

Research Paper

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A Prognostic Model Based on Circulating Tumour Cells is Useful for Identifying the Poorest Survival Outcome in Patients with Metastatic Colorectal Cancer

Wen-Chi Chou^{1,2*}, Min-Hsien Wu^{1,3,4*}, Pei-Hung Chang^{5,6*}, Hung-Chih Hsu¹, Gwo-jyh Chang², Wen-Kuan Huang^{1,7}, Chiao-En Wu^{1,8}, Jason Chia-Hsun Hsieh^{1,4}

- 1. Circulating Tumor Cell Lab, Division of Medical Oncology, Department of Internal Medicine, Chang Gung Memorial Hospital at Linkou, College of Medicine, Chang Gung University, Taoyuan, Taiwan;
- 2. Graduate Institute of Clinical Medical Sciences, College of Medicine, Chang Gung University, Taiwan;
- 3. Graduate Institute of Biochemical and Biomedical Engineering, Chang Gung University, Taoyuan City, Taiwan;
- 4. Department of Chemical and Materials Engineering, Chang Gung University, Taoyuan, Taiwan;
- 5. Division of Hematology-Oncology, Department of Internal Medicine, Chang Gung Memorial Hospital, Keelung and Chang Gung University, College of Medicine, Taiwan;
- 6. Cancer Center, Chang Gung Memorial Hospital, Keelung, Taiwan.
- Department of Oncology-Pathology, Karolinska Institutet, Stockholm, Sweden; Cancer Center Karolinska, Karolinska University Hospital, Stockholm SE-17176, Sweden.
- 8. Northern Institute for Cancer Research, School of Medicine, Newcastle University, Newcastle, United Kingdom.

*Drs. Chou, Wu and Chang contributed equally to this work.

🖂 Corresponding author: Dr. Jason Chia-Hsun Hsieh, MD. PhD. E-mail: wisdom2000@gmail.com; Tel.: +886-3-281200 ext. 2517

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Abstract

Background: There is an urgency to develop robust prognostic biomarkers for metastatic colorectal cancer (mCRC) patients receiving chemotherapy. The current study aimed to examine the prognostic significance of circulating tumour cells (CTCs) and to develop a prognostic model incorporating CTCs in predicting the outcomes of mCRC patients treated with chemotherapy.

Methods: Our study prospectively enrolled 55 mCRC patients who had undergone palliative chemotherapy between 2011 and 2014. Baseline CTCs and clinicopathological variables predictive of survival outcome were identified using univariate analysis. Negative selection-based protocol plus flow cytometry was used for CTC identification.

Results: The median overall survival (OS) and progression-free survival (PFS) were 24.2 months and 8.7 months, respectively. CTCs were detected in all the patients, and the median number of CTCs was 30.8/mL (range: 5.8-431.3/mL). The median OS and PFS were 37.1 and 13.3 months, respectively, for patients with CTC number ≤ 30 /mL, while the median OS and PFS were 14.9 months and 5.1 months, respectively, for patients with CTC number ≥ 30 /mL (both P<0.001). A prognostic model using CTCs in conjunction with other independent clinical variables further stratified patients into good and poor prognostic groups. The median OS and PFS were 32.4 and 11.5 months, respectively, in the good prognostic group and 5.4 and 2.7 months, respectively, in the poor prognostic group.

Conclusions: We developed a reliable CTC-based prognostic model for the prediction of clinical outcomes in mCRC patients treated with chemotherapy. This model may be used to assist clinicians in identifying those with the poorest prognosis before treatment.

Key words: metastatic colorectal cancer, circulating tumour cell, palliative chemotherapy, prognosis, model

Introduction

Colorectal cancer (CRC) is one of the most common and lethal cancers, accounting for 10% of

new cancer cases and 8.5% of total cancer-related deaths worldwide in 2012.[1] Systemic chemotherapy

remains the most effective therapy for patients with unresected metastatic colorectal cancer (mCRC), and the median survival duration is now approaching 30 months for patients treated with cytotoxic chemotherapy and biologic agents targeting angiogenesis and the epidermal growth factor receptor.[2, 3] Although the treatment response of these combinations has been fairly improved in the modern era, some patients who did not respond to these systemic treatments still have a dismal outcome. Indeed, approximately 28-44% of patients have no response to the first-line treatment with doublet chemotherapy plus either cetuximab or bevacizuamb.[2] Therefore, there is an urgency to develop robust prognostic biomarkers for mCRC patients receiving chemotherapy. Unfortunately, there has not yet been an accepted biologic or molecular marker of prognostic value for systemic chemotherapy.

In recent years, circulating tumour cells (CTCs) has been widely proposed to serve as biomarkers in various cancer types, including breast,[4, 5] prostate[6] and colorectal[7, 8] cancers. CTCs are cells migrating from solid tumours into the peripheral blood, which represent a disseminated disease and poor prognosis.[9-11] The CellSearch platform (Veridex, Raritan, NJ) is the only system approved by the US Food and Drug Administration to detect CTCs since 2004.[4] In 2008, Cohen et al first reported using the number of CTCs as a prognostic and predictive value for survival outcome and treatment response in mCRC patients treated with systemic chemotherapy. [7] However, the application of CTCs to predict the clinical outcomes of mCRC remains uncertain due to the relatively low efficiency of detection rate by the CellSearch platform[11] and no other device has been proven to be a standard method to date. In addition to CTCs, a number of clinical variables with predictive values have been identified in patients with mCRC receiving chemotherapy, including patient performance, [12] anatomic extent of the tumour, [13] histologic subtype, [14, 15] tumour grade, [16] serum tumour marker carcinoembryonic antigen (CEA),[17] and history of colectomy.[18] We hypothesized that a reliable prognostic tool combing clinicopathological features and CTCs for predicting the survival outcome and tumour response for mCRC patients would be most useful and possibly better to identify patients with the best and poorest prognoses rather than CTCs alone. To achieve this, the current study aimed to examine the prognostic significance of CTCs by using an easy-to-perform CTC detection platform using negative selection strategy and to develop a prognostic model that incorporates CTCs with other important clinical variables in predicting the survival

outcome and tumour response in mCRC patients treated with systemic chemotherapy, especially to identify those patients with the poorest survival outcome to systemic chemotherapy.

Methods

Patient selection and treatment of systemic chemotherapy

This study was designed as a prospective observational study at a single institute. In total, fifty-five patients who had a histological diagnosis of CRC and were receiving systemic chemotherapy treatment at Chang Gung Memorial Hospital (CGMH) Linkou branch between January 2011 and June 2014 were prospectively enrolled. The eligibility criteria included patients over 20 years of age, patients with unresected metastatic disease, and patients with a measurable lesion confirmed by image studies. Patients who received concurrent radiotherapy or palliative tumour resection within 28 days of palliative chemotherapy and those who received local ablation therapy during the enrolment period were excluded. All patients underwent a baseline evaluation that included demographic data, clinical pathological characteristics, history, computed tomography (CT) scan, and biochemical evaluation. Systemic antitumour therapy consisted of a doublet combination of cytotoxic chemotherapy with or without bevacizumab or cetuximab (limited in patients with KRAS wild type), depending on the physician's decision. Tumour response was evaluated by CT scan according to the response evaluation criteria in solid tumours (RECIST) 1.1.[19] Image studies were performed and interpreted by an institutional radiologist at baseline and were repeated every 8-12 weeks (the exact interval was at the physician's discretion) to evaluate tumour response. The overall survival (OS) and progression-free survival (PFS) lengths were calculated from the date of administration of the first cycle of palliative chemotherapy to the date of death or disease progression after chemotherapy. All patients were followed-up until death or June 30, 2016. The study protocol was approved by the Institutional Review Board of CGMH (IRB number 101-2161C and 102-2407B). Written informed consent to participate was obtained from all the patients.

Measurement of peripheral circulating tumour cells

CTC analysis was performed using a protocol of combined negative selection and positive detection strategies, which was designed and validated in our previous reports.[20, 21] Briefly, the methods were (1) a negative selection protocol for effective RBC and leukocyte depletion by red blood cell lysis and a CD45 depletion kit and (2) flow cytometry to quantitatively identify CTCs and calculate their numbers. The CTC identification was validated with a recovery rate of $44.6 \pm 9.1\%$ and a coefficient of variation of 20.4%.[20]

For CTC tests, 4 mL of peripheral blood was used after discarding the first 4 mL of blood to avoid epithelial contamination. Red blood cell lysis was performed within 72 hours after blood drawing. Samples were then negatively enriched by adding EasySep CD45 Depletion Cocktail (STEMCELL Technologies Inc., Vancouver, BC, Canada) at 25 µL/mL to the cells and EasySep Magnetic Nanoparticles (STEMCELL) at 50 μ L/mL to the cells. Immunomagnetically enriched samples containing spiked HCT116 cells were collected and labelled with Alexa Fluor® 488-conjugated anti-EpCAM monoantibody (1:400 dilution; Cell Signalling Technology, Inc., Danvers, MA, USA.) and a novel DNA-detecting far-red-fluorescing dye, Draq5 (50µM, Abcam, Cambridge, MA, USA), for nuclear staining. The isotype for EpCAM was also used for internal controls. In addition, we routinely used 4 mL of peripheral blood drawn from healthy individuals that was spiked with and without 1000 HCT116 cells (a colon cancer cell line, which was obtained from the Food Industry Research and Development Institute, Taiwan) for controls during the trial. The performance recovery, which was defined as the number of HCT116 cells detected by flow cytometry (BD FACSCalibur, BD Biosciences, San Jose, CA, USA) divided by the number of spiked HCT116 cells, and the coefficient of variation (CV) value have been calculated and reported to be stable in a previous report.[20] CTCs were defined as the cells that were positive both for EpCAM and Draq5. Serum CEA levels of enrolled patients were measured by standard kit purchased from Siemens ADVIA Centaur Chemiluminescent Immunoassay, Siemens Siemens Healthcare GmbH, Germany.

Statistical analysis

demographic The patients' data were summarized as the number (%) for categorical variables, and the median, 95% confidence interval (CI), and range for continuous variables. Eleven predefined variables, included age, gender, primary tumor site, histological grade, American Joint Committee on Cancer (AJCC) staging group, Eastern Cooperative Oncology Group (ECOG) performance scale, KRAS mutation status, CEA level, CTC numbers, previous history of colectomy, and chemotherapy regimens, recorded within 7 days before the first cycle of palliative chemotherapy in mCRC patients, were

then evaluated to ascertain their impact on OS and PFS. These potentially key prognostic variables were selected because minimal data were missing, and they could widely represent the available clinical data, which makes the findings broadly applicable. Variables with a P value < 0.10 in the univariate analysis were included for analysis in the multivariate model. Multivariate Cox's proportional hazard model analysis was performed to investigate the impact of independent factors on OS and PFS. A risk model was developed from a multivariate logistic regression. The β-coefficients from the risk model were used to generate the points of the prognostic score for calculating survival time. Akaike's information criterion (AIC) was used to develop different multivariate models by systematically removing predictors from a full model that contained all predictive factors that were found to be statistically significant in the multivariate analysis. Receiver operating characteristic (ROC) curves and the area under the curve (c-statistic) for the outcome of survival at 1, 2, and 3 years were calculated to determine the accuracy of the prognostic score. Patients were further stratified into two prognostic groups according to their total score obtained from the prognostic score. Fisher's exact test was performed to determine significant associations between the tumour responses of the chemotherapy and prognostic groups. OS and PFS among different prognostic categories were calculated according to the Kaplan-Meier method. Log-rank tests were used to determine significant differences among the survival curves. SPSS 17.0 software (SPSS Inc., Chicago, IL) was used for all the statistical analyses. All statistical assessments were 2-sided. A P value of less than 0.05 was considered significant.

Results

Table 1 shows the demographic data of the 55 patients. The median age was 60 years (range: 36-88 years), and 61.8% of the patients were male. Forty-two patients (76.4%) had a tumour that originated from the colon, and 13 patients (23.6%) had tumours that originated from the rectum. Moderately differentiated adenocarcinoma was the most common histological differentiation subtype (38 patients, 69.1%), followed by poorly (14 patients, 25.5%) and well-differentiated (3 patients, 5.5%) subtypes. Seventeen patients (30.9%) had a mutated KRAS gene, and 23 patients (41.8%) had previously undergone a colectomy for primary tumour resection. All the patients received 5-fluouracial based doublet chemotherapy. CTCs were detected in all the patients. The median number of CTCs was 30.8/mL (range: 5.8-431.3/mL). There was no significant difference of CTC numbers among

patients with different tumor stage and T- or Nclassification (**Figure 1**). Eighteen patients (32.7%) had a normal baseline carcinoembryonic antigen (CEA) level (≤5 ng/dL), and the median serum level of CEA was 10.5 ng/dL (range: 0.7 -2558 ng/dL) among all the patients.

Table 1. Patients' demographic data

Variable	Category	no	%
Age, year	Median (range)	60 (36-88)	
Gender	male	34	61.8
	female	21	38.2
Primary tumour site	colon	42	76.4
5	rectum	13	23.6
Histological grade	well	3	5.5
0 0	moderately	38	69.1
	poorly	14	25.5
AJCC Stage, 7th edition	4a	21	34.5
	4b	34	65.5
ECOG performance scale	0	11	20.0
-	1	25	45.5
	2~3	19	34.5
KRAS mutation status	no mutation	38	69.1
	mutation	17	30.9
CEA, ng/dL	median (range)	10.5 (0.7-2558)	
	≤5	18	32.6
	>5	37	67.4
CTC, number/mL	median (range)	36.8 (5.8-431.3)	
Previous history of colectomy	no	32	58.2
	yes	23	41.8
Line of chemotherapy	1	53	96.4
	2	1	1.8
	3	1	1.8
Chemotherapy regimen	Fluorouracil	55	100
	Irinotecan	39	70.9
	Oxaliplatin	16	29.1
	Bevacizumab	34	61.8
	Cetuximab	7	12.7

AJCC: American Joint Committee on Cancer; ECOG: Eastern Cooperative Oncology Group

CEA: carcinoembryonic antigen; CTC: circulating tumour cell

At the study end, 42 patients (76.4%) died during the study. The median OS and PFS were 24.2 months (95% CI: 14.0-34.4 months) and 8.7 months (95% CI: 6.1-11.3 months), respectively. The 1-, 2-, and 3-year survival rates were 74.5%, 50.2%, and 31.9%, respectively. The results of the univariate analyses for OS and PFS that were associated with the clinical variables are presented in Table 2. Based on the univariate analysis of the imputed data set, 6 of the 11 preselected variables, including primary tumour side, tumour stage, histological grade, previous history of colectomy, ECOG performance scale, and CTCs, showed a statistically significant effect on OS. However, the multivariate analysis identified tumour stage, histological grade, previous history of colectomy, ECOG performance score, and number of

CTCs as the only independent prognostic factors **(Figure 2)**. Regarding the PFS of all the patients, 3 of the 11 preselected variables, including previous history of colectomy, ECOG performance scale and number of CTCs, were significant variables in both the univariate analysis and multivariate analysis. Using the median number of CTCs as the cut-off point, the median OS and PFS were 37.1 months (95% CI: 22.6–51.6 months) and 13.3 months (95% CI: 11.7–14.9 months), respectively, for patients with CTC number \leq 30/mL, while the median OS and PFS were 14.9 months (95% CI: 8.6–21.2 months) and 5.1 months (95% CI: 3.5–6.7 months), respectively, for patients with CTC number \geq 30/mL (both P<0.001, **Figure 3**).

The risk model and scoring system of the prognostic score generated from β-coefficients of multivariate analysis for OS are shown in Table 3. The total prognostic scores ranged from 0 to 6. Using the prognostic score, patients were stratified into good (total score of 0-3) and poor (total score of 4-6) prognostic groups. This prognostic scoring system assigned 70.9% of the patients to the good prognostic group and 29.1% of the patients to the poor prognostic group. The median OS in the good and poor prognostic risk groups were 32.4 months (95% CI: 23.0-41.8 months) and 5.4 months (95% CI: 2.1-8.7 months), respectively (Figure 4A). The hazard ratio was 10.3 (95% CI: 3.4-23.5; P < 0.001) when comparing the poor and good prognostic groups. Accordingly, the median PFS in the good and poor prognostic risk groups was 11.5 months (range: 8.4-14.6 months) and 2.7 months (range: 1.3-4.1 months), respectively, with a hazard ratio of 5.15 (95% CI: 2.68-9.57; P < 0.001) (Figure 4B). The c-statistic that was obtained using ROC curve analysis for mortality within 1, 2, and 3 years was 0.82 (95% CI: 0.71-0.94), 0.83 (95% CI: 0.72-0.93), and 0.72 (95% CI: 0.58-0.86), respectively. The AIC was 230.6 for the full model and 257.3 after removal of the variable of CTCs from the full model. A smaller AIC value indicated a higher capacity for predictive stratification in the prognostic model. **Table 4** compared the performance of CTC numbers, clinical parameter scoring system except CTC and combined CTC numbers and clinical parameter scoring system (full score) for OS and PFS. The results showed the full score and CTC number performed well both for PFS and OS while clinical parameters did not. Therefore, our findings suggest CTC alone could have good performance for outcome prediction, but better performance was found when combined with clinical parameters.

Table 2. Univariate analyses for overall survival (OS) and progression-free survival (PFS)

			Univariate for OS				Univariate for PFS			
Variable	Category	No (%)	No of event (%)	median	95% CI	р	No of event (%)	median	95% CI	р
Overall		55 (100)	42 (76.4)	24.2	14.0-34.4		54 (98.2)	8.7	6.1-11.3	
Age	≤60	28 (50.9)	19 (67.9)	26.2	17.8-34.6	0.13	27 (96.4)	9.1	6.4-11.8	0.34
	> 60	27 (49.1)	23 (85.2)	19.8	4.4-35.2		27 (100)	7.5	5.5-9.5	
Gender	male	34 (61.8))	25 (73.5)	25.1	16.9-33.3	0.6	33 (97.1)	8.7	5.7-11.7	0.46
	female	21 (38.2)	17 (81.0)	19.8	8.6-30.9		21 (100)	7.5	4.7-10.3	
Primary site	colon	42 (76.4)	29 (69.0)	26.2	19.8-32.6	0.047	41 (97.6)	8.2	5.6-10.8	0.49
	rectum	13 (33.6)	13 (100)	15.7	10.1-21.3		13 (100)	9.7	2.3-17.1	
AJCC stage, 7th	IVa	21 (38.2)	14 (66.7)	34.6	22.6-46.4	0.026	20 (95.2)	12.9	6.8-19.0	0.17
	IVb	34 (61.8)	28 (82.4)	15.7	11.4-20.0		34 (100)	7.8	6.8-8.8	
Grade	well or moderately	41 (74.5)	31 (75.6)	26.3	20.3-32.3	0.008	40 (97.6)	9.1	6.5-11.7	0.69
	poorly	14 (25.5)	11 (78.6)	12.5	4.2-20.8		14 (100)	7.5	6.6-8.4	
Previous history of	yes	23 (41.8)	12 (52.2)	43.9	12.9-74.9	< 0.001	22 (95.7)	12.8	7.9–17.7	0.022
colectomy	no	32 (58.2)	30 (93.8)	15.7	9.6-21.8		32 (100)	7.2	4.9-9.5	
KRAS mutation	mutant	17 (30.9)	14 (82.4)	17.4	7.0-27.8	0.27	17 (100)	7.2	1.7-12.7	0.63
	no mutant	38 (69.1)	28 (73.7)	27.7	18.9-36.6		37 (97.4)	9.7	6.8-12.6	
Chemotherapy regimen	Oxaliplatin-based	17 (30.9)	13 (76.5)	15.7	8.6-22.8	0.48	17 (100)	5.4	0.0-9.8	0.14
	Irinotecan-based	38 (69.1)	29 (76.3)	26.3	17.6-34.9		37 (97.4)	10.6	8.3-12.9	
ECOG performance	0~1	36 (65.5)	24 (66.7)	31.8	21.0-42.6	< 0.001	35 (97.2)	11.5	8.0-15.0	< 0.001
	2~4	19 (45.5)	18 (94.7)	10.8	1.7-19.9		19 (100)	3.8	2.1-5.5	
CEA level, ng/dL	≤5	18 (32.7)	13 (72.2)	25.1	14.0-36.2	0.92	17 (94.4)	7.4	2.0-12.8	0.57
	>5	37 (67.3)	29 (78.4)	21.6	9.2-34.0		37 (100)	10	6.7-13.3	
CTC, numbers/	≤30	26 (47.3)	17 (65.4)	37.1	22.5-51.6	0.005	25 (96.2))	13.3	11.7-14.9	< 0.001
mL	>30	29 (52.7)	25 (86.2)	14.9	8.9-21.2		29 (100)	5.1	3.5-6.7	

AJCC: American Joint Committee on Cancer; ECOG: Eastern Cooperative Oncology Group

CEA: carcinoembryonic antigen; CTC: circulating tumour cell

Table	3.	Point	scoring	of	overall	survival	for	the	prognostic
models									

Variable	Point		
	0	1	2
AJCC 7th stage	IVa	IVb	
CTC number, per mL	≤30	>30	
ECOG performance	0-1		2-3
Histological grade	well or moderately	poorly	
Previous history of colectomy	yes	no	

AJCC: American Joint Committee on Cancer; CTC: circulating tumour cell; ECOG: Eastern Cooperative Oncology Group

Overall, chemotherapy resulted in tumours showing a partial response in 26 (47.3%) patients, stable disease in 13 (23.6%) patients, and progressive disease in the remaining 16 (29.1%) patients. The treatment response to a partial response, stable disease, and progressive disease of tumours were 25 (64.1%) patients, 11 (28.2%) patients, and 3 (7.7%) patients, respectively, in patients from the good risk group, while the treatment responses to a partial response, stable disease, and progressive disease of tumours were 1 (6.3%) patient, 2 (12.5%) patients, and 16 (81.3%) patients, respectively, in the poor risk group (Fisher's exact test P<0.001, **Figure 5**). **Figure 6** presents the correlations between Kras mutation status and CTC numbers (6a) as well as targeted agents and CTC numbers (6b). There was no significant difference of CTC numbers between patients with Kras mutant and wild type (Mann-Whitney U test p = 0.879) as well as CTC numbers between patients received bevacizumab-based for cetuximab-based chemotherapies (Mann-Whitney U test p = 0.933).

Table 4. The performance of CTC numbers, clinical parameterscoring system except CTC and combined CTC numbers andclinical parameter scoring system (full score) for OS and PFS

	Hazard Ratio (95% CI)	<i>p</i> value
For OS		
CTC alone (> 30 vs ≤ 30 cells/mL)	2.613(1.385-4.928)	0.003
Clinical parameter scoring except CTC	3.266(1.715-6.219)	< 0.001
Full Score	6.135(1.867-20.161)	0.003
For PFS		
CTC alone (> 30 vs ≤ 30 cells/mL)	4.942(2.600-9.392)	< 0.001
Clinical parameter scoring except CTC	1.830(1.041-3.217)	0.036
Full Score	3.741(1.631-8.581)	0.002







Figure 2. Multivariate analysis for overall survival and progression-free survival.



Figure 3. Kaplan-Meier overall survival (A) and progression-free survival (B) curves for patients, stratified according to the number of circulating tumour cells.



Figure 4. Kaplan-Meier overall survival (A) and progression-free survival (B) curves for patients, stratified according to the prognostic model.



Figure 5. Tumour response rate to chemotherapy, stratified according to the prognostic model



Discussion

mCRC is a common and lethal malignancy both in Taiwan and worldwide.^{1,} [22] Though the survival outcome and treatment response to chemotherapy plus a biologic agent had dramatic advances in the modern era, a small subset of patients still had a dismal outcome, which implies a stubbornness to antitumour treatment.[2, 3] Therefore, a prognostic model for picking up on those with the poorest outcomes in antitumour therapy may assist clinicians and patients in decision making. Our study developed a prognostic model that predicts survival outcome in patients with mCRCs, using CTCs in conjunction with other significant clinical variables. This model accurately predicted survival outcome with c-indices of 0.824, 0.828, and 0.719 for the probability of 1-, 2-, and 3-year mortalities. Our study showed that this model might be informative for clinicians and patients with mCRCs for estimating the survival outcome after systemic chemotherapy.

Tumour stage¹³, histological grade[16], previous colectomy[18], and ECOG scale[12] are the well-established prognostic factors in patients with mCRC. Advanced stage and poor histological grade represented aggressive tumour behaviour. A population-based study including 37,793 mCRC patients recently reported that palliative primary tumour resection was associated with improved overall and cancer-specific survival.[18] A lower ECOG scale represented a better ability of daily activity and general health status[23, 24], which were associated with the eligibility to receive lower chemotherapy and а incidence of treatment-related toxicity. These are all possible reasons for explaining why the ECOG scale contributed the most in our model. Our study identified that the baseline number of CTCs prior to chemotherapy was an independent prognostic factor for overall survival and a predictive factor for PFS and tumour response in mCRC patients. Furthermore, conjunction of the CTCs with other prognostic clinical variables increased the discrimination power of the model (AIC value decreased from 257.3 to 230.6 after the CTC variable was added into the full model). Our result encourages integrating CTCs and clinical variables for predicting the outcome of mCRC patients receiving systemic chemotherapy. Patients with a good prognosis may be encouraged to receive palliative chemotherapy to gain the maximal benefit of the treatment. In contrast, palliative chemotherapy may be futile in patients with a poor prognosis, given that the median OS, PFS, and tumour response of these patients were only 5.4 months, 2.7 months and 6.3%, respectively.

The number of CTCs is an emerging biomarker that has been shown to be an independent prognostic factor in numerous studies of various solid cancers and CRC.[4-7] In general, higher CTC counts have been consistently associated with disseminated stage and aggressive tumour biology.[10, 11, 25] Cohen et al first reported that patients with CTC number ≥3 per 7.5 mL of peripheral blood had a shorter median survival based on a prospective multicentre study of 430 mCRC patients.[7] In their report, the distribution of the number of CTCs per 7.5 mL of peripheral blood using the CellSearch system as undetectable, 1 to 2, and 3 or more were 48%, 26%, and 26%, respectively. Furthermore, two large-scale meta-analysis studies, which included 3094 patients in the first study and 1329 patients in the second study, showed that the detection of CTCs correlated with a poor prognosis for CRC patients.[26, 27] In addition, the optimal cut-off value of the number of CTCs for the prognostic value in patients positive for CTCs was undetermined [28]. Furthermore, only approximately one-third of the patients had detectable CTCs, and the discrimination of survival differences among those patients who were negative for CTCs was impossible. Therefore, it is imperative that new techniques be developed for CTC enrichment and identification in order for a wide utilization of CTCs in mCRC. This study used a negative selection method of CD45 depletion using magnetic beads followed by flow cytometry for CTC identification.[20, 21] CTCs were detected in all patients with mCRC, and the cut-off value of the median number of CTCs was analysed as an independent prognostic factor. Based on this simple and cheap method, we believed that CTCs may be widely available to use as a prognostic and predictive factor in mCRC patients who underwent palliative chemotherapy.

To the best of our knowledge, this is the largest prospective study using a negative selection method for CTC isolation to evaluate the prognostic role of CTCs in patients with mCRC. Based on this simple method, the clinical variables of the prognostic model accessible and available before palliative are chemotherapy. Therefore, the current model is widely applicable and clinically relevant to identifying a small subset of patients with the poorest outcome. However, there are several inherent limitations of this study. First, this study included small numbers of patients, selected from a single medical centre; therefore, the statistically significant impact of clinical variables on survival outcome might be limited due to the small sample size. Second, the chemotherapy regimen was not consistent in all patients, so analysis of the baseline CTCs relevant to PFS and tumour response might be confounded by different treatment

regimens. Finally, and most importantly, even though our prognostic model represented a good accuracy to predict 1-, 2-, and 3-year survival outcomes by the c-index, an external validation of the model is essential before it can be widely used. A prospective study is required to evaluate the prognostic and predictive values of this model in mCRC patients who underwent palliative chemotherapy.

In conclusion, we identified the number of CTCs at the baseline of palliative chemotherapy as an independent prognostic factor for predicting survival outcome, in terms of overall survival, progression-free survival, and tumour response, in patients with mCRC. In addition, we developed a reliable CTC-based prognostic model for the prediction of survival outcome and treatment response in mCRC patients, after systemic chemotherapy. This model may be used to assist clinicians in identifying those mCRC patients with the poorest prognosis and treatment outcome. Further large-scale prospective trials to confirm the role of CTC prognosis score guiding anticancer strategies for each mCRC patient were eagerly warranted.

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Competing Interests

The authors have declared that no competing interest exists.

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