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OPEN MTHFR C677T polymorphism interaction with heavy alcohol consumption increases head and neck carcinoma risk

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MTHFR C677T polymorphism has been indicated to be a risk factor for cancers, but its association with head and neck cancer (HNC) risk remains inconclusive. In the present study, we aimed to get a more precise estimation by performing a quantitative meta-analysis. Published papers up to Jun 2014 was searched and screened. Necessary information was rigorously extracted for data pooling and analyzing, and then, subgroup analyses on ethnicity, source of controls, sample size, tumor type, smoking and drinking status were also carried out. As a result, twenty-three case-control studies including 14298 subjects were included. The overall data failed to reveal a significant association between MTHFR C677T polymorphism and HNC risk (homozygote comparison model: OR = 1.16; 95%Cl = 0.93-1.45; dominant model: OR = 1.05; 95%Cl = 0.90-1.21; recessive model: OR = 1.14; 95%CI = 0.93-1.38). However, in the subgroup analysis about drinking status, increase risk was shown in the heavy drinking subgroup (TT vs CC: OR = 3.11; 95%Cl = 1.52-3.02). In conclusion, the results of the present study suggest that Homozygous TT alleles of MTHFR C677T polymorphism might be a risk factor for HNC among individuals who have a heavy drinking history. Further studies are needed to get a more definitive conclusion.

Head and neck carcinoma (HNC), the sixth most frequent kind of cancer worldwide, is a group of biologically similar cancers that originate from head and neck regions such as oral cavity, pharyngeal cavity, and larynx¹. Previous reports showed that life-style factors such as smoking, drinking, betel quid chewing, papilloma virus infection, and exposure to toxic substances are possible etiological risk factors for HNC^{2,3}. Nevertheless, though many individuals are exposed to these external factors, HNC develops only in a small proportion of the exposed people, indicating that intrinsic factors such as genetic polymorphism might play critical roles in its carcinogenic mechanisms.

Generally, folate plays a fundamental role of providing methyl groups for de novo deoxynucleoside synthesis and for intracellular methylation reactions in humans⁴. Low folate levels may result in uracil misincorporation during DNA synthesis, impaired DNA repair and chromosomal damage⁵. Several key enzymes were involved in the folate metabolism and their functions may have an effect on folate levels and DNA methylation. Methylenetetrahydrofolate reductase (MTHFR), a key enzyme for intracellular folate hemeostasis and metabolism, catalyses the irreversible conversion of 5,10-methylenetetrahydrofolate to 5-methyltetrahydrofolate that is the primary circulating form of folate, providing methyl groups for the methylation of homocysteine to methionine^{6,7}. A common variation of MTHFR (rs1801133), C677T in exon 4 (Ala222Val), may be implicated in tumorigenesis with alteration of MTHFR enzyme activity⁸. The

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Figure 1. A flow chart of the literature screening.

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homozygous genotype 677TT has been indicated to have only 30% of the *MTHFR* enzyme activity of the 677CC wild-type genotype⁹. Therefore, alteration of the enzyme activity resulted from the polymorphism has been thought to be associated with cancer progression and development¹⁰.

Published data on the association of *MTHFR* polymorphism with HNC have generated inconclusive results. Clarifying this association may help us better understand the possible risk of HNC and therefore contribute to its prevention. Previously, we assessed the relationship between *MTHFR* polymorphism and oral cancer risk¹¹. Recently, several original studies have been reported. Thus, in the present study, we performed an updated meta-analysis including the recent investigations that were conducted on head and neck cancers.

Materials and methods

Literature search strategy. The meta-analysis was presented according the PRISMA-MOOSE statement. We carried out a search in the internet covering well-known biomedical databases such as Medline, EMBASE, and CNKI without a language limitation. Papers published up to Jun 1, 2014 were included. The literature selection was performed from Jun 1, 2014 to Jun 30, 2014. The following keywords were used for searching: *methylenetetrahydrofolate reductase, MTHFR, head and neck, oral, pharynx, larynx, thyroid, mouth, neoplasm, tumor, cancer, variation* and *polymorphism*. All searched studies were retrieved and the bibliographies were checked for other relevant publications. Review articles and bibliographies of other relevant studies identified were hand searched to find additional eligible studies.

Inclusion criteria. The following criteria were used for the literature selection: First, studies should concern the association of *MTHFR* C677T polymorphism with HNC risk; second, studies should be case-control or cohort designed; third, papers must offer the size of the sample, odds ratios (ORs) and their 95% confidence intervals (CIs), the genetic distribution or the information that can help infer the results. After deliberate searching, we reviewed all papers in accordance with the criteria defined above for further analysis.

Data extraction. Information was carefully extracted from all eligible papers by two of the authors (XZ and JS) independently according to the inclusion criteria mentioned above. If their results were conflicting, an agreement was reached following a discussion. If a consensus could not be reached, another author (DL) was consulted to resolve the dispute and then a final decision was made by the majority of the votes. The extracted data were entered into a database.

Statistical analysis. The odds ratio (OR) of *MTHFR* C677T polymorphisms and HNC risk was estimated for each study. The pooled ORs were performed for a homozygote comparison model (TT versus CC), a dominant model (TT + CT versus CC) and a recessive model (TT versus CT + CC), respectively. For detection of any possible sample size biases, the OR and its 95% confidence interval (CI) to each study was plotted against the number of participants respectively. For assessment of the heterogeneity between studies, two indexes were calculated. One was I^2 metric, with I^2 =0-25% indicating no heterogeneity, I^2 =25-50% indicating moderate heterogeneity, and I^2 > 50% indicating large heterogeneity¹². The other index was a Chi-square based Q statistic test. If the result of the Q-test was P > 0.1 (indicating no heterogeneity), ORs were pooled according to the fixed-effect model (Mantel-Haenszel), Otherwise, the random-effect model (DerSimonian and laird) was used. The significance of the pooled ORs was determined by Z-test. The Hardy-Weinberg equilibrium (HWE) was assessed via Fisher's exact

| First Author | Publication Year | Number of Cases (male/ female) | Number of Controls (male/female) | Type of controls | Mean age, year Cases/ Controls | Racial decent | Types | Country | Ref No. |
|---------------|---------------------|--------------------------------------|--|--|---|-----------------------|-------------|--------------|------------|
| Weinstein | 2002 | 519 (NA/NA) | 629 (NA/NA) | Healthy controls (Age-, gender- matched; population-based) | 63.2/61.0 | Caucasian | Oral cavity | USA | 43 |
| Kureshi | 2004 | 50 (33/17) | 54 (30/24) | Healthy controls (Age-, gender- matched; population-based) | 51.3/50.2 | Mixed | Combined | Pakistan | 28 |
| Capaccio | 2005 | 65 (57/8) | 100 (88/12) | Healthy controls (Age-, gender-, drinking-, smoking status- matched; population-based) | 61.3/58.8 | Caucasian | Combined | Italy | 23 |
| Neumann | 2005 | 537 (411/126) | 545 (401/144) | Non-cancerous controls (Age-,sex-, smoking status-matched; hospital-based) | NA/NA | Caucasian Combined | | USA | 29 |
| Vairaktaris | 2006 | 110 (94/16) | 120 (102/18) | Healthy controls (Age-, sex, ethnicity- matched; population-based) | 52.1/51.5 | Caucasian Oral cavity | | GreekGerman | 41 |
| Hsiung | 2007 | 278 (193/85) | 526 (377/149) | Healthy controls (Age-, gender, town of residence-matched; population-based) | 60.1/61.0 | Mixed | Combined | USA | 26 |
| Reljic | 2007 | 81 (NA/NA) | 102 (37/65) | Healthy controls (Age-matched; population-based) | NA/NA | Caucasian | Combined | Croatia | 33 |
| Suzuki | 2007 | 237 (188/49) | 711 (564/147) | Non-cancerous controls (Age-,sex- matched; hospital-based) | 57.9/58.4 | Asian | Combined | Japan | 39 |
| Ni | 2008 | 207 (189/18) | 400 (362/38) | Healthy controls (Age-,sex-matched; population-based) | NA/NA | Asian | Larynx | China | 30 |
| Siraj | 2008 | 223 (NA/NA) | 513 (NANA) | Healthy controls (Age-matched; population-based) | NA/NA | Mixed | Thyroid | Saudi Arabia | 36 |
| Solomon | 2008 | 126 (NA/NA) | 100 (NA/NA) | Healthy individuals(population-based) | 54.7/55.4 | Asian | Oral cavity | India | 37 |
| Сао | 2010 | 529 (339/190) | 577 (367/210) | Healthy controls (Age-,sex-, ethnicity- matched; population-based) | 46.1/45.4 | Asian | Nasopharynx | China | 22 |
| Kruszyna | 2010 | 131 (131/0) | 250 (250/0) | Healthy controls (population-based) | 59.4/56.8 | Caucasian | Larynx | Poland | 27 |
| Rodrigues | 2010 | 100 (84/16) | 100 (76/24) | Non-cancerous controls (hospital-based) | 59.5/43.6 | Mixed | Combined | Brazil | 34 |
| Fard-Esfahani | 2011 | 154 (34/120) | 198 (50/148) | Non-cancerous controls (hospital-based) | NA/NA | Mixed | Thyroid | Iran | 24 |
| Prasad | 2011 | 97 (27/70) | 241 (116/125) | Healthy controls (population-based) | NA/NA | Asian | Thyroid | India | 32 |
| Sailasree | 2011 | 130 (88/42) | 139 (92/47) | Non-cancerous controls (Age-,gender- matched; hospital-based) | 58.0/57.0 | Asian | Oral cavity | India | 35 |
| Supic | 2011 | 96 (72/24) | 162 (130/32) | Non-cancerous controls(Age-, gender-, ethnicity-matched; population-based) | 58.0/58.0 | Caucasian | Oral cavity | Serbia | 38 |
| Tsai | 2011 | 620 (583/37) | 620 (572/48) | Healthy controls (Age-, gender-, habits- matched; population-based) | 65.5/63.5 | Asian | Oral cavity | China | 40 |
| Galbiatti | 2012 | 322 (280/42) | 531 (384/147) | Healthy controls (population-based) | NA/NA | Mixed | Combined | Brazil | 25 |
| Ozdemir | 2012 | 60 (11/49) | 50 (21/29) | Healthy controls (population-based) | 55.3/68.6 | Mixed | Thyroid | Turkey | 31 |
| Vylliotis | 2013 | 160 (NA/NA) | 168 (NA/NA) | Healthy controls (Age-, gender-, ethnicity-, working environment- matched; population-based) | 58.5/54.7 | Caucasian | Oral cavity | Greece | 42 |
| Kweon | 2014 | 2194 (404/1790) | 1669 (812/857) | Healthy controls (Age-, gender-, matched; population-based) | 50.6/52.2 | Asian | Thyroid | Korea | 44 |

Table 1. Characteristics of studies included in the meta-analysis. NA: not available

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test. Publication bias was assessed by visual inspection of funnel plots¹³, in which the standard error of log (OR) of each study was plotted against its log (OR). The symmetry of the funnel plot was further evaluated by Egger's linear regression test¹⁴. An asymmetric plot indicates a possible publication bias. Statistical analysis was conducted by using the STATA 11.0 software (Stata Corporation, Texas).

Results

Study characteristics. Publications relevant to the key words were retrieved and screened originally. A total of ninety-seven studies were searched and screened for retrieval, of which sixty-seven irrelevant studies were excluded. Then, three review articles¹⁵⁻¹⁷ were excluded. Next, three studies that were not case-control designed¹⁸⁻²⁰ (also not cohort designed) were discarded. Thus, a total of twenty-four papers were included for data extraction. Nevertheless, one duplicate publication²¹ was further excluded. Lastly, twenty-three case-control studies²²⁻⁴⁴ were selected (Fig. 1).

All the included studies were written in English except for one study in Chinese³⁰. We established a database according to the extracted information from each article. The relevant information was listed in Table 1. According to the lists, the first author and the number and characteristics of cases and controls

| | | | Cases | | | Co | ontrols | | HWE (control) | | |
|---------------|------|----------------------|---------|------|-----|---------|---------|-----|---------------|--------|--|
| First Author | Year | Genotyping method | TT | СТ | сс | TT | СТ | сс | Chi-squre | Р | |
| Weinstein | 2002 | PCR-RFLP | 15 | 53 | 67 | 15 | 62 | 69 | 0.038 | > 0.05 | |
| Kureshi | 2004 | PCR | 0 | 12 | 22 | 4 | 18 | 32 | 0.420 | > 0.05 | |
| Capaccio | 2005 | PCR | 14 | 33 | 18 | 18 | 46 | 36 | 0.242 | > 0.05 | |
| Neumann | 2005 | PCR | 35 | 244 | 258 | 51 | 216 | 278 | 0.914 | > 0.05 | |
| Vairaktaris | 2006 | PCR-RFLP | 6 | 76 | 28 | 10 | 65 | 45 | 4.065 | < 0.05 | |
| Hsiung | 2007 | PCR | 149 (a) | - | 128 | 306 (a) | - | 218 | - | - | |
| Reljic | 2007 | PCR-RFLP | 9 | 27 | 45 | 8 | 59 | 35 | 6.074 | < 0.05 | |
| Suzuki | 2007 | Taqman | 36 | 113 | 88 | 128 | 331 | 252 | 1.121 | > 0.05 | |
| Ni | 2008 | PCR-RFLP | 64 | 95 | 48 | 61 | 187 | 152 | 0.078 | > 0.05 | |
| Siraj | 2008 | PCR | 1 | 18 | 30 | 13 | 126 | 372 | 0.351 | > 0.05 | |
| Solomon | 2008 | PCR-RFLP | 23 | 55 | 48 | 10 | 42 | 48 | 0.033 | > 0.05 | |
| Cao | 2010 | PCR-RFLP | 32 | 169 | 310 | 30 | 188 | 334 | 0.275 | > 0.05 | |
| Kruszyna | 2010 | PCR-RFLP | 10 | 52 | 69 | 20 | 104 | 126 | 0.052 | > 0.05 | |
| Rodrigues | 2010 | PCR-RFLP | 13 | 43 | 44 | 14 | 40 | 46 | 1.182 | > 0.05 | |
| Fard-Esfahani | 2011 | Multiplex PCR | 14 | 71 | 69 | 8 | 108 | 82 | 14.224 | < 0.05 | |
| Prasad | 2011 | PCR-RFLP | 1 | 10 | 86 | 1 | 12 | 228 | 3.311 | > 0.05 | |
| Sailasree | 2011 | PCR-RFLP | 1 | 8 | 92 | 1 | 29 | 108 | 0.400 | > 0.05 | |
| Supic | 2011 | PCR-RFLP | 14 | 32 | 50 | 16 | 66 | 80 | 0.193 | > 0.05 | |
| Tsai | 2011 | PCR-RFLP | 43 | 186 | 391 | 62 | 236 | 322 | 3.606 | > 0.05 | |
| Galbiatti | 2012 | PCR-RFLP | 45 | 147 | 130 | 55 | 250 | 226 | 1.358 | > 0.05 | |
| Ozdemir | 2012 | Real-time PCR | 7 | 25 | 28 | 3 | 14 | 33 | 0.781 | > 0.05 | |
| Vylliotis | 2013 | PCR-RFLP | 6 | 76 | 28 | 10 | 65 | 45 | 4.065 | < 0.05 | |
| Kweon | 2014 | PCR-RFLP | 422 | 1050 | 722 | 291 | 851 | 527 | 2.748 | > 0.05 | |

Table 2. Distribution of MTHFR C677T genotype among cancer cases and controls. (a): TT+CT

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for each study as well as other necessary information were presented. As shown in this table, the selected studies involved cancers originated from oral cavity, pharynx, larynx and thyroid.

In the included studies, there were eight groups of Caucasians^{23,27,29,33,38,41-43}, eight of Asi ans^{22,30,32,35,37,39,40,44} and seven of mixed populations^{24–26,28,31,34,36}. Information about smoking could be extracted from five studies^{22,25,30,39,40} and data regarding drinking status were also available from five studies^{23,25,37-39}. The distributions of *MTHFR* genotype of the included studies were also presented (Table 2). Of note, the data about CT and TT in the study by Hsiung *et al.*²⁶ were combined as TT + CT, and therefore, the relevant data were only included in the dominant model for the whole evaluation. The genetic distributions of the control groups were consistent with HWE except for those in four studies^{24,33,41,42}.

Meta-analysis results. We evaluated the heterogeneity for the homozygote comparison model (TT versus CC), dominant model (TT + CT versus CC) and recessive model (TT versus CT + CC), respectively. As shown in Table 3, the heterogeneity for the overall data was significant in the three models, respectively, because the P value was less than 0.1 for Q-tests and I-squared values indicated more than moderate heterogeneity. Therefore, the random-effect models were used in the present meta-analysis.

Table 3 lists the main results of the meta-analysis. The overall data in the homozygote comparison model (OR = 1.16; 95%CI = 0.93-1.45), the dominant (OR = 1.05; 95%CI = 0.90-1.21) and the recessive models (OR = 1.14; 95%CI = 0.93-1.38) failed to reveal a marked association between *MTHFR* C677T polymorphism with HNC risk (Fig. 2).

Similar results were observed in the subgroups regarding ethnicity, source of controls, sample size, tumor type and smoking status. Nevertheless, in the subgroup analysis concerning drinking status, elevated risk was shown in the heavy drinking subgroup under the homozygote comparison model (OR = 3.11; 95%CI = 1.52-3.02), indicating that TT carriers who have a heavy drinking history might have an increased HNC risk.

Sensitivity analysis and Bias diagnostics. To test the stability of the overall results, we first excluded the studies whose genetic distributions in controls markedly deviated from HWE, given that the deviation might result in any bias⁴⁵. The significance of the overall data in the three models was also not statistically altered. Then, we also conducted one-way sensitivity analysis⁴⁶ by deleting a single study

| | | TT vs CC | | | (TT+CT) vs CC | | | | TT vs (CT+CC) | | | | |
|-------------------|----------------------|-------------------------|-------|---------------|----------------|-------------------------|-------|---------------|----------------|-------------------------|-------|---------------|----------------|
| | No. (cases/controls) | OR (95%CI) | Р | P (Q-test) | I ² | OR (95%CI) | Р | P (Q-test) | I ² | OR (95%CI) | Р | P (Q-test) | I ² |
| Total | 6354/7944 | 1.16 (0.93- 1.45) | 0.187 | 0.001 | 54.5% | 1.05 (0.90- 1.21) | 0.540 | 0.000 | 68.7% | 1.14 (0.93- 1.38) | 0.201 | 0.006 | 48.9% |
| Ethnicity | 1 | | | | | 1 | | 1 | | | | | |
| Caucasian | 1265/1545 | 0.95 (0.72- 1.26) | 0.742 | 0.860 | 0.0% | 1.05 (0.81- 1.38) | 0.702 | 0.013 | 60.7% | 0.92 (0.71- 1.20) | 0.533 | 0.511 | 0.0% |
| Mixed | 996/1968 | 1.41 (0.99- 1.99) | 0.054 | 0.489 | 0.0% | 1.06 (0.85- 1.33) | 0.597 | 0.143 | 37.5% | 1.38 (0.99- 1.92) | 0.056 | 0.415 | 0.2% |
| Asian | 4093/4431 | 1.24 (0.80- 1.91) | 0.331 | 0.000 | 80.3% | 1.02 (0.78- 1.34) | 0.878 | 0.000 | 82.7% | 1.21 (0.86- 1.70) | 0.279 | 0.001 | 73.1% |
| Source of control | Source of control | | | | | | | | | | | | |
| РВ | 5225/6252 | 1.24 (0.94- 1.63) | 0.121 | 0.001 | 59.0% | 1.10 (0.92- 1.32) | 0.280 | 0.000 | 72.6% | 1.21 (0.97- 1.51) | 0.086 | 0.019 | 46.4% |
| НВ | 1129/1692 | 0.88 (0.66- 1.17) | 0.377 | 0.392 | 2.5% | 0.92 (0.71- 1.19) | 0.512 | 0.084 | 51.3% | 0.91 (0.63- 1.34) | 0.646 | 0.181 | 36.1% |
| Sample size | | | | | | | | | | | | | |
| <500 | 1400/1881 | 1.28 (0.97- 1.68) | 0.083 | 0.794 | 0.0% | 1.07 (0.83- 1.38) | 0.613 | 0.001 | 62.2% | 1.24 (0.95- 1.60) | 0.111 | 0.637 | 0.0% |
| ≧500 | 4954/6063 | 1.09 (0.76- 1.55) | 0.641 | 0.000 | 80.6% | 1.03 (0.85- 1.23) | 0.785 | 0.000 | 77.1% | 1.08 (0.80- 1.45) | 0.630 | 0.000 | 76.5% |
| Smoking status | 1 | | , | 1 | | 1 | | 1 | 1 | | | 1 | |
| Never-smoking | 606/919 | 0.79 (0.45- 1.37) | 0.394 | 0.635 | 0.0% | 0.90 (0.71- 1.14) | 0.376 | 0.605 | 0.0% | 0.81 (0.52- 1.27) | 0.362 | 0.967 | 0.0% |
| Ever-smoking | 1223/1013 | 0.73 (0.49- 1.10) | 0.134 | 0.018 | 82.1% | 1.24 (0.59- 2.58) | 0.573 | 0.000 | 94.2% | 0.86 (0.62- 1.21) | 0.392 | 0.091 | 58.3% |
| Drinking status | · | | | | | | | | | | | | |
| Light-drinking | 201/474 | 0.78 (0.32- 1.90) | 0.588 | 0.455 | 0.0% | 1.05 (0.75- 1.46) | 0.795 | 0.431 | 0.0% | 0.99 (0.56- 1.73) | 0.961 | 0.717 | 0.0% |
| Heavy-drinking | 409/519 | 3.11 (1.52- 3.02) | 0.002 | 0.832 | 0.0% | 1.94 (0.88- 4.27) | 0.099 | 0.001 | 82.1% | 1.62 (0.46- 5.73) | 0.457 | 0.001 | 86.0% |
| Tumor type | | | | | | | | | | | | | |
| Oral cavity | 1298/1406 | 1.03 (0.66- 1.61) | 0.883 | 0.099 | 43.8% | 0.99 (0.67- 1.46) | 0.950 | 0.000 | 78.8% | 0.99 (0.67- 1.45) | 0.939 | 0.168 | 34.0% |
| Combined | 1653/2667 | 0.97 (0.74- 1.28) | 0.830 | 0.293 | 17.9% | 0.95 (0.79- 1.15) | 0.613 | 0.079 | 45.0% | 0.97 (0.73- 1.29) | 0.827 | 0.191 | 31.0% |
| Larynx | 338/650 | 1.82 (0.51- 6.44) | 0.353 | 0.007 | 86.1% | 1.37 (0.63- 3.00) | 0.432 | 0.006 | 86.8% | 1.64 (0.64- 4.17) | 0.302 | 0.033 | 77.9% |
| Thyroid | 2554/2669 | 1.11 (0.92- 1.32) | 0.275 | 0.416 | 0.0% | 1.28 (0.89- 1.84) | 0.185 | 0.020 | 65.8% | 1.16 (0.99- 1.37) | 0.063 | 0.457 | 0.0% |
| Nasopharynx | 511/552 | 1.15 (0.68- 1.94) | 0.601 | _ | - | 0.99 (0.78- 1.27) | 0.958 | _ | _ | 1.16 (0.70- 1.94) | 0.566 | _ | _ |

Table 3. Main results of the pooled data in the meta-analysis PB: population-based; HB: hospital-basedHeavy-drinking: >46 g ethanol/week



Figure 2. Meta-analysis for the association of HNC risk with *MTHFR* C677T polymorphism (TT + CT versus CC; overall data).





Figure 3. Publication bias test for the overall data (TT + CT versus CC). (a) Funnel plot; (b) Egger's linear regression test.

each time. The statistical significance of the results was not also altered (data not shown), confirming the stability and credibility of the results.

Funnel plots were created for assessment of possible publication bias. Then, Egger's linear regression tests were used to assess the symmetry of the plots. As a result, the data suggest that the funnel plots were symmetrical for the three models (homozygote comparison model: t = 0.13, P > 0.05; dominant model: t = -1.29, P > 0.05; recessive model: t = 0.87, P > 0.05), suggesting that the publication bias may have little influence on the results (Fig. 3).

Discussion

The effects of *MTHFR* C677T polymorphism on cancer risk were controversial. Recent published meta-analyses showed that 677 C > T variation might contribute to the development of breast and eso-phageal cancer^{47,48}; however, a decreased cancer risk was observed in colorectal cancer⁴⁹. Thus, *MTHFR* C677T polymorphism may play various roles in different carcinomas. In the present study, though the overall results failed to suggest a relationship between *MTHFR* C677T variation and HNC risk, the sub-group analysis indicated that homozygous 677TT might be associated with increased susceptibility to HNC in individuals who have a heavy drinking history.

Considering the possible effects of the confounding factors on the overall data, we tried to carry out subgroup analyses. In subgroup analysis according to ethnicity, no significant association between *MTHFR* 677T allele with HNC risk was observed among the three subgroups. The results were in line with the overall data, indicating that the effects of discrepancies among different ethnicities on the results were not evident.

When the data were stratified by source of controls and sample size, respectively, the results were not significantly different from the whole results, indicating that these two factors have little influence on the results. Likewise, in the subgroup stratified by tumor type, the results were in accordance with the overall data, implying that the dissimilarity of tumor types rarely affect the pooled results.

Whether chronic smoking interacts with folate status in the pathophysiologic process of disorders remains controversial⁵⁰. Evidence suggests that low serum folic acid concentrations were commonly detected in smokers⁵¹. Nevertheless, cytological damage, an early event of carcinogenesis, is evident in the mouths of smokers, but it dose not correlate with folate status⁵². In the present meta-analysis, data about smoking status could be extracted from five included studies. The results showed that elevated risk was observed in neither the never-smoking subgroup nor the ever-smoking subgroup. The data failed to reveal an obvious interaction of smoking and MTHFR variation for HNC risk. However, when the data were divided by drinking status, elevated risk was shown in the heavy drinking subgroup under the homozygote comparison model. The data were in line with our previous meta-analysis about oral cancer¹¹. The underlying mechanisms were unclear. MTHFR 677C \rightarrow T polymorphism might play different roles for cancer risk according to the folate levels. It might reduce cancer risk when folate status is normal⁵³ while lead to impaired stability and reduced activity of the enzyme under low folate conditions⁵⁴. Notably, exposure to alcohol consumption might result in low folate status because chronic alcohol exposure may interfere with folate absorption by suppressing the folate carrier expression, thus decreasing the expression of folate transport proteins and reducing the hepatic uptake and renal conservation of circulating folate^{55,56}. Also, alcohol might act as a folate antagonist that is responsible for abnormalities in folate-mediated one-carbon metabolism⁵⁷. Moreover, folate depletion in mitochondria caused by chronic alcohol exposure might lead to abnormal DNA synthesis and DNA methylation that has been thought to be involved in the carcinogenesis process^{58,59}. This might explain the interactions of alcohol with 677TT genotype in the genesis of HNC. However, only five of the selected studies concerned this issue. In the meantime, the data should be interpreted with care because of the limited sample sizes.

Evident heterogeneity was shown in every model for the overall data, and thus, random-effect models were used to pool the data. Nevertheless, significant heterogeneities were removed in some subgroups while observed in other subgroups, indicating that the heterogeneities may be multifactorial. Besides the confounding factors considered in the subgroup analysis, other factors such as age, gender and life-style factors might contribute to the heterogeneities.

Several limitations might be involved in this meta-analysis. First, only publications written in English and Chinese were searched and selected. It is possible that articles written in other languages that might meet the inclusion criteria were missed. Thus, inevitable publication biases might exist though the funnel plots were tested to be symmetrical. Second, subgroup analyses regarding age, gender and other risk factors such as virus infection status and betel quid chewing have not been conducted because the data in the primary literature were insufficient. Third, hospital-based controls were used in some of the included studies, leading to non-differential misclassification bias. Moreover, the controls in some studies were not strictly matched to the cases. Therefore, any selection bias might have an influence on the overall results and future investigations with large sample sizes are required.

Despite the limitations, though the overall data of the present meta-analysis did not reveal an association of *MTHFR* C677T polymorphism with HNC risk, the subgroup analysis indicated that *MTHFR* 677TT alleles might increase HNC risk in individuals who have a heavy drinking history.

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Author Contributions

X.Z. and Q.Z. planned and experiment designed; X.Z. and J.S. wrote the manuscript draft; D.L. and Y.W. processed data and prepared the figures and tables; All authors reviewed the manuscript.

Additional Information

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