text{th}}$  network member during the study period. In other words, the member's mood score was influenced by the diary keeper's mood score on the contact. Hence, the covariate of the average mood score of the  $i^{\text{th}}$  diary keeper when making contact with the  $j^{\text{th}}$  network member should be influential and must be included in the model for adjustment. We used the lme function from the R package 'nlme' to estimate the model parameters.<sup>20</sup>

### Participant involvement

The participants were involved in neither the development of the research questions nor the design of the study. None of the participants was involved in conducting the study nor was asked to provide input in the writing of this manuscript. There are no plans to disseminate the results of the research to the participants.

### RESULTS

Among the original 133 diary keepers, about 80% were female (106/133). Compared with a representative sample of general population, the group tended to be younger and better educated. Percentages of those under age 23 (college students), 23–39, 40–59 and 60 years or more were 33.8%, 36.1%, 21.8% and 8.3%, respectively. At least 82.7% had ever gone to college. Like previous contact diary studies, the study participants were over-represented by females and better-educated subpopulations.<sup>15</sup>

The high percentage of female diary keepers probably yielded more females on the contact lists (65.5%), although the latter was actually more balanced than the gender distribution among the diary keepers. The age distribution of the 12070 contacted persons was not very skewed, with age groups of 1–19, 20–29, 30–49, 50–59 and 60 years or more accounting for about 8.4%, 21.8%, 31.1%, 18.5% and 16.9%, respectively. About 53.3% of the contacted persons the diary keepers knew were known very well.

In fitting the mixed-effects model, we excluded about 12% of network members with any covariates missing, mainly those who had either no degree 1 or 2 neighbours. As a result, the following results were obtained from the remaining 10 581 network members with complete covariates. Table 1 shows model estimates of the parameters associated with individual mood variation. On average, a person's mood score increased about 0.13 ( $p < 0.0001$ ) for every additional point scored in the average mood of fellow members who were directly connected to the person (with one degree of separation). The average mood score of those members linked by two steps also contributed significantly ( $p = 0.002$ ), with an effect size of about 0.06, to the person. Such contribution diminishes, however, for those members at degree 3 and beyond (linked by at least three steps).

These key findings are noteworthy, because they have been adjusted by several highly relevant covariates in the mixed-effects model. Since the diary keepers rated the

**Table 1** Estimates of the effects associated with a network member's mood score in the mixed-effects models using diary data during May–November 2014

| Variable                           | Value   | SE     | T values | P values |
|------------------------------------|---------|--------|----------|----------|
| Male                               | 0.0032  | 0.0052 | 0.6236   | 0.5329   |
| Tie strength with diary keeper     |         |        |          |          |
| Knew well                          | 0.0163  | 0.0062 | 2.6177   | 0.0089   |
| (knew, not well)                   |         |        |          |          |
| Relationship with diary keeper     |         |        |          |          |
| Family member/relative             | 0.0003  | 0.0097 | 0.0330   | 0.9737   |
| Good friend                        | 0.0057  | 0.0101 | 0.5634   | 0.5731   |
| Coworker/trade partner             | 0.0036  | 0.0094 | 0.3887   | 0.6975   |
| Schoolmate/teacher/student         | 0.0301  | 0.0084 | 3.5603   | 0.0004   |
| (Others)                           |         |        |          |          |
| Contact mode (%)                   |         |        |          |          |
| Face-to-face                       | 0.0277  | 0.0097 | 2.8486   | 0.0044   |
| Voice only                         | 0.0202  | 0.0136 | 1.4877   | 0.1369   |
| (Text only)                        |         |        |          |          |
| Contact purpose (%)                |         |        |          |          |
| Work/school                        | −0.0142 | 0.0073 | −1.9522  | 0.0509   |
| Daily routine                      | −0.0205 | 0.0105 | −1.9516  | 0.0510   |
| (Others)                           |         |        |          |          |
| Contact duration (%)               |         |        |          |          |
| (<5 min)                           |         |        |          |          |
| 5–59 min                           | 0.0208  | 0.0079 | 2.6152   | 0.0089   |
| 1–4 hours                          | 0.0385  | 0.0101 | 3.8141   | 0.0001   |
| 4 hours or more                    | 0.0474  | 0.0149 | 3.1866   | 0.0014   |
| Average mood of diary keeper       | 0.7427  | 0.0074 | 100.9489 | 0.0000   |
| Average mood of network neighbours |         |        |          |          |
| One degree of separation           | 0.1326  | 0.0149 | 8.8671   | 0.0000   |
| Two degrees of separation          | 0.0590  | 0.0191 | 3.0962   | 0.0020   |
| All others                         | −0.0024 | 0.0167 | −0.1460  | 0.8839   |

mood of both parties after a contact, the two scores were often highly correlated. As shown in [table 1](#), a member's mood score was strongly associated, with a large effect size of 0.74, with the average of all the diary keepers' self-rated mood scores when making contact with that specific member. When the diary keeper knew a member well, that member had a better chance of receiving a higher score on personal mood. A network member also tended to receive a higher mood score for a face-to-face contact and a contact that lasted longer. When a contact was about work or school, or was part of the daily routine, the network member's mood was not as good as that of other contacts, at least based on the diary keeper's judgement.

The estimated SD of the three random components for one degree of separation, two degrees of separation and beyond two degrees was 0.0015, 0.0372 and 0.0037, respectively. It is clear that the strong mood associations between directly linked members were consistent across the 133 networks. The mood associations between members and

neighbours separated by two degrees varied a little large, compared with the estimated fixed effect, among the 133 networks. One possible reason for the large variation among the 133 networks may be the network sizes which ranged from 30 to 1399, with a median of 76. Members in a small network tended to have relatively few neighbours separated by two degrees. The average mood of smaller number of neighbours may incur large uncertainty on the estimated association effect.

### Robustness checks

To verify the effects of different degrees of separation, we tried a separate analysis of those members at three degrees of separation, using only 8505 network members with all valid variates. The alternative analytic design showed little change in both the coefficient estimates of the covariates and the coefficient estimates of one, two and three degrees of separation (0.13, 0.08 and −0.03, respectively). Thus, it would be unnecessary and unfruitful

to further determine whether concurrent mood could occur up to three, four or more degrees of separation. Unlike previous studies that have indicated how happiness is linked to more indirectly linked network members in the long run,<sup>4</sup> personal mood in everyday life appears to coexist among those separated by only one or two degrees in one's contact network. To check how robust it is to represent each network member's overall mood with the average of the person's mood scores during the whole study period, we also fitted the mixed-effects model by replacing the average with the median. The fitted models showed similar results regardless of whether we used the average or the median.

### Validation

To further verify the findings from the study, we applied the same mixed-effects model to a similar data set collected in a later study period of 7 months from April to October 2015 (online supplementary file 1). The results from the online supplementary table S1 indicate that the model estimates and significance levels are similar between the two study periods. In particular, while the average mood score of network members linked by only one step significantly contributed for about 0.22 to a given member's mood score, the coefficient estimate dropped to 0.12 for those separated by two degrees. Like the first study period, one's mood in this second period had little to do with those separated beyond two degrees.

We also conducted a subset validation to cross-check the accuracy from a network member's own answer with how well she actually knew each particular member in the diary keeper's network. This critical step of cross-checking alter-to-alter ties was possible because some of our participants joined the ClickDiary study as a group. Among 133 diary keepers, 74 also appeared on the lists of others' 'network members'. In total, 7310 individuals appeared on both contact lists of any two diary keepers who also appeared on each other's contact diary. In 6956 cases, when a first diary keeper believed that the second diary keeper (who happened to be the first diary keeper's network member, thus an alter here, or alter 1) knew one of these overlapped individuals (another alter, or alter 2), the second diary keeper (alter 1) also said she did indeed know this particular person (alter 2). Likewise, in 22 cases when a first diary keeper said that the second diary keeper did not know a network member they shared, the second diary keeper also confirmed that such a tie was absent. As

a result, the diary keepers had judged the alter-to-alter ties among the members in their contact networks with an accuracy rate at about 95.5%, which helps justify our strategy of using diary keeper's contact records to reconstruct part of their contact networks in everyday life.

To check the assumption that two connecting persons in a complete contact network have made contact with each other during the study period, we first identified 5249 individuals who knew those 'network members' (ie, those on the contact lists) who happened to be diary keepers as well. Then we went back to these diary keepers' own contact diaries and counted how many days each diary keeper actually made contact with each of these 5249 individuals during the 7-month period. The days of actual contacts between the pairs were well fitted to a negative binomial distribution, with a mean of 25.3 and a size of 0.63. That is, on average these 5249 pairs of network members contacted with each other on 25.3 days during the study period, although the range of contacts varies widely and is quite skewed (SD=32 days).

The results indicate that any two members in a contact network who knew each other had a 90.5% of chance to have at least one contact with each other during the 7 months. The finding further supports our underlying assumption that the network members who knew each other indeed had contact with each other during the 7-month period. Such interpersonal contacts, in turn, facilitated structural circumstances under which personal moods could disperse or emerge in parallel among network members.

### Sensitivity analysis

Our online diary platform allowed us to obtain contact information reported by diary keepers, not by the network members themselves, which raises a big concern about how accurately the diary keepers judged a network member's mood during a specific contact. To address this potential issue of the diary keeper's error in judging a network member's mood, we checked the extent to which such judged mood scores were accurate and reliable by matching part of them with the mood scores rated by the network members themselves, as with the case with cross-checking alter-to-alter ties. As a result, we were able to compare how other diary keepers estimated the mood of these 74 network members during 2368 contacts with how these 74 network members rated their own mood for each of the identical contacts in their own contact diaries.

**Table 2** Pairs of mood ratings from the diary keepers and 74 of their network members who also rated their own moods during the same contacts

|                                   |     | Member's moods rated by the diary keeper |            |               |               |
|-----------------------------------|-----|--|------------|---------------|---------------|
|                                   |     | (1) Poor                                 | (2) Good   | (3) Very good | (4) Excellent |
| Network member's self-rated moods | (1) | 0 (0.0%)                                 | 3 (0.1%)   | 1 (0.0%)      | 5 (0.2%)      |
|                                   | (2) | 5 (0.2%)                                 | 42 (1.8%)  | 203 (8.6%)    | 76 (3.2%)     |
|                                   | (3) | 5 (0.2%)                                 | 171 (7.2%) | 877 (37.0%)   | 331 (14.0%)   |
|                                   | (4) | 3 (0.1%)                                 | 75 (3.2%)  | 326 (13.8%)   | 245 (10.3%)   |

Counting the original answering categories, concordant pairs accounted for only 49.1% of all mood rating pairs (table 2). Of the 50.9% pairs that were discordant, however, 43.9% showed only a one-category difference (eg, while a diary keeper rated a network member's mood as 'excellent' during a specific contact, that member rated her own mood during that contact as 'very good', which accounted for 14.0% of all 2368 pairs). Therefore, about 93% of these score differences between the moods rated by diary keeper and network member on the identical contact fell between -1 and 1.

We performed a sensitivity analysis to check for the potential effects of the diary keeper's judgement error on the parameter estimates of the mixed-effects model, using a perturbation approach with the initial findings of such cross-checking. Specifically, we simulated 200 datasets of mood scores that deviated from the observed scores by 0, 1 and 2 with probabilities 0.5, 0.43 and 0.07, respectively. Fitting the same mixed-effects model with mood scores

from the  $b^{\text{th}}$  simulated dataset, we obtained estimates and SEs of the  $t^{\text{th}}$  model parameters, denoted by  $\theta_l^{(b)}$  and  $s_l^{(b)}$ , respectively.

Taking the influence of judgement error into account, we then estimated the parameter  $\theta_l$  by the average of these 200  $\theta_l^{(b)}$  with the SE equal to the square root of the sample mean of these 200  $s_l^{(b)}$  squared, plus the sample variance of these 200  $\theta_l^{(b)}$ . It is clear that the augmented judgement errors increased the SEs of the parameter estimates, as shown in table 3. Consequently, most of the covariate effects were no more significant. A person's mood score, however, still significantly increased about 0.088 ( $p=0.018$ ) for every additional point scored in the average mood of fellow network members with one degree of separation. The effect size was about 0.079 for those members separated by two degrees, although the strength reduced to marginal significance ( $p=0.049$ ). Such effects eventually diminished for those network

**Table 3** Combined effect estimates of the same mixed-effects models fitted with 200 different simulated datasets of mood using a perturbation approach

| Variable                               | Value   | SE     | T values | P values |
|--|---------|--------|----------|----------|
| Male                                   | 0.0027  | 0.0159 | 0.1720   | 0.8635   |
| Tie strength with diary keeper         |         |        |          |          |
| Knew well<br>(knew, not well)          | 0.0111  | 0.0199 | 0.5551   | 0.5788   |
| Relationship with diary keeper         |         |        |          |          |
| Family member/relative                 | 0.0008  | 0.0286 | 0.0266   | 0.9788   |
| Good friend                            | 0.0056  | 0.0299 | 0.1882   | 0.8507   |
| Coworker/trade partner                 | 0.0013  | 0.0264 | 0.0480   | 0.9617   |
| Schoolmate/teacher/student<br>(Others) | 0.0233  | 0.0243 | 0.9580   | 0.3381   |
| Contact mode (%)                       |         |        |          |          |
| Face-to-face                           | 0.0114  | 0.0306 | 0.3719   | 0.7099   |
| Voice only<br>(Text only)              | 0.0101  | 0.0452 | 0.2236   | 0.8231   |
| Contact purpose (%)                    |         |        |          |          |
| Work/school                            | -0.0163 | 0.0218 | 0.7465   | 0.4554   |
| Daily routine<br>(Others)              | -0.0123 | 0.0318 | 0.3879   | 0.6981   |
| Contact duration (%)                   |         |        |          |          |
| (<5 min)                               |         |        |          |          |
| 5-59 min                               | 0.0118  | 0.0257 | 0.4613   | 0.6446   |
| 1-4 hours                              | 0.0207  | 0.0314 | 0.6607   | 0.5088   |
| 4 hours or more                        | 0.0225  | 0.0455 | 0.4952   | 0.6205   |
| Average mood of diary keeper           | 0.5074  | 0.0300 | 16.9135  | 0.0000   |
| Average mood of network neighbours     |         |        |          |          |
| One degree of separation               | 0.0878  | 0.0371 | 2.3643   | 0.0181   |
| Two degrees of separation              | 0.0789  | 0.0401 | 1.9684   | 0.0490   |
| All others                             | 0.0217  | 0.0361 | 0.6026   | 0.5468   |



members at degree 3 and beyond. The average mood score of diary keepers remained highly influential, with an estimated size of 0.507 ( $p < 0.0001$ ), while other covariates were irrelevant to how network members' moods varied.

## DISCUSSION

With higher quality data collected through a new method (online diary) than those collected from one-shot survey data, our results about concurrent mood over interpersonal contacts in daily life are consistent with those of the Framingham Heart Study, which analysed 20 years of historical data.<sup>4</sup> Unlike other studies of egocentric networks, however, we analysed how personal moods of >10 000 network members were linked to one another in 133 complete contact networks, based on the information our diary keepers provided.

To be more consistent with how we collected and analysed the data, we focus on 'concurrent mood' in our subsequent discussions. We actually analysed a hybrid construct derived from the average mood for each ego-alter pair, even though conceptually, we have relied on ego's perceptions of both ego's and alter's moods during social interactions. Because each diary keeper judged her own mood and the mood of each contacted person at the same time, it is reasonable to assume that the two mood scores are strongly associated. To take potential 'raters' effects' into account, more precisely, we took the essential step to include the diary keepers' mood scores when in contact with specific members in the model, which were most influential, as shown in [table 1](#).

We tried various analytic strategies and model selections before achieving our final models. For instance, some network members were dropped out of the analysis mainly because they lacked any network neighbours separated by one or two degrees. We included the current covariates in the final mixed-effects model by standard variable selection procedures, which also took into account how factors on the individual, tie and contact levels might be linked to how diary keepers rated the mood scores. The conclusions from our analyses varied slightly when we used somewhat different criteria to choose network members and covariates for the modelling.

Among members' characteristics, age and gender turned out not to be significant in the exploratory data analysis. Since >65% of the members were female, we kept gender in the final model for adjustment. Because tie strength and the relationship between the member and the diary keeper were believed to be influential factors, we retained the relationship in the model even though it lacked strong significance. Compared with other factors, contact features were supposed to be more influential. We have included contact mode, purpose, duration and diary keeper's own mood during each contact, after excluding variables that were not statistically significant, such as when and where the contact took place. In addition to these covariates, it is possible that

some other observed or unobserved factors also might be relevant to diary keepers' features but not included in the model. To reduce such potential impacts on the estimated coefficients, we added random components representing the variation among diary keepers. Specifically, our mixed-effects model showed that the average mood of network members linked by only one step contributed significantly to a given member's mood, as did the average mood of those members separated by two degrees. Those members who were separated by three degrees or more did not show a clear association.

We reached the findings by a special longitudinal design that followed up 133 participants with online contact diaries for 7 months. With the advantage of a web application, the ClickDiary program offers a friendly interface to collect detailed information about an egocentric network, the estimated relationships among all network members and the mood status of both parties during each contact. To minimise recall bias, the program allowed diary keepers to record the main contacts that occurred with the same person only within the past 24 hours. That is, ClickDiary encouraged participants to enter the information about their daily contacts as soon as possible.

While there is no gold standard to evaluate the extent to which these diary entries are valid, it would help to cross-check the basic entries against similar studies of contact diaries. Participants in this study, for example, recorded an average of 12 contacts per day, which was very close to, although slightly fewer than, the number of contacts in compatible social surveys and more conventional paper-pencil diary studies in China, Hong Kong and Taiwan.<sup>21 22</sup> As discussed earlier, one major difference of the ClickDiary lies in the stricter criteria about what counts as a contact. Most previous diary studies either included all fleeting contacts or used 'two to three words in exchange' as the minimum requirement for enlisting contacts, whereas the ClickDiary asked for only the contacts that involved at least 'three sentences'. This last unique feature may also justify the validity of the basic profile of the findings from the ClickDiary.

By cross-checking the mutual ratings of a network member's mood, we were able to evaluate the judgement errors from the way diary keepers rated how others felt during a specific contact. The same rare data also enabled us to evaluate another major concern: how accurately the diary keepers judged the strength of ties among their network members. As in large probability sampling surveys on self-reported egocentric networks, which always involved a risk of informant inaccuracy, we asked the focal persons to judge the strength of ties among those surrounding them. Unlike those surveys where the respondents rated the degree of acquaintance among a small number of confidants,<sup>23-25</sup> however, our diary keepers tended to have a tough task, because their contact networks usually stretched far and beyond such core networks. As a result, they often needed to estimate how well any two individuals on their contact lists knew each other even though they did not know either

individual well enough in the first place. Although some diary keepers were likely to report 'do not know' under such a circumstance, it remains critical to seek a validity criterion to cross-check their responses.

Another option for conducting a concordant pairs analysis would be to use an intraclass correlation coefficient or Kappa coefficient. Since concordant pairs accounted for only 49.1% of all mood rating pairs, and only 74 of all diary keepers' contact persons also rated their own moods for the same contacts, the agreement is poor, with a small weighted Kappa value of about 0.096. Instead of relying solely on a summary index of agreement, we further investigated the disagreement structure. Given that among the 50.9% cases that showed disagreement, 43.9% revealed only a one-category difference, we assume that about 93% of these score differences (between the moods rated by diary keeper and network member on the identical contact) would fall between  $-1$  and  $1$ , and 7% of the differences would be 2. The cross-checking results provided us an opportunity for performing a sensitivity analysis to assess robustness of our main findings.

As in a recent similar diary study,<sup>26</sup> thus, we were able to validate the alter-to-alter ties of constructed contact networks from these 74 diary keepers. It is noteworthy that these data may not represent all the ties among the 12070 members in 133 contact networks. Thus, constructing 'complete contact networks' out of the diary keepers' evaluations on how their network members were tied to each other remains a limitation. Were all of these network members also involved in keeping a ClickDiary, the estimated ties among them could have been verified by the extent to which they actually contacted one another during the same study period. The task of asking 2070 network members to record every contact they made for 7 months, however, would have been too costly, enormous and unfeasible.

As a more realistic, alternative strategy, our design of constructing 'proxy' complete contact networks by relying on 133 diary keepers has facilitated a rare analysis of mood correlations among members in different positions in egocentric networks. Such an approach could be further justified if other conditions also satisfied the assumption that two linked members (ie, any two alters who knew each other according to diary keeper's judgement) indeed made contacts with each other during the study period. To check this assumption, we have also analysed our subset data, which showed that any two members in a contact network who knew each other had a 90.5% chance to actually contact each other at least once during the 7 months. If the chance to contact each other turned out to be lower in some contact networks, however, we would have overestimated network members' effect on concurrent mood. In addition, other unmeasured external factors also might affect the mood scores. We treated such unmeasured effects as random variations among different persons.

From the model estimates shown in [table 1](#), we found that diary keepers tended to rate the contact persons

with higher mood scores when the contacts were face-to-face or lasted for a longer period of time. According to the seminal study on 'emotional contagion', which we outlined earlier, the underlying mechanism could be facial mimicry during everyday encounters.<sup>8</sup> When people had a chance to contact each other for a longer time or in person, their concurrent mood would have been more obvious.

In addition, the tendency of mood spread also could have been linked to personality traits. Extraversion, for example, could induce positive mood, and neuroticism could help predict negative mood.<sup>27</sup> Even though our study measured diary keepers' Big-5 personality traits, we could not examine how such personality traits influence concurrent mood among the contact persons. In other words, we asked diary keepers to judge their own Big-5 personality traits, but not those of the contact persons, who were the main actors of the study. We assume that diary keepers' Big-5 traits would be correlated with their own mood and that these traits would have directly or indirectly affected how they perceived the mood of their contact persons (network members). To adjust for the effects due to various characteristics of these diary keepers, we included relevant covariates and random components in the mixed-effects models. The models indeed showed that the covariate of diary keeper's mood with a contact person had a very large effect on the contact person's mood score.

To address the issue of potential response bias from diary keepers, we have further polished our models. For instance, diary keepers' dispositional mood would have strong effects on their own mood scores and those of contacted persons (or network members). The network members' mood scores would be affected by the tie strength and relationships between the members and diary keepers. Diary keepers' ratings also may vary by contact attributes, such as contact mode and duration of each one-on-one contact. To take these potential effects into account, we included covariates of these tie and contact factors in the mixed-effects model. To account for the various effects among the diary keepers on network members' mood scores that are not fully adjusted by the covariates in the model, we further added random components to increase the accuracy of the estimates.

In this study, we demonstrated that two members connecting with each other had a high probability of making actual contact during the study period. Future studies should further benefit by recording when a pair of network members actually contacted each other. With such information about the exact timing of each contact among the network members, it would be more feasible to identify the direction of contagion or diffusion of personal mood within the network.

It would have been ideal to model how network members' moods associate with one another, if the temporal observations of the contacts among members had been more complete. Diary keepers, in practice, rated and recorded a member's mood when they actually

contacted the member, which occurred about eight times, on average, for each member during the study period. The limited number of observations of the network members' mood somewhat prevented us from directly analysing mood spread within a network. Under the circumstances, we were restricted to examine concurrent mood among neighbouring members, with the assumption that a pair of acquaintances had contacted each other at least once during the study period of 7 months.

We were able to identify about 2.12% of the network members who had contacts with multiple diary keepers. We believe that the real percentage would be small, even though that percentage might have been underestimated, because it is nearly impossible in real life to identify exactly how many network members actually came into contact with multiple diary keepers. In addition, two diary keepers may have assigned different names for identical persons, or know them by different names. That circumstance might pose another limitation, but the effect on our major results would be minimal. Furthermore, with the current information about the contacted persons' personal background, our model cannot fully adjust for the effects of homophily,<sup>28</sup> most notably the similarities in personality traits, as well as other relevant risk factors. In particular, some pairs of daily contacts tend to be those who resemble each other in that they systematically rate their own and other people's moods in a similar manner.

In our study, we retrieved contact records from 133 eligible diary keepers. During the study period, 259 other volunteers also registered but turned out to be ineligible because they failed to comply with the requirements of the diary keeping. About 60% of these ineligible volunteers were college students recruited from several classes, who quit after a few tries, while many others only visited the platform once. To align with the routine practices of empirical studies, we have treated these volunteers as ineligible or 'not applicable' cases, whose diary entries were largely incomplete or too scarcely completed to qualify for any network analysis.

Collecting diary data in this prospective study was not an easy task, because diary keeping has proved to be highly demanding for many participants, even with financial incentives. Such a heavy burden prevented some participants from recording online diaries as required, such as keeping diaries at least three times a week and at least 10 times a month. To have a better grasp of interpersonal contacts in everyday life, we have actually lowered our requirements for qualifications by including participants who had recorded online diaries for at least 30 days during the study period of 7 months. To achieve the minimum requirements for statistical analysis within an egocentric contact network, we also included only those who had contacted at least 30 unique individuals.

In our mixed-effects model, we removed network members who had either no one-degree separation neighbouring members or no two-degree neighbours. There was no extra information available to impute from neighbouring members. We also excluded strangers for

three reasons, as described earlier. Mainly, interactions with strangers carry very different implications in studies of network diffusion. Conceptually, strangers are by default not part of one's personal networks. Even though we asked participants to record contacts with all individuals, it was actually unusual or unnatural for participants to judge how a stranger was connected with their network members, which would require a somewhat different research framework and analytic strategies. It would be intriguing to explore, in some extended studies, whether and how interactions with strangers would bring about somewhat unique patterns of emotional contagion.

Like most other studies with a small sample size, the subjects who participated in the ClickDiary study volunteered without a strict sampling procedure. The resulting sample of diary keepers is thus skewed towards female, younger and better-educated subjects. As common in other diary studies that rely on a small sample of subjects, the main goal of our study was not using a representative sample to make an inference to the general population.<sup>11</sup> Rather, we used the detailed information about all contacts and ties to build 133 sophisticated complete contact networks, some of which intertwined with one another, which allowed us to examine how personal mood may occur concurrently in everyday life. After conducting detailed analyses of the complicated contact networks and cross-checking both parties' reports on personal moods involved in the identical contacts, nonetheless, we do observe clear patterns of concurrent mood in the first-order and second-order social ties. Our unique diary approach shows significant network autocorrelation of personal mood among the network members, even though we are unable to claim causal effects or a clear direction of mood association between ego and alters. Our approach, data and findings are particularly useful, in sum, given that observational studies can hardly provide sufficient empirical evidence as to how personal moods may spread to friends' friends through contacts. Furthermore, although our current observational data do not support claims of mood contagion, continuously improved designs in similar diary approaches could enhance the potential for addressing some of the hard questions about the causal effects of network contagion.

## CONCLUSIONS

In line with earlier studies about how emotions and moods emerge concurrently among network members, we aim to make a substantive contribution to the literature by extending the investigation to the mood averaged from a series of contacts between two individuals in everyday life. While the literature has focused on how emotions and moods transmit at the tie level, our study relies on a bottom-up approach that first scrutinises how such moods may vary at the contact level before aggregating the mood scores for each pair of a diary keeper and a contact person. We achieved this approach by collecting data with an improved version of contact diaries.

In addition to recording key contact features, as well as how each contact person was linked to the diary keeper, our participants also judged each contact person's mood during each specific contact and estimated how well each pair of contact persons knew each other. Such a version of contact diaries yield all ego-alter ties in an egocentric network, and generates nearly all alter-alter ties, which essentially enable us to construct comprehensive network structures surrounding each diary keeper. This methodological innovation, in turn, enhances our efforts to make the substantive contribution to the literature about social networking and emotional contagion.

As with most other social network studies, it is relatively easy to collect empirical data about the ties between a focal person and those surrounding him/her (or 'ego-alter ties'), which are key indicators to understand the structure of an egocentric network.<sup>29</sup> It becomes highly challenging, however, to collect or reconstruct helpful information about the relationships among network members, which allows researchers to analyse the structures of a complete network. One convenient and flexible design in this study relied on some incentives and the sampling strategy to help diary keepers evaluate and confirm how well any pairs of their network members knew each other. In particular, our system assigned an 'absent tie' as the default value of the alter-alter tie (meaning the pair did not know each other), which was the case in about 78.4% of all alter-alter ties. When rating these ties, the diary keeper only needed to either confirm such an absent tie or change the option to either 'knew each other well' or 'knew each other, but not well'. With a median of 76 alters per ego, an average diary keeper managed to evaluate the strength of  $76 \times 75 / 2 = 2850$  alter pairs within the study period. Being better motivated to report and confirm such alter-alter ties, as a result, diary keepers in this study completed and verified about 99.97% of all ties, which allowed us to analyse concurrent mood among nearly all network members in egocentric networks.

Using special study designs in ClickDiary, we have been able to cross-check both the network members' moods and the tie strength among members by matching the diary keepers' estimates and some of the network members' own ratings. Future studies could make the best use of all network members' own reports to reconfirm the strength of ties with one another in complete networks. Such ultimate validity criteria would further verify, in a more comprehensive manner, how accurately diary keepers had judged the ties among the members in their personal networks, even though the distributions of such estimates were similar to those of previous paper-pencil diary studies. Most notably, our findings imply that similar personal mood can occur simultaneously, to varying extents, among the friends, relatives and other acquaintances clustered around different locations within personal networks. Applying the core concepts of network diffusion and richly designed contact-by-contact data to the inquiries about personal well-being, the current study sheds new light on how social network perspectives can help explain the ways individuals express

their personal moods concurrently during social interactions in everyday life.

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## REFERENCES

- Hatfield E, Cacioppo JT, Rapson RL, *et al*. Emotional contagion. *Current Directions in Psychological Science* 1993;2:96–100.
- Hill EM, Griffiths FE, House T. Spreading of healthy mood in adolescent social networks. *Proc Biol Sci* 2015;282:20151180.
- Hill AL, Rand DG, Nowak MA, *et al*. Emotions as infectious diseases in a large social network: the SISa model. *P Roy Soc B-Biol Sci* 2010;277:3827–35.
- Fowler JH, Christakis NA. Dynamic spread of happiness in a large social network: longitudinal analysis over 20 years in the Framingham Heart Study. *BMJ* 2008;337:a2338.
- Rosenquist JN, Fowler JH, Christakis NA. Social network determinants of depression. *Mol Psychiatry* 2011;16:273–81.
- Centola D. The spread of behavior in an online social network experiment. *Science* 2010;329:1194–7.
- Coviello L, Sohn Y, Kramer AD, *et al*. Detecting emotional contagion in massive social networks. *PLoS One* 2014;9:e90315.
- Hatfield E, Cacioppo JT, Rapson RL. *Emotional contagion*. Cambridge: Cambridge University Press, 1994.
- Deng H, Hu P. Matching your face or appraising the situation: two paths to emotional contagion. *Front Psychol* 2017;8.
- Dezecache G, Conty L, Chadwick M, *et al*. Evidence for unintentional emotional contagion beyond dyads. *PLoS One* 2013;8:e67371.
- Kramer AD, Guillory JE, Hancock JT. Experimental evidence of massive-scale emotional contagion through social networks. *Proc Natl Acad Sci U S A* 2014;111:8788–90.
- Shalizi CR, Thomas AC. Homophily and contagion are generically confounded in observational social network studies. *Sociol Methods Res* 2011;40:211–39.
- Fu Yc, Ho HC, Chen HM. Weak ties and contact initiation in everyday life: Exploring contextual variations from contact diaries. *Social Networks* 2013;35:279–87.

14. Kahneman D, Krueger AB, Schkade DA, *et al.* A survey method for characterizing daily life experience: the day reconstruction method. *Science* 2004;306:1776–80.
15. Fu Yc. Contact diaries: building archives of actual and comprehensive personal networks. *Field Methods* 2007;19:194–217.
16. Chan TC, Yen TJ, Fu Yc, *et al.* Clickdiary: online tracking of health behaviors and mood. *J Med Internet Res* 2015;17:e147.
17. Gosling SD, Rentfrow PJ, Swann WB. A very brief measure of the big-five personality domains. *J Res Pers* 2003;37:504–28.
18. Chang YH, Fu Yc. The evolution of the Taiwan Social Change Survey. *JGSS Symposium 2003: Birth of JGSS and Its Fruit: Social Surveys in Different Countries and Areas and JGSS*. East Osaka City: Institute of Regional Studies, Osaka University of Commerce, 2004:149–60.
19. Read JM, Lessler J, Riley S, *et al.* Social mixing patterns in rural and urban areas of southern China. *P Roy Soc B-Biol Sci* 2014;281:20140268.
20. CRAN. *Nlme: linear and nonlinear mixed effects models [program], R package version 3.1-128 version*, 2016.
21. Fu Yc. Measuring personal networks with daily contacts: a single-item survey question and the contact diary. *Social Networks* 2005;27:169–86.
22. Fu Yc, Wang DW, Chuang JH. Representative contact diaries for modeling the spread of infectious diseases in Taiwan. *PLoS One* 2012;7.
23. Burt RS. Network items and the general social survey. *Soc Networks* 1984;6:293–339.
24. Marsden PV. Network data and measurement. *Annu Rev Sociol* 1990;16:435–63.
25. Marsden PV. Survey methods for network data. In: Scott J, Carrington PJ, eds. *The sage handbook of social network analysis*. London: Sage Publications, 2011:370–88.
26. Fu Yc, Velema TA, Hwang JS. Upward contacts in everyday life: Benefits of reaching hierarchical relations in ego-centered networks. *Social Networks* 2018;54:266–78.
27. Gomez R, Cooper A, Gomez A. Susceptibility to positive and negative mood states: test of Eysenck's, Gray's and Newman's theories. *Pers Individ Dif* 2000;29:351–65.
28. Bollen J, Gonçalves B, van de Leemput I, *et al.* The happiness paradox: your friends are happier than you. *EPJ Data Sci* 2017;6.
29. Kaptein M, Nass C, Parvinen P, *et al.* *Nice to know you: familiarity and influence in social networks*: System Sciences (HICSS), 2013 46th Hawaii International Conference on, 2013.