RESEARCH ARTICLE

A Meta-Analysis of Zilpaterol and Ractopamine Effects on Feedlot Performance, Carcass Traits and Shear Strength of Meat in Cattle

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Abstract

This study is a meta-analysis of the effects of the beta-agonists zilpaterol hydrochloride (ZH) and ractopamine hydrochloride (RAC) on feedlot performance, carcase characteristics of cattle and Warner Bratzler shear force (WBSF) of muscles. It was conducted to evaluate the effect of the use of these agents on beef production and meat quality and to provide data that would be useful in considerations on the effect of these agents on meat guality in Meat Standards Australia evaluations. We conducted a comprehensive literature search and study assessment using PubMed, Google Scholar, ScienceDirect, Scirus, and CAB and identification of other studies from reference lists in papers and searches. Searches were based on the key words: zilpaterol, zilmax, ractopamine, optaflexx, cattle and beef. Studies from theses obtained were included. Data were extracted from more than 50 comparisons for both agents and analysed using meta-analysis and metaregression. Both agents markedly increased weight gain, hot carcase weight and longissimus muscle area and increased the efficiency of gain:feed. These effects were particularly large for ZH, however, fat thickness was decreased by ZH, but not RAC. Zilpaterol also markedly increased WBSF by 1.2 standard deviations and more than 0.8 kg, while RAC increased WBSF by 0.43 standard deviations and 0.2 kg. There is evidence in the ZH studies, in particular, of profound re-partitioning of nutrients from fat to protein depots. This work has provided critically needed information on the effects of ZH and RAC on production, efficiency and meat quality.



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Data Availability: The authors confirm that all data underlying the findings are fully available without restriction. All relevant data are within the paper.

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while working with the Cooperative Research Centre for Beef Genetic Technologies several projects were funded by Pfizer Animal Health and Pfizer Animal Genetics. He is currently working on a project, in part funded by Elanco Animal Health. Frank R Dunshea has consulted for with Elanco Animal Health as a member of the Swine Nutrition Advisory Panel and has conducted research with ractopamine in swine funded through the Cooperative Research Centre for Internationally Competitive Pork of which Elanco Animal Health is a participating organisation. All three authors are members of the Meat Standards Australia Pathways committee. Meat and Livestock Australia played no role in the study design apart from facilitating and funding the study. It facilitated the study by funding an independent search that identified sufficient studies were present to conduct a meta-analysis. This literature search was made available to the authors but did not provide papers that were not identified by the subsequent systematic literature searches undertaken. The authors have no other competing interests in research funding, employment, patents, products in development or marketing products.

Introduction

Dietary additives, including β -adrenergic agonists (β -AA) are used by the beef industry to enhance the efficiency of gain and modify carcass characteristics and meat quality [<u>1</u>, <u>2</u>]. Both zilpaterol (ZH) and ractopamine (RAC) bind to β adrenergic receptors located in the cellular membranes and indirectly lead to decreased lipogenesis (fat synthesis and storage) and increased lipolysis (fat mobilization and hydrolysis) [<u>3</u>, <u>4</u>, <u>5</u>, <u>6</u>]. The magnitude of these changes is influenced by dose and duration of treatment with the β -AA, the type of β -AA, and the species of animal treated [<u>7</u>, <u>5</u>, <u>8</u>]. The β -AA also directly influence protein metabolism. Previous studies have shown that RAC improved growth performance [<u>1</u>, <u>2</u>], with the effects mediated through increased protein synthesis [<u>8</u>] and decreased protein degradation [<u>5</u>, <u>8</u>, <u>6</u>].

A review of the meat science literature [9] concluded that if improvements in animal growth were due in part to decreased protein degradation, this would reduce the rate of post-mortem proteolysis in muscle and result in tougher meat. Given that part of the β -AA effect is likely to be associated with protein degradation rates it's use may produce tougher meat. However the results on meat quality from individual studies appear equivocal with a range of individual studies varying in the magnitude and statistical significance of the responses [6].

The objectives of this study were to evaluate the effects of β -AA on feedlot performance and carcass traits of cattle, using meta-analytic methods. Metaanalysis integrates the results from many studies to provide a more robust estimate of the effects of β -AA (both ZH and RAC) on live and carcass traits. We also examined the heterogeneity or variability of responses in order to better target future research projects and resolve other hypotheses raised about the action of the products and gaps in current knowledge. An understanding of factors that give rise to heterogeneity between studies can also lead to more effective treatments or modifications of management.

Materials and Methods

Literature search

Our literature searches used PubMed, Google Scholar, ScienceDirect, Scirus, and CAB and identification of other studies from references lists in papers. Searches were based on the following key words: zilpaterol, zilmax, ractopamine, optaflexx, cattle and beef using the terms for the pharmaceutical agent or product brand names separately with the terms cattle and beef. Where possible, studies from theses obtained were also included. A list of published studies on ZH in beef cattle was provided by Meat and Livestock Australia (MLA). The papers were useful in the initial phase of the project, but did not yield additional studies after the study search was completed. In many cases, authors of articles were contacted to provide clarity in regard to results and for additional information.

Inclusion and exclusion criteria

Studies were included or excluded in this study based on a series of criteria developed by SBScibus and which are discussed in a review of the use of metaanalysis in animal and veterinary science [12]. Briefly, database searches were augmented with hand searches of library resources for relevant papers, books, abstracts, and conference proceedings. We pursued references in papers and contacted authors active in the field. Quality assessment criteria included randomization of study groups, appropriate analysis of data and comparability of treatment groups at entry to each trial. The extracted data were audited by five reviewers.

References [19–102] were evaluated for inclusion or exclusion in the study and details of these studies are provided in File S1. Trials were included in the analysis if they met the following criteria: full manuscripts from peer-reviewed journals, published after 1980, that evaluated use of ZH or RAC supplementation in cattle; had a description of randomization processes; reported the dose of β -AA used, animals studied were cattle; the paper contained sufficient data to determine the effect size for production outcomes (e.g., the number of cattle or pens, carcases or cuts of beef in each treatment and control group); a measure of effect amendable to effect size analysis for continuous data (e.g., standardized mean difference, **SMD**); a measure of variance (SE or SD) for each effect estimate or treatment and control comparisons. Studies were rejected that could not be adequately interpreted in regard to origin of muscles, that used non-representative sampling methods, or whose authors did not respond to clarify their approach.

Design criteria included the number of animals/pen, animals/group and pens/ group; experimental and analytical unit (animal, pen, carcass, muscle).Experimental details included days on feed, the number of days products were fed, dosage of products fed (per kg DMI or per kg/hd/d), diet and delivery methods of product and withdrawal period (days) of the test products prior to slaughter. Animal details included class of cattle (steers, heifers, bulls or cull cows); vaccination history, use of anthelmintics; other supplements and growth promoting products (hormonal implants); housing and feeding systems.

Output variables extracted for meta-analysis included final body weight (BW, kg), dry matter intake (DMI, kg/hd/day), average daily gain (ADG, kg/hd/day), gross feed efficiency (G:F ratio); hot carcass weight (HCW, kg), dressing percentage (D%, (kg HCW/kg BW)*100), ultimate muscle pH, longissimus muscle area (LMA, cm²),USDA muscle colour score (scale 1 to 9), USDA marbling score; 12th rib fat thickness (mm) and Warner-Bratzler shear force (WBSF, kg). The WBSF was predominantly assessed on striploin samples cooked to internal temperature of 70 to 71°C on a belt grill.

In many papers muscle, bone and overall maturity were reported as a proportion within the A maturity category, whilst in others it was reported as a continuous variable. As it was not possible to reconcile these units, maturity scores were not included in the current analyses. The ADG, DMI and G:F ratio were reported over the period of supplementation, generally 20–30 days, not over the full feeding period. There was also a lack of detail on sampling methods for

sensory samples for a number of ZH and RAC studies and it was considered there were insufficient studies to include this trait in the analyses.

Statistical analysis

We used Stata (Intercooled Stata v.13, USA) to analyze production and carcass data by standardized mean difference (SMD) which is also called effect size (ES) analysis, in which the difference between treatment and control groups means was standardized using the standard deviations of control and treatment groups. The SMD estimates were made for the fixed effect [10], and random effects [11] models that are reported. If the paper reported separate estimates of measure of variance (SE or SD) for each group, these were recorded as such. Many studies reported a common SE or SD and these estimates were used for both control and treatment groups. We also have provided a random effects weighted mean difference (WMD) between treated and control, with the weighting reflecting the inverse of the variance of the studies included according to non-standard method (Stata 13.0 Statacorp, Tx, USA).

Random effects models were conducted for each production outcome to estimate the effect size, 95% confidence intervals, and statistical significance of SMD. We recognize that there is a clustering effect that results from multiple comparisons to a single control group, but have determined that the variance inflation effect will be minor unless there are very large numbers of repeated comparisons. The statistical methods of meta-analytic procedures that were used in this paper have been previously published [12].

Efforts were made to clearly identify the units of interest used in the studies and to clarify the measures of dispersion reported in papers. If there was some uncertainty, authors were contacted to provide clarification of these measures and a number responded. If there was a lack of clarity in regards to the unit of measure, a more conservative measure was used. Specifically, if muscle characteristics were measured and evaluated as the unit of analysis, but the muscles were obtained from pen fed studies, pen was used in our analyses.

Forest plots

The effects of β -AA on production performance of beef cattle are displayed in the forest plots, using the estimated SMD of β -AA using random effects [11] models. Points to the left of the vertical line represent a reduction in the outcome, whereas points to the right of the line indicate an increase in the outcome variable. Each square represents the mean effect size for that study. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval (CI) for the effect size. Forest plots were only presented for selected output variables.

The weighting of a study is estimated by the inverse of the variance of the effect size. Box sizes are proportional to the inverse variance of the estimates. The size of the square box reflects the relative weighting of the study to the overall effect size estimate with larger squares representing greater weight. Boxes draw attention to the studies with the greatest weight. The grey vertical line represents the mean difference of zero or no effect.

Assessment of heterogeneity

Variations among the trial level SMD were assessed using a chi-squared (Q) test of heterogeneity. Heterogeneity in studies reflects underlying differences in clinical diversity of the herds and β -AA used, differences in study design and analytical methods and statistical variation around responses. Identifying the presence and sources of the heterogeneity improves understanding of the responses to β -AA. We used an α level of 0.10 because of the relatively poor power of the χ^2 test to detect heterogeneity among small numbers of trials [13]. Heterogeneity of results among the trials was quantified using the I^2 statistic [14], which quantifies the impact of heterogeneity on a meta-analysis, from mathematical criteria, that are independent of the number of studies and the treatment effect metric. I^2 is a transformation of the square root of the χ^2 heterogeneity statistic divided by its degrees of freedom and describes the proportion of total variation in study estimates that is due to heterogeneity. Negative values of I^2 were assigned a value of zero, consequently the value I^2 lies between 0 and 100%. An I^2 value greater than 50% indicates substantial heterogeneity.

Meta-regression

Meta-regression analyses were used to explore the source of heterogeneity of response, using the individual SMD for each trial as the outcome and the associated standard error as the measure of variance. Meta-regression can be used to explore sources of heterogeneity even if an initial overall test for heterogeneity is non-significant [14]. This also allows us to quantify the magnitude as a function of the *a priori* defined covariate changing and exploring reasons for heterogeneity (i.e. possible/probable study-level predictors). Meta-regression is also a technique that can formally test whether there is evidence of different effects in different subgroups of trials [15]. Inclusion of multiple covariates in the meta-regressions [15], and the use of the smoothed within-trial variance estimates were used to improve hypothesis testing with regard to the significance levels. The permutation test approach for assessing the statistical significance of meta-regression methods [16] used the programming methods [17, 18] to reduce the risk of type I error.

Meta-regression analysis was conducted by first screening individual variables using a P-value of ≤ 0.10 . All variables with P-value of ≤ 0.10 were entered into a forward stepwise weighted meta-regression, until all remaining variables were significant at P< 0.05. Factors identified a priori as factors that might influence responses to treatment were examined, including the unit of study (pen or animal or cut of meat), initial value of the study variable in the control group in each study (eg DMI of control), duration of feeding in the study (days on feed) and period or duration of treatment with each β -AA (eg RAC was fed for 30 days). Data were screened for plausible quadratic relationships for these variables by visual appraisal of univariable scatter plots between the covariate and SMD of each study.

Publication bias

We investigated the presence of publication bias using funnel plots which are a simple scatter plot of the intervention effect estimates from individual studies plotted against study precision. The name 'funnel plot' arises because precision of the estimated intervention effect increases as the size and precision of a study increases. Effect estimates from small studies will scatter more widely at the bottom of the graph and the spread narrows for larger studies. In the absence of bias the plot should approximately resemble a symmetrical (inverted) funnel. If there is bias, for example because smaller studies without statistically significant effects remain unpublished, this will lead to an asymmetrical appearance of the funnel plot and a gap will be evident in a bottom corner of the graph. In this situation, the effect calculated in a meta-analysis will tend to overestimate the intervention effect. The more pronounced the asymmetry, the more likely it is that the bias will be substantial. Individual studies with large or aberrant results were identified from the forest and funnel plots and evaluated for factors that may have influenced the results. This lead to the exclusion of some studies (see Tables 1 and 2).

Results and Discussion

Studies identified and information extracted

A comprehensive literature search using 5 search engines identified relevant peerreviewed papers, theses and proceedings that were published on ZH and RAC supplementation between 2000 and 2013. For both interventions, details were recorded, but are not provided, on sex of cattle, breed, treatments including vaccinations, parasite control, hormonal implants, feeding regimes, country in which the study was conducted and distance cattle were transported to slaughter.

A total of more than 50 studies on ZH, published between 2000 and 2013, were identified. Of these, 31 papers (with 83 sub-trials or comparisons on different outcomes) met the selection criteria. The required data and information from those studies that met the inclusion criteria were extracted and were tabulated and presented in <u>S1</u> to <u>S5 Tables</u> in <u>S1 File</u>. The data included information on animal performance, carcass characteristics and meat quality. <u>S6 Table</u> in <u>S1 File</u> provides details of studies that were excluded for various reasons.

Similar search, extraction and reporting methods were used for RAC and a total of 31 papers were identified with 68 sub-trials or comparisons on different outcomes. These trials are presented in <u>S7</u> to <u>S10 Tables</u> in <u>S1 File</u>. Contact with workers in the field provided valuable detail on studies and allowed re-analysis of some data. Again, three reviewers assessed, extracted and validated the data and 8 papers were excluded (<u>S11 Table</u> in <u>S1 File</u>). There were insufficient studies reporting some variables, especially those reporting sensory traits of meat, for meta-analytical evaluation.

The methods of meta-analysis used in this study have been described and published previously $[\underline{12}, \underline{103}, \underline{104}]$. A critical consideration in this study was to ensure that appropriate estimates of evaluative units (eg pen, animal or carcase)

Table 1. Effects of zilpaterol on feedlot performance and carcase characteristics.

Outcome	Number of studies (n)	Raw mean difference 95% Cl	Effect Size and 95% Cl	Weighted mean difference and 95% Cl	I ² estimated hetero-geneity	Significant meta- regression effects
Final Body weight (kg)	31	6.562	0.449	8.150	0.0	
		4.229 to 8.895	0.277 to 0.621	5.627 to 10.674		
Dry Matter Intake (kg/d)	26	-0.160	-0.470	-0.118	0.0	
		-0.243 to -0.077	−0.676 to −0.264	-0.167 to -0.070		
Average Daily Gain (kg/d)	29	0.210	0.884	0.153	29.1	
					(0.0) ¹	
		0.141 to 0.278	0.656 to 1.113	0.111 to 0.194		
Gain:feed (kg/kg ratio)	28	0.024	1.336	0.024	54.1	
		0.017 to 0.031	1.002 to 1.671	0.018 to 0.030		
Hot Carcase Weight (kg)	35	15.383	1.323	15.179	55.9	
		12.937 to 17.829	1.034 to 1.611	13.615 to 16.743		
Ultimate pH	5	0.022	0.102	-0.002	0.0	NA
		-0.033 to 0.077	−0.411 to 0.616	-0.017 to 0.013		
Longissimus muscle area (cm ²)	35	8.147	2.302	8.006	70.8	Unit (0.022) period (0.165) Multivariate ²
		7.115 to 9.180	1.898 to 2.705	7.052 to 8.959		
Objective measurement of 'redness' (Colour)	18	0.012	0.098	0.03	15.0	
		-0.122 to 0.098	−0.174 to 0.371	-0.004 to 0.066		
Fat thickness at the 12th rib (cm)	35	-0.106	-0.697	-0.11	47.8	Fat of control(0.120) and period (0.186) Multivariate ²
		-0.139 to -0.073	−0.940 to −0.453	-0.158 to -0.080		
Standardised USDA mar- bling score	34 (28)*	-23.004	-0.861	-22.947	37.3	Unit (0.030) period (0.050) Multivariate ²
		−31.245 to −14.763	-1.100 to -0.621	−30.330 to −15.564		
Dressing Percentage (%)	27	1.657	2.205	1.706	71.5	
		1.319 to 1.996	1.669 to 2.741	1.510 to 1.902		
Warner-Bratzler Shear Force (kg)	47	1.022	1.212	0.840	61	Age (0.001) unit (0.001) Multivariate ²
		0.854 to 1.191	1.024 to 1 401	0.720 to 0.960		

CI – Confidence interval. ¹Sensitivity with removal of Avendano-Reyes et al.

NA – insufficient studies to attempt. ²Higgins and Thompson (2004) method.

*6 trials removed with different units for Raw mean and weighted mean differences.

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Table 2. Effects of ractopamine on feedlot performance and carcase characteristics.

Outcome	Number of studies (n)	Raw mean difference 95% Cl	Effect Size and 95% Cl	Weighted mean difference and 95% Cl	I ² estimated hetero-geneity%	Significant meta- regression effects <0.05 (P)
Final Body weight (kg)	44	6.476	0.397	7.568	50	Body weight of control (0.01)
		3.200 to 9.752	0.238 to 0.557	5.584 to 9.553		
Dry Matter Intake (kg/d)	48	0.027	0.020	-0.003	26.0	
		-0.116 to 0.170	-0.122 to 0.161	-0.089 to 0.082		
Average Daily Gain (kg/ d)	49	0.244	0.76	0.193	54.4	Average Daily Gain control (0.035)
		0.150 to 0.337	0.564 to 0.957	0.149 to 0.237		
Gain:feed (kg/kg ratio)	41	0.019	0.772	0.018	47.8	
		0.012 to 0.026	0.583 to 0.961	0.014 to 0.022		
Hot Carcase Weight (kg)	54	7.376	0.47	6.182	46.5	
		3.475 to 11.277	0.312 to 0.628	4.551 to 7.812		
pН	5	-0.027	-0.326	-0.006	0.0	NA
		-0.079 to 0.024	-0.873 to 0.220	-0.023 to 0.011		
Longissimus muscle area (cm ²)	60	2.43	0.391	1.844	67.1	
		1.497 to 3.363	0.198 to 0.584	1.172 to 2.517*		
Fat thickness at the 12th rib (cm)	45	0.029	0.005	-0.003	43.1	Day on feed (0.012)
		-0.026 to 0.084	-0.171 to 0.182	-0.035 to 0.028		Decreases smd
USDA marbling score	53	-2.471	-0.108	-5.144	0.8	
		-10.216 to 5.274	-0.213 to -0.002	-9.615 to -0.674		
Dressing Percentage (%)	40	0.503	0.131	0.275	66	
		0.221 to 0.784	0.000 to 0.262	0.110 to 0.440	(0) ²	
Warner-Bratzler Shear Force (kg)	17	0.305	0.429	0.203	0	
		0.188 to 0.421	0.267 to 0.592	0.122 to 0.284		

CI - Confidence interval.

Estimated mean difference.

NA - insufficient studies to attempt.

Jenning et al and Vogel 3 dropped due to implausible standard errors.

*studies with odd values removed.

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were used for the analyses and that the estimates of standard deviation were also appropriate. Our decisions were influenced by advice received on design and analysis of pen studies that the analyses using SAS would report standard errors of the difference that were appropriate to providing estimates of standard deviation [105] Tempelman, pers com 2013). Studies that did not clearly provide the unit of analysis were evaluated at the highest level of unit identified, ie pen. This may provide a conservative bias in the analysis. Similarly, only random effects models were used, as previous work concluded that when there was uncertainty in the evaluative units caused by clustering of observations the random effects model was appropriate [106].

The effect of reported and assessed evaluative unit was examined by testing of this in meta-regression. Only three outcomes, indicated that the evaluative unit significantly influenced the results of meta-regression, LMA, marbling and WBSF for ZH studies (Table 1). However, in all cases the raw mean difference, an unweighted measure, differed for the outcomes for the pen and individual animal studies, suggesting that there may be a biological reason the differences in effect. In all cases, also, the expected effect of ZH, i.e. increased LMA, lower marbling and increased WBSF was accentuated in the group fed animals. The lack of other significant findings in regard to unit of evaluation in the study overall suggest the possibility that whether cattle had access to individual feeders or were fed in pens influenced the LMA, marbling and WBSF for ZH treated cattle, but had little effect on other output variables. It is possible that such results may reflect the differences between single access to feed and competing in the group-fed situation. Competition can increase stress and agonistic behaviours [107] which may heighten responsiveness to endogenous and exogenous adrenergic stimulation. While effects of group size on feed intake are variable, feed efficiency is generally poorer and a poorer nutritional state is also associated with heightened adrenergic responsiveness [108]. We conclude that the evaluative units were probably correctly identified and, consequently, had little influence on results.

Zilpaterol

The period of feeding cattle was 143.8 ± 56.1 days and the average period that cattle were exposed to ZH was 26.6 ± 9.0 days. The results of the number of studies analysed, raw mean differences for variables, SMD, weighted mean differences, I² estimated heterogeneity and results of meta-regressions evaluating the effects of ZH are provided in <u>Table 1</u>. These results showed that supplementation of cattle with ZH during the last period of feeding resulted in a substantial impact on live and carcass traits of the animal. The weighted mean differences in BW and ADG between ZH and control groups were substantial being of the order of 0.4 and 0.9 standard deviations, respectively. This effect was equivalent to approximately 8 kg BW and 0.15 kg ADG gain more than the controls. Cattle fed ZH also had a lower DMI of the order of 0.5 standard deviations which was equivalent to approximately 0.1 kg/hd/day. Given that ZH resulted in increased BW and ADG and decreased DMI there was a substantial improvement of 1.4 standard deviations in the G:F ratio which was equivalent to 0.02 improvement in G:F.

The increase in HCW due to ZH supplementation was1.3 standard deviations above controls and equivalent to an extra 15 kg of HCW, almost twice the increase in final BW (Table 1). The disproportionate increases in BW and HCW resulted in an improvement of dressing percentage of 2.2 standard deviations, which was equivalent to an increase of 1.7% in dressing percentage.

ZH acts as a repartitioning agent where there is a substantial redirection of nutrients from fat to protein deposition [6]. The magnitude of the metabolic changes in protein deposition was reflected in the large increase in LMA of 2.3 standard deviations, which was equivalent to an increase of ca. 8 cm² in LMA. This effect was not significantly explained by the difference in HCW between treatment and controls or the HCW of the treatment group. Similarly the reduced fat deposition was reflected in a decrease of 0.9 and 0.7 standard deviations in USDA marbling score and fat depth respectively. These differences were equivalent to a decrease in marbling score of 14 units and a decrease of 1 mm in fat depth. There were few data provided in the papers which examined the physiological and health impacts of this very substantial repartitioning of protein and fat in the treated cattle. It is possible that the profound effect of these agents on repartitioning of nutrients is associated with increased death rates reported for cattle treated with ZH and RAC [110]. Loneragan et al. [109], found that the cumulative risk and incidence rate of death was 75 to 90% greater in cattle fed the βAA compared to contemporaneous controls and that during the exposure period, 40 to 50% of deaths among groups fed the β AA were attributed to administration of the drug.

Forest Plots for HCW, LMA, marbling score and WBSF variables are presented in Figs. 1, 2, 3 and 4 respectively. These Forest Plots demonstrate the consistency of the ZH responses in key commercial carcass and quality traits. Funnel plots were produced, but are not presented. These plots did not show any marked asymmetry with the possible exception of gain:feed. Given the lack of evidence of missing studies in other variables derived from many of the same studies, we conclude that there is little or no evidence of publication bias.

Table 1 also reported an estimate of the heterogeneity for the output traits. Most traits had small estimates of heterogeneity with the exception of LMA and D%. Investigation of the sources of this heterogeneity using meta-regression identified that significant heterogeneity in LMA was attributable to the period of feeding of ZH, further emphasising the causal basis of the effect. However, the unit of evaluation, in this case animal or pen indicated that the SMD of LMA was greater in the pen fed studies. As shown in Fig. 5 this difference was evident in the raw, unweighted data, showing that the effect was not mediated through a difference in weighting of studies. It is therefore likely that this result and that for marbling in which the unit of measure was also a significant covariate and that similarly had a difference between pen and animal studies in the raw means reflect differences in expression of effect of ZH between pen fed animals and individually fed animals. The period of feeding of ZH was significant in increasing LMA and approached significance in reducing marbling, again indicating the biological significance of treatment. It is possible that the sensitivity to the ZH is increased under the possibly more stressful group fed conditions in which epinephrine levels may be higher.

<u>Figs. 5</u> and <u>6</u> show the difference in SMD and raw mean difference for individually fed animals and pen fed groups for LMA and for USDA marbling score, respectively. In studies where animals were individually fed, the response in

			%
Author	Year	SMD (95% CI)	Weight
Avendano-Reyes et al, 2006	2006	6.72 (3.57, 9.88)	0.73
Neill et al, 2008	2008	0.42 (-1.58, 2.42)	1.54
Neill et al, 2008	2008	0.26 (-1.72, 2.23)	1.57
Vasconcelos et al, 2008	2008	1.91 (0.61, 3.21)	2.70
Vasconcelos et al, 2008	2008	1.81 (0.54, 3.09)	2.75
Vasconcelos et al, 2008	2008	2.01 (0.69, 3.33)	2.65
Beckett et al, 2009	2009	1.82 (0.44, 3.21)	2.51
Beckett et al, 2009	2009	2.04 (0.60, 3.48)	2.39
Beckett et al, 2009	2009	2.70 (1.06, 4.34)	2.04
Beckett et al, 2009	2009	0.43 (-0.46, 1.32)	3.80
Beckett et al, 2009	2009	0.68 (-0.23, 1.58)	3.75
Beckett et al, 2009	2009	0.52 (-0.38, 1.41)	3.79
Elam et al, 2009	2009	2.04 (0.66, 3.43)	2.51
Elam et al, 2009	2009	2.45 (0.95, 3.95)	2.29
Elam et al, 2009	2009	2.74 (1.16, 4.32)	2.14
Kellermeir et al. 2009	2009	0.92 (-0.28, 2.12)	2.93
Kellermeir et al, 2009	2009	0.46 (-0.69, 1.61)	3.06
Montgomery et al, 2009a-(1)	2009	0.78 (-0.67, 2.24)	2.37
Montgomery et al, 2009a-(2)	2009	1.20 (-0.34, 2.74)	2.20
Montgomery et al, 2009a-(3)	2009	0.39 (-1.01, 1.80)	2.47
Montgomery et al, 2009a-(4)	2009	0.46 (-0.95, 1.87)	2.46
Montgomery et al, 2009b	2009	3.04 (1.72, 4.37)	2.64
Robles-Estrada et al, 2009	2009	1.89 (0.14, 3.64)	1.87
Strydom et al. 2009	2009	0.94 (0.05, 1.83)	3.80
Baxa et al, 2010	2010	1.47 (0.17, 2.77)	2.69
Baxa et al, 2010	2010	• 1.40 (0.11, 2.68)	2.73
Holland et al, 2010	2010	1 .30 (0.76, 1.84)	4.92
Parr et al, 2010	2010	1.67 (0.43, 2.91)	2.83
Scramlin et al, 2010	2010	3.00 (1.68, 4.31)	2.66
Garmyn et al, 2011	2011	0.31 (-0.75, 1.36)	3.31
Lawrence et al, 2011	2011	• 0.38 (-0.01, 0.78)	5.35
Lawrence et al, 2011	2011	0.97 (0.23, 1.70)	4.29
Peterson, 2011	2011	1.15 (-0.38, 2.68)	2.22
Rathmann et al, 2012	2012	0.98 (0.29, 1.67)	4.42
Rodas-Gonzales et al, 2012	2012	1.76 (0.81, 2.72)	3.60
Overall (I-squared = 55.9%, p	= 0.000)	\$ 1.32 (1.03, 1.61)	100.00
NOTE: Weights are from rand	om effects analysis		
	0.99		
	-9.88	U 9.88	

Fig. 1. Forest plot of Hot Carcase Weight responses for Zilpaterol studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of zilpaterol treatmenton hot standing carcase weight. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in hot carcase weight, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was heterogenous as indicated by the l² of 55.9%.

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LMA was smaller than for pen fed groups. In contrast, the reverse was evident for USDA marbling score where individually fed animals showed little response to ZH supplementation compared with group fed animals which showed a large decrease in marbling score. Obviously behavioural changes or competition effects between animals are driving these effects. The authors are not aware of these effects being

					%
Author	Year			SMD (95% CI)	Weight
Avendano-Reyes et al, 2006	2006		•	2.53 (0.94, 4.11)	2.76
Neill et al, 2008	2008			-0.06 (-2.02, 1.91)	2.25
Neill et al, 2008	2008			1.74 (-0.86, 4.33)	1.61
Vasconcelos et al, 2008	2008			3.87 (2.00, 5.74)	2.37
Vasconcelos et al, 2008	2008		•	4.29 (2.28, 6.31)	2.19
Vasconcelos et al, 2008	2008			3.91 (2.03, 5.80)	2.35
Beckett et al, 2009	2009		<u> </u>	1.60 (0.27, 2.93)	3.16
Beckett et al, 2009	2009	<u> </u>		2.79 (1.13, 4.46)	2.64
Beckett et al, 2009	2009	I —		2.67 (1.04, 4.30)	2.70
Beckett et al, 2009	2009	-	•	2.75 (1.49, 4.00)	3.29
Beckett et al, 2009	2009			3.23 (1.86, 4.60)	3.09
Beckett et al, 2009	2009		• • • • • • • • • • • • • • • • • • •	3.16 (1.81, 4.52)	3.12
Elam et al, 2009	2009			4.59 (2.38, 6.80)	1.97
Elam et al. 2009	2009			5.85 (3.17, 8.53)	1.54
Elam et al. 2009	2009			5.96 (3.24, 8.68)	1.51
Hilton et al. 2009	2009		<u> </u>	1.73 (0.69, 2.78)	3.64
Kellermeir et al. 2009	2009		<u>!</u>	1.71 (0.35, 3.06)	3.12
Kellermeir et al. 2009	2009			2.32 (0.80, 3.84)	2.86
Montgomery et al. 2009a-(1)	2009	I ——		2.32 (0.41, 4.23)	2.31
Montgomery et al. 2009a-(2)	2009	I —		2.53 (0.54, 4.52)	2.21
Montgomery et al. 2009a-(3)	2009		<u> </u>	1.71 (0.02, 3.39)	2.61
Aontgomery et al. 2009a-(4)	2009		•	2.03 (0.23, 3.83)	2.46
Aontgomery et al. 2009b	2009			3.50 (2.06, 4.94)	2.98
Robles-Estrada et al. 2009	2009			1.61 (-0.05, 3.27)	2.65
Strydom et al. 2009	2009			0.95 (0.06, 1.84)	3.90
Baxa et al. 2010	2010		•	3.51 (1.60, 5.42)	2.32
Baxa et al. 2010	2010			3.89 (1.84, 5.94)	2.15
Holland et al. 2010	2010		. –	1.13 (0.60, 1.66)	4.45
Scramlin et al. 2010	2010			2.77 (1.51, 4.02)	3.28
Garmyn et al. 2011	2011			0.74 (-0.35, 1.83)	3.57
Lawrence et al. 2011	2011			0.66 (0.26, 1.07)	4.60
Lawrence et al. 2011	2011		цi	1.55 (0.75, 2.34)	4.06
Peterson, 2011	2011		<u></u>	2.44 (0.48, 4.39)	2.26
Rathmann et al. 2012	2012		 	1.61 (0.85, 2.37)	4.12
Rodas-Gonzales et al. 2012	2012	-		1.37 (0.48, 2.27)	3.89
Overall (I-squared = 70.8%, p	= 0.000)		\diamond	2.30 (1.90, 2.70)	100.00
NOTE: Weights are from rando	om effects analysis		1		
	-8.68	Г О	8.	68	

Fig. 2. Forest plot of Longissimus muscle area (cm²) responses for Zilpaterol studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of zilpaterol treatmenton Longissimus muscle area. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in Longissimus muscle area, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was heterogenous as indicated by the l² of 70.8%.

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reported previously. The effect of feeding system on the magnitude of the responses highlights the need to evaluate products under the commercial conditions in which they will be used.

The effect of ZH on WBSF was large being increased by SMD 1.2 and a weighted mean difference increase of more than 0.8 kg in WBSF (<u>Table 1</u>). There

					%
Author	Year			SMD (95% CI)	Weight
Neill et al, 2008	2008			0.60 (-1.45, 2.64)	1.18
Neill et al, 2008	2008			-0.38 (-2.37, 1.62)	1.24
Vasconcelos et al, 2008	2008			-0.93 (-2.05, 0.18)	3.01
Vasconcelos et al, 2008	2008	•	-	-1.38 (-2.57, -0.20)	2.77
Vasconcelos et al, 2008	2008	• +		-1.63 (-2.87, -0.40)	2.62
Beckett et al, 2009	2009		-	-1.52 (-2.83, -0.20)	2.40
Beckett et al, 2009	2009	<u>+</u>		-2.30 (-3.81, -0.78)	1.94
Beckett et al, 2009	2009			-1.80 (-3.18, -0.42)	2.23
Beckett et al, 2009	2009	+ •		-0.60 (-1.50, 0.30)	3.89
Beckett et al, 2009	2009		-	-1.11 (-2.06, -0.17)	3.66
Beckett et al, 2009	2009 -	• 1	-	-1.18 (-2.14, -0.22)	3.62
Elam et al, 2009	2009			-0.48 (-1.59, 0.63)	3.03
Elam et al, 2009	2009 -	•	<u> </u>	-0.95 (-2.11, 0.21)	2.85
Elam et al, 2009	2009	•		-1.09 (-2.27, 0.09)	2.78
Hilton et al, 2009	2009	• +		-1.49 (-2.50, -0.49)	3.43
Kellermeir et al, 2009	2009	•		-0.81 (-2.00, 0.37)	2.77
Kellermeir et al, 2009	2009			-0.64 (-1.8 <mark>1</mark> , 0.52)	2.83
Montgomery et al, 2009a-(1)	2009 —			-0.85 (-2.32, 0.61)	2.04
Montgomery et al, 2009a-(2)	2009		-	-1.88 (-3.63, -0.13)	1.54
Montgomery et al, 2009a-(3)	2009			-0.09 (-1.47, 1.30)	2.22
Montgomery et al, 2009a-(4)	2009 -			-0.72 (-2.16, 0.73)	2.09
Montgomery et al, 2009b	2009	•!		-2.00 (-3.09, -0.91)	3.08
Robles-Estrada et al, 2009	2009			-0.52 (-1.94, 0.90)	2.15
Strydom et al. 2009	2009			-0.73 (-1.60, 0.14)	4.03
Baxa et al, 2010	2010	•		-1.85 (-3.24, -0.46)	2.20
Baxa et al, 2010	2010			-1.87 (-3.27, -0.47)	2.19
Holland et al, 2010	2010		•	-0.21 (-0.70, 0.28)	6.31
Parr et al, 2010	2010	· · · · · ·	•	0.51 (-0.56, 1.57)	3.18
Scramlin et al, 2010	2010	•		-0.78 (-1.69, 0.13)	3.83
Garmyn et al, 2011	2011		•	0. <mark>4</mark> 6 (-0.60, 1.53)	3.19
Lawrence et al, 2011	2011	+		-1.15 (-1.90, -0.40)	4.66
Peterson, 2011	2011		-	-2.03 (-3.83, -0.23)	1.47
Rathmann et al, 2012	2012	+	<u> </u>	-0.51 (-1.17, 0.16)	5.17
Rodas-Gonzales et al, 2012	2012	: =	•	0.16 (-0.64, 0.96)	4.38
Romero et al. 2009	2009			(Excluded)	0.00
Overall (I-squared = 37.3%, p	= 0.016)	\diamond		-0.86 (-1.10, -0.62)	100.00
NOTE: Weights are from rand	om effects analysis				
	-3.83		0	3.83	
	-0.00		v	0.00	

Fig. 3. Forest plot of Standardised USDA marbling score responses for Zilpaterol studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of zilpaterol treatmenton USDA Marbling Score. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in Standardised USDA marbling score, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was relatively homogenous as indicated by the l² of 37.3%.

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were several significant interactions identified by meta-regression. The magnitude of the ZH effect on WBSF decreased with days that the meat was aged. (Fig. 7). If dietary supplements or implants lower protein degradation rates in the live animal, these are likely to reduce post-mortem ageing rates of the carcase and produce tougher meat [9]. As the improvement in tenderness is exponential, it



			%
Author	Year	SMD (95% CI)	Weight
Avendano-Reves et al. 2006	2006	2 15 (0 68.3 62)	1.18
Hilton et al. 2009	2009	2.44 (1.25, 3.62)	1.59
Hilton et al. 2009	2009	1.07 (0.13, 2.01)	2.06
Hilton et al. 2009	2009	0.78 (-0.14, 1.69)	2.13
Holmer et al. 2009	2009	1.49 (0.49, 2.49)	1.93
Holmer et al. 2009	2009	2 23 (1 09 3 37)	1.66
Kellermeir et al. 2009	2009		0.43
Kellermeir et al. 2009	2009	4 13 (1 42 6 85)	0.43
Kellermeir et al. 2009	2009	3 11 (0 87, 5 35)	0.61
Leheska et al. 2009	2009	0.80 (-0.03, 1.63)	2.32
Leheska et al. 2009	2009	0.80 (-0.03, 1.63)	2.32
Leheska et al. 2000	2000	0.66 (-0.17, 1, 48)	2.35
Leheska et al. 2000	2003		2.35
Dathmann et al. 2009	2003		1.28
Pathmann et al. 2009	2003	3.57 (1.80, 5.34)	0.90
Pathmann et al. 2009	2009	3.57 (1.00, 5.54)	0.50
Pathmann et al. 2009	2009	3.00 (1.33, 3.04)	1 30
Ratimani et al. 2009	2009		1.50
Rathmann et al. 2009	2009	3.04 (1.43, 4.03)	1.04
Rathmann et al, 2009	2009	2.43 (1.00, 3.88)	1.23
Rathmann et al, 2009	2009		1.54
Rathmann et al, 2009	2009		1.21
Ratinnann et al. 2009	2009	2.00 (1.12, 4.07)	1.10
Strydom et al. 2009	2009		1.90
Brooks et al. 2010	2010	0.70 (0.18, 1.22)	3.21
Brooks et al. 2010	2010	0.97 (0.43, 1.50)	3.16
Brooks et al. 2010	2010	0.77 (0.25, 1.30)	3.20
Brooks et al. 2010	2010	0.76 (0.23, 1.28)	3.20
Brooks et al. 2010	2010		3.18
Brooks et al. 2010	2010		3.22
Brooks et al, 2010	2010	0.21 (-0.30, 0.72)	3.25
Brooks et al, 2010	2010	0.55 (0.03, 1.06)	3.22
Brooks et al, 2010	2010	0.24 (-0.27, 0.75)	3.25
Garmyn et al, 2010	2010		3.53
Garmyn et al, 2011	2011	1.67 (0.43, 2.91)	1.49
Garmyn et al, 2011	2011	1.79 (0.52, 3.06)	1.45
Garmyn et al, 2011	2011	1.49 (0.29, 2.70)	1.55
Garmyn et al, 2011	2011	0.93 (-0.19, 2.04)	1.72
Garmyn et al, 2011	2011	0.64 (-0.43, 1.72)	1.78
Rathmann et al, 2012	2012	1.44 (1.04, 1.84)	3.56
Rathmann et al, 2012	2012	1.34 (0.95, 1.74)	3.58
Rathmann et al, 2012	2012	 1.21 (0.82, 1.60)	3.60
Robles-Estrada et al, 2012	2012	0.91 (0.24, 1.57)	2.79
Robles-Estrada et al, 2012	2012	0.89 (0.23, 1.55)	2.79
Robles-Estrada et al, 2012	2012	0.63 (-0.01, 1.28)	2.84
Robles-Estrada et al, 2012	2012	1.13 (0.45, 1.81)	2.74
Hansen et al. 2012	2012	1.05 (0.10, 1.99)	2.06
Hansen et al. 2012	2012	1.70 (0.66, 2.74)	1.86
Hansen et al, 2011	2011	(Excluded)	0.00
Hansen et al, 2011	2011	(Excluded)	0.00
Overall (I-squared = 61.0%, p = 0	0.000)	0 1.21 (1.02, 1.40)	100.00
NOTE: Weights are from random	effects analysis		
	ا 22 هـ	I I 28.9 0	
	-0.00	v 0.00	

Fig. 4. Forest plot of Warner-Bratzler Shear Force responses for Zilpaterol studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of zilpaterol treatmenton Warner-Bratzler Shear Force. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in Warner-Bratzler Shear Force, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was heterogenous as indicated by the l² of 61.0%.

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Fig. 6. Raw mean differences for studies using animal or pen as the unit of interest for the effect of Zilpaterol on USDA marbling score. The dots in the Figures represent studies that are outliers.

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follows that as meat ages differences between ZH supplemented and control carcasses will decrease.

The type of study design, that is, whether the study was conducted at the animal level or the pen level also influenced WBSF. Fig. 8 showed an increase in the magnitude of the ZH effect for experiments where pen was the experimental unit. The implications drawn about the unit of evaluation for WBSF are difficult because animal was used as the unit of evaluation in some studies where carcases were randomly selected within a group of cattle fed in pens. However, all studies were in the same direction, indicating an increase in WBSF. The heterogeneity identified, however, is only in degree of increase of WBSF, reflecting quite substantial differences in study design among the studies displayed in Fig. 8.

Ractopamine

Cattle were fed on average 114.0 ± 44.8 days in the lot and the average period that cattle were exposed to RAC was 30.8 ± 5.3 days. The results of the studies, which pertain to the RAC exposure period, include the raw mean differences for variables, SMD, weighted mean differences, I² estimated heterogeneity and results of meta-regressions evaluating the effects of RAC and are provided in <u>Table 2</u>. Animals supplemented with RAC showed an increase of 0.4 and 0.8 standard deviations in BW and ADG, respectively, which were equivalent to about 8 kg and 0.19 kg/day. In contrast to ZH, the increase in weight gain with RAC was not mediated through increased DMI. The point direction was very close to a zero effect and the overall data showed a marked variation around the mean with some studies showing significant increases or significant decrease in DMI. There was no evidence in this case of missing data from the funnel plot and none of the covariates were significant in explaining the effects of the DMI.

While the increase in ADG during the period of supplementation was substantial, some caution should be expressed in regard to this result as there is some evidence of missing data on the funnel plot (Fig. 9). It appears that small negative studies may not be present in this particular data set. The covariates examined did not explain any of the variation in effect size. The gain:feed was significantly improved and quite similar in effect size to the response in regard to the ADG. This is unsurprising given the lack of significant increase in DMI observed. It clearly indicates that there are efficiency gains with treatment and the funnel plot was quite symmetric. It is possible, therefore that the funnel plots in the ADG data may be an artefact, reflecting the large number of investigations, of funnel plots in this study. Again none of the sources of variation investigated by meta-regression significantly influenced the gain to feed efficiency.

Forest plots for HCW, USDA marbling score and WBSF are presented in <u>Figs. 10, 11</u> and <u>12</u>. The Forest Plots for RAC supplementation show that the magnitudes of the RAC responses were generally smaller than those for ZH. Although not presented, funnel plots were calculated and showed little evidence of missing studies and consequent publication bias. However, there was substantial





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Fig. 8. Meta-regression of the effect of individually fed or pen fed on the standardised mean difference of studies examining Zilpaterol and Warner-Bratzler Shear Force. The regression is weighted by the effect size of studies which are indicated by the size of the marker. The larger the marker, the greater the effect size of the study.

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Fig. 9. A contour enhanced funnel plot of the effects of RAC on Average Daily Gain.

evidence of small study effects with these studies sometimes having quite large effects that differed from the larger studies.

The pH data were sparse and not significant. Evidence of repartitioning of nutrients was present in the carcase data. The SMD for LMA and dressing percentage was higher for the treated cattle and the SMD for marbling was significantly lower for the RAC treated cattle, as was the yield grade. The rib fat thickness was not significantly lower for RAC. These findings are consistent with the physiological action of RAC [5]. These results, in combination indicate that the response was to increase muscle, as indicated by increased LMA and HCW, but to decrease lipid as indicated by the marbling scores. The WBSF was also significantly increased by RAC treatment. The carcase data were less heterogenous than the feedlot data with only the HCW and LMA I² approaching or exceeding 50, respectively.

Conclusions

These findings support the previously identified physiological roles of the β -AA and provide strong evidence for producers and others to examine and consider the effects of ZH and RAC on beef cattle production. A large number of reviews and basic physiological studies have not clearly identified the mechanisms by which the actions of the specific β -AA are exerted, however, this study clearly demonstrates that the repartitioning effects are rapid, marked and highly integrated.

doi:10.1371/journal.pone.0115904.g009



				%
Author	Year		SMD (95% CI)	Weight
Cabletes (1001) (1)	1001		0.84 / 0.84 .4.04)	4.48
Schluter (1991) (1)	1331		0.04 (-0.04, 1.91)	1.10
Schluter (1991) (2)	1331	1	1.70 (0.20, 3.19)	0.91
Schuter (1991) (3)	1331		1.70 (0.20, 3.19)	0.91
Schroeder et al. 2003 (1)	2003		0.04 (-0.52, 0.59)	3.01
Schroeder et al. 2003 (2)	2003		0.10 (-0.46, 0.65)	3.01
Schroeder et al. 2003 (3)	2003		0.13 (-0.43, 0.68)	3.01
Laudert et al. 2004 (1)	2004		0.28 (-0.25, 0.77)	3.19
Laudert et al. 2004 (2)	2004		0.59 (0.09, 1.09)	3.23
Schroeder et al. 2005 (1)	2005		0.02 (-0.53, 0.58)	3.01
Schroeder et al. 2005 (2)	2005		0.09 (-0.47, 0.64)	3.01
Schroeder et al. 2005 (3)	2005		0.16 (-0.40, 0.71)	3.01
Avendano-Reyes et al. 2006	2006		4.17 (2.02, 8.33)	0.48
Walker et al. 2006 (1)	2006		0.88 (0.04, 1.73)	2.03
Waker et al. 2006 (2)	2006		0.66 (-0.16, 1.48)	2.08
Walker et al. 2006 (3)	2006		0.32 (-0.48, 1.13)	2.13
Talton, 2006	2006		1.89 (-0.18, 3.96)	0.52
Abney et al. 2007 (1)	2007	•	0.12 (-0.86, 1.10)	1.68
Abney et al. 2007 (2)	2007		0.35 (-0.64, 1.34)	1.66
Abney et al. 2007 (3)	2007	• • •	-0.05 (-1.03, 0.93)	1.68
Abney et al. 2007 (4)	2007		0.30 (-0.69, 1.28)	1.67
Abney et al. 2007 (5)	2007		0.49 (-0.51, 1.48)	1.65
Abney et al. 2007 (6)	2007		0.30 (-0.68, 1.29)	1.67
Gruber et al. 2007	2007		0.24 (-0.38, 0.85)	2.80
Sissom et al 2007 (1)	2007	T T	1.50 (-0.13, 3.13)	0.79
Sissom et al 2007 (2)	2007	-+	1.00 (-0.50, 2.50)	0.90
Sissom et al 2007 (3)	2007		0.35 (-0.64, 1.34)	1.66
Sissom et al 2007 (4)	2007		1.06 (0.00, 2.12)	1.52
Sissom et al 2007 (5)	2007	• <u>•</u>	-0.18 (-1.16, 0.81)	1.68
Winterholler et al. 2007	2007		1.15 (0.28, 2.02)	1.95
Quinn et al. 2008 Exp 1 (1)	2008		0.12 (-0.69, 0.92)	2.14
Quinn et al. 2008 Exp 2 (2)	2008		0.51 (-0.39, 1.40)	1.89
Quinn et al. 2008 Exp 2 (3)	2008		0.38 (-0.51, 1.26)	1.91
Quinn et al. 2008 Exp 2 (4)	2008		1.01 (0.08, 1.95)	1.78
Winterholler et al. 2008 (1)	2008		1.15 (-0.63, 2.94)	0.67
Winterholler et al. 2008 (2)	2008		0.00 (-1.60, 1.60)	0.81
Winterholler et al. 2008 (3)	2008	+	0.85 (-0.18, 1.88)	1.58
Winterholler et al. 2008 (4)	2008	<u>;</u>	-1.09 (-2.19, 0.01)	1.45
Allen et al. 2009	2009		-0.04 (-0.99, 0.91)	1.75
Holmer et al. 2009 (1)	2009	TI	3.97 (1.33, 6.61)	0.33
Holmer et al. 2009 (1)	2009		0.21 (-1.18, 1.60)	1.02
Vogel et al. 2009 (1)	2009	 →→	0.72 (0.09, 1.34)	2.73
Vogel et al. 2009 (2)	2009		0.78 (0.15, 1.41)	2.72
Vogel et al. 2009 (3)	2009	+	0.56 (-0.19, 1.32)	2.28
Bryant et al. 2010 (1)	2010	•	0.07 (-1.06, 1.20)	1.38
Bryant et al. 2010 (2)	2010		0.44 (-0.71, 1.58)	1.36
Bryant et al. 2010 (3)	2010		0.22 (-0.35, 0.79)	2.96
Scramlin et al. 2010	2010		1.23 (0.26, 2.19)	1.72
Woerner et al. 2011 (1)	2011	+	0.02 (-0.30, 0.35)	3.97
Woerner et al. 2011 (2)	2011		0.13 (-0.19, 0.46)	3.98
Jennings, 2011	2011		1.15 (0.14, 2.15)	1.63
Peterson, 2011	2011		0.43 (-0.98, 1.83)	1.00
Strydom et al	2011	+ + -	0.60 (-0.22, 1.42)	2.09
Boler et al. 2012 (1)	2012		4.35 (2.13, 6.57)	0.46
Boler et al. 2012 (2)	2012		4.91 (2.48, 7.34)	0.39
Overall (I-squared = 48.5%, p = 0.000)		a	0.47 (0.31, 0.63)	100.00
		1 T	,	
NOTE: Weights are from random effects	analysis			
	1	1 1		
	-7.34	0 7.3	34	

Fig. 10. Forest plot of hot carcass weight for ractopamine studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of ractopamine treatmenton hot standing carcase weight. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in hot carcass weight, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was heterogenous as indicated by the I² of 46.5%.

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			%
Author	Year	SMD (95% CI)	Weight
Schroeder et al. 2003 (1)	2003	-0.05 (-0.61, 0.50)	3.56
Schroeder et al. 2003 (2)	2003	-0.02 (-0.58, 0.53)	3.56
Schroeder et al. 2003 (3)	2003	-0.09 (-0.65, 0.46)	3.56
Laudert et al. 2004 (1)	2004	-0.00 (-0.51, 0.50)	4.24
Laudert et al. 2004 (2)	2004	-0.07 (-0.58, 0.42)	4.54
Walker et al. 2006 (1)	2008	0.08 (-0.72, 0.88)	1.72
Waker et al. 2006 (2)	2008	-0.68 (-1.51, 0.14)	1.62
Walker et al. 2006 (3)	2008	0.03 (-0.77, 0.83)	1.72
Talton, 2006	2008	0.17 (-1.44, 1.77)	0.43
Abney et al. 2007 (1)	2007	0.33 (-0.65, 1.32)	1.13
Abney et al. 2007 (2)	2007	0.88 (-0.15, 1.92)	1.04
Abney et al. 2007 (3)	2007	0.21 (-0.78, 1.19)	1.15
Abney et al. 2007 (4)	2007	-0.16 (-1.14, 0.83)	1.15
Abney et al. 2007 (5)	2007	0.55 (-0.45, 1.55)	1.10
Abney et al. 2007 (8)	2007	0.16 (-0.82, 1.14)	1.15
Gruber et al. 2007	2007	-0.42 (-1.03, 0.19)	2.93
Sissom et al 2007 (1)	2007	-1 00 (-2 50, 0 50)	0.50
Sissom et al 2007 (2)	2007	-0.38 (-1.78, 1.03)	0.56
Sissom et al 2007 (2)	2007	-0.29 (-1.28, 0.69)	1 14
Sissom et al 2007 (4)	2007	-0.35 (-1.34, 0.64)	1 13
Sissom et al 2007 (5)	2007	0.29 (-0.69, 1.29)	1.10
Wisterheller et al. 2007 (5)	2007	0.29 (-0.09, 1.26)	1.14
Quine et al. 2009 Eve 1 (1)	2009	0.38 (-0.42, 1.13)	1.05
Quint et al. 2008 Exp 1 (1)	2008	-0.29 (-1.09, 0.32)	1.71
Quinn et al. 2008 Exp 2 (2) Quinn et al. 2008 Exp 2 (2)	2008	-0.17 (-1.03, 0.71)	1.43
Quint et al. 2008 Exp 2 (3)	2008	-0.00 (-1.00, 0.20)	1.30
Quinn et al. 2008 Exp 2 (4)	2008	-0.27 (-1.15, 0.01)	1.43
Winternoller et al. 2008 (1)	2008	-2.18 (-4.38, 0.02)	0.23
Winternoller et al. 2008 (2)	2008	-1.03 (-2.78, 0.72)	0.30
Winternoller et al. 2008 (3)	2008	0.53 (-0.47, 1.53)	1.11
Winternoller et al. 2008 (4)	2008	0.72 (-0.33, 1.77)	1.00
Allen et al. 2009	2009	-0.10 (-1.05, 0.86)	1.22
Holmer et al. 2009 (1)	2009	2.11 (0.28, 3.94)	0.33
Holmer et al. 2009 (1)	2009	0.07 (-1.31, 1.46)	0.58
Vogel et al. 2009 (1)	2009	-0.18 (-0.79, 0.43)	2.99
Vogel et al. 2009 (2)	2009	-0.08 (-0.69, 0.52)	3.00
Vogel et al. 2009 (3)	2009	-0.15 (-0.89, 0.60)	2.00
Bryant et al. 2010 (1)	2010	0.15 (-0.98, 1.29)	0.88
Bryant et al. 2010 (2)	2010	-0.16 (-1.30, 0.97)	0.86
Bryant et al. 2010 (3)	2010	-0.26 (-0.83, 0.30)	3.39
Scramlin et al. 2010	2010	-0.47 (-1.38, 0.42)	1.40
Woerner et al. 2011 (1)	2011	-0.10 (-0.43, 0.23)	9.94
Woerner et al. 2011 (2)	2011	-0.16 (-0.48, 0.17)	10.13
Jennings, 2011	2011	-0.77 (-1.74, 0.19)	1.20
Peterson, 2011	2011	-1.77 (-3.49, -0.06)	0.38
Strydom et al	2011	0.39 (-0.42, 1.20)	1.69
Boler et al. 2012 (1)	2012	-0.50 (-1.65, 0.65)	0.83
Boler et al. 2012 (2)	2012	-0.78 (-1.96, 0.41)	0.79
Glanc, 2013 (1)	2013	0.92 (0.03, 1.80)	1.42
Glanc, 2013 (2)	2013	-1.25 (-2.18, -0.33)	1.30
Glanc, 2013 (3)	2013	-0.52 (-1.37, 0.33)	1.53
Glanc, 2013 (4)	2013	-0.18 (-1.02, 0.66)	1.58
Glanc, 2013 (5)	2013	0.15 (-0.69, 0.99)	1.58
Glanc, 2013 (8)	2013	-0.08 (-0.91, 0.76)	1.58
Overall (I-squared = 0.8%, p = 0.457)	Q	-0.11 (-0.21, -0.00)	100.00
NOTE: Weights are from random effect	is analysis		
	-4.38 0	4.38	
	·····		

Fig. 11. Forest plot of Standardised USDA marbling score for Ractopamine studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of ractopamine treatmenton USDA marbling score. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in Standardised USDA marbling score, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was homogenous as indicated by the l² of 0.8%.

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Fig. 12. Forestplot of Warner-Bratzler Shear Force for ractopamine studies. A Forest plot of the effect size or standardized mean difference (standardized using the z-statistic) and 95% confidence interval of the effect of ractopamine treatmenton Warner-Bratzler Shear Force. The weights that each study contributed are in the right hand column and are indicated by the size of the box. The larger the box, the greater the study contribution to the overall estimate. The solid vertical grey line represents a mean difference of zero or no effect. Points to the left of the line represent a reduction in Warner-Bratzler Shear Force, while points to the right of the line indicate an increase. The upper and lower limit of the line connected to the square represents the upper and lower 95% confidence interval for the effect size. The overall pooled effects size and 95% confidence interval is indicated by the diamond at the bottom. This effect was homogenous as indicated by the l² of 0%.

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Once these results have been critically reviewed by others, they can be immediately applied and used to formulate strategies to make best use of agents that markedly improve the efficiency of production. Registration is being sought for these products in countries including Australia. The adoption of these technologies will necessarily involve a consideration of the benefits in production costs and associated environmental benefits of improved efficiencies of resource use against the effects on shear force.

This work provided information on the effects of ZH and RAC on production, efficiency and meat quality. The meta-analysis provided more precise and robust estimates of the effects of the β -AA on efficiency of production and carcase quality measures. We identified, using meta-regression, that the method of feeding cattle

may influence responses to ZH and that aging of steaks can reduce the effects of ZH on WBSF.

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Supporting Information

S1 Checklist. PRISMA Checklist. doi:10.1371/journal.pone.0115904.s001 (DOC)

S1 File. Contains the following files: Table S1. Authors, experimental unit, number of cattle and pens per group, and means for control (Con) and treatment (ZH) groups for final body weight (BW), average daily gain (ADG), dry matter intake (DMI) and Gain/Feed (G/F) ratio. Table S2. Authors, number of cattle per group, experimental unit, number of cattle and carcasses per group for analysing carcass characteristics in experiments with control (Con) and treatment (ZH) groups. Table S3. Authors and study means for control (Con) and treatment (ZH) groups for hot carcass weight (HCW), longissimus muscle area (LMA), ultimate pH and muscle colour score. Table S4. Authors and study means for control (Con) and treatment (ZH) groups for USDA marbling score, fat thickness (mm) and dressing percentage (D%). Table S5. Author, experimental unit, days aged, number of animals, number of pens, number of muscles and study means for control (Con) and treatment (ZL) groups for Warner Bratzler shear force (WBSF) and cooking loss percentage (CL%). Table S6. Zilpaterol studies that did not meet the selection criteria, had insufficient data or data provided were not appropriate. Table S7. Authors, experimental unit, number of cattle and pens per group, and means for control (Con) and ractopamine treatment (RAC) groups for final body weight (BW), average daily gain (ADG), dry matter intake (DMI) and Gain/Feed (G:F) ratio. Table S8. Authors, \and means for control (Con) and ractopamine treatment (RAC) groups for hot carcass weight (HCW), longissimus muscle area (LMA) and ultimate muscle pH. Table S9. Authors and means for control (Con) and ractopaminetreatment (RAC) groups for USDA marbling score, fat thickness (mm) and dressing percentage (D%). Table S10. Authors, experimental unit, days aged, number of animals, number of pens, number of muscles and means for control (Con) and ractopamine treatment (RAC) groups for Warner –Bratzler shear force and cooking loss. Table S11. List of ractopamine studies that did not meet the selection criteria, insufficient data or data provided were not appropriate.

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Author Contributions

Conceived and designed the experiments: IJL JMT. Performed the experiments: IJL JMT. Analyzed the data: IJL. Wrote the paper: IJL JMT FRD.

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