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Short Communication

Alcohol dehydrogenase IB (ADHIB) genotype, alcohol consumption and breast cancer risk by age 50 years in a German case–control study

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In a population-based study of 613 cases and 1082 controls, alcohol dehydrogenase 1B (ADH1B) genotype was not an independent risk factor for breast cancer, athough the possibility was raised that it modifies risk associated with high levels of alcohol consumption (OR 1.1, 95% confidence interval (Cl) 0.8–1.6 for ADH1B*1/*1 genotype vs 0.2, 95% Cl 0.1–1.0 for ADH1B*2 carriers). British Journal of Cancer (2005) **92,** 2039–2041. doi:10.1038/sj.bjc.6602608 www.bjcancer.com Published online 10 May 2005 © 2005 Cancer Research UK

Keywords: alcohol dehydrogenase; polymorphism; gene-environment interaction; breast cancer

Alcohol consumption is one of the few modifiable risk factors for breast cancer (Singletary and Gapstur, 2001; Collaborative Group on Hormonal Factors in Breast Cancer, 2002), and we also observed a dose-dependent effect of alcohol intake on breast cancer risk in a case - control study of women up to age 50 years in Germany (Kropp et al, 2001). However, drinking behaviour as well as susceptibility to alcohol-induced carcinogenesis may be influenced by individual genetic make-up. In this context, class I alcohol dehydrogenases (ADHs), which are important enzymes in the major pathway of alcohol metabolism in vivo and are expressed in normal mammary epithelium (Triano et al, 2003), may be relevant. The alcohol dehydrogenase 1B (ADH1B) gene (formerly called ADH2) exhibits genetic polymorphisms resulting in altered functional and catalytic properties in vitro (Agarwal, 2001). A strongly increased oxidation capability has been associated with the *2 allele (Bosron et al, 1983; Eriksson et al, 2001), an Arg47His substitution with reported allele frequencies ranging between 0 and 6.8% in Europeans (Brennan et al, 2004).

Previously, a significant inverse association between the ADH1B*2 allele and frequency of alcohol consumption was observed in a case-only study (Sturmer *et al*, 2002). We employed a population-based case-control study of women up to age 50 years to examine the potential effect of ADH1B genotype on the association between alcohol consumption and breast cancer risk.

MATERIALS AND METHODS

Study subjects and data collection

A detailed description of the underlying case-control study has been provided elsewhere (Chang-Claude *et al*, 2000). Patients up to

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Received 10 January 2005; revised 30 March 2005; accepted 1 April 2005; published online 10 May 2005

age 50 years with a diagnosis of primary incident *in situ* or invasive breast cancer were recruited between 1 January 1992 and 31 December 1995. Controls were selected randomly from lists of residents supplied by population registries; for each patient, two controls were matched according to exact age and study region. Written informed consent from all participants was obtained. The study was approved by the ethics committee of the University of Heidelberg.

In total, 706 (70.2%) women of the 1005 breast cancer patients who were alive when identified completed the study questionnaire. Of all 2257 eligible controls, 1381 (61.2%) participated. Detailed information on demographic characteristics and various risk factors was elicited by means of a self-administered questionnaire. Alcohol consumption was assessed for three time periods, that is, 15-20, 20-30 and 30-50 years, and for different types of beverages (beer, wine, aperitifs, liquor and spirits). The detailed method for the calculation of average daily alcohol intake has been described previously (Kropp *et al*, 2001).

Menopausal status was defined as the reported state half a year before the reference date, which was the date of diagnosis for cases and the date of completion of the questionnaire for controls. The status of women with previous hysterectomy not accompanied by bilateral oophorectomy was classified as unknown.

The present analysis was restricted to 613 cases and 1082 controls with at least one parent of German nationality and who could be successfully genotyped (three and five failures, respectively).

Genotyping ADH1B by TaqMan PCR

Genomic DNA was extracted from peripheral blood using the QIAamp DNA Blood Kit (Qiagen, Hilden, Germany). Primer and hybridisation probes were designed with Primer Express. Sense primer 5'-CTCTTTATTCTGTAGATGGTGGCTGTAG-3' and antisense primer 5'-GGGTCACCAGGTTGCCACTA-3' were used to amplify a 76 bp fragment containing the G47A polymorphism of the *ADH1B* gene. Two minor-groove-binding (MGB) DNA probes were synthesised. The probe corresponding to the wild type (5'-FAM-TCTGTCGCACAGATG-MGB-3') was labelled with

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6-FAM, and the probe corresponding to the mutation (5'-VIC-AATCTGTCACACAGATGA-MGB-3') was labelled with VIC at the 5'-end. The genotypes were analysed in software 'Sequence Detector' version 1.7 by procedure allelic discrimination after PCR.

Amplification was performed in a final volume of $25 \,\mu$ l containing 40 ng of DNA, 300 nm of each primer, 200 nm of each probe and 12.5 μ l of TaqMan Universal PCR Master Mix (Perkin-Elmer, Weiterstadt, Germany).

In every assay, negative controls as well as controls for the wild type, mutant type and heterozygote were included. The PCR conditions were as follows: $2 \min \text{ at } 50^{\circ}\text{C}$ plus 95°C for $10 \min$, followed by 40 cycles of denaturation at 92°C for 15 s, annealing and extension in one step at 60°C for $1 \min$. Genotyping was performed blinded to case-control status of the blood sample providers.

Statistical analysis

 χ^2 tests were used to assess deviation from Hardy–Weinberg equilibrium. Multivariate conditional logistic regression analysis with 5-year age strata was carried out using the PHREG procedure of the statistical software package SAS release 8.2 (SAS Institute, Cary, NC, USA).

In the multivariate model, we included several relevant variables influencing breast cancer risk (see Table 2). Variables that did not alter the estimates substantially, such as study region, body mass index, use of oral contraceptives and age at menarche, were not included in the analyses presented. Statistical interaction between *ADH1B* genotype and alcohol consumption was tested by using multiplicative interaction terms and evaluated by the likelihood ratio test.

RESULTS

Selected characteristics of the study population are depicted in Table 1. In all, 65 breast cancer patients (10.6%) and 109 controls (10.1%) were heterozygous or homozygous carriers of the ADH1B*2 allele, corresponding to allele frequencies of 0.06 and 0.05, respectively. The distribution of ADH1B genotypes was in Hardy–Weinberg equilibrium (P=0.4 for cases and P=0.6 for controls).

We did not observe an association between *ADH1B* genotype and several risk factors, including first-degree family history of breast cancer, body mass index, parity, breastfeeding or smoking status (data not shown). However, alcohol consumption was found to differ significantly by *ADH1B* genotype among controls, with a mean average daily alcohol intake of 7.2 g (standard deviation (s.d.) 9.1) in women with *ADH1B*1/*1* genotype and 5.8 g (s.d. 7.8) in carriers of the *2 allele (Wilcoxon's rank-sum test P = 0.01). Results were virtually identical when all study subjects were considered (data not shown).

The analysis revealed no main effect of ADH1B genotype on breast cancer risk (adjusted OR for carriers of ADH1B*2 allele vs ADH1B*1/*1 genotype being 1.0, 95% confidence interval (CI) 0.7 -1.4). However, multivariate analysis separately for carriers and noncarriers of the ADH1B*2 allele yielded differences in breast cancer risk associated with increasing levels of alcohol intake (Table 2). Among carriers of the *2 allele, there was a significant decreasing trend in breast cancer risk with increasing alcohol consumption (P-value for linear trend 0.04), although the odds ratio (OR) was significant only for the category of 12 g or more alcohol per day. In contrast, among women with ADH1B*1/*1 genotype, breast cancer risk increased with increasing alcohol consumption (P = 0.03). The interaction between ADH1B genotype and alcohol intake was borderline statistically significant for the highest category of alcohol consumption (P = 0.05) but not for low or moderate levels.

Table I Selected characteristics of the study population

	Cases (n = 613)	%	Controls (n = 1082)	%	P-value ^a
Mean age (s.d.)	42.4 (5.8)		42.6 (5.7)		
<i>Menopausal status</i> Premenopausal Postmenopausal	483 34	78.8 5.6	874 67	80.8 6.2	
Unknown	96	15.7	4	13.0	0.30
First-degree family his No	tory 536	87.4	1024	94.6	
Yes	77	12.6	58	5.4	< 0.01
Parity 0 I−2 children 3+ children	36 4 4 63	22.2 67.5 10.3	228 680 174	21.1 62.9 16.1	< 0.01
Breastfeeding ^b 0	153	32.1	247	28.9	
I–I2 months I2+ months	289 35	60.6 7.3	501 106	58.7 12.4	0.01
Education Low Intermediate High	80 412 121	3. 67.2 9.7	149 679 254	13.8 62.8 23.5	0.15
0			201	2010	0110
Smoking Never active Former smoker Current smoker	271 132 210	44.2 21.5 34.3	515 258 309	47.6 23.8 28.6	0.05
Alcohol consumption 0 1-5 6-11	(g day on averag 117 233 113	ge ⁻¹) 19.1 38.0 18.4	176 454 239	16.3 42.0 22.1	
12+	150	24.5	213	19.7	0.02
ADH1B genotypes *1/*1 *1/*2	548 62	89.4 10.1	973 107	89.9 9.9	
*2/*2	3	0.5	2	0.2	0.53 ^c

s.d. = standard deviation; ADH1B = alcohol dehydrogenase 1B. $^{a}\chi^{2}$ test used. $^{b}Among$ parous women only. ^cFisher's exact test used.

DISCUSSION

Our data raise the possibility of an effect modification of the association between alcohol consumption and breast cancer risk by ADH1B genotype in our study population, which was apparent only at the highest consumption category of 12 g or more alcohol per day. We showed previously that breast cancer risk increased significantly with high daily alcohol intake of ≥ 19 g in this study population (Kropp et al, 2001). Owing to the small number of ADH1B*2 carriers, we were not able to further subdivide the highest alcohol intake category of $\ge 12 \text{ g day}^{-1}$. Corresponding to a previous case-only study (Sturmer et al, 2002), a case-only analysis of our data would yield a statistically significant interaction OR (OR 0.2, 95% CI 0.1–0.6 for ≥ 12 g alcohol day⁻¹). However, since ADH1B genotype and alcohol intake are not independent in our study population, the modifying effect of the ADH1B genotype on breast cancer risk associated with alcohol consumption is overestimated in the case-only analysis, partly due to residual confounding by differences in alcohol consumption caused by the genotype (Albert et al, 2001). Indeed, indications for an association between ADH1B genotype and alcohol consumption, alcoholism or

 Table 2
 ORs for breast cancer risk associated with alcohol intake stratified by ADH1B genotype

	ADH1B*1/*1 genotype			ADH1B*2 carrier		
	Cases/controls	OR ^a	95% CI	Cases/controls	OR ^a	95% CI
Alcohol consumption (g day	on average ^{-1})					
0	101/147	I (Ref)		16/29	I (Ref)	
I – 5	200/407	0.74	0.54-1.01	33/47	0.81	0.34-1.91
6-11	104/226	0.70	0.49-0.99	9/13	0.87	0.25-3.02
12+	143/193	1.12	0.78-1.59	7/20	0.29	0.08-1.00
P-value for trend ^b		0.03			0.04	

ADH1B = alcohol dehydrogenase 1B; OR = odds ratio; CI = confidence interval. ^aORs stratified for age by 5-year intervals; additional adjustment was made for first-degree family history (yes/no), parity (0, 1–2, 3+ children), menopausal status (pre-, postmenopausal, unknown), education (low, intermediate and high) and smoking status (never, former and current active smoking) as categorical variables and breastfeeding (total duration in months) as continuous variable; ^bTest for linear trend (mean of categories used).

adverse reactions such as flushing were observed in previous studies (Whitfield *et al*, 1998; Borras *et al*, 2000; Loew *et al*, 2003; Neumark *et al*, 2004).

There is still controversy in the literature regarding the effect of ADH1B genotype on alcohol pharmacokinetics *in vivo*. Most studies failed to detect differences in blood alcohol or acetaldehyde levels by ADH1B genotype (Yamamoto *et al*, 1993; Mizoi *et al*, 1994; Whitfield *et al*, 2001); only one recent study reported a significantly higher alcohol elimination rate in carriers of the ADH1B*2 allele (Neumark *et al*, 2004).

The reduction in breast cancer risk associated with high consumption levels in carriers of the ADH1B*2 allele in our study could therefore be explained by a higher alcohol elimination rate in these subjects. Owing to the low allele frequency of the ADH1B*2 allele in Caucasians, our study had limited power to detect a gene-environment interaction for high alcohol consump-

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tion levels and our findings about *ADH1B* genotype as an effect modifier of breast cancer risk associated with high alcohol consumption need confirmation in larger studies.

ACKNOWLEDGEMENTS

We wish to thank the many gynaecologists and oncologists of the 38 clinics of the study region; U Eilber for competent data management; M Schlaier and S Hees for technical assistance on genotyping. The epidemiologic study was supported by the Deutsche Krebshilfe and genotyping was funded by the Medical Faculty of the University of Ulm (P.685). C Lilla is recipient of a scholarship from the Deutsche Forschungsgemeinschaft (Graduiertenkolleg 793).

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