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Coffee consumption and hip fracture risk: a meta-analysis

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Abstract

To investigate the effect of coffee consumption on hip fracture risk, a meta-analysis was conducted. The PubMed database was screened for all published studies about coffee consumption and hip fracture through to November 2011. Reviews, PubMed option 'related articles' and references of retrieved papers were also searched for potentially relevant papers. Only studies that contained OR with 95 % CI for the association between coffee consumption and hip fracture risk were included. The summary risk estimates were calculated by fixed- and random-effects models. Subgroup analyses were carried out stratified by study designs and participant characteristics, respectively. A total of six prospective cohort studies and six case–control studies were included in the final analysis. The pooled OR displayed increased risk of hip fracture by 29·7 % (95 % CI 0·960, 1·751; P = 0.09) for the highest compared with the lowest coffee consumption significantly increased hip fracture risk by 54·7 % (95 % CI 1·152, 2·077; P = 0.004) among women, by 40·1 % (95 % CI 1·015, 1·935; P = 0.040) for elderly participants aged over 70 years, and by 68·3 % for Northern Americans (95 % CI 1·492, 1·899; P = 0.000). Other subgroup analyses according to data published before the year 2000 showed a positive association between coffee and hip fracture risk, and follow-up duration also positively affected hip fracture risk, especially when the follow-up length was less than 13 years. Although our meta-analysis has provided insufficient evidence that coffee consumption significantly increases hip fracture risk, coffee intake may increase hip fracture risk among women, elderly participants and Northern Americans. No dose–response pattern was observed.

Key words: Coffee consumption: Coffee intake: Hip fracture: Prospective cohort studies: Case-control studies: Meta-analyses

Accompanied by a rapid increase in incidence in recent decades, hip fracture, the most serious complication of osteoporosis, has become a major health problem in many countries⁽¹⁾, particularly among women. Individuals with hip fracture will become a huge burden to society and individuals^(2,3), and will be expected to experience poor life quality. Furthermore, it is estimated that there could be 4.5 million hip fractures by 2050⁽⁴⁾, so it is imperative to identify risk factors of hip fracture followed by developing preventive strategies.

Previous research has demonstrated that decreased bone mineral density in the proximal femur, low BMI⁽⁵⁾, low mobility⁽⁶⁾, impaired balance⁽⁷⁾ and low physical activity⁽⁸⁾ could

potentially influence fracture risk. Meanwhile, lifestyle-related extra skeletal risk factors, such as dietary factors, also have a tight relationship with hip fracture. As a common beverage, coffee consumption is high at present, and previous studies have suggested that coffee could increase excretion of urinary and faecal Ca⁽⁵⁾, decrease the efficiency of intestinal Ca absorption⁽⁹⁾ and result in Ca loss of about 4–6 mg per cup of coffee. It has been hypothesised that coffee consumption could increase individual susceptibility to fracture by decreasing bone mineral density and bone mechanical strength, but results from earlier research have been conflicting. A total of eight case–control studies^(5,10–16) and twelve prospective cohort studies^(6,17–27) have explored associations between coffee

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consumption and hip fracture risk, among which eight studies showed positive association^(6,11,18,21–24,27) and one study in Sweden found that coffee consumption could increase hip fracture risk for non-drinkers⁽¹⁹⁾. Another study in Norway showed that women with an intake of at least nine cups of coffee daily tended to have an increased risk of hip fracture, but only 6.8 % of women consumed so much coffee per d⁽²⁰⁾. However, results from three studies, conducted in the Netherlands⁽¹⁷⁾, Finland⁽²⁵⁾ and southern California⁽²⁶⁾, respectively, showed that coffee consumption had no association with hip fracture.

In view of the inconsistency of the published literatures, a meta-analysis was conducted to investigate the following: (1) to establish the current evidence concerning the relationship between coffee consumption and hip fracture risk; (2) to analyse the relationship between coffee consumption and the risk of hip fracture according to study design and characteristics of study populations; and (3) to determine the potential dose–response pattern between coffee consumption and hip fracture risk.

Methods

Literature search

The PubMed database (http://www.ncbi.nlm.nih.gov/entrez/ query.fcgi) was screened through to November 2011 by using the following key words of 'coffee consumption' or 'coffee intake' and 'hip fracture', with restriction for English-language articles only. Reference lists of retrieved articles, reviews and the PubMed option 'related articles' were also searched for potentially relevant papers. Only published studies with full text were included, while abstracts or unpublished studies were excluded. We did not contact authors to obtain additional data.

Inclusion criteria

Studies matching the following criteria were included in our final meta-analysis: (1) prospective cohort studies or casecontrol studies; (2) research factors were coffee consumption and hip fracture; (3) endpoint for prospective cohort studies was the incidence of hip fracture; (4) association of coffee consumption and hip fracture risk was evaluated, OR or relative risk, and the corresponding 95 % CI for the highest compared with the lowest category of coffee consumption were available.

Data extraction

Data were collected in a standardised data-collection form according to reference⁽²⁸⁾, such as: (1) key components of study designs, including first author's surname, publication year, study location, number and age of cases and controls for case–control studies, or number and age of cases and total participants, mean duration of follow-up for prospective cohort studies; (2) evaluation of exposure and outcome, such as dose of coffee consumption, OR or relative risk from the most fully adjusted model for the highest v. the lowest coffee consumption and their corresponding 95 % CI; and (3)



confounders adjusted for in multivariate analysis. Coffee consumption was expressed uniformly as cups/d.

Quality assessment of studies

As suggested by Stroup *et al.*⁽²⁸⁾, key components of study designs, but not the aggregate scores, were employed to assess the quality of individual studies in our meta-analysis, which included characteristics of study populations, exposure and outcome, duration of follow-up for prospective studies, and potential confounding factors. Literature search, study selection and data extraction were conducted by X.-L. L. and J.-H. X. independently. Discrepancies were discussed.

Statistical methods

OR was used to assess the association between coffee consumption and hip fracture risk, and relative risk was directly considered as OR when the incidence of hip fracture was low.

The heterogeneity assumption across studies was tested by a χ^2 -based Q test. A *P* value of greater than 0.10 indicated a lack of heterogeneity. The I^2 statistic, which describes the proportion of total variation in study estimates caused by heterogeneity⁽²⁹⁾, was also calculated. If the heterogeneity test was statistically significant, the random-effects model (DerSimonian and Laird method) was employed. Otherwise, the pooled estimation of the OR was calculated by the fixed-effects model (Mantel–Haenszel method)⁽³⁰⁾.

Effects of various study characteristics, including study design, case sex, case age, geographic regions, publication year of the paper, and length of follow-up for prospective cohort study were investigated in subgroup analyses. The cutpoint of follow-up duration was determined according to the median length of follow-up. As most of the datasets (eleven of fourteen) used in the present analysis were published before the year 2000, and the amount of caffeine in commercial coffee products, lifestyle, and other risk factors of hip fracture may be different from the past decades, the data were reanalysed on the basis of publication year 2000.

A sensitivity analysis was performed using both fixed- and random-effects models to evaluate the influence of single study on the overall risk, by omitting one study each time.

Potential publication bias was assessed by Begg's funnel plots and Egger's regression test $(P < 0.10)^{(31)}$, an asymmetric plot suggests possible publication bias.

All analyses were performed by using STATA version 11.0 (StataCorp LP). A *P* value < 0.05 was considered significant, except as where otherwise specified.

Results

Eligible studies

Our initial search by using 'coffee consumption' OR 'coffee intake' AND 'hip fracture' obtained twenty-one eligible studies in PubMed, and six studies were added after reviewing the reference lists of retrieved articles, reviews and the PubMed option 'related articles'. The full text of the total articles were reviewed, and fifteen studies were excluded for the following reasons: two studies were reviews^(32,33), eight studies could not provide the OR and their $CI^{(6,15,16,23-27)}$, and five studies had irrelevant endpoints and exposures^(34–38). The remaining twelve studies included six case–control studies^(5,10–14) and six prospective studies^(17–22), among which two studies^(20,22) reported their results separately by sex, and thus each study was regarded as two separate studies, so fourteen datasets were extracted. The selection progress of papers is displayed in Fig. 1.

Study characteristics

The characteristics of the included studies are displayed in Tables 1 and 2. All the studies were published between 1990 and 2011, and seven studies were conducted in Europe, four in North America and one in Asia. They were all population-based except the Nurses' Health Study⁽²¹⁾. The median length of follow-up for prospective studies was 13.3 years. Coffee consumption was assessed by using a semiquantitative dietary questionnaire. Case ascertainments were confirmed by using hospital discharge registers or individual medical records or telephone for eleven studies; only one was self-report fractures⁽²¹⁾. All the OR in each study were estimated based on the highest compared with the lowest coffee consumption. Most individual studies were adjusted for a wide range of potential confounders, including age, BMI, physical activity, smoking, alcohol use, energy intake, etc. except for the study of Johnell et al.⁽⁵⁾, which was only adjusted for age, research centre and BMI.



Meta-analysis results

The multivariable-adjusted OR for each study and the pooled OR for the highest compared with the lowest level of coffee consumption are presented in Fig. 2. Of the fourteen groups of datasets, eleven showed positive association and three were negatively associated. Of the eleven studies showing positive association, four studies reached statistical significance^(11,18,21,22). Of the three studies showing negative association, only one reached statistical significance⁽¹⁹⁾. Obvious heterogeneity was found across studies (P = 0.000; $I^2 = 84.0$ %), and therefore the random-effects model was employed; the pooled OR of hip fracture risk for the highest compared with the lowest coffee consumption was 1.297 (95 % CI 0.960, 1.751; P = 0.09).

Subgroup and sensitivity analyses

The results of subgroup analyses according to study design, case sex, case age, geographic region, and length of follow-up for prospective studies and publication year of the paper are presented in Table 3. Hip fracture risk increased significantly if the coffee drinkers were female, elderly subjects or North American, while study design had no effect on the pooled OR. Subgroup analysis according to data published before the year 2000 demonstrated that coffee drinking significantly increased hip fracture risk by 41.9 % (OR 1.419; 95 % CI 1.074, 1.874), while results from data published after the year 2000 showed a decreased risk of hip fracture by 9.7 % (OR 0.903; 95 % CI 0.473, 1.866). Results regarding the effect of follow-up length on hip fracture risk showed that when the

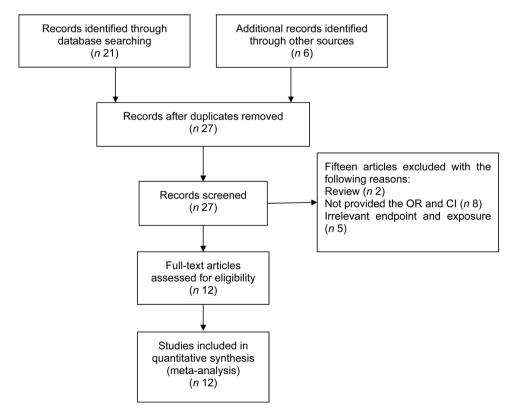


Table 1. Chara	Table 1. Characteristics of case-control studies of coffee consumption and hip fracture risk	rol studies of coffe	e consumption a	nd hip fracture	risk			
Study	Region	Mean age or age range (years)	Cases/ controls (<i>n</i>)	Sex	Exposure range (cups/d)	OR	95 % CI	Adjustment
Johnell <i>et al.</i> (1995) ⁽⁵⁾	Southern Europe	78.1	2086/3532	Women	≥3 v. never	1.01	1.01 0.82, 1.23	Age, centre, BMI
Kanis <i>et al.</i> (1999) ⁽¹⁰⁾	Southern Europe	73.9	730/1132	Men	≥3 v. never	0.78	0.55, 1.11	Age, BMI, centre, recreational physical activity, milk intake, consumption of tea, alcohol, cheese, smoking, sunlight exposure
Suzuki <i>et al.</i> (1997) ⁽¹¹⁾	Japan	78.6	249/498	Men and women	≥3 <i>v.</i> never	3.59	3.59 1.46, 8.85	Main work activity, BMI, sleep disturbance during the past 6 months, past history of stroke with hemiplegia and diabetes, consumption of coffee, alcohol, milk, fish, sun exposure, physical exercise, immobilised or bedridden condition, type of bedding, drinking Japanese green tea, and activity of daily living for bathing
Tavani <i>et al.</i> (1995) ⁽¹²⁾	Northern Italy	63	279/1061	Women	≥5 v. never	1.20	0.5, 2.7	Age, education, BMI, smoking status, total alcohol consumption. Ca intake, menopausal status and oestrogen replacement therapy use
Kreiger <i>et al.</i> (1992) ⁽¹³⁾	Toronto	73.6	102/277	Women	≥3 v. < 3	1.18	0.64, 2.18	Age, Quetelet index, ovariectomy, oestrogen replacement, cigarette smoking
Nieves <i>et al.</i> (1992) ⁽¹⁴⁾	New York and Philadelphia	>75	161/168	Women	≥3 <i>v.</i> never	1.87	0.74, 4.69	Hospital, age and BMI, smoking, education, and presence of chronic disease



length of follow-up was more than 13 years, hip fracture risk significantly decreased by 41.8 % (OR 0.582; 95 % CI 0.453, 0.747), while when the follow-up was no more than 13 years, the risk significantly increased by 71.2 % (OR 1.712; 95 % CI 1.517, 1.931).

The sensitivity analysis conducted by omitting one study each time and calculating the combined OR for the remaining studies was perfornmed. All other trials had no substantial effect on hip fracture risk (from 1.153 (95 % CI 0.853, 1.559) to 1.343 (95 % CI 0.944, 1.911)) except the study by Trimpou *et al.*⁽¹⁹⁾, which resulted in significantly increased hip fracture risk by 38.9 % (OR 1.389; 95 % CI 1.068, 1.806).

Dose-response analysis

Although several included studies^(5,10–12,14,17,19–22) had stratified their results by dose of coffee consumption, only four studies^(19–22) (five datasets) could be selected to conduct final dose–response analysis, while others were excluded for not providing participant number of each category of coffee consumption. The dose–response analysis showed a slight, but not significant, reduction in hip fracture risk by 0·3 % for every cup/d increment in coffee intake (OR 0·997; 95 % CI 0·964, 1·031; *P* for trend = 0·872).

Publication bias

Publication bias was tested by Begg's funnel plot and Egger's test. Result of Begg's funnel plot did not show significant asymmetry (Fig. 3), and Egger's test also had little evidence of publication bias (P = 0.891).

Discussion

Our present meta-analysis firstly evaluates the effect of coffee intake on hip fracture risk. Meta-analysis results showed a slightly, although not significantly, increased risk of hip fracture by 29·7 % (OR 1·297; 95 % CI 0·960, 1·751; P = 0.09) in response to coffee intake. Obvious heterogeneity was displayed. The dose–response relationship, only four studies included^(19–22), showed a slight reduction of hip fracture risk by 0·3 % (OR 0·997; 95 % CI 0·964, 1·031; P for trend = 0·872) in response to every cup/d increment of coffee consumption, but the results had no statistical significance.

The heterogeneity might be explained by the following. First, studies included in our meta-analysis were conducted in different geographic regions, where populations shared different genetic background, lifestyle and hip fracture incidence. Second, the sex and age of the participants varied narrowly.

Results from subgroup analyses demonstrated that female coffee drinker, elderly participants and North Americans were at higher risk of hip fracture, which is consistent with previous studies^(18,21,22,39). Elderly women have different hormone levels, which could affect the metabolism of Ca. The region of America had the highest incidence rate of hip fracture⁽³⁹⁾, which could be explained by the higher caffeine content of coffee than that of other regions, such as Italy⁽¹²⁾.

4



			Mean age							
Study	Region	Period	or age range (years)	Duration (years)	Cases/ participants (<i>n</i>)	Sex	Exposure range (cups/ d)	Relative risk	95 % CI	Adjustments
Van Lenthe <i>et al.</i> (2011) ⁽¹⁷⁾	Netherlands	1991–2003	>55	13	192/16 578	Men and	≥3 <i>v.</i> never	0.75	0.42, 1.36	Father's occupation, adult occupation, education, income proxy
(2010) ⁽¹⁸⁾ (2010) ⁽¹⁸⁾	Finland	1997–2007	72	10	21/1222	Women	>5 v.<5	2.58	1.01, 6.56	Biourb proxy Bioline physical activity, Ca intake, smoking, alcohol use, CVD
Trimpou <i>et al.</i> (2010) ⁽¹⁹⁾	Sweden	1974–2003	46–56	30	451/7495	Men	>5 v. never	0.55	0.42, 0.73	Age, physical activity, smoking, stature, occupation, alcohol abuse, intercurrent stroke, dementia, serum cholesterol concentrations, BMI, diabetes melittus, psvchological stress, svstolic blood pressure
Meyer <i>et al.</i> (1997) ⁽²⁰⁾ ∗	Norway	1977–1991	57.3	11.4	157/19 752	Women	≥9 v.<2	1.94	0.96, 3.91	Age, body height, BMI, self-reported physical activity at work and during leisure time, diabetes mellitus, disability pension. marital status, smoking
Meyer <i>et al.</i> (1997) ⁽²⁰⁾ †	Norway	1977–1991	55.7	11.4	55/19 938	Men	≥9 v.<2	1.04	0.37, 2.94	Age, body height, BMI, self-reaction of the provided activity at work and during leisure time, diabetes mellitus, disability pension. marifal status, smoking
Hernandez-Avila <i>et al.</i> (1991) ⁽²¹⁾	USA	1980–1986	34–59	9	59/84 484	Women	≥4 v. never	3.35	1.32, 8.49	Age, Quetelet index, menopause status, oestrogen-replacement therapy, Ca intake, alcohol intake
Kiel <i>et al.</i> (1990) ⁽²²⁾ *	Framingham 1971–1982	1971–1982	77.4	12	113/3170	Women	>2 v. 0–2	1.69	1.49, 1.92	Sex, age, Framingham examination number, metropolitan relative weight, postmenopausal oestrogen use, smoking, alcohol consumption
Kiel <i>et al.</i> (1990) ⁽²²⁾ †	Framingham 1971–1982	1971–1982	75.7	12	22/3170	Men	>2 v. 0–2	1.42	0.59, 3.38	Sex, age, Framingham examination number, metropolitan relative weight, postmenopausal oestrogen use, smoking, alcohol consumption
* Momen										

Table 2. Characteristics of prospective cohort studies of coffee consumption and hip fracture risk

* Women. † Men.

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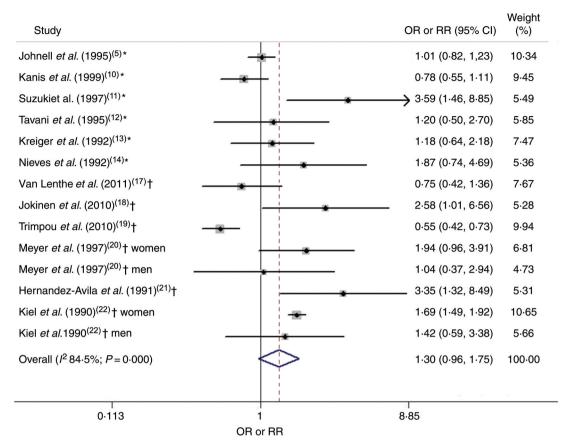


Fig. 2. Meta-analysis of studies that investigated association of coffee consumption with hip fracture risk (random-effects model). * OR. † Relative risk (RR).

Further subgroup analysis showed that follow-up duration was negatively associated with hip fracture risk. Among those studies where the duration of follow-up was less than 13 years, coffee consumption was commonly more than five cups/d, while other studies where the length of follow-up was more than 13 years, the median coffee consumption was less than two cups/d. Therefore, it might be the relatively higher coffee consumption not the shorter length of follow-up that induced the above results. Subgroup analysis stratified by the publication year of papers demonstrated that coffee intake increased hip fracture risk according to the data published before the year 2000, while decreased hip fracture risk using

Table 3.	Association	of coffee	e intake	and hip	fracture	stratified b	y certain	variables
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Variable	Total studies included (n)	Summary relative risk	95 % CI	P for heterogeneity	l ² (%)
All	14	1.297	0.960, 1.751	0	84.0
Study design					
Prospective cohort studies	8	1.357	0.822, 2.240	0	88.9
Case-control studies	6	1.181	0.889, 1.569	0.078	49.5
Sex					
Women	8	1.547	1.152, 2.077	0.001	70.5
Men	4	0.752	0.523, 1.080	0.102	51.7
Men and women	2	1.578	0.341, 7.309	0.004	87.7
Age					
>70 years	8	1.401	1.015, 1.935	0	80.3
50–70 years	6	1.145	0.645, 2.032	0	78.6
Geographic region					
Europe	8	0.970	0.710, 1.325	0.001	70.5
North America	5	1.683	1.492, 1.899	0.464	0
Asia	1	3.590	1.458, 8.839		
Length of follow-up					
>13 years	2	0.582	0.453, 0.747	0.349	0
≤13 years	6	1.712	1.517, 1.931	0.555	0
Publication year of the paper					
Before 2000	11	1.419	1.074, 1.874	0.001	79.9
After 2000	3	0.903	0.473, 1.866	0	74.1

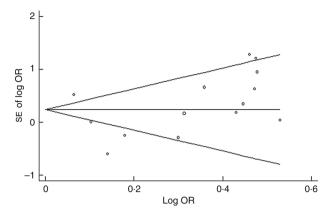


Fig. 3. Funnel plot of coffee consumption and risk of hip fracture.

data published after year 2000; the change in caffeine content caused by a change in method of coffee preparation or cup size might account for the above results⁽⁴⁰⁾.

In the sensitivity analysis, the study by Trimpou *et al.*⁽¹⁹⁾ had a significant effect on the pooled OR. Trimpou's study⁽¹⁹⁾ is the only one included in our final meta-analysis that showed a significant negative association between coffee intake and hip fracture, with the combined risk of 0.55 (95 % CI 0.42, 0.73). Its participants were all male, and were classified into two categories: non-drinkers and coffee drinkers. The author further testified that those non-drinkers, with a significantly increased risk (P = 0.0002) of hip fractures, often suffered from digestive tract problems; thus, it might be the digestive tract problems that increased hip fracture risk, not the coffee that protected the drinkers from hip fracture risk. Meanwhile, the participants in Trimpou's study⁽¹⁹⁾ were relatively young, which might be one reason for the increased pooled OR.

In addition to the variables included in the present study, there are other factors that should be considered. Method of coffee preparation, type of coffee, such as black coffee, coffee with milk, or coffee with cream determine the caffeine content. The menopausal status of women and differences in genetic background, such as genetic polymorphism in the NAT2 gene⁽⁴¹⁾, could affect the metabolism of caffeine and caffeine exposure of an individual. Different subtypes of hip fractures had different responses to coffee, and only trochanteric fracture was significantly associated with heavy coffee consumption⁽⁴²⁾. Dietary factors, such as intake of Ca and vitamin D, could affect Ca balance. Protein from non-dairy animal sources is likely to affect Ca excretion⁽⁴³⁾. Although several studies^(12,13,20) had demonstrated that low Ca intake and short sunlight exposure^(5,10,11) could increase hip fracture risk, only limited studies had stratified their results according to the level of Ca intake⁽¹²⁾ or menopausal status^(12,13). At present, as the selected papers in the present review could not provide more available evidence about the above factors, we failed to further analyse the effect of those factors on the relationship between coffee and hip fracture, and research concerning those factors is needed in the future.

The present review has several limitations that should be considered. First, obvious heterogeneity was observed. Therefore, the random-effects model was employed. Second, exposure misclassification should be considered. Coffee exposure is expressed by cups/d, and assessed using FFQ. This might introduce some degree of error in the exposure measurement. For example, the cup size and the caffeine content of the same commercial coffee products may differ materially between different studies. Nevertheless, the total cups of coffee intake per d at least gave us a hint about the relationship between coffee intake and hip fracture risk. Third, although most included studies adjusted for a number of potential confounders for hip fracture, it may be possible that other unmeasured or inadequately measured factors have confounded the true association of coffee intake and hip fracture. Fourth, all the studies except $one^{(11)}$ in our analysis were conducted in Western countries; additional research in other populations is warranted to generalise the findings.

However, there are several advantages of the present study. First, we had included twelve studies that were mainly population-based except for the study of Tavani *et al.*⁽¹²⁾, which was hospital-based; this could greatly reduce the likelihood of selection biases. Second, most of the trials (eleven of twelve) were conducted in Europe and America, where individuals shared a similar genetic sensitivity to caffeine, and the effect of genetic variation on the caffeine exposure of an individual was small. Third, the present study had collected a large number of cases, so the present results had a highly sufficient statistical power. The mean follow-up of 13.3 years of the prospective cohort studies was sufficiently long to detect an adequate number of hip fractures.

In Western countries, coffee consumption per capita is high, especially in Nordic countries. Meanwhile, hip fracture is considered a major health problem in many countries. Due to the fact that a substantial proportion of the population are daily consumers of coffee (and thus caffeine), and the effect of coffee intake on fracture risk has been conflicting, it is imperative to classify the exact effect of coffee consumption on hip fracture risk. Although our present meta-analysis showed that the role of coffee consumption in increasing hip fracture risk had no statistical significance, coffee intake may increase hip fracture risk for female, elderly coffee drinkers and Northern Americans. Further studies that consider interactions between coffee consumption and dietary factors, lifestyle factors and genetic polymorphisms are warranted to confirm our findings in the future.

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X.-L. L. and J.-H. X. were both involved in literature search, study selection and data extraction. J.-H. X. performed the data analyses; and X.-L. L. was involved in writing the manuscript. Both authors participated in the revision of the manuscript.

We declare no conflict of interest.

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